

1 **Reduction in risk of death among patients admitted with COVID-19 between first and**
2 **second epidemic waves in New York City**

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21 Using clinical and demographic data from COVID-19 hospitalizations at a tertiary New York

22 City medical center, we show that a reduction in mortality during the second epidemic wave was

23 associated with decreased strain on healthcare resources.

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41 **Abstract**

42 Many regions have experienced successive epidemic waves of COVID-19 since the emergence
43 of SARS-CoV-2 with heterogeneous differences in mortality. Elucidating factors differentially
44 associated with mortality between epidemic waves may inform clinical and public health
45 strategies. We examined clinical and demographic data among patients admitted with COVID-19
46 during the first (March-June 2020) and second (December 2020-March 2021) epidemic waves at
47 an academic medical center in New York City. Hospitalized patients (N=4631) had lower
48 mortality during the second wave (14%) than the first (23%). Patients in the second wave had a
49 lower 30-day mortality (Hazard Ratio (HR) 0.52, 95% CI 0.44, 0.61) than those in the first wave.
50 The mortality decrease persisted after adjusting for confounders except for the volume of
51 COVID-19 admissions (HR 0.88, 95% CI 0.70, 1.11), a measure of health system strain. Several
52 demographic and clinical patient factors were associated with an increased risk of mortality
53 independent of wave.

54
55 **Introduction**

56 By March 15, 2022, Severe Acute Respiratory Syndrome Coronavirus 2 (SARS-CoV-2),
57 which causes Coronavirus Disease 2019 (COVID-19), has led to over 460 million confirmed
58 infections and over 6 million deaths worldwide (1). New York City (NYC) experienced one of
59 the earliest and largest local epidemics with a peak of over 16,000 daily hospitalizations and 700
60 daily deaths in April 2020 (2). COVID-19 infections in New York City declined and remained
61 relatively low from July through November 2020 averaging fewer than 60 hospitalizations per
62 day and 15 deaths per day during this time period (2). A second epidemic wave occurred from

63 December 2020 through March 2021 resulting in a peak of nearly 400 daily hospitalizations and
64 90 daily deaths by February 2021(2).

65 Data from the US and Europe showed significant heterogeneity in mortality rates
66 between the first and subsequent waves of COVID-19 (3,4). While many regions have reported
67 lower case-fatality rates (CFRs) in the second epidemic wave compared to the first, some
68 countries have demonstrated the reverse pattern (5–7). Explanations for the frequently-observed
69 mortality reduction over time include the development and use of effective therapies, seasonal
70 effects, viral variant effects, age, race, ethnicity, and co-morbidity differences, but these
71 hypotheses have been under-explored (8). In the U.S., race, and ethnicity have been strong
72 correlates of COVID-19 mortality and may play a role in observed differences between epidemic
73 waves (9). Among regions with a trend toward increased mortality in the second wave, proposed
74 explanations include increased pressure on the healthcare system as well as the emergence of
75 new viral variants (7). Previous studies reporting CFRs between epidemic periods were not able
76 to examine the impact of related demographic, health system, or environmental factors.

77 Here, we investigated whether mortality differed by epidemic wave; and whether
78 individual-level demographic (e.g., age, race, and ethnicity) and clinical factors, as well as
79 markers of health system burden, affected mortality among COVID-19 patients admitted to an
80 academic medical center and an affiliated community hospital in New York City.

81

82 **Methods**

83 *Data Sources*

84 The study was conducted at a large quaternary academic medical center and an affiliated
85 community hospital in New York, NY, USA. Patients included were admitted with a positive or

86 presumed positive SARS-CoV-2 RT-PCR test within two days of hospital presentation between
87 March 1, 2020 and March 31, 2021. Data were extracted and cleaned from the medical center
88 clinical data warehouse and electronic health record (EHR) as previously described (10–12).
89 Patient demographics, anthropometric measurements, SARS-CoV-2 RT-PCR cycle threshold
90 (Ct) value, level of respiratory support, Intensive Care Unit (ICU) admission status, historical
91 and current medications, and discharge status were collected. Study approvals were obtained
92 from the Columbia University Irving Medical Center Institutional Review Board (IRB), New
93 York, NY, USA. The requirement for obtaining written informed consent was waived by the
94 IRB.

95

96 *Variables Assessed*

97 We defined COVID-19 cases in our cohort by three epidemic periods; the interval where
98 cases increased, peaked, and decreased, were called waves. The first wave was defined from
99 March 1, 2020, to June 30, 2020; the inter-epidemic period from July 1, 2020, to November 30,
100 2020; and the second wave from December 1, 2020, to March 31, 2021. Sex, age, race, and
101 ethnicity were self-reported. Body mass index (BMI) was collected on admission and calculated
102 by dividing weight in kilograms by the square of height in meters. BMI was categorized using a
103 $\geq 30\text{kg/m}^2$ cut point for obese and $< 30\text{kg/m}^2$ as normal. Viral load assessments based on SARS-
104 CoV-2 RT-PCR Ct values were reported for cobas (Roche Molecular Systems, Inc., Branchburg,
105 NJ), and Xpert Xpress assays (Cepheid, Inc., Sunnyvale, CA), but not for the BioFire
106 Respiratory Panel assay (BioFire Diagnostics, Salt Lake City, UT). The ORF1ab gene was
107 targeted for the cobas assay and the N2 gene was targeted for the Xpert Xpress assays.
108 Quantitative Ct values were converted to high, medium, and low viral load categories based on

109 tertiles. For the cobas assay, high, medium, and low viral load was defined by Ct values <25, 25-
110 30, and >30, respectively. For the Xpert Xpress assay, high, medium, and low viral load were
111 defined by Ct values <27, 27-32, and >32, respectively. Choice of viral load assay was based on
112 laboratory needs, resources, and timing. The level of respiratory support at hospital presentation
113 was recorded as: room air, nasal cannula, non-rebreather, non-invasive ventilation, or intubation.
114 We also recorded whether patients or their decision-makers elected for Do-Not-Intubate (DNI)
115 status. Patients admitted into an ICU within 24 hours of hospital admission were considered as
116 admitted to an ICU at presentation. Steroid usage was defined by documented receipt of
117 intravenous or oral formulations of prednisone, dexamethasone, or methylprednisolone. Use of
118 remdesivir, the first antiviral agent approved by the Food and Drug Administration (FDA) for
119 COVID-19, was also recorded (13). Underlying coronary artery disease (CAD), chronic kidney
120 disease (CKD), diabetes mellitus (DM), or hypertension (HTN) was defined by current or
121 historical International Classification of Disease (ICD-10) codes (Table S1). We calculated the
122 age-adjusted Charlson Comorbidity Index (CCI) using the EHR (14). Hospital Frailty Risk Score
123 (HFRS) was calculated among patients ≥ 75 years old (15). The weekly number of SARS-CoV-2
124 admissions were recorded as a proxy of the hospital COVID-19 burden. The primary outcome
125 was death or discharge to hospice. Survival time was calculated as days from the date of hospital
126 admission to the date of death or discharge to hospice for events and from admission to discharge
127 alive for the rest.

