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A systematic review and meta-analysis of sex differences in cannabis use disorder amongst people with comorbid mental illness

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Abstract

Background: While males are more likely diagnosed with cannabis use disorder (CUD), females are more susceptible to developing and maintaining CUD. Yet, for both sexes, CUD is associated with high rates of comorbid mental illness (MI).

Objectives: To identify and compare sex differences in the prevalence of comorbid CUD amongst individuals with/without MIs.

Methods: This systematic review generated pooled odds ratios (OR) and 95% confidence intervals (CI) from 37 studies (including clinical trials, cohort, and case-control studies) among individuals with and without MIs, quantifying sex differences in rates of comorbid CUD. A meta-analysis was also completed.

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Disclosure statement

Dr George reports that he is a consultant for Frutarom, Novartis, and the Canadian Centre for Substance Use and Addiction in the last 12 months. Dr Cooper reports that she served on the scientific advisory board of FSD Pharma in the last 12 months.

Results: In the CUD-only group, males were twice as likely to have CUD than females (OR = 2.0, CI = 1.9-2.1). Among MIs, males were more likely than females to have CUD comorbid with schizophrenia ($OR \sim 2.6, CI = 2.5-27$) and other psychotic, mood, and substance use disorders (1 > OR < 2.2, CI = 0.7-2.6). The reverse association (females > males) was observed for anxiety disorders and antisocial personality disorder (OR = 0.8, CI = 0.7-1.0). Among females. MIs increased the likelihood of having CUD, except for psychotic disorders and depression. A meta-analysis was inconclusive due to high heterogeneity across studies. Thus, comparisons across MI groups were not possible.

Conclusion: While males are more likely to be diagnosed with CUD, there are important sex differences in the prevalence of CUD across MI diagnoses that should be taken into account when approaching CUD prevention and determining treatment efficacy.

Keywords

Cannabis; cannabis use disorder; sex differences; mental illness; comorbidity; schizophrenia; depression; bipolar disorder; meta-analysis

Introduction

The prevalence of cannabis use and cannabis use disorders (CUD) has been rising worldwide, which may reflect the changing legal landscape surrounding cannabis use and reduced perceptions of risk (1). Globally, consumption of cannabis (annual prevalence) is estimated at 2.5% (2), while 0.2% of the global population is diagnosed with CUD (3). As recreational and medical use have become more widespread, the United States reported a 50–120% increase in 2016 in the annual prevalence of cannabis use, daily cannabis use, and mild CUD (4). In Canada, the prevalence of weekly cannabis use doubled and occasional use increased by 50% after legalization (5,6). Moreover, the number of individuals seeking treatment for CUD globally has been on an upward trajectory, with Europe, Australia, and the United States (US) recently reporting increases of 30% (7). Detrimental consequences associated with CUD include increased risk of psychosis and psychosocial impairment (8), and high healthcare costs, with CUD-related hospital costs estimated ~\$4.5 billion annually (9). Given that there are no approved pharmacological treatments, and behavioral treatments are only modestly effective for CUD, problematic cannabis use continues to be a significant public health concern (7).

It is well documented that CUD manifests differently between males and females in terms of development, severity, trajectory, and responsiveness to treatment (10,11). For instance, compared to females, males have an earlier age of onset of cannabis use (12) and a higher probability of initiating cannabis use (13). Males are also twice as likely to continue using cannabis than females (14) (past year prevalence 4.2% versus 1.7%, respectively), contributing to the higher rates of CUD observed among males (3.5%) compared to females (1.7%) (15). Social factors including cultural, familial, and socioeconomic may explain why males have a greater likelihood of CUD than females (10,16–18). For instance, studies show that males with CUD were more likely to be older than 45, have a high school education, or less and income over 20,000 USD and less likely to be unemployed or widowed/divorced compared to females (10,19). Yet, more recent data demonstrate that cannabis use is

increasing at a faster rate among females relative to males. The National Survey of Drug Use and Health (NSDUH) found that between 2015 and 2018 the prevalence of cannabis use among females (aged 18–25) increased by 3.8% compared to males who showed increases of only 0.8% (20), suggesting that the sex difference in cannabis use prevalence is narrowing. The increasing rates of cannabis use among females raise concern given that females may have enhanced susceptibility to develop and maintain problematic cannabis use. For example, compared to males, females show greater sensitivity to the reinforcing properties of cannabis, have an accelerated progression to CUD (e.g., telescoping effects), experience more severe withdrawal symptoms, and exhibit poorer treatment outcomes (see Cooper and Craft, 2018 (21)). Thus, increasing rates of cannabis use among females are of high clinical relevance.

A significant factor that complicates the course and treatment of CUD is that nearly 100% of individuals with CUD are diagnosed with a mental illness (MI) (e.g., schizophrenia, depression, anxiety, or substance use disorders (SUDs)) (22). Yet, current treatments for CUD that are sex-specific (23) do not typically consider comorbid MIs.

To date, the influence of sex on the rates of CUD comorbid with MIs remains largely unexplored. Therefore, this systematic review and meta-analysis aimed to identify and compare sex differences among individuals with and without MI on rates of CUD, measured by pooled odds ratios (ORs). We then discuss cannabis use and putative neurobiological mechanisms that may contribute to sex differences in CUD comorbid with MIs. As policies related to cannabis use are rapidly evolving across the globe leading to increased accessibility, availability, and reduced perception of risk (7), more individuals (with and without MIs) are at risk for developing CUD (24); this may be particularly true for females who were previously deterred from obtaining cannabis illegally (25). Therefore, a better understanding of the difference in the rates of CUD comorbid with MIs between males and females is necessary to guide the development of sex-specific prevention strategies and treatment approaches for these patients.

