# **Major depression disorder may causally associate with the increased breast cancer risk: Evidence from two-sample mendelian randomization analyses**

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### **Abstract**

**Introduction:** Major depression disorder (MDD) has been associated with increased breast cancer risk in epidemiological studies; however, it is still unknown whether this association is causal or not. The aim of this study is to determine the causal relationship between MDD and breast cancer risk.

**Methods:** Two-sample Mendelian randomization (MR) analyses with 92 single-nucleotide polymorphisms (SNPs) significantly associated with MDD as instrumental variables (IVs) were performed. Effects of these SNPs on breast cancer in women were estimated in the Breast Cancer Association Consortium (122,977 cases and 105,974 controls) using inverse variance weighted (IVW), weighted median and multivariable MR models. Heterogeneity and pleiotropy effects were assessed based on IVW and MR-Egger regression model, respectively. **Results:** An 8.7% increased risk of overall breast cancer [odds ratio (*OR*) = 1.087; 95% confidence interval (*CI*) 1.011–1.170; *P* = 0.025] per log-odds ratio increment of MDD risk based on the IVW model was noticed. Similar results were obtained with the multivariable MR model ( $OR = 1.118$ ,  $95\% CI = 1.010 - 1.237$ ;  $P = 0.031$ ). An increment but not statistically significant causality association was noticed between MDD and risk of ER+ (*OR* = 1.098, 95% *CI* = 0.984–1.227; *P* = 0.093) or ER-  $(OR = 1.129, 95\% CI = 0.982 - 1.297; P = 0.089)$  breast cancer under multivariable MR model. No significant pleiotropy effects were observed for the IVs in the two-sample MR studies.

**Conclusions:** The results suggested that a genetic predisposition of MDD is causally associated with overall breast cancer risk; however, the underlying biological mechanisms are worthy of further study.

### **KEYWORDS**

breast cancer, major depression disorder, mendelian randomization, single-nucleotide polymorphism

Qian Ren and Fangxiu Luo contributed equally to the current study.

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# **1** | **INTRODUCTION**

Breast cancer is one of the most common cancer and also the leading cause of cancer deaths in women.<sup>[1](#page-10-0)</sup> It was estimated a total of 2.1 million new breast cancer cases and 0.63 million deaths caused by breast cancer globally in [2](#page-10-1)018.<sup>2</sup> In a systemic analysis of the global burden of disease study, breast cancer played an important role in disability-adjusted life years (DALYs) along with a rising trend in most countries.<sup>[3](#page-10-2)</sup> Previous studies have reported that the genetic factors, physical activities, nutritional factors, and exposure to specific environmental pollutions are associated with the increased risk of breast cancer de-velopment<sup>[4](#page-10-3)</sup>; however, these analyses remain vulnerable to study design, and more studies are needed to identify the etiology factors for breast cancer.

Among the potential factors that may affect the risk of developing cancer, psychosocial factors have drawn an increasing attention.<sup>[5](#page-10-4)</sup> Depression is one of the most common psychosocial factors for patients with mood disorders, which was estimated to influence about 121 million people worldwide.<sup>[6](#page-10-5)</sup> A recent meta-analysis has suggested that depression was associated with increased mortality of breast cancer $7-9$ ; however, the results were not always consistent. $10$  Because of the short following-up time, incomplete or unreliable determination of depression status at baseline, and insufficient confounding control in the study design, the association between depression and breast cancer is still under debate.[11](#page-11-1) Therefore, more well-designed studies are needed to elucidate the causal associations between the depression and breast cancer risk.

The Mendelian randomization (MR) analysis, which uses germline genetic variation as the instrumental variable (IVs) of the potential exposure to evaluate the associations between the modifiable exposure and the outcomes, can be applied to make a causal inference. $^{12}$  $^{12}$  $^{12}$ Unlike traditional observational epidemiology, MR analysis could overcome the risk of bias caused by undetermined confoundings, reverse causality and measurement error. $^{13}$  $^{13}$  $^{13}$  MR study has been recognized as analogous with the randomized controlled trail (RCT) studies in assessing the causal relationship between the exposures and the outcomes, $14$  but it requires less time and expense and may address the questions that RCT studies are unable to ask.<sup>15</sup> A recent genome-wide association meta-analysis has been performed by Howard et al. who identified 102 independent single-nucleotide polymorphisms (SNPs) that significantly associated with major depression disease (MDD), and they evaluated the causal associations between MDD and 41 types of diseases or diseases-related traits using MR analysis meth $ods^{16}$ ; however, whether MDD is causally associated with

breast cancer risk has not yet determined. In the current study, we applied the genetic variations significantly associated with MDD as the genetic instrumental variables (IVs) to evaluate the causal relationship between MDD and breast cancer risk, which may provide novel intervention methods for breast cancer in future.

# **2** | **MATERIALS AND METHODS**

The study was performed based on the summary-level data of public genome-wide association studies (GWASs), and it is not necessary to be approved by the ethic committee in our institute. Consents from the participants were assumed to be obtained by the individual GWAS.

# **2.1** | **Study design and data sources**

We applied the two-sample MR study based on the public summary-level data derived from the GWASs to evaluated the causal association between MDD and breast cancer risk. Genetic IVs were derived from the GWAS performed by Howard et al., $^{16}$  and their associations with overall breast cancer, ER+ or ER- risk were assessed using the summary-level data from a GWAS study with 228,951 women including 122,977 breast cancer (69,501 ER+ and 21,468 ER- patients) cases and 105,974 controls of European ancestry from the Breast Cancer Association Consortium  $(BCAC).$ <sup>17</sup> To assess the influences of the potential confounders, we also performed the multivariable MR (MVMR) studies to assess the causal association between MDD and breast cancer with the adjustment for smoking initiation, alcoholic drinks per week, education (years of schooling), average household income (before tax), and age at menarche. Summary statistics data for smoking intiation and alcoholic drinks per week were extracted from meta-analysis of GWASs conducted by the GWAS and Sequencing Consortium of Alcohol and Nicotine use (GSCAN) consortium.<sup>18</sup> Summary level data of education attainment represented as the number of years of schooling were derived from a GWAS performed by Social Science Genetic Association Consortium  $(SSGAC).$ <sup>19</sup> SNPs significantly associated with the average household income before tax were extracted from the MRC-IEU UK Biobank GWAS of European ancestry with 397,751 response individuals.<sup>20,21</sup> Summary level data related to the associations between SNPs and age at menarche are extracted from the GWAS per-formed by ReproGen Consortium.<sup>[22](#page-11-11)</sup> Detailed information for the included GWASs was provided in Table [S1](#page-12-0), and the overall study design was shown in Figure [1](#page-2-0).



