

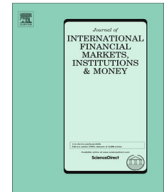


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## The effects of a “black swan” event (COVID-19) on herding behavior in cryptocurrency markets

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### ABSTRACT

This paper analyses herding in cryptocurrency markets in the time of the COVID-19 pandemic. We employ a combination of quantitative methods to hourly prices of the four most traded cryptocurrency markets - USD, EUR, JPY and KRW - for the period from 1st January 2019 to 13th March 2020. While there are several strong theoretical reasons to observe the “black swan” effect on cryptocurrency herding, our results suggest that COVID-19 does not amplify herding in cryptocurrency markets. In all markets studied, herding remains contingent on up or down markets days, but does not get stronger during the COVID-19. These results are important for cryptocurrency investors and regulators to enhance their understanding of cryptocurrency markets and the financial effects of the COVID-19 pandemic.

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## 1. Introduction

The spread of the COVID-19 pandemic generated strong contagion effect across financial markets around the globe, while the scale of its social and economic consequences is still hard to estimate and predict. There are ongoing debates regarding the nature of the continuing crisis with some experts comparing it to the Global Financial Crisis of 2008, and others likening it to war events, terror attacks, natural disasters and past epidemics. The COVID-19 has resulted in unprecedented measures to stop the spread of the virus, such as international and local travel restrictions, lockdowns and quarantines that have caused immediate and long-term damage to a vast majority of industries, and businesses of different sizes. Some may refer to this crisis as a “black swan” event, given that it was hard to predict and has never previously occurred. This makes a precise prediction of its impact rather challenging for all existing risk management models.

In the context of cryptocurrency markets, which are relatively new and unexplored financial assets, the COVID-19 proved to be an unprecedented shock. Barely a decade old, the Bitcoin has traditionally experienced periods of high volatility without being susceptible to any major systematic crisis. Cryptocurrencies as financial assets have not yet demonstrated their safe haven properties during any major economic crisis and recession, and early evidence suggests that Bitcoin failed to display hedging opportunities and flight to safety properties during the COVID-19 pandemic (Conlon and McGee, 2020; Corbet et al., 2020d). Given this finding, we assume that the COVID-19 can have a black swan effect on cryptocurrency, resulting in behavioral anomalies such as investor herding. To better understand the role of investor sentiment and panic in driving herding behaviour in the light of the COVID-19 pandemic, we study both unconditional herding, as well as that conditional on up and down market days (to capture investor optimism and pessimism).

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Herding behavior is a particularly interesting direction of research during the crisis periods, when investors may share similar fears and be susceptible to large-scale financial panic. However, till date, there is limited evidence available on herding behavior during the COVID-19 pandemic. [Espinosa-Mendez and Arias \(2020\)](#) analyze the impact of COVID-19 on herding in European equity markets and provide strong evidence of herding behavior because of the pandemic. [Chang et al \(2020\)](#) suggest that the increase in herding in energy markets during the COVID-19 pandemic can be explained by extremely low oil prices. The paper closest to ours in terms of contribution is the very recent [Philippas et al. \(2020\)](#) that provides a comprehensive analysis of herding behavior in cryptocurrency markets. However, our paper differs from them in that we focus specifically on herding behavior during the COVID-19 pandemic, that provides us with the unprecedented opportunity to add to this line of literature, providing novel evidence of the black swan effects on herding.

Therefore, in this paper, we aim to specifically answer the question: Does the COVID-19 pandemic amplify herding behavior in cryptocurrency markets? For this, we use hourly close prices for the main cryptocurrencies traded in USD, Euro, JPY and KRW, for the period starting 00 a.m. on 1st January 2019 till 8:00 p.m. of 13th March 2020. We analyze unconditional herding and herding conditional on up/down market days by means of the [Chang et al. \(2000\)](#) approach. The parameters are estimated using several methods.

First, we apply the Newey-West ([Newey and West, 1987](#)) Heteroscedasticity and Autocorrelation consistent (HAC) estimators to linear regressions using Bartlett kernel weights as described in [Newey and West \(1987, 1994\)](#). Applying these methods to estimate covariance matrices in regression analysis allows us to account for conditional heteroskedasticity of regression disturbances that may be of an unknown form. Statistical inference that rests on standard errors not robust to heteroskedasticity can be strongly misleading. Our choice of using estimators from the variance-covariance matrix addresses this issue.

To verify our results, we estimate a robust linear Bayesian model with priors estimated as in [Lewandowski et al. \(2009\)](#) and Markov-Switching regressions using the EM (Expectation-Maximization) algorithm proposed by [Dempster et al. \(1977\)](#), [Hamilton \(1989a, 1998b, 1994\)](#), and [Goldfeld and Quandt \(2005\)](#) to check for the presence of herding given different regimes. Markov-switching models are widely applied in literature starting with [Hamilton \(1989a, 1998b\)](#) and [Kim \(1994\)](#); and further developed by [Kim et al., 1998](#); [Guidolin and Timmermann, 2007](#), [Guidolin, 2009, 2011](#); [Alexander and Kaeck, 2008](#); [Hahn et al., 2010](#); [Liu, 2011](#); [Ang and Timmermann, 2011](#); etc. [Guidolin \(2012\)](#) shows that these type of models effectively capture volatility clustering, excess kurtosis, and heavy tails. The regimes allow for distinguishing Regime 1 (given the higher absolute value of the coefficient on  $x_2$ ), Regime 2 is more persistent in terms of the probability of switching to another regime.

Quantile regression ([Sim and Zhou, 2015](#)) is applied to test the behavior of the coefficients across quantiles. This approach allows us to address non-linearity in the relationship as well as to estimate the effects of the quantiles of one variable on those of another. This class of models provides more detailed results across different part of the distribution than the standard quantile regression ([Matkovskyy and Jalan, 2021](#)).

Time-Varying Regressions, TVR ([Bollerslev et al., 2016](#); [Casas et al., 2018](#)) are estimated to assess evolution of herding over time. These type of models, first introduced by [Robinson \(1989\)](#) for stationary processes were further generalized to nonstationary processes and correlated errors by [Chang and Martinez-Chombo \(2003a, 2003b\)](#), [Cai \(2007a, 2007b\)](#), [Corsi \(2009\)](#), [Chen et al. \(2017\)](#) etc. The TVR framework allows for a “natural” way of parameter estimation for the herding model. The time-varying coefficients are obtained by applying the local polynomial kernel estimator (e.g., Nadaraya-Watson estimator). This preserves the bias, variance, and automatic good boundary behavior properties of the local linear estimator, ensuring flexibility and robustness ([Cai, 2001](#)). Compared with the local linear method, these estimators are consistent and asymptotically normal. Comparing the TVR to state-space models, we see that the latter represent complex and nonlinear models that come with increased complexity and harder calibration. Application of the Kalman filter can provide an initial state estimate and covariance that is inconsistent with the true system state. Also, potential “outliers” can cause negative outcomes, leading to a non-positive semi-definite covariance matrix after update.

To detect the unknown structural break points under heteroskedasticity, the [Mumtaz, Gulfam & Asad \(2017\)](#) test is applied. This test has an advantage over other tests such as the sup F test that is widely used for structural change and assumes homoskedasticity. It checks simultaneously for breaks in regression coefficients as well as variance and its results help us detect the effects of the COVID-19 on changes in herding levels in the selected markets.

And finally, time-varying correlation among selected markets in terms of herding is estimated by means of the time-varying parameter copula models (GAS models with conditional multivariate Student- $t$  distribution and time-varying correlations as in [Creal et al., 2011, 2013](#); [Harvey, 2013](#)). The motivation to use Student- $t$  distribution is that the time-series of herding coefficients is relatively short. GAS models can be estimated in a rather straightforward manner and offer several advantages such as allowing for time-varying parameters for a great variety of nonlinear models and their ability to exploit the complete density structure ([Matkovskyy, 2019](#)).

This paper contributes to the literature in two main ways. First, it contributes to the growing body of literature on the financial effects of the COVID-19 pandemic ([Conlon and McGee, 2020](#); [Goodell, 2020](#); [Corbet et al., 2020a, 2020d](#); [Sharif et al., 2020](#)). Second, it contributes to the literature on herding behavior in cryptocurrency markets, providing novel evidence of the black swan effect on herding ([Bouri et al., 2019](#); [Vidal-Tomás et al., 2019](#); [Philippas et al., 2020](#), among others). Our results indicate unconditional herding in all selected markets, except the cryptocurrency KRW market using hourly data. In terms of mean values, we observe conditional herding on both up/down market days in the USD and JPY cryptocurrency markets, and on only up-market days in the Euro market. In a nutshell, our results show that for the selected cryptocurrency

markets, herding behavior exists largely during stable times. In addition, our results provide supporting evidence of the asymmetry in herding on up and down-market days, suggesting panic-driven herding on days with high value-drops in the cryptocurrency market.

