Estimates of the global burden of Congenital Rubella Syndrome, 1996-2019

Supplementary Material

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A. Vaccination coverage data

The vaccination coverage data that were used in the model are described in the main text. Missing SIA or routine coverage data were further supplemented from publications[1-22]. Data on coverage among adolescent girls came from publications where possible[23-25] (see main text for details).

Table 1: Countries and time periods for which RCV coverage for a given dose was estimated. One of four approaches was used: linear interpolation, constant gap, flat line and set value approaches.

The coverage is calculated so that the absolute difference in the coverage between the first and second doses in the given year is fixed at the level observed in the years in column 7, unless otherwise stated. Consistent with steady coverage of the first dose.

Adolescent

^{\$} Assumed to be at the level for the whole of Germany in 1990, multiplied by 0.797 (the factor by which the population in W Germany differed from the total population in Germany

* Calculated as the average annual difference between the first and second dose coverage for the period specified

& Provided for 7 year olds

B. Analyses of additional seroprevalence datasets

The methods used to estimate the force of infection for the seroprevalence datasets identified since the previous related analyses are described in [26]. In brief, four catalytic models (A, B, C, D) were fitted to the age-stratified seroprevalence data to estimate the average annual "force of infection" among <13 and ≥13 year olds (i.e. the rate at which susceptible <13 and ≥13 year olds are infected), and the sensitivity of the antibody assay. The pre-vaccination force of infection was allowed to either differ for those aged <13 and ≥13 years (models A and B) or be identical for all ages (models C and D). The sensitivity of the rubella serological (antibody) assay was either estimated (models A and C) or assumed to be 100% (models B and D).

The criteria for selecting the force of infection for further use are described in [26]. In brief, the force of infection estimates were selected in decreasing order of biological plausibility of the model, coming from model A, unless they met specific criteria (the force of infection was implausibly high (>600 per 1000 per year), zero in either age group, higher for older individuals than for children or its upper confidence limit was 100%). If this occurred, we used estimates from model B in preference to those from model C, and those from model C in preference to those from model D. If no model fitted the data convincingly, occurring when the best-fitting age-specific proportion seronegative passed through the 95% confidence limits of just one of the observed datapoints, the dataset was dropped from further analyses.

We included an additional criterion for countries for which the sensitivity of the assay was known to be high that model B was selected in preference to model A if all the other criteria were satisfied and the estimated sensitivity of the assay was 100%for model A, and the lower limit of the 95% confidence interval was implausibly low (less than 95%)[27].

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C. Overview of the transmission model

General structure and equations

The transmission model used is based on the one used in previous calculations of the global burden of CRS. We summarise the key features below; further details can be found elsewhere[26, 28].

[Figure S1](#page-6-1) summarises the general structure of the model in the absence of vaccination. The population is stratified into those who have maternal immunity, those who are susceptible, pre-infectious (infected but not yet infectious), infectious and immune. The demography in the model was described using a realistic age structured (RAS) population following Schenzle's approach[29], with individuals stratified by sex and into 99 age strata, corresponding to the ages <1 , $1 - <2$, $2 - <3$, ..., < 99 years.

Figure S1: General structure of the transmission model in the absence of vaccination.

For all countries, we used the country-and age-specific mortality rate, calculated from survival data for the period 2015-2020 from UN population databases[30]. Therefore, the number of people of age *a* at a given time $t N_a(t)$ depends on the mortality rate. The number of births in the model was calculated by multiplying model predictions of the population size in the given year by the crude annual per capita birth rate for the period 2015-2020, obtained from UN population databases[30]. Both the mortality and crude birth rates in the model were assumed to be fixed over time. Note that the absolute magnitude of the numbers of births in the model does not greatly influence the absolute numbers of CRS cases predicted for the global burden, since these were calculated by multiplying model predictions of the age-specific number of CRS cases per live birth by the observed numbers of live births by maternal age, as seen in UN population databases[30] and then summing the resulting numbers over all maternal ages.

Individuals are born into the first age stratum (stratum *a*=0) on the 31st August of each year and are assumed to have maternal immunity for 6 months. Following standard approaches[29] individuals in each age stratum move to the subsequent age stratum on the 31st August of each year, at the same time as vaccination occurs (see below), and leave the model once they reach age 100 years.

