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Establishment and Implementation of Evidence-Based Opioid Prescribing Guidelines in Cardiac Surgery

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Abstract

Background.—Despite the risk of new persistent opioid use after cardiac surgery, postdischarge opioid use has not been quantified and evidence-based prescribing guidelines have not been established.

Methods.—Opioid-naive patients undergoing primary cardiac surgery via median sternotomy between January and December 2019 at 10 hospitals participating in a statewide collaborative were selected. Clinical data were linked to patient-reported outcomes collected at 30-day follow-up. An opioid prescribing recommendation stratified by inpatient opioid use on the day before discharge (0, 1–3, or 4 pills) was implemented in July 2019. Interrupted time-series analyses were performed for prescription size and postdischarge opioid use before (January to June) and after (July to December) guideline implementation.

Results.—Among 1495 patients (729 prerecommendation and 766 postrecommendation), median prescription size decreased from 20 pills to 12 pills after recommendation release (P < .001), while opioid use decreased from 3 pills to 0 pills (P < .001). Change in prescription size over time was +0.6 pill/month before and -0.8 pill/month after the recommendation (difference = -1.4 pills/month; P = .036). Change in patient use was +0.6 pill/month before and -0.4 pill/month after the recommendation (difference = -1.0 pills/month; P = .017). Pain levels during the first week after surgery and refills were unchanged. Patients using 0 pills before discharge (n = 710) were prescribed a median of 0 pills and used 0 pills, while those using 1 to 3 pills (n = 536) were prescribed 20 pills and used 7 pills, and those using greater than or equal to 4 pills (n = 249) were prescribed 32 pills and used 24 pills.

Conclusions.—An opioid prescribing recommendation was effective, and prescribing after cardiac surgery should be guided by inpatient use.

Opioid overprescribing after surgery is common.^{1–6} A larger opioid prescription size has been associated with higher opioid use⁶ and an increased risk of new persistent opioid use 3 to 6 months after surgery in previously opioid-naive patients.⁸ New persistent opioid use has been reported in 3% to 8% of patients after minor and major surgery,^{9–11} over 10% of cancer patients undergoing curative surgery,¹² and up to 17% of patients undergoing cardiothoracic surgery.^{8,13–15} In addition, overprescribing has been shown to increase the risk of opioid overdose in patients' family members through opioid diversion.¹⁶

Despite high rates of new persistent opioid use and large prescription sizes at discharge,^{8,14,17} patient-reported opioid use after cardiac surgery has not been described and evidence-based prescribing guidelines have not been established. With the sharp rise

of opioidrelated drug overdoses¹⁷ prompting the declaration of the opioid crisis as a public health emergency, considerable efforts to reduce opioid use through evidence-based prescribing guidelines after general surgery procedures have been largely successful.^{1,5,18–20} However, many of these guidelines were established for outpatient procedures, and may not translate to cardiac surgery because patients undergo median sternotomy and spend a median 6 to 12 days in the hospital, depending on the pro-cedure.²¹ A single-center study of inpatient general surgery found that inpatient opioid use before discharge was associated with postdischarge use and could be utilized to guide prescribing through an individualized approach,²² which may be similarly beneficial in cardiac surgery patients. At present, appropriate opioid prescription quantities after cardiac surgery remain unknown, and the relationship between inpatient and postdischarge use has not been characterized. Over 300,000 cardiac surgery procedures are performed annually, and these procedures confer among the highest published rates of new persistent opioid use. Therefore, establishing evidence-based opioid prescribing recommendations should be prioritized to guide responsible prescribing and optimize patient care.

This prospective, multicenter collaboration was organized to (1) define opioid prescribing and patient-reported use after cardiac surgery, (2) assess inpatient opioid use as a predictor of postdischarge opioid use, (3) establish opioid prescribing recommendations for opioid-naive patients, and (4) evaluate the implementation of statewide recommendations by analyzing postintervention prescribing and patient-reported outcomes (PROs). We hypothesized that implementing recommendations would decrease both opioid prescribing and use and that incorporating inpatient opioid use would allow for more personalized prescribing.

Patients and Methods

This study was deemed exempt by the University of Michigan Institutional Review Board (HUM00156194).