Quantile regression estimates indicate stronger herding in higher quantiles of return variation in the euro and USD cryptocurrency markets. For the JPY and KRW markets, it is quite the opposite – herding is absent in higher quantiles. In general, we observe a decreasing trend in herding in the recent times, particularly in the USD and euro cryptocurrency markets. This can be attributed to shocks in conventional expansionary policy and non-standard policy supporting the hypothesis of [Krokida et al. \(2020\)](#). Generally, COVID-19 does not cause a significant shock to herding in the cryptocurrency markets except the JPY crypto market for which a structural break can be observed. Time-varying correlation in herding is noted. Herding in the USD –JPY and Euro-JPY cryptocurrency markets is cyclical, peaking twice a month. For the pair USD –Euro cryptocurrency markets, correlation increased during the first half of February 2020 and is rather constant over time.

The remainder of the paper is organized as follows. [Section 2](#) discusses the theoretical background of this paper. [Section 3](#) explains data and methodology. [Section 4](#) reports and discusses empirical results. [Section 5](#) concludes.

## 2. Background literature

Herding behavior among investors can explain some of the behavioral anomalies against the efficient market hypothesis. [Fama \(1970\)](#) suggests that in efficient markets, prices fully reflect all available information, making it impossible for an investor to generate abnormal returns using publicly available information. Many studies highlight the presence of herding behavior in financial markets, that tends to occur when some investors have access to private information and make investment decisions that are not in line with the general market trend. While it is hard to define what private information is and what it consists of, the actions of the investors in possession of such information can provide signals to other market participants, without any direct sharing or revelation. If market participants believe that others may be privy to useful private information, they may end up getting influenced by their decisions, consequently herding on their investment decisions, i.e., imitating the behavior of other investors. This can lead to deviation in prices away from fundamental values, resulting in high volatility and consequent destabilization of markets.

Early studies by [Morris and Shin \(1999\)](#), [Persaud \(2000\)](#), and [Shiller \(1990\)](#), to name but a few, find that herding and institutional risk management strategies may amplify volatility in financial markets. Herding behavior has been studied extensively for equity markets (e.g., [Christie and Huang, 1995](#); [Chang et al., 2000](#)). If investors herd, stock returns should correlate with market returns. This behavior, however, should be distinguished from “spurious herding”, where market participants facing similar information can make similar decisions. Evidence in favor of herding behavior in stock markets is still mixed and inconclusive, with many studies reporting the absence of herding ([Galariotis et al., 2016](#); [Lee, 2017](#)). Much less empirical evidence is available for herding in cryptocurrency markets (e.g., [Bouri et al., 2019](#); [Kaiser and Stöckl, 2019](#); [Kallinterakis and Wang, 2019](#)). The study of herding behavior in cryptocurrency markets is important given that cryptocurrencies have been in the limelight owing to their impressive historical returns and since their inception, have attracted the attention of many investors who never before participated in financial markets before their arrival.

The COVID-19 pandemic provides a unique opportunity to investigate herding behavior in cryptocurrency markets during this unprecedented “black swan” event. However, some may argue that for traditional financial markets, this event is not entirely “black swan”, given that there exist other historical events that have had similar impact on economies and markets in equity, commodity, and other financial derivatives. There are studies that have discussed the risk of pandemics and infectious diseases on the economy before the emergence of the COVID-19 crisis. For example, [Bloom et al. \(2018\)](#) discuss the economic risks of epidemics citing managerial and policy implications, while [Fan et al. \(2018\)](#) provide predictions of expected losses due to pandemics. Much earlier, [Saker et al. \(2004\)](#) discuss the impact of globalization on the spread of infectious diseases, highlighting that stronger economic ties between countries could affect the prevalence, spread, geographical range and control of many infections. Studies of previous epidemics, such as SARS, Ebola, Zika, and H1N1, or HIV/AIDS provide some empirical evidence on the impact of epidemics, associated risks and costs, and mitigation strategies ([Haacker, 2004](#); [Hoffman and Silverberg, 2018](#)). Furthermore, there is an emerging field of literature comparing the impact of COVID-19 to that of past pandemics ([Correia et al., 2020](#); [Eichenbaum et al., 2020](#); [Ma et al., 2020](#)), and the reaction of financial markets to its growth and spread ([Baker et al., 2020a, 2020b](#)).

While it might be expected to observe reduced risk-taking and flight-to-safety behavior among investors, early evidence from COVID-19 shows surprising trading patterns. [Ortmann et al. \(2020\)](#) report a significant increase in trading activities during the outbreak, where the average weekly trading intensity increases by 13.9% as the number of COVID-19 cases doubles. Furthermore, their results show an absence of flight-to-safety behavior among investors or a tendency to invest in more speculative assets, such as cryptocurrencies. [Chiah and Zhong \(2020\)](#) also document a surge in trading volumes in 37 equity markets behaviour, where trading activity increased the most in wealthier nations, and among markets with better corporate governance and legal systems. [Heo et al. \(2020\)](#) analyse risk tolerance during the pandemic and report the existence of two clusters of investors – the first with lower levels of both financial knowledge and risk tolerance, and the second comprising those with both higher financial knowledge and risk tolerance.

Bitcoin remains the cryptocurrency market leader and other cryptocurrencies often mimic its behavior ([Corbet et al., 2020b, 2020c](#)). In cryptocurrency literature, Bitcoin has often been compared to gold, though evidence of the safe haven

properties of this digital asset remains mixed (Corbet et al., 2019).<sup>1</sup> Specifically, in analyses of the COVID-19 crisis, Goodell and Goutte (2020), Le et al. (2020) suggest that Bitcoin can be considered as a safe haven asset during the first four months of the pandemic. These results contradict those of Corbet et al. (2020d) and Conlon and McGee (2020) who claim that Bitcoin did not act as a safe haven or diversifier during the early stage of the pandemic. Furthermore, Conlon et al. (2020) report that Bitcoin and other cryptocurrencies such as Ethereum and Tether, have failed to demonstrate safe haven properties for international equity markets during the same period. In contrast, Mariana et al. (2020) claim that cryptocurrencies are short-term safe havens during the first months of the COVID-19 pandemic, with Ethereum acting as a better safe haven than the Bitcoin.

While hedging and safe haven properties of cryptocurrencies remains a dominant theme of the early COVID-19 literature in finance (e.g., Corbet et al., 2020a, Conlon and McGee, 2020), substantial empirical evidence has also been collected to suggest financial contagion and spillover effect between various financial assets (e.g., Akhtaruzzaman et al., 2020; Okorie and Lin, 2020; Yarovaya et al., 2020a, 2020b), reaction and recovery of financial markets from the COVID-19 shock (e.g., Ashraf, 2020; Seven and Yilmaz, 2020; Heyden and Heyden, 2020; Mazur et al., 2020; Topcu and Gulal, 2020; Yarovaya et al., 2020c); predictability (Ciner, 2020); hedge funds performance (e.g. Yarovaya et al., 2021), among others. However, only a few papers have analyzed herding behavior during the COVID-19 pandemic (Aziz et al., 2020; Espinosa-Mendez and Arias, 2020; Chang et al., 2020), and to the best of our knowledge, our paper is the first to analyze herding in cryptocurrency markets during the COVID-19.

### 3. Theory development

Cryptocurrencies in general, and Bitcoin in particular, have attracted a huge amount of attention from investors (e.g., Urquhart, 2018; Philippas et al., 2019) due to its innovative Blockchain technology and the unprecedented opportunity to generate abnormal returns. For equity markets, it is evident that attention-based trading strategies are not always able to outperform well-diversified portfolios (Barber and Odean, 2008). However, evidence from cryptocurrency literature suggests that even small allocations to Bitcoin could substantially improve portfolio returns (Platanakis and Urquhart, 2019). Matkovskyy et al., 2021 show that the top 10 cryptocurrencies can enhance portfolio returns of the 10 worst-performing stocks in the S&P600, S&P400 and S&P100 indexes, to match those of the 10 best-performing stocks therein. Prices in cryptocurrency markets are also sensitive to macroeconomic news and FOMC announcements, though cryptocurrency investors may not always correctly interpret this news, resulting in differences between responses of cryptocurrency and traditional financial markets (Corbet et al., 2020b, 2020c).