<span id="page-2-0"></span>**FIGURE 1** The study assumptions of the two-sample Mendelian Randomization analysis between MDD and breast cancer. The assumptions including: (1) the genetic instrumental variables (IVs) should be significantly associated with MDD; (2) the genetic IVs should not be associated with other exposure-results confounding factors; (3) the genetic IVs must only be through MDD and not through any other confounders to influence the breast cancer risk. The dotted lines indicate violate of the assumptions. BCAC, Breast Cancer Association Consortium; ER, Estrogen Receptor; GSCAN, GWAS and Sequencing Consortium of Alcohol and Nicotine use; MRC-IEU, MRC Integrative Epidemiology Unit; PGC, Psychiatric Genomics Consortium; ReproGen, Reproductive Genetics Consortium; SSGAC, Social Science Genetic Association Consortium.

# **2.2** | **Instrumental variables selection and validation**

According to the MR study design, the IVs used to assess causal relationship between MDD and breast cancer risk should meet the following assumptions: (1) the IVs should be significantly associated with MDD risk; (2) the IVs should not be associated with other potential confounding factors; (3) there is no direct association between the IVs and breast cancer, or through any other way except the exposure (MDD) to correlate with breast cancer (Figure [1\)](#page-2-0). Here, single-nucleotide polymorphisms (SNPs) from a large genome-wide meta-analysis performed by Psychiatric Genomics Consortium (PGC) consortium were recognized as the potential genetic IVs. $^{16}$  As reported by Howard et al.,<sup>16</sup> 102 independent variants locate on 101 genomic loci that were associated with depression were identified in the meta-analysis with a total of 807,553 individuals including 246,363 cases and 561,190 controls that derived from three GWAS of depression in European ancestry.<sup>16</sup> All these SNPs have reached the significant genome-wide significance  $(P < 5 \times 10^{-8})$ , and the base pair positions of them were identified by clumping all associated variants across a 3 Mb window on chromosome (linkage disequilibrium  $r^2$  < 0.1) and have merged any overlapping clumps, indicating these IVs met the assumptions 1 and 2 of the two-sample MR study design. Of them, four SNPs including rs78337797, rs56314503, rs10774600,

and rs3213572 were not identified in the BCAC GWAS summary-level data for breast cancer risk or with high LD proxy SNPs  $(r^2 > 0.80)$ . Another six SNPs including rs12052908, rs1933802, rs2029865, rs2247523, rs263645, and rs2876520 were discarded for being palindromic with intermediate allele frequencies (Table [1](#page-3-0)). Thus, a total of 92 validated IVs was applied in the current two-sample MR study (Table [1\)](#page-3-0). To calculate  $R^2$ , we used the following formula:  $(2 \times EAF [1-EAF] \times \beta^2)$ , where EAF is the effect allele frequency, β is the estimated genetic effect on MDD. We calculated the *F*-statistic to assess the strength of each IVs using the formulation:  $F = (R^2 \times [n-1 \cdot k])$  $([1 - R^2] \times k)$  as reported,<sup>[23](#page-11-12)</sup> where  $R^2$  was calculated as proportion of variance in the phenotype explained by genetic variants, k is the number of IVs, and n is the sample size. An *F*-statistic >10 has been recognized as a robust IV and recommended to be used in MR study.[23](#page-11-12) The *F*-statistic value of these genetic IVs ranged from 147 to 532 indicating sufficient strength of the IVs in the two-sample MR studies (Table [1](#page-3-0)).

# **2.3** | **Genetic associations between IVs and breast cancer risk**

The summary-level data for the associations between each genetic IVs with total, ER+ and ER- breast cancer patients were retrieved through the MRC Integrative

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<span id="page-3-0"></span>**TABLE 1** Summary information on MDD-associated SNPs used as genetic instrumental variables (IVs) for the two-sample Mendelian randomization (MR) analyses