Methods

Study selection

This systematic review and meta-analysis examined rates of CUD in cannabis users with and without comorbid MIs, as a function of sex, and followed the guidelines for Preferred Reporting Items for Systematic Review and Meta-Analysis (PRISM-A) (26). This search strategy used Google Scholar, PubMed, Medline, and Psychlnfo databases to search for papers that included the keywords and were published in English from database inception to February 28th, 2020. The search included the keywords: 'cannabis use disorder,' 'gender,' and 'sex.' Keywords were also searched in combination (i.e., 'cannabis use disorder' AND 'prevalence' AND ['mental disorder' OR specific psychiatric comorbidity (e.g., 'schizophrenia,' 'psychotic disorder,' 'depression,' 'bipolar disorder,' 'anxiety disorder,' 'mood disorder,' 'posttraumatic stress disorder (PTSD),' 'personality disorder,' 'substance use disorder (SUD)')] AND ('gender' OR 'sex')). We included studies which referred to CUD as cannabis abuse and/or dependence. Three authors (KK, DJEL, RAR) initially assessed titles and abstracts identified by the search and reviewed the full text of the

remaining articles for inclusion. Any conflicts that arose were resolved by the senior author (TPG).

Studies were included if they met the following criteria: (1) included population-based data on individuals with and without CUD and with or without comorbid MI; (2) for the CUD-only group, had a diagnosis of past year or lifetime CUD; (3) for the comorbid MI groups, had a diagnosis of past year or lifetime CUD and the presence of one diagnosis for MI; and (4) the sex distribution within the study sample was provided. Exclusion criteria included: (1) studies examining recreational or medical cannabis use not meeting criteria for CUD; (2) for the CUD-only group had a diagnosis for MI other than CUD; (3) treatment trials, reviews, case series/reports, commentaries, opinion, unpublished studies, conference posters/abstracts; and (4) when multiple studies were found reporting on the same population cohort, only the most recent study was included.

Statistical approach

To provide an estimate of the extent to which sex is associated with having CUD comorbid with MIs, ORs for individual studies, and pooled ORs for each MI, along with 95% Confidence Intervals (CI) were calculated using SPSS version 26 for Windows. For each study, diagnostic outcome (with or without CUD) and sex (male or female) were represented by arranging observed counts into 2×2 tables to calculate individual ORs. The counts used in the table were either available in the original text or calculated from the data provided in the article. For each MI, study level counts were summated into a 2×2 table from which an overall comorbid MI pooled OR, and respective 95% CI was calculated. Comparisons of pooled ORs were performed by displaying pooled ORs and their respective CIs in a forest plot (see Figure 2). OR >1 indicates a greater likelihood of CUD in males in comparison to females, while OR <1 indicates a lower likelihood of CUD in males in comparison to females (OR = 1 indicates equal likelihood of CUD in both sexes). We avoided statistical pairwise testing between pooled ORs, since these analyses were exploratory, and such an approach would result in low power due to correcting for multiple comparisons.

In order to calculate ORs several assumptions were made when the sample sizes were not provided in the study article. (1) If the sex distribution for the MI sample was not provided, rates from the literature were used (see Supplementary Table 1); (2) Unless specifically stated, we made the assumption that participants without CUD had the MI that was being examined; (3) When studies used data from a larger dataset (e.g., NESARC-I), we used the total sample from the original (larger) dataset; (4) If there was a discrepancy regarding the sample size listed in a study's Methods section versus when calculated from sample data in their Results section, we used data calculated from data in the Results section. Of the 12 (of 37) studies where no sex distribution data for MI was provided, we attempted to contacted authors but were unable to obtain such data (see Supplementary Tables 1 and 4).

Meta-analysis—We used the METAFOR package in R statistical software (27) to estimate meta-analytic odds ratios for the association between sex and CUD, within groups of studies categorized by psychiatric diagnosis. We estimated random effects, multi-level meta-analytic models for each diagnosis group. Random effects meta-analysis is suitable when there

is a high degree of heterogeneity between estimates (28). Meta-analytic estimates from random-effects models are interpreted as an average across multiple sampled populations, rather than a representative estimate from a single sampled population. Multi-level modeling allowed us to account for non-independence when multiple estimates from the same study were included in a model (29). We examined heterogeneity using the Q-statistic and I^2 (30). The Q-statistic is interpreted as a hypothesis test with a null-hypothesis of no heterogeneity between study estimates. This statistic tends to be overpowered for large sample sizes; therefore, we also calculated $I^2.I^2$ is interpreted as the percent of heterogeneity between studies that is nonrandom. In other words, I^2 represents the proportion of heterogeneity between studies that results from explainable population/sample differences, rather than random heterogeneity expected when multiple samples are pulled from the same underlying population. Higher I^2 values indicate that random-effects modeling is the appropriate approach.

Results

Study selection results

See Figure 1 for the CONSORT diagram. A total of 5,493 studies were found in the overall search strategy, and 5,443 studies were excluded (3,866 were duplicates and content of 1,577 studies was not relevant). Fifty full-text articles were assessed for eligibility and 37 studies were included in this review (Table 1); see Supplementary Material, Table 2 for reasons of study exclusion. Diagnoses for MI included non-specified psychiatric disorder, schizophrenia, psychotic disorder, non-specified mood disorder (major depression, bipolar I and II), major depression, bipolar I and II, non-specified anxiety disorder, PTSD, non-specified personality disorder, antisocial disorder (ASPD), non-specified SUD, opioid use disorder (OUD), alcohol use disorder (AUD) and tobacco use disorder (TUD). Fourteen of these studies required at least one assumption to be made when calculating ORs for the individual studies and pooled ORs (see Supplementary Material, Table 4). The prevalence of CUD in males and females within each SMI for each study is reported in the Supplementary Material, Table 5.

Sex differences in the prevalence of CUD

Five studies were included in the CUD-only group (31-35) (Supplementary Material, Table 3). The pooled OR for the CUD-only males compared to CUD-only females was 2.0 (CI = 1.9-2.1, p < .01) as presented in Table 1. In the forest plot, the pooled OR for the CUD-only group is represented by a dashed vertical line (Figure 2).

