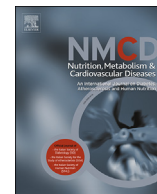




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## SYSTEMATIC REVIEWS AND META-ANALYSES

## Statin and outcomes of coronavirus disease 2019 (COVID-19): A systematic review, meta-analysis, and meta-regression

Timotius I. Hariyanto<sup>a</sup>, Andree Kurniawan<sup>b,\*</sup><sup>a</sup> Faculty of Medicine, Pelita Harapan University, Boulevard Jendral Sudirman street, Karawaci, Tangerang, 15811, Indonesia<sup>b</sup> Department of Internal Medicine, Faculty of Medicine, Pelita Harapan University, Boulevard Jendral Sudirman street, Karawaci, Tangerang, 15811, Indonesia

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**KEYWORDS**

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2019;  
COVID-19;  
Statin;  
Dyslipidemia;  
Treatment

**Abstract** *Aims:* One of the comorbidities associated with severe outcome and mortality of COVID-19 is dyslipidemia. Statin is one of the drugs which is most commonly used for the treatment of dyslipidemic patients. This study aims to analyze the association between statin use and composite poor outcomes of COVID-19.

*Data synthesis:* We systematically searched the PubMed and Europe PMC database using specific keywords related to our aims until November 25th, 2020. All articles published on COVID-19 and statin were retrieved. Statistical analysis was done using Review Manager 5.4 and Comprehensive Meta-Analysis 3 software.

A total of 35 studies with a total of 11,930,583 patients were included in our analysis. Our meta-analysis showed that statin use did not improve the composite poor outcomes of COVID-19 [OR 1.08 (95% CI 0.86–1.35),  $p = 0.50$ ,  $I^2 = 98\%$ , random-effect modelling]. Meta-regression showed that the association with composite poor outcomes of COVID-19 was influenced by age ( $p = 0.010$ ), gender ( $p = 0.045$ ), and cardiovascular disease ( $p = 0.012$ ). Subgroup analysis showed that the association was weaker in studies with median age  $\geq 60$  years-old (OR 0.94) compared to  $< 60$  years-old (OR 1.43), and in the prevalence of cardiovascular disease  $\geq 25\%$  (RR 0.94) compared to  $< 25\%$  (RR 1.24).

*Conclusion:* Statin use did not improve the composite poor outcomes of COVID-19. Patients with dyslipidemia should continue taking statin drugs despite COVID-19 infection status, given its beneficial effects on cardiovascular outcomes.

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**Introduction**

In March 2020, the World Health Organization (WHO) has declared the coronavirus disease 2019 (COVID-19) as global pandemic disease, and now nine months after that, the

number of positive and death cases from COVID-19 is still increasing. The reported manifestation of COVID-19 can range from mild respiratory symptoms such as fever and cough to severe and potentially lethal symptoms such as sepsis, arrhythmia, heart failure, and loss of consciousness [1]. Several meta-analysis have demonstrated that patients' comorbid conditions such as dyslipidemia was associated with the development of severe outcomes and mortality from COVID-19 [2–8]. Most of the patients with

\* Corresponding author. .

E-mail address: [andree.kurniawan@uph.edu](mailto:andree.kurniawan@uph.edu) (A. Kurniawan).

dyslipidemia will take statin as their daily medications. Statin belong to a HMG-CoA reductase inhibitor drugs that has been long known as an effective cholesterol-lowering agent. Previous studies have proposed that statin may be beneficial in improving the outcome of COVID-19 [9,10]. The beneficial effects of statin may be related to its pleiotropic properties. This pleiotropic property of statin is believed to reduce the burden of obesity, cardiovascular disease, and dyslipidemia which are associated with poor outcomes of COVID-19. Statin can impair the virus’s ability to infect cells and reducing its infectivity through down-regulation of CD147 in human cells, including pulmonary cells [9]. Statin can also prevent or reverse host cell lipid raft alterations induced by COVID-19 infection, which can reduce both cell infection and viral replication. Moreover, pleiotropic properties of statin can exert anti-inflammatory effects by inhibiting NLRP3 inflammasome through the TLR4/MyD88/NF-κB pathway, therefore restraining the uncontrolled inflammation which can be fatal in COVID-19 patients [9]. In silico study by Reiner et al. [10] showed that statin may be an effective SARS-CoV-2 Mpro inhibitors based on its binding energy which is higher than protease or polymerase inhibitors. However, all of these arguments are not yet supported by sufficient in human studies. This study aims to analyze the association between statin use and outcomes from COVID-19.