128

129 *Statistical Analyses*

130 Histogram plots were used to visualize the distribution of COVID-19 cases and
131 admissions. Descriptive statistics were reported including counts with percentages, medians and

132 their interquartile ranges (IQRs), and box-and-whisker plots. The Wilcoxon rank-sum test was
133 used to compare groups for continuous variables, and χ^2 test was used for categorical
134 variables. Unadjusted logistic regression analyses were used to estimate the associations between
135 epidemic wave and patient demographic, anthropometric, clinical, and viral load characteristics.
136 The wave was defined as a binary variable and the inter-epidemic period was excluded in
137 regression analyses.

138 Mortality was examined using Kaplan-Meier survival analysis and Cox Proportional
139 Hazards models. Deaths included those discharged to hospice and time to death was calculated
140 from the date of admission. Those who did not die were considered alive until March 31, 2021.
141 In Cox Proportional Hazards analyses, survival times were right censored on day 30 after
142 admission. Proportional hazards assumption was examined through graphical examination. Final
143 models focused on 30-day survival and investigated potential covariables in conjunction with the
144 epidemic wave. Models were also run separately for those ≥ 75 years old. All statistical analyses
145 were conducted using R Studio (Boston, MA, USA).

146

147 **Results**

148 Patients who met inclusion criteria (N=4631) were grouped by date of admission into the
149 first wave (March 1, 2020, to June 30, 2020; N=2846), an inter-epidemic period (July 1, 2020, to
150 November 30, 2020; N=366), and the second wave (December 1, 2020, to March 31, 2021;
151 N=1419).

152 The volume of SARS-CoV-2 cases and admissions (Figure 1A-B) vastly differed
153 between waves 1 and 2. The median length of hospitalization among patients who died or were
154 discharged to hospice (Figure 1C) was shorter in wave 1 than 2. The distribution of length of

155 hospital stay among patients who were discharged alive did not differ by epidemic wave (Figure
156 1D). Monthly mortality rate (per 100 inpatients) peaked at >25% in April 2020 during the first
157 wave, declined to 5-10% during the inter-epidemic period, and rose to 15% during the second
158 wave (Figure S1).

159 Table 1 shows patient characteristics by epidemic period and the unadjusted association
160 between each covariate and wave. Wave 2 patients were 0.56 (95% CI 0.47, 0.66) times as likely
161 to experience death, representing 23% of wave 1 patients and 14% of wave 2 patients. Patients in
162 the inter-epidemic period were 0.36 (95% CI 0.26, 0.52) times as likely to experience death as
163 those in wave 1. Age, history of co-morbidities, Charlson comorbidity index, and hospital frailty
164 risk score did not differ by wave. Patients during the second wave were less likely to identify as
165 male (51% versus 57%), identify as Non-Hispanic Black (11% versus 14%), to have DNI status
166 (25% versus 33%), and be admitted to the ICU at presentation (7% versus 11%). Wave 2 patients
167 were 1.31 (95% CI 1.02, 1.69) times as likely to have a low Ct value (high viral load). However,
168 only 23% of patients in the second wave had recorded Ct values compared with 97% during the
169 first wave due to the use of different testing assays. Patients during the second wave were less
170 likely to require supplemental oxygen (58% versus 65%), non-rebreather mask (5% versus 22%),
171 and invasive mechanical ventilation (4% versus 6%) at presentation. Patients in the second wave
172 were also more likely to receive supplemental oxygen via nasal cannula (48% versus 36%) and
173 non-invasive ventilation (2% versus 1%). Steroid and remdesivir use in wave 2 were 7.04 (95%
174 CI 6.12, 8.10) and 24.77-fold (95% CI 19.41, 31.60) higher than wave 1, respectively. Weekly
175 COVID-19 admissions divided by 50 were 0.54 times less in wave 2 compared to wave 1 (95%
176 CI 0.52, 0.57).

177 Figure 2A-D show Kaplan-Meier plots comparing survival between wave 1, inter-
178 epidemic period, and wave 2 (log-rank test, $p < 0.001$). For wave 1, the cumulative survival
179 probabilities declined from 0.87 on day 7 to 0.79 by day 30 after admission, worse than in wave
180 2 where these probabilities were 0.97 at day 7 and 0.88 by day 30 (Figure 2). Survival
181 probabilities were lower among patients ≥ 75 years old across both waves but the pattern of
182 improved survival in wave 2 persisted (Figure 2).

183 Unadjusted Cox regression for 30-day survival showed a 0.52-fold (95% CI 0.44, 0.61)
184 reduction in risk of death in wave 2 compared to wave 1 (Table 2). The lower risk of death
185 associated with wave 2 persisted after adjusting for potential demographic confounders. For
186 example, after adjusting for age, sex, and race individually, wave 2 was associated with 0.46
187 (95% CI 0.38, 0.54), 0.52 (0.44, 0.61), and 0.51-fold (0.43, 0.61) lower mortality rate,
188 respectively than wave 1. Oxygen level at presentation, a marker of the severity of disease,
189 attenuated the association between wave and mortality, although hazard ratios remained below 1
190 and statistically significant. After adjusting for the volume of weekly COVID-19 admissions, a
191 marker of health service strain, wave 2 was no longer associated with lower mortality (HR=0.88,
192 95% CI 0.70, 1.11).

193 Table 2 also shows increasing age, identifying as non-Hispanic white, lower Ct values,
194 DNI status, supplementary oxygen requirement at presentation, ICU admission at presentation,
195 and the volume of COVID-19 admissions were each associated with higher mortality rate after
196 adjusting for wave. Among patients ≥ 75 years of age, wave 2 had a 0.43-fold (95% CI 0.35,
197 0.53) reduced risk of death compared to wave 1 (Table S2), and the mortality pattern was similar
198 to that seen in the overall sample.

199 We expected steroid and remdesivir use to be associated with reduced mortality but
200 recognize that confounding by indication may produce results showing the opposite. Therefore,
201 we conducted stratified analyses by wave and ICU status (Table S3). These analyses suggested
202 that steroid and remdesivir effects were modified by wave i.e., lowered mortality risk in wave
203 one, ICU patients, and no benefit or slightly increased mortality risk in wave two. Particularly in
204 wave 1, there was a strong suggestion of confounding by indication for steroids.

205

206 **Discussion**

207 NYC, like many other parts of the world, has experienced multiple distinct epidemic
208 waves of COVID-19 (2–5,7). Our analysis of 4631 patients admitted with SARS-CoV-2 during
209 the first two epidemic waves and the inter-epidemic period revealed a decrease in risk of death or
210 discharge to hospice in wave 2 compared to the wave 1. The association between wave and
211 mortality persisted after covariate adjustment for several factors including age, sex, race, and
212 markers of disease severity. However, the association between wave and mortality disappeared
213 after adjusting for the volume of COVID-19 admissions suggesting that strain on hospital
214 resources may have been one of the factors accounting for the high mortality rate in epidemic
215 wave 1. Although the duration of hospital stay did not differ among patients who were
216 discharged alive, the median time to death in the second wave was nearly one week longer than
217 in wave 1.

