

Incidence and Prevalence of Coronavirus Disease 2019 Within a Healthcare Worker Cohort During the First Year of the Severe Acute Respiratory Syndrome Coronavirus 2 Pandemic

S[a](#page-0-1)rah B. Doernberg,^{1,a} Marisa Holubar,^{[2,](#page-0-2)a,©} Vivek Jain,^{[3](#page-0-3)[,a](#page-0-1)} Yingjie Weng,⁴ Di Lu,⁴ Jenna B. Bollyky,^{[5](#page-0-5)} Hannah Sample,^{[6](#page-0-6)} Beatrice Huang,^{[7](#page-0-7)} Charles S. Craik,^{[8](#page-0-8)} $^{\text{5, b}}$ $^{\text{5, b}}$ $^{\text{5, b}}$; the CHART Study Consortium^c Mandam and Yvonne Maldonado^{[5](#page-0-5)[,b](#page-0-10)}; the CHART Study Consortiu[m](#page-0-11)^c

¹Division of Infectious Diseases, University of California, San Francisco, San Francisco, California, USA; ²Division of Infectious Diseases and Geographic Medicine, Stanford University School of Medicine, Stanford, California, USA; ³Division of HIV, Infectious Diseases & Global Medicine, San Francisco General Hospital, University of California, San Francisco, San Francisco, California, USA; ⁴Quantitative Sciences Unit, Stanford University School of Medicine, California, USA; ⁵Division of Pediatric Infectious Diseases, Stanford University School of Medicine, California, USA;
⁶Department of Biochemis Department of Biochemistry and Biophysics, University of California, San Francisco, San Francisco, California, USA; ⁷Department of Family and Community Medicine, San Francisco General Hospital, University of California, San Francisco, San Francisco, California, USA; ⁸Department of Pharmaceutical Chemistry, University of California, San Francisco, San Francisco, California, USA; and ⁹Division of Infectious Disease and Global Epidemiology, Department of Epidemiology and Biostatistics, University of California, San Francisco, San Francisco, California, USA

Background. Preventing severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2_ infections in healthcare workers (HCWs) is critical for healthcare delivery. We aimed to estimate and characterize the prevalence and incidence of coronavirus disease 2019 (COVID-19) in a US HCW cohort and to identify risk factors associated with infection.

Methods. We conducted a longitudinal cohort study of HCWs at 3 Bay Area medical centers using serial surveys and SARS-CoV-2 viral and orthogonal serological testing, including measurement of neutralizing antibodies. We estimated baseline prevalence and cumulative incidence of COVID-19. We performed multivariable Cox proportional hazards models to estimate associations of baseline factors with incident infections and evaluated the impact of time-varying exposures on time to COVID-19 using marginal structural models.

Results. A total of 2435 HCWs contributed 768 person-years of follow-up time. We identifed 21 of 2435 individuals with prevalent infection, resulting in a baseline prevalence of 0.86% (95% confdence interval [CI], .53%–1.32%). We identifed 70 of 2414 incident infections (2.9%), yielding a cumulative incidence rate of 9.11 cases per 100 person-years (95% CI, 7.11–11.52). Community contact with a known COVID-19 case was most strongly correlated with increased hazard for infection (hazard ratio, 8.1 [95% CI, 3.8–17.5]). High-risk work-related exposures (ie, breach in protective measures) drove an association between work exposure and infection (hazard ratio, 2.5 [95% CI, 1.3–4.8). More cases were identifed in HCWs when community case rates were high.

Conclusions. We observed modest COVID-19 incidence despite consistent exposure at work. Community contact was strongly associated with infections, but contact at work was not unless accompanied by high-risk exposure.

Keywords. COVID-19; SARS-CoV-2; healthcare worker; healthcare personnel.

Many assume that healthcare workers (HCWs) acquire coronavirus disease 2019 (COVID-19) at work [[1](#page-10-0)[–3\]](#page-10-1). While early studies supported work-related risks, more recent studies have

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shown that other factors, including community-based exposures, race/ethnicity, and residential zip code, are associated with severe acute respiratory syndrome coronavirus 2 (SARS-CoV-2) acquisition and may be more consequential than workplace exposures [[4](#page-10-2)[–6](#page-10-3)].