Sex differences in the prevalence of CUD with comorbid MIs

Individual study ORs and pooled ORs for each MI are found in Table 1. Pooled ORs for CUD comorbid with MIs are compared in Figure 2. Except for schizophrenia/psychotic disorder and depression, we found that among females with MIs, the likelihood of CUD was lower (OR <2) than the CUD-only group.

Comorbid non-specified psychiatric disorders

Results from pooling seven studies (10,36–41) examining rates of CUD in non-specified comorbid psychiatric disorder produced an OR of 1.20 times greater for males than females (Cl = 1.19-1.22, p < .01).

Comorbid psychotic disorders

The pooled OR from seven studies examining psychotic disorders revealed that having a CUD was 2.1 times more likely for males than females (CI = 1.8–2.6, p < .01) (10,42–47). Similarly, the pooled OR from seven studies in schizophrenia was 2.6 times greater for males than females (36,40,47–51) (OR = 2.6, CI = 2.5–2.7, p < .01).

Comorbid mood disorders

The pooled OR from two studies in non-specified mood disorder revealed that the odds of having CUD were 1.83 times greater for males than females (CI = 1.80-1.86, p < .01) (33,40). From eight studies, the pooled OR of CUD and comorbid depression was 2.2 times greater for males compared to females (CI = 2.05 to 2.31, p < .01) (10,33,36,47,49,52–54). From 16 studies, the pooled OR of CUD and comorbid bipolar I and II was 1.7 times greater for males than females (CI = 1.6-1.8, p < .01) (10,33,36,47,49,55–65).

Comorbid anxiety disorders

The pooled OR from four studies of non-specified anxiety disorder revealed that the odds of having CUD were 1.6 times greater for males compared to females (CI = 1.5-1.7, p < .01) (10,33,36,40). Based on one study, the OR of CUD was 1.03 greater for males than females (CI = 0.8-1.4, p < .01) with comorbid PTSD (33).

Comorbid personality disorders

The pooled OR of four studies in non-specified personality disorder revealed that the odd of having CUD was 1.4 times greater for males than females (CI = 1.3-1.5, p < .01) (10,33,36,40). The pooled OR from the two studies examining CUD comorbid with conduct disorder was 0.94; however, the sex difference was not statistically significant (CI = 0.7-1.2, p = .65) (10,40). Based on one study, the OR was 0.84 greater for females than males with ASPD (CI = 0.7-1.0, p < .05) (10).

Comorbid SUDs

The result of pooling three studies (10,33,40) examining CUD comorbid with non-specified SUD was 1.31 greater for males than females (CI = 1.28-1.34, p < .01). The OR from the only study of CUD and comorbid OUD was 1.36 greater for males compared to females (CI = 1.11-1.68, p < .01) (66). The pooled OR from three studies (10,33,40) for CUD and comorbid AUD was 1.51 greater for males than females (CI = 1.46-1.55, p < .01). Lastly, the pooled OR from two studies (10,33) on TUD was 1.72 greater for males than females (CI = 1.58-1.87, p < .01).

Summary of meta-analysis—When we conducted our meta-analysis, we observed that our I^2 values were consistently very high. With the exception of one I^2 estimate of 68.8%

(psychotic disorders; presented in Supplementary Material, Table 6), all others were >85%. This high degree of heterogeneity indicates that a random-effects model and interpretation are appropriate; in other words, the meta-analytic ORs (Table 6) are best interpreted as an average across studies from multiple populations, rather than a representative estimate of the association for a single diagnostic population. These large values are likely the result of included studies having extremely large sample sizes and therefore very small standard errors; generating large heterogeneity statistics, which compared between study variation and within study variation. With such a high degree of heterogeneity evident, comparisons between diagnoses are not meaningful.

Discussion

Our systematic review demonstrated that the odds of having CUD comorbid with MIs do indeed differ by sex, and the magnitude of this effect varied by MI diagnosis. Among the CUD-only group, males were twice as likely as females to be diagnosed with CUD. Across disorders, the strongest sex effect was observed for schizophrenia, with males having nearly three times greater odds for CUD compared to females. Males were also more likely to have CUD comorbid with mood, psychotic, SUDs, anxiety, and non-specified personality disorders than females, although to a lesser extent than schizophrenia based on lower pooled ORs and non-overlapping CIs. We observed overlapping CIs between depression, psychotic disorders, and the CUD-only group, indicative of equivalent ORs. On the other hand, females demonstrated modestly greater odds of having CUD comorbid with ASPD compared to males. Conduct disorder showed no sex difference in the rate of CUD. Importantly, for MIs other than schizophrenia, psychotic disorders, and depression, females were at greater odds of having CUD compared to females without a comorbid SMI, suggesting that in females having a MI increased the likelihood of CUD, or vice versa, that having a CUD increased the likelihood of a MI.

In terms of our meta-analysis, the quantitative meta-analytic averages did not provide additional evidence of variability in the association between sex and CUD, by diagnosis, nor did the results indicate a lack of variability. Therefore, the meta-analysis results were inconclusive. Future research focused on the association between sex and CUD within specific diagnoses will benefit from understanding why there is such a high degree of variability within findings. Our investigation focused on differences between diagnoses, and therefore a deeper understanding within specific diagnoses is beyond the scope of this report.

Indeed, understanding the factors that contribute to the sex differences found in our systematic review will be essential to aid in the development of novel therapies for individuals with CUD and comorbid MIs. Below we discuss potential mechanisms involving the endocannabinoid (eCB) system and age of onset of MIs which may contribute to sex differences in the association between CUD and MI.

The endocannabinoid system

The eCB system is implicated in various regulatory functions critical for homeostasis. Homeostasis is maintained primarily through type 1 cannabinoid receptors (CB1R) that

are expressed in high concentrations in several brain regions, such as the basal ganglia, hippocampus, cerebral cortex, and cerebellum (67). Delta-9-tetrahydrocannabinol (THC), the primary psychoactive constituent of cannabis, is a CB1R partial agonist, similar to the endogenous cannabinoid ligands N-arachidonoylethanolamine (anandamide) and 2-arachidonoylglycerol (68). Sex differences in the eCB system are well documented, with sex hormones modulating eCB activity (69). The eCB system has also been hypothesized to play a role in the pathophysiology of various psychiatric disorders including mood disorders (70) and schizophrenia (71). Thus, eCB dysfunction may be associated with CUD and comorbid MIs in a sex-dependent manner.

