

Supplementary

1 Equations

Original CelFiE equations

Posterior distribution:

$$\begin{aligned} p_{ntm1}(\alpha, \beta) &:= p_{ntmc}(\alpha, \beta) && \text{if } x_{nmc} = 1 \\ &= \frac{\beta_{tm}\alpha_{nt}}{\sum_k \beta_{kt}\alpha_{nk}} \\ p_{ntm0}(\alpha, \beta) &:= p_{ntmc}(\alpha, \beta) && \text{if } x_{nmc} = 0 \\ &= \frac{(1 - \beta_{tm})\alpha_{nt}}{\sum_k (1 - \beta_{kt})\alpha_{nk}} \end{aligned} \tag{S1}$$

α and β update formula:

$$\alpha_{nt} = \frac{\sum_m (x_{nm}p_{ntm1} + (D_{nm}^X - x_{nm})p_{ntm0})}{\sum_{km} (x_{nm}p_{nkm1} + (D_{nm}^X - x_{nm})p_{nkm0})} \tag{S2}$$

$$\beta_{tm} = \frac{\sum_n p_{ntm1}X_{nm} + nY_{tm}}{\sum_n p_{ntm0}(D_{nm}^X - X_{nm}) + nD_{tm}^Y + \sum_n p_{ntm1}X_{nm}} \tag{S3}$$

Log-likelihood formulation:

$$\begin{aligned} Q(\alpha, \beta) &= \sum_{n,t,m} [(Y_{tm} + p_{ntm1}X_{nm}) \log(\beta_{tm}) + (D_{tm}^Y - Y_{tm} + p_{ntm0}(D_{nm}^X - X_{nm})) \log(1 - \beta_{tm})] \\ &\quad + \sum_{n,t,m} (X_{nm}p_{ntm1} + (D_{nm}^X - X_{nm})p_{ntm0}) \log \alpha_{nt} \end{aligned} \tag{S4}$$

Derivation of full data log-likelihood

$$\begin{aligned}
Q(\alpha, \hat{\beta}) &:= \mathbb{E}_{z|\hat{X}, \alpha, \hat{\beta}} \log P(\hat{X}, z, Y|\alpha, \beta) \\
&= \mathbb{E}_{z|\hat{X}, \alpha, \hat{\beta}} \left(\log P(\hat{X}|z, \hat{\beta}) + \log P(z|\alpha) + \log P(\hat{Y}|\hat{\beta}) \right) \\
&= \sum_{n,t,m,c} \mathbb{E}_{z|\hat{X}, \alpha, \hat{\beta}} \left[z_{ntmc} \sum_i \hat{x}_{nmci} \log \hat{\beta}_{tmi} + z_{ntmc} \log \alpha_{nt} \right] \\
&+ \sum_{n,t,m} \left(\log \left(\sum_i \hat{Y}_{tmi}! \right) - \sum_i \log(\hat{Y}_{tmi}!) + \sum_i \hat{Y}_{tmi} \log \hat{\beta}_{tmi} \right) \\
&= \sum_{n,t,m,c} \tilde{p}_{ntmc} \left[\sum_i \hat{x}_{nmci} \log \hat{\beta}_{tmi} + \log \alpha_{nt} \right] \\
&+ \sum_{n,t,m} \left(\log \left(\sum_i \hat{Y}_{tmi}! \right) - \sum_i \log(\hat{Y}_{tmi}!) + \sum_i \hat{Y}_{tmi} \log \hat{\beta}_{tmi} \right) \\
&= \sum_{n,t,m} \left[\sum_i p_{ntmi} \hat{x}_{nmi} \log \hat{\beta}_{tmi} + \sum_i p_{ntmi} \hat{x}_{nmi} \log \alpha_{nt} \right] \\
&+ \sum_{n,t,m} \left[\log \left(\sum_i \hat{Y}_{tmi}! \right) - \sum_i \log(\hat{Y}_{tmi}!) + \sum_i \hat{Y}_{tmi} \log \hat{\beta}_{tmi} \right] \\
&= \sum_{n,t,m,i} \left((p_{ntmi} \hat{x}_{nmi} + \hat{Y}_{tmi}) \log \hat{\beta}_{tmi} \right) + \sum_{n,t,m,i} p_{ntmi} \hat{x}_{nmi} \log \alpha_{nt} + n \sum_{t,m} \left[\log \left(\sum_i \hat{Y}_{tmi}! \right) - \sum_i \log(\hat{Y}_{tmi}!) \right]
\end{aligned} \tag{S5}$$

Derivation of α and $\hat{\beta}$ update formulas

Maximization of the log-likelihood w.r.t α and $\hat{\beta}$ can be done using the following fact that for a probability simplex $S_K \subset \mathbb{R}^K$ and any $a \in \mathbb{R}_{++}^K$:

$$\arg \max_{p \in S_K} \sum_k a_k \log p_k = (a_1, \dots, a_K) / \sum_{k=1}^K a_k$$

To derive α_t , we let $a_t = \sum_i p_{tmi} \hat{x}_i$ s.t.