			<b>Effect</b>	Other allele				P-value	$R^2$	
<b>SNP</b>	Chr	<b>Position</b>	allele		<b>EAF</b>	β	<b>SE</b>			<i>F</i> -statistic
rs301799	$\mathbf{1}$	8,489,302	T	$\mathcal{C}$	0.5694	$-0.0250$	0.0035	$1.36 \times 10^{-12}$	$1.08 \times 10^{-3}$	247.571
rs1002656	$\mathbf{1}$	37, 192, 741	$\mathbf T$	$\mathcal{C}$	0.7033	$-0.0266$	0.0038	$3.74 \times 10^{-12}$	$1.04 \times 10^{-3}$	238.531
rs1466887	1	37,709,328	T	C	0.5511	$-0.0199$	0.0036	$4.12 \times 10^{-8}$	$6.91 \times 10^{-4}$	158.258
rs11579246	$\mathbf{1}$	50,559,162	A	G	0.9067	0.0381	0.0061	$5.71 \times 10^{-10}$	$8.66 \times 10^{-4}$	198.380
rs1890946	$\mathbf{1}$	52,342,427	$\mathbf T$	$\mathsf C$	0.4671	$-0.0235$	0.0035	$2.68 \times 10^{-11}$	$9.69 \times 10^{-4}$	222.078
rs10789214	$\mathbf{1}$	67,146,817	T	$\mathsf C$	0.5661	0.0193	0.0035	$4.44 \times 10^{-8}$	$6.45 \times 10^{-4}$	147.799
rs2568958	$\mathbf{1}$	72,765,116	A	G	0.6156	0.0373	0.0036	$8.47 \times 10^{-25}$	$2.32 \times 10^{-3}$	532.085
rs10890020	$\mathbf{1}$	73,668,836	A	G	0.5156	$-0.0277$	0.0035	$4.03 \times 10^{-12}$	$1.35 \times 10^{-3}$	309.627
rs113188507	1	80,809,636	А	G	0.2838	0.0221	0.0039	$1.87 \times 10^{-8}$	$7.00 \times 10^{-4}$	160.366
rs10913112	$\mathbf{1}$	175,913,828	T	$\mathsf C$	0.3767	$-0.0264$	0.0036	$3.40 \times 10^{-13}$	$1.15 \times 10^{-3}$	264.386
rs72710803	$\mathbf{1}$	177,428,018	A	$\mathsf C$	0.9121	$-0.0410$	0.0062	$5.29 \times 10^{-11}$	$9.50 \times 10^{-4}$	217.726
rs169235	$\mathbf{1}$	181,740,924	A	G	0.753	$-0.0229$	0.0041	$2.98 \times 10^{-8}$	$6.88 \times 10^{-4}$	157.559
rs17641524	1	197,704,717	T	$\mathcal{C}$	0.2091	$-0.0320$	0.0043	$1.52 \times 10^{-13}$	$1.19 \times 10^{-3}$	273.601
rs12052908	$\overline{2}$	22,503,044	A	$\mathbf T$	0.5325	$-0.0220$	0.0035	$4.44 \times 10^{-10}$	$8.49 \times 10^{-4}$	194.646
rs1568452	2	58,012,833	T	$\mathsf{C}$	0.3851	0.0248	0.0036	$8.12 \times 10^{-12}$	$1.03 \times 10^{-3}$	235.290
rs7585722	$\overline{2}$	86,819,128	$\mathbf T$	$\mathsf C$	0.8458	$-0.0269$	0.0048	$2.68 \times 10^{-8}$	$6.65 \times 10^{-4}$	152.452
rs1226412	$\overline{c}$	157, 111, 313	$\mathbf T$	$\mathsf{C}$	0.7917	0.0256	0.0043	$3.46 \times 10^{-9}$	$7.62 \times 10^{-4}$	174.590
rs62188629	$\overline{c}$	208,044,470	A	G	0.3136	0.0236	0.0038	$7.13 \times 10^{-10}$	$8.45 \times 10^{-4}$	193.676
rs4346585	3	44,736,493	T	$\mathsf C$	0.696	$-0.0236$	0.0038	$7.13 \times 10^{-10}$	$8.31 \times 10^{-4}$	190.373
rs13084037	3	49,214,066	A	G	0.774	$-0.0245$	0.0042	$7.08 \times 10^{-9}$	$7.40 \times 10^{-4}$	169.616
rs7624336	3	53, 244, 151	$\mathbf T$	G	0.2087	0.0238	0.0043	$3.96 \times 10^{-8}$	$6.60 \times 10^{-4}$	151.110
rs141954845	3	61,192,911	A	G	0.388	0.0229	0.0037	$8.15 \times 10^{-10}$	$8.78 \times 10^{-4}$	201.167
rs6783233	3	117,509,984	T	$\mathsf{C}$	0.2833	0.0218	0.0039	$2.90 \times 10^{-8}$	$6.80 \times 10^{-4}$	155.875
rs1095626	3	157,977,962	$\mathbf T$	$\mathcal{C}$	0.5799	$-0.0264$	0.0035	$7.13 \times 10^{-14}$	$1.20 \times 10^{-3}$	274.319
rs7685686	4	3,207,142	A	G	0.5753	0.0202	0.0036	$2.57 \times 10^{-8}$	$7.03 \times 10^{-4}$	161.050
rs34937911	$\overline{4}$	42,110,353	T	$\mathcal{C}$	0.8838	0.0304	0.0055	$4.13 \times 10^{-8}$	$6.69 \times 10^{-4}$	153.315
rs45510091	4	123,186,393	A	G	0.9472	0.0448	0.0080	$1.83 \times 10^{-8}$	$7.08 \times 10^{-4}$	162.149
rs35553410	$\overline{4}$	131,237,381	T	$\mathsf C$	0.7462	$-0.0244$	0.0040	$1.42 \times 10^{-9}$	$7.95 \times 10^{-4}$	182.146
rs7659414	$\overline{4}$	177,350,956	A	$\mathsf{C}$	0.5782	$-0.0201$ 0.0035		$1.20 \times 10^{-8}$	$6.95 \times 10^{-4}$	159.169
rs60157091	$\sqrt{5}$	61,509,655	$\mathbf T$	$\mathsf C$	0.515	0.0200	0.0035	$1.42 \times 10^{-8}$	$7.04 \times 10^{-4}$	161.395
rs3099439	5	87,545,318	$\mathbf T$	$\mathbf C$	0.5288	$-0.0276$	0.0035	$5.05 \times 10^{-15}$	$1.34 \times 10^{-3}$	306.673
rs10061069	5	93,071,630	$\mathbf C$	${\bf G}$	0.2212	$-0.0275$ 0.0042		$8.15 \times 10^{-11}$	$9.18 \times 10^{-4}$	210.467
rs30266	5	103,972,357	A	${\bf G}$	0.3296	0.0308	0.0037	$1.45 \times 10^{-16}$	$1.48 \times 10^{-3}$	338.688
rs11135349	5	164,523,472	A	$\mathbf C$	0.4713	$-0.0295$	0.0035	$6.04 \times 10^{-17}$	$1.53 \times 10^{-3}$	350.376
rs200949	6	27,835,435	A	${\bf G}$	0.8744	0.0480	0.0053	$2.53 \times 10^{-19}$	$1.78 \times 10^{-3}$	408.881
rs9363467	6	66,565,703	$\mathbf T$	${\bf C}$	0.6035	0.0237	0.0036	$6.44 \times 10^{-11}$	$9.48 \times 10^{-4}$	217.135
rs7758630	6	101,387,304	A	T	0.4051	$-0.0225$	0.0036	$5.56 \times 10^{-10}$	$8.60 \times 10^{-4}$	197.094
rs1933802	6	105,365,891	$\mathbf C$	${\rm G}$	0.4536	$-0.0223$	0.0035	$2.57 \times 10^{-10}$	$8.69 \times 10^{-4}$	199.111
rs2876520	6	142,996,618	$\mathbf C$	G	0.5271	$-0.0230$	0.0036	$2.29 \times 10^{-10}$	$9.30 \times 10^{-4}$	213.024
rs725616	6	147,950,422	$\mathbf T$	$\mathsf{C}$	0.3644	0.0204	0.0036	$1.87 \times 10^{-8}$	$6.80 \times 10^{-4}$	155.705
rs2029865	6	165, 121, 844	A	T	0.4534	$-0.0201$	0.0035	$1.20 \times 10^{-8}$	$7.06 \times 10^{-4}$	161.743
rs3823624	$\overline{7}$	2,110,346	$\mathbf T$	${\bf C}$	0.8067	0.0272	0.0045	$1.99 \times 10^{-9}$	$8.13 \times 10^{-4}$	186.370
rs2043539	7	12,253,880	A	G	0.4177	0.0273	0.0035	$9.89 \times 10^{-15}$	$1.28 \times 10^{-3}$	292.880

### **TABLE 1** (Continued)



### **TABLE 1** (Continued)



Abbreviations: EAF, effect allele frequency; SE, standard error; SNP, single-nucleotide polymorphism.

