

# Defining the Lowest Threshold for Amyloid-PET to Predict Future Cognitive Decline and Amyloid Accumulation

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## Abstract

### Introduction

As clinical trials move toward earlier intervention, we sought to redefine the  $\beta$ -amyloid ( $A\beta$ )-PET threshold based on the lowest point in a baseline distribution that robustly predicts future  $A\beta$  accumulation and cognitive decline in 3 independent samples of clinically normal individuals.

### Methods

Sequential  $A\beta$  cutoffs were tested to identify the lowest cutoff associated with future change in cognition (Preclinical Alzheimer Cognitive Composite [PACC]) and  $A\beta$ -PET in clinically normal participants from the Harvard Aging Brain Study (n = 342), Australian Imaging, Biomarker and Lifestyle study of aging (n = 157), and Alzheimer's Disease Neuroimaging Initiative (n = 356).

### Results

Within samples, cutoffs derived from future  $A\beta$ -PET accumulation and PACC decline converged on the same inflection point, beyond which trajectories diverged from normal. Across samples, optimal cutoffs fell within a short range (Centiloid 15–18.5).

### Discussion

These optimized thresholds can help to inform future research and clinical trials targeting early  $A\beta$ . Threshold convergence raises the possibility of contemporaneous early changes in  $A\beta$  and cognition.

### Classification of Evidence

This study provides Class II evidence that among clinically normal individuals a specific  $A\beta$ -PET threshold is predictive of cognitive decline.

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## Glossary

**A $\beta$**  =  $\beta$ -amyloid; **AD** = Alzheimer disease; **ADNI** = Alzheimer's Disease Neuroimaging Initiative; **AIBL** = Australian Imaging, Biomarker and Lifestyle; **AIC** = Akaike information criterion; **BIC** = Bayesian information criterion; **CI** = confidence interval; **CL** = Centiloid; **CN** = clinically normal; **DVR** = distribution volume ratio; **FBP** =  $^{18}\text{F}$ -florbetapir; **FLR** = frontal, lateral temporal, and retrosplenial; **GMM** = gaussian mixture model; **HABS** = Harvard Aging Brain Study; **MCI** = mild cognitive impairment; **MMSE** = Mini-Mental State Examination; **PACC** = Preclinical Alzheimer Cognitive Composite; **PiB** = Pittsburgh compound B; **ROI** = region of interest; **SUVR** = standardized uptake value ratio.

Due to the largely disappointing results of anti-amyloid therapies in patients with mild cognitive impairment (MCI) and Alzheimer disease (AD), many clinical trials are pivoting to intervention earlier in the AD pathologic trajectory to slow or prevent disease progression.<sup>1,2</sup> Markers of  $\beta$ -amyloid (A $\beta$ ) can be used to target individuals earlier in the disease process<sup>3,4</sup> through the use of PET radioligands for A $\beta$  such as  $^{11}\text{C}$ -Pittsburgh compound B (PiB),<sup>5</sup>  $^{18}\text{F}$ -florbetapir (FBP),<sup>6</sup> and  $^{18}\text{F}$ -NAV4694.<sup>7</sup>

Elevated, or abnormal, levels of A $\beta$ -PET are most often defined on the basis of a threshold derived from cross-sectional A $\beta$ -PET<sup>8-13</sup> or from combined PET-neuropathology studies.<sup>14</sup> Some evidence indicates that current thresholds might be too high,<sup>10,15</sup> leading to the misclassification of some subthreshold individuals as A $\beta$ -. Recent neuropathologic studies in participants with PET before death suggest that far lower PET thresholds can be achieved.<sup>14</sup> However, without confirmation provided by autopsy on an individual basis, low PET signal is difficult to differentiate from noise, raising the possibility of introducing greater false positives.

Thus, our aim was to determine how low thresholds can go in the A $\beta$ -PET signal while remaining useful as a predictor of whether individuals are on an AD trajectory, signaled by continued A $\beta$  accumulation or cognitive decline. To evaluate reliability, we examined longitudinal data from 3 independent samples. We hypothesize that longitudinal cognitive and PET outcomes will provide justification for shifting to a threshold lower than that possible with cross-sectional bimodality-based approaches, thus potentially increasing potential benefit from anti-A $\beta$  therapy and increasing eligibility for trials.

## Methods

### Subjects

Clinically normal (CN) participants (Clinical Dementia Rating score = 0; Mini-Mental State Examination [MMSE] score  $\geq 24$ ) were included from 3 independent samples: the Harvard Aging Brain Study (HABS), the Australian Imaging, Biomarker and Lifestyle (AIBL) study of aging, and the Alzheimer's Disease Neuroimaging Initiative (ADNI). All studies included participants with a baseline Clinical Dementia Rating score of 0 and excluded participants with clinical depression, neurologic disorders, and head trauma.

### HABS Sample

HABS is focused on differentiating normal aging from AD and consequently recruited only individuals initially considered CN. Detailed inclusion criteria have been published previously.<sup>16</sup> For the primary analysis of baseline PiB predicting cognitive change, all HABS participants (n = 342) were included. Cognitive assessments were conducted annually, with a median follow-up length of 4.21 (SD = 2.34) years. Analyses relating to PiB accumulation focused on the subset with at least 2 PiB-PET scans (n = 214), 141 with 3 scans, and 39 with 4 scans. The median follow-up time was 2.82 (2.15) years.

### AIBL Sample

AIBL recruited individuals across the AD continuum and enriched its CN cohort for family history and memory complaints to ensure a high proportion of A $\beta$ + adults.<sup>17</sup> Detailed inclusion criteria for AIBL have been published previously.<sup>17</sup> One hundred fifty-seven CN participants were included from AIBL. Participants underwent cognitive testing approximately every 18 months, with a median cognitive follow-up of 7.48 (1.97) years. All participants had at least 1 cognitive follow-up visit; 123 had at least 2 PiB-PET scans, 87 had at least 3, and 54 had 4 PiB-PET scans, with a median PET follow-up of 3.06 (1.83) years.

### ADNI Sample

ADNI is a multisite study conducted across 63 sites in the United States and Canada, enrolling individuals across the AD continuum. In contrast to HABS and AIBL, the ADNI CN cohort was required not to have memory complaints at baseline but also was allowed a more lenient MMSE score of  $\geq 24$  (HABS and AIBL used an MMSE score  $\geq 25$ ). Detailed inclusion criteria for ADNI have been published previously.<sup>18</sup> A $\beta$ -PET scanning with FBP was introduced to ADNI at a later stage, and only longitudinal cognitive data collected prospectively from the first FBP-PET scan were included. Participants were restricted to those with a CN diagnosis at first PET scan (n = 356). All participants had at least 1 cognitive follow-up visit (median cognitive follow-up 2.97 [2.33] years); 284 had at least 2 scans, 124 had 3 scans, and 7 had 4 scans, with a median PET follow-up of 2.05 (1.60) years.

### Standard Protocol Approvals, Registrations, and Patient Consents

We conducted the procedures for this study under the Partners Human Research Committee, the Institutional Review Board for the Massachusetts General Hospital and Brigham

and Women's Hospital. All participants provided written consent.

## Cognition

Cognition was assessed with the Preclinical Alzheimer Cognitive Composite (PACC) 5 version<sup>19,20</sup> of the PACC,<sup>21</sup> which was honed to detect A $\beta$ -related cognitive decline. As previously reported,<sup>20,22</sup> the PACC5 was computed within each sample as the averaged *z* scores of 5 tests, with *z* scores standardized using the baseline mean and SD within each cohort. Of these 5 tests, 2 were overlapping across all samples: the MMSE and Logical Memory Delayed Recall. As a measure of executive function, Digit Symbol Substitution was used in HABS and AIBL, while ADNI used Trail Making Test Part B. As a measure of word recall, HABS used the Free and Cued Selective Reminding Test, AIBL used California Verbal Learning Test (second edition), and ADNI used Alzheimer's Disease Assessment Scale–Cognitive Battery Word Recall. For semantic memory, each study used versions of Categories, with HABS using Animals/Vegetables/Fruit, AIBL using Animals/Names, and ADNI using Animals.

All studies show similar distributions of baseline and slopes for these study-specific composites.<sup>20,22</sup>

## PET

### HABS PET

PiB-PET acquisition parameters for HABS have been published previously.<sup>16,23</sup> In brief, PET images were acquired on a Siemens (Munich, Germany) ECAT EXACT HR+ scanner with a 60-minute dynamic acquisition starting directly after injection. Distribution volume ratios (DVRs) were calculated via Logan plotting with a cerebellar gray reference tissue. At each time point, PET data were rigidly coregistered to the individual's closest magnetization-prepared rapid gradient echo with SPM12 (Wellcome Department of Cognitive Neurology, Function Imaging Laboratory, London). Cortical regions of interest (ROIs) were defined from the Desikan-Killiany atlas<sup>24</sup> via FreeSurfer version 6.0. Frontal, lateral temporal, and retrosplenial (FLR) regions were averaged into a global aggregate, as previously reported.<sup>8,20,23</sup> PET data were not partial volume corrected.

