# Neighborhood Socioeconomic Status and Stroke Incidence

# A Systematic Review

Yeonwoo Kim, PhD, Erica Twardzik, MS, Suzanne E. Judd, PhD, and Natalie Colabianchi, PhD Neurology® 2021;96:897-907. doi:10.1212/WNL.000000000011892

# **Abstract**

# **Objective**

To summarize overall patterns of the impact of neighborhood socioeconomic status (nSES) on stroke incidence and uncover potential gaps in the literature, we conducted a systematic review of studies examining the association between nSES and stroke incidence, independent of individual SES.

#### **Methods**

Four electronic databases and reference lists of included articles were searched, and corresponding authors were contacted to locate additional studies. A keyword search strategy included the 3 broad domains of neighborhood, SES, and stroke. Eight studies met our inclusion criteria (e.g., nSES as an exposure, individual SES as a covariate, and stroke incidence as an outcome). We coded study methodology and findings across the 8 studies.

#### Results

The results provide evidence for the overall nSES and stroke incidence association in Sweden and Japan, but not within the United States. Findings were inconclusive when examining the nSES-stroke incidence association stratified by race. We found evidence for the mediating role of biological factors in the nSES-stroke incidence association.

### **Conclusions**

Higher neighborhood disadvantage was found to be associated with higher stroke risk, but it was not significant in all the studies. The relationship between nSES and stroke risk within different racial groups in the United States was inconclusive. Inconsistencies may be driven by differences in covariate adjustment (e.g., individual-level sociodemographic characteristics and neighborhood-level racial composition). Additional research is needed to investigate potential intermediate and modifiable factors of the association between nSES and stroke incidence, which could serve as intervention points.

#### Correspondence

Dr. Colabianchi, colabian@umich.edu

#### **RELATED ARTICLE**

#### **Editorial**

Neighborhood Socioeconomic Differences in Stroke Risk? Highlighting the Need for Further Research

Page 879

From the Department of Kinesiology (Y.K.), University of Texas at Arlington, TX; School of Kinesiology (E.T.), University of Michigan, MI; Department of Biostatistics (S.E.J.), University of Alabama at Birmingham, AL; School of Kinesiology (N.C.), University of Michigan, MI.

Go to Neurology.org/N for full disclosures. Funding information and disclosures deemed relevant by the authors, if any, are provided at the end of the article.

Y. Kim and E. Twardzik contributed equally to this work.

# Glassary

**CHS** = Cardiovascular Health Study; **nSES** = neighborhood SES; **REGARDS** = REasons for Geographic and Racial Differences in Stroke; **SES** = socioeconomic status.

Although medical advances and prevention efforts have decreased stroke incidence over the long term, contemporary social disparities in stroke incidence remain marked.<sup>1</sup> In the United States, a decrease in stroke incidence is more notable among White Americans than Black Americans.<sup>2,3</sup> In addition, in the United States and European countries, adults from lower socioeconomic status (SES) have a higher risk of stroke incidence than those from higher SES.<sup>4-7</sup> Widening social disparities in stroke incidence may be related to contextual and structural factors.

Neighborhood SES (nSES) is an important predictor of individual health. Neighborhoods with low nSES often have fewer health-promoting amenities, such as community centers, physical activity facilities, and access to healthy groceries, as well as less social capital. Health-promoting environments may be associated with healthy behaviors and subsequent health outcomes, including stroke. As neighborhoodhealth research has become more prominent over the past 2 decades, recent empirical studies have tested the association between nSES and stroke. 6,12-18

Review studies<sup>19,20</sup> have documented the association between SES, both at the individual and neighborhood levels, and stroke incidence. However, the reviews included studies that examined the role of nSES in stroke incidence without controlling for individual-level SES. Omitting individual-level SES is likely to lead to biased estimates of the true relationship between nSES and stroke risk because individual-level SES is positively related to nSES and inversely related to stroke risk.<sup>21,22</sup> In addition, there has been no review study investigating mediators and moderators of the nSES-stroke incidence association, which could serve as points of policy and program interventions.

We conducted a systematic review to answer the following research questions: (1) is nSES associated with stroke incidence independent of individual SES? and (2) which moderating and mediating factors have been tested and found to be significant in the nSES-stroke incidence association?

# **Methods**

A systematic review was conducted to identify and synthesize all studies that had examined the effects of nSES on stroke incidence independent of individual SES. Preferred Reporting Items for Systematic Reviews and Meta-Analyses reporting guidelines for systematic reviews were followed and fulfilled. We predetermined specific inclusion and exclusion criteria, a comprehensive and explicit search strategy, a data extraction form, and a tool for quality assessment. 24-26

#### Inclusion/Exclusion Criteria

We included published and unpublished studies in English that met search criteria. Inclusion criteria were the following: (1) nSES as an exposure, (2) individual SES as a covariate, (3) stroke incidence (hemorrhagic or ischemic) as an outcome, (4) quantitative study, and (5) peer-reviewed article, dissertation, or government report. Exclusion criteria were the following: (1) articles focusing on a non-high-income country (defined by the World Bank)<sup>27</sup> because neighborhood effects may not be comparable to countries in different socioeconomic contexts, 28 (2) articles using only TIA as an outcome (because TIA represents a heterogeneous outcome often with mixed diagnosis and believed to be underreported in the population because of individuals missing the signs and symptoms of a TIA and, therefore, not getting a diagnosis of TIA), (3) articles examining associations within the stroke population (i.e., stroke incidence is not an outcome, but the population of interest), (4) review articles, or (5) ecological studies.

# Search Strategy

A total of 4 databases (i.e., ProQuest Dissertation & Theses, PsychInfo, PubMed, and Web of Science) were used to search for studies. Citations were searched and retrieved from all 4 databases on February 11, 2019, and May 19, 2020. A keyword search strategy included 3 broad domains of neighborhood, SES, and stroke. For example, the search strategy for PubMed consisted of the following: (Neighbor\* OR Communit\* OR Environ\* OR "Residence Characteristics") AND ("Socioeconomic Factors" OR socioeconomic OR SES OR poverty OR impoverished OR unemploy\* OR income OR affluen\* OR advantaged OR advantage OR deprivation OR occupation\* OR "Standard of Living "OR "Living Standard" OR "living standards" OR inequality OR wealth) AND ("stroke" OR stroke OR strokes OR "Cerebrovascular Accident" OR "cerebrovascular accidents" OR CVA OR "Cerebrovascular Apoplexy" OR "Vascular Accident" OR "vascular accidents" OR "Apoplexy" OR "Cerebral hemorrhage" OR "Subarachnoid hemorrhage"). The PubMed search strategy was translated to other research databases by (1) inputting our search strategy into a polyglot tool for translation (sraccelerator.com/#/polyglot), (2) subjecting the search strategy to independent review and revision by authors Y. Kim and E. Twardzik, and (3) comparing individual review and translation of the search strategy by Y. Kim and E. Twardzik; when disagreement arose, the authors would come to a consensus.