Epidemiology Unit (IEU) GWAS database [\(https://gwas.](https://gwas.mrcieu.ac.uk/) [mrcieu.ac.uk/\)](https://gwas.mrcieu.ac.uk/) using the TwoSampleMR package (ver-sion 0.5.6) of R [\(www.r-project.org\)](http://www.r-project.org).<sup>24</sup> Summary data for the associations between MDD associated genetic variants with breast cancer were obtained from a GWAS with 228,951 women (including 122,977 breast cancer cases and 105,974 controls) of European ancestry performed by the Breast Cancer Association Consortium (BCAC) with study ID "ieu-a-1126".<sup>[17](#page-11-7)</sup> Of the patients, 69,501 were diagnosed as ER+ breast cancer and 21,468 as ER- breast cancer patients, and summary-level data were extracted with the study ID "ieu-a-1127" and "ieu-a-1128", respectively. When SNP for the exposed phenotype is missing in summary statistics of breast cancer risk, it was replaced with another proxy SNPs that is in high linkage disequilibrium  $(r^2 > 0.80)$  with validated IVs as determined using the 1000 genomes reference panel in European ancestry individuls. None of the IVs is significantly associated with breast cancer at the genome-wide significance  $(P < 5 \times 10^{-8})$ ; Table [S2\)](#page-12-0), indicating that these IVs met the assumption 3 of two-sample MR study design.

# **2.4** | **Genetic instrumental variables for confounders**

To account for the confounders including smoking, alcohol, age at menarche, education attainment, and incoming in MVMR study, we extracted SNPs that significantly associated

 $(P<5\times10^{-8})$  with smoking initiation (*n* = 607,291, GSCAN),<sup>18</sup> alcoholic drinks per week ( $n = 335,394$ , GSCAN),<sup>[18](#page-11-8)</sup> years of schooling (n = 293,723, SSGAC),<sup>19</sup> average household income before tax ( $n = 397,751$ , MRC-IEU UK Biobank),<sup>[20,21](#page-11-10)</sup> and age at menarche ( $n = 182,416$ , ReproGen) through querying the the IEU GWAS database [\(https://gwas.mrcieu.ac.uk\)](https://gwas.mrcieu.ac.uk) using the TwoSampleMR package.<sup>24</sup> We further pruned the results to exclude all SNPs with a pairwise linkage disequirium (LD)  $r^2$  > 0.001 in a 10,000kb window of genomic region as final IVs.

# **2.5** | **Statistical power calculation**

The pos hoc statistical power was calculated using an online tool at [http://cnsgenomics.com/shiny/mRnd/.](http://cnsgenomics.com/shiny/mRnd/)<sup>25</sup> The 92 SNP MDD instruments explained an estimated 2.448% of phenotypic variability. Given a type 1 error of 5%, we had sufficient power (>80%) when the expected *OR* was 1.090 for overall breast cancer (122,977 cases and 105,974 controls) per log-odds ratio increase with MDD in women. In the stratification analysis, we noticed the statistical power was 0.46 and 0.55 when the expected *OR* was 1.059 and 1.101 for ER+ (69,501 cases and 105,974 controls) and ER- (21,468 cases and 105,974 controls) breast cancer per log-odds ratio increase with MDD in women, respectively. Statistical power estimates for the 92 genome-wide significant SNP instruments by breast cancer subtypes are presented in Table [S3.](#page-12-0)



<span id="page-6-0"></span>**TABLE 2**

Estimates of the two-sample MR analyses for causal associations between MDD and overall breast cancer risk

 $TABLE 2$  Estimates of the two-sample MR analyses for causal associations between MDD and overall breast cancer risk

#### **2.6** | **Statistical analyses**

In order to determine the causal relationship between MDD and breast cancer risk, we used the conventional in verse variance weighted (IVW) method, fixed-effects IVW model, as well as the weighted median method to generate the main MR estimates. To account for the potential con founders in assessing MDD and breast cancer risk by the genetic IVs, we also performed the MVMR analysis adjust ing for smoking, alcohol intake, education atttainment, in come, and age at menarche. The genome-wide significant signals of these exposures assessed independent of MDDrelated variants  $(R^2 \text{ LD} < 0.1 \text{ based on the 1000} \text{ Genome})$ panel reference) were included in the models (Table [S4\)](#page-12-0).<sup>[26](#page-11-15)</sup> Due to the missing of 23andMe data, assocaitions between the IVs of confounders and MDD were assessed in metaanalysis of UKBiobank and PGC with 170,756 cases and 329,443 cases. $16$  We applied the MR-Egger method to determine the evidence of pleiotropic effects for these genetic IVs, which occurs when the variants has an effect on dis ease outside of its effect on the exposure (MDD) in MR stud ies.<sup>27</sup> Intercepts that significantly deviate from the origin indicating an evidence of potential pleiotropy effects. We also performed the MR pleiotropy residual sum and out lier (MR-PRESSO) test to identify any outlier variants that challenge the pleiotropy effects assumption  $(P < 0.05)$ , and the effect estimates were recalculated after remove these outliers.<sup>28</sup> Heterogeneity between the causal estimates inferred by the IVs was assessed using Cochran's Q statistic under the conventional IVW method. The leave-one-out sensitivity method was performed to evaluate whether the causality estimate was affected by a single SNP under conventional IVW model. $^{29}$  We also performed MR-Steiger directionality test to determine whether the assumption that exposure (MDD) causes outcome (breast cancer risk) is valid.<sup>30</sup> The TwoSampleMR package (version 0.5.6) implemented in R (version 3.6.3) was used to obtain the data and perform all the two-sample MR analyses. All statistical tests were two-sided and considered statistical significance when  $P < 0.05$ .<sup>24,31</sup> For multiple testings, the Bonferronicorrected *P* values below 0.017 (where  $\alpha = 0.05/3$  breast cancer outcomes) were considered strong evidence of asso ciations and *P* values between 0.017 and 0.05 as suggestive evidence of associations.

# **3** | **RESULTS**

*P*-heterogneity; SE, standard error; SNP, single-nucleotide polymorphism.

P-heterogneity; SE, standard error; SNP, single-nucleotide polymorphism.

### **3.1** | **Causal relationship between MDD and risks of breast cancer**

Based on the IVW model of two-sample MR studies, we found that women were associated with an 8.7% increased

risk of overall breast cancer ( $OR = 1.087$ ,  $95\% CI = 1.011$ – 1.170;  $P = 0.025$  $P = 0.025$  $P = 0.025$ ; Table 2) per log-odds ratio increment of MDD. The causal estimate inferred by individual IVs was shown in funnel plot (Figure [2](#page-7-0)) as well as the scatterplot (Figure [S1\)](#page-12-1). Significant heterogeneity was noticed for the causal estimates between IVs based on the IVW model (*P*-heterogeneity <0.001; Table [2\)](#page-6-0). Similarly, the causal relationship between MDD and breast cancer was also established under weighted median model (*OR* = 1.091, 95% *CI* = 1.002–1.189; *P* = 0.045; Table [2](#page-6-0) and Figure [2](#page-7-0)) as suggested by the MR regression slopes in scatterplot (Figure [S1\)](#page-12-1). However, the association did not reach the statistical significance (*P*<0.017) for multiple testings. The leave-one-out analyses identified no individual IV largely affect the causal magnitude between MDD and the risks of total breast cancer under the IVW model (Figure [S2](#page-12-1)). The MR-Steiger directionality test suggested the causal direction by MDD on overall breast cancer risk is valid (*P*<0.001).

