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## Spanking and Child Outcomes: Old Controversies and New Meta-Analyses

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### Abstract

Whether spanking is helpful or harmful to children continues to be the source of considerable debate among both researchers and the public. This article addresses 2 persistent issues, namely whether effect sizes for spanking are distinct from those for physical abuse, and whether effect sizes for spanking are robust to study design differences. Meta-analyses focused specifically on spanking were conducted on a total of 111 unique effect sizes representing 160,927 children. Thirteen of 17 mean effect sizes were significantly different from zero and all indicated a link between spanking and increased risk for detrimental child outcomes. Effect sizes did not substantially differ between spanking and physical abuse or by study design characteristics.

### Keywords

spanking; physical punishment; discipline; meta-analysis

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Around the world, most children (80%) are spanked or otherwise physically punished by their parents (UNICEF, 2014). The question of whether parents should spank their children to correct misbehaviors sits at a nexus of arguments from ethical, religious, and human rights perspectives both in the U.S. and around the world (Gershoff, 2013). Several hundred studies have been conducted on the associations between parents' use of spanking or physical punishment and children's behavioral, emotional, cognitive, and physical outcomes, making spanking one of the most studied aspects of parenting. What has been learned from these hundreds of studies? Several efforts have been made to synthesize this large body of research, first in narrative form (Becker, 1964; Larzelere, 1996; Steinmetz, 1979; Straus, 2001) and later through meta-analyses (Ferguson, 2013; Gershoff, 2002; Larzelere & Kuhn, 2005; Paolucci & Violato, 2004). Each of these four meta-analyses included a different set of articles and came to varied conclusions, namely that physical punishment is largely ineffective and harmful (Gershoff, 2002), that physical punishment is effective under certain circumstances (Larzelere & Kuhn, 2005), and that physical punishment is linked with children's cognitive, emotional, and behavioral problems but only modestly (Ferguson, 2013; Paolucci & Violato, 2004). These competing conclusions have left both social science

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researchers and the public at large confused about what outcomes can and cannot be attributed to spanking.

As this body of work on spanking and physical punishment has accumulated, several nagging questions about the quality, consistency, and generalizability of the research have persisted. Two primary concerns that have been raised about past meta-analyses are that spanking has been confounded with potentially abusive parenting behaviors in some studies and that spanking has only been linked with detrimental outcomes in methodologically weak studies (Baumrind, Larzelere, & Cowan, 2002; Ferguson, 2013; Larzelere & Kuhn, 2005). The goal of the current article is to address these two concerns with a new set of meta-analyses using the most recent research studies to date. Because the social science theories regarding why spanking might be linked with child outcomes have been summarized extensively elsewhere (Donnelly & Straus, 2005; Gershoff, 2002), we will not repeat them here and instead will focus in this paper on key questions about the research conducted to date.

The terms “corporal punishment,” “physical punishment,” and “spanking” are largely synonymous in American culture. The majority of the studies discussed in our literature review use the term physical punishment which we define as *noninjurious, open-handed hitting with the intention of modifying child behavior*. In our meta-analyses, however, we focused on the most common form of physical punishment which is known in the U.S. as spanking, and which we define as *hitting a child on their buttocks or extremities using an open hand*.

## Previous Meta-Analyses of Physical Punishment and Spanking

The question of whether parents’ use of spanking or physical punishment is linked with children’s outcomes has been addressed in four published meta-analyses in the last 15 years. The first and most widely cited of the meta-analyses was by Gershoff (2002). This review included 88 studies used in separate meta-analyses of the associations between parents’ use of physical punishment and 11 child outcomes, four of which were measured in adulthood. Physical punishment was defined as “the use of physical force with the intention of causing a child to experience pain but not injury for the purposes of correction or control of the child’s behavior” (per Straus, 2001, p. 4) and excluded any methods that would “knowingly cause severe injury to the child” (Gershoff, 2002, p. 543). All 11 meta-analyses were significant and all but one indicated an undesirable association. Specifically, physical punishment was associated with more immediate compliance ( $d = 1.13$ ) but was also associated with lower levels of moral internalization ( $d = -.33$ ), quality of the parent–child relationship ( $d = -.58$ ), and mental health in childhood ( $d = -.49$ ) and adulthood ( $d = -.09$ ), as well as with higher levels of aggression in childhood ( $d = .36$ ) and adulthood ( $d = .57$ ), antisocial behavior in childhood ( $d = .42$ ) and adulthood ( $d = .42$ ), risk of being a victim of physical abuse ( $d = .69$ ), and risk of abusing own child or spouse as an adult ( $d = .13$ ).

The second meta-analytic article on the outcomes associated with physical punishment included 70 studies in three meta-analyses (Paolucci & Violato, 2004). Physical punishment was defined as “a form of nonabusive or customary physical punishment by a parent or adult



violence” (Ferguson, 2013, p. 199). The bivariate effect sizes for spanking and corporal punishment (cp) were significantly different from zero across all three outcomes: externalizing,  $d_{cp} = .18$  and  $d_{spanking} = .14$ ; internalizing,  $d_{cp} = .21$  and  $d_{spanking} = .12$ ; and cognitive performance  $d_{cp} = -.18$  and  $d_{spanking} = -.09$ . A secondary set of meta-analyses was conducted for studies that reported effect sizes controlling for children’s previous behavior; there were not sufficient numbers of studies for all possible comparisons, but reported effect sizes for externalizing behavior problems were  $d_{cp} = .08$  and  $d_{spanking} = .07$ , for internalizing was  $d_{spanking} = .10$ , and for cognitive performance was  $d_{cp} = -.11$ , all statistically significant at  $p < .05$ . The effect sizes for spanking were smaller than for corporal punishment, and the effect sizes for longitudinal associations controlling for the child’s previous behavior were smaller than basic longitudinal associations, yet all were significantly different from zero and all indicated detrimental outcomes associated with spanking or corporal punishment.

Taken together, these meta-analyses provide evidence that physical punishment is associated with negative child outcomes, particularly when the outcomes are divided into finer-grained categories (Ferguson, 2013; Gershoff, 2002) rather than when they are grouped into broad categories (Paolucci & Violato, 2004), and that harsher methods of physical punishment are more strongly associated with negative child outcomes than ordinary spanking (Ferguson, 2013; Larzelere & Kuhn, 2005).