$$\alpha_{nt} = \frac{\sum_{m,i} p_{ntmi} \hat{x}_{nmi}}{\sum_{m,t,i} p_{ntmi} \hat{x}_{nmi}} \tag{S6}$$

For $\hat{\beta}_{tmi}$ we let $a_i = p_{tmi} \hat{x}_i + \hat{Y}_{tmi}$ s.t.

$$\hat{\beta}_{tmi} = \frac{\sum_n (p_{ntmi} \hat{x}_{nmi} + \hat{Y}_{tmi})}{\sum_{n,i} (p_{ntmi} \hat{x}_{nmi} + \hat{Y}_{tmi})} \tag{S7}$$

2 Supplementary tables

Tab. S1: WGBS cell type data and sources

Cell type	Database	Sample 1	Sample 2
CD4-positive, alpha-beta T cell	Blueprint	S007G7	S007DD
CD8-positive, alpha-beta T cell	Blueprint	C003VO	C00256
endothelial cell of umbilical vein (resting)	Blueprint	S00DCS	S00BJM
monocyte	Blueprint	S01MAPA1	S01E03A1
erythroblast	Blueprint	S002S3	S002R5
macrophage	Blueprint	S0022I	S00390
mature eosinophil	Blueprint	S00V65	S006XE
memory B cell	Blueprint	C003N3	S017RE51
cytotoxic CD56-dim natural killer cell	Blueprint	C006G5	C002CT
mature neutrophil	Blueprint	C0010K	C000S5
conventional dendritic cell	Blueprint	S00CP651	S00D71
adipose	ENCODE	ENCFF318AMC	ENCFF477GKI
HepG2	ENCODE	ENCFF847OWL	ENCFF064GJQ
pancreas	ENCODE	ENCFF753ZMQ	ENCFF500DKA
small intestine	ENCODE	ENCFF266NGW	ENCFF122LEF
spleen	ENCODE	ENCFF550FZT	ENCFF333OHK
stomach	ENCODE	ENCFF435SPL	ENCFF497YOO
tibial nerve	ENCODE	ENCFF843SYR	ENCFF699KTW
skeletal muscle myoblast primary cell	ENCODE	ENCFF774GXJ	-
placenta	ENCODE	ENCFF239DTC	-

Tab. S2: Correlation between cell types proportions estimated by CelFEER and the true cell type proportions after adding simulated noise.

Probability of flipping methylation status	Correlation
0	0.9310 +- 0.0041
1/1000	0.9480 +- 0.0028
1/100	0.9591 +- 0.0017
1/50	0.9577 +- 0.0011
1/40	0.9485 +- 0.0004
1/30	0.9260 +- 0.0002
1/20	0.8093 +- 0.0001
1/10	0.1843 +- 0.0001

Tab. S3: Average distance between cell types for different genomic regions.

Marker preselection	CpG islands	MHB	None
Distance in reference data	0.21	0.43	0.58
Distance in input data	0.20	0.59	0.71

Tab. S4: Correlation (mean and standard deviation) between true and estimated cell type proportions for using MHBs and for using the entire genome as input for marker selection.

Marker preselection	MHB	None
Read based correlation	0.90 +- 0.07	0.94 +-0.04
Beta values correlation	0.82 +- 0.12	0.86 +- 0.09

Tab. S5: Correlation (mean and standard deviation) between the true simulated cell type proportions and those estimated by the compared methods.

Method	Correlation
CelFEER	0.94 +- 0.04
CelFiE	0.86 +- 0.09
Least squares regression	0.73 +- 0.16
Houseman et al.	0.69 +- 0.19

Tab. S6: Correlation of CelFiE (mean and standard deviation) and Houseman before and after using our improved method for finding markers.

	Original markers	Improved markers
CelFiE	0.24 +- 0.24	0.87 +- 0.05
Houseman	0.35	0.71

3 Supplementary figures

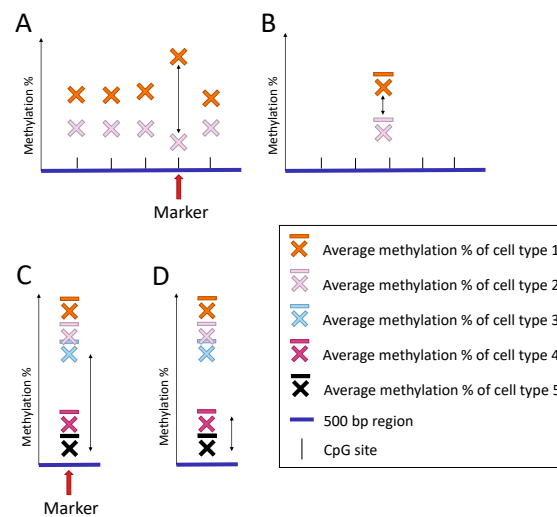


Fig. S1: Illustration of the two principal changes to the approach for finding markers in the genome. The arrows indicate the difference between cell types. Figures (A) and (C) illustrate how the markers are found originally, and Figures (B) and (D) how they are found after improvements. Figures (A) and (B) show how measuring the distance between single CpG sites (A) results in different markers than measuring the distance between 500 bp regions (B). Cell types 1 and 2 do not have a large distance when regarding their average over the entire region, making this region an unsuitable marker. Figures (C) and (D) show that the distance from the median cell type (C) is different from the distance from the min cell type (D). Using the median would result in a marker that does not differentiate well between cell types 4 and 5.