**Methods**

**Eligibility criteria**

Studies were included in this review if met the following inclusion criteria: representation for clinical questions (P: positive/confirmed cases of COVID-19; I: a group of patients who take statin as their medications; C: a group of patients who did not use statin; O: composite poor outcomes which comprise of risk COVID-19, severe COVID-19, and mortality), type of study was a randomized control trial, cohort, clinical trial, case-cohort, and cross-over design, and if the full-text article was available. The following types of articles were excluded: articles other than original research (e.g., review articles, letters, or commentaries); case reports; articles not in the English language; articles on research in pediatric populations (17 years of age or younger); and articles on research in pregnant women.

**Search strategy and study selection**

A systematic search of the literature was conducted on PubMed and Europe PMC using the keywords “statin” OR “lipid-lowering drugs” OR “lipid-lowering agents” AND “coronavirus disease 2019” OR “COVID-19”, between 2019 and present time (October 25th, 2020) with language restricted to English only. The title, abstract, and full text of all articles identified that matched the search criteria were assessed, and those reporting the rate of statin use in COVID-19 patients with a clinically validated definition of each component of the outcomes of interest were included in this

meta-analysis. The references of all identified studies were also analyzed (forward and backward citation tracking) to identify other potentially eligible articles. The study was carried out per the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines [11].

**Data extraction and quality assessment**

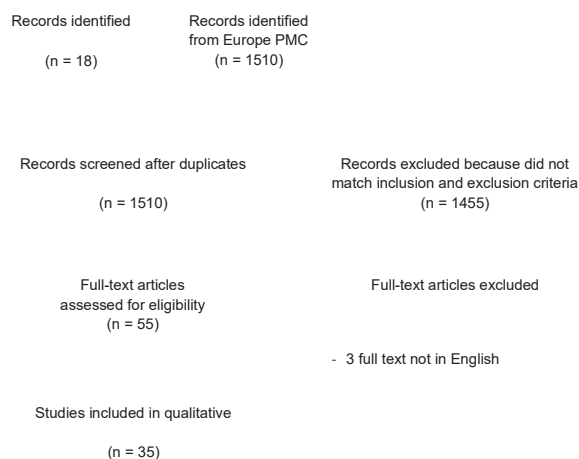
Data extraction was performed independently by two authors, we used standardized forms that include author, year, study design, number of participants, age, gender, hypertension, diabetes, cardiovascular disease, statin use, and proportion of patients with each outcome of COVID-19.

The outcome of interest was the composite poor outcomes that comprised of risk of COVID-19 infection, severe COVID-19, and mortality. The risk of COVID-19 infection was defined as the likelihood of someone getting contracted by COVID-19. Severe COVID-19 was defined as patients who had any of the following features at the time of, or after, admission: (1) respiratory distress ( $\geq 30$  breaths per min); (2) oxygen saturation at rest  $\leq 93\%$ ; (3) ratio of the partial pressure of arterial oxygen (PaO<sub>2</sub>) to a fractional concentration of oxygen inspired air (fiO<sub>2</sub>)  $\leq 300$  mmHg; or (4) critical complication (respiratory failure, septic shock, and or multiple organ dysfunction/failure) or admission into ICU. Mortality outcome from COVID-19 was defined as the number of patients who were dead because of COVID-19 infection.

Two investigators independently evaluated the quality of the included cohort and case–control studies using the Newcastle–Ottawa Scale (NOS) [12]. The selection, comparability, and exposure of each study were broadly assessed and studies were assigned a score from zero to nine. Studies with scores  $\geq 7$  were considered of good quality.

**Statistical analysis**

A meta-analysis was performed using Review Manager 5.4 (Cochrane Collaboration) and Comprehensive Meta-Analysis version 3 software. We used the Generic Inverse



