

## REVIEW PAPER

# Effect of dietary salt restriction on central blood pressure: A systematic review and meta-analysis of the intervention studies

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## Abstract

Central blood pressure (cBP) is highly associated with cardiovascular risk. Although reduction of salt intake leads to lower peripheral blood pressure (BP), the studies on cBP provided inconsistent results. Therefore, we performed a systematic review and a meta-analysis of the available intervention trials of salt reduction on cBP values to reach definitive conclusions. A systematic search of the online databases available (up to December 2018) was conducted including the intervention trials that reported non-invasively assessed cBP changes after two different salt intake regimens. For each study, the mean difference and 95% confidence intervals were pooled using a random-effect model. Sensitivity, heterogeneity, publication bias, subgroup, and meta-regression analyses were performed. Fourteen studies met the pre-defined inclusion criteria and provided 17 cohorts with 457 participants with 1-13 weeks of intervention time. In the pooled analysis, salt restriction was associated with a significant reduction in augmentation index (9.3%) as well as central systolic BP and central pulse pressure. There was a significant heterogeneity among studies ( $I^2 = 70\%$ ), but no evidence of publication bias. Peripheral BP changes seemed to partially interfere on the relationship between salt restriction and cBP. The results of this meta-analysis indicate that dietary salt restriction reduces cBP. This effect seems to be, at least in part, independent of the changes in peripheral BP.

## 1 | INTRODUCTION

Central blood pressure (cBP) indices and its derivatives are independent predictors of organ damage,<sup>1,2</sup> cardiovascular events, and all-cause mortality,<sup>3</sup> and several works indicated that they are more strongly related to cardiovascular risk than peripheral blood pressure (BP).<sup>4-6</sup> cBP can be non-invasively assessed by several devices using methods that evaluate central aortic pressure waveform.<sup>7</sup> This wave is composed by a forward traveling wave generated by left ventricular ejection and a backward traveling wave reflected from the periphery.<sup>7</sup> Clinical trials found that antihypertensive drugs may exert different effects on cBP compared with

brachial BP.<sup>8,9</sup> In addition, some studies tested the changes in cBP during lifestyle modifications with special emphasis for dietary intervention.<sup>10,11</sup>

Dietary salt (ie, sodium chloride) intake is an important determinant of BP, with clear evidence that a high-salt intake is associated with increased BP.<sup>12</sup> In addition, high-salt intake is related to cardiovascular events and cardiovascular organ damage both directly and through its BP effects.<sup>13-15</sup> Epidemiological and clinical studies indicated a direct association between habitual salt consumption and cBP.<sup>16,17</sup> This association is supported by experimental evidence of structural and functional alterations induced by high-salt intake on the arterial wall and also by the effect of

higher BP. Indeed, high-salt intake is associated with increased oxidative stress with a reduction of nitric oxide bioavailability<sup>18-21</sup> and increment in smooth muscle cell tone.<sup>22,23</sup> Several intervention studies in humans evaluated the effect of dietary salt changes on peripheral BP as well as on cBP: however, the evidence with respect to the effect on cBP is inconsistent<sup>24-37</sup> mainly because of the low statistical power of most studies, the heterogeneity of the participants' features, the short length of intervention, and the magnitude of salt restriction.

A number of intervention studies that assessed both pulse wave velocity and cBP<sup>27,28,30-35</sup> were the object of a recent meta-analysis assessing the effect of salt intake on vascular damage [ie, carotid-femoral pulse wave velocity].<sup>15</sup> However, given that pulse wave velocity is purely expression of aortic stiffness whereas cBP is determined by aortic stiffness and peripheral resistance, we performed a further systematic review and meta-analysis of the available intervention trials testing the effect of dietary salt intake reduction on cBP using additional data.<sup>24-26,29,36,37</sup>

## 2 | METHODS

### 2.1 | Data sources and search strategy

This meta-analysis was planned, conducted, and reported according to the PRISMA statement<sup>38</sup> (Table S1). We performed a systematic search of the available publications using MEDLINE, Scopus, WOS, and the Cochrane Library, up to December 2018. The search strategy, without restrictions, used the expressions "sodium intake/consumption" OR "salt intake/consumption" OR "dietary salt/sodium" AND "pulse wave analysis" OR "PWA" OR "central haemodynamic" OR "central blood pressure" OR "augmentation index," or combinations thereof, either in medical subject headings or in the title/abstract. Further information was retrieved through a manual search of references from recent reviews and relevant published original studies.

### 2.2 | Study selection and data extraction

The data selection and extraction was independently conducted and reported by two reviewers (LD, ELF). Discrepancies about inclusion of studies and interpretation of data were resolved in conference, and consensus was reached after discussion. To be included in the meta-analysis, a published study had to meet the following criteria: (a) original article, (b) adult population study, (c) intervention study, (d) indication of a difference in cBP parameters between two different salt intake regimens in one or more patients' cohorts, and (e) indication of the number of participants included in the exposed and control group for each cohort. The risk of bias of the studies included in the meta-analysis was assessed according to established criteria<sup>39</sup> and reported in Table S2.