Data Source

Clinical data were collected through the Michigan Society of Thoracic and Cardiovascular Surgeons(MSTCVS) Quality Collaborative, developed in 2001 as a cardiac surgeon–led quality collaborative embedded in the MSTCVS, including all 33 nonfederal hospitals perform-ing cardiac surgery in Michigan. The Collaborative receives standardized harvest files sent by each of the hospitals to The Society of Thoracic Surgeons national database.

Opioid naivite was defined as not using opioids at time of admission for surgery. PROs including postoperative pain scores, number of pills used postdischarge, number of days using an opioid, refills, and use of adjuncts were captured in an 11-item questionnaire prospectively administered at 30-day clinical follow-up (Supplemental Appendix 1). Inpatient opioid use during the 24-hour calendar day before discharg and prescription informa-tion were collected through chart review by data management team members at each hospital. Although perioperative patient instructions were not uniform across all 10 centers, each patient receiving a opioid prescription at all 10 centers reviewed the Opioid Start Talking form (Supplemental Appendix 2) with a provider according to state law.

Questionnaire responses were collected and managed using REDCap (Research Electronic Data Capture) tools hosted at the MSTCVS Coordinating Center.²³ Questionnaire responses were linked to clinical data by unique Society of Thoracic Surgeons record identification numbers and verified with dates of surgery and discharge.

Patient Population

Patients undergoing cardiac surgery via median sternotomy between January and December 2019 at 10 centers participating in the MSTCVS Quality Collaborative were identified (n = 3650). From these, operative mortalities (n = 153), those with postoperative length of stay greater than 30 days (n = 74), discharge to a location other than home (n = 612), and reoperations (n = 268) were sequentially excluded. From the resulting population (n = 2543), 1685 (66%) reported PROs, including 1495 (89%) opioid-naive patients (Figure 1). Clinical characteristics, inpatient opioid use, and prescription size were collected for all patients and compared between responders and nonresponders.

Opioid Prescribing Recommendations

Opioid prescribing and PROs from January to March 2019 were analyzed to form an initial recommendation. The final recommendation was established through multidisciplinary input from the Michigan Opioid Prescribing Engagement Network, hospital data managers, and surgeon leaders from all 10 centers (Figure 2). This recommendation was formed in June and disseminated in early July. Throughout the 1-year study period, all 10 hospitals collected day before discharge opioid use, prescribing data, and PROs at follow-up. The prerecommendation period was defined by operative dates January 1 to June 30, while the postrecommendation period included July 1 to December 31.

Statistical Analysis

Primary outcomes were prescription size and patient-reported, postdischarge opioid use. Secondary outcomes included inpatient opioid use before discharge, refills before follow-up, percent of patients not prescribed an opioid at discharge, and PROs including pain in the first week after surgery (with 1 denoting none, 2 denoting minimal, 3 denoting moderate, and 4 denoting severe) and use of nonopioid adjuncts.

Bivariate comparisons utilized paired, 2-tailed *t*-tests or Wilcoxon rank sum tests for continuous data and chi-square testing for categorical data, as appropriate. All opioid prescription and use data were initially converted to oral morphine equivalents and quantified in 5-mg oxycodone pills. Opioid prescribing and use was quantified in median (interquartile range [IQR]) and mean \pm SD number of pills. Prescription size and patient-reported use were evaluated overall and by hospital. Interrupted time-series analyses were performed comparing discharge opioid prescription size and postdischarge use prerecommendation and postrecommendation. The slope during each period represents the trend in number of pills per month, while both the differences in prescription size and postdischarge use (level change) and the difference in the trends of each before and after the recommendation release (trend change) were evaluated. The percent of patients not prescribed an opioid at discharge, opioid refill rate, and pain levels were displayed by 3-month periods to show trends related to the release of recommendations.

As in prior analyses,²² the unadjusted relationship between prescribing and postdischarge use was depicted after stratifying groups by opioid use on the day before discharge (0, 1–3, or 4 pills). Group definitions were consistent with the terciles of inpatient opioid use determined by the initial data used to develop the prescribing recommendation. Guideline adherence was measured by quantifying the number of patients in each tercile who were prescribed the recommended number of pills or fewer as outlined in Figure 2. Univariate analyses of postdischarge opioid use were performed with Wilcoxon rank sum tests for binary variables or else negative binomial regression. Multivariable mixed-effects negative binomial regression was performed to identify independent predictors of postdischarge opioid use. The primary exposure was inpatient opioid use during the day before discharge, while covariates included age, sex, race, ethnicity, primary insurance payor, diabetes, cerebrovascular disease, hypertension, smoking history (nonsmoker, former smoker, active smoker), peripheral vascular disease, heart failure, cancer within 5 years, depression, postoperative complication, operative status (elective, urgent, or emergent), pain during the first week after surgery, length of stay, surgery type, and hospital (random effect).

