Timing of renal replacement therapy for patients with acute kidney injury: A systematic review and meta-analysis

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Mark Andonovic¹, Richard Shemilt², Malcolm Sim^{1,3}, Jamie P Traynor², Martin Shaw¹, Patrick B Mark^{2,4} and Kathryn A Puxty^{1,5}

Abstract

Background: Acute kidney injury is associated with high mortality, and the optimal time to start renal replacement therapy for acute kidney injury is unknown despite several randomised controlled trials on the subject. We performed a systematic review and meta-analysis to assess the effect of earlier initiation of renal replacement therapy for acute kidney injury on mortality and reported secondary outcomes.

Methods: All literature in databases EMBASE, MEDLINE and CENTRAL was searched from January 1970 to March 2019 using terms related to renal replacement therapy, timing and randomised controlled trials. All randomised controlled trials with 25 or more adult participants suffering from acute kidney injury comparing timing of renal replacement therapy were included. The results of the selected studies were pooled and expressed in terms of risk ratios (RR) and 95% confidence intervals (95% CI) using a random effects model.

Results: A total of 7008 records were identified; 94 were selected for full text review of which 10 were included in the final meta-analysis. The 10 studies comprised 1956 participants (989 'early' group; 967 'late' group) with 918 total deaths; the analysis demonstrated no significant differences between the 'early' and 'late' renal replacement therapy groups (RR = 0.98 (95% CI = 0.84, 1.15)) for mortality. No significant differences between groups were evident for period-wise mortality; dialysis dependence; recovery of renal function; length of intensive care unit or hospital stay; or number of renal replacement therapies, mechanical ventilation and vasopressor-free days.

Conclusions: Current evidence does not support the use of early renal replacement therapy for patients with acute kidney injury. Data from ongoing and future randomised controlled trials are required to strengthen the evidence base in the area.

Keywords

Acute kidney injury, renal replacement therapy, timing, meta-analysis

Introduction

Acute kidney injury (AKI) is common within the critically ill and hospitalised patients. AKI, an evolution of the term acute renal failure (ARF), has been subjected to several classifications,^{1–3} making the reported incidence of AKI in patients admitted to intensive care units (ICUs) vary significantly (35%-67%).^{4–7} Owing to the high incidence within the critically ill, an increase in the severity of AKI is associated with increasing all-cause mortality of up to 57%.^{6–8}

Renal replacement therapy (RRT) is a key strategy in the treatment of severe AKI with life-threatening ¹Academic Unit of Anaesthesia, Pain and Critical Care, University of Glasgow, Glasgow, UK

²Department of Nephrology, Queen Elizabeth University Hospital, Glasgow, UK

³Department of Intensive Care, Queen Elizabeth University Hospital, Glasgow, UK

⁴Institute of Cardiovascular and Medical Sciences, University of Glasgow, Glasgow, UK

³Department of Intensive Care Medicine, Glasgow Royal Infirmary, Glasgow, UK

Corresponding author:

Mark Andonovic, Academic Department of Anaesthesia, Critical Care and Pain, University of Glasgow, Floor 2, New Lister Building, Glasgow Royal Infirmary, Glasgow, 8–16 Alexandra Parade, Glasgow G31 2ER, UK.

Email: mark.andonovic@glasgow.ac.uk

metabolic acidosis and volume overload unresponsive to medical therapy. While RRT is accepted as an impactful treatment, its implementation remains a matter of debate. Studies have compared differences between modalities of RRT, such as intermittent haemodialysis versus continuous RRT,^{9,10} haemofiltration versus haemodialysis¹¹ or dose.¹²

Further, the timing to initiate RRT for AKI remains a challenge. Many randomised controlled trials (RCTs) have been executed to determine whether 'early' compared to 'delayed' initiation is of benefit; two studies^{13,14} reported evidence on the subject in 2016, followed by several meta-analyses.^{15–17} However, a disparity between conclusions persists, with reports that no difference is evident between groups,^{16,17} while others conclude that earlier initiation of RRT conveys a decrease in mortality.^{18,19} Three subsequent RCTs published in 2018 added further data.^{20–22} A Cochrane Review was also published but excluded studies of patients not admitted to ICU.²³

Our objective was to conduct a systematic review and meta-analysis on all patients suffering from AKI who required RRT. Analysis would be carried out on studies comparing timing of the initiation of RRT in two groups of patients: the first group classified as 'early' and the second group classified as 'late', 'delayed' or 'standard treatment'. The studies must report on all-cause mortality to be included in the analysis. We specifically add the three RCTs published in 2018 to update previous meta-analyses and assess what these new data contribute to this area of study.

