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## Methylation subtypes of primary prostate cancer predict poor prognosis

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### Abstract

**Background**—Prostate cancer patients experience heterogeneous outcomes after radical prostatectomy. Genomic studies including The Cancer Genome Atlas (TCGA) have reported molecular signatures of prostate cancer, but few studies have assessed the prognostic effects of DNA methylation profiles.

**Methods**—We conducted the largest methylome subtyping analysis for primary prostate tumors to date, using methylome data from three patient populations: TCGA, a prostate cancer cohort study conducted at the Fred Hutchinson Cancer Research Center (FH), and the Canadian

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International Cancer Genome Consortium (ICGC) cohort. Four subtypes were detected in the TCGA dataset, then independently assigned to FH and ICGC cohort data. The identified methylation subtypes were assessed for association with cancer prognosis in the above three patient populations.

**Results**—Using a set of hypermethylated CpG sites, four methylation subtypes were identified in TCGA. Compared to Subtype 1, Subtype 4 had a hazard ratio (HR) of 2.09 ( $p=0.029$ ) for biochemical recurrence (BCR) in TCGA patients. HRs of 2.76 ( $p=0.002$ ) for recurrence and 9.73 ( $p=0.002$ ) for metastatic-lethal (metastasis or prostate cancer-specific death) outcomes were observed in the FH cohort. A similar pattern of association was noted in the Canadian ICGC cohort, though HRs were not statistically significant.

**Conclusions**—A hypermethylated subtype was associated with an increased hazard of recurrence and mortality in three studies with prostate tumor methylome data. Further molecular work is needed to understand the effect of methylation subtypes on cancer prognosis.

**Impact**—This study identified a DNA methylation subtype that was associated with worse prostate cancer prognosis after radical prostatectomy.

### Keywords

DNA methylation; molecular subtypes; prostate cancer; prognostic markers

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## Introduction

Prostate cancer accounts for over a quarter of all incident cancer diagnoses in U.S. men and ranks second as a cause of cancer-related deaths (1). The majority of prostate cancer patients present with localized disease at diagnosis. Some have indolent disease that can be cured or safely observed under active surveillance, while others may have highly variable clinical outcomes after radiotherapy or surgery – a portion of these will eventually relapse, with some developing metastasis that ultimately leads to prostate-cancer specific death. A critical unmet clinical need is to identify aggressive prostate cancers that confer a high risk for relapse and metastasis after primary treatment.

Existing risk stratification systems using clinical factors (Gleason score, prostate-specific antigen (PSA) levels, and TNM stage) have not been able to completely predict adverse outcomes (2–4). Over the past decade both genetic and genomic efforts have characterized molecular events contributing to prostate cancer initiation and progression (5–9). Somatic mutations in prostate cancer are less common than in other adult solid tumors (5). A varying degree of copy number aberrations has been observed, with higher burdens of aberrations and instability in aggressive cancers (6,10–12). The most common molecular alteration is the *TMPRSS2-ERG* gene fusion (13), though its clinical implication is not clear.

In addition to genetic alterations, epigenetic alterations play a crucial role in prostate cancer development. To date, most studies of DNA methylation and prostate cancer progression have been limited to candidate genes in relation to biochemical recurrence (14–19). More recently, genome-wide methylation studies for discovering prognostic markers have been conducted by arrays or pyrosequencing methods (20,21). Among recent studies,

TCGA confirmed substantial heterogeneity in DNA methylation and detected widespread changes in hyper- or hypo-methylation in 333 primary prostate tumors when compared to adjacent benign tissue obtained from radical prostatectomies (5,22,23). Using a set of hypermethylated CpG sites, TCGA defined four distinct epigenetic groups, one of which was exclusively associated with *ERG* fusion positive tumors. The clinical implications of these epigenetic subgroups have yet to be determined, in part due to the short follow-up for patients in TCGA not yielding an adequate number of adverse outcomes. This limitation also applies to characterizing the clinical significance of other molecular subtypes (24), since a disease with a long natural history like prostate cancer generally requires a long follow-up period to observe recurrence and mortality endpoints.

To delineate epigenetic subtypes and determine their clinical significance, we analyzed methylome data from three large primary prostate cancer patient populations who underwent radical prostatectomy (RP), all with clinical information for evaluating outcomes. We performed unsupervised clustering in one of the largest fresh-frozen prostate cancer methylome datasets currently available - the TCGA-PRAD, which includes ~500 samples; and we characterized the risk of recurrence associated with the derived subtypes using the recently curated TCGA clinical data (25). We then applied the classification rule to the other two tumor datasets and tested the resulting subtype associations with clinical factors, recurrence, and mortality.

## Materials and Methods

### Study participants and data collection

This study includes patients from three cohorts who had localized prostate cancer and treated with RP. CONSORT diagrams for these three cohorts are shown in Supplementary Figure 1.

**TCGA-PRAD**—The Cancer Genome Atlas Prostate Adenocarcinoma (TCGA-PRAD) dataset includes matched primary tumor and normal (adjacent) specimens from prostate cancer patients who underwent RP as primary therapy (5). All patients provided written informed consent and specimens were approved for collection and distribution by local Institutional Review boards (5). DNA methylation data generated using the Infinium HumanMethylation450 BeadChip (Illumina, San Diego, CA, USA) and RNA-seq data from freshly frozen tumor tissue were downloaded using the *TCGAbiolinks* R package (v3.1.4) (26); and the HTSeq count gene expression data was  $\log_2(\text{count} + 1)$  transformed. Tissue images were centrally reviewed by genitourinary pathologists to determine Gleason score. For clinical data, we used the TCGA Pan-Cancer Clinical Data Resource, which includes a progression-free interval (PFI) outcome (27). Prognosis outcomes after RP include non-recurrence (n=387), recurrence only (n=80), and the metastatic-lethal event, that a patient either had metastasis or died from prostate cancer (n=8). There were 475 tumors from patients of Caucasian, African American, and Asian ancestry that had methylation and clinical information available for this analysis.