## Remaining Concerns About the Research on Spanking and Child Outcomes

The meta-analyses in the present study were conducted in order to address two persistent questions about the research to date in order to clarify what is known about the potential impacts of parents’ use of physical punishment on children.

### Spanking Has Been Confounded With Harsher Forms of Physical Punishment

The main criticism of the Gershoff (2002) meta-analysis has been that it included harsh and potentially injurious behaviors, such as hitting children with objects, in its definition of physical punishment (Baumrind et al., 2002; Benjet & Kazdin, 2003; although note that this criticism applies to the Paolucci & Violato, 2004 meta-analysis as well). This broad definition of physical punishment included parent behaviors that most professionals and most parents would agree were abusive and that may be linked with negative outcomes while spanking is not (Kazdin & Benjet, 2003). Baumrind, Larzelere, and Cowan (2002) reanalyzed the data from Gershoff (2002), separating out what they deemed harsh or potentially abusive forms of physical punishment. They reported that the effect size for the studies using less severe physical punishment was significantly smaller than the effect size for harsh physical punishment ( $d_{less\ severe} = .30$  vs.  $d_{more\ severe} = .46$ ,  $\chi^2[1, n = 12,244] = 74.50$ ,  $p < .001$ ). They concluded that only severe methods of physical punishment are harmful. However, both effect sizes are significant and positive, indicating that both are associated with more undesirable child outcomes.

To help resolve this debate, our first research question was thus, are past findings that physical punishment is associated with detrimental child outcomes driven by the inclusion of harsh or abusive methods, or is spanking on its own associated with these detrimental

outcomes? We addressed this question using two strategies. First, we focused on “studies of parents’ behaviors labeled as “spanking” (see definition above) or as” synonymous terms for the same behavior (e.g., “smacking,” “slapping,” and “hitting”). This definition therefore excluded the use of objects, the use of methods that have a reasonable expectation of causing harm or injury (e.g., beating, burning, choking, whipping), and the use of methods that are gratuitous expressions of parent displeasure without a clear disciplinary component (e.g., pulling hair, shaking, shoving). By restricting our operationalization of physical punishment in this way, we were able to determine the extent to which ordinary spanking is linked with child outcomes.

Our second strategy was to examine the ways in which the strength and direction of the associations between spanking and child outcomes compare with the strength and direction of the associations between clearly abusive methods and child outcomes. We identified studies that assessed the same individuals for exposure to both ordinary spanking and to harsher methods in order to isolate the associations of one from the other. A comparison of studies of spanking to studies of abuse would not be helpful in this regard, because there could be many selection factors that distinguish the individuals reporting spanking from those reporting harsher methods. Some have argued that parents who use harsh or abusive methods are fundamentally different from parents who use only spanking (Baumrind et al., 2002) while some past research has found that genetic factors in the child elicit corporal punishment but not physical abuse (Jaffee et al., 2004). By focusing on studies that assessed the extent to which individuals experience both spanking and abuse, we compared the unique association of spanking with child outcomes to the unique association of abusive behaviors with child outcomes for the same samples of children.

### **Spanking Has Only Been Linked With Negative Child Outcomes in Cross-Sectional or Methodologically Weak Studies**

The primary standard for determining causal relations among variables has been the randomized controlled experiment because potentially confounding selection factors that might distinguish naturally occurring groups (e.g., spankers and nonspankers) are eliminated through randomization (Shadish, Cook, & Campbell, 2001). However, parents’ use of spanking is not easily or ethically studied through an experimental design, as children cannot be randomly assigned to parents with varying predispositions to spank, nor can parents typically be randomly assigned to spank or not spank. There are a small handful of experimental studies that examine whether children comply more in a laboratory setting when mothers use spanking (Bean & Roberts, 1981; Day & Roberts, 1983; Roberts, 1988; Roberts & Powers, 1990); we include these studies in the meta-analyses and discuss them more below. There also have been a few efforts to evaluate the effects of interventions designed to reduce spanking (e.g., Beauchaine, Webster-Stratton, & Reid, 2005), but these studies require a sample of parents who are willing to not spank and thus may be fundamentally different from most spankers in the population. The circumstances of experimentally manipulated spanking thus are likely to be unusual, leading to concern that experiments with parental spanking may suffer from a lack of external validity.

The next strongest approach to studying spanking are studies which examine whether it predicts changes in child outcomes over time. Such prospective longitudinal designs meet one of the key criteria for establishing causality, namely temporal precedence of the spanking independent variable (Shadish et al., 2001). Longitudinal effect sizes of the bivariate links between spanking and later child outcomes do not rule out the potential for a child elicitation effect; however, so few studies report a coefficient that controls only for initial child behavior (and not for a range of other covariates) that we are unable to meta-analyze them. Thus, while not a perfect solution, longitudinal bivariate coefficients are decidedly stronger methodologically than within-time coefficients.

Our second research question was thus: Are associations between spanking and child outcomes only found in methodologically weak studies? In order to address this question, we conducted moderator analyses that examined whether the direction and significance of the mean weighted effect sizes were similar across longitudinal, experimental, and cross-sectional studies. We also examined whether effect sizes varied according to five other dimensions of study design: measure of spanking, time period in which spanking was administered, index of spanking, whether the study assessed the associations of spanking with outcomes within a single group, or employed comparisons between two or more groups, and independence of raters of spanking and outcome. Using these dimensions of study quality as moderators allowed us to examine whether spanking is only associated with child outcomes in some types of studies and not others, a finding which would undermine the generalizability of spanking research.

## The Present Study

Given the pervasive use of spanking around the world, and in light of concerns raised about spanking by professional organizations (American Academy of Pediatrics, 2012) and intergovernmental and human rights organizations (Committee on the Rights of the Child, 2006), there is a need for definitive conclusions about the potential consequences of spanking for children. The purpose of the current study was to conduct a new set of meta-analyses to address the two unresolved debates described above and to do so while incorporating an additional 13 years of literature since the first meta-analysis was published (Gershoff, 2002). The present study is distinguished from the previous meta-analyses by focusing exclusively on parents' use of spanking, by including only peer-reviewed journal articles, by using random effects meta-analyses, and by incorporating several dozen new studies not included in previous meta-analyses.