First, 2 authors independently reviewed titles/abstracts. A study was included for full-text review in cases in which inclusion/exclusion could not be determined from its title or abstract. Second, the authors independently reviewed the full

text of all studies selected in the first step. If a study used nSES as a covariate and did not report the effect estimate within the article, the corresponding author for the study was contacted, and additional information was requested. When there were disagreements about a particular study, the authors discussed its eligibility until a consensus was reached. In addition, reference lists from included studies and review studies on neighborhood and stroke were manually searched for further relevant publications. Finally, we contacted through e-mail the corresponding authors of studies included in this review to find whether they could identify any other studies meeting our inclusion criteria and (if relevant statistical findings were not fully reported) to receive additional results or clarification. None of the studies that the corresponding authors recommended met our inclusion criteria. We received additional results for 1 study and confirmation of our interpretation for 3 studies.

# **Data Extraction and Quality Assessment**

For each study, 2 authors independently extracted and recorded data on research methodology and statistical results. Disagreement was resolved through discussion. Cohen kappa for an interreviewer agreement was  $0.827 \ (p < 0.001)$  across all categories. In addition, an adapted version of the Newcastle-Ottawa Assessment Scale was used to assess the risk of bias of the included studies.

# **Data Availability**

Data are available to investigators on request from the first author.

# Results

#### **Identified Studies**

Figure 1 shows the number of articles identified, included, and excluded, and the reasons for exclusion within this study. On screening the title, abstract, and full text, a total of 10 studies met the inclusion criteria; however, 2 studies were excluded because of missing effect estimate information for the nSES-stroke incidence relationship. As a result, 8 studies were included in the final sample.

# **Description of Included Studies**

All studies were published after January 2007, and 5 of 8 studies were conducted in the United States, as shown in table 1. Quality assessment of the study demonstrated satisfactory study quality and low risk of bias in ensuring a sufficient sample size of data (n = 8), clearly describing a measure of nSES (n = 8), controlling for important individual-level confounders (n = 8), clearly describing statistical testing and findings (n = 8), using a representative sample (n = 7), and assessing stroke incidence using record linkage or medical record review (n = 7). The only source of potential bias was no description of the response rate or a low response rate in the data (n = 4).

# Data

Table 2 presents the data source and exposure measures of included studies. The data sets used within each study varied.

Two studies in the United States used the data from the Cardiovascular Health Study (CHS). Three US studies each used the Chicago Health and Aging Project, Health and Retirement Study, Reasons for Geographic and Racial Differences in Stroke (REGARDS). Two studies in Sweden used data from the Nationwide Swedish Registers 13,33: the Total Population Register and the Patient Register in the study by Carlsson et al. and the Hospital Discharge Register and the Outpatient Register in the study by Forsberg et al. One study in Japan used the data from the Japan Public Health Center–based Prospective Study.

#### nSES Measure

All studies used an administrative entity created by governments to represent a neighborhood. Five studies in the United States  $^{12,17,18,31,32}$  defined neighborhoods using a census unit aggregation, including census tracts (n = 3) and block groups (n = 2). Census tracts and block groups are designed to be relatively homogeneous areas with respect to sociodemographic characteristics of the population. Block groups contain between 600 and 3,000 individuals, whereas census tracts contain about 4,000 individuals. Two studies in Sweden measured neighborhoods based on small area market statistics, which contain an average of 1,000 residents. One study in Japan  $^{16}$  defined neighborhoods based on Chocho-Aza units, a comparable size to a US block group.

Figure 1 PRISMA Flow Diagram for Study Selection

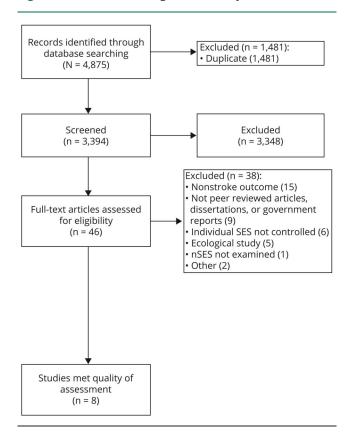


Table 1 Characteristics of the 8 Included Studies

Author (year)	Location	Sample	Age
Brown (2011) <sup>12</sup>	United States	4,619	Mean = 73
Carlsson (2017) <sup>13</sup>	Sweden	1,153,451	Mean = 40
Clark (2011) <sup>31</sup>	United States	5,789	Mean = 75
Forsberg (2018) <sup>33</sup>	Sweden	3,644,309	Range = 40-60
Freedman (2011) <sup>32</sup>	United States	15,374	Range = 55+
Honjo (2015) <sup>16</sup>	Japan	90,843	Mean = 50
Howard (2016) <sup>17</sup>	United States	24,875	Mean = 65
Yan (2013) <sup>18</sup>	United States	4,619	Mean = 73

All studies created a composite index of nSES using several indicators of neighborhood characteristics. However, the operationalization of nSES varied in the utilization of items. For example, 3 studies 12,17,18 derived an nSES index using a factor analysis with 6 items representing education, income, occupation, and housing values. Two studies 13,33 created an index of nSES using a principal components analysis with 4 items representing education, income, and employment, with higher values indicating higher levels of deprivation. One study<sup>31</sup> conducted a confirmatory factor analysis by including similar items as those above into indices. The remaining 2 studies 16,32 included items that do not directly indicate nSES conditions along with other items that directly measure nSES (e.g., a percentage of the total population in poverty, an unemployment rate, and a proportion of blue-collar workers) to create an nSES index; specifically, Honjo et al. 16 included a proportion of elderly couples or single households along with nSES items to create an area deprivation index, and Freedman et al.<sup>32</sup> included a percentage of the population that was Black along with nSES items when conducting an exploratory factor analysis.