Statistical significance was defined as a *P* value less than .05. All analyses were performed in Stata 16.0 (Sta-taCorp, College Station, TX).

Results

Patient Demographics and Outcomes

Among 2257 eligible opioid-naive patients, respondents (n = 1495) compared with nonrespondents (n = 762) were prescribed 2 less pills at discharge, had a mean 0.5day shorter hospital stay, and a lower proportion had a history of depression; patient demographics, clinical characteristics, and inpatient opioid use on the before discharge were otherwise similar (Supplemental Table 1). Among 1495 opioid-naive respondents across 10 centers, mean age was 64 ± 11 years and 26% were women. In total, 55% underwent isolated coronary artery bypass grafting (CABG), 30% underwent an isolated valve procedure, and 12% underwent CABG and valve procedures, most of which were elective (n = 892, 60%). A higher proportion of postrecommendation compared with prerecommendation patients had a history of hypertension, underwent isolated CABG, and were urgent rather than elective; preoperative characteristics otherwise did not statistically differ by study period (Table 1).

Mean postoperative length of stay was statistically shorter postrecommendation compared with prerecommendation (6.5 ± 3.0 days vs 6.9 ± 3.0 days; P = .027), whereas complication rate did not differ (post: n = 289 of 766 [38%] vs pre: n = 302 of 729 [41%]; P = .14). Inpatient opioid use on the day before discharge decreased from a median of 0.7 (IQR, 0–2.7) pill to 0.6 (IQR, 0–2.0) pill (P = .023) (Table 1).

Opioid Prescribing and Postdischarge Use

Median prescription size decreased from 20 (IQR, 0–37) pills prerecommendation to 12 (IQR, 0–24) pills postrecommendation (P<.001), while postdischarge use decreased from 3 (IQR, 0–16) pills to 0 (IQR, 0–11) pills (P<.001) and varied by hospital

(Supplemental Figure 1). Interrupted time-series analyses showed a significant difference between prerecommendation and postrecommendation prescribing (P = .003) and use (P = .002). The change in prescription size over time was +0.6 pill/ month (95% confidence interval [CI], -0.2 to +1.4) before and -0.8 pill/month (95% CI, -1.8 to +0.2) after the recommendation, representing a significant difference in slope of -1.4 pills/month (95% CI, -2.7 to -0.1; P = .036). The change in postdischarge use was +0.6 pill/month (95% CI, +0.1 to +1.2) before and -0.4 pill/month (95% CI, -1.0 to +0.2) after the recommendation, representing a significant difference in slope of -1.0 pills/month (95% CI, -1.8 to -0.2 pills; P = .017) (Figure 3).

Patient-Reported Outcomes

While a higher proportion of patients were not prescribed an opioid at discharge after the release of recommendations (pre: 25.2% vs post: 32.6%; P=.002), opioid refill rate (pre: 13.1% vs post: 13.8%; P for trend = .93) and patient-reported pain levels during the first week after surgery (P for trend = .89) did not significantly change (Figure 4). Additionally, patient-reported use of nonopioid adjuncts increased from 51% (n = 353 of 687) prerecommendation to 61% (n = 435 of 715) postrecommendation (P < .001). PROs before and after recommendation release varied by hospital (Supplemental Figure 2) and were collected at a median 29 (IQR, 25–36) days after discharge, which did not differ prerecommendation vs postrecommendation (29 [IQR, 25–35] days vs 29 [IQR, 25–37] days; P=.21).