According to MR-Egger test, there was no significant pleiotropy effects of the genetic IVs ( $P = 0.946$ ). When we applied the MR-PRESSO analysis, a stronger association was observed between the MDD and breast cancer risk after excluding potential outlying pleiotropic IVs (*OR* = 1.090, 95%  $CI = 1.017 - 1.168$ ), which reached the multiple testing significance threshold  $(P = 0.016)$ . Further, the results from the MV analyses adjusting for potential confounders including regulatory smoking, alcoholic drinks per week, education attainment, household income (before tax), and age at menarche also suggested that MDD is associated with increased breast cancer ( $OR = 1.118$ , 95%  $CI = 1.010 1.237; P = 0.031$ ).

# **3.2** | **Subtype stratification analysis the causal effects of MDD on ER+ or ERbreast cancer**

To further explore the causal relationships between MDD and the risks of breast cancer of different pathological subtype, we performed the two-sample MR studies based on the summary level data form the GWASs performed in ER+ and ER- breast cancer patients. No statistically significant association was observed between MDD and ER+ breast cancer as suggested by the IVW methods ( $OR = 1.059$ , 95%  $CI = 0.969 - 1.157$ ;  $P = 0.208$ , weighted median methods and MR-PRESSO

REN ET AL. **1991**<br>**Cancer Medicine \_\_\_\_WII FY**  $\frac{1991}{2}$ 

corrected testing (Table [2\)](#page-6-0). The causal estimate inferred by individual IVs was shown in scatterplot (Figure [S3](#page-12-1)) as well as the funnel plot (Figure [S4](#page-12-1)). There was a significant heterogeneity for the causal estimates between IVs for ER+ breast cancer based on the conventional IVW model (*P*-heterogeneity <0.001) and no significant pleiotropy effects of the IVs in assessing causality association between MDD and ER+ breast cancer risk were noticed. The leave-one-out analyses identified no individual IV largely affect the causal relationship between MDD and ER+ breast cancer under the IVW model (Figure [S5\)](#page-12-1). Multivariable MR analysis adjusting for potential confounders found a marginal causality relationship between MDD and  $ER+$  breast cancer ( $OR = 1.098$ , 95% *CI* = 0.984-1.227;  $P = 0.093$ .

For ER- breast cancer, we noticed a statistically marginal causality relationship with MDD as suggested by IVW (*OR* = 1.163, 95% *CI* = 0.994–1.361; *P* = 0.060), weighted median (*OR* = 1.100, 95% *CI* = 0.989–1.224;  $P = 0.078$ ), and MR-PRESSO analyses ( $OR = 1.101, 95\%$  $CI = 0.990 - 1.224$ ;  $P = 0.081$ ; Table [2\)](#page-6-0). Causal estimates inferred by individual IVs was shown in scatterplot (Figure [S6](#page-12-1)) as well as the funnel plot (Figure [S7\)](#page-12-1). No significant heterogeneity was noticed between the IVs under the IVW model ( $P$ -heterogeneity = 0.303) and no evidence for the directional horizontal pleiotropy was identified in the MR-Egger regression analysis (Table [2\)](#page-6-0). We performed leave-one-out sensitivity analysis based on the conventional IVW model, and no individual IV largely affect the causal estimates between MDD and the risk of ER- breast cancer was noticed (Figure [S8\)](#page-12-1). The multivariable MR also suggesting a statistically marginal causality between MDD and ER- breast cancer  $(OR = 1.129, 95\% CI = 0.982 - 1.297; P = 0.089).$ 

# **4** | **DISCUSSION**

By using the summarized statistical data from public GWASs, we performed a two-sample MR study to evaluate the causal associations between MDD and breast cancer risk. We found a statistical causal association between a genetic predisposition of MDD and the risk of overall breast cancer. As breast cancer with distinct pathological status may caused by different etiology factors and show varied responses to clinical treatments, $32$  we performed stratification analysis between MDD and ER+ or

<span id="page-7-0"></span>**FIGURE 2** Forest plot of the two-sample MR analysis results between MDD and the risk of breast cancer using conventional inverse variance weighted (IVW) and weighted median methods. Each black dot and horizontal line represented the causal effect obtained by Wald ratio method and 95% *CI* for individual IV, respectively. The red dots and horizontal lines at the bottom represent the overall causal effects obtained by different methods and 95% *CI*, respectively. X-axis is shown in natural logarithmic scale.



MR effect size for MDD on overall breast cancer, OR (95% CI).

ER- breast cancer; however, no statistically significant casuality was noticed due to relatively lower statistical power  $(<80\%)$  in stratification studies. Whether MDD has different influences on ER+ or ER- breast cancer also needs to be elucidated with more studies.

Previous observational studies have suggested that MDD might be associated with increased breast cancer  $risk^{11,33,34}$ ; however, the results were not always consistent.<sup>10</sup> Discrepancy results may be due to the different diagnostic criteria of depression, the different sample size of people observed in the cohort, the different following-up time, $35,36$  and other confounding factors.<sup>11,37</sup> As most cancers have a longer incubation period, it is difficult for traditional epidemiology to establish the causal relationship between MDD and breast cancer risk. The loss of following-up may vary depending on the state of cancer or depression, which may lead to biased estimates of association and standard error. Wrong misspecification in risk models may be another limitation, as other important confounding factors such as genetic susceptibility or congenital, and unmeasured psychological characteristics that have not been explained. Measurement methods of depression may influent the assessment of the exposure factors. The most commonly used Diagnostic Interview Schedule (DIS) test for severe depression has a relatively high specificity (more than 95% of test negatives are true negatives), but the estimated sensitivity is often difficult to exceed 40% (about 30% to 40%).<sup>[38,39](#page-11-22)</sup> Therefore, misclassification of depression may contribute to the null association results in specific studies. Some studies often use the self-reported cancer status of study participants as the outcome inclusion criterion. This measurement error will also weaken the estimation of association between MDD and breast cancer risk. $11$  At least but not the last, due to high median age at onset of breast cancer, related research is trapped in personnel, funding, and other factors often make it difficult to track the entire cohort or until the onset of cancer, which may distort the associations be-tween MDD and breast cancer risk.<sup>[40](#page-11-23)</sup> Unlike these traditional epidemiological studies, we applied the two-sample MR, which used genetic variants do not influence the outcomes via a different biological pathway from the exposure of interest, to control confounders. We conducted multiple sensitivity analyses to test for the influences of pleiotropy on our causal estimates, and robust results were observed according to these various tests. Moreover, design of MR analysis along with the MR Steiger test of directionality suggest a low probability of reverse causality in assessing the influences of MDD on breast cancer risk.<sup>30</sup>