### AIBL PET

Detailed acquisition parameters have been published previously.<sup>9</sup> PiB-PET images were processed in-house with a modified version of the HABS pipeline to account for the lack of full dynamic data. Fifty- to 70-minute frames were summed and coregistered to their MRI at each time point. Subject-specific cortical parcellations from FreeSurfer were unavailable, so instead each individual's PET was normalized to Montreal Neurological Institute space, and the same HABS FLR ROI was selected from a customized probabilistic template space atlas<sup>25</sup> based on the Desikan-Killiany atlas. The standardized uptake value ratio (SUVR) was computed with FLR referenced to cerebellar gray.

### ADNI PET

Detailed acquisition parameters were published previously.<sup>26</sup> FBP-PET images were processed with the same pipeline used to process AIBL data, but the FBP SUVR was normalized to whole cerebellum.

### Centiloid Scale

Due to our interests in identifying thresholds within the uncertain range between clearly A $\beta$ - and A $\beta$ + positive, we prioritized using in-house methods that would maximize measurement accuracy within each sample. To compare between samples and to provide generalizable estimates, we then converted each sample to the Centiloid (CL) scale.<sup>27</sup> Notably, the CL transformation is expected to reduce but not completely eliminate between-sample differences in acquisition and processing.<sup>28</sup> CL values should therefore be considered approximate.

As recommended<sup>27</sup> and published previously,<sup>20,25</sup> we ran the GAAIN dataset using our in-house pipeline and ROIs to compute the linear transformation between the in-house and CL scales and then applied the CL transformation. An additional linear transformation was required to transform the ADNI data to translate from FBP SUVR to PiB SUVR using the GAAIN dataset, similar to previous recommendations.<sup>29</sup> Two additional steps were used to transform the HABS data to CL due to differences in acquisition time from the CL method and the expression of PiB as a DVR.<sup>25</sup> The resulting equations were as follows: HABS: CL = 143.06  $\times$  HABS DVR – 145.60; AIBL: CL = 96.94  $\times$  AIBL SUVR – 105.20; ADNI: CL = 180.20  $\times$  ADNI FBP SUVR – 179.70.

## Statistical Analysis

All analyses were conducted in R version 3.6.0. For sample descriptives and comparison with optimal thresholds, thresholds were generated within each sample at baseline by fitting the data to a gaussian mixture model (GMM).<sup>8,20,23,30</sup> In all samples, a bimodal 2-class model fit best, exhibiting the highest *Bayesian information criterion* (BIC) and significantly higher likelihood than 1-class or 3-class models in bootstrapped (*n* = 1,000) likelihood ratio tests. Mixed-effects models were used to account for missing data and loss to follow-up.

The first objective was to identify a cognitively derived PET cutoff separately in each sample, at the lowest point in baseline A $\beta$ -PET distribution that predicted subsequent cognitive decline. This has been classified Class II evidence that among CN individuals a specific A $\beta$ -PET threshold is predictive of cognitive decline. To do so, we conducted a series of 2-level linear mixture models to predict PACC5 performance over time iterating through a range of possible cutoffs (*c*) of A $\beta$ -PET (i.e., 1.1–1.8 by 0.01), adjusting for sex, age at baseline, *APOE*  $\epsilon$ 4 status, and years of education. The first level of the model was as follows:

$$\begin{aligned} \text{PACC5} \sim & \text{time} \times \text{Random slope model} \\ & + \text{time} \times \text{covariates} \\ & + \text{random intercept} \end{aligned}$$

The second level was as follows:

$$\begin{aligned} \text{Random slope model} = & \text{BL-PET} \times I(\text{BL PET} > c) \\ & + \text{random intercept} \end{aligned}$$

The nested random slope term accounts for change in PACC5 over time as a function of an threshold indicator  $I$  (baseline PET  $> c$ ) that assigns a 0 or 1 depending on whether individuals fell above or below  $c$  and the continuous effect of tracer retention (baseline PET) above the cutoff. The rationale for this nested slope model was based on the assumption that below the optimal minimal cutoff the PACC5 should not relate to A $\beta$ -PET tracer retention, whereas above the cutoff the magnitude of PACC5 decline is expected to track with the magnitude of A $\beta$ -PET retention.<sup>12</sup> The random intercept embedded within the nested random slope model represents an average measure of the baseline PET for all individuals in the sample regardless of the threshold value  $c$ , with random effects added to allow variations between individuals. The Akaike information criterion (AIC)/BIC was used to select the optimal cognitively derived PIB cutoff by comparing model fits under different baseline thresholds.<sup>31,32</sup> Both are reported but identified the same cutoffs in all models. Sensitivity analysis was carried out with model expansion in both the PACC model and the PET model below to ensure that the optimal models were selected.

We also aimed to identify an accumulation-derived cutoff based on subsequent A $\beta$ -PET accumulation. Here, we iterated through a series of linear mixed-effects model of the effect of time, A $\beta$ -PET cutoff, the same set of covariates used above, and the random intercept on A $\beta$ -PET with a range of possible cutoffs  $c$ :

$$\text{PET} \sim \text{time} \times (\text{BL PET} > c) + \text{time} \times \text{covariates} + \text{random intercept}$$

The rationale for this model was based on previous evidence<sup>33–35</sup> that at low A $\beta$  levels roughly equivalent proportions of negative and positive PET changes are observed and are presumed to largely reflect fluctuations in nonspecific binding of the tracer, but as baseline PET increases, the observed changes become predominately positive and reflect contributions from both nonspecific fluctuations and A $\beta$  accumulation. We sought to identify the threshold beyond which changes are predominately positive and thus maximize the difference between A $\beta$ +/- groups. To do so, we iteratively compared models with different baseline A $\beta$ -PET thresholds and selected the model that maximized the group differences in change<sup>36</sup> as reflected by the greatest effect size (standardized  $\beta$ ) for the A $\beta$  group  $\times$  time term. This selected criterion differs from that used for the PACC model. Goodness of fit was still tested with AIC/BIC, but due to the greater dynamic range in DVR/SUVR compared with change over time, overall model fit is dominated by the between-participant differences in DVR/SUVR (the main effect

of A $\beta$  group) rather than within-participant change (time  $\times$  A $\beta$  group). To identify a cutoff based on optimal change over time, we selected the cutoff associated with the maximal effect size for the A $\beta$  group  $\times$  time term.

To further describe the additional individuals who would be identified as A $\beta$ + with the optimal cutoffs, participants were assigned into low (below optimal cutoff), intermediate (between low and GMM cutoffs), and high (above GMM cutoff) groups within each sample and then pooled into 1 large sample ( $n = 855$ ). These groups were then compared on age, sex, APOE  $\epsilon 4$  status, A $\beta$  slope (CL/year), and PACC5 score slope (SD/year) with independent-sample  $t$  tests and  $\chi^2$  tests.

Finally, the optimal threshold was used to conduct power analyses to project the sample size needed to detect A $\beta$ -PET accumulation and PACC score decline over a 5-year clinical trial. In the pooled sample, 5 possible CL range groups were selected with the lower bound for inclusion set at the optimal CL threshold and the CL upper bound varying from 35, 50, 70, and 100 to the maximum observed in our samples (140). Within each group, we computed the sample size needed to detect different possible percent changes (10%–100%) in A $\beta$ -PET accumulation and PACC score decline over 5 years with 80% power. Power to detect A $\beta$ -PET accumulation was based on estimates from a linear mixed-effect model of CL  $\sim$  time + random intercept. For PACC score decline, estimates were based on a linear mixed effect of PACC score  $\sim$  time + random intercept + random slope.

## Data Availability

All data for ADNI and PET data for AIBL are publicly available online at [ida.loni.usc.edu](http://ida.loni.usc.edu). AIBL cognitive data and HABS data are available on request.

## Results

### Sample Differences

Table 1 displays descriptive statistics for the HABS, AIBL, and ADNI samples, dichotomized within sample using GMM into A $\beta$ +/- groups. Between samples, AIBL had a numerically higher proportion of APOE  $\epsilon 4$  carriers due to their enriched recruitment strategy,<sup>17</sup> although it did not reach statistical significance compared with HABS ( $M_{\text{Difference}} = 0.07$ ,  $p = 0.15$ , 95% confidence interval [CI]  $-0.03$  to  $0.17$ ) or ADNI ( $M_D = 0.08$ ,  $p = 0.08$ , 95% CI  $0.01$ – $0.18$ ). ADNI participants were older ( $M_D = 2.91$ ,  $p < 0.001$ , 95% CI  $1.83$ – $3.98$ ) and more highly educated than HABS participants ( $M_D = 0.70$ ,  $p < 0.001$ , 95% CI  $0.28$ – $1.11$ ) and older than AIBL participants ( $M_D = 2.17$ ,  $p < 0.001$ , 95% CI  $0.92$ – $3.43$ ), although HABS and AIBL participants did not differ by age. AIBL participants had a longer cognitive follow-up than HABS ( $M_D = 2.99$ ,  $p < 0.001$ , 95% CI  $2.57$ – $3.41$ ) and ADNI ( $M_D = 3.31$ ,  $p < 0.001$ , 95% CI  $2.89$ – $3.73$ ) participants.



