

Cochrane Database of Systematic Reviews

Glucocorticoids for croup in children (Review)

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[Intervention Review]

Glucocorticoids for croup in children

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ABSTRACT

Background

Glucocorticoids are the mainstay for the treatment of croup. The existing evidence demonstrates that glucocorticoids are effective in the treatment of croup in children. However, updating the evidence on their clinical relevance in croup is imperative. This is an update to a review first published in 1999, and updated in 2004, 2011, and 2018.

Objectives

To investigate the effects and safety of glucocorticoids in the treatment of croup in children aged 18 years and below.

Search methods

We searched the Cochrane Library, which includes the Cochrane Central Register of Controlled Trials (CENTRAL; 2022 Issue 9), Ovid MEDLINE Epub Ahead of Print, In-Process & Other Non-Indexed Citations and Ovid MEDLINE (1946 to 4 March 2022), Embase (Ovid) (1974 to 4 March 2022). We also searched the WHO ICTRP and ClinicalTrials.gov on 4 March 2022.

Selection criteria

We included randomised controlled trials (RCTs) in children (aged 18 years and below) with croup. We assessed the effect of glucocorticoids compared to the following: placebo, any other pharmacologic agents, any other glucocorticoids, any combination of other glucocorticoids, given by different modes of administration, or given in different doses. The included studies must have assessed at least one of our primary outcomes (defined as the change in croup score or return visits, (re)admissions to the hospital or both) or secondary outcomes (defined as the length of stay in hospital or emergency departments, patient improvement, use of additional treatments, or adverse events).

Data collection and analysis

Review authors independently extracted data, with another review author verified. We entered the data into Review Manager 5 for metaanalysis. Two review authors independently assessed studies for risk of bias using the Cochrane risk of bias tool. Two review authors assessed the certainty of the evidence for the primary outcomes using the GRADE approach.

Main results

This updated review includes 45 RCTs with a total of 5888 children, an increase of two RCTs with 1323 children since the last update. We also identified one ongoing study and one study awaiting classification. We assessed most studies (98%) as at high or unclear risk of bias.

Any glucocorticoid compared to placebo



Compared to placebo, glucocorticoids may result in greater reductions in croup score after two hours (standardised mean difference (SMD) –0.65, 95% confidence interval (CI) –1.13 to –0.18; 7 RCTs, 426 children; low-certainty evidence); six hours (SMD –0.76, 95% CI –1.12 to –0.40; 11 RCTs, 959 children; low-certainty evidence); and 12 hours (SMD –1.03, 95% CI –1.53 to -0.53; 8 RCTs, 571 children; low-certainty evidence). The evidence for change in croup score after 24 hours is very uncertain (SMD –0.86, 95% CI –1.40 to –0.31; 8 RCTs, 351 children; very low-certainty evidence).

One glucocorticoid compared to another glucocorticoid

There was little to no difference between prednisolone and dexamethasone for reduction in croup score at two-hour post-baseline score (SMD 0.06, 95% CI –0.06 to 0.18; 1 RCT, 1231 children; high-certainty evidence). There was likely little to no difference between prednisolone and dexamethasone for reduction in croup score at six-hour post-baseline score (SMD 0.21, 95% CI –0.21 to 0.62; 1 RCT, 99 children; moderate-certainty evidence). However, dexamethasone probably reduced the return visits or (re)admissions for croup by almost half (risk ratio (RR) 0.55, 95% CI 0.28 to 1.11; 4 RCTs, 1537 children; moderate-certainty evidence), and showed a 28% reduction in the use of supplemental glucocorticoids as an additional treatment (RR 0.72, 95% CI 0.53 to 0.97; 2 RCTs, 926 children).

Dexamethasone given in different doses

Compared to 0.15 mg/kg, 0.60 mg/kg dexamethasone probably reduced the severity of croup as assessed by the croup scoring scale at 24-hour postbaseline score (SMD 0.63, 95% CI 0.16 to 1.10; 1 RCT, 72 children; moderate-certainty evidence); however, this was not the case at two hours (SMD -0.27, 95% CI -0.76 to 0.22; 2 RCTs, 861 children; high-certainty evidence). There was probably no reduction at six hours (SMD -0.45, 95% CI -1.26 to 0.35; 3 RCTs, 178 children; moderate-certainty evidence), and the evidence at 12 hours is very uncertain (SMD -0.60, 95% CI -4.39 to 3.19; 2 RCTs, 113 children; very low-certainty evidence). There was little to no difference between doses of dexamethasone in return visits or (re)admissions of children or both (RR 0.91, 95% CI 0.71 to 1.17; 3 RCTs, 949 children; high-certainty evidence) or length of stay in the hospital or emergency department (mean difference 0.12, 95% CI -0.32 to 0.56; 2 RCTs, 892 children). The need for additional treatments, such as epinephrine (RR 0.78, 95% CI 0.34 to 1.75; 2 RCTs, 885 children); intubation (risk difference 0.00, 95% CI -0.00 to 0.00; 2 RCTs, 861 children); or use of supplemental glucocorticoids (RR 0.77, 95% CI 0.51 to 1.15; 2 RCTs, 617 children), also did not differ between doses of dexamethasone.

There were moderate to high levels of heterogeneity in the analyses for most comparisons. Adverse events were observed for some of the comparisons reported in the review.

Authors' conclusions

The evidence that glucocorticoids reduce symptoms of croup at two hours, shorten hospital stays, and reduce the rate of return visits or (re)admissions has not changed in this update. A smaller dose of 0.15 mg/kg of dexamethasone may be as effective as the standard dose of 0.60 mg/kg. More RCTs are needed to strengthen the evidence for effectiveness of low-dose dexamethasone at 0.15 mg/kg to treat croup.

PLAIN LANGUAGE SUMMARY

Glucocorticoids for croup in children

Review question

What is the effectiveness and safety of glucocorticoids when treating children with croup?

Background

Respiratory viruses are the main cause of croup in children. Croup leads to a swelling of the throat and airway, which can make breathing difficult. Children also present with a special type of cough called a barking cough. Glucocorticoids are types of steroids that help reduce the swelling, thereby making it easier for children with croup to breathe.

This is an update of a review first published in 1999 and updated in 2004, 2011, and 2018.

Search date

The evidence is current to 4 March 2022.

Study characteristics

We included 2 new studies with 1323 children, for a total of 45 studies with 5888 children aged 0 to 18 years published between 1964 and 2021. The three types of glucocorticoids used in the new studies were budesonide, dexamethasone, and prednisolone. The most recent study compared the effectiveness of budesonide and dexamethasone. The other new study compared the effectiveness of dexamethasone and prednisolone, as well as a small dose of dexamethasone (0.15 mg/kg) versus 0.60 mg/kg dexamethasone. We added the data from the new study that compared the doses of dexamethasone to previously included studies looking at the same comparison.

Study funding sources



Funding sources included government (11%), academic or research institute (7%), industry (18%), or foundations (9%). More than half of the studies (55%) did not report funding sources.

Key results

Compared to prednisolone, dexamethasone showed no improvement in croup score at two and six hours after presenting to the hospital or emergency department, and probably reduced return visits or (re)admissions for croup by almost half. The addition of supplemental glucocorticoid favoured dexamethasone versus prednisolone. Compared to 0.15 mg/kg dexamethasone, the standard dose of 0.60 mg/kg probably reduced the severity of croup as assessed by the croup scoring scale at 24 hours after presenting to the hospital or emergency department. However, we did not find any important difference between groups in croup scoring scale at 2, 6, or 12 hours, return visits or (re)admissions of children, or length of stay in the hospital or emergency department. The need for additional treatments such as the use of other drugs like epinephrine, supplemental glucocorticoid, or the use of a tube to help breathing did not differ between 0.15 mg/kg and 0.60 mg/kg dexamethasone. No serious adverse events from the use of the glucocorticoids were reported in the newly included studies.

Conclusions

The evidence has not changed that glucocorticoids reduce symptoms of croup at two hours, shorten hospital stays, and reduce the rate of return visits or (re)admissions compared to placebo (dummy treatment). A small dose of dexamethasone at 0.15 mg/kg may be as effective as the standard dose of 0.60 mg/kg. More studies are needed to strengthen the evidence for the effectiveness of low-dose dexamethasone at 0.15 mg/kg to treat croup. We conclude that glucocorticoids are effective in the treatment of croup in children.

Certainty of evidence

Most studies (98%) had problems related to their methods, reporting issues, or both. For any glucocorticoid compared to placebo, we downgraded the certainty of the evidence for change in croup score after 2, 6, 12, and 24 hours and return visits or (re)admissions due to study variability, imprecision and inconsistency of study results, and risk of bias. There is little evidence that reporting bias influenced our results for return visits or (re)admissions, or both. Similar threats to the certainty of the evidence were present in the other comparisons in this review, including concerns related to risk of bias and inconsistency and imprecision of study results.

SUMMARY OF FINDINGS

Summary of findings 1. Any glucocorticoid compared to placebo for croup

Any glucocorticoid compared to placebo for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: any glucocorticoid

Comparison: placebo

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect (95% CI)	№ of partici- pants	Certainty of the evidence	Comments**
	Placebo	Placebo Any glucocorticoid		(studies)	(GRADE)	
Change in croup score. Assessed with different scores in different stud- ies. Lower scores mean fewer symptoms. (Follow-up: 2 hours)	The mean change in croup score was -1.50 to -0.81 .	The mean change in croup score was 0.65 standard deviations in favour (1.13 more to 0.18 more).	-	426 (7 RCTs)	⊕⊕⊝⊝ Lowa,b	A standard deviation of 0.65 represents a moderate difference between groups.
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 6 hours)	The mean change in croup score was -3.23 to -0.65 .	The mean change in croup score was 0.76 standard deviations in favour (1.12 more to 0.40 more).	-	959 (11 RCTs)	⊕⊕⊝⊝ Lowc,d	A standard deviation of 0.76 represents a large difference between groups.
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 12 hours)	The mean change in croup score was -7.62 to -1.00 .	The mean change in croup score was 1.03 standard deviations in favour (1.53 more to 0.53 more).	-	571 (8 RCTs)	⊕⊕⊝⊝ Lowe,f	A standard deviation of 1.03 represents a large difference between groups.
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms.	The mean change in croup score was −2.56 to −1.05 .	The mean change in croup score was 0.86 standard deviations in favour (1.40 more to 0.31 more).	-	351 (8 RCTs)	⊕⊙⊙⊝ Very lowg,h	A standard deviation of 0.86 represents a large difference between groups.

(Follow-up: 24 hours)				
Return visits or (re)ad- missions or both	204 per 1000 106 per 1000 (74 to 153)	RR 0.52 (0.36 to 0.75)	1679 (10 RCTs)	⊕⊕⊝⊝ Low ^{i,j}
Adverse events	13/26 (50%) studies reported collecting adverse events no serious adverse events. Bjornson 2004 reported 7 in 0.83% in the dexamethasone group and 4/361, 1.11% in 1996 reported 1 child with neutropenia consistent with amethasone group (1/28, 3.57%). Kuusela 1988 reported tions (pneumonia, sinusitis, otitis media) requiring antidexamethasone group and 2/16, 12.5% in the placebo schild with pneumonitis in the placebo group (1/13, 7.79 nia in the dexamethasone group (2/16, 12.5%). Roberts exacerbated symptoms, 5 children with emotional dist stance of eye irritation in the budesonide group (9/42, 2 acerbated symptoms, 6 children with emotional distress and 1 instance each of eye irritation and tongue irritation 40%).	stances of pneumonia (3/359, the placebo group). Johnson bacterial tracheitis in the dexd 7 secondary bacterial infection therapy: 5/35, 14% in the group. Super 1989 reported 1 (6) and 2 children with pneumo-1999 reported 1 instance of tess, 2 with vomiting, and 1 in-1.4%), and 3 instances of exs, 3 with vomiting, 2 rashes,	1399 (13 RCTs)	⊕⊕⊝⊝ Lowk,l

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

 q We downgraded by one level for inconsistency. There was considerable heterogeneity ($l^{2} = 81\%$), and variation in point estimates.

bWe downgraded by one level for risk of bias. The contributing studies were at high (n = 3) and unclear (n = 4) risk of bias.

cWe downgraded by one level for inconsistency. There was considerable heterogeneity (12 = 83%), and variation in point estimates and in direction of effects for one study.

dWe downgraded by one level for risk of bias. The contributing studies were at high (n = 3) and unclear (n = 8) risk of bias.

eWe downgraded by one level for inconsistency. There was considerable heterogeneity (1² = 86%), and variation in point estimates.

fWe downgraded by one level for risk of bias. The contributing studies were at high (n = 2) and unclear (n = 6) risk of bias.

gwe downgraded by two levels for inconsistency. There was considerable heterogeneity ($l^2 = 81\%$), and variation in point estimates. The confidence intervals did not overlap for some studies. There was variation in the direction of effects.

hWe downgraded by one level for risk of bias. The contributing studies were at high (n = 2) and unclear (n = 6) risk of bias.

We downgraded by one level for inconsistency. There was substantial heterogeneity ($I^2 = 52\%$), and variation in point estimates.

iWe downgraded by one level for risk of bias. The contributing studies were at high (n = 3) and unclear (n = 7) risk of bias.

^{**}We used Cohen's interpretation of effect sizes to determine the magnitude of the difference between groups (0.2 represents a small effect, 0.5 represents a medium effect, 0.8 represents a large effect).

kWe downgraded by one level for imprecision. Narrative synthesis conducted, estimates are not precise. We downgraded by one level for risk of bias. The contributing studies were at high (n = 2) and unclear (n = 11) risk of bias.

Summary of findings 2. Any glucocorticoid compared to epinephrine for croup

Any glucocorticoid compared to epinephrine for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: any glucocorticoid Comparison: epinephrine

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect (95% CI)	№ of partici- pants	Certainty of the evidence	Comments**	
	Epinephrine	Any glucocorticoid	(33 /0 Ci)	(studies)	(GRADE)		
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 2 hours)	The mean change in croup score was -4.24 to -3.74.	The mean change in croup score was 0.77 standard deviations not in favour (0.24 more to 1.77 less).	-	130 (2 RCTs)	⊕⊝⊝⊝ Very low ^a ,b,c	A standard deviation of 0.77 represents a large difference between groups.	
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 6 hours)	The mean change in croup score was -1.25 to -1.10.	The mean change in croup score was 0.10 standard deviations in favour (1.18 more to 0.97 less).	-	63 (2 RCTs)	⊕⊝⊝⊝ Very low ^d ,e,f	A standard deviation of 0.10 represents a minimal difference between groups.	
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 12 hours)	The mean change in croup score was -3.86 to -1.45.	The mean change in croup score was 0.07 standard deviations in favour (0.57 more to 0.43 less).	-	129 (3 RCTs)	⊕⊕⊙⊝ Lowg,h	A standard deviation of 0.07 represents a minimal difference between groups.	
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 24 hours)	The mean change in croup score was -4.40 to -2.01 .	The mean change in croup score was 0.17 standard deviations not in favour (0.18 more to 0.51 less).	-	129 (3 RCTs)	⊕⊕⊙⊝ Lowg, ⁱ	A standard deviation of 0.17 represents a small difference between groups.	

Return visits or (re)admissions or both	0 per 1000 0 per 1000 (0 to 0)	RD 0.00 (-0.04 to 0.04)	130 (2 RCTs)	⊕⊕⊙⊝ Lowg,j
Adverse events	3/4 (75%) studies reported collecting ac 1996 reported no serious adverse event of secondary bacterial infections (pneur quiring antibiotic therapy in the dexame riadou 2010 reported 4 cases of tremor a epinephrine group.	s. Kuusela 1988 reported 5 cases monia, sinusitis, otitis media) re- ethasone group (5/16, 31.3%). Ebo-	162 (3 RCTs)	⊕⊕⊙⊝ Low ^k ,l,

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RD: risk difference

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^aWe downgraded by one level for inconsistency. There was considerable heterogeneity (1² = 87%), and variation in point estimates. There was minimal overlap of the confidence intervals.

bWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size). The effect estimate included both the null effect and a clinically important benefit for epinephrine compared to glucocorticoids.

cWe downgraded by one level for risk of bias. The contributing studies were at high risk of bias (n = 2).

dWe downgraded by two levels for inconsistency. There was considerable heterogeneity (I² = 78%), and variation in point estimates and in the direction of effects.

eWe downgraded by one level for imprecision. The sample size was small (did not meet optimal information size). The effect estimate included both the null effect and a clinically important effect for glucocorticoids compared to epinephrine.

fWe downgraded by one level for risk of bias. The contributing studies were at unclear risk of bias (n = 2).

gWe downgraded by one level for imprecision. The sample size was small (did not meet optimal information size).

hWe downgraded by one level for risk of bias. The contributing studies were at high (n = 1) and unclear (n = 2) risk of bias.

 $^{\text{i}}$ We downgraded by one level for risk of bias. The contributing studies were at high (n = 1) and unclear (n = 2) risk of bias.

JWe downgraded by one level for risk of bias. The contributing studies were at high risk of bias (n = 2).

kWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

We downgraded by one level for risk of bias. The contributing studies were at high (n = 2) and unclear (n = 1) risk of bias.

^{**}We used Cohen's interpretation of effect sizes to determine the magnitude of the difference between groups (0.2 represents a small effect, 0.5 represents a medium effect, 0.8 represents a large effect).



BACKGROUND

Description of the condition

Croup is a common childhood respiratory disease that leads to frequent emergency department often visits (Bjornson 2008). It is a spectrum of diseases including laryngotracheitis, laryngotracheobronchitis, laryngotracheobronchopneumonitis (Sizar 2021). Patients may present with sudden onset of a seal-like barking cough, often accompanied by stridor, voice hoarseness, and respiratory distress (Bjornson 2008). As with many other acute respiratory diseases, croup can be mild, moderate, or severe in presentation. In brief, the pathophysiology of croup involves upper-airway obstruction due to generalised inflammation of the airways, triggered by viral infection (especially the parainfluenza virus, which accounts for over 75% of infections) (Bjornson 2013). Whilst croup is a selflimiting viral infection, the burden of frequent hospitalisation contributes significantly to healthcare utilisation (Bjornson 2013; Denny 1983). Croup accounts for 7% and 3% of hospitalisation in children under five and children between six months and three years in North America (Johnson 2014; Weinberg 2009). Likewise, one European study found that 16% of children aged five to eight years old had suffered from croup at least once, and 5% had experienced recurrent croup (Van Bever 1999).

Description of the intervention

The clinical benefits of glucocorticoids in the management of croup are well documented in the literature (Griffin 2000; Kairys 1989). Unlike the controversies that existed in the 1970s concerning the treatment of croup (Cherry 1979), many clinical guidelines now support the use of glucocorticoids (Alberta Medical Association 2008). Glucocorticoids have also been shown to decrease the rate and length of hospitalisation, return visits, and admission to intensive care unit in children with croup (Brown 2002; Geelhoed 1996b; Kairys 1989). Studies have also continued to highlight the effectiveness of glucocorticoids in reducing the severity of croup (Brown 2002).

How the intervention might work

One of the cardinal features of inflammation is oedema or swelling. Whilst there are associated generalised swellings of the airway in croup, inflammation and oedema of the subglottic larynx (the narrowest part of the paediatric airway) and trachea, especially near the cricoid cartilage, are most clinically significant (Cherry 2008). Glucocorticoids have anti-inflammatory properties through which they reduce croup-related mucosal oedema and inflammation and as such reduce the associated difficulty in breathing (Cherry 2008).

Why it is important to do this review

Systematic reviews of randomised controlled trials (RCTs) on the use of glucocorticoid for the treatment of croup have contributed significantly to the evidence around the management of croup to date. The first Cochrane Review on this study question included 24 RCTs that examined the effectiveness of treating croup with glucocorticoids (Ausejo 2000). A few other reviews have been conducted since to update the existing evidence (Gates 2018; Russell 2004; Russell 2011). The current review is necessary to incorporate new evidence to help strengthen or refute the findings of previous reviews on this study question. As there is a

growing debate about the lowest effective dose of glucocorticoid in the management of croup (Alshehr 2005; Chub-Uppakarn 2007; Dobrovoljac 2009), this review aimed to address this, and to update the existing evidence on the effect of glucocorticoids on croup.

OBJECTIVES

To investigate the effects and safety of glucocorticoids in the treatment of croup in children aged 18 years and below.

METHODS

Criteria for considering studies for this review

Types of studies

We included RCTs in child health research that met our inclusion criteria irrespective of language, publication status, trial conduct and reporting quality, or risk of bias. We excluded all other study designs.

Types of participants

We included RCTs on children aged 18 years and below diagnosed with croup, pseudo croup, or laryngotracheitis. We defined croup as a syndrome consisting of hoarseness, barking cough, and stridor, where an alternative diagnosis of acute stridor had been excluded. We included both inpatients and outpatients, and defined children admitted to the emergency department as outpatients.

Types of interventions

We included studies where the intervention was the use of one or more glucocorticoids via any route of drug administration. There were no restrictions on the type or dose of glucocorticoid administered. We defined the control as the use of a placebo or any other active pharmacologic agent. We considered the following scenarios: the use of any glucocorticoid compared to placebo, glucocorticoid compared to epinephrine, or one glucocorticoid compared to one or a combination of other glucocorticoids, or glucocorticoids given by different modes of administration, or glucocorticoids given in different doses. We excluded studies if none of the treatment groups received one or more glucocorticoids.

Types of outcome measures

We included RCTs that measured on one or more of our primary or secondary outcomes. We excluded studies that failed to meet all of our inclusion criteria.

Primary outcomes

- 1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours.
- 2. Return visits or (re)admissions to the hospital, or both.

Secondary outcomes

- 1. Length of stay in the hospital or emergency department.
- 2. Patient improvement at 2, 6, 12, and/or 24 hours (yes or no, as reported in the individual studies).
- 3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids.
- 4. Any adverse events.



Search methods for identification of studies

Electronic searches

We adopted the search strategy developed by a research librarian in the previous review (Gates 2018) on 4 March 2022 (Appendix 1). The update searches were conducted by the librarian Mê-Linh Lê. We included subject headings and keywords for croup and glucocorticoids and restricted the search to RCTs. We searched the Cochrane Library, which includes the Cochrane Central Register of Controlled Trials (CENTRAL; 2022, Issue 9), Ovid MEDLINE Epub Ahead of Print, In-Process & Other Non-Indexed Citations and Ovid MEDLINE (1946 to 4 March 2022), and Embase (Ovid) (1974 to 4 March 2022).

Searching other resources

We searched the World Health Organization International Clinical Trials Registry Platform (WHO ICTRP) (trialsearch.who.int/) and ClinicalTrials.gov (clinicaltrials.gov) on 4 March 2022 (Appendix 1). We scanned the reference lists of relevant systematic reviews identified during screening and the included studies to identify additional relevant primary studies.

Data collection and analysis

Selection of studies

We transferred the citations identified via the search to Rayyan software after de-duplication (Ouzzani 2016). Three review authors (CT, AK, MR) independently screened the identified citations for eligibility using a two-stage sifting approach to review the title, abstract, and full-text article. Any disagreements were resolved by discussion or by involving another review author (AA) when necessary.

Data extraction and management

Three review authors (CT, AK, MR) independently extracted the data, which were all in the English language. We used Microsoft Excel to manage data extraction (Microsoft Excel). We leveraged the data extraction form used in our previous review (Gates 2018). The details of the data extracted based on participant characteristics, experimental and control interventions, and primary and secondary outcomes have all been previously published (Gates 2018). Any disagreements during data extraction were resolved by discussion or by involving another review author (AA) when necessary.

Assessment of risk of bias in included studies

We used the Cochrane risk of bias tool to assess risk of bias of the included studies (Higgins 2011b). We judged the risk of bias for each study as low, high, or unclear for seven domains: random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessors, incomplete outcome data, selective reporting, and other bias. We determined the overall risk of bias as follows: low when all domains were judged as low risk; unclear when one or more domains were judged as unclear risk; and high when one or more domains were judged as high risk. Two review authors (CT, MR) independently assessed risk of bias, resolving any disagreements by discussion or by involving another review author (AA) when necessary.

Measures of treatment effect

We added relevant data from the included studies into Review Manager 5 for analysis (Review Manager 2020). We computed the effect of treatment using the random-effects model.

Croup scores were reported as the Westley score (Westley 1978), the telephone outpatient (TOP) score (Bjornson 2016), the Downes and Raphaelly score (Downes 1975), or various author-created scales. We therefore used standardised mean differences (SMDs) to combine the outcome for any croup score. A treatment effect (difference between treatment means) divided by its measurement variation (e.g. a pooled standard deviation) gives the SMD. We did not find effect estimates to be significantly different between Westley and other croup scores, so we included studies that reported any croup score in the subgroup analyses. Of note, a decrease in Westley score of one point from baseline is thought to be a clinically important change.

We expressed length of stay as mean differences (MDs) and calculated an overall MD. We calculated risk ratios (RRs) for binary data (i.e. return visits or (re)admissions (or both), patient improvement, use of additional treatments). We calculated risk differences (RDs) where outcomes had zero events in both groups. For return visits or (re)admissions (or both), we calculated the number needed to treat for an additional beneficial outcome (NNTB) for significant results. Because there was substantial variation in control group event rates between studies, we reported the NNTB for the mean control group rate, as well as for the smallest and largest control group rate observed.

We reported data on adverse events narratively.

Unit of analysis issues

As reported in Gates 2018, we calculated the change from baseline croup score in 28 (62%) studies where the change from baseline measures was not reported directly (Alshehr 2005; Amir 2006; Cetinkaya 2004; Chub-Uppakarn 2007; Dobrovoljac 2012; Duman 2005; Eboriadou 2010; Fifoot 2007; Fitzgerald 1996; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 2005; Godden 1997; Husby 1993; Johnson 1996; Klassen 1994; Klassen 1998; Kuusela 1988; Leipzig 1979; Martinez Fernandez 1993; Massicotte 1973; Rittichier 2000; Roberts 1999; Roorda 1998; Super 1989; Vad Pedersen 1998; Von Mühlendahl 1982).

We pooled counts, means, and variances using standard formulae for seven (15%) studies that contained more than one experimental treatment group (Cetinkaya 2004; Eboriadou 2010; Fifoot 2007; Geelhoed 1995c; Johnson 1998; Luria 2001; Parker 2019). One study by Geelhoed (Geelhoed 1995a; Geelhoed 1995b), and another by Skowron (Skowron 1966a; Skowron 1966a and b; Skowron 1966b), presented the results of two individual trials in one publication. We treated these as separate comparisons in the analyses and used pooled counts only when they were reported as such in the publications.

Dealing with missing data

When they were not directly reported, we estimated the variances for continuous data in accordance with the work of Abrams 2005 and Follmann 1992. Using standard formulae, we imputed standard deviations from standard errors in three (7%) studies (Alshehr 2005; Johnson 1998; Von Mühlendahl 1982), ranges in



three (7%) studies (Alshehr 2005; Roorda 1998; Super 1989), 95% confidence intervals (CIs) in two (4%) studies (Fitzgerald 1996; Klassen 1998), and interquartile ranges (IQRs) in three (7%) studies (Johnson 1996; Klassen 1994; Klassen 1998). When the change in croup score from baseline was not directly reported (n = 14, 31%), we derived the variance of the change assuming a correlation of 0.5 between pre- and post-treatment scores (Alshehr 2005; Amir 2006; Chub-Uppakarn 2007; Fitzgerald 1996; Johnson 1996; Klassen 1994; Klassen 1998; Kuusela 1988; Leipzig 1979; Martinez Fernandez 1993; Roorda 1998; Super 1989; Vad Pedersen 1998; Von Mühlendahl 1982)

In 11 (26%) studies, data from which to impute variances for change in croup score or length of stay were inadequate; for these studies we substituted average variances from other studies in the main analysis (Cetinkaya 2004; Dobrovoljac 2012; Eboriadou 2010; Geelhoed 1995c; Geelhoed 2005; Godden 1997; Husby 1993; Kuusela 1988; Massicotte 1973; Roberts 1999; Skowron 1966a; Skowron 1966b). Furukawa and colleagues assert that when the number of studies with imputed data within a meta-analysis is relatively small, variance data can be safely borrowed from other studies and still provide accurate results (Furukawa 2006). For certain outcomes only one study was included in the comparison, and that study did not report a variance estimate; in such a case we did not calculate a point estimate of effect (Cetinkaya 2004; Duman 2005; Fifoot 2007; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 2005; Rittichier 2000).

We substituted medians for means in nine (20%) studies (Alshehr 2005; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Johnson 1996; Klassen 1994; Klassen 1998; Parker 2019; Super 1989; Von Mühlendahl 1982). When data for our prespecified time points (2, 6, 12, and 24 hours from baseline) were not reported, we used time points close to these if available. We substituted one hour for two hours in one study (Dobrovoljac 2012); four hours for six hours in 12 (28%) studies (Alshehr 2005; Amir 2006; Fifoot 2007; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 2005; Godden 1997; Johnson 1996; Klassen 1994; Klassen 1996; Klassen 1998; Massicotte 1973); five hours or discharge for six hours in one study (Johnson 1998); and 14 hours for 12 hours in one study (Massicotte 1973).

Assessment of heterogeneity

In keeping with Gates 2018, we assessed heterogeneity quantitatively with the Chi^2 test for heterogeneity and the I^2 statistic (Higgins 2002). The I^2 statistic indicates the per cent variability due to between-study (or interstudy) variability as opposed to within-study (or intrastudy) variability. We considered an I^2 of less than 40% to be low (potentially unimportant), 30% to 60% to be moderate, 50% to 90% to be substantial, and 75% to 100% to be considerable (Higgins 2011a, Section 9.5.2).

Assessment of reporting biases

In addition to visually inspecting the funnel plots, we used the rank correlation test and weighted regression for the detection of publication bias (Begg 1994; Egger 1997; Light 1984). We used more than one method because the relative merits of the methods are not well established.

Data synthesis

We used random-effects models to combine treatment effects regardless of quantified heterogeneity for the analyses of all outcomes.

Subgroup analysis and investigation of heterogeneity

We explored heterogeneity between studies using subgroup analyses for the primary outcomes of change in croup score from baseline to 2, 6, 12, and 24 hours, and return visits or (re)admissions or both, using the Chi² test for subgroup differences in meta-analysis. We explored heterogeneity by croup score, by inpatient or outpatient status, and by glucocorticoid.

Sensitivity analysis

In some analyses, we imputed variance data for most of the included RCTs (e.g. any glucocorticoid compared to placebo, change in croup score after two hours). We undertook sensitivity analyses for these and all other analyses containing imputed variance data using the largest, smallest, and average variances from the other included RCTs. As per protocol, we did not undertake any additional sensitivity analyses.

Summary of findings and assessment of the certainty of the evidence

We created summary of findings tables for our two main comparisons (any glucocorticoid compared to placebo and any glucocorticoid compared to epinephrine) for the primary outcomes: change in croup score at 2, 6, 12, and 24 hours from baseline, and return visits or (re)admissions or both. The findings for the two main comparisons have not changed since the previous version of the review, as no new data were identified in the current update (Gates 2018). As per protocol, we created summary of findings tables for the remaining comparisons; however, in order not to detract from the two main comparisons, these are included in the Additional tables section. We used the five GRADE considerations (study limitations, consistency of effect, imprecision, indirectness, and publication bias) to assess the certainty of the body of evidence as it relates to the studies that contributed data to the meta-analyses (Atkins 2004). We used the methods and recommendations described in Section 8.5 and Chapter 12 of the Cochrane Handbook for Systematic Reviews of Interventions (Higgins 2011a), employing GRADEpro GDT software (GRADEpro GDT). We justified all decisions to downgrade the certainty of the evidence using footnotes, and made comments to aid the reader's understanding where necessary.

RESULTS

Description of studies

Results of the search

We identified 100 records in the 2022 update search (Figure 1). We retrieved 83 citations from the database searches and 17 records from trial registers, from which we identified and removed 41 duplicates. We screened 59 records by title and abstract and excluded 49 citations. We screened 10 full-text articles of which six were excluded, with reasons for their exclusion provided. A flow diagram illustrating the 2022 update selection process is shown in Figure 1. We added two new RCTs with 1323 children (Huang 2021; Parker 2019), one ongoing study (IRCT20190914044765N1),



and one study awaiting classification (Chen 2018). This updated review includes 45 RCTs with a total of 5888 children.



Figure 1. Flow diagram of study selection for this review.

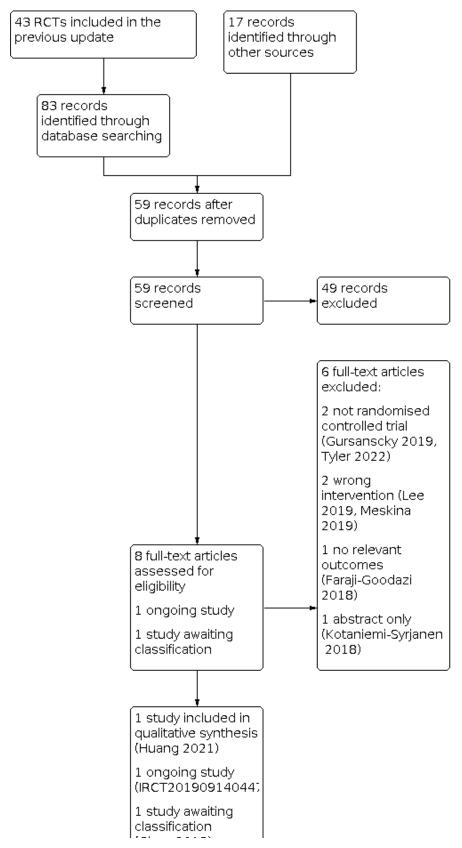
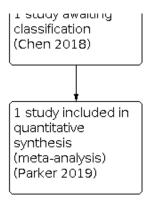




Figure 1. (Continued)



Included studies

Participant and trial characteristics

We identified 42 studies (93%) published in English, and one each in French (Massicotte 1973), Spanish (Martinez Fernandez 1993), and Danish (Vad Pedersen 1998). Four studies (9%) included children with mild croup (Bjornson 2004; Geelhoed 1996a; Luria 2001, Parker 2019). Twenty-three studies (51%) assessed outpatient children (n = 22 emergency department visits, n = 1 physician office visits) (Alshehr 2005; Amir 2006; Bjornson 2004; Cetinkaya 2004; Cruz 1995; Dobrovoljac 2012; Donaldson 2003; Duman 2005; Eboriadou 2010; Fifoot 2007; Garbutt 2013; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1996a; Johnson 1996; Johnson 1998; Klassen 1994; Klassen 1996; Klassen 1998; Luria 2001; Parker 2019; Rittichier 2000; Soleimani 2013; Sparrow 2006). Twenty-three studies (51%) assessed hospitalised children (Chub-Uppakarn 2007; Eden 1964; Eden 1967; Fitzgerald 1996; Geelhoed 1995c; Geelhoed 2005; Godden 1997; Huang 2021; Husby 1993; James 1969; Koren 1983; Kuusela 1988; Leipzig 1979; Martinez Fernandez 1993; Massicotte 1973; Parker 2019; Roberts 1999; Roorda 1998; Skowron 1966a; Skowron 1966b; Super 1989; Tibballs 1992; Vad Pedersen 1998; Von Mühlendahl 1982).

Thirty-two studies (71%) were two-armed trials (Alshehr 2005; Amir 2006; Bjornson 2004; Chub-Uppakarn 2007; Cruz 1995; Dobrovoljac 2012; Donaldson 2003; Eden 1964; Eden 1967; Fitzgerald 1996; Garbutt 2013; Geelhoed 1996a; Geelhoed 2005; Godden 1997; Huang 2021; Husby 1993; James 1969; Johnson 1996; Klassen 1994; Klassen 1996; Koren 1983; Leipzig 1979; Massicotte 1973; Rittichier 2000; Roberts 1999; Roorda 1998; Soleimani 2013; Sparrow 2006; Super 1989; Tibballs 1992; Vad Pedersen 1998; Von Mühlendahl 1982); eight studies (18%) were three-armed trials (Duman 2005; Eboriadou 2010; Fifoot 2007; Geelhoed 1995c; Johnson 1998; Klassen 1998; Luria 2001; Parker 2019); and three studies (7%) were four-armed trials (Cetinkaya 2004; Kuusela 1988; Martinez Fernandez 1993). Two studies (4%) included two individual two-armed trials each (Geelhoed 1995a; Geelhoed 1995b; Skowron 1966a; Skowron 1966b).

Characteristics of the comparisons

Twenty-six studies (58%) investigated any glucocorticoid compared to placebo. Of these, 15 (58%) investigated dexamethasone (Bjornson 2004; Cruz 1995; Dobrovoljac 2012; Eden 1967; Geelhoed 1996a; James 1969; Johnson 1996; Koren 1983; Kuusela 1988; Leipzig 1979; Luria 2001; Martinez Fernandez 1993; Skowron 1966a and b; Super 1989; Von Mühlendahl 1982); four (15%)

investigated budesonide (Godden 1997; Husby 1993; Klassen 1994; Roberts 1999); three (12%) investigated prednisolone (Eden 1964; Massicotte 1973; Tibballs 1992); one (4%) investigated fluticasone (Roorda 1998); and three (12%) investigated both dexamethasone and budesonide (Cetinkaya 2004; Geelhoed 1995c; Johnson 1998). Four studies (10%) investigated any glucocorticoid compared to epinephrine. Of these, one investigated budesonide (Fitzgerald 1996); two investigated dexamethasone (Kuusela 1988; Martinez Fernandez 1993); and one investigated both dexamethasone and beclomethasone (Eboriadou 2010).

Thirteen studies (29%) investigated one glucocorticoid compared to another glucocorticoid. Of these, one investigated budesonide compared to dexamethasone (Huang 2021); six investigated dexamethasone compared to budesonide (Cetinkaya 2004; Duman 2005; Geelhoed 1995c; Johnson 1998; Klassen 1998; Vad Pedersen 1998); one investigated dexamethasone compared to betamethasone (Amir 2006); one investigated dexamethasone compared to beclomethasone (Eboriadou 2010); and four investigated dexamethasone compared to prednisolone (Fifoot 2007; Garbutt 2013; Parker 2019; Sparrow 2006). Three studies investigated one glucocorticoid compared to a combination of glucocorticoids. Of these, one investigated dexamethasone and budesonide compared to a combination of dexamethasone and budesonide (Klassen 1998), and two investigated dexamethasone compared to a combination of dexamethasone and budesonide (Geelhoed 2005; Klassen 1996).

Five studies (11%) investigated dexamethasone using different modes of administration. Of these, four investigated oral compared to intramuscular dexamethasone (Cetinkaya 2004; Donaldson 2003; Rittichier 2000; Soleimani 2013), and one investigated oral compared to nebulised dexamethasone (Luria 2001). Four studies investigated dexamethasone given in different doses. Of these, three investigated 0.60 mg/kg compared to 0.15 mg/kg dexamethasone (Alshehr 2005; Chub-Uppakarn 2007; Fifoot 2007), and one investigated both 0.60 mg/kg compared to 0.30 mg/kg and 0.30 mg/kg compared to 0.15 mg/kg dexamethasone (Geelhoed 1995a; Geelhoed 1995b).

Reported outcomes: primary outcomes

Sixteen studies (35%) reported a two-hour change in croup score (Amir 2006; Chub-Uppakarn 2007; Dobrovoljac 2012; Duman 2005; Eboriadou 2010; Fifoot 2007; Fitzgerald 1996; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 2005; Godden 1997; Husby 1993; Johnson 1996; Parker 2019; Roberts 1999; Roorda



1998); 20 studies (44%) reported a six-hour change in croup score (Alshehr 2005; Amir 2006; Chub-Uppakarn 2007; Fifoot 2007; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Godden 1997; Johnson 1996; Johnson 1998; Klassen 1994; Klassen 1996; Klassen 1998; Kuusela 1988; Martinez Fernandez 1993; Massicotte 1973; Roberts 1999; Roorda 1998; Vad Pedersen 1998; Von Mühlendahl 1982); 12 studies (27%) reported a 12-hour change in croup score (Alshehr 2005; Chub-Uppakarn 2007; Fitzgerald 1996; Geelhoed 1995c; Godden 1997; Kuusela 1988; Martinez Fernandez 1993; Massicotte 1973; Roberts 1999; Super 1989; Vad Pedersen 1998; Von Mühlendahl 1982); and 11 studies (24%) reported a 24-hour change in croup score (Alshehr 2005; Cetinkaya 2004; Fitzgerald 1996; Godden 1997; Kuusela 1988; Leipzig 1979; Martinez Fernandez 1993; Rittichier 2000; Roberts 1999; Roorda 1998; Super 1989). Of the 30 studies (67%) that reported a change in croup score, 18 (60%) used a validated score (the Westley score or a modified Westley score) (Alshehr 2005; Amir 2006; Cetinkaya 2004; Chub-Uppakarn 2007; Dobrovoljac 2012; Duman 2005; Fifoot 2007; Godden 1997; Husby 1993; Johnson 1996; Johnson 1998; Klassen 1994; Klassen 1996; Klassen 1998; Parker 2019; Rittichier 2000; Roorda 1998; Super 1989); 11 (37%) used author-created scales (Fitzgerald 1996; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 2005; Kuusela 1988; Leipzig 1979; Martinez Fernandez 1993; Massicotte 1973; Roberts 1999; Vad Pedersen 1998; Von Mühlendahl 1982); and one used the score by Downes 1975 (Eboriadou 2010). The studies by Bjornson 2004 and Garbutt 2013 used another validated score, the telephone outpatient (TOP) score, to measure clinical improvement. The TOP score is a two-item, three-point score used to assess the presence of stridor and barky cough by asking parents about their child's symptoms in the previous 24 hours (Bjornson 2016). Twentyseven studies (60%) reported return visits or (re)admissions to the hospital or both (Alshehr 2005; Amir 2006; Bjornson 2004; Cruz 1995; Donaldson 2003; Duman 2005; Eboriadou 2010; Fifoot 2007; Fitzgerald 1996; Garbutt 2013; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 1996a; Geelhoed 2005; Johnson 1996; Johnson 1998; Klassen 1994; Klassen 1996; Klassen 1998; Luria 2001; Parker 2019; Rittichier 2000; Roberts 1999; Skowron 1966a; Skowron 1966a and b; Skowron 1966b; Soleimani 2013; Sparrow 2006; Vad Pedersen 1998).

Reported outcomes: secondary outcomes

A total of 13 studies (29%) reported length of stay in the hospital or emergency department (Alshehr 2005; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 2005; Godden 1997; Klassen 1998; Kuusela 1988; Leipzig 1979; Parker 2019; Roorda 1998; Skowron 1966a; Skowron 1966a and b; Skowron 1966b; Sparrow 2006; Super 1989). Twelve studies (27%) reported patient improvement; of these, one reported improvement after two hours (Roberts 1999); eight reported improvement after six hours (Eden 1964; Eden 1967; Johnson 1996; Klassen 1994; Klassen 1996; Klassen 1998; Massicotte 1973; Roberts 1999); six reported improvement after 12 hours (Eden 1964; Eden 1967; James 1969; Massicotte 1973; Roberts 1999; Super 1989); and seven reported improvement after 24 hours (Cruz 1995; Donaldson 2003; Eden 1964; Eden 1967; James 1969; Roberts 1999; Super 1989). About two-thirds of the included studies (n = 30) reported the use of additional treatments; of these, 12 reported intubation/tracheotomies (Chub-Uppakarn 2007; Eden 1967; Fitzgerald 1996; Geelhoed 1995c; Godden 1997; James 1969; Johnson 1996; Johnson 1998; Leipzig 1979; Parker 2019; Roorda 1998; Skowron 1966a; Skowron 1966a and b; Skowron 1966b);

four reported the use of antibiotics (Husby 1993; James 1969; Koren 1983; Rittichier 2000); 14 reported the use of supplemental glucocorticoids (Dobrovoljac 2012; Fifoot 2007; Fitzgerald 1996; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Johnson 1996; Klassen 1994; Klassen 1996; Klassen 1998; Parker 2019; Rittichier 2000; Roorda 1998; Super 1989; Vad Pedersen 1998); 22 reported the use of epinephrine (Amir 2006; Dobrovoljac 2012; Donaldson 2003; Duman 2005; Fifoot 2007; Fitzgerald 1996; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 2005; Godden 1997; Johnson 1996; Johnson 1998; Klassen 1994; Klassen 1996; Klassen 1998; Koren 1983; Parker 2019; Rittichier 2000; Roberts 1999; Sparrow 2006; Super 1989; Tibballs 1992); and five reported the use of a mist tent (Alshehr 2005; Johnson 1996; Klassen 1996; Rittichier 2000; Super 1989). Twenty-four studies reported collecting adverse events data, of which eight reported serious adverse events following the administration of glucocorticoids (namely secondary bacterial infections, e.g. pneumonia, otitis media) (Alshehr 2005; Bjornson 2004; Johnson 1996; Klassen 1998; Kuusela 1988; Parker 2019; Roberts 1999; Super 1989), and 16 reported no serious adverse events (Chub-Uppakarn 2007; Duman 2005; Eden 1967; Fifoot 2007; Fitzgerald 1996; Garbutt 2013; Huang 2021; Husby 1993; James 1969; Johnson 1998; Klassen 1994; Leipzig 1979; Roorda 1998; Sparrow 2006; Tibballs 1992; Vad Pedersen 1998).

Funding

The included studies received funding from government (11%), academic (7%), industry (18%), and foundations (9%) sources. However, more than half (55%) of the included studies did not report any funding sources.

Excluded studies

We excluded six studies following the searches in 2022 (Figure 1). Gursanscky 2019 and Tyler 2022 were not randomised trials; Lee 2019 and Meskina 2019 were randomised trials that did not investigate glucocorticoids; and Faraji-Goodarzi 2018 was a randomised trial that did not report any relevant outcomes. See Characteristics of excluded studies table.

We edited the excluded studies list to remove legacy excluded studies that evidently did not meet the inclusion criteria (e.g. letters, commentaries, summaries, case studies). We made this change to comply with current Cochrane standards for methods and reporting. We excluded 38 studies in this 2022 updated review.

Ongoing studies

We identified one ongoing study, IRCT20190914044765N1, and one study awaiting classification, Chen 2018 (Figure 1). We will assess these studies for inclusion in a future update.

Risk of bias in included studies

We presented the risk of bias of all included studies as assessed using the Cochrane risk of bias tool in Figure 2 and Figure 3. We judged the overall risk of bias to be low in one study (Garbutt 2013), unclear in 32 studies (Alshehr 2005; Bjornson 2004; Chub-Uppakarn 2007; Cruz 1995; Donaldson 2003; Eden 1964; Eden 1967; Fifoot 2007; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1996a; Geelhoed 2005; Godden 1997; Huang 2021; Husby 1993; James 1969; Johnson 1996; Johnson 1998; Klassen 1996; Klassen 1998; Koren 1983; Kuusela 1988; Leipzig 1979; Luria 2001; Martinez Fernandez 1993; Massicotte 1973; Parker 2019; Roorda 1998; Skowron 1966a and b; Sparrow 2006; Super 1989; Tibballs 1992; Von Mühlendahl 1982),



and high in 12 studies (Amir 2006; Cetinkaya 2004; Dobrovoljac 2012; Duman 2005; Eboriadou 2010; Fitzgerald 1996; Geelhoed 1995c; Klassen 1994; Rittichier 2000; Roberts 1999; Soleimani 2013;

Vad Pedersen 1998). Rationales for our risk of bias judgements are provided in the risk of bias tables in the Characteristics of included studies table.

Figure 2. Risk of bias graph for studies included in the 2022 update synthesis: review authors' judgements about each risk of bias item presented as percentages across all included studies.

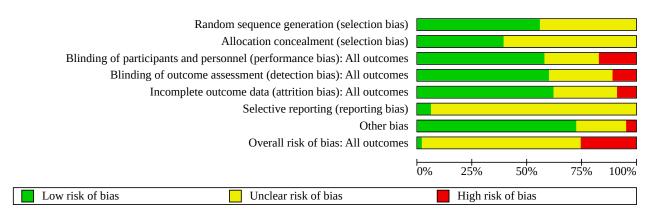




Figure 3. Risk of bias summary for studies included in the 2022 update synthesis: review authors' judgements about each risk of bias item for each included study.

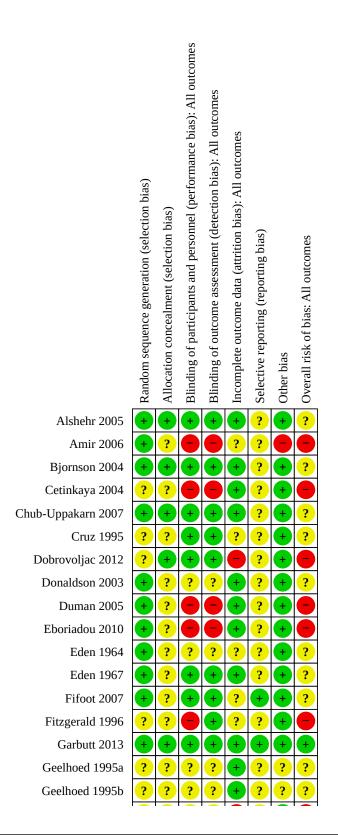
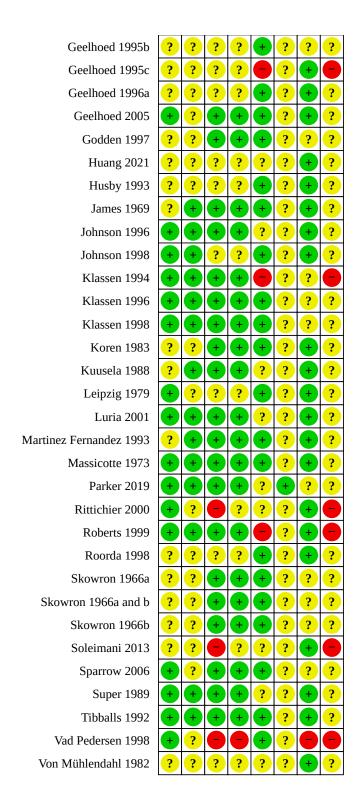




Figure 3. (Continued)



Allocation

We judged risk of bias for random sequence generation to be low in 27 studies (60%) and unclear in 18 studies (40%). The 17 studies at unclear risk of bias were described as randomised; however, the method for generating the randomisation sequence was unclear

or not reported (Cetinkaya 2004; Cruz 1995; Dobrovoljac 2012; Fitzgerald 1996; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 1996a; Godden 1997; Huang 2021; Husby 1993; James 1969; Koren 1983; Kuusela 1988; Martinez Fernandez 1993; Roorda 1998; Skowron 1966a and b; Soleimani 2013; Von Mühlendahl



1982). Randomisation was adequately described in the remaining 27 studies. We judged risk of bias for allocation concealment to be low in 19 studies (42%) and unclear in 26 studies (58%); in the latter studies, there was insufficient information reported in the publication to determine whether or not the groups to which the children were allocated could have been foreseen (Amir 2006; Cetinkaya 2004; Cruz 1995; Donaldson 2003; Duman 2005; Eboriadou 2010; Eden 1964; Eden 1967; Fifoot 2007; Fitzgerald 1996; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 1996a; Geelhoed 2005; Godden 1997; Huang 2021; Husby 1993; Koren 1983; Leipzig 1979; Rittichier 2000; Roorda 1998; Skowron 1966a and b; Soleimani 2013; Sparrow 2006; Vad Pedersen 1998; Von Mühlendahl 1982). Allocation concealment was adequate in the remaining 19 studies.

Blinding

We judged risk of bias for blinding of participants and personnel to be low in 26 studies (58%), unclear in 11 studies (24%), and high in eight studies (18%). Of the eight studies at high risk of bias, four appeared to be open-label (Amir 2006; Duman 2005; Rittichier 2000; Vad Pedersen 1998). Cetinkaya 2004 did not explicitly describe any measures taken to blind participants and personnel from treatment assignment, and any blinding could have been broken. Personnel were not blinded in Fitzgerald 1996. In Eboriadou 2010, the treatments were clearly distinguishable, and the method for blinding was not described even though the study was termed "double-blind". In Soleimani 2013, only the outcome assessor was blinded. Of the 11 studies assessed as at unclear risk of bias, seven were described as double-blind without any further details regarding who was blinded or how blinding was achieved (Eden 1964; Geelhoed 1996a; Huang 2021; Husby 1993; Leipzig 1979; Roorda 1998; Von Mühlendahl 1982). In Donaldson 2003, Geelhoed 1995a, Geelhoed 1995c, and Johnson 1998, blinding was attempted, but we judged that the blinding could have been broken; however, it was unclear how often this could have occurred. The remaining studies included satisfactory descriptions of how participants and personnel were blinded.

We judged risk of bias for blinding of outcome assessment to be low in 27 studies (60%), unclear in 13 studies (29%), and high in five studies (11%). For 22 studies (49%), there was no mention of a third-party outcome assessor, so the judgement for outcome assessment was carried over from blinding of participants and personnel (Cetinkaya 2004; Chub-Uppakarn 2007; Cruz 1995; Dobrovoljac 2012; Duman 2005; Eboriadou 2010; Eden 1964; Eden 1967; Geelhoed 1995a; Geelhoed 1995b; Geelhoed 1995c; Geelhoed 1996a; Geelhoed 2005; Godden 1997; Huang 2021; Husby 1993; Koren 1983; Kuusela 1988; Luria 2001; Martinez Fernandez 1993; Massicotte 1973; Sparrow 2006; Tibballs 1992). Of the remaining studies, we judged two as at high risk of bias because outcome assessors were not blinded (Amir 2006; Vad Pedersen 1998). We judged seven studies as at unclear risk of bias: in Donaldson 2003, Johnson 1998, and Rittichier 2000, blinding of the outcome assessors was attempted, but we judged that the blinding could have been broken, although it was unclear how often this could have occurred; the studies by Leipzig 1979, Roorda 1998, and Von Mühlendahl 1982 were described as double-blind, but it was unclear if the outcome assessors were blinded; and in Soleimani 2013, the outcome assessor was described as blinded, but it was unclear how or if the blinding could have been broken. The remaining studies provided satisfactory descriptions of how outcome assessors were blinded.

