



# Mapping partner drug resistance to guide antimalarial combination therapy policies in sub-Saharan Africa

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**Resistance to artemisinin-based combination therapies (ACTs) threatens the global control of *Plasmodium falciparum* malaria. ACTs combine artemisinin-derived compounds with partner drugs to enable multiple mechanisms of clearance. Although ACTs remain widely effective in sub-Saharan Africa, long-standing circulation of parasite alleles associated with reduced partner drug susceptibility may contribute to the development of clinical resistance. We fitted a hierarchical Bayesian spatial model to data from over 500 molecular surveys to predict the prevalence and frequency of four key markers in transporter genes (*pfcr* 76T and *pfmdr1* 86Y, 184F, and 1246Y) in first-level administrative divisions in sub-Saharan Africa from the uptake of ACTs (2004 to 2009) to their widespread usage (2010 to 2018). Our models estimated that the *pfcr* 76T mutation decreased in prevalence in 90% of regions; the *pfmdr1* N86 and D1246 wild-type genotypes increased in prevalence in 96% and 82% of regions, respectively; and there was no significant directional selection at the *pfmdr1* Y184F locus. Rainfall seasonality was the strongest predictor of the prevalence of wild-type genotypes, with other covariates, including first-line drug policy and transmission intensity more weakly associated. We lastly identified regions of high priority for enhanced surveillance that could signify decreased susceptibility to the local first-line ACT. Our results can be used to infer the degree of molecular resistance and magnitude of wild-type reversion in regions without survey data to inform therapeutic policy decisions.**

malaria | drug resistance | surveillance

Artemisinin-based combination therapies (ACTs) are recommended as the first-line treatment for uncomplicated *Plasmodium falciparum* malaria worldwide (1). ACTs combine potent artemisinin-derived compounds and partner drugs with longer half-lives, allowing for dual mechanisms of action and prolonged antiparasitic action. Of the ACTs currently recommended by the World Health Organization, artemether-lumefantrine (AL) and artesunate-amodiaquine (AS-AQ) are the most commonly adopted first-line ACTs in sub-Saharan Africa (1). While efficacy remains high, recent reports of mutations associated with reduced artemisinin sensitivity in Rwanda and long-standing circulation of alleles associated with bidirectional impacts on partner drug susceptibility raise concern that ACT resistance may arise and spread in sub-Saharan Africa (2–4).

Molecular markers are single nucleotide polymorphisms (SNPs), deletions, or copy number variations in the parasite genome that are associated with reduced susceptibility to antimalarial drugs. SNPs along two key transporter genes, *P. falciparum* chloroquine resistance transporter (*pfcr*) and *P. falciparum* multidrug resistance 1 (*pfmdr1*), were first discovered to confer resistance to antimalarial drugs in the 1990s, primarily *pfcr* 76T to chloroquine, and have since been implicated in impacting susceptibility to AL and AS-AQ (5). Notably, the partner drugs lumefantrine and AQ exert opposing selective pressures: parasites with genotype *pfmdr1* 86Y, Y184, 1246Y, and *pfcr* 76T have reduced susceptibility to AQ, while *pfmdr1* N86, 184F, D1246, and *pfcr* K76 confer reduced susceptibility to lumefantrine (5–7). While data are strongest for

partner drugs, artemisinins may also directly exert a selective pressure at *pfmdr1* and other loci (8, 9).

In southeast Asia, partner drug resistance arose rapidly in the wake of, and in some regions independently of, artemisinin resistance (10). In sub-Saharan Africa, it has been suggested that partner drug failure is expected to result in greater increases in malaria morbidity than would be observed from artemisinin resistance alone (4). The population prevalence of markers in *pfcr* and *pfmdr1* can indicate the level of clinical resistance or tolerance to the partner drugs AQ and lumefantrine (11). When used as a surveillance tool, molecular marker data can rapidly provide data to support modifications of local partner drugs and to sustain artemisinin-based treatment and prevention options. Despite broad consensus to scale up molecular surveillance for antimalarial resistance, few studies have proposed systematically designed approaches such as where sentinel sites should be located or how often sampling should be conducted (12, 13).

We previously developed a database of studies assessing prevalence of partner drug resistance-associated markers in sub-Saharan Africa, supplementing an existing database by the Worldwide Antimalarial Resistance Network (14, 15). That work revealed important gaps in the geographic coverage of surveillance. We hypothesized that the geographic landscape of mutations associated

## Significance

Antimalarial resistance has emerged and spread with every antimalarial deployed to date. Currently, parasite genotypes associated with reduced artemisinin and partner drug sensitivity have been reported in Asia, South America, and, most recently, Africa. Analyzing spatial-temporal trends in molecular markers can help policymakers choose efficacious partner drugs and slow the emergence of artemisinin resistance and spread of multidrug-resistant parasites. We display evidence of a continent-wide increase in molecular markers associated with reduced lumefantrine susceptibility, the partner drug of the most widely used artemisinin-based combination therapy in sub-Saharan Africa. We also generate hypotheses for large-scale demographic and environmental risk factors implicated in the spread of antimalarial resistance. Our results can help identify regions of developing parasite resistance that may require enhanced surveillance.

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with partner drug resistance exhibit spatial patterns that can be mapped and leveraged to determine sites needing additional surveillance. In this study, we use a hierarchical Bayesian spatial method to map the geographical distribution of key molecular markers in *pfprt* and *pfmdr1* in sub-Saharan Africa since the uptake of ACTs and analyze potential demographic and environmental determinants of resistance landscapes. We also identify potential regions for enhanced surveillance based on the magnitude and direction of changes in genotypes over time, suggesting areas that may pose an increased risk of artemisinin emergence or spread (16). Our work will assist in strategic surveillance efforts to sustain first-line ACT partner drug efficacy in sub-Saharan Africa and inform future surveillance as new antimalarials are deployed and additional molecular markers are identified.