## Method

### Identification of Potential Studies for Inclusion

The studies for the present meta-analyses were identified from two main sources. The primary source for studies was a comprehensive literature review of articles listed in four academic abstracting databases (ERIC, Medline, PsycInfo, and Sociological Abstracts) that had been published before June 1, 2014. Each database was searched using six terms for physical punishment, namely "spank\*," "corporal punishment," "physical punishment," "physical discipline," "harsh punishment," and "harsh discipline." In addition, all of the

studies used in the previously published meta-analyses (Ferguson, 2013; Gershoff, 2002; Larzelere & Kuhn, 2005; Paolucci & Violato, 2004) were considered for inclusion. These two methods yielded a total of 1,574 unique articles to be considered for inclusion in the current meta-analyses.

### **Coding of Studies for Inclusion or Exclusion**

Coding of studies involved a two-step process. In the initial step, the titles, abstracts, or full text of the 1,574 studies identified through the sources above were subjected to an initial screening. Studies were excluded at this stage if they were not relevant to or usable in the meta-analyses; examples of studies excluded at this stage were literature reviews, studies of beliefs about rather than use of spanking, and studies that were not available in English. This initial screening process eliminated 1,016 studies and retained 558 potential studies.

In the second step of coding, each of these 558 potential studies was coded independently by each of the authors. Any disagreements in coding were resolved through follow-up discussion. Studies were coded as to whether they met several criteria: (a) the study was published in a peer-reviewed journal; all book chapters, unpublished dissertations, and unpublished conference papers were excluded, even if they had been included in any of the previously published meta-analyses; (b) the study included a measure of parents' use of customary, noninjurious spanking (or slapping or hitting) that was intended to be a correction of a child's misbehavior. The terms "spank" or "smack" were used alone or in combination with other general terms (e.g., slap) in 63% of studies. The remaining studies measured corporal punishment as "physical punishment" or "physical discipline" (19%), "corporal punishment" (10%), and "slap or hit" (8%); (c) the study reported a bivariate association between parents' spanking and the child outcome of interest; and (d) The study included appropriate statistics for calculating effect sizes. The reasons for exclusion of all 1,499 studies are listed the Appendix. Only 75 studies met all four criteria and were retained for the meta-analyses.

### **Inclusion and Exclusion of Studies From Past Meta-Analyses**

All of the 162 unique studies used in the four previously published meta-analyses were considered for inclusion, but only 36 met all of our criteria. Of the 88 studies in Gershoff (2002), 23 were included in the present study. Paolucci and Violato (2004) analyzed 70 studies; 16 were included here. Of the 26 studies in Larzelere and Kuhn (2005), 11 were included. Ferguson (2013) analyzed 45 studies; of these, 11 were included in the current meta-analyses. Reasons for study exclusion are available from the first author. Thus, 39 of the 75 studies included in the current meta-analyses (52%) have not been included in previous meta-analyses.

### **Coding of Effect Sizes**

All study-level effect sizes were calculated independently by each of the authors; for all effect sizes, agreement was achieved to at least the third decimal place. When discrepancies occurred in effect size calculations, the discrepancy was discussed, and then each author independently recalculated the effect size. This process was repeated, if necessary, until consensus was achieved. Study-level effect sizes were transformed into standardized mean

difference effect sizes to allow combination across effect sizes using Cohen's formula for  $d$  (Cohen, 1988; Sterne, 2009)

$$\text{Cohen's } d = \frac{\text{mean}_{\text{treatment}} - \text{mean}_{\text{comparison}}}{sd_{\text{pooled}}}$$

where  $sd_{\text{pooled}}$  was calculated as

$$sd_{\text{pooled}} = \sqrt{\frac{(n_1 - 1) * sd_1^2 + (n_2 - 1) * sd_2^2}{n_1 + n_2 - 2}}$$

Calculation of Cohen's  $d$  was straightforward when an article reported the sample size, mean and standard deviation of a group exposed to spanking and one that had never been spanked. For articles that did not report effects as group comparisons, we utilized formulas found in Borenstein, Hedges, Higgins, and Rothstein (2009) and Johnson (1993) to convert quantitative measures of association such as correlations and differences of proportions to Cohen's  $d$  effect sizes. For each study, we also calculated the standard error of the estimate of Cohen's  $d$  utilizing formulas given in Sterne (2009).

### Selection or Aggregation of Single Effect Sizes From Studies

Because meta-analyses are focused on simple effects, only bivariate comparisons or correlations can be used (Borenstein et al., 2009); thus, bivariate associations such as standardized differences of means or correlations were selected over adjusted coefficients from multivariate models. When both longitudinal and cross-sectional results were available, the appropriate longitudinal effect sizes were used in the meta-analyses in order to obtain the most methodologically robust effect size. If a study reported multiple effect sizes for the same outcome, such as when bivariate associations were reported for subgroups but not the whole sample, the weighted average of these subgroup effect sizes was used as the effect size for that study for that outcome. We allowed studies that reported effect sizes for more than one of our target outcomes to contribute to each appropriate meta-analysis; however, each study (or dataset, in the case of multiple articles from one dataset) was permitted to contribute only one effect size to each analysis for a specific outcome, so that a single individual was only counted once in any given meta-analysis for a specific outcome.

### Coding of Study-Level Moderators

Seven study characteristics were coded for each study to be used in moderator analyses: (a) study design (experimental, longitudinal, cross-sectional, or retrospective); (b) measure of spanking (observation, parent report, child report, child retrospective, or both parent and child reports); (c) index of spanking (when used [either observed or in an experiment], frequency, frequency and severity, ever in time period, or ever in life); (d) independence of the raters of spanking and the child outcome (same rater or different raters); (e) time period in which spanking was administered (observed, last week, last month, last year, ever, hypothetical, specific time period, or not specified); (f) the country in which the study was conducted (U.S. or other than U.S.); and (g) the age range of children at the time of spanking



(less than 2-years-old, 2- to 5-years-old, 6- to 10-years-old, and 11- to 15-years-old). The authors independently coded these characteristics for each study. Any discrepancies were resolved through discussion.

### Meta-Analytic Procedure

Once all study effect sizes had been converted to the metric of Cohen's  $d$ , effect sizes were combined in a meta-analysis. Each study was entered into the model, weighted by its precision ( $1/se_d$ ), and combined into a weighted average of effect sizes for the respective outcome domain. The meta-analyses reported in this paper utilized the random effects model (Borenstein et al., 2009; DerSimonian & Laird, 1986) using the Stata command *metan* (Bradburn, Deeks, & Altman, 2009). The random effects model for meta-analysis does not assume that there is a single underlying effect size of the studies being analyzed and rather allows effect sizes to differ across studies to account for the fact that study samples differ by characteristics such as age, gender, race, ethnicity and nationality. The random effects model calculates the mean effect size, an estimate of statistical significance, and a measure of the heterogeneity of effect sizes in terms of their variation around the estimated mean effect size. We conducted a separate meta-analysis for each child outcome as well as an overall meta-analysis for all of the studies together.

