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Computers & Industrial Engineering





Inventory pooling decisions under demand scenarios in times of COVID-19



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ARTICLE INFO ABSTRACT Governments have been challenged to provide timely medical care to face the COVID-19 pandemic. The aim of Keywords. Resource allocation this research is to propose a novel inventory pooling model to help determine order sizes and safety inventories Inventories in local hospital warehouses. The current study attempts to portray the availability of pharmaceutical items Hospital in public hospitals facing COVID-19 challenges. Different from previous studies, this research builds upon Models the consecrated theory of inventory pooling, extending it to pandemic circumstances where the intractability Statistical of kurtosis and skewness in inventory models are critical issues for making sure that medicines have high availability at a low cost. These effects on the total cost of inventory are explored and compared to a supply system with no consolidation. A continuous-review model is assumed with allocation rules for centralization and regular transshipment given different skewness and kurtosis structures for the demand, describing them by the copula method. This method models a multivariate demand considering that the marginal distributions of the demand can be specified by the Generalized Additive Model for Location, Scale and Shape, which offers advantages to model demands considering virtually any marginal statistical distribution. Numerical simulations and an illustrative example show that distributions of demands with more negative skewness and high kurtosis favor to a greater extent obtaining lower total costs with regular supply transshipment systems. Our study points out important considerations for supply chain decision makers when having demands with skewness and kurtosis patterns.

1. Introduction

The health care industry provides a vital service for modern societies (Kochan, Nowicki, Sauser, & Randall, 2018). Analogously to what happens in other sectors, the health care industry competes based on time and quality. With the outbreak of COVID-19, several nations deployed a war effort in delivering goods and services to help control the pandemic, including the hospital care industry. In Chile, for instance, efforts have been made to increase the installed capacity for diagnoses, beds for critical patient care, as well as hospital equipment and supplies, including medication. Therefore, responsive and timely health care operations play a significant contribution to socio-economic integrity, promoting patient survival in rural and urban areas. What Newhouse (1970) said: "hospital services seem to be desirable in some ethical sense, which justifies the claim that consumers have a right to medical care", applies nowadays more than ever. We thus posit that the importance of hospitals in times of disaster provides leverage to consumer rights with respect to medical care, which in turn presents an indirect significant impact on the entire economy in the support and

https://doi.org/10.1016/j.cie.2021.107591

Available online 9 August 2021 0360-8352/© 2021 Elsevier Ltd. All rights reserved. provision of rural and urban populations. This being the case, improved availability of pharmaceutical items at low cost is deemed necessary to prevent such disasters (Hamzah & See, 2019), being the ultimate means of survival for rural and urban communities.

In this regard, the inventory consolidation effect (or portfolio effect) has been a highly analyzed topic in logistics literature (Askin, Baffo, & Xia, 2014; Ballou, 1981; Ballou & Burnetas, 2003; Dolati Neghabadi, Evrard Samuel, & Espinouse, 2019). Companies often satisfy the demand per unit of time (DPUT) with an independent system (IS) of supply, in which each point of sale (or demand area) is exclusively served by dedicated facilities; see Ballou and Burnetas (2003), Dolgui, Tiwari, Sinjana, Kumar, and Son (2018) and Tyagi and Das (1998). Note that there is no consolidation effect in an IS of supply. When the variability of the DPUT and the lead time (LT) are high, companies can reduce them by using grouped orders; see Wanke (2009). In that case, the DPUT may be satisfied with an inventory pooling that uses a centralized supply or a regular supply employing transshipment. Observe that transshipment assumes that a part of the demand is supplied from locations that supply different points of departure, even

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though there may not be a quantity of inventory on hand at the original supply location; see Evers (1996) and Rayat, Musavi, and Bozorgi-Amiri (2017). When both systems are compared, the inventory pooling often provides lower costs than when using the IS of supply.

The consolidated total cost (TC_c) of an inventory pooling and the total cost of an IS of supply (TC_i) are composed of various costs related to safety stocks and average quantities of inventories stored, ordered, and distributed during a cycle; see Knofius, van der Heijden, and Zijm (2019) and Wanke and Saliby (2009). The TC_c is subject to changes according to the allocation rules that determine if it is more convenient to centralize the warehouses to supply the demand or regularly occupy a transshipment policy; see Wen, Choi, and Chung (2019). The decision of which allocation rule is more convenient for inventory pooling is made after comparing the corresponding TC_c .

To the best of our knowledge, the consolidation effect has been mostly studied assuming independent and normally distributed DPUTs, but when satisfying each point of sale or delivery, DPUTs are random variables (RVs) that may show any shape (Sadeghi & Niaki, 2015). For the statistical distributions of RV independently and identically distributed (IID), such as the DPUTs, a statistical moment is a particular calculable dimension of the shape of its probability density functions (PDFs). The zero-th moment is always 1, the 1-th moment is the mean, the 2-th central moment is the variance, the 3-th standardized moment is the skewness, while the 4-th standardized moment is the kurtosis; see Casella and Berger (2002), Deng, Miao, Ma, Wei, and Feng (2020) and Lin, Sun, and Yu (2020). As we will see later in the background of this paper, all these moments are related, and for this paper we will postulate that they can influence inventory consolidation decisions.

Often an RV data set has a joint statistical distribution based on marginal distributions with different skewness and kurtosis, as noted in Chan, Lim, and McAleer (2005) and Escribano and Pfann (1998). This structure of dependence of the DPUT can be described by the copula method through the parametric specification and association of the marginal statistical distributions that will make up a multivariate joint distribution; see Autchariyapanitkul, Chanaim, and Sriboonchitta (2014). Although theoretical foundations of copulas are complex, its practical treatment is made simple using open source software such as R (Kopczewski, Sobolewski, & Miernik, 2018). This approach is made simpler than others, such as simulation (Qian, Li, & Hu, 2017). In this context, it is essential to have excellent goodness of fit to actual data for a theoretical description of the marginal statistical distributions (Alavifard, 2019; Zhi, Wang, & Xu, 2020).

Generalized Additive Model for Location, Scale and Shape (GAMLSS) is a semiparametric regression type model introduced by Stasinopoulos and Rigby (2007) that allows great versatility in modeling random variables, which can be used in describing multivariate copulas of joint statistical distributions (Rohmer & Gehl, 2020).

In the current COVID-19 contingency scenario, drug demand patterns have undergone major changes. This is materialized in the fact that some products suffer large increases in the quantities demanded (demand shock), while others suffer an abrupt drop in these quantities requested for the treatment of patients. These changes in demand patterns can be shown as an increase in kurtosis and skewness to the right, in the first case, and as a decrease in kurtosis and left skewness in the second case. In both cases, these demand patterns are critical in determining the optimal lots to order and total inventory costs. In this study we extend the possibility of treating the kurtosis and asymmetry of the demand for critical items in the availability of pharmaceutical products in public hospitals that face COVID-19 contingency issues. In this context, the main contribution of this paper is to help understand how the kurtosis and skewness of the demand for items affect the supply systems of pooling and independent inventory policies and their total costs. So, the general objective of this paper is to propose a new methodology for studying how the skewness and kurtosis of joint distributions of DPUTs affect the TCc of an inventory pooling by using different allocation rules. The specific objectives are twofold:

(i) compare the total costs between systems with inventory pooling and independent supply under a joint statistical distribution composed of marginal distributions with different moment measures; and (ii) analyze the changes in the indicators related to the inventory total costs under different DPUT scenarios with different moment measures.

The remainder of the paper is organized as follows. In Section 2 we review literature regarding Healthcare inventory management and Inventory pooling with skewness and kurtosis scenarios in uncertain product demand. In the methodology of Section 3 we show how the multivariate statistical distribution of DPUTs based on copulas with different skewness and kurtosis of the marginal distributions affect the TC_c of inventories with different allocation rules. Also in this section we design two simulation studies and an illustrative case with their respective statistical analyses. Results are shown in Section 4. Section 5 we provide managerial guidelines regarding the topic under study. Section 6 shows the discussion, limitations, and possible future research, and finally we presented our conclusions in Section 7.

2. Literature review

2.1. Healthcare inventory management

The literature has been investigating the issue of inventory management in the healthcare industry from different perspectives during recent years. Table 1 summarizes the related studies over the past decade.