While all major theories in economics and finance assume that investors are rational, fully informed, and that their decisions are based on all publicly available information, empirical evidence seems to suggest that investors often behave irrationally, thereby creating noise in financial markets with their decisions (Shleifer and Summers, 1990). This could be especially pronounced for new and immature cryptocurrency markets, and more so in times of increased uncertainty (Kahneman and Tversky, 1973), such as that created by the COVID-19 crisis.

Sharif et al. (2020) analyze the impact of the COVID-19 pandemic on the Economic Policy Uncertainty (EPU) index, oil prices, and the US stock market and find that the COVID-19 had the most pronounced impact on EPU, increasing uncertainty to unprecedented levels. Knowing that rationality is bounded to the extent of available information and cognitive abilities of the individual (Simon, 1997), we can assume that limited information on the COVID-19 virus, and limited understanding of its effects, coupled with potentially low computing capacity to estimate its impact using standard forecasting models, the behavior of Bitcoin investors during the period from January 2020 to March 2020 could be irrational. This forms the key motivation of our paper.

To examine irrationality and consequent herding behavior in cryptocurrency markets following the uncertainty induced by the COVID-19, we specify and test the following hypotheses:

**Hypothesis 1.** There is herding in cryptocurrency markets during the COVID-19 pandemic.

Furthermore, taking into account that cryptocurrency markets are highly volatile, it is important to investigate whether herding is impacted by investor optimism and pessimism, i.e., by up and down markets days. This leads to our second hypothesis:

**Hypothesis 2.** This herding behavior is contingent upon up and down-market days.

A confirmation of Hypothesis 1 will provide evidence of herding in cryptocurrency markets during the early months of COVID-19 crisis, while that of Hypothesis 2 will highlight the relationship between herding and investor sentiment, i.e., optimism and pessimism following market periods with rising and declining prices.

It is worth mentioning here that although in most cases herding behavior is considered irrational, herding could be rational too. Rational herding occurs when investors simply mimic each other's decisions, even when doing so contradicts their own beliefs, expectations and interpretation of available information (Bikhchandani et al., 1992), resulting in an information cascade. Thus, even if Bitcoin investors believe that the cryptocurrency market will remain relatively unaffected by the COVID-19 shock, they could mimic the behavior of larger investors, or follow expert opinion suggesting continuing

<sup>1</sup> Please see Corbet et al. (2019) for a systematic review of cryptocurrency literature.



consequences of the pandemic on cryptocurrency prices. This could prompt them to choose to cash out, rather than remain invested. This strategy could be considered as rational herding during the early months of the COVID-19 pandemic.

There are several approaches to examine herding in the cryptocurrency market during the COVID-19 pandemic. Two primary models of investor herding behavior have been employed in recent publications. The first approach, the [Lakonishok et al. \(1992\)](#) (LSV), accounts for the number of transactions by investors with respect to a specific security (e.g., [Lakonishok et al., 1992](#); [Grinblatt et al., 1995](#); [Wermers, 1999](#); [Wylie, 2005](#)). These studies empirically document the existence of herding in the stock market, with more evidence of herding in buying stocks. The second popular approach is the cross-sectional standard dispersion (CSSD) model of [Christie and Huang \(1995\)](#) and its improved version, the CSAD model of [Chang et al. \(2000\)](#). This approach can be considered as the benchmark in the literature on stock market herding ([Fang et al., 2017](#)).

Herding in traditional financial markets across countries show mixed and inconclusive results. Using both daily and monthly returns, [Christie and Huang \(1995\)](#) analyse market participants in the U.S equity market during periods of market stress and document inconsistency in herding during periods of large price movements. Using intraday NYSE stock data during 1998–2000, [Patterson and Sharma \(2007\)](#) provide evidence of some level of herding. [Chang et al. \(2000\)](#) use cross-sectional absolute deviation and document significant evidence of herding for South Korea and Taiwan, partial evidence for Japan, and no evidence for US and Hong Kong. [Hwang and Salmon \(2004\)](#) find significant movements and persistence of herding in US and South Korea. [Litimi et al. \(2016a, 2016b\)](#) analyse American companies listed on NYSE/AMEX/NASDAQ from 1985 to 2013 and document the presence of herding behaviour in the U.S stock market. They also show that while market volatility decreases with increases in herding behaviour, the latter contributes to different financial crises and bubbles. [Choe et al. \(1999\)](#) document herding in the Korean financial market.

For other financial assets, results point towards the presence of herding behaviour. For instance, [Galarotis et al. \(2016\)](#) document herding in the European financial market, particularly in bond trading during the crisis. [Demirer et al. \(2015\)](#) find significant evidence of herding behavior in grains only during the high volatility state. [De Souza Raimundo Júnior et al. \(in press\)](#) document a high degree of herding in the commodities market. [Bernales et al. \(2020\)](#) analyse equity option contracts traded in the US between 1996 and 2012 and report herding behavior during periods of market stress.

In cryptocurrency markets, research on herding in is its infancy, with only a few papers available till date. [Bouri et al. \(2019\)](#) employ the [Chang et al. \(2000\)](#) approach and find evidence of insignificant herding or statistically significant anti-herding behaviour in the static model. Mild herding activity is observed in the second half of 2016, which the authors attribute to an increase in economic uncertainty. Applying cross-sectional absolute standard deviations (CSAD) and cross-sectional standard deviation of returns (CSSD) to cryptocurrency markets from January 2015 to February 2017, [Vidal-Tomás et al. \(2019\)](#) document herding. Using cross-sectional absolute deviation (CSAD) and state-space models over the longer period January 2015 to March 2019, [Kaiser and Stöckl \(2019\)](#) also confirm the existence of herding behavior in the cryptocurrency market.

A more extensive analysis of herding in cryptocurrency markets is undertaken by [Kallinterakis and Wang \(2019\)](#) who use daily prices, market capitalization and volume of the top 296 cryptocurrencies for the 12/2013–07/2018 window. Results indicate significant herding in the cryptocurrency market (even without the Bitcoin), which is stronger during up-market, low volatility and high-volume days. [Da Gama Silva et al. \(2019\)](#) apply cross-sectional absolute deviation (CSAD) and cross-sectional standard deviation (CSSD) tests to daily data for the 50 most-liquid and capitalized currencies from March 2015 to November 2018 and find that herding behavior is present mainly in down market days.

[Philippas et al. \(2020\)](#) examine how informative signals from exogenous factors contribute to herding intensity in the cryptocurrency market. They use the following main groups of indicators that generate information signals: (i) benchmark market-based indices; (ii) risk (volatility) indicators as expectation of risk attitude (the volatility index, VIX, the treasury yields volatility index, TYVIX, the volatility risk premium, VP, as a proxy for market sentiments); (iii) uncertainty indicators (the Economic Policy Uncertainty, EPU, index, the global equity markets' and global foreign exchange markets' connectedness measures); (iv) media attention indicators to capture information demand and supply and its cumulative sentimental influence (the Google Trends daily and the daily volume of Twitter hashtag 'btc'); and (v) commodities (the returns on gold and crude oil due to their safe haven characteristics). They document herding behavior of cryptocurrency investors and attribute it to the fact that contrary to equities or fixed income securities, cryptocurrency prices are more likely to be influenced by market sentiment owing to lack of a fundamental basis. They also show that higher the Bitcoin returns, higher the motivation for investors to be independent from the market. They also find that the major cryptocurrencies do not herd with the minor ones.

Although the period considered by [Philippas et al. \(2020\)](#) from January 2016 to May 2018-end includes some periods of increased volatility in Bitcoin market, it cannot directly be compared to the COVID-19 crisis, given the unprecedented levels of panic and uncertainty that investors are currently exposed to. Furthermore, previous literature on herding during crises is ambiguous, for example, [Chiang and Zheng \(2010\)](#) find that in most cases, there are no differences in herding coefficients during crisis and tranquil periods, except for the US and Latin America. [Economou et al. \(2011\)](#) and [Mobarek et al. \(2014\)](#) document that herding behavior is more prominent during crisis periods. Therefore, there is need for further research on herding in cryptocurrency markets during the crisis, and the COVID-19 pandemic offers a unique opportunity to investigate this question and contribute to this growing stream of literature.

#### 4. Data and methodology

We collect hourly observations of close prices for the main cryptocurrencies traded in USD, Euro, JPY and KRW, which are the top 4 currencies by trading volume, for the period from 00 a.m. 1st January 2019 to 8:00 p.m. of 13th March 2020. Since the primary objective of the paper is to study herding during the COVID-19 pandemic and given that its first wave is not long enough in terms of adequacy of observations, we face the issue of selecting an appropriate sampling frequency between the noisier higher-frequency data and the much less informative, low frequency data.