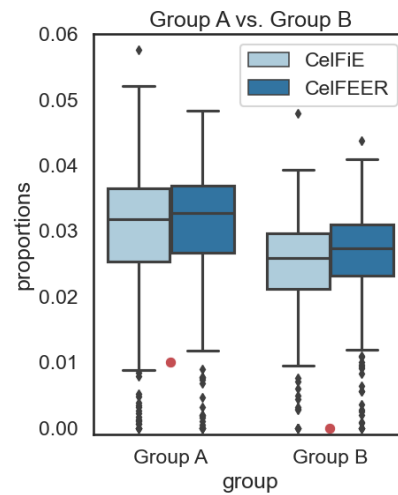


Fig. S2: Estimates of the proportion of a rare cell type (1%) that is present in group A but not in group B, estimated over 50 replicate runs using CelFiE and CelFEER. Only the estimated and true proportions of this rare cell type are plotted. The true proportions are represented by the red dots.

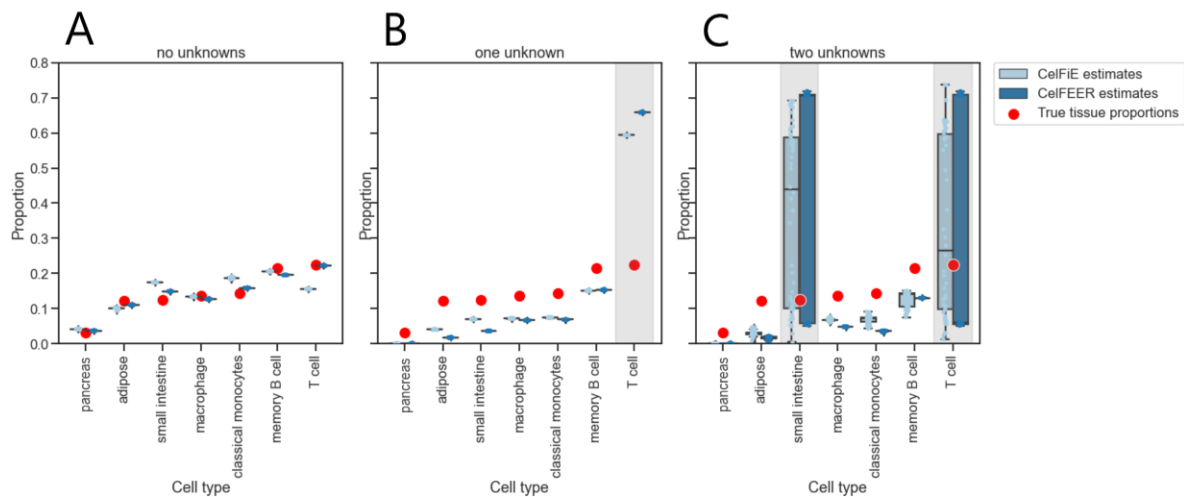


Fig. S3: Cell type proportions estimated by CelFiE and CelFEER for zero, one and two unknowns respectively. The boxplots visualize the estimated proportions of 50 replicates for a randomly chosen individual. On top of the boxplots, the individual datapoints are plotted.

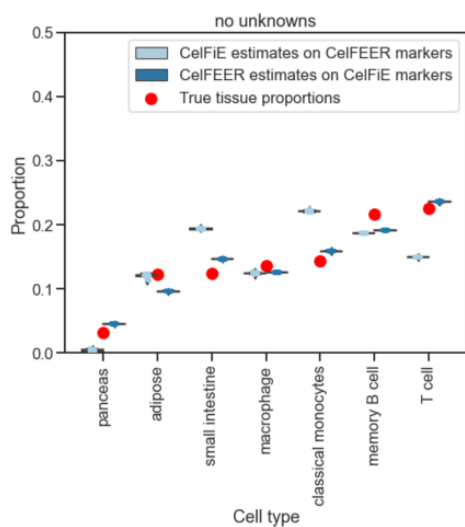
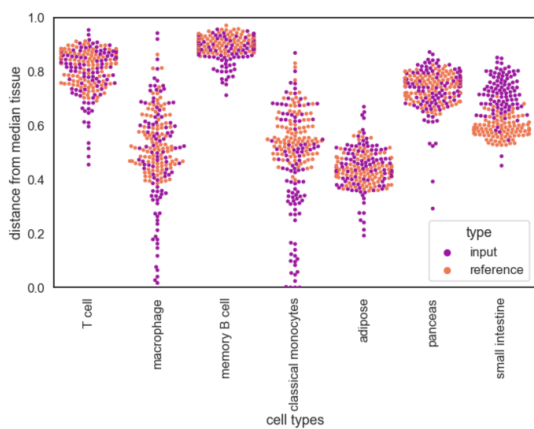
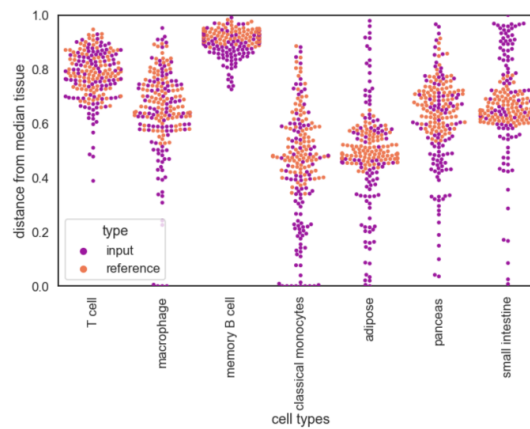