Stratification by Inpatient Opioid Use and Guideline Adherence

Among the 47% (n = 710 of 1495) of patients who used 0 pills on the day before discharge, nearly half (n = 316 of 710, 45%) were still prescribed a median 19 (IQR, 12–28) pills. Utilizing the same groupings of inpatient use as in the prescribing recommendations (0 pills [n = 710], 1–3 pills [n = 536], and 4 pills [n = 249]), both prescription size and postdischarge use increased significantly with each group of increasing inpatient use (Figure 5). Median prescription size was greater than postdischarge use for patients with inpatient use before discharge in all 3 groups: 0 pills (prescription size: 0 [IQR, 0–18] pills vs use: 0 [IQR, 0–0] pills; *P*<.001), 1–3 pills (prescription size: 20 [IQR, 13–28] pills vs use: 7 [IQR, 1–17] pills; *P*<.001), and greater than or equal to 4 pills before discharge (prescription size: 32 [IQR, 24–40] pills vs use: 24 [IQR, 10–40] pills; *P*<.001). Median postoperative length of stay for each group was 6 (IQR, 5–8) days, 6 (IQR, 5–7) days, and 5 (IQR, 5–7) days, respectively. Among postrecommendation patients, 52% (n = 397 of 766) were prescribed a number of opioid pills adherent to the prescribing recommendations, whereas 35% (n = 254 of 729) of prerecommendation patients were prescribed quantities adherent to the recommendation (*P*<.001).

Predictors of Postdischarge Use

Multivariable analysis indicated that inpatient opioid use the day before discharge was independently predictive of postdischarge opioid use. Additional predictors of postdischarge use included younger age, absence of cerebrovascular disease, absence of cancer within the past 5 years, patients with Medicare rather than no insurance, higher pain levels, and shorter postoperative length of stay (Table 2).

Comment

In this multicenter collaboration to establish, implement, and evaluate opioid prescribing recommendations after cardiac surgery, both opioid prescribing and patient-reported use decreased significantly, while patient-reported pain levels and opioid refills remained unchanged. In addition, use of nonopioid adjuncts and the proportion of patients not prescribed an opioid at discharge increased after release of the recommendations. Collectively, these findings suggest that inpatient opioid use should guide prescribing and that the proposed opioid prescribing recommendations for opioid-naive patients undergoing cardiac surgery (Figure 2) were effective. Moreover, postoperative opioid use continued to decline, suggesting that future efforts could employ more aggressive opioid-sparing approaches.

Opioid prescribing guidelines have been successfully implemented for a wide range of procedures.^{5,18–20} However, many of these procedures are outpatient or involve a short hospital stay, whereas cardiac surgery patients may spend a week or more as an inpatient. Administrative claims analyses have described new persistent opioid use after cardiac^{7,10,13,16} and general thoracic^{7,9,11,12,14} procedures, and efforts have been initiated to establish thoracic surgery prescribing guidelines.4 However, the implementation of these thoracic guidelines has not been evaluated and only recommendations based on single-center expert opinion have been proposed for cardiac surgery.²⁴ A study of inpatient general surgery concluded that inpatient opioid use before discharge should guide prescribing, rather than a single prescription quantity.²² Our study reinforces these findings in a cardiac surgery population and expands these data by implementing a stratified, 4-part prescribing and pain management recommendation across 10 hospitals. Interestingly, only 11% of patients undergoing thoracic⁴ and 15% of patients undergoing inpatient general surgerv²² were not prescribed an opioid at discharge, vs 29% of cardiac patients in this study. However, we also found that among the 47% of patients not using opioids on the day before discharge, 45% were still prescribed an opioid. Future efforts at reductions in prescribing should focus on patients not using opioids before discharge, as this cohort appears to require few, if any, opioids after discharge.

Importantly, patient-reported pain scores during the first week after surgery remained stable after guideline release, at a mean between 1 (none) and 2 (minimal), qualitatively lower than prior general surgery studies utilizing the same scale.¹⁹ Although opioid refill rates also did not change, the overall rate was 13.4% (hospital range, 3.4%-23.3%) (Supplemental Figure 2D), substantially higher than published rates of 0.4%,¹⁸ 3.2%,⁵ and 6.5%²² in major general surgery series. The high overall refill rate and substantial hospital variation in conjunction with a 29% rate of patients not prescribed an opioid at discharge together highlight the wide variability in opioid requirements among cardiac surgery patients, supporting a more individualized approach to prescribing compared with other surgical procedures. Variation in these data across hospitals may also reflect differences in implementation, as well as the overall adherence to these recommendations for only 52% of patients overall. Given the effectiveness demonstrated in this study, future efforts examining the barriers and facilitators of implementation will be critical for sustainable change. For example, hospital practices in handling refill requests varied widely among

collaborating centers, depending on factors such as patient driving distance to the hospital or resources to support in-person clinical assessments. The emergence of the coronavirus disease 2019 (COVID-19) pandemic could further impact and exacerbate existing barriers to in-person assessments for postoperative pain. Patients may feel hesitant or fearful to return for clinical assessment due to COVID-19 concerns. Providers may increase prescription sizes to counteract the increased barriers to in-person follow-up during the pandemic or more readily prescribe refills. These practice patterns and challenges should be continually assessed, depending on local factors including COVID-19 prevalence, hospital resources, and patient preferences.