Methods

Registration

The review is registered with PROSPERO's Register of Systematic Reviews, ID Number: CRD42019145074.

https://www.crd.york.ac.uk/prospero/display_ record.php?ID=CRD42019145074

Eligibility criteria

The inclusion criteria followed agreed RCT guidelines on reporting differences in the timing of initiating RRT (early vs. late, standard vs. early, early vs. delayed). Non-RCTs, the paediatric population and patient population without AKI, were excluded. No guidelines as to defining RRT timing exist; therefore, the definition of 'early' and 'late' is according to the individual studies' interpretation unless the definition of 'late' was outwith that considered a 'standard' RRT initiation which has resulted in two 'early' group classifications. Studies that defined the 'late' group as initiation within 12 h of diagnosis with any stage AKI were also excluded.

(EMBASE, MEDLINE and Three databases Cochrane Central Register of Controlled Trials (CENTRAL)) were interrogated for the period January 1974 to March 2019. The search strategy was as broad as possible to capture all RCTs conducted on the subject; the only filter applied was to restrict results to English language. MEDLINE and CENTRAL searches used MeSH terms (supplementary Figure s1.1). A near identical search interrogated EMBASE, but with certain terms altered to match Emtree headings (Figure s1.2). In order to identify ongoing or not published completed trials, the International Trials Registry (https://www.who.int/ ictrp/en/) and the National Institutes of Health's registry (https://www.clinicaltrials.gov) were searched.

Study selection

Two authors (MA, RS) independently compiled a list of citations gathered from the three sources. Obvious duplicate citations were removed by databases when merging; however, if any two citations had discrepancies, they were both retained for title review. Both authors reviewed the titles independently and selected eligible studies for abstract review; a thorough abstract review was then conducted to select studies eligible for full text review. A concluding, full text review was then executed and any differences between the two reviewers were referred to a third reviewer (KP) to make a final decision on eligibility.

Data extraction

The papers were each initially assessed for time-period mortality reported on and then the data were recorded independently using a pre-defined form. Two independent reviewers (MA, RS) extracted key data including the number of patients recruited, definition of 'early' and 'late' RRT groups and measured outcomes. After consolidation, data on the number of events and the total for both 'early' and 'late' groups were collected and outcomes in terms of mean, median, mode and interquartile ranges were extracted as reported.

Outcome measures

The following outcomes were extracted:

Primary outcomes. Overall mortality rate, in-ICU, in-hospital, 28-, 60- and 90-day mortality rates.

Secondary outcomes. Dialysis dependence at 28, 60 and 90 days, recovery of renal function (return to baseline) at 90 days, adverse events, length of ICU stay, length of hospital stay, number of RRT days, number of

RRT-free days, number of mechanical ventilation-free days and number of vasopressor-free days.

Risk of bias

Each study was assessed independently by two authors (MA, RS) for potential risk of bias using the seven domains cited in the Cochrane Collaboration's tool²⁴; a funnel plot categorised the risk of publication bias across the studies. The quality of evidence for the primary outcomes was assessed independently using the GRADE tool.²⁵

Data synthesis

The results were expressed in terms of risk ratio (RR) and 95% confidence intervals (95% CI) for mortality and secondary outcomes. Heterogeneity between studies was determined through the I^2 statistic; a value of >40% was interpreted as a significant degree of heterogeneity. RR for each outcome was estimated using both fixed and random effects to identify high degrees of heterogeneity between studies. Statistical comparison was captured as a *P*-value for each analysis; a value of <0.05 was considered statistically significant. Any outcome reported in terms of continuous data was expressed in terms of pooled raw differences between the two groups medians (a negative difference favouring early RRT) and 95% CIs. This has been previously described as comparing favourably to methods which transform medians and IQRs to mean and standard deviation.²⁶ All data were analysed using the software R (R version 3.5.1, The R Foundation).

The following pre-defined sub-groups were analysed for overall mortality to assess possible sources of heterogeneity including risk of bias, RRT modality, severity of illness and patient population.

- Low risk versus high or unclear risk of bias
- Intermittent haemodialysis versus continuous RRT versus mixed
- ICU-only population versus mixed population
- Medical versus surgical versus mixed patients

Results

Selected studies

The literature search returned a total 7008 references after duplicate removal (Figure 1). The features of the 10 studies selected for inclusion in the review are



Figure 1. PRISMA flow diagram.