218 There were several other variables correlated with decreased mortality. Patients
219 presenting in the second wave were less likely to require oxygen at presentation. Among those
220 who did require oxygen, patients in the second wave were more likely to require nasal cannula,
221 suggesting that their disease was less severe at the time of presentation. It is likely that during the

222 first wave, when hospitals in NYC were widely reported to be overwhelmed, patients may have
223 been more reluctant to present to the hospital until they developed a greater degree of respiratory
224 distress, resulting in a higher chance of intubation on arrival. We did not observe differences in
225 frailty among patients age ≥ 75 years old, individual co-morbidities or Charlson comorbidity
226 index over time.

227 Interventions may also account for differences in the mortality between the two waves.
228 Non-invasive ventilation was commonly avoided early in the pandemic due to concerns about
229 aerosolizing the virus as well as the theory that early intubation would lead to less risk of lung
230 injury (16,17). This approach was later questioned, and subsequent studies demonstrated no
231 benefit to early intubation for COVID-19, leading to increased use of non-invasive ventilation
232 during the second epidemic wave (16–18). We observed effect modification by epidemic wave,
233 and a paradoxical effect of COVID-19 therapies on mortality in wave two. Therapeutic use of
234 low-dose corticosteroid and remdesivir were associated with lower mortality in wave 1. We used
235 ICU status as a proxy for disease severity and it partially explained the association between
236 steroid and remdesivir use and increased mortality observed in wave 2. Early in the COVID-19
237 pandemic, corticosteroids were proposed as a potential intervention to counteract progression to
238 acute respiratory distress syndrome (ARDS). Still, their use was not routine in many centers in
239 part due to a lack of supportive data for corticosteroids in ARDS due to influenza (19). In July
240 2020, the RECOVERY group published data showing a reduction in mortality with the use of
241 dexamethasone among patients receiving supplemental oxygen, resulting in widespread adoption
242 of corticosteroid use for patients admitted with COVID-19 (20). We observed a much higher rate
243 of steroid use among patients in the second wave, which may have contributed to that group's
244 lower overall mortality rate, and associated with mortality due to the residual confounding effect

245 of use by disease severity that we could not measure or control for in this analysis. It is likely
246 that other changes in the clinical management of COVID-19 patients based on accumulated data
247 throughout the pandemic similarly contributed to the lower mortality. Remdesivir became a
248 widely used anti-viral for admitted patients requiring oxygen and the first anti-viral to be FDA-
249 approved for COVID-19. However, it was only shown to shorten time to recovery rather than
250 reduce mortality (13,21). Monoclonal antibody therapies also became available for patients early
251 in the course of infection with mild symptoms (22,23). Early proning of patients requiring
252 respiratory support was also associated with improved ventilation and outcomes (24,25). As data
253 were gathered throughout the early months of the pandemic, clinicians and hospitals rapidly
254 assembled and distributed COVID-19 management guidelines to delineate the most evidence-
255 based and proven interventions. These protocols undoubtedly led to better uniformity in clinical
256 practice, decreased use of unproven or ineffective therapies, and greater use of treatments with
257 the potential to reduce mortality.

258 In December 2020, the FDA granted emergency use authorization (EUA) to mRNA-
259 based COVID-19 vaccines developed by Pfizer/BioNTech and Moderna (26–29). Increasing
260 prevalence of vaccination during the second epidemic wave in NYC may have contributed to
261 decreases in COVID-19 admissions, especially among high-risk groups. Baseline patient
262 characteristics including age, sex, race/ethnicity, BMI, and the presence of several co-morbidities
263 were similar between the two epidemic waves, suggesting that the availability of vaccines did
264 not alter the overall demographics of patients admitted with COVID-19 through March 31, 2021.
265 We suspect that vaccination had a limited impact on mortality in the second wave since vaccine
266 uptake in the population at risk by March 31, 2021 was still highly limited.

267 Rapid increases in COVID-19 cases during epidemic waves put substantial pressure on
268 healthcare systems worldwide. During the first wave in NYC, many hospitals were overwhelmed
269 with the rapid influx of patients combined with staff and equipment shortages, including limited
270 ventilators, personal protective equipment (PPE), and certain essential medications. Studies have
271 shown the critical importance of adequate medical resources with COVID-19 mortality
272 inversely-correlated with available hospital beds and healthcare workers (30). In our analysis, we
273 see a significant association between COVID-19 mortality and the rate of COVID-19
274 admissions. This relationship may be explained by the strain placed on hospital resources with
275 increasing COVID-19 cases. We note that the second wave in NYC reached a lower peak
276 number of cases with a more even distribution of admissions over the same period (2). In our
277 analysis, over twice as many patients were admitted with COVID-19 during the four-month first
278 wave compared to the four-month second wave period. This result is in line with studies
279 associating efforts that flatten the curve of COVID-19 cases with reduced case fatality (31).

280 Lastly, differences in mortality by wave may be affected by evolution of SARS-CoV-2
281 and the prevalence of different viral genotypes. Wave 2 in NYC was primarily driven by
282 multiple variants of the ancestral SARS-CoV-2 lineage (32). Multiple subtypes of the Iota
283 (B.1.526) lineage were characterized in NYC during the second wave, with a high prevalence of
284 the E484K mutation, which is associated with resistance to therapeutic monoclonal antibodies as
285 well as convalescent and vaccinee sera (32). The Iota lineage was subsequently outpaced by the
286 Alpha (B.1.1.7) variant of concern in NYC, which several studies have associated with both
287 increased transmissibility and mortality compared to the ancestral virus (33,34).

288

289 *Limitations*

290 Our study has several limitations. Cases of COVID-19 included in this analysis are likely
291 to be undercounted from the first wave. All cases admitted with positive tests during this period
292 would be included. Still, testing capacity was limited at the time resulting in tests being
293 prioritized for patients with high clinical suspicion of COVID-19 or underlying comorbidities.
294 Detection of incidental COVID-19 likely increased in the second wave when routine testing was
295 widely available. Information bias in the EHR resulted in inadequate information to accurately
296 characterize patients with co-morbid conditions. RT-PCR Ct data were also missing in a
297 differential way that could have biased findings in either direction. The inclusion of patients
298 admitted to our institution may also not wholly reflect NYC-wide cases since some individuals
299 likely decided to avoid presentation to the hospital, especially during the first pandemic wave. In
300 addition, it is possible that pre-existing immunity had a differential impact on infections and
301 severe illness during the second wave. Our conclusions are limited to hospitalized patients.
302 Extrapolating to the general population can increase the likelihood of Berkson's bias in
303 identifying spurious correlations not present outside the hospital setting. Our estimates of
304 hospital capacity are based on COVID-19 admissions due to difficulties accurately estimating
305 total hospital admissions from our database. Patients admitted with COVID-19, however, utilize
306 specific hospital resources that would be expected to impact the care of other COVID-19
307 patients, including oxygen, ventilators, and ICU beds and staff. While the global population of
308 admitted patients may not be expected to utilize the same hospital resources to the same degree,
309 some conditions such as other respiratory viral infections, bacterial pneumonia, asthma, chronic
310 obstructive pulmonary disease, interstitial lung disease, and heart failure may be expected to
311 utilize similar resources and would not be accounted for in our analysis. These data would be
312 strengthened by estimating ICU bed capacity. We reduced selection bias in our sample and

313 model specification by right censoring patients after 30 days since the proportional odds
314 assumption did not hold. Patients observed for longer than 30 days were a small subset and
315 excluded from regression analyses.