Incomplete outcome data

We judged risk of bias for incomplete outcome data to be low in 27 studies (60%), unclear in 14 studies (31%), and high in four studies (9%). The four studies at high risk of bias reported large losses to follow-up that were imbalanced between groups (Dobrovoljac 2012; Geelhoed 1995c; Klassen 1994; Roberts 1999). Dobrovoljac 2012 and Roberts 1999 used the last observation carried forward (LOCF) method to estimate endpoint outcome values. Regarding the studies at unclear risk of bias, in one study the number of children analysed was not reported (Amir 2006), and in seven studies it was either unclear to which group the children who were lost to follow-up had been allocated, or whether or not the losses to follow-up were balanced between groups (Cruz 1995; Eden 1964; Johnson 1996; Huang 2021; Kuusela 1988; Rittichier 2000; Soleimani 2013; Von Mühlendahl 1982). In four studies, losses to follow-up ranged from 13% to 17% (Fifoot 2007; Luria 2001; Soleimani 2013; Super 1989). In Fitzgerald 1996, loss to follow-up was 5%, and the LOCF method was used to estimate endpoint outcome values. In Parker 2019, 11% of participants were missing at one-hour croup assessment with unexplained exclusion reasons. We judged risk of bias due to incomplete outcome data not a concern for the remaining studies.

Selective reporting

We judged risk of bias for selective reporting to be low in three studies (7%) and unclear in 42 studies (93%). In the three studies at low risk of bias, the outcomes in the trial registers matched those reported in the publications (Fifoot 2007; Garbutt 2013; Parker 2019). For the remaining 42 studies, no protocol or trial registry was cited in the publication or located via online searches. In all cases, the outcomes reported in the methods matched those reported in the results section of the publications.

Other potential sources of bias

We judged risk of bias from other sources to be low in 35 studies (78%), unclear in eight studies (18%), and high in two studies (4%). In the two studies at high risk of bias, there was a baseline imbalance in croup score (Amir 2006; Vad Pedersen 1998). For six studies at unclear risk of bias, there was the potential for bias in participant selection because some children were not enrolled due to manpower constraints, failure of the emergency department to contact the research team, or because the emergency department was busy (Geelhoed 1995a; Geelhoed 1995b; Godden 1997; Klassen 1994; Klassen 1996; Klassen 1998; Sparrow 2006). In one study at unclear risk of bias, baseline data were not presented, therefore it was not possible to estimate whether or not baseline imbalances existed between groups (Skowron 1966a and b). For the remaining study at unclear risk of bias, participants were enrolled more than once (Parker 2019).

Effects of interventions

See: **Summary of findings 1** Any glucocorticoid compared to placebo for croup; **Summary of findings 2** Any glucocorticoid compared to epinephrine for croup

See Summary of findings 1 and Summary of findings 2.

Comparison 1: Any glucocorticoid compared to placebo

See Summary of findings 1.



Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

Compared to placebo, glucocorticoids may have resulted in greater reductions in croup score after two hours (standardised mean difference (SMD) -0.65, 95% confidence interval (CI) -1.13 to -0.18; P = 0.007, I² = 81%; 7 RCTs, 426 children; low-certainty evidence; Analysis 1.1); six hours (SMD -0.76, 95% CI -1.12 to -0.40; P < 0.001, I² = 83%; 11 RCTs, 959 children; low-certainty evidence; Analysis 1.2); and 12 hours (SMD -1.03, 95% CI -1.53 to -0.53; P < 0.001, I² = 86%; 8 RCTs, 571 children; low-certainty evidence; Analysis 1.3). The evidence for change in croup score after 24 hours is very uncertain (SMD -0.86, 95% CI -1.40 to -0.31; P = 0.002, I² = 81%; 8 RCTs, 351 children; very low-certainty evidence; Analysis 1.4).

There were no subgroup differences in reductions in croup score by score (Westley 1978 or otherwise) (Analysis 1.1; Analysis 1.2; Analysis 1.3; Analysis 1.4) or by inpatient or outpatient status (Analysis 1.5; Analysis 1.6; Analysis 1.7) at any time point. At two hours, there was no subgroup difference in effect by glucocorticoid $(Chi^2 = 5.65, P = 0.06, I^2 = 64.6\%; Analysis 1.8)$. At six hours, there was a subgroup difference in effect by glucocorticoid (Chi² = 11.46, P = 0.009, I² = 73.8%; Analysis 1.9), accounted for by the larger reduction in croup score for prednisolone (SMD -1.87, 95% CI -2.62 to -1.13; P < 0.001; 1 RCT, 42 children) compared to budesonide (SMD -0.81, 95% CI -1.04 to -0.58; P < 0.001, I² = 0%; 5 RCTs, 333 children) and dexamethasone (SMD -0.62, 95% CI -1.17 to -0.08; P = 0.03, $I^2 = 85\%$; 6 RCTs, 567 children). Fluticasone did not show an effect (SMD 0.06, 95% CI -0.89 to 1.02; P = 0.90; 1 RCT, 17 children). At 12 hours, there was a subgroup difference in effect by glucocorticoid $(Chi^2 = 10.08, P = 0.006, I^2 = 80.2\%; Analysis 1.10), accounted for$ by the larger reduction in croup score for prednisolone (SMD -2.40, 95% CI -3.26 to -1.55; P < 0.001; 1 RCT, 39 children) compared to budesonide (SMD -0.97, 95% CI -1.26 to -0.68; P < $0.001, I^2 = 0\%$; 3 RCTs, 209 children) and dexamethasone (SMD -0.85, 95% CI -1.55 to -0.15; P = 0.02, I^2 = 84%; 5 RCTs, 323 children). At 24 hours, there was a subgroup difference in effect by glucocorticoid (Chi² = 9.02, P = 0.01, I² = 77.8%; Analysis 1.11). Although larger reductions in croup score were observed with budesonide (SMD -1.40, 95% CI -1.88 to -0.93; P < 0.001, I² = 0%; 2 RCTs, 89 children) and dexamethasone (SMD -0.89, 95% CI -1.55 to -0.22; P = 0.009, I² = 81%; 6 RCTs, 245 children) compared to placebo, fluticasone did not show an effect (SMD 0.21, 95% CI -0.75 to 1.17; P = 0.67; 1 RCT, 17 children).

2. Return visits or (re)admissions to the hospital or both

Compared to placebo, glucocorticoids may have reduced the rate of return visits or (re)admissions to the hospital or both by almost half (risk ratio (RR) 0.52, 95% CI 0.36 to 0.75; P < 0.001, $I^2 = 52\%$; 10 RCTs, 1679 children; low-certainty evidence; Analysis 1.12). There were no subgroup differences in effect by glucocorticoid (budesonide or dexamethasone, Analysis 1.13); by inpatient or outpatient status (Analysis 1.12); or by croup severity (mild or moderate croup, Analysis 1.14).

The number needed to treat for an additional beneficial outcome (NNTB) is presented in Table 1. The NNTB was 7 children (95% CI 5 to 12) for the mean placebo group rate (30.62%). The NNTB was 102 children (95% CI 78 to 179) for the smallest placebo group rate (2.06%). Lastly, the NNTB was 3 children (95% CI 2 to 5) for the largest placebo group rate (72.00%).

Secondary outcomes

1. Length of stay in the hospital or emergency department

Compared to those given a placebo, children treated with glucocorticoids spent fewer hours in the hospital (mean difference (MD) -14.90, 95% CI -23.58 to -6.22; P < 0.001, I² = 54%; 8 RCTs, 476 children; Analysis 1.15). All of the included studies investigated inpatients. There was no subgroup difference in effect by glucocorticoid (budesonide, dexamethasone, or fluticasone; Analysis 1.16).

2. Patient improvement at 2, 6, 12, and/or 24 hours

Only one study investigated patient improvement two hours after the administration of glucocorticoids compared to placebo. Roberts 1999 studied 82 hospitalised children aged six months to eight years with moderate to severe croup who were given budesonide or placebo, and observed no difference in improvement after two hours (RR 1.81, 95% CI 0.96 to 3.40; P = 0.07; 1 RCT, 82 children; Analysis 1.17). Compared to placebo, glucocorticoids were associated with improvement in a greater proportion of children after six hours (RR 1.45, 95% CI 1.12 to 1.88; P = 0.005, $I^2 = 34\%$; 6 RCTs, 332 children; Analysis 1.18); 12 hours (RR 1.33, 95% CI 1.09 to 1.62; P = 0.005, $I^2 = 53\%$; 6 RCTs, 340 children; Analysis 1.19); and 24 hours (RR 1.28, 95% CI 1.01 to 1.61; P = 0.04, $I^2 = 75\%$; 5 RCTs, 251 children; Analysis 1.20).

Only inpatients were included in the 12-hour analysis (Analysis 1.19). There were no subgroup differences in estimates of effect by inpatient or outpatient status at six or 24 hours (Analysis 1.18; Analysis 1.20). There were no subgroup differences in effect by glucocorticoid at six hours (budesonide, dexamethasone, or prednisolone; Analysis 1.21), 12 hours (budesonide, dexamethasone, or prednisolone; Analysis 1.22), or 24 hours (dexamethasone or prednisolone; Analysis 1.23).

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

There was no difference between children treated with glucocorticoids and those given placebo in the use of antibiotics (risk difference (RD) 0.00, 95% CI –0.04 to 0.04; P = 1.00, I² = 0%; 3 RCTs, 202 children; Analysis 1.24); the use of epinephrine (RD –0.03, 95% CI –0.08 to 0.01; P = 0.16, I² = 45%; 9 RCTs, 709 children; Analysis 1.25); the rate of intubation/tracheostomy (RD 0.00, 95% CI –0.01 to 0.01; P = 0.79, I² = 0%; 11 RCTs, 1090 children; Analysis 1.26); the use of a mist tent (RD –0.20, 95% CI –0.87 to 0.47; P = 0.55, I² = 95%; 2 RCTs, 84 children; Analysis 1.27); or the use of supplemental glucocorticoids (RR 0.61, 95% CI 0.36 to 1.03; P = 0.07, I² = 10%; 6 RCTs, 305 children; Analysis 1.28).

4. Any adverse events

Of the 26 studies that investigated any glucocorticoid compared to placebo, 13 reported collecting adverse events data. Of these, eight reported no serious adverse events (Eden 1967; Husby 1993; James 1969; Johnson 1998; Klassen 1994; Leipzig 1979; Roorda 1998; Tibballs 1992). Bjornson 2004 reported seven instances of pneumonia (3/359, 0.83% in the dexamethasone group and 4/361, 1.11% in the placebo group). Johnson 1996 reported one child with neutropenia consistent with bacterial tracheitis in the dexamethasone group (1/28, 3.57%). Kuusela 1988 reported seven secondary bacterial infections (pneumonia, sinusitis, otitis media) requiring antibiotic therapy: 5/35, 14% in the dexamethasone



group and 2/16, 12.5% in the placebo group. Super 1989 reported one child with pneumonitis in the placebo group (1/13, 7.7%) and two children with pneumonia in the dexamethasone group (2/16, 12.5%). Roberts 1999 reported one instance of exacerbated symptoms, five children with emotional distress, two with vomiting, and one instance of eye irritation in the budesonide group (9/42, 21.4%), and three instances of exacerbated symptoms, six children with emotional distress, three with vomiting, two rashes, and one instance each of eye irritation and tongue irritation in the placebo group (16/40, 40%).

Comparison 2: Any glucocorticoid compared to epinephrine

See Summary of findings 2.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24

Compared to epinephrine, we do not know if there was no difference in the change in croup score following treatment with glucocorticoids after two hours (SMD 0.77, 95% CI -0.24 to 1.77; P = 0.13, I² = 87%; 2 RCTs, 130 children; very low-certainty evidence; Analysis 2.1) and six hours (SMD -0.10, 95% CI -1.18 to 0.97; P = 0.85, I² = 78%; 2 RCTs, 63 children; very low-certainty evidence; Analysis 2.2). There may be no difference between groups at 12 hours (SMD -0.07, 95% CI -0.57 to 0.43; P = 0.78, I² = 47%; 3 RCTs, 129 children; low-certainty evidence; Analysis 2.3) or 24 hours (SMD 0.17, 95% CI -0.18 to 0.51; P = 0.35, I² = 0%; 3 RCTs, 129 children; low-certainty evidence; Analysis 2.4).

The analyses at six hours (Analysis 2.2), 12 hours (Analysis 2.3), and 24 hours (Analysis 2.4) included only inpatients. At two hours, there was a subgroup difference in effect by inpatient or outpatient status (Chi² = 7.44, P = 0.006, I² = 86.6%; Analysis 2.1). For outpatients, glucocorticoids were less effective at reducing croup score compared to epinephrine after two hours (SMD 1.29, 95% CI 0.73 to 1.84; P < 0.001; 1 RCT, 64 children). No difference was detected between the two treatments for inpatients (SMD 0.26, 95% CI –0.22 to 0.75; P = 0.29; 1 RCT, 66 children).

At two hours, there was a subgroup difference in effect by glucocorticoid ($\text{Chi}^2=7.37,\ P=0.03,\ I^2=72.9\%;\ \text{Analysis}\ 2.5$). Epinephrine was more effective at reducing croup score compared to beclomethasone (SMD 1.41, 95% CI 0.62 to 2.19; P < 0.001; 1 RCT, 33 children) and dexamethasone (SMD 1.13, 95% CI 0.35 to 1.91; P = 0.005; 1 RCT, 31 children). At this time point, there was no difference in reduction in croup score between budesonide and epinephrine (SMD 0.26, 95% CI –0.22 to 0.75; P = 0.29; 1 RCT, 66 children). The 12- and 24-hour analyses investigated budesonide and dexamethasone, and there were no subgroup differences in effect (Analysis 2.6; Analysis 2.7).

2. Return visits or (re)admissions to the hospital or both

Eboriadou 2010 and Fitzgerald 1996 investigated return visits and (re)admissions, respectively, following the administration of glucocorticoids (dexamethasone and beclomethasone, and budesonide, respectively) compared to epinephrine. Both studies may not have reported any events (RD 0.00, 95% CI -0.04 to 0.04; P = 1.00, I² = 0%; 2 RCTs, 130 children; low-certainty evidence; Analysis 2.8).

Secondary outcomes

1. Length of stay in the hospital or emergency department

Kuusela 1988 investigated length of stay for 32 children hospitalised with croup who were treated with dexamethasone, epinephrine, a combination of dexamethasone and epinephrine, or placebo. There was no difference in hours spent in the hospital between children treated with dexamethasone and those treated with epinephrine (MD -10.00, 95% CI -33.89 to 13.89; P = 0.41; 1 RCT, 32 children; Analysis 2.9).

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies reported on patient improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

Fitzgerald 1996 investigated the use of additional treatments for children aged six months to six years admitted to the hospital with croup who were treated with budesonide or epinephrine. There was no difference in the proportion of children who required additional epinephrine between groups (RR 0.30, 95% CI 0.03 to 2.69; P = 0.28; 1 RCT, 66 children; Analysis 2.10). No child was intubated (RD 0.00, 95% CI -0.06 to 0.06; P = 1.00; 1 RCT, 66 children; Analysis 2.11). There was no difference between groups in the proportion of children who required supplemental glucocorticoids (RR 0.83, 95% CI 0.48 to 1.43; P = 0.49; 1 RCT, 66 children; Analysis 2.12).

4. Any adverse events

Of the four studies that investigated glucocorticoids compared to epinephrine, three reported collecting adverse events data. Fitzgerald 1996 reported no serious adverse events. Kuusela 1988 reported five cases of secondary bacterial infections (pneumonia, sinusitis, otitis media) requiring antibiotic therapy in the dexamethasone group (5/16, 31.3%). Eboriadou 2010 reported four cases of tremor and tachycardia (4/25, 16%) in the epinephrine group.

Comparison 3: Dexamethasone compared to budesonide

See Table 2.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

Compared to budesonide, dexamethasone may have resulted in a greater reduction in croup score after six hours (SMD -0.46, 95% CI -0.79 to -0.13; P = 0.006, I² = 51%; 4 RCTs, 326 children; low-certainty evidence; Analysis 3.1) and 12 hours (SMD -0.75, 95% CI -1.19 to -0.30; P = 0.001, I² = 0%; 2 RCTs, 84 children; low-certainty evidence; Analysis 3.2). The analysis at 12 hours included only inpatients (Analysis 3.2). At six hours, there was no subgroup difference in effect by inpatient or outpatient status (Analysis 3.1).

2. Return visits or (re)admissions to the hospital or both

There was probably no difference in the rate of return visits or (re)admissions to the hospital (or both) between dexamethasone and budesonide groups (RR 0.69, 95% CI 0.40 to 1.22; P = 0.20, I² = 0%; 5 RCTs, 374 children; moderate-certainty evidence; Analysis



3.3). There were no subgroup differences in effect by inpatient or outpatient status (Analysis 3.3).

Secondary outcomes

1. Length of stay in the hospital or emergency department

There was no difference in hours spent in the hospital or emergency department between children treated with dexamethasone and those treated with budesonide (MD -0.51, 95% CI -1.28 to 0.25; P = 0.15, I² = 51%; 2 RCTs, 184 children; Analysis 3.4). There was no subgroup difference in effect by inpatient or outpatient status (Analysis 3.4).

2. Patient improvement at 2, 6, 12, and/or 24 hours

Klassen 1998 investigated response to treatment, defined as a two-point improvement in croup score, amongst 198 children aged three months to five years who were treated with budesonide, dexamethasone, or a combination of budesonide and dexamethasone in the emergency department for croup. There was no difference in response to treatment between those treated with dexamethasone and those treated with budesonide (RR 1.12, 95% CI 0.93 to 1.34; P = 0.22; 1 RCT, 134 children; Analysis 3.5).

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

Compared to those treated with budesonide, children treated with dexamethasone were at a reduced risk of needing treatment with epinephrine (RR 0.45, 95% CI 0.21 to 0.96; P = 0.04, I^2 = 0%; 4 RCTs, 321 children; Analysis 3.6). Geelhoed 1995c and Johnson 1998 investigated the need for intubation/tracheostomy amongst children treated with dexamethasone or budesonide for croup. There were no events in either study (RD 0.00, 95% CI -0.04 to 0.04; P = 1.00, I^2 = 0%; 2 RCTs, 145 children; Analysis 3.7). There was no difference in the need for additional glucocorticoids between children treated with dexamethasone and those treated with budesonide (RR 0.48, 95% CI 0.18 to 1.32; P = 0.15, I^2 = 0%; 3 RCTs, 240 children; Analysis 3.8).

4. Any adverse events

Of the six studies investigating dexamethasone compared to budesonide, three (50%) reported no serious adverse events (Duman 2005; Johnson 1998; Vad Pedersen 1998). Klassen 1998 reported one case of oral thrush in the budesonide group (1/65, 1.5%) and one case each of hives and violent behaviour in the dexamethasone group (2/69, 2.9%).

Comparison 4: Dexamethasone compared to beclomethasone

See Table 3.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

No studies investigated the change in croup score for this comparison.

2. Return visits or (re)admissions to the hospital or both

Eboriadou 2010 investigated return visits to the emergency department amongst 39 children aged six months to five years treated with dexamethasone or beclomethasone for croup. There

were probably no children that returned for additional care (RD 0.00, 95% CI -0.09 to 0.09; P = 1.00; 1 RCT, 39 children; moderate-certainty evidence; Analysis 4.1).

Secondary outcomes

1. Length of stay in the hospital or emergency department

No studies investigated length of stay in the hospital or emergency department for this comparison.

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies investigated clinical improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

No studies investigated the use of additional treatments for this comparison.

4. Any adverse events

Eboriadou 2010 investigated this comparison and reported no adverse events related to the glucocorticoids.

Comparison 5: Dexamethasone compared to betamethasone

See Table 4.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

Amir 2006 investigated reduction in croup score for 52 children aged six months to six years were treated with dexamethasone or betamethasone in the emergency department for croup. Compared to betamethasone, dexamethasone may have resulted in a greater reduction in croup score after two hours (SMD -0.62, 95% CI -1.17 to -0.06; P = 0.03; 1 RCT, 52 children; low-certainty evidence; Analysis 5.1) and six hours (SMD -0.67, 95% CI -1.23 to -0.11; P = 0.02; 1 RCT, 52 children; low-certainty evidence; Analysis 5.2).

2. Return visits or (re)admissions to the hospital or both

Amir 2006 investigated re-examinations by a primary care physician amongst 52 children aged six months to six years treated with dexamethasone or betamethasone in the emergency department for croup. There may have been no difference in the rate of re-examinations between dexamethasone and betamethasone groups (RR 0.95, 95% CI 0.67 to 1.34; P = 0.76; 1 RCT, 52 children; low-certainty evidence; Analysis 5.3).

Secondary outcomes

1. Length of stay in the hospital or emergency department

No studies investigated length of stay in the hospital or emergency department for this comparison.

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies investigated clinical improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

Amir 2006 investigated the need for epinephrine amongst 52 children aged six months to six years treated with dexamethasone



or betamethasone in the emergency department for croup. The risk for needing epinephrine was higher for those treated with dexamethasone compared to those treated with betamethasone (RR 2.11, 95% CI 1.18 to 3.76; P = 0.01; 1 RCT, 52 children; Analysis 5.4).

4. Any adverse events

No studies investigated adverse events for this comparison.

Comparison 6: Dexamethasone compared to prednisolone

See Table 5.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

Two trials compared the effect of dexamethasone versus prednisolone. Parker 2019 found little to no difference at two-hour postbaseline croup score assessment (SMD 0.06, 95% CI –0.06 to 0.18; P = 0.32; 1 RCT, 1231 children; high-certainty evidence; Analysis 6.1). Fifoot 2007 found likely little to no difference between groups at six-hour postbaseline croup score assessment (SMD 0.21, 95% CI –0.21 to 0.62; P = 0.33; 1 RCT, 99 children; moderate-certainty evidence; Analysis 6.2).

2. Return visits or (re)admissions to the hospital or both

Dexamethasone probably reduced the rate of return visits or readmissions for croup by about half when compared to prednisolone (RR 0.55, 95% CI 0.28 to 1.11; P = 0.09, I² = 59%; 4 RCTs, 1537 children; moderate-certainty evidence; Analysis 6.3).

Secondary outcomes

1. Length of stay in the hospital or emergency department

There was little to no difference between groups in length of stay in the hospital or emergency department (MD -0.02, 95% CI -0.42 to 0.39; P = 0.94, I² = 12%; 2 RCTs, 1363 children; Analysis 6.4).

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies reported on patient improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

We found no difference in the addition of epinephrine to the treatment received by children treated with dexamethasone versus those treated with prednisolone (RR 0.90, 95% CI 0.50 to 1.64; P = 0.74, I² = 0%; 3 RCTs, 1463 children; high-certainty evidence; Analysis 6.5). No child required intubation (RD 0.00, 95% CI –0.00 to 0.00; P = 1.00; 1 RCT, 1231 children; Analysis 6.6). However, there was a 28% reduction in the use of supplemental glucocorticoids as an additional treatment between children who received dexamethasone and those who received prednisolone (RR 0.72, 95% CI 0.53 to 0.97; P = 0.03, I² = 0%; 2 RCTs, 926 children; Analysis 6.7).

4. Any adverse events

Although not specific to this comparison, Parker 2019 reported a few adverse events in four participants: one participant in the dexamethasone group had a febrile convulsion, and one participant in the prednisolone group had insomnia. Unlike Parker

2019, three trials did not report serious adverse events (Fifoot 2007; Garbutt 2013; Sparrow 2006).

Comparison 7: Budesonide compared to dexamethasone

See Table 6.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

No studies reported on change in clinical croup score for this comparison.

2. Return visits or (re)admissions to the hospital or both

No studies reported on return visits or (re)admission to the hospital (or both) for this comparison.

Secondary outcomes

1. Length of stay in the hospital or emergency department

No studies reported on length of stay in the hospital or emergency department for this comparison.

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies reported on patient improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

No studies reported on the use of additional treatments for this comparison.

4. Any adverse events

Huang 2021 investigated the effect of inhaled budesonide versus dexamethasone in children with acute infectious laryngitis and found no adverse condition following treatment. The study authors did not report on any other outcomes relevant to this review.

Comparison 8: Budesonide and dexamethasone compared to dexamethasone

See Table 7.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

There was probably no difference in reduction in croup score after six hours between children treated with combined dexamethasone and budesonide versus those treated with dexamethasone alone (SMD 0.05, 95% CI -0.19 to 0.30; P = 0.67, I² = 0%; 3 RCTs, 255 children; moderate-certainty evidence; Analysis 7.1). There was no difference in effect by inpatient or outpatient status (Analysis 7.1).

2. Return visits or (re)admissions to the hospital or both

There may have been no difference in the rate of admissions or return visits between children treated with combined dexamethasone and budesonide versus those treated with dexamethasone alone (RR 0.91, 95% CI 0.45 to 1.83; P = 0.79, I² = 0%; 3 RCTs, 254 children; low-certainty evidence; Analysis 7.2). There was no subgroup difference in effect by inpatient or outpatient status (Analysis 7.2).



Secondary outcomes

1. Length of stay in the hospital or emergency department

There was no difference in hours spent in the hospital or emergency department amongst children treated with dexamethasone and budesonide versus those treated with dexamethasone alone (MD 0.44, 95% CI -0.05 to 0.92; P = 0.08, I^2 = 0%; 2 RCTs, 204 children; Analysis 7.3). There were no subgroup differences in effect by inpatient or outpatient status (Analysis 7.3).

2. Patient improvement at 2, 6, 12, and/or 24 hours

After six hours, there was no difference in the clinical improvement of children treated with dexamethasone and budesonide versus those treated with dexamethasone alone (RR 1.11, 95% CI 0.65 to 1.90; P = 0.70; 2 RCTs, 183 children; Analysis 7.4). This analysis only included outpatients (Analysis 7.4).

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

There was no difference in the need for epinephrine (RR 1.42, 95% CI 0.27 to 7.39; P = 0.67, I² = 0%; 2 RCTs, 183 children; Analysis 7.5); a mist tent (RR 1.07, 95% CI 0.69 to 1.65; P = 0.77; 1 RCT, 50 children; Analysis 7.6); or supplemental glucocorticoids (RR 1.10, 95% CI 0.07 to 16.66; P = 0.95, I² = 66%; 2 RCTs, 182 children; Analysis 7.7) amongst children treated with dexamethasone and budesonide versus those treated with dexamethasone alone.

4. Any adverse events

Klassen 1998 reported no adverse events in either the dexamethasone group or the dexamethasone and budesonide group.

Comparison 9: Budesonide and dexamethasone compared to budesonide

See Table 8.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24

Klassen 1998 investigated children aged three months to five years treated in the emergency department with dexamethasone, budesonide, or a combination of the two for croup. There was probably no difference in reduction in croup score after six hours amongst children treated with combined dexamethasone and budesonide versus those treated with budesonide alone (SMD -0.18, 95% CI -0.52 to 0.17; P = 0.32; 1 RCT, 129 children; moderate-certainty evidence; Analysis 8.1).

2. Return visits or (re)admissions to the hospital or both

Klassen 1998 investigated return visits to the emergency department amongst children aged three months to five years treated with dexamethasone, budesonide, or a combination of the two for croup. There were probably no events in either treatment group (RD 0.00, 95% CI -0.03 to 0.03; P = 1.00; 1 RCT, 129 children; moderate-certainty evidence; Analysis 8.2).

Secondary outcomes

1. Length of stay in the hospital or emergency department

Klassen 1998 investigated hours spent in the emergency department amongst children aged three months to five years treated with dexamethasone, budesonide, or a combination of the two for croup. There was no difference in length of stay amongst children treated with dexamethasone and budesonide versus those treated with budesonide alone (MD 0.25, 95% CI –0.36 to 0.86; P = 0.42; 1 RCT, 129 children; Analysis 8.3).

2. Patient improvement at 2, 6, 12, and/or 24 hours

Klassen 1998 investigated response to treatment, defined as a two-point reduction in croup score, amongst children aged three months to five years treated in the emergency department with dexamethasone, budesonide, or a combination of the two for croup. There was no difference in response to treatment amongst children treated with dexamethasone and budesonide versus those treated with budesonide alone (RR 0.97, 95% CI 0.79 to 1.20; P = 0.80; 1 RCT, 129 children; Analysis 8.4).

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

Klassen 1998 investigated the need for additional treatments amongst children aged three months to five years treated in the emergency department with dexamethasone, budesonide, or a combination of the two for croup. There was no difference between groups in the need for epinephrine (RR 1.02, 95% CI 0.15 to 6.99; P = 0.99; 1 RCT, 129 children; Analysis 8.5) or supplemental glucocorticoids (RR 1.31, 95% CI 0.52 to 3.29; P = 0.57; 1 RCT, 129 children; Analysis 8.6).

4. Any adverse events

Klassen 1998 reported one case of oral thrush in the budesonide group (1/65, 1.5%) and no adverse events in the dexamethasone and budesonide group.

Comparison 10: Oral compared to intramuscular dexamethasone

See Table 9.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

No studies investigated change in croup score for this comparison.

2. Return visits or (re)admissions to the hospital or both

There was probably no difference in the rate of return visits or admissions following treatment with oral dexamethasone compared to intramuscular dexamethasone (RR 0.81, 95% CI 0.58 to 1.12; P=0.21, $I^2=0\%$; 3 RCTs, 440 children; moderate-certainty evidence; Analysis 9.1). The analysis only included outpatients (Analysis 9.1).

Secondary outcomes

1. Length of stay in the hospital or emergency department

No studies investigated length of stay in the hospital or emergency department for this comparison.



2. Patient improvement at 2, 6, 12, and/or 24 hours

Donaldson 2003 investigated clinical improvement, defined as parents' assessment that their child's condition had improved at least somewhat after 24 hours, amongst children aged three to 84 months treated in the emergency department with oral or intramuscular dexamethasone for croup. There was no difference between groups in rate of clinical improvement (RR 1.07, 95% CI 0.95 to 1.19; P = 0.27; 1 RCT, 95 children; Analysis 9.2).

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

There was no difference in the need for antibiotics (RR 0.14, 95% CI 0.02 to 1.15; P = 0.07; 1 RCT, 277 children; Analysis 9.3); epinephrine (RR 0.94, 95% CI 0.71 to 1.24; P = 0.64, I 2 = 0%; 2 RCTs, 372 children; Analysis 9.4); a mist tent (RR 1.34, 95% CI 0.31 to 5.89; P = 0.70; 1 RCT, 277 children; Analysis 9.5); or supplemental glucocorticoids (RR 1.10, 95% CI 0.50 to 2.41; P = 0.81; 1 RCT, 277 children; Analysis 9.6) amongst children treated with oral dexamethasone versus those treated with intramuscular dexamethasone.

4. Any adverse events

No studies investigated adverse events for this comparison.

Comparison 11: Oral compared to nebulised dexamethasone

See Table 10.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

No studies investigated change in croup score for this comparison.

2. Return visits or (re)admissions to the hospital or both

Luria 2001 investigated returns to medical care of children aged six months to six years following treatment with oral or nebulised dexamethasone in the emergency department for croup. There were probably fewer return visits to medical care amongst those treated with oral dexamethasone versus those treated with nebulised dexamethasone (RR 0.39, 95% CI 0.17 to 0.89; P = 0.03; 1 RCT, 176 children; moderate-certainty evidence; Analysis 10.1).

Secondary outcomes

1. Length of stay in the hospital or emergency department

No studies investigated length of stay in the hospital or emergency department for this comparison.

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies investigated clinical improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

No studies investigated the use of additional treatments for this comparison.

4. Any adverse events

No studies investigated adverse events for this comparison.

Comparison 12: Dexamethasone 0.30 mg/kg compared to 0.15 mg/kg

See Table 11.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

No studies investigated change in croup score for this comparison.

2. Return visits or (re)admissions to the hospital or both

Geelhoed 1995b investigated re-presentations to medical care for croup amongst children aged greater than three months treated in the emergency department with 0.30 mg/kg or 0.15 mg/kg dexamethasone. There may have been no difference in the rate of re-presentations to medical care between groups (RR 0.94, 95% CI 0.06 to 14.27; P = 0.96; 1 RCT, 60 children; low-certainty evidence; Analysis 11.1).

Secondary outcomes

1. Length of stay in the hospital or emergency department

No studies investigated length of stay in the hospital or emergency department for this comparison.

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies investigated clinical improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

Geelhoed 1995b investigated the need for additional treatments amongst children aged greater than three months treated in the emergency department with 0.30 mg/kg or 0.15 mg/kg dexamethasone for croup. There was no difference between the two treatments in the need for epinephrine (RR 0.43, 95% CI 0.19 to 0.98; P = 0.05; 1 RCT, 60 children; Analysis 11.2). No child required supplemental glucocorticoids (RD 0.00, 95% CI -0.06 to 0.06; P = 1.00; 1 RCT, 60 children; Analysis 11.3).

4. Any adverse events

No studies investigated adverse events for this comparison.

Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.30 mg/kg $\,$

See Table 12.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

No studies investigated change in croup score for this comparison.

2. Return visits or (re)admissions to the hospital or both

Geelhoed 1995a investigated re-presentations to medical care for croup amongst children aged greater than three months treated in the emergency department with 0.60 mg/kg or 0.30 mg/kg dexamethasone. There may have been no difference in the rate of re-presentations to medical care amongst children treated with 0.60 mg/kg versus 0.30 mg/kg dexamethasone (RR 1.40, 95% CI 0.25



to 7.81; P = 0.70; 1 RCT, 60 children; low-certainty evidence; Analysis 12.1).

Secondary outcomes

1. Length of stay in the hospital or emergency department

No studies investigated length of stay in the hospital or emergency department for this comparison.

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies investigated clinical improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

Geelhoed 1995a investigated the need for additional treatments amongst children aged greater than three months treated in the emergency department with 0.60 mg/kg or 0.30 mg/kg dexamethasone for croup. There was no difference between the two treatments in the need for epinephrine (RR 0.78, 95% CI 0.27 to 2.28; P = 0.65; 1 RCT, 60 children; Analysis 12.2) or supplemental glucocorticoids (RR 2.81, 95% CI 0.12 to 66.40; P = 0.52; 1 RCT, 60 children; Analysis 12.3).

4. Any adverse events

No studies investigated adverse events for this comparison.

Comparison 14: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg

See Table 13.

Primary outcomes

1. Change in clinical croup score from baseline to 2, 6, 12, and/or 24 hours

There was no reduction at two hours (SMD -0.27, 95% CI -0.76 to 0.22; P = 0.28, I² = 62%; 2 RCTs, 861 children; high-certainty evidence; Analysis 13.1). There was probably no difference at six hours (SMD -0.45, 95% CI -1.26 to 0.35; P = 0.27, I² = 85%; 3 RCTs, 178 children; moderate-certainty evidence; Analysis 13.2). We do not know if there was no difference at 12 hours (SMD -0.60, 95% CI -4.39 to 3.19; P = 0.76, I² = 98%; 2 RCTs, 113 children; very low-certainty evidence; Analysis 13.3). We found that treating children with 0.60 mg/kg versus 0.15 mg/kg dose of dexamethasone probably reduced the severity of croup scores at 24-hour postbaseline score (SMD 0.63, 95% CI 0.16 to 1.10; P = 0.009; 1 RCT, 72 children; moderate-certainty evidence; Analysis 13.4).

2. Return visits or (re)admissions to the hospital or both

There was little to no difference in return visits or (re)admissions or both between children treated with dexamethasone 0.60 mg/kg versus 0.15 mg/kg (RR 0.91, 95% CI 0.71 to 1.17; P = 0.48, $I^2 = 0\%$; 3 RCTs, 949 children; high-certainty evidence; Analysis 13.5).

Secondary outcomes

1. Length of stay in the hospital or emergency department

There was little to no difference in hours spent at the outpatient clinic between children treated with dexamethasone 0.60 mg/kg versus 0.15 mg/kg (MD 0.12, 95% CI -0.32 to 0.56; P = 0.59, I² = 0%; 2 RCTs, 892 children; Analysis 13.6).

2. Patient improvement at 2, 6, 12, and/or 24 hours

No studies reported on patient improvement for this comparison.

3. The use of additional treatments, including: antibiotics, epinephrine, intubation/tracheostomy, mist tent, and/or supplemental glucocorticoids

We found no difference in the need of additional treatments between treating children with croup with dexamethasone at 0.60 mg/kg versus 0.15 mg/kg in cases of epinephrine use (RR 0.78, 95% CI 0.34 to 1.75; P = 0.54, I² = 0%; 2 RCTs, 885 children; Analysis 13.7); intubation (RD 0.00, 95% CI -0.00 to 0.00; P = 1.00, I² = 0%; 2 RCTs, 861 children; Analysis 13.8); or use of supplemental glucocorticoids (RR 0.77, 95% CI 0.51 to 1.15; P = 0.19, I² = 0%; 2 RCTs, 617 children; Analysis 13.9).

4. Any adverse events

Of the four studies investigating dexamethasone at 0.60 mg/kg versus dexamethasone at 0.15 mg/kg, two (50%) reported no adverse events in either treatment group (Chub-Uppakarn 2007; Fifoot 2007). Parker 2019 reported 16 cases of vomiting (16/410, 4.0%) and one case of 30 seconds of febrile convulsion (1/410, 0.2%) in the 0.60 mg/kg dexamethasone group, and 13 cases of vomiting (13/410, 3.3%), one case of stridor (1/410, 0.2%), and one case of hyperactivity 30 minutes after the dose (1/410, 0.2%) in the 0.15 mg/kg dexamethasone group. Alshehr 2005 reported one case of bacterial tracheitis and two cases of bronchopneumonia in the 0.60 mg/kg dexamethasone group (3/36, 8.3%), and no adverse events in the 0.15 mg/kg dexamethasone group.

Publication bias

The publication bias for change in croup score (at six hours), and return visits or (re)admissions to the hospital (or both), for glucocorticoids compared to placebo remains the same as in the previous version of this review (Gates 2018). Insufficient numbers of included studies precluded testing for publication bias for any of the other comparisons or outcomes.

DISCUSSION

This updated review includes 45 RCTs with a total of 5888 children. This is an increase of two RCTs with 1323 children from the last update (Huang 2021; Parker 2019). Parker 2019 reported relevant data to update the existing evidence on the effects of glucocorticoid compared to prednisolone, and the optimal dosage of glucocorticoid for the treatment of croup. Huang 2021 reported only on adverse events related to the use of budesonide compared to dexamethasone for croup.

Summary of main results

Any glucocorticoid compared to placebo

Twenty-six studies investigated glucocorticoids compared to placebo. Glucocorticoids may have reduced the symptoms of croup within two hours of treatment, with the effect lasting at least 24 hours. The effect was dependent on the glucocorticoid administered. Budesonide and dexamethasone reduced the symptoms of croup within two hours of treatment, with the effect lasting at least 24 hours. One trial showed that prednisolone reduced the symptoms of croup within six hours, with the effect lasting at least 12 hours. One trial showed that fluticasone did not reduce the symptoms of croup after two, six, or 24 hours



compared to placebo. The certainty of the evidence for the effect of glucocorticoids compared to placebo for reducing the symptoms of croup from two to 12 hours was low, as there was considerable between-study heterogeneity in effect estimates (Summary of findings 1). The certainty of the evidence for the effect of glucocorticoids compared to placebo for reducing the symptoms of croup after 24 hours was very low, as there was considerable between-study heterogeneity in the magnitude and direction of the effect (Summary of findings 1).

Compared to placebo, both budesonide and dexamethasone may have reduced the rate of return visits and/or (re)admissions to the hospital or emergency department. The certainty of the evidence for the effect of glucocorticoids in reducing the rate of return visits or (re)admissions or both was low, as there was considerable between-study heterogeneity in the effect estimates (Summary of findings 1).

Compared to placebo, glucocorticoids reduced length of stay in the hospital by approximately 15 hours and resulted in clinical improvement in a greater proportion of children after six hours. The effect lasted at least 24 hours. There was little to no difference in the need for additional treatments between children treated with glucocorticoids and those treated with placebo. Treatment with glucocorticoids was infrequently associated with serious adverse events.

Any glucocorticoid compared to epinephrine

Four studies investigated glucocorticoids compared to epinephrine. We do not know if there was no difference in the reduction in symptoms of croup for children treated with epinephrine compared to those treated with glucocorticoids at two and six hours. There may not have been a reduction in croup symptoms at 12 or 24 hours following administration of glucocorticoids or epinephrine. After two hours, the effect was dependent on the glucocorticoid administered. Epinephrine resulted in greater reductions in symptoms of croup compared to beclomethasone and dexamethasone. There was little no difference in reduction in croup symptoms between epinephrine and budesonide two hours after treatment. The certainty of the evidence for the effect of glucocorticoids compared to epinephrine for reducing the symptoms of croup was very low to low. The sample sizes for the six-, 12-, and 24-hour analyses were small, and there was considerable between-study heterogeneity in effect estimates for the six-hour analysis. There was considerable between-study heterogeneity in the magnitude and direction of the effect estimates for the two-hour analysis; the sample size for the comparison was small; and the pooled effect estimate was imprecise (Summary of findings 2).

There may have been no difference in the rate of return visits or (re)admissions or both following treatment with glucocorticoids compared with epinephrine. The certainty of the evidence for the effect of glucocorticoids compared to epinephrine for reducing the rate of return visits or (re)admissions or both was low, as the sample size did not meet the optimal information size, and the contributing studies reported no events (Summary of findings 2).

There was little to no difference in length of stay in the hospital for children treated with glucocorticoids compared to those treated with epinephrine, nor were there any differences between groups in the need for additional treatments. One study reported a 31.3% rate

of secondary bacterial infections amongst children treated with dexamethasone. Another study reported a 16% rate of tremor and tachycardia amongst children treated with epinephrine.

One glucocorticoid compared to another glucocorticoid

Thirteen studies investigated one glucocorticoid compared to another glucocorticoid. Compared to budesonide, dexamethasone may have resulted in greater reductions in symptoms of croup after six and 12 hours. The certainty of the evidence for the effect of dexamethasone compared to budesonide for reducing the symptoms of croup was low, as the contributing studies were all at high or unclear risk of bias; there was substantial between-study heterogeneity in effect estimates for the sixhour analysis; and the sample size did not meet the optimal information size for the 12-hour analysis (Table 2). Compared to betamethasone, dexamethasone may have resulted in greater reductions in symptoms of croup after two and six hours. The certainty of the evidence for the effect of dexamethasone compared to betamethasone for reducing the symptoms of croup was low, as the only study contributing to the analysis was at high risk of bias and had a small sample size (Table 4). There was no difference in the reduction in symptoms of croup two hours following treatment with dexamethasone or prednisolone, and likely no difference at six hours. The certainty of the evidence for the effect of dexamethasone compared to prednisolone for reducing the symptoms of croup was moderate, as the only study that contributed to the analysis had a small sample size (Table 5).

There was probably no difference between dexamethasone and budesonide in the rate of return visits or (re)admissions or both. The certainty of the evidence for the effect of dexamethasone compared to budesonide for reducing the rate of return visits or (re)admissions or both was moderate, as few events were reported, and the effect estimate included the null effect as well as considerable benefit for dexamethasone compared to budesonide (Table 2). There was probably no difference between dexamethasone and beclomethasone, and there may have been no difference between dexamethasone and betamethasone in the rate of return visits or (re)admissions or both. The certainty of the evidence for the effect of dexamethasone compared to beclomethasone for reducing the rate of return visits or (re)admissions or both was moderate, as the only study that contributed to the analysis had a small sample size and reported no events (Table 3). The certainty of the evidence for the effect of dexamethasone compared to betamethasone for reducing the rate of return visits or (re)admissions or both was low, as only one small study contributed to the analysis, and the effect estimate included the null effect as well as appreciable benefit and harm (Table 4). Compared to prednisolone, dexamethasone probably reduced the rate of return visits or (re)admissions or both by about 50%. The certainty of the evidence for the effect of dexamethasone compared to prednisolone for reducing the rate of return visits or (re)admissions or both was moderate, as the sample size did not reach the optimal information size (Table 5). The addition of data from Parker 2019 attenuated the magnitude of the difference (61%) previously reported in this comparison category.

There was no difference in length of stay in the hospital or emergency department between children treated with dexamethasone compared to budesonide, or dexamethasone compared to prednisolone. One study showed no difference in clinical improvement between children treated with



dexamethasone and those treated with budesonide. Compared to those treated with budesonide, children treated with dexamethasone were at a reduced risk for needing epinephrine. There was no difference between children treated with dexamethasone and those treated with budesonide in need for intubation or supplemental glucocorticoids. Compared to those treated with betamethasone, children treated with dexamethasone were at a increased risk for needing epinephrine. There was no difference between children treated with dexamethasone and those treated with prednisolone in the need for epinephrine. However, there was a 28% reduction in the use of supplemental glucocorticoids as an additional treatment. Adverse events were infrequently reported.

Huang 2021 was the only study that investigated the effect of inhaled budesonide versus dexamethasone. The study authors found no adverse condition following the treatment of infectious laryngitis with inhaled budesonide or dexamethasone. They did not report on change in clinical croup scores between baseline and 2, 6, 12, and/or 24 hours, or any other outcomes relevant to this review.

One glucocorticoid compared to a combination of glucocorticoids

Three studies investigated one glucocorticoid compared to a combination of glucocorticoids. There was probably no difference in reduction in symptoms of croup for children treated with dexamethasone compared to combined dexamethasone and budesonide, and probably no difference for children treated with budesonide compared to combined budesonide and dexamethasone. The certainty of the evidence for the effect of dexamethasone compared to dexamethasone and budesonide for reducing the symptoms of croup was moderate, as the sample size for the analysis did not meet the optimal information size (Table 8). The certainty of the evidence for the effect of budesonide compared to budesonide and dexamethasone for reducing the symptoms of croup was moderate (Table 8), as only one small study contributed to the analysis.

There may have been no difference in rate of return visits or (re)admissions to the hospital or both following treatment with dexamethasone compared to combined dexamethasone and budesonide, and there was probably no difference for this outcome following treatment with budesonide compared to combined budesonide and dexamethasone. The certainty of the evidence for the effect of dexamethasone compared to dexamethasone and budesonide for reducing the rate of return visits or (re)admissions or both was low (Table 7), as the sample size for the analysis did not meet the optimal information size; there were few events; and the estimate was imprecise. The certainty of the evidence for the effect of budesonide compared to budesonide and dexamethasone for reducing the rate of return visits or (re)admissions or both was moderate (Table 8), as only one small study contributed to the analysis.

There was no difference in hours spent in the hospital or emergency department, clinical improvement, or the need for additional treatments for children treated with dexamethasone compared to those treated with combined dexamethasone and budesonide, nor for children treated with budesonide compared to combined budesonide and dexamethasone. Only one study collected adverse events data, which included one case (1.5%) of oral thrush in

the budesonide group and no events in the budesonide and dexamethasone group (Klassen 1998).

Glucocorticoids given by different modes of administration

Five studies investigated dexamethasone given by different modes of administration. There was probably no difference in the rate of return visits or (re)admissions or both for children treated with oral dexamethasone compared to those treated with intramuscular dexamethasone. There was probably a reduced rate of return visits or (re)admissions or both for children treated with oral dexamethasone compared to those treated with nebulised dexamethasone. The certainty of the evidence for the effect of oral compared to intramuscular dexamethasone for reducing the rate of return visits or (re)admissions or both was moderate, as the contributing studies reported few events, and the estimate was imprecise (Table 9). The certainty of the evidence for the effect of oral compared to nebulised dexamethasone for reducing the rate of return visits or (re) admissions or both was moderate, as only one study contributed to the analysis, and the sample size did not meet the optimal information size (Table 10).

There was no difference in clinical improvement or in the need for additional treatments between children treated with oral dexamethasone and those treated with intramuscular dexamethasone. None of the studies comparing dexamethasone given by different modes of administration reported collecting adverse events data.

Dexamethasone given in different doses

Five studies investigated dexamethasone given in different doses. There was no reduction in croup score after two hours for inpatients treated with 0.60 mg/kg dexamethasone compared to those treated with 0.15 mg/kg dexamethasone. There was probably no reduction in croup score after six hours for children treated with 0.60 mg/kg dexamethasone compared to those treated with 0.15 mg/kg dexamethasone. After 12 hours, we do not know if there was no difference in the change in croup score amongst children treated with 0.60 mg/kg compared to 0.15 mg/kg dexamethasone. The effect differed by inpatient and outpatient status. One study showed that there was probably a reduction in croup score with 0.60 mg/kg after 24 hours (Alshehr 2005).

In inpatients, the 0.60 mg/kg dose resulted in a greater reduction in croup score after 12 hours, whereas in outpatients, the 0.15 mg/kg dose was more effective. One study investigated change in croup score after 24 hours for inpatients treated with 0.60 mg/ kg or 0.15 mg/kg dexamethasone (Alshehr 2005). Children treated with 0.15 mg/kg probably experienced greater reductions in croup score after 24 hours compared to those treated with 0.60 mg/kg dexamethasone. The certainty of the evidence for the effect of 0.60 mg/kg dexamethasone compared to 0.15 mg/kg dexamethasone for reducing croup score was very low to high (Table 13). The sixhour analysis included three studies, but the sample size did not meet the optimal information size. In the 12-hour analysis, there was considerable between-study heterogeneity in effect estimates, and the sample sizes did not meet the optimal information size. In the 12-hour analysis, the pooled effect estimate included the null effect as well as appreciable benefit and harm. The 24-hour analysis included only one study with a small sample size.

There may have been no difference in the rate of return visits or (re)admissions or both for children treated with 0.30 mg/



kg compared to 0.15 mg/kg dexamethasone and 0.60 mg/kg compared to 0.30 mg/kg dexamethasone. There was no difference in the rate of return visits or (re)admissions for children treated with 0.60 mg/kg compared to 0.15 mg/kg dexamethasone. The certainty of the evidence was low for the effect of 0.30 mg/kg compared to 0.15 mg/kg dexamethasone (Table 11), and 0.60 mg/kg compared to 0.30 mg/kg dexamethasone (Table 12), for reducing the rate of return visits or (re)admissions or both for croup, as the analysis included only one small study that reported few events, and the effect estimate included benefit, the null effect, and potential for harm. The certainty of the evidence for the effect of 0.60 mg/kg compared to 0.15 mg/kg dexamethasone on return visits or (re)admissions or both was high (Table 13).

Likewise, we found no difference in length of stay in the hospital or emergency department for children treated with 0.60 mg/kg compared to 0.15 mg/kg dexamethasone. There was no difference in the need for additional treatments between children treated with 0.30 mg/kg compared to 0.15 mg/kg dexamethasone; 0.60 mg/kg compared to 0.30 mg/kg dexamethasone; or 0.60 mg/kg compared to 0.15 mg/kg dexamethasone. Adverse events were infrequently reported for the 0.15 mg/kg and 0.60 mg/kg doses of dexamethasone.

Overall completeness and applicability of evidence

We searched for RCTs that compared glucocorticoids to placebo, or any other active pharmacologic treatment for croup. However, in this update we found only two new studies that investigated one glucocorticoid compared to another glucocorticoid (Huang 2021; Parker 2019), and one study that investigated glucocorticoids given in different doses (Parker 2019). Overall, the number of included studies was large (n = 45), of which 26 (58%) investigated glucocorticoids compared to placebo. Only four studies investigated glucocorticoids compared to epinephrine; 13 investigated one glucocorticoid compared to another glucocorticoid; three investigated one glucocorticoid compared to a combination of glucocorticoids; five investigated glucocorticoids given by different modes of administration; and five investigated glucocorticoids given in different doses. Most studies (67%) reported a change in croup score for at least one time point, and 58% used the Westley croup score (Westley 1978), which has been shown to be a valid and reliable measure of croup severity. Most studies (51%) investigated outpatients presenting to emergency departments or outpatient clinics, generally with mild to moderate croup. In a study conducted by Rosychuk 2010 that described the epidemiology of croup presentation to emergency departments within the Alberta, Canada emergency databases, less than 6% of children presenting to the emergency department with croup symptoms required hospitalisation. We have therefore presented subgroup analyses by inpatient or outpatient setting as a form of sensitivity analysis because of the possible overrepresentation of studies with inpatient cases of croup. However, the findings from these subgroup analyses should be interpreted with caution.

We found no evidence of publication bias for our two primary outcomes: change in croup score (at six hours), and return visits or (re)admissions to the hospital or both for glucocorticoids compared to placebo.

Certainty of the evidence

This systematic review included 45 RCTs of 5888 children. Most studies were at unclear or high overall risk of bias (98%). We assessed risk of bias for random sequence generation as low in 60% of studies. The allocation sequence was adequately concealed between randomisation and assignment to treatment groups in 42% of studies. We were unable to ascertain whether the conduct of these studies was methodologically flawed. However, based on the information provided in the publications, we cannot exclude the possibility of selection bias. Empirically, selection bias has been associated with exaggerated estimates of treatment effects (Jüni 2001; Wood 2008). Inadequate allocation concealment is more likely to result in biased estimates of treatment effects when the outcomes of a study are subjective (Wood 2008). Croup score, one of our primary outcomes, is typically assessed by the healthcare provider, and interobserver variability has been reported to be fair to moderate (Chan 2001). Hartling 2014 demonstrated that the association between selection bias and the estimate of treatment effects may not hold true for RCTs in child health. We are therefore uncertain as to how selection bias may have impacted our results.