The force of infection in the model at a given time t (λ _{*y*}(*t*) and λ _{*o*}(*t*)) depends on age-specific contact between people and the prevalence of infectious people, with the contact parameters calculated from average annual force of infection estimates from seroprevalence data (see below). For convenience, vaccination is implemented on a single day each year in the model, except for children that are vaccinated when aged 9 months (see below).

[Table 2](#page-8-0) and [Table 3](#page-9-0) give definitions of the variables and parameters respectively that are used in the model. Throughout the description, we use the subscript "*y"* to refer to younger individuals (aged <13 years) and the subscript "*o"* to refer to older individuals (aged ≥13

years). Where necessary, the subscript "w" is used to denote females.

The differential equations in age stratum *a* (a=0, 1, 2, ...,99 years) are provided below (see [Table 2](#page-8-0) and [Table 3](#page-9-0) for the definitions of variables and parameters).

$$
\frac{dM_g(t)}{dt} = -m_{0,g}M_g(t) \qquad 0 < t < 182 \text{ mod } 365
$$
\n
$$
\frac{dS_{a,g}(t)}{dt} = -\lambda_a(t)S_{a,g}(t) - m_{a,g}S_{a,g}(t)
$$
\n
$$
\frac{dE_{a,g}(t)}{dt} = \lambda_a(t)S_{a,g}(t) - m_{a,g}E_{a,g}(t) - fE_{a,g}(t)
$$
\n
$$
\frac{dI_{a,g}(t)}{dt} = fE_{a,g}(t) - m_{a,g}I_{a,g}(t) - rI_{a,g}(t)
$$
\n
$$
\frac{dR_{a,g}(t)}{dt} = rI_{a,g}(t) - m_{a,g}R_{a,g}(t)
$$

The equations for the transitions occurring on 31st August each year are as follows:

$$
M_g(t) = B_g(t)
$$

\n
$$
S_{a,g}(t) = S_{a-1,g}(t-\delta t)(1 - v_e v_{a,g}(t) - m_{a-1,g} - \lambda_{a-1}(t-\delta t))
$$
 for $0 < a \le a_f$ years
\n
$$
E_{a,g}(t) = E_{a-1,g}(t-\delta t)(1 - m_{a-1,g} - t) + \lambda_{a-1}(t-\delta t)S_{a-1,g}(t-\delta t)
$$
 for $0 < a \le a_f$ years
\n
$$
I_{a,g}(t) = I_{a-1,g}(t-\delta t)(1 - m_{a-1,g} - t) + t E_{a-1,g}(t-\delta t)
$$
 for $0 < a \le a_f$ years
\n
$$
R_{a,g}(t) = R_{a-1,g}(t-\delta t)(1 - m_{a-1,g} - t) + rI_{a-1,g}(t-\delta t) + v_{a,g}(t)v_e S_{a-1,g}(t-\delta t)
$$
 for $0 < a \le a_f$ years

The equations for the transitions occurring 6 months after the 31st August (or equivalently, 28th February), when individuals in the first year of life lose their maternal immunity are: *S0,g(t) = Mg(t-δt)(1-m0,g)*

$$
M_g(T)=0
$$

Vaccination for children in their first year of life that is scheduled before 12 months of age is implemented at age 9 months, using the following equations:

$$
S_{a,g}(t) = S_{a,g}(t-\delta t)(1 - V_e V_{a,g}(t) - m_{a,g} - \lambda_a(t-\delta t))
$$
 for a=0 years for

$$
R_{a,g}(t) = R_{a,g}(t-\delta t)(1-m_{a,g}-t) + rI_{a,g}(t-\delta t) + V_{a,g}(t)V_e S_{a,g}(t-\delta t)
$$
 for $0 < a \le a_f$ years

The equations were solved using a specially written C-program, using the Euler method with a time step, *δt,* of 0.25 day. The model was run for 170 simulated years before the earliest possible introduction of RCV (1970), starting from the equilibrium numbers of individuals in each compartment, and with a population size of 750,000 (*N(T0)*), with equal numbers of males and females.