316

317 *Public Health Implications*

318 We noted a distinct reduction in COVID-19 mortality between the first and second
319 epidemic waves in NYC associated with several covariates. The explanation for this reduction is
320 multifactorial and likely includes standardization of COVID-19 management, availability, and
321 knowledge of effective therapies, knowledge of ineffective treatments and interventions, as well
322 as reduced strain on critical healthcare resources. Public health interventions are also likely to be
323 critical contributors to the observed mortality differences given changes in lockdown policies,
324 mask guidance, social distancing behavior, availability and speed of SARS-CoV-2 testing, and
325 availability of vaccines for high-risk groups. A focus on the specific variables associated with
326 reduced and increased mortality in this analysis may help prepare for future potential epidemic
327 waves by improving the accuracy of COVID-19 projections, demographic impact, policy
328 decisions, and public health preparations. Furthermore, plans to address future potential
329 pandemics may benefit from prioritizing rapid, systematic methods of studying and developing
330 treatment standards and plans to rapidly adjust hospital capacity and scale up necessary
331 resources.

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461

462 **Figures and Tables**

463 Table 1: Demographic and clinical characteristics of patients hospitalized with COVID-19 in the first
464 wave, second wave, and inter-epidemic period.

	N	Overall N=4631	First Wave* N=2846	Inter- epidemic* N=366	Second wave* N=1419	OR (95% CI)**
Discharge status, N (%)	4631					
Alive		3738 (81%)	2191 (77%)	330 (90%)	1217 (86%)	Ref
Death/Hospice		893 (19%)	655 (23%)	36 (10%)	202 (14%)	0.56 (0.47, 0.66)
Status at 30 days, N (%)	4631					
Alive or discharged		3835 (83%)	2246 (79%)	338 (92%)	1251 (88%)	Ref
Death/Hospice		796 (17%)	600 (21%)	28 (8%)	168 (12%)	0.50 (0.42, 0.60)
Age (years), mean (SD)	4631	65.5 (17.4)	65.4 (17.1)	63.0 (18.5)	66.1 (17.6)	1.00 (1.00, 1.01)
Age (years), N (%)	4631					
18-50		869 (19%)	517 (18%)	87 (24%)	265 (19%)	Ref
50-65		1208 (26%)	776 (27%)	94 (26%)	338 (24%)	0.85 (0.70, 1.03)
65-75		1060 (23%)	662 (23%)	86 (23%)	312 (22%)	0.92 (0.75, 1.12)
75+		1494 (32%)	891 (31%)	99 (27%)	504 (36%)	1.10 (0.92, 1.33)
Sex, N (%)	4631					
Female		2101 (45%)	1235 (43%)	166 (45%)	700 (49%)	Ref
Male		2530 (55%)	1611 (57%)	200 (55%)	719 (51%)	0.79 (0.69, 0.89)
Race/Ethnicity, N (%)	4631					
Hispanic/Latino		2398 (52%)	1474 (52%)	197 (54%)	727 (51%)	Ref
Non-Hispanic Black		586 (13%)	400 (14%)	32 (9%)	154 (11%)	0.78 (0.63, 0.96)
Non-Hispanic White		536 (12%)	308 (11%)	48 (13%)	180 (13%)	1.18 (0.96, 1.45)
Other		1111 (24%)	664 (23%)	89 (24%)	358 (25%)	1.09 (0.93, 1.28)
BMI (kg/m ²), N (%)	4301					
<30		2766 (64%)	1666 (64%)	233 (67%)	867 (63%)	Ref
≥30		1535 (36%)	920 (36%)	113 (33%)	502 (37%)	1.05 (0.91, 1.20)
Ct value, median (IQR)	3324	28 (23, 33)	29 (24, 33)	26 (17, 33)	27 (22, 33)	0.99 (0.97, 1.00)
Viral load categories***, N (%)	3319					
Low (Ct >32 or 30)		1400 (42%)	1195 (43%)	77 (33%)	128 (39%)	Ref
Medium (Ct 27-32 or 25-30)		778 (23%)	667 (24%)	38 (16%)	73 (22%)	1.02 (0.75, 1.38)
High (Ct <27 or 25)		1141 (34%)	893 (32%)	120 (51%)	128 (39%)	1.34 (1.03, 1.74)
Ever DNI, N (%)	4631					
Yes		1344 (29%)	941 (33%)	53 (14%)	350 (25%)	0.66 (0.57, 0.76)
No		3287 (71%)	1905 (67%)	313 (86%)	1069 (75%)	Ref
Oxygen level at presentation, N (%)	4631					
Room Air		1816 (39%)	998 (35%)	219 (60%)	599 (42%)	Ref
Nasal Cannula		1835 (40%)	1036 (36%)	123 (34%)	676 (48%)	1.09 (0.94, 1.25)
Non-rebreather		699 (15%)	622 (22%)	9 (2%)	68 (5%)	0.18 (0.14, 0.24)
Non-invasive ventilation		50 (1.1%)	19 (1%)	7 (2%)	24 (2%)	2.10 (1.15, 3.92)
Intubation		231 (5.0%)	171 (6%)	8 (2%)	52 (4%)	0.51 (0.36, 0.70)
ICU admission by time, N (%)	4631					
Non-ICU		3769 (81%)	2262 (79%)	302 (83%)	1205 (85%)	Ref
ICU at presentation		437 (10%)	309 (11%)	34 (9%)	94 (7%)	0.57 (0.45, 0.72)
ICU after presentation		425 (9%)	275 (10%)	30 (8%)	120 (8%)	0.82 (0.65, 1.02)
Steroid use, N (%)	4501					
Yes		1913 (43%)	712 (26%)	191 (53%)	1010 (72%)	7.32 (6.34, 8.46)
No		2588 (57%)	2026 (74%)	169 (47%)	393 (28%)	Ref
Remdesivir use, N (%)	4631					
Yes		817 (18%)	82 (3%)	134 (37%)	601 (42%)	24.8 (19.5, 31.8)
No		3814 (82%)	2764 (97%)	232 (63%)	818 (58%)	Ref
History of Coronary Artery Disease, N (%)	4631	730 (16%)	433 (15%)	65 (18%)	232 (16%)	1.09 (0.91, 1.29)
History of Chronic Kidney Disease, N (%)	4631	798 (17%)	484 (17%)	64 (17%)	250 (18%)	1.04 (0.88, 1.23)
History of Diabetes, N (%)	4631	1824 (39%)	1139 (40%)	148 (40%)	537 (38%)	0.91 (0.80, 1.04)
History of Hypertension, N (%)	4631	2741 (59%)	1689 (59%)	217 (59%)	835 (59%)	0.98 (0.86, 1.12)
Age adjusted Charlson comorbidity score, median (IQR)	4582	4 (2, 5)	4 (2, 5)	3 (1, 5)	4 (2, 5)	1.02 (0.99, 1.04)
Age adjusted Charlson comorbidity	4582					

index, N (%)						
0-1		931 (20%)	546 (19%)	94 (26%)	291 (21%)	Ref
2-3		1318 (29%)	838 (30%)	93 (26%)	387 (27%)	0.87 (0.72, 1.04)
4-5		1358 (30%)	843 (30%)	107 (29%)	408 (29%)	0.91 (0.75, 1.09)
6+		975 (21%)	581 (21%)	70 (19%)	324 (23%)	1.05 (0.86, 1.27)
Hospital Frailty Risk Score among age \geq 75, mean (SD)	1476	6.3 (5.9)	6.4 (6)	7.1 (6.3)	6.0 (5.5)	0.99 (0.97, 1.01)
SARS-Cov-2 hospital admission volume per week, mean (SD)	4631	320 (245)	458 (218)	31 (17)	119 (34)	0.99 (0.98, 0.99)
SARS-Cov-2 hospital admission volume per week divided by 50, mean (SD)	4631	6.6 (4.8)	9.3 (4.3)	0.7 (0.4)	2.9 (0.6)	0.54 (0.52, 0.57)