More than half (58%) of the included studies were at low risk of bias for blinding of participants and personnel, and 60% were at low risk for blinding of outcome assessors. Many of the studies judged as at unclear risk of bias for the blinding domains were described as "blind" or "double-blind". However, details about who was blinded or how (or both) were omitted from the publications. Whilst it is possible that these studies were well conducted but inadequately reported, we cannot confidently exclude the potential for performance and detection bias. In eight (18%) studies, participants and personnel were not blinded. All of these studies but one investigated glucocorticoids given via different modes of administration (e.g. orally, intramuscularly, nebulised), therefore blinding participants and personnel to the treatment assignment would not have been feasible. Studies that are not blinded or that are inadequately blinded can result in exaggerated estimates of treatment effects (Wood 2008). This association may not be true for RCTs in child health (Hartling 2014), therefore we are uncertain as to how the inclusion of unblinded or inadequately blinded trials may have impacted our results.

Most studies (60%) were at low risk of bias for incomplete outcome data. Although most studies (93%) were at unclear risk of bias for selective reporting, the outcomes reported in the results matched those reported in the methods sections of the publications in most cases

For the comparison any glucocorticoid versus placebo, we detected between-study heterogeneity in point estimates of effect as well as heterogeneity in the pooled estimates of effect by glucocorticoid for change in croup score. For this reason, we downgraded the certainty of the evidence for change in croup score after 2, 6, 12, and 24 hours. With respect to the estimates for individual glucocorticoids, after two hours the between-study estimates for budesonide were heterogeneous. Two studies showed a clear benefit for dexamethasone, whilst Johnson 1996 showed the potential for no difference in effect between dexamethasone and placebo. Between-study estimates for the effectiveness of budesonide compared to placebo after 6, 12, and 24 hours showed a consistent beneficial effect. For dexamethasone, between-study estimates were highly heterogeneous at all time points and included the potential for benefit, no effect, or harm compared



to placebo. In future updates of this review, we may use metaregression analyses to explore factors that could explain at least some of the observed heterogeneity (e.g. the 'effective' dosage of the active comparator). If such an analysis is deemed important to clinicians and researchers, it should be planned and documented a priori before future updates of this review. Only one very small study (N = 17) investigated croup score for fluticasone compared to placebo 2, 6, and 24 hours after treatment (Roorda 1998). Another single study (N = 42) investigated croup score for prednisolone compared to placebo 6 and 12 hours after treatment (Massicotte 1973). We caution against drawing any conclusions based on the evidence from these small, single studies.

Accounting for the pooled estimates of effect by glucocorticoid, the test for subgroup differences between the effects of budesonide, dexamethasone, and fluticasone two hours following their administration indicated marginal differences in croup scores (P = 0.06). Whilst fluticasone (based on one study) compared to placebo showed no reduction in croup scores (P = 0.36), the pooled effect estimate for budesonide indicated a reduction in croup scores (P = 0.005), and a marginal reduction for dexamethasone (P = 0.06). There was a subgroup difference in effect between budesonide, dexamethasone, fluticasone, and prednisolone six hours following their administration (P = 0.009). This was accounted for by the fact that the effect estimate for prednisolone (based on one study) was substantially larger compared to the pooled estimates for budesonide and dexamethasone, and fluticasone (based on one study) had no effect (P = 0.90). There was a subgroup difference in effect between budesonide, dexamethasone, and prednisolone 12 hours following their administration (P = 0.006). This was accounted for by the fact that the effect estimate for prednisolone (based on one study) was substantially larger compared to the pooled estimates for budesonide and dexamethasone. There was a subgroup difference in effect between budesonide, dexamethasone, and fluticasone 24 hours following their administration (P = 0.01). This was accounted for by the fact that the effect estimate for fluticasone (based on one study) indicated no effect (P = 0.36), whilst the pooled estimates for budesonide and dexamethasone both showed beneficial effects.

For the comparison any glucocorticoid versus placebo, we also downgraded the certainty of the evidence for return visits or (re)admissions or both due to inconsistency. There was little evidence that publication bias influenced our results for return visits or (re)admissions or both.

Similar threats to the certainty of the evidence were present in the other 12 comparisons in this review, including concerns regarding risk of bias, inconsistency, and imprecision. Aside from the comparison of glucocorticoids versus placebo, for which seven to 11 RCTs made up the analyses for the primary outcomes, all of the other comparisons included between one and five studies. Combined with the fact that the studies mostly included small samples of children (median n = 72, interquartile range 54 to 99), many analyses had to be downgraded due to imprecision, as the optimal information size criteria were not met. Since many of the analyses contained only one or two small RCTs, we caution against drawing any firm conclusions from the results of these few small studies. There exist very few within-study comparisons of one glucocorticoid compared to another, of glucocorticoids given by different modes of administration, or of different doses of the same glucocorticoid.

Potential biases in the review process

We know that the risk of bias of studies used in a meta-analysis is crucial to the certainty of evidence produced. This may affect the translation and uptake of research evidence to practice. We used the Cochrane risk of bias tool to assess the risk of bias of all included studies, and judged most of the studies as at unclear or high risk of bias. We were unable to update some of the other comparisons due to a lack of new data. These comparison categories are highlighted in the Results section. The lack of new data in these areas may signal areas where more RCTs on glucocorticoids and croup are needed. We also know that some comparisons may not warrant new RCTs at this time because of the considerable high-certainty evidence that is available. To the best of our knowledge, 26 RCTs have compared any glucocorticoids versus placebo, and the findings have been consistently in support of the use of glucocorticoids in the treatment of croup (Gates 2018). That said, other areas that may benefit from more RCTs are comparisons around the lowest optimal glucocorticoids dose, and the most effective mode of administration.

Agreements and disagreements with other studies or reviews

This update strengthens evidence from previous review publications that supported the use of a lower dose of glucocorticoids (Chub-Uppakarn 2007; Geelhoed 1995). It would be preferable to treat croup with the lowest effective dose of glucocorticoids to limit steroid exposure in these children. Data from Parker 2019 and Chub-Uppakarn 2007 suggest that 0.15 mg/kg is as effective in the treatment of croup as 0.60 mg/kg. Their findings are also supported by Geelhoed 1995, which found no difference in return visits or (re)admissions amongst children treated for croup using 0.30 mg/kg compared to 0.15 mg/kg. Likewise, there was no difference between the two treatment groups in the need for additional treatment with epinephrine or supplemental glucocorticoid.

AUTHORS' CONCLUSIONS

Implications for practice

The evidence has not changed that glucocorticoids reduce symptoms of croup at two hours, which may last up until 24 hours; shorten hospital stays; and reduce the rate of return visits or (re)admission.

Apart from dexamethasone and prednisolone, we found insufficient data to draw conclusions about the role of other glucocorticoids (e.g. fluticasone, beclomethasone) for reducing the symptoms of croup. Adverse events were reported from the use of glucocorticoids in some of the included studies.

Implications for research

This update further strengthens the evidence base for the effectiveness of glucocorticoids in the treatment of croup. Dexamethasone reduces return visits or (re)admission of croup by about half. A small dose of dexamethasone at 0.15 mg/kg may be as effective as the current standard dose of 0.60 mg/kg. More randomised controlled trials are needed to strengthen the evidence for the effectiveness of low-dose dexamethasone at 0.15 mg/kg to treat croup.



The findings of this update review are in keeping with previous versions of the review which asserted that additional trials assessing the effectiveness of dexamethasone and budesonide compared to placebo are not warranted. The cumulative metagraph by year for change in croup score six hours after treatment shows that the standardised mean difference for the effect of

glucocorticoids compared to placebo has been stable (Figure 4). Accordingly, we also found no new studies published since 1999 that reported on this outcome for this comparison. For return visits or (re)admissions or both, the cumulative meta-graph by year indicates that the pooled risk ratio has also been relatively constant (Figure 5). No new trials reporting on this outcome for this comparison have been published since 2004.

Figure 4. Cumulative meta-graph by year for change in croup score six hours after treatment for any glucocorticoid compared to placebo.

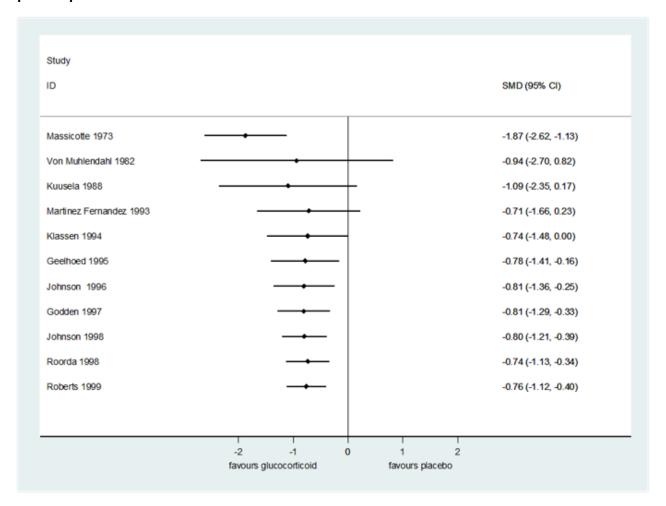
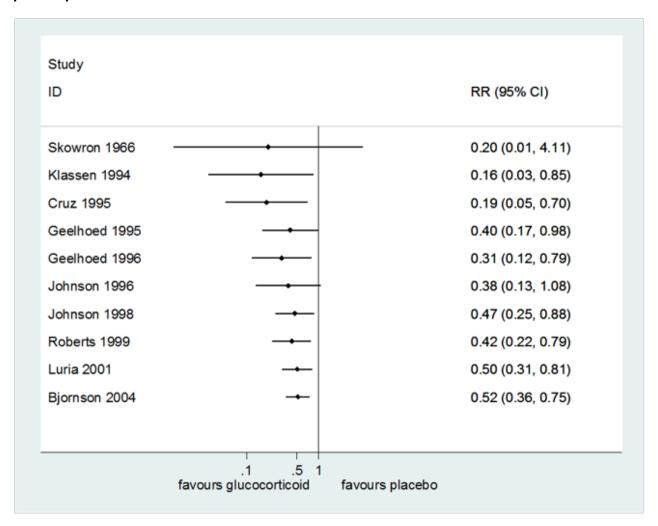




Figure 5. Cumulative meta-graph by year for return visits or (re)admissions or both for any glucocorticoid compared to placebo.



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The following people conducted the editorial process for this update.

- Sign-off Editors (final editorial decision): Mark Jones (Bond University, Australia); Mieke van Driel (The University of Queensland, Australia).
- Managing Editors (provided editorial guidance to authors, edited the review, selected peer reviewers, and collated peerreviewer comments): Liz Dooley (Bond University, Australia); Fiona Russell (Bond University, Australia).

- Contact Editor (assessed peer-review comments and recommended an editorial decision): Lubna Al-Ansary (Department of Family and Community Medicine, College of Medicine, King Saud University, Riyadh, Saudi Arabia).
- Statistical Editor (provided comments): Robert S Ware (Menzies Health Institute Queensland, Griffith University, Australia).
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Peer reviewers (provided comments and recommended an editorial decision):

- Clinical/content review: Dr Gina Neto (University of Ottawa, Canada).
- Consumer review: A Fraiz (Registered Nurse and Consumer).
- Methods review: Rachel Richardson (Associate Editor, Cochrane).
- Search review: Justin Clark (Institute for Evidence-Based Healthcare, Bond University, Australia); Liz Doney (Cochrane Skin, University of Nottingham, UK).



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* Indicates the major publication for the study

CHARACTERISTICS OF STUDIES

Characteristics of included studies [ordered by study ID]

Alshehr 2005

Study characteristics			
Methods	Randomised, double-blind trial		
Participants	Study period: September 1998 to December 2002		
	Setting: emergency rooms and outpatient clinics in 3 medical institutes, Abha City, Saudi Arabia		
		dren aged 3 months to 9 years who had been given a diagnosis of croup and had severe respiratory distress (Westley croup score > 3)	
	Exclusion criteria: symptoms or signs suggesting another cause of stridor; history of chronic pulmonary disease; severe systemic disease; immune dysfunction; stridor or intubation for more than month; glucocorticoid therapy in the last 4 weeks before study entry		
	Baseline characterist	ics (N = 72):	
	• mean (SD) age, mo	reatment: 56%; comparator: 53% onths: treatment: 16.8 (12); comparator: 17.6 (13) estley croup score: treatment: 5.0 (3 to 6); comparator: 4.5 (3 to 6)	
Interventions	Treatment (N = 36): single dose 0.15 mg/kg oral dexamethasone Comparator (N = 36): single dose 0.6 mg/kg oral dexamethasone		
Outcomes	Change in Westley croup score from baseline to 4, 12, and 24 hours; hospitalisation; length of stay in hospital; use of mist tent		
Notes	All children received mist therapy throughout the observation period.		
	Funding source: not reported.		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "A blocked randomization code was produced by random-number generating software"	
Allocation concealment (selection bias)	Low risk	Quote: "To make the study drugs indistinguishable from each other, they were packaged in opaque containers and diluted on the same amount of solution." "A blocked randomization code was produced and the code was not broken until after the study ended and all decisions regarding data analysis were finalized"	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "double-blind" "To make the study drugs indistinguishable from each other, they were packaged in opaque containers and diluted on the same amount of solution."	



Alshehr 2005 (Continued)		
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "double-blind"; "to make the study drugs indistinguishable from each other, they were packaged in opaque containers and diluted on the same amount of solution"; "the code was not broken until after the study ended and all the decisions regarding data analysis were finalized"
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 14% (N = 12) of children recruited were excluded prior to randomisation. All randomised children were followed up.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Amir 2006

Study characteristics	
Methods	Randomised controlled trial
Participants	Study period: November 2002 to March 2003
	Setting: emergency department of Schneider Children's Medical Center, Israel
	Inclusion criteria: children aged 6 months to 6 years with a clinical picture of mild to moderate acute laryngotracheitis presenting to the emergency department. Mild to moderate croup defined as a West ley croup score of 1 to 11
	Exclusion criteria: spasmodic croup; acute epiglottitis; bacterial tracheitis; pneumonia; foreign body aspiration; chronic lung disease; congenital or acquired anatomical airway anomalies; immunosuppressed or immunocompromised; treated before arrival at the emergency department with inhaled bronchodilators or corticosteroids in any form; exposed to varicella in the previous 28 days; contradic tions to corticosteroid treatment
	Baseline demographics (N = 52):
	• proportion male: treatment: 73%; comparator: 27%
	 mean (SD) age in months: treatment: 31.1 (20.4); comparator: 26.7 (16.8)
	• mean (SD) Westley croup score: treatment: 3.6 (2.6); comparator: 2.0 (2.5)
Interventions	Treatment (N = 26): single 0.60 mg/kg (maximum 10 mg) dose of intramuscular dexamethasone
	Comparator (N = 26): single 0.40 mg/kg dose of oral betamethasone
	Treatments were of equivalent potency; supplemental oxygen was provided to children in whom air saturation was $<$ 95%.
Outcomes	Change in Westley croup score from baseline to 2 and 4 hours; return visits to the emergency department; use of epinephrine
Notes	Funding source: no external funding
Risk of bias	



Amir 2006 (Continued)

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "Study participants were assigned by a random numbers table"
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Comment: no blinding described. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	High risk	Comment: used a third-party outcome assessor. No blinding. Subjective outcomes
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: the number of children analysed is not reported
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	High risk	Comment: baseline imbalance in croup score
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk

Bjornson 2004

Study characteristics

Methods	Randomised, double-blind controlled trial
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Participants

Study period: September 2001 to April 2002 and September 2002 to February 2003

Setting: 4 paediatric emergency departments in Canada (Alberta Children's Hospital in Calgary, Stollery Children's Health Centre in Edmonton, Winnipeg Children's Hospital in Winnipeg, or Children's Hospital of Eastern Ontario in Ottawa)

Inclusion criteria: children aged 3 months to 9 years with mild croup based on an initial medical evaluation. Mild croup was defined as onset within the past 72 hours of a seal-like, barking cough and a Westley croup score of ≤ 2.

Exclusion criteria: symptoms or signs of another cause of stridor; history of congenital or acquired stridor, asthma, exposure to varicella within the previous 21 days, chronic pulmonary disease, severe systemic disease, or known immune dysfunction; treatment with corticosteroids within the past 2 weeks; treatment of respiratory distress with epinephrine prior to enrolment; those enrolled in another clinical trial in the past 4 weeks; parents unable to speak either English or French; lack of a telephone at home; a prior visit to the emergency department because of croup during this episode of the disease

Baseline demographics (N = 720):

- proportion male: treatment: 61%; control: 61%
- mean (SD) age in months: treatment: 35 (23); control: 35 (23)



Bjornson 2004 (Continued)	• Westley croup score: treatment: 38% score of 0, 38% score of 1, 24% score of 2; control: 38% score of 0, 43% score of 1, 19% score of 2		
Interventions	Treatment (N = 359): single 0.60 mg/kg (maximum of 20 mg) dose of oral dexamethasone		
	Control (N = 361): single dose of oral placebo (10 mL distilled water)		
	Both treatment and placebo included 50 mL of wild cherry-flavoured syrup.		
Outcomes	Return visits to any healthcare provider within 7 days		
Notes	Funding source: Canadian Institutes of Health Research; Alberta Children's Hospital Foundation; Children's Hospital of Eastern Ontario Research Institute; Stollery Children's Hospital Foundation; Cumberland Pharmaceuticals		

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "computer-generated randomization scheme, stratified by centre, used random permuted blocks of 6-20 children to ensure the comparable assignment of eligible patients."
Allocation concealment (selection bias)	Low risk	Quote: "Codes were secured at each center's pharmacy until enrolment and all decisions regarding data analysis had been finalized." The preparations were "not distinguishable" and "packaged in sequentially numbered, sealed bags".
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "double-blind"; "Parents were unable to determine which preparation their child had received"; "Preparations were not distinguishable by appearance, volume, weight, taste, or smell."
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "double blind"; "Preparations were not distinguishable by appearance, volume, weight, taste, or smell and were packaged in identical syringes in sequentially numbered, sealed bags"; "biostatistician who was not otherwise involved in the study performed the data analysis."
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: missing data imputed using intention-to-treat analysis. 5% (N = 37) protocol deviations equally distributed between groups. 2% (N = 13) had incomplete follow-up.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Cetinkaya 2004

Study characteristics	
Methods	Randomised controlled trial
Participants	Study period: not reported



Cetinkaya 2004 (Continued)

Setting: paediatric emergency department at the Şişlu Etfal Education and Research Hospital in Istanbul, Turkey

Inclusion criteria: children aged 6 to 36 months who were admitted to the paediatric emergency clinic with a diagnosis of croup, defined as the acute (< 48 h) onset of stridor, chest wall retraction, barking cough, and hoarse voice

Exclusion criteria: epiglottitis; reactive airway exacerbation; foreign body aspiration; acute bacterial pneumonia, acquired or congenital upper airway anomalies; intubated in the previous month; received steroids within the preceding 2 weeks

Baseline characteristics (N = 60):

• proportion males: 67%

• age: not reported

Westley croup score: not reported

Interventions

All children received 5 to 6 L/min of moisturised oxygen for 20 minutes upon arriving at the hospital, along with a single dose of 0.16 mg/kg of salbutamol.

Treatment 1 (N = 15): 500 µg nebulised budesonide, a single dose of oral placebo (multivitamin syrup), and 2 mL intramuscular placebo (saline)

Treatment 2 (N = 15): 2 mL nebulised placebo (saline), a single dose oral placebo (multivitamin syrup), and 0.60 mg/kg (maximum 8 mg) intramuscular dexamethasone

Treatment 3 (N = 15): 2 mL nebulised placebo (saline), a single dose of 0.60 mg/kg (maximum 8 mg) oral dexamethasone, and 2 mL intramuscular placebo (saline)

Control (N = 15): 2 mL intramuscular placebo (saline), 2 mL nebulised salbutamol solution with saline, and oral placebo (multivitamin syrup)

Outcomes

Change in Westley croup score from baseline to 24 hours

Notes

Funding source: not reported

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Comment: blinding not explicitly described; it appears there was an attempt to blind parents and children by using similar-appearing oral, nebulised, and intramuscular placebos. It is unclear if personnel were blinded. Subjective outcomes
Blinding of outcome as- sessment (detection bias) All outcomes	High risk	Comment: no description of blinding or third-party outcome assessors. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data



Cetinkaya 2004 (Continued)		
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk

Chub-Uppakarn 2007

Study characteristics	
Methods	Randomised, double-blind controlled trial
Participants	Study period: March 2001 to October 2003
	Setting: paediatric ward of Hatyai Hospital in the southern part of Thailand
	Inclusion criteria: children aged 6 months to 5 years who were admitted to the paediatric ward with moderate to severe croup. Westley croup score 4 to 7
	Exclusion criteria: history of contact with chicken pox within the preceding 3 weeks; history of congenital or acquired stridor; chronic pulmonary disease; asthma; severe systemic disease or known immune dysfunction; treatment with corticosteroids within the preceding 2 weeks; treatment with epinephrine for respiratory distress before enrolment
	Baseline demographics (N = 41):
	 proportion male: treatment 1: 55%; treatment 2: 86% mean (SE) age in months: treatment 1: 16.9 (2.0); treatment 2: 18.8 (2.6) mean (SD) Westley croup score: treatment 1: 4.26 (0.22); treatment 2: 4.60 (0.25)
Interventions	Treatment 1 (N = 21): single 0.15 mg/kg dose (maximum 3 mg) of intramuscular dexamethasone
	Treatment 2 (N = 20): single 0.60 mg/kg dose (maximum 12 mg) of intramuscular dexamethasone
Outcomes	Change in Westley croup score at 2, 6, and 12 hours; intubations; adverse events
Notes	All children were treated with a single nebulisation of epinephrine (1:1000) 1 mL in 0.9% saline 3 mL at baseline.
	Funding source: not reported

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "computer-generated randomization scheme used random permuted blocks of four children"
Allocation concealment (selection bias)	Low risk	Quote: "codes were secured at the hospital pharmacy until enrolment and all decisions regarding data analysis had been finalized"
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "the preparations of dexamethasone suspension consisted of 10 mL of dexamethasone phosphate injection in concentrations of 1.2 and 0.3 mg/mL. The preparations were packaged in identical containers by a hospital pharmacist and were not distinguishable in appearance"



Chub-Uppakarn 2007 (Continued)			
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel	
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data	
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.	
Other bias	Low risk	Comment: no other sources of bias identified	
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk	

Cruz 1995

Study characteristics		
Methods	Randomised, double-blind controlled trial	
Participants	Study period: November 1992 to December 1993	
	Setting: emergency de	epartment at the Children's Hospital of Michigan, USA
	Inclusion criteria: children aged 6 months to 5 years reporting to the emergency department we clinical diagnosis of acute laryngotracheitis or viral croup. Westley croup score of at least 2 and be managed as outpatients	
	Exclusion criteria: history of prior intubation; structural airway anomalies; those requiring more than 1 racaemic epinephrine treatment; hospitalisation; β -agonist therapy; received steroids in the past 24 hours	
	Baseline demographics (N = 45):	
	 median (SD) age in 	reatment: 74%; control: 63% months: treatment: 18.0 (19.0); control: 21.0 (8.0) estley croup score: treatment: 3 (2 to 5); control: 3 (2 to 5)
Interventions	Treatment (N = 19): single 0.60 mg/kg (maximum 10 mg) dose of intramuscular dexamethasone	
	Control (N = 19): equal volume of placebo (normal saline)	
Outcomes	Return visits to the emergency department; patient improvement 24 hours after discharge	
Notes	Funding source: not reported	
Risk of bias		
Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement



Cruz 1995 (Continued)		
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "double blind"; "Both the patients and the investigators were blinded to the content of the syringe"; "the drug code was broken only after the last patient had completed the study"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 16% (N = 7) were excluded or lost to follow-up; it is unclear if losses were balanced between groups
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Dobrovoljac 2012

Study characteristics	
Methods	Randomised, double-blind controlled trial
Participants	Study period: not reported
	Setting: emergency department of the Princess Margaret Hospital for Children, Perth, Australia
	Inclusion criteria: children over 6 months of age presenting to the emergency department with mild to moderate croup (harsh cough with or without stridor, Westley croup score 1 to 6)
	Exclusion criteria: Westley croup score < 1; severe croup (Westley croup score > 6) requiring epinephrine upon arrival; received steroids within the last week; other significant co-existing illnesses
	Baseline demographics (N = 70):
	 proportion male: treatment: 69%; control: 66% mean (SD) age in months: treatment: 37.1 (22.6); control: 27.4 (25.7) mean (SD) Westley croup score: treatment: 2.7 (0.8); control: 2.8 (1.0)
Interventions	Treatment (N = 35): single 0.15 mg/kg (1 mg/mL solution) dose of oral dexamethasone
	Control (N = 35): same volume of placebo solution
	All children also received 0.15 mg/kg oral dexamethasone at 60 minutes (after study completion).
Outcomes	Change in Westley croup score from baseline to 2 hours; use of epinephrine and supplemental gluco-corticoids
Notes	Funding source: not reported



Dobrovoljac 2012 (Continued)

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Low risk	Quote: "Randomization was performed by the pharmacy department and the randomization list was kept concealed from emergency physicians, nurses and parents until the end of the trial."
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "the PMH pharmacy ensured that the two preparations could not be differentiated"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	High risk	Comment: 6% loss to follow-up due to worsening condition, all in the place-bo group (N = 4, 11%). LOCF method used, which could have biased results in favour of the treatment group.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk

Donaldson 2003

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Study characteristics	5
Methods	Randomised, double-blind trial
Participants	Study period: January 1999 to December 1999
	Setting: emergency department of William Beaumont Hospital, USA
	Inclusion criteria: children aged 3 to 84 months with history of inspiratory stridor or a barky cough and a Westley croup score of ≥ 2 after 10 to 15 minutes of cool mist therapy in the emergency department
	Exclusion criteria: Westley croup score < 2; signs suggesting another cause for stridor such as epiglottitis, bacterial tracheitis, foreign body, chronic lung disease; severe comorbidities; inability of parents to give informed consent; glucocorticoid therapy within 4 weeks of presenting
	Baseline demographics (N = 96):
	• proportion male: treatment 1: 73%; treatment 2: 57%
	 mean (SD) age in months: treatment 1: 23.2 (17.9); treatment 2: 28.9 (17.7)
	• mean (SD) Westley croup score: treatment 1: 3.5 (1.8); treatment 2: 3.5 (1.7)
Interventions	Treatment 1 (N = 49): 0.60 mg/kg intramuscular dexamethasone and oral placebo (syrup)



Outcomes	Unscheduled revisits; parent-reported symptom relief after 24 hours; use of epinephrine
Donaldson 2003 (Continued)	Treatment 2 (N = 46): 0.60 mg/kg oral dexamethasone and intramuscular placebo (direct pressure with hub of syringe on thigh)

Notes Funding source: not reported

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "block randomization method from a random number generator performed by the department of Pharmacy"
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Quote: "In both groups, neither the parents nor the treating physicians were present in the treatment room during the administration of medications"; "The emergency medicine faculty were blinded to the route of administration of the drug"; "If the child vomited while in the ED, the treatment given was unblinded"
		Comment: blinding was attempted but could be broken if the child vomited whilst in the emergency department. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: third-party outcome assessor described as blinded. Because blinding of children and parents could have been broken, the assessors could have become unblinded during conversation with parents. Subjective outcomes
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: intention-to-treat analysis used. 1% (N = 1) loss to follow-up, unclear from which group
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Duman 2005

Study	chara	ctoric	ticc

Methods	Randomised controlled trial
Participants	Study period: September 2002 to September 2003
	Setting: paediatric emergency department of Dokuz Eylül University Faculty of Medicine, Izmir, Turkey
	Inclusion criteria: children aged more than 6 months presenting to the emergency department with a history of inspiratory stridor, barking cough, hoarseness and signs of inspiratory distress, and a Westley croup score of ≥ 2



Duman 2005 (Continued)

Exclusion criteria: Westley croup score < 2, and with other suggested causes for stridor (epiglottitis, bacterial tracheitis, foreign body aspiration); past history of laryngoscopy, tracheal intubation, chronic lung disease, or severe comorbidities; immediate intubation or transfer to intensive care; corticosteroid therapy within 4 weeks of presentation; history of tuberculosis personally or in the family; chickenpox within the preceding 21 days; known immunodeficiency

Baseline demographics (N = 76):

- **proportion male:** treatment 1: 77%; treatment 2: 90%; comparator: 77%
- mean (SD) age in months: treatment 1: 41.5 (25.5); treatment 2: 43.3 (24.7); comparator: 34.8 (22.4)
- mean (SD) Westley croup score: treatment 1: 5.3 (1.2); treatment 2: 5.5 (1.8); comparator: 5.0 (1.3)

Interventions

Treatment 1 (N = 31): 2.5 mL (0 to 20 kg) or 5.0 mL (20 to 40 kg) nebulised epinephrine with same volume normal saline and 0.60 mg/kg intramuscular dexamethasone

Treatment 2 (N = 19): 2.5 mL (0 to 20 kg) or 5.0 mL (20 to 40 kg) nebulised epinephrine with same volume normal saline and 2 mg nebulised budesonide

Comparator (N = 26): cool mist therapy and 0.60 mg/kg intramuscular dexamethasone

In all groups, the drug was administered for a period of 20 minutes using a nebuliser with oxygen at a rate of 6 to 7 L/min through a face mask.

Outcomes

Change in Westley croup score from baseline to 2 hours; admissions from the emergency department; use of epinephrine

Notes

Funding source: no external funding

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "Patients were allocated to treatment according to a randomisation list produced at the beginning of the study. Patients were randomised in blocks of three."
		Comment: assumed that the blocked randomisation was computer-generated
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Comment: no blinding. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	High risk	Comment: no blinding. No description of a third-party outcome assessor. Subjective outcomes
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias	High risk	Comment: at least 1 domain judged as high risk



Duman 2005 (Continued) All outcomes

Eboriadou 2010

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: January 2000 to December 2001		
	Setting: emergency department at a hospital in Greece		
	Inclusion criteria: children aged 6 months to 5 years presenting to the emergency department with a clinical diagnosis of viral croup by a history of a short period of viral upper respiratory symptoms followed by hoarseness or barking cough and clinical evidence of hoarseness, barking cough, or stridor in the emergency department; modified Downes and Raphaelly croup score ≥ 2; could be managed as outpatients		
	Exclusion criteria: known structural airways anomalies; acute epiglottitis; bacterial tracheitis; monia; foreign body aspiration or past history of laryngoscopy; history of prior intubation; chroway obstruction; received steroids in the past 24 hours; required more than 1 treatment with ne L-epinephrine or hospitalisation		
	Baseline characterist	ics (N = 64):	
	 proportion males: treatment 1: 80%; treatment 2: 58%; treatment 3: 65% mean age in years: treatment 1: 2.6; treatment 2: 3.2; treatment 3: 3.4 mean modified Downes and Raphaelly croup score: treatment 1: 5.13; treatment 2: 5.89; treatment 3: 3.95 		
Interventions	Treatment 1 (N = 25): s oxygen at a rate of 5 L/	ingle 5 mL (1:1000 mg/mL) dose of nebulised L-epinephrine via nebuliser with minute	
Treatment 2 (N = 1		single 0.60 mg/kg (maximum 8 mg) dose of intramuscular dexamethasone	
	Treatment 3 (N = 20): single 200 μg dose of inhaled beclomethasone dipropionate delivered via AeroChamber device Supplemental oxygen was used for children with oxygen saturation values < 95%.		
Outcomes	Change in modified Downes and Raphaelly croup score from baseline to 2 hours; return visits to the emergency department		
Notes	Funding source: not reported		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "patients were assigned by a random numbers table to receive"	
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Blinding of participants and personnel (perfor- mance bias)	High risk	Comment: described as "double-blind", though the interventions were clearly distinguishable, and the mechanism of blinding was not described	



Eboriadou 2010 (Continued) All outcomes		
Blinding of outcome assessment (detection bias) All outcomes	High risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk

Eden 1964

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: not reported		
	Setting: hospital in the	e USA	
	Inclusion criteria: children hospitalised for treatment of acute croup, including all children with acute respiratory infections characterised by hoarseness, inspiratory stridor, and a barking cough		
	Exclusion criteria: not reported		
	Baseline characteristics (N = 50):		
	proportion male: not reported		
	age: not reported		
	croup score: not reported		
Interventions	All children received as routine therapy oxygen, increased humidity, and tetracycline.		
	Treatment (N = 25): 1 n	ng/kg intramuscular methyl prednisolone every 6 hours for 24 hours	
	Control (N = 25): 1 mg/kg placebo preparation every 6 hours for 24 hours		
Outcomes	Patient improvement at 6, 12, 24 hours		
Notes	Funding source: Upjohn Company (supplied drugs for the trial)		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "The patients were divided into two groups according to a table of random sampling"	



Eden 1964 (Continued)		
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor-	Unclear risk	Quote: "The composition of each preparation was unknown to the investigators until the end of the study"
mance bias) All outcomes		Comment: described as double-blind. Investigators blinded, but it is unclear if participants or personnel (or both) were blinded because who administered the treatments is not stated. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: no description of a third-party outcome assessor, unclear who performed the measurements. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 6% (N = 3) lost to follow-up due to inadequate evaluation. All losses were in 1 group, but it is unclear which group.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Eden 1967

Study characteristics	s		
Methods	Randomised, double-blind controlled trial		
Participants	Study period: not reported		
	Setting: hospital in the USA		
	Inclusion criteria: children hospitalised with acute croup, including those presenting with acute respiratory infections characterised by barking cough, hoarseness, sternal retractions, and respiratory stridor		
	Exclusion criteria: not reported		
	Baseline demographics (N = 50):		
	 proportion male: treatment: 60%; control: 80% age in years (%): treatment: 68% 0 to 2 years, 28% 2 to 4 years, 4% > 4 years; control: 56% 0 to 2 years, 28% 2 to 4 years, 16% > 4 years croup score: not reported 		
Interventions	Treatment (N = 25): 0.1 mL/kg per dose schedule (0.1 mg/kg) intramuscular dexamethasone every 6 hours for 48 hours (total daily dose of 0.40 mg/kg)		
	Control (N = 25): volume of 0.1 mL/kg per dose schedule of placebo preparation intramuscularly every 6 hours for 48 hours		
Outcomes	Patient improvement at 6, 12, 24 hours; tracheostomy		



Eden 1967 (Continued)

Notes

Funding source: dexamethasone (Decadron) by Merck Sharp & Dohme

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "The patients were divided into two groups according to a table of random sampling"
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "double blind" "The composition of each preparation was unknown to the investigators until the end of the study"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Fifoot 2007

Study characteristics

Participants	Study period: July 2004 to October 2005	
Methods	Randomised, double-blind trial	

Setting: emergency department of Mater Childrens' Hospital, Brisbane, Australia

Inclusion criteria: children aged 6 months to 6 years presenting to the emergency department with croup (Westley croup score \geq 2) with parents available for telephone follow-up 1 week following enrolment

Exclusion criteria: chronic respiratory disease (excluding asthma); severe croup (Westley croup score > 7); known allergy or relative contraindication to steroids (varicella or exposure to varicella within the past 3 weeks, history of tuberculosis, diabetes, or hypertension, known immunodeficiency); treatment with steroids in the preceding week or with nebulised adrenaline en route or immediately on arrival in the emergency department

Baseline demographics (N = 99):

- proportion male: treatment 1: 79%; treatment 2: 65%; treatment 3: 80%
- mean (SD) age in years: treatment 1: 1.76 (1.52); treatment 2: 1.53 (1.31); treatment 3: 1.74 (1.61)



Fifoot 2007 (Continued)	• mean (SD) Westley croup score: treatment 1: 3.15 (0.89); treatment 2: 2.71 (0.84); treatment 3: 2.81 (0.87)
Interventions	Treatment 1 (N = 34): single 1 mg/kg dose of oral prednisolone
	Treatment 2 (N = 34): single 0.15 mg/kg dose of oral dexamethasone
	Treatment 3 (N = 31): single 0.60 mg/kg dose of oral dexamethasone
	Children who did not tolerate oral therapy after 2 attempts received nebulised budesonide.
Outcomes	Change in Westley croup score from baseline to 2 and 4 hours; re-presentations with croup; use of epinephrine and use of supplemental glucocorticoids
Notes	Funding source: no external funding

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "Subjects were randomized (using a computer-generated sequence)"
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "Recruiting staff and study investigators were blinded to treatment assignments. MCH pharmacists prepared each steroid agent as a solution, such that each child would receive an identical volume of preparation flavoured to standardize taste and palatability"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "Recruiting staff and study investigators were blinded to treatment assignments. MCH pharmacists prepared each steroid agent as a solution, such that each child would receive an identical volume of preparation flavoured to standardize taste and palatability"
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 13% (N = 13) loss to follow-up. Losses balanced between groups. Did not use intention-to-treat analysis for the telephone follow-up outcomes
Selective reporting (reporting bias)	Low risk	Comment: no deviations from protocol detected
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Fitzgerald 1996

Study characteristics	S	
Methods	Randomised, double-blind controlled trial	
Participants	Study period: not reported	
	Setting: emergency departments of 3 paediatric tertiary referral hospitals, Sydney, Australia	



Fitzgerald 1996 (Continued)

Inclusion criteria: children aged 6 months to 6 years admitted to the emergency department based on a decision by the medical staff regarding croup severity. To be included in study children were required to have acute (viral) or spasmodic croup and a minimum Westley croup score of 6.

Exclusion criteria: significant past or present systemic disease; pre-existing known airway abnormalities; confirmed hypersensitivity to budesonide or L-adrenaline; suspected epiglottitis; foreign body aspiration; bronchiolitis or asthma; need for immediate intubation or transfer to intensive care; treated with glucocorticoids in the 4 weeks prior to the study

Baseline demographics (N = 67):

- · proportion males: not reported
- mean (SD) age in months: treatment 1: 20.9 (12.7); treatment 2: 24.9 (12.5)
- mean (SD) modified Westley croup score: treatment 1: 7.1 (1.2); treatment 2: 7.7 (1.1)

Interventions	Treatment 1 (N = 35): single 2 mg/4 mL dose of nebulised budesonide		
	Treatment 2 (N = 31): single 4 mg/4 mL dose of nebulised L-epinephrine		
Outcomes	Change in modified Westley croup score from baseline to 2, 12, and 24 hours; readmission to the hospital; intubation, use of epinephrine, and use of additional steroids		
Notes	Funding source: Astra Pharmaceuticals		

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Comment: described as double-blind, but nursing staff (personnel) were unblinded. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "The study medication was administered by nursing staff and the investigator was not present when the medication was placed in the opaque nebulizer bowl"
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 5% (N = 3) loss to follow-up. All-patients-treated analysis excluded N = 1 child (1.5%). 13 children who received medications appropriately did not remain for the entire 24 hours. Last value extended for those who recovered before the 24-hour period; however, no children returned or were readmitted; it is unclear how this may have affected the findings.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk



Garbutt 2013

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: 26 October 2009 to 16 April 2010, and 6 September 2010 to 29 April 2011		
	Setting: 10 offices of primary care practitioners in St Louis, MO, USA		
	Inclusion criteria: children aged 1 to 8 years with croup symptoms for \leq 48 hours and a clinical diagnosis of mild or moderate croup at an office visit, based on symptoms in the past 24 to 36 hours		
	Exclusion criteria: diagnosis of severe croup or impending respiratory failure; prior treatment inephrine or oral corticosteroids for this croup episode; symptoms or signs suggesting other stridor; chronic respiratory disease including asthma; known contraindication to steroid us not in the same household as the child in the subsequent 4 days, could not participate in the terviews, or was not English speaking		
	Baseline demographi	cs (N = 87):	
	• proportion male: treatment: 61%; comparator: 68%		
	· · · · · · · · · · · · · · · · · · ·	ears: treatment: 2.67 (1.43); comparator: 3.11 (1.58)	
	• mean (SD) Westley croup score: treatment: 0.4 (0.7); comparator: 0.6 (0.8)		
	• mean (SD) Telepho	one Outpatient Score: treatment: 2.2 (0.9); comparator: 2.0 (0.9)	
Interventions	Treatment (N = 41): sin days	gle 2 mg/kg (maximum 60 mg/day) dose of oral prednisolone once per day for 3	
	Comparator (N = 46): single 0.60 mg/kg (maximum 18 mg) dose of oral dexamethasone, followed by days of placebo comparable in appearance, smell, and taste		
Outcomes	Additional health care for croup within 11 days of randomisation		
Notes	Funding source: National Center for Research Resources (National Institutes of Health); National Institutes of Health Roadmap for Medical Research		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "randomized blocks were used to assign subjects to treatment groups, with randomization stratified by site. Computer generated random numbers determined how the two treatments were allocated"	
Allocation concealment (selection bias)	Low risk	Quote: "Study drug packages were prepared offsite by the pharmacist"; "The pharmacist packaged the bottles in a sealed opaque envelope"; "For allocation concealment, the drug formulation ensured the volume of the weight-based dose was equivalent for each medication."	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "The pharmacist packaged the bottles in a sealed opaque envelope"; "comparable in appearance, smell and taste"; "patients, parents, PCPs, and study team members were blinded to treatment assignments"	
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "The pharmacist packaged the bottles in a sealed opaque envelope"; "comparable in appearance, smell and taste"; "patients, parents, PCPs, and study team members were blinded to treatment assignments"	



Garbutt 2013 (Continued)		
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: intention-to-treat analysis used. Losses to follow-up on days 1, 2, 3, 4, and 11 were 7%, 9%, 9%, 3%, and 2% (non-cumulative). Participation in follow-up interviews was balanced between groups.
Selective reporting (reporting bias)	Low risk	Comment: no deviations from protocol detected
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Low risk	Comment: all domains judged as low risk

Geelhoed 1995a

Study characteristics			
Methods	Randomised, double-blind controlled trial (trial A; see Geelhoed 1995b for trial B)		
Participants	Study period: July 1994 to August 1994		
	Setting: Princess Margaret Hospital for Children, Perth, Australia		
		ldren older than 3 months admitted to the hospital with a diagnosis of croup tory stridor, chest wall retractions, barking cough, and hoarse voice) and a minigroup score of 3	
	croup); families withou apy in the past week; p	ner acute or chronic medical problems; modified Westley croup score < 3 (mild at a telephone or with limited English language abilities; any kind of steroid therere-existing upper airway condition; history of prolonged stridor; those presentere suggesting a diagnosis other than croup; admitted directly to the intensive roup	
	Baseline demographics (N = 60):		
	• proportion male: treatment: 62%; comparator: 81%		
	• mean (SD) age in months: treatment: 35 (19); comparator: 42 (27)		
	mean modified we	stley croup score: treatment: 3.8; comparator: 3.7	
Interventions Treatment (N = 29): single 0.30		gle 0.30 mg/kg (maximum 6 mg) dose of oral dexamethasone	
	Comparator (N = 31): si	ingle 0.60 mg/kg (maximum 12 mg) of oral dexamethasone	
Outcomes	Change in modified Westley croup score from baseline to 2 and 4 hours; re-presentations with croup; length of hospital stay; use of epinephrine and use of additional glucocorticoids		
Notes Study reports on 2 comparisons Funding source: not reported		nparisons.	
		ported	
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	



Geelhoed 1995a (Continued)		
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor-	Unclear risk	Quote: "treatments were administered in a double blind fashion" "If the child was withdrawn from the study, their study code was broken."
mance bias) All outcomes		Comment: described as double-blind. Unclear who was blinded. Code could be broken, but it is unclear how frequently this occurred. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: intention-to-treat analysis used. 2% (N = 1) in trial A withdrew, 5% (N = 3) in trial A lost to follow-up.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: some children were not enrolled when the emergency department was busy, potential to bias participant selection
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Geelhoed 1995b

Study characteristics	
Methods	Randomised, double-blind controlled trial (trial B; see Geelhoed 1995a for trial A)
Participants	Study period: August 1994 to December 1994
	Setting: Princess Margaret Hospital for Children, Perth, Australia
	Inclusion criteria: children older than 3 months admitted to the hospital with a diagnosis of croup (acute onset of inspiratory stridor, chest wall retractions, barking cough, and hoarse voice) and a minimum modified Westley croup score of 3
	Exclusion criteria: other acute or chronic medical problems; modified Westley croup score < 3 (mild croup); families without a telephone or with limited English language abilities; any kind of steroid ther apy in the past week; pre-existing upper airway condition; history of prolonged stridor; those presenting with a clinical picture suggesting a diagnosis other than croup; admitted directly to the intensive care unit with severe croup
	Baseline demographics (N = 60):
	 proportion male: treatment: 90%; control: 74% mean (SD) age in months: treatment: 38 (34); comparator: 32 (23) mean modified Westley croup score: treatment: 4.0; comparator: 3.7
Interventions	Treatment (N = 29): single 0.15 mg/kg (maximum 3 mg) dose of oral dexamethasone
	Comparator (N = 31): single 0.30 mg/kg (maximum 6 mg) dose of oral dexamethasone



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Outcomes Change in modified Westley croup score from baseline to 2 and 4 hours; re-presentations with croup; length of hospital stay; use of epinephrine and use of additional glucocorticoids

Notes Study reports on 2 comparisons.

Funding source: not reported

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias)	Unclear risk	Quote: "treatments were administered in a double blind fashion" "If the child was withdrawn from the study, their study code was broken."
All outcomes		Comment: described as double-blind. Unclear who was blinded. Code could be broken, but it is unclear how frequently this occurred. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: intention-to-treat analysis used. 8% (N = 5) in trial B were lost to follow-up.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: some children were not enrolled when the emergency department was busy, potential to bias participant selection
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Geelhoed 1995c

Study	characteristics
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Methods	Randomised, double-blind controlled trial

Participants **Study period:** not reported

Setting: Princess Margaret Hospital for Children, Perth, Australia

Inclusion criteria: children aged 3 months and older admitted to hospital with a diagnosis of croup (acute onset of inspiratory stridor, chest wall retractions, barking cough, hoarse voice) and a minimum croup score of 3

Exclusion criteria: croup score < 3 (mild croup); caregivers did not consent; family did not have a telephone; caregivers had limited English language abilities; received steroid therapy in the past week; pre-



Geelhoed 1995c (Continued)

existing conditions of the upper airway or prolonged stridor; clinical examination suggested a diagnosis other than croup; admitted directly to the intensive care unit with severe croup

Baseline demographics (N = 80):

- proportion male: treatment 1: 52%; treatment 2: 85%; control: 80%
- mean (SD) age in months: treatment 1: 35 (35); treatment 2: 33 (30); control: 30 (23)
- mean croup score: treatment 1: 3.8; treatment 2: 3.7; control: 3.8

Interventions Treatment 1 (N = 23): single dose (0.60 mg/kg) oral dexamethasone and 4 mL nebulised saline

Treatment 2 (N = 27): single 2 mg (4 mL) dose nebulised budesonide and placebo

Control (N = 30): single dose oral placebo and 4 mL of nebulised saline

Outcomes Change in croup score from baseline to 2, 4, and 12 hours; re-presentations with croup; length of hospital stay; use of epinephrine, use of supplemental glucocorticoids, and intubations

Funding source: not reported

Risk of bias

Notes

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Quote: "treatments were administered in a double blind fashion" "If the attending doctors considered patients to be severely ill or failing to improve, they could be withdrawn at any time to receive steroids, at which time their study code was broken."
		Comment: described as double-blind. Unclear who was blinded. Code could be broken. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	High risk	Comment: did not use intention-to-treat analysis for the telephone follow-up. 17% (N = 9) withdrew, more in placebo than in treatment group (23% compared to 9%). An additional 10% (N = 8) were lost to follow-up, unclear from which group.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk



Geelhoed 1996a

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: not reported		
	Setting: emergency department of Princess Margaret Hospital for Children, Perth, Australia		
	Inclusion criteria: children older than 3 months presenting to the emergency department with a diagnosis of croup (acute onset of inspiratory stridor, barking cough, hoarseness, and chest wall retractions) not severe enough to warrant admission		
	Exclusion criteria: other acute or chronic medical problems; families that did not have a telephone or had limited English language abilities; received any type of steroids in the preceding week; pre-existing upper airway condition; history of prolonged stridor; clinical picture that suggested a diagnosis other than croup		
	Baseline demographi	cs (N = 100):	
		reatment: 68%; control: 72%	
	 mean (SD) age in months: treatment: 37 (23); control: 45 (26) mean croup score: treatment: 0.9; control: 0.9 		
	Treatment (N = 50): single 0.15 mg/kg dose of oral dexamethasone		
	Control (N = 50): single dose of oral placebo		
Outcomes	Reattendance at the emergency department with croup		
Notes	Funding source: no external funding		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Blinding of participants	Unclear risk	Quote: "Treatments were given double blind"	
and personnel (perfor- mance bias) All outcomes		Comment: no further explanation given. Unclear who was blinded. Subjective outcomes	
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel	
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 4% (N = 4) lost to follow-up. Equal between groups	
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.	
Other bias	Low risk	Comment: no other sources of bias identified	



Geelhoed 1996a (Continued)

Overall risk of bias All outcomes Unclear risk

Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Geelhoed 2005

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: not reported		
	Setting: emergency department of Princess Margaret Hospital for Children, Perth, Australia		
	Inclusion criteria: children older than 3 months who presented to the emergency department with a diagnosis of croup (acute onset of inspiratory stridor, chest wall retractions, barking cough, hoarse voice) and no other acute or chronic medical problems requiring admission. Croup was defined as the acute onset of inspiratory stridor, barking cough, hoarse voice, and chest wall retractions.		
	Exclusion criteria: other acute or chronic medical problems; families that did not have a telephone or had limited English language abilities; received any type of steroids in the preceding week; pre-existing upper airway condition; history of prolonged stridor; clinical picture that suggested a diagnosis other than croup		
	Baseline demographi	ics (N = 72):	
	 proportion male: treatment: 72%; control: 67% mean (SD) age in months: treatment: 35 (22); control: 36 (30) mean (SD) croup score: treatment: 4.1 (0.8); control: 4.1 (0.8) 		
Interventions	Treatment (N = 36): sin	Treatment (N = 36): single 0.15 mg/kg dose of oral dexamethasone and 2 mg of nebulised budesonide	
	Control (N = 36): single 0.15 mg/kg dose of oral dexamethasone and a dose of equivalent v nebulised placebo (saline)		
Outcomes	Change in croup score from baseline to 2 and 4 hours; readmissions for croup; length of stay in hospital		
Notes	Funding source: not reported		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "Subjects were randomized to 1 of 2 groups based on hospital pharmacy computer-generated numbers."	
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "a nurse who had no further part in the management of the child administered treatment. Other staff and subjects were blinded to the nebulized treatments given."	
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel	



Geelhoed 2005 (Continued)			
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 1% (N = 1) lost to follow-up	
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.	
Other bias	Low risk	Comment: no other sources of bias identified	
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk	

Godden 1997

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: November 1993 to April 1995		
	Setting: paediatric wards of Poole NHS Trust Hospital, Dorset, England		
	Inclusion criteria: children admitted to hospital with a clinical diagnosis of croup based on the modified Westley croup score		
	Exclusion criteria: receiving bronchodilators or received systemic steroids within the previous month		
	Baseline demographics (N = 89):		
	 proportion male: treatment: 72%; control: 64% mean (range) age in months: treatment: 35.7 (7 to 116); control: 37.4 (7 to 93) mean (SD) modified Westley croup score (N = 87): treatment: 5.30 (3.44); control: 5.15 (3.70) 		
Interventions	Treatment (N = 47): initial 2 mg (4 mL) dose of nebulised budesonide, followed by a repeating dose of 1 mg every 12 hours		
	Control (N = 42): initial 4 mL dose of nebulised placebo (normal saline), followed by a repeating dose of 2 mL placebo (normal saline) every 12 hours		
	Both treatment and placebo were delivered via an opaque nebuliser chamber, driven by wall oxygen at a rate of 8 L/min.		
Outcomes	Change in modified Westley croup score from baseline to 2, 4, 12, and 24 hours; length of stay in hospital; use of epinephrine and intubation		
Notes Baseline croup score not presented for 2 children due to prior treatment with nebulise			
	Funding source: not reported		
Risk of bias			
Bias	Authors' judgement Support for judgement		
Random sequence generation (selection bias)	Unclear risk Comment: insufficient information provided to permit a judgement		



Godden 1997 (Continued)		
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "trial solution was supplied in an opaque respule within a sealed silver foil packet." "The patient initially received 4 mL of a solution containing either normal saline vehicle or 4mg (4mL) of budesonide, via an opaque nebuliser chamber."
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 8% (N = 7) withdrew, equal between groups (9% in study group, 7% in control group)
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: some potentially eligible children were not enrolled due to man- power constraints, which could have biased participant selection
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Huang 2021

Study characteristics	5		
Methods	Randomised controlled trial		
Participants	Study period: January 2019 to January 2020		
	Setting: Wuhan Children's Hospital, Wuhan, China		
	Inclusion criteria: children meeting the diagnostic criteria of acute infectious laryngitis in Zhufutang practical paediatrics confirmed by laryngoscope examination and diagnosed for the first time; had complete clinical data; written informed consent obtained from all guardians of patients		
	Exclusion criteria: children with systemic diseases, congenital larynx and other systemic serious diseases, allergic to drugs used in this study, and presence of bronchial foreign bodies		
	Baseline demographics (N = 92):		
	 proportion male: treatment 1: 63%; treatment 2: 56% mean (SD) age in years: treatment 1: 3.48 (0.28); treatment 2: 3.51 (0.31) 		
Interventions	Treatment 1 (N = 46): 1.00 mg/kg of dexamethasone injection twice a day		
	Treatment 2 (N = 46): budesonide inhalation therapy and 2.00 mL of budesonide added 3.00 mL normal saline twice a day		
Outcomes	Time for clinical symptoms to disappear after 3 days of treatment; therapeutic effects; adverse events		
Notes	No sources of funding reported.		
Risk of bias			



Huang 2021 (Continued)

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Quote: "patients were randomly allocated to either the study and control group with 46 cases in each group"
		Comment: stated only as randomly allocated. No details provided on the method of sequence generation.
Allocation concealment (selection bias)	Unclear risk	Comment: no description provided on the method of allocation concealment
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Comment: no description provided
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: no description provided
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Quote: "46 cases in each group; control group involved 18 males and 14 females"
		Comment: no participant flow diagram provided. Unclear if any participants were lost to follow-up. Baseline characteristics of proportion of males and females do not add up in the control group.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified
Other bias	Low risk	Quote: "general data of the two groups were comparable (P > 0.05)"
		Comment: baseline characteristics between groups were comparable with no significant differences
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Husby 1993

Study	, chai	acte	ristics
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Study characteristic	S
Methods	Randomised, double-blind controlled trial
Participants	Study period: October 1990 to December 1991
	Setting: Department of Paediatrics, Kolding Hospital, Denmark
	Inclusion criteria: children admitted to hospital with croup (inspiratory stridor, cough, and respiratory distress) with a modified Westley croup score > 5 and informed parental consent
	Exclusion criteria: clinical condition consistent with epiglottitis, foreign body aspiration, bronchiolitis, or asthma; received local or systemic steroid treatment or epinephrine
	Baseline demographics (N = 36) (1 child excluded before placebo was administered):
	 proportion male: treatment: 80%; control: 75% median (range) age in years: treatment: 1.6 (0.6 to 4.9); control: 1.1 (0.4 to 4.2)



Husby 1993 (Continued)	• median (range) mo	odified Westley croup score: treatment: 8 (6 to 10); control: 8 (6 to 12)	
Interventions	Treatment (N = 20): 2, 1000 μg (2 mL 500 μg/mL) doses of nebulised budesonide, 30 minutes apart		
	Control (N = 16): 2, 2 m	L doses of placebo (0.9% saline), 30 minutes apart	
	Both treatment and pla	acebo were given with a dynamic flow rate of 8 L/min.	
Outcomes	Change in modified We	estley croup score from baseline to 2 hours; use of antibiotics	
Notes	Funding source: not re	ported	
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Comment: described as double-blind, no further explanation. Insufficient information provided to permit a judgement.	
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel	
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 1 child omitted as did not receive treatment due to technical prob- lems. No other missing outcome data	
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.	
Other bias	Low risk	Comment: no other sources of bias identified	
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk	

James 1969

Study characteristic	s
Methods	Randomised, double-blind controlled trial
Participants Study period: September 1965 to November 1966 (excluding 2 periods during May and Aug total of 6 weeks)	
	Setting: children's division of the Los Angeles County-University of Southern California Medical Center, USA



James 1969 (Continued)

Inclusion criteria: children admitted to hospital with a diagnosis of croup or laryngotracheobronchitis (dyspnoea with inspiratory stridor, subcostal, suprasternal, or sternal retractions, and a barking, seal-like cough)

Exclusion criteria: very mild stridor at admission; history of persistent or congenital stridor; suspected diagnosis of acute epiglottitis; clinical or roentgenographic evidence of an associated pneumonitis

Baseline demographics (N = 88):

proportion male: treatment: 76%; control: 66%
median age in months: treatment: 17; control: 12
modified Westley croup score: not reported

Interventions

 $Treatment \ (N=45): single \ 4 \ mg/mL \ dose \ of intramuscular \ dexame thas one \ so dium \ phosphate$

Control (N = 43): single 4 mg/mL dose of placebo solution identical in appearance

Both treatment and placebo were administered 0 to 3 hours after admission.