Neurodevelopmental trajectories are sexually dimorphic with females undergoing neural maturation earlier (~ 10 years old) than males (~ 14 years old) (72). This period is characterized by dramatic changes in brain growth and connectivity orchestrated by the eCB system (73). Importantly, for males, these changes coincide with cannabis initiation given that cannabis experimentation often occurs in mid-to-late adolescence (irrespective of sex) (74). Cannabinoid-induced changes in eCB activity at this sensitive time may lead to critical neurobiological aberrations that alter brain function and behavior (75). Importantly, these changes have been implicated in the pathophysiology of schizophrenia (71). Furthermore, the same genetic variants (e.g., single nucleotide polymorphisms) that confer risk for schizophrenia in adolescence have been found to increase the risk of CUD (76). For example, evidence suggests an increased probability of early psychosis among individuals with polymorphisms in the catechol-O-methyltransferase (COMT) gene who use cannabis during adolescence (77). Notably, the prevalence of this polymorphism has been more strongly associated with the development of schizophrenia in males than females (78). Converging lines of evidence also suggest that heavier exposure to cannabis is associated with higher risk for psychotic outcomes (79,80), which is more common among males than females (81). Males are less sensitive to the effects of cannabis and thus may consume more cannabis than females in an effort to achieve similar euphoric and mood-enhancing effects (82). Because males consume more cannabis than females and at a higher frequency (81), they may be more susceptible to the negative neurobiological consequences that prime the brain for the development of an MI. This has been repeatedly shown for schizophrenia (83), and similar mechanisms may be at play for the development of depression in males (18). In contrast to males, females are less likely to consume cannabis (84), and use lower quantities of cannabis (81) and their earlier neurodevelopmental window may 'protect' them against some of the neurobiological consequences associated with cannabis exposure in adolescence. As such, sex-specific cannabis-onset and consumption patterns may underlie the greater odds of CUD and comorbid schizophrenia and psychotic disorders, and to a lesser extent depression, among males than females.

Reduced CB1R availability has been observed in several cortical and subcortical regions (e.g., hippocampus, insula) in males with schizophrenia (with no CUD/minimal cannabis exposure) compared to healthy control males (85) and in cannabis users without schizophrenia (86). CB1R downregulation may underlie increased tolerance to the effects of cannabis in heavy cannabis smokers (87), including the rewarding effects of THC (88). Thus, reduced CB1R availability may contribute to a greater likelihood of males having a CUD comorbid with schizophrenia/psychosis compared to females. However, at this point,

there are no data related to CB1R levels in female schizophrenia patients. Sex-dependent research in the area of CBIRs is fundamental to our understanding of the role of the eCB system and its activity in CUD and comorbid MIs and should be urgently addressed.

Preclinical studies demonstrate that hormonal factors, specifically estradiol, strongly influence the functioning of the eCB system (89). Estradiol regulates CB1R expression in a region-dependent manner rendering sex differences in various brain structures including the amygdala (90). Given the amygdala's role in mediating anxiety-like responses (91), higher CB1R density in the amygdala of females, compared to males, may be one mechanism contributing to their greater sensitivity to the anxiogenic effects of cannabis (90). In addition, anandamide levels are lower in females than males (92), which may reflect higher concentrations of fatty acid amide hydrolase in females (93), the metabolic enzyme responsible for degrading anandamide. Decreased whole-brain anandamide levels are predictive of anxiety-like behaviors (94) and are evident in female patients with anxiety disorders (92). Notably, lower levels of anandamide have also been documented in the cerebrospinal fluid of chronic cannabis users compared to infrequent cannabis users (95). Consistent with this reduction in anandamide levels in CUD, there is increased in vivo fatty acid amide hydrolase (FAAH, the degradative enzyme for anandamide) binding in humans with CUD measured with the PET tracer [11 C]CURB (96). Taken together, this suggests that a putative pathway may be implicated in both the development of CUD and anxiety disorders in females.

Age of onset of MIs

Conceivably, MIs that are diagnosed early in adolescence coinciding with initiation and peak cannabis use may potentiate the risk of developing comorbid CUD. Indeed, females tend to have an early onset of certain SMIs. For example, clinical studies report that the onset of generalized anxiety disorder in females occurs around age 15 (97) and increased rates of social anxiety in females are associated with transitioning to adolescence (98). Early onset of anxiety disorders may motivate and drive cannabis use in females, given their lower life experience and maturity, and lack of coping skills at this young age. As a result, females may use cannabis to manage stress and other negative symptoms associated with their primary disorder (99). While at low doses cannabis exerts anxiolytic-type effects, at high doses cannabis has the opposite effect and may exacerbate anxiety (100), leading to escalating cannabis use.

Further, adolescents with early exposure to addictive substances are also more likely to have conduct problems (101) and notably, ASPD onset is around age 13. In line with this, strong associations have been found between conduct problems in adolescence (e.g., verbal bullying (101), antisocial behavior (103), norm-violating behavior, aggression, and alcohol drinking (103)) and early cannabis initiation (101) as well as progression to daily use (103).

Thus, there is evidence that early onset of some MIs (anxiety disorders, SUDs, ASPD) in females is associated with cannabis use. Since females progress quickly from use to CUD (i.e., telescoping) (104), the likelihood of comorbidity among them is high and thus lends support for our finding of females at higher odds for CUD comorbid with selected SMIs compared to females with out SMI.