## Results

### Main Meta-Analyses

A total of 111 unique effect sizes were derived from data representing 204,410 child measurement occasions; these studies included data from a total of 160,927 unique children. The study-level effect sizes, confidence intervals, and sample sizes are listed in Table 1. For between-subjects designs, the subsample sizes for the subgroup that were spanked and the subgroup that was not spanked are presented, whereas for within-subjects designs a single sample size is presented. As a means of graphically representing the effect sizes, this table also includes bar graphs of the effect sizes and their corresponding confidence intervals both for the individual studies and for the random effects mean effect size for each outcome category. For the purposes of comparison and aggregation across meta-analyses, all of the study-level effect sizes were coded so that larger positive values corresponded to more detrimental child outcomes. This meant that for studies in which the outcome variable was a beneficial outcome (e.g., conscience), the effect sizes were recoded so that higher values reflected adverse outcomes rather than beneficial outcomes (e.g., low conscience).

As the effect sizes and bar graphs in Table 1 indicate, the findings across studies were highly consistent. Of the 111 individual effect sizes, 102 were in the direction of a detrimental outcome with 78 of these statistically significant. In contrast, nine of the effect sizes were in the direction of a beneficial outcome but only one (Tennant, Detels, & Clark, 1975) was statistically different from zero. Thus, among the 79 statistically significant effect sizes, 99% indicated an association between spanking and a detrimental child outcome.

Table 2 summarizes the mean weighted effect sizes and confidence intervals for each outcome along with a  $Z$  test for significant difference from zero and an  $I^2$  statistic that

estimates the amount of variation in the mean weight effect size that was attributable to underlying study heterogeneity. Spanking was significantly associated with 13 of the 17 outcomes examined. In each case, spanking was associated with a greater likelihood of detrimental child outcomes. In childhood, parental use of spanking was associated with low moral internalization, aggression, antisocial behavior, externalizing behavior problems, internalizing behavior problems, mental health problems, negative parent–child relationships, impaired cognitive ability, low self-esteem, and risk of physical abuse from parents. In adulthood, prior experiences of parental use of spanking were significantly associated with adult antisocial behavior, adult mental health problems, and with positive attitudes about spanking. The remaining four meta-analyses were not significantly different from zero. The 13 statistically significant mean effect sizes ranged in size from .15 to .64. The overall mean weighted effect size across all of the 111 study-level effect sizes was  $d = .33$ , with a 95% confidence interval of .29 to .38; this mean effect was statistically different from zero,  $Z = 14.84$ ,  $p < .001$ .

### Moderator Analyses Comparing Spanking With Physical Abuse

To address the concern that the findings of negative outcomes associated with spanking in past research were a result of the confounding of spanking with overly harsh or potentially abusive methods, we identified seven studies that reported bivariate associations for both spanking and physical abuse. The latter was defined variously as “hitting with fist or object, beating up, kicking, or biting” (Bugental, Martorell, & Barraza, 2003), “beaten to injury” (Lau, Chan, Lam, Choi, & Lai, 2003), “been injured from a beating” (Lau et al., 2005), “frequent or severe physical punishment” (Fergusson, Boden, & Horwood, 2008), “use of a weapon, punching, or kicking” (Lynch et al., 2006), “severe physical assault” (Miller-Perrin, Perrin, & Kocur, 2009), and “physical abuse leading to bruising” (Schweitzer, Zavar, Pavlicova, & Fallon, 2011). Each of these studies employed a within-subjects design; in each case, the same respondent (either a parent or the adult child recalling the behavior) reported both how often the parent used spanking and, in a separate question, how often the parent used abusive methods of discipline. Two of the studies contributed more than one effect size, yielding a total of 10 pairs of effect sizes for spanking and physical abuse. The effect sizes are presented in Table 3. In three cases, the effect size for spanking was larger than that for physical abuse. The weighted mean effect size for spanking was  $d = .25$ , while for physical abuse it was  $d = .38$ . Both were significantly different from zero and both were positive in sign, indicating that both spanking and physical abuse were associated with greater levels of detrimental child outcomes. The magnitude of the mean effect size for spanking was 65% of the magnitude of the mean effect size for physical abuse.

### Moderator Analyses by Study Characteristics

We examined whether study-level effect sizes varied across seven study-level characteristics using meta-regression to calculate average effect sizes by study subgroup (Borenstein et al., 2009; Harbord & Higgins, 2009). The results from these moderator analyses are presented in Table 4. All of the comparisons were nonsignificant, indicating that the effect sizes did not vary by study characteristic. The finding that the average effect size for longitudinal studies was the same as that for cross-sectional studies ( $\beta = -.07$ ,  $ns$ ) is important in light of the criticism that previous meta-analyses were overly influenced by cross-sectional studies

(Baumrind et al., 2002); for the studies examine here, no evidence was found that the magnitude or direction of effect sizes was smaller in longitudinal than cross-sectional studies. Average effect sizes also did not significantly vary based on how spanking was measured, how it was indexed, whether the raters of spanking and outcome were independent, the time period over which spanking was measured, the country of the study, or the age group of the children studied.

### Tests for Publication Bias

One potential threat to the validity of meta-analyses is what is referred to as publication bias, or commonly “the file drawer effect.” Namely, how likely is it that there are many studies with contradictory findings that were not published that would undermine the conclusions of the meta-analysis? For each meta-analysis, we conducted the publication bias test developed by Egger, Davey Smith, Schneider, and Minder (1997) and implemented in Stata (Harbord, Harris, & Sterne, 2009; Steichen, 1998). None of the tests were significant, indicating that the risk of publication bias for any of our mean effect sizes is small.

### Discussion

The goal of this article was to address two major concerns about past meta-analyses of the association of parents’ use of spanking and a range of child outcomes. We will discuss each in turn, but begin with a summary of the overall findings from this new set of meta-analyses.