To address this issue, we apply a volatility signature to test for overall volatility for different frequencies of Bitcoin close prices. This is in line with Andersen and Bollerslev (1998), who argue that for the purpose of calculating realized volatility of asset returns, the ideal frequency is that which minimises both microstructural bias and sampling error. This is made possible with the help of a volatility signature plot that represents average realized volatility against various sampling intervals, with the ideal frequency being the one at which one observes a relative stabilization of overall volatility.

Since their introduction, volatility signature plots have been used widely to address data frequency issues (see for instance, Corsi et al., 2008, Degiannakis and Floros, 2013 and for cryptocurrencies, Akyildirim et al., 2020; Jalan et al., 2020).

Our estimated volatility signature using one-minute data shows that the variance is stabilized at 1-hour intervals. Particularly:

- during the first 10 min, bitcoin volatility decreases from 0.54 to 0.053 (about 90%);
- between 11 and 60 min, it decreases to 0.015 (70% decrease);
- between 61 and 180 min, volatility drops to 0.004 (20% decrease), but the variance itself is not significant.

On this basis, we conclude that in terms of data frequency, hourly intervals are the best-suited for this study.

Table 1 presents information on cryptocurrencies considered and data sources, i.e., exchanges used to collect data across the four markets.

Figs. A1, A5, A9, A13 display price dynamics for selected cryptocurrencies during the observation period. Here one can see that post-COVID, the worst 24-hour drop in cryptocurrencies is observed on 12.03.2020 (average drop in value – 38%) (see the graphs in Appendix). Compared to its maximum value in 2020, the Bitcoin suffers a value drop of about 53%, while Ethereum and Litecoin lose 61 and 64% of their highest values, respectively. This highlights the general impact of the COVID-19 on the otherwise resilient cryptocurrency market.

We estimate herding behavior by means of the Chang et al. (2000) approach. Despite the availability of alternative models (Bohl et al., 2013; Lee, 2017; Clements et al., 2017), we choose this approach given its wide use in prior literature to be able to ensure comparability of our results with those of prior studies.

Thus, the following specification for estimates is used:

$$CSAD_{m,t} = \beta_0 + \beta_1 |R_{m,t}| + \beta_2 R_{m,t}^2 + e_t \quad (1)$$

where  $R_{m,t}$  is the average absolute market return of all actively traded selected cryptocurrencies in each currency market, i.e., cryptocurrency USD, euro, JPY and KRW at time  $t$ ,  $CSAD_{m,t}$  is the Cross-Sectional Absolute Deviation of returns and is calculated as follows:

$$CSAD_{m,t} = \frac{\sum_{i=1}^n |R_{i,t} - R_{m,t}|}{n} \quad (2)$$

where  $R_{i,t}$  is the first logarithmic difference of closing prices for cryptocurrency  $i$  at time  $t$

$$R_{i,t} = \ln P_t - \ln P_{t-1} \quad (3)$$

If herding is not present in a market, the relationship between the cross-sectional return dispersion,  $CSAD_{m,t}$ , and absolute market returns,  $|R_{m,t}|$ , would be expected to be positive and linear, implying that  $\beta_1$  would be expected to be significantly positive, while  $\beta_2$  insignificant. On the contrary, in the presence of herding, when values of  $|R_{m,t}|$  are high and thus substantial market movements are observed, the relationship between  $CSAD_{m,t}$  and  $|R_{m,t}|$  would be non-linear, implying that  $\beta_2$

**Table 1**  
Data used in the study.

	USD cryptocurrency market	Euro cryptocurrency market	JPY cryptocurrency market	KRW cryptocurrency market
<b>Exchanges</b>	Binance, Bitbay, BitFinex, Bitstamp, Bittrex, Cexio, CoinBase, Gemini, Kraken, Poloniex	Bitbay, Bitstamp, Exmo, Kraken, CoinBase	OKOIN, Zaif, Bitflyer, Kraken	Bithumb
<b>Cryptocurrencies</b>	BTC, LTC, ETH	BTC, LTC, ETH	BTC, BCH, LTC, ETH, MONA, XEM, ZAIF	BTC, ETH, LTC, BTG, XMR, XRP

Note: Data source: <http://www.cryptodatadownload.com/>.

would be negative and significant. Thus, herding lowers cross-sectional dispersion of returns compared to the case of rational pricing. Bernales et al., 2020 postulate that herding is stronger when  $\beta_1$  is negative, implying a negative relationship between the cross-sectional deviation of the cryptocurrency's return and the magnitude of respective market returns.

To assess herding on up/down market days, Eq. (1) is extended following Cui et al. (2019), to:

$$CSAD_{m,t} = \beta_0 + \beta_1 D^{up} |R_{m,t}| + \beta_2 (1 - D^{up}) |R_{m,t}| + \beta_3 D^{up} R_{m,t}^2 + \beta_4 (1 - D^{up}) R_{m,t}^2 + e_t \tag{4}$$

where  $D^{up}$  is equal to one (zero) on days with positive (negative) values of  $R_{m,t}$ . Significantly negative values of  $\beta_3$  ( $\beta_4$ ) would indicate the presence of herding on days of positive (negative) average cryptocurrency market performance.

These parameters are estimated using several methods. First, we apply the classic Newey-West (Newey & West, 1987) Heteroscedasticity and Autocorrelation consistent (HAC) estimators to linear regressions using Bartlett kernel weights as described in Newey and West (1987, 1994). To corroborate these estimates, we estimate a robust linear Bayesian model with priors estimated as in Lewandowski et al. (2009). We also estimate Markov-Switching regressions using the EM algorithm as in Hamilton (1989a, 1998b, 1994), Goldfeld and Quantd (2005) to check for the presence of herding given different regimes.

Quantile regression (Sim and Zhou, 2015) is applied to test the behaviour of the coefficients across quantiles. Eqs. (3) and (4) are modified as follows:

$$CSAD_{m,t} = \beta_0^\theta + \beta_1^\theta |R_{m,t}| + \beta_2^\theta R_{m,t}^2 + e_t^\theta \tag{5}$$

$$CSAD_{m,t} = \beta_0^\theta + \beta_1^\theta D^{up} |R_{m,t}| + \beta_2^\theta (1 - D^{up}) |R_{m,t}| + \beta_3^\theta D^{up} R_{m,t}^2 + \beta_4^\theta (1 - D^{up}) R_{m,t}^2 + e_t^\theta \tag{6}$$

where  $\theta$  is the  $\theta$ th quantile of the conditional distribution of the average absolute market return of all actively traded selected cryptocurrencies per currency market,  $e_t^\theta$  is the error term with a zero  $\theta$ -quantile.

Then, Time-Varying Regressions, TVR (Bollerslev et al., 2016; Casas et al., 2018) are estimated to assess evolution of herding over time. Given that a classical linear model can be expressed as  $y_t = x_t^T \beta + u_t$ , where  $t = 1, \dots, T$ ,  $y_t$  is a dependent variable  $CSAD_{m,t}$ ,  $x_t = (x_{1t}, x_{2t}, \dots, x_{dt})^T$  is a vector of regressors at time  $t$ ,  $\beta = (\beta_0, \beta_1, \dots, \beta_d)^T$  is a vector of coefficients and  $u_t$  is the error term. If the coefficients are allowed to vary over time, the time-varying coefficient model (TV-LM) can be specified as follows:

$$y_t = x_t^T \beta(z_t) + u_t, t = 1, \dots, T \tag{7}$$

where  $z_t$  is the smoothing variable, transforming coefficients to be a function of  $z_t$ :  $\beta(z_t) = (\beta_0(z_t), \beta_1(z_t), \dots, \beta_d(z_t))^T$ .  $z_t$  can be defined in two ways. First, as an unknown function of time,  $\beta(z_t) = f(\tau)$ , as proposed in Robinson (1989), and further developed by Cai (2007a, 2007b) and Chen et al. (2017). Second, this variable can be defined as an unknown function of a random variable,  $\beta(z_t) = f(z_t)$ , developed by Hastie and Tibshirani (1993), Cai et al. (2000); Chang and Martinez-Chombo (2003a, 2003b), Cail et al. (2009), and Gao and Phillips (2013). The estimation is done by combining OLS and the local polynomial kernel estimator (Fan and Gijbels, 1996). Given that  $\beta(\cdot)$  is twice differentiable, an approximation of  $\beta(z_t)$  can be expressed by means of the Taylor rule,  $\beta(z_t) = \beta(z) + \beta'(z)^{(1)}(z_t - z)$ , where  $\beta^{(1)}(z) = d\beta(z)/dz$  is the first derivative. The following minimization problem should then be solved:

$$\arg \min_{\theta_0, \theta_1} \sum_{t=1}^T [y_t - x_t^T \theta_0 - (z_t - z) x_t^T \theta_1]^2 K_b(z_t - z) \tag{8}$$