Limitations

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This systematic review and meta-analysis have several limitations. The severity of CUD may have varied between studies (e.g., mild, moderate, severe, not specified). The severity of MI may have also varied between studies given both inpatient and outpatient populations were studied (40). Studies did not assess the potency of cannabis (i.e., THC or cannabidiol content), which may have differential effects on clinical outcomes (105). Many MIs are comorbid with each other (106); thus, it is possible that individuals with CUD and comorbid MI also met criteria for another MI that was not reported (e.g., TUD). In addition, assumptions were made for 14/37 studies where we were not able to procure the sex distribution data for MI or a discrepancy existed in sample size (see Supplementary Table 1, Table 4). Finally, we were not able to quantitate sex differences in CUD across psychiatric disorders using a meta-analytic approach due to the high heterogeneity of patient characteristics between studies.

Conclusions

We found that the odds of having CUD differ by sex, and the magnitude of this effect varies by MI. We found that males were more likely than females to have CUD with all MIs, except for anxiety disorders, ASPD, and conduct disorder. Importantly, females with selected MIs were found to be at increased odds of having CUD compared to females without an MI. To date, treatments (pharmacological, behavioral, and combined) for CUD and comorbid MI have not addressed the importance of sex as a factor in disease trajectory or clinical care. Our results suggest that females may benefit from increased cannabis screening when presenting with MI. For males, psychoeducation focusing on the consequences of early and heavy use may help attenuate or at least delay use to a time when they are less neurobiologically sensitive. Our meta-analysis was inconclusive due to high levels of heterogeneity between studies, indicating that future studies should attempt to recruit more homogeneous samples with explicit inclusion and exclusion criteria. The underrepresentation of females inclinical studies may be hampering our full understanding of sex effects associated with CUD and MI. Future research needs to be mindful of potential sex differences in the rates of CUD comorbid with MIs, which may ultimately aid in developing novel, personalized, and sex-specific treatments for these comorbid disorders.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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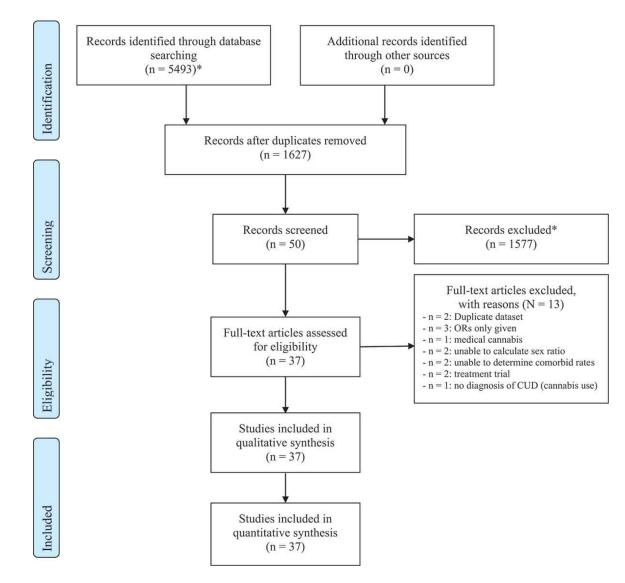


Figure 1. Study selection PRISMA flow diagram.

*Included animal studies, non-biological science studies, reviews, editorials, and commentaries

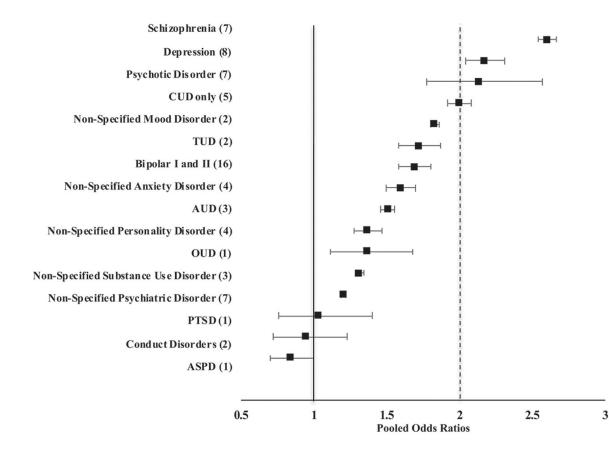


Figure 2. Forest plot of pooled ORs and 95% CIs for CUD comorbid with SMIs in males compared to females.

Numbers in brackets indicate the number of studies included in the pooled odds ratios (ORs). Studies were sorted by increasing ORs, with squares indicating ORs and horizontal lines indicating 95% confidence intervals (CIs). ASPD, antisocial disorder; AUD, alcohol use disorder; CUD, cannabis use disorder; OUD, opiate use disorder; PTSD, post-traumatic stress disorder; TUD, tobacco use disorder. Solid vertical line at 1 is the reference point where no sex difference exists. Dashed vertical line indicates the OR of the CUD-only group.

Table 1.

Odds and Pooled Odds Ratios for CUD in Samples with and without a Comorbid SMI as a Function of Sex