2.2. Inventory pooling with skewness and kurtosis scenarios in uncertain product demand

TC_c usually corresponds to the sum of four components: (C1) consolidated safety stock (SS_c) multiplied by the holding cost (HC), (C2) consolidated cycle stock (CS_c) multiplied by HC, (C3) consolidated distribution cost (DC_c), and (C4) consolidated order cost (OC_c); see Wanke (2009) for details about these components. The uncertainty in the DPUT and lead time (LT) of the items is important in determining the sum of safety stocks that are consolidated in inventory pooling; see Tallon (1993). Evers and Beier (1993) extended the SS_c to multiple stocking points (locations or facilities) that serve the demand, while Wanke (2014) did the same with the CS_c. In the context of reducing the TC_c, Wanke (2009, 2014) and Wanke and Saliby (2009) carried out a mathematical treatment of the consolidation effect considering three assumptions on the same safety-stock for the desired service level in all locations, which are as follows: 1. continuously review the model for lot-sizing; 2. define a reorder point as inventory control under uncertainty of the lead time demand (LTD), so that the consolidation does not affect the total average demand of the system; 3. variables DPUT, LT, and LTD must be independent and identically distributed in Gaussian form.

The allocation rules for the demand are an important aspect to be considered in the consolidation effect. A first allocation rule corresponds to a given centralized location, which supplies the same proportion of demand to each decentralized location, known as the Tyagi and Das' allocation rule. Under this rule, and considering the DPUT as an IID RV, the maximum consolidation effect depends on the LT conditions between locations. Note that the consolidation effect is maximized when the TC_c and inventory level are minimized. If both centralized locations present equal LT means and standard deviations (SDs), then the proportions of demand to be supplied in each decentralized location should be equal. Otherwise, that is under different LT means and/or SDs, a maximum consolidation effect is obtained when inventories are centralized into a single location; see Wanke (2009). In a second allocation rule, a location can supply a primary demand area and also other demand areas, which in turn can be supplied by other locations. This is known as Ballou and Burnetas' allocation rule and here one must ask whether a proportion of the demand should

Summary of previous studies in inventory management in the healthcare industry.

Authors	Aim	Method/data	Results/implications
De Vries (2011)	Investigate the shaping of inventory management system	A qualitative exploratory case study	 Inventory management system formation is significantly affected by the interactive relationships between different stakeholders involved in the project How to coordinate and balance the different interests held by various stakeholders is the challenge faced by top management
Bhakoo, Singh, and Sohal (2012)	Investigate the inventory management along this supply chain	Semi-structured interviews	Inventory management across the supply chain can be improved by collaborative arrangement between manufacturers and wholesalers/distributors
Zepeda, Nyaga, and Young (2016)	Examine the effects of horizontal inter-organizational arrangements on inventory costs for hospitals	A linear mixed effects model (LMM) with random intercepts	Affiliation with local, regional, and national systems has mitigating effects under weak logistics services infrastructure with the mitigating effect being greatest for affiliation in local systems.
Gebicki, Mooney, Chen, and Mazur (2014)	Evaluate the medication inventory management	Event-driven simulation	The trade off between patient safety and cost can be addressed by incorporating drug characteristics in the ordering decisions.
Wang, Cheng, Tseng, and Liu (2015)	Examine hospital inventory management	A dynamic drum-buffer-rope replenishment model	The optimal replenishment timing and quantity of total inventory cost with no stock-out occurrence can be effectively determined by the model.
Niakan and Rahimi (2015)	Examine the healthcare inventory routing problem	A multi-objective mathematical model and possibilistic fuzzy approach	The model proposed is superior in handling uncertain parameters.
Forcina, Petrillo, Bona, Felice, and Silvestri (2017)	Target stocking level which minimizes the total cost	Dynamic models the system while satisfying the service level constraint	The proposed model has characteristics of generality that allow the application in other areas
Timajchi, Al-e Hashem, and Rekik (2019)	Examine the healthcare inventory routing problem	Bi-objective mixed integer mathematical programming	Risk routes can be avoided and the economic supply network performance can be increased by the transshipment option.
Saedi, Kundakcioglu, and Henry (2016)	Investigate the optimal inventory policy for a healthcare facility	A stochastic model	Hospitals would benefit from inventory pooling since products in the healthcare industry, in particular in hospitals, have the characteristics that they have limited shelf-life.
Chen, Xiao, Wang, and Lei (2020)	Investigate the optimal order policies and the inventory adjustment quantity for a perishable products with a two-period half-life	A stochastic optimization model	 The expedited order plan performs better in terms of product wastage risk control, but it has a higher level of shortage risk compared to the returns plan. The risk of wastage and shortage in inventory management can be effectively controlled by combining the expedited order plan and returns plan

be supplied by a primary location and the remainder by secondary locations or not. In the assumption of IID DPUTs being considered, regular transshipment (RT) offers a good alternative for positively correlated DPUTs from all centralized locations with the possibility of balancing high/low values of the LT mean and DPUT SD at different centralized locations. It also seems to be the best system when the HC is very low; see Wanke and Saliby (2009).

As mentioned, previous works on the consolidation effect topic are based on the continuous review model under the assumption of IID normal DPUTs. These works provided evidence that the corresponding TC_c decreased in different scenarios that considered the following indicators: (i) the correlation level between the DPUTs of the service points; (ii) the mean and SD of DPUTs and LTs; (iii) the security factor for the LTD; and (iv) holding, order, and distribution costs. For more details on these indicators and scenarios; see Wanke (2009) and Wanke and Saliby (2009). We can point out the following main conclusions generated from previous works: (a) negative correlations between DPUTs and high values for SDs of DPUT are associated with a low TC_c obtained by centralized systems; (b) positive correlations and low values for SDs of DPUT are associated with an IS of supply; (c) different SD (heterogeneity) for the LT is better handled by an RT because this system combines and balances the DPUTs that are served from different decentralized locations. The above findings are confirmed theoretically for non-Gaussian distributions by Corbett and Rajaram (2006). However, our study differs from the previous one since we will explore the effect of left and right asymmetries and degree of kurtosis on DPUTs distributions to conclude how these influence the total costs between systems with inventory pooling and independent supply obtained by allocation rules. Skewness is a measure that reflects when a distribution or data set is symmetric or is shifted to the left or right of the central data. Kurtosis is a measure of data and outliers called tails, which can be heavy or light. Data with high kurtosis tend to have heavy tails or outliers, while a dataset with low kurtosis tends to have light tails or no outliers, where distributions with positive kurtosis are called leptokurtic. Those with kurtosis around zero are called mesokurtics and those with negative kurtosis are denominated platykurtics. Fig. 1 illustrates these differences. Note that the leptokurtic distributions concentrate higher probabilities of occurrence in the values close to the mean, while in the platykurtic the probabilities are more distributed in the tails of the distribution, while the mesokurtic distributions express the moderate cases of probabilistic distribution. On the other hand, the right skewness implies that there is a shift of



Fig. 1. Kurtosis and skewness of a distribution: (a) Leptokurtic, (b) Mesokurtic, (c) Platykurtic, (d) Right skewness, (e) Normal skewness, (f) Left skewness.

the probabilities towards the data higher than the average, otherwise with the left skew, while the normal symmetry corresponds well to a Gaussian or normal distribution.

Nelsen (1999) introduced the copula method to describe the dependence between RVs where the copula is a multivariate probability distribution with uniform marginal statistical distribution for each variable. Then, for the case of IID DPUTs, the copula method allows us to obtain the joint probabilistic distribution of correlated RVs without needing to know the joint distribution, which is easily obtainable from real DPUT data. As mentioned in the introduction, it is possible to use the GAMLSS model to model the marginal distributions of random variables that make up the joint multivariate distribution. Unlike generalized linear models (GLM), the GAMLSS considers a family of generalized, discrete, or continuous statistical distributions, which can have varying degrees of skewness and kurtosis. Thanks to its formulation, it is possible to model any parameter of these statistical distributions for the response variable linearly or not in an additive parametric or non-parametric form of covariates with known or random values; see Stasinopoulos, Rigby, and Akantziliotou (2008). The advantage of using this type of statistical modeling is that GAMLSS is a regression toolbox appropriate for a big dataset of response variables that can consider linear or smoothing functions of predictive covariates to model any parameter of location, scale, or shape of the statistical distribution. The current packages available in R software (Stasinopoulos, Rigby, Voudouris, Heller, & De Bastiani, 2015) allow working with continuous (any type of skewness or kurtosis), discrete (including zero inflated data), and mixture statistical distributions. Models can be selected according to criteria of goodness of fit to the real data, as well as by generating random numbers with arbitrary distributions of interest for theoretical or empirical research (Rojas & Ibacache-Quiroga, 2020; Rojas, Leiva, Wanke, Lillo, & Pascual, 2019).