This approach can be fit to a set of weighted local regressions with an optimally chosen window size, defined by the bandwidth  $b$ . Using the weights derived from the kernel  $K_b(z_t - z) = b^{-1} K(\frac{z_t - z}{b})$ , yields the following local estimator:

$$\begin{pmatrix} \hat{\beta}_t \\ \hat{\beta}_t^{(1)} \end{pmatrix} = \begin{pmatrix} S_{T,0}(z_t) & S_{T,1}^T(z_t) \\ S_{T,1}(z_t) & S_{T,2}(z_t) \end{pmatrix}^{-1} \begin{pmatrix} T_{T,0}(z_t) \\ T_{T,1}(z_t) \end{pmatrix} \tag{9}$$

where,

$$S_{T,s}(z_t) = \frac{1}{T} \sum_{i=1}^T X_i X_i (z_i - z_t)^s K\left(\frac{z_i - z_t}{h}\right) \tag{10}$$

$$T_{T,s}(z_t) = \frac{1}{T} \sum_{i=1}^T X_i' (z_i - z_t)^s K\left(\frac{z_i - z_t}{h}\right) y_i \tag{11}$$

To detect the unknown structural break points under heteroskedasticity, the Mumtaz, Gulfam & Asad (2017) test is applied:

$$MZ = (T - k) \log \hat{\sigma}_0^2 - \left\{ (T_1 - k) \log \hat{\sigma}_1^2 + (T_2 - k) \log \hat{\sigma}_2^2 \right\} \tag{12}$$



$$\sup MZ = \max_{a \leq j \leq b} MZ_j, k < a \leq j \leq b < T - k \tag{13}$$

where there are two subgroups with  $T_1$  and  $T_2$  observations respectively,  $k$  is the number of parameters  $(\beta, \sigma^2)$ ,  $\hat{\sigma}_i^2 = Y_i - \frac{X_i \hat{\beta}_i}{T_i - k}$ ,  $i = 1, 2$ , indicating periods before and after the structural break, respectively, while  $a$  and  $b$  are time periods before and after the breakpoint that occurs at  $t = j$ .

Time-varying correlation among the selected markets in terms of herding is estimated by means of the time-varying parameter copula models, i.e., GAS models with conditional multivariate Student-t distribution and time-varying correlations (Creal et al., 2011, 2013; Harvey, 2013). These models allow for time-varying parameters in copulas and thus help exploit the complete density structure of the data, rather than merely means and higher moments. The time-varying coefficients from time-varying regressions are used as inputs. We therefore estimate how these parameters are correlated across markets.

These parameters are updated over time by applying the scaled score of the likelihood function. The evolution in the time-varying parameter vector  $\theta_t$  can be specified as follows (see Creal et al., 2011, 2013 for technical details):

$$\theta_{t+1} = \kappa + As_t + B\theta_t \tag{14}$$

where  $\kappa, A$  and  $B$  are matrices that contains coefficients, particularly,  $\kappa$  and  $A$  control for the level and the persistence of the mean reverting process for  $\theta_t$ ,  $s_t$  is a vector proportional to the score of  $y_t | y_{1:t-1} p(y_t; \theta_t)$ , where  $y_{1:t-1} \equiv (y_1, \dots, y_{t-1})$  and  $\theta_t \in \Theta \subseteq \mathfrak{R}^J$  and is defined as

$$s_t \equiv S_t(\theta_t) \nabla_t(y_t, \theta_t) \tag{15}$$

where  $S_t$  is a  $J \times J$  positive defined scaling matrix known at time  $t$ ,  $\nabla_t(y_t, \theta_t)$  is the score of  $y_t | y_{1:t-1} p(y_t; \theta_t)$  estimated as  $\nabla_t(y_t, \theta_t) \equiv \frac{\partial \log p(y_t, \theta_t)}{\partial \theta_t}$ .

The updated equation for  $\theta_t$  is specified as:

$$\theta_t \equiv \Lambda(\theta_t) \tag{16}$$

$$\theta_t \equiv \kappa + As_t + B\theta_{t-1} \tag{17}$$

where  $s_t \equiv S_t(\theta_t) \nabla_t(y_t, \theta_t)$ . Jacobian matrix estimated at  $\theta_t$  is  $\mathcal{J}(\theta_t) \equiv \frac{\partial \Lambda(\theta_t)}{\partial \theta_t}$ . Finally, the matrices  $\kappa, A$  and  $B$  are used to estimate  $\xi$ , by means of maximum likelihood (ML) approach:

$$\xi \equiv \arg \max_{\xi} \mathcal{L}(\xi; y_{1:T}) \tag{18}$$

where  $\mathcal{L}(\xi; y_{1:T}) \equiv \log p(y_1; \theta_1) + \sum_{t=2}^T \log p(y_t; \theta_t)$ ,  $\theta_1 \equiv (I - B)^{-1} \kappa$ , and  $\theta_t \equiv \theta(y_{1:t-1}, \xi)$ .

## 5. Empirical results and discussion

### 5.1. Helicopter view

The results for unconditional herding behavior across the four markets are presented in Table 2 below. Here one observes positive and significant (at 1%) coefficients for  $\beta_1$  across all markets. These are not directly interpreted to assess herding behavior. However, negative and significant  $\beta_2$  values indicate strong herding behavior. Coefficients on  $\beta_2$  are negative

**Table 2**  
Herding behavior estimates.

	(Intercept)	$\beta_1$	$\beta_2$
<b>Cryptocurrency</b> USD market (std.error, t-stat)	0.001*** (0.0005, 25.5881)	0.2527*** (0.011, 22.0579)	-0.891*** (0.130, -6.8484)
<b>Cryptocurrency</b> JPY market (std.error, t-stat)	0.0033*** (0.0001, 32.7129)	0.502*** (0.025, 19.82)	-1.703*** (0.389, -4.3784)
<b>Cryptocurrency</b> Euro market (std.error, t-stat)	0.0014 *** (0.000044, 32.495)	0.23*** (0.01, 19.383)	-0.35*** (0.155, -2.272)
<b>Cryptocurrency</b> KRW market (std.error, t-stat)	0.002*** (0.000, 23.519)	0.287*** (0.038, 7.620)	2.223(.) (1.279, 1.738)

Note: 1. The Newey-West Heteroscedasticity and Autocorrelation consistent (HAC) estimators are provided for linear regressions; Signif. codes: 0 '\*\*\*' 0.001 '\*\*\*' 0.01 '\*\*\*' 0.05 '.' 0.1 ' . ' 1.

2. We also estimated using aggregated daily data, but the results are statistically insignificant.

3. The models were also estimated by means of the robust linear Bayesian model with the priors estimated as in Lewandowski et al. (2009). The coefficients have the same sign and magnitude as those reported above.

and significant at 1% for all markets except for KRW, which is positive with a significance level of 95%. This seems to indicate the presence of unconditional herding for all markets except the KRW, providing supporting evidence for [Hypothesis 1](#). While our results support the finding of [Ballis and Drakos \(2019\)](#), [Bouri et al. \(2019\)](#), [Vidal-Tomás et al. \(2019\)](#), among others, who report the presence of herding in cryptocurrency markets, they contradict those of [Stavroyiannis and Babalos \(2019\)](#), who document lack of herding. Herding behavior in cryptocurrency markets can be explained by the dominating role of irrational individual investors in the most traded cryptocurrencies ([Kaiser and Stöckl, 2019](#)).

As [Bernales et al., 2020](#) note, herding would be even stronger if the relationship between the cross-sectional deviation of asset returns and the magnitude of market returns would be negative. This would require  $\beta_1$  to be negative. However,  $\beta_1$  is found to be positive in our study across various specifications, which implies that strong herding as defined by [Bernales et al., 2020](#), is not detected in our sample.

Detection of the unknown structural break points under Heteroskedasticity ([Mumtaz et al., 2017](#)) reveals only one day for each of the USD, JPY and KRW cryptocurrency markets. The break is detected on 2019-07-15 at 01:00:00 for the cryptocurrency USD. Similarly, for the JPY cryptocurrency market, the structural break occurs on 2019-04-21 at 01:00:00 and on 2019-01-29 at 04:00:00 for the cryptocurrency KRW market. A careful look at the dates for structural breaks above suggests that they lie before the COVID-19 struck and hence, the pandemic did not cause any significant change in herding levels in these markets.