Independent predictors of high postdischarge opioid use included younger age, shorter postoperative length of stay, higher inpatient opioid use, and higher patient-reported pain levels. Younger age was also identified as an independent predictor of postdischarge opioid use after general surgery^{6,22} and has additionally been identified as an independent predictor of new persistent opioid use 3 to 6 months after thoracic, 7,10,12 cardiac, 7,10 and other types of major elective surgery.¹⁰ Shorter length of stay as a predictor of higher postdischarge opioid use reinforces the principle that cardiac patients should not be treated uniformly, but rather stratified depending on inpatient requirements before discharge. In doing so, we identified distinct cohorts, with 47% of patients not using opioids the day before discharge and a median 0 pills after discharge, whereas patients using greater than or equal to 4 pills before discharge used a median 24 pills postdischarge (Table 2 and Figure 5). Because opioid-naive status was defined at admission, the cohort with increased opioid requirements may have had recent or remote opioid use, which may have altered their opioid tolerance and thresholds for adequate pain relief. In addition, the group with higher use could be due to a shorter length of stay, although the significance of this difference between groups (median 6 days vs 5 days) is unclear without collecting data on earlier inpatient use.

The clinical significance of a 0.1-pill decrease in the median opioid use on the da0y before discharge is unclear, but the significant trend of decreasing inpatient opioid use during the study may represent improved inpatient opioid weaning. Although not measured in this study, perioperative opioid-sparing analgesia and early inpatient opioid weaning may further reduce prescribing and postdischarge use and have been described as components of cardiac enhanced recovery after surgery pathways.^{25–27} Enhanced recovery after surgery programs comprehensively address perioperative management but remain difficult to formally evaluate due to inconsistent compliance and multiple competing interventions, and have not consistently addressed opioid prescribing.^{25–27} Because this prescribing recommendation stratifies patients by inpatient use before discharge rather than a single quantity, it could be integrated into any existing perioperative pain management program to meet the needs of a diverse spectrum of cardiac surgery patients. Finally, this recommendation should be continually reassessed through learning health system principles of quality improvement based on routine practice data.²⁸

This study has several limitations. First, PROs are subject to recall bias. To partially mitigate this, outcomes were collected in a standardized questionnaire at a relatively consistent interval after discharge (median 29 [IQR, 25–36] days). Second, PROs were provided by 66% of the total population, which may impact the generalizability of these data.

However, clinical characteristics, inpatient opioid use, and prescribing data were collected for 100% of patients, and differences between responders vs nonresponders were minimal. Furthermore, a 66% response rate compares favorably to the 56% rate reported in prior work utilizing similar methodology.¹⁹ Finally, inpatient nonnarcotic pain protocols were not standardized across hospitals during the study period and may have influenced pain management. However, the impact of differing inpatient protocols likely remained consistent before and after guideline implementation at individual centers.

In conclusion, opioid prescribing decreased by 40% and postdischarge opioid use decreased from a median of 3 to 0 pills, while pain levels and refills did not change after implementation of an opioid prescribing recommendation based on inpatient opioid use. The proposed guidelines appeared effective and should be continually reassessed through a learning health system framework.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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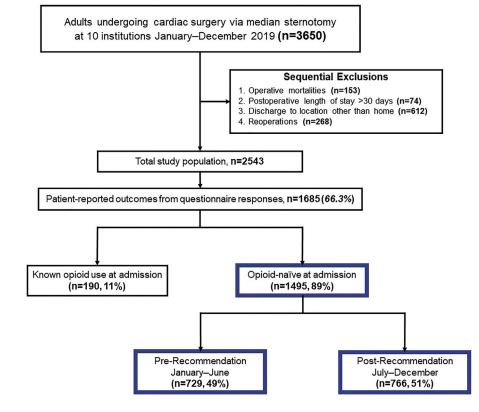


Figure 1.

Patient population STROBE diagram.