**Figure 1** PRISMA flowchart.

**Table 1** Characteristics of included studies.

Study and country	Sample size	Design	Overall age mean $\pm$ SD	Male n (%)	Hypertension n (%)	Diabetes n (%)	Cardiovascular disease n (%)	Statin use n (%)
Alamdari NM et al. [12] 2020 (Iran)	459	Retrospective cohort	61.7 $\pm$ 11.8	320 (69.7%)	214 (46.6%)	119 (25.1%)	185 (40.3%)	117 (25.5%)
An C et al. [13] 2020 (Korea)	10,237	Retrospective cohort	44.9 $\pm$ 19.7	4088 (39.9%)	7090 (69.3%)	1021 (10%)	511 (5%)	1074 (10.5%)
Argenziano MG et al. [14] 2020 (USA)	1000	Retrospective cohort	62.6 $\pm$ 18.5	596 (59.6%)	601 (60.1%)	372 (37.2%)	233 (23.3%)	361 (36.1%)
Ayed M et al. [15] 2020 (Kuwait)	103	Retrospective cohort	53.3 $\pm$ 14	88 (85.5%)	36 (35%)	40 (39.2%)	12 (11.8%)	10 (9.8%)
Bifulco M et al. [16] 2020 (Italy)	541	Retrospective cohort	65 $\pm$ 11.3	341 (63%)	273 (50.4%)	130 (24%)	143 (26.4%)	117 (21.6%)
Cariou B et al. [17] 2020 (France)	1317	Retrospective cohort	69.8 $\pm$ 13	855 (64.9%)	1003 (77.2%)	1317 (100%)	140 (11.6%)	627 (47.6%)
Daniels LB et al. [18] 2020 (USA)	170	Retrospective cohort	59 $\pm$ 19	98 (57.6%)	75 (44.1%)	34 (20%)	56 (32.9%)	46 (27%)
De Spiegeleer AD et al. [19] 2020 (Belgium)	154	Retrospective cohort	85.9 $\pm$ 7.2	51 (33.1%)	39 (25.3%)	28 (18.2%)	N/A	31 (20.1%)
Dreher M et al. [20] 2020 (Germany)	50	Retrospective cohort	66.3 $\pm$ 13.3	33 (66%)	35 (70%)	29 (58%)	7 (14%)	18 (36%)
Gupta A et al. [21] 2020 (USA)	2626	Retrospective cohort	62.3 $\pm$ 20	1497 (57%)	1430 (54.4%)	968 (36.8%)	1052 (40%)	876 (33.3%)
Higuchi T et al. [22] 2020 (Japan)	57	Retrospective cohort	52.1 $\pm$ 25.5	32 (56.1%)	16 (28.1%)	13 (22.8%)	5 (8.8%)	12 (21.1%)
Hippisley-Cox J et al. [23] 2020 (England)	8,275,949	Prospective cohort	48.4 $\pm$ 18.4	4,115,973 (49.7%)	1,414,021 (17.1%)	575,610 (6.9%)	433,631 (5.24%)	1,073,039 (12.9%)
Ho F et al. [24] 2020 (England)	285,817	Prospective cohort	57.6 $\pm$ 8.4	131,589 (46%)	N/A	13,658 (4.7%)	13,819 (4.8%)	44,250 (15.4%)
Holman N et al. [25] 2020 (England)	3,138,410	Retrospective cohort	65.8 $\pm$ 14.3	1,756,110 (55.9%)	734,965 (23.4%)	3,138,410 (100%)	196,305 (6.2%)	2,218,500 (70.6%)
Huh K et al. [26] 2020 (Korea)	65,149	Case-control	48.3 $\pm$ 15.3	32,183 (49.4%)	21,368 (32.8%)	17,981 (27.6%)	13,533 (20.7%)	11,103 (17%)
Inciardi RM et al. [27] 2020 (Italy)	99	Retrospective cohort	67 $\pm$ 12	80 (81%)	63 (64%)	30 (31%)	53 (54%)	25 (26%)
Israel A et al. [28] 2020 (Israel)	20,757	Retrospective cohort	59 $\pm$ 19.1	10,473 (50.4%)	8511 (41%)	11,076 (53.3%)	6873 (33.1%)	937 (4.5%)
Izzi-Engbeaya C et al. [29] 2020	889	Retrospective cohort	65.8 $\pm$ 17.5	534 (60%)	418 (47%)	337 (38%)	373 (42%)	180.9 $\pm$ 99.8
Kibler M et al. [30] 2020 (France)	702	Retrospective cohort	82 $\pm$ 6.9	313 (44%)	587 (83.6%)	213 (30.3%)	318 (45.3%)	344 (50.2%)
Lala A et al. [31] 2020 (USA)	2736	Retrospective cohort	66.4 $\pm$ 15.8	1630 (59.6%)	1065 (38.9%)	719 (26.3%)	935 (34.1%)	984 (36%)
Luo P et al. [32] 2020 (China)	283	Retrospective cohort	64.5 $\pm$ 10	156 (55.1%)	164 (57.9%)	283 (100%)	43 (15.1%)	55 (19.4%)
Maddaloni E et al. [33] 2020 (Italy)	237	Case-control	75 $\pm$ 12.5	151 (63.7%)	N/A	237 (100%)	31 (13%)	150 (63.2%)
Masana L et al. [34] 2020 (Spain)	2157	Retrospective cohort	66.3 $\pm$ 17.7	1234 (57.3%)	1081 (50.1%)	501 (23.2%)	620 (28.7%)	581 (26.9%)
McCarthy CP et al. [35] 2020 (USA)	247	Retrospective cohort	62.3 $\pm$ 19.2	143 (57.9%)	128 (51.8%)	68 (27.5%)	114 (46.1%)	107 (43.3%)
Ramachandran P et al. [36] 2020 (USA)	295	Retrospective cohort	64.3 $\pm$ 14.8	162 (54.9%)	209 (70.8%)	132 (44.7%)	45 (15.2%)	114 (38.6%)
Rodriguez-Nava G et al. [37] 2020 (USA)	87	Retrospective cohort	67 $\pm$ 12.5	56 (64.4%)	N/A	N/A	N/A	47 (54%)
Saeed O et al. [38] 2020 (USA)	4252	Retrospective cohort	65 $\pm$ 16	2255 (53%)	3060 (72%)	2266 (53%)	1111 (26%)	1355 (31.8%)
Song SL et al. [39] 2020 (USA)	249	Retrospective cohort	62.6 $\pm$ 17.7	142 (57%)	122 (49%)	83 (33.3%)	70 (28.1%)	123 (49.3%)
Tan WYT et al. [40] 2020 (Singapore)	717	Retrospective cohort	40.6 $\pm$ 28.1	410 (57.2%)	139 (19.3%)	76 (10.6%)	50 (6.9%)	151 (21%)
Ullah AZMD et al. [41] 2020 (England)	15,586	Retrospective cohort	57.1 $\pm$ 18.2	6840 (43.9%)	10,167 (65.2%)	6047 (38.8%)	4421 (28.4%)	5221 (33.5%)
Vila-Corcoles A et al. [42] 2020 (Spain)	34,936	Retrospective cohort	70.9 $\pm$ 11.3	16,805 (48.1%)	34,936 (100%)	9829 (28.1%)	10,097 (28.9%)	11,328 (32.4%)
Wang B et al. [43] 2020 (USA)	58	Retrospective cohort	67 $\pm$ 12.5	30 (52%)	37 (64%)	16 (28%)	20 (34%)	27 (47%)
Yan H et al. [44] 2020 (China)	49,245	Retrospective cohort	49.9 $\pm$ 16.6	23,799 (48.3%)	9985 (20.2%)	2977 (6%)	637 (2.1%)	458 (1.5%)
Zeng H et al. [45] 2020 (China)	1031	Retrospective cohort	60.3 $\pm$ 14.3	538 (52.2%)	384 (37.2%)	189 (18.3%)	84 (8.1%)	38 (3.6%)
Zhang XJ et al. [46] 2020 (China)	13,981	Retrospective cohort	56.3 $\pm$ 16.2	6830 (48.8%)	4860 (34.7%)	2282 (16.3%)	1171 (8.3%)	1219 (8.7%)