#### **Outcome Measure**

Table 3 shows outcome measures, covariates, and analysis of included studies. The primary outcome was measured as follows: stroke incidence in 4 studies. <sup>16,17,31,32</sup> and ischemic stroke incidence in 4 studies. <sup>12,13,18,33</sup> Of note, although 7 studies <sup>12,13,16-18,31,33</sup> used a time to event outcome, Freedman et al. 32 did not consider the time of onset throughout a 2-year study period. Four studies identified stroke cases within their study population using a single measure of stroke, with 3 studies using record linkage 13,31,33 and 1 study using selfreported stroke.<sup>32</sup> Three studies identified stroke cases using 2 measures of stroke. Honjo et al. 16 used self-reported stroke and linked this report back to records, and 2 studies used record linkage and adjudication of the stroke through medical record review. 12,18 Only 1 study used 3 measures of stroke incidence.<sup>17</sup> Howard et al.<sup>17</sup> had participants self-report stroke incidence through follow-up calls every 6 months, obtained medical records from the stroke event, and confirmed stroke cases through review of participant medical records by a stroke neurologist.

### **Individual SES and Covariates**

Studies included a variable of individual SES to examine the nSES-stroke incidence associations, theoretically independent of individual SES. Of 8 studies, 4 studies<sup>12,17,18,33</sup> controlled for both education and income, 2 studies<sup>13,31</sup> controlled only for education, 1 study<sup>32</sup> controlled for education, income, and household wealth, and the other study<sup>16</sup> controlled only for occupation. All included studies controlled for individual SES, age, and sex as covariates within the analysis. Beyond these 3 variables, covariate adjustment varied by individual study (column 4, table 3).

# **Analysis Methods**

Seven studies conducted Cox proportional hazards <sup>13,16-18,31,33</sup> or multilevel Cox proportional hazards <sup>12</sup> using a time-to-event outcome. On the other hand, the remaining study, which measured whether the respondents had stroke during a 2-year study period, conducted a multilevel logistic analysis.<sup>32</sup>

# **Findings of Multivariate Analysis**

Table 4 and figure 2 present summarized findings of the included studies. Below these results are summarized among studies examining (1) an overall effect within their study sample, (2) effects within subgroups, and (3) mediating mechanisms.

#### **Overall Effects**

Four of the 8 studies tested and presented an overall effect of nSES on stroke incidence within their full study sample.  $^{13,16,17,31}$  Although the 4 studies  $^{13,16,17,31}$  showed the expected direction of effect, namely a positive effect of nSES disadvantage on stroke risk, this relationship was statistically significant in only the 2 studies conducted in non-US countries.  $^{13,16}$  Specifically, a higher risk of ischemic stroke incidence was associated with increasing nSES disadvantage (Carlsson et al.  $^{13}$ : Hazard Ratio (HR) = 1.19 [95% Confidence Interval (CI) = 1.08–1.30]; Honjo et al.  $^{16}$ : HR = 1.19 [95% CI = 1.01–1.41]), after controlling for education,  $^{13}$  and occupation  $^{16}$  as indicators of individual SES. The other 2 studies conducted in the United States showed an insignificant nSES-stroke incidence association after controlling

Table 2 Data Sources and Exposure Measures of the 8 Included Studies

Author (year)	Data source	Spatial scale of exposure	Exposure measure	Operationalization of exposure measure
Brown (2011) <sup>12</sup>	Cardiovascular Health Study (CHS)	Census Tract <sup>a</sup>	Neighborhood socioeconomic index [reference: the most advantaged neighborhoods]	Index of 6 items: (1) median household income; (2) median value of housing units; (3) percent households with interest, dividend, or rental income; (4) percent of residents ≥25 with a high school degree; (5) percent of residents ≥25 with a college degree; and (6) percent of residents in executive, managerial, or professional specialty occupations
Carlsson (2017) <sup>13</sup>	Nationwide Swedish Registers	Small Area Market Statistics <sup>b</sup>	Neighborhood-level deprivation [higher values indicate higher levels of neighborhood advantage]	Index of 4 items: (1) low education level (<10 years of formal education); (2) low income (income from all sources, including interest and dividends, i.e., <50% of the median individual income); (3) unemployment (excluding full-time students, those completing military service, and early retirees); and (4) receipt of social welfare
Clark (2011) <sup>31</sup>	Chicago Health and Aging Project (CHAP)	Census Block Group	Neighborhood socioeconomic status [higher values indicate higher levels of neighborhood advantage]	Index of 4 items: (1) percent on public assistance; (2) percent of household earning ≤ \$25,000 per year; (3) percent with college degree or more; and (4) percent o owner-occupied dwellings valued at ≥ \$200,000
Forsberg (2018) <sup>33</sup>	Nationwide Swedish Registers	Small Area Market Statistics <sup>b</sup>	Neighborhood socioeconomic status [higher values indicate higher levels of neighborhood deprivation]	Index of 4 items: (1) low education level; (2) low income (3) unemployment; and (4) receipt of social welfare
Freedman (2011) <sup>32</sup>	Health and Retirement Study (HRS)	Census Tract <sup>a</sup>	Neighborhood deprivation [higher values indicate higher levels of neighborhood deprivation]	Index of 6 items: (1) percentage of the total population in poverty; (2) the percentage of the population aged 65 years or older in poverty; (3) the percentage of households receiving public assistance income; (4) the unemployment rate among persons aged 16 years or older; (5) the percentage of housing units without a vehicle; and (6) percentage of the population that was Black
			Neighborhood economic advantage [higher values indicate higher levels of neighborhood advantage]	Index of 3 items: (1) upper quartile of the percentage of owner-occupied housing units in the tract; (2) the percentage of families with a total annual income of \$75,000 or more; and (3) the percentage of adults with a college degree
Honjo (2015) <sup>16</sup>	Japan Public Health Center-based Prospective Study (PHC)	Chocho- aza <sup>d</sup>	Area deprivation index [reference: the least deprived neighborhoods]	Index of 8 items: (1) proportion of elderly couple households; (2) proportion of elderly single households (3) proportion of single-mother households; (4) proportion of rent houses; (5) proportion of sales and service workers; (6) proportion of agricultural workers; (7) proportion of blue-collar workers; and (8) unemployment rate
Howard (2016) <sup>17</sup>	Reasons for Geographic and Racial Differences in Stroke (REGARDS)	Census Block Group <sup>c</sup>	Neighborhood socioeconomic status [reference: the most advantaged neighborhoods]	Index of 6 items: (1) log of the median household income, (2) log of the median value of owner-occupied housing units, (3) proportion of households receiving interest, dividend, or net rental income, (4) proportion o adults aged ≥25 years with a high school diploma, (5) proportion of adults aged ≥25 years with college degree and (6) proportion of people employed in executive, managerial, or professional occupations
Yan (2013) <sup>18</sup>	Cardiovascular Health Study (CHS)	Census Tract <sup>a</sup>	Neighborhood socioeconomic status [reference: the most advantaged neighborhoods]	Index of 6 items: (1) log of the median household income, (2) log of the median value of owner-occupied housing units, (3) proportion of households receiving interest, dividend, or net rental income, (4) proportion o adults aged ≥25 years with a high school diploma, (5) proportion of adults aged ≥25 years with college degree and (6) proportion of people employed in executive, managerial, or professional occupations

<sup>&</sup>lt;sup>a</sup> Census Tracts are created by the United States Census Bureau. Each Census Tract is designed to be relatively homogeneous units with respect to population characteristics, economic status, and living conditions and averages about 4,000 inhabitants.
<sup>b</sup> Small area market statistic (SAMS) is created by the Swedish Government-owned statistics bureau Statistics Sweden. Each SAMS contains an average of

<sup>&</sup>lt;sup>o</sup> Small area market statistic (SAMS) is created by the Swedish Government-owned statistics bureau Statistics Sweden. Each SAMS contains an average o 1,000 residents.