465 *Epidemic waves were defined as: First wave, March 1, 2020 to June 30, 2020; Inter-epidemic,
 466 July 1, 2020 to November 30, 2020; Second wave, December 1, 2020 to March 31, 2021. **The
 467 odds ratio (OR) of the characteristic in the second wave relative to the first wave. ***Viral load
 468 categories: For the cobas assay, high, medium and low viral load was defined by Ct <25, 25-30,
 469 and >30, respectively. For the Xpert Xpress assay, high, medium and low viral load was defined
 470 by Ct <27, 27-32, >32, respectively.

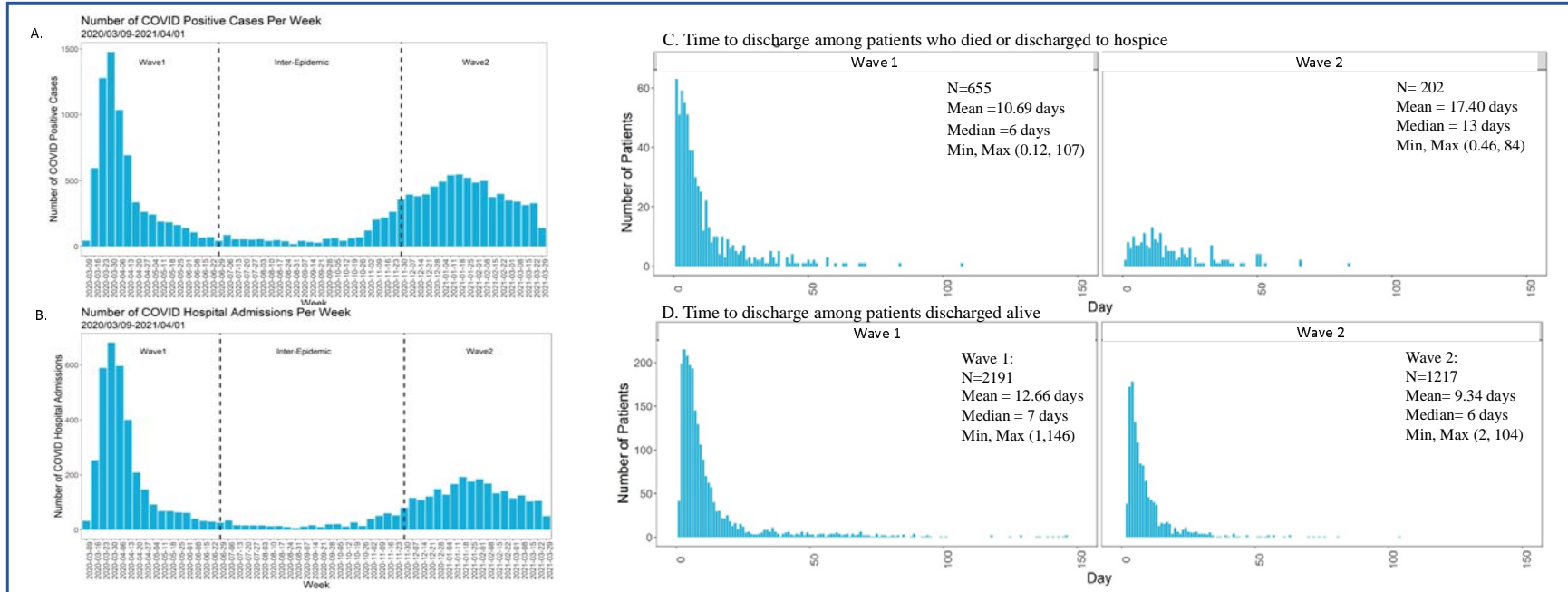
471 Table 2: Unadjusted and adjusted Cox proportional hazards model of the association between
 472 epidemic wave and death by 30 days after admission adjusted for potentially confounding
 473 factors, among all patients.

	Unadjusted	Adjusted for epi-period	
	HR (95% CIs)	HR for covariate (95% CIs) adjusted for Epidemic period	HR for epidemic period (95% CIs) adjusted each covariate individually
Epidemic Period			
First	Ref	NA	NA
Second	0.52 (0.44, 0.61)		
Age (years)			
18-50	Ref	Ref	Ref
50-65	3.14 (2.00, 4.94)	3.08 (1.96, 4.84)	0.46 (0.38, 0.54)
65-75	6.37 (4.12, 9.85)	6.36 (4.12, 9.84)	
75+	14.3 (9.40, 21.7)	15.02 (9.88, 22.83)	
Sex			
Female	Ref	Ref	Ref
Male	1.00 (0.87, 1.15)	0.97 (0.84, 1.12)	0.52 (0.44, 0.61)
Race/Ethnicity			
Hispanic/Latino	Ref	Ref	Ref
Non-Hispanic Black	1.03 (0.83, 1.29)	1.00 (0.80, 1.25)	0.51 (0.43, 0.61)
Non-Hispanic White	1.25 (1.01, 1.56)	1.29 (1.04, 1.60)	
Other	1.03 (0.86, 1.23)	1.04 (0.87, 1.24)	
BMI (kg/m ²)			
<30	Ref	Ref	Ref
≥30	0.87 (0.74, 1.02)	0.87 (0.74, 1.02)	0.56 (0.47, 0.67)
Ct value, median (IQR)	0.97 (0.96, 0.98)	0.97 (0.96, 0.98)	0.43 (0.30, 0.61)
Viral load categories			
Low (Ct >32 or 30)	Ref	Ref	Ref
Medium (Ct 27-32 or 25-30)	1.97 (1.59, 2.44)	1.98 (1.60, 2.45)	0.39 (0.27, 0.55)
High (Ct <27 or 25)	2.53 (2.09, 3.07)	2.60 (2.15, 3.16)	
Ever DNI			
Yes	13.2 (11.0, 15.9)	12.91 (10.76, 15.48)	0.58 (0.50, 0.70)
No	Ref	Ref	Ref
Oxygen level at presentation			
Room Air	Ref	Ref	Ref
Nasal Cannula	2.34 (1.86, 2.95)	2.36 (1.87, 2.97)	0.73 (0.61, 0.87)
Non-rebreather	9.32 (7.45, 11.7)	8.62 (6.87, 10.83)	
Non-invasive ventilation	8.39 (5.08, 13.9)	8.95 (5.42, 14.79)	
Intubation	8.11 (6.12, 10.7)	7.84 (5.91, 10.40)	
ICU admission by time, N (%)			
Non-ICU	Ref	Ref	Ref
ICU at presentation	3.55 (2.96, 4.26)	3.36 (2.80, 4.04)	0.56 (0.47, 0.67)
ICU after presentation	3.25 (2.71, 3.90)	3.16 (2.64, 3.79)	
Steroid use, N (%)			
No	Ref	Ref	Ref
Yes	1.15 (1.00, 1.33)	1.56 (1.34, 1.83)	0.43 (0.36, 0.52)
Remdesivir use, N (%)			
No	Ref	Ref	Ref
Yes	0.81 (0.66, 0.99)	1.34 (1.05, 1.71)	0.46 (0.37, 0.56)
Hospital admission volume divided by 50 (per week)	1.09 (1.07, 1.11)	1.08 (1.06, 1.10)	0.88 (0.70, 1.11)