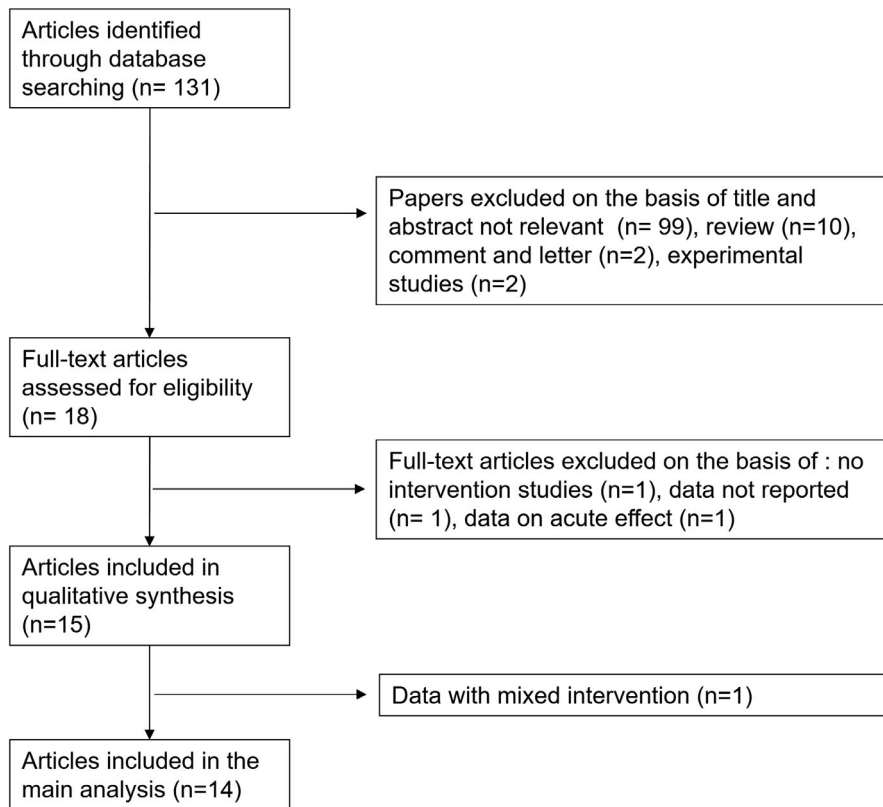
### 2.3 | Statistical analysis

Weighted mean differences (MD)—and standard error of the mean (SEM)—of the defined outcomes were extracted from the selected publications. If these were not available, MD and SEM were calculated from the comparison of the outcomes at low and high-salt regimens. The conversion to percentages was used in the analysis to calculate the between-regimen changes in the outcomes. The pooled MD and 95% confidence interval (CI) were estimated using a random-effect model.<sup>40</sup> The influence of the individual cohorts or of a particular study was estimated by sensitivity analysis. The Cochrane Q test and the  $I^2$  statistic were used to evaluate statistical heterogeneity across the studies. Funnel plots were constructed and visually assessed for possible publication bias.<sup>41</sup> Egger's and Begg's tests were also used to explore potential publication bias. Subgroup and meta-regression analyses were used to identify associations between changes in central hemodynamic parameters and relevant study's or patients' characteristics as possible sources of heterogeneity. The analyses were carried out for augmentation index (Aix), Aix adjusted for heart rate, central systolic BP (cSBP), and central pulse pressure (cPP), since adequate data were available for these outcomes. The statistical analyses were performed using the Stata Corp. software (version 11.2).

## 3 | RESULTS

### 3.1 | Characteristics of the studies included in the meta-analysis

Of a total of 131 publications retrieved, 15 studies were identified that met the inclusion criteria (Figure 1). However, one of them was excluded because it was based on mixed intervention.<sup>42</sup> Thus, eventually 14 studies were used for the analysis.<sup>24-37</sup> The main characteristics of the identified studies and of the respective study populations were recorded and reported in Table 1. Overall, the meta-analysis involved 457 participants from six countries. Three studies provided multiple cohorts including different categories of patients, one study recruited patients retrospectively stratified by BP salt sensitivity, another one included healthy normotensive participants stratified by age, and another one a group of healthy women stratified by history of pre-eclampsia. With respect to the comparison of the effects of higher vs lower salt intake, all but three studies were randomized controlled trials with a crossover design. Almost all studies used 24 hours urinary sodium excretion as a proxy for sodium intake during intervention, while two studies utilized 8 hours overnight urine specimens and one both 24 hours and 8 hours overnight collections. cBP was assessed by different methods: in the majority of studies by applanation tonometry, in two studies by a pressure transducer, and in two other studies by Doppler transducer.



**FIGURE 1** Stepwise procedure for selection of the studies. Flowchart indicating the results of the systematic review with inclusions and exclusions

Most of the studies assessed peripheral BP by an automated or semi-automated device, three by a mercury sphygmomanometer,<sup>26,28,37</sup> and only one by ambulatory blood pressure monitoring procedure.<sup>32</sup> All but two studies<sup>35,36</sup> described a careful standardization of the peripheral BP assessment method (ie, average of measurements and total number of measurements). The length of intervention ranged from 1 to 13 weeks. The evaluation of the “risk of bias” indicated that all but four studies were at low risk (Table S2).

### 3.2 | Effect of salt reduction on augmentation index

Detailed features of 14 studies (17 cohorts, 457 total participants) included in this analysis are given in Table 1.<sup>24-37</sup> In the pooled analysis, lower salt intake (average reduction of 24-hour urinary sodium excretion = 64%, ranged from 27% to 88%) was associated with significantly lower average AIx compared with the higher salt regimen (MD = -9.3%, 95% CI: -15.5 to -3.0,  $P = .003$ ). There was significant between-study heterogeneity ( $P < .01$ ;  $I^2 = 70\%$ ) (Figure 2A). Visual analysis of the funnel plot indicated little asymmetry (Figure S1), whereas Egger's and Begg's tests did not find significant evidence of publication bias (Egger:  $P = .8$ , Begg:  $P = .4$ ). A trend toward a direct association between lower salt intake and reduction in AIx was detected in almost all of the cohorts included in the analysis, and it was statistically significant in four of them, while there was a non-significant opposite trend in only two

cohorts (Figure 2A). Sensitivity analysis showed that the average change in AIx did not vary substantially with the exclusion of any individual study (Table S3).