After Cardiac Surgery via Median Sternotomy

P	rescribe around-the-cloc Inpatient: 1000mg oral ace Post-discharge: 1000mg e	etaminophen every 8 hours	5	# Pills Used on Day	Prescribe
				Before Discharge:	(# Pills):
	In-hospital opioid use 5mg oxycodone to be t	•	0.0	0	0 – 4
	Sing oxycouone to be t	aken every o nours as	needed.	1 – 3	15
				≥ 4	25
	Recommend clinic visit prior to opioid refills	t for assessment			MARK US
		Distribute opioid edu including instructions		• •	
	MADIOVASCULE			iaus	OPEN

Figure 2.

Opioid prescribing recommendations for opioid-naive patients undergoing cardiac surgery.

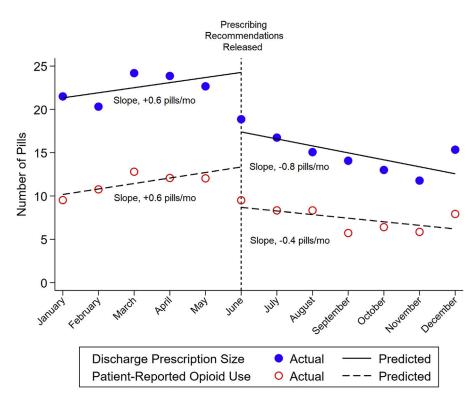


Figure 3.

Interrupted time-series analyses of opioid prescription size and patient-reported postdischarge opioid use before and after release of prescribing guidelines.

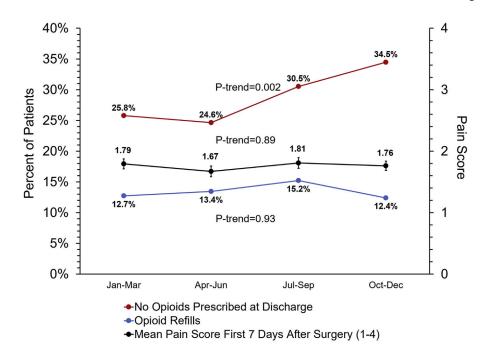


Figure 4.

Percent of patients without an opioid prescription at discharge, refill rate, and patient-reported pain scores.

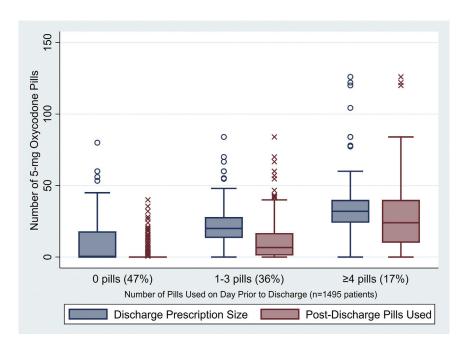


Figure 5.

Opioid prescription size (blue) and patient-reported postdischarge opioid use (red) stratified by the amount of opioid use during the inpatient day before discharge (0, 1-3, or 4 pills).

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Table 1.

Patient Characteristics

	(ceraii (n = 1493))	Prerecommendation (January to June) ($n = 729$)	T OSH COMPUTERINGHOUT (AULY W DECEMPTED) ($\Pi = 100$)	6) P Value
Age y	64 ± 11	63 ± 11	64 ± 11	.07
Female	382 (26)	173 (24)	209 (27)	.12
Body mass index, kg/m ²	29.7 ± 5.7	29.8 ± 5.7	29.6 ± 5.7	.54
Race				
Caucasian	1296 (87)	630 (86)	666 (87)	.84
Black	103 (7)	54 (7)	49 (6)	
Non-Caucasian, non-Black	47(3)	23 (3)	24 (3)	
Not stated	49(3)	22 (3)	27 (4)	
Diabetes	555 (37)	256 (35)	299 (39)	.11
Hypertension	1246 (83)	591 (81)	655 (86)	.021
Cerebrovascular disease	310 (21)	137 (19)	173 (23)	.07
Peripheral vascular disease	157 (11)	79 (11)	78 (10)	69.
Congestive heart failure	525 (35)	270 (37)	255 (34)	.15
Cancer within 5 y	95 (6)	54(7)	41 (5)	11.
Depression	277 (19)	128 (18)	149 (20)	.33
Smoking status				
Never smoker	613 (41)	307 (42)	306 (40)	.61
Former smoker	625 (42)	303 (42)	322 (42)	
Current smoker	256 (17)	119 (16)	137 (18)	
Procedure				
Isolated CABG	822 (55)	371 (51)	451 (59)	.005
Isolated valve	446 (30)	228 (31)	218 (28)	
CABG/valve	181 (12)	101 (14)	80 (10)	
Other	46(3)	29 (4)	17 (2)	
Operative acuity				
Elective	892 (60)	456 (63)	436 (57)	.024
Urgent	561 (38)	249 (34)	312 (41)	
Emergent	42(3)	24 (3)	18 (2)	

