1 Supplementary Note

Supplementary Methods

Comparison with left atrial biplane measurements

 The UK Biobank's cardiovascular MRI imaging protocol did not include a volumetric short-axis 57 stack throughout the left atrium¹, so left atrial measurements represent estimates of an unmeasured true left atrial volume. To assess quality, we compared the Poisson surface reconstruction approach with biplane measurements and tested each for association with prevalent atrial fibrillation. Using the R function *cor.test*, we correlated the Poisson surface reconstruction algorithm-based left atrial volume measurements with biplane-based volumes manually measured by experts².

GWAS sensitivity analysis: LVEDV-indexing

 A sensitivity analysis was conducted to assess the consequence of accounting for body size based on each individual's LVEDV (rather than BSA). In addition to functioning as a sensitivity analysis for that purpose, accounting for left ventricular volume could, in principle, help to identify loci whose effects have the opposite effect direction between atrium and ventricle. However, adjusting for heritable covariates in GWAS can also induce associations via collider bias³ . Like the primary analyses, the LVEDV-indexed sensitivity analyses were conducted with BOLT-LMM with the same covariates and settings (**Online Methods**). To attempt to identify LVEDV-indexed associations that were likely attributable to the adjustment for LVEDV, we also conducted a GWAS of LVEDV in the same participants with the same settings, and then tested each of the LVEDV-indexed lead SNPs for independent association with LVEDV.

GWAS sensitivity analysis: no exclusion for abnormal cardiac filling

patterns

A sensitivity analysis was conducted to assess the consequence of retaining participants

identified by the deep learning model as having apparently abnormal cardiac filling. For this

sensitivity analysis, only LAmin and BSA-indexed LAmin were evaluated. Like the primary

analyses, BOLT-LMM was used for this analysis with the same covariates and settings (**Online**

Methods).

GWAS sensitivity analysis: genetic diversity

The primary analyses permitted the inclusion of all participants with LA measurements,

regardless of genetic identity (**Supplementary Figure 9**). As a sensitivity analysis, individuals

were analyzed within genetic inlier groups instead of jointly. To accomplish this, first self-

reported ethnicity—which is only informally correlated with genetic identity—was aggregated

into European (British, Irish, and Other European), African (African,

87 "Any other Black background", "White and Black African", and

88 "White and Black Caribbean"), South Asian (Bangladeshi, Indian, Pakistani), and East Asian.

Individuals with self-reported ancestry of "Any_other_mixed_background", "Mixed",

"White_and_Asian", "Any_other_Asian_background", "Caribbean", "Do_not_know",

91 "Other ethnic group", or "Prefer not to answer" were not analyzed further. Then, for each

group of participants, the R package *aberrant* was run on the centrally computed genetic

principal components of ancestry using a 40 standard deviation window similar to the approach

94 of Bycroft, *et al^{4,5}*. Inliers for each genetic identity group were retained. Individuals that were not

part of an inlier genetic identity group were excluded. The genetic identity inlier groups were

termed EUR, AFR, SAS, and EAS.

 The sample sizes for the AFR, SAS, and EAS subsets were all well below the threshold recommended for the use of BOLT-LMM ("*We recommend BOLT-LMM for analyses of human genetic data sets containing more than 5,000 samples*", BOLT-LMM v2.4.1 User Manual https://alkesgroup.broadinstitute.org/BOLT-LMM/BOLT-LMM_manual.html). Therefore, for each of the four genetic inlier groups, a GWAS was conducted with REGENIE v2.2.4 which does not 103 have the same limitation⁶. All models were adjusted for sex, age and age2 at the time of MRI, the first 10 principal components of ancestry, the genotyping array, and the MRI scanner's unique identifier. Fixed-effect meta-analysis was then conducted with METAL (release version $2020 - 05 - 05$) ⁷.

 Two additional GWAS were conducted in BOLT-LMM v2.3.4 using the same covariates as the primary GWAS: one for the inlier EUR population, and another where an equivalent number of individuals were dropped at random from the original GWAS cohort (without regard for genetic inlier grouping) to yield a sample size that was the same as the inlier EUR population.

 GWAS loci from the primary analysis were fetched from the meta-analysis, the EUR-specific GWAS, and the GWAS in which individuals were dropped at random.