Describing transmission between people in the model

The contact parameters in the model were assumed to differ between younger and older individuals according to the following matrix of "Who Acquired Infection From Whom":

$$
\begin{pmatrix} \beta_1 & 0.7\beta_2\\ 0.7\beta_2 & \beta_2 \end{pmatrix}
$$

The contact parameters in the model for each country were calculated for each bootstrap estimate for the force of infection for younger and older individuals, before the introduction of vaccination using methods described previously[26]. For a given assumption about contact between individuals, the force of infection at time *t* for individuals among younger and older individuals (*λy(t)* and *λo(t)* respectively*),* is given by the following equations:

$$
\lambda_{y}(t) = \frac{C_{yy}I_{y}(t) + C_{yo}I_{o}(t)}{N_{y}(t)}
$$

$$
\lambda_{o}(t) = \frac{C_{oy}I_{y}(t) + C_{oo}I_{o}(t)}{N_{o}(t)}
$$

cyy, cyo, coy and *coo* are related to *βyy, βyo, βoy* and *βoo* through the following equations, where *T⁰* is the start of the model runs:

$$
c_{yy} = \beta_{yy} N_y(T_0)
$$

\n
$$
c_{yo} = \beta_{yo} N_y(T_0)
$$

\n
$$
c_{oy} = \beta_{oy} N_o(T_0)
$$

\n
$$
c_{oo} = \beta_{oo} N_o(T_0)
$$

The parameters, *β¹* and *β²* in the WAIFW matrix for given values for the average force of infection before the introduction of vaccination among younger and older individuals for a given country (denoted by $\bar{\lambda}_y$ and $\bar{\lambda}_o$ respectively) were calculated using the following equations:

$$
\beta_1 = \frac{\bar{\lambda}_y - 0.7\beta_2\bar{I}_o}{\bar{I}_y}
$$

$$
\beta_2 = \frac{0.7\bar{\lambda}_o}{0.7\bar{I}_y + \bar{I}_o}
$$

Where \bar{I}_y and \bar{I}_o are the average numbers of infectious people (males and females combined) for younger and older people, respectively. \bar{I}_y and \bar{I}_o are calculated using the approximations $\bar{I}_y \approx \bar{\lambda}_y \bar{S}_y D$ and $\bar{I}_o \approx \bar{\lambda}_o \bar{S}_o D$, where D is the duration of infectiousness and \bar{S}_y and \bar{S}_o are the average numbers of susceptible children and older individuals respectively, calculated as follows:

$$
\bar{S}_y = \sum_{a=1}^{a_y} N_a(T_0) e^{-\overline{\lambda}_y(a-0.5)}
$$

$$
\bar{S}_o = \sum_{a=a_y+1}^{a_f} N_a(T_0) e^{-12.5\overline{\lambda}_y} e^{-\overline{\lambda}_o(a-a_y)}
$$

Here, the number of people of age *a* was calculated using the following equation, namely by multiplying the population size at the start $(N_a(T_o))$ by the proportion of the population at equilibrium that was of age *a:*

$$
N_a(T_0) = \sum_{a=a_y+1}^{a_f} N(T_0) \frac{N_a(T_E)}{\sum_{a=0}^{a_f} N_a(T_E)}
$$

The equilibrium numbers of people in each age group were obtained by running the model until 2019 in the absence of vaccination.

D. Values of parameters that were varied in the transmission model

Table 4: Summary of the basecase and ranges of the parameters used in the transmission model.

Vaccine efficacy

The vaccine efficacy assumed in our analyses is consistent with estimates of the vaccine effectiveness from a recent systematic review and meta analysis[32], which identified four suitable studies for calculating a pooled estimate of the vaccine effectiveness for the most widely-used strain (RA27/3). The pooled estimate was found to be 97% (95% CI: 92-99%) with the estimate from the study with the greatest weight (41.54 vs 22.55, 22.53 and 13.38) in the meta analysis being 95% (95% CI: 84-98%). The beta distribution used in our analyses results in a median vaccine efficacy of 95% [\(Figure S2\)](#page-14-0), and the range is consistent with the pooled vaccine effectiveness estimate and the range from the study with the greatest weight.