Outcomes

Patient improvement at 12 and 24 hours; use of antibiotics and tracheostomy

Notes

Funding source: Merck Sharp & Dohme

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Low risk	Quote: "dexamethasone and a placebo solution of identical appearance were provided in randomly numbered vials. As each patient was admitted to the study, he received the predetermined dose of medication from the next bottle in the series, after which the bottle was discarded."
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "solution of identical appearance were provided in randomly numbered unlabeled vials" "All of the evaluations were completed before the drug code was broken"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "solution of identical appearance were provided in randomly numbered unlabeled vials" "All of the evaluations were completed before the drug code was broken"
		Comment: outcome measures assessed by staff and in some cases also by the investigator. Both staff and investigator appear to have been blinded.
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk



Johnson 1996

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: October 1989 to November 1991		
	Setting: emergency department of The Hospital for Sick Children, Toronto, Canada		
	Inclusion criteria: children presenting to the emergency department between 8:00 a.m. and me with signs and symptoms consistent with acute laryngotracheitis, including seal-like barking consistent most and inspiratory stridor that had developed 2 to 72 hours before they were seen; persistent most respiratory distress (modified Westley croup score 2.5 to 5) after having been treated for at least minutes with humidified oxygen provided by plastic tubing aimed toward the nose and mouth parental informed consent		
	Exclusion criteria: history of congenital stridor or endotracheal intubation longer than 1 month; sign or symptoms suggesting another cause of stridor (e.g. bacterial tracheitis, epiglottitis, supraglottic feign body, spasmodic croup); presence of marked expiratory wheeze that responded to treatment with bronchodilators; history of chronic respiratory problems such as asthma requiring routine daily treatment with bronchodilators; presence of a severe systemic disease that would affect the decision to mit or discharge; history of receiving corticosteroids in the last 2 weeks or racaemic epinephrine hydrochloride in the last 4 hours		
	Baseline demographi	cs (N = 55):	
	 proportion male: treatment: 71%; control: 56% median (25th, 75th percentile) age in months: treatment: 15 (11, 29); control: 17 (9, 22) median (25th, 75th percentile) modified Westley croup score: treatment: 4 (3, 4); control: 		
Interventions	Treatment (N = 28): single 10 mg (< 8 kg body weight), 15 mg (8 to 12 kg body weight), or 20 mg (> 1 kg body weight) dose of nebulised parenteral dexamethasone sodium phosphate solution (10 mg/mixed with normal saline to make 4 mL provided with 100% oxygen at a flow rate of 6 to 7 L/min Control (N = 27): placebo (normal saline) provided in the same fashion All children were also treated with humidified oxygen supplied by plastic tubing aimed towards the nose and mouth throughout the study period.		
Outcomes	Change in modified Westley croup score from baseline to 2 and 4 hours; hospitalisation rate; improvement at 4 hours; use of epinephrine, use of mist tent, intubation, and use of additional glucocorticoids		
Notes	Funding source: Pediatric Consultants, The Hospital for Sick Children, Toronto, Canada		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence genera-	Low risk	Quote: "Patients were randomized in blocks of 10"	
tion (selection bias)		Comment: block randomisation assumed to be computer-generated	
Allocation concealment (selection bias)	Low risk	Quote: "the randomized code was generated and held by the hospital pharmacy until after enrolment of all patients."	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "The pharmacy dispensed the appropriate medication in a vial specially prepared for this study, based on a randomization schedule" "To ensure blinding of investigators, staff, and parents, nebulizer containers were covered during and after nebulization"	



Johnson 1996 (Continued)		
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "The pharmacy dispensed the appropriate medication in a vial specially prepared for this study, based on a randomization schedule" "To ensure blinding of investigators, staff, and parents, nebulizer containers were covered during and after nebulization"
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 16% (N = 9) protocol deviations not included in the analysis. More losses in the intervention group, but unclear if this was significant
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Johnson 1998

Johnson 1998	
Study characteristic	s
Methods	Randomised, double-blind controlled trial
Participants	Study period: September 1993 to May 1996
	Setting: emergency departments of The Hospital for Sick Children in Toronto and Alberta Children's Hospital in Calgary, Canada
	Inclusion criteria: children aged from 3 months to 9 years presenting to the emergency department between 3:00 p.m. and 6:00 a.m. who were given a diagnosis of croup (acute onset of inspiratory stridor associated with a seal-like barking cough) with persistent moderately severe respiratory distress (modified Westley croup score of 3 to 6) after being treated with humidified oxygen for 30 minutes
	Exclusion criteria: signs and symptoms suggesting another cause of stridor (e.g. epiglottitis, bacterial tracheitis, supraglottic foreign body); parents who were unable to speak English well enough to give informed consent; history of chronic pulmonary disease, severe systemic disease, immune dysfunction, stridor, intubation for more than 1 month; glucocorticoid therapy in the 4 weeks prior to entering the study
	Baseline demographics (N = 144):
	 proportion males: 69% mean (SD) age in months: 24 (18) mean (SD) modified Westley croup score: treatment 1: 3.8 (0.9); treatment 2: 4.0 (0.9); control: 3.8 (0.8)
Interventions	Treatment 1 (N = 48): single 4 mg dose of nebulised budesonide
	Treatment 2 (N = 47): single 0.6 mg/kg dose of intramuscular dexamethasone
	Control (N = 49): single dose of nebulised placebo suspension
	All children received 0.5 mL of 2.25% racepinephrine and normal saline combined with either treatment or placebo (total volume 8 mL) via nebuliser with oxygen from a wall outlet at a rate of 6 to 7 L/min through a face mask held tightly to the child's face over a period of 20 minutes.



Jo	hnson	1998	(Continued)
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Outcomes Change in modified Westley croup score from baseline to 5 hours or discharge; rate of hospitalisation; use of epinephrine, use of supplemental glucocorticoids, intubations

Notes Funding source: Astra Pharma

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "A blocked randomization code was produced by random-number generating software."
Allocation concealment (selection bias)	Low risk	Quote: "A blocked randomization code provided only to the pharmacy at the hospital" "The pharmacies prepared sequential patient packets containing study drugs that were sealed and were identical in appearance and weight."
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Quote: "To make the nebulized study drugs indistinguishable from each other, they were packaged in opaque containers and discharged directly into a colored nebulizer" "to maintain masking, the study nurse temporarily took the parents away from their child while an emergency staff nurse not otherwise involved in the care of the child injected the dexamethasone into the child's thigh, placed a bandage over the injection site (all children received a bandage whether or not they received dexamethasone), and initiated nebulization." Comment: blinding was attempted, but it could have been broken. Unmasking occurred in 3 cases. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: outcome assessor described as blinded, but blinding could have been broken. Unmasking occurred in 3 cases. Subjective outcomes
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: intention-to-treat analysis used, including 17% (N = 25) with protocol deviations. 2% (N = 1) from the treatment group were lost to follow-up because parents could not be reached.
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Klassen 1994

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Studv	chara	ıcter	ristics

otady characteristic	•
Methods	Randomised, double-blind controlled trial
Participants	Study period: October 1992 to October 1993
	Setting: emergency department at the Children's Hospital of Eastern Ontario, Canada
	Inclusion criteria: children aged from 3 months to 5 years presenting to the emergency department between 9:00 a.m. and midnight (except holidays) with mild to moderate croup consisting of hoarseness, inspiratory stridor, and barking cough, and a Westley croup score ≥ 2 after breathing humidified oxygen for at least 15 minutes



Klassen 1994 (Continued)

Exclusion criteria: diagnosis of epiglottitis or chronic upper or lower airway disease (not including asthma); corticosteroids administered within the past 2 weeks; severe croup (defined as a Westley croup score of 8 or higher or requiring treatment with racaemic epinephrine immediately on arrival)

Baseline demographics (N = 54):

- **proportion males:** treatment: 63%; control: 74%
- mean (SD) age in years: treatment: 1.8 (1.2); control: 2.2 (1.4)
- median (25th, 75th percentile) Westley croup score: treatment: 4 (3, 5); control: 4 (3, 5)

Interventions

Treatment (N = 27): single 2 mg (4 mL) dose of nebulised budesonide

Control (N = 27): single 4 mL dose of nebulised placebo (0.9% saline solution)

Both treatment and placebo administered by an updraft nebuliser with a continuous flow of oxygen at 5 to 6 L/min.

Outcomes

Change in Westley croup score from baseline to 4 hours; admissions to the hospital; 2-point improvement in croup score at 4 hours; use of epinephrine, use of supplemental glucocorticoids

Notes

Funding source: Ontario Ministry of Health

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "Randomization was performed in blocks of 10 by the pharmacy department, with a random number table"
Allocation concealment (selection bias)	Low risk	Quote: "The randomization list was kept concealed from the research assistants, parents and emergency physicians and from the child's regular physician until the end of the trial." "the pharmacy provided both budesonide and normal saline in opaque brown syringes"
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "the pharmacy provided both budesonide and normal saline in opaque brown syringes to ensure blinding. The research assistants then placed the study drug directly into an opaque nebulizer reservoir."
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "the pharmacy provided both budesonide and normal saline in opaque brown syringes to ensure blinding. The research assistants then placed the study drug directly into an opaque nebulizer reservoir."
		Comment: research assistants blinded
Incomplete outcome data (attrition bias) All outcomes	High risk	Did not use intention-to-treat analysis. 7% (N = 4) lost, all from the placebo group (15%) due to worsening condition or lack of satisfaction with treatment. Unbalanced between groups
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: 24 children were not enrolled because the emergency department failed to contact the study team; this could potentially have biased participant selection
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk



Klassen 1996

Study characteristics			
Methods	Randomised, double-b	olind controlled trial	
Participants	Study period: October 1993 to April 1994		
	Setting: emergency department of the Children's Hospital of Eastern Ontario, Canada		
	Inclusion criteria: children aged 3 months to 5 years presenting to the emergency department with mild to moderate croup (hoarseness, inspiratory stridor, barking cough) and a modified Westley croup score ≥ 3 after at least 15 minutes of mist therapy		
	Exclusion criteria: diagnosis of epiglottitis, chronic upper or lower airway disease (excluding asthma), severe croup (modified Westley croup score ≥ 8); received glucocorticoids within the previous 2 weeks; needed immediate racaemic epinephrine on arrival		
	Baseline demographi	cs (N = 50):	
	 proportion males: 	treatment: 68%; control: 76%	
	· · · · · · ·	ears: treatment: 1.2 (0.7); control: 1.8 (1.3)	
	 mean (SD) modifie 	d Westley croup score: treatment: 4.4 (1.1); control: 4.1 (0.9)	
Interventions	All children received a	single 0.60 mg/kg dose of oral dexamethasone upon entry into the study.	
	Treatment (N = 25): single 2 mg (4 mL) dose of nebulised budesonide		
	Control (N = 25): single 4 mL dose of placebo (0.9% saline solution)		
	Both treatment and control delivered by an updraft nebuliser with a continuous flow of oxygen at a rate of 5 to 6 L/min.		
Outcomes	Change in modified Westley croup score from baseline to 4 hours; admissions to the hospital; 2-point improvement in croup score at 4 hours; use of epinephrine, use of supplemental glucocorticoids, and use of mist tent		
Notes	Funding source: Ontario Ministry of Health		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "Randomization was performed in blocks of 10 by the pharmacy department, using a random numbers table"	
Allocation concealment (selection bias)	Low risk	Quote: "Randomization was performed by the pharmacy department"; "the pharmacy provided both budesonide and normal saline in opaque, brown syringes."; "the randomization code was revealed only after all patients had completed the trial"	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "Budesonide is slightly opaque; therefore, to conceal its identity, the pharmacy provided both budesonide and normal saline in opaque, brown syringes. The research assistants placed the drugs directly into an opaque nebulizer reservoir. Once nebulized, the drugs were indistinguishable by sight and smell." "Both the research assistants and the physicians caring for the patients in the emergency department were blinded to treatment assignment"	
Blinding of outcome assessment (detection bias)	Low risk	Quote: "Budesonide is slightly opaque; therefore, to conceal its identity, the pharmacy provided both budesonide and normal saline in opaque, brown sy-	



Klassen 1996 (Continued) All outcomes		ringes. The research assistants placed the drugs directly into an opaque nebulizer reservoir. Once nebulized, the drugs were indistinguishable by sight and smell." "Both the research assistants and the physicians caring for the patients in the emergency department were blinded to treatment assignment"
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 2% (N = 1) in the placebo group required racaemic epinephrine and was excluded; 2% (N = 1) lost to follow-up in the treatment group because parent could not be contacted for follow-up
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: 33 children were not enrolled because the study team was not contacted; this could potentially have biased participant selection
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Klassen 1998

Study characteristics	3
Methods	Randomised, double-blind controlled trial
Participants	Study period : October 1995 to April 1996 and October 1996 to January 1997
	Setting: emergency departments of the Children's Hospital of Eastern Ontario, Ottawa, or the Winnipeg Children's Hospital, Winnipeg, Canada
	Inclusion criteria: children aged 3 months to 5 years who presented to the emergency department with croup (hoarseness, inspiratory stridor, and barking cough) and Westley croup score of ≥ 2 following at least 15 minutes of mist therapy; parents available for telephone follow-up a week after enrolling in the study
	Exclusion criteria: epiglottitis; chronic respiratory disease (except asthma); severe croup (Westley croup score ≥ 8); racaemic epinephrine treatment upon arriving at the emergency department; glucocorticoids in the last 2 weeks; history of tuberculosis in child or household; chickenpox or exposure to it within the past 21 days; known immunodeficiency
	Baseline demographics (N = 198):
	• proportion males: treatment 1: 77%; treatment 2: 62%; treatment 3: 64%
	• median (25th, 75th percentile) age in years: treatment 1: 1.5 (1.0, 2.2); treatment 2: 1.3 (0.8, 2.1); treatment 3: 1.6 (1.0, 2.5)
	• mean (95% CI) Westley croup score: treatment 1: 3.5 (3.2 to 3.7); treatment 2: 3.6 (3.3 to 3.8); treatment 3: 3.8 (3.5 to 4.0)
Interventions	Treatment 1 (N = 65): single 2 mg (4 mL) dose of nebulised budesonide plus the appropriate volume of oral placebo (clear syrup solution)
	Treatment 2 (N = 69): single 4 mL dose of nebulised placebo (saline solution) plus 0.6 mg/kg oral dexamethasone
	Treatment 3 (N = 64): single 4 mL dose of nebulised budesonide plus 0.6 mg/kg of oral dexamethasone
Outcomes	Change in Westley croup score from baseline to 4 hours (or until discharge); admissions to hospital; length of stay in the emergency department; 2-point improvement in croup score at 4 hours; use of epinephrine and use of additional glucocorticoids



Klassen 1998 (Continued)

Notes

Funding source: Ontario Ministry of Health; Emergency Health Services, Toronto, Ontario; Manitoba Medical Services Foundation

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "A central pharmacy randomized individuals to the 3 groups, using computer-generated random numbers in random blocks of 6 or 9"
Allocation concealment (selection bias)	Low risk	Quote: "the list was kept at a central pharmacy until the end of the study to ensure allocation concealment. Because the drugs were packaged identically and identified only by a sequential study number, the research assistant who administered the intervention remained unaware of the next group assignment"
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "Dexamethasone syrup and placebo dexamethasone syrup were identical in taste and appearance All solutions were packaged in brown syringes and the research assistant instilled either solution directly into an opaque nebulizer reservoir."
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "Dexamethasone syrup and placebo dexamethasone syrup were identical in taste and appearance All solutions were packaged in brown syringes and the research assistant instilled either solution directly into an opaque nebulizer reservoir."
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: intention-to-treat analysis used. 1% (N = 1) loss to follow-up in the treatment group
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: 10 children were not enrolled because the study team was not contacted; this could potentially have biased participant selection
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Koren 1983

Studv	chara	cteristics

Study Characteristics	
Methods	Randomised, double-blind controlled trial
Participants	Study period: January 1979 to January 1980
	Setting: paediatric division of the Chaim Sheba Medical Center, Tel Hashomer, Israel
	Inclusion criteria: children aged 8 months to 8 years hospitalised with croup (all with inspiratory stridor, dyspnoea, subcostal and/or suprasternal retraction, and a barking cough); informed consent given by the parents
	Exclusion criteria: evidence of associated bronchitis, bronchopneumonia, and acute epiglottitis
	Baseline demographics (N = 78):



Koren 1983	(Continued)
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- **proportion males:** laryngotracheitis (LT) group: treatment: 59%; control: 56%; spasmodic croup (SC) group: treatment: 57%; control: 64%
- mean age in years, months: LT group: treatment: 2 years, 5 months; control: 2 years, 7 months; SC group: treatment: 2 years, 6 months; control: 2 years, 8 months
- croup score: not measured

Interventions

Children in all groups were sedated via rectal administration of 75 mg/kg of body weight of chloral hydrate

Treatment (N = 40): single 0.60 mg/kg dose of intramuscular dexamethasone sodium phosphate (4 mg/ mL)

Control (N = 38): single dose of intramuscular placebo solution of identical appearance

Outcomes Use of epinephrine and use of antibiotics

Notes Funding source: not reported

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "Dexamethasone sodium phosphate and a placebo solution of identical appearance were administered" "All evaluations were completed before the study code was opened"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Kuusela 1988

Study	char	actoi	rictics

Methods	Randomised, double-blind controlled trial
Participants	Study period: October 1984 to March 1985 and January 1986 to April 1986



Kuusela 1988 (Continued)

Setting: Department of Paediatrics, Tampere University Central Hospital, Finland

Inclusion criteria: children diagnosed and admitted for acute laryngitis (croup)

Exclusion criteria: not reported

Baseline demographics (N = 72):

- proportion males: treatment 1: 74%; treatment 2: 88%; treatment 3: 88%; control: 67%
- mean (SD) age in years: treatment 1: 2.9 (1.5); treatment 2: 2.3 (1.7); treatment 3: 2.8 (1.8); control: 3.2 (2.7)
- · croup score: not measured

Interventions

All children received initial treatment in the emergency department, then were transferred to a ward where they were placed in a humid room and given oral fluids at a minimum of 20 mL/kg body weight over the next 4 hours.

Treatment 1 (N = 19): single 0.6 mg/kg (0.12 mL/kg, maximum 2 mL) dose of intramuscular dexamethasone plus at least 1, 0.25 mL/5 kg (maximum 1.5 mL) dose of nebulised L-epinephrine (2.25% epinephrine base with 0.5% chlorobutanol as preservative). Additional doses of nebulised L-epinephrine every 2 hours as needed

Treatment 2 (N = 16): single 0.6 mg/kg (0.12 mL/kg, maximum 2 mL) dose of intramuscular dexamethasone plus at least 1, 0.25 mL/5 kg (maximum 1.5 mL) dose of nebulised placebo solution. Additional doses of nebulised placebo every 2 hours as needed

Treatment 3 (N = 16): single 0.6 mg/kg (0.12 mL/kg, maximum 2 mL) dose of intramuscular placebo plus at least 1, 0.25 mL/5 kg (maximum 1.5 mL) dose of nebulised L-epinephrine (2.25% epinephrine base with 0.5% chlorobutanol as preservative). Additional doses of nebulised L-epinephrine every 2 hours as needed

Control (N = 21): single 0.6 mg/kg ($0.12 \, \text{mL/kg}$, maximum 2 mL) dose of intramuscular placebo plus at least 1, $0.25 \, \text{mL/5}$ kg (maximum 1.5 mL) dose of nebulised placebo solution. Additional doses of nebulised placebo every 2 hours as needed

Outcomes

Change in clinical score based on dyspnoea and cough scale from baseline to 6, 12, and 24 hours; length of stay in hospital

Notes

Funding source: not reported

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Low risk	Quote: "the ampules were unlabelled, numbered, and randomized; the code was not available for investigators until the end of the study"
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "The placebo preparation consisted of the corresponding diluent supplied in similar ampules."; "the code was not available for the investigators until the end of the study"; "The active solution and the placebo preparation were identically packed in individual randomized vials containing 10 ml"
Blinding of outcome as- sessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias)	Unclear risk	Comment: did not use intention-to-treat analysis. 8% (N = 6) excluded due to protocol violations, unclear what group they were in



K	Luusela 1988 (Continued)		
	All outcomes		
-			

Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Leipzig 1979

Study characteristics	

Methods	Randomised, double-blind controlled trial
Participants	Study period: November 1976 to March 1978

Setting: Pediatric Service of the State University Hospital or the Crouse-Irving Memorial Hospital, Syracuse, NY, USA

Inclusion criteria: all children admitted to hospital with a diagnosis of croup with disease of sufficient severity on a predetermined scoring system; consent of the child's physician and parents

Exclusion criteria: not reported

Baseline demographics (N = 30):

- proportion males: not reported
- mean age in months: treatment: 21.3; control: 21.0
- mean (SD) croup score: treatment: 8.46 (1.45); control: 8.14 (1.46)

Interventions

Treatment (N = 16): 2, 0.30 mg/kg doses of intramuscular dexamethasone (4 mg/mL)

Control (N = 14): 2 doses of intramuscular placebo (sterile saline) (1 dose initially and another 2 hours later)

Outcomes

Change in croup score from baseline to 24 hours; length of stay at hospital; intubation

Notes

Funding source: not reported

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "assigned from a table of random numbers"
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias)	Unclear risk	Quote: "Vials had been previously prepared containing either dexamethasone (4 mg/L) or sterile saline. They were marked only with a number, assigned from a table of random numbers"
All outcomes		Comment: described as double-blind. Unclear who was blinded and who prepared the vials. Subjective outcomes



Leipzig 1979 (Continued)		
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Quote: "Vials had been previously prepared containing either dexamethasone (4 mg/L) or sterile saline. They were marked only with a number, assigned from a table of random numbers"
		Comment: described as double-blind. Unclear who was blinded and who prepared the vials. Subjective outcomes
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Luria 2001

Study characteristics	
Methods	Randomised, double-blind controlled trial
Participants	Study period: September 1995 to December 1997
	Setting: emergency departments at either Children's Hospital Medical Center in Cincinnati or Children's Hospital in Columbus, OH, USA
	Inclusion criteria: children aged 6 months to 6 years presenting to the emergency department with mild croup (barky cough, stridor and/or hoarseness for < 48 hours) and having a viral prodrome consisting of fever, cough, or rhinorrhoea (in an attempt to exclude children with spasmodic croup)
	Exclusion criteria: treated with corticosteroids 14 days prior to enrolling in the study; a clinical picture consistent with spasmodic croup; history of prolonged endotracheal intubation; history of chronic respiratory illness (i.e. asthma or cystic fibrosis); a condition associated with airway abnormalities; those without a working telephone or with a severe disease (i.e. received nebulised racaemic epinephrine or corticosteroids at the order of the emergency department physician or had < 94% oxygen saturation)
	Baseline demographics (N = 264):
	• proportion males: treatment 1: 72%; treatment 2: 64%; control: 65%
	• mean (range) age in months: treatment 1: 28 (6 to 70); treatment 2: 31 (6 to 71); control: 26 (6 to 71)
	• mean (range) modified Westley croup score: treatment 1: 1.6 (0 to 6); treatment 2: 1.6 (0 to 5); control: 1.7 (0 to 5)
Interventions	Treatment group 1 (N = 85): single 0.60 mg/kg (maximum 10 mg) dose of oral dexamethasone (1 mg/mL) plus nebulised placebo solution
	Treatment group 2 (N = 91): single dose of oral placebo plus a 160 μg dose of nebulised dexamethasone sodium phosphate
	Control (N = 88): single dose of oral placebo plus nebulised placebo solution
	Nebulised study preparations were delivered with a nebuliser that had a fill volume of 3 mL and the oxygen flow set at 5 to 6 L/min.



Luria 2001	(Continued)
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Outcomes	Return visits to the emergency department	
Notes	Funding source: the Bremer Foundation at the Ohio State University, Columbus, OH, USA	

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "The randomization was performed in blocks of 15 by the study pharmacist at each enrolling site with the use of a random number generator"
Allocation concealment (selection bias)	Low risk	Quote: "The study pharmacist then assembled numbered 'croup kits' containing study preparations that reflected the results of the randomization." "The study physician retrieved the lowest numbered kit when enrolling a new subject to maintain the randomization order. Only pharmacists knew the results of the randomization."
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "The kits were sealed to prevent any tampering and were kept in the EDs Only the study pharmacists knew the results of the randomization All oral study preparations were mixed 1:1 with a commercially available grape flavouring to minimize taste bias."
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: intention-to-treat analysis used, including N = 9 protocol deviations. 16% (N = 43) loss on day 7 for the telephone follow-up (14% in the oral treatment group, 15% in the nebulised treatment group, 19% in the placebo group)
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Martinez Fernandez 1993

Studv char	racteristics
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Study Characteristic	
Methods	Randomised, double-blind controlled trial
Participants	Study period: October 1989 to September 1990
	Setting: children's hospital in Spain
	Inclusion criteria: children hospitalised with symptoms suggestive of croup (acute laryngitis, laryngotracheobronchitis, spasmodic croup)
	Exclusion criteria: child's croup judged by the physician to be too severe
	Baseline characteristics (N = 66):
	proportion males: not reported



Martinez Fernandez 1993 (Continued)

- · age: not reported
- **mean (SD) croup score:** treatment 1: 3.5 (1.7); treatment 2: 2.9 (1.4); treatment 3: 3.3 (1.1); control: 3.2 (1.5)

Interventions

Treatment 1 (N = 15): single dose of intramuscular placebo, plus 0.14% nebulised L-epinephrine initially and every 4 hours as needed

Treatment 2 (N = 16): single 0.5 mg/kg dose of intramuscular dexamethasone, plus nebulised placebo (saline) initially and every 4 hours as needed

Treatment 3 (N = 18): single 0.5 mg/kg dose of intramuscular dexamethasone, plus 0.14% nebulised Lepinephrine initially and every 4 hours as needed

Control (N = 17): single dose of intramuscular placebo, plus nebulised placebo (saline) initially and every 4 hours as needed

All children received humidified oxygen and fluid therapy.

Outcomes

Change in croup score from baseline to 6, 12, and 24 hours

Notes

Written in Spanish

Funding source: not reported

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Low risk	Comment: treatments shipped in pre-numbered ampoules, unlabelled and randomly ordered by the pharmacy
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Comment: described as double-blind. Treatments shipped in pre-numbered ampoules, unlabelled and randomly ordered by the pharmacy.
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk



Massicotte 1973

Study characteristics			
Methods	Randomised, double-b	lind controlled trial	
Participants	Study period: December 1971 to March 1972		
	Setting: L'Hôpital Sain	nte-Justine, Montreal, Quebec, Canada	
	Inclusion criteria: chil	dren admitted to hospital with severe croup (croup score ≥ 9)	
	tered consciousness, p	Id or moderate croup (croup score < 9); require tracheotomy on admission (aleribuccal cyanosis); those with epiglottitis, foreign body aspiration, diphtheria, cute or chronic medical conditions; corticosteroids in the past 48 hours, allergy lin	
	Baseline characterist	ics (N = 42):	
	 proportion males: 	treatment: 80%; control: 80%	
	 age: not reported, r 	nost < 4 years	
	• mean croup score: control: 12.27	acute onset: treatment: 12.55; control: 12.33; progressive onset: treatment: 12.38;	
Interventions	All children were placed in a humidified room and received intravenous saline, and additionally received 100 mg ampicillin/24 hours in 4 doses (1 dose every 6 hours) over the course of 10 days.		
	Treatment (N = 25): single 4 mg/kg dose of intravenous methyl-prednisolone initially (40 mg for 6 to 8 kg; 60 mg for 9 to 12 kg; 80 mg for 13 to 16 kg; 120 mg for 17 to 20 kg), followed by repeated doses at 4 and 8 hours, if needed		
	Control (N = 17): single doses at 4 and 8 hours,	4 mg/kg dose of intravenous placebo (lactose) initially, followed by repeated if needed	
Outcomes	Change in croup score from baseline to 4 and 14 hours; patient improvement at 4 and 14 hours		
Notes	Written in French Funding source: Upjohn Canada		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Comment: random numbers table used	
Allocation concealment (selection bias)	Low risk	Comment: the treatments were identical in appearance, containing either methyl-prednisolone or placebo. Administration was double-blind, and the code was not broken until the last child had completed the study.	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Comment: the treatments were identical in appearance, containing either methyl-prednisolone or placebo. Administration was double-blind, and the code was not broken until the last child had completed the study.	
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel	
Incomplete outcome data (attrition bias)	Low risk	Comment: no missing outcome data	



Massicotte 1973 (Continued)

All outcomes

Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Parker 2019

Study characteristics			
Methods	Randomised, double-b	olind controlled trial	
Participants	Study period: March 2	2009 to July 2012	
	<u> </u>	nergency department of Princess Margaret Hospital for Children and emergency Ilup Health Campus, Perth, Western Australia	
		ldren aged > 6 months, contactable by telephone, and English-speaking careight of 20 kg was imposed to limit the maximum possible dexamethasone dose	
	treatment, steroid the ical suspicion of an alt	own prednisolone or dexamethasone allergy, immunosuppressive disease or rapy or enrolment in the same study within the previous 14 days, and a high clinernative diagnosis, with specific prompts to include bacterial tracheitis, inhaled tryngeal abscess, epiglottitis, angio-oedema, vascular ring, and subglottic steno-	
	Baseline demographi	cs (N = 1231):	
	• mean (SD) age in n	reatment 1: 61%; treatment 2: 62%; treatment 3: 63% nonths: treatment 1: 29.2 (17.3); treatment 2: 30.5 (16.3); treatment 3: 30.4 (16.2) ed Westley croup score: treatment 1: 1.4 (1.4); treatment 2: 1.5 (1.4); treatment 3:	
Interventions	Treatment 1 (N = 410): single 0.60 mg/kg dose of oral dexamethasone		
	Treatment 2 (N = 410): single 0.15 mg/kg dose of oral dexamethasone		
	Treatment 3 (N = 411): single 1.00 mg/kg dose of prednisolone		
Outcomes	Change in Westley croup score from baseline to 2 hours; reattendance with croup; length of stay; use of epinephrine; intubation; use of supplemental glucocorticoids; any adverse events		
Notes	Funding source: Princess Margaret Hospital Foundation		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "randomization list was computer-generated at www.randomization.com by using randomly permuted block sizes in the ratio of 4 patients from each group, with block randomization by center"	



Parker 2019 (Continued)		
Allocation concealment (selection bias)	Low risk	Quote: "All medications were prepared, randomly assigned, and labeled by a clinical trials pharmacist at Princess Margaret Hospital."
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "Both staff (administering and assessing treatments) and patients were therefore blinded to the treatment allocation."
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "Both staff (administering and assessing treatments) and patients were therefore blinded to the treatment allocation."
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Quote: "The analysis was conducted per protocol via intention to treat" "The distribution of repeat enrollments and patients meeting an exclusion criterion was relatively balanced across treatment groups (dexamethasone/low-dose dexamethasone/prednisolone distributed at 1:1.48:1.14 and 1:1.29: 1.71, respectively)"
		Comment: 141/1231 (11%) participants were missing at hour 1 of croup assessment with reasons for their exclusion unexplained
Selective reporting (reporting bias)	Low risk	Comment: all prespecified study outcomes of interest have been reported in the results
Other bias	Unclear risk	Comment: several children were enrolled in this study more than once
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain was judged as unclear risk

Rittichier 2000

Study characteristics	
Methods	Randomised controlled trial

Participants

Study period: October 1996 to June 1999

Setting: emergency department at children's hospital in Denver, CO, USA

Inclusion criteria: children aged 3 months to 12 years presenting to the emergency department with moderate croup (hoarseness and barky cough associated with either a history/presence of stridor at rest, and/or retractions) of < 48 hours onset of illness (defined as onset of barky cough)

Exclusion criteria: children with epiglottitis; foreign body aspiration; reactive airway exacerbation; acute bacterial pneumonia; acquired or congenital upper airway anomalies such a tracheomalacia; immunocompromised; history of steroid exposure in the previous 2 weeks; children with mild croup (history or presence of a barky cough without the presence/history of the associated stridor or retractions); children with severe croup who had altered mental status, severe retractions, cyanosis associated with their croup; admitted to the hospital during the initial emergency department visit, either before consideration or after being enrolled in the study

Baseline demographics (N = 277):

- proportion male: treatment 1: 70%; treatment 2: 69%
- median (SD) age in years: treatment 1: 2.03 (1.81); treatment 2: 2.01 (1.84)
- mean croup score: treatment 1: 2.09; treatment 2: 1.95



Rittichier 2000 (Continued)			
Interventions	All children were given	a cool mist therapy per emergency department protocol.	
	Treatment 1 (N = 139):	single 0.60 mg/kg (maximum 8 mg) dose of intramuscular dexamethasone	
	Treatment 2 (N = 138):	single 0.60 mg/kg (maximum 8 mg) dose of oral dexamethasone	
Outcomes		Change in Westley croup score from baseline to 24 hours; unscheduled return visits to the emergency department; use of epinephrine, use of additional glucocorticoids, use of mist tent, and use of antibiotics	
Notes	Funding source: not re	ported	
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "using a random allocation chart based on a table of random numbers"	
Allocation concealment	Unclear risk	Quote: "The randomization code was held by the nursing staff in the ED"	
(selection bias)		Comment: randomisation code was held by nurses in the ED, and enrolment was performed by physicians, fellows, and residents. Unclear whether they could have determined the allocation sequence	
Blinding of participants and personnel (perfor-	High risk	Quote: "Nurses administered the dexamethasone either orally or intramuscularly per hospital protocol."	
mance bias) All outcomes		Comment: no blinding of participants and personnel. Subjective outcomes	
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Quote: "To help with blinding of the physician Band-Aids were placed on all patients whether they received PO or IM medicine." "Caretakers were contacted by a caller who was blinded to the route of administration Caretakers were instructed to not disclose the route of administration of the medicine to the caller."	
		Comment: blinding was attempted but could have been broken. Subjective outcomes	
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 16% lost to follow-up (N = 13 protocol deviations; N = 27 could not be reached, of which data for N = 10 were lost in a hospital move). Unclear if losses were balanced between groups	
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.	
Other bias	Low risk	Comment: no other sources of bias identified	
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk	

Roberts 1999

Study characteristics	
Methods	Randomised, double-blind controlled trial



Roberts 1999 (Continued)

Participants

Study period: April 1994 to April 1996

Setting: infectious diseases ward of the Women's and Children's Hospital in North Adelaide, Australia

Inclusion criteria: children aged 6 months to 8 years admitted with croup (inspiratory stridor, barking cough, hoarse voice, and respiratory distress) and a croup score of ≥ 4; stable croup (2 scores taken 15 minutes apart within 1 point of each other); written informed consent from the parent

Exclusion criteria: suspected epiglottitis, foreign body aspiration, bronchiolitis, pneumonia, or active asthma; intubation due to airways disease in the previous 12 months; acute wheezing; treatment with corticosteroids in the previous 4 weeks; treatment with adrenaline in the previous week; significant past or present pulmonary, cardiovascular, renal, hepatic, gastrointestinal, neurological, or endocrine disease that could interfere with the study; children unable to inhale the nebuliser mist for at least 1.5 minutes

Baseline demographics (N = 82):

- proportion males: treatment: 76%; control: 78%
- mean (SD) age in years: treatment: 2.3 (1.4); control: 2.2 (1.0)
- mean (SD) croup score: treatment: 6.4 (1.5); control: 6.3 (1.4)

Interventions

Treatment (N = 42): 2 mg/4 mL dose of nebulised budesonide every 12 hours for a maximum of 4 doses

Control (N = 40): dose of nebulised placebo (same formulation but without budesonide) every 12 hours for a maximum of 4 doses

Both treatment and placebo driven by an air or oxygen flow of 6 L/min.

Outcomes

Change in croup score from baseline to 2, 6, 12, 24 hours; revisits to the hospital for follow-up; 2-point improvement in croup score at 2, 6, and 12 hours; use of epinephrine

Notes

Used a croup score similar to Leipzig 1979 with alterations to stridor assessment, oxygen saturation, and temperature

Funding source: Astra Draco

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "Computer-generated randomization in blocks of six were performed by Astra Draco"
Allocation concealment (selection bias)	Low risk	Quote: "All study medication was packaged identically, identified only by a study number the randomisation code was kept at Astra Draco and broken only after study completion, hence all treatment decisions were made without awareness of study allocation."
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "All respules and all ventstreams used in the study were made of an opaque plastic to conceal any differences between the active and placebo doses" "The randomization code was kept at Astra Draco and only broken after study completion"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "All respules and all ventstreams used in the study were made of an opaque plastic to conceal any differences between the active and placebo doses" "The randomization code was kept at Astra Draco and only broken after study completion"
Incomplete outcome data (attrition bias)	High risk	Comment: 20% (N = 10) lost due to withdrawals, protocol deviations, inability to contact parents. Used the last value extended principle for analysis



Roberts 1999	(Continued)
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All outcomes

Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk

Roorda 1998

Stuay	cnaracteristics	
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Methods	Randomised, double-blind controlled trial		
Participants	Study period: October 1995 to March 1997		
	Setting: hospital in the Netherlands		
	Inclusion criteria: children aged 4 to 52 months hospitalised with moderate croup		
	Exclusion criteria: received systemic steroids, any bronchodilators, or antibiotics in the previous 48 hours		
	Baseline demographics (N = 17):		
	• proportion male: treatment: 89%; control: 63%		
	 mean (range) age in months: treatment: 29 (6 to 44); control: 38 (4 to 52) 		
	• mean (range) modified Westley croup score: treatment: 3.1 (1 to 5); control: 2.9 (1 to 8)		
Interventions	Treatment (N = 9): 2 doses of 1000 μg of fluticasone propionate administered by metred dose inhaler (4 puffs of 250 μg), 30 minutes apart		
	Control (N = 8): placebo administered in a similar fashion		
Outcomes	Change in modified Westley croup score from baseline to 2, 6, and 24 hours; length of stay in hospital; use of additional glucocorticoids and intubation		
Notes	Funding source: not reported		

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Comment: described as double-blind. Unclear who was blinded. Subjective outcomes



Roorda 1998 (Continued)		
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: described as double-blind. Unclear who was blinded. Subjective outcomes
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Skowron 1966a

Study characteristics	•
Methods	Randomised, double-blind controlled trial
Participants	Study period: December 1964 to March 1965
	Setting: tracheitis ward of The Hospital for Sick Children, Toronto, Canada
	Inclusion criteria: children hospitalised with croup
	Exclusion criteria: not reported
	Baseline demographics (N = 200 in total, N = 94 for 1.0 mL dexamethasone compared to placebo, 6 excluded):
	• proportion males: 77%
	• mean age in years: 2.3
	croup score: not measured
Interventions	$1.0\mathrm{mL}$ (4 mg) dexamethas one compared to placebo (see Skowron 1966b for 1.5 mL (6 mg) dexame thasone compared to placebo)
	All children were placed in a croupette with moist air, given twice-daily intramuscular crystalline sodium penicillin (825,000 IU/day) and streptomycin sulphate (0.5 g), as well as secobarbital, 3/4 grain per rectum on admission for children over 6 months.
	Treatment (N = 41): 1.0 mL (4 mg, based on approximately 0.4 mg/kg) subcutaneous dexamethasone every 6 hours for a total of 4 doses
	Control (N = 53): 1.0 mL placebo every 6 hours for a total of 4 doses
Outcomes	Readmissions to the hospital; length of stay in the hospital; tracheotomy
Notes	Funding source: not reported
Risk of bias	
Bias	Authors' judgement Support for judgement



Skowron 1966a (Continued)		
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Quote: "The test material was provided in labelled vials, the content of which was unknown to the investigators. As each child was admitted to the series, he received subcutaneously either material A or material B, according to a random selection code"
		Comment: bottles were not sequentially numbered, but instead labelled A or B. Unclear where the random selection code was held
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "The test material was provided in labelled vials, the content of which was unknown to the investigators." "After the results were documented, the code was broken"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "The test material was provided in labelled vials, the content of which was unknown to the investigators." "After the results were documented, the code was broken"
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 3% (N = 6) lost due to protocol deviations
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: no baseline data presented, impossible to judge if baseline imbalances existed
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Skowron 1966a and b

Skowioli 1966a aliu b				
Study characteristics				
Methods	See Skowron 1966a and	See Skowron 1966a and Skowron 1966b		
Participants	See Skowron 1966a and	See Skowron 1966a and Skowron 1966b		
Interventions	See Skowron 1966a and Skowron 1966b			
Outcomes	See Skowron 1966a and Skowron 1966b			
Notes	See Skowron 1966a and Skowron 1966b			
Risk of bias				
Bias	Authors' judgement	Support for judgement		
Random sequence generation (selection bias)	Unclear risk	See Skowron 1966a and Skowron 1966b		



Skowron 1966a and b (Continued)			
Allocation concealment (selection bias)	Unclear risk	See Skowron 1966a and Skowron 1966b	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	See Skowron 1966a and Skowron 1966b	
Blinding of outcome assessment (detection bias) All outcomes	Low risk	See Skowron 1966a and Skowron 1966b	
Incomplete outcome data (attrition bias) All outcomes	Low risk	See Skowron 1966a and Skowron 1966b	
Selective reporting (reporting bias)	Unclear risk	See Skowron 1966a and Skowron 1966b	
Other bias	Unclear risk	See Skowron 1966a and Skowron 1966b	
Overall risk of bias All outcomes	Unclear risk	See Skowron 1966a and Skowron 1966b	

Skowron 1966b

Study characteristics	<u> </u>
Methods	Randomised, double-blind controlled trial
Participants	Study period: December 1964 to March 1965
	Setting: tracheitis ward of The Hospital for Sick Children, Toronto, Canada
	Inclusion criteria: children hospitalised with croup
	Exclusion criteria: not reported
	Baseline demographics (N = 200 in total, N = 100 for 1.5 mL dexamethasone compared to place-bo):
	 proportion males: 77% mean age in years: 2.3 croup score: not measured
Interventions	1.5 mL (6 mg) dexamethasone compared to placebo (see Skowron 1966a for 1.0 mL dexamethasone compared to placebo)
	All children were placed in a croupette with moist air, given twice-daily intramuscular crystalline sodium penicillin (825,000 IU/day) and streptomycin sulphate (0.5 g), as well as secobarbital, 3/4 grain per rectum on admission for children over 6 months.
	Treatment (N = 56): 1.5 mL (6 mg, based on approximately 0.5 mg/kg) subcutaneous dexamethas one every 6 hours for a total of 4 doses
	Control (N = 44): 1.5 mL placebo every 6 hours for a total of 4 doses



Skowron 1966b (Continued)

Outcomes	Readmissions to the hospital; length of stay in the hospital; tracheotomy	
Notes	Funding source: not reported	

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Allocation concealment (selection bias)	Unclear risk	Quote: "The test material was provided in labelled vials, the content of which was unknown to the investigators. As each child was admitted to the series, he received subcutaneously either material A or material B, according to a random selection code"
		Comment: bottles were not sequentially numbered, but instead labelled A or B. Unclear where the random selection code was held
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "The test material was provided in labelled vials, the content of which was unknown to the investigators." "After the results were documented, the code was broken"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "The test material was provided in labelled vials, the content of which was unknown to the investigators." "After the results were documented, the code was broken"
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 3% (N = 6) lost due to protocol deviations
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: no baseline data presented, impossible to judge if baseline imbalances existed
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Soleimani 2013

Study	chara	ctoristics

Stuay cnaracteristic	S
Methods	Randomised, single-blind controlled trial
Participants	Study period: January 2009 to March 2010
	Setting: emergency department at Ali-Ebne Abitaleb Hospital, Iran
	Inclusion criteria: children aged 6 months to 6 years admitted to the emergency department with barking cough, stridor, hoarseness, and respiratory distress
	Exclusion criteria: chronic pulmonary disease; severe croup (croup score > 7); recurrent croup; allergy to corticosteroids; contraindication of corticosteroid (history of tuberous sclerosis, history of varicella



Soleimani 2013 (Continued)

infection during the past 3 weeks); history of corticosteroid administration during the last 4 weeks; foreign body; epiglottitis; bacterial tracheitis; immune deficiency

Baseline demographics (N = 68):

- proportion males: 53%
- mean (SD) age in months: 26.3 (1.5)
- mean (SD) croup score: treatment 1: 1.81 (0.59); treatment 2: 2.03 (0.47)

Interventions	Treatment 1 (N = 36): single dose 0.60 mg/kg intramuscular dexamethasone	
	Treatment 2 (N = 32): single dose 0.60 mg/kg oral dexamethasone	
Outcomes	Return visits or (re)admissions or both	
Notes	Not indexed in the databases searched (located via a search of Google.ca)	
	Funding source: Zahedan University of Medical Sciences, Zahedan, Iran	

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	Comment: described as "randomly divided"; unclear how the randomisation sequence was generated
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Comment: participants and personnel were not blinded. Treatments were clearly distinguishable.
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: outcome assessor described as blinded. Unclear how they were blinded and if the blinding could have been broken
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 15% loss to follow-up; unclear from which group children were lost. No intention-to-treat analysis
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk

Sparrow 2006

Study characteristics	
Methods	Randomised, double-blind controlled trial
Participants	Study period: not reported



Sparrow 2006 (Continued)

Setting: emergency department of Princess Margaret Hospital for Children, Perth, Australia

Inclusion criteria: children older than 3 months who presented to the emergency department with mild to moderate croup defined by clinical symptoms (acute onset of inspiratory stridor and a hoarse voice accompanied by a barking cough) and a modified Taussig croup score of < 5

Exclusion criteria: children whose families did not have a telephone; limited English language proficiency; received steroids

Baseline demographics (N = 133):

- proportion males: treatment: 74%; comparator: 63%
- mean (SD) age in months: treatment: 45 (31.6); comparator: 37 (28.8)
- mean (SD) modified Taussig croup score: treatment: 2.0 (1.2); comparator: 2.0 (1.3)

Notes	Funding source: not reported	
Outcomes	Unscheduled re-presentations to medical care; time spent in the emergency department; use of epinephrine	
	Comparator (N = 68): single 0.15 mg/kg dose of oral dexamethasone	
Interventions	Treatment (N = 65): single 1 mg/kg dose of oral prednisolone	

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Quote: "computer generated orders randomised into blocks of 10"
Allocation concealment (selection bias)	Unclear risk	Quote: "The PMH pharmacy ensured that the two steroid preparations could not be differentiated, the code being held by the pharmacy. Bottles were simply labelled solution A or solution B" Comment: bottles were not sequentially numbered, but instead labelled A or B
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "Bottles were simply labelled solution A or solution B, and following randomization given by a nurse who took no further part in the child's care."; "the code was not broken until all data were collected and all follow up completed"; "similar taste and appearance"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: no missing outcome data
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Unclear risk	Comment: many children were not approached because the emergency department was busy in the winter; this could potentially have biased participant selection
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk



Super 1989

Study characteristics			
Methods	Randomised, double-blind controlled trial		
Participants	Study period: October	r 1983 to April 1985	
	Setting: Cleveland Me	tropolitan General Hospital or Rainbow Babies and Children's Hospital, OH, USA	
	Inclusion criteria: children admitted to hospital with a diagnosis of croup (barking cough, inspiratory stridor, hoarseness) with a viral prodrome consisting of rhinorrhoea, cough, or fever and a modified Westley croup score ≥ 3 after 30 minutes of mist therapy		
		nical picture consistent with acute epiglottitis, spasmodic croup, or pneumonia; ess except asthma; history of tracheal intubation or laryngoscopy	
	Baseline demographi	cs (N = 29):	
	• proportion male: t	reatment: 68%; control: 62%	
		nonths: treatment: 15.5 (5); control: 15.8 (12)	
	 median (range) mo 	odified Westley croup score: treatment: 4.5 (3 to 7); control: 5.0 (3 to 6)	
Interventions	All children received mist therapy for at least 30 minutes.		
	Treatment (N = 16): single 0.6 mg/kg dose of parenteral dexamethasone		
	Control (N = 13): single dose of parenteral placebo (saline)		
Outcomes	Change in modified Westley croup score from baseline to 12 and 24 hours; length of stay in hospital; 2-unit improvement in croup score at 12 and 24 hours; use of supplemental glucocorticoids and use of mist tent		
Notes	Funding source: not reported		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "Patients were randomly assigned, by a table of random numbers"	
Allocation concealment (selection bias)	Low risk	Quote: "either parenterally administered dexamethasone or saline solution of the same color, volume and consistency as the dexamethasone. Randomization and drug preparation were done in the pharmacy" "The drug code was broken only after the last patient completed the study"	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "Neither the patient nor the investigators knew whether the patient received dexamethasone or placebo. The drug code was broken only after the last patient completed the study."	
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Quote: "Neither the patient nor the investigators knew whether the patient received dexamethasone or placebo. The drug code was broken only after the last patient completed the study." "All decisions regarding the data analysis were made before the drug code was broken. Whenever possible these decisions were made in favour of the null hypothesis"	



Super 1989 (Continued)		
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: 17% (N = 5) lost to follow-up (19% in treatment group due to early discharge and protocol deviation, 15% in placebo group due to early discharge or missed observations)
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Tibballs 1992

Study characteristics			
Methods	Randomised, double-b	lind controlled trial	
Participants	Study period: not reported		
	Setting: Royal Children	n's Hospital, Melbourne, Australia	
	tion for upper airway o	pitalised children aged 6 months or older who required endotracheal intubabstruction caused by croup (defined as coryzal symptoms, fever, barking cough, n, inspiratory stridor, or cyanosis developing over several days)	
		ldren younger than 6 months old; congenital airway anomalies; previous intuba- o (sudden onset without preceding fever or symptoms of upper respiratory tract	
	Baseline demographi	cs (N = 70, 3 excluded):	
		treatment: 63%; control: 66% n months: treatment: 19 (6 to 99); control: 19 (6 to 83) easured	
Interventions	All children received endotracheal intubation under inhalational anaesthesia with halothane, first with an oral endotracheal tube, in order to secure the airway rapidly and assess the diameter, and second substituted with a nasal tube. Humidification was provided with heat and moisture exchangers, with oxygen added as required. The tube was aspirated routinely every 1 to 2 hours to remove secretions.		
	Treatment (N = 38): 1 n hours until 24 hours af	ng/kg nasogastric prednisolone within 24 hours of intubation and then every 12 ter extubation	
	Control (N = 32, 3 excluuntil 24 hours after ext	ided): 1 mg/kg of placebo within 24 hours of intubation and then every 12 hours ubation	
Outcomes	Use of epinephrine		
Notes	Funding source: not reported		
Risk of bias			
Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Low risk	Quote: "order determined by a table of random numbers"	