Separate analysis after exclusion of the cohorts at high risk of bias<sup>24,26,35,37</sup> confirmed a beneficial effect of salt restriction (MD = -4.3%; 95% CI: -8.1 to -0.4). The analysis stratified by countries suggested a stronger effect of salt restriction in studies performed in United States compared to those carried out in Oceania, Europe, or China ( $P$  for interaction = .001) (Table 2).

The change in peripheral BP during the intervention was a significant source of heterogeneity (1% lower systolic or diastolic BP being associated with a decrease in AIx of 2.5% and 1.8%, respectively) (Table 2). The percentage of residual variation due to heterogeneity was reduced from 70% to 20% (systolic BP) and to 47% (diastolic BP). However, the influence of changes in BP disappeared when the cohorts of relatively younger and older people were separately analyzed. Indeed, in the analysis of 11 cohorts of relatively younger participants the reduction in systolic BP and diastolic BP did not affect the changes in AIx (reduction in systolic BP(%):  $\beta = 1.8$ , -0.3-3.9; reduction in diastolic BP(%):  $\beta = 1.4$ , 1.0-3.8). Likewise, also in the cohorts of older participants changes in peripheral BP were not sources of heterogeneity (reduction in systolic BP (%):  $\beta = 2.3$ , -6.0-10.6; reduction in diastolic BP (%):  $\beta = 1.0$ , -8.6-10.6). BP during high- and low-salt regimen did not affect the relationship between salt intake and AIx (Table 2). Also, AIx values during high- and low-salt intake were not a significant source of heterogeneity (Table 2).

Although meta-regression analysis did not identify the urinary sodium excretion at high- and low-salt intake as a significant source of heterogeneity (Table 2), the stratification by a cutoff of 90 mmol/24 hours (median of excretion changes) in urinary sodium excretion changes indicated a significantly greater reduction of the Alx in cohorts with lower change (<90 mmol/24 hours: MD = -13.7%, -24.7 to -2.8, vs > 90 mmol/24 hours: MD = -4.3%, -7.9 to -0.8;  $P$  for interaction < .01). Of note, these cohorts also had a higher average length of intervention.

A significant reduction in Alx after low-salt diet was more pronounced in older subjects (more than 60 years) as compared to younger participants (Table 2). Also, the instrumental method was a source of heterogeneity, a significantly higher effect of salt restriction being found in studies that assessed cBP by pressure transducers ( $P$  for interaction = .001) (Table 2). On the contrary, the different methods utilized to measure peripheral BP were not a significant source of heterogeneity (Table 2). As well as, the different description of its assessment did not affect the effect of the salt restriction on cBP ( $P$  for interaction = .97).

Meta-regression analysis indicated no influence of BMI, total number of participants, and gender on the association between dietary salt restriction and Alx (Table 2). Subgroup analysis did not detect the hypertensive status, antihypertensive treatment, and study design as significant sources of heterogeneity (Table 2). A similar result was obtained by subgroup and meta-regression analysis in relation to length of intervention (Table 2).

### 3.3 | Effect of salt reduction on augmentation index adjusted for heart rate

The separate analysis of 6 cohorts including Alx adjusted for heart rate<sup>34-37</sup> also showed a significant reduction of Alx on dietary salt restriction (MD = -6.3%; -10.9 to -1.8%;  $P$  < .01) (Figure S2). There was no significant heterogeneity among the studies ( $P$  = .6,  $I^2$  = 0%). The funnel plot for the effect of salt restriction on Alx suggested no significant publication bias, which was confirmed by Egger's and Begg's tests (Egger:  $P$  = .2, Begg:  $P$  = 1.0). A trend toward a direct association between lower salt intake and reduction in Alx was detected in all but one cohort included in the analysis. Sensitivity analysis showed that the average change in Alx did not vary substantially with the exclusion of any individual study. As expected, no features affected the effect of salt restriction.

### 3.4 | Effects of salt reduction on central systolic blood pressure

In total, six studies with eight cohorts and 261 total participants were included in the assessment of the association between salt intake reduction and changes in cSBP.<sup>25,26,29,34,36,37</sup> In the pooled analysis of eight cohorts, reduced salt intake (average difference

in 24-hour urinary sodium excretion = 67%, ranged from 48% to 88%) was associated with significantly lower cSBP (MD = -5.0%; -8.0 to -2.1%,  $P$  = .001), with significant heterogeneity between studies ( $P$  = .001,  $I^2$  = 73%) (Figure 2B). The funnel plot for the effect of salt restriction on cSBP suggested no significant publication bias, which was confirmed by Egger's and Begg's tests (Egger:  $P$  = .97, Begg:  $P$  = 1.0) (Figure S3). The evaluation of individual studies showed a trend toward a favorable association between salt reduction and the changes in cSBP in all but one cohort, with significantly lower MD in five of them, whereas a non-significant opposite trend was observed only in one small cohort. Sensitivity analysis indicated that the decrease in cSBP during salt reduction did not differ after the exclusion of any individual study.