**Fred Hutch cohort**—The Fred Hutch (FH) cohort has been previously described in detail (26,27). Briefly, it includes patients with newly diagnosed prostate cancer in 1993–1996

and 2002–2005 who enrolled in population-based studies in King County, WA, and who were identified using the Seattle-Puget Sound Surveillance, Epidemiology, and End Results (SEER) cancer registry. Our study was restricted to 458 men who had RP as the primary therapy for clinically localized adenocarcinoma of the prostate and had genome-wide methylation data and clinical data available. Gleason grade, diagnostic PSA and pathologic tumor stage were collected and centrally coded by Puget Sound SEER (20,30). Pathological stage data were used to define two stage groups: localized stage (pT2, N0/NX, M0); regional stage (pT3–pT4 and/or N1, M0). Time to recurrence or censoring was determined using data from follow-up surveys and review of medical records (n=445). There were 13 patients who did not complete surveys or died of prostate cancer (n=13) prior to mailing of the surveys. Their recurrence time was imputed by a linear regression model including curative therapy and Gleason score and the model was trained to account for the fact that BCR occurs before metastasis or death from prostate cancer. Prognosis outcomes after RP include non-recurrence (n=315), recurrence only (n=114), and metastatic-lethal events (n=29). DNA methylation data were measured using the Infinium HumanMethylation450 BeadChip (Illumina, San Diego, CA, USA) (29) on tumor samples from FFPE blocks and have been deposited to dbGaP (Accession: phs001921.v1.p1). More information about clinical and genomics data collection and quality control procedures have been described previously (20,30). The Fred Hutchinson Cancer Research Center Institutional Review Board approved the study, and all participants provided written informed consent statements.

**Canadian ICGC cohort**—The Canadian ICGC cohort includes intermediate-risk patients with pathologically confirmed localized prostate cancer as described previously (23). All patients (n=236) were defined as N0M0 and treated with RP. Fresh-frozen RP samples were stored at the University Health Network (UHN) Pathology BioBank and the Centre Hospitalier Universitaire de Quebec-Universit Laval (CHUQ) Genito-Urinary BioBank prior to analysis. All participants provided written informed consent statements, and tumor tissues were utilized based on UHN Research Ethics Board-approved study protocols (UHN 06–0822-CE, UHN 11–0024-CE and CHUQ 2012–913:H12–03–192). Prognosis outcomes after RP include non-recurrence (n=171), recurrence only (n=42), and metastatic-lethal events (n=23). Gleason scores were evaluated by two genitourinary pathologists using hematoxylin and eosin-stained slides (23). Serum PSA levels were measured at the time of diagnosis.

### DNA methylation assays and pre-processing

Level 3 TCGA-PRAD data contain  $\beta$ -values, where the  $\beta$ -value is the ratio of the methylated signal over the total signal. Probes that had a common SNP within 10 bp of the interrogated CpG site, were located within 15bp of a repetitive element, aligned to multiple sites on the human genome, or had detection p-values greater than 0.05 for a specific data point were masked (5).

Methylome data from the FH cohort were profiled using the Infinium HumanMethylation450 BeadChip (Illumina, San Diego, CA, USA). We excluded four low quality samples that clustered together with low signal intensities. None of the samples was found to have at least 5% of probes with a detection p-value greater than 0.05. We also excluded non-CpG sites and CpG sites that had a detection p-value greater than 0.05 in

at least 10% of samples, CpG sites in any SNP or within 10 base pairs of a SNP with a minor allele frequency greater than 1% in any 1000 Genomes population, or CpG sites classified as cross-reactive probes. We then performed background correction and applied Noob normalization method (31) on the methylation data.

Global DNA methylation for the Canadian ICGC cohort was evaluated using Illumina Infinium HumanMethylation450 BeadChip kits as described previously (23). The raw methylation microarray data were preprocessed using the *dasen* method from the R package *wateRmelon* (32). Methylation intensities were filtered based on the detectability over noise level. Probe positions and chromosome locations were annotated using the *IlluminaHumanMethylation450kanno.ilmn12.hg19* package (version 0.6.0). All methylation data processing was performed using the R language (version 4.0.5) (33).

### Tumor clustering

The TCGA tumors were clustered based on the following criteria for selecting informative hypermethylated CpG sites: 1) having low mean methylation in adjacent benign samples  $<0.1$ ; 2) uncorrelated with ABSOLUTE (34) tumor purity estimates based on copy number changes; 3) having  $\beta$ -values greater than 0.3 in at least 5% of tumors, a criterion commonly used to detect reliable methylation sites and guard against measurement errors in methylation array; 4) among the top 5,000 most variable probes in tumors; 5) showing differential variability between ~500 primary prostate tumors and 50 adjacent benign prostate samples by the Levene test (with a family-wise error rate  $<0.05$  by Bonferroni correction); and 6) mean methylation  $<0.1$  in adjacent benign samples. These criteria were motivated by the TCGA clustering analysis (5), but made more stringent by requiring lower methylation in adjacent benign samples and significantly increased variability in tumor samples. The final set contained 1306 CpG sites (Supplementary Table 1). Hierarchical clustering was performed using the function *eclust* in the R package *factoextra* (v1.0.7) (35) based on these CpGs with dichotomized  $\beta$ -values ( $>0.3$  vs.  $0.3$ ). Binary distance metric and the agglomeration method of “ward.D2” (36) were used when performing hierarchical clustering. In the FH and Canadian ICGC studies, dichotomized  $\beta$ -values for the same set of probes were used.

