

---

## The serotype case-case design: a direct comparison of a novel methodology with a case-control study in a national *Salmonella* Enteritidis PT14b outbreak in England and Wales

---

D. ZENNER\*, K. JANMOHAMED, C. LANE, C. LITTLE, A. CHARLETT,  
G. K. ADAK AND D. MORGAN

Health Protection Services Colindale, Health Protection Agency, London, UK

Received 22 May 2012; Final revision 24 October 2012; Accepted 3 December 2012;  
first published online 16 January 2013

### SUMMARY

Societal and technological changes render traditional study designs less feasible for investigation of outbreaks. We compared results obtained from case-case and case-control designs during the investigation of a *Salmonella* Enteritidis PT14b (SE14b) outbreak in Britain to provide support for validation of this approach. Exposures of cases were compared to concurrent non-Enteritidis *Salmonella* cases and population controls recruited through systematic digit phone dialling. Infection with SE14b was associated with eating in oriental restaurants [odds ratio (OR) 35·8, 95% confidence interval (CI) 4·4–290·9] and consuming eggs away from home (OR 13·8, 95% CI 1·5–124·5) in the case-case study and was confirmed through a concurrent case-control study with similar effect estimates and microbiological findings of SE14b in eggs from a specific chicken flock on a Spanish farm. We found that the case-case design was feasible, quick and inexpensive, potentially minimized recall bias and made use of already interviewed cases with subtyping results. This approach has potential for use in future investigations.

**Key words:** Epidemiological methods, gastrointestinal outbreaks, study design, validation.

### INTRODUCTION

Gastrointestinal illness caused by *Salmonella enterica* can result in geographically widespread outbreaks because sales of implicated vehicles are often widely distributed [1]. Investigation of outbreaks of infection requires microbiological and/or epidemiological evidence to conduct appropriate control measures and inform public health action. The case-control method is often used in analytical studies in outbreak settings to examine the direction, strength and significance of association between exposure and outcome [2–4].

Recruiting controls for national outbreaks using systematic/random number dialling, has become increasingly difficult due to a wide range of societal and technological changes such as greater social mobility and use of mobile phones.

This means that certain demographic groups are under-represented when contacted by landline and this may introduce significant selection biases, and concern that controls may not have been at risk of exposure (e.g. less eating out) [5]. The use of General Practitioner (GP)-nominated or case-nominated controls has also become more difficult due to increased work demands on GPs and increasing public scepticism.

Alternative ways to produce valid epidemiological evidence have been explored and have included

\* Author for correspondence: Dr D. Zenner, Consultant Epidemiologist, Health Protection Services Colindale, Health Protection Agency, 61 Colindale Avenue, London NW9 5EQ, UK.  
(Email: dominik.zenner@hpa.org.uk)

case-case studies comparing severe cases with milder ones [6], or case-case studies comparing infections of different microbial subtypes [5, 7–9]. The potential advantages of a serotype case-case design have been discussed in the literature [10] and authors have trialled this method for a variety of research questions with aggregate [9] or individual-level datasets [7]. The serotype case-case method is potentially quicker and less expensive because data from cases with infectious intestinal disease are readily available as they are routinely collected with standardized surveillance questionnaires in the UK. This method also has the potential to minimize important biases, such as recall or response bias, often associated with case-control designs [5, 10]. However, there is a paucity of literature attempting to validate this approach and to our knowledge no published study has directly compared this type of analysis concurrently with conventional approaches in a national outbreak situation.

We report the results of our outbreak investigations into a national outbreak of *Salmonella* Enteritidis phage type (PT) 14b with antimicrobial resistance to nalidixic acid and reduced susceptibility to ciprofloxacin (SE14b) from 1 September to 31 December 2009 in England and Wales. The results of the case-control study had been reported previously [11]; here we directly compared the epidemiological findings from a national case-control study to that of a serotype case-case study in order to inform further use and evaluation of this alternative epidemiological tool.