## Results

**Geographic Distribution of Mutant Genotypes.** We generated maps of marker prevalence and uncertainty interpolated across the malaria-endemic subcontinent from 2004 to 2009 and from 2010 to 2018 (Fig. 1). The mutations at *pfprt* 76T and *pfmdr1* 86Y were extensively prevalent in 2004 to 2009 and much more sparsely present in 2010 to 2018 (Table 1), with *pfmdr1* 86Y prevalent only in the Gulf of Guinea region. The *pfmdr1* 1246Y genotype exhibited the highest degree of prevalence in central and east Africa, as well as Côte d'Ivoire and South Africa, compared to all other regions from 2004 to 2009, but low prevalence across the continent in the latter time period. Meanwhile, *pfmdr1* 184F exhibited more randomness over time and space than the other markers, concentrated in the northwest from 2004 to 2009 and the northeast from 2010 to 2018 (Table 1). *Pfprt* 76T estimates displayed the highest levels of uncertainty compared to other markers in the later time period, with isolated hotspots of uncertainty on the mainland and high uncertainty in Madagascar. Because of the similarity in regional trends and contributions of covariates between frequency and prevalence estimates (SI Appendix, Figs. S4 and S5) and superior prevalence data coverage, additional analyses involved prevalence estimates only.

**Temporal Reversion to Wild-Type Genotypes.** The mean estimated prevalence of mutant genotypes decreased significantly from 2004 to 2009 to 2010 to 2018 for *pfmdr1* 86Y and 1246Y and *pfprt* 76T [ $P(\mu_{p_{2004-2009}} > \mu_{p_{2010-2018}} | data) = 1.0$  for all three markers (SI Appendix, Fig. S6)], with the greatest magnitude of change for *pfmdr1* 86Y (Table 1). West and southern Africa exhibited the smallest estimated changes in prevalence of all four markers. There were noticeable exceptions to wild-type reversion: For example, *pfprt* 76T prevalence increased, according to our predictions, in parts of Niger, Nigeria, Djibouti, and Madagascar, although such areas tended to exhibit higher uncertainty. The distribution of *pfmdr1* 184F changed markedly over the two time periods, yet we found no evidence of directional selection toward the wild-type *pfmdr1* Y184, suggesting more random and less spatially explicit patterns of selection [ $P(\mu_{p_{2004-2009}} > \mu_{p_{2010-2018}} | data) = 0.51$ ]. The magnitude and direction of temporal changes for each administrative division can be visualized in SI Appendix, Fig. S6. The prevalence of *pfprt* 76T was positively correlated with *pfmdr1* 86Y, which, in turn, was associated with *pfmdr1* Y184 and *pfmdr1* 1246Y. Correlations between changes in prevalence for each pair of markers can be found in SI Appendix, Fig. S7.

**Contributions of Covariates to Genotypic Variability.** Seasonality was consistently the strongest variable associated with the presence of wild-type alleles *pfprt* K76, *pfmdr1* N86, and *pfmdr1* D1246 and of the mutant allele *pfmdr1* 184F (Fig. 2 and SI Appendix, Fig. S4) in both time periods. From 2004 to 2009, transmission intensity of *P. falciparum* was associated with prevalence of markers *pfprt* K76 and *pfmdr1* N86, 184F, and D1246, while the transmission

intensity of *Plasmodium vivax* was associated with *pfprt* 76T. From 2010 to 2018, city accessibility and first-line drug policy were correlated with the prevalence of *pfprt* K76, *pfmdr1* N86, *pfmdr1* D1246, and *pfmdr1* 184F alleles, with AS-AQ and both/other drug regimens generally associated with the presence of opposing alleles compared to AL (Fig. 2).

We also estimated the impacts of covariates on the magnitude and direction of changes in marker prevalence over time (SI Appendix, Fig. S8). Areas with higher rainfall seasonality exhibited smaller changes in prevalence over time for all four markers after controlling for first-line drug policy and other variables. First-line drug policy was associated with the magnitude of change of most markers, although with much greater bounds of uncertainty: Compared to AL, AS-AQ was associated with a greater magnitude of change toward mutant genotypes *pfprt* 76T and *pfmdr1* 86Y, and the use of multiple or other first-line therapies was associated with smaller changes toward wild-type genotypes *pfmdr1* Y184 and *pfmdr1* D1246. When stratifying the results by first-line drug regimen, trends were more difficult to parse out (SI Appendix, Table S2 and Fig. S9).

**Regions for Enhanced Surveillance.** We identified numerous regions that could be prioritized for enhanced surveillance (Fig. 3) based on the relative strength and direction of change in prevalence over time and the degree of uncertainty in model predictions. Higher-priority regions were located in parts of the Central African Republic, Ethiopia, Tanzania, Gabon, Equatorial Guinea, Sierra Leone, Nigeria, Togo, and Mali. A specific set of administrative divisions demonstrated the highest posterior probability of selection (surveillance prioritization) across all model iterations, suggesting greater certainty in the relative strength of directional selection over time (SI Appendix, Fig. S10). These administrative divisions include Luapula Province, Zambia; Litoral Province, Equatorial Guinea; Southern Nations, Nationalities, and People's Region, Ethiopia; Pointe-Noire Department, Republic of the Congo; and Edo State, Nigeria, with Pointe-Noire having the highest probability of selection for enhanced surveillance compared to any other region. These results may be partly reflective of higher sampling density in these subregions.

## Discussion

Our study demonstrates the striking change in prevalence of key partner drug susceptibility-associated molecular markers in sub-Saharan Africa since the introduction of ACTs and provides an attempt to identify focal regions to monitor for decreased partner drug susceptibility. We found a strong wild-type resurgence of *pfprt* K76, likely reflecting selection due to the combined impacts of the uptake of AL and withdrawal of chloroquine monotherapy, as evidenced by many studies at smaller spatial scales or in imported parasites (17, 18). We also developed spatially smoothed maps demonstrating wild-type selection of *pfmdr1* N86 and D1246 across the continent, supporting a study by Okell et al. (19) that also found evidence of widespread wild-type reversion of these markers. Together, our results suggest that surveillance can be more effective if not only coordinated at a national scale but also taking into account regional spatial dependencies.