Figure S2: Illustration of the truncated beta distribution with α =33 and ß=2 used to sample values for the vaccine efficacy.

The risk of a child being born with CRS if the mother is infected during the first 16 weeks of pregnancy

A systematic review by Thompson et al[36] identified three studies as being the most reliable for estimating the risk of a child being born with CRS, given rubella infection in the mother during the first 16 weeks of pregnancy. [Table 5](#page-15-0) summarizes the characteristics of these studies. The study of Miller et al[33] is considered to be the most reliable of these studies, considering 1016 pregnant women with laboratory-confirmed rubella infection in England and Wales during January 1976 – September 1978. The other studies considered 491 (Grillner et al[34]) and 32 (Hahne et al[35]) pregnant women respectively. The percentage of infected women who continued with their pregnancies to term ranged from 42% or 378/966 (Miller et al[33]) to 64% (315/491) and 94% (30/32) in Grillner et al[34] and Hahne et[35] al respectively.

Table 5: Characteristics of the three studies used in these analyses for defining the risk of a child being born with CRS if the mother is infected with rubella during the first 16 weeks of pregnancy

All three studies provided estimates of the risk of congenital infection and the risk of a child being born with adverse outcomes relating to rubella infection (referred to here, for simplicity, as probable CRS), given infection in the mother at different stages of pregnancy [\(Table 6\)](#page-16-0). For a given study, the overall risk of probable CRS given rubella infection in the mother at a given stage can be computed as the product of the risk of congenital infection at that stage and the risk of probable CRS, given congenital rubella infection.

Table 6: Data from the three studies[33-35] on the risk of congenital rubella infection, risk of probable CRS given congenital rubella infection and the risk of probable CRS following rubella infection in the mother during the first 16 weeks of pregnancy

* Calculated as the product of the risks in columns 4 and 5 for all periods.

** Grillner et al mention that two mothers may have been infected with rubella before conception. If this occurred, two children would have been at reduced risk of rubella infection from the mother, compared to children infected at other stages of pregnancy. When calculating the risk that the children were infected, these two children are included in the denominator for assumption 1 and are excluded for assumption 2.

***The data for Hahne et al are taken from Table 1 in the paper, which includes infections in both The Netherlands and Canada, which were part of the same outbreak

For each study, as elsewhere[26, 37], we calculated the risk of probable CRS given infection

in the mother during the first 16 weeks of pregnancy as the weighted average of the risk of

probable CRS given infection in the mother during the different stages for the first 16 weeks

of pregnancy. We used bootstrapping to compute 95% confidence intervals and distribution

of the risk of probable CRS given rubella infection in the mother during the first 16 weeks of

pregnancy for each study, as described below.

For each period of infection *g* in the study, we used the risk of probable CRS given infection

in the mother at that time (r_a) to generate 10,000 bootstrap samples for the number of

children who were born with probable CRS, denoted as $c_{i,g}$ for the f^h bootstrap replicate.

For a given bootstrap *i*, we used the following equation to compute the weighted average of the risk (w_i) that a child was born with an outcome consistent with CRS following infection in the mother during the first 16 weeks of pregnancy:

$$
w_i = \sum_{g=1}^G \frac{c_{i,g}}{n_g} d_g \bigg/ \sum_{g=1}^G d_g
$$

Equation 1

Here d_q is the duration of period g in the study and n_q is the number of children with congenital rubella infection who had been infected during period *g* in the pregnancy that were followed up. [Table](#page-17-0) 7 summarizes the values for d_g , n_g and r_g for the three studies.

Table 7: Data from the three studies that were used to generate the distribution of the risk of probable CRS given maternal rubella infection during the first 16 weeks of pregnancy

For the study of Hahne et al[35], the risk, as calculated using [Equation 1,](#page-17-1) reflects the average during the first 18 weeks of pregnancy, excluding weeks 1-3 and 13, since the study had no data on pregnant women infected during those weeks. Also data for weeks 14-18 were not broken down into data for weeks 14-16 and >17 weeks. In sensitivity analyses

(denoted as assumption 2 for Hahne et al), we recompute the CRS risk assuming that the risks during weeks 1-3 and 13 were the same as those in weeks 4-8 and 11-12 respectively.