Polygenic score sensitivity analyses

 In addition to the primary LA polygenic scores produced with PRScs, an additional set of LA 117 polygenic scores was created as a weighted allelic sum based on the lead variants for each trait. That is, for each tested participant, at each of the lead variant alleles, the number of effect alleles possessed by the participant was multiplied by the effect estimate; these were then summed for all alleles for each phenotype. They were tested for association with diseases in the same way as the PRScs scores.

Supplementary Results

Semantic segmentation model quality assessment

 In a held-out test set of 20 manually annotated images from the two-chamber short axis view that were not used in training or validation, the average Dice coefficient was 0.89 (SD 0.06) for the left atrial blood pool. For 20 held-out images from the three-chamber view, the Dice score was 0.88 (SD 0.07). For 40 held-out images from the four-chamber view, the Dice score was 0.94 (SD 0.03).

 The short axis imaging sequence was not designed to capture the atria: the atrial short axis 131 sequence was eliminated from the acquisition protocol to save acquisition time¹. The left atrium was nevertheless recognizable in the basal-most segments of images obtained in the short axis view. In the short axis view, the average Dice score for the left atrium was 0.78 (SD 0.35) when weighted by the total number of pixels assigned to the left atrium by the cardiologist or the model, or 0.90 (SD 0.28) when considering images correctly identified by the model as having no left atrial pixels to have a Dice score of 1.

 In the two-chamber view, the average Hausdorff distance was 6.7mm (SD 4.0mm). In the three- chamber view, the average Hausdorff distance was 8.8mm (SD 8.5mm). In the four-chamber view, the average Hausdorff distance was 5.2mm (SD 4.1mm). In the short-axis view, the average Hausdorff distance was 5.8mm (SD 4.2mm).

 In the two chamber view, the average mean contour distance was 1.8mm (SD 0.6mm). In the three-chamber view, the average mean contour distance was 2.3mm (SD 2.2mm). In the four-chamber view, the average mean contour distance was 1.3mm (SD 0.90mm). In the short-axis view, the average mean contour distance was 1.7mm (SD 1.7mm). The mean contour distance 147 for the automated left atrial segmentation in each of these views was less than the in-plane pixel spacing of 1.83mm.

Segmentation and reconstruction quality control

 QC-flagged samples (due to more than 1 connected component, frame-to-frame pixel changes greater than 5 standard deviations above the mean, the absence of left atrial pixels, or an abnormal number of CINE images as detailed in the **Online Methods**) were significantly more likely to fail to achieve a successful Poisson reconstruction (OR 1.4, P=1.3E-19). Among the left atria that were successfully reconstructed, we tested whether the presence or absence of any of the QC flags was associated with volumetric measurements. However, the distribution was similar regardless of QC status (**Supplementary Figure 10**); the presence of QC flags was statistically non-significant for LAmin (0.020 SD greater with a flag, P=0.06) and had a similarly small effect estimate for LAmax (0.036 SD greater with a flag, P=5E-04). Therefore, all samples that were successfully reconstructed were retained for analysis.

Comparison with left atrial biplane measurements

We correlated the Poisson surface reconstruction algorithm-based left atrial volume

measurements with biplane-based volumes manually measured by experts in 3,401

163 participants². When limiting the inputs into the Poisson surface reconstruction algorithm to only

the two- and four-chamber long axis views ("Poisson biplane"), which are the two views used to

- 165 calculate the biplane volume, the correlation improved for both LAmax (from r=0.814, 95% CI
- 0.802 to 0.825, P=2.9E-804 with the full reconstruction to r=0.887, 95% CI 0.880 to 0.894,
- P=4.5E-1143 with the Poisson biplane) and LAmin (from r=0.768, 95% CI 0.754 to 0.781,
- P=1.1E-659 to r=0.860, 95% CI 0.851 to 0.868, P=6.9E-994). We interpreted these results as

 supporting the notion that, when presented with the same input information, the modeling approach yields estimates that are similar to the standard biplane estimation.

 We then used logistic regression to recapitulate prior observations that individuals with pre-173 existing atrial fibrillation have larger atrial volumes^{8,9}. In a subset of 39,148 participants, of whom 808 had atrial fibrillation, both the full Poisson reconstruction and the Poisson biplane reconstruction could be performed. Although the Poisson biplane better correlated with the manual measurements in the previous analysis, the full Poisson reconstruction was more 177 strongly associated with prevalent atrial fibrillation (LAmax OR 1.72, P=1.3E-78 and LAmin OR 1.86, P=1.0E-132) compared to the Poisson biplane model (LAmax OR 1.65, P=6.3E-66 and LAmin OR 1.80, P=2.8E-130).