We identified intriguing areas of remaining *pfprt* 76T and *pfmdr1* 86Y genotype presence, particularly in Madagascar, Ethiopia, the Gulf of Guinea, and parts of West Africa. One potential reason for the persistence in Madagascar and Ethiopia may be related to the usage of chloroquine to treat *P. vivax* where the species remains endemic (1, 20). In the Gulf of Guinea, a region where seasonal malaria chemoprevention (SMC) with AQ and sulfadoxine-pyrimethamine (SP) is used in children and higher rates of AS-AQ are used for treatment than in many other parts of Africa, the added pressure of AQ may explain the higher prevalence of these alleles (1). However, because the widespread administration of SMC began in the latter half of the second time interval, SMC was

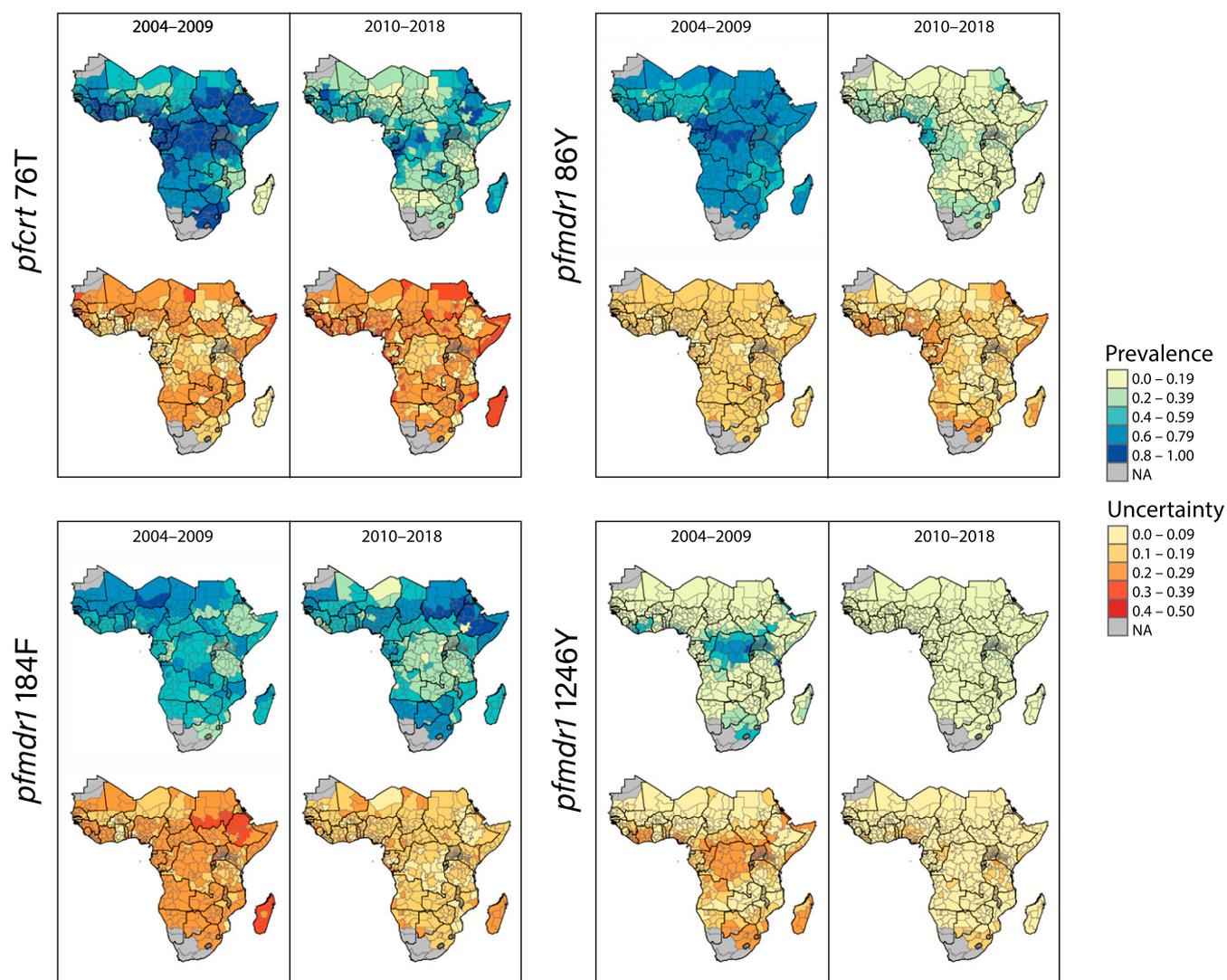


Fig. 1. Maps displaying model estimates and predictions for molecular marker prevalence in first-level administrative divisions from 2004 to 2009 and 2010 to 2018 and respective posterior SDs representing uncertainty.

not included as a covariate. In all of these regions, the continued use of chloroquine to treat *P. falciparum* may also drive mutation persistence, but rigorous data on consumption of chloroquine are not available at this scale (21–23).

We also assessed the contribution of socioecological factors that may help explain drug resistance heterogeneity, aided by the additional power in scale of this analysis. Wild-type reversion of *pfct* K76, *pfmdr1* N86, and *pfmdr1* D1246 occurred across sub-Saharan Africa regardless of first-line therapy, yet countries using AS-AQ or multiple first-line therapies tended to exhibit a smaller magnitude of reversion overall. Areas with high seasonality strongly favored wild-type genotypes, but again, the extent of the reversion to wild type was partially lessened by the use of AQ. Our results support strong evidence that *pfct* 76T confers a fitness cost to the parasite and add to a growing body of evidence of a fitness cost of *pfmdr1* 86Y and 1246Y (24–26). In the dry season, these mutations may lose their competitive advantage in the absence of sustained drug pressure (27, 28). We also found that higher transmission intensity favored *pfct* K76, *pfmdr1* N86, and *pfmdr1* D1246 in the earlier time period, which may be related to higher rates of recombination in circulating parasites, increased levels of immunity in high transmission regions, and heterogeneity in the proportion of

infections that encounter treatments (29–31). Our results ultimately suggest that environmental and epidemiologic factors may be of increased importance, in addition to national policy, in delineating factors to consider when conducting surveillance.