AKI: acute kidney injury; CKD: chronic kidney disease; RRT: renal replacement therapy.

detailed in Table 1; studies varied in size from 28 patients²⁹ to 488 patients.²² Of the 10, 8 included ICU patients only.

Overall mortality

The 10 studies comprising 1956 patients reported on overall all-cause mortality at varying times: 989 into the 'early' and 967 into the 'late' groups. A total of 918 deaths were reported; 459 in the 'early' and 459 in the 'late' group, corresponding to a mortality rate of 46.4% for patients receiving early and 47.5% for those receiving conventional/ late RRT.

Figure 2 illustrates results from the 10 studies depicting no significant difference between 'early' or 'late' initiation of RRT for mortality rates: RR = 0.98 (95% CI = 0.84,1.15 (random effects modelling)). A marked heterogeneity between studies was evident with an I^2 of 46% (P = 0.05). Pre-defined subgroup analyses were carried out to further explore the possible cause.

Impact on mortality after accounting for risk of bias

Two studies were assessed to have either a high or unclear risk of bias (supplementary Table s1.1), and their pooled results suggested a mortality benefit for 'early' RRT (Figure s1.3); RR = 0.37 (95% CI = 0.08,1.65). The remaining eight studies were assessed as low risk of bias, with pooled results showing no statistically significant difference between groups; RR = 1.00 (95% CI = 0.89,1.13). The heterogeneity in the low risk of bias group decreased to I^2 of 23% (from the overall analysis value of 46%).

Impact on mortality after accounting for RRT modality

The RRT modality used to deliver the intervention and its impact on mortality is presented in supplementary Figure s1.4; two studies used intermittent haemodialysis with no significant difference between the 'early' and 'late' arms: RR = 1.30 (95% CI = 0.63,2.70); four used only continuous RRT with no significant difference between groups: RR = 0.91 (95% CI = 0.57,1.46) and the remaining four studies utilised a mixture of these two modalities and also found no significant difference between groups: RR = 0.95 (95% CI = 0.81,1.11).

Impact on mortality after consideration of critical illness

Two studies included all inpatients (supplementary Figure s1.5). The difference between the 'early' and 'late' groups was not statistically significant; RR = 1.30 (95% CI = 0.63, 2.70). The remaining

eight studies included ICU patients only with no observable difference between the two groups: RR = 0.95 (95% CI = 0.80,1.12).

Impact on mortality by admission type: Medical versus surgical versus mixed population

Two studies only involved patients from a medical cohort (supplementary Figure s1.6); no differences in mortality between the two RRT groups were observed: RR = 1.30 (95% CI = 0.63, 2.70). Only one study used participants from the surgical cohort, the result indicating a mortality benefit in the early RRT group: RR = 0.17 (95% CI = 0.05, 0.61). The remaining seven studies contained a mixed population of patients and no statistical difference existed between groups: RR = 0.98 (95% CI = 0.88, 1.10).

Time-based mortality

All studies reported mortality numbers over differing time periods, sub-categorised into in-ICU, in-hospital, 28, 60 and 90 days (Figure 3). ICU mortality was reported by two studies; no statistical difference was evident between the two RRT treatment groups: RR = 1.02 (95% CI = 0.66, 1.58). In-hospital mortality was reported by three studies with no significant difference between groups: RR = 1.16(95%) CI = 0.84, 1.60). Six reported 28-day mortality with no significant difference found: RR = 0.99 (95%) CI = 0.88, 1.11). Two reported 60-day mortality with no significant difference observed: RR = 0.89, 95% CI = 0.71, 1.12). Three reported 90-day mortality with no statically significant difference between early and late groups: RR = 0.93 (95% CI = 0.69, 1.23).

Secondary outcomes

Dialysis dependence was reported at 28, 60 and 90 days. Four studies reported on rates of dialysis dependence in surviving patients after 90 days (supplementary Figure s1.7). The pooled data demonstrated no significant differences between 'early' and 'late' groups: 16/279 versus 18/289 patients (RR = 0.87). Four studies reported on rates of dialysis dependence after 28 days with no statistically significant difference (supplementary Figure s1.8): 65/423 versus 76/425 patients (RR = 0.84). Dialysis dependence at day 60 was reported by two studies with a benefit suggested in the early RRT group (supplementary Figure s1.5): 14/226 versus 22/214 patients (RR = 0.59).