## METHODOLOGY

Cases of laboratory-confirmed SE14b occurring between 1 September and 31 December 2009 in England and Wales were compared to cases of confirmed non-Enteritidis *Salmonella* during the same time period (case-case study). The results of this study were compared to those of a national case-control study of SE14b, which used the same cases and specifically recruited healthy controls by systematic digit dialling. We did not include cases with known epidemiological links to local outbreaks, but descriptive epidemiological evidence from these discrete outbreaks showed tentative links to the consumption of eggs and a national analytical study was performed to support this evidence on a national level in cases not linked to the outbreaks.

## Definitions

A *case* of SE14b was defined as a person with diarrhoea and/or vomiting, whose stool sample was positive for SE14b in England and Wales, where the sample was received by the Laboratory of Gastrointestinal Pathogens, Health Protection Agency (HPA), between 1 September and 31 December 2009.

A *control-case* was defined as a person with diarrhoea and/or vomiting, whose stool sample was positive for any non-Enteritidis *Salmonella* serotype in England and Wales, where the sample was received between 1 September and 31 December 2009.

A *control* was defined as a person with no abdominal symptoms who agreed to be interviewed as a result of systematic phone number dialling.

## Exclusions

Those who had travelled outside of the UK within 5 days prior to symptom onset or were contacts of other cases of gastrointestinal illness were excluded. Non-PT14b serotypes of *S. Enteritidis* were excluded from the study, because an *a-priori* association between this serotype with poultry and egg exposures is more common than with other serotypes [12]. Only cases without known epidemiological links to local outbreaks (i.e. associated with a local caterer) were included, those associated with 16 concurrent discrete outbreaks of SE14b were excluded from the study.

## Cases, controls and control-cases

Cases and control-cases were identified through the interrogation of standardized surveillance data, and questionnaires were completed for all confirmed cases prior to their microbiological subtyping results. Standardized amended routine questionnaires were used for cases, control-cases and controls and covered clinical and demographic information, food exposures, grocery shopping habits and animal contacts.

Controls were recruited systematically by altering the last 1–3 digits of the cases' landline telephone numbers (ascending or descending by five). They were therefore chosen from the same telephone exchange and likely to live in the same geographical area as the cases. All interviews were conducted over 5 week-day evenings between 2 October and 2 December 2009. Work numbers and faxes were excluded and all interviewees gave consent by telephone prior to interview.

For the case-case study the sample size was limited by the natural occurrence of *Salmonella* spp. infections during the observation period (about 1:1.2 ratio). The study was powered to detect an odds ratio (OR) of 3 in common exposures (25% in control-cases) and an OR of 5.5 in rarer exposures (5% in control-cases) with 80% power to be significant at the 5% level. In the case-control study, 60 cases and 120 controls were sufficient to enable the detection of an OR of 3 (for 50% of the controls exposed) to around 4 (for 10% of the controls exposed) with 90% power and 5% significance.

The results of the case-control study are reported in detail elsewhere [11] and are presented here in abbreviated format to facilitate comparison.

### Data analysis

Descriptive and statistical analyses of the data were undertaken using MS Access 2007 (Microsoft Corp., USA) and Stata v. 11 (StataCorp., USA). For both studies (case-case and case-control), association with possible exposures were explored in single variable analysis using estimated ORs and 95% confidence intervals (CIs) and  $\chi^2$  or Fisher's exact tests as appropriate. Exposures with  $P < 0.2$  in the single variable analysis were deemed eligible for inclusion in the multivariable analysis. A logistic regression model was built in forward fashion, using likelihood ratio tests (LRTs) to determine for each exposure variable whether it was significantly associated with being a case, while adjusting for the potential confounding caused by other exposures in the model (LRT,  $P < 0.05$ ). Variables, where effect modification was plausible were tested for interaction. A complete case analysis was performed.