As shown in [Table 8,](#page-19-0) for Miller et al the estimated risk of a child being born with probable CRS if the mother had been infected during the first 16 weeks of pregnancy was 65% (95% CI: 49-77%). For Grillner et al it was lower at 37% (assumption 1) or 49% (assumption 2), although the 95% confidence intervals were very wide (6-72% and 9-77% respectively) and overlapped with those of Miller et al. Similarly, for Hahne et al, the estimates for assumption 2 were consistent with those for Miller et al, and the confidence interval for assumption 1 overlapped with those for Miller et al. The actual values in Hahne et al may be an underestimate of the risk during 16 weeks, given that it considered 18 weeks, and the risk after 16 weeks has been found to be small in many studies[38].

The resulting distribution of bootstrap replicates for the risk of CRS following rubella infection in the mother during the first 16 weeks of pregnancy for all three studies is summarized in [Figure S3](#page-19-1) The distribution of bootstrap values of this CRS risk from Miller et al was generally consistent with that for the gamma distribution (shape and scale parameters of 37 and 56 respectively) assumed in our analyses, with the values from Hahne et al and over 50% of the values from Grillner et al (assumption 2) also falling within the range of the assumed Gamma distribution.

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Table 8: Estimated risk of a child being born with probable CRS following rubella infection in the first 16 (Miller et al and Grillner et al) or 18 (Hahne et al) weeks of pregnancy

Study	Estimated risk (%) of probable CRS (95% CI)
Miller et al	65.5 (48.7,76.6)
Grillner et al, assumption 1	37.5 (6.3,72.3)
Grillner et al, assumption 2	48.7 (9.4,77.2)
Hahne et al, assumption 1	55.6 (41.9,69.2)
Hahne et al, assumption 2	61.1 (45.8,76.5)

Figure S3: Distribution of the bootstrap replicates of the risk of a child being born with probable CRS following rubella infection in the mother during the first 16 (Miller et al and Grillner et al) or 18 (Hahne et al) weeks of pregnancy. See [Table 6,](#page-16-0) [Table 7](#page-17-0) and [Table 8](#page-19-0) for further details about the assumptions for Grillner et al and Hahne et al. The top left figure also shows the outcome of taking 1000 random samples from the Gamma distribution with shape and scale parameters 37 and 56 respectively.

Table 9: Method for compiling 1000 values for the pre-vaccination force of infection for a given country according to the number of seroprevalence datasets available for that country. Further details of the actual datasets used are in [Table 10](#page-21-0) and [Table 11.](#page-22-0)

Table 10: Datasets used to set up 1000 force of infection bootstrap files for the WHO Regions. These bootstrap files were used to generate 1000 contact parameters for use in the transmission model to calculate the median and 95% range of the CRS burden for countries in a given region which did not have any pre-vaccination seroprevalence data (see main text). Note that these datasets had been accepted after performing the selection procedure described in section B. [Table 11](#page-22-0) provides the datasets for each country.

Table 11: Summary of the bootstrap datasets used to establish the pre-vaccination force of infection for each country using the transmission model, using the WHO regional grouping to assign datasets for countries without serological datasets from before the introduction of RCV. See [Table 10](#page-21-0) for the datasets used to make up the bootstrap datasets.

E: Equations for the CRS incidence and annual numbers of CRS cases

For a given model run, *j*, out of the 1000 model runs, the number of CRS cases per 100,000 livebirths for a given country, *c*, for each year *y* during 1996-2019 was calculated using the following equation

$$
I_{CRS,c,j}^B(A_{15-49}, y) = \frac{N_{CRS,c,j}(A_{15-49}, y)}{\sum_{a=15}^{49} f(a, y)N_w(a, y)} \times 100,000
$$