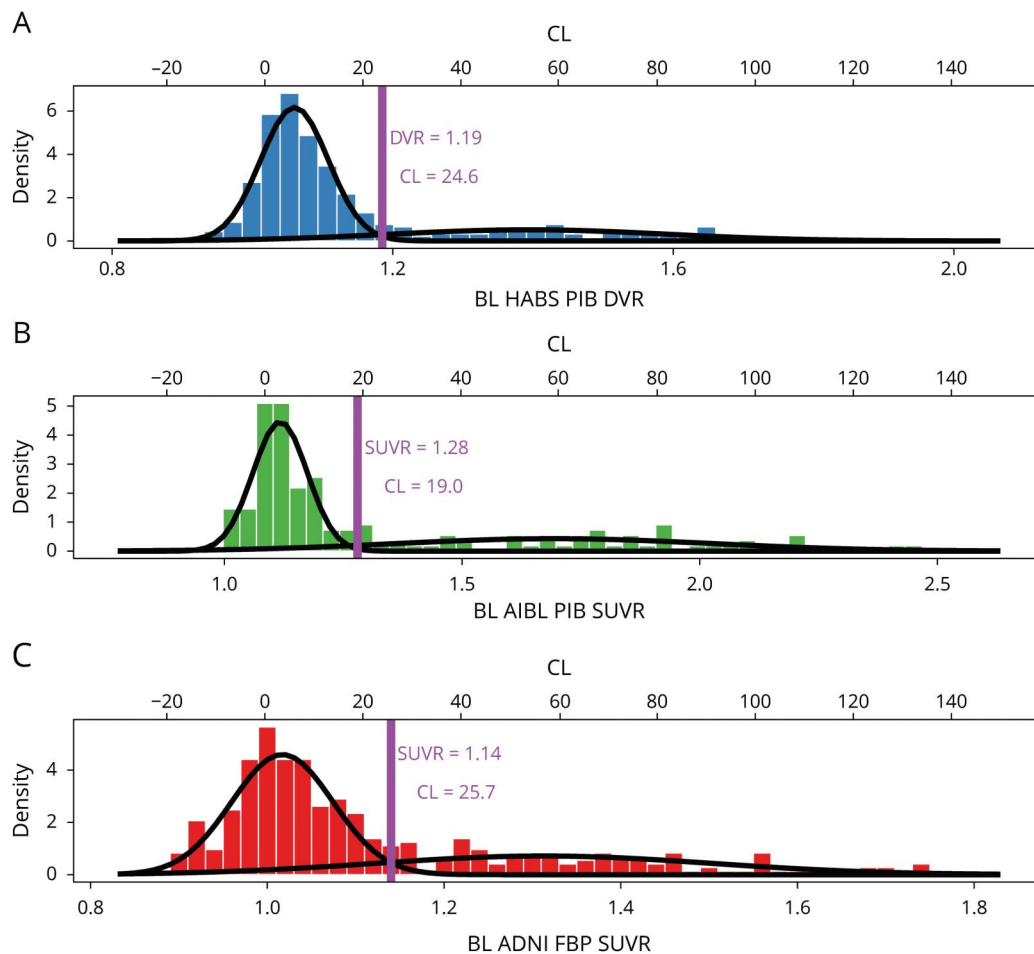
**Table 1** Sample Descriptives for HABS, AIBL, and ADNI

	HABS			AIBL			ADNI			Sample Difference
	All	A $\beta$ -	A $\beta$ +	All	A $\beta$ -	A $\beta$ +	All	A $\beta$ -	A $\beta$ +	F or $\chi^2$ ( <i>p</i> )
<b>No.</b>	342	265	77	157	106	51	356	252	104	6.74 (0.03)
<b>Age, y</b>	71.7 (8.00)	70.8 (8.05)	74.8 (6.48)	72.5 (6.72)	70.8 (6.34)	75.8 (6.28)	74.6 (6.50)	73.9 (6.60)	76.5 (5.88)	14.6 (<0.001)
<b>Education, y</b>	15.78 (3.01)	15.8 (2.97)	16.1 (2.94)	14.54 (3.1)	14.2 (3.2)	14.7 (3.0)	16.54 (2.59)	16.7 (2.56)	16.2 (2.61)	NA
<b>Sex, n female (%)</b>	206 (60)	160 (60)	46 (60)	86 (55)	59 (56)	27 (53)	189 (56)	125 (50)	64 (62)	3.79 (0.15)
<b>APOE <math>\epsilon</math>4, n (%)</b>	89/324 (27)	44/251 (18)	45/73 (62)	54 (34)	27 (25)	27 (53)	94 (26)	50 (20)	44 (42)	3.61 (0.16)
<b>Baseline PET</b>	1.14 DVR (0.18)	1.06 DVR (0.05)	1.43 DVR (0.17)	1.32 SUVR (0.35)	1.12 SUVR (0.06)	1.75 SUVR (0.30)	1.12 SUVR (0.18)	1.02 SUVR (0.06)	1.36 SUVR (0.161)	NA
<b>Baseline CL</b>	18.15 (25.90)	6.10 (6.94)	59.6 (24.4)	23.19 (33.51)	3.35 (5.81)	64.4 (29.4)	21.65 (33.28)	3.64 (10.3)	65.3 (29.1)	1.85 (0.16)
<b>Baseline PACC5 score</b>	0.06 (0.68)	0.07 (0.68)	0.01 (0.68)	0.08 (0.62)	0.12 (0.63)	0.01 (0.60)	-0.05 (0.69)	-0.00 (0.71)	-0.15 (0.62)	3.14 (0.04)
<b>Median PACC5 score follow-up, y</b>	4.21 (2.34)	4.15 (2.40)	4.43 (2.07)	7.48 (1.97)	7.49 (1.65)	6.33 (2.38)	2.97 (2.33)	2.85 (2.35)	4.02 1.98	215 (<0.001)
<b>Median PiB follow-up, y</b>	2.81 (2.14)	2.80 (2.16)	2.89 (2.09)	3.06 (1.83)	3.16 (1.75)	1.79 (1.94)	2.05 (1.60)	2.07 (1.60)	2.02 (1.59)	21.8 (<0.001)
<b>Progression to MCI/dementia, n (%)</b>	21 (6)	6 (2)	15 (21)	22 (14)	12 (11)	10 (20)	59 (18)	29 (12.4)	30 (30.6)	18.9 (<0.001)

Abbreviations: A $\beta$  =  $\beta$ -amyloid; ADNI = Alzheimer's Disease Neuroimaging Initiative; AIBL = Australian Imaging, Biomarker and Lifestyle; CL = Centiloid; DVR = distribution volume ratio; HABS = Harvard Aging Brain Study; MCI = mild cognitive impairment; NA = not applicable; PACC5 = Preclinical Alzheimer Cognitive Composite 5 version; PiB = Pittsburgh compound B; SUVR = standardized uptake value ratio.

Descriptives are shown for both each full sample and dichotomized into A $\beta$ +/- groups using gaussian mixture model (figure 1). Means (SDs) are displayed for continuous variables and numbers (percents) for categorical variables. To demonstrate which variables varied across sample, 1-way analysis of variance *F* statistics are reported in the final column for continuous variables and  $\chi$  statistics for categorical variables, as well as *p* values. Some participants in HABS did not have *APOE* data available, so total with genetic data are also displayed. Education was measured differently in AIBL and is not directly comparable to education in HABS or ADNI. Baseline PET measures are provided for within-sample description and baseline CL for between-sample comparison.

**Figure 1** Bimodal Gaussian Distribution of PET Signal in HABS, AIBL, and ADNI



A–C) Histograms of the baseline (BL) PET data in each sample are plotted and fitted to a bimodal gaussian distribution using gaussian mixture model (GMM). Data are plotted on the tracer for each sample and processing-specific scale on the bottom x-axis and transformed to the Centiloid (CL) scale on the top x-axis. The GMM threshold (purple line) was identified within each sample as the point at which an individual had an equal probability of being in the lower  $\beta$ -amyloid ( $A\beta$ )– and higher  $A\beta$ + gaussian. ADNI = Alzheimer’s Disease Neuroimaging Initiative; AIBL = Australian Imaging, Biomarker and Lifestyle; DVR = distribution volume ratio; HABS = Harvard Aging Brain Study; PiB = Pittsburgh compound B; SUVR = standardized uptake value ratio.