A non-significant more pronounced reduction in cSBP was detected in the cohorts of old-aged subjects (MD = -6.7%, -9.5 to -3.9, vs -4.6%, -8.4 to -0.8,  $P$  for interaction .2), while meta-regression analysis indicated the changes in BP as significant sources of heterogeneity (SBP:  $\beta$  = 1.14, 0.57-1.70; DBP:  $\beta$  = 0.86, 0.18-1.53), but only in younger participants ( $P$  < .05).

### 3.5 | Effects of salt reduction on central pulse pressure

In the pooled analysis of five studies (seven cohorts, 268 total participants),<sup>25,34-37</sup> there was a direct relationship between reduction in salt intake (average difference in 24 hours urinary sodium excretion = 76%, ranged from 48% to 88%) and changes in cPP (MD = -7.6%; 95% CI: -11.9 to -3%-3%,  $P$  = .001) (Figure 2C). There was low heterogeneity among studies ( $P$  = .13,  $I^2$  = 40%) and no evidence of publication bias by inspection of funnel plot and by Egger's ( $P$  = .8) and Begg's tests ( $P$  = .7) (Figure S4). The evaluation of individual studies showed a trend toward an association between reduction in salt intake and cPP in all cohorts, with a significant change in three of them. Sensitivity analysis showed that the average change in cPP did not vary substantially with the exclusion of any individual study. Meta-regression analysis did not detect any significant sources of heterogeneity ( $P$  > .05).

### 3.6 | Effects of salt reduction on peripheral blood pressure

A meta-analysis of the effects of salt restriction on brachial BP in the same cohorts was performed.<sup>24-34,36,37</sup> Pooled analyses showed a significant reduction of both systolic (MD = -4.9%, -6.6 to -3.2) and diastolic BP (MD = -3.3%, -5.2 to -1.4) upon reduction of salt intake. There was high heterogeneity among studies (systolic BP:  $I^2$  = 73%,  $P$  < .01; diastolic BP:  $I^2$  = 80%,  $P$  < .01), whereas no evidence of publication bias was detected (Egger's test, systolic BP:  $P$  = .8, diastolic BP:  $P$  = .3; Begg's tests, systolic BP:  $P$  = .6, diastolic BP:  $P$  = .5).

**TABLE 1** Characteristics of the studies included in the meta-analysis

First Author, year (ref)	Country	Cohort (n. of participants)	Selected features of the study participants	Mean Age (yrs)	Mean BMI (Kg/m <sup>2</sup> )
Seals, 2001 <sup>24</sup>	USA	Post-menopausal female participants (W17)	Post-menopausal status, SBP 130-159, DBP ≤99 mm Hg, no treatment, not smoking, absence of chronic diseases	65	28.1
Gates, 2004 <sup>25</sup>	USA	White hypertensive older participants (6M, 6W)	Untreated stage 1 systolic hypertension	64	24.7
Al-Solaiman, 2009 <sup>26</sup>	USA	SR participants (3M, 7W)	Lean normotensive without metabolic syndrome and obese hypertensive subjects with metabolic syndrome	34.3	30.1
		SS participants (2M, 7W)		44.1	26.5
Dickinson, 2009 <sup>27</sup>	Australia	Overweight/obese participants (7M, 22W)	SBP < 160 mm Hg, No CVD, no antihypertensive therapy, BMI > 27 and <40 Kg/m <sup>2</sup>	52.7	31.6
Pimenta, 2009 <sup>28</sup>	USA	Hypertensive White and Black patients (4M, 8W)	Resistant Hypertension (with HCT and RAAS-blocking treatment)	55.5	32.9
Starmans-Kool, 2011 <sup>29</sup>	UK	Healthy young subjects (10M)	SBP/DBP < 140/90 mm Hg, no treatment; no CVD, no diabetes	32	-
Todd, 2010 <sup>30</sup>	New Zealand	(Pre)Hypertensive or hypertensive participants (13M, 21W)	SBP/DBP > 130/85 mm Hg or treatment	51.8	25.7
Todd, 2012 <sup>31</sup>	New Zealand	Healthy Caucasian subjects (5M, 18W)	SBP/DBP < 130/85 mm Hg, no treatment; no CVD, BMI <30 Kg/m <sup>2</sup>	43.7	25.3
McMahon, 2013 <sup>32</sup>	Australia	CKD patients (15M, 5W)	Hypertension (SBP 130-169, DBP ≥ 70 mm Hg), CKD stage 3 or 4 (not transplanted).	68.5	29.3
Dickinson, 2014 <sup>33</sup>	Australia	Overweight/obese normotensive participants (8M, 17W)	BMI: 27-40 Kg/m <sup>2</sup> , no CVD, SBP/DBP < 140/90 mm Hg, no treatment	-	-
Gijsbers, 2015 <sup>34</sup>	Netherlands	(Pre)Hypertensive Caucasian participants (24M, 12W)	No smoking, SBP 130-159 mm Hg, no treatment, no CVD, no diabetes	65.8	27.2
Van der Graaf, 2016 <sup>35</sup>	Netherlands	Subjects with history of NP (18W)	SBP/DBP < 140/90 mm Hg, non-smokers	36	22.6
		Subjects with history of PP (18W)		36	25.3
Muth, 2017 <sup>36</sup>	USA	Healthy normotensive participants—Young (30M, 19W)	SBP/DBP < 140/90 mm Hg, no CVD, no diabetes, non-smokers, non-obese	27	23.6
		Healthy normotensive participants—Middle-aged (13M, 23W)		52	25.1
Xing, 2018 <sup>37</sup>	China	(Pre)Hypertensive participants (60M, 39W)	SBP 130-159, DBP 85-100 mm Hg, no treatment, absence of chronic diseases	53.4	25.1

Abbreviations: CKD, chronic kidney disease; CVD, cardiovascular disease; DASH, dietary approaches to stop hypertension; DBP, diastolic blood pressure; HCT, hydrochlorothiazide; LS, low salt; M, men; MBP, mean blood pressure; NP, normotensive pregnancy; PP, pre-eclamptic pregnancy; PWV, pulse wave velocity; RAAS, renin-angiotensin-aldosterone system; SBP, systolic blood pressure; SR, salt-resistant participants; SS, salt-sensitive participants; W, women.