Fig. S4: CelFiE and CelFEER run on different markers.



(a) CelFiE markers



(b) CelFEER markers

Fig. S5: Markers found by (a) CelFiE and (b) CelFEER for seven different cell types.

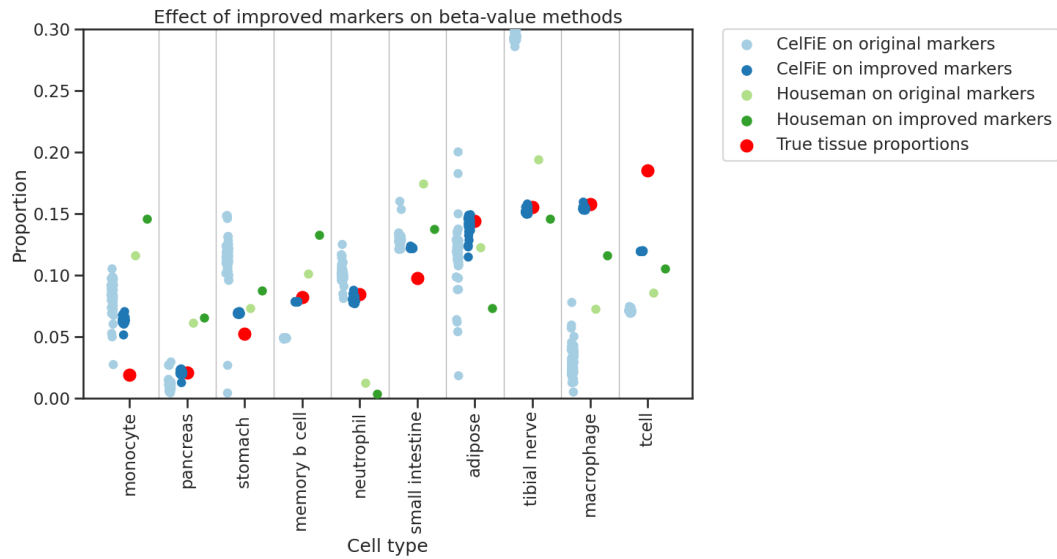


Fig. S6: CelFiE's and Houseman's estimated cell type proportions on a simulated cfDNA mixture using WGBS cell type data, using both the markers found as described in [1] and the markers found using our improved method. The results of 50 replicate runs on a randomly chosen individual are displayed.

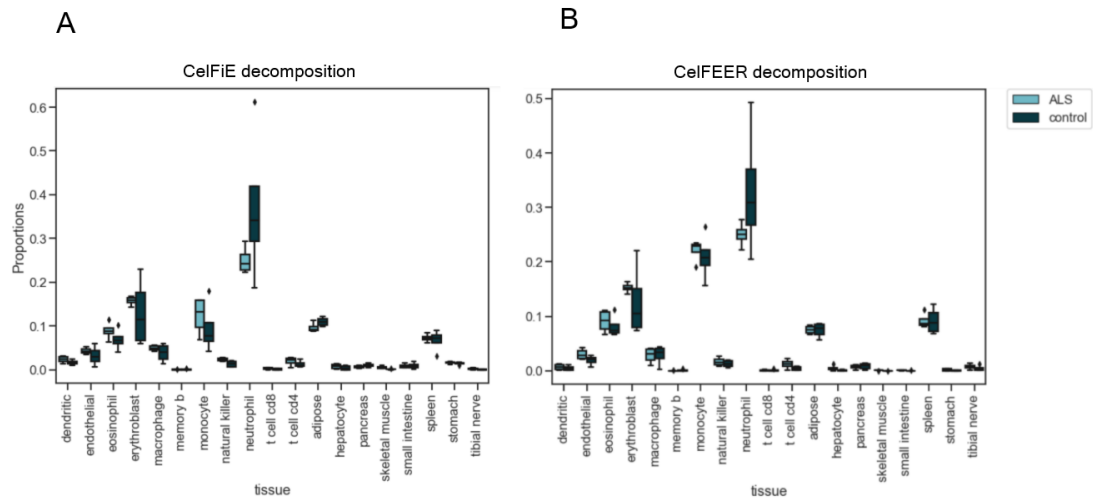


Fig. S7: Estimated proportions of cfDNA in ALS patients ($n = 4$) and a control group ($n = 4$) of (A) CelFiE and (B) CelFEER.