Overall, each sample exhibited typical bimodal  $A\beta$ -PET distributions (figure 1) that mapped onto a similar range on the CL scale, but some small differences are noted. The AIBL sample had a greater SUVR range than HABS and ADNI participants. Notably, the GMM in AIBL also provided a lower CL threshold for positivity ( $CL_{GMM} = 19.0$ ) than in HABS ( $CL_{GMM} = 24.6$ ) or ADNI ( $CL_{GMM} = 25.7$ ). In ADNI, a larger variance was observed in the  $A\beta$ – group, but this variance translated into more negative CL values rather than a higher GMM threshold. The proportion of  $A\beta$ + participants based on the GMM threshold varied by sample, with a lower proportion in HABS (22.5%) than in AIBL (32.5%,  $\chi^2 = 5.61$ ,  $p = 0.02$ ) and ADNI (29.0%  $\chi^2 = 4.08$ ,  $p = 0.04$ ). The varied mixing proportions may contribute to slight differences in how the distributions transform to the CL scale.

### Cognitively Derived Cutoff

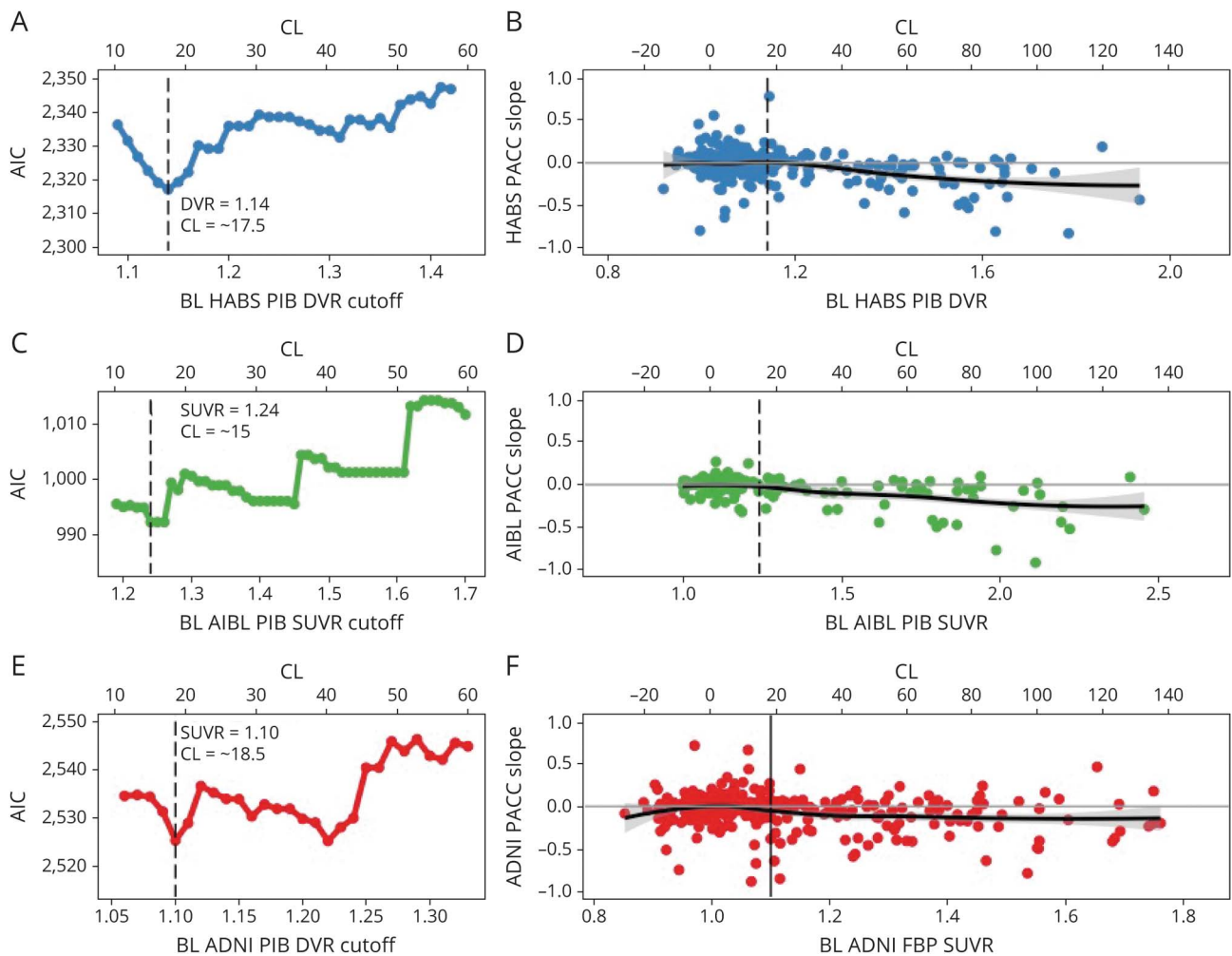
We detected similar optimal cutoffs associated with cognitive decline across the 3 samples: HABS at a PiB DVR of 1.14 (CL = 17.5), AIBL at a PiB SUVR of 1.24 (CL = 15.0), and ADNI

at an FBP SUVR of 1.1 (CL = 18.5) (figure 2). The optimal cutoff was selected from the model with absolute lowest AIC/BIC in HABS, and the lower of 2 thresholds associated with similar AIC/BIC in AIBL and ADNI (table 2). Across samples, the optimal cutoff fell 0.04 to 0.05 SUVR/DVR below their respective GMM cutoff. As shown in figure 1, the cutoff corresponds to an approximate inflection point in the baseline  $A\beta$ -PET distribution beyond which cognitive slopes begin to negatively deviate from zero.

### Accumulation-Derived Cutoff

Analyses of  $A\beta$ -PET accumulation–derived optimal thresholds converged on the same threshold as with those derived from cognitive outcomes in HABS (PiB DVR 1.14/CL = 17.5), and AIBL (PiB SUVR 1.24/CL = 15.0) and nearly converged in ADNI (FBP SUVR 1.09/CL = 16.7). Across samples, the standardized  $\beta$  criterion provided clear cutoffs for maximizing group differences in accumulation (figure 3). As shown in figure 3, these cutoffs corresponded to an inflection point in baseline  $A\beta$ -PET distribution beyond which PET

**Figure 2** Optimal Cutoffs Within HABS, AIBL, and ADNI Based on Longitudinal Cognitive Outcomes



For each sample, the *Akaike information criterion* (AIC) demonstrating the model fit for a range of possible cutoffs are shown (A, C, E). In Harvard Aging Brain Study (HABS) (A), an optimal cutoff of Pittsburgh compound B (PiB) distribution volume ratio (DVR) 1.14/Centiloid (CL) 17.5 is derived from a clear lowest AIC (least information loss). In Australian Imaging, Biomarker and Lifestyle (AIBL), similarly low AIC was achieved twice, and the lower cutoff at a PiB standardized uptake value ratio (SUVR) 1.24/CL 15.0 was selected. Likewise, in Alzheimer's Disease Neuroimaging Initiative (ADNI) an optimal cutoff of  $^{18}\text{F}$ -florbetapir (FBP) SUVR 1.1/CL 18.5 was selected as the lower cutoff from 2 similarly well-fitting models with nearly equivalent AIC. Notably, across samples, the optimal cutoffs based on cognitive decline fit in a tight CL range of 15.0 to 18.5. (B, D, F) Preclinical Alzheimer Cognitive Composite 5 version (PACC5) score slope over time is plotted as a function of baseline (BL) PET within each sample using a loess curve to demonstrate the shift in trajectories of change with increasing baseline  $\beta$ -amyloid ( $\text{A}\beta$ )-PET tracer retention. PACC5 score slope for each participant was extracted from the slope of the linear regression of PACC5 score over time. Data are unadjusted for covariates. In each sample, the magnitude of tracer retention is unrelated to PACC score slope until an inflection point is reached, beyond which greater levels of  $\text{A}\beta$ -tracer retention are associated with increasing rates of cognitive decline on the PACC5. The inflection point corresponds to the optimal cutoff derived from the iterative models of PACC5 score over time.

slopes shift from a roughly equal mix of negative and positive PET slopes to predominately positive slopes. In contrast, the GMM cutoff had lower standardized  $\beta$  values (table 2) and failed to identify several individuals with low baseline but high accumulation as  $\text{A}\beta^+$ .