<sup>a</sup>8-h overnight urine.

<sup>b</sup>Assessed by ABPM (ambulatory blood pressure monitoring).

The reduction in BP after salt restriction was detected both in younger and older participants, but it was more pronounced in old-aged subjects (systolic BP, MD=-6.9%, -9.6 to -4.3 vs -4.2%, -6.4 to -2.1; *p* for interaction = 0.01; diastolic BP, MD=-5.2%, -8.0 to -2.3 vs -2.8%, -5.3 to -0.2, *p* for interaction = 0.02).

Additional analyses did not detect significant difference between changes in brachial BP and cBP (*p* for interaction > 0.1).

## 4 | DISCUSSION

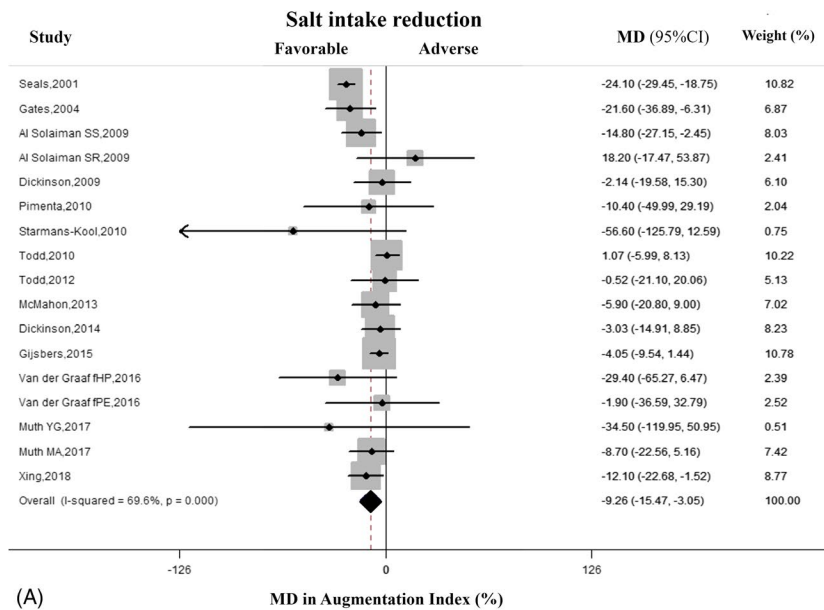
The results of this meta-analysis indicate that reduction of dietary salt intake is associated with lowering cBP. The effect of lower salt intake was more pronounced after prolonged salt reduction, and in fact, greater reduction in cBP was achieved in the cohorts with relatively lower levels of salt reduction but higher length of intervention.

Duration of intervention (days)	Assessment Method	Low vs High Sodium Comparison (mmol/24 h)	Changes in SBP/DBP (mm Hg)	Study Design
90	Pressure trasducer (TCB-500, Millar Instruments)	86 vs 124	-16/-7	Intervention study single-blind (High-salt diet vs low-salt diet)
28	Pressure trasducer (TCB-500, Millar Instruments)	54 vs 135	-7/-1.6	Crossover Double-blind (slow-sodium suppl. vs placebo)
21	Applanation tonometry (Sphygmocor)	34 vs 97	2.8/3.2	Intervention study (DASH vs LS-DASH)
14	Doppler trasducer	51 vs 104 64 vs 156	-7.1/-3.2 -5/-1	Crossover (sodium restriction vs regular sodium)
7	Applanation tonometry (Sphygmocor)	46 vs 252	-22.8/-9.1	Crossover (slow-sodium suppl. vs low-salt diet)
14	Applanation tonometry (SPT 301; Millar Instruments, Houston, TX)	94 vs 191	-2/0	Crossover Double-blind (slow-sodium suppl. vs placebo)
28	Applanation tonometry (Sphygmocor)	60 vs 200 <sup>a</sup>	-5.8/-3.4	Crossover Single-blind (sodium suppl. vs low-salt diet)
28	Applanation tonometry (Sphygmocor)	60 vs 200 <sup>a</sup>	0.1/0.4	Crossover Single-blind (sodium suppl. vs low-salt diet)
42	Applanation tonometry (Sphygmocor)	75 vs 168	-9.7/-3.9 <sup>b</sup>	Crossover Double-blind (slow-sodium suppl. vs placebo)
42	Doppler trasducer	113 vs 155	-3/-1	Crossover Single-blind (slow-sodium suppl. vs placebo)
28	Applanation tonometry (Sphygmocor)	105 vs 203	-7.5/-3.3	Crossover Double-blind (slow-sodium suppl. vs placebo)
7	Applanation tonometry (Sphygmocor)	39 vs 221	-/-	Crossover (sodium restriction vs high sodium)
7	Applanation tonometry (Sphygmocor)	45 vs 258	-/-	
7	Applanation tonometry (Sphygmocor)	32 vs 243 28 vs 243	-3/-3 -8/-1	Crossover (slow-sodium suppl. vs placebo)
7	Applanation tonometry (Sphygmocor)	234 vs 55 <sup>a</sup>	-10.1/-5.6	Intervention study (High-salt diet vs low-salt diet)