Our proposal differs from what has been addressed in the literature on the consolidation effect of inventories. We describe IID DPUT as a marginal statistical distribution described by GAMLSS models, making it possible to generate copulas to build joint multivariate with different skewness and kurtosis from the marginal distributions and to explore the TC_c that leads to an inventory pooling by using different allocation rules.

3. Methodology

3.1. How to describe the DPUT for an inventory item

GAMLSS formulation. Let *Y* be an IID RV corresponding to the DPUT. If *Y* be the DPUT of an inventory item, we considered that μ is the expected value of a response variable. Consider to *d* as a covariate. If $f(y|\theta)$ be a conditional PDF on parameters θ ($F_{y|\theta}$ is the conditional cumulative distribution function (CDF)), where $\theta = (\mu, \sigma, \nu, \tau)^{\top} = (\theta_1, \theta_2, \theta_3, \theta_4)^{\top}$ is a vector of four distribution parameters. In the GAMLSS formulation, only μ is a function of the covariates, while μ and

 σ are location and scale parameters, and v and τ are shape parameters. If $\{y_i\}, i = 1, ..., n$ is an $n \times 1$ vector of the response variable to model, considering k = 1, 2, 3, 4 as parameters, then g_k is a link functions related to the *k*th parameter θ_k and to covariates by the following additive models:

$$g_1(\boldsymbol{\mu}) = \boldsymbol{\eta}_1 = \boldsymbol{D}_1 \boldsymbol{\beta}_1 + \sum_{j=1}^{J_1} h_{j1}(\boldsymbol{d}_{j1}), \tag{1}$$

$$g_2(\sigma) = \eta_2 = \sum_{j=1}^{J_2} h_{j2}(d_{j2}),$$
 (2)

$$g_3(\mathbf{v}) = \eta_3 = \sum_{j=1}^{J_3} h_{j3}(\mathbf{d}_{j3}), \text{ and}$$
 (3)

$$g_4(\tau) = \eta_4 = \sum_{j=1}^{J_4} h_{j4}(d_{j4}), \tag{4}$$

where $\mu, \sigma, \nu, \tau, \eta_t$ and d_{j1} , for $j = 1, ..., J_k$ and k = 1, 2, 3, 4, are $n \times 1$ vectors. D_1 is an $n \times J_1$ known matrix of variables and the regression coefficients β_1 to be estimated is a $J_1 \times 1$ vector. h_{jk} is a semi-parametric additive function for the covariate D_{jk} evaluated at the vector d_{jk} , which is assumed fixed and known.

For details of parameter estimate, diagnostic, and good fit on the data see Stasinopoulos and Rigby (2007).

Moments, skewness, and kurtosis. If F is a CDF of any statistical distribution, considering the Riemann–Stieltjes integral (Liu, 2004), the *n*th moment of the statistical distribution is expressed by:

$$\mu'_n = \mathbb{E}\left[Y^n\right] = \int_{-\infty}^{\infty} y^n \,\mathrm{d}F(y)$$

with E as an expectation operator for the mean.

The zero-th moment of any PDF is 1.

The first raw moment is the mean:

1

$$\mu \equiv E[Y].$$

The second central moment is the variance, and the square root of the variance is the SD:

$$SD \equiv \left(E\left[(y-\mu)^2 \right] \right)^{\frac{1}{2}}$$

The normalized *n*th central moment of the RV Y is

$$\frac{\mu_n}{SD^n} = \frac{\mathrm{E}\left[(Y-\mu)^n\right]}{SD^n},$$

and represents the distribution.

The normalized third central moment is called the skewness. A distribution that is skewed to the left has a negative skewness, and vice versa. Zero values indicate symmetry of the distribution. The Fisher coefficient of skewness (CSk) is defined as:

$$CSk = \frac{\mu_3}{SD^3}$$

where μ'_3 is the third moment centered.

The fourth central moment is a measure of outliers values far from the average distribution values and is denominated kurtosis. Statistical distributions with kurtosis less than 3 are said to be "platykurtic", while distributions with kurtosis greater than 3 are said to be "leptokurtic". The Fisher coefficient of kurtosis (CK) is defined as:

$$CK = \frac{\mu_4'}{SD^4},$$

where μ'_{4} is the fourth moment centered.

Joint CDF F_{Y_1,Y_2} modeling by copulas. Now, let Y_1 and Y_2 be IID RVs with their unknown joint CDF F_{Y_1,Y_2} . Then, we have the relationship

$$C(a,b;\rho) = F_{Y_1,Y_2}(F_{Y_1}^{-1}(a),F_{Y_2}^{-1}(b)), \quad a,b \in [0,1],$$
(5)

where ρ is a parameter of dependence between Y_1 and Y_2 , which is denoted as $\rho = \rho_{Y_1,Y_2}$. The copula defined in (5) is assumed to be continuous and twice differentiable. Thus, from (5) and Corbett and Rajaram (2006), we obtain that the probability density function (PDF) associated with $C(u, v, \rho)$, which can be expressed as

$$c(F_{Y_2}(y_2), F_{Y_1}(y_1), \rho_{Y_1, Y_2}) = \frac{\partial^2 C}{\partial F_{Y_2} \partial F_{Y_1}}$$

For Y RV, the PDF f_Y defined on $[0,\infty)$ (non-negative support), is a CDF

$$F_Y(y) = \int_0^y f_Y(v) \,\mathrm{d}v$$

and a quantile function (QF) $y(q) = F_Y^{-1}(q)$, for 0 < q < 1. Note that $F_{Y_1}^{-1}(a)$ and $F_{Y_2}^{-1}(b)$, which can be obtained when marginal statistical distributions are known or adjusted by the GAMLSS model from actual data.

Readers interested in using copulas to describe correlated DPUTs in consolidated effects are referred to Wanke (2014).

3.2. Consolidation effect

As mentioned in the introduction section, the TC_c usually considers the components C1, C2, C3, and C4 that contain, respectively, the following elements:

$$SS_{c} = \sum_{f=1}^{m} \sqrt{\left(E\left(\sum_{i=1}^{n} w_{i,f}Y_{i}\right)\right)^{2} Var(T_{f}) + Var\left(\sum_{i=1}^{n} w_{i,f}Y_{i}\right) E(T_{f})}, \quad (6)$$

$$CS_{c} = \frac{1}{\sqrt{2HC}} \sum_{f=1}^{m} \sqrt{OC_{f}E\left(\sum_{i=1}^{n} w_{i,f}Y_{i}\right)},$$
(7)

$$OC_{c} = \frac{1}{2CS_{c}} \sum_{f=1}^{m} OC_{f} E\left(\sum_{i=1}^{n} w_{i,f} Y_{i}\right), \qquad (8)$$

$$DC_{c} = \sum_{f=1}^{m} \sum_{i=1}^{n} DC_{T_{f,i}} w_{i,f} Y_{i},$$
(9)

where, for the centralized location f, OC_f is the order cost (expressed in USD\$/order); $HC_f = HC$ is the holding cost (expressed in USD\$/unit); $w_{i,f}$ is the proportion of the mean DPUT transferred from the decentralized location i to the centralized location f; $E(T_f)$ is the LT mean; $Var(T_f)$ is the LT variance; and $DC_{T_{f,i}}$ is the unitary transportation distribution cost to move a single item from a centralized location f to a decentralized location i. Note that $0 \le w_{i,f} \le 1$, for all i = 1, ..., n and f = 1, ..., m, whereas $\sum_{f=1}^{m} w_{i,f} = 1$, for all i, with m being the number of centralized locations and n the number of decentralized locations, for $1 \le m \le n$.