#### Spanking Is Associated With Higher Risk for Detrimental Outcomes

Thirteen of the 17 child outcomes examined were found to be significantly associated with parents’ use of spanking. Among the outcomes in childhood, spanking was associated with more aggression, more antisocial behavior, more externalizing problems, more internalizing problems, more mental health problems, and more negative relationships with parents. Spanking was also significantly associated with lower moral internalization, lower cognitive ability, and lower self-esteem. The largest effect size was for physical abuse; the more children are spanked, the greater the risk that they will be physically abused by their parents.

Three of the four adult outcomes were significantly associated with a history of spanking from parents: adult antisocial behavior, adult mental health problems, and adult support for physical punishment. While these findings suggest that there may be lasting impacts of spanking that reach into adulthood, they are only suggestive, as adults who engage in antisocial behavior or who are experiencing mental health problems may focus on negative memories of their childhoods and report more spanking than they actually received. The finding that a history of received spanking is linked with more support for spanking of children as an adult may be an example of intergenerational transmission of spanking, or it may be an example of adults selectively remembering their past as a way of rationalizing their current beliefs. Only one of the 20 effect sizes for outcomes in adulthood was from a prospective longitudinal study (McCord, 1991). More longitudinal studies are needed to confirm the direction of effect.

An important observation about the meta-analyses is that the individual studies are highly consistent: 71% of all of the effect sizes, and 99% of the significant effect sizes, indicated a

significant association between parental spanking and detrimental child outcomes. The only study that found a significant association with a beneficial outcome, Tennant, Detels, and Clark (1975), had a unique sample (U.S. Army soldiers stationed in West Germany in 1971 and 1972, most of whom were White [77%]), and found that soldiers who recalled being spanked were less likely to report using amphetamines or opiates. While this is clearly a beneficial outcome, the uniqueness of the sample limits the generalizability of this finding and may explain why this study is an outlier, as it was in the Larzelere and Kuhn (2005) meta-analyses.

Three outcomes in childhood were not significantly associated with spanking in the meta-analyses: immediate defiance, child alcohol or substance abuse, and low-self regulation. The failure to reach significance for immediate defiance appears to result from the small  $n$  (150), while for child alcohol or substance abuse and low self-regulation the cause appears to be heterogeneity in effect sizes. The finding that spanking was not linked with immediate defiance was unexpected given the opposite findings in the Gershoff (2002) meta-analyses. The disparity arose because we coded the effect sizes from the three experimental studies of compliance (Bean & Roberts, 1981; Day & Roberts, 1983; Roberts & Powers, 1990) differently. Unlike Gershoff (2002) in which effect sizes for each study were calculated by subtracting the rate of compliance among children in the spanking condition from the rate of compliance among children in the comparison condition, we calculated the within condition difference in pre- and postintervention compliance rates for the spank and no-spank groups and then subtracted these two difference scores from each other. Because there were baseline differences between the treatment and control groups in each study, our effect sizes thus captured the extent to which spanking was associated with decreases in immediate defiance over baseline. From these five studies, it appears that children are as likely to defy their parents when they spank as comply with them, but future research will be needed to substantiate this conclusion.

Taken together, these meta-analyses support the conclusion that parents' use of spanking is associated with detrimental child outcomes. As most of the included studies were correlational or retrospective (72%), causal links between spanking and child outcomes cannot be established by these meta-analyses. That said, given that a correlational association is a necessary condition for a causal relationship (Shadish et al., 2001), we can conclude that the data are consistent with a conclusion that spanking is associated with undesirable outcomes.

### **Spanking Alone Is Associated With Detrimental Outcomes, but in Similar Ways to Physical Abuse**

Our first research question was whether spanking would be associated with detrimental child outcomes when studies relying on harsh and potentially abusive methods were removed. The answer to this question is: Yes, it is. As noted above, all of the mean effect sizes indicated that even when a restricted definition of spanking is used, spanking is associated with detrimental child outcomes. The mean effect size across all studies,  $d = .33$ , was smaller than the overall mean effect size reported by Gershoff (2002),  $d = .40$ , but still statistically significant.

In order to better compare the findings for spanking with those for abuse, we identified seven studies that reported effect sizes for spanking and for physical abuse for the same child outcomes and conducted a meta-analysis of this set of studies. Spanking in these studies was significantly associated with detrimental outcomes with an effect size  $d = .25$ , while the mean effect size for physical abuse for these studies was significant and  $d = .38$  (see Table 3). However, the mean effect size for all studies from Table 2,  $d = .33$ , is closer to the mean effect size for physical abuse in Table 3 than it is to the mean effect size for spanking from these 10 select effect sizes, indicating that spanking and physical abuse have relations with child outcomes that are similar in magnitude and identical in direction.

That spanking and physical abuse may have similar associations with child outcomes is consistent with previous literature. Both behaviors involve parents intentionally hitting (and hurting) children, albeit to different degrees (Gershoff, 2013), and most instances of substantiated physical abuse (75%), like all instances of spanking, begin as responses to children's misbehavior (Durrant et al., 2006). In addition, many researchers have argued that spanking and physical abuse are on a continuum of violence against children, and that spanking can escalate into physical abuse (Straus, 2001), an argument supported by our finding that spanking was significantly associated with physical abuse (Table 2;  $d = .64$ ). Clearly not all parents who spank their children also administer more severe punishment; as with all of the meta-analyses presented, the association only indicates that milder and more severe corporal punishment are linked, and that the former may increase the risk that children will also be physically abused.

### Spanking Effect Sizes Are Similar Across Study Characteristics

A major concern raised about the spanking literature in general and previous meta-analyses in particular is that their reliance on cross-sectional designs may mask what are truly child-elicitation effects (Baumrind et al., 2002; Ferguson, 2013; Larzelere et al., 2004). In other words, associations between spanking and problematic behavior may reflect the fact that difficult children elicit more spanking from parents, not that spanking causes the problematic behavior in the first place (Baumrind et al., 2002; Larzelere et al., 2004). Longitudinal or experimental designs are needed to isolate the direction of effect, and several were available for inclusion in the meta-regression moderation analyses. While it was indeed true that the majority of studies (70%) were cross-sectional or retrospective in nature, the effect sizes for the longitudinal and experimental studies were not significantly different from the effect sizes for the cross-sectional studies (see Table 4). This finding indicates that methodologically stronger studies did not find significantly smaller effect sizes than methodologically weaker studies, lending more confidence to the findings from the main meta-analyses that include both. The mean effect size for spanking also did not vary by any of the other six study characteristic moderators. The association between spanking and detrimental child outcomes did not depend on how spanking was assessed, who reported the spanking, the country where the study was conducted, or what age children were the focus of the study. Across all categories, methodologically stronger study designs identified the same risk for negative outcomes as did weaker study designs, suggesting that the associations between spanking and child outcomes are robust to study design.