Our two-sample MR study supports the hypothesis that depression can lead to higher risk of breast cancer.<sup>[11](#page-11-1)</sup> Multiple biological mechanisms are hypothesized to mediate the potential beneficial role of MDD on breast

 **|** REN et al. **1993**

cancer development. MDD is associated with higher circulating levels of inflammatory cytokines, such as interleukin (IL)-6, IL-8, IL-1β, TNF-α, soluble IL-2 receptor (sIL-2R), and C-reactive protein (CRP), which promote cellular proliferation in breast tissue and have also been linked to development of breast cancer. $41,42$ Continuous chronic inflammation status may underlie the cancer development caused by  $MDD.<sup>43</sup> MDD$ have also been associated with a dysfunctional activation of the hypothalamic–pituitary adrenal (HPA) axis, which are risk factors for breast cancer development.<sup>[44](#page-11-26)</sup> As a stress hormone, cortisol is released from the adrenal glands when the adrenocorticotropic hormone in the blood is elevated. When cortisol is under continuous chronic stress and abnormal activation of the HPA axis, its metabolic level may be dysregulated. Therefore, the normal cortisol response pattern is destroyed under a long-term stressor and the ability to respond to stress is destroyed.<sup>[45](#page-12-2)</sup> Cortisol also plays important roles in the activation of cell growth and cell cycle regulation signaling.[46–48](#page-12-3) Related studies indicated that flattening cortisol levels throughout the day is significantly associated with an increased risk of breast cancer.<sup>49,50</sup> Moreover, MDD has also been associated with a decreased immunosurveillance with increased risk of breast cancer. The depression augments the production of autoantibodies that against 16α-OHE1-ER through the generation of inflammatory conditions.<sup>51</sup> Finally, emerging evidence suggests that the gut microbiome may play an important role in the MDD and breast cancer relationship. MDD patients have less *Faecalibacterium*, which inversely correlated with the severity of the depression,<sup>[52](#page-12-6)</sup> and the dysbiosis of the gut microbiome has been associated with increased risks of breast and cancer.<sup>[53](#page-12-7)</sup>

An important assumption of MR is that the genetic IVs should not influence the outcome through a different biological pathway from the exposure of tested. Previous epidemiology studies suggested that MDD may lead to higher rates of smoking and alcohol intake. $54,55$ Meta-analysis of the observational epidemiological studies had suggested that smoking may weakly increase the breast cancer risk.<sup>[56](#page-12-9)</sup> Dimou et al. have observed a genetic predisposition to higher lifetime amount of smoking was positively associated with overall breast cancer  $(OR = 1.13, 95\% CI = 1.00 - 1.26; P = 0.04)$ . MR study performed by Howard et al. $^{16}$  reported that MDD is causally associated with breast cancer risk factors including smoking and age at menarche. Although the MR study did not identify an association between genetically predicted alcohol consumption with breast cancer risk, the prospective cohort studies have suggested that higher alcohol intake was associated increased breast cancer risk.[57](#page-12-10) Moreover, socio-economic status (SES) such as education level and household income were reported to be negatively associated with depressive symptoms,  $58,59$ and SES is associated with breast cancer susceptibility and prognosis of breast cancer patients in clinic. $60$  To account for such potential confounders that may mediate the causality between MDD and breast cancer, we performed MVMR to evaluate the causality relationship between MDD and breast cancer risk with the adjustments for regulatory smoking, alcohol intake, years of schooling, household income, and age at menarche. The MVMR results also observed an increment of overall breast cancer risk for women with a genetic predisposition, suggesting that the causality between MDD and breast cancer risk is robust and it was unlikely caused by confounders.

There were several advantages for the current twosample MR studies, first, because alleles follow the principle of random distribution when forming gametes during meiosis, the causal relationship between genotype and disease in MR studies will not be distorted by confounding factors, which is the main limitation of traditional observation studies. Second, it is much easier to use aggregated level data of public GWAS genetic consortia data to assess the causality relationships between MDD and breast cancer risk rather than the prospective cohorts studies or RCTs. The two-sample MR study performed here required less time and expenses compare with common epidemiological studies. However, several limitations in the current studies need to be acknowledged in current study. First, the genetic IVs used in MR analysis are usually weak instruments, and only a small proportion of exposure variance was identified by the genetic IVs. The non-causal connection between MDD and breast cancer in the stratification studies may be due to the low statistical power. Second, other confounding factors (such as population stratification and exposure time) may also affect the results. Third, because this study relies on aggregated GWAS data from people of European descent, the promotion of this result should be validated in other races.

In conclusion, the current MR analysis showed that MDD was causally associated with the increase of overall breast cancer. This result support the view that early intervention should be conducted on the women with MDD to reduce the risk of breast cancer; however, the results need to be confirmed using other study designs, including prospective cohort studies and large-scale intervention trials.

### **AUTHOR CONTRIBUTIONS**

QR, FL, SG, and PC designed, researched, and revised the manuscript strictly. QR, FL, and PC analyzed the data set and wrote the manuscript. All authors approved the final version to be published.

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## **CONFLICT OF INTEREST**

The authors declare that the research was conducted in the absence of any commercial or financial relationships that could be construed as a potential conflict of interest.

### **DATA AVAILABILITY STATEMENT**

The data that support the findings of this study are openly available in the IEU GWAS database [\(https://gwas.mr](https://gwas.mrcieu.ac.uk)[cieu.ac.uk\)](https://gwas.mrcieu.ac.uk) that comprises the harmonised complete GWAS summary datasets, which can be queried using the R package of TwoSampleMR.

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### **REFERENCE**

- <span id="page-10-0"></span>1. Veronesi U, Boyle P, Goldhirsch A, Orecchia R, Viale G. Breast cancer. *Lancet*. 2005;365(9472):1727-1741.
- <span id="page-10-1"></span>2. Bray F, Ferlay J, Soerjomataram I, Siegel RL, Torre LA, Jemal A. Global cancer statistics 2018: GLOBOCAN estimates of incidence and mortality worldwide for 36 cancers in 185 countries. *CA Cancer J Clin*. 2018;68(6):394-424.
- <span id="page-10-2"></span>3. GBD DALYs, Collaborators HALE. Global, regional, and national disability-adjusted life-years (DALYs) for 359 diseases and injuries and healthy life expectancy (HALE) for 195 countries and territories, 1990-2017: a systematic analysis for the global burden of disease Study 2017. *Lancet*. 2018;392(10159):1859-1922.
- <span id="page-10-3"></span>4. Buja A, Pierbon M, Lago L, Grotto G, Baldo V. Breast cancer primary prevention and diet: an umbrella review. *Int J Environ Res Public Health*. 2020;17(13):4731.
- <span id="page-10-4"></span>5. Smith K. Mental health: a world of depression. *Nature*. 2014;515(7526):181.
- <span id="page-10-5"></span>6. Malhi GS, Mann JJ. Depression. *Lancet*. 2018;392(10161): 2299-2312.
- <span id="page-10-6"></span>7. Wang YH, Li JQ, Shi JF, et al. Depression and anxiety in relation to cancer incidence and mortality: a systematic review and metaanalysis of cohort studies. *Mol Psychiatry*. 2020;25(7):1487-1499.
- 8. Gallo JJ, Armenian HK, Ford DE, Eaton WW, Khachaturian AS. Major depression and cancer: the 13-year follow-up of the Baltimore epidemiologic catchment area sample (United States). *Cancer Causes Control*. 2000;11(8):751-758.
- 9. Brewer JK. Behavioral genetics of the depression/cancer correlation: a look at the Ras oncogene family and the 'cerebral diabetes paradigm'. *J Mol Neurosci*. 2008;35(3):307-322.