The covariates in our models and their relationship to drug resistance have important caveats. Our results are subject to ecological fallacy due to the large geographic scale of the analysis. In particular, because of limited data on antimalarial consumption patterns available across the subcontinent, multiple variables acted as potential proxies for detailed drug consumption patterns at the level of the administrative division, including first-line drug policy, ACT coverage, insecticide-treated bed net (ITN) coverage, and city accessibility. For example, we found that ITN coverage and city accessibility were correlated with *pfmdr1* N86, 184F, and D1246, but these factors likely acted as proxies for drug quality, access, and mobility patterns (32). More studies are needed to determine the explicit association between these covariates, drug consumption patterns, and molecular marker prevalence (21, 22, 32). Our modeling approach cannot speak to many of the biological aspects of malaria transmission or mechanisms by which these covariates impact drug resistance. Future field and experimental studies,

**Table 1. Fitted and predicted estimates of molecular marker prevalence**

Marker	Time period	East (mean, SD)	West (mean, SD)	Central (mean, SD)	Southern (mean, SD)	Aggregate prevalence	No. of ADs with decreases in mutation prevalence (%)
<i>pfprt</i> 76T	2004 to 2009	0.76 (0.19)	0.67 (0.17)	0.74 (0.15)	0.50 (0.27)	0.75 (0.21)	544 (89.5%)
	2010 to 2018	0.52 (0.19)	0.44 (0.14)	0.45 (0.18)	0.33 (0.13)	0.45 (0.18)	
<i>pfmdr1</i> 86Y	2004 to 2009	0.70 (0.09)	0.54 (0.12)	0.69 (0.12)	0.58 (0.12)	0.62 (0.15)	585 (96.2%)
	2010 to 2018	0.34 (0.11)	0.25 (0.14)	0.24 (0.17)	0.18 (0.14)	0.23 (0.16)	
<i>pfmdr1</i> 184F	2004 to 2009	0.43 (0.13)	0.62 (0.11)	0.55 (0.12)	0.49 (0.11)	0.52 (0.21)	313 (51.4%)
	2010 to 2018	0.54 (0.19)	0.55 (0.16)	0.49 (0.15)	0.49 (0.16)	0.51 (0.17)	
<i>pfmdr1</i> 1246Y	2004 to 2009	0.36 (0.26)	0.14 (0.16)	0.21 (0.20)	0.13 (0.13)	0.23 (0.15)	496 (81.6%)
	2010 to 2018	0.09 (0.11)	0.07 (0.07)	0.05 (0.04)	0.05 (0.03)	0.08 (0.08)	

The table shows the average fitted/estimated prevalence and SD for all administrative divisions in the respective time period and region of sub-Saharan Africa, the overall average for all administrative divisions (ADs), and the overall number of ADs with decreased prevalence of mutant alleles in the latter time period.

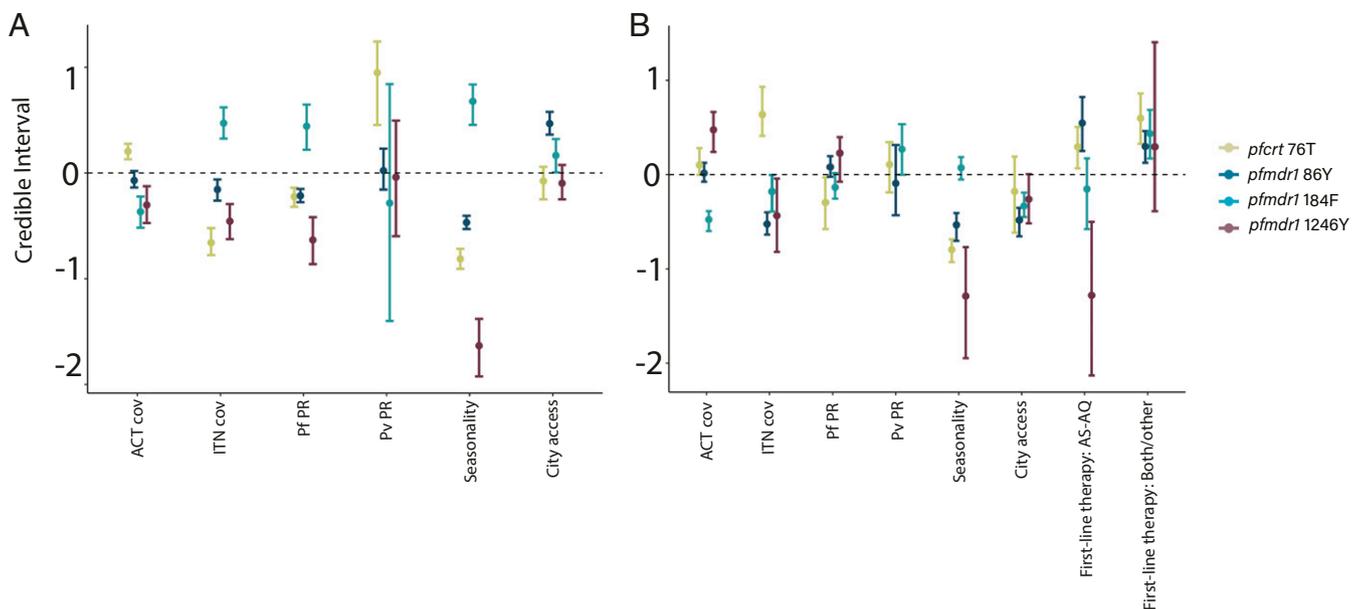
perhaps at more focal scales at which biological-level data can be obtained, are also needed.

Our analyses are also limited by gaps in spatial, temporal, and genetic coverage of data. To increase the predictive capacity of our models, we aggregated data into two time units, and our results may not precisely match localized reports of resistance or trends where more continuous data are available. To develop the largest possible dataset, we examined four well-known, commonly assessed molecular markers, but we were not able to analyze haplotypes or incorporate more detailed sequencing data. Although copy number variation in *pfmdr1* has been found in parts of Africa, limited reporting and large heterogeneity in polygenomic infections made these data challenging to model and interpret (15, 33). Our framework allows for incorporation of existing markers as well as novel mutations as they become available and more widely assessed, such as those conferring phenotypic resistance to artemisinin in Africa, as was the case in southeast Asia (34). Finally, systematic acquisition and storage of samples can also serve as a repository for retrospective studies of newly identified markers of interest.