 We interpreted these findings as indicating that (1) the Poisson-based measurements were well correlated with manual measurements, and (2) while full volumetric imaging stacks through the atria were not available to adjudicate correctness, the Poisson-based measurements that incorporated all available views (2ch, 3ch, 4ch, and SAX) were more strongly correlated with atrial fibrillation than the Poisson biplane measurements.

Quality control for the deep learning model for abnormal cardiac filling

patterns

 Among 200 participants whose MRIs were manually reviewed (100 flagged as having abnormal cardiac filling patterns and 100 flagged as having normal cardiac filling patterns), manual review determined that 164 were normal and 36 were abnormal. The sensitivity of the model for identifying abnormal cardiac filling patterns was 100% (95% CI 90.3-100.0%) and the specificity was 61% (95% CI 53.1-68.5%). These findings suggested that the model may have overdetected abnormal cardiac filling—leading to the exclusion of more participants than

necessary—but had little evidence for false negatives.

 Relationship between cardiac filling patterns and left atrial volume Among the 40,558 participants with LA measurements whose filling patterns could be analyzed, we identified 1,013 participants whose patterns did not appear to be consistent with normal cardiac filling patterns. Of these, 376 (37%) had a pre-existing history of AF or atrial flutter. The same 376 participants represented 32% of all 1,189 participants with a history of AF or atrial flutter. The remaining 637 participants with abnormal cardiac filling patterns did not have a history of AF or flutter, representing only 1.6% of the 39,369 participants without such history. Among participants with no history of AF or atrial flutter, those with an abnormal atrial cardiac filling patterns had significantly elevated LA volumes (**Figure 3**; N = 637; LAmin: +1.3 standard deviations [SD] compared to the 38,732 with no AF history and normal cardiac filling patterns, P 206 = 3.1E-321; LAmax: $+0.8$ SD, P = 3.7E-103). The most extreme volumes were observed in participants with a history of AF or atrial flutter who also had an abnormal cardiac filling pattern (N = 376; LAmin: +4.3 SD compared to those with no AF history and normal cardiac filling patterns, P = 1.6E-1937; LAmax: +2.5 SD, P = 8.9E-623). The 813 participants with a history of AF and normal cardiac filling patterns had larger volumes than those with normal cardiac filling 211 patterns and no AF history (LAmin: $+0.6$ SD, P = 2.4E-101).

 Atrial size was associated with AF, stroke, hypertension, and heart failure After excluding participants with abnormal cardiac filling patterns, we conducted analyses in the remaining 39,545 participants. First, we confirmed previous reports of the relationship between prevalent diseases and atrial size and function. Compared to the 38,732 UK Biobank participants without a diagnosis of AF or atrial flutter prior to MRI, the 813 with a pre-existing

 diagnosis had larger LA volumes (LAmin: +8.8mL, P = 9.2E-117; LAmax: +10.1mL, P = 1.5E-218 61) and a reduced LAEF (-4.6%, $P = 9.7E-68$). Participants with a history of heart failure, hypertension, or stroke also had elevated LA volumes (**Figure 3, left panel**; **Supplementary Table 1**).

 We then examined the relationship between LA measurements and incident cardiovascular diseases. We excluded an additional 1,114 participants with prevalent AF, heart failure, or stroke diagnosed prior to MRI, and 1,525 with missing height, weight, or body mass index (BMI) measurements at the time of MRI. Only a brief period of follow-up time of 2.2 +/- 1.5 years after the MRI assessment center visit was available for most participants. Nevertheless, participants with a larger LA had a greater risk of subsequently being diagnosed with AF (293 incident AF diagnoses; hazard ratio [HR] 1.73 per standard deviation [SD] increase in LAmin; 95% CI 1.60- 1.88; P = 4.0E-39; **Figure 3, right panel**). The LAmin was also associated with an increased risk of incident ischemic stroke (98 cases; HR 1.32 per SD; 95% CI 1.11-1.57; P = 2.0E-03) and heart failure (125 cases; HR 1.69; 95% CI 1.48-1.92; P = 1.3E-15). The associations between other LA measurements and these diseases are detailed in **Supplementary Table 2**. We performed a sensitivity analysis that accounted for ECG features and left ventricular structure and function; this yielded a similar point estimate for LAmin as a marker of incident AF 236 risk (HR 1.89 per SD; 95% CI 1.66-2.15; P = 4.5E-22). In this sensitivity analysis, LAmin 237 remained a significant predictor of incident heart failure (HR 1.51 per SD; 95% CI 1.23-1.86; P = 8.1E-05) but not of incident ischemic stroke (HR 1.10 per SD; 95% CI 0.84-1.43; P = 0.48; **Supplementary Table 3**).