We were surprised that none of the moderators was significant given that most of the  $I^2$  values in Table 2 indicate high levels of heterogeneity. There is little guidance in the literature about how to interpret significant  $I^2$  values when paired with significant and consistent mean effect sizes. We suspect that the  $I^2$  tests are picking up on expected variability in our independent variable. Unlike clinical trials, for which the  $I^2$  was developed, in which treatment is systematized, it is nearly impossible to manipulate the amount or frequency of exposure to spanking and thus every participant's experience with our independent variable is different. Given that 13 of the 17 mean effect sizes were significantly different from zero and that nearly three quarters (71%) of the studies yielded effect sizes in the direction of detrimental outcomes, including nearly all of the significant effect sizes, we suspect that the  $I^2$  is picking up on heterogeneity in spanking itself rather than in its associations with child outcomes.

### Limitations

The primary limitation of these meta-analyses is their inability to causally link spanking with child outcomes. This is problematic because there is selection bias in who gets spanked—children with more behavior problems elicit more discipline generally and spanking in particular (Larzelere, Kuhn, & Johnson, 2004). Cross-sectional designs do not allow the temporal ordering of spanking and child outcomes that could help rule out the selection bias explanation. As noted above, randomized experiments of spanking are difficult if not ethically impossible to conduct, and thus this shortcoming of the literature will be difficult to correct through future studies.

The main viable strategy for doing so is through the use of analytic methods which increase our confidence that the causal direction is as hypothesized. Whenever such strategies have been employed, they have confirmed that spanking is associated with detriments to children. A series of cross-lagged studies (Berlin et al., 2009; Gershoff, Lansford, Sexton, Davis-Kean, & Sameroff, 2012; McLeod, Kruttschnitt, & Dornfeld, 1994; Sheehan & Watson, 2008) has demonstrated that spanking predicts changes in children's behavior, over and above their initial levels and the child effect of early problem behavior on later spanking.

Another statistical method that has been employed to strengthen conclusions is fixed effects regressions, which control for time-invariant unobserved characteristics that may account for observed relationships between spanking and child outcomes, such as children's initial levels of problem behavior. Using fixed effects models with the National Longitudinal Study of Youth (NLSY), Grogan-Kaylor (2005) found that increases in parents' use of spanking predicted increases in children's externalizing behaviors over time.

A third method is to establish spanking as a significant mediator of treatment effects on children for interventions that include a focus on reducing parents' use of spanking. In one example, an evaluation of the Incredible Years intervention for young children with behavior problems (Beauchaine et al., 2005) found that treatment effects on a reduction in conduct problems were significantly mediated through a reduction in parents' use of spanking. Similarly, analysis of data from a national randomized controlled trial of the federal Head Start program for low income children found that parents in the program significantly reduced their spanking, which was in turn linked with decreases in child aggression over

time (Gershoff, Ansari, Purtell, & Sexton, 2016). More studies capitalizing on experimental designs are needed.

By looking at change over time and by accounting for potential alternative explanations through statistical methods or by capitalizing on data from experimental designs, such studies support the conclusion that there is a significant association of parents' use of spanking with later child outcomes, over and above children's initial behavior and child elicitation effects. None of these studies using advanced statistical methods found evidence that the pathway is entirely one of selection or child elicitation, or that spanking predicts improvements in children's behavior over time, as critics of the literature on spanking have contended (Baumrind et al., 2002; Ferguson, 2013; Kazdin & Benjet, 2003; Larzelere et al., 2004). Rather, these studies with strong designs provide more, not less, support for a potential causal link between spanking and detrimental child outcomes.

## Conclusion

Spanking children to correct misbehavior is a widespread practice, yet one shrouded in debate about its effectiveness and even its appropriateness. The meta-analyses presented here found no evidence that spanking is associated with improved child behavior and rather found spanking to be associated with increased risk of 13 detrimental outcomes. These analyses did not find any support for the contentions that spanking is only associated with detrimental outcomes when it is combined with abusive methods or that spanking is only associated with such outcomes in methodologically weak studies. Across study designs, countries, and age groups, spanking has been linked with detrimental outcomes for children, a fact supported by several key methodologically strong studies that isolate the ability of spanking to predict child outcomes over time. Although the magnitude of the observed associations may be small, when extrapolated to the population in which 80% of children are being spanked, such small effects can translate into large societal impacts. Parents who use spanking, practitioners who recommend it, and policymakers who allow it might reconsider doing so given that there is no evidence that spanking does any good for children and all evidence points to the risk of it doing harm.

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## Appendix

### Appendix

Number and Percent of Studies Excluded from the Meta-Analyses by Exclusion Code

Reason for exclusion from meta-analyses	Number of studies excluded	Percent
Spanking not linked with child outcomes (e.g., prevalence only).	238	16
Not an empirical article (e.g., a literature review).	221	15
Definition of physical punishment included harsh methods of physical punishment beyond spanking, slapping, or hitting.	194	13

Reason for exclusion from meta-analyses	Number of studies excluded	Percent
Spanking was not measured in the study.	171	11
Study was an unpublished dissertation.	104	7
Article was not relevant.	85	6
Attitudes toward, and not use of, physical punishment was assessed.	82	5
Study was of physical punishment in schools or other institutions.	73	5
Study did not include a bivariate association between spanking and the child outcome.	61	4
Study was of an intervention to reduce physical punishment.	47	3
Available statistics were unclear, insufficient, or inappropriate for the meta-analyses.	46	3
Spanking was combined with yelling or some form of psychological aggression.	44	3
Study was not available in English.	32	2
Spanking was combined with other types of discipline.	30	2
Study was published as a book chapter or conference presentation.	23	2
Study used same dataset as another study in the meta-analysis.	23	2
Dependent variable did not fit into other outcome categories.	11	1
Spanking was of animals, not children.	5	<1
Article was unavailable through interlibrary loan.	3	<1
Spanking measure included threats of spanking.	3	<1
Physical punishment measure was nontraditional (i.e., aversive noise; washing mouth out with soap).	2	<1
Study involved a special population of children (chromosomal abnormality).	1	<1
Total number of excluded studies	1,499	100%