With *m* and *n* defined above, the case m = n = 2 can be extended to any value of *m*, *n*, with $m \le n$. This is based on the previous works of Ballou and Burnetas (2003), Tyagi and Das (1998), Wanke (2009, 2014) and Wanke and Saliby (2009), which allows us to have a benchmark for comparing our results. When n = m = 2, Tyagi and

Das' allocation rule provides $w_{1,f} = w_{2,f} = w_f$, for $0 \le w_f \le 1$, and $\sum_{f=1}^{m} w_f = 1$, with $w_2 = 1 - w_1$. However, under Ballou and Burnetas' allocation rule, it is known that $w_{1,1} = w_{2,2} = w_p$, and $w_{1,2} = w_{2,1} = w_{2,1}$ $1 - w_p$, where the subindex p denotes the primary location, that is, the location assigned with highest proportion of demand. Observe that the optimal solution under Ballou and Burnetas' allocation rule does not only behave differently from the solution obtained with Tyagi and Das' allocation rule, but it also implies in different inventory pooling models; see Wanke (2009). Note that, under Ballou and Burnetas' allocation rule, both demand points are supplied exclusively by a warehouse, under an optimal solution where w_p is zero or one (IS of supply), but under Tyagi and Das' allocation rule, if w_1 is zero or one, both decentralized locations share one single serving facility. Nevertheless, when the optimal values of w_p are greater than zero and less than one, RT takes place under Ballou and Burnetas' allocation rule. The above implies that both demand points are supplied by all the warehouses. 0 < 0 $w_1 < 1$ produces the same pattern as that Tyagi and Das' allocation rule, but they do not correspond to optimal solutions; see Wanke (2014). The use of these allocation rules results in a consolidated inventory cost, allowing the decisions of both supply policies to be compared and to opt for the one with the lowest total cost. The differences and how these allocation rules operate are illustrated inFig. 2.

Next we describe how to calculate CS_c , SS_c , OC_c , and DC_c under both allocation rules. Then, under Tyagi and Das' allocation rule in the decentralized location *i*, expressions defined in (6)–(9) acquire the functional forms

$$\begin{split} \mathrm{SS}_{\mathrm{c}} &= \left(\sqrt{w_{1}^{2} \left(\sum_{i=1}^{2} \mu_{i} \right)^{2} \mathrm{Var}(T_{1}) + w_{1}^{2} \mathrm{Var}\left(\sum_{i=1}^{2} Y_{i} \right) \mathrm{E}(T_{1})} \\ &+ \sqrt{(1 - w_{1})^{2} \left(\sum_{i=1}^{2} \mu_{i} \right)^{2} \mathrm{Var}(T_{2}) + (1 - w_{1})^{2} \mathrm{Var}\left(\sum_{i=1}^{2} Y_{i} \right) \mathrm{E}(T_{2})} \right), \\ \mathrm{CS}_{\mathrm{c}} &= \sqrt{\frac{\sum_{i=1}^{2} \mu_{i}}{2\mathrm{HC}}} \left(\sqrt{\mathrm{OC}_{1} w_{1}} + \sqrt{\mathrm{OC}_{2}(1 - w_{1})} \right), \\ \mathrm{OC}_{\mathrm{c}} &= \frac{\mathrm{OC}_{1} w_{1}(\mu_{1} + \mu_{2})}{2\mathrm{CS}_{\mathrm{c}}} + \frac{\mathrm{OC}_{2}(1 - w_{1})(\mu_{1,} + \mu_{2})}{2\mathrm{CS}_{\mathrm{c}}}, \\ \mathrm{DC}_{\mathrm{c}} &= \mathrm{DC}_{\mathrm{T}_{1,1}} \mu_{1} + \mathrm{DC}_{\mathrm{T}_{1,2}} \mu_{2}, \end{split}$$

where μ_i , for i = 1, 2, are defined as in (1), and

1

$$\operatorname{Var}\left(\sum_{i=1}^{2} Y_{i}\right) = \operatorname{Var}(Y_{1}) + \operatorname{Var}(Y_{2}) + 2\rho_{Y_{2},Y_{1}}\sqrt{\operatorname{Var}(Y_{1})\operatorname{Var}(Y_{2})}.$$

In addition, to calculate DC_c , we assume (without loss of generality) that inventories are consolidated at location 1. Under Ballou and Burnetas' allocation rule, and considering Y_i in the decentralized location *i*, (6), (7), (8), and (9) acquire the following functional forms, respectively.

$$\begin{split} \mathrm{SS}_{\mathrm{c}} &= \left(((w_{p}\mu_{1} + (1 - w_{p})\mu_{2})^{2}\mathrm{Var}(T_{1}) + (w_{p}^{2}\mathrm{Var}(Y_{1}) + (1 - w_{p})^{2}\mathrm{Var}(Y_{2}) \\ &+ 2\rho_{Y_{2},Y_{1}}\sqrt{\mathrm{Var}(Y_{1})\mathrm{Var}(Y_{2})}w_{p}(1 - w_{p}))\mathrm{E}(T_{1})) \right)^{1/2} \\ &+ \left((((1 - w_{p})\mu_{1} + w_{p}\mu_{2})^{2}\mathrm{Var}(T_{1}) + ((1 - w_{p})^{2}\mathrm{Var}(Y_{1}) \\ &+ w_{p}^{2}\mathrm{Var}(Y_{2})2\rho_{Y_{2},Y_{1}}\sqrt{\mathrm{Var}(Y_{1})\mathrm{Var}(Y_{2})}w_{p}(1 - w_{p}))\mathrm{E}(T_{2})) \right)^{1/2}, \\ \mathrm{CS}_{\mathrm{c}} &= \frac{1}{\sqrt{2\mathrm{HC}}} \left(\sqrt{\mathrm{OC}_{1}(w_{p}\mu_{1} + (1 - w_{p})\mu_{2})} \\ &+ \sqrt{\mathrm{OC}_{2}\left((1 - w_{p})\mu_{1} + w_{p}\mu_{2}\right)} \right), \\ \mathrm{OC}_{\mathrm{c}} &= \frac{\mathrm{OC}_{1} \left(\mu_{1}w_{p} + \mu_{2}(1 - w_{p}) \right)}{2\mathrm{CS}_{\mathrm{c}}} + \frac{\mathrm{OC}_{2} \left((1 - w_{p})\mu_{1} + w_{p}\mu_{2}\right)}{2\mathrm{CS}_{\mathrm{c}}}, \\ \mathrm{DC}_{\mathrm{c}} &= \left(w_{p}\mathrm{DC}_{\mathrm{T}_{1,1}} + (1 - w_{p})\mathrm{DC}_{\mathrm{T}_{2,1}} \right) \mu_{1} \\ &+ \left(w_{p}\mathrm{DC}_{\mathrm{T}_{2,2}} + (1 - w_{p})\mathrm{DC}_{\mathrm{T}_{1,2}} \right) \mu_{2}. \end{split}$$



Fig. 2. (a) Tyagi and Das (1998) allocation rule, (b) Ballou and Burnetas (2003) allocation rule.

Note that, in (10), one must suppose that the highest proportions w_p of μ_i are supplied with less $DC_{T_{1,1}}$ and $DC_{T_{2,2}}$. In contrast, for the decentralized location *i* considered under an IS of supply, the safety stock, cycle stock, order cost, and distribution cost are respectively given by

$$SS_i = \sum_{f=1}^m \sqrt{E(T_f) \operatorname{Var}\left(\sum_{i=1}^n Y_i\right) + \sum_{i=1}^n \mu_i^2 \operatorname{Var}(T_f)},$$
(11)

$$CS_i = \sqrt{\frac{OC_i \mu_i}{2HC}}, \quad OC_i = \sum_{f=1}^m \frac{OC_f \mu_i}{2CS_i}, \quad DC_i = \mu_i DC_{T_{f,i}}.$$
(12)

Following Tallon (1993), it is possible to define a consolidation effect indicator, which corresponds to "the percentage reduction in aggregate safety stock made possible by consolidation effect of inventory from multiple locations into one location". This indicator of consolidation effect of inventory into location i is defined as

$$CE_i = 1 - \frac{CS_c + SS_c}{CS_i + SS_i}, \quad i = 1, ..., n.$$
 (13)

As shown the SD of the DPUTs is an important parameter for determining both SS_c and SS_i as defined in (6) and (11), respectively. In turn, in Section 3.1 we have reviewed that the SD is an important input for calculating CSk and CK. Therefore, it is expected that these indicators influence SS_c and SS_i , and in turn the total costs reached in supply systems with inventory pooling under the allocation rules of Ballou and Burnetas (2003) and Tyagi and Das (1998) or independent systems.

3.3. Computational framework and simulation scenarios

We present details of the computational framework utilized. We implemented our proposal in a non-commercial software named R; see http://www.r-project.org. See Rojas, Leiva, Wanke, and Marchant (2015), Rojas, Wanke, Coluccio, Vega-Vargas, and Huerta-Canepa (2020), Wanke, Ewbank, Leiva, and Rojas (2016) and Wanke and Leiva (2015) to visualize R applications in supply models.