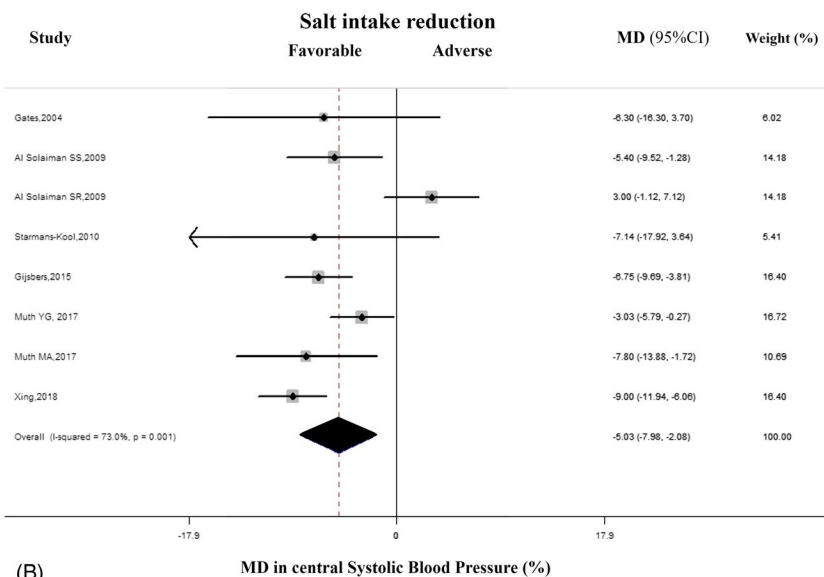
The effect of low-salt intake on cBP was not limited to hypertensive patients as indeed a significant reduction in cBP was detected both in the cohorts including pre-hypertensive and/or hypertensive participants and in the cohorts that enrolled only normotensive individuals.

By contrast, age was an important cause of heterogeneity in the response of cBP to salt restriction in as much as the effect of the

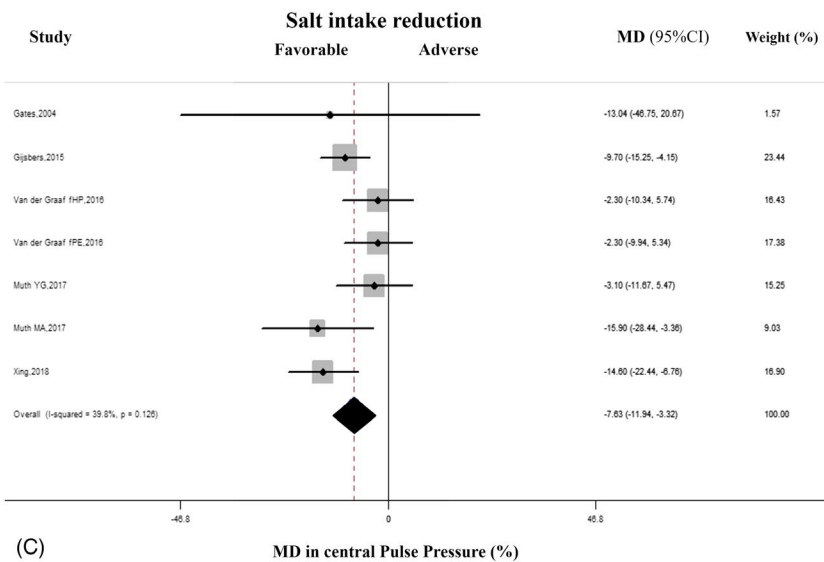
latter was greater in the cohorts including older participants. This result may be at least partly explained by the greater salt sensitivity of older compared with younger subjects.<sup>7,43</sup> On the other hand, gender, BMI, and heart rate did not seem to play a role in the response of cBP to reduction in salt intake. With regard to the study design, there was not a significant difference between results of randomized controlled trials and not.



(A) MD in Augmentation Index (%)



(B) MD in central Systolic Blood Pressure (%)



(C) MD in central Pulse Pressure (%)

**FIGURE 2** A, Effect of lower sodium intake on augmentation index (AIx). Forest plot of the effect of lower dietary sodium intake on AIx in 17 population cohorts from 14 published studies. B, Effect of lower sodium intake on central systolic blood pressure (cSBP). Forest plot of the effect of lower dietary salt intake on cSBP in eight population cohorts from six published studies. C, Effect of lower sodium intake on central pulse pressure (cPP). Forest plot of the effect of lower dietary sodium intake on cPP in seven population cohorts from five published studies. Results are expressed as mean difference (MD) and 95% confidence intervals (95% CI). Squares indicate study-specific relative risk estimates (size of the square reflects the study-specific statistical weight); horizontal lines indicate 95% CI; and diamond indicates the overall relative risk with its 95% CI. fHP, formerly healthy pregnant women; fPE, formerly pre-eclamptic women; SR, salt-resistant participants; YG, young participants; MA, middle-aged participants



