

# Interleukin-6, C-reactive protein, fibrinogen, and risk of recurrence after ischaemic stroke: Systematic review and meta-analysis

European Stroke Journal  
2021, Vol. 6(1) 62–71  
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DOI: 10.1177/2396987320984003  
journals.sagepub.com/home/eso



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## Abstract

**Background:** Recent randomised trials showed benefit for anti-inflammatory therapies in coronary disease but excluded stroke. The prognostic value of blood inflammatory markers after stroke is uncertain and guidelines do not recommend their routine measurement for risk stratification.

**Methods:** We performed a systematic review and meta-analysis of studies investigating the association of C-reactive protein (CRP), interleukin-6 (IL-6) and fibrinogen and risk of recurrent stroke or major vascular events (MVEs). We searched EMBASE and Ovid Medline until 10/1/19. Random-effects meta-analysis was performed for studies reporting comparable effect measures.

**Results:** Of 2,515 reports identified, 39 met eligibility criteria (IL-6,  $n = 10$ ; CRP,  $n = 33$ ; fibrinogen,  $n = 16$ ). An association with recurrent stroke was reported in 12/26 studies (CRP), 2/11 (fibrinogen) and 3/6 (IL-6). On random-effects meta-analysis of comparable studies, CRP was associated with an increased risk of recurrent stroke [pooled hazard ratio (HR) per 1 standard-deviation (SD) increase in  $\log_e$ -CRP (1.14, 95% CI 1.06–1.22,  $p < 0.01$ )] and MVEs (pooled HR 1.21, CI 1.10–1.34,  $p < 0.01$ ). Fibrinogen was also associated with recurrent stroke (HR 1.26, CI 1.07–1.47,  $p < 0.01$ ) and MVEs (HR 1.31, 95% CI 1.15–1.49,  $p < 0.01$ ). Trends were identified for IL-6 for recurrent stroke (HR per 1-SD increase 1.17, CI 0.97–1.41,  $p = 0.10$ ) and MVEs (HR 1.22, CI 0.96–1.55,  $p = 0.10$ ).

**Conclusion:** Despite evidence suggesting an association between inflammatory markers and post-stroke vascular recurrence, substantial methodological heterogeneity was apparent between studies. Individual-patient pooled analysis and standardisation of methods are needed to determine the prognostic role of blood inflammatory markers and to improve patient selection for randomised trials of inflammatory therapies.

## Keywords

Ischaemic stroke, transient ischaemic attack, recurrence, interleukin-6, C-reactive protein, fibrinogen, prognosis, inflammation

Date received: 20 July 2020; accepted: 6 December 2020

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## Introduction

Despite the widespread use of modern secondary prevention therapy, there is a significant residual vascular risk in ischaemic stroke survivors, with rates of recurrent stroke and MVEs of 25–30% at 5 years.<sup>1</sup> There is an urgent need for new therapeutic targets to reduce this residual risk after stroke.

Inflammation has a central role in atherosclerosis pathogenesis, stroke, and coronary events.<sup>2–4</sup> Plaque inflammation imaged by positron emission tomography (PET) predicts early recurrent stroke in patients with symptomatic carotid stenosis.<sup>5</sup> CRP and fibrinogen are associated with increased risk of first stroke and coronary events in apparently-healthy individuals.<sup>6,7</sup> Mendelian randomisation and genetic epidemiological studies indicate increased risk of first stroke associated with higher expression of monocyte chemoattractant protein-1 (MCP-1) and enhanced IL-6 receptor signalling.<sup>8,9</sup> In randomised trials (RCTs) of patients with coronary disease, the risk of major vascular events was reduced in patients treated with canakinumab (an interleukin-1 $\beta$  inhibitor) and colchicine (a tubulin inhibitor with pleiotropic anti-inflammatory effects).<sup>10–12</sup>

Despite the growing evidence linking inflammation to vascular risk, the prognostic role of inflammatory markers after stroke remains unclear. Studies which have investigated this question have reported contradictory findings and current guidelines do not recommend routine measurement of inflammatory markers for risk prediction after stroke. Improved data are needed to guide clinical practice and improve selection of patients for clinical trials of anti-inflammatory therapy. We performed a systematic review of studies investigating the association between IL-6, CRP, and fibrinogen, measured after ischaemic stroke or transient ischaemic attack (TIA), and the risk of recurrent stroke and major vascular events.