REN ET AL. **1995**<br>**Cancer Medicine \_\_\_\_\_WILEY\_\_\_<sup>1995</sup>** 

- 
- <span id="page-11-0"></span>10. Sun HL, Dong XX, Cong YJ, et al. Depression and the risk of breast cancer: a meta-analysis of cohort studies. *Asian Pac J*
- <span id="page-11-1"></span>*Cancer Prev*. 2015;16(8):3233-3239. 11. Gross AL, Gallo JJ, Eaton WW. Depression and cancer risk: 24 years of follow-up of the Baltimore epidemiologic catchment area sample. *Cancer Causes Control*. 2010;21(2):191-199.
- <span id="page-11-2"></span>12. Smith GD, Ebrahim S. 'Mendelian randomization': can genetic epidemiology contribute to understanding environmental determinants of disease? *Int J Epidemiol*. 2003;32(1):1-22.
- <span id="page-11-3"></span>13. Lawlor DA, Harbord RM, Sterne JA, Timpson N, Davey SG. Mendelian randomization: using genes as instruments for making causal inferences in epidemiology. *Stat Med*. 2008;27(8):1133-1163.
- <span id="page-11-4"></span>14. Bowden J, Holmes MV. Meta-analysis and mendelian randomization: a review. *Res Synth Methods*. 2019;10(4):486-496.
- <span id="page-11-5"></span>15. Kawashiri MA, Tada H, Nomura A, Yamagishi M. Mendelian randomization: its impact on cardiovascular disease. *J Cardiol*. 2018;72(4):307-313.
- <span id="page-11-6"></span>16. Howard DM, Adams MJ, Clarke TK, et al. Genome-wide metaanalysis of depression identifies 102 independent variants and highlights the importance of the prefrontal brain regions. *Nat Neurosci*. 2019;22(3):343-352.
- <span id="page-11-7"></span>17. Michailidou K, Lindstrom S, Dennis J, et al. Association analysis identifies 65 new breast cancer risk loci. *Nature*. 2017;551(7678):92-94.
- <span id="page-11-8"></span>18. Liu M, Jiang Y, Wedow R, et al. Association studies of up to 1.2 million individuals yield new insights into the genetic etiology of tobacco and alcohol use. *Nat Genet*. 2019;51(2):237-244.
- <span id="page-11-9"></span>19. Okbay A, Beauchamp JP, Fontana MA, et al. Genome-wide association study identifies 74 loci associated with educational attainment. *Nature*. 2016;533(7604):539-542.
- <span id="page-11-10"></span>20. Sudlow C, Gallacher J, Allen N, et al. UKbiobank: an open access resource for identifying the causes of a wide range of complex diseases of middle and old age. *PLoS Med*. 2015;12(3):e1001779.
- 21. Rosoff DB, Kaminsky ZA, McIntosh AM, Davey Smith G, Lohoff FW. Educational attainment reduces the risk of suicide attempt among individuals with and without psychiatric disorders independent of cognition: a bidirectional and multivariable mendelian randomization study with more than 815,000 participants. *Transl Psychiatry*. 2020;10(1):388.
- <span id="page-11-11"></span>22. Perry JR, Day F, Elks CE, et al. Parent-of-origin-specific allelic associations among 106 genomic loci for age at menarche. *Nature*. 2014;514(7520):92-97.
- <span id="page-11-12"></span>23. Palmer TM, Lawlor DA, Harbord RM, et al. Using multiple genetic variants as instrumental variables for modifiable risk factors. *Stat Methods Med Res*. 2012;21(3):223-242.
- <span id="page-11-13"></span>24. Hemani G, Zheng J, Elsworth B, et al. The MR-base platform supports systematic causal inference across the human phenome. *Elife*. 2018;7:e34408.
- <span id="page-11-14"></span>25. Brion MJ, Shakhbazov K, Visscher PM. Calculating statistical power in mendelian randomization studies. *Int J Epidemiol*. 2013;42(5):1497-1501.
- <span id="page-11-15"></span>26. Dimou N, Yarmolinsky J, Bouras E, et al. Causal effects of lifetime smoking on breast and colorectal cancer risk: mendelian randomization study. *Cancer Epidemiol Biomarkers Prev*. 2021;30(5):953-964.
- <span id="page-11-16"></span>27. Burgess S, Thompson SG. Interpreting findings from mendelian randomization using the MR-egger method. *Eur J Epidemiol*. 2017;32(5):377-389.
- <span id="page-11-17"></span>28. Verbanck M, Chen CY, Neale B, Do R. Detection of widespread horizontal pleiotropy in causal relationships inferred from mendelian randomization between complex traits and diseases. *Nat Genet*. 2018;50(5):693-698.
- <span id="page-11-18"></span>29. Burgess S, Bowden J, Fall T, Ingelsson E, Thompson SG. Sensitivity analyses for robust causal inference from mendelian randomization analyses with multiple genetic variants. *Epidemiology*. 2017;28(1):30-42.
- <span id="page-11-19"></span>30. Hemani G, Tilling K, Davey SG. Orienting the causal relationship between imprecisely measured traits using GWAS summary data. *PLoS Genet*. 2017;13(11):e1007081.
- 31. Yavorska OO, Burgess S. MendelianRandomization: an R package for performing mendelian randomization analyses using summarized data. *Int J Epidemiol*. 2017;46(6):1734-1739.
- <span id="page-11-20"></span>32. Hui L, Huang T, Lian J, et al. Potential prognostic value of clinical characteristics, hormone status and major depressive disorder in breast cancer. *Future Oncol*. 2017;13(17):1493-1503.
- 33. Chen YH, Lin HC. Increased risk of cancer subsequent to severe depression--a nationwide population-based study. *J Affect Disord* 2011; 131(1–3): 200–6, Increased risk of cancer subsequent to severe depression—a nationwide population-based study.
- 34. McGee R, Williams S, Elwood M. Depression and the development of cancer: a meta-analysis. *Soc Sci Med*. 1994;38(1):187-192.
- <span id="page-11-21"></span>35. Oerlemans ME, van den Akker M, Schuurman AG, Kellen E, Buntinx F. A meta-analysis on depression and subsequent cancer risk. *Clin Pract Epidemiol Ment Health*. 2007;3:29.
- 36. Jacobs JR, Bovasso GB. Early and chronic stress and their relation to breast cancer. *Psychol Med*. 2000;30(3):669-678.
- 37. Reeves KW, Okereke OI, Qian J, Tamimi RM, Eliassen AH, Hankinson SE. Depression, antidepressant use, and breast cancer risk in pre- and postmenopausal women: a prospective Cohort Study. *Cancer Epidemiol Biomarkers Prev*. 2018;27(3):306-314.
- <span id="page-11-22"></span>38. Eaton WW, Neufeld K, Chen LS, Cai G. A comparison of self-report and clinical diagnostic interviews for depression: diagnostic interview schedule and schedules for clinical assessment in neuropsychiatry in the Baltimore epidemiologic catchment area follow-up. *Arch Gen Psychiatry*. 2000;57(3):217-222.
- 39. Anthony JC, Folstein M, Romanoski AJ, et al. Comparison of the lay diagnostic interview schedule and a standardized psychiatric diagnosis. Experience in eastern Baltimore. *Arch Gen Psychiatry*. 1985;42(7):667-675.
- <span id="page-11-23"></span>40. Brenner H. Long-term survival rates of cancer patients achieved by the end of the 20th century: a period analysis. *Lancet*. 2002;360(9340):1131-1135.
- <span id="page-11-24"></span>41. Liu Y, Ho RC, Mak A. Interleukin (IL)-6, tumour necrosis factor alpha (TNF-alpha) and soluble interleukin-2 receptors (sIL-2R) are elevated in patients with major depressive disorder: a meta-analysis and meta-regression. *J Affect Disord*. 2012;139(3):230-239.
- 42. Dowlati Y, Herrmann N, Swardfager W, et al. A metaanalysis of cytokines in major depression. *Biol Psychiatry*. 2010;67(5):446-457.
- <span id="page-11-25"></span>43. Reuter S, Gupta SC, Chaturvedi MM, Aggarwal BB. Oxidative stress, inflammation, and cancer: how are they linked? *Free Radic Biol Med*. 2010;49(11):1603-1616.
- <span id="page-11-26"></span>44. Kennis M, Gerritsen L, van Dalen M, Williams A, Cuijpers P, Bockting C. Prospective biomarkers of major depressive