**Table 2** Newcastle–Ottawa quality assessment of observational studies.

First author, year	Study design	Selection	Comparability	Outcome	Total score	Result
Alamdari NM et al. [12] 2020	Cohort	***	**	***	8	Good
An C et al. [13] 2020	Cohort	***	**	***	8	Good
Argenziano MG et al. [14] 2020	Cohort	***	**	***	8	Good
Ayed M et al. [15] 2020	Cohort	***	**	***	8	Good
Bifulco M et al. [16] 2020	Cohort	***	**	**	7	Good
Cariou B et al. [17] 2020	Cohort	***	**	****	9	Good
Daniels LB et al. [18] 2020	Cohort	***	**	***	8	Good
De Spiegeleer AD et al. [19] 2020	Cohort	***	**	***	8	Good
Dreher et al. [20] 2020	Cohort	**	**	***	7	Good
Gupta A et al. [21] 2020	Cohort	***	**	****	9	Good
Higuchi T et al. [22] 2020	Cohort	***	**	***	8	Good
Hippisley-Cox J et al. [23] 2020	Cohort	***	**	***	8	Good
Ho F et al. [24] 2020	Cohort	***	**	***	8	Good
Holman N et al. [25] 2020	Cohort	****	**	***	9	Good
Huh K et al. [26] 2020	Case-control	**	**	***	8	Good
Inciardi RM et al. [27] 2020	Cohort	***	**	***	8	Good
Israel A et al. [28] 2020	Cohort	***	**	***	8	Good
Izzi-Engbeaya C et al. [29] 2020	Cohort	***	**	***	8	Good
Kibler M et al. [30] 2020	Cohort	***	**	***	8	Good
Lala A et al. [31] 2020	Cohort	***	**	****	9	Good
Luo P et al. [32] 2020	Cohort	**	**	***	7	Good
Maddaloni E et al. [33] 2020	Case-control	***	**	**	7	Good
Masana L et al. [34] 2020	Cohort	****	**	***	9	Good
McCarthy CP et al. [35] 2020	Cohort	****	**	***	9	Good
Ramachandran P et al. [36] 2020	Cohort	***	**	***	8	Good
Rodríguez-Nava G et al. [37] 2020	Cohort	***	**	**	7	Good
Saeed O et al. [38] 2020	Cohort	***	**	****	9	Good
Song SL et al. [39] 2020	Cohort	***	**	***	8	Good
Tan WYT et al. [40] 2020	Cohort	**	**	***	7	Good
Ullah AZMD et al. [41] 2020	Cohort	***	**	***	8	Good
Vila-Corcoles A et al. [42] 2020	Cohort	***	**	**	7	Good
Wang B et al. [43] 2020	Cohort	***	**	***	8	Good
Yan H et al. [44] 2020	Cohort	***	**	***	8	Good
Zeng H et al. [45] 2020	Cohort	***	**	***	8	Good
Zhang XJ et al. [46] 2020	Cohort	***	**	****	9	Good