A direct comparison between the case-control and case-case analysis was performed. To compare the estimates of significant risk factors in the two control groups, multinomial logistic regression was performed using an outcome consisting of the three categories (cases, controls, control-cases). We estimated the ratio of relative risks using a linear combination of the regression model coefficients, comparing relative exposures of control-cases and controls.

## RESULTS

### Descriptive analysis

A total of 489 SE14b cases were confirmed during the outbreak. Excluding cases associated with discrete

outbreaks ( $n = 101$ ), early non-interviewed cases and those used for the trawling exercise employed to generate hypotheses for the source of infection, a total of 81 cases were confirmed during the study period. Of these, 63 were eligible for analysis as cases. Eighteen cases were excluded due to travel history ( $n = 4$ ) or were contacts of other cases ( $n = 14$ ). There were 75 potential control-cases who fulfilled the control-case definition after exclusion of 17 who had travel histories and 14 who had indicated contact with other symptomatic cases with gastrointestinal illness in the 5 days prior to onset of their symptoms. Of the control-cases, 28 were infected with *S. Typhimurium* and 47 with other *Salmonella enterica* serotypes.

An average of 3.6 telephone calls (range 1–32) were made to recruit one control. A total of 108 controls were enrolled. Controls were more likely to be female (67% vs. 44%,  $P = 0.004$ ) and older compared to cases (average age 52.5 vs. 36.8 years, respectively,  $P < 0.0001$ ) and control-cases (Table 1). Control-cases were much more similar to cases than controls; age and sex were found to be significant confounders in the case-control study, but not in the case-case study.

Onset dates of cases were between 1 September and 16 November 2009. The average duration of illness was 6 days and the predominant symptoms were diarrhoea, abdominal pain, fever, nausea, headaches, and to a lesser degree vomiting. About 27% of the cases reported blood in their stool and 19% were admitted to hospital (Table 1). No deaths were reported in the study cases. The clinical and epidemiological details have been reported in more detail elsewhere [11].

### The case-case study

Single variable analysis of the case-case study demonstrated a strong association between eating out and symptomatic infection with SE14b ( $P = 0.001$ ), particularly in oriental restaurants ( $P < 0.0001$ ), Indian restaurants ( $P = 0.026$ ) and Italian restaurants ( $P = 0.037$ ). Exposure to fish ( $P = 0.006$ ), salads ( $P = 0.006$ ), eggs ( $P = 0.007$ ), pre-prepared sandwiches ( $P = 0.007$ ), cold meats ( $P = 0.012$ ), fruits ( $P = 0.017$ ), vegetables ( $P = 0.025$ ) and bacon ( $P = 0.036$ ) were also significantly associated with SE14b infection in the single variable analysis (Table 2). There was no strong evidence of association between SE14b infection and

Table 1. *Basic demographics for cases, control-cases and controls*

	Cases	Control-cases	Controls
Sex			
Female	28 (44.4)	39 (52.0)	72 (66.7)
Male	35 (55.6)	36 (48.0)	36 (33.3)
Age group (years)			
< 10	9 (14.3)	23 (30.7)	2 (1.9)
10–29	19 (30.2)	10 (13.3)	8 (7.4)
30–49	14 (22.2)	18 (24.0)	38 (35.2)
50–69	11 (17.5)	14 (18.7)	38 (35.2)
≥ 70	10 (15.9)	10 (13.3)	22 (20.4)
<b>Totals</b>	<b>63</b>	<b>75</b>	<b>108</b>
Symptoms			
Diarrhoea	59 (93.6)	73 (97.3)	—
Bloody stools	15 (23.8)	21 (28.0)	—
Vomiting	20 (31.7)	33 (44.0)	—
Fever	32 (50.8)	44 (58.7)	—
Abdominal pain	49 (77.8)	57 (76.0)	—
Nausea	29 (46.0)	32 (42.7)	—
Headaches	26 (41.3)	30 (40.0)	—
Hospital admission	12 (19.0)	21 (28.0)	—

Values given are *n* (%).

age ( $P=0.07$ ) or sex ( $P=0.4$ ) in the single variable analysis.