### Participants Between the Optimal and GMM Cutoffs

A total of 54 individuals were identified with intermediate  $\text{A}\beta$  between the optimal and GMM cutoffs across samples (table 3). This group was also intermediate between the low  $\text{A}\beta$  and high  $\text{A}\beta$  group in terms of age ( $M_{D\_low\_int} = 2.36$ ,  $p = 0.01$ , 95% CI 0.52–4.20;  $M_{D\_int\_high} = -1.58$ ,  $p = 0.10$ , 95% CI -3.49 to 0.33), number of *APOE*  $\epsilon 4$  carriers ( $M_{D\_low\_int} = 0.113$ ,  $p = 0.07$ , 95%

CI -0.03 to 0.25;  $M_{D\_int\_high} = -0.21$ ,  $p = 0.01$ , 95% CI -0.36 to -0.06), and PACC5 score slope ( $M_{D\_low\_int} = -0.07$ ,  $p = 0.05$ , 95% CI -0.001 to 0.12;  $M_{D\_int\_high} = -0.17$ ,  $p = 0.01$ , 95% CI -0.31 to -0.03). Notably, the intermediate  $\text{A}\beta$  group had CL slopes comparable to those of the high  $\text{A}\beta$  group and higher CL slopes than the low  $\text{A}\beta$  group ( $M_{D\_low\_int} = 2.26$ ,  $p < 0.001$ , 95% CI 1.09–3.43;  $M_{D\_int\_high} = -0.64$ ,  $p = 0.40$ , 95% CI -0.86 to 2.14).

### Power to Detect Reduced $\text{A}\beta$ -PET Accumulation and PACC Score Decline

Power analyses (figure 4) demonstrated that a CL in the 18 to 35 range would require the smallest sample size to detect reduced  $\text{A}\beta$ -PET accumulation ( $n = 64$  for 20% reduction over 5 years).

**Table 2** Model Fit and Parameter Estimates at Optimal and GMM cutoffs in HABS, AIBL, and ADNI

On PACC5								
	Model at Optimal Cutoff				Model at GMM Cutoff			
	Full Model		Time × Aβ Cutoff × BL DVR/SUVR		Full Model		Time × Aβ Cutoff × BL DVR/SUVR	
	AIC	BIC	β	SE	AIC	BIC	β	SE
HABS	2,317	2,422	-1.14	0.31	2,291	2,434	-1.00	0.38
AIBL	991	1,079	-0.18	0.15	995	1,084	-0.08	0.10
ADNI	2,519	2,614	-0.47	0.30	2,527	2,621	-0.52	0.34

On Aβ-PET Accumulation								
	Model at Optimal Cutoff				Model at GMM Cutoff			
	Full Model		Time × Aβ Cutoff <sup>a</sup>		Full Model		Time × Aβ Cutoff	
	AIC	BIC	β Value	SE	AIC	BIC	β Value	SE
HABS	-1,490	-1,398	0.12	0.01	-1,580	-1,489	0.10	0.02
AIBL	-593	-508	0.14	0.02	-615	-530	0.13	0.02
ADNI	-1,551	-1,457	0.08	0.01	-1,622	-1,529	0.06	0.01

Abbreviations: Aβ = β-amyloid; ADNI = Alzheimer's Disease Neuroimaging Initiative; AIBL = Australian Imaging, Biomarker and Lifestyle; AIC = Akaike information criterion; BIC = Bayesian information criterion; BL = baseline; DVR = distribution volume ratio; GMM = gaussian mixture model; HABS = Harvard Aging Brain Study; PACC5 = Preclinical Alzheimer Cognitive Composite 5 version; SUVR = standardized uptake value ratio. AIC and BIC are reported describing model fit for the models predicting PACC5 decline (top) and Aβ Accumulation (bottom) at the optimal cutoff and GMM cutoff within each sample. In addition, standardized β values, standard error, and p values are shown as a measure of effect size on the term of interest (time × Aβ cutoff × BL DVR/SUVR in the PACC5 model; time × Aβ cutoff in the Aβ accumulation model).  
<sup>a</sup> Model thresholds were selected by iterating through multiple possible models, reducing the interpretability of the nominal p values associated with individual models.<sup>36</sup> As a result, p values are not reported.

In contrast, trials with a primary cognitive outcome would exhibit optimal power using a wider CL range of 18 to 70. Targeting individuals in the 18 to 50 CL range for an early intervention trial would provide a balance between minimizing sample size needed to detect PACC score decline while maintaining high power to detect reduced Aβ-PET accumulation (n = 886).

## Discussion

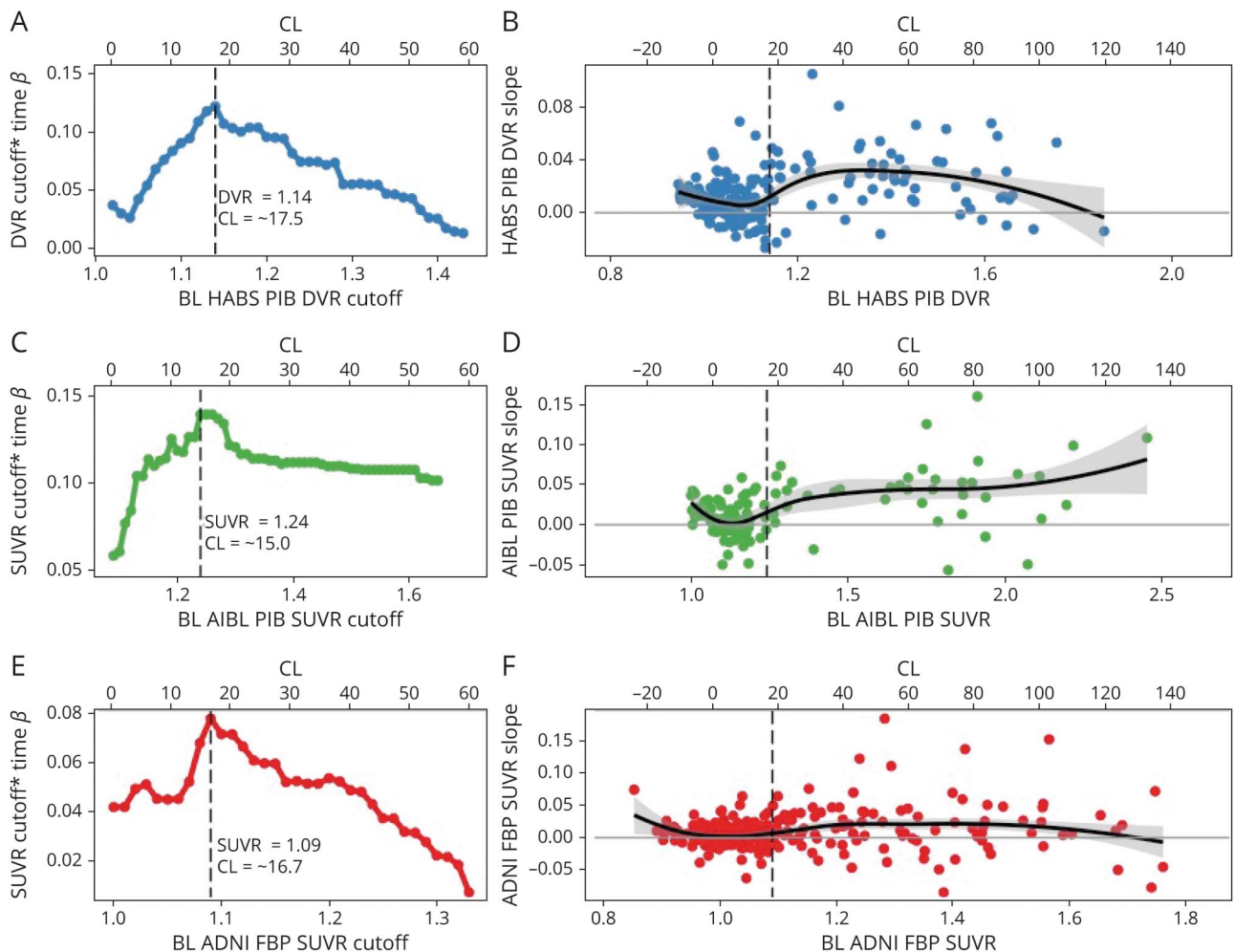
As the field moves toward intervening earlier in the course of the disease, it is important to ascertain the lowest point in the AD-related biomarker cascade that can robustly identify individuals who are embarking on the AD pathologic trajectory and effectively exclude those unlikely to benefit from intervention. Here, we focused on Aβ-PET markers because they are currently the primary method for recruiting at-risk CN individuals in many early intervention trials.<sup>1,37</sup> Independently deriving baseline Aβ cutoffs for PET on the basis of the prediction of both future cognitive decline and Aβ accumulation converged on the same or very nearly the same point in the distribution across 3 independent samples. Furthermore, the 3 samples indicated an optimal threshold in the short CL range of 15.0 to 18.5. In each sample, the optimal cutoff fell at an inflection point beyond which Aβ-PET and cognitive trajectories diverged from normal. This point was consistently below the GMM-derived cutoff that relies primarily on the bimodality of the Aβ distribution but was higher than cutoffs previously reported and based on autopsy

data.<sup>14</sup> Individuals in the intermediate range between the optimal and GMM cutoffs were older than those in the lowest PET tracer uptake range (below the optimal cutoff) and exhibited Aβ accumulation comparable to that displayed in those with the highest PET tracer uptake (above the GMM cutoff). The lower optimal cutoff allows excellent power to detect reductions in Aβ accumulation and sufficient power to detect slowing of cognitive decline in preclinical individuals with low to moderate Aβ burden. Consequently, this threshold may prove useful for research and clinical trials targeting individuals in the early stages of amyloidosis.