Existing literature and media reports have noted widely varying estimates of prevalence, incidence, and risk factors for infection [[7](#page-10-4)–9]. Few studies have assessed HCW infection and risk longitudinally, despite changes in community prevalence of COVID-19 over time, workplace infection prevention eforts, and dynamic individual adherence to public health measures outside of work. Two European groups reported data from longitudinal HCW screening programs and estimated the prevalence and incidence of infection [\[2,](#page-10-5) [10\]](#page-10-6), but, to date, a similar granular approach to describing the prevalence and incidence of COVID-19 in US HCWs has not been reported to our knowledge.

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^bG. W. R. and Y. M. contributed equally to this work.

^cConsortium collaborators are listed in the Acknowledgments.

Correspondence: Marisa Holubar, Division of Infectious Diseases and Geographic Medicine, Stanford University School of Medicine 300 Pasteur Dr, Lane Bldg 145, Stanford, CA 94305 ([mholubar@stanford.edu\)](mailto:mholubar@stanford.edu?subject=).

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In addition, most seroprevalence studies of HCWs—both cross-sectional and longitudinal—have used a single unconfrmed serological test result without orthogonal confrmation (ie, using a diferent test) as their main outcome measure. Most studies have also not reported neutralizing antibody titers [[11](#page-10-7)– [14](#page-10-8)]. Orthogonal antibody testing increases specifcity, which is critical when testing populations with low disease prevalence [\[15](#page-10-9)]. Furthermore, neutralizing antibody titers provide a functional assessment of immune responses.

To address these current gaps in our understanding of SARS-CoV-2, we sought to estimate and characterize the prevalence and incidence of COVID-19 using both reverse-transcription polymerase chain reaction (RT-PCR) and orthogonal antibody testing in a large longitudinal cohort of HCWs during a dynamic phase of the US epidemic and to identify risk factors associated with infection among HCWs.

METHODS

Ethics Statement

The COVID-19 Healthcare Worker Antibody and RT-PCR Tracking (CHART) Study was approved by the Committee on Human Subjects Research at the University of California, San Francisco (UCSF), and the Panel on Human Subjects in Medical Research at Stanford University School of Medicine.

Study Population and Setting

From May to September 2020, we recruited HCWs from Stanford Health Care (SHC), UCSF Health (UCSF), and Zuckerberg San Francisco General Hospital for this longitudinal, prospective cohort study. These 3 medical centers serve large, mostly nonoverlapping catchment populations in the San Francisco Bay Area and implemented similar mitigation policies over time ([Supplementary Table 1\)](http://academic.oup.com/cid/article-lookup/doi/10.1093/cid/ciac210#supplementary-data). Recruitment included medical center-wide email and verbal announcements, targeted email notifications to department leaders, and recruitment flyers.

HCWs completed an electronic screening questionnaire (Supplementary Material). Inclusion criteria were (1) age ≥18 years, (2) employment at any of the 3 medical centers, and (3) not anticipating ending employment or taking leave in the next 6 months. Eligible HCWs provided consent electronically. We collected study data using Research Electronic Data Capture (REDcap) electronic data capture tools hosted at Stanford University [[16,](#page-10-10) [17](#page-10-11)].

Schedule of Evaluations

The study was conducted from July 2020 to January 2021. Participants completed up to 10 visits: 7 visits at 2-week intervals (\pm 7 days), followed by 3 visits at 4-week intervals (\pm 7 days) up to completion or the end of the study. At all visits, participants completed an electronic survey, and study staff collected nasopharyngeal (NP) swab samples; swab samples were

optional for the final 3 visits. Participants underwent phlebotomy monthly for anti–SARS-CoV-2 antibody testing. For individuals who tested positive with either RT-PCR or serology, 4 additional weekly visits were scheduled for serological testing only. Participants received no incentives or compensation for joining the study.