### Statistical analysis

CpGs with different variability between tumor samples and benign samples were detected by the Levene method implemented in R package *Lawstat* (v3.4) (37). Hierarchical clustering of TCGA samples was performed by the R package *factoextra* (v1.0.7) (35). The heatmap of TCGA methylation data was visualized using the R package *ComplexHeatmap* (v2.6.2) (38). The two-dimensional t-distributed stochastic neighbor embedding (tsne) plot was visualized using the R package *Rtsne* (v0.15) (39). Assignment of validation samples (FH and ICGC) to the subtypes was determined by the k-Nearest Neighbor Classification (KNN) method, in which  $k=21$  was determined by the square root of the number of complete cases, which is rounded to the nearest odd integer (40). The binary distance metric was used for cluster assignment. In survival analysis, deaths due to other causes than prostate cancer have been accounted for using the cumulative incidence function plots and cause-specific hazard regression (41,42). Cause-specific Cox proportional hazards regression was used

to compute hazard ratios (HRs), 95% confidence intervals (CIs), and p-values. Cumulative incidence functions (CIFs) were plotted to evaluate associations between subtypes and prostate cancer recurrence or metastatic-lethal events, as implemented in the R package *cmprsk* (v2.2–10) (43). Pathway analysis was done by the *gometh* function of the R package *missMethyl* (v1.24.0) (44). Logistic regression models were used to construct the prediction models with predictors including age, race, Gleason score, methylation subtypes, diagnostic PSA level, and tumor stage. The area under curve measure (AUC) of the receiver operating characteristic (ROC) performance measure was computed for prediction models. The ROC analysis was carried out using the R package *pROC* (v1.17.0.1) (45).

## Results

### Characteristics of primary cancer patients in the three studies

Table 1 shows the demographic and clinical characteristics of the three primary prostate cancer datasets analyzed for methylation subtypes. Prostate cancer patients from TCGA were older and had more aggressive disease with a higher PSA, a higher tumor grade and a higher T category when compared to FH, and Canadian ICGC studies. On the other hand, the ICGC cohort has fewer patients with Gleason 6 and fewer patients with PSA 20.

### Clustering TCGA primary prostate tumor samples by cancer-specific methylation

CpGs with increased variability in TCGA primary tumors relative to adjacent benign samples were selected for clustering analysis. Among 216,605 CpGs passing quality control filters, the volcano plot shows that over 95% of the differential variability CpGs are hypervariable, and 20,446 CpGs show statistically significant increased variability (family wise error rate <0.05 for Levene test, red dots in Figure 1A). Aforementioned filters were applied to identify cancer-specific methylation and to limit the impact of tumor purity on clustering. The final set of 1,306 CpGs (Supplementary Table 1) was used in hierarchical clustering. The vast majority of these CpGs do not fall into copy number alteration (CNA) regions previously reported for the prostate cancer genome (6,11,12), therefore the subtypes defined by these CpGs are unlikely to reflect prostate cancer genomic instability due to copy number changes. Four subtypes were identified using the selected 1,306 CpGs (Figure 1b); mean beta values were 0.128, 0.127, 0.277, 0.249, respectively, from Subtype 1 to Subtype 4. The main differences are between Subtype 1–2 versus 3–4. Subtypes 1 and 2 had relatively smaller methylation differences with a subset of CpGs showing differences between these groups (Supplementary Table 1). A similar pattern was seen for Subtypes 3 and 4. In Figure 1c, the two-dimensional tSNE plot shows that samples in Subtype 1 are surrounded by the other three subtypes; there is a clear separation of Subtypes 3 and 4 relative to Subtypes 1 and 2, with the latter two subtypes slightly overlapping. The heatmaps of assigned subtypes for FH cohort and Canadian ICGC cohort are shown in Supplementary Figure 2.

Table 2 shows the top 20 CpGs that were associated with the 4 Subtypes using the ANOVA test. Nearly all of them were linked to prostate cancer development in the literature, including *ESR1* – a gene encoding estrogen receptor 1. Most of them had low methylation in Subtypes 1 and 2, but hypermethylation in Subtypes 3 and 4, consistent with the hierarchical

clustering process. Notably, pathway analysis of the 1,306 CpGs showed that the leading enriched pathway is “neuroactive ligand-receptor interaction”, which was previously identified by gene expression data in prostate cancer (46).

Tables 3 and 4 show the characteristics of the four subtypes in TCGA patients and FH patients. Consistently in both study cohorts, Subtypes 3 and 4 are more likely to include older men, men having higher Gleason scores, higher PSA levels, and presenting with later-stage disease at diagnosis (Tables 3 and 4). Consistent with the previous paper clustering TCGA-PRAD samples (5), there is a cluster with a substantially higher proportion of patients with the *TMPRSS2-ERG* fusion transcript (Subtype 2, 88%). Subtype 3 also has a higher portion of patients with the fusion transcript than Subtypes 1 and 4, with Subtype 4 having the lowest frequency of the fusion transcript (6% in TCGA and 11% in FH).