Tibballs 1992 (Continued)		
Allocation concealment (selection bias)	Low risk	Quote: "Unidentified placebo and prednisolone were supplied by the pharmacy in an order determined by a table of random numbers"
Blinding of participants and personnel (perfor- mance bias) All outcomes	Low risk	Quote: "double blind" "Unidentified placebo and prednisolone were supplied by the pharmacy"
Blinding of outcome assessment (detection bias) All outcomes	Low risk	Comment: no description of a third-party outcome assessor. Carried over judgement from blinding of participants and personnel
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 4% (N = 3) lost because of exclusion due to bacterial infection or protocol deviations, all in the placebo group
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	Low risk	Comment: no other sources of bias identified
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk

Vad Pedersen 1998

Study characteristics	
Methods	Randomised controlled trial
Participants	Study period: October 1989 to September 1990
	Setting: paediatric department of Esbjerg Centralsygehus (Central Hospital), Denmark
	Inclusion criteria: children hospitalised with croup based on a modified Westley croup score ≥ 3 (including chest wall retractions, barking cough, respiratory frequency, and stridor)
	Exclusion criteria: required immediate intensive care; clinical suspicion of epiglottitis; cyanosis; croup recurrence; had received local or systemic steroid treatment; being treated with carbamazepine, phenobarbital, or rifampicin
	Baseline demographics (N = 59, 2 excluded):
	 proportion males: 63% mean age in months: treatment 1: 23.8; treatment 2: 24.1 mean (SD) croup score: treatment 1: 3.67 (1.02); treatment 2: 4.17 (0.99)
Interventions	Treatment 1 (N = 27): 2, 1000 μg doses of inhaled budesonide at 30-minute intervals
	Treatment 2 (N = 29): single 0.6 mg/kg (0.15 mL/kg) dose of intramuscular dexamethasone
Outcomes	Change in croup score from baseline to 6 and 12 hours; return visits to the hospital; use of supplemental glucocorticoids
Notes	Written in Danish
	Funding source: not reported



Vad Pedersen 1998 (Continued)

Risk of bias

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	Comment: block randomisation with varying block size
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement
Blinding of participants and personnel (perfor- mance bias) All outcomes	High risk	Comment: unblinded. Subjective outcomes
Blinding of outcome assessment (detection bias) All outcomes	High risk	Comment: unblinded. Subjective outcomes
Incomplete outcome data (attrition bias) All outcomes	Low risk	Comment: 3% (N = 2) who were randomised were excluded due to protocol deviations
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.
Other bias	High risk	Comment: baseline imbalance in croup score
Overall risk of bias All outcomes	High risk	Comment: at least 1 domain judged as high risk

Von Mühlendahl 1982

Study characterist	ics
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Methods	Randomised, double-blind controlled trial	
Participants	Study period: January 1979 to April 1980	
	Setting: 3 paediatric clinics in West Berlin, Germany	

Inclusion criteria: children admitted to hospital with a diagnosis of pseudo-croup to 1 of 3 paediatric clinics in West Berlin

Exclusion criteria: children who were already somnolent or cyanotic at admission (stage III or IV or pseudo-croup)

Baseline demographics (N = 406; 349 included in the evaluation):

- **proportion males:** not reported
- age distribution:
 - $\circ~$ treatment: 15 were < 1 year; 50 were 1 to 1 11/12 years; 96 were 2 to 5 11/12 years; 15 were 6 to 10 11/12 years
 - o control: 11 were < 1 year; 44 were 1 to 1 11/12 years; 107 were 2 to 5 11/12 years; 11 were 6 to 10 11/12 years
- croup score:
 - o treatment: 77 had a score of 1 to 3; 99 had a score ≥ 4



Von Mü	hlend	lah	l 1982	(Continued)
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0 (ontrol: 67	had a	score of	1 to 3	: 106 had	a score ≥ 4
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Interventions	Treatment (N = 176): single dose 6 mg oral dexamethasone
	Control (N = 173): single dose 6 mg oral placebo
Outcomes	Change in croup score from baseline to 6 and 12 hours
Notes	Funding source: not reported

Risk of bias

Bias	Authors' judgement	Support for judgement	
Random sequence generation (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Allocation concealment (selection bias)	Unclear risk	Comment: insufficient information provided to permit a judgement	
Blinding of participants and personnel (perfor- mance bias) All outcomes	Unclear risk	Comment: described as double-blind. Unclear who was blinded. Subjective outcomes	
Blinding of outcome assessment (detection bias) All outcomes	Unclear risk	Comment: described as double-blind. Unclear who was blinded. Subjective outcomes	
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	Comment: of 406 children, 57 (14%) failed to complete the study. 24 (7%) were eliminated due to protocol violation; 19 (5%) received further doses of steroids; and 4 (1%) developed measles. Unclear what group the lost children were in. Did not use intention-to-treat analysis	
Selective reporting (reporting bias)	Unclear risk	Comment: no protocol identified. All prespecified outcomes in the methods appeared in the results.	
Other bias	Low risk	Comment: no other sources of bias identified	
Overall risk of bias All outcomes	Unclear risk	Comment: at least 1 domain judged as unclear risk, and no domains judged as high risk	

CI: confidence interval ED: emergency department IU: international units

LOCF: last observation carried forward

NHS: National Health Service SD: standard deviation SE: standard error

Characteristics of excluded studies [ordered by study ID]

Study	Reason for exclusion	
Anene 1996	Randomised controlled trial; children were not diagnosed with croup	



Study	Reason for exclusion
Bollobas 1965	Not a randomised controlled trial; children were not diagnosed with croup
Cichy 1983	Not a randomised controlled trial
Connolly 1969	Randomised controlled trial; children were not diagnosed with croup
Couser 1992	Randomised controlled trial; children were not diagnosed with croup
Eghbali 2016	Randomised controlled trial; the intervention was L-epinephrine
Faghihinia 2007	Randomised controlled trial; no usable results were presented (unclear how many children were in each group)
Faraji-Goodarzi 2018	Randomised controlled trial; no relevant outcomes
Flisberg 1973	Not a randomised controlled trial
Freezer 1990	Not a randomised controlled trial; study was a retrospective chart review
Gill 2017	Not a randomised controlled trial
Goddard 1967	Randomised controlled trial; children were not diagnosed with croup
Gursanscky 2019	Not a randomised controlled trial; review article
Haque 1981	Not a randomised controlled trial; children in the control group received no treatment
Havaldar 1997	Not a randomised controlled trial; children were not diagnosed with croup
Kelley 1992	Not a randomised controlled trial; study was a retrospective chart review
Kotaniemi-Syrjanen 2018	Abstract only
Kunkel 1996	Not a randomised controlled trial; intervention was epinephrine
Ledwith 1995	Not a randomised controlled trial; intervention was epinephrine
Lee 2019	Wrong intervention; intervention was epinephrine
Martensson 1960	Not a randomised controlled trial
McDonogh 1994	Not a randomised controlled trial; study was a retrospective chart review
Meskina 2019	Wrong intervention; intervention was homeopathic drug
Mohammadzadeh 2014	Randomised controlled trial; intervention was epinephrine
NCT01748162	Randomised controlled trial; did not report any relevant outcomes
Novik 1960	Not a randomised controlled trial
Osváth 1994	Not a randomised controlled trial
Prendergast 1994	Not a randomised controlled trial; intervention was epinephrine



Study	Reason for exclusion
Rizos 1998	Not a randomised controlled trial
Roked 2015	Not a randomised controlled trial; study was a retrospective chart review
Ross 1969	Not a randomised controlled trial; study was a retrospective chart review
Serra 1997	Not a randomised controlled trial
Sumboonnanonda 1997	Randomised controlled trial; children in the control group received no treatment
Sussman 1964	Randomised controlled trial; children were not diagnosed with croup
Tal 1983	Randomised controlled trial; children were not diagnosed with croup
Tellez 1991	Randomised controlled trial; children were not diagnosed with croup
Tyler 2022	Not a randomised controlled trial; nested case-control study
Wilhelmi 1976	Not a randomised controlled trial

Characteristics of studies awaiting classification [ordered by study ID]

Chen 2018

Methods	Randomised controlled trial		
Participants	Study period: November 2016 to April 2017		
	Participants: children with acute laryngitis		
Interventions	Treatment (N = 40): 1.0 mg budesonide inhalation, and single injection of 0.3 to 0.5 mg/kg dexamethasone		
	Comparator (N = 38): single injection of 0.3 to 0.5 mg/kg dexamethasone		
Outcomes	Change in croup score from baseline to 12 and 24 hours		
Notes	Full text could not be found via library.		

Characteristics of ongoing studies [ordered by study ID]

IRCT20190914044765N1

Study name	Comparison the effect of oral and intravenous dexamethasone effect on the mild and moderate croup treatment in children
Methods	Randomised controlled trial (parallel)
Participants	Inclusion criteria: age 6 months until 72 months; temperature < 38 °C; no lung disease; no asthma; no recurrent use of bronchodilator in last week; not Westley croup score 2 until 7; no steroid use in last week



IRCT20190914044765N1 (Continued	Exclusion criteria: lung disease; severe croup; recurrent croup; allergy to steroid; prohibition to steroid use; steroid in last week; epiglottitis; tracheitis; temperature > 38 °C
Interventions	Treatment 1: oral dexamethasone
	Treatment 2: intravenous dexamethasone
Outcomes	Westley croup score; respiratory rate; heart rate; oxygen saturation; follow-up 3-day relapse
Starting date	27 August 2019
Contact information	Dr Ehsan Khoshnejad Afkham Kashan University of Medical Sciences, No. 65, Rajaii Ave, Somayeh Ave, Qom, Ghoum, 3715815548, Iran email: dr.ehkhaf@yahoo.com
Notes	Trial registration at the Iranian Registry of Clinical Trials (IRCT) (IRCT20190914044765N1)
	Funding source: no external funding specified

DATA AND ANALYSES

Comparison 1. Any glucocorticoid compared to placebo

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1.1 Croup score (change baseline - 2 hours) by score	7	426	Std. Mean Difference (IV, Random, 95% CI)	-0.65 [-1.13, -0.18]
1.1.1 Westley score	5	264	Std. Mean Difference (IV, Random, 95% CI)	-0.72 [-1.44, 0.01]
1.1.2 Non-Westley score	2	162	Std. Mean Difference (IV, Random, 95% CI)	-0.51 [-0.93, -0.10]
1.2 Croup score (change baseline - 6 hours) by score	11	959	Std. Mean Difference (IV, Random, 95% CI)	-0.76 [-1.12, -0.40]
1.2.1 Westley score	5	336	Std. Mean Difference (IV, Random, 95% CI)	-0.79 [-1.02, -0.56]
1.2.2 Non-Westley score	6	623	Std. Mean Difference (IV, Random, 95% CI)	-0.81 [-1.43, -0.18]
1.3 Croup score (change baseline - 12 hours) by score	8	571	Std. Mean Difference (IV, Random, 95% CI)	-1.03 [-1.53, -0.53]
1.3.1 Westley score	2	113	Std. Mean Difference (IV, Random, 95% CI)	-1.54 [-2.56, -0.53]
1.3.2 Non-Westley score	6	458	Std. Mean Difference (IV, Random, 95% CI)	-0.87 [-1.45, -0.30]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1.4 Croup score (change baseline - 24 hours) by score	8	351	Std. Mean Difference (IV, Random, 95% CI)	-0.86 [-1.40, -0.31]
1.4.1 Westley score	4	169	Std. Mean Difference (IV, Random, 95% CI)	-1.05 [-1.72, -0.37]
1.4.2 Non-Westley score	4	182	Std. Mean Difference (IV, Random, 95% CI)	-0.70 [-1.56, 0.16]
1.5 Croup score (change baseline - 2 hours) by inpa- tient/outpatient	7	426	Std. Mean Difference (IV, Random, 95% CI)	-0.65 [-1.13, -0.18]
1.5.1 Inpatient	5	301	Std. Mean Difference (IV, Random, 95% CI)	-0.80 [-1.44, -0.16]
1.5.2 Outpatient	2	125	Std. Mean Difference (IV, Random, 95% CI)	-0.32 [-0.93, 0.29]
1.6 Croup score (change baseline - 6 hours) by inpa- tient/outpatient	11	959	Std. Mean Difference (IV, Random, 95% CI)	-0.76 [-1.12, -0.40]
1.6.1 Inpatient	8	723	Std. Mean Difference (IV, Random, 95% CI)	-0.72 [-1.22, -0.23]
1.6.2 Outpatient	3	236	Std. Mean Difference (IV, Random, 95% CI)	-0.84 [-1.11, -0.56]
1.7 Croup score (change baseline - 24 hours) by inpa- tient/outpatient	8	351	Std. Mean Difference (IV, Random, 95% CI)	-0.86 [-1.40, -0.31]
1.7.1 Inpatient	7	291	Std. Mean Difference (IV, Random, 95% CI)	-0.82 [-1.46, -0.19]
1.7.2 Outpatient	1	60	Std. Mean Difference (IV, Random, 95% CI)	-1.09 [-1.71, -0.48]
1.8 Croup score (change baseline - 2 hours) by gluco- corticoid	7	426	Std. Mean Difference (IV, Random, 95% CI)	-0.66 [-1.10, -0.22]
1.8.1 Budesonide	4	246	Std. Mean Difference (IV, Random, 95% CI)	-1.01 [-1.71, -0.30]
1.8.2 Dexamethasone	3	163	Std. Mean Difference (IV, Random, 95% CI)	-0.49 [-1.00, 0.03]
1.8.3 Fluticasone	1	17	Std. Mean Difference (IV, Random, 95% CI)	0.45 [-0.52, 1.42]
1.9 Croup score (change baseline - 6 hours) by gluco- corticoid	11	959	Std. Mean Difference (IV, Random, 95% CI)	-0.74 [-1.07, -0.41]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1.9.1 Budesonide	5	333	Std. Mean Difference (IV, Random, 95% CI)	-0.81 [-1.04, -0.58]
1.9.2 Dexamethasone	6	567	Std. Mean Difference (IV, Random, 95% CI)	-0.62 [-1.17, -0.08]
1.9.3 Fluticasone	1	17	Std. Mean Difference (IV, Random, 95% CI)	0.06 [-0.89, 1.02]
1.9.4 Prednisolone	1	42	Std. Mean Difference (IV, Random, 95% CI)	-1.87 [-2.62, -1.13]
1.10 Croup score (change baseline - 12 hours) by gluco- corticoid	8	571	Std. Mean Difference (IV, Random, 95% CI)	-1.04 [-1.51, -0.56]
1.10.1 Budesonide	3	209	Std. Mean Difference (IV, Random, 95% CI)	-0.97 [-1.26, -0.68]
1.10.2 Dexamethasone	5	323	Std. Mean Difference (IV, Random, 95% CI)	-0.85 [-1.55, -0.15]
1.10.3 Prednisolone	1	39	Std. Mean Difference (IV, Random, 95% CI)	-2.40 [-3.26, -1.55]
1.11 Croup score (change baseline - 24 hours) by gluco- corticoid	8	351	Std. Mean Difference (IV, Random, 95% CI)	-0.89 [-1.41, -0.37]
1.11.1 Budesonide	2	89	Std. Mean Difference (IV, Random, 95% CI)	-1.40 [-1.88, -0.93]
1.11.2 Dexamethasone	6	245	Std. Mean Difference (IV, Random, 95% CI)	-0.89 [-1.55, -0.22]
1.11.3 Fluticasone	1	17	Std. Mean Difference (IV, Random, 95% CI)	0.21 [-0.75, 1.17]
1.12 Return visits or (re)ad- missions or both by inpa- tient/outpatient	10	1679	Risk Ratio (M-H, Random, 95% CI)	0.52 [0.36, 0.75]
1.12.1 Inpatient	3	323	Risk Ratio (M-H, Random, 95% CI)	0.39 [0.12, 1.30]
1.12.2 Outpatient	7	1356	Risk Ratio (M-H, Random, 95% CI)	0.53 [0.35, 0.80]
1.13 Return visits or (re)admissions or both by glucocorticoid	10	1679	Risk Ratio (M-H, Random, 95% CI)	0.51 [0.36, 0.72]
1.13.1 Budesonide	4	225	Risk Ratio (M-H, Random, 95% CI)	0.42 [0.19, 0.90]
1.13.2 Dexamethasone	8	1454	Risk Ratio (M-H, Random, 95% CI)	0.53 [0.34, 0.81]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1.14 Return visits or (re)ad- missions or both by croup severity	10	1679	Risk Ratio (M-H, Random, 95% CI)	0.52 [0.36, 0.76]
1.14.1 Mild croup	3	1068	Risk Ratio (M-H, Random, 95% CI)	0.54 [0.30, 0.95]
1.14.2 Moderate croup	7	611	Risk Ratio (M-H, Random, 95% CI)	0.48 [0.26, 0.86]
1.15 Length of stay by inpa- tient	8	476	Mean Difference (IV, Random, 95% CI)	-14.90 [-23.58, -6.22]
1.15.1 Inpatient	8	476	Mean Difference (IV, Random, 95% CI)	-14.90 [-23.58, -6.22]
1.16 Length of stay by gluco- corticoid	8	476	Mean Difference (IV, Random, 95% CI)	-14.55 [-22.70, -6.41]
1.16.1 Budesonide	2	131	Mean Difference (IV, Random, 95% CI)	-15.29 [-26.89, -3.69]
1.16.2 Dexamethasone	6	328	Mean Difference (IV, Random, 95% CI)	-18.25 [-27.87, -8.62]
1.16.3 Fluticasone	1	17	Mean Difference (IV, Random, 95% CI)	4.80 [-12.34, 21.94]
1.17 Improvement (at 2 hours) by inpatient	1	82	Risk Ratio (M-H, Random, 95% CI)	1.81 [0.96, 3.40]
1.17.1 Inpatient	1	82	Risk Ratio (M-H, Random, 95% CI)	1.81 [0.96, 3.40]
1.18 Improvement (at 6 hours) by inpatient/outpatient	6	332	Risk Ratio (M-H, Random, 95% CI)	1.45 [1.12, 1.88]
1.18.1 Inpatient	4	224	Risk Ratio (M-H, Random, 95% CI)	1.35 [0.96, 1.90]
1.18.2 Outpatient	2	108	Risk Ratio (M-H, Random, 95% CI)	1.78 [1.16, 2.74]
1.19 Improvement (at 12 hours) by inpatient	6	340	Risk Ratio (M-H, Random, 95% CI)	1.33 [1.09, 1.62]
1.19.1 Inpatient	6	340	Risk Ratio (M-H, Random, 95% CI)	1.33 [1.09, 1.62]
1.20 Improvement (at 24 hours) by inpatient/outpatient	5	251	Risk Ratio (M-H, Random, 95% CI)	1.28 [1.01, 1.61]
1.20.1 Inpatient	4	213	Risk Ratio (M-H, Random, 95% CI)	1.18 [0.98, 1.43]
1.20.2 Outpatient	1	38	Risk Ratio (M-H, Random, 95% CI)	2.00 [1.14, 3.51]
1.21 Improvement (at 6 hours) by glucocorticoid	6	332	Risk Ratio (M-H, Random, 95% CI)	1.45 [1.12, 1.88]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
1.21.1 Budesonide	2	135	Risk Ratio (M-H, Random, 95% CI)	1.66 [1.19, 2.32]
1.21.2 Dexamethasone	2	105	Risk Ratio (M-H, Random, 95% CI)	1.43 [0.76, 2.72]
1.21.3 Prednisolone	2	92	Risk Ratio (M-H, Random, 95% CI)	1.34 [0.69, 2.62]
1.22 Improvement (at 12 hours) by glucocorticoid	6	340	Risk Ratio (M-H, Random, 95% CI)	1.33 [1.09, 1.62]
1.22.1 Budesonide	1	82	Risk Ratio (M-H, Random, 95% CI)	1.41 [1.08, 1.84]
1.22.2 Dexamethasone	3	166	Risk Ratio (M-H, Random, 95% CI)	1.52 [1.06, 2.18]
1.22.3 Prednisolone	2	92	Risk Ratio (M-H, Random, 95% CI)	1.15 [0.85, 1.55]
1.23 Improvement (at 24 hours) by glucocorticoid	5	251	Risk Ratio (M-H, Random, 95% CI)	1.28 [1.01, 1.61]
1.23.1 Dexamethasone	4	201	Risk Ratio (M-H, Random, 95% CI)	1.39 [1.05, 1.84]
1.23.2 Prednisolone	1	50	Risk Ratio (M-H, Random, 95% CI)	1.04 [0.91, 1.20]
1.24 Additional treatments: antibiotics	3	202	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.04, 0.04]
1.25 Additional treatments: epinephrine	9	709	Risk Difference (M-H, Random, 95% CI)	-0.03 [-0.08, 0.01]
1.26 Additional treatments: intubation/tracheostomy	11	1090	Risk Difference (M-H, Random, 95% CI)	-0.00 [-0.01, 0.01]
1.27 Additional treatments: mist tent	2	84	Risk Difference (M-H, Random, 95% CI)	-0.20 [-0.87, 0.47]
1.28 Additional treatments: supplemental glucocorti- coids	6	305	Risk Ratio (M-H, Random, 95% CI)	0.61 [0.36, 1.03]



Analysis 1.1. Comparison 1: Any glucocorticoid compared to placebo, Outcome 1: Croup score (change baseline - 2 hours) by score

	Glu	cocorticoi	d	Placebo				Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.1.1 Westley score										
Dobrovoljac 2012	-1.7	1.14	35	-0.9	1.4	35	15.6%	-0.62 [-1.10 , -0.14]		
Godden 1997	-2.58	1.14	46	-1.17	1.4	40	15.8%	-1.10 [-1.56 , -0.65]		
Husby 1993	-3.86	1.14	20	-0.88	1.4	16	11.4%	-2.31 [-3.18 , -1.44]		
Johnson 1996	-1	1.28	28	-1	1.28	27	15.1%	0.00 [-0.53, 0.53]		
Roorda 1998	-0.9	1	9	-1.5	1.52	8	10.4%	0.45 [-0.52 , 1.42]		
Subtotal (95% CI)			138			126	68.3%	-0.72 [-1.44, 0.01]		
Heterogeneity: Tau ² = 0	.56; Chi ² = 28	3.49, df = 4	4 (P < 0.00	001); I ² = 8	36%				•	
Test for overall effect: Z	Z = 1.94 (P =	0.05)								
1.1.2 Non-Westley scor	re									
Geelhoed 1995c	-1.73	1.14	50	-0.81	1.4	30	15.7%	-0.73 [-1.20, -0.27]		
Roberts 1999	-1.41	1.14	42	-1.01	1.4	40	16.0%	-0.31 [-0.75, 0.12]		
Subtotal (95% CI)			92			70	31.7%	-0.51 [-0.93 , -0.10]		
Heterogeneity: Tau ² = 0	.04; Chi ² = 1.	67, df = 1	(P = 0.20)	$I^2 = 40\%$					~	
Test for overall effect: Z	Z = 2.44 (P =	0.01)								
Total (95% CI)			230			196	100.0%	-0.65 [-1.13 , -0.18]		
Heterogeneity: Tau ² = 0	.31; Chi ² = 3	1.01, df = 0	6 (P < 0.00	01); I ² = 81	1%				•	
Test for overall effect: 2	Z = 2.70 (P = 0.00)	0.007)	•	**					-2 -1 0 1 2	
Test for subgroup differ	•	,	1 (P = 0.6)	3) I ² = 0%				Favou	rs glucocorticoid Favours plac	

Analysis 1.2. Comparison 1: Any glucocorticoid compared to placebo, Outcome 2: Croup score (change baseline - 6 hours) by score

	Glu	cocortico	id	Placebo				Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.2.1 Westley score										
Godden 1997	-2.96	1.37	44	-1.74	1.49	39	10.0%	-0.85 [-1.30 , -0.40]		
Johnson 1996	-2	1.28	17	-1	0.74	21	8.4%	-0.96 [-1.64, -0.28]		
Johnson 1998	-2.44	1.44	95	-1.3	1.4	49	10.6%	-0.79 [-1.15 , -0.44]	•	
Klassen 1994	-3	1.96	27	-1	2.56	27	9.2%	-0.86 [-1.42, -0.30]	-	
Roorda 1998	-1.1	1	9	-1.2	1.89	8	6.5%	0.06 [-0.89, 1.02]		
Subtotal (95% CI)			192			144	44.8%	-0.79 [-1.02, -0.56]	&	
Heterogeneity: Tau ² = 0.00; C	Chi ² = 3.47, df	= 4 (P = 0)	.48); I ² = 0	%					*	
Test for overall effect: $Z = 6.7$	78 (P < 0.0000	1)								
1.2.2 Non-Westley score										
Geelhoed 1995c	-2.51	1.37	50	-1.05	1.49	30	9.8%	-1.02 [-1.50 , -0.54]	-	
Kuusela 1988	-1.65	0.7	16	-0.65	0.66	21	7.9%	-1.44 [-2.18, -0.71]		
Martinez Fernandez 1993	-0.4	1.45	16	-1	1.37	17	8.3%	0.42 [-0.28 , 1.11]	 -	
Massicotte 1973	-5.94	1.37	25	-3.23	1.49	17	7.9%	-1.87 [-2.62 , -1.13]		
Roberts 1999	-2.26	1.37	42	-0.81	1.49	40	9.9%	-1.00 [-1.47, -0.54]	-	
Von Mühlendahl 1982	-3.14	1.79	176	-3	1.78	173	11.4%	-0.08 [-0.29, 0.13]		
Subtotal (95% CI)			325			298	55.2%	-0.81 [-1.43 , -0.18]	▲	
Heterogeneity: Tau ² = 0.52; C	Chi ² = 49.96, d	f = 5 (P <	0.00001); 1	2 = 90%					•	
Test for overall effect: $Z = 2$.	54 (P = 0.01)									
Total (95% CI)			517			442	100.0%	-0.76 [-1.12 , -0.40]	•	
Heterogeneity: Tau ² = 0.29; C	Chi ² = 59.76, d	f = 10 (P <	< 0.00001);	$I^2 = 83\%$					*	
Test for overall effect: $Z = 4$.	12 (P < 0.0001	.)							-4 -2 0 2 4	
Test for subgroup differences	: Chi ² = 0.00,	df = 1 (P =	0.96), I ² =	0%				Favou	rs glucocorticoid Favours placel	



Analysis 1.3. Comparison 1: Any glucocorticoid compared to placebo, Outcome 3: Croup score (change baseline - 12 hours) by score

	Glu	cocorticoi	id	Placebo				Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.3.1 Westley score										
Godden 1997	-3.41	1.23	44	-1.99	1.31	41	13.7%	-1.11 [-1.57 , -0.65]		
Super 1989	-3.5	1.15	16	-1	1.09	12	9.8%	-2.16 [-3.12 , -1.19]		
Subtotal (95% CI)			60			53	23.5%	-1.54 [-2.56 , -0.53]		
Heterogeneity: Tau ² = 0.40; Ch	$i^2 = 3.70$, df	= 1 (P = 0)	.05); I ² = 7	3%					•	
Test for overall effect: $Z = 2.99$	P = 0.003									
1.3.2 Non-Westley score										
Geelhoed 1995c	-2.59	1.23	50	-1.1	1.31	30	13.5%	-1.17 [-1.66 , -0.68]		
Kuusela 1988	-1.9	0.63	16	-1.2	0.77	21	12.0%	-0.96 [-1.65 , -0.27]		
Martinez Fernandez 1993	-1.1	1.51	16	-1.1	1.55	17	12.0%	0.00 [-0.68, 0.68]		
Massicotte 1973	-10.71	1.23	24	-7.62	1.31	15	10.7%	-2.40 [-3.26 , -1.55]		
Roberts 1999	-2.41	1.23	42	-1.32	1.31	40	13.7%	-0.85 [-1.30 , -0.40]		
Von Mühlendahl 1982	-3.31	1.62	75	-3.05	1.82	112	14.7%	-0.15 [-0.44, 0.14]	4	
Subtotal (95% CI)			223			235	76.5%	-0.87 [-1.45 , -0.30]	•	
Heterogeneity: Tau ² = 0.43; Ch	i ² = 36.55, d	f = 5 (P < 1)	0.00001); 1	$1^2 = 86\%$					•	
Test for overall effect: $Z = 2.97$	7 (P = 0.003)									
Total (95% CI)			283			288	100.0%	-1.03 [-1.53 , -0.53]	•	
Heterogeneity: Tau ² = 0.43; Ch	i ² = 49.02, d	f = 7 (P < 0)	0.00001); 1	$1^2 = 86\%$					~	
Test for overall effect: $Z = 4.01$	(P < 0.0001)							-4 -2 0 2 4	
Test for subgroup differences:	Chi ² = 1.28, o	df = 1 (P =	0.26), I ² =	21.6%				Favou	rs glucocorticoid Favours placeb	

Analysis 1.4. Comparison 1: Any glucocorticoid compared to placebo, Outcome 4: Croup score (change baseline - 24 hours) by score

	Glu	cocortico	id	Placebo				Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.4.1 Westley score										
Cetinkaya 2004	-2.53	1.23	45	-1.05	1.62	15	13.3%	-1.09 [-1.71 , -0.48]		
Godden 1997	-4.14	1.23	35	-2.11	1.62	32	13.9%	-1.40 [-1.94 , -0.87]		
Roorda 1998	-2.1	1.15	9	-2.4	1.56	8	10.8%	0.21 [-0.75 , 1.17]		
Super 1989	-3.5	1.15	13	-1.5	1.09	12	10.9%	-1.72 [-2.67, -0.78]		
Subtotal (95% CI)			102			67	48.9%	-1.05 [-1.72 , -0.37]	•	
Heterogeneity: Tau ² = 0.32; C	$Chi^2 = 10.08, d$	f = 3 (P =	0.02); I ² =	70%					~	
Test for overall effect: $Z = 3.0$	05 (P = 0.002)									
1.4.2 Non-Westley score										
Kuusela 1988	-2.09	1.23	16	-1.49	1.62	21	13.0%	-0.40 [-1.06, 0.26]		
Leipzig 1979	-7.27	1.39	16	-2.56	2.5	14	10.8%	-2.31 [-3.26, -1.36]		
Martinez Fernandez 1993	-1.9	1.23	16	-2.4	1.32	17	12.8%	0.38 [-0.31, 1.07]		
Roberts 1999	-2.51	1.23	42	-1.52	1.62	40	14.5%	-0.68 [-1.13 , -0.24]		
Subtotal (95% CI)			90			92	51.1%	-0.70 [-1.56 , 0.16]		
Heterogeneity: Tau ² = 0.65; C	Chi ² = 20.63, d	f = 3 (P =	0.0001); I ²	= 85%						
Test for overall effect: $Z = 1.5$	59 (P = 0.11)									
Total (95% CI)			192			159	100.0%	-0.86 [-1.40 , -0.31]	•	
Heterogeneity: Tau ² = 0.48; C	Chi ² = 36.25, d	f = 7 (P <	0.00001);	$I^2 = 81\%$					•	
Test for overall effect: $Z = 3.0$	08 (P = 0.002)								-4 -2 0 2 4	
Test for subgroup differences	: Chi ² = 0.38,	df = 1 (P =	0.54), I ² =	= 0%				Favou	urs glucocorticoid Favours placeb	



Analysis 1.5. Comparison 1: Any glucocorticoid compared to placebo, Outcome 5: Croup score (change baseline - 2 hours) by inpatient/outpatient

	Glu	cocortico	id	Placebo				Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.5.1 Inpatient										
Geelhoed 1995c	-1.73	1.14	50	-0.81	1.4	30	15.7%	-0.73 [-1.20 , -0.27]		
Godden 1997	-2.58	1.14	46	-1.17	1.4	40	15.8%	-1.10 [-1.56 , -0.65]		
Husby 1993	-3.86	1.14	20	-0.88	1.4	16	11.4%	-2.31 [-3.18 , -1.44]		
Roberts 1999	-1.41	1.14	42	-1.01	1.4	40	16.0%	-0.31 [-0.75 , 0.12]		
Roorda 1998	-0.9	1	9	-1.5	1.52	8	10.4%	0.45 [-0.52 , 1.42]		
Subtotal (95% CI)			167			134	69.4%	-0.80 [-1.44, -0.16]		
Heterogeneity: Tau ² = ().42; Chi ² = 24	4.52, df =	4 (P < 0.00	001); I ² = 84	1%				•	
Test for overall effect: 2	Z = 2.47 (P =	0.01)								
1.5.2 Outpatient										
Dobrovoljac 2012	-1.7	1.14	35	-0.9	1.4	35	15.6%	-0.62 [-1.10 , -0.14]		
Johnson 1996	-1	1.28	28	-1	1.28	27	15.1%	0.00 [-0.53 , 0.53]		
Subtotal (95% CI)			63			62	30.6%	-0.32 [-0.93 , 0.29]		
Heterogeneity: Tau ² = ().13; Chi ² = 2.	89, df = 1	(P = 0.09)	; I ² = 65%				. , ,	\blacksquare	
Test for overall effect: 2	Z = 1.03 (P =	0.30)	,							
Total (95% CI)			230			196	100.0%	-0.65 [-1.13 , -0.18]		
Heterogeneity: Tau ² = (31 · Chi ² = 3	1 01 df =		01)· I² = 81	1%	150	103.070	1115, 0110]	—	
Test for overall effect: 7	*	,	0 (1 \ 0.00	,01), 1 01	. 70					
Test for overall effect: A	•		1 (D = 0.3	0) 12 - 12	40/			Eavo	-4 -2 0 2 4 urs glucocorticoid Favours place	
rest for snogroup differ	rences: Cill ² =	1.15, uI =	· 1 (P – 0.2	0), 1 13.4	+ 70			Favoi	urs glucocorticoid Favours placel	

Analysis 1.6. Comparison 1: Any glucocorticoid compared to placebo, Outcome 6: Croup score (change baseline - 6 hours) by inpatient/outpatient

	Glu	cocortico	id		Placebo			Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.6.1 Inpatient										
Geelhoed 1995c	-2.51	1.37	50	-1.05	1.49	30	9.8%	-1.02 [-1.50 , -0.54]	_ -	
Godden 1997	-2.96	1.37	44	-1.74	1.49	39	10.0%	-0.85 [-1.30 , -0.40]	_ 	
Kuusela 1988	-1.65	0.7	16	-0.65	0.66	21	7.9%	-1.44 [-2.18 , -0.71]		
Martinez Fernandez 1993	-0.4	1.45	16	-1	1.37	17	8.3%	0.42 [-0.28 , 1.11]	 	
Massicotte 1973	-5.94	1.37	25	-3.23	1.49	17	7.9%	-1.87 [-2.62 , -1.13]		
Roberts 1999	-2.26	1.37	42	-0.81	1.49	40	9.9%	-1.00 [-1.47 , -0.54]		
Roorda 1998	-1.1	1	9	-1.2	1.89	8	6.5%	0.06 [-0.89, 1.02]		
Von Mühlendahl 1982	-3.14	1.79	176	-3	1.78	173	11.4%	-0.08 [-0.29, 0.13]	4	
Subtotal (95% CI)			378			345	71.8%	-0.72 [-1.22 , -0.23]		
Heterogeneity: Tau ² = 0.41; C	hi ² = 54.08, d	f = 7 (P <	0.00001);	$I^2 = 87\%$					•	
Test for overall effect: $Z = 2.8$	88 (P = 0.004)									
1.6.2 Outpatient										
Johnson 1996	-2	1.28	17	-1	0.74	21	8.4%	-0.96 [-1.64 , -0.28]		
ohnson 1998	-2.44	1.44	95	-1.3	1.4	49	10.6%	-0.79 [-1.15 , -0.44]		
Klassen 1994	-3	1.96	27	-1	2.56	27	9.2%	-0.86 [-1.42, -0.30]		
Subtotal (95% CI)			139			97	28.2%	-0.84 [-1.11 , -0.56]	•	
Heterogeneity: Tau ² = 0.00; C	$2hi^2 = 0.20$, df	= 2 (P = 0)	.91); I ² = 0	1%					•	
Test for overall effect: $Z = 5.9$	98 (P < 0.0000	1)								
Total (95% CI)			517			442	100.0%	-0.76 [-1.12 , -0.40]	•	
Heterogeneity: Tau ² = 0.29; C	hi ² = 59.76, d	f = 10 (P <	< 0.00001);	; I ² = 83%					•	
Test for overall effect: $Z = 4.1$	12 (P < 0.0001	.)							-2 -1 0 1 2	
Test for subgroup differences:	Chi ² = 0.16,	df = 1 (P =	= 0.69), I ² =	= 0%				Favou	rs glucocorticoid Favours plac	



Analysis 1.7. Comparison 1: Any glucocorticoid compared to placebo, Outcome 7: Croup score (change baseline - 24 hours) by inpatient/outpatient

	Glu	cocortico	id	Placebo				Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.7.1 Inpatient										
Godden 1997	-4.14	1.23	35	-2.11	1.62	32	13.9%	-1.40 [-1.94 , -0.87]		
Kuusela 1988	-2.09	1.23	16	-1.49	1.62	21	13.0%	-0.40 [-1.06, 0.26]		
Leipzig 1979	-7.27	1.39	16	-2.56	2.5	14	10.8%	-2.31 [-3.26 , -1.36]		
Martinez Fernandez 1993	-1.9	1.23	16	-2.4	1.32	17	12.8%	0.38 [-0.31, 1.07]		
Roberts 1999	-2.51	1.23	42	-1.52	1.62	40	14.5%	-0.68 [-1.13, -0.24]		
Roorda 1998	-2.1	1.15	9	-2.4	1.56	8	10.8%	0.21 [-0.75 , 1.17]		
Super 1989	-3.5	1.15	13	-1.5	1.09	12	10.9%	-1.72 [-2.67, -0.78]		
Subtotal (95% CI)			147			144	86.7%	-0.82 [-1.46 , -0.19]		
Heterogeneity: $Tau^2 = 0.58$; Chi Test for overall effect: $Z = 2.55$		f = 6 (P <	0.00001); l	[2 = 83%						
1.7.2 Outpatient										
Cetinkaya 2004	-2.53	1.23	45	-1.05	1.62	15	13.3%	-1.09 [-1.71 , -0.48]		
Subtotal (95% CI) Heterogeneity: Not applicable Test for overall effect: Z = 3.47	(P = 0.0005)	45			15	13.3%	-1.09 [-1.71 , -0.48]	•	
Total (95% CI)		,	192			159	100.0%	-0.86 [-1.40 , -0.31]		
Heterogeneity: $Tau^2 = 0.48$; Chi Test for overall effect: $Z = 3.08$ Test for subgroup differences: C	(P = 0.002)	`	,,					Favour	-2 -1 0 1 2 s glucocorticoid Favours	

Analysis 1.8. Comparison 1: Any glucocorticoid compared to placebo, Outcome 8: Croup score (change baseline - 2 hours) by glucocorticoid

	Glu	cocortico	id	Placebo				Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.8.1 Budesonide										
Geelhoed 1995c	-1.51	1.14	27	-0.81	1.4	15	12.4%	-0.56 [-1.20, 0.09]		
Godden 1997	-2.58	1.14	46	-1.17	1.4	40	14.2%	-1.10 [-1.56 , -0.65]		
Husby 1993	-3.86	1.14	20	-0.88	1.4	16	10.2%	-2.31 [-3.18, -1.44]		
Roberts 1999	-1.41	1.14	42	-1.01	1.4	40	14.4%	-0.31 [-0.75, 0.12]	-	
ubtotal (95% CI)			135			111	51.2%	-1.01 [-1.71 , -0.30]		
Heterogeneity: Tau ² = 0	0.43; Chi ² = 18	3.71, df =	3(P = 0.00)	003); I ² = 84	%				•	
Test for overall effect: 2	Z = 2.78 (P =	0.005)								
.8.2 Dexamethasone										
Oobrovoljac 2012	-1.7	1.14	35	-0.9	1.4	35	14.0%	-0.62 [-1.10 , -0.14]		
Geelhoed 1995c	-1.99	1.14	23	-0.81	1.4	15	11.9%	-0.93 [-1.61 , -0.24]		
ohnson 1996	-1	1.28	28	-1	1.28	27	13.5%	0.00 [-0.53, 0.53]		
Subtotal (95% CI)			86			77	39.5%	-0.49 [-1.00, 0.03]		
Heterogeneity: Tau ² = 0	0.13; Chi ² = 5.	10, df = 2	(P = 0.08)	; I ² = 61%					~	
Test for overall effect: 2	Z = 1.86 (P =	0.06)								
1.8.3 Fluticasone										
Roorda 1998	-0.9	1	9	-1.5	1.52	8	9.3%	0.45 [-0.52 , 1.42]		
Subtotal (95% CI)			9			8	9.3%	0.45 [-0.52 , 1.42]		
Heterogeneity: Not app	licable									
Test for overall effect: 2	Z = 0.91 (P =	0.36)								
Total (95% CI)			230			196	100.0%	-0.66 [-1.10 , -0.22]		
Heterogeneity: Tau ² = 0	0.30; Chi ² = 3	1.59, df =	7 (P < 0.00	001); I ² = 78	1%				•	
Test for overall effect: 2	Z = 2.93 (P =	0.003)							-2 -1 0 1 2	
Test for subgroup differ	rences: Chi ² =	5.65, df =	2 (P = 0.0	6), I ² = 64.6	5%			Favou	rs glucocorticoid Favours place	



Analysis 1.9. Comparison 1: Any glucocorticoid compared to placebo, Outcome 9: Croup score (change baseline - 6 hours) by glucocorticoid

	Glu	cocortico	id		Placebo			Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.9.1 Budesonide										
Geelhoed 1995c	-2.25	1.37	27	-1.05	1.49	15	7.3%	-0.83 [-1.49, -0.17]		
Godden 1997	-2.96	1.37	44	-1.74	1.49	39	8.7%	-0.85 [-1.30 , -0.40]		
Johnson 1998	-2	1.39	48	-1.3	1.4	24	8.4%	-0.50 [-0.99, 0.00]		
Klassen 1994	-3	1.96	27	-1	2.56	27	8.0%	-0.86 [-1.42 , -0.30]		
Roberts 1999	-2.26	1.37	42	-0.81	1.49	40	8.6%	-1.00 [-1.47, -0.54]		
Subtotal (95% CI)			188			145	41.0%	-0.81 [-1.04, -0.58]	•	
Heterogeneity: Tau ² = 0.00; Ch	$i^2 = 2.27$, df	= 4 (P = 0)	.69); I ² = 0	1%					•	
Test for overall effect: $Z = 6.98$	(P < 0.0000	1)								
1.9.2 Dexamethasone										
Geelhoed 1995c	-1.99	1.37	23	-1.05	1.49	15	7.3%	-0.65 [-1.32, 0.02]		
Johnson 1996	-2	1.28	17	-1	0.74	21	7.2%	-0.96 [-1.64, -0.28]		
Johnson 1998	-2.9	1.37	47	-1.3	1.4	25	8.2%	-1.15 [-1.67, -0.62]		
Kuusela 1988	-1.65	0.7	16	-0.65	0.66	21	6.8%	-1.44 [-2.18, -0.71]		
Martinez Fernandez 1993	-0.4	1.45	16	-1	1.37	17	7.1%	0.42 [-0.28 , 1.11]		
Von Mühlendahl 1982	-3.14	1.79	176	-3	1.78	173	9.9%	-0.08 [-0.29, 0.13]	_	
Subtotal (95% CI)			295			272	46.6%	-0.62 [-1.17, -0.08]		
Heterogeneity: Tau ² = 0.38; Ch	i ² = 32.38, d	f = 5 (P <	0.00001);	$I^2 = 85\%$					•	
Test for overall effect: $Z = 2.23$	(P = 0.03)									
1.9.3 Fluticasone										
Roorda 1998	-1.1	1	9	-1.2	1.89	8	5.6%	0.06 [-0.89, 1.02]		
Subtotal (95% CI)			9			8	5.6%	0.06 [-0.89, 1.02]		
Heterogeneity: Not applicable										
Test for overall effect: $Z = 0.13$	(P = 0.90)									
1.9.4 Prednisolone										
Massicotte 1973	-5.94	1.37	25	-3.23	1.49	17	6.8%	-1.87 [-2.62 , -1.13]		
Subtotal (95% CI)			25			17	6.8%	-1.87 [-2.62 , -1.13]		
Heterogeneity: Not applicable										
Test for overall effect: $Z = 4.92$	(P < 0.0000)	1)								
Total (95% CI)			517			442	100.0%	-0.74 [-1.07 , -0.41]	•	
Heterogeneity: Tau ² = 0.27; Ch	i ² = 60.05, d	f = 12 (P <	0.00001);	$I^2 = 80\%$				-	•	
Test for overall effect: $Z = 4.41$	(P < 0.0001)							-2 -1 0 1 2	
Test for subgroup differences: (Chi ² = 11.46	df = 3 P	= 0,009). 1	2 = 73.8%				Favour	s glucocorticoid Favours pla	



Analysis 1.10. Comparison 1: Any glucocorticoid compared to placebo, Outcome 10: Croup score (change baseline - 12 hours) by glucocorticoid

	Glu	cocortico	id		Placebo			Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.10.1 Budesonide										
Geelhoed 1995c	-2.33	1.23	27	-1.1	1.31	15	11.0%	-0.96 [-1.63, -0.29]		
Godden 1997	-3.41	1.23	44	-1.99	1.31	41	12.4%	-1.11 [-1.57 , -0.65]		
Roberts 1999	-2.41	1.23	42	-1.32	1.31	40	12.5%	-0.85 [-1.30 , -0.40]		
Subtotal (95% CI)			113			96	35.9%	-0.97 [-1.26 , -0.68]	♦	
Heterogeneity: Tau ² = 0.00; Ch	$ni^2 = 0.62$, df	= 2 (P = 0)	.73); I ² = 0	%					*	
Test for overall effect: $Z = 6.58$	B (P < 0.0000	1)								
1.10.2 Dexamethasone										
Geelhoed 1995c	-2.82	1.23	23	-1.1	1.31	15	10.6%	-1.33 [-2.06 , -0.61]		
Kuusela 1988	-1.9	0.63	16	-1.2	0.77	21	10.8%	-0.96 [-1.65 , -0.27]		
Martinez Fernandez 1993	-1.1	1.51	16	-1.1	1.55	17	10.9%	0.00 [-0.68, 0.68]		
Super 1989	-3.5	1.15	16	-1	1.09	12	8.9%	-2.16 [-3.12 , -1.19]		
Von Mühlendahl 1982	-3.31	1.62	75	-3.05	1.82	112	13.3%	-0.15 [-0.44, 0.14]	4	
Subtotal (95% CI)			146			177	54.5%	-0.85 [-1.55 , -0.15]		
Heterogeneity: Tau ² = 0.51; Ch	ni ² = 25.60, d	f = 4 (P <	0.0001); I ²	= 84%					•	
Test for overall effect: $Z = 2.39$	θ (P = 0.02)									
1.10.3 Prednisolone										
Massicotte 1973	-10.71	1.23	24	-7.62	1.31	15	9.6%	-2.40 [-3.26 , -1.55]		
Subtotal (95% CI)			24			15	9.6%	-2.40 [-3.26 , -1.55]	•	
Heterogeneity: Not applicable									•	
Test for overall effect: $Z = 5.50$	P < 0.0000	1)								
Total (95% CI)			283			288	100.0%	-1.04 [-1.51 , -0.56]	•	
Heterogeneity: Tau ² = 0.41; Ch	ni ² = 49.04, d	f = 8 (P <	0.00001);	$[^2 = 84\%]$					•	
Test for overall effect: $Z = 4.30$) (P < 0.0001	.)							-4 -2 0 2 4	
Test for subgroup differences:	Chi ² = 10.08.	df = 2 (P)	= 0.006), 1	2 = 80.2%				Favou	rs glucocorticoid Favours pla	



Analysis 1.11. Comparison 1: Any glucocorticoid compared to placebo, Outcome 11: Croup score (change baseline - 24 hours) by glucocorticoid

	Glu	cocortico	id		Placebo			Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.11.1 Budesonide										
Cetinkaya 2004	-3.03	1.23	15	-1.05	1.62	7	9.5%	-1.40 [-2.41, -0.39]		
Godden 1997	-4.14	1.23	35	-2.11	1.62	32	12.8%	-1.40 [-1.94, -0.87]		
Subtotal (95% CI)			50			39	22.3%	-1.40 [-1.88 , -0.93]	•	
Heterogeneity: Tau ² = 0.00; Ch	$i^2 = 0.00$, df	= 1 (P = 1	.00); $I^2 = 0$	%					•	
Test for overall effect: $Z = 5.79$	(P < 0.0000	1)								
1.11.2 Dexamethasone										
Cetinkaya 2004	-2.27	1.23	30	-1	1.62	8	10.9%	-0.95 [-1.76 , -0.13]		
Kuusela 1988	-2.09	1.23	16	-1.49	1.62	21	12.0%	-0.40 [-1.06, 0.26]		
Leipzig 1979	-7.27	1.39	16	-2.56	2.5	14	9.9%	-2.31 [-3.26, -1.36]		
Martinez Fernandez 1993	-1.9	1.23	16	-2.4	1.32	17	11.7%	0.38 [-0.31, 1.07]		
Roberts 1999	-2.51	1.23	42	-1.52	1.62	40	13.4%	-0.68 [-1.13, -0.24]		
Super 1989	-3.5	1.15	13	-1.5	1.09	12	10.0%	-1.72 [-2.67, -0.78]		
Subtotal (95% CI)			133			112	67.8%	-0.89 [-1.55 , -0.22]		
Heterogeneity: Tau ² = 0.54; Ch	i ² = 26.10, d	f = 5 (P <	0.0001); I ²	= 81%					•	
Test for overall effect: $Z = 2.62$	P = 0.009									
1.11.3 Fluticasone										
Roorda 1998	-2.1	1.15	9	-2.4	1.56	8	9.9%	0.21 [-0.75 , 1.17]		
Subtotal (95% CI)			9			8	9.9%	0.21 [-0.75, 1.17]		
Heterogeneity: Not applicable										
Test for overall effect: $Z = 0.43$	8 (P = 0.67)									
Total (95% CI)			192			159	100.0%	-0.89 [-1.41 , -0.37]	•	
Heterogeneity: Tau ² = 0.47; Ch	i ² = 36.87, d	f = 8 (P <	0.0001); I ²	= 78%					•	
Test for overall effect: $Z = 3.36$	6 (P = 0.0008)							-4 -2 0 2	
Test for subgroup differences: 0	$Chi^2 = 9.02, \sigma$	df = 2 (P =	0.01), I ² =	77.8%				Favou	irs glucocorticoid Favours place	



Analysis 1.12. Comparison 1: Any glucocorticoid compared to placebo, Outcome 12: Return visits or (re)admissions or both by inpatient/outpatient

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI	
1.12.1 Inpatient								
Geelhoed 1995c	5	39	4	24	7.1%	0.77 [0.23, 2.59]		
Roberts 1999	1	34	7	32	3.0%	0.13 [0.02, 1.03]		
Skowron 1966a and b	0	97	2	97	1.5%	0.20 [0.01, 4.11]	•	
Subtotal (95% CI)		170		153	11.6%	0.39 [0.12, 1.30]		
Total events:	6		13					
Heterogeneity: Tau ² = 0.28	; Chi ² = 2.5	5, df = 2 (l	P = 0.28); I	2 = 22%				
Test for overall effect: Z =	1.53 (P = 0.	13)						
1.12.2 Outpatient								
Bjornson 2004	26	354	54	354	19.9%	0.48 [0.31, 0.75]	-	
Cruz 1995	1	19	4	19	2.9%	0.25 [0.03, 2.04]		
Geelhoed 1996a	0	48	8	48	1.7%	0.06 [0.00, 0.99]		
Johnson 1996	17	27	17	25	21.1%	0.93 [0.62 , 1.37]	-	
Johnson 1998	29	95	35	49	22.2%	0.43 [0.30, 0.61]		
Klassen 1994	1	27	7	27	3.1%	0.14 [0.02 , 1.08]		
Luria 2001	26	176	18	88	17.5%	0.72 [0.42 , 1.24]		
Subtotal (95% CI)		746		610	88.4%	0.53 [0.35, 0.80]	•	
Total events:	100		143				•	
Heterogeneity: Tau ² = 0.14	; Chi ² = 15.0	66, df = 6	(P = 0.02);	$I^2 = 62\%$				
Test for overall effect: $Z = \frac{1}{2}$	3.02 (P = 0.	003)						
Total (95% CI)		916		763	100.0%	0.52 [0.36, 0.75]	•	
Total events:	106		156				•	
Heterogeneity: Tau ² = 0.13	; Chi ² = 18.0	64, df = 9	(P = 0.03);	$I^2 = 52\%$			0.02 0.1 1 10 50	
Test for overall effect: $Z = \frac{1}{2}$	3.42 (P = 0.	0006)					urs glucocorticoid Favours placebo	
Test for subgroup difference	es: Chi² = 0	.23, df = 1	1 (P = 0.63)	$I^2 = 0\%$				



Analysis 1.13. Comparison 1: Any glucocorticoid compared to placebo, Outcome 13: Return visits or (re)admissions or both by glucocorticoid