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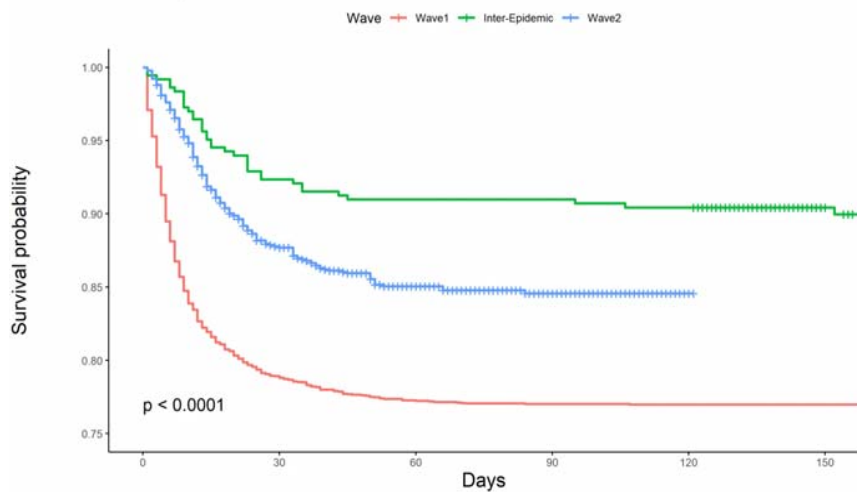
478 Figure 1. Distribution of all SARS-Cov-2 cases (A), SARS-CoV-2 admissions (B), time to death among patients who died or were

479 discharged to hospice (C), and hospital length of stay among patients who were discharged alive (D).

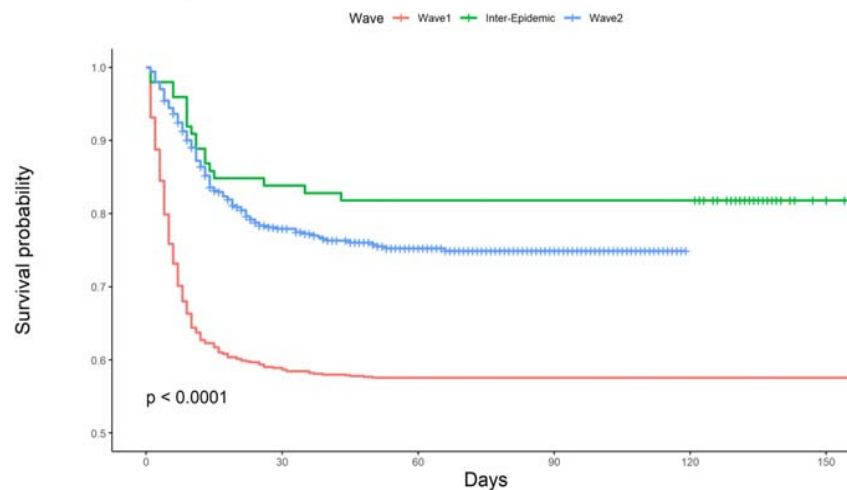
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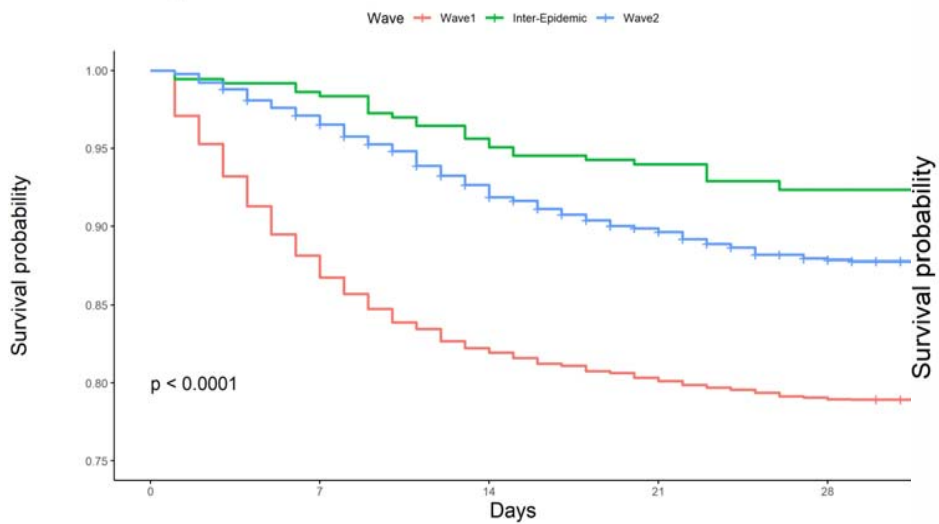
A 150-Day Survival Curve for All Patients



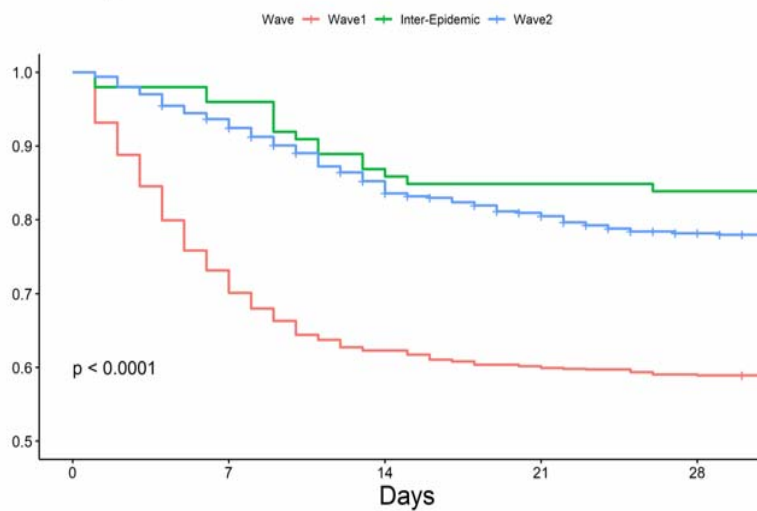
B 150-Day Survival Curve for Patients Over 75