## Patients and methods

### *Systematic review protocol and manuscript preparation*

The systematic review protocol was published in advance on PROSPERO (registration number: CRD42018116190) and the manuscript was written according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) guidelines.<sup>13</sup>

### *Search strategy and study eligibility*

Eligible studies: 1) Had 50 patients or more with ischaemic stroke or TIA; 2) Blood samples analysed

for CRP and/or IL-6 and/or fibrinogen, measured after the index event; 3) Reported associations with recurrent stroke (ischaemic or haemorrhagic stroke) or MVEs (defined as any combination of myocardial infarction, unstable angina, coronary revascularisation, or vascular death); 4) Included cohort studies, case-control studies, or observational data taken from RCTs. Key study exclusion criteria were animal or in vitro studies, patients <18 years, non-stroke participants, biomarkers measured prior to event, and other study designs (crossover studies, cross-sectional studies, editorials, case reports, review articles, or duplicate studies).

Without language restrictions, we searched Ovid Medline (1/1/1946–10/1/2019) and Embase (1/1/1970–10/1/2019) using a combination of medical subject headings and free text search terms (web-supplement). Abstract screening and full text review were performed independently by 2 reviewers (JM, EO) with disagreements regarding eligibility resolved by consensus with the participation of a third reviewer (PK).

### *Assessment of study quality*

The Quality In Prognosis Studies (QUIPS) tool was used to assess the risk of bias of included studies and was adjudicated by 2 authors (JM, EO).<sup>14</sup> QUIPS evaluates six domains of potential bias in prognostic research: (i) study participation; (ii) study attrition; (iii) prognostic factor measurement; (iv) confounder measurement; (v) outcome measurement; (vi) analysis/reporting. Additionally we ascertained whether 9 specific measures of methodological quality were present: (i) reference to the presence of any study protocol; (ii) prior publication of a biomarker protocol; (iii) assay specified; (iv) blinded biomarker measurement; (v) blinded outcome assessment; (vi) outcome explicitly defined; (vii) screening for infection/inflammatory conditions; (viii/ix) measurement and adjustment of cardiovascular risk factors (age, smoking, diabetes mellitus, lipids/statin therapy, blood pressure, obesity).

### *Data extraction*

Details of the data extraction protocol are provided in the web-supplement. Briefly, data extraction was performed by 2 reviewers (JM, EO) using a standardized template. All measures of association between inflammatory markers and outcome were recorded (eg. odds ratio (OR), hazard ratio (HR), mean/median differences between marker levels in groups with/without the specified outcome event). Where multiple effect estimates were provided, the most adjusted measure of effect was chosen for the final analysis.

## Statistical analyses

Baseline study characteristics were summarized using mean/SD or median/interquartile range (IQR). Forest plots were used to visually summarize study effect estimates and their corresponding 95% CIs, for each biomarker and each pre-specified outcome. A pooled meta-analysis using a random effects model<sup>15</sup> was restricted to studies reporting homogenous effect measures in accordance with best-practice recommendations.<sup>16,17</sup> Between-study heterogeneity was assessed by the  $I^2$  statistic.<sup>18</sup> We investigated for small study effects and publication bias using contour-enhanced funnel plots and tested for the presence of funnel plot asymmetry using Egger's test where appropriate.<sup>19,20</sup> STATA 15.0 (StataCorp, College Station, TX) software was used for statistical analyses. A  $p$  value  $<0.05$  was considered significant.

## Results

### Search results

The original search identified 2,515 non-duplicate records. After title and abstract screening, 130 studies were selected for full text review. Thirty-nine articles met eligibility criteria (37 individual study cohorts). There were 33 reports from 30 studies for CRP,<sup>21–53</sup> 14 reports from 16 studies for fibrinogen<sup>21,36–39,42–44,47,54–58</sup> and 10 studies for IL-6 (Figure 1).<sup>21,24,37–40,43,47,48,59</sup>

### Quality assessment

Using the QUIPS tool, study attrition, study confounding and outcome assessment were the 3 most frequently identified domains for risk of bias [Supplementary Figure 1(a-c)]. The 9 studies included in the pooled analysis had a lower risk of bias across most domains when compared with studies in the narrative synthesis. Only one study was deemed a high risk of bias in 2 domains [Supplementary Table 1]. The methodological quality assessment found that most studies did not publish a study protocol, screen for infectious/inflammatory conditions, specify blinded outcome adjudication or measure/adjust for all cardiovascular risk factors [Supplementary Figure 2(a-c)].