Variance formula with random-effects models to calculate each outcome’s risk. The heterogeneity was assessed by using the  $I^2$  statistic with a value of <25%, 26–50%, and >50% were considered as low, moderate, and high degrees of heterogeneity, respectively. The effect estimate was reported as odds ratio (OR) along with its 95% confidence intervals (CIs). P-value was two-tailed, and the statistical significance was set at  $\leq 0.05$ . Random effects meta-regression was performed using a restricted-maximum likelihood for pre-specified variables including age, gender, hypertension, diabetes, and cardiovascular disease. Subgroup analysis was performed for each component of composite poor outcomes. We performed Begg’s funnel-plot analysis to qualitatively assess the risk of publication bias.

**Results**

**Study selection and characteristics**

A total of 1528 records were obtained through systematic electronic searches. After the removal of duplicates, 1510 records remained. A total of 1455 records were excluded after screening the titles/abstracts because they did not

match our inclusion and exclusion criteria. After evaluating 55 full-texts for eligibility, 10 full-text articles were excluded because they do not have the outcome of interest (risk of COVID-19, severe COVID-19, and mortality), 7 full-text articles were excluded because they do not have the control/comparison group, 3 full-text articles were excluded because the articles were not in English, and finally, 35 studies [12–46] with a total of 11, 930, 583 sample sizes were included in the meta-analysis (Fig. 1). Of a total of 35 included studies, 31 were retrospective cohort, 2 studies were prospective cohort, while the remaining 2 study were case–control study. The essential characteristics of the included studies are summarized in Table 1.