<sup>&</sup>lt;sup>c</sup> Census Block Groups are created by the United States Census Bureau. Each Census Block Group typically has a population of 600 to 3,000 people.

<sup>&</sup>lt;sup>d</sup> The Chocho-aza (CA) unit is the smallest administrative unit in Japan, which is roughly comparable to a European parish or a US block-group-size neighborhood.

Table 3 Outcome, Covariates, and Analysis of the 8 Included Studies

Author (year)	Outcome	Individual socioeconomic status	Covariates	Analysis
Brown (2011) <sup>12</sup>	Incident ischemic stroke (record linkage and adjudication through medical record review)	Educational level and income	Age, sex, and race	Multilevel Cox proportional hazards
Carlsson (2017) <sup>13</sup>	Incident ischemic stroke (record linkage)	Educational level	Age, sex, marital status, immigrant status, region of residence, and neighborhood goods and services	Cox proportional hazards
Clark (2011) <sup>31</sup>	Incident stroke (record linkage)	Educational level	Age, sex, race, systolic blood pressure, physical activity, smoking status, diabetes, chronic conditions, history of stroke, and neighborhood cohesion	Cox proportional hazards
Forsberg (2018) <sup>33</sup>	Incident ischemic stroke (record linkage)	Educational level and family income	Age, sex, and marital status	Cox proportional hazards
Freedman (2011) <sup>32</sup> , <sup>a</sup>	Incident stroke over the course of 2 years (self-report)	Educational level, mean assets, income, and childhood socioeconomic status	Age, sex, race/ethnicity, marital status, region, interview mode, smoking, childhood health, region of birth, high immigrant area, high crime and more segregation, residential stability, connectivity, air pollution, and density	Multilevel logistic
Honjo (2015) <sup>16</sup>	Stroke incidence (self-report and record linkage)	Occupation	Age, sex, Public Health Center district, and population density	Cox proportional hazards
Howard (2016) <sup>17</sup>	Incident stroke (self-report, medical record, and review by a stroke neurologist)	Educational level and household income	Age, sex, race, and region	Cox proportional hazards
Yan (2013)	Incident ischemic stroke (record linkage and adjudication through medical record review)	Educational level and income	Age, sex, race, and marital status	Cox proportional hazards

<sup>&</sup>lt;sup>a</sup> The study did not consider the time of onset throughout those 2 years.

for individual SES (education<sup>31</sup> and both education and income<sup>17</sup>).

#### **Effects Within Subgroups**

The fourth column of table 4 presents summarized findings of multivariate analyses stratified by sex, age, and/or race. Six of the 8 studies<sup>12,13,17,18,32,33</sup> examined the effects of nSES on stroke risk within subgroups of sex (n = 2), sex and age (n = 2)1), and race (n = 3). Of the 6 studies, 3 studies <sup>13,32,33</sup> stratified the data by sex. Carlsson et al. 13 reported that neighborhood disadvantage increased stroke risk among females (HR = 1.31 [95% CI = 1.14-1.51]) but not among males (HR = 1.09 [95% CI = 0.97-1.24]) but that neighborhood advantage decreased stroke risk among males (HR = 0.87 [95% CI = 0.76-0.99]) but not among females (HR = 0.87 [95% CI = 0.74–1.01]). The study by Forsberg et al.<sup>33</sup> stratified the data by sex and age groups of 40, 50, and 60 years to examine the nSES-stroke incidence association within age subgroups. The study found the same direction of association between nSES deprivation and stroke incidence across all the age groups and for both sexes with a greater risk of stroke incidence with greater nSES deprivation (men: HRs = 1.02-1.04; women: HRs = 1.04-1.07). Freedman et al.<sup>32</sup> observed no significant relationship between nSES disadvantage and stroke incidence among men (OR = 1.02 [95% CI = 0.85-1.22]) or women (OR = 1.11 [95% CI = 0.97-1.27]). Of note, Carlsson et al. <sup>13</sup>

and Forsberg et al.<sup>33</sup> controlled only for individual-level sociodemographic characteristics and/or region of residence, whereas Freedman et al.<sup>32</sup> controlled for several neighborhood measures (i.e., immigrant concentration, crime, segregation, residential stability, street connectivity, air pollution, and population density) along with individual-level demographic and socioeconomic characteristics. The neighborhood-level covariates might serve as mediators in the nSES-stroke incidence association and subsequently make nSES insignificant in the nSES-stroke incidence association. Overall, these results suggest that across these studies, there was no significant effect modification by sex in the relationship between nSES effects and stroke incidence.

Three studies <sup>12,17,18</sup> used the same indicators to measure nSES and stratified the data by race. Figure 2 shows that 2 studies <sup>12,18</sup> found that the nSES-stroke incidence association was in a different direction for Black and White Americans within each study, whereas the other <sup>17</sup> found the same direction observed among Black and White Americans. Specifically, Brown et al. <sup>12</sup> and Yan et al. <sup>18</sup> reported a significantly positive association between nSES deprivation and ischemic stroke incidence among White Americans (Brown et al. <sup>12</sup>: HR = 1.32 [95% CI = 1.01–1.72]; Yan et al. <sup>18</sup>: HR = 1.32 [95% CI = 1.01–1.73]) but reported an insignificant inverse association between nSES deprivation and ischemic stroke

**Table 4** Summarized Statistical Results About the Association Between nSES and Stroke Incidence in the 8 Included Studies