Laboratory RT-PCR and Serology

The UCSF Clinical Laboratories and Chan Zuckerberg Biohub analyzed samples from the UCSF and Zuckerberg San Francisco General Hospital subcohorts. Serology was performed using an assay to detect antinucleocapsid immunoglobulin (Ig) G antibodies (Abbott Architect; Abbott Laboratories) [18]. The Stanford Clinical Virology Laboratory analyzed samples using an assay to detect antispike IgG antibodies (Eurimmune Medizinische Labordiagnostika) [19] as well as a laboratory-derived assay to detect anti–receptor-binding domain IgG [\[20\]](#page-10-12). Samples that were positive at one laboratory underwent confirmatory testing at the other laboratory. Serum samples that were positive for antibodies to either spike, nucleocapsid, or both proteins were assayed for the presence of neutralizing antibodies at UCSF or at Vitalant Research Institute (San Francisco, California) by optimizing a lentivirus-based pseudotype neutralization assay [[21\]](#page-11-0).

At UCSF, RT-PCR testing was performed using (1) the M2000 Abbott RealTime SARS-CoV-2 assay [22], amplifying the RdRP and N genes; (2) the MAGPIX Luminex NxTag CoV Extended Panel assay [23], amplifying the N, the Orf1ab, and E genes; or (3) a Clinical Laboratory Improvement Amendments–validated laboratory-derived test modifed from the Centers for Disease Control and Prevention, amplifying the N and E genes [24, [25](#page-11-1)]. At SHC, RT-PCR testing was performed using an SHC laboratory–derived test amplifying the E gene or the Panther Fusion SARS-CoV-2 assay [26, 27].

Definitions of COVID-19 Exposures, Positive Test Results, and Cases

We defined a low-risk work exposure as providing direct care to or being within 6 feet of a patient with COVID-19, directly interacting with the environment where a patient with COVID-19 received care, or processing laboratory samples from a patient with COVID-19. We defined a high-risk exposure at work as ever interacting with a patient who had COVID-19 without wearing full personal protective equipment (PPE)—the institutionally recommended PPE for care of patients with COVID-19—or while having a breach in PPE (eg, tears or accidental removal).

We defned an RT-PCR result as positive if the result was (1) detected or (2) indeterminate (positive RT-PCR result followed by negative subsequent confrmatory RT-PCR test results(s) obtained according to medical center occupational health protocols).

We defned a confrmed positive serological result as an initial positive serological result (antinucleocapsid or antispike antibody) followed by confrmation with a second positive serological result using a diferent target (antinucleocapsid, antispike, or neutralizing antibodies). A confrmed positive serological result represented prior COVID-19 infection. We defned an unconfrmed positive serological result as an isolated positive antinucleocapsid or antispike antibody result (ie, a negative result on confrmatory testing) in the absence of RT-PCR positivity.

We defned baseline prevalent cases as those in participants with a positive RT-PCR or confrmed positive serological result at their initial visit. Participants who did not have baseline infection entered an incident cohort. We defned incident cases among this cohort as those in participants with a positive RT-PCR or confrmed positive serological result at any subsequent visit. The date of incident infection was the first date on which either the RT-PCR or the frst serological result was positive (if confrmatory testing occurred within 4 weeks).

Statistical Analyses

We estimated the prevalence as the proportion of cases at baseline among the total number of enrolled participants who completed baseline visits. We estimated the cumulative incidence as the number of incident cases divided by the total follow-up time per 100 person-years and assumed a uniform incidence distribution across the 6-month follow-up time. We censored person-time when a participant met the case definition, completed or withdrew from the study, or received a first dose of any COVID-19 vaccine. We calculated the confidence intervals (CIs) using a nonparametric bootstrapping method. We conducted a sensitivity analysis to assess the impact of different case definitions on estimates considering (1) all unconfirmed positive serological results as cases and (2) all individuals with a single positive RT-PCR result, no positive serological result, and ≥1 serological measurement ≥4 weeks after the positive RT-PCR result as potential false-positives and removing them from case counts. We obtained community-wide data on COVID-19 incidence in the 6 Bay Area counties from the California Department of Public Health [\[28\]](#page-11-2).

We compared characteristics of prevalent and incident cases with those of noncases. For binary time-varying exposures, we used participant self-report at the most recent visit before censoring. For continuous time-varying exposures, we computed median responses across all visits before censoring. We reported symptoms using the most recent reported status at the visit at which infection was identifed. We reported standardized mean diference to describe the magnitude of diferences in characteristics between incident cases and noncases. The magnitude of efect is considered small, medium, or large with a standardized mean diference of 0.2, 0.5, or 0.8, respectively.