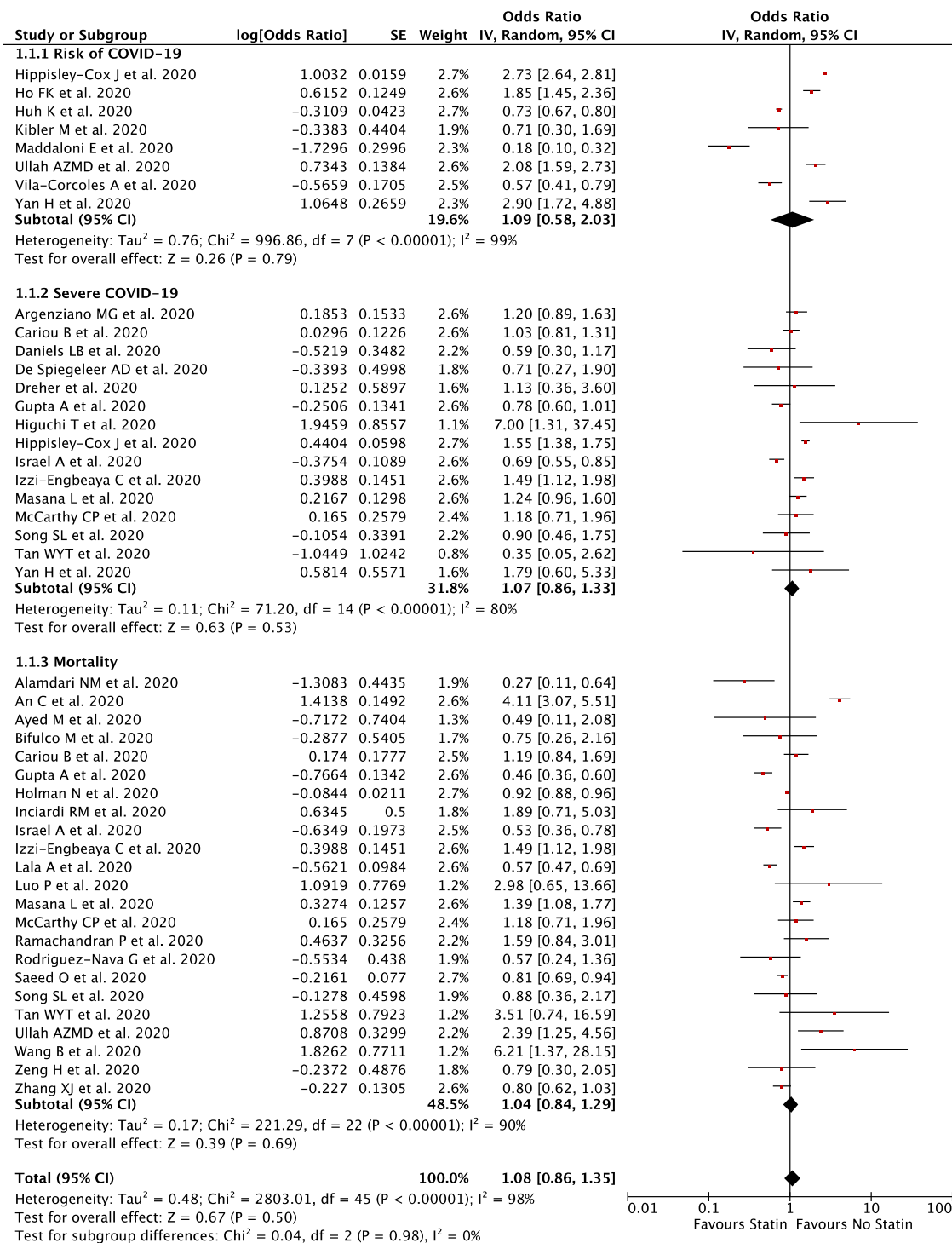
**Quality of study assessment**

Studies with various study designs including cohort and case–control were included in this review and assessed accordingly with the appropriate scale or tool. Newcastle Ottawa Scales (NOS) were used to assess the cohort and case–control studies (Table 2). All included studies were rated ‘good’. In conclusion, all studies were seemed fit to be included in the meta-analysis.

**Statin and outcomes**

Our pooled analysis showed that statin was not associated with composite poor outcome [OR 1.08 (95% CI 0.86–1.35),  $p = 0.50$ ,  $I^2 = 98%$ , random-effect modelling] (Fig. 2). Subgroup analysis showed that statin was not associated

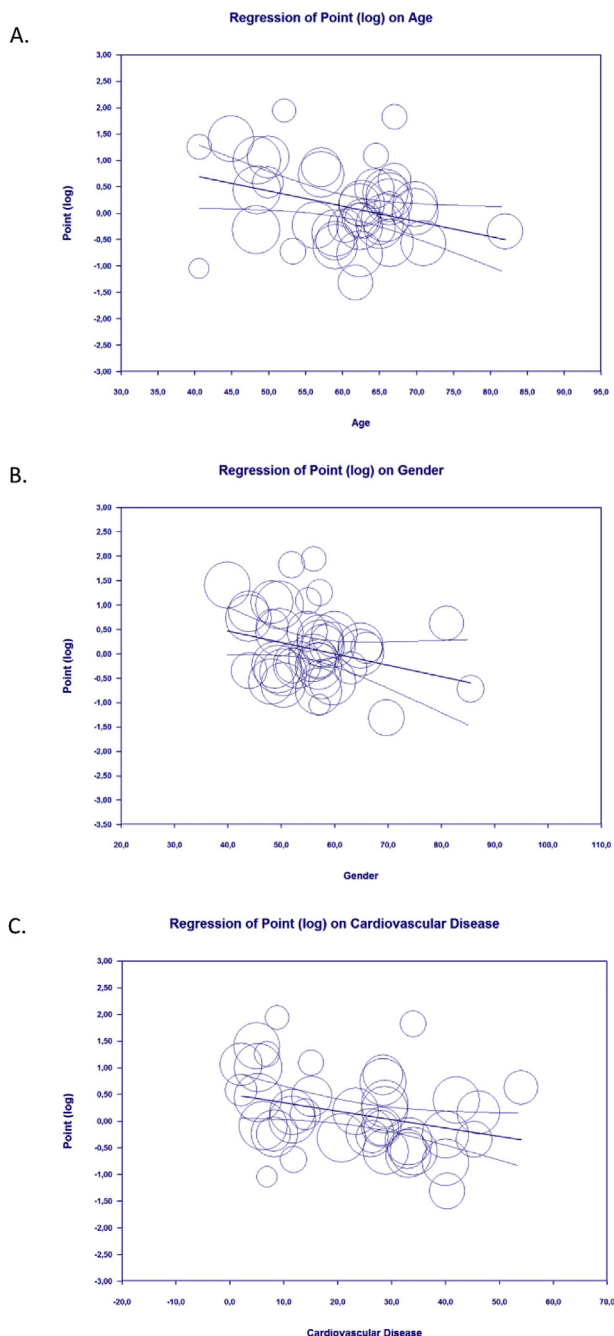
with risk of COVID-19 [OR 1.09 (95% CI 0.58–2.03),  $p = 0.79$ ,  $I^2 = 99%$ , random-effect modelling], severe COVID-19 [OR 1.07 (95% CI 0.86–1.33),  $p = 0.53$ ,  $I^2 = 80%$ , random-effect modelling], and mortality from COVID-19 [OR 1.04 (95% CI 0.84–1.29),  $p = 0.69$ ,  $I^2 = 90%$ , random-effect modelling].