**TABLE 2** Subgroup and meta-regression analysis of the effect of salt restriction on augmentation index

Subgroup analysis	Variables (n. of cohorts)	Pooled mean (%) reduction Alx	95% CI	P for interaction
Age	<60 (13)	-5.8	-10.7 to -0.9	0.001
	>60 y(4)	-13.9	-26.8 to -1.0	
Length of intervention	1 wk (6)	-11.4	-19.2 to -3.7	0.9
	>1 wk (11)	-8.3	-16.2 to -0.4	
Country of origin	USA (7)	-16.6	-24.6 to -8.6	0.001
	Oceania (5)	-0.9	-6.1 to 4.2	
	Europe (4)	-9.6	-24.5 to 5.3	
	Asia (1)	-12.1	-22.7 to 1.5	
(Pre)Hypertension status	Yes (7)	-10.9	-20.3 to -1.4	0.3
	No (10)	-7.3	-13.4 to -1.2	
Antihypertensive treatment <sup>a</sup>	Yes (2)	-6.5	-20.4 to 7.5	0.4
	No (14)	-10.8	-17.7 to -4.0	
Assessment device of central BP	Sphygmocor (13)	-5.6	-9.8 to -1.5	0.001
	Pressure transducer (2)	-23.8	-28.9 to -18.8	
	Doppler transducer (2)	-2.7	-12.6 to 7.1	
Assessment device of peripheral BP <sup>b</sup>	Automated/semi-automated sphygmomanometer (12)	-9.7	-17.8 to -1.5	0.7
	Mercury sphygmomanometer (4)	-11.7	-19.4 to -4.0	
Study design	Randomized controlled trials (13)	-4.5	-8.2 to -0.7	0.06
	Non-randomized controlled trials (4)	-15.2	-25.6 to -4.8	
<b>Meta-regression analysis</b>				
<b>Variables (n. of cohorts)</b>		<b>Reduction in Alx (%) (coefficient)</b>		<b>95% CI</b>
Age (y)(16)		-0.18		-0.81 to 0.45
BMI (Kg/m <sup>2</sup> )(15)		1.08		-1.94 to 4.11
Length of intervention (week)		-0.17		-0.37 to 0.04
Number of participants (n) (17)		0.05		-0.21 to 0.31
Gender (% men) (17)		0.11		-0.12 to 0.35
SBP at low-salt intake (mm Hg) (15)		-0.19		-0.86 to 0.48
SBP at high-salt intake (mm Hg) (15)		-0.27		-0.77 to 0.23
DBP at low-salt intake (mm Hg) (15)		0.36		-1.26 to 1.98
DBP at high-salt intake (mm Hg) (15)		-0.09		-1.30 to 1.13
SBP difference (reduction in %) (15)		2.47		1.21 to 3.72
DBP difference (reduction in %) (15)		1.80		0.04 to 3.56
Alx at low-salt intake (%) (17)		0.51		-0.54 to 1.56
Alx at high-salt intake (%) (17)		-0.13		-1.00 to 0.74
Urinary Na at low-salt intake (mmol/24 h) (17)		0.04		-0.19 to 0.27
Urinary Na at high-salt intake (mmol/24 h) (17)		0.03		-0.01 to 0.13
Urinary Na difference (reduction in %) (17)		-0.06		-0.38 to 0.25

Abbreviations: Alx, augmentation index; DBP, diastolic blood pressure; SBP, systolic blood pressure.

<sup>a</sup>Analysis did not include the study by Todd et al (ref. 30), because participants with and without antihypertensive therapy were included.

<sup>b</sup>Analysis did not include the study by McMahon et al (ref. 32) because only ambulatory blood pressure monitoring values were reported.

#### 4.1 | Potential mechanisms involved

Although our study had no specific potential to address the mechanisms of the effect of salt intake restriction on cBP, a few considerations can be made as to the possible explanations of our study

results. In the first place, because of the calibration method of cBP assessment, that includes brachial BP values,<sup>7</sup> it might be hypothesized that the decrease in cBP upon reduction in salt intake is determined by the decrease in peripheral BP. Our analyses indicate that the peripheral BP at high and at low-salt intake did not affect the



relationship between salt intake and cBP. However, our results suggest that changes in peripheral BP did not fully explain the reduction in cBP. Of note, although the effect of salt restriction on changes in cBP seems greater than those on peripheral BP, this difference is not statistically significant.

There is experimental evidence in support of the relationship between salt intake and cBP. In particular, a large body of evidence indicated a reduced nitric oxide bioavailability after sodium loading due to increase in reactive oxygen species and reduction of endothelial nitric oxide synthase activity.<sup>18-21,44-47</sup> The effects of salt loading were especially seen in smooth muscle cells, which featured decreased availability of and reduced responsiveness to nitric oxide,<sup>22,23</sup> increased sympathetic nerve activity,<sup>48</sup> and increased renin-angiotensin-aldosterone system (RAAS) activity.<sup>49-53</sup> Also, salt sensitivity of BP might play a role,<sup>43</sup> but only one study with a weak methodology among those included in our meta-analysis evaluated this condition.<sup>26</sup>

## 4.2 | Study strengths and limitations

Major strengths of our meta-analysis are the following: (a) the finding of a trend to reduction of cBP upon dietary salt restriction in the majority of the cohorts examined; (b) the “low risk” of bias in the majority of the studies; (c) the lack of detectable publication bias; (d) the inclusion of only intervention trials with exclusive evaluation of the salt effect; (e) the inclusion as outcomes of only non-invasively assessed cBP parameters; and (f) the measurement of 24 hours urinary sodium excretion, a recognized gold standard for monitoring salt intake,<sup>54</sup> in all but three studies.