In the incident cohort, we frst assessed associations between time to infection and baseline characteristics using multivariable

Cox proportional hazards models. We evaluated the impact of prespecifed time-varying exposures on time to infection via marginal structural models [[29–](#page-11-3)[32](#page-11-4)]. We implemented a 2-step marginal structural model for each time-varying exposure by frst estimating the inverse probability of treatment weights, in which exposure probability was estimated for each participant at each visit, conditioning on fxed and other time-varying exposures up to that time. To stabilize weights, we excluded correlated time-varying variables. Each participant was weighted with the inverse predicted probability of exposure to simulate a counterfactual participant. Second, we applied an extended Cox proportional hazard model with inverse probability of treatment weights and reported hazard ratios (HRs) for the impact of time-varying exposures on time to infection. For all regression analyses, we imputed missing laboratory data using a last observation carried forward method and missing time-invariant or time-varying data using multiple imputation. We controlled family-wise type I error at 0.05 and used the signifcance level of .05 in hypothesis tests. All analysis were conducted (by Y. W., D. L., and M. D.) using SAS 9.4.3 sofware (SAS Institute) and R sofware, version 4.5.3. (R Project for Statistical Computing).

RESULTS

HCW Demographics

Of 3918 individuals screened, 2435 provided consent and completed the first study visit, contributing 768 total person-years of follow-up time [\(Figure 1\)](#page-3-0). Baseline demographics are presented in [Table 1.](#page-4-0) Overall, participants' mean age was 40.4 years (standard deviation, 10.1 years), 1923 of 2435 (79%) were female, and most participants (1921 of 2435 [79%]) reported providing direct patient care, including 701 of 1921 (36%) who performed aerosol-generating procedure . Many participants reported work-related COVID-19 exposure (1477 of 2419 [61%]) with 797 of 1477 (54%) reporting high-risk exposure. Only 176 of 2419 participants (7%) overall reported contact with a COVID-19–positive person outside of work.

Overall, demographic and behavior characteristics of participants with prevalent and incident COVID-19 refected overall cohort characteristics [\(Table 1\)](#page-4-0), including 73 of 91 (80%) providing direct patient care, mostly as nurses (42 of 91 [46%]) or clinicians (MD, MD-equivalent, APP, or trainee) (24 of 91 [26%]). During the course of the study, time spent in the healthcare environment and work-related exposures to COVID-19 were both stable for HCWs ([Figure 2A](#page-7-0) and [2B](#page-7-0)).

Prevalence and Incidence of COVID-19

We identified 21 of 2435 individuals with evidence of COVID-19 at baseline and estimated a prevalence of 0.86% (95% CI, .53%–1.32%). We identified 70 of 2414 individuals (2.9%) with incident COVID-19 during follow-up and estimated a cumulative incidence rate of 9.11 cases per 100 person-years (95%

Figure 1. Participant flow diagram. A confirmed positive serological result was defined as an initial positive result (antinucleocapsid or antispike antibody [Ab] result), followed by confirmation with a second positive serological result using a different target (antinucleocapsid, antispike, or neutralizing Abs). A positive unconfirmed serological result was defined as an isolated positive antinucleocapsid or antispike Ab result (ie, a negative result with confirmatory testing) in the absence of reverse-transcription polymerase chain reaction (RT-PCR) positivity. Prevalent coronavirus disease 2019 (COVID-19) cases were defned as those in participants with a positive RT-PCR or a confirmed positive serological result at baseline, and incident COVID-19 cases as those in participants with a positive RT-PCR or a confirmed positive serological result at any subsequent visit.

CI, 7.11–11.52). The number of incident cases increased with rising prevalence of COVID-19 in the 8-county region in which the study was conducted ([Figure 2\)](#page-7-0). Incidence rate estimates did not differ by subgroups of sex, race/ethnicity, or job role [\(Supplementary Figure 1](http://academic.oup.com/cid/article-lookup/doi/10.1093/cid/ciac210#supplementary-data)).