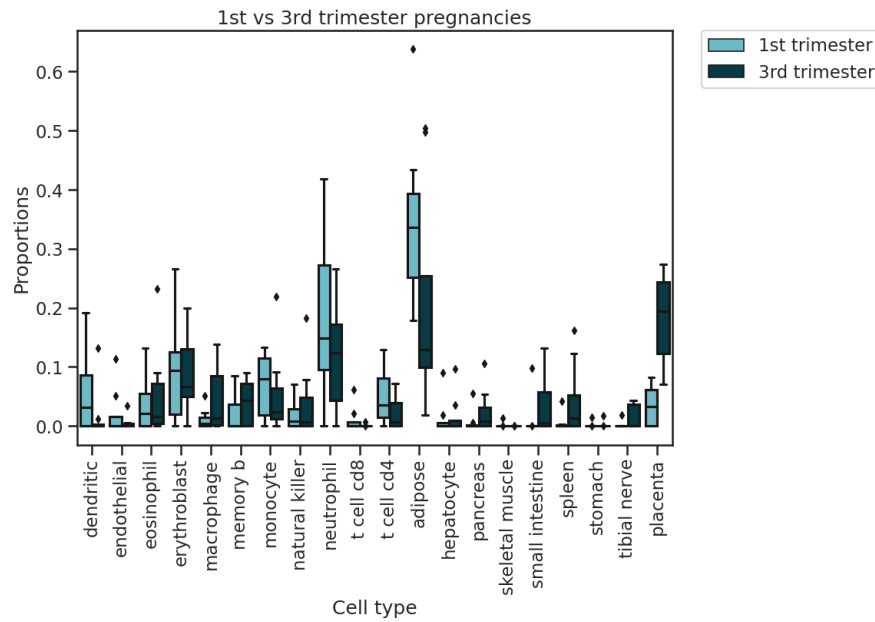


Fig. S8: Full cell type decomposition of first trimester ($n = 8$) and third trimester ($n = 8$) pregnancies cfDNA.

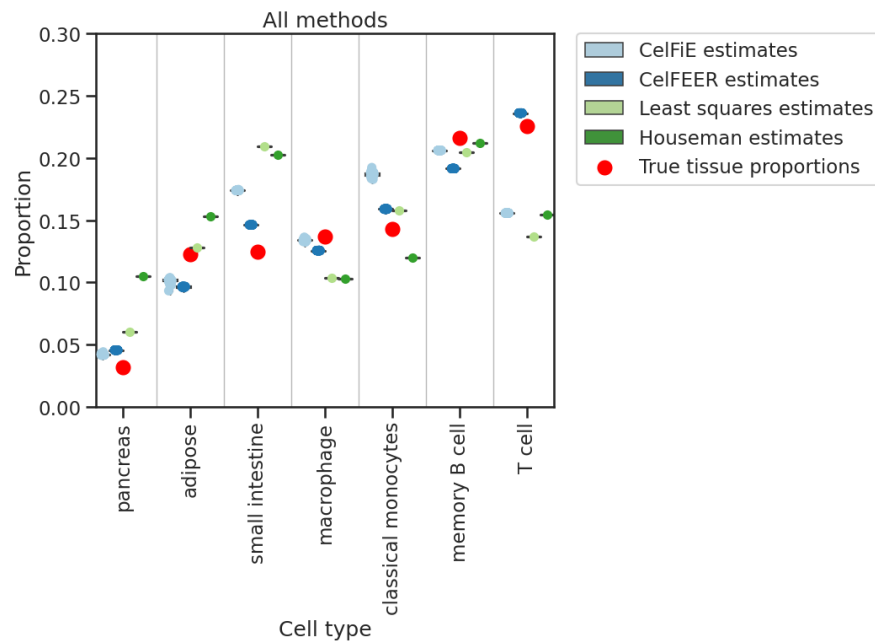


Fig. S9: Cell type proportions estimated by CelFiE, CelFEER, least squares regression and Houseman. The boxplots visualize the estimated proportions for a randomly chosen individual, using 50 replicates for CelFiE and CelFEER.

4 Selection of cell type informative markers

A crucial step in predicting the cell type of origin is selecting markers in the genome that represent the cell types. Not only does a set of distinct markers improve prediction, it can make sequencing of cfDNA less expensive since only the DNA overlapping the markers needs to be sequenced. Methylation markers that span multiple CpG sites are in literature often referred to as differentially methylated regions. To find cell type informative markers, we started by analyzing the markers found using the

method created by Caggiano et al. [1], which is described in the materials and methods. This method was then improved to find more informative markers. In this section and the following we refer to the absolute counts of methylated CpG sites as methylation values, and to the fraction of methylated to unmethylated CpG sites as methylation percentages.