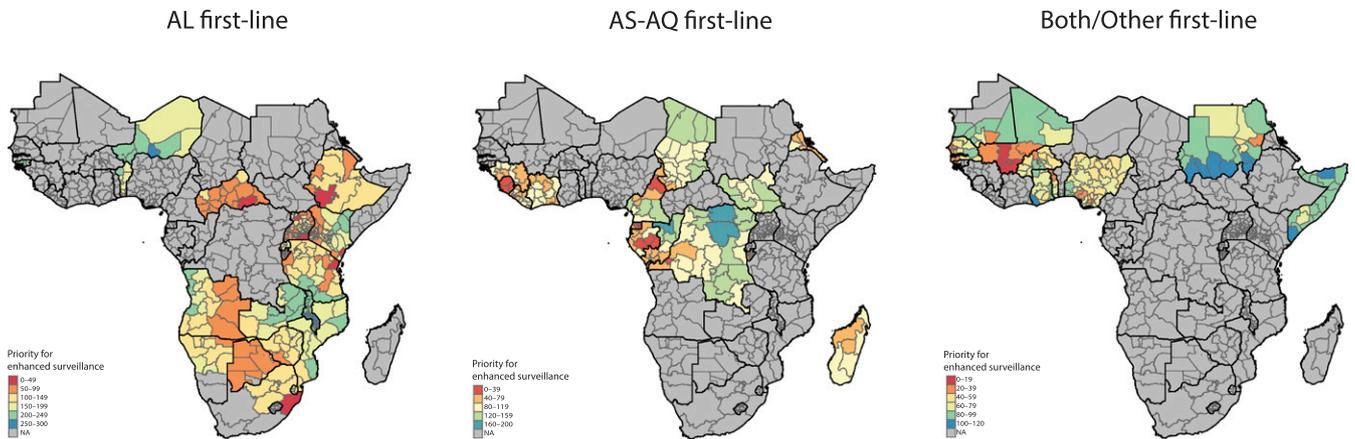
A number of previous studies have mapped molecular markers associated with antimalarial resistance at regional scales. Geospatial

maps were first developed for markers implicated in resistance to SP, a former first-line *P. falciparum* treatment, in sub-Saharan Africa (13, 35–37). Additional studies by Tun et al. (38) and Grist et al. (39) mapped polymorphisms in *pfK13* conferring artemisinin resistance in southeast Asia. Finally, Okell et al. (19) analyzed trends in *pfmdr1* markers using point prevalence surveys across sub-Saharan Africa. We add to this body of work by including another critical locus, *pfprt* K76T, and by developing spatially smoothed maps of partner drug resistance prevalence and uncertainty. These maps enable researchers to estimate marker prevalence in areas without sampling and to leverage uncertainty for strategic surveillance (14). This study also helps determine continent-wide geographic and temporal trends in key molecular markers and their association with environmental and policy landscapes.

In the final section of this study, we provide maps (Fig. 3) delineating regions undergoing potentially strong selection toward genotypes associated with reduced drug susceptibility, providing a genetic background within which artemisinin resistance may be more likely to emerge or spread (16). We recommend enhanced surveillance in these regions, which could entail more frequent surveys to establish the real-time frequency of molecular markers



**Fig. 2.** Contributions of covariates in the periods of 2004–2009 (A) and 2010–2018 (B). Posterior regression parameter estimates (mean, points, and 95% credible interval, whiskers) for covariates used in spatial models estimating the prevalence of individual molecular markers and time periods. Continuous variables are scaled for ease of comparison. The reference category for first-line therapy was AL. First-line drug policies were included only as covariates in the latter time period because they were assumed to have little effect during the period of ACT uptake (details in *SI Appendix, Table S1*).



**Fig. 3.** Prioritized sites for enhanced molecular surveillance. First-level administrative divisions were comparatively ranked according to the relative magnitude of change over time, dependent on first-line drug regimen to account for directional selection, after accounting for uncertainty in model results. Red regions are ranked highest priority for surveillance; blue regions are lower priority.

as well as patterns of drug consumption and quality. Additional focused studies in these areas may include treatment efficacy and pharmacokinetic studies as well as deep sequencing of local parasite strains to delineate haplotypes and novel variants (11, 12, 40, 41). The utility of our surveillance maps will improve with additional data coverage from surveillance conducted throughout a country/region, ideally under the guidance of national malaria control programs, and future versions should be optimized to support local research and decision making (42, 43). We intend for this work to help catalyze global agencies and regional networks alike to allocate resources urgently and strategically for drug resistance surveillance.

## Materials and Methods

**Prevalence Data Assembly and Aggregation.** We aggregated the results of 254 published studies in 35 countries in sub-Saharan Africa and Sudan, as previously described, encompassing 501 surveys assessing molecular markers in a specific location and year (14). In brief, all surveys assessed *P. falciparum* samples collected between 2004 and 2018 in sub-Saharan Africa for genotypes *pfmdr1* N86Y, Y184F, and D1246Y and/or *pfcr1* K76T. We excluded studies that reported prevalence data aggregated across multiples sites more than 300 km apart and whose results could not be disaggregated ( $n = 23$  surveys) (35). From each survey, we extracted the midpoint year of sampling, geographic coordinates of the study site, the number of samples tested at each locus, and the number of samples with wild-type, mutant, or mixed genotypes (in the case of polygenomic infections) at each locus. Prevalence was calculated as the number of individuals with mutant or mixed infections out of the total surveyed (35).

The study region was divided into the 608 first-level administrative divisions having endemic malaria transmission during the study period (44); the island nation of Comoros was excluded because administrative divisions did not have spatially contiguous neighbors, a requirement for our spatial models. The study period was divided into two time units representing the uptake of ACTs, 2004 to 2009, and widespread usage of ACTs, 2010 to 2018 (19). The observed prevalence of each marker in each administrative division was calculated as the weighted average of all respective survey estimates within each administrative division and time period (SI Appendix, Fig. S1).

While we define prevalence as the number of individuals with mutant or mixed infections out of the total number of *P. falciparum*-infected individuals surveyed, frequency is the proportion of parasite clones with the molecular marker in the parasite population, which may matter in highly endemic regions in which individuals often have polygenomic infections (35, 45). For example, say an individual is infected with four clones of *P. falciparum*, one of which has the mutant genotype of interest: If only that single person were surveyed, the frequency of the mutation would be 25%, whereas the prevalence would be 100%. Because the majority of studies ( $n = 209/254$  studies) did not report the average number of clones within the surveyed population, we used the subset of data that reported the prevalence of mixed and mutant

infections separately ( $n = 245/501$  surveys) and inferred the proportion of multiclonal infections using the *P. falciparum* parasite rate obtained from the Malaria Atlas Project (44). We then applied methodology developed by Okell et al. (35) to estimate the observed frequency of each marker in each survey and aggregated surveys in the same method as described above.