C 30-Day Survival Curve for All Patients



D 30-Day Survival Curve for Patients Over 75



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E. Cumulative survival probabilities of all patients

Wave 1	Num at risk	2846	2508	2340	2288	2250	2199	2192	2191	2191
	Num of Cumulative events	0	377	514	566	599	648	654	655	655
	Survival Prob	100	0.8675	0.8194	0.8011	0.7881	0.7723	0.7702	0.7699	0.7699
	95% CI (LL, UL)		(0.8552, 0.8552)	(0.8194, 0.8337)	(0.7866, 0.8159)	(0.7733, 0.8033)	(0.7571, 0.7879)	(0.7549, 0.7858)	(0.7545, 0.7854)	(0.7545, 0.7854)
Inter-epidemic	Num at risk	366	361	350	344	338	333	333	331	191
	Num of Cumulative events	0	6	18	22	28	33	33	35	35
	Survival Prob	100	0.9836	0.9508	0.9399	0.9235	0.9098	0.9098	0.9044	0.9044
	95% CI		(0.9707, 0.9967)	(0.9289, 0.9732)	(0.9158, 0.9646)	(0.8967, 0.9511)	(0.8810, 0.9397)	(0.8810, 0.9397)	(0.8747, 0.9350)	(0.8747, 0.9350)
Wave 2	Num at risk	1419	1371	1274	1190	1112	1097	729	9	0
	Num of Cumulative events	0	49	114	144	167	199	202	202	202
	Survival Prob	100	0.9654	0.9186	0.8964	0.8771	0.8505	0.8456	0.8456	0.8456
	95% CI		(0.9559, 0.9750)	(0.9044, 0.9331)	(0.8805, 0.9126)	(0.8599, 0.8946)	(0.8315, 0.8701)	(0.8258, 0.8659)	(0.8258, 0.8659)	
	Time (Days)	0	7	14	21	30	60	90	120	150

F. Cumulative survival probabilities of patients > 75 years

Wave 1	Num at risk	891	652	555	536	526	513	513	513	513
	Num of Cumulative events	0	266	336	357	366	378	378	378	378
	Survival Prob	100	0.7015	0.6173	0.5993	0.5870	0.5758	0.5758	0.5758	0.5758
	95% CI		(0.6720, 0.7322)	(0.5862, 0.6500)	(0.5680, 0.6324)	(0.5555, 0.6202)	(0.5442, 0.6091)	(0.5442, 0.6091)	(0.5442, 0.6091)	(0.5442, 0.6091)
Inter-epidemic	Num at risk	99	95	86	84	83	81	81	81	42
	Num of Cumulative events	0	4	14	15	16	18	18	18	18
	Survival Prob	100	0.9596	0.8586	0.8485	0.8384	0.8182	0.8182	0.8182	0.8182
	95% CI		(0.9216, 0.9992)	(0.7926, 0.9300)	(0.7807, 0.9221)	(0.7689, 0.9141)	(0.7456, 0.8978)	(0.7456, 0.8978)	(0.7456, 0.8978)	(0.7456, 0.8978)
Wave 2	Num at risk	504	470	419	387	359	233	94	0	0
	Num of Cumulative events	0	38	82	97	108	120	121	121	121
	Survival Prob	100	0.9245	0.8357	0.8048	0.7792	0.7525	0.7489		
	95% CI		(0.9479, 0.9017)	(0.8038, 0.8689)	(0.7707, 0.8404)	(0.7434, 0.8167)	(0.7149, 0.7922)	(0.7107, 0.7809)		
	Time (Days)	0	7	14	21	30	60	90	120	150

485 Figure 2. Kaplan-Meier survival plots of all patients (A) and patients ≥ 75 years old (B) hospitalized in the first and second wave of
486 COVID-19 in New York City censored on March 31, 2021, all patients (C) and patients ≥ 75 years old (D) censored at 30 days, and
487 cumulative survival probabilities for all patients (E), and patients ≥ 75 years old (F).

488 **Supplemental material**

489 Table S1: ICD-10 codes used to query patient charts to determine pre-existing conditions for
490 each of the listed diagnoses. Electronic medical record also queried for listed text strings to
491 determine pre-existing conditions.

HTN*	ICD-10 codes:	I10, I11, I12, I13, I15, I16, O10.1, O10.2, O10.3, O10.4, O10.9
	String matches:	"htn", "hyperten"
DM*	ICD-10 codes:	E08, E09, E10, E11, E13, O24.4
	String matches:	"dm", "diabetes"
CAD*	ICD-10 codes:	I21, I22, I23, I24, I25, Z98.61, Z95.1
	String matches:	"cad", "coronary"
CKD*	ICD-10 codes:	N03, N07, N08, N11, N14, N18, N19, N29, I12, I13, Z99.2, E10.22, E11.22, E13.22, E08.22, O10.3, D63.1
	String matches:	"ckd", "chronic kidney"

492 *Abbreviations: Hypertension (HTN), diabetes mellitus (DM), coronary artery disease (CAD),
493 chronic kidney disease (CKD)

494 Table S2: Unadjusted and adjusted Cox proportional hazards model of the association between
 495 epidemic wave and death by 30 days after admission adjusted for potentially confounding
 496 factors, among those ≥ 75 years old.