Regions are more robust markers than single sites

Caggiano et al. [1] use the traditional approach of using single CpG sites as markers. This method, however, decreases the ability to differentiate between different cell types as it is sensitive to both biological and technical noise. In order to reduce noise, the CpG sites 250 bp upstream and 250 bp downstream of the markers are added to the markers' methylation counts. The authors showed that their method only returns sensible results when the methylation values are thus summed into regions. It nonetheless happens that the 500 bp surrounding the markers contain little CpG sites. This method does not exploit earlier findings that the methylation status is highly coupled between adjacent CpG sites [2]. Moreover, regions where CpG sites are clustered in high numbers, called CpG islands (CGIs), are known to be epigenetic regulatory regions that can be cell type specific [3].

According to these findings, it makes more sense to compare regions containing multiple CpG sites instead of single CpG sites to find differential markers. To test this hypothesis, CpG sites were grouped in a simple fashion: CpG sites were summed if they were in a 500 bp vicinity of each other. The starting location of each 500 bp window was set to be the first CpG site which contained measurements and did not fall in a previous bin. This strategy has the downside that it may split clusters in two, but if this is the case and if this cluster is differential, it is not harmful for the method to use both parts of the cluster as markers.

In addition to summing over 500 bp windows, we also summed over 10 bp windows with the idea of removing noise while still looking at mostly local methylation. After finding markers on the 10 bp windows, the surrounding CpG sites were summed to nevertheless obtain a total window of 500 bp. In order to compare the markers' ability to differentiate between cell types, we looked at the absolute difference between the methylation percentage of each marker's cell type and the median methylation percentage of all cell types. Accordingly, the distance for tissue k is defined as:

$$\text{distance}(t_k) = \frac{\sum_m \beta_{m,t_k}}{M} - \text{median} \left\{ \frac{\sum_m \beta_{m,t_1}}{M}, \dots, \frac{\sum_m \beta_{m,t_T}}{M} \right\} \quad (\text{S8})$$

,where t indicates the tissue, T the total number of tissues, m the marker, and M the total number of markers.

To test the generalizability of the markers, we did this for both the reference data (which was used to find the markers) and for the input data. As can be seen in Figure S10, the markers are most differential when they are first summed in 500 bp windows, and the variance in distance has substantially decreased. This strategy also seems to result in markers that generalize relatively well to unseen data, as the input and reference data have a similar distance to the median of other cell types. Although summing in 500 bp windows seems to return better markers than summing in 10 bp windows, it is remarkable how much improvement can be seen compared to the original method, especially for the tibial nerve cells. This is probably the effect of the decrease in noise which appears even if we sum over such small intervals. The results confirm the belief that markers are more differentiable when CpG sites are first summed compared to when they are summed after selecting individual sites. For this reason, all future experiments on markers are done on sites summed in 500 bp regions. In this section, we used only hypomethylated markers as they promised to be most distinguishing between cell types.

We additionally investigated the effect of using different window sizes. Firstly, we looked at the distance between tissues in the reference dataset (which was used to find the markers). The distance is measured by calculating the difference with the median methylation across all cell types of the average methylation in the cell type (Equation S8). For window sizes of 10 bp, 100 bp, 500 bp, 700 bp and 1000 bp, we found an average distance of 0.40, 0.61, 0.58, 0.55 and 0.52, respectively. To find how well the markers generalize to the input data, we measured the distance in the reference data as well. In the same order, the different window sizes resulted in distances of 0.65, 0.56, 0.78, 0.63 and 0.41, showing that a 500 bp window allows for best generalization.