The multivariable analysis showed a significant association between eating out (OR 3.4, 95% CI 1.4–8.5,  $P=0.006$ ) and consuming eggs outside the home (OR 5.2, 95% CI 1.1–25.3,  $P=0.02$ ), and this exposure occurred in 83% of the cases. However eating out can be entirely explained by eating in oriental restaurants in a three-level model, and hence the final multivariate model of the case-case study shows significant associations between symptomatic infection with SE14b and eating in oriental restaurants (OR 35.8, 95% CI 4.4–290.9,  $P<0.0001$ ) and consuming eggs away from home (OR 13.8, 95% CI 1.5–124.5,  $P=0.005$ , Table 3*a*).

Eating vegetables at home was associated with infection in the logistic regression model; however, further analysis demonstrated that this variable summarized a wide variety of different vegetables from different locations. The variable was not considered epidemiologically relevant and excluded from further analysis. Age and sex were examined for inclusion in the multivariable regression analysis, but not found to be significant (LRT  $P=0.96$  for age and  $P=0.58$  for sex) and the magnitude of the associations with exposures included in the model did not alter when these were included. Interactions were not found.

#### Comparison with the case-control study

Multivariable analysis of the case-control study (Table 3*b*) also demonstrated a significant association between eating out and becoming a case (OR 2.7, 95% CI 1.1–6.9,  $P=0.03$ ), particularly eating in an oriental restaurant ( $P=0.005$ ), and the significant association of eating out can be entirely explained by eating in oriental restaurants. Among food exposures, eating eggs away from home ( $P=0.011$ ) and vegetarian foods eaten away from home ( $P=0.002$ ) were identified as significant risk factors for becoming a case. Associations between infection and eating cold meats away from home or barbecued foods at home were not included in the final logistic regression models, based on lack of explanatory power (non-significant LRT). Results were adjusted for age and sex.

The results of multivariable analyses of case-case and case-control studies gave different effect estimates, because persons in the control group were less likely to eat at oriental restaurants compared with the control-cases. Although age and sex are included in this model, the disparity between the cases and controls in these variables could result in residual confounding due to unmeasured confounders that impact on the controls eating habits. However, a direct comparison of key exposure variables did not

Table 2. *Single variable analysis of exposure variables in the case-case study*

Exposure	OR	95% CI	P value
Any eating out	3.52	1.6–7.7	0.001
Eating out at parties	2.01	0.75–5.4	0.172
Eating out at			
Oriental restaurants	36.64	5.85–∞	<0.001
Indian restaurants	5.39	1.19–∞	0.026
Italian restaurants	7.5	1.1–∞	0.037
Public houses	1.21	0.33–4.47	0.785
Kebab houses	1.16	0.38–3.59	0.797
Fried chicken bars	0.93	0.4–9.2	0.941
Burger bars	1.03	0.32–3.31	0.968
Food exposures			
Fish	2.93	1.36–6.31	0.006
Salads	2.91	1.35–6.29	0.006
Eggs	2.92	1.33–6.44	0.007
Pre-prepared sandwiches	3.37	1.39–8.16	0.007
Cold meats	2.64	1.23–5.67	0.012
Fruit	2.81	1.2–6.59	0.017
Vegetables	2.95	1.14–7.58	0.025
Bacon	2.33	1.05–5.13	0.036
Barbecued food	6.44	0.95–∞	0.058
Beef	2.07	0.98–4.38	0.058
Vegetarian food	2.52	0.93–6.84	0.071
Cakes and desserts	2.02	0.9–4.55	0.089
Game birds	—	0.62–∞	0.121
Pate	4.87	0.7–∞	0.127
Lamb	2.15	0.78–5.89	0.142
Other poultry	2.68	0.69–10.36	0.167
Shellfish	2.29	0.67–7.82	0.199
Pork	1.42	0.64–3.17	0.393
Cheese	1.37	0.63–2.98	0.431
Chicken	1.22	0.55–2.69	0.634
Sausages	1.18	0.55–2.51	0.672
Halal meat	0.6	0.4–7.9	0.682
Offal	1.19	0.2–6.99	0.866
Visiting farms	3.4	0.72–∞	0.133
Contact with animals	0.92	0.47–1.82	0.813
Living on farm/smallholding	1.28	0–∞	0.862