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Variable	Overall (N = 1495)	Prerecommendation (January to June) $(n = 729)$	Overall (N = 1495) Prerecommendation (January to June) (n = 729) Postrecommendation (July to December) (n = 766) P Value (N = 1495) P VALUE (N =	P Value
Postoperative length of stay, d	6.7 ± 3.0	6.9 ± 3.0	6.5 ± 3.0	.027
Postoperative complication	591 (40)	302 (41)	289 (38)	.14
Inpatient opioid use on day before discharge, in 5-mg oxycodone pills	0.7 (0–2.4)	0.7 (0–2.7)	0.6 (0–2.0)	.023

Values are mean \pm SD, n (%), or median (interquartile range).

CABG, coronary artery bypass grafting.

Table 2.

Surgery
Cardiac
After
Use
Opioid
of Postdischarge
nt Predictors of F
Independent

$\begin{split} & Ag^4 \\ & (6) \ (n = 510) & (201) & (14 \pm 20) & (201) & (201) \\ & (6) \ (n = 48) & (10 - 10) & (10 - 10) & (14 \pm 20) & (201) & (201) \\ & (6) \ (n = 48) & (10 - 10) & (10 - 10) & (14 \pm 2) & (201) & (201) \\ & (2) \ (n = 120) & (10 - 12) & (10 - 12) & (21 + 1) & (21 + 1) & (21 + 1) \\ & (2) \ (n = 120) & (10 - 12) & (10 - 12) & (21 + 1) &$	Age^{a} < 60 y (n = 510)				
η (α = 510) $5 (\alpha^{-1})$ $5 (\alpha^{-1})$ $1 (\alpha^{-1} + \alpha^{-1})$ $(-\alpha^{0})$ $9 (\alpha = 48)$ $1 (\alpha^{-1} + \alpha^{-1})$ $1 (\alpha^{-1} + \alpha^{-1})$ $(-\alpha^{-1} + \alpha^{-1})$ $(-\alpha^{-1} + \alpha^{-1})$ $\gamma (\alpha = 48)$ $\alpha = 48$ $\alpha (\alpha^{-1} + \alpha^{-1})$ $(-\alpha^{-1} + \alpha^{-1})$ $(-\alpha^{-1} + \alpha^{-1})$ $(-\alpha^{-1} + \alpha^{-1})$ $\gamma (\alpha = 48)$ $\alpha (\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $\gamma (\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $\alpha^{-1} - \alpha^{-1} - \alpha^{-1}$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $\alpha^{-1} - \alpha^{-1} - \alpha^{-1}$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $\alpha^{-1} - \alpha^{-1} - \alpha^{-1} - \alpha^{-1} - \alpha^{-1} - \alpha^{-1}$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $(-\alpha^{-1} - \alpha^{-1})$ $\alpha^{-1} - \alpha^{-1} - $	<60 y (n = 510)				
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$		5 (0–21)	14 ± 20	<.001	<.001
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	60–69 y (n = 487)	1 (0–14)	9 ± 15		
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	70 y (n = 498)	0 (0-4)	4 ± 9		
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	Cerebrovascular disease				
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	Yes $(n = 286)$	0 (0–12)	9 ± 15	0.23	.013
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	No (n = 1120)	1 (0–13)	9 ± 16		
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	Cancer within the past 5 y				
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	Yes $(n = 91)$	0 (0–10)	7 ± 13	.24	.049
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	No $(n = 1308)$	1 (0–13)	9 ± 16		