The joint distribution when facing IID DPUT is simulated by using the copula method, occupying and R package denominated copula; see Hofert, Kojadinovic, Maechler, and Yan (2014). The copula package provides classes of commonly used Archimedean, elliptical, and extreme value copulas, as well as other copula families such Gaussian. This package also contains methods for computing values related to PDF and CDF, generating a random number, plotting tools, fitting of copula models and goodness-of-fit tests.

The mathematical programming of inventory pooling models is performed by an R package named nloptr; see Johnson (2014).

Recall that *m* is the number of centralized locations and *n* the number of decentralized locations, but the results presented here consider m = n = 2, because the results of this situation can be extended to any value of *m*, *n*, with $m \le n$.

The simulation study 1 of scenarios establish different: (i) inventory policies (supply and allocation), (ii) statistical models for the DPUT, and (iii) mathematical programming to minimize costs. We used three types of supplies for inventory policies (IS/RT/inventory centralization –IC–) and two allocation rules (Tyagi and Das/Ballou and Burnetas). The statistical modeling is based on IID DPUTs assuming a Weibull type 3 (WEI3) statistical distribution, see Stasinopoulos et al. (2020). We can estimate the parameters of the WEI statistical distribution proposed for the modeling of Y_1 and Y_2 , by using the histDist command of the gamlss package. We can also calculate the variances as:

$$\begin{aligned} \mathrm{sd}_{Y_1}^2 &= \mu_{Y_1}^2 \left\{ \frac{\Gamma(\frac{1}{\sigma_{Y_1}+1})}{\Gamma(\frac{1}{\sigma_{Y_1}+1})^2} - 1 \right\}, \text{ and} \\ \mathrm{sd}_{Y_2}^2 &= \mu_{Y_2}^2 \left\{ \frac{\Gamma(\frac{2}{\sigma_{Y_2}+1})}{\Gamma(\frac{1}{\sigma_{Y_2}+1})^2} - 1 \right\}, \text{ where } \Gamma(\cdot) \text{ is the gamma function.} \end{aligned}$$

We assumed several structures of skewness and kurtosis with bivariate WEI3 marginal statistical distributions for the DPUT. We used the following indicators with WEI3 statistical distribution to generate these 10,000 scenarios:

Statistical parameters

- $\begin{array}{l} \mu_{Y_1}, \mu_{Y_2} \sim {\rm U}(80, 120), \\ \sigma_{Y_1}, \sigma_{Y_2} \sim {\rm U}(0.8, 20), \end{array}$
- **Correlation parameter**
- $\rho_{Y_1,Y_2} \sim U(-1,1)$, Costs
- $DC_{T_{1,1}}, DC_{T_{2,2}} \sim U(0.20, 0.25),$
- $DC_{T_{1,2}}^{-1,1}, DC_{T_{2,1}}^{-2,2} \sim U(0.25, 0.30),$
- HC $\sim U(1, 1000),$
- OC₁ ∼ U(17, 67) and OC₂ ∼ U(20, 140);
- The choice of these values is based on previous studies on the topic; see Wanke (2009, 2014) and Wanke and Saliby (2009).

Mathematical programming is performed to minimize the total cost of inventory by using an objective function

$$TC_{c} = HC(CS_{c} + SS_{c}) + DC_{c} + OC_{c},$$
(14)

under Tyagi and Das' allocation rule, that is, when considering an IC. If the IS of supply is utilized, the mathematical programming for minimizing the total cost of inventory uses the objective function

$$TC_i = HC(CS_i + SS_i) + DC_i + OC_i, \quad i = 1, 2.$$
 (15)

Thus, such as Wanke (2009), we can compare IC and IS of supply based on (14) and (15), respectively, but now assuming WEI3 statistical distributions for the DPUTs. We call this Case 1. Furthermore, in this simulation study, we compare IC, RT, and IS of supply, such as in Wanke and Saliby (2009), employing the total cost of inventory expressed as

$$TC_{c} = HC(CS_{c} + SS_{c}) + DC_{c},$$
(16)

under both Tyagi and Das' and Ballou and Burnetas' allocation rules. Note that (16) is considered as the objective function to be minimized by the corresponding mathematical programming and compared to the optimization of (15). This is Case 2.

Recall that the TC_c defined in (14) is formed by the components C1, C2, C3, and C4, while, the TC_c defined in (16) only considers components C1, C2, and C3. Analogously for the TC_i defined in (15), where the cost components considered are the same. Then, for m = n = 2, we considered 10,000 different scenarios to minimize the corresponding total cost, by using mathematical programming for obtaining the optimum values of C1, C2, C3, and C4.

Our simulation study 2 is based on Wanke (2009), who showed that the behavior of the consolidation effect indicator given in (13) can be treated as a response variable described by a linear function according to its relationship with the ratio $R = Var(T_2)/Var(T_1)$. Then, we explored the relationships of the consolidated effect in respect to different parameters of inventory model, for segments of *R*.

3.4. Illustrative actual case

We illustrate our methodology with a real demand data set of two identical products with bivariate demand obtained from a mix of inventories that will allow us to show the methodology proposed. We first performed a descriptive statistical analysis for the demand data. Considering a goodness-of-fit analysis of the statistical distribution WEI3, we postulated it to theoretically describe the data and then estimated parameters that theoretically define their marginal statistical distribution and correlation. Second, we used inventory management models of centralization under the rules of Ballou and Burnetas (2003) and Tyagi and Das (1998) and compared their results of TCs with IS of supply, considering parameters, indicators, and the diverse costs involucrate.

In this illustration, we performed a statistical analysis of the demand for two identical products used in palliative COVID-19 treatments in anonymous Chilean public hospital pharmacies. Our proposal is to be able to optimize an inventory system for these services and reduce costs associated with the centralization or decentralization of decisions. The data to be analyzed corresponded to a monthly demand of salbutamol inhalers from two hospital pharmacies. The actual supply system of this product is of the type IS.

3.5. Statistical analysis

In simulation study 1, groups of scenarios were formed according to the grouping policy (IC, RT or IS) that achieves the lowest TC by allocation rules. We analyzed these results through the non-parametric Kruskall–Wallis test. This test makes it possible to relate the median of the different parameters of the inventory model used, in each of the groups (IC, RT, or IS) and to verify if these show statistically significant differences.

For statistical analysis of simulation study 2, we explored the relationships of the consolidated effect with respect to different parameters of inventory model, through B-Spline regression, as this approach provides a way to estimate the values of a response variable between fixed points called knots. A polynomial regression between nodes is calculated, where the splines are a series of chained polynomial segments that are joined into knots. The choice of the degree of freedom (minus one if there is an intercept) generates knots at suitable quantiles of the variable, which will ignore missing values. An analogous regression analysis considering normal statistical distribution for DPUTs can be consulted in Table 7 of Wanke (2009).

We apply descriptive statistics to the DPUTs data for the statistical analysis of our illustrative actual case, then we estimate the parameters of the WEI 3 statistical distribution of these variables, by using the histDist command of the gamlss package in R software. We check the goodness of fit by examining their quantile–quantile (QQ) plots, which allows us to verify differences between the probability distribution of a population from which a random sample has been drawn and a theoretical distribution such as WEI3 used for comparison. Its interpretation is simple, since the graph should follow a diagonal within confidence intervals. Finally, by means of the optimization methods indicated in the preceding subsections, we conclude the best method of grouping inventories for the actual case.

4. Results

4.1. Results simulation study I

Firstly, in the 10,000 scenarios proposed for both Cases 1 and 2 and their skewness and kurtosis frameworks for the DPUTs, we analyzed which supply system would provide the minimum total cost. The results show that the total costs described by (14) for Case 1 or (16) for Case 2 are smaller when generated by an IC. For Case 1, we get 7590 optimum scenarios under IC and 2410 with an IS of supply. The number of scenarios with IC, RT, and IS of supply for Case 2 are 6270, 1830, and 1900, respectively.

Secondly, using the 10,000 scenarios with IS, IC, and RT, we carried out a statistical analysis that is composed of two parts: (i) to identify the more relevant component (C1-C4) of the total cost from a quantitative perspective; and (ii) to determine whether statistically significant differences are presented for the association of the statistical indicators CSk, CK, and costs with the supply systems in Cases 1 and 2. See Fig. 3 to examine the shape of the statistical distribution and define the most important component of each total cost. Note that in all cases the values of the distribution median lead to the conclusion that C1 is the component with the highest proportion in total costs. We computed the number of scenarios where C1 is the larger component of total costs for Cases 1 and 2.