### CRP and risk of recurrent stroke and vascular events

The search identified 30 studies for CRP with 25,363 patients (Supplementary Table 2). Where stated, CRP measurement was in the acute phase (0–14 days after the index event) in 61% of studies and in the subacute or convalescent phase in 39%. 12 studies explicitly stated that participants were screened for

infectious/inflammatory conditions. There was marked variability in the definition of CRP exposure across studies, including risk analysed per increase in standard deviation, per tertile/quartile/decile, per measured unit (mg/dL) or by arbitrary cut-points on log-transformed and untransformed scales [Supplementary Figure 3(a-b), Supplementary Table 3]. Only 6/30 studies adjusted for all pre-specified vascular risk factors.

The association of CRP with recurrent stroke was reported in 26 studies, of which 12 (46%) described a positive association. A meta-analysis was limited to 5 studies (7,284 patients) which uniformly analyzed the risk of recurrent stroke per 1-SD rise in  $\log_e$ -CRP.<sup>28,29,32,44,48</sup> Using a random-effects model, the pooled HR of recurrent stroke per 1-SD rise in  $\log_e$ -CRP was 1.14 [95% CI 1.06–1.22,  $p < 0.01$ ; Figure 2 (a)], with low heterogeneity ( $I^2 = 0\%$ ).

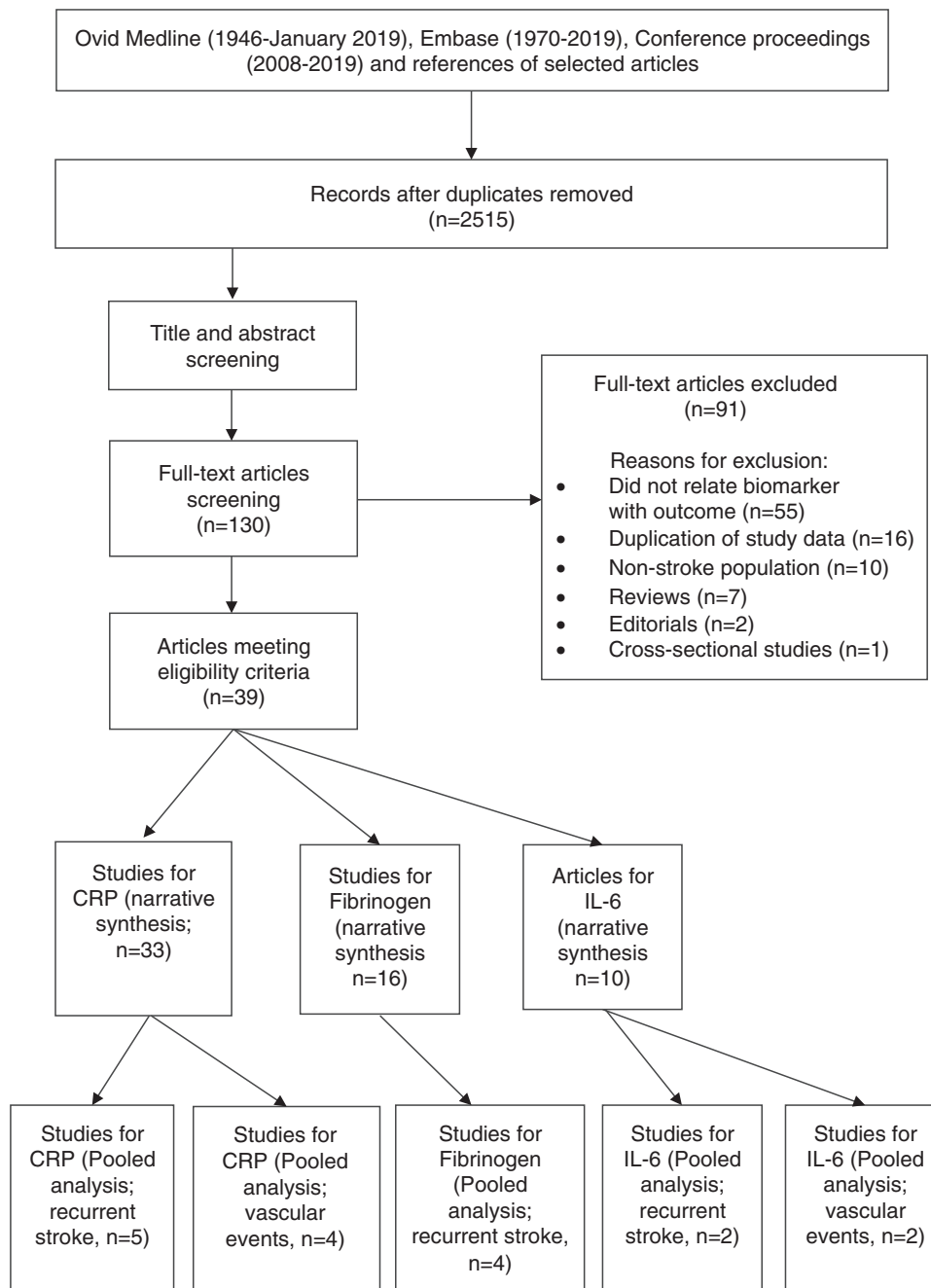
The association of CRP with recurrent MVEs was reported in 17 studies, of which 10 (59%) reported a positive association. Due to heterogeneity in CRP measures between studies, meta-analysis was confined to 4 studies (3,633 patients) that reported risk per SD rise in  $\log_e$ -CRP.<sup>28,32,45,48</sup> The HR of MVEs per 1-SD increase in  $\log_e$ -CRP was 1.21 [95% CI 1.10–1.34,  $p < 0.01$ ; Figure 2(b)], with low heterogeneity ( $I^2 = 10.5\%$ ).