All 21 prevalent COVID-19 cases met the case defnition with a positive serological result; only 3 also had a positive RT-PCR result. Most of the 70 incident cases were identifed by a positive RT-PCR result (53 of 70 [76%]), with or without a positive serological result. Of the 17 of 70 participants (24%) meeting the case defnition by positive serological results alone, only 2 (12%) had a positive RT-PCR result at a later visit (2 or 5 weeks afer the positive serological result).

We performed a sensitivity analysis using an alternative incident COVID-19 case defnition that included all unconfrmed positive serological results as cases, resulting in 26 prevalent and 71 incident cases. This slightly increased the baseline prevalence to 1.07% (95% CI, .79%–1.56%) and increased the cumulative incidence rate to 9.26 cases per 100 person-years (95% CI, 7.24–11.69).

To examine the impact of potential false-positive RT-PCR results, we performed a second sensitivity analysis using a second alternative case defnition that excluded 7 cases meeting this definition. This decreased the cumulative incidence rate to 8.18 cases per 100 person-years (95% CI, 6.29–10.4). Overall, the testing yield of the incident cohort was relatively low: only 30 of 12 007 RT-PCR tests (0.25%) performed in asymptomatic participants had positive results, and 7 of the 30 (23%) met the false-positive case defnition.

[Figure 3](#page-8-0) demonstrates the participant-level temporal sequence of testing results for all baseline prevalent cases and all incident cases. We found substantial evolution of antibody responses over time: of the 56 cases initially diagnosed using RT-PCR, 11 had ≥1 positive antibody at diagnosis. By the end of

Table 1. Continued

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bThe SMD is a comparative measure of effect size between groups. The magnitude of effect is considered small if the SMD is 0.2, medium at 0.5, and large at 0.8.

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follow-up, this rose to 27 individuals with ≥1 positive antibody test result. Furthermore, 11 individuals who had antibodies de tected with one assay subsequently tested positive using another assay.

COVID-19 Symptoms

Of the 70 incident cases, 36 (51.4%) were asymptomatic at diagnosis (30 with positive RT-PCR results and 6 with posi tive serological results only). Of the 36 participants who were asymptomatic at diagnosis, 14 of 22 (64%) who completed a fol low-up symptom assessment had remained asymptomatic. Of the 25 participants who were symptomatic at diagnosis, 19 had positive RT-PCR results and 6 had positive serological results only.

Among the 1170 participants who reported symptoms at any visit, 58 (5%) were confrmed as having prevalent or incident cases. Among 1252 participants who never reported symptoms, 32 (3%) were confrmed as having prevalent or incident cases.

While participants with incident cases more commonly re ported ever having symptoms (48 of 70 [69%]), many categor ized as noncases (1112 of 2344 [48%]) reported symptoms at least once ([Supplementary Table 2](http://academic.oup.com/cid/article-lookup/doi/10.1093/cid/ciac210#supplementary-data)). The most common symptoms reported by those categorized as noncases were fatigue (326 [14%]), headache (46 [20%]), nasal congestion (325 [14%]), and rhinorrhea (412 [17%]); those categorized as noncases in frequently reported fever, chills, or decreased sense of taste or smell, while case patients reported them more commonly.

Predictors of COVID-19 Infection

In a multivariable Cox proportional hazards model, we did not find an association of incident COVID-19 with fixed vari ables, including baseline age, sex, race, ethnicity, household size, role, and work category [\(Supplementary Table 3\)](http://academic.oup.com/cid/article-lookup/doi/10.1093/cid/ciac210#supplementary-data). In marginal structural models of self-reported time-varying variables, community contact with a known COVID-19 case was strongly correlated with increased hazard for COVID-19 (HR, 8.1 [95% CI, 3.8–17.5]; [Table 2\)](#page-9-0). Self-reported exposure to a patient with COVID-19 at work was associated with infection $(P = .01)$, but this appeared to be primarily driven by high-risk exposures (ie, a PPE failure or breach or an exposure to patient biolog ical material; HR, 2.5 [95% CI, 1.3–4.8]). Increasing community COVID-19 case rate showed a trend toward elevated adjusted hazard of HCW infection, but this finding did not reach sta tistical significance (HR, 1.3 [95% CI, .97–1.8]). Time spent in the healthcare workplace, time spent providing direct patientfacing care, and adherence to community mitigation strategies were not associated with COVID-19 infection.