For Case 1 with DPUTs, C1 corresponds to the highest proportion of the total cost in 8950 and 1050 scenarios under IC and IS of supply, respectively. For Case 2 with DPUTs, C1 is the most relevant component in 4570, 5310, and 120 scenarios under modalities RT, IC, and IS of supply, respectively. We compared the value of cost associated with component C1 regarding the sum of costs related to the remaining components (C2, C3, and C4) using the χ^2 test for the difference of proportions and detected statistically significant differences at 1% in favor of C1 for both Cases 1 and 2 under IC, RT, and IS of supply. This allows us to conclude that C1 is the most relevant component of the total cost.

We used the Kruskal–Wallis (KW) test to compare medians of indicators related to statistical parameters, symmetry coefficient, kurtosis coefficient, and costs associated with IC, RT, and IS of supply. This test is carried out when minimum total costs are reached in Cases 1 (IC and

Table 2

Medians and corresponding Kruskal–Wallis *p*-value when comparing them for the model, indicator, and supply system mentioned based on the total cost for Case 1.

Indicator	IC	IS	KW p-value
μ_{Y_1}	100.80	100.08	0.015
μ_{Y_2}	100.78	99.98	0.868
σ_{Y_1}	10.78	9.12	< 0.001
σ_{Y_2}	11.05	8.94	< 0.001
sd_{Y_1}	221.11	205.64	< 0.001
sd_{Y_2}	226.61	201.33	< 0.001
CSk ₁	-0.63	-0.55	< 0.001
CSk ₂	-0.64	-0.57	< 0.001
CK1	3.57	3.52	0.009
CK ₂	3.56	3.56	0.094
ρ_{Y_1,Y_2}	-0.14	0.47	< 0.001
μ_{T_1}	2.96	2.66	< 0.001
μ_{T_2}	2.93	2.81	< 0.001
σ_{T_1}	1.26	1.26	0.338
σ_{T_2}	1.29	1.18	< 0.001
DC _T	0.124	0.13	0.020
DCT	0.126	0.12	< 0.001
DC _{T12}	0.374	0.37	< 0.001
DC _{T21}	0.375	0.38	0.036
HC	0.36	0.14	< 0.001
OC ₁	41.08	43.10	0.052
OC ₂	83.78	81.03	< 0.001

IS of supply) and 2 (IC, RT and IS of supply). Tables 2 and 3 present the results for Cases 1 and 2, respectively.

Note that the medians of ρ_{Y_1,Y_2} , σ_{Y_1} , σ_{Y_2} , CSk₁, CSk₂, CK₁, and CK₂ are associated with a decrease in the total cost under Case 2, as indicated in Table 3. Positive coefficients of correlation and small demand variability are associated with the IS of supply, confirming the findings of Wanke and Saliby (2009). It is also possible to see that minors CSk1, CSk2, CK1, and CK2, are associated with the IS of supply, indicating that the skewness and kurtosis of the statistical distribution of the data of DPUTs are factors to consider in the decision of the best supply system (centralized or non-centralized). As pointed out in Section 3.1, these parameters directly impact to the variances, correlations, skewness, and kurtosis of DPUTs , which are elements of SS_c and SS_i. Table 4 provides the medians of a selection of parameters that favor IC, RT, and IS of supply, considering only C1 or the HC of the safety stock (C1 = HC \times SS). We realize that only the highest kurtosis and correlations favor lower C1 with RT systems and the effect of the other indicators is significant, but not always with the same type of association in relation to the total costs under (16), which balances the variances, skewness, and kurtosis of the better DPUTs. Furthermore, in comparison to total costs under (16), ρ_{Y_1,Y_2} does not need to be so negative for inducing lower C1 in an IC. However, it must increase to highly positive values to favor lower C1 in the IS of supply and to positive values close to zero to induce lower C1 in the RT.

4.2. Results simulation study II

Table 5 shows the estimates of the regression coefficients obtained for the conditional expectance of the response variable related to consolidated effect (CE), standard errors (SE), and *p*-value of the corresponding *t*-test, with knots at suitable quantiles of R < 1. The adjusted R-squared obtained for the B spline regression model was 0.7762. Note that not all the indicators turned out to be statistically significant as predictors of CE, and some of them changed the direction of the relationship by changing the knots at suitable quantile of CE related with *R*. sd_{Y_1} , sd_{Y_2} and ρ_{Y_1,Y_2} , are always positive or negative predictors of CE, respectively, for both the indicators and their B splines. However, CSk_{Y_1} , CSk_{Y_2} , CK_{Y_1} and CK_{Y_2} , change the direction of the relationship with respect to CE according to the indicators or its B splines occupied as predictors. Table 3

Medians and corresponding Kruskal–Wallis p-value when comparing them for the model mentioned, indicator, and supply system, based on the total cost for Case 2.

Indicator	indicator IC IS RT		RT	KW p-value
μ_{Y_1}	100.92	101.98	98.27	< 0.001
μ_{Y_2}	100.66	99.93	101.21	0.797
σ_{Y_1}	10.47	9.60	11.27	< 0.001
σ_{Y_2}	10.93	8.88	10.93	< 0.001
sd_{Y_1}	218.77	207.73	223.35	0.002
sd_{Y_2}	224.62	199.46	232.23	< 0.001
CSk ₁	-0.63	-0.58	-0.62	< 0.001
CSk ₂	-0.64	-0.56	-0.64	< 0.001
CK1	3.52	3.50	3.68	< 0.001
CK ₂	3.56	3.61	3.56	0.009
ρ_{Y_1,Y_2}	-0.15	0.44	0.09	< 0.001
μ_{T_1}	2.96	2.68	2.83	0.002
μ_{T_2}	2.98	2.80	2.62	< 0.001
σ_{T_1}	1.24	1.29	1.37	< 0.001
σ_{T_2}	1.29	1.17	1.24	< 0.001
DCT	0.124	0.126	0.126	< 0.001
DCT	0.126	0.123	0.127	< 0.001
DCT	0.374	0.373	0.374	0.105
$DC_{T_{2}}$	0.375	0.376	0.374	0.026
HC	0.33	0.10	0.47	< 0.001
OC ₁	41.42	43.51	38.61	< 0.001
OC ₂	87.01	82.03	68.85	< 0.001

Table 4

Medians and corresponding Kruskal–Wallis *p*-value when comparing them for the indicator mentioned and supply system based on the component C1.

Indicator	IC	IS	RT	KW p-value
μ_{Y_1}	100.75	104.63	100.28	< 0.001
μ_{Y_2}	100.75	93.11	100.71	< 0.001
σ_{Y_1}	9.75	6.99	10.99	< 0.001
σ_{Y_2}	10.80	8.65	10.67	0.394
sd_{Y_1}	216.97	244.95	222.33	0.015
sd_{Y_2}	221.21	195.57	224.20	0.082
ρ_{Y_1,Y_2}	-0.09	0.90	0.05	< 0.001
CSk ₁	-0.61	-0.19	-0.63	< 0.001
CSk ₂	-0.63	-0.60	-0.63	0.169
CK1	3.52	3.74	3.58	< 0.001
CK ₂	3.56	3.42	3.57	0.546

4.3. Results from illustrative actual case

The bivariate data set from the illustrative actual case has a length of 36 months (see Section 3.4), and its descriptive statistical measures can be consulted in Table 6.

Table 7 shows estimated parameters for the modeling by WEI3 statistical distribution of Y_1 and Y_2 by GAMLSS.

In Fig. 4, we show an empirical quantile–quantile (QQ) plot to check the good standing of our proposed statistical distribution (WEI3) to model the DPUT. Since almost every values is within the confidence bands, we conclude that the statistical distribution proposed fits the data perfectly.

The rest of the coefficients required for carrying out the nonlinear optimizations under the centralization rules of Ballou and Burnetas (2003) and Tyagi and Das (1998) can be found in Table 8 with the results of TCs with respect to IS of supply compared.

Now we can compare the TC obtained under the centralization rules of Ballou and Burnetas (2003) and Tyagi and Das (1998) with the IS supply system. Under the Tyagi and Das (1998) allocation rule, $w_1 = 0.999$ (indicating centralization in localization 1) obtained a TC = 77.04 USD\$. Under the Ballou and Burnetas (2003) allocation rule, $w_p = 0.624$ (it is indicating an RT system) and obtained a TC = 99.42 USD\$, while the IS of supply obtained a TC = 87.53 USD\$, concluding that the best supply policy in this actual case is that of centralization in localization 1.