### **Methylation subtypes predict recurrence and metastasis outcomes**

The association of the four methylation subtypes with cancer recurrence was assessed using cause-specific proportional hazards models (Table 5 and Supplementary Table 2). Patients with metastasis or prostate cancer death were also compared to patients without recurrence whenever such data were available. A base (minimally adjusted) model was first assessed with age and race to predict BCR; Gleason score was additionally adjusted for to evaluate whether the subtypes were an independent prognostic factor (Supplementary Table 2). In TCGA data, Subtypes 2–4 showed increasing hazards for cancer progression relative to Subtype 1, reaching a hazard ratio of 2.09 (p-value=0.029) for Subtype 4 in the base model (Table 5). Adjusting for Gleason score, stage and PSA increased p-values for the associations, to a different degree for the four subtypes in TCGA, e.g., the p-value for Subtype 4 increased to 0.28 (HR=1.47). The association between methylation subtypes and the metastatic-lethal outcome was not assessed because there is a limited number of TCGA patients developed metastatic-lethal event (n=8).

Using the KNN method, samples from independent datasets (FH and Canadian ICGC) were assigned to the four subtypes using the classification rule developed in TCGA, and were then evaluated in relation to cancer recurrence or metastatic/death events. Due to quality control criteria, a small number of CpGs among the 1,306 CpGs used for clustering were missing in the FH cohort (90 missing CpGs) and the Canadian ICGC cohort (14 missing CpGs), and were therefore omitted when computing distances. The fractions of the four subtypes were similar in TCGA and FH (Table 1), though FH has fewer Subtype 4 samples; there were proportionally less Subtype 1 and more Subtype 3 samples in the Canadian ICGC cohort, which may reflect its intermediate-risk patient population. Consistent with the TCGA associations, FH patients in Subtypes 2–4 showed increasing HRs (Table 5), reaching 2.76 for Subtype 4 (p-value=0.002) in the base model. The association with metastatic-lethal prostate cancer as compared to no recurrence was even more pronounced, with patients in Subtype 4 showing a HR=9.73 in the base model (p-value=0.002). Adjusting for Gleason score, PSA, and tumor stage diminished the significance, yet the association of Subtype 4 with metastatic-lethal prostate cancer remained significant (p-value=0.043, HR=5.90). The BCR association results in the Canada ICGC cohort are not significant for either of the subtypes, nor the associations with metastasis/death events. In Supplementary Table 2, the

associations between subtypes and prognosis were also assessed when adjusting for age and Gleason score, Subtype 4 remained significantly associated with metastatic-lethal prostate cancer in the FH cohort. The global association of 4 subtypes with prognosis outcomes are statistically significant for the FH cohort. A similar trend of hazard ratios was observed for stratified analysis in the three risk groups, though with less significance (Supplementary Table 3).

Figure 2 shows the cumulative incidence function plots for the four subtypes in TCGA and FH, accounting for deaths due to other reasons and separated by BCR events and metastasis/death events. From Subtype 1 to Subtype 4, the gradual increase of hazards for cancer recurrence and for metastasis/death events was more evidently seen with statistical significance in the FH cohort than in TCGA patients, due to a larger number of events in the FH cohort. Adding the 4 subtypes to a logistic regression model with age, Gleason score, PSA, and stage for predicting metastatic-lethal prostate cancer increased the AUC for ROC from 0.8 to 0.831, though the increment is borderline statistically significant ( $p=0.12$ , Supplementary Figure 3). Similarly, adding the 4 subtypes to a logistic regression model with age, Gleason score, PSA, and stage for predicting BCR outcomes increased the AUC for ROC from 0.721 to 0.733, with a  $p$ -value of 0.12. This is largely due to the observation that among the 4 subtypes, subtype 4 is the one that is significantly associated with prognosis, less so for subtype 2 and 3.

## Discussion

Using a set of differentially-variable hypermethylated CpG sites, we have identified four subtypes of primary prostate tumors from patients with localized disease at diagnosis and had been treated with RP, which predict gradually worsening prognosis from Subtype 1 to 4, i.e., increased hazard of biochemical relapse and eventual metastasis and cancer-specific mortality. The subtypes were detected in the TCGA dataset using a clustering algorithm, then the classification rule was independently applied to the FH and Canadian ICGC studies. One of the most striking findings is that, when compared to the first subtype and adjusting for age and race, the fourth subtype had a HR of 2.09 for recurrence ( $p=0.029$ ) and of 6.05 for metastatic-lethal outcomes ( $p=0.13$ ) in the TCGA dataset; a HR of 2.76 for recurrence ( $p=0.002$ ) and a HR of 9.73 for metastatic-lethal outcomes ( $p=0.002$ ) in the FH dataset. Though these results failed to replicate in the ICGC cohort, the consistency across three studies of an increased risk of recurrence and cancer-specific mortality supports the clinical significance of Subtype 4.

Nearly all previous methylation studies for prostate cancer prognosis were supervised analysis comparing methylation between prognosis groups (14–19,21). To our knowledge, this study is the only analysis that used an unsupervised clustering algorithm to define methylation subtypes first, then identified subtype-prognosis associations. The association between subtype 4 and prognosis was validated in the Fred Hutch cohort. We suspect that the inconsistency in the Canadian ICGC dataset is driven by cohort heterogeneity and the potential batch effects from different methylation experiments. TCGA patients typically presents more aggressive diseases because of the large tumor size TCGA needed to perform multi-omics analyses. The Fred Hutch cohort contains patients with less aggressive disease



presentations drawn from a SEER population. The ICGC cohort is an intermediate risk group with fewer patients with Gleason score  $\leq 6$ , and fewer patients with high PSA  $>20$ , and so there might be fewer high-risk patients that can be identified to have adverse outcomes (Tables 1 and 5). For example, there were only two metastasis/death events in Subtype 4 in the ICGC cohort. It is also possible that the relatively small proportion of patients assigned to Subtype 1, which was the basis of the comparison, affected our ability to detect a significant association.