No statistically significant differences between the 'early' and 'late' groups was observable for all adverse events except catheter-related complications; the results of these can be found in Table 2 with forest plots presented in supplementary Figures s1.9–s1.17. Analysis of the six studies reporting catheter-related complications (supplementary Figure s1.18) suggested

omes	all mortality h of hospital stay 'se events	hospital and day mortality rery of renal ction at 90 days ion of ICU and spital stay se events	y mortality ges in BP, urine put and atinine	spital mortality is dependence at days oer of RRT days 'se events	hospital and day mortality is dependence at days h of ICU and spital stay 'se events
Outc	Over? Lengtl Adver	ICU, I 28- 28- 28- 28- 28- 10- 10- 10- 10- 10- 10- 10- 10- 10- 10	I4-da Chang out	In hos Dialys 90 Numb Adver	ICU, I 90- 70- 90 90 Length host Adver
Late definition for RRT	Conservative management	Plasma Urea level >40 mmol/l, potassium >6.5 mmol/l or severe pulmonary oedema	Urine output <20 ml/h for two consecutive hours (or daily urinary output 500 ml or less)	Treatment refractory hyperkalaemia, volume overload, acidosis. Uremic nausea and anorexia with inability to maintain oral intake	Potassium >6.0 mmol/l, serum bicarbonate <10 mmol/l, PaO2/FiO2 <200 with infiltrates on chest radiograph sug- gestive of pulmonary oedema
Early definition for RRT	Early hemodialysis (as soon as met eligibility criteria)	RRT started within 12 h of inclusion	Urine output <30 ml/h for three consecu- tive hours (or daily urinary output 750 ml or less)	Serum urea >70 mg/dl and/or creatinine level >7 mg/dl	RRT started within 12 h of fulfilling eli- gibility criteria
RRT modality	ОHI	CRRT	CRRT	ЧH	Mixed
Numbers	Total = 35 Early = 18 Late = 17	Total = 106 Early = 70 Late = 36	Total = 28 Early = 14 Late = 14	Total = 208 Early = 102 Late = 106	Total = 100 Early = 48 Late = 52
Inclusion criteria	Diagnosis of acute tubular necrosis with serum creatinine <7 mg% and blood urea <120 mg%	Urine output <30 ml/h for >6 h and creatinine clearance <20 ml/min	Post-CABG patients. Hourly urine output <30 ml/h and serum creatinine increased at rate of 0.5 mg/dl/day or more	Severe AKI with increas- ing serum urea and creatinine levels	Volume repletes severe AKI with two criteria from three: creatinine doubled from baseline, urine output <6 ml/kg in last 12 h or whole blood NGAL >400 ng/ ml. Absence of urgent indications for RRT.
Setting/patient group	Single centre; all inpa- tients; medical only	Two centres, single country; ICU only; mixed patients	Single centre; ICU only; surgical only	Single centre; all inpa- tients; medical only	Multiple centres, single country; ICU only; mixed patients
Study	Pursnani et al. ²⁷	Bouman et al. ^{28.a}	Sugahara and Suzuki ²⁹	Jamale et al. ³⁰	Wald et al. ³¹

Table I. Study characteristics.