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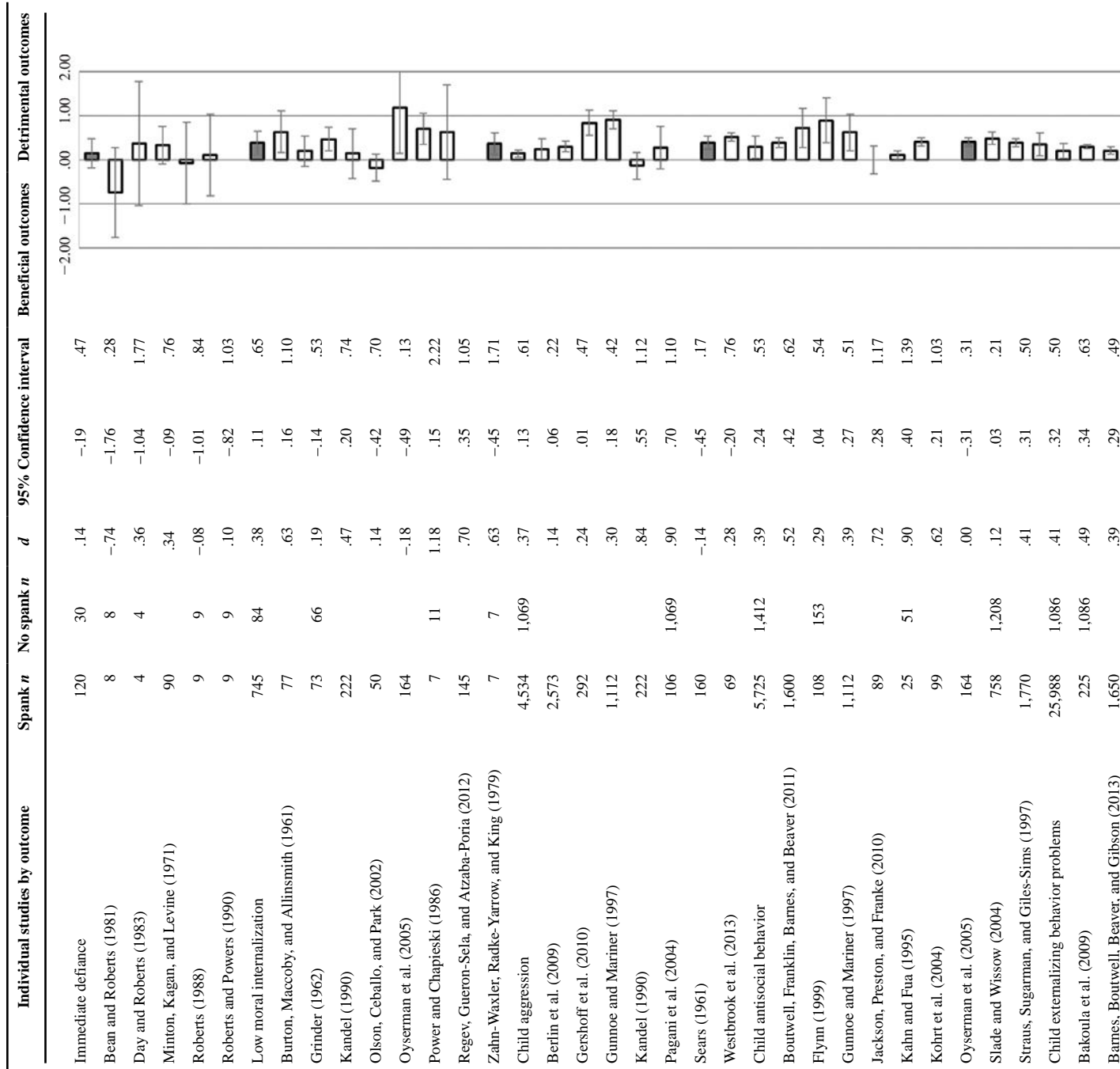
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**Table 1**

Study-Level Effect Sizes for Spanking by Child Outcome





Individual studies by outcome	Spank <i>n</i>	No spank <i>n</i>	<i>d</i>	95% Confidence interval	Beneficial outcomes	Detrimental outcomes
Deb and Adak (2006)	34	66	.88	.45	1.31	
Durrant (1993)	463	63	.39	.13	.66	
Graziano et al. (1992) U.S. sample	95		.45	.04	.87	
Graziano et al. (1992) India sample	160		.20	-.11	.51	
Hemenway, Solnick, and Carter (1994)	264	48	.12	-.18	.43	
Overall	89,638	114,772	.33	.29	.38	

*Note.* "Spank *n*" refers to the subsample in between-subjects designs that reported spanking, or to the entire sample in within-subjects designs. "No spank *n*" refers to the subsample in between-subjects designs that did not report spanking.

**Table 2**

Summary of Spanking Meta-Analyses by Outcome

Detrimental child outcome	<i>K</i>	Spank <i>n</i>	No Spank <i>n</i>	<i>d</i>	95% CI	<i>Z</i>	<i>I</i> <sup>2</sup>
Immediate defiance	5	120	30	.14	-.19 .47	.85	.80%
Low moral internalization	8	745	84	<b>.38</b>	.11 .65	2.76***	67.40%
Child aggression	7	4,534	1,069	<b>.37</b>	.13 .61	3.07***	91.40%
Child antisocial behavior	9	5,725	1,086	<b>.39</b>	.24 .53	5.28***	84.50%
Child externalizing behavior problems	14	25,988	1,086	<b>.41</b>	.32 .50	9.19***	88.40%
Child internalizing behavior problems	8	12,413	3,486	<b>.24</b>	.13 .35	4.36***	88.50%
Child mental health problems	10	5,122	1,313	<b>.53</b>	.42 .64	5.17***	76.00%
Child alcohol or substance abuse	3	6,621	90,359	.09	-.11 .29	.90	91.30%
Negative parent-child relationship	5	755	0	<b>.51</b>	.36 .66	6.76***	.00%
Impaired cognitive ability	8	8,358	11	<b>.17</b>	.01 .32	2.13*	84.30%
Low self-esteem	3	766	990	<b>.15</b>	.04 .26	2.76**	100%
Low self-regulation	3	2,525	0	.30	-.07 .67	1.58	94.30%
Victim of physical abuse	8	3,334	996	<b>.64</b>	.39 1.74	4.07***	93.30%
Adult antisocial behavior	3	985	4,206	<b>.36</b>	.06 .65	2.35**	92.00%
Adult mental health problems	8	1,855	4,707	<b>.24</b>	.09 .40	3.05**	73.20%
Adult alcohol or substance abuse	4	2,596	4,796	.13	-.08 .35	1.21	91.90%
Adult support for physical punishment	5	1,016	177	<b>.38</b>	.15 .61	3.28***	55.50%
Overall effect size	111	89,638	114,722	<b>.33</b>	.29 .38	14.84***	88.80%

Note. *K* = number of effect sizes in the meta-analysis; *d* = mean weighted effect size; *Z* = significance test that *d* differs from zero; *I*<sup>2</sup> = the variation in the mean effect size attributable to heterogeneity. Bolded effect sizes are significantly different from zero.