		Effect es	timate	Statistical test of the moderation effect	
Author (year)	Exposure	Total sample	Stratified sample	Moderator	Findings
Brown (2011) <sup>12</sup>	Neighborhood socioeconomic index [reference: the most advantaged neighborhoods]		Null (Black Americans) Positive (White Americans)		
Carlsson (2017) <sup>13</sup>	Neighborhood-level socioeconomic status [higher values indicate more advantaged neighborhoods]	Inverse	Inverse (men) Inverse (women)	Sex	Null
Clark (2011) <sup>31</sup>	nSES [higher values indicate more advantaged neighborhoods]	Null			
Forsberg (2018) <sup>33</sup>	nSES [higher values indicate more deprived neighborhoods]		Positive (men, age 40 y) Positive (men, age 50 y) Positive (men, age 60 y) Positive (women, age 40 y) Positive (women, age 50 y) Positive (women, age 60 y)		
Freedman (2011) <sup>32</sup>	Neighborhood deprivation [higher values indicate more deprived neighborhoods] Neighborhood economic advantage [higher values indicate more advantaged neighborhoods]		Null (men) Null (women) Null (men) Null (women)		
Honjo (2015) <sup>16</sup>	Area deprivation index [reference: the least deprived neighborhoods]	Positive			
Howard (2016) <sup>17</sup>	nSES [reference: the most advantaged neighborhoods]	Null	Null (Black Americans) Null (White Americans)	(1) Race (2) Race × sex	(1) Null (2) Null
Yan (2013)	nSES [reference: the most advantaged neighborhoods]		Null (Black Americans) Positive (White Americans)		

Abbreviation: nSES, neighborhood socioeconomic status

incidence among Black Americans (Brown et al.  $^{12}$ : HR = 0.60 [95% CI = 0.33–1.07]; Yan et al.  $^{18}$ : HR = 0.60 [95% CI = 0.35–1.05]). Of note, these 2 studies used the same data (the CHS) and had a smaller sample size for Black Americans (n = 785) than White Americans (n = 3,834), which may have led to the insignificant association in Black Americans. On the other hand, Howard et al.  $^{17}$  reported an insignificant positive association between neighborhood disadvantage and stroke incidence among Black (HR = 1.17 [95% CI = 0.76–1.80]) and White Americans (HR = 1.26 [95% CI = 0.94–1.70]). Of note, in the study by Howard et al.,  $^{17}$  the sample size for Black Americans was 10,274, and there was considerable variability in nSES within the study sample.

#### **Mediating Mechanisms**

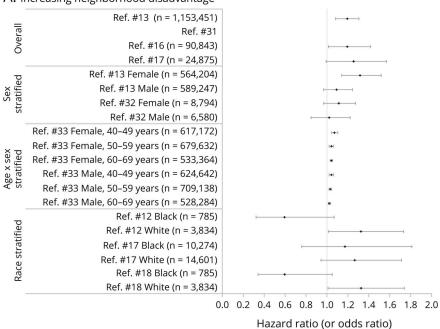
Table 5 shows summarized findings of mediation tests. Three studies <sup>12,16,18</sup> examined the mediation effect of biological factors. The studies <sup>12,16,18</sup> found that biological factors (e.g., subclinical cardiovascular disease) mediated the relationship between nSES and stroke risk. Of note, Brown et al. <sup>12</sup> used the method of Baron and Kenny<sup>34</sup> to decompose the total effect into indirect and direct effects, whereas Honjo et al. <sup>16</sup> and Yan et al. <sup>18</sup> assessed mediation through the inclusion of biological risk factors as covariates in the model and observing whether the association between nSES deprivation and stroke incidence was attenuated, which is a less desirable method of testing the mediation effect.

All 3 studies  $^{12,16,18}$  also examined the mediation effect of behavioral factors (n = 1), psychological factors (n = 1), or both (n = 1). Brown et al.  $^{12}$  and Yan et al.,  $^{18}$  who used the method of Baron and Kenny,  $^{34}$  found no significant mediation effect of behavioral or psychological factors. Honjo et al.  $^{16}$  found a possibility of the mediation effect of behavioral and psychological factors. Of note, as described above, Honjo et al.  $^{16}$  assessed mediation by observing whether the association between nSES deprivation and stroke incidence was attenuated with the inclusion of behavioral and psychological factors in the model.

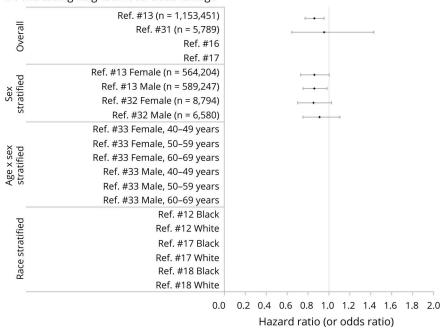
# Discussion

This review summarizes empirical evidence examining the relationship between neighborhood socioeconomic conditions and stroke incidence, along with mediating and moderating mechanisms of this relationship. Overall, neighborhood socioeconomic disadvantage is associated with increased stroke risk; however, the association was not significant in all the studies.  $^{13,16,17,31}$  Differences in significant findings are likely driven by differences in the sample size and socioenvironmental contexts. For example, there were large differences in the sample size between studies finding a significant association (N = 1,153,451 for Carlsson et al.  $^{13}$  and 90,843 for Honjo et al.  $^{16}$ ) and studies finding a nonsignificant





# B. Increasing neighborhood disadvantage



(A) Presents (1) the coefficients of least advantaged neighborhoods (vs. moderately advantaged neighborhoods) from the article by Carlsson et al., <sup>13</sup> (2) the coefficients of most deprived neighborhoods (vs least deprived neighborhoods) from Honjo et al., 16 and (3) the coefficients of least advantaged neighborhoods (vs most advantaged neighborhoods) from Howard et al., <sup>17</sup> Brown et al., <sup>12</sup> and Yan et al. <sup>18</sup> From the studies by Freedman et al.32 and Forsberg et al., 33 we presented coefficients of a continuous variable of neighborhood disadvantage, and its higher scores indicate more disadvantaged neighborhoods. (B) Presents (1) the coefficients of the most advantaged neighborhoods (vs moderately advantaged neighborhoods) from the article by Carlsson et al. 13 and (2) the coefficients of a continuous variable of neighborhood advantage, and its higher scores indicate more advantaged neighbor-hoods. From the study by Clark et al.,<sup>31</sup> we presented coefficients of a continuous variable of neighborhood advantage, and its higher scores indicate more advantaged neighborhoods.

association (N = 24,875 for Howard et al.<sup>17</sup> and 5,789 for Clark et al.<sup>31</sup>). It may be that the studies with larger numbers of participants were better powered to find a significant result. In addition, Carlsson et al.<sup>13</sup> and Honjo et al.<sup>16</sup> conducted their studies in the socioenvironmental contexts of Sweden and Japan, respectively, where all residents have health insurance coverage under a universal health care system. On the other hand, Clark et al.<sup>31</sup> and Howard et al.<sup>17</sup> conducted their studies in the US socioenvironmental contexts, where health

insurance is made up by a complex market of public and private providers. Different health insurance coverage and utilization of health services may be uncontrolled covariates in the United States, which should be considered in future research.