Within each sample, the optimal cutoff derived from cognitive outcomes converged on the identical or nearly identical cutoff derived from Aβ accumulation. This may suggest a more contemporaneous onset of change in Aβ and cognition than would be expected from the amyloid cascade hypothesis,<sup>38,39</sup> which posits a cognitive lag behind Aβ accumulation. It is important to note, however, that these optimal thresholds remain thresholds of detection and not necessarily initial onset of Aβ. Undoubtedly, toxic Aβ species exist, both soluble and insoluble isoforms, below these lower PET-detection thresholds, because PET tracers primarily have been found to bind to aggregated fibrillar Aβ.<sup>40</sup> The converging cutoffs based on Aβ-PET accumulation and cognitive decline may instead reflect the lowest point in the PET distribution where signal exceeds noise and is therefore likely to predict future outcomes. While additional studies are needed to probe the



**Figure 3** Optimal Cutoffs Within HABS, AIBL, and ADNI Based on Longitudinal A $\beta$ -PET Accumulation



For each sample, the standardized  $\beta$  for the cutoff  $\times$  time interaction term from iterative models across a range of possible cutoffs is shown (A, C, E). The optimal cutoff based on A $\beta$ -PET accumulation was set at the cutoff that gave the highest standardized  $\beta$  in each sample. In Harvard Aging Brain Study (HABS) and Australian Imaging, Biomarker and Lifestyle (AIBL), the  $\beta$ -amyloid (A $\beta$ ) accumulation-derived optimal cutoff is identical to the cognitively derived optimal cutoff: (A) HABS: Pittsburgh compound B (PiB) distribution volume ratio (DVR) 1.14/Centiloid (CL) 17.5; (C) AIBL: PiB standardized uptake value ratio (SUVR) 1.24/CL 15.0. In Alzheimer's Disease Neuroimaging Initiative (ADNI), the A $\beta$  accumulation-derived optimal cutoff ( $^{18}$ F-florbetapir [FBP] SUVR 1.09, CL 16.7) was very slightly lower than the cognitively derived cutoff (FBP SUVR 1.1, CL 18.5). (B, D, F) DVR/SUVR slope over time is plotted as a function of baseline DVR/SUVR within each sample using a loess curve to demonstrate the shift in trajectories of change as function of baseline A $\beta$ -PiB tracer retention. DVR/SUVR slope for each participant was extracted from the slope of the linear regression of DVR/SUVR over time. Data are unadjusted for covariates. In each sample, the slopes below the optimal cutoff consist of a roughly equal proportion of both negative and positive slopes, presumed to reflect random fluctuations in signal noise. Increasing baseline DVR/SUVR is associated with a small negative trend in this range suggestive of regression to the mean. The optimal cutoff appears to mark a shift toward more positive slopes presumed to reflect A $\beta$  accumulation.

possibility of contemporaneous changes in A $\beta$  and cognition, the present findings add to mounting evidence that even very low levels of A $\beta$ -PET signal are associated with subtle cognitive changes<sup>35,41,42</sup> in advance of more severe decline observed at higher levels of A $\beta$  pathology. Furthermore, while the observed PACC changes are subtle, it has recently been demonstrated in the same 3 cohorts that subtle decline over 3 years is predictive of future progression to MCI.<sup>22</sup> While the present study was focused earlier in the AD spectrum, it is noted that individuals with high baseline A $\beta$  and declining cognition were the most likely to progress to MCI.

The optimal thresholds remained in a short CL range despite the use of different tracers (PiB and FBP) across the 3 unique

cohorts. Although FBP is known to exhibit greater nonspecific binding than PiB and consequently requires a more conservative whole cerebellum reference region,<sup>6</sup> both the GMM and optimal thresholds remained in a similar CL range. FBP translated primarily into a wider A $\beta$  range for ADNI, with CL values extending further into the negatives. Furthermore, despite the differences in sample populations, the robustness of the pattern of A $\beta$  accumulation and cognitive decline across samples yielded highly similar thresholds. Exclusive recruitment of CN adults in HABS resulted in a lower proportion of A $\beta$ + participants but did not hinder detection of a lower optimal threshold. The multisite ADNI exhibited greater variance in measurement over time, leading to a more modest relationship between A $\beta$  and cognitive decline, but

**Table 3** Descriptives for Participants Between Optimal and GMM Cutoffs Pooled Across Samples

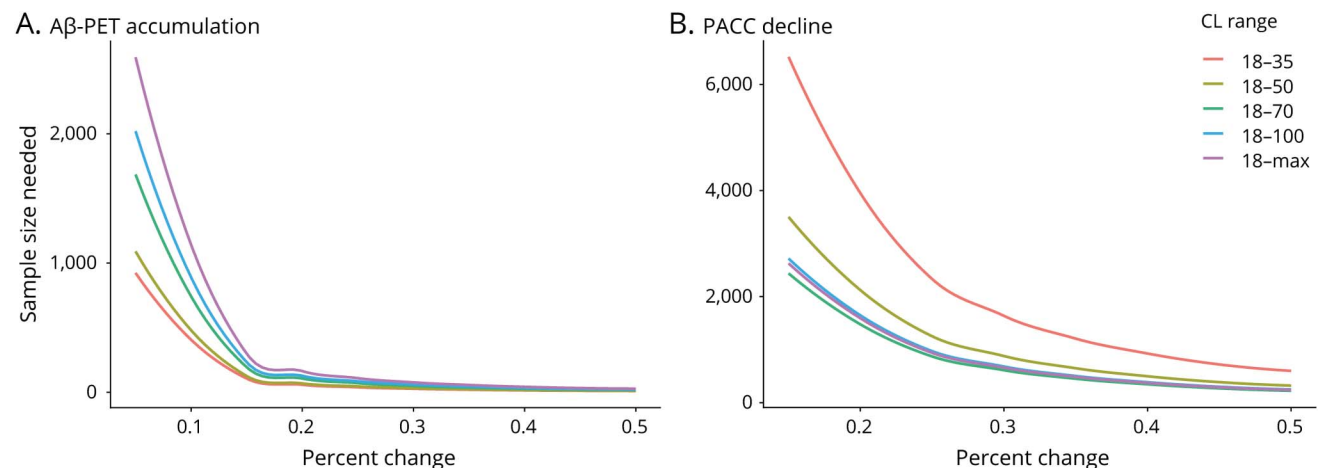
	A $\beta$ Group			Low vs Intermediate		Intermediate vs High		Low vs High	
	Low: <Optimal	Intermediate: >Optimal <GMM	High: >GMM	t Value	p Value	t Value	p Value	t Value	p Value
<b>Total, n (%)</b>	569 (67)	54 (6)	232 (27)	—	—	—	—	—	—
<b>HABS, n (%)</b>	246 (72)	19 (5)	77 (23)	—	—	—	—	—	—
<b>AIBL, n (%)</b>	102 (65)	4 (3)	51 (32)	—	—	—	—	—	—
<b>ADNI, n (%)</b>	221 (62)	31 (9)	104 (29)	—	—	—	—	—	—
<b>Age, y</b>	71.9 (7.41)	74.2 (6.38)	75.8 (6.19)	-2.26	0.02	-1.68	0.09	-7.18	<0.001
<b>Sex, n female (%)</b>	316 (56)	28 (52)	137 (59)	0.14	0.7	0.66	0.42	0.69	0.41
<b>APOE <math>\epsilon</math>4+, n (%)</b>	82/556 (19)	16/53 (30)	116/228 (51)	3.21	0.07	6.58	0.01	80.18	<0.001
<b>CL slope</b>	0.791 (2.48)	3.05 (3.44)	3.70 (6.45)	-5.14	<0.001	-0.59	0.56	-7.91	<0.001
<b>PACC5 score slope</b>	-0.01 (0.21)	-0.07 (0.29)	-0.24 (0.49)	-1.93	0.05	-2.45	0.01	-8.93	<0.001

Abbreviations: A $\beta$  =  $\beta$ -amyloid; ADNI = Alzheimer's Disease Neuroimaging Initiative; AIBL = Australian Imaging, Biomarker and Lifestyle; CL = Centiloid; GMM = gaussian mixture model; HABS = Harvard Aging Brain Study; PACC5 = Preclinical Alzheimer Cognitive Composite 5 version.

Individuals were grouped into low (below optimal cutoff), intermediate (between low and GMM cutoffs) and high (above GMM cutoff) groups within each sample and pooled to describe the additional individuals identified by the optimal threshold relative the low and high A $\beta$  groups. The total for the pooled-sample number and within-sample number is reported, along with the percentage within sample. In the pooled sample, mean (SD) is reported for age, CL slope, and PACC5 score slope and number (percent) for sex and APOE.

still arrived at a comparable threshold. In AIBL, enrichment for family history and memory complaints led to fewer individuals with low to moderate A $\beta$  and a greater dynamic range in the PiB SUVR distribution. Despite the resulting slight discrepancy in the translation of AIBL to the CL scale, the observed 15 CL optimal threshold remained very close to 16.7 to 18.5 CL observed in HABS and ADNI.