### Fibrinogen and risk of recurrent stroke and vascular events

There were 16 studies for fibrinogen ( $n = 12,961$  patients, Supplementary Table 4). Thirty-eight percent of studies measured fibrinogen in the acute phase after the index event and in the subacute or convalescent phase in the remainder. Just 2 studies screened for infectious/inflammatory conditions. There was also a marked variability in the definition of fibrinogen exposure and measures of association reported across studies [Supplementary Figure 4(a-b), Supplementary Table 5]. Only 1 study fully adjusted for all vascular risk factors.

The association between fibrinogen and recurrent stroke was reported in 11 studies, of which 2 reported a positive association. A pooled analysis was restricted to 4 studies ( $n = 6,995$ ) which reported risk of supra-median versus sub-median fibrinogen.<sup>44,57</sup> The pooled HR of recurrent stroke associated with elevated fibrinogen was 1.26 (95% CI 1.07–1.47,  $p < 0.01$ ; Figure 3), without statistical heterogeneity ( $I^2 = 0\%$ ).

Five of the 10 studies investigating fibrinogen and risk of MVEs reported a positive association. A meta-analysis was possible for 3 comparable studies which reported HRs for supra- versus sub-median fibrinogen. This analysis was previously published by Rothwell et al. and reported increased risk of MVEs with