GWAS sensitivity analysis - LVEDV-indexing

 We are not aware of a general solution to the interpretation of GWAS signals that incorporate adjustment for heritable covariates. However, we observed the LVEDV-indexed lead SNPs to fall into three patterns: first, some SNP associations appeared to be driven largely by the LVEDV indexing rather than LA volume. As an example of this pattern, the LVEDV-indexed LAmax association with *BAG3* (P=3.5E-10) was comparable to that for the LVEDV association 246 with *BAG3* alone (P=2.1E-10), while the unadjusted LAmax measurement was not associated $(P > 1E-3)$. At each of these loci, the effect direction in LVEDV was opposite to that in the respective LVEDV-indexed LA volume GWAS, which was expected. Practically, these signals appeared to be driven by the LVEDV values, with the LA measurements acting as noise. Second, some SNP associations appeared to be driven by the LAmax association alone, with only minimal contribution from the LVEDV adjustment. For example, the LVEDV-indexed LAmax association with *IRAK1BP1* (P=2.0E-8) was similar to that for the LAmax association (P=2.7E- 11), while the SNP was not associated with LVEDV (P > 1E-3). Third, some SNP associations appeared to be driven by the interplay between LA volumes and the LVEDV adjustment. For example, the *NEDD4L* locus was associated with LVEDV-indexed LAmax (P=4.7E-8) despite 256 not being strongly associated with either LVEDV or LAmax alone (P > 1E-3 for both).

 For the LVEDV-indexed LA volumes, 11 loci reached genome-wide significance for LAmax, 12 for LAmin, and four for LASV. Of these, six of the LVEDV-indexed LAmax loci had association P < 1E-3 with LVEDV, as did nine of the LAmin loci and two of the LASV loci. Novel loci that were not associated at genome-wide significance in the unadjusted GWAS, and which were not associated with LVEDV at 1E-3 or stronger, included *BLK*, *ANKRD1*, *MYH7*, and *NEDD4L* for LAmax; *CASQ2*, *DHX15*, *PROB1*, *UQCRB*, *ANKRD1*, and *MYH7* for LAmin, and *TNKS* and *HNRNPM* for LASV. Most of these loci were identified in the BSA-indexed GWAS as well.

 GWAS sensitivity analysis: no filtering for abnormal cardiac filling patterns Given the high sensitivity but low specificity of the model detecting abnormal cardiac filling patterns, sensitivity analysis retained the 615 participants who were not identified as having a normal cardiac filling pattern for GWAS of LAmin and BSA-indexed LAmin, yielding a total sample size of N=35,664 participants. (Because some participants are excluded by other criteria downstream of this filter in the primary GWAS, this number is smaller than the 1,013 noted in **Supplementary Figure 3**.) The lead SNPs are recorded in **Supplementary Data 11**. Compared with the main analysis of 35,049 participants, some loci with marginal P-values were lost while others were gained; net, an additional two loci (10 in total) were identified for LAmin and an unchanged number of loci (13) were significant for BSA-indexed LAmin. For example, the association signal for *PITX2* variant rs2466455 for LAmin increased in significance from P=4.6E-06 to P=3.10E-08 in this sensitivity analysis. Similarly the strongest associated variant near *PITX2* for BSA-indexed LAmin in this analysis (rs2723334, P=1.70E-10) had stronger evidence for association than in the primary analysis (P=2.2E-08).

GWAS sensitivity analysis: genetic diversity

 Data from all participants were used for the primary GWAS, incorporating a diversity of genetic identities (**Supplementary Figure 9**). In a sensitivity analysis, only individuals with inlier genetic identities for one of four inlier groups were retained and analyzed separately (EUR, AFR, SAS, or EAS; **Supplementary Figure 11**). In this analysis, the largest inlier group was that for EUR, with 31,878 participants (9.9% smaller than the primary analysis). The second largest group was comprised of the 2,655 participants (7.6%) who were not genetic inliers for any group and were therefore not included in these sensitivity analyses. This was followed by SAS (N=284), AFR (N=133), and EAS (N=99), together comprising about 1.5% of the primary GWAS sample

 size. GWAS were separately conducted for EUR, SAS, AFR, and EAS, and then meta- analyzed. Because of the loss of the participants who were included in the joint analysis but were not inliers for any genetic identity group, the multi-ancestry meta-analytic approach represented a loss of 7.6% of the total sample size compared to the primary analysis. These meta-analytic P-values were fetched for the lead variants from the primary analysis and are displayed in **Supplementary Data 2** as the "*P_META*" column.