**Figure 2** Forest plot that demonstrates the association of statin with composite poor outcome and its subgroup which comprises of risk of COVID-19, severe COVID-19, and mortality.

**Meta-regression**

Meta-regression showed that the association between statin and composite poor outcome was affected by age ( $p = 0.010$ ) (Fig. 3A), gender ( $p = 0.045$ ) (Fig. 3B), and cardiovascular disease ( $p = 0.012$ ) (Fig. 3C), but not affected by hypertension ( $p = 0.610$ ) and diabetes mellitus ( $p = 0.246$ ).



**Figure 3** Bubble-plot for Meta-regression. Meta-regression analysis showed that the association between statin and composite poor outcome was affected by age [A], gender [B], and cardiovascular disease [C].

**Subgroup analysis**

Subgroup analysis for studies with median age  $\geq 60$  years [OR 0.94 (95% CI 0.80–1.09),  $p = 0.39$ ,  $I^2 = 88\%$ , random-effect modelling] showed a lower OR for composite poor outcome compared to  $< 60$  years [OR 1.43 (95% CI 0.91–2.24),  $p = 0.12$ ,  $I^2 = 99\%$ , random-effect modelling].

Subgroup analysis for studies with prevalence of male gender  $\geq 50\%$  [OR 0.91 (95% CI 0.79–1.04),  $p = 0.17$ ,  $I^2 = 82\%$ , random-effect modelling] showed a lower OR for composite poor outcome compared to  $< 50\%$  [OR 1.49 (95% CI 0.97–2.27),  $p = 0.07$ ,  $I^2 = 99\%$ , random-effect modelling].

Subgroup analysis for studies with prevalence of cardiovascular disease  $\geq 25\%$  [OR 0.94 (95% CI 0.76–1.16),  $p = 0.58$ ,  $I^2 = 88\%$ , random-effect modelling] showed a lower OR for composite poor outcome compared to  $< 25\%$  [OR 1.24 (95% CI 0.89–1.72),  $p = 0.20$ ,  $I^2 = 99\%$ , random-effect modelling].

Subgroup analysis for studies with USA population [OR 0.85 (95% CI 0.68–1.06),  $p = 0.15$ ,  $I^2 = 77\%$ , random-effect modelling] showed a lower OR for composite poor outcome compared to Asia population [OR 1.12 (95% CI 0.75–1.69),  $p = 0.57$ ,  $I^2 = 93\%$ , random-effect modelling] and Europe population [OR 1.16 (95% CI 0.81–1.68),  $p = 0.42$ ,  $I^2 = 99\%$ , random-effect modelling].