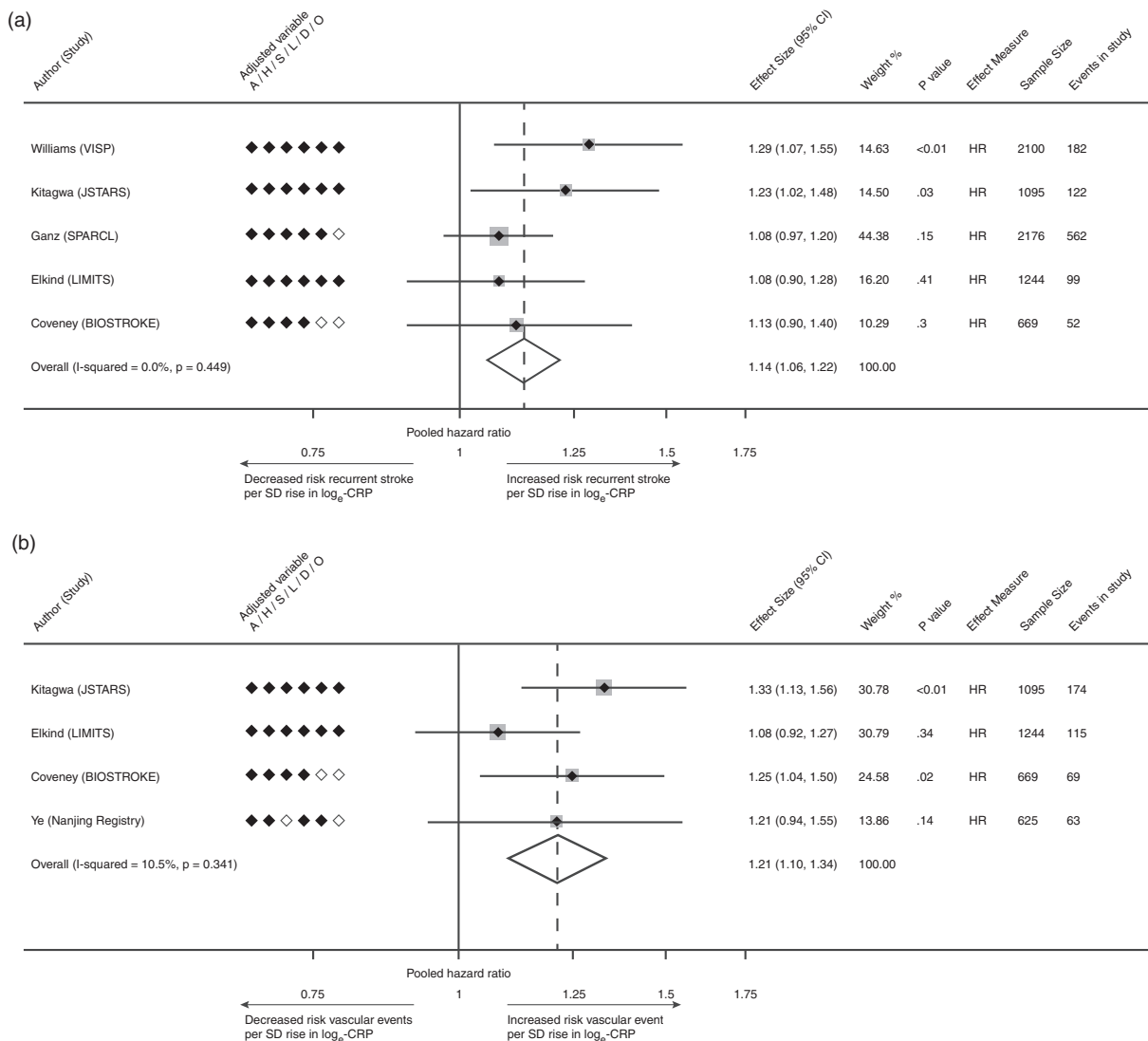
**Figure 1.** PRISMA Flowchart for the selection of eligible articles.

elevated fibrinogen (HR 1.31, 95% CI 1.15–1.49,  $p < 0.01$ ,  $I^2 = 0\%$ ).<sup>57</sup> We independently repeated this analysis and found identical results.