OR, Odds ratio; CI, confidence interval.

Food item categories are ordered in order of statistical significance. Calculations based on small cell values (e.g. zero) were not estimated here (e.g. indefinites). This is denoted by a symbol ( $\infty$ ).

show any significant differences for eating at oriental restaurants (OR 6.1, 95% CI 0.8–48.2,  $P=0.09$ ) or eating eggs away from home (OR 1.5, 95% CI 0.2–14.5,  $P=0.7$ ) between persons in control or control-case groups (Table 4). This analysis demonstrates that eating habits of control-cases were more similar to cases compared to eating habits of controls, and this is in keeping with demographic similarities.

## DISCUSSION

### Findings of this study

To our knowledge this is the first study which directly compared case-case with case-control designs for a large national outbreak investigation. We showed the feasibility of the case-case method in an outbreak situation and its comparability with the conventional case-control method.

Table 3. *Multivariable models of (a) case-case analysis and (b) case-control studies. These are two separate models*

	OR	95% CI	P value (Wald test)	P value (LRT)
<i>(a) Case-case study</i>				
Eaten foods from oriental restaurants	35.78	4.40–290.91	0.001	<0.0001
Eating eggs away from home	13.80	1.53–124.47	0.02	0.005
<i>(b) Case-control study</i>				
Eaten foods from oriental restaurants	4.45	1.57–12.64	0.005	0.005
Eating eggs away from home	7.66	1.58–37.05	0.01	0.008
Eating vegetarian food away from home	21.28	2.95–153.76	0.002	0.002
Age (modelled continuous)	0.96	0.94–0.98	0.001	0.0003
Sex (male vs. female)	2.95	1.17–7.46	0.02	0.02

OR, Odds ratio; CI, confidence interval; LRT, likelihood ratio test.

The results of the case-control analysis are adjusted for age and sex. Age and sex were not found to be significant confounders in the case-case study and hence not included.

Table 4. *Analysis of key exposure variables, comparing effects in the case-control and case-case study (unadjusted for age and sex)*

	OR	95% CI	P value
<i>Case-control study</i>			
Eaten foods from oriental restaurants	5.54	2.21–13.89	<0.00001
Eating eggs away from home	8.25	1.96–34.68	0.004
<i>Case-case study</i>			
Eaten foods from oriental restaurants	35.78	4.40–290.91	0.001
Eating eggs away from home	13.80	1.53–124.47	0.02
<i>Comparing exposures of controls and control-cases</i>			
Eaten foods from oriental restaurants	6.06	0.76–48.25	0.09
Eating eggs away from home	1.46	0.15–14.47	0.7

Exposures in the control group were compared to those in the control-cases using multinomial regression and estimating the ratio of relative risks using a linear combination of the regression model coefficients. The comparison study analyses relative exposures of controls and control-cases and the base group are the controls.

Both studies showed significant associations between symptomatic infection with SE14b and eating out in oriental restaurants as well as eating eggs away from home, albeit with non-significantly different effect sizes. These findings are entirely consistent with the microbiological and environmental investigations of this outbreak, which identified the same strain in eggs sourced from a specific chicken flock in a particular farm in Spain, and there were numerous outbreaks of infection mainly associated with oriental restaurants. Our study supports the findings of an increased gastrointestinal risk from pooling eggs for multiple uses [13]. As a result of these investigations, and the resulting multi-agency public health action,

eggs from this flock were taken out of circulation and the outbreak was contained.