Table I. Continued

Outcomes	28- and 60-day mortality Dialysis dependence at 28 and 60 days Length of ICU and hospital stay Number of RRT, mechanical ventilation and vasopressor free days	28-, 60- and 90-day mortality Dialysis dependence at 28, 60 and 90 days Length of ICU and hospital stay Length of mechanical ventilation and RRT Adverse events	28-day mortality Dialysis dependence at 28 days Mechanical ventilation free days ICU-free days ICU-free days Renal recovery at 28 days Balance of input and output fluid
Late definition for RRT	Urea >40 mmol/l, potassium >6 mmol/l (or >5.5 mmol/l des- pite medical treat- ment), $pH < 7.15$, pulmonary oedema due to fluid overload requiring oxygen >5 l/ or FiO ₂ >50%, oliguria or anuria >72 h	Commenced within 12 h of diagnosis of stage 3 AKI, or if urea >100 mg/dL, potassium >6.0 mmol/l and or ECG changes, urine output <200 ml in 12 h or organ oedema resistant to diuretic treatment	Severe refractory acidosis (pH $<$ 7.2 or HCO ₃ $<$ 15), severe peripheral oedema, pulmonary oedema, no response to diuretics, refractory hyperkalaemia (K $>$ 6.2 or ECG changes), anuria or oliguria or high BUN (>60)
Early definition for RRT	RRT commenced within 6 h after documentation of KDIGO ³ stage 3 AKI	RRT started within 8 h of diagnosis of KDIGO stage 2 AKI	RRT started within 12 h of randomization
RRT modality	Mixed	Mixed	CRRT
Numbers	Total = 619 Early = 311 Late = 308	Total = 231 Early = 112 Late = 119	Total = 40 Early = 20 Late = 20
Inclusion criteria	KDIGO ³ stage 3 AKI compatible with a diag- nosis of ischaemic or toxic acute tubular necrosis and receiving mechanical ventilation and/or catecholamine infusion.	KDIGO stage 2 AKI (baseline creatinine doubled or urinary output <0.5 ml/kg/h for >12 h) despite optimal resuscitation, NGAL >150 ng/ml and one of: severe sepsis, use of vasopressors, refrac- tory fluid overload and progression of non- renal organ dysfunction	Patients aged 18 or older diagnosed with AKI by RIFLE criteria ¹
Setting/patient group	Multiple centres, single country; ICU only; mixed patients	Single centre; ICU only; mixed patients	Single centre; ICU only; mixed patients
Study	Gaudry et al. ¹³	Zarbock et al. ¹⁴	Srisawat et al. ^{20,b}

Table I. Continued

Study	Setting/patient group	Inclusion criteria	Numbers	RRT modality	Early definition for RRT	Late definition for RRT	Outcomes
Lumlertgul et al. ^{21.c}	Multiple centres, single country; ICU only; mixed patients	AKI with diagnosis of Acute Tubular Necrosis, clinically resuscitated and euvolaemic, no urgent indication or contra- indications for RRT.	Total = 118 Early = 58 Late = 60	CRRT	RRT was started in the early group within 6 h of randomisation	Urea > 100 mg/dl, potas- sium >6 mmol/l, serum bicarbonate < 12 mmol/ l, pH <7.15, PaO_2/FiO_2 ratio <200 or chest radiographs compatible with pulmonary oedema	28-day mortality Dialysis dependence and renal recovery at 28 days Length of ICU and hospital stay Number of RRT and mechanical ventila- tion-free days
Barbar et al. ²²	Multiple centres, single country; ICU only; mixed patients	Early phase of septic shock (within 48 h of start of vasopressor therapy) developing AKI with at least one criterion of the failure stage of the RIFLE classification system ¹	Total = 488 Early = 246 Late = 242	Mixed	RRT commenced within 12 h of documentation of 'failure' stage AKI'	RRT commenced 48 h after diagnosis of AKI or if prior to this: serum potassium >6.5 mmol/l, pH <7.15 or fluid overload with pulmonary oedema	 28, 90 and 180 day mortality Dialysis dependence at 28 and 90 days Length of ICU and hospital stay RRT, mechanical venti- lation and vasopres- sor-free days
AKI: acute kidney injury; E <digo: disease="" in<="" kidney="" td=""><td>3UN: blood urea nitrogen; CA nproving global outcomes; NG</td><td>ABG: coronary artery bypass graft; 3AL: neutrophil gelatinase-associat</td><td>CRRT: continuou ed lipocalin; RIFLE</td><td>is renal replacen E: isk, injury, failu</td><td>nent therapy; ECG: electrocardure, loss, end-stage kidney dise</td><td>diogram; ICU: intensive care unit; ease; RRT: renal replacement ther</td><td>IHD: ischaemic heart disease: apy.</td></digo:>	3UN: blood urea nitrogen; CA nproving global outcomes; NG	ABG: coronary artery bypass graft; 3AL: neutrophil gelatinase-associat	CRRT: continuou ed lipocalin; RIFLE	is renal replacen E: isk, injury, failu	nent therapy; ECG: electrocardure, loss, end-stage kidney dise	diogram; ICU: intensive care unit; ease; RRT: renal replacement ther	IHD: ischaemic heart disease: apy.

^aPatients in early group split into low-volume (*n* = 35) and high-volume (*n* = 35) hemofiltration. ^bPatients were tested for plasma neutrophil gelatinase associated lipocalin (pNGAL) levels after recruitment. Patients with pNGAL level greater than or equal to 400ng/ml were randomized into early or late groups. ^cPatients underwent a furosemide stress test first. If they were non-responsive they were randomised into early or late groups.