Fig. 3. (a) C1 in IC, RT, and IS, (b) C2 in IC, RT, and IS (c) C3 in IC, RT, and IS, (d) C4 in IC, RT, and IS.



Fig. 4. QQ plot for statistical distributions proposal for Y_1 and Y_2 .

Table 5

Estimate, its SE and p-value of the *t*-test of the parameter associated with the covariate indicated in the B-spline regression of the consolidated effect using a WEI3 model for DPUT, for the *R* knots.

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.1518	0.0468	3.24	0.0012
μ_{Y_1}	0.0001	0.0002	0.74	0.4580
$bs(\mu_{Y_1})1$	0.0005	0.0075	0.07	0.9440
$bs(\mu_{Y_1})2$	-0.0189	0.0074	-2.54	0.0110
μ_{Y_2}	0.0002	0.0001	1.50	0.1336
$bs(\mu_{Y_2})$	0.0119	0.0073	1.63	0.1036
$bs(\mu_{Y_2})2$	-0.0120	0.0072	-1.66	0.0965
σ_{Y_1}	-0.0072	0.0041	-1.70	0.0786
$bs(\sigma_{Y_1})$ $bs(\sigma_{Y_1})$	-0.0321	0.0143	-2.25	0.0248
σ_{v}	-0.0051	0.0047	-1.10	0.2727
$bs(\sigma_{Y_{1}})1$	-0.1165	0.0467	-2.50	0.0126
$bs(\sigma_{Y_2})2$	-0.0036	0.0156	-0.23	0.8169
CSk_{Y_1}	0.0404	0.0101	4.01	0.0001
$bs(CSk_{Y_1})1$	-0.0367	0.0225	-1.63	0.1035
$bs(CSk_{Y_1})2$	-0.1637	0.0367	-4.46	0.0000
CSk_{Y_2}	-0.0303	0.0119	-2.55	0.0108
$bs(CSk_{Y_2})$	0.0070	0.0209	0.33	0.7386
$\text{DS}(\text{CSK}_{Y_2})$	0.0965	0.0424	2.28	0.0227
SO_{Y_1}	0.0004	0.0001	4.10	0.0000
$bs(sd_{Y_1})$	0.1431	0.0308	3.94 4 44	0.0001
sd_y	0.0003	0.0001	2.95	0.0032
$bs(sd_y)1$	0.1111	0.0321	3.46	0.0005
$bs(sd_y)2$	0.0391	0.0127	3.09	0.0020
ρ_{Y_1,Y_2}	-0.2379	0.0023	-102.90	0.0000
$bs(\rho_{Y_1,Y_2})1$	-0.1173	0.0073	-16.16	0.0000
$bs(\rho_{Y_1,Y_2})2$	-0.0227	0.0073	-3.10	0.0019
CK_{Y_1}	-0.0204	0.0069	-2.96	0.0031
$bs(CK_{Y_1})1$	-0.0192	0.0142	-1.35	0.1757
$bs(CK_{Y_1})2$	0.1443	0.0443	3.26	0.0011
CK_{Y_2}	0.0115	0.0058	1.97	0.0491
$DS(CK_{\gamma_2})$	0.0430	0.0151	2.85	0.0043
$DS(CK_{Y_2})Z$	-0.1477	0.0459	-3.21	0.0013
μ_{T_1} bs(μ_{T_1})1	0.0266	0.0074	3.60	0.0003
$bs(\mu_T)^2$	-0.0111	0.0073	-1.52	0.1285
μ_{T_2}	0.0166	0.0012	14.02	0.0000
$bs(\mu_{T_2})1$	0.0219	0.0072	3.05	0.0023
$bs(\mu_{T_2})2$	0.0030	0.0076	0.39	0.6959
σ_{T_1}	-0.0496	0.0031	-16.05	0.0000
$bs(\sigma_{T_1})1$	0.0155	0.0074	2.09	0.0368
$bs(\sigma_{T_1})2$	-0.0260	0.0073	-3.58	0.0003
σ_{T_2}	-0.0059	0.0032	-1.84	0.0654
$DS(\sigma_{T_2})$	-0.0381	0.0075	-5.07	0.0000
$DS(\delta_{T_2})$	-0.0210	0.0074	-2.93	0.0034
$bc_{T_{1,1}}$	-0.0379	0.0930	-0.40	0.0000
$b_{T_{1,1}}$	-0.0119	0.0073	-1.62	0.0035
$DC_{T_{1,1}}$	0.5270	0.0941	5.60	0.0000
$bs(DC_{m})1$	0.0158	0.0074	2.14	0.0324
$bs(DC_{T_{2,2}})^2$	-0.0434	0.0074	-5.91	0.0000
DC _T	-0.2831	0.0963	-2.94	0.0033
$bs(DC_T)$ 1	-0.0119	0.0072	-1.64	0.1004
$bs(DC_T)$	0.0035	0.0076	0.46	0.6463
DCT	0.0527	0.0914	0.58	0.5645
bs(DC _{Tat})1	0.0043	0.0072	0.59	0.5550
bs(DC _{T21})2	-0.0186	0.0074	-2.52	0.0117
HC	-0.0714	0.0069	-10.34	0.0000
bs(HC)1	-0.0388	0.0073	-5.31	0.0000
bs(HC)2	0.0060	0.0073	0.83	0.4093
OC ₁	-0.0003	0.0001	-3.54	0.0004
bs(OC ₁)1	0.0008	0.0074	0.11	0.9130
bs(OC ₁)2	-0.0043	0.0075	-0.57	0.5686
	0.0003	0.0000	8.77	0.0000
$DS(OC_2)$	0.0000	0.0075	0.00	0.9994
DS(UC, JZ	-0.0038	0.0073	-0.52	0.0041

Table 6

Descriptive measures for the monthly d	demand bi	oivariate (data se	et indicated.
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Measure	Y_1 (units/month)	Y_2 (units/month)
Min. (units/month)	16.14	84.01
1st Qu. (units/month)	58.36	106.34
Median (units/month)	79.89	113.39
Mean (units/month)	97.82	112.29
3rd Qu. (units/month)	127.77	121.29
Max. (units/month)	263.21	129.95
CSk (dimensionless)	0.62	-0.64
CK (dimensionless)	3.06	3.75
ρ_{Y_1,Y_2} (dimensionless)	-0.35	p < 0.05

Table 7

Estimated parameters for modeling by WEI3 statistical distribution of Y_1 and Y_2 by GAMLSS.

Estimated parameters	Y_1 (units/month)	Y_2 (units/month)
μ (units/month)	98.4	112
σ (dimensionless)	1.7	12.3
sd ² (units/month) ²	7926.34	71 588.35
AIC (dimensionless)	392	278

5. Management implications in decision making

Next we discuss the implications of our results for decision making management in the field of logistics by first referring to the importance of considering skewness and kurtosis patterns that DPUT data can have when inventory consolidation is desirable. With the results obtained, we conclude that these skewness and kurtosis patterns have a direct influence on parameters related to the variance and correlation of DPUTs. These parameters directly impact the component C1 of the total costs of inventory, which is the most relevant component in all the methods of inventory pooling and IS of supply; see Fig. 1. A simple way to detect skewness and kurtosis patterns of the DPUT is by using histograms. For an interpretation of these graphs, the interested reader is referred to Brown (1997). A selection of the main implications is provided next:

- In general, DPUT distributions with a more negative skewness (CSk < 0) and more leptokurtic, that is to say more pointed and with thicker tails than normal (CK > 3), favor to a greater extent obtaining lower total costs with RT systems. The RT of supply is favored given that it is associated with an increase in the variances and correlations of DPUTs in these conditions of symmetry and kurtosis.
- In general, the above conditions of skewness and kurtosis favor a lower C1 with RT systems. On the contrary, more positive skewness of DPUTs and less leptokurtic, meaning tails not so heavy (CK tending to 3), favor to a greater extent the obtaining of a lower C1 with IS systems.
- The statistical indicators with the highest direct impact on the percentages of consolidated effect are ordered as: $DC_{T_{2,2}} > bs(sd_{Y_1})1 > bs(CK_{Y_1})2 > bs(sd_{Y_2})1 > bs(sd_{Y_1})2 > bs(sd_{Y_2})2 > bs(\mu_{T_1})$ 1 > bs(μ_{T_1})2 > μ_{T_1} .
- The statistical indicators with the highest inverse impact on the percentages of consolidated effect are ordered as: $DC_{T_{1,2}} > \rho_{Y_1,Y_2}$ > $bs(CSk_{Y_1})2 > bs(CK_{Y_2})2 > bs(\rho_{Y_1,Y_2})1 > HC > \sigma_{T_2} > bs(DC_{T_{2,2}})2$.