| CUD-only Wu et al. 2014 2.044 1.940 to 2.153 <0.000 Kerridge et al. 2018 2.322 2.034 to 2.651 <0.000 Haberstick et al. 2014 1.581 1.408 to 1.776 <0.000 Haberstick et al. 2014 1.581 1.408 to 1.776 <0.000 Non-Specified Psychiatric Disorder Lai and Sitharthan 2012 1.012 0.931 to 1.100 0.7841 Khan et al. 2013 1.735 1.609 to 1.870 <0.000 0.317 to 3.151 1.0000 Swift et al. 2001 4.072 2.967 to 5.590 <0.000 Mueser et al. 2002 2.727 0.993 to 7.491 0.0516 Mueser et al. 2002 2.727 0.993 to 7.491 0.0042 Schimmelmann et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.030 1.219 to 7.533 0.017 Psychotic Disorder Machielsen et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.030 1.219 to 7.533 0.017 Psychotic Disorder Machielsmere al. 2018 2.376 <td< th=""><th>Psychiatric Comorbidity</th><th>Study</th><th>Odds Ratio</th><th>95% CI</th><th>p =</th></td<> | Psychiatric Comorbidity | Study | Odds Ratio | 95% CI | p = |
|--|------------------------------------|--------------------------|-------------------|-----------------|------------|
| CUD-only Kerridge et al. 2018 2.322 2.034 to 2.651 <0.000 | | Stinson et al. 2006 | 2.756 | 2.312 to 3.284 | < 0.000 |
| CUD-only Farmer et al. 2016 1.761 1.256 to 2.470 0.0010 Haberstick et al. 2014 1.581 1.408 to 1.776 <0.000 | CUD-only | Wu et al. 2014 | 2.044 | 1.940 to 2.153 | < 0.000 |
| Painter et al. 2010 1.701 1.236 00 2.470 0.0000 Haberstick et al. 2014 1.581 1.408 to 1.776 <0.000 Pooled studies 1.996 1.913 to 2.083 <0.000 Non-Specified Psychiatric Disorder Lai and Sitharthan 2012 1.012 0.931 to 1.100 0.7841 Non-Specified Psychiatric Disorder Lai and Sitharthan 2012 1.000 0.317 to 3.151 1.0000 Mueser et al. 2010 4.072 2.967 to 5.590 <0.000 Mueser et al. 2000 2.254 1.387 to 3.663 0.0010 Zhu and Wu 2017 1.987 1.964 to 2.011 <0.000 Karam et al. 2002 2.727 0.993 to 7.491 0.0516 Pooled studies 1.201 1.188 to 1.215 <0.000 Karam et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.731 1.513 to 9.201 0.0042 Schimmelmann et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 2.392 1.656 to 3.455 <0.000 Rabinowitz et al. 1998 < | | Kerridge et al. 2018 | 2.322 | 2.034 to 2.651 | < 0.000 |
| Pooled studies 1.996 1.913 to 2.083 <0.000 Non-Specified Psychiatric Disorder Lai and Sitharthan 2012 1.012 0.931 to 1.100 0.7841 Khan et al. 2013 1.735 1.609 to 1.870 <0.000 | | Farmer et al. 2016 | 1.761 | 1.256 to 2.470 | 0.0010 |
| Non-Specified Psychiatric Disorder Lai and Sitharthan 2012 1.012 0.931 to 1.100 0.7844 Khan et al. 2013 1.735 1.609 to 1.870 <0.000 | | Haberstick et al. 2014 | 1.581 | 1.408 to 1.776 | < 0.000 |
| Non-Specified Psychiatric Disorder Khan et al. 2013 1.735 1.609 to 1.870 <0.000 | | Pooled studies | 1.996 | 1.913 to 2.083 | <0.000 |
| Non-Specified Psychiatric Disorder Teesson et al. 2012 1.000 0.317 to 3.151 1.000 Swift et al. 2001 4.072 2.967 to 5.590 <0.000 | | Lai and Sitharthan 2012 | 1.012 | 0.931 to 1.100 | 0.7841 |
| Non-Specified Psychiatric Disorder Swift et al. 2001 4.072 2.967 to 5.590 <0.000 Mueser et al. 2000 2.254 1.387 to 3.663 0.0010 Karam et al. 2002 2.727 0.993 to 7.491 0.0516 Pooled studies 1.201 1.188 to 1.215 <0.000 Karam et al. 2002 2.727 0.993 to 7.491 0.0516 Pooled studies 1.201 1.188 to 1.215 <0.000 Schimmelmann et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.030 1.219 to 7.533 0.0176 Brunette et al. 2018 2.376 1.428 to 3.952 0.0003 Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.822 Kavanagh et al. 2013 1.086 0.712 to 1.657 0.7012 Kavanagh et al. 2013 1.086 0.712 to 1.657 0.7012 Kavanagh et al. 2015 4.046 3.209 to 5.101 <0.000 | | Khan et al. 2013 | 1.735 | 1.609 to 1.870 | < 0.000 |
| Non-Specified Psychiatric Disorder Mueser et al. 2000 2.254 1.387 to 3.663 0.0010 Zhu and Wu 2017 1.987 1.964 to 2.011 <0.000 | | Teesson et al. 2012 | 1.000 | 0.317 to 3.151 | 1.0000 |
| Nuesel et al. 2000 2.2.34 1.53 to 3.003 0.001 Zhu and Wu 2017 1.987 1.964 to 2.011 <0.000 | | Swift et al. 2001 | 4.072 | 2.967 to 5.590 | < 0.000 |
| Karam et al. 2002 2.727 0.993 to 7.491 0.0516 Pooled studies 1.201 1.188 to 1.215 <0.0004 Machielsen et al. 2010 3.731 1.513 to 9.201 0.0044 Schimmelmann et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.030 1.219 to 7.533 0.0174 Brunette et al. 2018 2.376 1.428 to 3.952 0.0006 Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.0821 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | Non-Specified Psychiatric Disorder | Mueser et al. 2000 | 2.254 | 1.387 to 3.663 | 0.0010 |
| Pooled studies 1.201 1.188 to 1.215 <0.000 Machielsen et al. 2010 3.731 1.513 to 9.201 0.0042 Schimmelmann et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.030 1.219 to 7.533 0.0176 Brunette et al. 2018 2.376 1.428 to 3.952 0.0009 Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.0821 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | | Zhu and Wu 2017 | 1.987 | 1.964 to 2.011 | < 0.000 |
| Machielsen et al. 2010 3.731 1.513 to 9.201 0.0042 Schimmelmann et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.030 1.219 to 7.533 0.017 Brunette et al. 2018 2.376 1.428 to 3.952 0.0002 Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.0821 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | | Karam et al. 2002 | 2.727 | 0.993 to 7.491 | 0.0516 |
| Schimmelmann et al. 2012 1.717 0.773 to 3.814 0.1843 Lange et al. 2014 3.030 1.219 to 7.533 0.0176 Brunette et al. 2018 2.376 1.428 to 3.952 0.0009 Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.0821 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | | Pooled studies | 1.201 | 1.188 to 1.215 | <0.000 |
| Psychotic Disorder Lange et al. 2014 3.030 1.219 to 7.533 0.0170 Psychotic Disorder Brunette et al. 2018 2.376 1.428 to 3.952 0.0009 Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.0821 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | | Machielsen et al. 2010 | 3.731 | 1.513 to 9.201 | 0.0042 |
| Psychotic Disorder Brunette et al. 2018 2.376 1.428 to 3.952 0.0009 Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.0821 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | | Schimmelmann et al. 2012 | 1.717 | 0.773 to 3.814 | 0.1843 |
| Psychotic Disorder Rabinowitz et al. 1998 0.301 0.078 to 1.165 0.0821 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | | Lange et al. 2014 | 3.030 | 1.219 to 7.533 | 0.0170 |
| Kabinowitz et al. 1998 0.301 0.078 to 1.103 0.0021 Kavanagh et al. 2004 2.392 1.656 to 3.455 <0.000 | | Brunette et al. 2018 | 2.376 | 1.428 to 3.952 | 0.0009 |
| Khan et al. 2013 1.086 0.712 to 1.657 0.7012 Pooled studies 2.135 1.773 to 2.570 <0.000 Zhu and Wu 2017 2.676 2.611 to 2.742 <0.000 | Psychotic Disorder | Rabinowitz et al. 1998 | 0.301 | 0.078 to 1.165 | 0.0821 |
| Pooled studies 2.135 1.773 to 2.570 <0.000 Zhu and Wu 2017 2.676 2.611 to 2.742 <0.000 | | Kavanagh et al. 2004 | 2.392 | 1.656 to 3.455 | < 0.000 |
| Zhu and Wu 2017 2.676 2.611 to 2.742 <0.000 DeRosse et al. 2010 4.301 2.589 to 7.147 <0.000 | | Khan et al. 2013 | 1.086 | 0.712 to 1.657 | 0.7012 |
| DeRosse et al. 2010 4.301 2.589 to 7.147 <0.000 Nesvag et al. 2015 4.046 3.209 to 5.101 <0.000 | | Pooled studies | 2.135 | 1.773 to 2.570 | <0.000 |
| Nesvag et al. 2015 4.046 3.209 to 5.101 <0.000 Lai and Sitharthan 2012 0.9995 0.885 to 1.129 0.9939 Libuy et al. 2018 0.983 0.603 to 1.603 0.9441 Rabinowitz et al. 1998 3.317 1.708 to 6.442 0.0004 Dubertret et al. 2006 3.778 1.779 to 8.026 0.0005 Pooled studies 2.600 2.539 to 2.662 <0.0006 Non-Specified Mood Disorder Zhu and Wu 2017 1.825 1.799 to 1.852 <0.000 | | Zhu and Wu 2017 | 2.676 | 2.611 to 2.742 | < 0.000 |
| Schizophrenia Lai and Sitharthan 2012 0.9995 0.885 to 1.129 0.9935 Libuy et al. 2018 0.983 0.603 to 1.603 0.944 Rabinowitz et al. 1998 3.317 1.708 to 6.442 0.0004 Dubertret et al. 2006 3.778 1.779 to 8.026 0.0003 Pooled studies 2.600 2.539 to 2.662 <0.0003 | | DeRosse et al. 2010 | 4.301 | 2.589 to 7.147 | < 0.000 |
| Schizophrenia Libuy et al. 2018 0.983 0.603 to 1.603 0.944 Rabinowitz et al. 1998 3.317 1.708 to 6.442 0.0004 Dubertret et al. 2006 3.778 1.779 to 8.026 0.0005 Pooled studies 2.600 2.539 to 2.662 <0.0006 Non-Specified Mood Disorder Zhu and Wu 2017 1.825 1.799 to 1.852 <0.000 | Schizophrenia | Nesvag et al. 2015 | 4.046 | 3.209 to 5.101 | < 0.000 |
| Libuy et al. 2018 0.963 0.000 it 1.003 0.944 Rabinowitz et al. 1998 3.317 1.708 to 6.442 0.0004 Dubertret et al. 2006 3.778 1.779 to 8.026 0.0005 Pooled studies 2.600 2.539 to 2.662 <0.000 Non-Specified Mood Disorder Zhu and Wu 2017 1.825 1.799 to 1.852 <0.000 | | Lai and Sitharthan 2012 | 0.9995 | 0.885 to 1.129 | 0.9939 |
| Dubertret et al. 2006 3.778 1.779 to 8.026 0.0005 Pooled studies 2.600 2.539 to 2.662 <0.000 Non-Specified Mood Disorder Zhu and Wu 2017 1.825 1.799 to 1.852 <0.000 | | Libuy et al. 2018 | 0.983 | 0.603 to 1.603 | 0.9441 |
| Dubertret et al. 2006 3.778 1.779 to 8.026 0.0005 Pooled studies 2.600 2.539 to 2.662 <0.000 Non-Specified Mood Disorder Zhu and Wu 2017 1.825 1.799 to 1.852 <0.000 | | Rabinowitz et al. 1998 | 3.317 | 1.708 to 6.442 | 0.0004 |
| Non-Specified Mood Disorder Zhu and Wu 2017 1.825 1.799 to 1.852 <0.000 Mon-Specified Mood Disorder Kerridge et al. 2018 12.216 7.867 to 18.970 <0.000 | | | 3.778 | 1.779 to 8.026 | 0.0005 |
| Non-Specified Mood Disorder Kerridge et al. 2018 12.216 7.867 to 18.970 <0.000 Pooled studies 1.830 1.804 to 1.857 <0.000 | | Pooled studies | 2.600 | 2.539 to 2.662 | <0.000 |
| Pooled studies 1.830 1.804 to 1.857 <0.000 | Non-Specified Mood Disorder | Zhu and Wu 2017 | 1.825 | 1.799 to 1.852 | < 0.000 |
| | | Kerridge et al. 2018 | 12.216 | 7.867 to 18.970 | < 0.000 |
| Depression Khan et al. 2013 2.111 1.819 to 2.452 < 0.000 | | Pooled studies | 1.830 | 1.804 to 1.857 | <0.000 |
| | Depression | Khan et al. 2013 | 2.111 | 1.819 to 2.452 | < 0.000 |