Overall Mortal	ity								
Study	Experim Events	ental Total	Co Events	ntrol Total	Risk Ratio	RR	95%-CI	Weight (fixed)	Weight (random)
Pursnani et al. (1997) Bouman et al. (2002) Sugahara et al. (2004) Jamale et al. (2013) Wald et al. (2015) Gaudry et al. (2016) Zarbock et al. (2016) Srisawat et al. (2018) Barbar et al. (2018)	4 31 2 21 18 150 44 10 36 143	18 70 14 102 48 311 112 20 58 236	5 14 12 13 19 153 65 9 35 134	17 36 14 106 52 308 119 20 60 235		0.76 1.14 0.17 1.68 1.03 0.97 0.72 1.11 1.06 1.06	$\begin{matrix} [0.24; 2.35] \\ [0.70; 1.85] \\ [0.05; 0.61] \\ [0.89; 3.17] \\ [0.62; 1.71] \\ [0.83; 1.14] \\ [0.54; 0.95] \\ [0.58; 2.14] \\ [0.79; 1.43] \\ [0.91; 1.24] \end{matrix}$	1.1% 4.0% 2.6% 2.8% 4.0% 33.3% 13.7% 2.0% 7.5% 29.1%	1.8% 7.5% 1.4% 4.9% 7.0% 21.7% 14.6% 4.7% 14.1% 22.3%
Fixed effect model Random effects mode Heterogeneity: $I^2 = 46\%$, 2	l t ² = 0.0222	989 2, p = 0	.05	967	0.1 0.5 1 2 10 Favours Early Favours Late	0.98 0.98	[0.89; 1.07] [0.84; 1.15]	100.0% 	 100.0%



Mortality					
Study	Experimental Events Total Eve	Control nts Total	Risk Ratio	Weight V RR 95%−Cl (fixed) (rar	Veight ndom)
In ICU Mortality Bouman et al. (2002) Wald et al. (2015) Fixed effect model Random effects model Heterogeneity: I ² = 0%, t	$23 70 \\ 13 48 \\ 118 \\ 18 \\ 2^2 = 0, p = 0.51$	10 36 16 52 88		1.18 [0.63; 2.21] 1.6% 0.88 [0.47; 1.63] 1.8% 1.02 [0.66; 1.58] 3.4% 1.02 [0.66; 1.58]	1.4% 1.5% 2.9%
In Hospital Mortality Bouman et al. (2002) Jamale et al. (2013) Wald et al. (2015) Fixed effect model Random effects model Heterogeneity: J ² = 5%, t	$\begin{array}{ccc} 31 & 70 \\ 21 & 102 \\ 16 & 48 \\ & 220 \\ 2 \\ = 0.0041, \ \rho = 0.35 \end{array}$	14 36 13 106 19 52 194	*	1.14 [0.70; 1.85] 2.2% - 1.68 [0.89; 3.17] 1.5% 0.91 [0.53; 1.56] 2.2% 1.19 [0.87; 1.64] 5.9% 1.16 [0.84; 1.60]	2.4% 1.4% 1.9% 5.7%
28-Day Mortality Bouman et al. (2002) Gaudry et al. (2016) Zarbock et al. (2016) Srisawat et al. (2018) Lumlertgul et al. (2018) Barbar et al. (2018) Fixed effect model Random effects model Heterogeneity: J ² = 0%, t	$\begin{array}{cccccccccccccccccccccccccccccccccccc$	9 36 134 308 48 119 9 20 35 60 102 242 785		1.14 [0.58; 2.25] 1.4% 0.95 [0.79; 1.15] 16.1% 0.75 [0.53; 1.07] 5.5% 1.11 [0.58; 2.14] 1.1% 1.06 [0.79; 1.43] 4.1% 1.07 [0.87; 1.31] 12.3% 0.98 [0.88; 1.10] 40.5%	1.2% 14.6% 4.3% 1.3% 6.2% 12.3% 40.0%
60-Day Mortality Gaudry et al. (2016) Zarbock et al. (2016) Fixed effect model Random effects model Heterogeneity: I ² = 50%,	$\begin{array}{cccc} 150 & 311 \\ 43 & 112 \\ & 423 \\ \tau^2 = 0.0149, p = 0.16 \end{array}$	153 308 60 119 427		0.97 [0.83; 1.14] 18.3% 0.76 [0.57; 1.02] 6.9% 0.91 [0.79; 1.05] 25.3% 0.89 [0.71; 1.12]	18.2% 6.2% 24.3%
90-Day Mortality Wald et al. (2015) Zarbock et al. (2016) Barbar et al. (2018) Fixed effect model Random effects model Heterogeneity: I ² = 66%,	$\begin{array}{cccc} 18 & 48 \\ 44 & 112 \\ 138 & 239 \\ 399 \\ \tau^2 = 0.0413, \ \rho = 0.05 \end{array}$	19 52 65 119 128 238 409		1.03 [0.62; 1.71] 2.2% 0.72 [0.54; 0.95] 7.5% 1.07 [0.91; 1.26] 15.3% 0.96 [0.84; 1.10] 25.0% 0.93 [0.69; 1.23]	2.1% 6.7% 18.3% 27.1%
Fixed effect model Random effects mode Heterogeneity: $l^2 = 6\%$, τ Residual heterogeneity: l	1977 ² = 0.0014, <i>p</i> = 0.39 ² = 22%, <i>p</i> = 0.23	1903	0.5 1 2 Favours Early Favours Late	0.97 [0.91; 1.05] 100.0% 0.97 [0.90; 1.05] — 1	 00.0%