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.13.1 Budesonide							
Geelhoed 1995c	3	21	2	12	3.6%	0.86 [0.17, 4.43]	
Johnson 1998	18	48	17	24	17.4%	0.53 [0.34, 0.83]	
Klassen 1994	1	27	7	27	2.5%	0.14 [0.02, 1.08]	—
Roberts 1999	1	34	7	32	2.5%	0.13 [0.02, 1.03]	-
Subtotal (95% CI)		130		95	26.1%	0.42 [0.19, 0.90]	
Total events:	23		33				
Heterogeneity: Tau ² = 0.2	21; Chi ² = 4.2	3, df = 3 (1)	$P = 0.24$); I^2	2 = 29%			
Test for overall effect: Z							
1.13.2 Dexamethasone							
Bjornson 2004	26	354	54	354	17.5%	0.48 [0.31, 0.75]	
Cruz 1995	1	19	4	19	2.4%	0.25 [0.03, 2.04]	
Geelhoed 1995c	2	18	2	12	3.1%	0.67 [0.11, 4.11]	
Geelhoed 1996a	0	48	8	48	1.4%	0.06 [0.00, 0.99]	—
Johnson 1996	17	27	17	25	18.7%	0.93 [0.62 , 1.37]	
Johnson 1998	11	47	18	25	14.6%	0.33 [0.18, 0.58]	_
Luria 2001	26	176	18	88	15.2%	0.72 [0.42 , 1.24]	
Skowron 1966a and b	0	97	2	97	1.2%	0.20 [0.01, 4.11]	—
Subtotal (95% CI)		786		668	73.9%	0.53 [0.34, 0.81]	•
Total events:	83		123				•
Heterogeneity: Tau ² = 0.1	6; Chi ² = 15.	54, df = 7	(P = 0.03);	$I^2 = 55\%$			
Test for overall effect: Z	= 2.90 (P = 0.	004)					
Total (95% CI)		916		763	100.0%	0.51 [0.36 , 0.72]	•
Total events:	106		156				· · · · · · · · · · · · · · · · · · ·
Heterogeneity: $Tau^2 = 0.1$	2; Chi ² = 20.	19, df = 11	(P = 0.04)	$I^2 = 46\%$			0.05 0.2 1 5 20
Test for overall effect: Z	= 3.90 (P < 0.	0001)				Favoi	urs glucocorticoid Favours place
Гest for subgroup differer	nces: $Chi^2 = 0$.28, df = 1	(P = 0.60)	$I^2 = 0\%$			



Analysis 1.14. Comparison 1: Any glucocorticoid compared to placebo, Outcome 14: Return visits or (re)admissions or both by croup severity

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.14.1 Mild croup							
Bjornson 2004	26	354	53	354	19.9%	0.49 [0.31, 0.77]	-
Geelhoed 1996a	0	48	8	48	1.7%	0.06 [0.00, 0.99]	-
Luria 2001	26	176	18	88	17.5%	0.72 [0.42 , 1.24]	-
Subtotal (95% CI)		578		490	39.1%	0.54 [0.30, 0.95]	
Total events:	52		79				~
Heterogeneity: $Tau^2 = 0.12$	1; Chi ² = 3.82	2, df = 2 (I	P = 0.15); I ²	= 48%			
Test for overall effect: Z =	2.14 (P = 0.	03)					
1.14.2 Moderate croup							
Cruz 1995	1	19	4	19	2.9%	0.25 [0.03, 2.04]	
Geelhoed 1995c	5	39	4	24	7.1%	0.77 [0.23, 2.59]	
Johnson 1996	17	27	17	25	21.2%	0.93 [0.62, 1.37]	-
Johnson 1998	29	95	35	49	22.3%	0.43 [0.30, 0.61]	-
Klassen 1994	1	27	7	27	3.0%	0.14 [0.02 , 1.08]	
Roberts 1999	1	34	7	32	3.0%	0.13 [0.02 , 1.03]	
Skowron 1966a and b	0	97	2	97	1.5%	0.20 [0.01, 4.11]	
Subtotal (95% CI)		338		273	60.9%	0.48 [0.26, 0.86]	•
Total events:	54		76				•
Heterogeneity: Tau ² = 0.25	5; Chi ² = 15.0	00, df = 6	(P = 0.02);	$I^2 = 60\%$			
Test for overall effect: Z =	2.47 (P = 0.	01)					
Total (95% CI)		916		763	100.0%	0.52 [0.36 , 0.76]	•
Total events:	106		155				•
Heterogeneity: Tau ² = 0.13	3; Chi ² = 18.4	49, df = 9	(P = 0.03);	$I^2 = 51\%$			0.01 0.1 1 10 100
Test for overall effect: Z =	3.41 (P = 0.	0006)				Favoi	urs glucocorticoid Favours placeb
Test for subgroup differen	ces: Chi ² = 0	.08, df = 1	(P = 0.77)	$I^2 = 0\%$			

Analysis 1.15. Comparison 1: Any glucocorticoid compared to placebo, Outcome 15: Length of stay by inpatient

	Glucocorticoid		d	1	Placebo			Mean Difference	Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
1.15.1 Inpatient										
Geelhoed 1995c	12.54	35.7	50	20	31.5	30	14.1%	-7.46 [-22.46 , 7.54]		
Godden 1997	36	35.7	47	55	31.5	42	15.0%	-19.00 [-32.96 , -5.04]		
Kuusela 1988	49	23	16	91	40	21	10.3%	-42.00 [-62.49 , -21.51]		
Leipzig 1979	34.4	17.7	16	55.4	22.3	14	14.5%	-21.00 [-35.55 , -6.45]		
Roorda 1998	62.4	18	9	57.6	18	8	12.5%	4.80 [-12.34 , 21.94]	_ _ _	
Skowron 1966a	38.4	35.7	41	55.2	31.5	53	15.1%	-16.80 [-30.63, -2.97]		
Skowron 1966b	48	35.7	56	61.44	31.5	44	15.6%	-13.44 [-26.63, -0.25]		
Super 1989	86.4	84	16	72	45.6	13	2.9%	14.40 [-33.65, 62.45]		
Subtotal (95% CI)			251			225	100.0%	-14.90 [-23.58, -6.22]	•	
Heterogeneity: Tau ² = 8	30.42; Chi ² = 1	15.29, df =	7 (P = 0.0	3); I ² = 549	6				•	
Test for overall effect: 2	Z = 3.36 (P =	0.0008)								
Total (95% CI)			251			225	100.0%	-14.90 [-23.58 , -6.22]	•	
Heterogeneity: Tau ² = 8	30.42; Chi ² = 1	15.29, df =	7 (P = 0.0	3); I ² = 54%	6					
Test for overall effect: 2	Z = 3.36 (P =	0.0008)							-50 -25 0 25 50	
Test for subgroup differ	rences: Not ap	plicable						Favou	rs glucocorticoid Favours pla	



Analysis 1.16. Comparison 1: Any glucocorticoid compared to placebo, Outcome 16: Length of stay by glucocorticoid

	Glu	cocortico	id		Placebo			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
1.16.1 Budesonide									
Geelhoed 1995c	13	35.7	27	20	31.5	15	9.4%	-7.00 [-27.87 , 13.87]	
Godden 1997	36	35.7	47	55	31.5	42	14.3%	-19.00 [-32.96 , -5.04]	
Subtotal (95% CI)			74			57	23.7%	-15.29 [-26.89 , -3.69]	
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0.	.88, df = 1	(P = 0.35)	$I^2 = 0\%$					•
Test for overall effect: 2	Z = 2.58 (P =	0.010)							
1.16.2 Dexamethasone	e								
Geelhoed 1995c	12	35.7	23	20	31.5	15	9.0%	-8.00 [-29.61, 13.61]	
Kuusela 1988	49	23	16	91	40	21	9.6%	-42.00 [-62.49 , -21.51]	
Leipzig 1979	34.4	17.7	16	55.4	22.3	14	13.8%	-21.00 [-35.55, -6.45]	
Skowron 1966a	38.4	35.7	41	55.2	31.5	53	14.4%	-16.80 [-30.63, -2.97]	
Skowron 1966b	48	35.7	56	61.44	31.5	44	15.0%	-13.44 [-26.63, -0.25]	-
Super 1989	86.4	84	16	72	45.6	13	2.6%	14.40 [-33.65, 62.45]	
Subtotal (95% CI)			168			160	64.5%	-18.25 [-27.87 , -8.62]	•
Heterogeneity: Tau ² = 5	56.28; Chi ² = 8	8.49, df =	5 (P = 0.13	s); I ² = 41%					~
Test for overall effect: 2	Z = 3.71 (P =	0.0002)							
1.16.3 Fluticasone									
Roorda 1998	62.4	18	9	57.6	18	8	11.8%	4.80 [-12.34, 21.94]	
Subtotal (95% CI)			9			8	11.8%	4.80 [-12.34, 21.94]	•
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.55 (P =	0.58)							
Total (95% CI)			251			225	100.0%	-14.55 [-22.70 , -6.41]	•
Heterogeneity: Tau ² = 6	59.89; Chi ² = 1	15.29, df =	= 8 (P = 0.0	5); I ² = 489	6				•
Test for overall effect: 2	Z = 3.50 (P =	0.0005)							-50 -25 0 25 50
Test for subgroup differ	rences: Chi ² =	5.40, df =	= 2 (P = 0.0)	7), $I^2 = 63.0$)%			Favou	rs glucocorticoid Favours place

Analysis 1.17. Comparison 1: Any glucocorticoid compared to placebo, Outcome 17: Improvement (at 2 hours) by inpatient

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.17.1 Inpatient							
Roberts 1999	19	42	10	40	100.0%	1.81 [0.96, 3.40]	
Subtotal (95% CI)		42		40	100.0%	1.81 [0.96, 3.40]	
Total events:	19		10				
Heterogeneity: Not appl	licable						
Test for overall effect: Z	Z = 1.84 (P =	0.07)					
Total (95% CI)		42		40	100.0%	1.81 [0.96 , 3.40]	
Total events:	19		10				
Heterogeneity: Not appl	licable						$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
Test for overall effect: Z	Z = 1.84 (P =	0.07)					Favours placebo Favours glucocorticoid
Test for subgroup differ	ences: Not a	pplicable					



Analysis 1.18. Comparison 1: Any glucocorticoid compared to placebo, Outcome 18: Improvement (at 6 hours) by inpatient/outpatient

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.18.1 Inpatient							
Eden 1964	17	25	17	25	24.2%	1.00 [0.68 , 1.46]	-
Eden 1967	10	25	10	25	11.4%	1.00 [0.51 , 1.97]	
Massicotte 1973	20	25	7	17	13.7%	1.94 [1.07, 3.54]	
Roberts 1999	32	42	18	40	24.1%	1.69 [1.16, 2.48]	_
Subtotal (95% CI)		117		107	73.4%	1.35 [0.96, 1.90]	
Total events:	79		52				
Heterogeneity: Tau ² = 0	0.06; Chi ² = 5	5.98, df = 3	P = 0.11	$I^2 = 50\%$			
Test for overall effect:	Z = 1.70 (P =	0.09)					
1.18.2 Outpatient							
Johnson 1996	20	28	10	27	15.7%	1.93 [1.12 , 3.33]	_
Klassen 1994	13	27	8	26	10.9%	1.56 [0.78, 3.14]	
Subtotal (95% CI)		55		53	26.6%	1.78 [1.16, 2.74]	
Total events:	33		18				
Heterogeneity: Tau ² = 0	0.00; Chi ² = ().22, df = 1	(P = 0.64)	$I^2 = 0\%$			
Test for overall effect:	Z = 2.64 (P =	(800.0					
Total (95% CI)		172		160	100.0%	1.45 [1.12 , 1.88]	•
Total events:	112		70				•
Heterogeneity: Tau ² = 0	0.04; Chi ² = 7	7.56, df = 5	6(P = 0.18)	$I^2 = 34\%$			$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
Test for overall effect:	Z = 2.78 (P =	0.005)					Favours placebo Favours glucocorticoic
Test for subgroup differ	rences: Chi ²	= 0.99, df =	= 1 (P = 0.3	2), I ² = 0%	, D		

Analysis 1.19. Comparison 1: Any glucocorticoid compared to placebo, Outcome 19: Improvement (at 12 hours) by inpatient

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.19.1 Inpatient							
Eden 1964	21	25	21	25	22.4%	1.00 [0.79, 1.27]	
Eden 1967	20	25	17	25	17.3%	1.18 [0.84, 1.64]	
James 1969	35	45	20	43	16.2%	1.67 [1.17, 2.39]	
Massicotte 1973	24	25	12	17	18.1%	1.36 [0.99, 1.87]	
Roberts 1999	37	42	25	40	21.1%	1.41 [1.08, 1.84]	-
Super 1989	13	16	4	12	4.9%	2.44 [1.06, 5.61]	
Subtotal (95% CI)		178		162	100.0%	1.33 [1.09, 1.62]	•
Total events:	150		99				•
Heterogeneity: Tau ² = 0	0.03; Chi ² = 1	0.59, df =	5 (P = 0.06); I ² = 53%	ó		
Test for overall effect: 2	Z = 2.78 (P =	0.005)					
Total (95% CI)		178		162	100.0%	1.33 [1.09 , 1.62]	•
Total events:	150		99				•
Heterogeneity: Tau ² = 0	0.03; Chi ² = 1	0.59, df =	5 (P = 0.06); I ² = 53%	ó		0.5 0.7 1 1.5 2
Test for overall effect: 2	Z = 2.78 (P =	0.005)					Favours placebo Favours glucocorticoid
Test for subgroup differ	ences: Not a	pplicable					-



Analysis 1.20. Comparison 1: Any glucocorticoid compared to placebo, Outcome 20: Improvement (at 24 hours) by inpatient/outpatient

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.20.1 Inpatient							
Eden 1964	24	25	23	25	29.1%	1.04 [0.91 , 1.20]	•
Eden 1967	24	25	21	25	26.8%	1.14 [0.95 , 1.38]	-
James 1969	42	45	32	43	26.6%	1.25 [1.04 , 1.52]	-
Super 1989	11	13	4	12	6.3%	2.54 [1.10, 5.84]	
Subtotal (95% CI)		108		105	88.8%	1.18 [0.98, 1.43]	•
Total events:	101		80				•
Heterogeneity: Tau ² = 0.0	02; Chi ² = 8	.86, df = 3	P = 0.03	$I^2 = 66\%$			
Test for overall effect: Z	= 1.71 (P =	0.09)					
1.20.2 Outpatient							
Cruz 1995	16	19	8	19	11.2%	2.00 [1.14, 3.51]	
Subtotal (95% CI)		19		19	11.2%	2.00 [1.14, 3.51]	
Total events:	16		8				
Heterogeneity: Not appli	cable						
Test for overall effect: Z	= 2.42 (P =	0.02)					
Total (95% CI)		127		124	100.0%	1.28 [1.01 , 1.61]	•
Total events:	117		88				•
Heterogeneity: Tau ² = 0.0	04; Chi ² = 1	6.09, df =	4 (P = 0.00	3); I ² = 75 ⁶	%		0.05 0.2 1 5 20
Test for overall effect: Z	= 2.07 (P =	0.04)					Favours placebo Favours glucocortico
Test for subgroup differe	nces: Chi² =	3.02, df =	= 1 (P = 0.0	8), I ² = 66.	8%		



Analysis 1.21. Comparison 1: Any glucocorticoid compared to placebo, Outcome 21: Improvement (at 6 hours) by glucocorticoid

	Glucoco	rticoid	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.21.1 Budesonide							
Klassen 1994	13	27	8	26	10.9%	1.56 [0.78, 3.14]	
Roberts 1999	32	42	18	40	24.1%	1.69 [1.16, 2.48]	_ _
Subtotal (95% CI)		69		66	35.0%	1.66 [1.19, 2.32]	•
Total events:	45		26				•
Heterogeneity: Tau ² = 0.0	00; Chi ² = 0	0.04, df = 1	(P = 0.84)	$I^2 = 0\%$			
Test for overall effect: Z	= 2.97 (P =	0.003)					
1,21.2 Dexamethasone							
Eden 1967	10	25	10	25	11.4%	1.00 [0.51 , 1.97]	
Johnson 1996	20	28	10	27	15.7%	1.93 [1.12, 3.33]	
Subtotal (95% CI)		53		52	27.1%	1.43 [0.76, 2.72]	
Total events:	30		20				
Heterogeneity: Tau ² = 0.	12; Chi ² = 2	2.19, df = 1	(P = 0.14)	$I^2 = 54\%$			
Test for overall effect: Z	= 1.10 (P =	0.27)					
1.21.3 Prednisolone							
Eden 1964	17	25	17	25	24.2%	1.00 [0.68, 1.46]	<u> </u>
Massicotte 1973	20	25	7	17	13.7%	1.94 [1.07, 3.54]	
Subtotal (95% CI)		50		42	37.9%	1.34 [0.69, 2.62]	
Total events:	37		24				
Heterogeneity: Tau ² = 0.1	17; Chi ² = 3	3.59, df = 1	(P = 0.06)	$I^2 = 72\%$			
Test for overall effect: Z	= 0.86 (P =	0.39)					
Total (95% CI)		172		160	100.0%	1.45 [1.12 , 1.88]	•
Total events:	112		70				\
Heterogeneity: Tau ² = 0.0	04; Chi ² = 7	7.56, df = 5	(P = 0.18)	$I^2 = 34\%$			$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
Test for overall effect: Z	= 2.78 (P =	0.005)					Favours placebo Favours glucocort
Test for subgroup differe	nces: Chi ²	= 0.40, df =	2 (P = 0.8	2), $I^2 = 0\%$	ó		



Analysis 1.22. Comparison 1: Any glucocorticoid compared to placebo, Outcome 22: Improvement (at 12 hours) by glucocorticoid

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.22.1 Budesonide							
Roberts 1999	37	42	25	40	21.1%	1.41 [1.08, 1.84]	
Subtotal (95% CI)		42		40	21.1%	1.41 [1.08, 1.84]	•
Total events:	37		25				
Heterogeneity: Not appli	cable						
Test for overall effect: Z	= 2.54 (P =	0.01)					
1.22.2 Dexamethasone							
Eden 1967	20	25	17	25	17.3%	1.18 [0.84, 1.64]	
James 1969	35	45	20	43	16.2%	1.67 [1.17, 2.39]	
Super 1989	13	16	4	12	4.9%	2.44 [1.06, 5.61]	
Subtotal (95% CI)		86		80	38.4%	1.52 [1.06, 2.18]	
Total events:	68		41				
Heterogeneity: Tau ² = 0.0	05; Chi ² = 3	3.97, df = 2	P = 0.14	$I^2 = 50\%$			
Test for overall effect: Z	= 2.25 (P =	0.02)					
1.22.3 Prednisolone							
Eden 1964	21	25	21	25	22.4%	1.00 [0.79, 1.27]	- - -
Massicotte 1973	24	25	12	17	18.1%	1.36 [0.99, 1.87]	-
Subtotal (95% CI)		50		42	40.6%	1.15 [0.85 , 1.55]	•
Total events:	45		33				
Heterogeneity: Tau ² = 0.0	03; Chi ² = 2	2.34, df = 1	(P = 0.13)	$I^2 = 57\%$			
Test for overall effect: Z	= 0.88 (P =	0.38)					
Гotal (95% СІ)		178		162	100.0%	1.33 [1.09, 1.62]	•
Total events:	150		99				
Heterogeneity: Tau ² = 0.0	03; Chi ² = 1	0.59, df =	5 (P = 0.06); I ² = 53%	ó		0.1 0.2 0.5 1 2 5 10
Test for overall effect: Z	= 2.78 (P =	0.005)					Favours placebo Favours glucocor
Гest for subgroup differe	nces: Chi2 =	= 1.61 df =	= 2 (P = 0.4)	5) $I^2 = 0\%$,		



Analysis 1.23. Comparison 1: Any glucocorticoid compared to placebo, Outcome 23: Improvement (at 24 hours) by glucocorticoid

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.23.1 Dexamethasone							
Cruz 1995	16	19	8	19	11.2%	2.00 [1.14, 3.51]	_
Eden 1967	24	25	21	25	26.8%	1.14 [0.95 , 1.38]	-
James 1969	42	45	32	43	26.6%	1.25 [1.04, 1.52]	-
Super 1989	11	13	4	12	6.3%	2.54 [1.10, 5.84]	
Subtotal (95% CI)		102		99	70.9%	1.39 [1.05, 1.84]	•
Total events:	93		65				ľ
Heterogeneity: Tau ² = 0.0	04; $Chi^2 = 8$.86, df = 3	(P = 0.03)	$I^2 = 66\%$			
Test for overall effect: Z	= 2.31 (P =	0.02)					
1.23.2 Prednisolone							
Eden 1964	24	25	23	25	29.1%	1.04 [0.91, 1.20]	•
Subtotal (95% CI)		25		25	29.1%	1.04 [0.91, 1.20]	•
Total events:	24		23				ľ
Heterogeneity: Not appli	cable						
Test for overall effect: Z	= 0.59 (P =	0.55)					
Total (95% CI)		127		124	100.0%	1.28 [1.01 , 1.61]	
Total events:	117		88				▼
Heterogeneity: Tau ² = 0.0	04; Chi ² = 1	6.09, df =	4 (P = 0.00	3); I ² = 75 ⁹	%		0.1 0.2 0.5 1 2 5 10
Test for overall effect: Z	= 2.07 (P =	0.04)					Favours placebo Favours glucocorticoid
Test for subgroup differe	nces: Chi² =	3.24, df =	= 1 (P = 0.0	7), I ² = 69.	1%		-

Analysis 1.24. Comparison 1: Any glucocorticoid compared to placebo, Outcome 24: Additional treatments: antibiotics

	Glucoco	rticoid	Place	ebo		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Husby 1993	2	20	1	16	6.2%	0.04 [-0.14 , 0.21]	
James 1969	40	45	39	43	12.1%	-0.02 [-0.14 , 0.11]	
Koren 1983	0	40	0	38	81.7%	0.00 [-0.05, 0.05]	•
Total (95% CI)		105		97	100.0%	0.00 [-0.04, 0.04]	
Total events:	42		40				T
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0	.25, df = 2	P = 0.88	$I^2 = 0\%$		-	-0.2-0.1 0 0.1 0.2
Test for overall effect:	Z = 0.01 (P =	1.00)				Favours	glucocorticoid Favours placebo

Test for subgroup differences: Not applicable



Analysis 1.25. Comparison 1: Any glucocorticoid compared to placebo, Outcome 25: Additional treatments: epinephrine

	Glucoco	rticoid	Place	ebo		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Dobrovoljac 2012	0	35	2	35	13.7%	-0.06 [-0.15 , 0.03]	
Geelhoed 1995c	8	42	5	25	4.8%	-0.01 [-0.21 , 0.19]	
Godden 1997	3	47	4	42	10.8%	-0.03 [-0.14, 0.08]	
Johnson 1996	5	28	1	27	6.8%	0.14 [-0.02, 0.30]	
Johnson 1998	4	95	4	49	14.5%	-0.04 [-0.13, 0.05]	
Klassen 1994	0	27	2	27	10.4%	-0.07 [-0.19, 0.04]	
Koren 1983	0	40	0	38	21.5%	0.00 [-0.05, 0.05]	
Roberts 1999	1	42	3	40	13.4%	-0.05 [-0.14, 0.04]	
Tibballs 1992	8	38	16	32	4.1%	-0.29 [-0.51 , -0.07]	
Total (95% CI)		394		315	100.0%	-0.03 [-0.08, 0.01]	
Total events:	29		37				
Heterogeneity: Tau ² = 0	0.00; Chi ² = 1	4.48, df =	8 (P = 0.07); I ² = 45%	ó		-0.2 -0.1 0 0.1 0.2
Test for overall effect:	Z = 1.41 (P =	0.16)				Favoi	urs glucocorticoid Favours place

Test for overall effect: Z = 1.41 (P = 0.16) Test for subgroup differences: Not applicable

Analysis 1.26. Comparison 1: Any glucocorticoid compared to placebo, Outcome 26: Additional treatments: intubation/tracheostomy

	Glucoco	rticoid	Place	ebo		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Eden 1967	0	25	0	25	2.4%	0.00 [-0.07 , 0.07]	
Geelhoed 1995c	0	44	0	30	4.7%	0.00 [-0.05, 0.05]	
Godden 1997	0	47	2	42	2.4%	-0.05 [-0.12, 0.03]	
James 1969	0	45	0	43	7.2%	0.00 [-0.04, 0.04]	
Johnson 1996	2	28	0	27	1.1%	0.07 [-0.04, 0.18]	
Johnson 1998	0	95	0	49	14.0%	0.00 [-0.03, 0.03]	
Leipzig 1979	0	16	1	14	0.5%	-0.07 [-0.24 , 0.10]	(-
Roorda 1998	0	9	0	8	0.3%	0.00 [-0.20 , 0.20]	←
Skowron 1966a	0	41	1	53	4.5%	-0.02 [-0.07, 0.04]	-
Skowron 1966b	0	44	0	56	8.9%	0.00 [-0.04, 0.04]	
Von Mühlendahl 1982	1	176	1	173	54.0%	-0.00 [-0.02 , 0.02]	-
Total (95% CI)		570		520	100.0%	-0.00 [-0.01 , 0.01]	
Total events:	3		5				Ť
Heterogeneity: Tau ² = 0.0	0; Chi ² = 4.25	5, df = 10 ((P = 0.94);	$I^2 = 0\%$			-0.05-0.025 0 0.025 0.05
Test for overall effect: Z =	= 0.27 (P = 0.5)	79)				Favou	urs glucocorticoid Favours placebo

Test for overall effect: Z = 0.27 (P = 0.79) Test for subgroup differences: Not applicable



Analysis 1.27. Comparison 1: Any glucocorticoid compared to placebo, Outcome 27: Additional treatments: mist tent

	Glucoco	rticoid	Place	ebo		Risk Difference	Risk Diff	erence
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	m, 95% CI
Johnson 1996	0	28	0	27	52.1%	0.00 [-0.07 , 0.07]		
Super 1989	8	16	12	13	47.9%	-0.42 [-0.71 , -0.14]	 -	
Total (95% CI)		44		40	100.0%	-0.20 [-0.87 , 0.47]		
Total events:	8		12					
Heterogeneity: Tau ² = 0	0.22; Chi ² = 2	21.05, df =	1 (P < 0.00	001); I ² =	95%		-1 -0.5 0	0.5 1
Test for overall effect:	Z = 0.59 (P =	0.55)				Favou	rs glucocorticoid	Favours placebo

Test for subgroup differences: Not applicable

Analysis 1.28. Comparison 1: Any glucocorticoid compared to placebo, Outcome 28: Additional treatments: supplemental glucocorticoids

	Glucoco	rticoid	Place	ebo		Risk Ratio	Risk R	latio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	m, 95% CI
Dobrovoljac 2012	0	35	4	35	3.3%	0.11 [0.01 , 1.99]		_
Geelhoed 1995c	4	50	7	30	18.1%	0.34 [0.11, 1.07]	_	
Johnson 1996	0	28	2	27	3.1%	0.19 [0.01, 3.85]		
Klassen 1994	15	27	21	27	69.9%	0.71 [0.48, 1.06]		
Roorda 1998	1	9	0	8	2.9%	2.70 [0.13, 58.24]		
Super 1989	1	16	0	13	2.8%	2.47 [0.11, 56.03]		•
Total (95% CI)		165		140	100.0%	0.61 [0.36 , 1.03]		
Total events:	21		34				•	
Heterogeneity: Tau ² = 0	0.06; Chi ² = 5	.53, df = 5	6(P = 0.35)	$I^2 = 10\%$			0.005 0.1 1	10 200
Test for overall effect:	Z = 1.84 (P =	0.07)					urs glucocorticoid	Favours placebo

Test for subgroup differences: Not applicable

Comparison 2. Any glucocorticoid compared to epinephrine

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
2.1 Croup score (change baseline - 2 hours) by inpatient/outpatient	2	130	Std. Mean Difference (IV, Random, 95% CI)	0.77 [-0.24, 1.77]
2.1.1 Inpatient	1	66	Std. Mean Difference (IV, Random, 95% CI)	0.26 [-0.22, 0.75]
2.1.2 Outpatient	1	64	Std. Mean Difference (IV, Random, 95% CI)	1.29 [0.73, 1.84]
2.2 Croup score (change base- line - 6 hours) by inpatient	2	63	Std. Mean Difference (IV, Random, 95% CI)	-0.10 [-1.18, 0.97]
2.2.1 Inpatient	2	63	Std. Mean Difference (IV, Random, 95% CI)	-0.10 [-1.18, 0.97]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
2.3 Croup score (change base- line - 12 hours) by inpatient	3	129	Std. Mean Difference (IV, Random, 95% CI)	-0.07 [-0.57, 0.43]
2.3.1 Inpatient	3	129	Std. Mean Difference (IV, Random, 95% CI)	-0.07 [-0.57, 0.43]
2.4 Croup score (change base- line - 24 hours) by inpatient	3	129	Std. Mean Difference (IV, Random, 95% CI)	0.17 [-0.18, 0.51]
2.4.1 Inpatient	3	129	Std. Mean Difference (IV, Random, 95% CI)	0.17 [-0.18, 0.51]
2.5 Croup score (change base- line - 2 hours) by glucocorti- coid	2	130	Std. Mean Difference (IV, Random, 95% CI)	0.88 [0.13, 1.63]
2.5.1 Budesonide	1	66	Std. Mean Difference (IV, Random, 95% CI)	0.26 [-0.22, 0.75]
2.5.2 Dexamethasone	1	31	Std. Mean Difference (IV, Random, 95% CI)	1.13 [0.35, 1.91]
2.5.3 Beclomethasone	1	33	Std. Mean Difference (IV, Random, 95% CI)	1.41 [0.62, 2.19]
2.6 Croup score (change base- line - 12 hours) by glucocorti- coid	3	129	Std. Mean Difference (IV, Random, 95% CI)	-0.07 [-0.57, 0.43]
2.6.1 Budesonide	1	66	Std. Mean Difference (IV, Random, 95% CI)	0.02 [-0.47, 0.50]
2.6.2 Dexamethasone	2	63	Std. Mean Difference (IV, Random, 95% CI)	-0.14 [-1.09, 0.82]
2.7 Croup score (change base- line - 24 hours) by glucocorti- coid	3	129	Std. Mean Difference (IV, Random, 95% CI)	0.17 [-0.18, 0.51]
2.7.1 Budesonide	1	66	Std. Mean Difference (IV, Random, 95% CI)	0.21 [-0.27, 0.70]
2.7.2 Dexamethasone	2	63	Std. Mean Difference (IV, Random, 95% CI)	0.12 [-0.38, 0.61]
2.8 Return visits or (re)admissions or both by inpatient/outpatient	2	130	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.04, 0.04]
2.8.1 Inpatient	1	66	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.06, 0.06]
2.8.2 Outpatient	1	64	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.06, 0.06]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
2.9 Length of stay by inpatient	1	32	Mean Difference (IV, Random, 95% CI)	-10.00 [-33.89, 13.89]
2.9.1 Inpatient	1	32	Mean Difference (IV, Random, 95% CI)	-10.00 [-33.89, 13.89]
2.10 Additional treatments: epinephrine	1	66	Risk Ratio (M-H, Random, 95% CI)	0.30 [0.03, 2.69]
2.11 Additional treatments: intubation/tracheostomy	1	66	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.06, 0.06]
2.12 Additional treatments: supplemental glucocorticoids	1	66	Risk Ratio (M-H, Random, 95% CI)	0.83 [0.48, 1.43]

Analysis 2.1. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 1: Croup score (change baseline - 2 hours) by inpatient/outpatient

	Glu	cocortico	id	Ep	inephrine	<u>.</u>		Std. Mean Difference	Std. Mea	an Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Rand	lom, 95% CI
2.1.1 Inpatient										
Fitzgerald 1996	-3.2	1.9	35	-3.74	2.18	31	50.9%	0.26 [-0.22, 0.75]		—
Subtotal (95% CI)			35			31	50.9%	0.26 [-0.22, 0.75]		
Heterogeneity: Not app	licable									
Test for overall effect: 2	Z = 1.06 (P =	0.29)								
2.1.2 Outpatient										
Eboriadou 2010	-1.62	1.9	39	-4.24	2.18	25	49.1%	1.29 [0.73, 1.84]		
Subtotal (95% CI)			39			25	49.1%	1.29 [0.73, 1.84]		
Heterogeneity: Not app	licable									
Test for overall effect: 2	Z = 4.56 (P < 0.00)	0.00001)								
Total (95% CI)			74			56	100.0%	0.77 [-0.24 , 1.77]		
Heterogeneity: Tau ² = 0).45; Chi ² = 7.	44, df = 1	(P = 0.006)	5); I ² = 87%						
Test for overall effect: 2	Z = 1.50 (P =	0.13)						-	-2 -1	0 1 2
Test for subgroup differ	rences: Chi ² =	7.44, df =	1 (P = 0.0	06), I ² = 86	5.6%			Favours	glucocorticoid	Favours epinephrii

Analysis 2.2. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 2: Croup score (change baseline - 6 hours) by inpatient

	Glu	cocorticoi	id	Ep	inephrine	2		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
2.2.1 Inpatient									
Kuusela 1988	-1.65	0.7	16	-1.25	0.47	16	50.0%	-0.65 [-1.37, 0.06]	
Martinez Fernandez 1993	-0.4	1.45	16	-1.1	1.61	15	50.0%	0.45 [-0.27, 1.16]	
Subtotal (95% CI)			32			31	100.0%	-0.10 [-1.18 , 0.97]	
Heterogeneity: Tau ² = 0.47; C	Chi ² = 4.56, df	= 1 (P = 0)	.03); I ² = 7	'8%					\top
Test for overall effect: $Z = 0.1$	19 (P = 0.85)								
Total (95% CI)			32			31	100.0%	-0.10 [-1.18 , 0.97]	
Heterogeneity: Tau ² = 0.47; C	$Chi^2 = 4.56$, df	= 1 (P = 0)	.03); I ² = 7	'8%					T
Test for overall effect: $Z = 0.1$	19 (P = 0.85)								-4 -2 0 2 4
Test for subgroup differences	: Not applicab	le						Favou	urs glucocorticoid Favours epinephi



Analysis 2.3. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 3: Croup score (change baseline - 12 hours) by inpatient

	Glu	cocorticoi	id	Ep	inephrine	<u>!</u>		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
2.3.1 Inpatient									
Fitzgerald 1996	-3.82	1.84	35	-3.86	2.86	31	42.2%	0.02 [-0.47, 0.50]	
Kuusela 1988	-1.9	0.63	16	-1.45	0.77	16	28.8%	-0.62 [-1.34, 0.09]	
Martinez Fernandez 1993	-1.1	1.51	16	-1.7	1.81	15	28.9%	0.35 [-0.36 , 1.06]	
Subtotal (95% CI)			67			62	100.0%	-0.07 [-0.57, 0.43]	
Heterogeneity: Tau ² = 0.09; C	$hi^2 = 3.79$, df	= 2 (P = 0)	.15); I ² = 4	7%					\top
Test for overall effect: $Z = 0.2$	P = 0.78								
Total (95% CI)			67			62	100.0%	-0.07 [-0.57 , 0.43]	
Heterogeneity: Tau ² = 0.09; C	$hi^2 = 3.79$, df	= 2 (P = 0)	.15); I ² = 4	7%					$\overline{}$
Test for overall effect: $Z = 0.2$	28 (P = 0.78)								-1 -0.5 0 0.5 1
Test for subgroup differences:	Not applicab	le						Favou	rs glucocorticoid Favours epinephrii

Analysis 2.4. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 4: Croup score (change baseline - 24 hours) by inpatient

	Glu	cocorticoi	id	Ep	inephrine	!		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
2.4.1 Inpatient									
Fitzgerald 1996	-3.84	2.38	35	-4.4	2.83	31	51.1%	0.21 [-0.27, 0.70]	
Kuusela 1988	-2.09	1.81	16	-2.01	2.2	16	25.0%	-0.04 [-0.73, 0.65]	
Martinez Fernandez 1993	-1.9	1.23	16	-2.3	1.57	15	23.9%	0.28 [-0.43, 0.99]	
Subtotal (95% CI)			67			62	100.0%	0.17 [-0.18, 0.51]	
Heterogeneity: Tau ² = 0.00; C	Chi ² = 0.47, df	= 2 (P = 0)	.79); I ² = 0	%					
Test for overall effect: $Z = 0.9$	94 (P = 0.35)								
Total (95% CI)			67			62	100.0%	0.17 [-0.18 , 0.51]	
Heterogeneity: Tau ² = 0.00; C	Chi ² = 0.47, df	= 2 (P = 0)	.79); I ² = 0	%					
Test for overall effect: $Z = 0.9$	94 (P = 0.35)								-1 -0.5 0 0.5 1
Test for subgroup differences:	: Not applicabl	le						Favou	rs glucocorticoid Favours epiner



Analysis 2.5. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 5: Croup score (change baseline - 2 hours) by glucocorticoid

	Glu	cocorticoi	id	Ep	inephrine	•		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
2.5.1 Budesonide									
Fitzgerald 1996	-3.2	1.9	35	-3.74	2.18	31	38.7%	0.26 [-0.22, 0.75]	-
Subtotal (95% CI)			35			31	38.7%	0.26 [-0.22, 0.75]	•
Heterogeneity: Not appli	icable								_
Test for overall effect: Z	= 1.06 (P =	0.29)							
2.5.2 Dexamethasone									
Eboriadou 2010	-1.91	1.9	19	-4.24	2.18	12	30.7%	1.13 [0.35 , 1.91]	
Subtotal (95% CI)			19			12	30.7%	1.13 [0.35, 1.91]	•
Heterogeneity: Not appli	icable								
Test for overall effect: Z	= 2.83 (P =	0.005)							
2.5.3 Beclomethasone									
Eboriadou 2010	-1.34	1.9	20	-4.24	2.18	13	30.6%	1.41 [0.62, 2.19]	
Subtotal (95% CI)			20			13	30.6%	1.41 [0.62, 2.19]	•
Heterogeneity: Not appli	icable								•
Test for overall effect: Z	= 3.50 (P =	0.0005)							
Total (95% CI)			74			56	100.0%	0.88 [0.13, 1.63]	
Heterogeneity: Tau ² = 0.	32; Chi ² = 7.	.37, df = 2	(P = 0.03)	; I ² = 73%					•
Test for overall effect: Z	= 2.29 (P =	0.02)						_	-4 -2 0 2 4
Test for subgroup differe	ences: Chi² =	7.37, df =	2 (P = 0.0	3), I ² = 72.9	9%			Favours §	glucocorticoid Favours epinepl

Analysis 2.6. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 6: Croup score (change baseline - 12 hours) by glucocorticoid

	Glu	cocorticoi	id	Ep	inephrine	<u>:</u>		Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
2.6.1 Budesonide										
Fitzgerald 1996	-3.82	1.84	35	-3.86	2.86	31	42.2%	0.02 [-0.47, 0.50]	-	
Subtotal (95% CI)			35			31	42.2%	0.02 [-0.47, 0.50]	•	
Heterogeneity: Not applicable									Ť	
Test for overall effect: $Z = 0.07$	(P = 0.95)									
2.6.2 Dexamethasone										
Kuusela 1988	-1.9	0.63	16	-1.45	0.77	16	28.8%	-0.62 [-1.34, 0.09]		
Martinez Fernandez 1993	-1.1	1.51	16	-1.7	1.81	15	28.9%	0.35 [-0.36 , 1.06]		
Subtotal (95% CI)			32			31	57.8%	-0.14 [-1.09, 0.82]	—	
Heterogeneity: Tau ² = 0.34; Ch	$i^2 = 3.61$, df	= 1 (P = 0)	.06); I ² = 7	2%					\top	
Test for overall effect: $Z = 0.28$	(P = 0.78)									
Total (95% CI)			67			62	100.0%	-0.07 [-0.57 , 0.43]		
Heterogeneity: Tau ² = 0.09; Ch	$i^2 = 3.79$, df	= 2 (P = 0)	.15); I ² = 4	7%					Ť	
Test for overall effect: $Z = 0.28$	(P = 0.78)								-4 -2 0 2 4	
Test for subgroup differences: O	$Chi^2 = 0.08$.	df = 1 (P =	0.78), I ² =	= 0%				Favou	rs glucocorticoid Favours epinephi	



Analysis 2.7. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 7: Croup score (change baseline - 24 hours) by glucocorticoid

	Glu	cocortico	id	Ep	Epinephrine			Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
2.7.1 Budesonide										
Fitzgerald 1996	-3.84	2.38	35	-4.4	2.83	31	51.1%	0.21 [-0.27, 0.70]		
Subtotal (95% CI)			35			31	51.1%	0.21 [-0.27, 0.70]		
Heterogeneity: Not applicable										
Test for overall effect: $Z = 0.86$	(P = 0.39)									
2.7.2 Dexamethasone										
Kuusela 1988	-2.09	1.81	16	-2.01	2.2	16	25.0%	-0.04 [-0.73, 0.65]		
Martinez Fernandez 1993	-1.9	1.23	16	-2.3	1.57	15	23.9%	0.28 [-0.43, 0.99]		
Subtotal (95% CI)			32			31	48.9%	0.12 [-0.38, 0.61]		
Heterogeneity: Tau ² = 0.00; Chi	$i^2 = 0.39$, df	= 1 (P = 0)	.53); I ² = 0	%						
Test for overall effect: $Z = 0.46$	(P = 0.65)									
Total (95% CI)			67			62	100.0%	0.17 [-0.18 , 0.51]		
Heterogeneity: Tau ² = 0.00; Chi	$i^2 = 0.47$, df	= 2 (P = 0)	.79); I ² = 0	%						
Test for overall effect: Z = 0.94	(P = 0.35)								-2 -1 0 1 2	
Test for subgroup differences: C	$Chi^2 = 0.08$,	df = 1 (P =	0.78), I ² =	= 0%				Favou	rs glucocorticoid Favours epinephr	

Analysis 2.8. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 8: Return visits or (re)admissions or both by inpatient/outpatient

	Glucoco	rticoid	Epinep	hrine		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
2.8.1 Inpatient							
Fitzgerald 1996	0	35	0	31	54.6%	0.00 [-0.06, 0.06]	_ _
Subtotal (95% CI)		35		31	54.6%	0.00 [-0.06, 0.06]	
Total events:	0		0				T
Heterogeneity: Not applie	cable						
Test for overall effect: Z	= 0.00 (P =	1.00)					
2.8.2 Outpatient							
Eboriadou 2010	0	39	0	25	45.4%	0.00 [-0.06, 0.06]	
Subtotal (95% CI)		39		25	45.4%	0.00 [-0.06, 0.06]	
Total events:	0		0				
Heterogeneity: Not applie	cable						
Test for overall effect: Z	= 0.00 (P =	1.00)					
Total (95% CI)		74		56	100.0%	0.00 [-0.04 , 0.04]	
Total events:	0		0				Ť
Heterogeneity: $Tau^2 = 0.0$ Test for overall effect: $Z = 0.0$			(P = 1.00)	$I^2 = 0\%$		-0 Favours	.2 -0.1 0 0.1 0.2 glucocorticoid Favours epinephrir

Test for subgroup differences: Chi² = 0.00, df = 1 (P = 1.00), I^2 = 0%



Analysis 2.9. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 9: Length of stay by inpatient

	Glu	cocortico	id	$\mathbf{E}_{\mathbf{I}}$	inephrine	2		Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
2.9.1 Inpatient									
Kuusela 1988	49	23	16	59	43	16	100.0%	-10.00 [-33.89 , 13.89]	
Subtotal (95% CI)			16			16	100.0%	-10.00 [-33.89 , 13.89]	
Heterogeneity: Not appl	icable								
Test for overall effect: Z	L = 0.82 (P = 0.82)	0.41)							
Total (95% CI)			16			16	100.0%	-10.00 [-33.89 , 13.89]	
Heterogeneity: Not appl	icable								
Test for overall effect: Z	= 0.82 (P = 0.82)	0.41)							-100 -50 0 50 100
Test for subgroup differen							urs glucocorticoid Favours epinephrin		

Analysis 2.10. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 10: Additional treatments: epinephrine

Study or Subgroup	Glucoco Events	rticoid Total	Epinep Events	hrine Total	Weight	Risk Ratio M-H, Random, 95% CI	Risk R M-H, Rando	
Fitzgerald 1996	1	35	3	31	100.0%	0.30 [0.03 , 2.69]	-	_
Total (95% CI) Total events:	1	35	3	31	100.0%	0.30 [0.03 , 2.69]		-
Heterogeneity: Not appl	licable						0.01 0.1 1	10 100
Test for overall effect: $Z = 1.08 (P = 0.28)$							ars glucocorticoid	Favours epinephrine
Test for subgroup differ	ences: Not a	pplicable						

Analysis 2.11. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 11: Additional treatments: intubation/tracheostomy

	Glucoco	rticoid	Epinep	hrine		Risk Difference	Risk Difference		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random,	95% CI	
Fitzgerald 1996	0	35	0	31	100.0%	0.00 [-0.06 , 0.06]	-		
Total (95% CI)		35		31	100.0%	0.00 [-0.06, 0.06]			
Total events:	0		0				Ť		
Heterogeneity: Not app	licable						-0.2 -0.1 0 (0.1 0.2	
Test for overall effect: 2	Z = 0.00 (P =	1.00)				Favour		Favours epinephrine	
Test for subgroup differ	ences: Not a	pplicable							



Analysis 2.12. Comparison 2: Any glucocorticoid compared to epinephrine, Outcome 12: Additional treatments: supplemental glucocorticoids

	Glucoco	Glucocorticoid		Epinephrine		Risk Ratio	Risk Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Randon	n, 95% CI
Fitzgerald 1996	14	35	15	31	100.0%	0.83 [0.48 , 1.43]	•	
Total (95% CI)		35		31	100.0%	0.83 [0.48 , 1.43]		
Total events:	14		15				7	
Heterogeneity: Not app	licable					0.	01 0.1 1	10 100
Test for overall effect: 2	Z = 0.68 (P =	0.49)					glucocorticoid	Favours epinephrine
Test for subgroup differ	ences: Not a	pplicable						

Comparison 3. Dexamethasone compared to budesonide

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
3.1 Croup score (change baseline - 6 hours) by inpatient/outpatient	4	326	Std. Mean Difference (IV, Random, 95% CI)	-0.46 [-0.79, -0.13]
3.1.1 Inpatient	2	97	Std. Mean Difference (IV, Random, 95% CI)	-0.63 [-1.04, -0.22]
3.1.2 Outpatient	2	229	Std. Mean Difference (IV, Random, 95% CI)	-0.36 [-0.90, 0.18]
3.2 Croup score (change baseline - 12 hours) by inpatient	2	84	Std. Mean Difference (IV, Random, 95% CI)	-0.75 [-1.19, -0.30]
3.2.1 Inpatient	2	84	Std. Mean Difference (IV, Random, 95% CI)	-0.75 [-1.19, -0.30]
3.3 Return visits or (re)admissions or both by inpatient/outpatient	5	374	Risk Ratio (M-H, Random, 95% CI)	0.69 [0.40, 1.22]
3.3.1 Inpatient	2	95	Risk Ratio (M-H, Random, 95% CI)	0.64 [0.14, 2.79]
3.3.2 Outpatient	3	279	Risk Ratio (M-H, Random, 95% CI)	0.71 [0.38, 1.30]
3.4 Length of stay by inpatient/outpatient	2	184	Mean Difference (IV, Random, 95% CI)	-0.51 [-1.28, 0.25]
3.4.1 Inpatient	1	50	Mean Difference (IV, Random, 95% CI)	-1.00 [-1.93, -0.07]
3.4.2 Outpatient	1	134	Mean Difference (IV, Random, 95% CI)	-0.20 [-0.78, 0.38]
3.5 Improvement (at 6 hours) 1 by outpatient		134	Risk Ratio (M-H, Random, 95% CI)	1.12 [0.93, 1.34]
3.5.1 Outpatient	1	134	Risk Ratio (M-H, Random, 95% CI)	1.12 [0.93, 1.34]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
3.6 Additional treatments: epinephrine	4	321	Risk Ratio (M-H, Random, 95% CI)	0.45 [0.21, 0.96]
3.7 Additional treatments: intubation/tracheostomy	2	145	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.04, 0.04]
3.8 Additional treatments: supplemental glucocorticoids	3	240	Risk Ratio (M-H, Random, 95% CI)	0.48 [0.18, 1.32]

Analysis 3.1. Comparison 3: Dexamethasone compared to budesonide, Outcome 1: Croup score (change baseline - 6 hours) by inpatient/outpatient

	Dexa	amethaso	1e	Bı	ıdesonide			Std. Mean Difference	Std. Mean Difference		
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI		
3.1.1 Inpatient											
Geelhoed 1995c	-2.82	1.07	23	-2.25	1.18	27	20.3%	-0.50 [-1.06, 0.07]			
Vad Pedersen 1998	-3.67	0.87	24	-2.93	1.01	23	19.0%	-0.77 [-1.37 , -0.18]			
Subtotal (95% CI)			47			50	39.3%	-0.63 [-1.04, -0.22]			
Heterogeneity: Tau ² = 0	.00; Chi ² = 0.	44, df = 1	(P = 0.51)	; I ² = 0%					•		
Test for overall effect: 2	Z = 3.00 (P = 0.00)	0.003)									
3.1.2 Outpatient											
Johnson 1998	-2.9	1.37	47	-2	1.39	48	28.0%	-0.65 [-1.06, -0.23]			
Klassen 1998	-2.4	0.97	69	-2.3	1.14	65	32.7%	-0.09 [-0.43, 0.24]			
Subtotal (95% CI)			116			113	60.7%	-0.36 [-0.90 , 0.18]			
Heterogeneity: Tau ² = 0	.12; Chi ² = 4.	11, df = 1	(P = 0.04)	; I ² = 76%							
Test for overall effect: 2	Z = 1.29 (P = 0)	0.20)									
Total (95% CI)			163			163	100.0%	-0.46 [-0.79 , -0.13]			
Heterogeneity: Tau ² = 0	.06; Chi ² = 6.	12, df = 3	(P = 0.11)	; I ² = 51%					~		
Test for overall effect: 2	Z = 2.73 (P =	0.006)							-2 -1 0 1 2		
Test for subgroup differ	ences: Chi ² =	0.61, df =	1 (P = 0.4)	4), I ² = 0%				Favours	s dexamethasone Favours budeso		

Analysis 3.2. Comparison 3: Dexamethasone compared to budesonide, Outcome 2: Croup score (change baseline - 12 hours) by inpatient

	Dexam		ne	Bı	ıdesonide			Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random	, 95% CI
3.2.1 Inpatient										_
Geelhoed 1995c	-2.91	0.86	23	-2.33	0.93	27	61.1%	-0.64 [-1.21 , -0.06]	-	
Vad Pedersen 1998	-3.91	0.86	19	-3.07	0.93	15	38.9%	-0.92 [-1.64 , -0.20]		
Subtotal (95% CI)			42			42	100.0%	-0.75 [-1.19 , -0.30]	•	
Heterogeneity: Tau ² = 0	.00; Chi ² = 0.	37, df = 1	(P = 0.54)	$I^2 = 0\%$					•	
Test for overall effect: Z	Z = 3.28 (P = 0.00)	0.001)								
Total (95% CI)			42			42	100.0%	-0.75 [-1.19 , -0.30]	•	
Heterogeneity: Tau ² = 0	.00; Chi ² = 0.	37, df = 1	(P = 0.54)	$I^2 = 0\%$					•	
Test for overall effect: Z	Z = 3.28 (P =	0.001)							-4 -2 0	2 4
Test for subgroup differences: Not applicable								Favour	s dexamethasone	Favours budesonide



Analysis 3.3. Comparison 3: Dexamethasone compared to budesonide, Outcome 3: Return visits or (re)admissions or both by inpatient/outpatient

	Dexamet	hasone	Budes	onide		Risk Ratio	Risk I	Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	m, 95% CI
3.3.1 Inpatient								
Geelhoed 1995c	2	18	3	21	11.3%	0.78 [0.15 , 4.15]		
Vad Pedersen 1998	0	29	1	27	3.2%	0.31 [0.01, 7.33]		
Subtotal (95% CI)		47		48	14.4%	0.64 [0.14, 2.79]		-
Total events:	2		4					
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0.00	.26, df = 1	(P = 0.61);	$I^2 = 0\%$				
Test for overall effect: 2	Z = 0.60 (P =	0.55)						
3.3.2 Outpatient								
Duman 2005	2	31	0	19	3.6%	3.13 [0.16, 61.80]		
Johnson 1998	11	47	18	48	78.9%	0.62 [0.33 , 1.18]	-	
Klassen 1998	1	69	0	65	3.1%	2.83 [0.12, 68.22]		
Subtotal (95% CI)		147		132	85.6%	0.71 [0.38, 1.30]		•
Total events:	14		18					
Heterogeneity: Tau ² = 0	0.00; Chi ² = 1.	.90, df = 2	(P = 0.39);	$I^2 = 0\%$				
Test for overall effect: 2	Z = 1.13 (P =	0.26)						
Total (95% CI)		194		180	100.0%	0.69 [0.40 , 1.22]		•
Total events:	16		22					
Heterogeneity: Tau ² = 0	0.00; Chi ² = 2.	.14, df = 4	(P = 0.71);	$I^2 = 0\%$			0.01 0.1 1	10 100
Test for overall effect: 2	Z = 1.27 (P =	0.20)					rs dexamethasone	Favours budesoni
Test for subgroup differ	rences: Chi² =	0.02. df =	1 (P = 0.90)), $I^2 = 0\%$				

Analysis 3.4. Comparison 3: Dexamethasone compared to budesonide, Outcome 4: Length of stay by inpatient/outpatient

	Dexa	amethaso	ne	Bı	ıdesonide			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
3.4.1 Inpatient									
Geelhoed 1995c	12	1.35	23	13	1.98	27	39.2%	-1.00 [-1.93, -0.07]	
Subtotal (95% CI)			23			27	39.2%	-1.00 [-1.93 , -0.07]	
Heterogeneity: Not app	licable								
Test for overall effect: Z	Z = 2.11 (P = 0)	0.03)							
3.4.2 Outpatient									
Klassen 1998	2.13	1.35	69	2.33	1.98	65	60.8%	-0.20 [-0.78, 0.38]	
Subtotal (95% CI)			69			65	60.8%	-0.20 [-0.78, 0.38]	
Heterogeneity: Not app	licable								
Test for overall effect: Z	Z = 0.68 (P =	0.50)							
Total (95% CI)			92			92	100.0%	-0.51 [-1.28 , 0.25]	
Heterogeneity: Tau ² = 0	.16; Chi ² = 2.	06, df = 1	(P = 0.15)	; I ² = 51%					•
Test for overall effect: Z	Z = 1.32 (P =	0.19)							-2 -1 0 1 2
Test for subgroup differ	ences: Chi ² =	2.06, df =	1 (P = 0.1	.5), I ² = 51.4	4%			Favours	s dexamethasone Favours budeso