Numerous studies included in this review presented stratified results, presenting differences in the impact of nSES on stroke incidence by sex, <sup>13,32,33</sup> sex and age, <sup>33</sup> and race. <sup>12,17,18</sup> Overall,

**Table 5** Summarized Statistical Results of Mediation Tests in 3 Studies Examining the Mediating Role of Biological and/or Behavioral Factors in the Association Between nSES and Stroke Incidence

Author (year)	Mediators	Findings
Brown (2011) <sup>12</sup>	(1) Biologic risk factors: subclinical cardiovascular disease, atrial fibrillation, hypertension, diabetes, and total cholesterol/HDL (2) Behavioral risk factors: smoking, alcohol use, and physical activity	(1) Higher incidence of ischemic stroke among Whites in the lowest-SES neighborhoods was mediated through higher levels of biological risk among Whites (2) Null
Honjo (2015) <sup>16</sup>	(1) Cardiovascular disease biological risk factors: overweight and medical history of hypertension/diabetes/hyperlipidemia (2) Behavioral and psychosocial factors: smoking, ethanol intake, physical activity, perceived psychological stress, and marital status	(1) Relationship between neighborhood deprivation and incidence stroke was attenuated when including biological risk factors in the model (2) Relationship between neighborhood deprivation and incidence stroke was attenuated when behavioral and psychosocial factors were included in the model
Yan (2013) <sup>18</sup>	(1) Biological risk factors: atrial fibrillation, subclinical CVD, hypertension, diabetes, and hyperlipidemia (2) Psychosocial factors: depressive symptoms, social support, and social networks	(1) Relationship between neighborhood deprivation and incidence stroke was attenuated and became insignificant when biological risk factors were added (2) Null

To assess the mediating mechanism, Brown et al.<sup>12</sup> used the Baron and Kenny<sup>34</sup> approach to decompose the total effect into indirect and direct effects. Honjo et al.<sup>16</sup> included potential mediating variables as covariates in their final model and observed whether the association between nSES deprivation and stroke incidence was attenuated. Yan et al.<sup>18</sup> used the Baron and Kenny<sup>34</sup> approach to test the mediating effect of psychosocial factors; however, the study<sup>18</sup> observed whether the association between nSES deprivation and stroke incidence attenuated with the inclusion of behavioral and psychological factor in the model.

Abbreviation: nSES, neighborhood socioeconomic status

the studies included found no differential effect by sex and age. Regarding racial differences, findings across studies are inconsistent. Studies 12,18 using CHS data found racial differences in the nSES-stroke incidence association, whereas a study<sup>17</sup> using REGARDS data found no racial difference in the association. All 3 of these studies used the same operationalization of nSES exposure but differed in the spatial scale in which it was measured and the data set it was examined within. Brown et al. 12 and Yan et al. 18 both measured nSES at the census tract level and observed that less advantaged neighborhoods increased stroke risk among Whites but not Black Americans in the CHS. Howard et al. 17 measured nSES at the block group level and did not observe differential effects by race in the REGARDS study. Although previous studies suggested the differences in census tract and block groups measures to be relatively small, 35,36 the role of spatial scale in this relationship is still largely unknown. Further investigation is needed to determine whether the relationship between nSES and stroke incidence is differential by race.

The inconsistency in findings by race could be further explored by considering the role of various features of neighborhood environments in the racial disparity in stroke incidence. The literature demonstrated neighborhood racial/ethnic composition as a potential explanation for racial disparities in health. The example, Bond Huie et al. The found that neighborhood racial/ethnic composition was a stronger driver of Black-White disparity in mortality than neighborhood socioeconomic characteristics. In addition, the literature suggests that systemic racism and stress from daily experiences of racism increase the risk of adverse health outcomes and racial disparities in health. Thus, the direct effects of neighborhood racial/ethnic composition and racially discriminatory experiences warrant further examination.

This review identified methodological limitations in the 8 studies and suggestions for future research. Measurement of nSES was inconsistent across studies. Studies used different composite measures of nSES identified through factor analysis, 12,17,18 principal components analysis, 13,33 confirmatory factor analysis, 31 exploratory factor analysis, 32 or an author-derived index. 16 In addition, there were nSES indices that included measures from area-level indicators that do not directly indicate or represent nSES conditions (e.g., age 16 or race<sup>32</sup> of resident) along with other items that directly measure nSES. Finally, all 8 studies captured absolute measures of nSES resource, rather than a relative nSES measure such as within-neighborhood income inequality and betweenneighborhood income segregation. A relative nSES measure assesses the distribution of socioeconomic resource opportunities 41-44; thus, examining relative nSES measures may capture a different aspect of nSES conditions relative to an absolute nSES measure, such as psychological stress from highly unequal resource distribution.

Considering the residential mobility of survey participants after the baseline (i.e., moving to a new home) is warranted. Statistics showed that 18% of US adults aged 55–74 years<sup>45</sup> and 18–24% of Swedish adults aged 57–66 years<sup>46</sup> moved at least once in 5 years. However, all studies included in this review did not account for individuals' residential mobility and, thus, implicitly assumed equal neighborhood effects for the differential durations lived in the baseline neighborhood.<sup>47</sup> For example, Brown et al.<sup>12</sup> and Yan et al.<sup>18</sup> obtained nSES information at the baseline (1989–1993) and examined its effect on stroke incidence collected in the follow-ups (until 2006, every 6 months). Thus, a longitudinal lens on neighborhood circumstances is warranted to estimate neighborhood impacts on stroke incidence.

More than half of the studies relied on 1 measure of stroke incidence. <sup>12,13,31-33</sup> In particular, 1 study relied solely on self-report stroke incidence. <sup>32</sup> Self-reported stroke may result in information bias of the outcome, whereby those with stroke-like symptoms may self-report stroke, resulting in a false-positive report. In addition, participants who experience a stroke may have severe cognitive impairment and could be lost to follow-up or report stroke status incorrectly. <sup>48</sup> Because of these potential sources of bias, we urge additional caution when interpreting the effect estimates relying solely on self-report stroke incidence. Finally, some demographic groups have not been investigated. Future research needs to examine diverse demographic subgroups at risk such as Hispanics and transgender individuals.