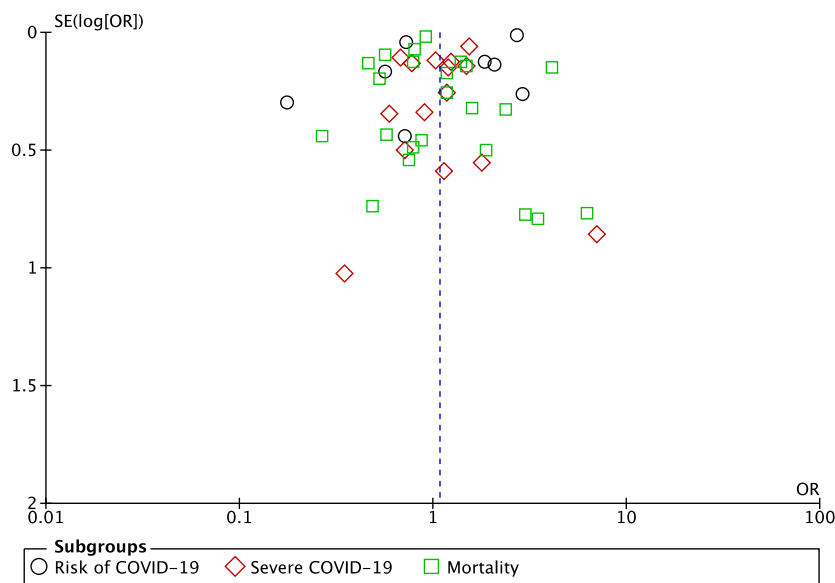
**Publication bias**

The funnel-plot analysis showed a qualitatively symmetrical inverted funnel-plot for the association between statin and composite poor outcome (Fig. 4), showing no indication of publication bias.

**Discussion**

Previous meta-analysis that showed the benefit of statin in reducing the disease severity and/or mortality from COVID-19 only involved 4 studies and did not elaborate the effects of the confounding factors which can affect the relationship between statin use and outcome of COVID-19, such as age, gender, and comorbid conditions, resulting in weak and preliminary conclusions [47]. Our systematic review and meta-analysis already involved 35 included studies and not only analyzes the association between statin use and composite poor outcomes of COVID-19, but also elaborate the effect of the confounding factors such as age, gender, and comorbid conditions.

This comprehensive meta-analysis of 35 studies showed that statin use was not associated with composite poor outcomes of COVID-19 which comprised of risk of COVID-19, severe COVID-19, and mortality from COVID-19. This association was influenced by age, gender, and cardiovascular disease. Further analysis based on meta-regression showed that magnitude of risk linked to statin use as a single factor was greater in studies with younger, non-male dominant, and low prevalence of cardiovascular disease patients, which is yet to be addressed by the existing literature.



**Figure 4** Funnel plot analysis for the association of statin with composite poor outcome of COVID-19.

Several reasons can be proposed to explain why statin use did not improve the outcomes from COVID-19. First, apart from its beneficial effects through pleiotropic properties, the action of statin on TLR and NF- $\kappa$ B signaling also carry the potential risk of exacerbating compensatory immune signals and poor disease outcome [48]. This theory is supported by a retrospective analysis of the findings from a multicenter clinical trial on the efficacy of rosuvastatin against infection-induced ARDS that showed higher IL-18 levels and mortality in statin-treated patients [49]. Second, based on the experimental studies in animal models, statin therapy may upregulate the expression of angiotensin-converting enzyme 2 (ACE2), which mediates the SARS-CoV-2 entry into host cells, thus can increase its viral load and infectivity, resulting in severe outcome of the disease [50,51]. Third, statin may cause hepatotoxicity and myotoxicity in some patients that can cause acute kidney injury. Moreover, this adverse effect of statin is higher in older patients and may be exacerbated when given concomitantly with antiviral agents for COVID-19 such as lopinavir and darunavir [52,53]. Statins are a substrate for cytochrome P450 (CYP) system, especially 3A isoenzymes and P-glycoproteins (P-gp), while protease inhibitors, such as lopinavir and darunavir are potent inhibitors of both CYP3A and P-gp, so their concomitant administration will result in markedly statin exposure and adverse effects [48,54]. Therefore, the beneficial effects of statin may be counterbalanced by its potentially harmful effects and causing a neutral effect toward COVID-19. Finally, the anti-inflammatory effects of statin are relatively lower than corticosteroids [55], making statin cannot make a

significant alteration in the inflammation or cytokine storm that happened in COVID-19, thus the composite poor outcomes of COVID-19 was not altered by statin administration.

Patients with dyslipidemia and cardiovascular disease should hence be advised to still continue taking statin during the COVID-19 pandemic given its pleiotropic effects which offer benefit in reducing the cardiovascular adverse events and its neutral effects toward COVID-19 outcomes. Physicians should also still consider giving statin into the treatment regime of their patients who have dyslipidemia or cardiovascular disease if the patients have not been given it yet. However, in older patients, the benefit-risk balance of statin therapy should be carefully evaluated, either discontinuation of statin therapy and switching into other lipid-lowering therapies or continuation with caution and at lower doses is possible options.

This study has several limitations. First, data on the dosage and duration of statin therapy were lacking in the included studies, hence, cannot be analyzed. Second, we include some pre-print studies to minimize the risk of publication bias, however, the authors have made exhaustive efforts to ensure that only sound studies were included and we expect that most of those studies currently available in pre-print form will eventually be published and that we will identify them through ongoing electronic literature surveillances. We hope that this study can give further insight into statin therapy in COVID-19 patients.

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## Declaration of competing interest

The authors declare no conflict of interest.

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