Analysis 3.5. Comparison 3: Dexamethasone compared to budesonide, Outcome 5: Improvement (at 6 hours) by outpatient

	Dexamet	hasone	Budeso	onide		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
3.5.1 Outpatient							
Klassen 1998	57	69	48	65	100.0%	1.12 [0.93 , 1.34]	
Subtotal (95% CI)		69		65	100.0%	1.12 [0.93 , 1.34]	
Total events:	57		48				
Heterogeneity: Not app	licable						
Test for overall effect: 2	Z = 1.22 (P = 0)	0.22)					
Total (95% CI)		69		65	100.0%	1.12 [0.93 , 1.34]	
Total events:	57		48				
Heterogeneity: Not app	licable						0.7 0.85 1 1.2 1.5
Test for overall effect: 2	Z = 1.22 (P = 0)	0.22)				Favo	ours budesonide Favours dexamethasone
Test for subgroup differ	ences: Not ap	plicable					

Analysis 3.6. Comparison 3: Dexamethasone compared to budesonide, Outcome 6: Additional treatments: epinephrine

	Dexamet	hasone	Budeso	onide		Risk Ratio	Risk Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI	
Duman 2005	1	31	2	19	10.7%	0.31 [0.03, 3.15]		
Geelhoed 1995c	2	21	6	21	26.4%	0.33 [0.08, 1.47]		
Johnson 1998	4	47	9	48	47.3%	0.45 [0.15, 1.37]		
Klassen 1998	2	69	2	65	15.6%	0.94 [0.14 , 6.49]		
Total (95% CI)		168		153	100.0%	0.45 [0.21, 0.96]		
Total events:	9		19				•	
Heterogeneity: Tau ² = 0	.00; Chi ² = 0.	82, df = 3	(P = 0.84);	$I^2 = 0\%$		0.0	1 0.1 1 10 100)
Test for overall effect: 2	Z = 2.06 (P = 0)	0.04)					examethasone Favours budeson	
Test for subgroup differ	ences: Not ap	plicable						

Analysis 3.7. Comparison 3: Dexamethasone compared to budesonide, Outcome 7: Additional treatments: intubation/tracheostomy

	Dexamet	hasone	Budeso	onide		Risk Difference	Risk Diff	erence
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Randoi	m, 95% CI
Geelhoed 1995c	0	23	0	27	22.2%	0.00 [-0.08, 0.08]		
Johnson 1998	0	47	0	48	77.8%	0.00 [-0.04 , 0.04]		
Total (95% CI)		70		75	100.0%	0.00 [-0.04, 0.04]		-
Total events:	0		0				T	
Heterogeneity: Tau ² = 0	.00; Chi ² = 0.	00, df = 1	(P = 1.00);	$I^2 = 0\%$			-0.1 -0.05 0	0.05 0.1
Test for overall effect: 2	Z = 0.00 (P = 1)	1.00)				Favour	s dexamethasone	Favours budesonide
Test for subgroup differ	ences: Not ap	plicable						



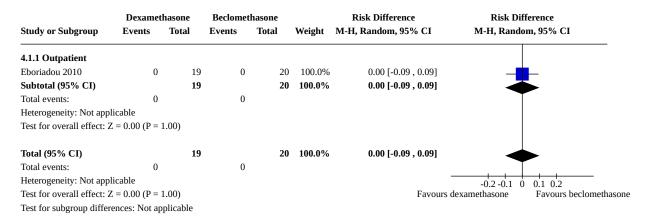
Analysis 3.8. Comparison 3: Dexamethasone compared to budesonide, Outcome 8: Additional treatments: supplemental glucocorticoids

	Dexamet	hasone	Budeso	onide		Risk Ratio	Risk R	atio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Randor	n, 95% CI
Geelhoed 1995c	2	23	2	27	28.8%	1.17 [0.18 , 7.69]		
Klassen 1998	3	69	7	65	59.3%	0.40 [0.11, 1.50]		
Vad Pedersen 1998	0	29	3	27	11.9%	0.13 [0.01, 2.47]	· -	_
Total (95% CI)		121		119	100.0%	0.48 [0.18 , 1.32]		
Total events:	5		12					
Heterogeneity: Tau ² = 0	0.00; Chi ² = 1.	71, df = 2	(P = 0.43);	$I^2 = 0\%$		(0.01 0.1 1	10 100
Test for overall effect: 2	Z = 1.42 (P =	0.15)					s dexamethasone	Favours budesonide
Test for subgroup differ	rences: Not ap	plicable						

Comparison 4. Dexamethasone compared to beclomethasone

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
4.1 Return visits or (re)admissions or both by outpatient	1	39	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.09, 0.09]
4.1.1 Outpatient	1	39	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.09, 0.09]

Analysis 4.1. Comparison 4: Dexamethasone compared to beclomethasone, Outcome 1: Return visits or (re)admissions or both by outpatient



Comparison 5. Dexamethasone compared to betamethasone

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
5.1 Croup score (change baseline - 2 hours) by outpatient	1	52	Std. Mean Difference (IV, Random, 95% CI)	-0.62 [-1.17, -0.06]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
5.1.1 Outpatient	1	52	Std. Mean Difference (IV, Random, 95% CI)	-0.62 [-1.17, -0.06]
5.2 Croup score (change baseline - 6 hours) by outpatient	1	52	Std. Mean Difference (IV, Random, 95% CI)	-0.67 [-1.23, -0.11]
5.2.1 Outpatient	1	52	Std. Mean Difference (IV, Random, 95% CI)	-0.67 [-1.23, -0.11]
5.3 Return visits or (re)admissions or both by outpatient	1	52	Risk Ratio (M-H, Random, 95% CI)	0.95 [0.67, 1.34]
5.3.1 Outpatient	1	52	Risk Ratio (M-H, Random, 95% CI)	0.95 [0.67, 1.34]
5.4 Additional treatments: epinephrine	1	52	Risk Ratio (M-H, Random, 95% CI)	2.11 [1.18, 3.76]

Analysis 5.1. Comparison 5: Dexamethasone compared to betamethasone, Outcome 1: Croup score (change baseline - 2 hours) by outpatient

	Dexa	amethaso	ne	Beta	amethasor	1e		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
5.1.1 Outpatient									
Amir 2006	-3.06	2.27	26	-1.68	2.14	26	100.0%	-0.62 [-1.17 , -0.06]	
Subtotal (95% CI)			26			26	100.0%	-0.62 [-1.17 , -0.06]	•
Heterogeneity: Not app	licable								•
Test for overall effect: 2	Z = 2.17 (P = 0)	0.03)							
Total (95% CI)			26			26	100.0%	-0.62 [-1.17 , -0.06]	
Heterogeneity: Not app	licable								•
Test for overall effect: 2	Z = 2.17 (P = 0)	0.03)							-2 -1 0 1 2
Test for subgroup differ	ences: Not ap	plicable						Favours	s dexamethasone Favours betametha

Analysis 5.2. Comparison 5: Dexamethasone compared to betamethasone, Outcome 2: Croup score (change baseline - 6 hours) by outpatient

	Dexa	amethaso	ne	Beta	ımethasoı	1e		Std. Mean Difference	Std. Mean I	Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Randon	ı, 95% CI
5.2.1 Outpatient										
Amir 2006	-3.42	2.26	26	-1.89	2.23	26	100.0%	-0.67 [-1.23 , -0.11]	-	
Subtotal (95% CI)			26			26	100.0%	-0.67 [-1.23 , -0.11]	_	
Heterogeneity: Not app	licable								•	
Test for overall effect: 2	Z = 2.35 (P = 0)	0.02)								
Total (95% CI)			26			26	100.0%	-0.67 [-1.23 , -0.11]		
Heterogeneity: Not app	licable								•	
Test for overall effect: 2	Z = 2.35 (P = 0)	0.02)							-4 -2 0	2 4
Test for subgroup differ	ences: Not ap	plicable						Favour	rs dexamethasone	Favours betamethasone



Analysis 5.3. Comparison 5: Dexamethasone compared to betamethasone, Outcome 3: Return visits or (re)admissions or both by outpatient

	Dexamet	hasone	Betamet	hasone		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
5.3.1 Outpatient							
Amir 2006	18	26	19	26	100.0%	0.95 [0.67, 1.34]	_
Subtotal (95% CI)		26		26	100.0%	0.95 [0.67, 1.34]	
Total events:	18		19				
Heterogeneity: Not appl	licable						
Test for overall effect: Z	Z = 0.31 (P = 0)).76)					
Total (95% CI)		26		26	100.0%	0.95 [0.67 , 1.34]	
Total events:	18		19				\top
Heterogeneity: Not appl	licable					(0.2 0.5 1 2 5
Test for overall effect: Z	Z = 0.31 (P = 0.31)	0.76)					dexamethasone Favours betamethasone
Test for subgroup differ	ences: Not ap	plicable					

Analysis 5.4. Comparison 5: Dexamethasone compared to betamethasone, Outcome 4: Additional treatments: epinephrine

Study or Subgroup	Dexamet Events	hasone Total	Betamet Events	hasone Total	Weight	Risk Ratio M-H, Random, 95% CI	Risk Ratio M-H, Random, 95% CI
Amir 2006	19	26	9	26	100.0%	2.11 [1.18 , 3.76]	-
Total (95% CI)	10	26	0	26	100.0%	2.11 [1.18 , 3.76]	•
Total events: Heterogeneity: Not appl Test for overall effect: 2		0.01)	9			0.01 Favours de	0.1 1 10 100 examethasone Favours betamethasone
Test for subgroup differ	ences: Not ap	plicable					

Comparison 6. Dexamethasone compared to prednisolone

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
6.1 Croup score (change baseline - 2 hours) by outpatient	1	1231	Std. Mean Difference (IV, Random, 95% CI)	0.06 [-0.06, 0.18]
6.1.1 Outpatient	1	1231	Std. Mean Difference (IV, Random, 95% CI)	0.06 [-0.06, 0.18]
6.2 Croup score (change baseline - 6 hours) by outpatient	1	99	Std. Mean Difference (IV, Random, 95% CI)	0.21 [-0.21, 0.62]
6.2.1 Outpatient	1	99	Std. Mean Difference (IV, Random, 95% CI)	0.21 [-0.21, 0.62]
6.3 Return visits or (re)admissions or both by outpatient	4	1537	Risk Ratio (M-H, Random, 95% CI)	0.55 [0.28, 1.11]
6.3.1 Outpatient	4	1537	Risk Ratio (M-H, Random, 95% CI)	0.55 [0.28, 1.11]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
6.4 Length of stay by outpatient	2	1363	Mean Difference (IV, Random, 95% CI)	-0.02 [-0.42, 0.39]
6.4.1 Outpatients	2	1363	Mean Difference (IV, Random, 95% CI)	-0.02 [-0.42, 0.39]
6.5 Additional treatments: epinephrine	3	1463	Risk Ratio (M-H, Random, 95% CI)	0.90 [0.50, 1.64]
6.6 Additional treatments: intubation/tracheotomy	1	1231	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.00, 0.00]
6.7 Additional treatments: supplemental glucocorticoids	2	926	Risk Ratio (M-H, Random, 95% CI)	0.72 [0.53, 0.97]

Analysis 6.1. Comparison 6: Dexamethasone compared to prednisolone, Outcome 1: Croup score (change baseline - 2 hours) by outpatient

	Dexa	amethaso	ne	Pro	ednisolon	e		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
6.1.1 Outpatient									
Parker 2019	-0.81	1.36	820	-0.89	1.27	411	100.0%	0.06 [-0.06, 0.18]	
Subtotal (95% CI)			820			411	100.0%	0.06 [-0.06, 0.18]	<u> </u>
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.99 (P = 0.00)	0.32)							
Total (95% CI)			820			411	100.0%	0.06 [-0.06 , 0.18]	
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.99 (P =	0.32)							-0.5 -0.25 0 0.25 0.5
Test for subgroup differ	ences: Not ap	plicable						Favour	rs dexamethasone Favours prednisolon

Analysis 6.2. Comparison 6: Dexamethasone compared to prednisolone, Outcome 2: Croup score (change baseline - 6 hours) by outpatient

	Dexa	amethaso	ne	Pr	ednisolon	e		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
6.2.1 Outpatient									
Fifoot 2007	-2.16	0.97	65	-2.35	0.81	34	100.0%	0.21 [-0.21, 0.62]	
Subtotal (95% CI)			65			34	100.0%	0.21 [-0.21, 0.62]	
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.97 (P =	0.33)							
Total (95% CI)			65			34	100.0%	0.21 [-0.21 , 0.62]	
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.97 (P = 0.97)	0.33)						-	-1 -0.5 0 0.5 1
Test for subgroup differ	rences: Not ap	plicable						Favours d	lexamethasone Favours prednisolone



Test for subgroup differences: Not applicable

Analysis 6.3. Comparison 6: Dexamethasone compared to prednisolone, Outcome 3: Return visits or (re)admissions or both by outpatient

	Dexamet	hasone	Prediso	olone		Risk Ratio	Risk Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 9	5% CI
6.3.1 Outpatient								
Fifoot 2007	7	57	5	29	22.3%	0.71 [0.25, 2.05]		
Garbutt 2013	1	46	3	41	8.0%	0.30 [0.03, 2.75]		
Parker 2019	153	820	89	411	44.3%	0.86 [0.68, 1.09]	_	
Sparrow 2006	5	68	19	65	25.4%	0.25 [0.10, 0.63]		
Subtotal (95% CI)		991		546	100.0%	0.55 [0.28, 1.11]		
Total events:	166		116					
Heterogeneity: Tau ² = 0	0.27; Chi ² = 7.	29, df = 3 ((P = 0.06); 1	$1^2 = 59\%$				
Test for overall effect: 2	Z = 1.67 (P = 0)	0.09)						
Total (95% CI)		991		546	100.0%	0.55 [0.28 , 1.11]		
Total events:	166		116				~	
Heterogeneity: Tau ² = 0	0.27; Chi ² = 7.	29, df = 3 ((P = 0.06); 1	$[^2 = 59\%]$			0.02 0.1 1	10 50
Test for overall effect: 2	Z = 1.67 (P = 0)	0.09)				Favour	****	vours prednisolone

Analysis 6.4. Comparison 6: Dexamethasone compared to prednisolone, Outcome 4: Length of stay by outpatient

	Dexa	amethaso	ne	Pr	ednisolone	2		Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
6.4.1 Outpatients									
Parker 2019	2	3.2	820	2.1	2.6	410	86.1%	-0.10 [-0.43, 0.23]	
Sparrow 2006	1.9	4	68	1.4	1.87	65	13.9%	0.50 [-0.55 , 1.55]	
Subtotal (95% CI)			888			475	100.0%	-0.02 [-0.42, 0.39]	
Heterogeneity: Tau ² = 0	0.02; Chi ² = 1.	13, df = 1	(P = 0.29)	; I ² = 12%					\perp
Test for overall effect: 2	Z = 0.08 (P = 0.08)	0.94)							
Total (95% CI)			888			475	100.0%	-0.02 [-0.42 , 0.39]	
Heterogeneity: Tau ² = 0	0.02; Chi ² = 1.	13, df = 1	(P = 0.29)	; I ² = 12%					\top
Test for overall effect: 2	Z = 0.08 (P = 0.08)	0.94)							-1 -0.5 0 0.5 1
Test for subgroup differ	rences: Not ap	plicable						Favours	dexamethasone Favours prednisol

Analysis 6.5. Comparison 6: Dexamethasone compared to prednisolone, Outcome 5: Additional treatments: epinephrine

	Dexamet	hasone	Prednis	olone		Risk Ratio	Risk R	atio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	m, 95% CI
Fifoot 2007	2	65	3	34	11.6%	0.35 [0.06 , 1.99]		_
Parker 2019	21	820	10	411	63.6%	1.05 [0.50 , 2.21]	-	-
Sparrow 2006	5	68	5	65	24.8%	0.96 [0.29 , 3.15]	-	<u>-</u>
Total (95% CI)		953		510	100.0%	0.90 [0.50 , 1.64]		
Total events:	28		18				Ť	
Heterogeneity: Tau ² = 0	.00; Chi ² = 1.	32, df = 2	(P = 0.52);	$I^2 = 0\%$		0.0	02 0.1 1	10 500
Test for overall effect: 2	Z = 0.33 (P = 0.33)	0.74)				Favours de	examethasone	Favours prednisolone
Test for subgroup differ	ences: Not ap	plicable						



Analysis 6.6. Comparison 6: Dexamethasone compared to prednisolone, Outcome 6: Additional treatments: intubation/tracheotomy

	Dexamet	hasone	Prednis	olone		Risk Difference	Risk Difference	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI	
Parker 2019	0	820	0	411	100.0%	0.00 [-0.00 , 0.00]	-	
Total (95% CI)		820		411	100.0%	0.00 [-0.00, 0.00]		
Total events:	0		0					
Heterogeneity: Not appl	icable					-0.01	-0.005 0 0.005 0.	1 01
Test for overall effect: Z	L = 0.00 (P = 1)	1.00)				Favours dex	amethasone Favours predni	
Test for subgroup differ	ences: Not ap	plicable						

Analysis 6.7. Comparison 6: Dexamethasone compared to prednisolone, Outcome 7: Additional treatments: supplemental glucocorticoids

	Dexamet	hasone	Prednis	olone		Risk Ratio	Risk R	atio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	n, 95% CI
Fifoot 2007	9	57	5	29	9.4%	0.92 [0.34 , 2.48]		
Parker 2019	74	560	53	280	90.6%	0.70 [0.51, 0.96]	-	
Total (95% CI)		617		309	100.0%	0.72 [0.53, 0.97]		
Total events:	83		58				•	
Heterogeneity: Tau ² = 0	0.00; Chi ² = $0.$	26, df = 1	(P = 0.61);	$I^2 = 0\%$		+ 0.:	2 0.5 1	2 5
Test for overall effect:	Z = 2.13 (P = 0)	0.03)					examethasone	Favours prednisolone
Test for subgroup differ	rences. Not an	nlicable						

Test for subgroup differences: Not applicable

Comparison 7. Budesonide and dexamethasone compared to dexamethasone

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
7.1 Croup score (change baseline - 6 hours) by inpatient/outpatient	3	255	Std. Mean Difference (IV, Random, 95% CI)	0.05 [-0.19, 0.30]
7.1.1 Inpatient	1	72	Std. Mean Difference (IV, Random, 95% CI)	0.16 [-0.30, 0.63]
7.1.2 Outpatient	2	183	Std. Mean Difference (IV, Random, 95% CI)	0.03 [-0.32, 0.39]
7.2 Return visits or (re)admissions or both by inpatient/outpatient	3	254	Risk Ratio (M-H, Random, 95% CI)	0.91 [0.45, 1.83]
7.2.1 Inpatient	1	71	Risk Ratio (M-H, Random, 95% CI)	1.03 [0.46, 2.29]
7.2.2 Outpatient	2	183	Risk Ratio (M-H, Random, 95% CI)	0.58 [0.13, 2.60]
7.3 Length of stay by inpatient/outpatient	2	204	Mean Difference (IV, Random, 95% CI)	0.44 [-0.05, 0.92]



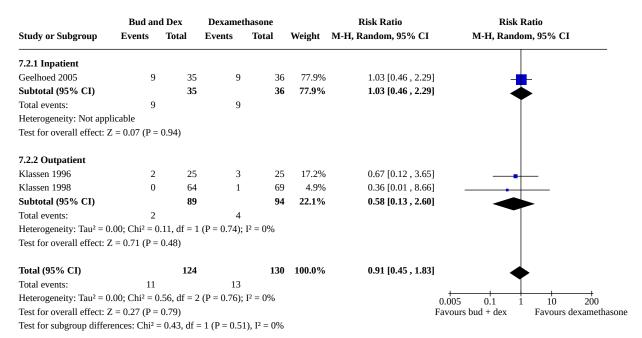
Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size	
7.3.1 Inpatient	1	71	Mean Difference (IV, Random, 95% CI)	-1.30 [-6.75, 4.15]	
7.3.2 Outpatient	1	133	Mean Difference (IV, Random, 95% CI)	0.45 [-0.04, 0.94]	
7.4 Improvement (at 6 hours) by outpatient	2	183	Risk Ratio (M-H, Random, 95% CI)	1.11 [0.65, 1.90]	
7.4.1 Outpatient	2	183	Risk Ratio (M-H, Random, 95% CI)	1.11 [0.65, 1.90]	
7.5 Additional treatments: epinephrine	2	183	Risk Ratio (M-H, Random, 95% CI)	1.42 [0.27, 7.39]	
7.6 Additional treatments: mist tent	1	50	Risk Ratio (M-H, Random, 95% CI)	1.07 [0.69, 1.65]	
7.7 Additional treatments: supplemental glucocorticoids	2	182	Risk Ratio (M-H, Random, 95% CI)	1.10 [0.07, 16.66]	

Analysis 7.1. Comparison 7: Budesonide and dexamethasone compared to dexamethasone, Outcome 1: Croup score (change baseline - 6 hours) by inpatient/outpatient

	Bu	d and De	x	Dex	amethaso	ne		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
7.1.1 Inpatient									
Geelhoed 2005	-3.04	1.22	36	-3.24	1.19	36	28.3%	0.16 [-0.30, 0.63]	l
Subtotal (95% CI)			36			36	28.3%	0.16 [-0.30, 0.63]	
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.70 (P =	0.49)							
7.1.2 Outpatient									
Klassen 1996	-1.4	1.3	25	-1.8	1.4	25	19.5%	0.29 [-0.27, 0.85]	1
Klassen 1998	-2.5	1.13	64	-2.4	0.97	69	52.3%	-0.09 [-0.44, 0.25]	l <u>-</u>
Subtotal (95% CI)			89			94	71.7%	0.03 [-0.32, 0.39]	
Heterogeneity: Tau ² = 0	0.02; Chi ² = 1.	34, df = 1	(P = 0.25)	; I ² = 26%					T
Test for overall effect: 2	Z = 0.18 (P =	0.86)							
Total (95% CI)			125			130	100.0%	0.05 [-0.19 , 0.30]	ı 📥
Heterogeneity: Tau ² = 0	0.00; Chi ² = 1.	65, df = 2	(P = 0.44)	; I ² = 0%					
Test for overall effect: 2	Z = 0.43 (P =	0.67)							-1 -0.5 0 0.5 1
Test for subgroup differ	ences: Chi ² =	0.20 df =	= 1 (P = 0.6)	$I^2 = 0\%$					Favours bud + dex Favours dexamethase



Analysis 7.2. Comparison 7: Budesonide and dexamethasone compared to dexamethasone, Outcome 2: Return visits or (re)admissions or both by inpatient/outpatient



Analysis 7.3. Comparison 7: Budesonide and dexamethasone compared to dexamethasone, Outcome 3: Length of stay by inpatient/outpatient

	Bud and Dex			Dexamethasone			Mean Difference		Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
7.3.1 Inpatient										
Geelhoed 2005	11.1	11.9	35	12.4	11.5	36	0.8%	-1.30 [-6.75 , 4.15]		
Subtotal (95% CI)			35			36	0.8%	-1.30 [-6.75 , 4.15]		
Heterogeneity: Not app	licable									
Test for overall effect: 2	Z = 0.47 (P = 0.47)	0.64)								
7.3.2 Outpatient										
Klassen 1998	2.58	1.51	64	2.13	1.35	69	99.2%	0.45 [-0.04, 0.94]	—	
Subtotal (95% CI)			64			69	99.2%	0.45 [-0.04, 0.94]	~	
Heterogeneity: Not app	licable								Y	
Test for overall effect: 2	Z = 1.81 (P = 0)	0.07)								
Total (95% CI)			99			105	100.0%	0.44 [-0.05 , 0.92]	•	
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0.	39, df = 1	(P = 0.53)	; I ² = 0%					Y	
Test for overall effect: 2	Z = 1.76 (P =	0.08)							-4 -2 0 2 4	
Test for subgroup differences: $Chi^2 = 0.39$, $df = 1$ (P = 0.53), $I^2 = 0\%$								F	Favours bud + dex Favours de:	xamethaso



Analysis 7.4. Comparison 7: Budesonide and dexamethasone compared to dexamethasone, Outcome 4: Improvement (at 6 hours) by outpatient

	Bud an	d Dex	Dexamet	hasone		Risk Ratio	Ris	k Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Ran	dom, 95% CI
7.4.1 Outpatient								
Klassen 1996	21	25	14	25	45.1%	1.50 [1.02, 2.21	l]	-
Klassen 1998	46	64	57	69	54.9%	0.87 [0.72 , 1.05	5]	
Subtotal (95% CI)		89		94	100.0%	1.11 [0.65 , 1.90)]	
Total events:	67		71					
Heterogeneity: Tau ² = 0	0.13; Chi ² = 6	.31, df = 1	(P = 0.01);	$I^2 = 84\%$				
Test for overall effect: 2	Z = 0.39 (P =	0.70)						
Total (95% CI)		89		94	100.0%	1.11 [0.65 , 1.90)]	
Total events:	67		71					
Heterogeneity: Tau ² = 0	0.13; Chi ² = 6	.31, df = 1	(P = 0.01);	$I^2 = 84\%$			0.01 0.1	1 10 100
Test for overall effect: 2	Z = 0.39 (P =	0.70)				Fave	ours dexamethasone	Favours bud + dex
Test for subgroup differ	ences: Not a	pplicable						

Analysis 7.5. Comparison 7: Budesonide and dexamethasone compared to dexamethasone, Outcome 5: Additional treatments: epinephrine

	Bud an	d Dex	Dexamet	hasone		Risk Ratio	Risk	Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	om, 95% CI
Klassen 1996	1	25	0	25	27.2%	3.00 [0.13 , 70.30]	
Klassen 1998	2	64	2	69	72.8%	1.08 [0.16 , 7.43] —	
Total (95% CI)		89		94	100.0%	1.42 [0.27 , 7.39	1	
Total events:	3		2					
Heterogeneity: Tau ² = 0	.00; Chi ² = 0	.30, df = 1	(P = 0.59);	$I^2 = 0\%$			0.01 0.1 1	10 100
Test for overall effect: Z	L = 0.42 (P =	0.67)					Favours bud + dex	Favours dexamethasone
Test for subgroup differ	ences: Not a _l	plicable						

Analysis 7.6. Comparison 7: Budesonide and dexamethasone compared to dexamethasone, Outcome 6: Additional treatments: mist tent

	Bud an	d Dex	Dexamet	hasone		Risk Ratio	Risk R	atio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Randor	n, 95% CI
Klassen 1996	16	25	15	25	100.0%	1.07 [0.69 , 1.65]	-	 }
Total (95% CI)		25		25	100.0%	1.07 [0.69 , 1.65]		•
Total events:	16		15				T	
Heterogeneity: Not app	olicable						0.1 0.2 0.5 1	2 5 10
Test for overall effect:	Z = 0.29 (P =	0.77)					Favours bud + dex	Favours dexamethasone
Test for subgroup diffe	rences: Not a	pplicable						



Analysis 7.7. Comparison 7: Budesonide and dexamethasone compared to dexamethasone, Outcome 7: Additional treatments: supplemental glucocorticoids

	Bud an	d Dex	Dexamet	hasone		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Klassen 1996	0	25	2	24	38.3%	0.19 [0.01 , 3.81]	
Klassen 1998	9	64	3	69	61.7%	3.23 [0.92 , 11.42]	l
Total (95% CI)		89		93	100.0%	1.10 [0.07, 16.66]	
Total events:	9		5				
Heterogeneity: Tau ² = 2	2.71; Chi ² = 2	.98, df = 1	(P = 0.08);	$I^2 = 66\%$			0.005 0.1 1 10 200
Test for overall effect:	Z = 0.07 (P =	0.95)					Favours bud + dex Favours dexamethason
Test for subgroup diffe	rences: Not a	pplicable					

Comparison 8. Budesonide and dexamethasone compared to budesonide

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
8.1 Croup score (change baseline - 6 hours) by outpatient	1	129	Std. Mean Difference (IV, Random, 95% CI)	-0.18 [-0.52, 0.17]
8.1.1 Outpatient	1	129	Std. Mean Difference (IV, Random, 95% CI)	-0.18 [-0.52, 0.17]
8.2 Return visits or (re)admissions or both by outpatient	1	129	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.03, 0.03]
8.2.1 Outpatient	1	129	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.03, 0.03]
8.3 Length of stay by outpatient	1	129	Mean Difference (IV, Random, 95% CI)	0.25 [-0.36, 0.86]
8.3.1 Outpatient	1	129	Mean Difference (IV, Random, 95% CI)	0.25 [-0.36, 0.86]
8.4 Improvement (at 6 hours) by outpatient	1	129	Risk Ratio (M-H, Random, 95% CI)	0.97 [0.79, 1.20]
8.4.1 Outpatient	1	129	Risk Ratio (M-H, Random, 95% CI)	0.97 [0.79, 1.20]
8.5 Additional treatments: epinephrine	1	129	Risk Ratio (M-H, Random, 95% CI)	1.02 [0.15, 6.99]
8.6 Additional treatments: supplemental glucocorticoids	1	129	Risk Ratio (M-H, Random, 95% CI)	1.31 [0.52, 3.29]



Analysis 8.1. Comparison 8: Budesonide and dexamethasone compared to budesonide, Outcome 1: Croup score (change baseline - 6 hours) by outpatient

	Bu	Bud and Dex			Budesonide			Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
8.1.1 Outpatient									
Klassen 1998	-2.5	1.13	64	-2.3	1.14	65	100.0%	-0.18 [-0.52 , 0.17	′1 —
Subtotal (95% CI)			64			65	100.0%	-0.18 [-0.52 , 0.17	7]
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.99 (P = 0.00)	0.32)							
Total (95% CI)			64			65	100.0%	-0.18 [-0.52 , 0.17	7]
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.99 (P = 0.00)	0.32)							-1 -0.5 0 0.5 1
Test for subgroup differ	rences: Not ap	plicable							Favours bud + dex Favours budeson

Analysis 8.2. Comparison 8: Budesonide and dexamethasone compared to budesonide, Outcome 2: Return visits or (re)admissions or both by outpatient

	Bud an	d Dex	Budes	onide		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
8.2.1 Outpatient							
Klassen 1998	0	64	0	65	100.0%	0.00 [-0.03, 0.03]]
Subtotal (95% CI)		64		65	100.0%	0.00 [-0.03, 0.03	
Total events:	0		0				\top
Heterogeneity: Not app	licable						
Test for overall effect: 2	Z = 0.00 (P =	1.00)					
Total (95% CI)		64		65	100.0%	0.00 [-0.03 , 0.03	
Total events:	0		0				\top
Heterogeneity: Not app	licable						$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
Test for overall effect: 2	Z = 0.00 (P =	1.00)					Favours bud + dex Favours budesonide
Test for subgroup differ	rences: Not a	pplicable					

Analysis 8.3. Comparison 8: Budesonide and dexamethasone compared to budesonide, Outcome 3: Length of stay by outpatient

	Bu	d and Dex	(Bı	udesonide			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
8.3.1 Outpatient									
Klassen 1998	2.58	1.51	64	2.33	1.98	65	100.0%	0.25 [-0.36 , 0.86	<u> </u>
Subtotal (95% CI)			64			65	100.0%	0.25 [-0.36, 0.86	5]
Heterogeneity: Not app	licable								
Test for overall effect: Z	Z = 0.81 (P =	0.42)							
Total (95% CI)			64			65	100.0%	0.25 [-0.36 , 0.86	5]
Heterogeneity: Not app	licable								
Test for overall effect: 2	Z = 0.81 (P =	0.42)							-2 -1 0 1 2
Test for subgroup differ	ences: Not ap	plicable							Favours bud + dex Favours budesonide



Analysis 8.4. Comparison 8: Budesonide and dexamethasone compared to budesonide, Outcome 4: Improvement (at 6 hours) by outpatient

	Bud an	d Dex	Budeso	onide		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
8.4.1 Outpatient							
Klassen 1998	46	64	48	65	100.0%	0.97 [0.79, 1.20]	_
Subtotal (95% CI)		64		65	100.0%	0.97 [0.79, 1.20]	
Total events:	46		48				
Heterogeneity: Not appl	licable						
Test for overall effect: Z	Z = 0.25 (P =	0.80)					
Total (95% CI)		64		65	100.0%	0.97 [0.79 , 1.20]	
Total events:	46		48				\top
Heterogeneity: Not appl	licable						0.5 0.7 1 1.5 2
Test for overall effect: Z	Z = 0.25 (P =	0.80)				Fav	ours budesonide Favours bud + dex
Test for subgroup differ	ences: Not aj	pplicable					

Analysis 8.5. Comparison 8: Budesonide and dexamethasone compared to budesonide, Outcome 5: Additional treatments: epinephrine

	Bud an	d Dex	Budes	onide		Risk Ratio	Risk l	Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	om, 95% CI
Klassen 1998	2	64	2	65	100.0%	1.02 [0.15 , 6.99] —	<u> </u>
Total (95% CI)		64		65	100.0%	1.02 [0.15 , 6.99]	
Total events:	2		2					
Heterogeneity: Not app	licable						0.005 0.1 1	10 200
Test for overall effect: 2	Z = 0.02 (P =	0.99)					Favours bud + dex	Favours budenoside
Test for subgroup differ	rences: Not a	pplicable						

Analysis 8.6. Comparison 8: Budesonide and dexamethasone compared to budesonide, Outcome 6: Additional treatments: supplemental glucocorticoids

	Bud an	d Dex	Budes	onide		Risk Ratio	Risk R	latio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	m, 95% CI
Klassen 1998	9	64	7	65	100.0%	1.31 [0.52 , 3.29] -	<u> </u>
Total (95% CI)		64		65	100.0%	1.31 [0.52 , 3.29	1	•
Total events:	9		7					
Heterogeneity: Not app	olicable						0.01 0.1 1	10 100
Test for overall effect: 2	Z = 0.57 (P =	0.57)					Favours bud + dex	Favours budenoside
Test for subgroup differ	rences: Not a	pplicable						

Comparison 9. Oral compared to intramuscular dexamethasone

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
9.1 Return visits or (re)admissions or both by outpatient	3	440	Risk Ratio (M-H, Random, 95% CI)	0.81 [0.58, 1.12]



Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
9.1.1 Outpatient	3	440	Risk Ratio (M-H, Random, 95% CI)	0.81 [0.58, 1.12]
9.2 Improvement (at 24 hours) by outpatient	1	95	Risk Ratio (M-H, Random, 95% CI)	1.07 [0.95, 1.19]
9.2.1 Outpatient	1	95	Risk Ratio (M-H, Random, 95% CI)	1.07 [0.95, 1.19]
9.3 Additional treatments: antibiotics	1	277	Risk Ratio (M-H, Random, 95% CI)	0.14 [0.02, 1.15]
9.4 Additional treatments: epinephrine	2	372	Risk Ratio (M-H, Random, 95% CI)	0.94 [0.71, 1.24]
9.5 Additional treatments: mist tent	1	277	Risk Ratio (M-H, Random, 95% CI)	1.34 [0.31, 5.89]
9.6 Additional treatments: supplemental glucocorticoids	1	277	Risk Ratio (M-H, Random, 95% CI)	1.10 [0.50, 2.41]

Analysis 9.1. Comparison 9: Oral compared to intramuscular dexamethasone, Outcome 1: Return visits or (re)admissions or both by outpatient

	Oral	Dex	Intramusc	ular Dex		Risk Ratio	Risk Ratio
Study or Subgroup	Events Total		Events Total		Weight M-H, Random, 95% CI		M-H, Random, 95% CI
9.1.1 Outpatient							
Donaldson 2003	10	46	12	49	20.2%	0.89 [0.43 , 1.85]	
Rittichier 2000	35	138	45	139	78.4%	0.78 [0.54 , 1.14]	-
Soleimani 2013	1	32	1	36	1.5%	1.13 [0.07, 17.26]	
Subtotal (95% CI)		216		224	100.0%	0.81 [0.58, 1.12]	
Total events:	46		58				*
Heterogeneity: $Tau^2 = 0$.	.00; $Chi^2 = 0$.15, df = 2	$(P = 0.93); I^2$	$^{2} = 0\%$			
Test for overall effect: Z	= 1.27 (P =	0.21)					
Total (95% CI)		216		224	100.0%	0.81 [0.58 , 1.12]	
Total events:	46		58				_
Heterogeneity: $Tau^2 = 0$.	.00; $Chi^2 = 0$.15, df = 2	$(P = 0.93); I^2$	$^{2} = 0\%$			0.1 0.2 0.5 1 2 5 10
Test for overall effect: Z	= 1.27 (P =	0.21)					Favours oral dex Favours intramuscula
Test for subgroup differe	ences: Not a	pplicable					



Analysis 9.2. Comparison 9: Oral compared to intramuscular dexamethasone, Outcome 2: Improvement (at 24 hours) by outpatient

	Oral	Dex	Intramuscu	ılar Dex		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
9.2.1 Outpatient							
Donaldson 2003	44	46	44	49	100.0%	1.07 [0.95, 1.19]	
Subtotal (95% CI)		46		49	100.0%	1.07 [0.95, 1.19]	
Total events:	44		44				
Heterogeneity: Not appl	icable						
Test for overall effect: Z	Z = 1.10 (P =	0.27)					
Total (95% CI)		46		49	100.0%	1.07 [0.95, 1.19]	
Total events:	44		44				
Heterogeneity: Not appl	icable					•	0.7 0.85 1 1.2 1.5
Test for overall effect: Z	Z = 1.10 (P =	0.27)				Favours into	ramuscular dex Favours oral dex
Test for subgroup differ	ences: Not a	pplicable					

Analysis 9.3. Comparison 9: Oral compared to intramuscular dexamethasone, Outcome 3: Additional treatments: antibiotics

Study or Subgroup	Oral : Events	Dex Total	Intramuscu Events	ılar Dex Total	Weight	Risk Ratio M-H, Random, 95% CI	Risk Ratio M-H, Random, 95% CI	
Rittichier 2000	1	138	7	139	100.0%	0.14 [0.02 , 1.15]	-	_
Total (95% CI)		138		139	100.0%	0.14 [0.02 , 1.15]		
Total events:	1		7					
Heterogeneity: Not app	licable						0.005 0.1 1 10 200	
Test for overall effect: Z	Z = 1.83 (P =	0.07)					Favours oral dex Favours intramus	scular
Test for subgroup differences: Not applicable								

Analysis 9.4. Comparison 9: Oral compared to intramuscular dexamethasone, Outcome 4: Additional treatments: epinephrine

	Oral 1	Dex	Intramusc	ular Dex		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Donaldson 2003	25	46	28	49	60.4%	0.95 [0.66 , 1.36]	
Rittichier 2000	29	138	32	139	39.6%	0.91 [0.59 , 1.42]	_
Total (95% CI)		184		188	100.0%	0.94 [0.71 , 1.24]	•
Total events:	54		60				7
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0	.02, df = 1	$(P = 0.88); I^2$	$^{2} = 0\%$			0.2 0.5 1 2 5
Test for overall effect: $Z = 0.47$ ($P = 0.64$)							Favours oral dex Favours intramuscular dex
Test for subgroup differ	ences: Not a	oplicable					



Analysis 9.5. Comparison 9: Oral compared to intramuscular dexamethasone, Outcome 5: Additional treatments: mist tent

	Oral	Dex	Intramuscu	ılar Dex		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Rittichier 2000	4	138	3	139	100.0%	1.34 [0.31 , 5.89]	-
Total (95% CI)		138		139	100.0%	1.34 [0.31, 5.89]	
Total events:	4		3				
Heterogeneity: Not app	licable						0.005 0.1 1 10 200
Test for overall effect: $Z = 0.39$ ($P = 0.70$)						Favours oral dex Favours intramuscular de	
Test for subgroup differences: Not applicable							

Analysis 9.6. Comparison 9: Oral compared to intramuscular dexamethasone, Outcome 6: Additional treatments: supplemental glucocorticoids

	Oral	Dex	Intramusc	ular Dex		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Rittichier 2000	12	138	11	139	100.0%	1.10 [0.50 , 2.41]	•
Total (95% CI)		138		139	100.0%	1.10 [0.50, 2.41]	•
Total events:	12		11				T
Heterogeneity: Not app	olicable						0.005 0.1 1 10 200
Test for overall effect:	Z = 0.24 (P =	0.81)					Favours oral dex Favours intramuscular of
Test for subgroup diffe	rences: Not a	pplicable					

Comparison 10. Oral compared to nebulised dexamethasone

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
10.1 Return visits or (re)admissions or both by outpatient	1	176	Risk Ratio (M-H, Random, 95% CI)	0.39 [0.17, 0.89]
10.1.1 Outpatient	1	176	Risk Ratio (M-H, Random, 95% CI)	0.39 [0.17, 0.89]



Analysis 10.1. Comparison 10: Oral compared to nebulised dexamethasone, Outcome 1: Return visits or (re)admissions or both by outpatient

	Oral	Dex	Nebulis	ed Dex		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
10.1.1 Outpatient							
Luria 2001	7	85	19	91	100.0%	0.39 [0.17, 0.89]	
Subtotal (95% CI)		85		91	100.0%	0.39 [0.17, 0.89]	
Total events:	7		19				•
Heterogeneity: Not app	licable						
Test for overall effect:	Z = 2.24 (P =	0.03)					
Total (95% CI)		85		91	100.0%	0.39 [0.17, 0.89]	
Total events:	7		19				•
Heterogeneity: Not app	licable						0.01 0.1 1 10 100
Test for overall effect:	Z = 2.24 (P =	0.03)					Favours oral dex Favours nebulised dex
Test for subgroup differ	rences: Not a	pplicable					

Comparison 11. Dexamethasone 0.30 mg/kg compared to 0.15 mg/kg

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
11.1 Return visits or (re)admissions or both by outpatient	1	60	Risk Ratio (M-H, Random, 95% CI)	0.94 [0.06, 14.27]
11.1.1 Outpatient	1	60	Risk Ratio (M-H, Random, 95% CI)	0.94 [0.06, 14.27]
11.2 Additional treatments: epinephrine	1	60	Risk Ratio (M-H, Random, 95% CI)	0.43 [0.19, 0.98]
11.3 Additional treatments: supplemental glucocorticoids	1	60	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.06, 0.06]

Analysis 11.1. Comparison 11: Dexamethasone 0.30 mg/kg compared to 0.15 mg/kg, Outcome 1: Return visits or (re)admissions or both by outpatient

	Dex 0.30	mk/kg	Dex 0.15	mg/kg		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
11.1.1.0							
11.1.1 Outpatient							
Geelhoed 1995b	1	31	1	29	100.0%	0.94 [0.06 , 14.27]	
Subtotal (95% CI)		31		29	100.0%	0.94 [0.06, 14.27]	
Total events:	1		1				\top
Heterogeneity: Not appl	icable						
Test for overall effect: Z	L = 0.05 (P =	0.96)					
Total (95% CI)		31		29	100.0%	0.94 [0.06 , 14.27]	
Total events:	1		1				. T
Heterogeneity: Not appl	icable					0.00	01 0.1 1 10 1000
Test for overall effect: Z	Z = 0.05 (P =	0.96)				Favours d	ex 0.30 mg/kg Favours dex 0.15 mg/kg
Test for subgroup differen	ences: Not a _l	pplicable					



Analysis 11.2. Comparison 11: Dexamethasone 0.30 mg/kg compared to 0.15 mg/kg, Outcome 2: Additional treatments: epinephrine

	Dex 0.30	mg/kg	Dex 0.15	mg/kg		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Geelhoed 1995b	6	31	13	29	100.0%	0.43 [0.19, 0.98]	-
Total (95% CI)		31		29	100.0%	0.43 [0.19, 0.98]	
Total events:	6		13				
Heterogeneity: Not app	licable					0.01	0.1 1 10 100
Test for overall effect: $Z = 2.00 (P = 0.05)$					Favours dex	0.30 mg/kg Favours dex 0.15 mg/kg	
Test for subgroup differ	rences: Not a	pplicable					

Analysis 11.3. Comparison 11: Dexamethasone 0.30 mg/kg compared to 0.15 mg/kg, Outcome 3: Additional treatments: supplemental glucocorticoids

	Dex 0.30	mg/kg	Dex 0.15	mg/kg		Risk Difference	Risk Difference	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI	
Geelhoed 1995b	0	31	0	29	100.0%	0.00 [-0.06 , 0.06]	•	_
Total (95% CI)		31		29	100.0%	0.00 [-0.06, 0.06]	•	
Total events:	0		0				Ť	
Heterogeneity: Not app	licable						-0.5 -0.25 0 0.25 0.5	
Test for overall effect: 2	Z = 0.00 (P =	1.00)				Favour	s dex 0.30 mg/kg Favours dex 0.15	mg/kg
Test for subgroup differ	rences: Not a	pplicable						

Comparison 12. Dexamethasone 0.60 mg/kg compared to 0.30 mg/kg

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
12.1 Return visits or (re)admissions or both by outpatient	1	60	Risk Ratio (M-H, Random, 95% CI)	1.40 [0.25, 7.81]
12.1.1 Outpatient	1	60	Risk Ratio (M-H, Random, 95% CI)	1.40 [0.25, 7.81]
12.2 Additional treatments: epinephrine	1	60	Risk Ratio (M-H, Random, 95% CI)	0.78 [0.27, 2.28]
12.3 Additional treatments: supplemental glucocorticoids	1	60	Risk Ratio (M-H, Random, 95% CI)	2.81 [0.12, 66.40]



Analysis 12.1. Comparison 12: Dexamethasone 0.60 mg/kg compared to 0.30 mg/kg, Outcome 1: Return visits or (re)admissions or both by outpatient

	Dex 0.60	mg/kg	Dex 0.30	mg/kg		Risk Ratio	Risk Ratio
Study or Subgroup	Events Total		Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
12.1.1 Outpatient							
Geelhoed 1995a	3	31	2	29	100.0%	1.40 [0.25, 7.81]	
Subtotal (95% CI)		31		29	100.0%	1.40 [0.25, 7.81]	
Total events:	3		2				
Heterogeneity: Not applica	able						
Test for overall effect: Z =	0.39 (P =	0.70)					
Total (95% CI)		31		29	100.0%	1.40 [0.25 , 7.81]	
Total events:	3		2				
Heterogeneity: Not applica	able					0.0	1 0.1 1 10 100
Test for overall effect: Z =	0.39 (P =	0.70)					ex 0.60 mg/kg Favours dex 0.30 mg/k
Test for subgroup difference	ces: Not ap	plicable					

Analysis 12.2. Comparison 12: Dexamethasone 0.60 mg/kg compared to 0.30 mg/kg, Outcome 2: Additional treatments: epinephrine

Study or Subgroup	Dex 0.60 Events	mg/kg Total	Dex 0.30 Events	Dex 0.30 mg/kg events Total		Risk Ratio M-H, Random, 95% CI	Risk Ratio M-H, Random, 95% CI
Geelhoed 1995a	5	31	6	29	100.0%	0.78 [0.27 , 2.28]	-
Total (95% CI)		31		29	100.0%	0.78 [0.27, 2.28]	
Total events:	5		6				\neg
Heterogeneity: Not appl	icable					0.0	1 0.1 1 10 100
Test for overall effect: Z	L = 0.45 (P =	0.65)				Favours d	ex 0.60 mg/kg Favours dex 0.30 mg/kg
Test for subgroup differ	ences: Not a	pplicable					

Analysis 12.3. Comparison 12: Dexamethasone 0.60 mg/kg compared to 0.30 mg/kg, Outcome 3: Additional treatments: supplemental glucocorticoids

	Dex 0.60	mg/kg	Dex 0.30	mg/kg		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Geelhoed 1995a	1	31	0	29	100.0%	2.81 [0.12 , 66.40]	
Total (95% CI)		31		29	100.0%	2.81 [0.12, 66.40]	
Total events:	1		0				
Heterogeneity: Not appl	icable						0.002 0.1 1 10 500
Test for overall effect: Z	= 0.64 (P =	0.52)				Favou	irs dex 0.60 mg/kg Favours dex 0.30 mg/kg
Test for subgroup differen	ences: Not a	pplicable					

Comparison 13. Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg

Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
13.1 Croup score (Westley) (change baseline - 2 hours) by in- patient/outpatient	2	861	Std. Mean Difference (IV, Random, 95% CI)	-0.27 [-0.76, 0.22]



Outcome or subgroup title	No. of studies	No. of partici- pants	Statistical method	Effect size
13.1.1 Inpatient	1	41	Std. Mean Difference (IV, Random, 95% CI)	-0.63 [-1.25, 0.00]
13.1.2 Outpatient	1	820	Std. Mean Difference (IV, Random, 95% CI)	-0.10 [-0.23, 0.04]
13.2 Croup score (change base- line - 6 hours) by inpatient/outpa- tient	3	178	Std. Mean Difference (IV, Random, 95% CI)	-0.45 [-1.26, 0.35]
13.2.1 Inpatient	1	41	Std. Mean Difference (IV, Random, 95% CI)	-1.43 [-2.13, -0.74]
13.2.2 Outpatient	2	137	Std. Mean Difference (IV, Random, 95% CI)	-0.02 [-0.35, 0.32]
13.3 Croup score (change base- line - 12 hours) by inpatient/out- patient	2	113	Std. Mean Difference (IV, Random, 95% CI)	-0.60 [-4.39, 3.19]
13.3.1 Inpatient	1	41	Std. Mean Difference (IV, Random, 95% CI)	-2.55 [-3.39, -1.71]
13.3.2 Outpatient	1	72	Std. Mean Difference (IV, Random, 95% CI)	1.32 [0.81, 1.83]
13.4 Croup score (change base- line - 24 hours) by outpatient	1	72	Std. Mean Difference (IV, Random, 95% CI)	0.63 [0.16, 1.10]
13.4.1 Outpatient	1	72	Std. Mean Difference (IV, Random, 95% CI)	0.63 [0.16, 1.10]
13.5 Return visits or (re)admissions or both by outpatient	3	949	Risk Ratio (M-H, Random, 95% CI)	0.91 [0.71, 1.17]
13.5.1 Outpatient	3	949	Risk Ratio (M-H, Random, 95% CI)	0.91 [0.71, 1.17]
13.6 Length of stay by outpatient	2	892	Mean Difference (IV, Random, 95% CI)	0.12 [-0.32, 0.56]
13.6.1 Outpatient	2	892	Mean Difference (IV, Random, 95% CI)	0.12 [-0.32, 0.56]
13.7 Additional treatments: epi- nephrine	2	885	Risk Ratio (M-H, Random, 95% CI)	0.78 [0.34, 1.75]
13.8 Additional treatments: intu- bation/tracheotomy	2	861	Risk Difference (M-H, Random, 95% CI)	0.00 [-0.00, 0.00]
13.9 Additional treatments: sup- plemental glucocorticoids	2	617	Risk Ratio (M-H, Random, 95% CI)	0.77 [0.51, 1.15]



Analysis 13.1. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 1: Croup score (Westley) (change baseline - 2 hours) by inpatient/outpatient

	Dex	Dex 0.60 mg/kg			0.15 mg/k	ιg		Std. Mean Difference	Std. Mean Difference	
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	
13.1.1 Inpatient										
Chub-Uppakarn 2007	-1.2	0.24	20	-1.05	0.23	21	32.6%	-0.63 [-1.25, 0.00]		
Subtotal (95% CI)			20			21	32.6%	-0.63 [-1.25, 0.00]		
Heterogeneity: Not applic	able								•	
Test for overall effect: Z =	= 1.95 (P = 0.	05)								
13.1.2 Outpatient										
Parker 2019	-0.88	1.42	410	-0.75	1.3	410	67.4%	-0.10 [-0.23, 0.04]	•	
Subtotal (95% CI)			410			410	67.4%	-0.10 [-0.23, 0.04]	<u> </u>	
Heterogeneity: Not applic	able								1	
Test for overall effect: Z =	= 1.37 (P = 0.	17)								
Total (95% CI)			430			431	100.0%	-0.27 [-0.76 , 0.22]		
Heterogeneity: Tau ² = 0.0	9; Chi ² = 2.6	1, df = 1 (I	P = 0.11); I	$^{2} = 62\%$						
Test for overall effect: Z =	= 1.08 (P = 0.	28)							-2 -1 0 1 2	
Test for subgroup differen	ices: Chi ² = 2	.61, df = 1	(P = 0.11)	$I^2 = 61.7\%$	ó			Favo	ours 0.60 mg/kg Favours 0.15 m	

Analysis 13.2. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 2: Croup score (change baseline - 6 hours) by inpatient/outpatient

	Dex	0.60 mg/l	kg	Dex	0.15 mg/l	kg		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
13.2.1 Inpatient									
Chub-Uppakarn 2007	-2.75	0.27	20	-2.37	0.25	21	30.5%	-1.43 [-2.13 , -0.74]	
Subtotal (95% CI)			20			21	30.5%	-1.43 [-2.13 , -0.74]	•
Heterogeneity: Not applic	able								
Test for overall effect: Z =	= 4.05 (P < 0.	0001)							
13.2.2 Outpatient									
Alshehr 2005	-2.9	2.4	36	-3.1	1.8	36	35.0%	0.09 [-0.37, 0.56]	-
Fifoot 2007	-2.23	0.99	31	-2.09	0.97	34	34.5%	-0.14 [-0.63, 0.35]	_ _
Subtotal (95% CI)			67			70	69.5%	-0.02 [-0.35 , 0.32]	•
Heterogeneity: Tau ² = 0.0	0; Chi ² = 0.4	7, df = 1 (l	P = 0.49); I	2 = 0%					Ţ
Test for overall effect: Z =	= 0.10 (P = 0.	92)							
Total (95% CI)			87			91	100.0%	-0.45 [-1.26 , 0.35]	
Heterogeneity: Tau ² = 0.4	3; Chi ² = 13.4	43, df = 2	(P = 0.001)); I ² = 85%					
Test for overall effect: Z =	= 1.10 (P = 0.	27)						-	-2 -1 0 1 2
Test for subgroup differer	nces: Chi² = 1	2.96, df =	1 (P = 0.00	1003), $I^2 = 92$	2.3%			Favours d	dex 0.60 mg/kg Favours dex 0.15 mg/kg