Figure 3. Impact of early versus late RRT on mortality rates at various time periods. ICU: intensive care unit.

Outcome	Number of participants (studies)	Risk ratio	95% Confidence intervals
Recovery of renal function to baseline at 90 days	181 (2 studies)	1.00	0.94–1.06
Bleeding events	1905 (8 studies)	0.80	0.56-1.15
Arrhythmias	1591 (6 studies)	1.11	0.84–1.45
Dialysis-related hypotension	1080 (6 studies)	1.14	0.82-1.57
Hypokalaemia	737 (2 studies)	1.04	0.77-1.40
Thrombocytopenia	725 (2 studies)	1.03	0.89-1.19
Hypocalcaemia	449 (3 studies)	1.12	0.92-1.36
Hypophosphatemia	737 (2 studies)	2.68	0.62-11.58
Hyperkalaemia	1107 (2 studies)	0.27	0.01-5.85

Table 2. Summary of seconda	ry outcomes related to adverse events and	I recovery of renal function at 90 days
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an increase in complications within the 'early' group: RR = 1.85 (95% CI = 1.18, 2.88).

Two studies^{13,28} reported median and interquartile values as two separate classes for early RRT. In Bouman et al.,²⁸ the 'early' group was segmented into high- and low-volume haemofiltration; in Gaudry et al.,¹³ values were given for survivors/non-survivors in both 'early' and 'late' groups; these two studies were excluded from the analysis since no composite values were reported. In the remaining four studies,^{14,21,22,31} medians and interquartile ranges were pooled, showing no statistically significant difference between the 'early' and 'late' groups for either length of ICU stay (estimated difference in length of stay = 0.34 days (95% CI = -1.60,2.28, P = 0.73)), or length of hospital stay (estimated difference in length of stay = -1.75 days (95% CI -5.84,2.34, P = 0.40)).

Three studies reported on the impact of the number of RRT days. One study³⁰ reported in terms of mean, \pm SD and therefore was excluded; the other two reporting in terms of median and interquartile ranges.^{14,22} Although a large estimated difference in medians was evident, they were considered statistically insignificant; estimated difference = -5.99 (95%) CI = -23.52, 11.53, P = 0.50; this was also the case for number of mechanical ventilation-free days (estimated difference in length of stay = 6.94 days (95%CI = -4.59, 18.48, P = 0.24)). No clear difference between groups in terms of the number of RRT-free days (estimated difference in length of stay = -1.33days (95% CI = -3.66, 1.01, P = 0.27)), or vasopressor-free days (estimated difference in length of stay = -0.45 days (95% CI = -3.22, 2.32, P = 0.75)) was observable.

Risk of bias across studies

The risk of bias was estimated through a funnel plot using the overall mortality as an outcome. The inverted standard error against the RR is shown in supplementary Figure s1.19, where the 'dotted' lines signify the expected distribution of the studies. One study²⁹ is a significant outlier; otherwise distributions corroborate a reduced risk of bias across the selected studies.

Discussion

The systematic literature review identified a total of 10 studies that describe the impact of early versus conventional/late-RRT on mortality. While the time period for follow-up varied throughout, the analysis showed no statistically significant difference in terms of overall, in-ICU, in-hospital, 28-, 60- and 90-day mortality. Further, subgroup analyses detected no significant differences between modality of RRT, or general hospital inpatients versus ICU patients only.