In TCGA, Subtypes 3 and 4 had higher Gleason scores, though adjusting for Gleason score did not remove the association between subtypes and outcomes. In the only prospective cohort study (FH), Subtype 4 contained 24 (86%) patients with Gleason score  $\leq 7$  (3+4), yet showed a 5 to 10 fold increase in the risk for metastasis and death. Indeed, when the four subtypes were added to a prediction model with age, Gleason score, PSA and tumor stage, the AUC for ROC increased from 0.8 to 0.831 (p-value=0.12) in the FH dataset (Supplementary Figure 3). These results suggest that the methylation subtypes may be an independent predictor of prognosis, likely occurring earlier in the carcinogenesis pathway than before the pathologic indicators at diagnosis such as grade and stage.

An interesting observation is that Subtype 4 had the lowest proportion of tumors with the *TMPRSS2-ERG* fusion (6% in TCGA and 9% in FH, Tables 3 and 4). Since much higher proportions of the fusion were found in TCGA (31%) and FH (42%), Subtypes 3 and 4 appear to be distinct subgroups with different genomic features. Despite being the most common genomic alteration (~50%) in primary prostate tumors, the clinical implication of *TMPRSS2-ERG* fusion is not clear (47,48). While there is some evidence that the *TMPRSS2-ERG* fusion is associated with shorter survival among men managed conservatively or by watchful waiting (49,50), it has not been consistently associated with recurrence or survival among men treated with RP (32,51). Our results add to the existing knowledge of prostate cancer genomics by identifying a high-risk subtype for adverse cancer outcomes that has a low frequency of the *TMPRSS2-ERG* fusion.

In a similar clinical context, prostate cancer DNA copy number alterations have been shown to predict patient outcome after primary treatment (11). Though DNA methylation may be intertwined with copy number changes, we found no evidence to support that these subtypes are driven by DNA copy number alterations. We examined the overlap of the 1,306 selected CpGs and the most recurrently aberrant regions (95 genes) reported in Table 2 of (11), and we only found four CpGs among the 1,306 CpGs used for clustering, namely cg10722846, cg10795666, cg26559315, and cg26699292, which reside in 4 genes (*TRAPPC9*, *MLNR*, *LYNX1*, *WWOX*). Similar efforts were made for another two larger sets of CNA regions reported for the prostate cancer genome (6,12): for CNA segments reported in (6), there are 52 CpGs (4%) out of the 1,306 CpGs falling into the reported CNA regions; for CNA segments reported in (12), there are 183 CpGs (14%) falling into the reported CNA regions.

As a major strength of our analysis, we were able to perform methylation subtyping using three large methylome datasets with well-annotated clinical outcomes, including a prospective cohort of prostate cancer cases recruited from a population-based SEER cancer registry (FH study). We detected the methylation subtypes in TCGA and evaluated them

in two other studies, diminishing potential confounding in the initial TCGA discovery set. Another strength of our analysis is that it improves on the initial TCGA clustering (5), since we used a subset of differentially variable CpGs that has significantly lower methylation in matched adjacent benign samples to define the subtypes.