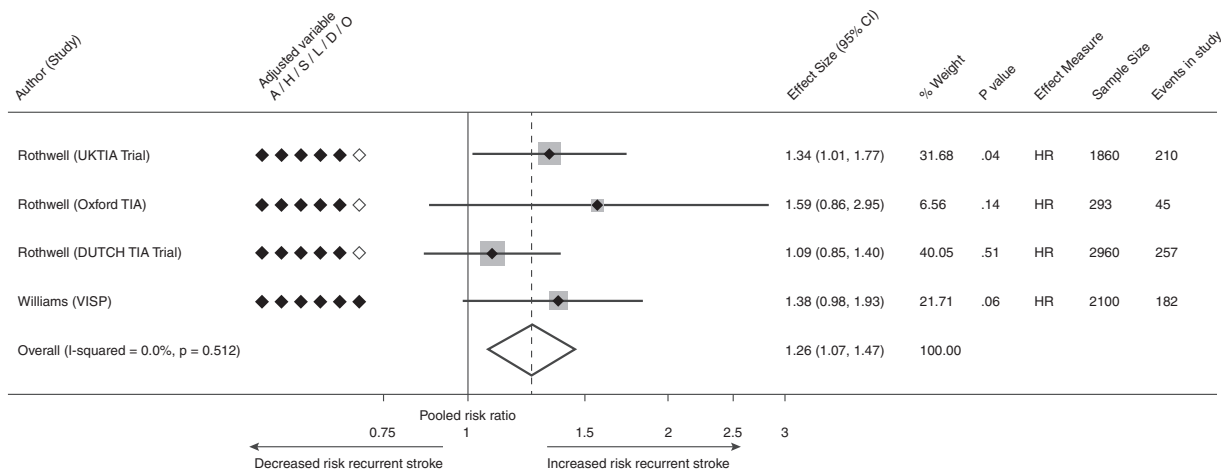
#### ***IL-6 And risk of recurrent stroke and vascular events***

Ten studies were identified for IL-6 ( $n = 7,934$ , Supplementary Table 6) with marked variability in definitions of IL-6 exposure, timing of phlebotomy, screening for infection and measures of association

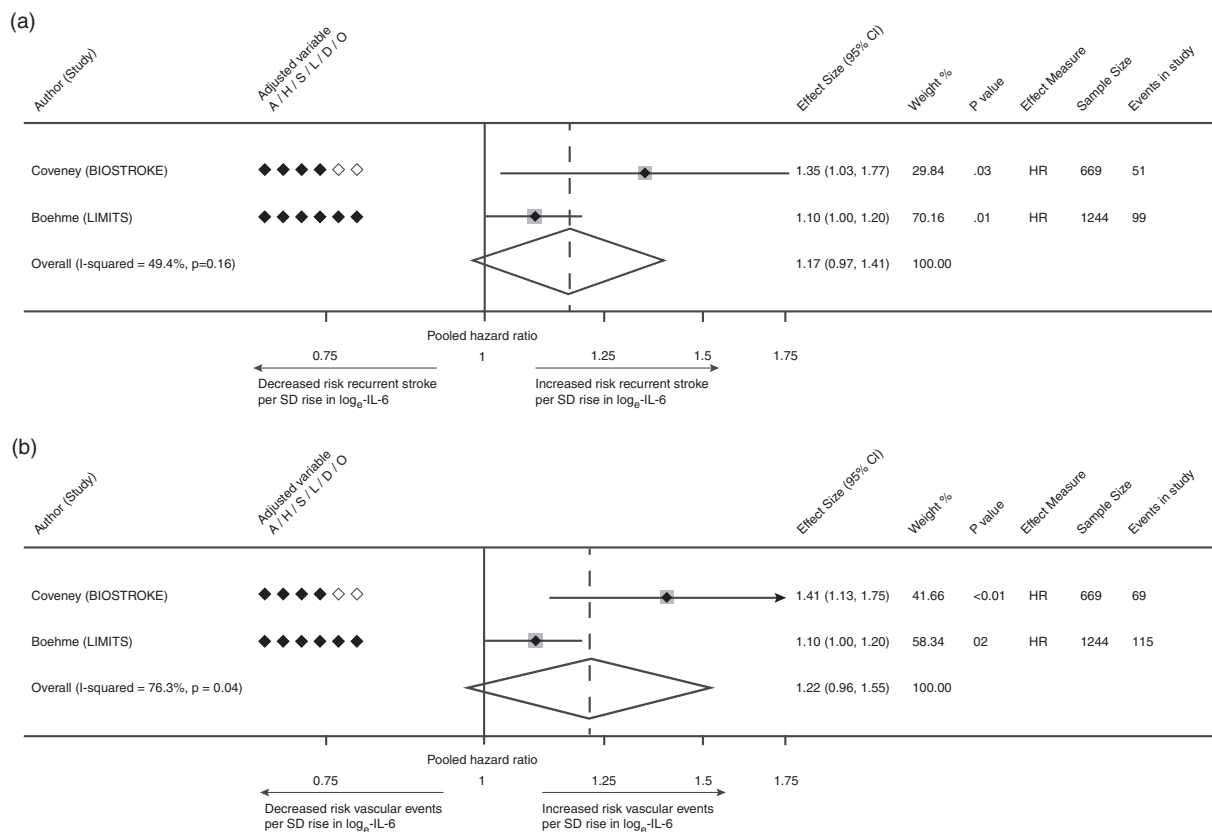
[Supplementary Figure 5(a-b), Supplementary Table 7]. Three of six studies reported a positive association with recurrent stroke. Only 1 study fully adjusted for all vascular risk factors. Due to varying methodology, only two comparable studies ( $n = 1,913$ ) were meta-analysed which analysed risk per 1-SD rise in  $\log_e$ -IL-6.<sup>48,59</sup> A non-significant statistical trend was identified for the association between IL-6 and recurrent stroke (HR per 1-SD increase 1.17 [95% CI 0.97–1.41,  $p = 0.10$ ; Figure 4(a)] with moderate heterogeneity ( $I^2 = 49.4\%$ ).