$= 761$) $0 (0-9)$ 7 ± 13 Reference $= 15$) $0 (0-0)$ 2 ± 7 08 $= 15$) $0 (0-0)$ 2 ± 7 08 $12 (0-37)$ 19 ± 21 <001 $\operatorname{intal}(\operatorname{n}=35)$ $0 (0-10)$ 9 ± 18 <01 $\operatorname{intal}(\operatorname{n}=35)$ $0 (0-10)$ 9 ± 18 <01 atr $4 (0-17)$ 11 ± 16 001 atr 10 ± 15 001 0 atr 0 2 ± 5 <01 atr 0 2 ± 5 <01 atr $24 (10-40)$ 28 ± 23 <001 atr 0 2 ± 5 <01 atr 0 2 ± 7 <01 atr 0 2 ± 7 <01 atr $0 (0-3)$ 24 ± 9 <01 atr 10 ± 15 10 ± 15 <01	Primary insurance payor ^b				
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Medicare (reference) $(n = 761)$	(6-0) 0	7 ± 13	Re	ference
12 (0-37) 19 ± 21 < 001 0 (0-10) 9 ± 18 $.51$ 4 (0-17) 11 ± 16 $.001$ 3 (0-14) 10 ± 15 $.001$ 7 (1-17) 10 ± 15 $.09$ 7 (1-17) 11 ± 13 $.091$ 24 (10-40) 28 ± 23 $< .001$ 0 2 ± 7 $.601$ 0 2 ± 7 $.601$ 3 (0-14) 10 ± 15 $.001$ 14 (2-28) 19 ± 22 $< .001$	Self-pay/no insurance $(n = 15)$	0 (0-0)	2 ± 7	.08	.011
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Medicaid $(n = 113)$	12 (0–37)	19 ± 21	<.001	.50
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Military/other governmental $(n = 35)$	0 (0–10)	9 ± 18	.51	89.
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	Commercial $(n = 437)$	4 (0–17)	11 ± 16	.001	.48
$ \begin{array}{cccccc} 0 & & 2\pm5 & <001 \\ 7 (1-17) & & 11\pm13 & \\ 24 (10-40) & & 28\pm23 & \\ 0 & & & 2\pm7 & & \\ 0 & & & & & & \\ 0 & & & & & & & \\ 0 & & & &$	Health maintenance organization $(n = 134)$	3 (0–14)	10 ± 15	60.	.61
$\begin{array}{cccccccc} 0 & & & & & & & & & & & & & & & & & & $	Number of pills taken on day before discharge a				
7 (1-17)11 \pm 1324 (10-40) 28 ± 23 24 (10-40) 28 ± 23 0 2 ± 7 8 (0-3) 2 ± 7 10 (0-3) 4 ± 9 3 (0-14) 10 ± 15 10 (12-28) 19 ± 22 < 001	0 (n = 710)	0	2 ± 5	<.001	<.001
$\begin{array}{cccc} 24 \ (10 - 40) & 28 \pm 23 \\ & & \\ 0 & & \\ 0 & & \\ 0 & & \\ 0 & & \\ 0 & & \\ 0 & & \\ 0 & & \\ 0 & & \\ 1 & & \\ 10 & $	1-3 (n = 536)	7 (1–17)	11 ± 13		
0 2 ± 7 Reference 0 (0-3) 4 ± 9 .001 3 (0-14) 10 ± 15 <.001	4 (n = 249)	24 (10-40)	28 ± 23		
0 2 ± 7 Reference 0 (0-3) 4 ± 9 .001 3 (0-14) 10 ± 15 <001	Pain during the first week after surgery b				
$0 (0-3)$ 4 ± 9 .001 $3 (0-14)$ 10 ± 15 <.001	None $(n = 80)$	0	2 ± 7	Re	ference
$3 (0-14)$ 10 ± 15 <001 $14 (2-28)$ 19 ± 22 <.001	Minimal $(n = 418)$	0 (0–3)	4 ± 9	.001	.001
14 (2–28) 19 ± 22 <:001	Moderate $(n = 683)$	3 (0–14)	10 ± 15	<.001	<.001
	Severe $(n = 238)$	14 (2–28)	19 ± 22	<.001	<.001

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Variable	Pills Taken After Discharge	Pills Taken After Discharge Pills Taken After Discharge Univariate P Value Multivariable P Value	Univariate P Value	Multivariable P Value
3–5 d (n = 615)	3 (0–17)	11 ± 17	<.001	.004
6–7 d (n = 487)	1 (0–12)	9 ± 15		
8 d (n = 393)	0 (0–0)	7 ± 14		

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 $^{a}\mathrm{Analyzed}$ as a continuous variable in multivariable regression

 $b_{\mbox{Analyzed}}$ as a categorical variable in multivariable regression.

Values are median (interquartile range) or mean \pm SD. The full model is shown in in Supplemental Table 2.