DISCUSSION

In this large observational cohort of HCWs, we observed modest COVID-19 infection rates despite consistent COVID-19

B. Composite COVID-19 exposures at work

C. COVID-19 exposures outside of work

A. Time spent in healthcare setting

D. Community and healthcare worker COVID-19 infections over time

Figure 2. Work and community-related coronavirus disease 2019 (COVID-19) exposures and incident cases among healthcare workers (HCWs) over time. *A–C,* Selfreported work and home exposures over time. Each line depicts the 7-day smoothed median responses of each self-reported home or community behavior or exposure. Gray shading represents 95% confdence intervals around the average. *D,* Incident cases in the context of surrounding community caseload. Boxes represent unique incident cases and are color coded by how they met case defnitions, and the line represents the 14-day smoothed average of community-reported cases from the 6 San Francisco Bay Area Counties surrounding the 3 medical centers. Abbreviation: RT-PCR, reverse-transcription polymerase chain reaction.

exposure at work. Changes in COVID-19 incidence tracked most closely with community infection rates and self-reported community contact with known COVID-19 cases rather than work-related factors, except when breaches in standard safety protocols or PPE occurred [\[5\]](#page-10-13). Our data provide evidence of the overall safety of standard healthcare work environment protocols and PPE guidelines and are concordant with emerging literature showing that the main COVID-19 related risks to HCWs are those coming from home- and community-based factors [\[33](#page-11-5)].

By combining longitudinal and orthogonal RT-PCR and serological testing, our study allowed for a robust granular estimation of the true incidence of COVID-19 infection among HCWs. Unlike many studies based on a single serological test, we used confrmatory serology and also measured neutralizing antibody responses [[34\]](#page-11-6). As our data show, the serological response to infection is multifaceted and evolves over time; measuring a single antibody response to one target may result in inaccurate estimates of true infection rates [[35\]](#page-11-7). By testing serially and confrming antibody responses, we captured

Baseline COVID-19 cases

Incident COVID-19 cases

Figure 3. Timing and sequence of positive tests among healthcare workers with coronavirus disease 2019 (COVID-19). Each row represents all test results for each prevalent and incident case over the study period. Gray shading indicates each participant's follow-up time. Dots represent reverse-transcription polymerase chain reaction (RT-PCR) results and boxes represent serological results. Blue coloring represents negative RT-PCR or serological results; red coloring, positive RT-PCR or confrmed positive serological results. Orange boxes represent unconfirmed positive serological results. The thickness of the red boxes is correlated with the number of confirmed positive serological results (eg, 2 or 3 positive antibody test results).

Table 2. Marginal Structural Model of Variables Associated With Incident Coronavirus Disease 2019[a](#page-9-1)

Abbreviations: CI, confdence interval; COVID-19, coronavirus disease 2019; HR, hazard ratio.

^aFor each marginal structural model, we estimated the inverse probability of treatment weights in which exposure probability was estimated for each participant conditioning on fixed variables (age, sex, race, ethnicity, household size, role, and work category) and the time-varying variables shown in the table

bLow-risk exposure at work included interacting with a patient who had COVID-19, with no reported breach in personal protective equipment (PPE) or other safety protocols. High-risk exposure at work was defined as ever interacting with a patient with COVID-19 without wearing full personal protective equipment (PPE) or while having a breach in PPE (eg, tears or accidental removal).

some COVID-19 cases that would have likely been missed with a single test in time and excluded others that were likely false-positives.

In our sensitivity analysis accounting for potential falsepositive COVID-19 cases, our incidence estimates were 10% lower. This may have been an underestimate because many cases were diagnosed at the end of the study and lacked follow-up time to diferentiate true-positives from false-positives. Misclassifying false-positive test results as true cases can afect the ability of a healthcare system to operate by limiting critical staffing and can also have adverse implications for household contacts of HCWs. COVID-19 screening programs for HCWs must balance the value of prompt diagnosis with the downside of potential false-positive results.