This review has several strengths. First, we used several strategies to identify eligible studies: a search of 4 databases, an extensive search of reference lists, and contacting the corresponding authors of studies included in this review. Second, we reviewed a wide range of nSES factors from education, income, and occupation to welfare receipt and wealth. We also examined the nSES-stroke incidence association independent of individual SES and its mediating and moderating mechanism, which have not been explored in previous review studies. <sup>19,20</sup>

However, our study is not without limitations. First, we included only quantitative studies. Qualitative research can give a better understanding of how people perceive and experience nSES affecting their stroke risk. Study publication bias and outcome reporting bias might have occurred. In addition, this review focused on nSES; however, there are other important attributes of neighborhood environments affecting individual health, such as racial/ethnic composition, racial residential segregation, and built and social environments. Finally, we were unable to conduct a meta-analysis because the included studies measured nSES using different items, adjusted for a different set of covariates, and conducted different analyses ranging from multilevel logistic regression to Cox proportional hazards. As we move the field forward, it is advantageous for researchers to have consistency in the measurement of the neighborhood socioeconomic context. Because of the lack of consistency in study methodology (e.g., measurement and covariate adjustment), we were unable to explain differences between study findings. To facilitate comparisons in the United States in future studies, we recommend that research on nSES uses the 6-item measure at the census tract level validated through factor analysis. <sup>49</sup> Testing the validity of the 6-item measure outside the United States context is also warranted. Consistency in operationalization will allow for summarization of effect estimates across studies. Furthermore, we suggest a step-wise model building strategy that includes covariate adjustment for age, sex, and race/ethnicity, followed by adjustment for individual SES and, finally, inclusion of the 6-item nSES measure. Finally, we suggest future studies to use Cox proportional hazards and present an overall effect estimate in the full sample, which would allow for comparisons of effect estimates across studies.

In conclusion, this review summarized 8 articles examining the nSES-stroke risk association with adjustment for individual SES. The association between neighborhood disadvantage and stroke risk was found above and beyond the individual SES, but it was not significant in all the studies. <sup>13,16,17,31</sup> Research examining the moderating effect of race on the nSES-stroke risk association was inconsistent. <sup>12,17,18</sup> Future research needs to use a standard operational definition of neighborhood and nSES measures to allow for better comparability across studies. Furthermore, additional research is needed to investigate other neighborhood characteristics and biological, behavioral, and psychological factors as potential intermediate and modifiable mediators of the nSES-stroke risk relationship that could serve as policy and program interventions.

# **Study Funding**

Research reported in this publication was supported by the National Institute of Neurological Disorders and Stroke of the NIH under award number 1R01NS092706 and the Eunice Kennedy Shriver National Institute of Child Health and Human Development of the NIH under award number F31HD098870. The content is solely the responsibility of the authors and does not necessarily represent the official views of the NIH.

#### **Disclosure**

Y. Kim reports funding from the Eunice Kennedy Shriver National Institute of Child Health & Human Development of the US NIH (R03HD101752). E. Twardzik reports funding from the Eunice Kennedy Shriver National Institute of Child Health & Human Development of the NIH (F31HD098870, E. Twardzik, PI). S.E. Judd reports funding from the US NIH (1 U01 HL146382, 2 U01 NS041588, 1R01NS09270 1AG 057540, and 5R01HL136666) and the Centers for Disease Control and Prevention 1U01DP006302-01. N. Colabianchi reports funding from the US NIH (R03HD101752, R01HL137731, R01NS092706, R01AG057540, and R01 HL132979). Go to Neurology.org/N for full disclosures.

# **Publication History**

Received by Neurology July 15, 2020. Accepted in final form February 2, 2021.

Aр	pend	ix /	∖uth	nors
----	------	------	------	------

Author	Location	Contribution
Yeonwoo Kim, PhD	University of Texas at Arlington, TX, USA	Designed and conceptualized the study, interpreted the data, and drafted the manuscript for intellectual content
Erica Twardzik, MS	University of Michigan, MI, USA	Designed and conceptualized the study, interpreted the data, and drafted the manuscript for intellectual content

#### Appendix (continued)

Author	Location	Contribution
Suzanne E. Judd, PhD	University of Alabama at Birmingham, AL, USA	Revised the manuscript for intellectual content
Natalie Colabianchi, PhD	University of Michigan, MI, USA	Interpreted the data and revised the manuscript for intellectual content

#### References

- Marshall IJ, Wang Y, Crichton S, McKevitt C, Rudd AG, Wolfe CD. The effects of socioeconomic status on stroke risk and outcomes. Lancet Neurol 2015;14: 1206–1218
- Howard VJ, Kleindorfer DO, Judd SE, et al. Disparities in stroke incidence contributing to disparities in stroke mortality. Ann Neurol 2011;69:619–627.
- Kleindorfer D, Khoury J, Moomaw CJ, et al. Stroke incidence is decreasing in whites, but not in blacks: a population-based estimate of temporal trends in stroke incidence from the Greater Cincinnati/Northern Kentucky stroke study. Stroke 2010;41:1326.
- Agyemang C, van Oeffelen AA, Norredam M, et al. Socioeconomic inequalities in stroke incidence among migrant groups: analysis of nationwide data. Stroke 2014;45: 2397–2403.
- Avendano M, Kawachi I, Van Lenthe F, et al. Socioeconomic status and stroke incidence in the US elderly: the role of risk factors in the EPESE study. Stroke 2006; 37:1368–1373.
- Cesaroni G, Agabiti N, Forastiere F, Perucci CA. Socioeconomic differences in stroke incidence and prognosis under a universal healthcare system. Stroke 2009;40: 2812–2819
- Bray BD, Paley L, Hoffman A, et al. Socioeconomic disparities in first stroke incidence, quality of care, and survival: a nationwide registry-based cohort study of 44 million adults in England. Lancet Public Health 2018;3:e185–e193.
- Bowling A, Stafford M. How do objective and subjective assessments of neighbourhood influence social and physical functioning in older age? Findings from a British survey of ageing. Soc Sci Med 2007;64:2533–2549.
- Franzini L, Caughy M, Spears W, Esquer MEF. Neighborhood economic conditions, social processes, and self-rated health in low-income neighborhoods in Texas: a multilevel latent variables model. Soc Sci Med 2005;61:1135–1150.
- Sampson RJ, Morenoff JD, Gannon-Rowley T. Assessing "neighborhood effects": social processes and new directions in research. Annu Rev Sociol 2002;28:443