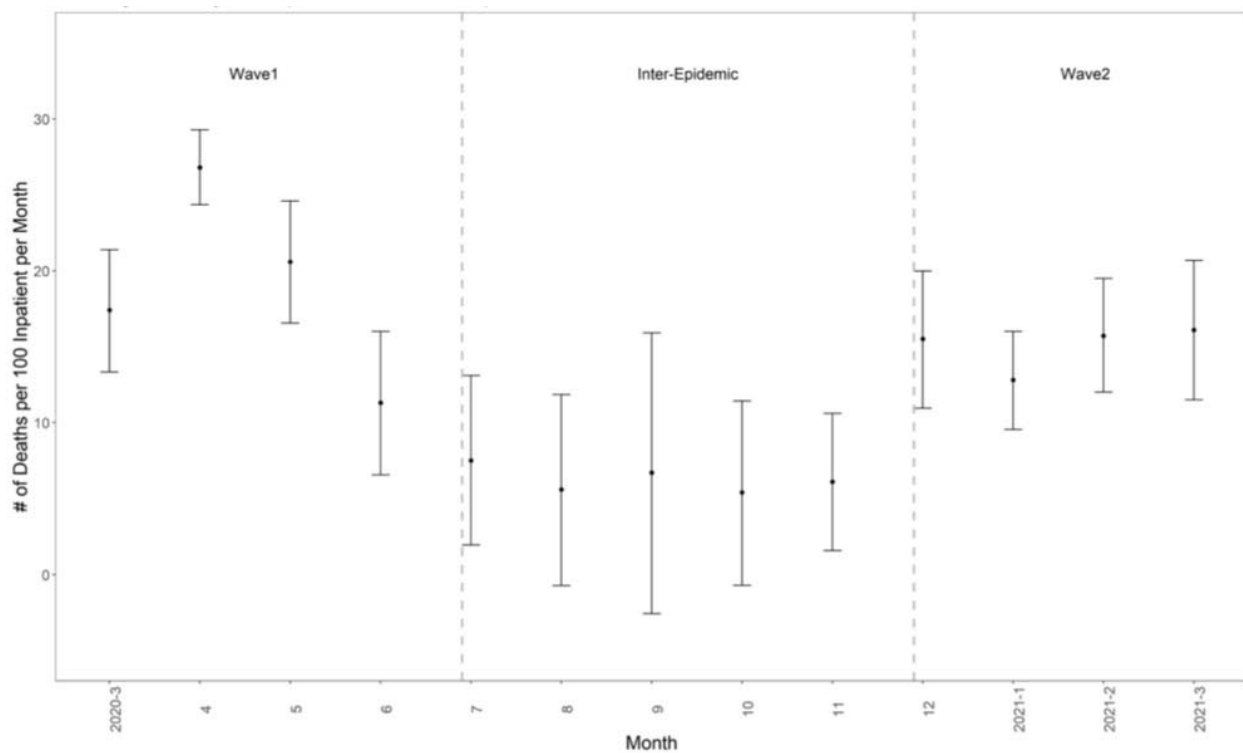
	Unadjusted	Adjusted for epi-period	
	HR (95% CIs)	HR for covariate (95% CIs) adjusted for Epidemic period	HR for epidemic period (95% CIs) adjusted each covariate individually
Epidemic Period			
First	Ref	NA	Ref
Second	0.43 (0.35, 0.53)		0.43 (0.35, 0.53)
Sex			
Female	Ref	Ref	Ref
Male	1.14 (0.95, 1.36)	1.11 (0.93, 1.33)	0.43 (0.35, 0.54)
Race/Ethnicity			
Hispanic/Latino	Ref	Ref	Ref
Non-Hispanic Black	1.14 (0.84, 1.54)	1.10 (0.81, 1.49)	0.43 (0.35, 0.53)
Non-Hispanic White	1.08 (0.83, 1.40)	1.08 (0.84, 1.41)	
Other	0.97 (0.77, 1.22)	0.98 (0.78, 1.23)	
BMI (kg/m ²)			
<30	Ref	Ref	Ref
≥ 30	1.23 (0.99, 1.53)	1.20 (0.96, 1.50)	0.48 (0.38, 0.59)
Ct value, median (IQR)	0.98 (0.97, 0.99)	0.98 (0.96, 0.99)	0.35 (0.22, 0.55)
Viral load categories			
Low (Ct >32 or 30)	Ref	Ref	Ref
Medium (Ct 27-32 or 25-30)	1.61 (1.22, 2.12)	1.62 (1.23, 2.14)	0.35 (0.22, 0.56)
High (Ct <27 or 25)	1.72 (1.35, 2.20)	1.77 (1.39, 2.26)	
Ever DNI			
Yes	6.81 (5.11, 9.06)	6.22 (4.66, 8.29)	Ref
No	Ref	Ref	0.54 (0.44, 0.68)
Oxygen level at presentation			
Room Air	Ref	Ref	Ref
Nasal Cannula	2.61 (1.93, 3.54)	2.55 (1.88, 3.46)	0.66 (0.52, 0.82)
Non-rebreather	8.65 (6.44, 11.6)	7.58 (5.60, 10.25)	
Non-invasive ventilation	4.79 (2.44, 9.39)	4.78 (2.44, 9.37)	
Intubation	7.82 (5.18, 11.9)	6.97 (4.60, 10.57)	
ICU admission by time, N (%)			
Non-ICU	Ref	Ref	Ref
ICU at presentation	2.54 (1.97, 3.27)	2.36 (1.83, 3.04)	0.44 (0.35, 0.54)
ICU after presentation	2.02 (1.51, 2.71)	2.14 (1.60, 2.86)	
Steroid use, N (%)			
No	Ref	Ref	Ref
Yes	1.00 (0.83, 1.20)	1.49 (1.22, 1.83)	0.37 (0.29, 0.47)
Remdesivir use, N (%)			
No	Ref	Ref	Ref
Yes	0.76 (0.58, 0.98)	1.69 (1.20, 2.37)	0.34 (0.26, 0.45)
Hospital admission volume divided by 50 (per week)	1.10 (1.08, 1.12)	1.08 (1.05, 1.11)	0.72 (0.53, 0.97)

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498 Table S3: Cox proportional hazards model of the association between steroid use and 30-day
 499 mortality stratified to examine confounding by disease severity.

Stratifying variable	Treatment Variable	Dead N	Alive N	HR (95% CI)
	Steroid ⁺	338	1384	1.15 (1.00, 1.33)
	No Steroid	405	2052	
Age < 75 years	Steroid ⁺	149	1034	1.47 (1.17, 1.85)
	No Steroid	139	1491	
Age ≥ 75 years	Steroid ⁺	189	350	1.00 (0.92, 1.20)
	No Steroid	266	561	
Wave 1	Steroid ⁺	185	527	1.34 (1.12, 1.59)
	No Steroid	390	1668	
Wave 2	Steroid ⁺	153	857	4.2 (2.51, 7.26)
	No Steroid	15	384	
Wave 1, ICU	Steroid ⁺	94	202	0.53 (0.41, 0.70)
	No Steroid	130	144	
Wave 1, Non-ICU	Steroid ⁺	91	325	2.22 (0.97, 5.09)
	No Steroid	260	1524	
Wave 2, ICU	Steroid ⁺	79	107	1.48 (1.17, 1.88)
	No Steroid	6	22	
Wave 2, Non-ICU	Steroid ⁺	74	750	3.83 (1.92, 7.64)
	No Steroid	9	362	
	Remdesivir ⁺	108	575	0.81 (0.66, 0.99)
	No Remdesivir	660	2922	
Age < 75 years	Remdesivir ⁺	42	420	0.85 (0.62, 1.17)
	No Remdesivir	251	2157	
Age ≥ 75 years	Remdesivir ⁺	66	155	0.76 (0.58, 0.98)
	No Remdesivir	409	765	
Wave 1	Remdesivir ⁺	8	74	0.42 (0.21, 0.84)
	No Remdesivir	592	2172	
Wave 2	Remdesivir ⁺	100	501	2.06 (1.51, 2.80)
	No Remdesivir	68	750	
Wave 1, ICU	Remdesivir ⁺	6	34	0.30 (0.13, 0.67)
	No Remdesivir	225	319	
Wave 1, Non-ICU	Remdesivir ⁺	2	40	0.27 (0.07, 1.07)
	No Remdesivir	367	1853	
Wave 2, ICU	Remdesivir ⁺	45	68	0.95 (0.62, 1.45)
	No Remdesivir	40	61	
Wave 2, Non-ICU	Remdesivir ⁺	55	433	2.96 (1.88, 4.67)
	No Remdesivir	28	689	

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Month	2020-3	4	5	6	7	8	9	10	11	12	2021-1	2021-2	2021-3
Deaths (N)	71	458	100	22	7	3	2	3	7	45	60	68	47
Admissions (N)	409	1708	486	195	93	54	30	56	115	291	470	432	292

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503 Figure S1. In-hospital monthly mortality rates per 100 patients and 95% confidence bounds

504 among 4132 patients admitted to the medical center between March 2020 and March 2021.