**Figure 2.** (a) Pooled analysis of studies reporting risk of recurrent stroke per 1-SD rise in log<sub>e</sub>-CRP; (b) Pooled estimate for risk of recurrent MVEs per 1-SD rise in log<sub>e</sub>-CRP. Legend: A, age; H, hypertension; S, smoking; L, hyperlipidemia; D, diabetes mellitus; O, obesity; HR, hazard ratio.



**Figure 3.** Pooled estimate for risk of recurrent stroke for supra-median compared with sub-median fibrinogen level.



**Figure 4.** (a) Pooled estimate for risk of recurrent stroke per 1-SD rise in  $\log_e$ -IL-6; (b) Pooled estimate for risk of recurrent MVEs per 1-SD rise in  $\log_e$ -IL-6.

Of 7 studies which reported the association of IL-6 with MVEs, 5 reported a positive association. Meta-analysis was again restricted to 2 comparable studies ( $n = 1,913$ ).<sup>48,59</sup> A non-significant trend was identified for the association of IL-6 with MVEs [pooled HR 1.22, 95% CI 0.96–1.55,  $p = 0.10$ ; Figure 4(b)], with considerable statistical heterogeneity ( $I^2 = 76.3\%$ ).

### Small study effects

There was evidence of funnel-plot asymmetry for outcomes of recurrent stroke (Egger's  $p < 0.01$ ) and MVEs (Egger's  $p < 0.01$ ) for CRP, indicating possible publication bias. Asymmetry was also apparent on funnel-plot inspection of fibrinogen studies, but statistical assessment was not possible due to few available studies [Supplementary Figure 6(a-d)]. Funnel-plot analysis was not possible for IL-6 due to the small number of studies.

## Discussion

Atherosclerotic inflammation has emerged as a therapeutic target for prevention of coronary events and stroke. The central inflammatory pathway in

atherosclerotic plaque development involves upstream interleukin-1 $\beta$  (IL-1 $\beta$ ) activation by the NOD-like receptor family pyrin domain-containing 3 (NLRP3) inflammasome in response to cholesterol crystal deposition and other pro-inflammatory stimuli. The pleiotropic cytokine IL-6 is expressed by T-cells and monocyte-macrophages in response to IL-1 $\beta$ , and has multiple pro-atherogenic and pro-inflammatory effects including hepatocyte production of CRP and fibrinogen.<sup>60,61</sup> To inform the design of randomised trials of anti-inflammatory therapies for secondary prevention after stroke, robust data on inflammatory biomarkers are needed.