| Psychiatric Comorbidity | Study | Odds Ratio | 95% CI | p = |
|------------------------------------|-------------------------|------------|-----------------|------------|
| | Rabinowitz et al. 1998 | 1.825 | 0.680 to 4.901 | 0.2327 |
| | Harder et al. 2008 | 3.363 | 2.467 to 4.586 | < 0.000 |
| | Feingold et al. 2017 | 2.150 | 1.487 to 3.110 | < 0.000 |
| | Nesvag et al. 2015 | 3.941 | 3.555 to 4.369 | < 0.000 |
| | Pacek et al. 2013 | 1.779 | 1.457 to 2.171 | < 0.000 |
| | Lai and Sitharthan 2012 | 1.000 | 0.876 to 1.140 | 0.996 |
| | Kerridge et al. 2018 | 1.291 | 1.008 to 1.655 | 0.043 |
| | Pooled studies | 2.174 | 2.046 to 2.310 | < 0.000 |
| | Khan et al. 2013 | 1.084 | 0.908 to 1.295 | 0.371 |
| | Hunt et al. 2016 | 2.206 | 1.956 to 2.489 | < 0.000 |
| | Cassidy et al. 2001 | 2.500 | 1.635 to 3.824 | < 0.000 |
| | Kawa et al. 2005 | 2.431 | 1.112 to 5.314 | 0.026 |
| | Morgan et al. 2005 | 3.886 | 1.320 to 11.430 | 0.013 |
| | Strakowski et al. 2007 | 1.631 | 0.756 to 3.518 | 0.212 |
| | VanRossum et al. 2009 | 3.017 | 2.436 to 3.736 | < 0.000 |
| | Altshuler et al. 2010 | 1.238 | 0.743 to 2.064 | 0.413 |
| Bipolar I and II | Braga et al. 2012 | 1.741 | 0.741 to 4.087 | 0.203 |
| | Lev-Ran et al. 2013 | 1.644 | 1.122 to 2.409 | 0.010 |
| | Weinstock et al. 2016 | 1.487 | 0.729 to 3.036 | 0.275 |
| | Rabinowitz et al. 1998 | 2.589 | 1.386 to 4.838 | 0.002 |
| | Kerridge et al. 2018 | 1.937 | 1.298 to 2.892 | 0.001 |
| | Nesvag et al. 2015 | 2.686 | 2.238 to 3.225 | < 0.000 |
| | Lai and Sitharthan 2012 | 0.998 | 0.827 to 1.204 | 0.981 |
| | Cotton et al. 2013 | 2.151 | 1.009 to 4.584 | 0.047 |
| | Pooled studies | 1.688 | 1.585 to 1.799 | <0.000 |
| | Khan et al. 2013 | 1.856 | 1.639 to 2.102 | < 0.000 |
| | Zhu and Wu 2017 | 1.686 | 1.549 to 1.835 | < 0.000 |
| Non-Specified Anxiety Disorder | Lai and Sitharthan 2012 | 1.007 | 0.791 to 1.283 | 0.953 |
| | Kerridge et al. 2018 | 1.476 | 1.162 to 1.875 | 0.001 |
| | Pooled studies | 1.596 | 1.497 to 1.700 | <0.000 |
| PTSD | Kerridge et al. 2018 | 1.031 | 0.762 to 1.396 | 0.841 |
| | Khan et al. 2013 | 1.749 | 1.551 to 1.972 | < 0.000 |
| Non-Specified Personality Disorder | Zhu and Wu 2017 | 1.812 | 1.523 to 2.157 | < 0.000 |
| | Lai and Sitharthan 2012 | 1.000 | 0.857 to 1.167 | 0.997 |
| | Kerridge et al. 2018 | 1.888 | 1.579 to 2.257 | <0.000 |
| | Pooled studies | 1.367 | 1.272 to 1.469 | <0.000 |
| | | | | |