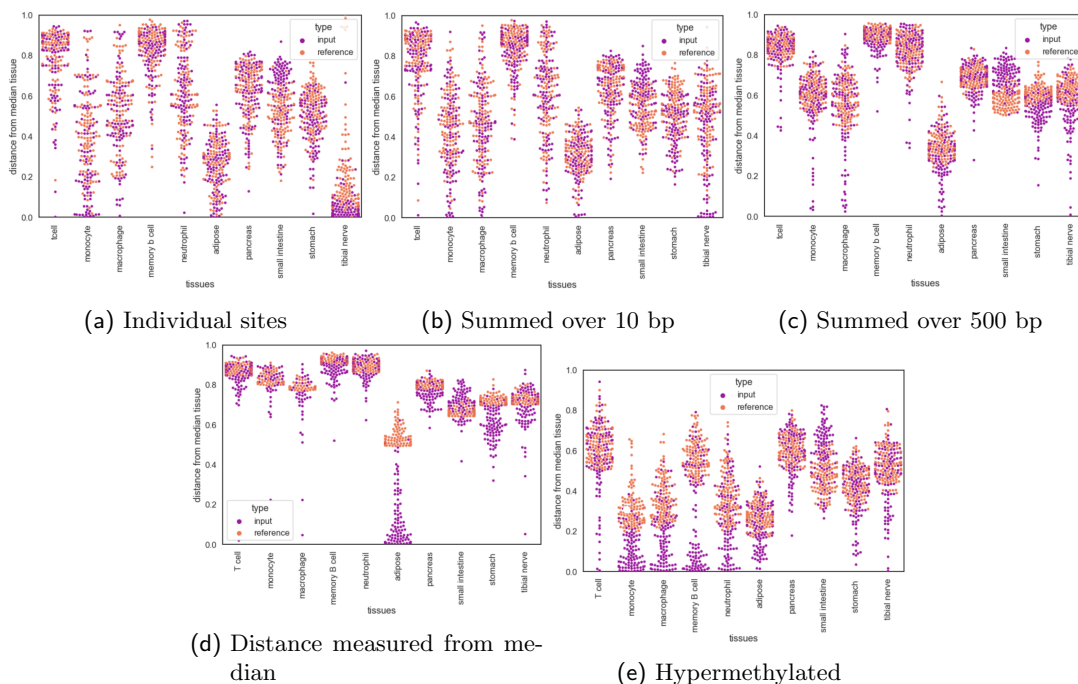


Fig. S10: Distance from median methylation percentage for three different strategies; Purple dots represent the input at different marker locations and orange dots represent the reference at the same marker locations. The reference data was used to find the marker locations.

Row 1: Comparison between single CpG site markers which are summed with their 500 bp neighbouring sites (S10a), 10 bp markers which are summed with their 490 bp neighbouring sites (S10b) and 500 bp markers (S10c).

Row 2: Comparison between markers defined by their distance from the median methylation percentage (S10d), distance from the maximum (S10e) and distance from the minimum (S10c).

All figures in row one use hypomethylated markers, and all figures in row two are first summed over 500 bp.

Hypomethylated sites are easier to differentiate than hypermethylated sites or than a mixture of both

Caggiano et al. [1] originally determined the best markers for each cell type by comparing the distances between the methylation percentages of each individual cell type to the median methylation percentage of all cell types. This should, in theory, result in a mixture of hypo- and hypermethylated markers. A sufficiently large distance to the median is, however, not a very strict requirement as it does not remove the probability of having two or more cell types with a very similar methylation percentage (especially as the number of cell types in the reference grows). Moreover, in practice almost all of the markers found using this method are hypomethylated, so there is little benefit in also allowing for hypermethylated markers.

To make the markers more differential, we measured the distance between the methylation percentage of each cell type and the minimum methylation percentage of all other cell types. This approach was compared to the original approach (where the distance from the median is measured instead) as well as to a similar approach where we looked only for hypermethylated markers (and thus compared to the maximum of all other cell types). When comparing the markers' distances from the median, the original method seems to result in the best markers for all cell types except adipose (Figure S10d). Hypomethylated markers, on the other hand, have a slightly smaller distance from the median for all cell types except for adipose, for which the distance is larger (Figure S10e). Hypermethylated markers have overall the smallest distance from the median (Figure S10c).

However, as reasoned above, the distance from the median may not be the best metric for defining the ability to differentiate between cell types. Therefore, we can not assume that the distance from

the median also translates to the best cell type deconvolution results. For this reason, we looked at the results on a simulated mixture of the WGBS data of 10 cell types and measured the Pearson's correlation between the true and estimated cell type proportions of 50 replicate runs for 10 individuals. We set the true cell type proportions to a linearly incrementing array that sums to one. While the hypomethylated markers resulted in a correlation of $r^2 = 0.86 \pm 0.01$, the hypermethylated and original method resulted in a correlation of $r^2 = 0.68 \pm 0.04$ and $r^2 = 0.58 \pm 0.03$ respectively. This confirms the idea that the distance from the median is not the best metric for obtaining differentiable markers.

This can additionally be observed from the amount of markers found by each metric. The method for finding markers works in such way that it first finds the 100 best markers for each cell type and then removes the markers that are overlapping multiple cell types. As can be seen in Figure S11, the original method finds less markers which means that the markers it finds have a high amount of overlap between cell types. Especially monocytes and macrophage cells seem to have much overlap, which makes sense given the fact that macrophage cells are differentiated monocyte cells [4]. Hypo- and hypermethylated markers are nevertheless able to differentiate these two cell types. To test whether the markers found using the original method would result in better performance if more markers were included, we first tested for uniqueness of the 200 best markers of each cell type and then included the 100 best markers. This way each cell type had 100 markers. This resulted in a negligible increase in performance.

As the hypomethylated markers seem to give the best results, all experiments in this section, including the previous section, use hypomethylated markers.