### Comparison with other studies

The case-case design has been discussed methodologically [5, 10] and employed for the study of risk factors using surveillance data, covering a wide range of topics [7–9, 14, 15]. Comparisons between the case-case method and more established methodologies have rarely been reported. One study compared the results from a case-case with a case-control design in a small local outbreak and found compatible results in both studies, but a direct comparison was not possible

as the exposure data were different in the studies and the lack of statistical power precluded multi-variable analysis [8]. Another study investigated a regional increase of *S. Enteritidis* infections in a 3-year period using non-*Enteritidis* serotypes and healthy individuals as controls [16]. They identified eggs as a generic risk factor for *S. Enteritidis* infection during this period. However, the length of the observation period and the lack of microbiological phage typing and molecular profiling do not allow a great level of detail on risk exposures and it is possible that their data captured more than one outbreak and more than one source. Our study used a case-case as well as the case-control design in a parallel investigation as a 'field testing' exercise in a national outbreak situation.

### Strengths and limitations

We found that the use of readily available data in the case-case design is advantageous; because it reduces costs and increases the speed of outbreak investigations (cases are routinely interviewed). In addition this design has a number of methodological advantages, some of which have been discussed already [5, 7, 10, 16]. Recall bias is a major issue in case-control designs; we believe that the potential for recall is reduced [10], because cases and control-cases both had symptoms and were interviewed contemporaneously with their illness and before serotyping results became available. It is possible that response bias might be of lesser concern compared to conventional case-control studies, since the (unknown) serotype is unlikely to influence the decision to participate in interviews. In this study we showed that the age and sex profiles were more similar to the control-cases compared to controls recruited via systematic digit telephone dialling.

By using the same telephone exchange for cases, controls were chosen from the same geographical area in the case-control study; the same was not possible for the case-case design. However, in areas where cases were clustered, greater awareness may have led to a greater proportion of control-cases being interviewed and this may have led to some geographical clustering of control-cases, even without matching. In general, geographical clustering is beneficial, because it makes geographically clustered exposures (e.g. using the same supermarket) more similar, but less important in our study as we investigated widely distributed food items.

Control-cases were chosen, because they became infected with similar, albeit different microbial subtypes contemporaneously, and the case-case design is therefore limited to pathogens where different strain types are associated with different exposures, as is the case with *Salmonella* spp. Although recorded exposure prevalences of control-cases can be larger than in the general population at risk [7], this is more relevant for surveillance studies and less for outbreak investigations. However there is the potential for underestimating effect sizes for common exposures (between cases and control-cases) [5, 7]. *S. Enteritidis* infection is frequently associated with eggs and poultry, and inclusion of *S. Enteritidis* infection as control-cases would have prevented the findings of a significant association between eggs and illness in this study design as a likely common exposure [12]. The similarity of the results compared to the case-control study (non-significant differences of effect estimates), and the positive outbreak microbiology add to the validity of our findings for this particular outbreak investigation. Contrary to earlier studies [16] we also used microbial subtyping data (e.g. phage typing and resistance profiling) to delineate case and control-case definitions and to develop an appropriate study design. The method could be used in other pathogen outbreaks, providing the control-case group can be defined clearly and there is no known association between the control pathogen and the exposure of interest.

The number of control-cases was limited by the natural occurrence of *Salmonella* infections during our observation period. It is possible that the association between infection and other rarer exposures may have been missed. However, rarer exposures are unlikely to be relevant in our study and the evidence of this study is supported by the large effect sizes for the main exposures in the case-case design, even when compared to the case-control design and the other epidemiological and microbiological findings discussed above.