On removal of studies with a high or unclear risk of bias, the heterogeneity reduced (I^2 from 46% to 23%), but with no impact on the difference in mortality between groups, suggesting these studies are likely to have influenced the consistency of the overall analysis.

The only subgroup that identified a difference in outcome as a function of RRT initiation was the surgical only population where 'early' RRT resulted in an improvement in mortality. However, it must be noted that the conclusion was based on a single, small study,²⁹ which reported vastly different mortality rates between the 'early' and 'late' groups (14.29% vs. 85.71%). The study was the smallest included in the present meta-analysis (n=28) and owing to its limited extent, the impact of a few additional patients will markedly alter the statistical significance between groups. In addition, the study was also assessed to have an overall unclear risk of bias as well as high risk of reporting incomplete outcome data; therefore, it is likely that the study with a purely surgical population has skewed results significantly. However, it should be noted that while limited conclusions can be drawn, this may indeed represent a difference based on patient population and that further studies may provide better understanding.

The meta-analysis did not identify any association between timing of RRT for AKI and dialysis dependence at 28 or 90 days, but it should be noted that absolute numbers were small. Although results from two studies^{13,14} investigating dependence at day 60 suggested a benefit in the early group, fewer studies reported day 60 compared to days 28 and 90. In both studies, the absolute numbers of dialysis-dependent patients at 60 days were relatively small which potentially skew the conclusions drawn. Further, Zarbock et al.¹⁴ also reported on dialysis dependence at day 90 with no significant difference between the groups. Other reported secondary outcomes such as renal recovery, length of ICU stay, length of hospital stay, number of RRT days, RRT-free days, mechanical ventilation-free days and vasopressor-free days also showed no statistically significant differences between groups.

The pooled results of the majority of adverse events showed no significant difference between groups with the exception of one: higher rates of catheter related complications were seen in the 'early' group which is likely due to the increased number of catheters inserted compared to the 'late' group.

The variability in the classification of 'early' and 'late' contribute to increasing the difficulty in pooling data for direct comparisons. Recent studies for the early group^{13,14,20,21,22,31} have adopted a time frame from eligibility while others utilised physiological variables to determine the initiation of RRT. Timeframes ranged from commencement within 6- to 12-h window from meeting eligibility criteria, whereas physiological criteria ranged from varying urine outputs to serum creatinine or urea levels. In addition to the difference between timing versus physiological factors, studies utilising international guidelines for either inclusion or to determine commencement of early RRT used varying classifications.^{1–3} While a known factor prior to devising the search strategy, it was nevertheless deemed that a systematic comparison of differing strategies would be informative despite the paucity of available data.

The value of initiating RRT earlier has been subjected to extensive debate, and while theoretical benefits have been postulated such as limiting fluid overload and organ dysfunction as well as removal of inflammatory mediators,³² the hypothesis has not been supported through an assessment of measured patient outcomes. Initiation of RRT at an earlier stage will also result in a higher proportion of patients receiving RRT which may, in turn, result in higher rates of complications as well as significant increases to cost.

Previous meta-analyses have reached differing conclusions; two conducted prior to the RCTs from 2013, suggested that 'early' RRT may convey a mortality benefit.^{18,19} In contrast, more recent RCTs concluded that there was no difference in mortality between groups.^{16,17} In 2018, three further RCTs^{20–22} concluded no difference of note in mortality; the large IDEAL-ICU trial²² was stopped early due to futility.

Evidence drawn from the pooling of studies tends to indicate little significant differences exist between early and late initiation of RRT for AKI. In addition, with the exception of 28-day mortality which was found to be of moderate quality, all pooled primary outcomes assessed using the GRADE tool were found to be of low quality (supplementary Table s1.2). The currently ongoing STARRT-AKI trial³³ will add valuable data in an area where there is still a paucity of contextualised data which will, in turn, fuel significant debate.

Conclusions

The systematic review and meta-analysis revealed no significant difference between early and late initiation of RRT for AKI with regard to the primary outcome of overall mortality and multiple secondary outcomes such as length of ICU and hospital stay and dialysis dependence at 90 days. This agrees with recent previous meta-analyses that current evidence does not support the use of early RRT for patients with AKI. Additional data from ongoing and future RCTs are necessary to strengthen the evidence base.

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ORCID iDs

Mark Andonovic D https://orcid.org/0000-0002-5290-4680 Jamie P Traynor D https://orcid.org/0000-0002-4339-0366

Supplemental material

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