\* *p* < .05.

\*\* *p* < .01.

\*\*\* *p* < .001.

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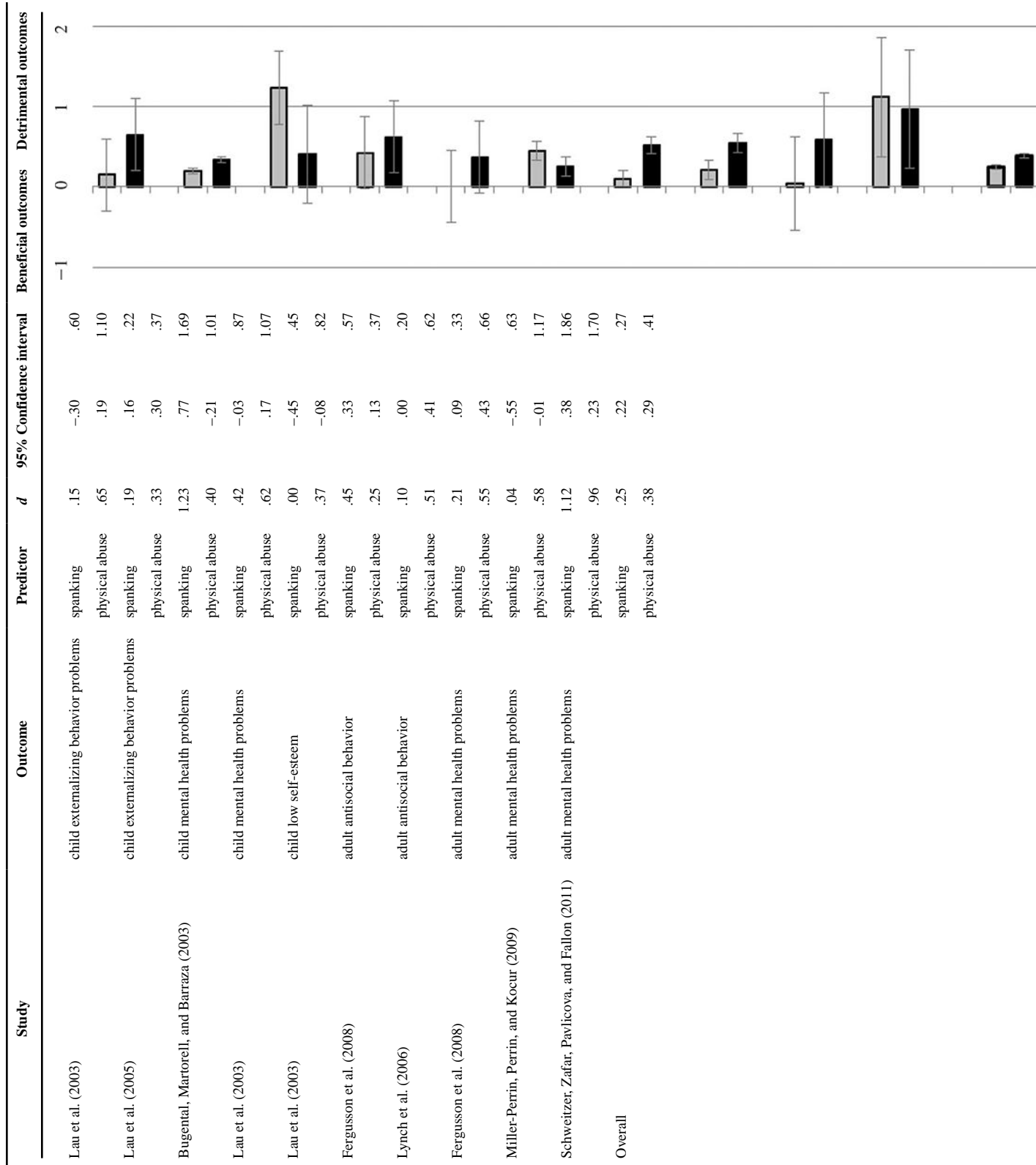
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**Table 3**

Effect Sizes for Studies That Reported Effect Sizes Separately for Spanking and Physical Abuse



**Table 4**

Effect Sizes Moderated by Study Characteristics

Moderator	N in each subcategory	$\beta$ for difference from referent	Significant differences among subgroups
Study design (referent = cross-sectional)	61		
Longitudinal	23	-.07	
Retrospective	23	-.03	
Experimental	4	-.50	
Measure (referent = parent report)	65		
Observation	6	-.10	
Child report	11	.03	
Child retrospective	26	-.03	
Both parent and child	3	-.10	
Index of spanking (referent = frequency)	79		
Frequency and severity	8	-.11	
Ever in life	15	-.06	
When used	5	-.29	
Ever in time period	4	-.18	
Raters of spanking and outcome (referent = same rater)	42		
Independent raters	67	-.04	
Time period (referent = not specified)	51		
Observation	6	-.10	
Last week	22	.01	
Last month	3	-.26	
Last year	8	.16	
Ever	8	-.17	
Hypothetical	1	-.40	
Specific time period	12	-.04	
Country (referent = U.S.)	77		
Other than U.S.	34	.04	
Age range of children at time of spanking (referent = 2- to 5-years-old)	36		
Less than 2-years-old	15	.17	

Moderator	N in each subcategory	$\beta$ for difference from referent	Significant differences among subgroups
6- to 10-years-old	30	.05	
11- to 15-years-old	27	.01	

Note.  $\beta$ s in third column represent difference from the  $\beta$  for the referent category. None were significantly different from the  $\beta$  for the referent category.