**Covariate Data Assembly and Selection.** Variables that we hypothesized to be associated with drug resistance were considered for inclusion (SI Appendix, Table S1 and Supplementary Methods) (12, 46). The final selected set of variables were *P. falciparum* parasite rate in children ages 2 to 10 y (*P. falciparum* parasite rate [PfPR]<sub>2-10</sub>), *P. vivax* parasite rate in individuals ages 0 to 99 y (*P. vivax* parasite rate [PvPR]<sub>0-99</sub>), seasonality of annual rainfall, ACT coverage, ITN coverage, city accessibility, and national first-line drug policy (AL, AS-AQ, or both/other; SI Appendix, Fig. S2) for the treatment of *P. falciparum*. For all covariates, excluding those only available at the national level, we averaged pixel values from rasterized data to generate administrative division-level estimates. Covariates were extracted for the midpoint year of both time periods, when possible, and for the equivalent study region (SI Appendix, Table S1 and Fig. S3).

**Bayesian Hierarchical Spatial Model.** We used a logistic regression framework with spatially correlated random effects to identify factors that explain variability in the prevalence of molecular markers associated with partner drug resistance by first-level administrative divisions and to predict prevalence in administrative divisions without observed data. We previously found evidence of spatial autocorrelation in the data, suggesting utility of spatial analysis (14). Here, we specify a conditional autoregressive (CAR) model for the random effects which allows for the possibility that neighboring administrative divisions exhibit similar behavior, leading to localized smoothness of estimated prevalence, potentially improved predictions in unobserved regions, and accurate statistical inference for the regression associations of interest. The model for the aggregated survey data during a selected time period  $t$  ( $t = 1$ : 2004 to 2009,  $t = 2$ : 2010 to 2018) is given as

$$Y_{kt} \sim \text{Binomial}(N_{kt}, p_{kt}) \text{ for } k = 1, \dots, K$$

and

$$\ln\left(\frac{p_{kt}}{1-p_{kt}}\right) = \mathbf{x}_{kt}^T \boldsymbol{\beta} + \psi_{kt}$$

where  $Y_{kt}$  is the number of mutant or mixed genotypes observed in administrative division  $k$  during time period  $t$ ,  $N_{kt}$  is the total number of individuals tested, and  $p_{kt}$  is the true prevalence of resistance. Logistic regression is used to link  $p_{kt}$  with the covariates and random effect, where  $\mathbf{x}_{kt}$  represents a vector of previously described covariates specific to administrative division  $k$  during time period  $t$  and  $\psi_{kt}$  is the spatially correlated random effect.

We specify the Leroux version of the CAR model for the spatial random effects (47), with an intuitive conditional form such that

$$\psi_{kt} \mid \Psi_{-kt}, \tau^2, \rho \sim N \left( \frac{\rho \sum_{j=1}^K w_{kj} \psi_{jt}}{\rho \sum_{j=1}^K w_{kj} + 1 - \rho}, \frac{\tau^2}{\rho \sum_{j=1}^K w_{kj} + 1 - \rho} \right),$$

where  $\Psi_{-kt}$  is the full vector of spatial random effects from time period  $t$  with  $\psi_{kt}$  removed and  $w_{kj}$  is equal to one if administrative division  $k$  and administrative division  $j$  are neighbors (i.e., share a common border) and is equal to zero otherwise;  $w_{kk}$  is equal to zero for all  $k$  (4). The CAR model includes two hyperparameters that control the total variability of the effects ( $\tau^2$ ) and whether this variability is spatially smooth ( $\rho \approx 1$ ) or independent ( $\rho \approx 0$ ), both of which are assigned prior distributions so that the data inform these quantities. This model specifies that the prior mean of a particular random effect is a weighted average of its neighbors' random effect values. As  $\rho$  approaches one, more emphasis is placed on the neighbors' values, whereas the random effect becomes centered closer to zero as  $\rho$  approaches zero.

Models were fit separately for each marker and time period. We selected weakly informative priors for model parameters to allow the data to drive the inference, such that  $\beta_j \sim N(0, 100^2)$ ,  $\rho \sim \text{Uniform}(0, 1)$  and  $\tau^2 \sim \text{Inverse gamma}(0.01, 0.01)$ . We collected 1,000,000 samples (after discarding 200,000 during a burn-in period) from the joint posterior distribution of the parameters using a Markov chain Monte Carlo sampling algorithm. We thinned these samples by a factor of 100 to reduce posterior autocorrelation, resulting in 10,000 nearly independent samples with which to make posterior inference. Convergence was assessed through visual inspection of trace plots and calculation of the Geweke diagnostic for all monitored parameters, with no obvious signs of nonconvergence across all fitted models. All analyses were conducted in R version 3.5.1. Spatial models were implemented with the CARBayes package (48).

**Estimation of Marker Prevalence and Changes over Time.** In administrative divisions with observed data, we extracted the posterior mean of  $p_{kt}$ . In administrative divisions without observed data, we used the CAR structure

of the spatial model to predict the prevalence and uncertainty by collecting samples from the corresponding posterior predictive distributions of the random effects. Hereafter, prevalence represents the posterior mean or SD of the predicted or fitted estimates for each administrative division. For each marker, we compared the change in prevalence over time by calculating the posterior probability that the prevalence from 2004 to 2009 was larger than the prevalence from 2010 to 2018. We also fitted a linear regression model with CAR random effects to determine which covariates explained variability in the estimated differences in prevalence over time while also accounting for spatial correlation in the data (SI Appendix, Supplementary Methods).

**Identification of Potential Regions for Enhanced Surveillance.** We considered regions of high priority for enhanced surveillance to be those with large relative changes toward the genotypes that confer reduced susceptibility to the first-line regimen used in that region, or a large relative change in any direction in the case of multiple first-line therapies. For each administrative division, we calculated the relative change in marker prevalence over the two time periods for each posterior sample collected, yielding 10,000 difference estimates. We then subset administrative divisions by the first-line drug regimen used in that region (AL, AS-AQ, or both/other) to account for opposing selective pressures of drugs on markers in *pfcr*t and *pfmdr*1. Maps were created with the tmap package (49).