### CONCLUSION

In conclusion we present a direct comparison of a case-case design and case-control design, with similar results observed in the two study designs. Acknowledging the limitations of a case-case approach, our study is an important step to fully validate this study design. It is often more timely and requires less additional resources, because the information on

control-cases is often readily available as part of routine enhanced surveillance in many countries including the UK. The potential reduction in both selection and recall biases using contemporaneous case-controls provides an additional advantage over traditional case-control study designs. The serotype case-case design could therefore significantly augment the epidemiological outbreak investigation study designs available to the field epidemiologist.

## ACKNOWLEDGEMENTS

We acknowledge the help and support of Elisabeth de Pinna, Head of the Salmonella Reference Unit at the Laboratory of Gastrointestinal Pathogens, Microbiology Services Colindale, Health Protection Agency, UK.

## DECLARATION OF INTEREST

None.

## REFERENCES

1. **Behravesh CB, et al.** Multistate outbreak of Salmonella serotype Typhimurium infections associated with consumption of restaurant tomatoes, USA, 2006: hypothesis generation through case exposures in multiple restaurant clusters. *Epidemiology and Infection* 2012; **140**: 2053–2061.
2. **Fonseca MG, Armenian HK.** Use of the case-control method in outbreak investigations. *American Journal of Epidemiology* 1991; **133**: 748–752.
3. **Palmer SR.** Outbreak investigation: the need for 'quick and clean' epidemiology. *International Journal of Epidemiology* 1995; **24** (Suppl. 1): S34–S38.
4. **Salmon RL, et al.** How is the source of food poisoning outbreaks established? The example of three consecutive Salmonella Enteritidis PT4 outbreaks linked to eggs. *Journal of Epidemiology and Community Health* 1991; **45**: 266–269.
5. **McCarthy N, Giesecke J.** Case-case comparisons to study causation of common infectious diseases. *International Journal of Epidemiology* 1999; **28**: 764–768.
6. **Giraudon I, et al.** Large outbreak of Salmonella phage type 1 infection with high infection rate and severe illness associated with fast food premises. *Public Health* 2009; **123**: 444–447.
7. **Gillespie IA, et al.** A case-case comparison of *Campylobacter coli* and *Campylobacter jejuni* infection: a tool for generating hypotheses. *Emerging Infectious Diseases* 2002; **8**: 937–942.
8. **Krumkamp R, Reintjes R, Dirksen-Fischer M.** Case-case study of a Salmonella outbreak: an epidemiologic method to analyse surveillance data. *International Journal of Hygiene and Environmental Health* 2008; **211**: 163–167.
9. **Wilson N, et al.** Case-case analysis of enteric diseases with routine surveillance data: potential use and example results. *Epidemiologic Perspectives & Innovations* 2008; **5**: 6.
10. **Rosenbaum PR.** Attributable effects in case 2 studies. *Biometrics* 2005; **61**: 246–253.
11. **Janmohamed K, et al.** National outbreak of Salmonella Enteritidis phage type 14b in England, September to December 2009: case-control study. *Eurosurveillance* 2011; **16**.
12. **Gantois I, et al.** Mechanisms of egg contamination by Salmonella Enteritidis. *FEMS Microbiology Reviews* 2009; **33**: 718–738.
13. **Gormley FJ, Rawal N, Little CL.** Choose your menu wisely: cuisine-associated food-poisoning risks in restaurants in England and Wales. *Epidemiology and Infection* 2011; 1–11.
14. **de Valk H, et al.** Two consecutive nationwide outbreaks of listeriosis in France, October 1999–February 2000. *American Journal of Epidemiology* 2001; **154**: 944–950.
15. **Zock JP, et al.** Asthma Characteristics in cleaning workers, workers in other risk jobs and office workers. *European Respiratory Journal* 2002; **20**: 679–685.
16. **Kist MJ, Freitag S.** Serovar specific risk factors and clinical features of Salmonella Enterica ssp. *Enterica* serovar Enteritidis: a study in South-West Germany. *Epidemiology and Infection* 2000; **124**: 383–392.