Much attention in recent years has been paid to the question of whether existing thresholds are too high.<sup>10,43</sup> Our CL range of 15 to 18.5 in HABS, ADNI, and AIBL was closely corroborated by a fourth PiB sample (Mayo Clinic Study of Aging) that demonstrated reliable worsening of PiB accumulation beyond a CL 19.<sup>15</sup> In each of our 3 samples, the optimal cutoff fell approximately 0.04 to 0.05 DVR/SUVR

**Figure 4** Sample Sizes Needed to Detect A $\beta$ -PET Accumulation and PACC Score Decline

(A and B) Sample size needed per arm to detect varying levels of change over 5 years with 80% power in (A)  $\beta$ -amyloid (A $\beta$ )-PET accumulation and (B) Preclinical Alzheimer Cognitive Composite (PACC) score decline for groups with different ranges of A $\beta$  burden measured in Centiloid (CL) at baseline. Power analyses were estimated from a combined Harvard Aging Brain Study (HABS), Alzheimer's Disease Neuroimaging Initiative (ADNI), and Australian Imaging, Biomarker and Lifestyle (AIBL) dataset measured in CL. As shown in the red line, an early intervention trial targeting individuals in the low 18 to 35 CL range would require the smallest sample to detect A $\beta$ -PET accumulation because individuals in this range typically exhibit the highest rates of accumulation. However, detecting PACC decline in the low 18 to 35 CL range would necessitate very large sample sizes. Targeting individuals in the 18 to 50 CL range would minimize the sample size needed to detect PACC score decline while maximizing power to detect reduced A $\beta$ -PET accumulation.

below their respective GMM cutoffs (CL 20–26) and below the CL 24 to 35 range most often implicated in studies using similar approaches.<sup>29,44–46</sup>

These thresholds do, however, sit slightly higher than the CL 12 cutoff that was recently suggested by autopsy and antemortem PET data from La Joie et al.<sup>14</sup> Due to the lag between PET and autopsy, it is likely that if PET were acquired closer to death, A $\beta$  accumulation in the intervening years would result in a higher CL threshold. However, a recent combined CSF-PET study<sup>47</sup> found a CL 12 threshold based on CSF A $\beta_{42}$  levels. Thus, it may instead be that autopsy and CSF allow earlier detection of A $\beta$  than is possible with PET due to their sensitivity to earlier forms of A $\beta$ , including diffuse plaques and soluble A $\beta$  oligomers. Our findings (figures 2 and 3) suggest that a CL of 12 is likely to introduce false positives in PET. A CL in the 15 to 18.5 range therefore provides a good compromise, minimizing false positives that could weaken outcomes and introduce ethical considerations while increasing inclusion of those at higher risk for future deleterious changes along the AD trajectory.

The 15 to 18.5 CL range as a lower bound may thus be a useful target threshold for clinical trials, particularly those aiming to get closer to primary prevention. As posited by the amyloid cascade hypothesis<sup>38</sup> and supported with recent evidence,<sup>48,49</sup> A $\beta$  accumulates very gradually over 2 decades or more<sup>34</sup> and leads to neocortical tau proliferation, neurodegeneration, and cognitive decline. The success of some anti-A $\beta$  therapies may depend on very early intervention at low to moderate A $\beta$  levels, before individuals exhibit neocortical tau and neurodegeneration.<sup>1,37</sup> Our analyses demonstrated that when focusing on individuals with low to moderate A $\beta$  (18–50 CL), clinical trials will be well powered to detect changes in A $\beta$  accumulation. Furthermore, despite the subtle rate of change in cognition in this earlier stage of preclinical AD, our power analyses indicate that cognition may be still be a useful outcome measure with larger sample sizes or longer follow-up. The exact threshold chosen for a given trial will depend on the specific goals of the study and the length of follow-up available to detect change; there are design trade-offs with targeting earlier vs later stages of preclinical AD and potential sample size considerations with increased heterogeneity in baseline A $\beta$  burden.

By modeling possible cutoffs separately, we limit our ability to estimate uncertainty in the cutoff. However, by using 3 independent sample with different tracers and scanners, we were able to demonstrate a short CL range (15–18.5) in which an optimal cutoff likely falls. Researchers should be cautious in selecting a single cutoff within this range because linear transformations to CL cannot be expected to completely account for all differences in scanner, acquisition, processing, tracer, and sample.<sup>25,28</sup> Evidence suggests that linear transformation to a common scale may be particularly problematic at the low levels of tracer retention around which the threshold falls.<sup>25</sup> Furthermore, while the CL recommendations<sup>27</sup> for translating in-house pipelines and ROIs to the CL scale were followed, the

additional transformations likely confer some additional error to the CL estimation. As a result, we elected to select a more conservative 18CL in our power analyses to minimize the inclusion of false positives due to measurement error. Clinical trials aimed at later AD stages should continue using more conservative existing thresholds to minimize false positives further.

In addition, ADNI had a higher proportion of individuals with only 2 PET scans, which reduces our ability to more robustly estimate A $\beta$  change. While  $\geq 3$  scans is optimal for measuring change, we opted to use all data presently available. As these studies continue to gather larger samples with  $\geq 3$  scans available, it is possible that more refined PET change measures will result in even lower optimal thresholds. Participants in the HABS, AIBL, and ADNI samples are predominantly White, highly educated individuals with low cardiovascular risk, and future studies with more representative samples are needed. Finally, testing specific aspects of episodic memory such as free recall and more targeted regional composites<sup>50–52</sup> may be even more sensitive to early detection.

Across 3 independent samples of CN older adults, we found evidence that lower thresholds for A $\beta$ -PET in the CL range of 15.0 to 18.5 are optimally predictive of future A $\beta$  accumulation and cognitive decline. This range appears to correspond to an inflection point in the A $\beta$ -PET distribution beyond which A $\beta$  and cognitive trajectories diverge from normal. Such thresholds may be useful in future studies and clinical trials using A $\beta$ -PET imaging to detect and track the earliest stages of AD. Thresholds derived from future changes in A $\beta$  and cognition converged across 3 samples, and future studies are needed to determine whether this phenomenon suggests that early changes in A $\beta$  and cognition are contemporaneous or if this convergence reflects signal-to-noise properties of the PET tracers.

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## Disclosure

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<b>Aaron P. Schultz, PhD</b>	Massachusetts General Hospital, Boston	Design, data processing, interpretation
<b>Michael J. Properzi, BEng/ BCompSci</b>	Massachusetts General Hospital, Boston	Data processing, analysis, interpretation
<b>Julie C. Price, PhD</b>	Massachusetts General Hospital, Boston	Interpretation and editing of manuscript
<b>J. Alex Becker, PhD</b>	Massachusetts General Hospital, Boston	Interpretation and editing of manuscript
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<b>Rebecca A. Betensky, PhD</b>	New York University, NY	Design, Interpretation and editing of manuscript
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## Appendix (continued)