A limitation of our analysis is that the number of patients with metastatic-lethal endpoints, the most clinically relevant outcomes, were limited (8 in TCGA, 29 in FH, 23 in Canada ICGC). Though the effect sizes for Subtype 4 were encouraging, the small number of events requires interpreting these associations with caution. Future analyses of larger cohorts with methylation and survival data will be required to further evaluate the prognostic potential of these methylation cluster-defined prostate cancer subtypes.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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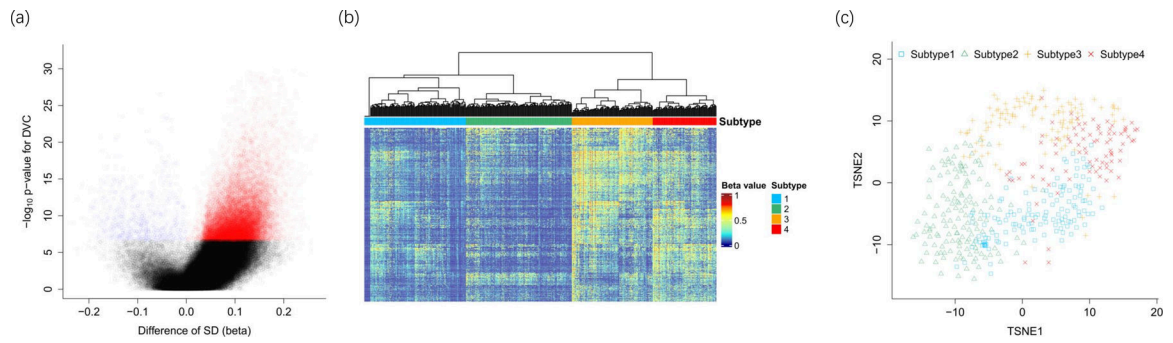
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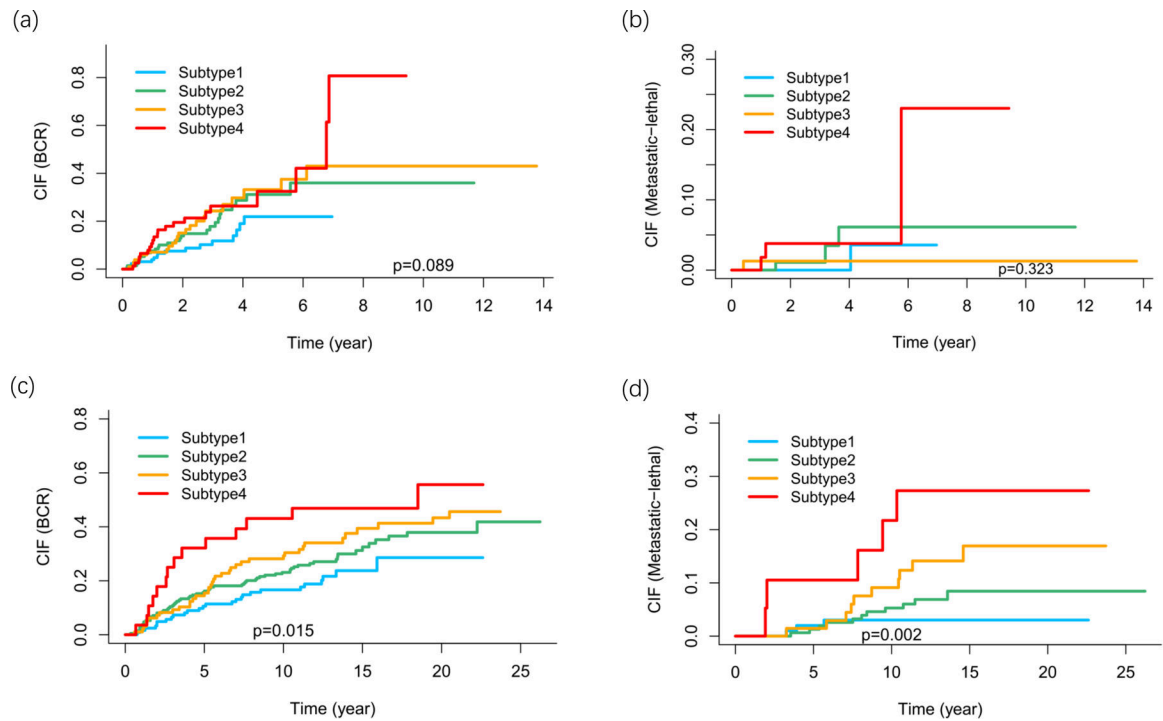
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**Figure 1.** TCGA-PRAD DNA methylation. (a) Volcano plot for testing differentially variable CpGs between cancers and normal samples. (b) DNA methylation subtypes. (c) TSNE plot for tumor DNA methylation subtypes



**Figure 2.** Cumulative incidence functions (CIFs) for developing biochemical recurrence (BCR) or metastatic-lethal events, after accounting for deaths due to other reasons. (a) TCGA, recurrence. (b) TCGA, metastatic-lethal (BCR excluded). (c) FH, recurrence. (d) FH, metastatic-lethal (BCR excluded).

**Table 1.**

Demographic and clinical characteristics of prostate cancer patients in the three studies.

	TCGA (n= 475)	Fred Hutch (n= 458)	Canada ICGC (n= 236)	p-value
Age	Median (25%, 75%)	59 (53, 63)	62 (59, 67)	$4.82 \times 10^{-16}$
Race	Caucasian	405 (85%)	430 (94%)	226 (95%)
	African American	58 (12%)	28 (6%)	6 (3%)
	Asian	12 (3%)	0 (0%)	4 (2%)
Gleason Score	<= 6	41 (9%)	217 (47%)	21 (9%)
	7 (3+4)	146 (31%)	166 (36%)	137 (58%)
	7 (4+3)	97 (20%)	41 (9%)	58 (25%)
PSA (ng/mL)	8-10	191 (40%)	34 (7%)	17 (7%)
	Missing	0 (0%)	0	3 (1%)
	<4	50 (11%)	70 (15%)	24 (10%)
Stage	4-10	264 (56%)	274 (60%)	155 (66%)
	10-20	93 (20%)	60 (13%)	56 (24%)
	20	53 (11%)	26 (6%)	1 (0.4%)
Subtype	Missing	15 (3%)	28 (6%)	0 (0%)
	Local/T2	178 (37%)	312 (68%)	114 (48%)
	Regional/T3	211 (44%)	146 (32%)	102 (43%)
Subtype	T4/N1	80 (17%)	0 (0%)	0 (0%)
	Missing	6 (1%)	0 (0%)	20 (8%)
	1	137 (29%)	123 (27%)	21 (9%)
Subtype	2	143 (30%)	210 (46%)	77 (33%)
	3	109 (23%)	97 (21%)	97 (41%)
	4	86 (18%)	28 (6%)	41 (17%)



Table 2.

Top 20 CpGs associated with the four methylation subtypes.