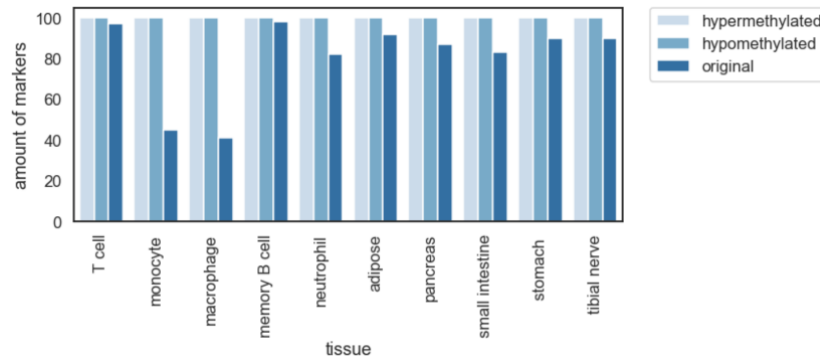


Fig. S11: The bar chart shows the amount of markers found for each cell type using each of the three different ways to measure the distance between cell types.

Restricting the genome for CpG islands or Methylation Haplotype Blocks results in less differentiable markers

We create regions by binning the genome into 500 bp windows. Restricting the search for markers to a subset of regions in the genome which are known to be cell type specific might improve the cell type specificity of the markers. Methylation patterns at CpG islands are known to correlate with cell type [5]. Alternatively, methylation haplotype blocks (MHBs), invented by Guo et al., are regions that aid in tissue deconvolution [2]. We investigated whether replacing our approach of binning the genome in regions of 500 bp by either of those two approaches could enhance the performance of CelFEER.

We rely on beta values to compare different markers to prevent optimizing the markers for our own method. We used the same seven cell types used for comparing CelFiE and CelFEER in the main text. We measured how good the new sets of markers are in differentiating between different cell types by measuring the distance between the cell types (Equation S8). The resulting distances can be seen in Table S3.

We additionally ran CelFiE and CelFEER on markers that were restricted to MHBs. As implied by the distance, the average correlation between the true and estimated cell types suffers from restricting the genome. The results can be seen in Table S4.

Additional improvements for increased differentiation between cell types

In addition to the improvements discussed in the previous two sections, there were two possible unwanted outcomes in the original method for finding markers. The first of which is that the authors introduced only a requirement for the median read depth of all cell types at a candidate marker site. This means that if one cell type is covered by one single read only at a candidate CpG site, this CpG site can still become a marker for that cell type as long as all other cell types have sufficient coverage. A simple adjustment was made to the method by setting a minimum depth threshold for cell types at their potential marker sites. This threshold was set equal to the median depth threshold.

The second possible undesirable behaviour is caused by the manner of checking for the uniqueness of the markers. As only the top 100 markers of all cell types is checked for overlapping markers, it is possible that the same site is the 100th best marker for cell type x and the 101st best marker for cell type y . This situation was prevented by keeping a list of the 150 best markers for each cell type which are all checked for uniqueness, such that the 100th best marker for cell type x could not even be the 150th best marker for cell type y .

The effects of both changes were measured by calculating the Pearson's correlation between the true and estimated cell type proportions for 10 individuals and 10 cell types of 50 replicate runs. The true cell type proportions were drawn from a uniform distribution and made to sum to one. Using no improvements, the correlation between the true and estimated cell types was $r^2 = 0.87 \pm 0.09$. Using only the additional uniqueness criterion did not change the results, and resulted in the same amount of correlation. The stricter depth criterion, however, improved the correlation to $r^2 = 0.91 \pm 0.06$. Combining both improvements resulted in the same correlation. This means that the situation described above does not occur, and the markers are already sufficiently unique. This is perhaps a consequence of using hypomethylated markers only.

Other methods also benefit from our proposed marker selection approach

Since CelFiE benefits from using our enhanced markers as input, we used this set of markers as input for the array-based method by Houseman et al. [11] (Figure S6). We find that using the proposed marker selection procedure significantly improves the results of this method compared to Caggiano et al.'s procedure [1] (Table S6).

References

- [1] Caggiano, C., Celona, B., Garton, F., Mefford, J., Black, B. L., Henderson, R., Lomen-Hoerth, C., Dahl, A., and Zaitlen, N. (2021) Comprehensive cell type decomposition of circulating cell-free DNA with CelFiE. *Nature communications*, **12**(1), 1–13.
- [2] Guo, S., Diep, D., Plongthongkum, N., Fung, H.-L., Zhang, K., and Zhang, K. (2017) Identification of methylation haplotype blocks aids in deconvolution of heterogeneous tissue samples and tumor tissue-of-origin mapping from plasma DNA. *Nature genetics*, **49**(4), 635–642.
- [3] Tahir, R. A., Zheng, D., Nazir, A., and Qing, H. (2019) A review of computational algorithms for CpG islands detection. *Journal of biosciences*, **44**(6), 1–11.
- [4] Yang, J., Zhang, L., Yu, C., Yang, X.-F., and Wang, H. (2014) Monocyte and macrophage differentiation: circulation inflammatory monocyte as biomarker for inflammatory diseases. *Biomarker research*, **2**(1), 1–9.
- [5] Ghosh, S., Yates, A. J., Frühwald, M. C., Miecznikowski, J. C., Plass, C., and Smiraglia, D. (2010) Tissue specific DNA methylation of CpG islands in normal human adult somatic tissues distinguishes neural from non-neural tissues. *Epigenetics*, **5**(6), 527–538.