**Data Availability.** Data, code, and maps generated by this study are publicly available in GitHub at [https://github.com/hannaehrlich/maldrugres\\_SSA](https://github.com/hannaehrlich/maldrugres_SSA) (50). Molecular marker survey data can also be viewed and deposited in the Worldwide Antimalarial Resistance Network at <https://www.wwarn.org/tracking-resistance/act-partner-drug-molecular-survevor>. Please direct any requests for tailored mapping modifications to the corresponding author.

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- World Health Organization, *World Malaria Report* (World Health Organization, Geneva, Switzerland, 2020).
- A. Uwimana *et al.*, Emergence and clonal expansion of in vitro artemisinin-resistant *Plasmodium falciparum* kelch13 R561H mutant parasites in Rwanda. *Nat. Med.* **26**, 1602–1608 (2020).
- A. Uwimana *et al.*, Association of *Plasmodium falciparum* kelch13 R561H genotypes with delayed parasite clearance in Rwanda: An open-label, single-arm, multicentre, therapeutic efficacy study. *Lancet Infect. Dis.* S1473-3099(21)00142-0 (2021).
- H. C. Slater, J. T. Griffin, A. C. Ghani, L. C. Okell, Assessing the potential impact of artemisinin and partner drug resistance in sub-Saharan Africa. *Malar. J.* **15**, 10 (2016).
- M. Venkatesan *et al.*, Polymorphisms in *Plasmodium falciparum* chloroquine resistance transporter and multidrug resistance 1 genes: Parasite risk factors that affect treatment outcomes for *P. falciparum* malaria after artemether-lumefantrine and artesunate-amodiaquine. *Am. J. Trop. Med. Hyg.* **91**, 833–843 (2014).
- S. Picot *et al.*, A systematic review and meta-analysis of evidence for correlation between molecular markers of parasite resistance and treatment outcome in falciparum malaria. *Malar. J.* **8**, 89 (2009).
- A. Arya, L. P. Kojom Foko, S. Chaudhry, A. Sharma, V. Singh, Artemisinin-based combination therapy (ACT) and drug resistance molecular markers: A systematic review of clinical studies from two malaria endemic regions—India and sub-Saharan Africa. *Int. J. Parasitol. Drugs Drug Resist.* **15**, 43–56 (2021).
- G. Henriques *et al.*, Directional selection at the *pfmdr*1, *pfcr*t, *pfubp*1, and *pfap2mu* loci of *Plasmodium falciparum* in Kenyan children treated with ACT. *J. Infect. Dis.* **210**, 2001–2008 (2014).
- M. T. Duraisingh, C. Roper, D. Walliker, D. C. Warhurst, Increased sensitivity to the antimalarials mefloquine and artemisinin is conferred by mutations in the *pfmdr*1 gene of *Plasmodium falciparum*. *Mol. Microbiol.* **36**, 955–961 (2000).
- M. Imwong *et al.*, The spread of artemisinin-resistant *Plasmodium falciparum* in the Greater Mekong subregion: A molecular epidemiology observational study. *Lancet Infect. Dis.* **17**, 491–497 (2017).
- C. Nsanjabana, F. Arie, H.-P. Beck, *et al.*, Molecular assays for antimalarial drug resistance surveillance: A target product profile. *PLoS One* **13**, e0204347 (2018).
- A. O. Talisuna *et al.*, Mitigating the threat of artemisinin resistance in Africa: Improvement of drug-resistance surveillance and response systems. *Lancet Infect. Dis.* **12**, 888–896 (2012).
- I. Naidoo, C. Roper, Mapping 'partially resistant', 'fully resistant', and 'super resistant' malaria. *Trends Parasitol.* **29**, 505–515 (2013).
- H. Y. Ehrlich, J. Jones, S. Parikh, Molecular surveillance of antimalarial partner drug resistance in sub-Saharan Africa: A spatial-temporal evidence mapping study. *Lancet Microbe* **1**, e209–e217 (2020).
- S. D. Otienoburu *et al.*, An online mapping database of molecular markers of drug resistance in *Plasmodium falciparum*: The ACT Partner Drug Molecular Surveyor. *Malar. J.* **18**, 12 (2019).
- O. J. Watson *et al.*, Pre-existing partner-drug resistance facilitates the emergence and spread of artemisinin resistance: A consensus modelling study. *bioRxiv* [Preprint] (2021). <https://www.biorxiv.org/content/10.1101/2021.04.08.437876v1> (Accessed 25 April 2021).
- M. K. Lafer *et al.*, Return of chloroquine antimalarial efficacy in Malawi. *N. Engl. J. Med.* **355**, 1959–1966 (2006).
- F. Lu *et al.*, Return of chloroquine sensitivity to Africa? Surveillance of African *Plasmodium falciparum* chloroquine resistance through malaria imported to China. *Parasit. Vectors* **10**, 355 (2017).
- L. C. Okell, L. M. Reiter, L. S. Ebbe, *et al.*, Emerging implications of policies on malaria treatment: Genetic changes in the *Pfmdr*-1 gene affecting susceptibility to artemether-lumefantrine and artesunate-amodiaquine in Africa. *BMJ Glob. Health* **3**, e000999 (2018).
- E. Hailemeskel *et al.*, Prevalence of *Plasmodium falciparum* *Pfcr*t and *Pfmdr*1 alleles in settings with different levels of *Plasmodium vivax* co-endemicity in Ethiopia. *Int. J. Parasitol. Drugs Drug Resist.* **11**, 8–12 (2019).
- A. E. Frosch, M. Venkatesan, M. K. Lafer, Patterns of chloroquine use and resistance in sub-Saharan Africa: A systematic review of household survey and molecular data. *Malar. J.* **10**, 116 (2011).
- M. Ocan *et al.*, Persistence of chloroquine resistance alleles in malaria endemic countries: A systematic review of burden and risk factors. *Malar. J.* **18**, 76 (2019).
- J. A. Flegg *et al.*, Trends in antimalarial drug use in Africa. *Am. J. Trop. Med. Hyg.* **89**, 857–865 (2013).
- A. Ecker, A. M. Lehane, J. Clain, D. A. Fidock, *PfCR*T and its role in antimalarial drug resistance. *Trends Parasitol.* **28**, 504–514 (2012).
- R. Hayward, K. J. Saliba, K. Kirk, *pfmdr*1 mutations associated with chloroquine resistance incur a fitness cost in *Plasmodium falciparum*. *Mol. Microbiol.* **55**, 1285–1295 (2005).
- E. Ochong, P. K. Tumwebaze, O. Byaruhanga, B. Greenhouse, P. J. Rosenthal, Fitness consequences of *Plasmodium falciparum* *pfmdr*1 polymorphisms inferred from ex vivo culture of Ugandan parasites. *Antimicrob. Agents Chemother.* **57**, 4245 (2013).
- H. A. Babiker, A. A. Gadalla, L. C. Ranford-Cartwright, The role of asymptomatic *P. falciparum* parasitaemia in the evolution of antimalarial drug resistance in areas of seasonal transmission. *Drug Resist. Updat.* **16**, 1–9 (2013).
- R. Ord *et al.*, Seasonal carriage of *pfcr*t and *pfmdr*1 alleles in Gambian *Plasmodium falciparum* imply reduced fitness of chloroquine-resistant parasites. *J. Infect. Dis.* **196**, 1613–1619 (2007).
- E. Y. Klein, D. L. Smith, R. Laxminarayan, S. Levin, Superinfection and the evolution of resistance to antimalarial drugs. *Proc. Biol. Sci.* **279**, 3834–3842 (2012).