For the cryptocurrency euro market, a break in hedging behavior is detected on 2020-03-09 at 10:00:00. This makes the cryptocurrency euro market the only one in our sample that experienced a significant change in herding patterns due to the COVID-19 pandemic. This result is in line with the general finding in literature that herding increases with the level of uncertainty, in addition to being consistent with papers that have analyzed herding in financial markets during the previous crisis shocks (e.g. [Bernales et al., 2020](#); [Demirer et al., 2015](#); [Galarotis et al., 2016](#)). Furthermore, Markov-Switching regressions for two regimes also suggest statistically significant herding behavior in the selected cryptocurrency markets (see Appendix).

For the cryptocurrency USD market, one notes that while herding is stronger in Regime 1 (given the higher absolute value of the coefficient on  $x_2$ ), Regime 2 is more persistent in terms of the probability of switching to another regime (88% versus 61% for Regime 1). For the cryptocurrency euro market, Regime 2 is not only more persistent (just like the USD), but unlike the latter, also depicts stronger herding behavior. The cryptocurrency JPY market shows patterns similar to the cryptocurrency euro market except that herding in Regime 2 is found to be significantly higher than that in Regime 1 based on the size of the coefficient on  $x_2$ . For the cryptocurrency KRW markets, even when Regime 2 is more persistent, herding behavior therein is not significant. For the USD, JPY and KRW markets, even when herding behavior is significantly different across regimes, we can make no inference with respect to the impact of COVID-19 thereon. Studies by [Balcilar and Demirer \(2015\)](#) and [Bouri et al. \(2019\)](#) report that herding in cryptocurrency markets tends to occur when uncertainty increases. This, however, is not evident in our results for all cryptocurrency markets analyzed, and there is no evidence of increase in herding that can be attributed to the COVID-19 uncertainty.

To better understand the dynamics of herding, we employ quantile regressions to examine the impact of various quantiles of return variation on those of herding behavior. Results for unconditional herding demonstrate significantly higher levels of herding behavior at higher quantiles of return variation for the USD and Euro cryptocurrency markets.

Conditional herding on the other hand, is highly responsive to higher levels of return variation and increases drastically on low-market days for the USD and Euro cryptocurrency market. This seems to suggest panic-driven herding on days with high value-drops in the cryptocurrency market. Previous studies, for example, by [Vidal-Tomás et al. \(2019\)](#), document herding only during down-market days, while [Kallinterakis and Wang \(2019\)](#) find that herding is stronger during up-market days. [Ballis and Drakos \(2019\)](#) further demonstrate that the up-events market dispersion follows market movements at a faster pace compared to the down events. For up-market days, however, herding does increase with higher quantiles of return variation, but remains much less sensitive compared to down-market days. This difference in behavior seems to suggest that herding behavior is most likely a response to market panic. [Gurdgiev and O'Loughlin \(2020\)](#) discuss herding from the perspective of fear and uncertainty, and document that bullishness/bearishness of markets has an asymmetric impact on cryptocurrency prices. In that sense, our results provide additional evidence in support of the asymmetry in herding on up and down-market days. For the JPY and KRW cryptocurrency markets, one observes that herding increases significantly at higher quantiles of return variation, though at very high levels of variation in return (94th percentile and above for JPY and roughly the 70th percentile for KRW), herding behavior stops. This lack of herding in very high quantiles for the JPY and KRW Bitcoin is observed for both up and down-market states.

[Fig. 1](#) below displays unconditional herding behavior using time-varying coefficients for the period 01/01/2019-13/03/2020 using hourly data.

The objective is to identify trends in herding during and the pre-COVID period. The USD market shows high volatility in herding behaviour during the period studied. Very high herding is observed at the beginning of 2019, which gradually dissipates in April-May, only to increase again starting July 2019. Herding remains absent between November 6th-December 8th, 2019, after which it increases again and can be observed till the end of the sample period. However, declining herding levels are noted after middle of February 2020 on an average.

An examination of cryptocurrency price behavior in the USD for the same period suggests an enormous increase in bitcoin prices from \$5265 to \$8900 (increase of 67%) for the first time in May 2019. Prices continue to rise, attaining their maximum of \$18,175 on 10 July 2019, after which a decline sets in. During July-August 2019, bitcoin was being traded in the range of \$9,462.50 - \$11,815.04, when its biggest drop is observed on October 24, taking the price down to \$7421.20. In merely three

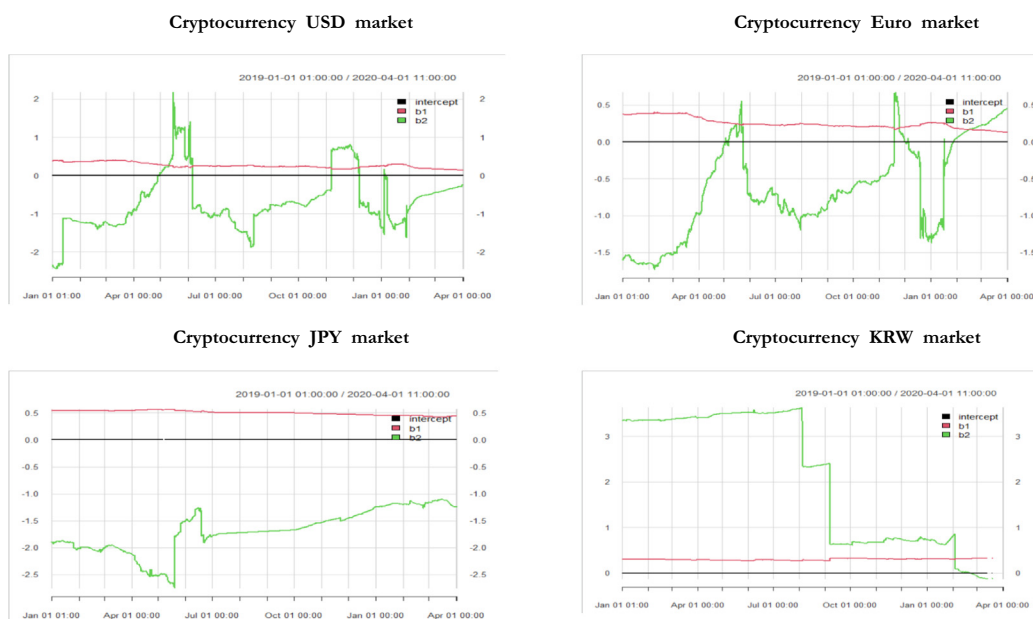


Fig. 1. Unconditional Herding in cryptocurrency markets, time-varying regression.

days however, Bitcoin prices bounce back to attain both the month and quarter high at \$9595.34. Bitcoin remained relatively stable during the first few days of November 2019 with prices tumbling again from \$9,396.19 to \$8,771.30 from November 4 to November 9, reaching \$7,026.83 on November 24. From mid-November to mid-December, the cryptocurrency market remained stressed with issues experienced by several exchanges that eventually led to wiping out the October rally. Correlating herding behavior in cryptocurrency USD market with volatility swings, one can infer that periods of high price volatility generally resulted in a decrease in herding levels in the market. In other words, the market exhibited herding behavior largely during stable times, implying general rationality in market participants' behavior. Therefore, for the cryptocurrency USD market, while our findings contradict prior studies that attribute herding in cryptocurrency markets to a large number of irrational investors and high explosivity of these markets (e.g. Kaiser and Stöckl, 2019; Kallinterakis and Wang, 2019), they support those of Stavroyiannis and Babalos (2019).

Similar trends are observed in the cryptocurrency euro market, except that herding seems to disappear after February 2020 and continues to decrease. Herding in the cryptocurrency JPY market is observed throughout the sample period but herding levels seem to decrease after July 2019. For the KRW cryptocurrency however, no significant herding is observed in 2019. In fact, small levels of herding are observed only around February 2020.