CpG	CHR	Position	Gene group	Gene	Subtype1 mean	Subtype2 mean	Subtype3 mean	Subtype4 mean	p-value *
cg26415547	12	66583048	1stExon 5'UTR	<i>IRAK3</i>	0.12	0.08	0.34	0.34	$2.58 \times 10^{-45}$
cg13009111	11	71350975		<i>KRTAP5-1//FAM86C</i>	0.17	0.1	0.3	0.35	$6.75 \times 10^{-41}$
cg21884421	15	65648103	Body	<i>IGDCC3</i>	0.13	0.16	0.48	0.32	$1.59 \times 10^{-38}$
cg07462540	7	8473279	TSS1500	<i>NXPHI</i>	0.12	0.11	0.24	0.35	$2.07 \times 10^{-36}$
cg20750832	5	170742345		<i>TLX3/NPM1</i>	0.21	0.22	0.49	0.49	$2.88 \times 10^{-36}$
cg16332256	4	9534388		<i>DEFB131/MIRS4812</i>	0.14	0.11	0.27	0.42	$4.83 \times 10^{-36}$
cg02145932	20	13200969	TSS1500	<i>ISMI</i>	0.12	0.05	0.3	0.36	$1.10 \times 10^{-35}$
cg15720669	3	138666040	TSS200	<i>FOXL2/C3orf72</i>	0.09	0.13	0.37	0.3	$1.56 \times 10^{-35}$
cg12072560	18	12254173	TSS1500 TSS200	<i>CIDEA</i>	0.13	0.11	0.38	0.32	$9.4 \times 10^{-35}$
cg17802942	3	138666050	TSS200	<i>FOXL2/C3orf72</i>	0.08	0.13	0.36	0.28	$1.61 \times 10^{-34}$
cg00114029	1	35351407	Body	<i>DLGAP3</i>	0.14	0.13	0.31	0.37	$1.19 \times 10^{-33}$
cg18932798	10	105037503	1stExon	<i>INA</i>	0.12	0.06	0.28	0.37	$1.61 \times 10^{-32}$
cg15980539	6	152128865	5'UTR 1stExon	<i>ESR1</i>	0.12	0.11	0.29	0.32	$1.64 \times 10^{-32}$
cg11120927	2	239072674		<i>KLHL30/ILKAP</i>	0.1	0.19	0.42	0.33	$1.82 \times 10^{-32}$
cg15041550	6	39016590	1stExon 5'UTR	<i>GLPIR</i>	0.16	0.16	0.41	0.33	$2.95 \times 10^{-32}$
cg12664209	20	13200954	TSS1500	<i>ISMI</i>	0.21	0.07	0.42	0.49	$7.97 \times 10^{-32}$
cg11850773	18	904963	5'UTR TSS1500 1stExon	<i>ADCYAPI</i>	0.13	0.17	0.24	0.32	$2.23 \times 10^{-31}$
cg00851770	18	12254175	TSS1500 TSS200	<i>CIDEA</i>	0.09	0.08	0.29	0.24	$2.30 \times 10^{-31}$
cg27109129	4	74864165	Body	<i>CXCL5</i>	0.23	0.28	0.63	0.5	$3.20 \times 10^{-31}$
cg03382304	21	27012176	1stExon	<i>JAM2</i>	0.13	0.07	0.21	0.29	$3.22 \times 10^{-31}$

\* p-value is computed using ANOVA for association with 4 Subtypes.

**Table 3.**

Distribution of relevant demographic and clinical characteristics by subtypes in the TCGA-PRAD cohort.

	Subtype 1	Subtype 2	Subtype 3	Subtype 4	p-value	
Age	Median (25%, 75%)	60 (56, 66)	60 (55, 65)	64 (59, 68)	64 (58, 67)	1.02×10 <sup>-5</sup>
Race						
	Caucasian	112 (82%)	131 (92%)	86 (79%)	76 (88%)	
	African American	19 (14%)	9 (6%)	21 (19%)	9 (10%)	0.037
	Asian	6 (4%)	3 (2%)	2 (2%)	1 (1%)	
Gleason Score						
	≤6	19 (14%)	15 (10%)	3 (3%)	4 (5%)	
	7 (3+4)	49 (36%)	56 (39%)	21 (19%)	20 (23%)	8.32×10 <sup>-5</sup>
	7 (4+3)	23 (17%)	23 (16%)	28 (26%)	23 (27%)	
	8–10	46 (34%)	49 (34%)	57 (52%)	39 (45%)	
PSA (ng/mL) at diagnosis						
	<4	18 (13%)	19 (13%)	6 (6%)	7 (8%)	
	4–10	76 (55%)	88 (62%)	55 (50%)	45 (52%)	
	10–20	31 (23%)	25 (17%)	23 (21%)	14 (16%)	7.38×10 <sup>-3</sup>
	20	11 (8%)	7 (5%)	19 (17%)	16 (19%)	
	Missing	1 (1%)	4 (3%)	6 (6%)	4 (5%)	
Stage						
	Local/T2	60 (44%)	60 (42%)	29 (27%)	29 (34%)	
	Regional/T3	54 (39%)	59 (41%)	62 (57%)	36 (42%)	0.024
	T4/N1	21 (15%)	22 (15%)	16 (15%)	21 (24%)	
	Missing	2 (1%)	2 (1%)	2 (2%)	0 (0%)	
<i>TMPRSS2-ERG</i> gene fusion						
	No fusion	119 (87%)	17 (12%)	75 (69%)	81 (94%)	5.9×10 <sup>-49</sup>
	Fusion	18 (13%)	126 (88%)	34 (31%)	5 (6%)	