Analysis 13.3. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 3: Croup score (change baseline - 12 hours) by inpatient/outpatient

	Dex	0.60 mg/l	κg	Dex	0.15 mg/l	ιg		Std. Mean Difference	Std. Mea	n Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Rand	om, 95% CI
13.3.1 Inpatient										
Chub-Uppakarn 2007	-3.6	0.25	20	-2.95	0.25	21	49.6%	-2.55 [-3.39 , -1.71]	-	
Subtotal (95% CI)			20			21	49.6%	-2.55 [-3.39 , -1.71]	•	
Heterogeneity: Not applicab	ole								•	
Test for overall effect: $Z = 5$	5.92 (P < 0.	00001)								
13.3.2 Outpatient										
Alshehr 2005	-2.5	0.75	36	-3.5	0.75	36	50.4%	1.32 [0.81, 1.83]		
Subtotal (95% CI)			36			36	50.4%	1.32 [0.81, 1.83]		
Heterogeneity: Not applicab	ole									\
Test for overall effect: $Z = 5$	5.05 (P < 0.	00001)								
Total (95% CI)			56			57	100.0%	-0.60 [-4.39 , 3.19]		
Heterogeneity: $Tau^2 = 7.36$;	Chi ² = 59.0	00, df = 1	(P < 0.000	01); I ² = 98 ⁴	%					
Test for overall effect: $Z = 0$	0.31 (P = 0.	76)							-10 -5	0 5 10
Test for subgroup difference	es: Chi ² = 5	9.00, df =	1 (P < 0.00	0001), I ² = 9	98.3%			Favour	s dex 0.60 mg/kg	Favours dex 0.15 mg/l

Analysis 13.4. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 4: Croup score (change baseline - 24 hours) by outpatient

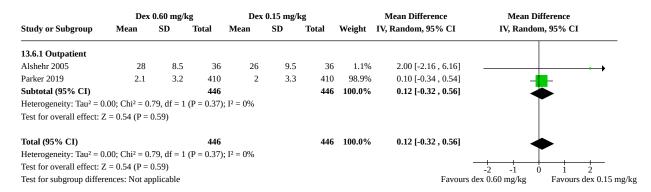
	Dex	0.60 mg/l	kg	Dex	0.15 mg/k	ιg		Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
13.4.1 Outpatient									
Alshehr 2005	-3.5	0.75	36	-4	0.82	36	100.0%	0.63 [0.16, 1.10]	ı _ _
Subtotal (95% CI)			36			36	100.0%	0.63 [0.16, 1.10]	
Heterogeneity: Not appl	licable								
Test for overall effect: Z	Z = 2.60 (P = 0.00)	0.009)							
Total (95% CI)			36			36	100.0%	0.63 [0.16 , 1.10]	
Heterogeneity: Not appl	licable								
Test for overall effect: Z	Z = 2.60 (P = 0.00)	0.009)							$\begin{array}{c ccccccccccccccccccccccccccccccccccc$
Test for subgroup differ	ences: Not ap	plicable						Favor	urs dex 0.60 mg/kg Favours dex 0.15 m

Analysis 13.5. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 5: Return visits or (re)admissions or both by outpatient

	Dex 0.60	mg/kg	Dex 0.15	mg/kg		Risk Ratio	Risk Ratio	1
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 9	5% CI
13.5.1 Outpatient								_
Alshehr 2005	14	36	15	36	19.9%	0.93 [0.53, 1.64]		_
Fifoot 2007	3	27	4	30	3.2%	0.83 [0.20, 3.39]	•	
Parker 2019	73	410	80	410	76.9%	0.91 [0.69, 1.21]		
Subtotal (95% CI)		473		476	100.0%	0.91 [0.71, 1.17]		
Total events:	90		99				\blacksquare	
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0	0.02, df = 2	(P = 0.99);	$I^2 = 0\%$				
Test for overall effect:	Z = 0.70 (P =	0.48)						
Total (95% CI)		473		476	100.0%	0.91 [0.71 , 1.17]		
Total events:	90		99				\blacksquare	
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0	0.02, df = 2	(P = 0.99);	$I^2 = 0\%$			0.5 0.7 1	1.5 2
Test for overall effect:	Z = 0.70 (P =	0.48)				Favour		avours dex 0.15 mg/k
Test for subgroup diffe	rences: Not a	pplicable						



Analysis 13.6. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 6: Length of stay by outpatient



Analysis 13.7. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 7: Additional treatments: epinephrine

	Dex 0.60	mg/kg	Dex 0.15	mg/kg		Risk Ratio	Risk F	Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Rando	m, 95% CI
Fifoot 2007	1	31	1	34	8.9%	1.10 [0.07 , 16.80]]	
Parker 2019	9	410	12	410	91.1%	0.75 [0.32 , 1.76]	l —	-
Total (95% CI)		441		444	100.0%	0.78 [0.34 , 1.75]		•
Total events:	10		13				J	
Heterogeneity: Tau ² = 0	0.00; Chi ² = 0	.07, df = 1	(P = 0.79);	$I^2 = 0\%$			0.005 0.1 1	10 200
Test for overall effect: 2	Z = 0.61 (P =	0.54)				Favo	urs dex 0.60 mg/kg	Favours dex 0.15 mg/kg
Test for subgroup differ	rences: Not a _l	pplicable						

Analysis 13.8. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 8: Additional treatments: intubation/tracheotomy

	Dex 0.60	mg/kg	Dex 0.15	mg/kg		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Chub-Uppakarn 2007	0	20	0	21	0.3%	0.00 [-0.09 , 0.09]	
Parker 2019	0	410	0	410	99.7%	0.00 [-0.00 , 0.00]	•
Total (95% CI)		430		431	100.0%	0.00 [-0.00, 0.00]	•
Total events:	0		0				Ť
Heterogeneity: Tau ² = 0.00); $Chi^2 = 0.0$	0, df = 1	P = 1.00); I ²	$^{2} = 0\%$		•	-0.020.01 0 0.010.02
Test for overall effect: Z =	0.00 (P = 1.	.00)				Favours of	dex 0.60 mg/kg Favours dex 0.15 mg/kg
Test for subgroup differen	ces: Not app	licable					



Analysis 13.9. Comparison 13: Dexamethasone 0.60 mg/kg compared to 0.15 mg/kg, Outcome 9: Additional treatments: supplemental glucocorticoids

	Dex 0.60	mg/kg	Dex 0.15	mg/kg		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Fifoot 2007	4	27	5	30	11.2%	0.89 [0.27 , 2.97]	
Parker 2019	32	282	42	278	88.8%	0.75 [0.49 , 1.15]	
Total (95% CI)		309		308	100.0%	0.77 [0.51 , 1.15]	
Total events:	36		47				
Heterogeneity: Tau ² = 0	.00; Chi ² = 0	.07, df = 1	(P = 0.80);	$I^2=0\%$			0.2 0.5 1 2 5
Test for overall effect: Z	Z = 1.30 (P =	0.19)				Favours	s dex 0.60 mg/kg Favours dex 0.15 mg/l

Test for subgroup differences: Not applicable

ADDITIONAL TABLES

Table 1. Number needed to treat for an additional beneficial outcome for return visits or (re)admissions or both for any glucocorticoid compared to placebo

Baseline rate (%)	NNTB (95% CI)
Mean baseline rate	
30.62	7 (5 to 12)
Smallest baseline rate	
2.06	102 (78 to 179)
Largest baseline rate	
72.00	3 (2 to 5)

NNTB: number needed to treat for an additional beneficial outcome

Table 2. Dexamethasone compared to budesonide for croup

Dexamethasone compared to budesonide for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: dexamethasone **Comparison:** budesonide

Outcomes	Anticipated abso	Relative - effect	№ of par- ticipants	Certainty of the evi-	Com- ments**	
	Budesonide	Dexamethasone	(95% CI)	(studies)	dence (GRADE)	ments
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms.	The mean change in croup score was -2.93 to -2.00 .	The mean change in croup score was 0.46 standard deviations in favour (0.79 more to 0.13 more).	-	326 (4 RCTs)	⊕⊕⊙⊝ Low ^{a,b}	A standard deviation of 0.46 rep- resents a moderate difference

Table 2. Dexamethasone compared to budesonide for croup (Continued)



(Follow-up: 6 hours)					between groups.
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms.	The mean change in croup score was -3.07 to -2.33.	The mean change in croup score was 0.75 standard deviations in favour (1.19 more to 0.30 more).	- 84 (2 RCTs)	⊕⊕⊙⊝ Low ^c ,d	A standard deviation of 0.75 represents a large difference between

Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 12 hours)	change in croup score was -3.07 to -2.33.	score was 0.75 standard deviations in favour (1.19 more to 0.30 more).	-	84 (2 RCTs)	⊕⊕⊙⊝ Low ^c ,d	A standard deviation of 0.75 represents a large difference between groups.
Return visits or (re)ad-	Study population		RR 0.69	374 (5.DCT-)	⊕⊕⊕⊚	
missions or both	122 per 1000 84 per 1000		- (0.40 to 1.22)	(5 RCTs)	Moderate ^e	
		(49 to 149)				
Adverse events	3/4 (75%) studies 2005; Johnson 199	reported collecting adverse ever reported no serious adverse events and the serious adverse events (Name 1998). Klasser Brush in the budesonide group	rents (Duman n 1998 report-	335 (4 RCTs)	⊕⊕⊝⊝ Low ^{f,} g	

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^qWe downgraded by one level for risk of bias. The contributing studies were at high (n = 2) and unclear (n = 2) risk of bias. Allocation concealment was unclear in two studies; blinding was unclear in two studies; and one study was unblinded. There was a baseline imbalance in croup score in one study.

bWe downgraded by one level for inconsistency. There was substantial heterogeneity (I² = 51%), and variation in point estimates.

cWe downgraded by one level for risk of bias. The contributing studies were at high risk of bias. Allocation concealment was unclear in both studies; blinding was unclear in one study, and the other study was unblinded. There was a baseline imbalance in croup score in one study. ^dWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

eWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size). The effect estimate included a null effect as well as considerable benefit for dexamethasone compared to budesonide.

fWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

EWe downgraded by one level for risk of bias. The contributing studies were at high (n = 2) and unclear (n = 2) risk of bias.

Table 3. Dexamethasone compared to beclomethasone for croup

Dexamethasone compared to beclomethasone for croup

Patient or population: children with croup

^{**}We used Cohen's interpretation of effect sizes to determine the magnitude of the difference between groups (0.2 represents a small effect, 0.5 represents a medium effect, 0.8 represents a large effect).



Table 3. Dexamethasone compared to beclomethasone for croup (Continued)

Setting: emergency department, inpatients and outpatients

Intervention: dexamethasone **Comparison:** beclomethasone

Outcomes	Anticipated absolut	e effects* (95% CI) Rela (95%	tive effect	№ of partici- pants	Certainty of the evidence (GRADE)
	Beclomethasone	Dexamethasone (337)	o Ci)	(studies)	
Return visits or (re)admissions or	Study population	RD 0	0.00 09 to 0.09)	39 (1 RCT)	⊕⊕⊕⊝ Moderate ^a
both	0 per 1000	0 per 1000	75 (0 0.03)	(21(01)	Moderate.
		(0 to 0)			
Adverse events (no	Eboriadou 2010 repo	rted no adverse events related to the g	d to the glucocorticoids.	39	00 00
events)				(1 RCT)	Low ^{b,c}

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RD: risk difference

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

Table 4. Dexamethasone compared to betamethasone for croup

Dexamethasone compared to betamethasone for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: dexamethasone **Comparison:** betamethasone

			Relative ef- — fect (95% CI)	№ of partici- pants (studies)	Certainty of the evidence (GRADE)
В	Setamethasone	Dexamethasone			
sessed with the Westley croup in score. Lower scores mean few-	The mean change in croup score from 1 study was 1.68.	The mean change in croup score was 0.62 units in favour (1.17 more to 0.06 more).	-	52 (1 RCT)	⊕⊕⊝⊝ Low ^{a,b}

^aWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

^bWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

cWe downgraded by one level for risk of bias. The one contributing study was at high risk of bias.



Table 4. Dexamethasone compared to betamethasone for croup (Continued)

Change in croup score. Assessed with the Westley croup score. Lower scores mean fewer symptoms.

The mean change in croup score from 1 study was **-1.89**.

The mean change in croup score was **0.67 units in favour** (1.23 more to 0.11 more).

52 ⊕⊕⊙⊝ (1 RCT) Lowb,c

(Follow-up: 6 hours)

Return visits or (re)admissions or both	Study population	1	RR 0.95 —— (0.67 to 1.34)	52 (1 RCT)	⊕⊕⊝⊝ Low ^d
	731 per 1000	694 per 1000	(6.6. 65 2.6.)	(2.1.0.)	2011
		(490 to 979)			
Adverse events	Amir 2006 did no	t report collecting adverse	events data.	52 (1 RCT)	⊕⊕⊝⊝ Low ^{a,e}

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^qWe downgraded by one level for risk of bias. The one contributing study was at high risk of bias. Allocation concealment was unclear, and the study was not blinded. There was a baseline imbalance in croup score.

bWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

^cWe downgraded by one level for risk of bias. The one contributing study was at high risk of bias. Allocation concealment was unclear, and the study was not blinded. There was a baseline imbalance in croup score.

^dWe downgraded by two levels for imprecision. The sample size was small (did not meet the optimal information size). The effect estimate included both the null effect and appreciable benefit or harm for dexamethasone compared to betamethasone.

eWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

Table 5. Dexamethasone compared to prednisolone for croup

Dexamethasone compared to prednisolone for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: dexamethasone **Comparison:** prednisolone

Outcomes	Anticipated absolute	Relative ef- - fect	№ of partici- pants	Certainty of the evidence	
	Prednisolone	Dexamethasone	(95% CI)	(studies)	(GRADE)
Change in croup score. Assessed with the Westley croup score. Lower scores mean fewer symptoms.	The mean change in croup score from 1 study was -0.89 .	The mean change in croup score was 0.06 units not in favour (0.06 more to 0.18 less).	-	1231 (1 RCT)	⊕⊕⊕⊕ High



Table 5. Dexamethasone compared to prednisolone for croup (Continued)

(Follow-up: 2 hours)

Change in croup score. Assessed with the Westley croup score. Lower scores mean fewer symptoms. (Follow-up: 6 hours)	The mean change in croup score from 1 study was -2.35 .	The mean change in croup score was 0.21 units not in favour (0.21 more to 0.62 less).	-	99 (1 RCT)	⊕⊕⊕⊝ Moderate ^a
Return visits or (re)ad- missions or both	Study population	RR 0.55 - (0.28 to 1.11)	1537 (4 RCTs)	⊕⊕⊕⊝ Moderate ^b	
inissions of both	212 per 1000	117 per 1000 (59 to 236)	(0.20 to 1.11)	(111013)	Moderate-
Adverse events	Fifoot 2007, Garbutt 202 verse events related to a of insomnia (1/411, 0.24 the prednisolone group of 30-second febrile cor ty (1/820, 0.1%) in the d	reported 1 case /411, 3.3%) in 20, 3.5%), 1 case	1550 (4 RCTs)	⊕⊕⊕⊙ Moderate ^c	

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

Table 6. Budesonide compared to dexamethasone for croup

Budesonide compared to dexamethasone for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: budesonide **Comparison:** dexamethasone

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect —— (95% CI)	№ of partici- pants	Certainty of the evi- dence	
	Dexamethasone	Budesonide	(33 % Ci)	(studies)	(GRADE)	
Adverse events	Huang 2021 reported i	no adverse events.		92 (1 RCT)	⊕⊕⊕⊝ Moderate ^a	

^qWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

bWe downgraded by one level for inconsistency. There was substantial heterogeneity (I² = 59%), and variation in point estimates.

^cWe downgraded by one level for imprecision. Narrative synthesis conducted, estimates are not precise.



Table 6. Budesonide compared to dexamethasone for croup (Continued)

*The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^aWe downgraded one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

Table 7. Budesonide and dexamethasone compared to dexamethasone

Budesonide and dexamethasone compared to dexamethasone for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: budesonide and dexamethasone

Comparison: dexamethasone

Outcomes	Anticipated absolute effects* (95% CI)		Relative - effect	№ of par- ticipants	Certainty of the evi-	Com- ments**
	Dexamethasone	Budesonide and dexamethasone	(95% CI)	(studies)	dence (GRADE)	illents
Change in croup score. Assessed with different scores in different studies. Lower scores mean fewer symptoms. (Follow-up: 6 hours)	The mean change in croup score was -3.24 to -1.80.	The mean change in croup score was 0.05 standard deviations not in favour (0.19 more to 0.30 less).	-	255 (3 RCTs)	⊕⊕⊕⊝ Moderate ^a	A standard deviation of 0.05 represents a minimal difference between groups.
Return visits or (re)admissions or	Study population	dy population		254 (3 RCTs)	⊕⊕⊝⊝ Lowb	
both	100 per 1000	91 per 1000 (45 to 183)	- (0.45 to 1.83)	(5 110 15)	LOWS	
Adverse events	lassen 1998 reporte	reported collecting adverse events data. K- orted no adverse events in either the dexam-		133 (1 RCTs)	⊕⊕⊕⊝ Moderate ^c	
	ethasone group or the dexamethasone and budesonide group.			(11010)		

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

^{**}We used Cohen's interpretation of effect sizes to determine the magnitude of the difference between groups (0.2 represents a small effect, 0.5 represents a medium effect, 0.8 represents a large effect).



Table 7. Budesonide and dexamethasone compared to dexamethasone (Continued)

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^aWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

bWe downgraded by two levels for imprecision. The sample size was small (did not meet the optimal information size). The effect estimate included both the null effect and a significant benefit or harm for dexamethasone and budesonide compared to dexamethasone alone. cWe downgraded by one level for imprecision. Narrative sythesis was conducted, estimates are not precise.

Table 8. Budesonide and dexamethasone compared to budesonide

Budesonide and dexamethasone compared to budesonide for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: budesonide and dexamethasone

Comparison: budesonide

Outcomes	Anticipated absolute	effects* (95% CI)	Relative ef- — fect	№ of partici- pants	Certainty of the evidence	
	Budesonide	Budesonide and dexamethasone	(95% CI)	(studies)	(GRADE)	
Change in croup score. Assessed with the West- ley croup score. Low- er scores mean fewer symptoms.	The mean change in croup score from 1 study was -2.30 .	The mean change in croup score was 0.18 units in favour (0.52 more to 0.17 less).	-	129 (1 RCT)	⊕⊕⊕⊝ Moderate ^a	
(Follow-up: 6 hours)						
Return visits or (re)ad- missions or both	Study population		RD 0.00 - (-0.03 to 0.03)	129 (1 RCT)	⊕⊕⊕⊝ Moderate ^a	
missions of both	0 per 1000	0 per 1000 (0 to 0)	= (0.03 to 0.03)	(TRCI)	Moderate	
Adverse events Klassen 1998 reported 1 case of oral thrush in the budesonide group (1/65, 1.5%) and no adverse events in the dexamethasone and budesonide group.		129 (1 RCT)	⊕⊕⊕⊝ Moderate ^b			

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RD: risk difference

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.



 $\it a$ We downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

bWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

Table 9. Oral dexamethasone compared to intramuscular dexamethasone for croup

Oral dexamethasone compared to intramuscular dexamethasone for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: oral dexamethasone

Comparison: intramuscular dexamethasone

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect — (95% CI)	№ of partici- pants	Certainty of the evidence	
	Intramuscular dexamethasone	Oral dexamethasone	— (33 % Ci)	(studies)	(GRADE)	
Return visits or (re)admissions or	Study population		RR 0.81 — (0.58 to 1.12)	440 (3 RCTs)	⊕⊕⊕⊝ Moderate ^a	
both	259 per 1000	210 per 1000 (150 to 290)	— (0.38 to 1.12)	(3 NC13)	Model area	
Adverse events	None of the studies repor	ted collecting adverse ever	its data.			

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^qWe downgraded by one level for imprecision. The effect estimate included both a null effect and substantial benefit for oral compared to intramuscular dexamethasone.

Table 10. Oral dexamethasone compared to nebulised dexamethasone for croup

Oral dexamethasone compared to nebulised dexamethasone for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: oral dexamethasone **Comparison:** nebulised dexamethasone

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect — (95% CI)	№ of partici- pants	Certainty of the evidence
	Nebulised dexamethasone	Oral dexamethasone	— (95% CI)	(studies)	(GRADE)
Return visits or (re)admissions or	Study population		RR 0.39 — (0.17 to 0.89)	176 (1 RCT)	⊕⊕⊕⊝ Moderate ^a
both	209 per 1000	81 per 1000	(0.11 to 0.03)	(11.01)	Moderates



Table 10. Oral dexamethasone compared to nebulised dexamethasone for croup (Continued)

(35 to 186)

Adverse events None of the studies reported collecting adverse events data.

*The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^aWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

Table 11. Dexamethasone 0.30 mg/kg compared to dexamethasone 0.15 mg/kg for croup

Dexamethasone 0.30 mg/kg compared to dexamethasone 0.15 mg/kg for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: dexamethasone 0.30 mg/kg **Comparison:** dexamethasone 0.15 mg/kg

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect - (95% CI)	№ of partici- pants	Certainty of the evidence
	Dexamethasone 0.15 mg/kg	Dexamethasone 0.30 mg/kg	- (33% CI)	(studies)	(GRADE)
Return visits or (re)admissions	Study population		RR 0.94 – (0.06 to 14.27)	60 (1 RCT)	⊕⊕⊝⊝ Low ^a
or both	34 per 1000	32 per 1000 (2 to 492)	= (0.00 to 14.27)	(I NCI)	LOW
Adverse events	Geelhoed 1995b did not	ot report collecting adverse events data.		60 (1 RCT)	⊕⊕⊝⊝ Lowa,b

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.



^aWe downgraded by two levels for imprecision. The sample size was small (did not meet the optimal information size). The effect estimate included significant benefit, the null effect, and potential harm for 0.30 mg/kg compared to 0.15 mg/kg dexamethasone.

bWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

Table 12. Dexamethasone 0.60 mg/kg compared to dexamethasone 0.30 mg/kg for croup

Dexamethasone 0.60 mg/kg compared to dexamethasone 0.30 mg/kg for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: dexamethasone 0.60 mg/kg **Comparison:** dexamethasone 0.30 mg/kg

Outcomes	Anticipated absolute effects* (95% CI)		Relative effect — (95% CI)	№ of partici- pants	Certainty of the evidence
	Dexamethasone 0.30 mg/kg	Dexamethasone 0.60 mg/kg	- (95% CI)	(studies)	(GRADE)
Return visits or (re)admissions	Study population		RR 1.40 — (0.25 to 7.81)	60 (1 RCT)	⊕⊕⊝⊝ Low ^a
or both	69 per 1000	97 per 1000 (17 to 539)	_ (0.23 to 7.61)	(I KCI)	LOW
Adverse events	Geelhoed 1995a did not	report collecting adverse events da	ata.	60 (1 RCT)	⊕⊝⊝⊝ Very low ^{a,b}

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^aWe downgraded by two levels for imprecision. The sample size was small (did not meet the optimal information size). The effect estimate included significant benefit, the null effect, and potential for harm for 0.60 mg/kg compared to 0.30 mg/kg dexamethasone.

^bWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

Table 13. Dexamethasone 0.60 mg/kg compared to dexamethasone 0.15 mg/kg for croup

Dexamethasone 0.60 mg/kg compared to dexamethasone 0.15 mg/kg for croup

Patient or population: children with croup

Setting: emergency department, inpatients and outpatients

Intervention: dexamethasone 0.60 mg/kg **Comparison:** dexamethasone 0.15 mg/kg

Outcomes	Anticipated absolut	te effects* (95% CI)	Relative — effect	№ of par- ticipants (studies)	Certainty of the evi-	Com- ments**
	Dexamethasone 0.15 mg/kg	Dexamethasone 0.60 mg/kg	(95% CI)		dence (GRADE)	menes



Гable 13. Dexameth	nasone 0.60 mg/kg c	ompared to dexamethas	one 0.15 mg	/kg for croup	(Continued)	
Change in croup score. Assessed with the Westley croup score. Lower scores mean fewer symp- toms. (Follow-up: 2 hours)	The mean change in croup score was -1.05 to -0.75 .	The mean change in croup score was 0.27 standard deviations in favour (0.76 more to 0.22 less).	-	861 (2 RCTs)	⊕⊕⊕⊕ High	A standard deviation of 0.14 rep- resents a small dif- ference between groups.
Change in croup score. Assessed with the Westley croup score. Lower scores mean fewer symptoms.	The mean change in croup score was -3.10 to -2.09.	The mean change in croup score was 0.45 units in favour (1.26 more to 0.35 less).	-	178 (3 RCTs)	⊕⊕⊕⊝ Moderate ^a	
(Follow-up: 6 hours)						
Change in croup score. Assessed with the Westley croup score. Lower scores mean fewer symp- toms.	The mean change in croup score was -3.50 to -2.95 .	The mean change in croup score was 0.60 units in favour (4.39 more to 3.19 less).	-	113 (2 RCTs)	⊕⊝⊝⊝ Very Low ^{b,c}	
(Follow-up: 12 hours)						
Change in croup score. Assessed with the Westley croup score. Lower scores mean fewer symp- toms.	The mean change in croup score from 1 study was -4.00 .	The mean change in croup score was 0.63 units not in favour (0.16 less to 1.10 less).	-	72 (1 RCT)	⊕⊕⊕⊙ Moderate ^a	
(Follow-up: 24 hours)						
Return visits or (re)admissions or	Study population		RR 0.91 - (0.71 to	949 (3 RCTs)	⊕⊕⊕⊕ High	
both	208 per 1000	189 per 1000 (148 to 243)	1.17)	(5.1.5.15)	riigii	
Adverse events	Parker 2019 reported 1 case of 30 seconds of 0.60 mg/kg dexameth (13/410 (3.3%), 1 case hyperactivity (1/410, 1/20) group. Alshehr 2005 r 2 cases of bronchopn sone group (3/36, 8.30) kg dexamethasone gr 2007 reported no adv	0.2%) in the of vomiting I 1 case of methasone racheitis and lexametha- the 0.15 mg/ nd Fifoot	170 (3 RCTs)	⊕⊕⊕⊙ Moderate ^d		

^{*}The risk in the intervention group (and its 95% confidence interval) is based on the assumed risk in the comparison group and the relative effect of the intervention (and its 95% CI).

CI: confidence interval; RCT: randomised controlled trial; RR: risk ratio

GRADE Working Group grades of evidence

^{**}We used Cohen's interpretation of effect sizes to determine the magnitude of the difference between groups (0.2 represents a small effect, 0.5 represents a medium effect, 0.8 represents a large effect).



Table 13. Dexamethasone 0.60 mg/kg compared to dexamethasone 0.15 mg/kg for croup (Continued)

High certainty: We are very confident that the true effect lies close to that of the estimate of the effect.

Moderate certainty: We are moderately confident in the effect estimate: the true effect is likely to be close to the estimate of the effect, but there is a possibility that it is substantially different.

Low certainty: Our confidence in the effect estimate is limited: the true effect may be substantially different from the estimate of the effect.

Very low certainty: We have very little confidence in the effect estimate: the true effect is likely to be substantially different from the estimate of effect.

^qWe downgraded by one level for imprecision. The sample size was small (did not meet the optimal information size).

^bWe downgraded by two levels level for inconsistency. There was considerable heterogeneity (I² = 99%), and variation in point estimates. The 95% confidence intervals did not overlap.

cWe downgraded by two levels for imprecision. The sample size was small (did not meet the optimal information size). The effect estimate included both the null effect and appreciable benefit and harm for 0.60 mg/kg compared to 0.15 mg/kg dexamethasone.

dWe downgraded by one level for imprecision. Narrative synthesis was conducted, estimates are not precise.

APPENDICES

Appendix 1. Search strategies for the 2022 update

Ovid MEDLINE(R) and Epub Ahead of Print, In-Process, In-Data-Review & Other Non-Indexed Citations and Daily 1946 to 4 March 2022

Search Strategy:		
#	Searches	Results
1	Adrenal Cortex Hormones/	68009
2	Beclomethasone/	3062
3	exp Betamethasone/	7720
4	Budesonide/	4755
5	Cortisone/	19687
6	Corticosterone/	26436
7	Cortodoxone/	893
8	Dexamethasone/	53872
9	exp Glucocorticoids/	201487
10	Hydrocortisone/	74655
11	Hydroxycorticosteroids/	375
12	exp Methylprednisolone/	20351
13	Prednisolone/	33795
14	Prednisone/	40638



(Continued)		
15	Pregnenolone/	4436
16	Pregnenediones/	2273
17	Tetrahydrocortisol/	269
18	Triamcinolone/	3962
19	adrenal cortex hormone*.tw,kf,nm.	68324
20	becl?met*.tw,kf,nm.	4015
21	betamet?asone*.tw,kf,nm.	8204
22	budesonide*.tw,kf,nm.	6808
23	clobetasol*.tw,kf,nm.	1961
24	corticoid*.tw,kf,nm.	6443
25	corticosteroid*.tw,kf,nm.	114283
26	corticosterone*.tw,kf,nm.	35897
27	cortisone*.tw,kf,nm.	23717
28	cortodoxone*.tw,kf,nm.	893
29	dexamet?asone*.tw,kf,nm.	76779
30	glucocortico*.tw,kf,nm.	122905
31	hydrocortisone*.tw,kf,nm.	80392
32	hydroxycorticosteroid*.tw,kf,nm.	6735
33	hydroxypregnenolone*.tw,kf,nm.	1008
34	methylprednisolone*.tw,kf,nm.	28760
35	prednisolone*.tw,kf,nm.	48416
36	prednisone*.tw,kf,nm.	55854
37	pregnenedione*.tw,kf,nm.	2279
38	pregnenolone*.tw,kf,nm.	7389
39	tetrahydrocortisol*.tw,kf,nm.	508
40	triamcinolone*.tw,kf,nm.	12442
41	or/1-40 [Glucocorticoids Concept]	513682
42	exp Laryngitis/	4066



(Continued)		
43	(croup* or pseudocroup*).tw,kf.	1883
44	(laryngo tracheo bronch* or laryngotracheobronch*).tw,kf.	558
45	(laryngo tracheit* or laryngotracheit*).tw,kf.	951
46	laryngit*.tw,kf.	2167
47	or/42-46 [Croup Concept]	6844
48	and/41,47 [Glucocorticoids and Croup Concepts]	632
49	randomized controlled trial.pt.	560028
50	controlled clinical trial.pt.	94719
51	randomized.ab.	552306
52	placebo.ab.	225982
53	drug therapy.fs.	2450860
54	randomly.ab.	377029
55	trial.ab.	589312
56	groups.ab.	2317503
57	or/49-56 [RCT Filter]	5275938
58	exp animals/ not humans.sh.	4966780
59	57 not 58 [Cochrane Highly Sensitive Search Strategy for identifying randomized trials in MEDLINE: sensitivity- and precision-maximizing version (2008 revision); Ovid format, from section 3.6.1 in Lefebvre C, Glanville J, Briscoe S, Littlewood A, Marshall C, Metzendorf M-I, et al. Technical Supplement to Chapter 4: Searching for and selecting studies. In: Higgins JPT, Thomas J, Chandler J, Cumpston MS, Li T, Page MJ, Welch VA (eds). Cochrane Handbook for Systematic Reviews of Interventions Version 6. Cochrane, 2019. Available from: www.training.cochrane.org/handbook]	4590783
60	and/48,59 [RCT filter applied to glucocorticoid & croup results]	396
61	(201804* or 2019* or 2020*).ed. [Update Limit]	2108579
62	and/60-61 [Update Limit Applied]	17
63	remove duplicates from 62	17
64	(202009* or 2021* or 2022*).ed. [Update Limit]	1811972
65	60 and 64	11
66	remove duplicates from 65	11



Embase 1974 to 2022 Week 8

Search Strategy:		
#	Searches	Results
1	beclomethasone/	7662
2	betamethasone/	18188
3	budesonide/	22808
4	corticosteroid/	257015
5	corticosterone/	32159
6	cortisone/	13015
7	cortodoxone/	2123
8	dexamethasone/	167648
9	exp glucocorticoid/	778646
10	hydrocortisone/	133565
11	hydroxycorticosteroid/	304
12	methylprednisolone/	107667
13	prednisolone/	135821
14	prednisone/	184211
15	pregnane derivative/	2049
16	pregnenolone/	4757
17	steroid hormone/	12245
18	tetrahydrocortisol/	539
19	adrenal cortex hormone*.tw,kw.	438
20	becl?met*.tw,kw.	4671
21	betamet?asone*.tw,kw.	7424
22	budesonide*.tw,kw.	9804
23	clobetasol*.tw,kw.	2089
24	corticoid*.tw,kw.	7327



(Continued)		
25	corticosteroid*.tw,kw.	172123
26	corticosterone*.tw,kw.	32198
27	cortisone*.tw,kw.	8488
28	cortodoxone*.tw,kw.	8
29	dexamet?asone*.tw,kw.	86461
30	glucocortico*.tw,kw.	105249
31	hydrocortisone*.tw,kw.	20385
32	hydroxycorticosteroid*.tw,kw.	758
33	hydroxypregnenolone*.tw,kw.	733
34	methylprednisolone*.tw,kw.	29433
35	prednisolone*.tw,kw.	42598
36	prednisone*.tw,kw.	53777
37	pregnenedione*.tw,kw.	6
38	pregnenolone*.tw,kw.	5931
39	tetrahydrocortisol*.tw,kw.	461
40	triamcinolone*.tw,kw.	10669
41	or/1-40 [Glucocorticoids Concept]	1062940
42	croup/	1994
43	laryngitis/	3965
44	laryngotracheobronchitis/	702
45	pseudocroup/	252
46	(croup* or pseudocroup*).tw,kw.	2154
47	laryngit*.tw,kw.	2246
48	(laryngo tracheit* or laryngotracheit*).tw,kw.	854
49	(laryngo tracheo bronch* or laryngotracheobronch*).tw,kw.	544
50	or/42-49 [Croup Concept]	8520
51	and/41,50 [Combined Glucocorticoids and Croup Concepts]	1494
52	crossover procedure/	69541



(Continued)		
53	double blind procedure/	192593
54	randomized controlled trial/	696988
55	single blind procedure/	45282
56	allocat*.tw,kw.	178563
57	assign*.tw,kw.	443634
58	(cross over* or crossover*).tw,kw.	116961
59	(doubl* adj blind*).tw,kw.	227957
60	factorial*.tw,kw.	43563
61	placebo*.tw,kw.	339624
62	random*.tw,kw.	1763875
63	(singl* adj blind*).tw,kw.	28391
64	volunteer*.tw,kw.	277008
65	or/52-64 [Recommended terms to limit to trials in Embase]	2632950
66	exp animals/ not exp humans/	4907469
67	65 not 66	2387867
68	and/51,67 [RCT Filter and main search concept]	202
69	("2018" or "2019" or "2020").yr.	4984140
70	and/68-69 [Update Date Applied]	26
71	remove duplicates from 70	26
72	("2020" or "2021" or "2022").yr.	3880081
73	and/68,72	14

Cochrane Library (via Wiley) 4 March 2022

Search Strategy	Search Strategy:		
ID	Search	Hits	
#1	[mh ^"Adrenal Cortex Hormones"]	2502	
#2	[mh ^Beclomethasone]	1141	



### [mh Paudesonide] 1519 ### [mh Paudesonide] 1861 ### [mh Paudesonide] 1861 ### [mh Pocritione] 160 ### [mh Pocritione] 41 ### [mh Pocritione] 32 ### [mh Pocritione] 4938 ### [mh Pocritione] 4938 ### [mh Paudenthasone] 4938 ### [mh Paudenthasone] 4938 ### [mh Paudenthasone] 4154 ### [mh Paudenthasone] 6154 ### [mh Paudenthasone] 10 ### [mh Paudenthasone] 2851 ### [mh Paudenthasone] 2851 ### [mh Paudenthasone] 4135 ### [mh Paudenthasone] 4135 ### [mh Paudenthasone] 22 ### [mh Paudenthasone] 566 ### [mh Paudenthasone] 566 ### [mh Paudenthasone] 684 ### [mh Paud	(Continued)		
#5	#3	[mh Betamethasone]	1519
## [mh ^Corticosterone] ## ## ## ## ## ## ## ## ## ## ## ## ##	#4	[mh ^Budesonide]	1861
#7	#5	[mh ^Cortisone]	160
#8	#6	[mh ^Corticosterone]	41
#9 [mh Glucocrticoids] 4764 #10 [mh ^Hydrocortisone] 6154 #11 [mh ^Hydrocycorticosteroids] 10 #12 [mh Methylprednisolone] 2851 #13 [mh ^Prednisolone] 3157 #14 [mh ^Prednisone] 4135 #15 [mh ^Pregnenolone] 22 #16 [mh ^Pregnenolone] 566 #17 [mh ^Pregnenolone] 12 #18 [mh ^Triamcinolone] 684 #19 adrenal cortex hormone*:ti,ab,kw 0 #20 becl?met*:ti,ab,kw 2650 #21 (betametasone* or betamethasone*):ti,ab,kw 5030 #22 budesonide*:ti,ab,kw 5030 #23 clobetasol*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 23957 #25 corticosteroid*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 585 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#7	[mh ^Cortodoxone]	32
#10	#8	[mh ^Dexamethasone]	4938
#11 [mh ^Hydroxycorticosteroids] 10 #12 [mh Methylprednisolone] 2851 #13 [mh ^Prednisolone] 3157 #14 [mh ^Prednisolone] 4135 #15 [mh ^Pregnenolone] 22 #16 [mh ^Pregnenolone] 566 #17 [mh ^Tetrahydrocortisol] 12 #18 [mh ^Triamcinolone] 684 #19 adrenal cortex hormone*:ti,ab,kw 0 #20 becl?met*:ti,ab,kw 2650 #21 (betametasone* or betamethasone*):ti,ab,kw 5030 #22 budesonide*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 739 #25 corticosteroid*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#9	[mh Glucocorticoids]	4764
#12	#10	[mh ^Hydrocortisone]	6154
#13	#11	[mh ^Hydroxycorticosteroids]	10
#14 [mh ^Prednisone] 4135 #15 [mh ^Pregnenolone] 22 #16 [mh ^Pregnenediones] 566 #17 [mh ^Tetrahydrocortisol] 12 #18 [mh ^Triamcinolone] 684 #19 adrenal cortex hormone*:ti,ab,kw 0 #20 becl?met*:ti,ab,kw 2650 #21 (betametasone* or betamethasone*):ti,ab,kw 5030 #22 budesonide*:ti,ab,kw 5030 #23 clobetasol*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 739 #25 corticosteroid*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 40	#12	[mh Methylprednisolone]	2851
#15	#13	[mh ^Prednisolone]	3157
#16	#14	[mh ^Prednisone]	4135
#17	#15	[mh ^Pregnenolone]	22
#18	#16	[mh ^Pregnenediones]	566
#19 adrenal cortex hormone*:ti,ab,kw 0 #20 becl?met*:ti,ab,kw 2650 #21 (betametasone* or betamethasone*):ti,ab,kw 2617 #22 budesonide*:ti,ab,kw 5030 #23 clobetasol*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 739 #25 corticosteroid*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#17	[mh ^Tetrahydrocortisol]	12
#20 becl?met*:ti,ab,kw 2650 #21 (betametasone* or betamethasone*):ti,ab,kw 2617 #22 budesonide*:ti,ab,kw 5030 #23 clobetasol*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 739 #25 corticosteroid*:ti,ab,kw 23957 #26 corticosterone*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#18	[mh ^Triamcinolone]	684
#21 (betametasone* or betamethasone*):ti,ab,kw 2617 #22 budesonide*:ti,ab,kw 5030 #23 clobetasol*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 739 #25 corticosteroid*:ti,ab,kw 23957 #26 corticosterone*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#19	adrenal cortex hormone*:ti,ab,kw	0
#22 budesonide*:ti,ab,kw 5030 #23 clobetasol*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 739 #25 corticosteroid*:ti,ab,kw 23957 #26 corticosterone*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#20	becl?met*:ti,ab,kw	2650
#23 clobetasol*:ti,ab,kw 746 #24 corticoid*:ti,ab,kw 739 #25 corticosteroid*:ti,ab,kw 23957 #26 corticosterone*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#21	(betametasone* or betamethasone*):ti,ab,kw	2617
#24corticoid*:ti,ab,kw739#25corticosteroid*:ti,ab,kw23957#26corticosterone*:ti,ab,kw154#27cortisone*:ti,ab,kw585#28cortodoxone*:ti,ab,kw40#29(dexametasone* or dexamethasone*):ti,ab,kw13356	#22	budesonide*:ti,ab,kw	5030
#25 corticosteroid*:ti,ab,kw 23957 #26 corticosterone*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#23	clobetasol*:ti,ab,kw	746
#26 corticosterone*:ti,ab,kw 154 #27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#24	corticoid*:ti,ab,kw	739
#27 cortisone*:ti,ab,kw 585 #28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#25	corticosteroid*:ti,ab,kw	23957
#28 cortodoxone*:ti,ab,kw 40 #29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#26	corticosterone*:ti,ab,kw	154
#29 (dexametasone* or dexamethasone*):ti,ab,kw 13356	#27	cortisone*:ti,ab,kw	585
	#28	cortodoxone*:ti,ab,kw	40
#30 glucocortico*:ti,ab,kw 9736	#29	(dexametasone* or dexamethasone*):ti,ab,kw	13356
	#30	glucocortico*:ti,ab,kw	9736



(Continued)		
#31	hydrocortisone*:ti,ab,kw	9741
#32	hydroxycorticosteroid*:ti,ab,kw	107
#33	hydroxypregnenolone*:ti,ab,kw	8
#34	methylprednisolone*:ti,ab,kw	5679
#35	prednisolone*:ti,ab,kw	7576
#36	prednisone*:ti,ab,kw	10186
#37	pregnenedione*:ti,ab,kw	566
#38	pregnenolone*:ti,ab,kw	113
#39	tetrahydrocortisol*:ti,ab,kw	40
#40	triamcinolone*:ti,ab,kw	3418
#41	#1 or #2 or #3 or #4 or #5 or #6 or #7 or #8 or #9 or #10 or #11 or #12 or #13 or #14 or #15 or #16 or #17 or #18 or #19 or #20 or #21 or #22 or #23 or #24 or #25 or #26 or #27 or #28 or #29 or #30 or #31 or #32 or #33 or #34 or #35 or #36 or #37 or #38 or #39 or #40	71548
#42	[mh Laryngitis]	128
#43	(croup* or pseudocroup*):ti,ab,kw	237
#44	("laryngo tracheo bronch*" or laryngotracheobronch*):ti,ab,kw	28
#45	("laryngo tracheit*" or laryngotracheit*):ti,ab,kw	15
#46	laryngit*:ti,ab,kw	250
#47	#43 or #43 or #44 or #45 or #46	478
#48	#41 and #47	159
#49	#41 and #47 with Publication Year from 2020 to 2022, in Trials	6

Trial registry: ClinicalTrials.gov (https://clinicaltrials.gov/)

Search strategy

Advanced search >

 $Search\ Terms:\ croup\ OR\ laryngit is\ OR\ laryngot rache obronchit is\ OR\ laryngot rache it is$

Age: Child

Intervention: Anti-Inflammatory Agents



World Health Organization International Clinical Trials Registry Platform (http://apps.who.int.uml.idm.oclc.org/trialsearch/)

Search strategy

Advanced search >

Title: croup OR laryngitis OR laryngotracheobronchitis OR laryngotracheitis

Search for clinical trials in children

Recruitment Status is: ALL

Date Selection: 2018-2022

FEEDBACK

Taste of oral steroids may be a problem,

Summary

A recent letter in the Lancet has questioned the results of a study on oral prednisolone for wheeze in young children on the basis that (amongst other things) oral prednisolone tastes very bitter and may not have been taken well by the children in the study.(1)

Whilst the authors have replied that they overcame the problem by asking parents to mix the powder with the child's favorite juice, I have had comments from parents in the past that their children did not like the taste of soluble prednisolone tablets, and I gather that dexamethasone solution is also very bitter.

For this reason I have abandoned the use of prednisolone and dexamethasone in children with croup or acute asthma, and use soluble betamethasone tablets instead. Betamethasone and dexamethasone are equal in potency and both are more potent than oral prednisolone; the British National Formulary states that the equivalent dose is that 5 mg of prednisolone is equivalent to 750 μ g betamethasone (which equates to one and a half 500 μ g tablets). It should also be noted that dexamethasone oral solution costs about 10 times as much as betamethasone tablets!

My extrapolation of the results of this review to the use of betamethasone in primary care is based on two assumptions. Firstly that betamethasone is equivalent to dexamethasone, and secondly that the outpatient trials in secondary care contain patients that are similar to those presenting in primary care. I wonder if the authors agree that this is reasonable?

Reference

1. Weinberger M, Ahrens R. Oral prednisolone for viral wheeze in young children. Lancet 2004;363(9405):330

I certify that I have no affiliations with or involvement in any organisation or entity with a direct financial interest in the subject matter of my criticisms.

Reply

In response to Dr. Cates' comment regarding the use of betamethasone for the treatment of croup, we are unable to conclude that betamethasone is efficacious for the treatment of croup.

Among the included studies, only Klassen et al (1998) reported the results of the blinding methodology. Children were randomized to identically tasting and appearing budesonide, dexamethasone or both treatments. Research assistants and parents were asked to identify which study medication the child received. The responses were similar and this indicates that blinding was successful. In addition, Klassen has conducted RCTs using intravenous dexamethasone with a 70% sucrose solution. This has been very well-tolerated with a very low incidence of vomiting. Paediatric croup and asthma trials have shown that when compared to prednisone/prednisolone, oral dexamethasone combined with flavoured syrup is both well-tolerated and an inexpensive treatment.

To date, we are not aware of any RCTs in children with croup that compared betamethasone to placebo or an active treatment, such as dexamethasone. Although betamethasone is theoretically as potent as dexamethasone, there is no actual empirical data to prove this. Therefore, we cannot judge the equivalency, or the tolerability, of betamethasone versus dexamethasone. Perhaps a randomized controlled trial should be conducted that directly compares betamethasone to dexamethasone so the palatability and equivalency can be assessed.

In response to the second stated assumption, there are guidelines for generalising results of trials to clinical practice and physicians need to carefully consider the comparability of participants in any one study to their own patients.¹



¹Guyatt G, Haynes B, Jaeschke R, Cook D, Greenhalgh T, Meade M, Green L, Naylor C, Wilson M, McAlister F, Richardson M. Introduction: the philosophy of evidence-based medicine. In: Guyatt G, Rennie D, editors. Users' guides to the medical literature: a manual for evidence-based clinical practice. Chicago: AMA Press; 2002. pp. 3-12.

Kelly Russell Terry Klassen David Johnson

I certify that I have no affiliations with or involvement in any organisation or entity with a direct financial interest in the subject matter of my criticisms.

Contributors

Chris Cates

WHAT'S NEW

Date	Event	Description
4 March 2022	New search has been performed	The following new authors joined the review group for this 2022 update: Alex Aregbesola, Clara Tam, Asha Kothari, Mê-Linh Lê, and Mirna Ragheb. As in the previous review, we added an age range for children in our inclusion criteria (0 to 18 years). We did not extract data for some outcomes prespecified in the protocol because the new included studies did not report these outcomes. We used the Cochrane risk of bias tool to assess risk of bias for all studies included in the meta-analysis. We used the GRADE approach to assess the certainty of the evidence (GRADE-pro GDT). We added summary of findings tables for two comparisons.
4 March 2022	New citation required and conclusions have changed	We included two new trials (Huang 2021; Parker 2019), and identified one ongoing trial (IRCT20190914044765N1), and one trial awaiting classification (Chen 2018). We excluded six new trials (Faraji-Goodarzi 2018; Gursanscky 2019; Kotaniemi-Syrjanen 2018; Lee 2019; Meskina 2019; Tyler 2022). The previous version of this review concluded that glucocorticoids reduced symptoms of croup at two hours, shortened hospital stays, and reduced the rate of return visits to care. We concluded that dexamethasone probably reduces revisits or readmission of croup by about half. A smaller dose of 0.15 mg/kg of dexamethasone may be as effective as the standard dose of 0.60 mg/kg. More randomised controlled trials are needed to strengthen the evidence for effectiveness of low-dose dexamethasone at 0.15 mg/kg to treat croup.

HISTORY

Protocol first published: Issue 3, 1997 Review first published: Issue 1, 2000

Date	Event	Description
3 April 2018	New search has been performed	New authors joined the team to update the review. We updated the searches and included five new trials. Three were newly identified trials (Dobrovoljac 2012; Garbutt 2013; Soleimani 2013); one was a previously excluded trial (Chub-Uppakarn 2007); and one trial was previously awaiting classification (Eboriadou 2010). We excluded six new trials, Eghbali 2016; Faghihinia 2007; Fara-



Date	Event	Description
		ji-Goodarzi 2018; Gill 2017; Mohammadzadeh 2014; Roked 2015, and one ongoing trial (NCT01748162). We included one new ongoing trial (ACTRN12609000290291). We assessed risk of bias of the included studies and the certainty of the evidence. We added two new primary outcomes: change in croup score after two hours and patient improvement after two hours. We added adverse events as a secondary outcome and summary of findings tables. We included two new comparisons: oral compared to nebulised dexamethasone, and dexamethasone compared to beclomethasone.
3 April 2018	New citation required and conclusions have changed	Our conclusions have changed. The previous version of this review concluded that when compared to placebo, glucocorticoids reduce croup symptoms within six hours and that the effect lasts 12 hours. In this update we concluded that when compared to placebo, glucocorticoids reduce croup symptoms within two hours and that the effect lasts at least 24 hours.
16 September 2014	New search has been performed	Searches updated. We included one new trial, Dobrovoljac 2012, and excluded one new trial (Faghihinia 2007). We added a two-hour croup score and a two-hour improvement outcome.
16 September 2014	New citation required but conclusions have not changed	Review updated; conclusions remain unchanged.
1 December 2011	Amended	Grammatical correction made to the Plain language summary.
18 July 2011	Amended	Analysis 5.2 contained an error, as the negative signs for the change in croup scores at six hours were not included. The mean difference remains non-significant.
23 July 2010	New search has been performed	Searches conducted. We added seven new trials since the 2004 publication (Alshehr 2005; Amir 2006; Cetinkaya 2004; Duman 2005; Fifoot 2007; Geelhoed 2005; Sparrow 2006). We excluded three new trials (Chub-Uppakarn 2007; Custer 2005; Schooff 2005).
20 May 2010	New citation required but conclusions have not changed	New authors joined the team to update the review. The conclusions remain unchanged.
16 August 2008	Amended	Converted to new review format
3 February 2004	Feedback has been incorporated	Feedback incorporated.
7 April 2003	New search has been performed	Searches conducted.
17 August 1997	New search has been performed	Searches conducted.

CONTRIBUTIONS OF AUTHORS

Alex Aregbesola (AA): study selection, data extraction and verification, risk of bias assessment, GRADE assessment, statistical analyses, manuscript preparation.

Clara Tam (CT): data extraction and verification, risk of bias assessment, GRADE assessment, statistical analyses, manuscript preparation. Asha Kothari (AK): data extraction and verification, risk of bias assessment, GRADE assessment, manuscript preparation. Mê-Linh Lê (ML): study selection, contribution to the manuscript.

Mirna Ragheb (MR): data extraction and verification, risk of bias assessment, contribution to the manuscript.



Terry P Klassen (TPK): clinical adviser, contribution to the manuscript.

DECLARATIONS OF INTEREST

Alex Aregbesola: declared that they have no conflict of interest. Clara Tam: declared that they have no conflict of interest. Asha Kothari: declared that they have no conflict of interest. Mê-Linh Lê: declared that they have no conflict of interest. Mirna Ragheb: declared that they have no conflict of interest.

Terry P Klassen: is an author of four of the included studies (Bjornson 2004; Klassen 1994; Klassen 1996; Klassen 1998).

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· Alberta Research Centre for Health Evidence (ARCHE), University of Alberta, Canada

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Funding for previous versions of this review

External sources

· Alberta Heritage Foundation for Medical Research, Canada

Funding for previous versions of this review

DIFFERENCES BETWEEN PROTOCOL AND REVIEW

The following new authors joined the review group for the 2022 update: Alex Aregbesola, Clara Tam, Asha Kothari, Mê-Linh Lê, and Mirna Ragheb. As in the previous version of this review, we added an age range for children in our inclusion criteria (0 to 18 years). We did not extract data for some of the outcomes in the protocol because the newly included studies did not report these outcomes. We used GRADEpro GDT software to assess the certainty of the body of evidence (GRADEpro GDT). We updated the summary of findings tables for two comparisons and added a summary of findings table for one new comparison in the Additional tables section.

INDEX TERMS

Medical Subject Headings (MeSH)

*Croup [drug therapy]; Dexamethasone [therapeutic use]; Epinephrine [therapeutic use]; Glucocorticoids [therapeutic use]; Prednisolone [therapeutic use]; Randomized Controlled Trials as Topic; *Respiratory Tract Infections

MeSH check words

Adolescent; Child; Humans