Here, the denominator is the number of births in year *y* for all women aged 15-49 years, calculated using the fertility rate ($f(a, y)$) and population size ($N_w(a, y)$) of women aged *a* in year *y* in the UN population data and the numerator $(N_{CRS,c,j}(A_{15-49}, y))$ is the estimated humber of CRS cases born to women aged 15-49 years in year y for the fth set of parameter values. The latter number ($N_{CRS, c,j}(A_{15-49}, y)$) was calculated by summing the daily number of CRS cases born to women aged 15-49 years, as follows:

$$
N_{CRS,c,j}(A_{15-49}, y) = \sum_{t=1}^{365} \sum_{a=15}^{49} \frac{r_{CRS,j} s_{w,j}(a,t) f(a,y) N_w(a,y) (1 - e^{-112\lambda_o(t)})}{365}
$$

Here, *sw,j(a,t)* is the modelled proportion of women aged *a* on day *t* that are susceptible, *λo,j(t)* is the daily model-generated force of infection among women on day t, and *rCRS,j* is the risk of a newborn of a mother infected during the first 16 weeks of pregnancy having CRS for the fth bootstrap set of values. As elsewhere[26, 131, 132], we assume that infection during the first 16 weeks of pregnancy carries an average 65% (95% range: (47,88%)) risk of the newborn having CRS.

Regional and global estimates

The regional median and 95% CI of the CRS incidence per 100,000 live births was calculated using country-specific estimates (see above), weighted by the population size.

The equation for the jth model run of the average CRS incidence per 100,000 live births for the *N* countries in a given region was as follows:

$$
\frac{\sum_{c=1}^{N} I_{CRS,c,j}^{B}(A_{15-49}, y)P_{c}(y)}{\sum_{c=1}^{N} P_{c}(y)}
$$

where $I_{CRS,c,j}^B(A_{15-49}, y)$ is defined above and $P_c(y)$ is the population size of country *c* in year *y*. As previously, given China's large population size, the regional incidence for WPR was calculated with and without excluding China, .

We summed the annual numbers for the jth set of bootstrap parameter values for each country in the region to obtain the corresponding regional totals, which were summed to obtain the global burden (i.e. including all countries). These calculations were repeated for each of the 1000 combinations of parameter values.

The 95% CI of the national, regional and global numbers of CRS cases were approximated by the 95% range of the corresponding 1000 values. The country-specific central value was taken as the median from 1000 model runs if the country had either no or >1 seroprevalence dataset. Otherwise, it was taken as that derived from the estimated prevaccination force of infection from the observed data.

F. Estimates of the pre-vaccination force of infection

Table 12: Summary of the additional datasets that were identified since the previous systematic review, best-fitting values for the force of infection and (where appropriate) the sensitivity of the antibody assay for each catalytic model before the introduction of RCV. The values in parentheses reflect the 95% confidence intervals, obtained by bootstrapping. Several of the estimates were published in an interim update of the literature review[28] and are included here for completeness.

* The methods for calculating the force of infection for Laas are identical to those presented in [121], except that, for consistency with the estimates for the other countries, the force of infection is assumed to differ between those aged \leq 13 years and those aged >13 years.

G. Estimates of the CRS incidence

Figure S4: Average estimates of the number of CRS cases per 100,000 live births in 2019 for all countries.

Table 13: The median CRS incidence per 100,000 live births and number of CRS cases born in each WHO region and worldwide in 1996, 2000, 2010 and 2019 and the percentage of the regional live births occurring in countries which had introduced RCV by these years. The numbers in parentheses reflect 95% confidence limits.

Figure S5: Estimates of the number of CRS cases born annually in each WHO region during 1996-2019. The error bars show the 95% range (CI).

Figure S6: Estimates of the average number of CRS cases born annually in each country in 2019

Table 14: Comparison between the estimated number of CRS cases in 2010, 2019 and during 2011-2019 with vaccination as implemented, and the numbers that might have been expected if vaccination had not been introduced after 2010. The final column shows the estimated number of CRS cases averted during 2011-2019 by the introduction of RCV after 2010.

Figure S7: Estimates of the annual number of CRS cases during 1996-2019 globally, in the African, Eastern Mediterranean, South East Asian regions, as calculated using the RCV coverage as implemented and the number that might have been seen if there had been no new introductions of RCV after 2010. The shaded areas show the 95% confidence intervals.

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