**Table 4.** Distribution of relevant demographic and clinical characteristics by subtypes in the FH cohort.

	Subtype 1	Subtype 2	Subtype 3	Subtype 4	p-value	
Age	Median (25%, 75%)	60 (53, 64)	57 (51, 61)	61 (56, 64)	60 (56, 65)	6.28×10 <sup>-6</sup>
Race						
	Caucasian	106 (86%)	203 (97%)	93 (96%)	28 (100%)	1.36×10 <sup>-3</sup>
	African American	17 (14%)	7 (3%)	4 (4%)	0 (0%)	
Gleason Score						
	<= 6	67 (54%)	110 (52%)	34 (35%)	6 (21%)	
	7 (3+4)	42 (34%)	75 (36%)	35 (36%)	14 (51%)	1.08×10 <sup>-3</sup>
	7 (4+3)	8 (7%)	13 (6%)	16 (16%)	4 (14%)	
	8-10	6 (5%)	12 (6%)	12 (12%)	4 (14%)	
PSA (ng/mL) at diagnosis						
	<4	18 (15%)	36 (17%)	14 (14%)	2 (7%)	
	4-10	79 (64%)	132 (63%)	47 (48%)	16 (57%)	
	10-20	14 (11%)	20 (10%)	20 (21%)	6 (21%)	0.031
	20	3 (2%)	11 (5%)	9 (9%)	3 (11%)	
	Missing	9 (7%)	11 (5%)	7 (7%)	1 (4%)	
Stage						
	Local/T2	94 (76%)	141 (67%)	63 (65%)	14 (50%)	0.034
	Regional/T3	29 (24%)	69 (33%)	34 (35%)	14 (50%)	
<i>TMPRSS2-ERG</i> gene fusion						
	No fusion	95 (77%)	27 (13%)	53 (55%)	23 (82%)	
	Fusion	19 (15%)	172 (82%)	41 (42%)	3 (11%)	1.8×10 <sup>-35</sup>
	Missing	9 (7%)	11 (5%)	3 (3%)	2 (7%)	

Associations between methylation subtypes and biochemical recurrence (BCR) or metastatic-lethal events. Cause-specific Cox proportional hazards regression models are used. In the base models, BCR is adjusted for age and race; the metastatic-lethal events models are adjusted for age. In the full models, BCR is adjusted for age, race, Gleason score, PSA, and tumor stage; metastatic-lethal events are adjusted for age, Gleason score, PSA, and tumor stage. (a) TCGA-PRAD, results of metastatic-lethal events are not shown due to the limited sample size. (b) FH, (c) Canadian ICGC.

**Table 5.**

	Base model				Full model			
	No. patients (No. events)	HR (95% CI)	p-value	Global p-value <sup>a</sup>	No. patients (No. events)	HR (95% CI)	p-value	Global p-value <sup>a</sup>
<b>a)</b>								
BCR								
Subtype 1	137(16)	REF			134(16)	REF		
Subtype 2	143(28)	1.59 (0.86–2.95)	0.14		137(27)	1.55 (0.83–2.91)	0.17	
Subtype 3	109(23)	1.89 (0.98–3.62)	0.057	0.07	101(22)	1.40 (0.72–2.74)	0.32	0.25
Subtype 4	86(21)	2.09 (1.08–4.03)	0.029		82(18)	1.47 (0.73–2.98)	0.28	
<b>b)</b>								
BCR								
Subtype 1	123(26)	REF			114(25)	REF		
Subtype 2	210(65)	1.44 (0.91–2.29)	0.12		199(59)	1.14 (0.71–1.86)	0.57	
Subtype 3	97(38)	1.80 (1.09–2.99)	0.03	0.007	90(37)	1.09 (0.64–1.86)	0.75	0.36
Subtype 4	28(14)	2.76 (1.43–5.31)	0.002		27(14)	1.41 (0.70–2.82)	0.33	
Metastatic-lethal								
Subtype 1	100(3)	REF			91(2)	REF		
Subtype 2	156(11)	2.19 (0.60–7.94)	0.23		149(9)	2.06 (0.43–9.80)	0.36	
Subtype 3	69(10)	4.75 (1.30–17.33)	0.02	0.004	63(10)	3.35 (0.69–16.38)	0.14	0.12
Subtype 4	19(5)	9.73 (2.32–40.75)	0.002		18(5)	5.90 (1.06–32.92)	0.043	
<b>c)</b>								

a)

	Base model					Full model						
	No. patients (No. events)	HR (95% CI)	p-value	Global p-value <sup>a</sup>	No. patients (No. events)	HR (95% CI)	p-value	Global p-value <sup>a</sup>	No. patients (No. events)	HR (95% CI)	p-value	Global p-value <sup>a</sup>
<b>BCR</b>												
Subtype 1	21(4)	REF			21(4)	REF			21(4)	REF		
Subtype 2	77(19)	1.60 (0.54–4.73)	0.39		73(16)	0.84 (0.27–2.63)	0.77		73(16)	0.84 (0.27–2.63)	0.77	
Subtype 3	97(34)	1.93 (0.69–5.34)	0.18	0.19	86(28)	1.15 (0.39–3.40)	0.81	0.81	86(28)	1.15 (0.39–3.40)	0.81	0.81
Subtype 4	41(8)	1.06 (0.32–3.53)	0.93		33(4)	0.47 (0.12–1.90)	0.29		33(4)	0.47 (0.12–1.90)	0.29	
<b>Metastatic-lethal</b>												
Subtype 1	18(1)	REF			18(1)	REF			18(1)	REF		
Subtype 2	65(7)	2.67 (0.32–22.23)	0.36		63(6)	1.50 (0.16–14.14)	0.72		63(6)	1.50 (0.16–14.14)	0.72	
Subtype 3	76(13)	3.30 (0.43–25.49)	0.25	0.25	68(10)	1.66 (0.20–14.02)	0.64	0.59	68(10)	1.66 (0.20–14.02)	0.64	0.59
Subtype 4	35(2)	1.46 (0.13–16.41)	0.76		30(1)	0.59 (0.04–10.08)	0.72		30(1)	0.59 (0.04–10.08)	0.72	

<sup>a</sup>Global p-value is for testing association of all 4 subtypes.