disorder: a systematic review and meta-analysis. *Mol Psychiatry*. 2020;25(2):321-338.

- <span id="page-12-2"></span>45. McEwen BS. The neurobiology of stress: from serendipity to clinical relevance. *Brain Res*. 2000;886(1–2):172-189.
- <span id="page-12-3"></span>46. Lutgendorf SK, Cole S, Costanzo E, et al. Stress-related mediators stimulate vascular endothelial growth factor secretion by two ovarian cancer cell lines. *Clin Cancer Res*. 2003;9(12):4514-4521.
- 47. Pulaski L, Kania K, Ratajewski M, Uchiumi T, Kuwano M, Bartosz G. Differential regulation of the human MRP2 and MRP3 gene expression by glucocorticoids. *J Steroid Biochem Mol Biol*. 2005;96(3–4):229-234.
- 48. Thaker PH, Han LY, Kamat AA, et al. Chronic stress promotes tumor growth and angiogenesis in a mouse model of ovarian carcinoma. *Nat Med*. 2006;12(8):939-944.
- <span id="page-12-4"></span>49. Ticher A, Haus E, Ron IG, Sackett-Lundeen L, Ashkenazi IE. The pattern of hormonal circadian time structure (acrophase) as an assessor of breast-cancer risk. *Int J Cancer*. 1996;65(5):591-593.
- 50. Su F, Ouyang N, Zhu P, et al. Psychological stress induces chemoresistance in breast cancer by upregulating mdr1. *Biochem Biophys Res Commun*. 2005;329(3):888-897.
- <span id="page-12-5"></span>51. Khan WA, Khan MWA, Sherwani S, Siddiqui WA. Depression enhanced the production of autoantibodies against 16alphahydroxyestrone-estrogen receptor adduct in breast cancer. *Int Immunopharmacol*. 2019;66:251-259.
- <span id="page-12-6"></span>52. Jiang H, Ling Z, Zhang Y, et al. Altered fecal microbiota composition in patients with major depressive disorder. *Brain Behav Immun*. 2015;48:186-194.
- <span id="page-12-7"></span>53. Helmink BA, Khan MAW, Hermann A, Gopalakrishnan V, Wargo JA. The microbiome, cancer, and cancer therapy. *Nat Med*. 2019;25(3):377-388.
- <span id="page-12-8"></span>54. Sarris J, Thomson R, Hargraves F, et al. Multiple lifestyle factors and depressed mood: a cross-sectional and longitudinal analysis of the UKbiobank (N = 84,860). *BMC Med*. 2020;18(1):354.
- 55. Lu W, Kim JC, Yoon AS, Yun KK, Solomon P. Trends and disparities in unmet treatment needs for co-occurring depression
- <span id="page-12-9"></span>56. Macacu A, Autier P, Boniol M, Boyle P. Active and passive smoking and risk of breast cancer: a meta-analysis. *Breast Cancer Res Treat*. 2015;154(2):213-224.
- <span id="page-12-10"></span>57. Ong JS, Derks EM, Eriksson M, et al. Evaluating the role of alcohol consumption in breast and ovarian cancer susceptibility using population-based cohort studies and twosample mendelian randomization analyses. *Int J Cancer*. 2021;148(6):1338-1350.
- <span id="page-12-11"></span>58. Li L, Sun W, Luo J, Huang H. Associations between education levels and prevalence of depressive symptoms: NHANES (2005- 2018). *J Affect Disord*. 2022;301:360-367.
- 59. Hinata A, Kabasawa K, Watanabe Y, et al. Education, household income, and depressive symptoms in middle-aged and older Japanese adults. *BMC Public Health*. 2021;21(1):2120.
- <span id="page-12-12"></span>60. Klassen AC, Smith KC. The enduring and evolving relationship between social class and breast cancer burden: a review of the literature. *Cancer Epidemiol*. 2011;35(3):217-234.

## <span id="page-12-1"></span>**SUPPORTING INFORMATION**

<span id="page-12-0"></span>Additional supporting information can be found online in the Supporting Information section at the end of this article.

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