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## References

- Sperling R, Mormino E, Johnson K. The evolution of preclinical Alzheimer's disease: implications for prevention trials. *Neuron* 2014;84:608–622.
- Sperling RA, Rentz DM, Johnson KA, et al. The A4 study: stopping AD before symptoms begin? *Sci Transl Med* 2014;6:228fs213.
- Sperling RA, Aisen PS, Beckett LA, et al. Toward defining the preclinical stages of Alzheimer's disease: recommendations from the National Institute on Aging-Alzheimer's Association workgroups on diagnostic guidelines for Alzheimer's disease. *Alzheimers Dement* 2011;7:280–292.
- Jack CR Jr, Bennett DA, Blennow K, et al. NIA-AA Research Framework: toward a biological definition of Alzheimer's disease. *Alzheimers Dement* 2018;14:535–562.
- Klunk WE, Engler H, Nordberg A, et al. Imaging brain amyloid in Alzheimer's disease with Pittsburgh compound-B. *Ann Neurol* 2004;55:306–319.
- Wong DF, Rosenberg PB, Zhou Y, et al. In vivo imaging of amyloid deposition in Alzheimer disease using the radioligand 18F-AV-45 (florbetapir [corrected] F 18). *J Nucl Med* 2010;51:913–920.
- Jureus A, Swahn BM, Sandell J, et al. Characterization of AZD4694, a novel fluorinated Abeta plaque neuroimaging PET radioligand. *J Neurochem* 2010;114:784–794.
- Mormino EC, Betensky RA, Hedden T, et al. Amyloid and APOE epsilon4 interact to influence short-term decline in preclinical Alzheimer disease. *Neurology* 2014a;82:1760–1767.
- Rowe CC, Ellis KA, Rimajova M, et al. Amyloid imaging results from the Australian Imaging, Biomarkers and Lifestyle (AIBL) study of aging. *Neurobiol Aging* 2010;31:1275–1283.
- Villeneuve S, Rabinovici GD, Cohn-Sheehy BI, et al. Existing Pittsburgh compound-B positron emission tomography thresholds are too high: statistical and pathological evaluation. *Brain* 2015;138:2020–2033.
- Mormino EC, Brandel MG, Madison CM, et al. Not quite PiB-positive, not quite PiB-negative: slight PiB elevations in elderly normal control subjects are biologically relevant. *Neuroimage* 2012;59:1152–1160.
- Farrell ME, Kennedy KM, Rodrigue KM, et al. Association of longitudinal cognitive decline with amyloid burden in middle-aged and older adults: evidence for a dose-response relationship. *JAMA Neurol* 2017;74:830–838.
- Cohen AD, Mowrey W, Weissfeld LA, et al. Classification of amyloid-positivity in controls: comparison of visual read and quantitative approaches. *Neuroimage* 2013;71:207–215.
- La Joie R, Ayakta N, Seeley WW, et al. Multisite study of the relationships between antemortem [(11C)PiB-PET Centiloid values and postmortem measures of Alzheimer's disease neuropathology. *Alzheimers Dement* 2019;15:205–216.
- Jack CR, Jr., Wiste HJ, Weigand SD, et al. Defining imaging biomarker cut points for brain aging and Alzheimer's disease. *Alzheimers Dement* 2017;13:205–216.
- Dagley A, LaPoint M, Huijbers W, et al. Harvard Aging Brain Study: dataset and accessibility. *Neuroimage* 2017;144:255–258.
- Ellis KA, Bush AI, Darby D, et al. The Australian Imaging, Biomarkers and Lifestyle (AIBL) study of aging: methodology and baseline characteristics of 1112 individuals recruited for a longitudinal study of Alzheimer's disease. *Int Psychogeriatr* 2009;21:672–687.
- Aisen PS, Petersen RC, Donohue MC, et al. Clinical core of the Alzheimer's Disease Neuroimaging Initiative: progress and plans. *Alzheimers Dement* 2010;6:239–246.
- Papp KV, Rentz DM, Orlovsky I, Sperling RA, Mormino EC. Optimizing the preclinical Alzheimer's cognitive composite with semantic processing: the PACCs. *Alzheimers Dement* 2017;3:668–677.
- Buckley RF, Mormino EC, Amariglio RE, et al. Sex, amyloid, and APOE epsilon4 and risk of cognitive decline in preclinical Alzheimer's disease: findings from three well-characterized cohorts. *Alzheimers Dement* 2018;14:1193–1203.
- Donohue MC, Sperling RA, Salmon DP, et al. The preclinical Alzheimer cognitive composite: measuring amyloid-related decline. *JAMA Neurol* 2014;71:961–970.
- Papp KV, Buckley R, Mormino E, et al. Clinical meaningfulness of subtle cognitive decline on longitudinal testing in preclinical AD. *Alzheimers Dement* 2019;16:552–560.
- Johnson KA, Schultz A, Betensky RA, et al. Tau positron emission tomographic imaging in aging and early Alzheimer disease. *Ann Neurol* 2016;79:110–119.
- Desikan RS, Segonne F, Fischl B, et al. An automated labeling system for subdividing the human cerebral cortex on MRI scans into gyral based regions of interest. *Neuroimage* 2006;31:968–980.
- Properzi MJ, Buckley RF, Chhatwal JP, et al. Nonlinear distributional mapping (NoDiM) for harmonization across amyloid-PET radiotracers. *Neuroimage* 2019;186:446–454.
- Landau SM, Mintun MA, Joshi AD, et al. Amyloid deposition, hypometabolism, and longitudinal cognitive decline. *Ann Neurol* 2012;72:578–586.
- Klunk WE, Koeppe RA, Price JC, et al. The Centiloid Project: standardizing quantitative amyloid plaque estimation by PET. *Alzheimers Dement* 2015;11:1–15.e1–e4.
- Su Y, Flores S, Hornbeck RC, et al. Utilizing the Centiloid scale in cross-sectional and longitudinal PiB PET studies. *Neuroimage Clin* 2018;19:406–416.
- Navitsky M, Joshi AD, Kennedy I, et al. Standardization of amyloid quantitation with florbetapir standardized uptake value ratios to the Centiloid scale. *Alzheimers Dement* 2018;14:1565–1571.
- Mormino EC, Betensky RA, Hedden T, et al. Synergistic effect of beta-amyloid and neurodegeneration on cognitive decline in clinically normal individuals. *JAMA Neurol* 2014b;71:1379–1385.
- Akaike H. Information theory and an extension of the maximum likelihood principle. *Proceedings of the 2nd International Symposium on Information Theory* ;1973; Budapest: 267–281.
- Schwarz G. Estimating the dimension of a model. *Ann Stat* 1978;6:461–464.
- Jack CR, Jr., Wiste HJ, Lesnick TG, et al. Brain beta-amyloid load approaches a plateau. *Neurology* 2013;80:890–896.
- Villemagne VL, Burnham S, Bourgeat P, et al. Amyloid beta deposition, neurodegeneration, and cognitive decline in sporadic Alzheimer's disease: a prospective cohort study. *Lancet Neurol* 2013;12:357–367.
- Farrell ME, Chen X, Rundle MM, Chan MY, Wig GS, Park DC. Regional amyloid accumulation and cognitive decline in initially amyloid-negative adults. *Neurology* 2018;91:e1809–e1821.
- Miller R, Siegmund D. Maximally selected chi square statistics. *Biometrics* 1982;38:1011–1016.
- Sperling RA, Jack CR, Jr., Aisen PS. Testing the right target and right drug at the right stage. *Sci Transl Med* 2011;3:111cm133.
- Hardy J, Selkoe DJ. The amyloid hypothesis of Alzheimer's disease: progress and problems on the road to therapeutics. *Science* 2002;297:353–356.
- Jack CR, Jr., Knopman DS, Jagust WJ, et al. Hypothetical model of dynamic biomarkers of the Alzheimer's pathological cascade. *Lancet Neurol* 2010;9:119–128.
- Ikonomic MD, Klunk WE, Abrahamson EE, et al. Post-mortem correlates of in vivo PiB-PET amyloid imaging in a typical case of Alzheimer's disease. *Brain* 2008;131:1630–1645.
- Leal SL, Lockhart SN, Maass A, Bell RK, Jagust WJ. Subthreshold amyloid predicts tau deposition in aging. *J Neurosci* 2018;38:4482–4489.
- Landau SM, Horgn A, Jagust WJ, Alzheimer's Disease Neuroimaging Initiative: memory decline accompanies subthreshold amyloid accumulation. *Neurology* 2018;90:e1452–e1460.
- Bischof GN, Jacobs HIL. Subthreshold amyloid and its biological and clinical meaning: long way ahead. *Neurology* 2019;93:72–79.
- Leuzy A, Chiotis K, Hasselbalch SG, et al. Pittsburgh compound B imaging and cerebrospinal fluid amyloid-beta in a multicentre European memory clinic study. *Brain* 2016;139:2540–2553.
- Rowe CC, Dore V, Jones G, et al. (18F)-florbetaben PET beta-amyloid binding expressed in Centiloids. *Eur J Nucl Med Mol Imaging* 2017;44:2053–2059.
- Battle MR, Pillay LC, Lowe VJ, et al. Centiloid scaling for quantification of brain amyloid with [(18F)]flutemetamol using multiple processing methods. *EJNMMI Res* 2018;8:107.
- Salvado G, Molinuevo JL, Brugulat-Serrat A, et al. Centiloid cut-off values for optimal agreement between PET and CSF core AD biomarkers. *Alzheimers Res Ther* 2019;11:27.
- Hanseeuw BJ, Betensky RA, Jacobs HIL, et al. Association of amyloid and tau with cognition in preclinical Alzheimer disease: a longitudinal study. *JAMA Neurol* 2019;76:915–924.
- Jack CR Jr, Bennett DA, Blennow K, et al. A/T/N: an unbiased descriptive classification scheme for Alzheimer disease biomarkers. *Neurology* 2016;87:539–547.
- Insel PS, Mormino EC, Aisen PS, Thompson WK, Donohue MC. Neuroanatomical spread of amyloid beta and tau in Alzheimer's disease: implications for primary prevention. *Brain Commun* 2020;2:fcaa007.
- Palmqvist S, Scholl M, Strandberg O, et al. Earliest accumulation of beta-amyloid occurs within the default-mode network and concurrently affects brain connectivity. *Nat Commun* 2017;8:1214.
- Grothe MJ, Barthel H, Sepulcre J, et al. In vivo staging of regional amyloid deposition. *Neurology* 2017;89:2031–2038.