Fig. 5 shows a decision-making scheme developed to facilitate managing the supply decision of centralized or decentralized inventories according to skewness and kurtosis of the demand. To determine which type of inventory to pick, both skewness (CSk) and kurtosis (CK) must be evaluated simultaneously. The blue arrow indicates the direction to take that favors IC, the red arrow indicates the direction to take toward an IS, and the yellow arrow indicates the direction to take that favors RT.



Fig. 5. Decision-making scheme developed to facilitate managing the supply decision of centralized or decentralized inventories according to skewness and kurtosis of the demand.

Table 8

Coefficients for carrying out optimization in actual case.										
μ_{T_1}	μ_{T_2}	σ_{T_1}	σ_{T_1}	OC	OC ₂	DCT	DC _{T22}	DC _{T12}	DC _{T21}	HC
(Days)	(Days)	(Days)	(Days)	(USD\$/order)	(USD\$/order)	(USD\$)	(USD\$)	(USD\$)	(USD\$)	(USD\$/units)
1.82	1.11	0.76	1.09	45.64	128.98	0.11	0.14	0.40	0.38	0.18

6. Discussion

Coronavirus is not only a disease that has a devastating effect on people's lives by attacking the respiratory system, but it has a very fast spread rate, which has posed a big challenge to the related medicine production. The production process does not only need to focus on a timely medicine availability to meet a growing volume of demand, but also on how to simultaneously control the cost, which is the issue faced by all the producers. Since the outbreak of coronavirus, efforts have been made to investigate the issues of resource allocation, medical cost as well as supply of medical related resources (Bartsch et al., 2020; Zhang, Yao, Wang, Long, & Fu, 2020). The above challenge was not only faced by the medical producers, but most importantly the hospitals that deal with this virus by directly contacting and treating the patients. Hospitals play a key role in slowing down the spread of the virus and there was a dilemma that was faced by the hospitals which was the significant increase in the number of patient's but limited amounts of resources in terms of medical related staffs, equipment, and medicine. Keeping these issues in mind, the investigation of pooling inventory policies would be of particularly relevant, useful, and important for policy making purposes by the government, through which a better decision-making process can be facilitated during the Covid-19 crisis.

The aim of this study is to provide solutions to the policy makers in terms of how to mitigate inventory stock-out of critical supplies and significantly improve their ability by applying a novel stochastic inventory optimization model. Although this type of model has been proposed by a few studies in different contexts recently (Jackson, Tolujevs, & Reggelin, 2018; Pirhooshyaran & Snyder, 2020), we are the first piece of research that uses this model in the context of COVID-19. The development of these practices would be of very paramount not only for the government as well as health care ministries in Chile, but it is very useful to develop a dynamic communication mechanism between hospitals transnationally, so it is recommended that the health managers should be actively engaged in this inventory policy approach and relevant trainings should be provided to them in terms of further improving their technical abilities of using R and Phyton, so as to better understand the policies.

The usefulness of the current study also lies in the fact that the topic area is relevant not only to academic research in the area of operations research, but also very practical in the real world in terms of management of decisions. In the literature, there are a number of studies that investigate the issues of inventory control. Demand per unit of time is one of the areas of interest.Cobb, Rumí, and Salmerón (2013) investigated the optimal policies in inventory management when demand per unit of time follows a log-normal distribution. In comparison, while an inventory model is formulated by Pando, Garci, San-José, Sicilia, et al. (2012) in which the demand per unit of time is a concave potential function of the inventory level. The stochastic inventory model was also examined by Hayya, Harrison, and Chatfield (2009) in which the demand per unit of time and the lead time random variables that are distributed independently and identically. Our results suggest new decision-making patterns such as considering the skewness and kurtosis of data related to demands per unit of time since it has shown its effect on the variance, which interacts with the best allocation rule for inventory pooling in terms of the consolidated safety stock multiplied by the holding cost and by the total cost. More precisely, in the presence of negative skewness and high kurtosis for the demands per unit of time, there must be fewer safety stocks and less total costs under the Ballou and Burnetas (2003) allocation rule that leads to inventory pooling for an RT system, implying that a contrary case favors dedicated facilities for the supply in terms of total cost. For the cost component related to the consolidated safety stock multiplied by the holding cost, the regular transshipment had a better statistical performance than the independent system of supply. This is in accordance with Reyes and Meade (2006). We propose the methodology, which was implemented using the R software, facilitated by relevant R code . The methodology adopted in the current study possesses the advantages of being able to undertake a simulation study based on data with Skewness and Kurtosis.

Regarding future research, this work is expandable to more general models with heteroscedasticity of variance, such as the widely known variants of generalized lineal models, which are very flexible allowing linear and non-linear functional structures. In addition, in line with this work, the methodology to model a time series, such as the generalized autoregressive and moving average (GARMA) model can be explored and its effects over CE. Theoretically, GARMA has been discussed by relevant studies (Albarracin, Alencar, & Ho, 2019; Gomes, Morettin, Cordeiro, & Taddeo, 2018), however little effort has been made to apply this model to the real world practice, in particular in the area of inventory management. Furthermore, multivariate time series may be also considered. It is important to consider these new probabilistic approaches in cost reduction studies by potential locations for consolidation facilities that combine shipments, thus improving the level of service. Some of these issues are being analyzed by the authors whose findings will be reported in future articles.

Due to the global outbreak of coronavirus at the beginning of this year, there has been an unprecedented pressure and challenge faced by hospitals across all countries. This also posed important questions in inventory management in the healthcare industry. Previously the empirical research focused on the cost perspective and in particular how to balance the cost and availability in the inventory management (Moons, Waeyenbergh, & Pintelon, 2019; Saedi et al., 2016), while nowadays and in the future, we think the focus of the research in the area will be oriented to investigating how to manage the inventory during unexpected events. In other words, the future research trend will focus more on cost savings, inventory availability, uninterrupted supply, as well as on dealing with unexpected/unpredicted change in demand. Our study can be regarded as a pioneer in this research perspective by investigating inventory pooling decisions by considering the skewness and kurtosis of data.

7. Conclusions

This research developed a novel inventory pooling model to assist in determining order sizes and safety inventories in public hospital warehouses, which have been dramatically impacted by COVID-19 challenges as regards availability of distinct items in distinct warehousing locations. Besides, the model developed here contributes to the current body of literature on inventory management, not only by overcoming the intractability of kurtosis and skewness in classic inventory models, but also by devising a framework to assess inventory consolidation gains. As long as all the moments of the distribution of the demand during the lead-time are related, and thus impact on inventory consolidation decisions, the copula method - through the parametric specification and association of the marginal statistical distributions that make-up a multivariate joint distribution - was employed for the first time to address inventory pooling issues. This paper is also helpful to hospital inventory managers and practitioners, ascertaining, under more realistic assumptions, that medicines have high availability at a low cost in different hospital warehouses. Therefore, in summary, our paper makes significant contributions from both theoretical and practical perspectives.

Numerical simulations and an illustrative example were considered under the assumption of a continuous-review model while testing for distinct inventory allocation rules as regards inventory centralization and regular transshipment. Overall results indicated that distributions of demands with more negative skewness and high kurtosis favor are keys for obtaining lower total costs with regular supply transshipment systems. Based on the results, we also provided relevant practical policy implications, which was clearly presented in a good detail in the discussion section. Future studies should address not only distinct business environments as regards inventory cost parameters but should also consider different modeling assumptions, such as the periodic-review and the order-up-to-level regime.

As regards the limitations of the proposed model, they are mainly related to modeling multivariate demand considering that the marginal distributions of the warehouse demands can be specified by the Generalized Additive Model for Location, Scale and Shape. While such approach offers advantages to model demands considering virtually any marginal statistical distribution, it falls short in addressing autocorrelated demands at each warehousing location. This could also constitute a research topic for future studies.

CRediT authorship contribution statement

Fernando Rojas: Conceptualization, Methodology, Software, Writing – original draft. **Peter Wanke:** Visualization, Investigation, Supervision, Software, Validation. **Fernando Bravo:** Data curation, Writing – review & editing. **Yong Tan:** Writing – review & editing.

Acknowledgment

This research was carried out thanks to the funding of the Fondecyt initiation project code: 11190004, Chile.

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