In this context, we report the first systematic review and meta-analysis to investigate the association of IL-6, fibrinogen and CRP and vascular recurrence after stroke. Our study provides new data in several ways: First, we found important variability in methodological aspects such as inclusion of patients with severe stroke or other confounding pro-inflammatory diseases, adjustment for confounders, and definition of exposure variables. This variability limits the interpretation of the available evidence and indicates that greater standardisation should be a priority in the design and reporting of future studies.<sup>16</sup> Second, on pooled analysis of

the sub-set of comparable studies we found that each SD increase in log<sub>e</sub>-CRP was associated with a one-fifth increase in risk of future major vascular events in stroke survivors. These findings were consistent across studies, without evidence of heterogeneity. Based on individual-participant data available, 1-SD rise in log<sub>e</sub>-CRP corresponds to an approximate four-fold difference in mg/L on the original scale of CRP measurement. Although caution is needed due to the unavoidable selection of a sub-group of available studies, the results are biologically-plausible and consistent with those reported for patients with coronary disease.<sup>6</sup>

Third, similar results were found for fibrinogen, where pooled risks were increased by 26% (recurrent stroke) and 31% (all vascular events) when patients with supra- versus sub-median levels were compared. This finding is consistent with other data showing an association of fibrinogen with risk of first stroke and coronary events.<sup>7</sup> Fourth, although only 2 studies could be pooled, analysis for IL-6 also showed non-significant but consistent directions of association for both recurrent stroke and MVEs. Overall, the available data suggested a consistent pattern of increased vascular risk in stroke patients with elevated blood inflammatory markers.

Our findings add to other evidence suggesting that atherosclerotic inflammation is an independent risk factor for ischaemic vascular events. In epidemiological studies, CRP, fibrinogen and MCP-1 are associated with first stroke.<sup>6-8</sup> Genetic variants predisposing to higher MCP-1 levels are associated with increased stroke risk, whilst genetic predisposition to reduced IL-6 receptor signalling is protective against stroke and coronary events.<sup>8,9,62</sup> In PET-imaging studies, carotid plaque 18-fluorodeoxyglucose (<sup>18</sup>F-FDG) uptake predicts early recurrent stroke, independent of stenosis severity.<sup>5</sup> In randomised placebo-controlled trials, IL-1 $\beta$  inhibition with canakinumab and tubulin/NLRP3 inhibition with colchicine reduce recurrent events in patients with coronary disease.<sup>10-12</sup>

Strengths of our study are its comprehensiveness, use of recommended methods for systematic reviews of observational studies, and assessment of key inflammatory biomarkers which are most likely to have utility in randomised trials and clinical practice. Due to variability in the definition of exposure variables, we deliberately restricted pooled analysis to a sub-set of comparable studies, which allows valid interpretation of these results.

We acknowledge several limitations. We cannot exclude the possibility of selection bias in studies included for pooled analyses. Because individual patient data were unavailable, pooled risk estimates were not fully adjusted for all possible confounding variables. As with all systematic reviews, the

methodological quality of included studies and possibility of publication bias need to be considered in the interpretation of our findings. Our selection of studies for pooled analysis based on exposure variables defined as continuous measures was not pre-specified. However, we believe that this approach is appropriate, as analysis on continuous scales is recommended in prognostic research.<sup>16</sup>

## Conclusion

Our review provides new evidence suggesting that inflammation is associated with increased risk of future vascular events in stroke patients. An individual-participant data meta-analysis is needed to determine whether an association remains after adjustment for residual confounding, and to address other knowledge gaps such as optimal choice of biomarker, relationship with stroke subtypes, optimal timing of measurement, and impact of important factors such as infarct volume and stroke severity. Our findings support the rationale for randomised trials of anti-inflammatory therapies after stroke. One such trial, CONVINCE (ClinicalTrials.gov Identifier: NCT02898610) is currently investigating low-dose colchicine after ischaemic stroke or TIA.

## Data availability

Data are available on reasonable request.

## Declaration of conflicting interests

The author(s) declared no potential conflicts of interest with respect to the research, authorship, and/or publication of this article.

## Funding

The author(s) disclosed receipt of the following financial support for the research, authorship, and/or publication of this article: PJK – funded from Health Research Board Ireland and Irish Heart Foundation.

## Ethical approval

The data included in this analysis was extracted from previously published work. Therefore, no additional ethical approval or consent was required.

## Guarantor

PJK.


## Contributorship


JJM/EOR/PJK – literature search. JJM/PJK/EOR/SC/RC/LH/JM/RM/BM/TC/SM/MOD contributed to study conception, protocol design and drafting of the manuscript. BW/FH provided additional unpublished data.


## Acknowledgements

Stroke Clinical Trials Network Ireland (SCTNI) investigators.

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## Supplemental Material

Supplemental material for this article is available online.

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