We further report estimates of herding behaviour conditional on up/down market days (see Table 3 below), and our results provide evidence of significant herding behaviour on up-market days in the USD, Euro and JPY cryptocurrency markets, but on down-market days, herding is observed for USD and JPY markets only, supporting the asymmetry in herding behaviour during bullish and bearish markets (Gurdgiev and O'Loughlin, 2020)

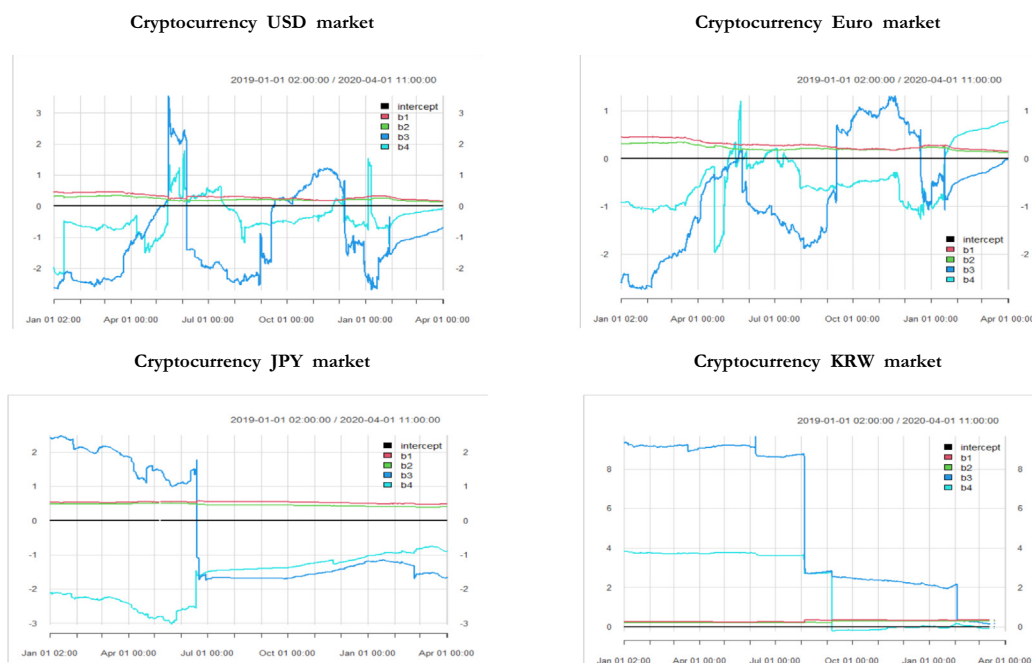
**Table 3**  
Conditional on up/down market days Herding behavior estimates.

	(Intercept)	$\beta_1$	$\beta_2$	$\beta_3$	$\beta_4$
<b>Cryptocurrency USD</b>	0.001***	0.300***	0.215***	-1.56*** (0.106,-14.658)	-0.499***
(std.error, t-stat)	(0.000, 25.058)	(0.014, 21.5)	(0.011, 18.837)		(0.079,-6.306)
<b>Cryptocurrency JPY</b>	0.003***	0.549***	0.452**	-1.774(.)	-1.388***
(std.error, t-stat)	(0.000, 33.614)	(0.031, 17.978)	(0.022, 20.430)	(1.044,-1.699)	(0.375,-3.701)
<b>Cryptocurrency Euro</b>	0.001***	0.271***	0.193***	-0.916***	0.101
(std.error, t-stat)	(0.000, 33.055)	(0.014, 19.433)	(0.012, 16.666)	(0.161,-5.678)	(0.198, 0.510)
<b>Cryptocurrency KRW</b>	0.002***	0.352***	0.217***	2.095	2.683*
(std.error, t-stat)	(0.000, 22.757)	(0.060, 5.896)	(0.039, 5.539)	(2.935, 0.714)	(1.264, 2.123)

Note: 1. The Newey-West Heteroscedasticity and Autocorrelation consistent (HAC) estimators are provided for linear regressions; Signif. codes: 0 '\*\*\*' 0.001 '\*\*\*' 0.01 '\*\*\*' 0.05 '.' 0.1 ' ' 1.

2. We also estimated the parameters by using aggregated daily data, but the results are statistical insignificant.

3. The models were also estimated by means of the robust linear Bayesian model with the priors estimated as in Lewandowski et al. (2009). The coefficients have the same sign and magnitude.



**Fig. 2.** Conditional Herding on up/down market days, cryptocurrency USD market (time-varying regression).

A significantly negative value of  $\beta_3$  ( $\beta_4$ ) can suggest the presence of herding on days of positive (negative) average performance for the cryptocurrency market. In the case of the Bitcoin USD market, both  $\beta_3$  and  $\beta_4$  are negative and statistically significant, implying herding on both positive and negative average market performance, though that on positive market performance is stronger (absolute value of  $\beta_3 >$  absolute value of  $\beta_4$ ). Applying a quantile regression to test the results, one can observe that absolute values of  $\beta_3$  are higher for all quantiles, implying that herding is more pervasive during positive average market performance. During February, maximum herding conditional on market performance is observed, after which it starts to decrease.

Fig. 2 below displays results for conditional herding using time-varying coefficients for the period from 01/01/2019 to 13/03/2020, using hourly data. A close analysis of the USD market in Fig. 2 suggests higher levels of herding on up-market days. Even when herding is observed for both up and down-market days, the magnitude of herding levels on up-market days remains significantly higher. The only exception to this observation is May 2019 when down-market herding seems to dominate and October-December 2019, when only down-market herding exists without any up-market herding. Both up and down-market herding are observed for the COVID-19 period starting January 2020, though up-market herding dominates. A similar trend is observed for the euro Bitcoin market except that starting February 2020, no significant down-market herding exists and even up-market herding shows a declining trend. The JPY cryptocurrency market shows consistent up and down-market herding from July 2019 before which only down-market herding seems to exist. Post July 2019, levels of both types of herding remain very close to each other till the beginning of March 2020, after which up-market herding begins to dominate. Interestingly, no significant up-market herding is noted for the KRW market throughout the sample period. Down-market herding becomes mildly significant only after September 2019 after which it remains stable through the end of the sample period.

## 5.2. A closer look at 2020

The objective is to analyse trends in conditional and unconditional herding during the uncertainty prevalent in the COVID-19 period, we plot patterns in herding behaviour in the four geographical markets exclusively for the year starting January 1, 2020. Fig. 3 presents results for unconditional herding. Here we can observe high volatility in unconditional herding for the USD cryptocurrency market. Herding levels increase significantly after roughly the end of the first week of January and thereafter continue to decline, remaining significant. Herding levels in the JPY remain rather stable across 2020 but increase slightly after mid-March. The cryptocurrency euro market opens with significant levels of herding in 2020, which begin to decline starting roughly mid-January. This decline continues until herding levels become insignificant in the start of February. In the KRW market, herding is observed after mid-February after which it continues to increase.

Fig. 4 presents results for herding conditional on up and down-market days, respectively. The USD cryptocurrency market opens with a significantly high level of up-market herding at the start of 2020, but levels continue to decline throughout, still

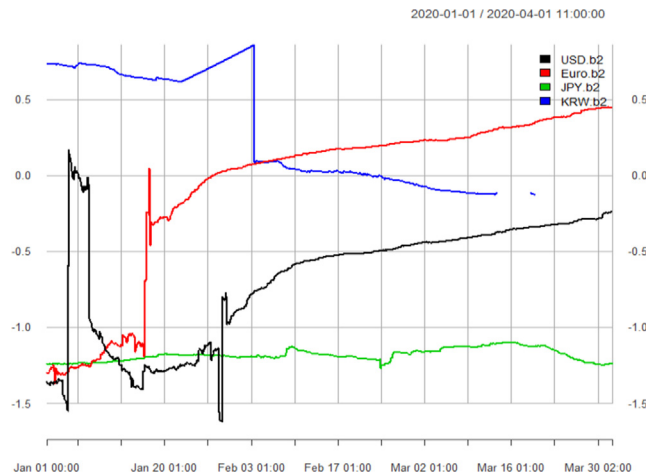


Fig. 3.  $\beta_2$  coefficients of unconditional herding in 2020 for the selected cryptocurrency markets.