  –478.
- Wen M, Browning CR, Cagney KA. Poverty, affluence, and income inequality: neighborhood economic structure and its implications for health. Soc Sci Med 2003; 57:843–860.
- Brown AF, Liang L-J, Vassar SD, et al. Neighborhood disadvantage and ischemic stroke: the cardiovascular health study (CHS). Stroke 2011;42:3363–3368.
- Carlsson AC, Li X, Holzmann MJ, et al. Neighborhood socioeconomic status at the age of 40 years and ischemic stroke before the age of 50 years: a nationwide cohort study from Sweden. Int J Stroke 2017;12:815–826.
- Grimaud O, Béjot Y, Heritage Z, et al. Incidence of stroke and socioeconomic neighborhood characteristics: an ecological analysis of Dijon stroke registry. Stroke 2011;42:1201–1206.
- Heeley EL, Wei JW, Carter K, et al. Socioeconomic disparities in stroke rates and outcome: pooled analysis of stroke incidence studies in Australia and New Zealand. Med J Aust 2011;195:10–14.
- Honjo K, Iso H, Nakaya T, et al. Impact of neighborhood socioeconomic conditions on the risk of stroke in Japan. J Epidemiol 2015;25:254–260.
- Howard VJ, McClure LA, Kleindorfer DO, et al. Neighborhood socioeconomic index and stroke incidence in a national cohort of blacks and whites. Neurology 2016;87: 2340–2347.
- Yan T, Escarce JJ, Liang LJ, et al. Exploring psychosocial pathways between neighbourhood characteristics and stroke in older adults: the cardiovascular health study. Age and ageing 2013;42:391–397.
- Addo J, Ayerbe L, Mohan KM, et al. Socioeconomic status and stroke: an updated review. Stroke 2012;43:1186–1191.

- Backholer K, Peters SA, Bots SH, Peeters A, Huxley RR, Woodward M. Sex differences in the relationship between socioeconomic status and cardiovascular disease: a systematic review and meta-analysis. J Epidemiol Community Health 2017;71:550–557.
- Hedman L, van Ham M. Understanding Neighbourhood Effects: Selection Bias and Residential Mobility. Springer Netherlands; 2011:79–99.
- Diez Roux AV. Investigating neighborhood and area effects on health. Am J Public Health 2001;91:1783–1789.
- Shamseer L, Moher D, Clarke M, et al. Preferred reporting items for systematic review and meta-analysis protocols (PRISMA-P) 2015: elaboration and explanation. Bmj 2015;350:g7647.
- Higgins JP, Thomas J, Chandler J, et al. Cochrane Handbook for Systematic Reviews of Interventions: John Wiley & Sons, 2019.
- Littell JH, Corcoran J, Pillai V. Systematic Reviews and Meta-Analysis: Oxford University Press; 2008.
- Slavin RE. Best evidence synthesis: an intelligent alternative to meta-analysis. J Clin Epidemiol 1995;48:9–18.
- World Bank. World Bank Country and Lending Groups [online]. Available at: data. worldbank.org/about/country-and-lending-groups#High\_income. Accessed May 19, 2020.
- Minh A, Muhajarine N, Janus M, Brownell M, Guhn M. A review of neighborhood effects and early child development: how, where, and for whom, do neighborhoods matter? Health place 2017;46:155–174.
- Calling S, Li X, Kawakami N, Hamano T, Sundquist K. Impact of neighborhood resources on cardiovascular disease: a nationwide six-year follow-up. BMC public health 2016;16:634.
- Kulick ER, Wellenius GA, Boehme AK, Sacco RL, Elkind MS. Residential proximity to major roadways and risk of incident ischemic stroke in NOMAS (The Northern Manhattan Study). Stroke 2018;49:835–841.
- Clark CJ, Guo H, Lunos S, et al. Neighborhood cohesion is associated with reduced risk of stroke mortality. Stroke 2011;42:1212–1217.
- Freedman VA, Grafova IB, Rogowski J. Neighborhoods and chronic disease onset in later life. Am J Public Health 2011;101:79–86.
- Forsberg P-O, Ohlsson H, Sundquist K. Causal nature of neighborhood deprivation on individual risk of coronary heart disease or ischemic stroke: a prospective national Swedish co-relative control study in men and women. Health Place 2018;50:1–5.
- Baron RM, Kenny DA. The moderator-mediator variable distinction in social psychological research: conceptual, strategic, and statistical considerations. J Personal Soc Psychol 1986;51:1173.
- Krieger N, Chen JT, Waterman PD, Rehkopf DH, Subramanian SV. Race/ethnicity, gender, and monitoring socioeconomic gradients in health: a comparison of areabased socioeconomic measures—the public health disparities geocoding Project. Am J Public Health (1971) 2003;93:1655–1671.
- Lovasi GS, Moudon AV, Smith NL, et al. Evaluating options for measurement of neighborhood socioeconomic context: evidence from a myocardial infarction casecontrol study. Health place 2008;14:453–467.
- Anderson NB, Bulatao RA, Cohen B, on Race P, Council NR. What makes a place healthy? Neighborhood influences on racial/ethnic disparities in health over the life course. In: Critical Perspectives on Racial and Ethnic Differences in Health in Late Life: National Academies Press (US); 2004.
- Huie B, Hummer RA, Rogers RG. Individual and contextual risks of death among race and ethnic groups in the United States. J Health Soc Behav 2002;359–381.
- Shaw G. It's a public health crisis: how systemic racism can Be neurotoxic for black Americans. Neurol Today 2020;20:1–24.
- Phelan JC, Link BG. Is racism a fundamental cause of inequalities in health? Annu Rev Sociol 2015;41:311–330.
- Bernstein M, Crosby F. An empirical examination of relative deprivation theory. J Exp Soc Psychol 1980;16:442–456.
- Guimond S, Dubé-Simard L. Relative deprivation theory and the Quebec nationalist movement: the cognition–emotion distinction and the personal–group deprivation issue. J Personal Soc Psychol 1983;44:526.
- Podder N. Relative deprivation, envy and economic inequality. Kyklos 1996;49: 353–376.
- 44. Yitzhaki S. Relative deprivation and the gini coefficient. Q J Econ 1979;93:321-324.
- 45. U.S. Bureau of the Census. Geographical Mobility: 2010 to 2015 [data File]; 2018.
- Andersson E, Abramsson M. Changing residential mobility rates of older people in Sweden. Ageing Soc 2012;32:963.
- Sheehan CM, Cantu PA, Powers DA, Margerison-Zilko CE, Cubbin C. Long-term neighborhood poverty trajectories and obesity in a sample of California mothers. Health place 2017;46:49–57.
- Kalaria RN, Akinyemi R, Ihara M. Stroke injury, cognitive impairment and vascular dementia. Biochim Biophys Acta Mol basis Dis 2016;1862:915–925.
- Diez-Roux AV, Merkin SS, Arnett D, et al. Neighborhood of residence and incidence of coronary heart disease. N Engl J Med 2001;345:99–106.