Our study is subject to several limitations. We enrolled volunteer participants and had a high fraction of MD, MD-equivalent, and RN practitioners. This cohort composition did not comprehensively refect the occupational diversity within our medical centers. Thirty-eight percent of those screened did not enroll in the study; because we did not assess reasons for nonparticipation, it is unclear to what degree this may have introduced any bias in our study population. We relied on self-reporting of COVID-19–related risks both at work and home, which may have resulted in overreporting of adherence to protective measures. In addition, our institutional PPE

recommendations changed over time; as such, not all breaches in PPE are considered equivalent. However, unlike many studies that have used information from employee health and safety offces, our study was independent of the medical centers in order to foster confdential no-fault reporting.

Because sequencing of virus was beyond the scope of the study, the association between self-reported breach in PPE and incident COVID-19 cases remains solely an association and not proof that the breach itself led to the incident infection. In addition, we did not perform orthogonal SARS-CoV-2 antibody testing on samples that were initially antibody negative, and we thus could have missed certain incident cases. We also did not include confirmatory testing of RT-PCR results so could have inadvertently included false-positive RT-PCR results in incidence rate estimates. We addressed this with a sensitivity analysis and found that incidence rates were minimally affected. Finally, our study was conducted before more recent variants of concern with increased transmissibility and immune escape emerged. One key strength of our study was our use of marginal structural modeling, using detailed longitudinal data to better estimate risks.

Within a large group of frontline HCWs, our data indicate that healthcare workplaces pursuing comprehensive mitigation strategies can operate safely despite facing sequential waves of COVID-19 cases. However, HCWs do face community-based risks for acquiring COVID-19. Medical center infection control practices, vaccination programs, and community mitigation approaches should be sustained and maximized to protect HCWs and health systems during periods of future risk related to rising caseloads and emerging SARS-CoV-2 variants.

Supplementary Data

Supplementary materials are available at *Clinical Infectious Diseases* online. Consisting of data provided by the authors to beneft the reader, the posted materials are not copyedited and are the sole responsibility of the authors, so questions or comments should be addressed to the corresponding author.

Notes

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REDCap platform. The Stanford Research Electronic Data Capture (REDCap) platform [\(http://redcap.stanford.edu\)](http://redcap.stanford.edu) is developed and operated by the Stanford Medicine Research IT team, and the platform's services at Stanford are subsidized by the Stanford School of Medicine Research Office and the National Center for Research Resources and National Center for Advancing Translational Sciences, National Institutes of Health (grant UL1 TR001085). REDCap is a secure, web-based sofware platform designed to support data capture for research studies, providing an intuitive interface for validated data capture; audit trails for tracking data manipulation and export procedures; automated export procedures for seamless data downloads to common statistical packages; and procedures for data integration and interoperability with external sources.

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CHART Study Consortium collaborators. University of California, San Francisco (UCSF): Parul Bhargava, Markus Bohn, Jessica Chao, Charles S. Craik, Sarah B. Doernberg, Jacob Ghahremani, David Glidden, Ralph Gonzales, Beatrice Huang, Sravya Jaladanki, Aida Julien, Daniel Lowenstein, Steve Miller, Audrey Mustoe, Marcus Paoletti, George W. Rutherford,

Hannah Sample, Rodolfo Villa, Emerald Wan, and Aimee Williams. San Francisco General Hospital, UCSF: Lillian Brown, Jessica Chuang, Vivek Jain, Carina Marquez, Guntas Padda, Luis Rubio, and Daisy Valdivieso. Stanford Medicine/Stanford Health Care: Rosebay Abad, Anthony Bet, Jenna Bollyky, Manisha Desai, Jefrey Fung, Anna Graber, Cole Holderman, Marisa Holubar, Hannah Kelley, Amanda Kempema, Christina Kong, Christopher Leung, Joseph Lohmann, Di Lu, Yvonne Maldonado, Lloyd Minor, Lorena Orozco, Benjamin A. Pinsky, Jamie Saxeena, Matthew Sklar, Hilary Tang, Jasmine Wiese, and Yingjie Weng. Chan Zuckerberg Biohub: Emily Crawford, Joe DeRisi, and all members of the CLIAHUB Consortium.

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