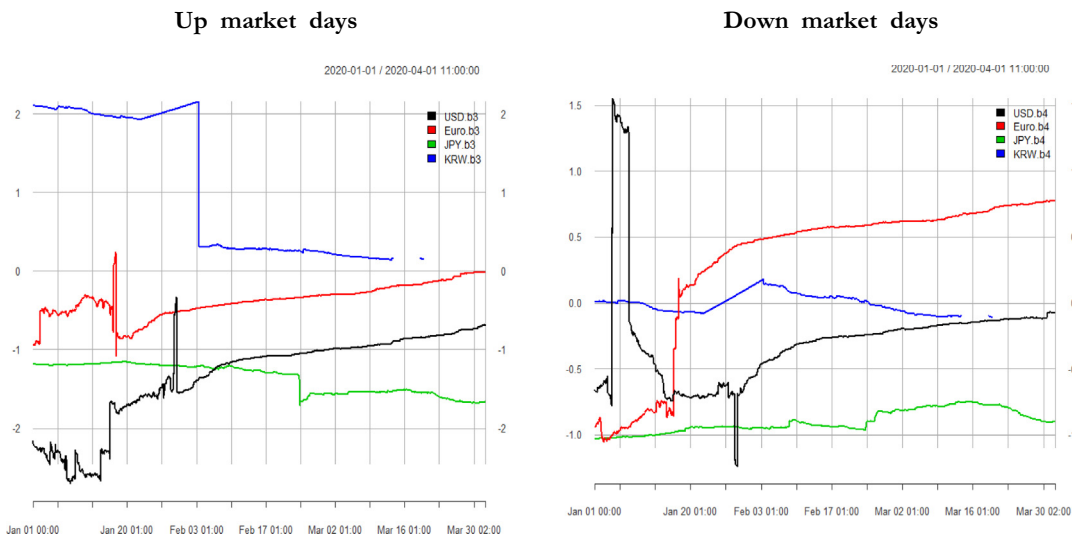


Fig. 4. Coefficients of Conditional herding on up and down-market days ( $\beta_3$ ) in 2020 for the selected cryptocurrency markets.

remaining significant. A similar pattern is observed for the euro cryptocurrency market except that at the end of the sample period, up-market herding levels become almost insignificant. The JPY cryptocurrency market shows consistently stable levels of up-market herding since the start of 2020, though levels tend to rise after roughly the third week of February 2020. The KRW cryptocurrency market shows virtual absence of any up-market herding in 2020.

The USD cryptocurrency market shows significant down-market herding starting roughly January 10 after which levels first increase and starting February, begin to decline. This trend continues till the end of the sample period though herding levels remain significant. For the euro cryptocurrency market, down-market herding continuously declines in 2020 until it turns insignificant in mid-January. Down-market herding for the JPY cryptocurrency market remains largely consistent and significant throughout 2020, though a small decline in levels is noted after end of February 2020. For the KRW, no significant down-market herding is observed before end of February 2020. Levels however, remain small.

To summarize, the USD cryptocurrency market exhibits both up and down-market herding in 2020, with levels falling throughout. The Euro cryptocurrency market shows declining levels of both up and down-market herding that end up reaching insignificant levels. The JPY cryptocurrency market shows rather stable levels of up and down-market herding in 2020 with the former increasing and the latter decreasing towards the end of February. Only small levels of down-market herding are observed for the KRW starting after February-end with no evidence of any up-market herding.

Fig. 5 shows the time evolution of the pair-wise correlation in unconditional herding across markets. Here one notices that the correlation in unconditional herding behaviour between the euro-USD is consistently high and increases even further starting end of January. This seems to suggest a rather symmetrical herding reaction to the COVID-19 in these two markets.



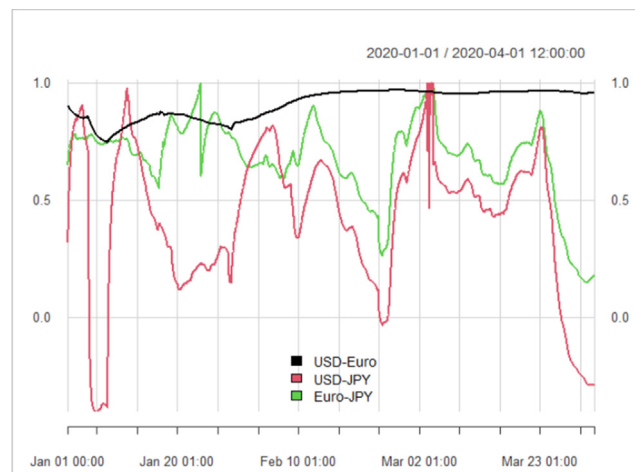


Fig. 5. TV correlation of the unconditional herding ( $\beta_2$ ), Time-varying copula model.

The pairs of the USD-JPY and euro-JPY cryptocurrency markets reveal a cyclical pattern of correlation, with alternate peaks and troughs throughout the sample period that demonstrate agreement in herding-related sentiments. The only disagreement is observed in the period which falls right after the introduction of the CME options on 14/01/2020. The USD-JPY correlation falls sharply starting March 23 and reaches a value of 0 in merely 4 days, after which it becomes negative.

## 6. Conclusion

This paper provides novel empirical evidence on cryptocurrency market herding during the COVID-19 pandemic. We analyse hourly closing prices for the four highest-traded cryptocurrencies in USD, EURO, JPY and KRW, that represent the highest trading volume for the period from 00 a.m. 1st January 2019 to 8:00 p.m. of 13th March 2020, using the well-known CSAD measure, estimated using several approaches to ensure robustness of our results.

While daily data shows insignificant herding, using hourly data we find significant evidence of unconditional herding in all selected markets, except the cryptocurrency KRW market. The results also indicate the existence of conditional herding on both up/down market days in the USD market. While herding is observed only for up-market days for the euro cryptocurrency market, for the JPY and KRW cryptocurrency markets, herding is observed only for down-market days. These findings are consistent with previous literature that suggests asymmetry in herding during up and down-market states (e.g. Phillipas et al., 2020; Gurdgiev and O'Loughlin, 2020; Kaiser and Stöckl, 2019; Kallinterakis and Wang, 2019; Vidal-Tomás et al., 2019). Particularly, we document the highest herding activity during the negative average performance in the KRW market. Application of quantile regressions reveals stronger herding in higher quantiles of return variation in the Euro and USD cryptocurrency markets, while absence of herding is noted for higher quantiles in the JPY and KRW markets.

While the COVID-19 pandemic increased volatility in cryptocurrency markets, we observe a decreasing trend in herding in the recent times, particularly in the USD and euro cryptocurrency market. This can be attributed to shocks in conventional expansionary policy and non-standard policy supporting the hypothesis of Krokida et al. (2020). Furthermore, these findings contradict the popular belief that herding is stronger during times of heightened uncertainty. Thus, we report that COVID-19 does not significantly amplify herding in the cryptocurrency markets. A notable exception is the euro cryptocurrency market for which we observe a structural break during the COVID-19 crisis. Finally, our results show that correlation in herding is time-varying. Specifically, herding in the USD and JPY cryptocurrency markets is cyclical, peaking twice a month, whereas for the USD and Euro cryptocurrency markets, correlation increased during the first half of February 2020.

The results of this study have important implications for policy makers, academics, and investors in cryptocurrency markets. The first implication arises in the form of a better understanding of the rather nascent crypto markets, which continue to be opaque in terms of investor composition. Hsin and Tseng (2012) suggest that herding propensity is likely to decrease when the number of informed investors relatively to uninformed investors in a market is large. In that sense, our results of plummeting herding levels despite the heightened uncertainty caused by the pandemic highlight the dominance of informed and probably institutional investors in the crypto market. Our results are in line with Feng et al. (2018) who find evidence of informed trading in the Bitcoin market prior to large events.

Second, our results highlight the importance of a careful analysis of liquidity in the cryptocurrency market that could potentially explain the lower propensity to herd. Liquidity of crypto markets is a well-researched area but it may be worthwhile to investigate the link between lack of liquidity in these markets to the peculiar herding patterns observed in this study. This thought is motivated from the results of Galariotis et al. (2016) who document evidence of herd behaviour only for highly liquid stocks in mature markets. A recent paper by Jalan et al., 2020 document a statistically significant positive

effect of the introduction of Bitcoin futures on USD Bitcoin spot market liquidity, though not large enough to stabilize the highly illiquid market.

Third, given that herding behaviour is an extremely complex psychological phenomenon influenced by several factors such as uncertainty (e.g. Lin, 2018) and fear (e.g. Economou et al., 2018) and the fact that herding has been held responsible for different financial bubbles and crises over time (for instance, Litimi et al., 2016a, 2016b), our results in the context of cryptocurrency markets can be useful for asset allocation, diversification and potential spill-over effect analysis for these markets. This can help not only for portfolio management using cryptos but also predict the impact of future crises and black swan events on the cryptocurrency market.

We would like to acknowledge that this paper is one of the first papers that examines herding behaviour during the COVID-19 pandemic and therefore suggest that the presented results be interpreted with caution taking into account the early stage of the pandemic and amount of data available to date. Despite this, the paper builds a strong foundation for further research in this area, which will be of immense benefit to regulators and policy makers as new events in the COVID story unfold.

## Appendix A. Supplementary material

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.intfin.2021.101321>.

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