Nonetheless, this meta-analysis has some limitations: the first one is the inability to rationally assess a dose-effect relationship between salt intake reduction and decrease in cBP. Second, this study does not allow to draw definitive conclusions about the long-term effects of dietary salt restriction on cBP, given that only one trial included in the meta-analysis had an intervention period longer than 6 weeks. Third, our meta-analysis was conducted based on aggregated data and not on individual data, so limiting the possibility to carry out additional potential analyses. Fourth, some characteristics of the studies included in the meta-analysis represent inherent limitations of the study. In particular, the potential influence of the concomitant antihypertensive drug treatment cannot be definitely ruled out also because of the small number of studies with antihypertensive treatment included and the mixed therapy considered. Although there was a greater reduction of cBP in the participants without antihypertensive therapy, the subgroup analysis did not detect significant difference. On the other hand, previous meta-analyses showed a greater effect of salt restriction on subclinical organ damage during concomitant administration of renin-angiotensin-aldosterone blockers.<sup>14,15</sup> However, they included a greater number of studies with antihypertensive treatment.

Likewise, the heterogeneity of method used in included studies may be a limitation. Although there are no studies directly comparing

the methods, there may be differences between the output of the methods, especially cPP and Alx, because of high-frequency signals. Indeed, carotid applanation tonometry does not use transfer function, while radial tonometry uses a transfer function with calibration from brachial BP. Therefore, use of brachial BP may introduce some errors. However, the potential influence of the measurement device used in the different studies cannot be correctly assessed since the majority of the studies utilized applanation tonometry in respect to small number of studies that used other devices. On the other hand, the different methods utilized to evaluate peripheral BP did not affect the results. Likewise, also a careful description of the peripheral BP assessment method was not a significant source of heterogeneity.

Also, the evaluation of race differences cannot be assessed, although subgroup analysis indicated country of origin as potential source of heterogeneity.

Another possible limitation was given by the small sample size of most of the available studies, high heterogeneity of studies' characteristics, for example, length of intervention, magnitude of salt restriction, and participants' features.

Finally, although the cBP has been debated in its usefulness in the risk prediction compared to brachial BP, some evidence suggested that cBP is more closely associated with cardiovascular organ damage and a better predictor of future cardiovascular events than brachial BP.<sup>4-7</sup>

## 4.3 | Implications for public health

In keeping with previous demonstrations of a favorable effect of salt restriction on cardiovascular damage,<sup>14,15</sup> the results of this meta-analysis support the concept of a protective role of lower dietary salt in individuals both with (pre)hypertension and not. This concept is in line with the recognized beneficial effect of moderate dietary salt restriction on BP.

The results of the present study have important implications for public health: cBP (expressed as Alx or cSBP or cPP) is a recognized predictor of cardiovascular events.<sup>3</sup> Based on previous studies indicating increments of 38% in all-cause mortality and of 32% in cardiovascular events for 10% increase in Alx,<sup>3</sup> a decrease in Alx upon reduction of salt intake is expected to translate into a substantial reduction in cardiovascular risk. Likewise, in consideration of the increase in cardiovascular events of 13% for an increase of cSBP or cPP by 10 mm Hg,<sup>3</sup> a substantial reduction of this risk is expected as a result of dietary salt restriction.

Based on our results, benefit from salt intake restriction may be expected in younger and older people, during moderate long-term salt restriction, and independently of baseline risk.

As the habitual salt intake in most countries in the world is close to 10 g per day,<sup>55-58</sup> an average reduction of 60% per day, as in our meta-analysis, would lead to the achievement of the recommended target of 5 g or lower per day for the population<sup>55</sup>: this observation suggests that the results of our study could be applicable to real life conditions and are relevant to population-based strategies for reduction of salt intake.<sup>59</sup>

## 5 | CONCLUSIONS AND PERSPECTIVES

The results of this study show that dietary salt restriction reduces cBP parameters, at least in part independently from the concomitant changes in peripheral BP. In consideration of the importance of cBP as a predictor of cardiovascular morbidity and mortality and that the non-invasive assessment underestimates the actual cBP values,<sup>60</sup> this effect of salt restriction significantly adds to its recognized value in cardiovascular disease prevention. In addition, these findings indicate cBP as a possible additional parameter to clinically evaluate the response to salt reduction. Our results support the recommendations in favor of moderate and long-term reduction in dietary salt intake to decrease the risk of cardiovascular diseases. Future powered randomized controlled trials should be carried out to focus on the effect of long-term moderate dietary salt reduction on central hemodynamics, to further support the conclusions of our review and extend current knowledge in this field.

### CONFLICT OF INTEREST

LD was a technical advisor to the World Health Organization and is a member of the scientific committee of the Italian Society of Human Nutrition. PS is an unpaid member of WASH, scientific coordinator of the Interdisciplinary Working Group for Reduction of Salt Intake in Italy (GIRCSI), and member of the committee for the preparation of the Italian Nutritional Guidelines. The remaining authors do not disclose any conflict of interest.

### AUTHORS' CONTRIBUTIONS

Lanfranco D'Elia, Ersilia La Fata, and Ferruccio Galletti designed the study; Lanfranco D'Elia, Ersilia La Fata, and Alfonso Giaquinto conducted the systematic literature review and extracted the data; Lanfranco D'Elia conducted the statistical analysis; Lanfranco D'Elia and Ferruccio Galletti wrote the manuscript; Ersilia La Fata, Alfonso Giaquinto, and Pasquale Strazzullo contributed to the manuscript.

### DISCLOSURES

Part of the preliminary study data was previously presented at the British and Irish Hypertension Society 2019 Meeting.

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