30. R. Ataide *et al.*, Host immunity to *Plasmodium falciparum* and the assessment of emerging artemisinin resistance in a multinational cohort. *Proc. Natl. Acad. Sci. U.S.A.* **114**, 3515–3520 (2017).
31. P. J. Rosenthal, The interplay between drug resistance and fitness in malaria parasites. *Mol. Microbiol.* **89**, 1025–1038 (2013).
32. G. Rathmes *et al.*, Global estimation of anti-malarial drug effectiveness for the treatment of uncomplicated *Plasmodium falciparum* malaria 1991–2019. *Malar. J.* **19**, 374 (2020).
33. N. B. Gadalla *et al.*, Increased *pfmdr1* copy number and sequence polymorphisms in *Plasmodium falciparum* isolates from Sudanese malaria patients treated with artemether-lumefantrine. *Antimicrob. Agents Chemother.* **55**, 5408–5411 (2011).
34. F. Ariey *et al.*, A molecular marker of artemisinin-resistant *Plasmodium falciparum* malaria. *Nature* **505**, 50–55 (2014).
35. L. C. Okell, J. T. Griffin, C. Roper, Mapping sulphadoxine-pyrimethamine-resistant *Plasmodium falciparum* malaria in infected humans and in parasite populations in Africa. *Sci. Rep.* **7**, 7389 (2017).
36. J. A. Flegg *et al.*, Spatiotemporal mathematical modelling of mutations of the *dhps* gene in African *Plasmodium falciparum*. *Malar. J.* **12**, 249 (2013).
37. S. Sridaran *et al.*, Anti-folate drug resistance in Africa: Meta-analysis of reported *dhfr* and *dhps* mutant genotype frequencies in African *Plasmodium falciparum* parasite populations. *Malar. J.* **9**, 247 (2010).
38. K. M. Tun *et al.*, Spread of artemisinin-resistant *Plasmodium falciparum* in Myanmar: A cross-sectional survey of the K13 molecular marker. *Lancet Infect. Dis.* **15**, 415–421 (2015).
39. E. P. Grist *et al.*, Optimal health and disease management using spatial uncertainty: A geographic characterization of emergent artemisinin-resistant *Plasmodium falciparum* distributions in Southeast Asia. *Int. J. Health Geogr.* **15**, 37 (2016).
40. T. O. Apinjoh, A. Ouattara, V. P. K. Titanji, A. Djimde, A. Amambua-Ngwa, Genetic diversity and drug resistance surveillance of *Plasmodium falciparum* for malaria elimination: Is there an ideal tool for resource-limited sub-Saharan Africa? *Malar. J.* **18**, 217 (2019).
41. C. Nsanjabana, Resistance to artemisinin combination therapies (ACTs): Do not forget the partner Drug! *Trop. Med. Infect. Dis.* **4**, 26 (2019).
42. J. L. Warren, C. Perez-Heydrich, M. Yunus, Bayesian spatial design of optimal deep tube well locations in Matlab, Bangladesh. *Environmetrics* **24**, 377–386 (2013).
43. P. J. Diggle *et al.*, Spatial modelling and the prediction of Loa loa risk: Decision making under uncertainty. *Ann. Trop. Med. Parasitol.* **101**, 499–509 (2007).
44. D. J. Weiss *et al.*, Mapping the global prevalence, incidence, and mortality of *Plasmodium falciparum*, 2000–17: A spatial and temporal modelling study. *Lancet* **394**, 322–331 (2019).
45. T. Robinson, S. G. Campino, S. Auburn, *et al.*, Drug-resistant genotypes and multi-clonality in *Plasmodium falciparum* analysed by direct genome sequencing from peripheral blood of malaria patients. *PLoS One* **6**, e23204 (2011).
46. D. A. Pfeffer *et al.*, malariaAtlas: An R interface to global malariometric data hosted by the Malaria Atlas Project. *Malar. J.* **17**, 352 (2018).
47. B. G. Leroux, X. Lei, N. Breslow, "Estimation of disease rates in small areas: A new mixed model for spatial dependence" in *Statistical Models in Epidemiology, the Environment, and Clinical Trials*, M. E. Halloran, D. Berry, Eds. (Springer, New York, NY, 2000), pp. 179–191.
48. D. Lee, CARBayes version 4.6: An R package for spatial areal unit modelling with conditional autoregressive priors. *J. Stat. Softw.* **55**, 1–24 (2013).
49. M. Tennekes, tmap: Thematic Maps in R. *J. Stat. Softw.* **84**, 39 (2018).
50. H. Y. Ehrlich, Sunil Parikh, maldrugres\_SSA. GitHub. [https://github.com/hannaehrich/maldrugres\\_SSA](https://github.com/hannaehrich/maldrugres_SSA). Deposited 27 June 2021.