| Psychiatric Comorbidity | Study | Odds Ratio | 95% CI | $\mathbf{p} =$ |
|-------------------------|----------------------|-------------------|----------------|----------------|
| Conduct Disorders | Khan et al. 2013 | 0.444 | 0.277 to 0.714 | 0.0008 |
| | Zhu and Wu 2017 | 2.357 | 1.667 to 3.332 | < 0.0001 |
| | Pooled studies | 0.941 | 0.723 to 1.224 | 0.6493 |
| Non-Specified SUD | Khan et al. 2013 | 1.188 | 1.100 to 1.284 | < 0.0001 |
| | Zhu and Wu 2017 | 1.212 | 1.186 to 1.239 | < 0.0001 |
| | Kerridge et al. 2018 | 2.330 | 2.016 to 2.694 | < 0.0001 |
| | Pooled studies | 1.311 | 1.284 to 1.338 | <0.0001 |
| OUD | Shand et al. 2011 | 1.363 | 1.108 to 1.677 | 0.0034 |
| AUD | Khan et al. 2013 | 1.299 | 1.196 to 1.411 | < 0.0001 |
| | Zhu and Wu 2017 | 1.320 | 1.278 to 1.365 | < 0.0001 |
| | Kerridge et al. 2018 | 4.097 | 3.351 to 5.009 | < 0.0001 |
| | Pooled studies | 1.505 | 1.460 to 1.551 | <0.0001 |
| TUD | Khan et al. 2013 | 1.374 | 1.244 to 1.519 | < 0.0001 |
| | Kerridge et al. 2018 | 2.253 | 1.911 to 2.656 | < 0.0001 |
| | Pooled studies | 1.720 | 1.579 to 1.873 | <0.0001 |

ASPD, antisocial disorder; AUD, alcohol use disorder; OUD, opiate use disorder; PTSD, post-traumatic stress disorder; SMI, serious mental illness; TUD, tobacco use disorder.