 Two additional sensitivity analyses were performed using BOLT-LMM: a EUR-specific GWAS, and an analysis in which individuals were dropped at random to achieve the same sample size as the EUR-specific GWAS. The P-values for the primary analysis's lead variants are also displayed in **Supplementary Data 2** with the "*P_EUR*" and "*P_RANDOMDROP*" columns, respectively.

The weakest association signal occurred for the BSA-indexed LAmin phenotype in the multi-

ancestry meta-analysis at the *GOSR2* locus (P=2.5E-06), which was an order of magnitude

weaker than the evidence for the EUR subgroup without meta-analysis (P=2.0E-07).

Nevertheless, across these sensitivity analyses, we largely observed minor variation in

association signal without clear evidence for population stratification.

311 Supplementary Tables

312 Supplementary Table 1: relationship between left atrial measurements and

313 prevalent disease

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315 Association between left atrial measurements (dependent variables) and condition present at

316 the time of imaging (independent variables). Models were adjusted for age, sex and the

317 magnetic resonance imaging device serial number. Effect estimates and standard errors are

318 displayed in standard deviation units. P values are two-tailed. SE: standard error.

320 Supplementary Table 2: relationship between left atrial measurements and

321 incident disease

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Association between left atrial measurements (independent variables) and incidence of

324 conditions subsequent to imaging (dependent variables) based on Cox proportional hazards
325 models. Models were adjusted for age, sex, the magnetic resonance imaging device serial models. Models were adjusted for age, sex, the magnetic resonance imaging device serial

326 number, height, weight, and body mass index. P values are two-tailed. Effect estimates ("Coef")

327 are exponentiated to hazard ratios ("HR"). SE: standard error.

329 Supplementary Table 3: relationship between left atrial measurements and 330 incident disease after adjustment for left ventricular features

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332 Association between left atrial measurements (independent variables) and incidence of

333 conditions subsequent to imaging (dependent variables) based on Cox proportional hazards
334 models. Models were adjusted for age, sex, the magnetic resonance imaging device serial models. Models were adjusted for age, sex, the magnetic resonance imaging device serial 335 number, height, weight, body mass index, heart rate, electrocardiographic features (P-wave

336 duration, QRS duration, PQ interval, QTc interval) and left ventricular features (end-systolic

337 volume, end-diastolic volume, and ejection fraction). P values are two-tailed. Effect estimates

338 ("Coef") are exponentiated to hazard ratios ("HR"). SE: standard error.

340 Supplementary Table 4: REML heritability and genetic correlation

- REML-based heritability of each trait and genetic correlation between trait pairs are depicted.
- Corr represents the point estimate for each estimate. SE: standard error.

345 Supplementary Table 5: *ldsc* heritability and intercept

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347 Ldsc-based heritability, lambda GC, mean χ^2 , and *ldsc* intercepts are depicted.

350 measurements, atrial fibrillation, and stroke

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LA: left atrium. SE: standard error. P values are two-tailed.

354 Supplementary Table 7: Relationship between atrial fibrillation polygenic 355 score and left atrial measurements

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358 Low and High represent the lower and upper bounds of the 95% confidence interval for the

359 Estimate. SE: standard error. Units are in standard deviations of the left atrial measurements 360 per standard deviation of the atrial fibrillation polygenic score. P values are two-tailed from a 361 linear model.

 Supplementary Table 8: relationship between left atrial polygenic scores and atrial fibrillation risk

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HR: hazard ratio in atrial fibrillation risk per standard deviation in the left atrial polygenic score.

SE: standard error. P values are two-tailed from a Cox model.

Supplementary Figures

Supplementary Figure 1 - Measurement distributions

Trait distributions for the left atrial phenotypes without adjustment and after adjustment for body

surface area (BSA). LAEF is dimensionless and is therefore not adjusted for BSA.

 Supplementary Figure 2 - Normal and abnormal cardiac filling patterns Curves depicting the data used in the abnormal filling pattern detector are displayed for one individual with a normal pattern (**top panel**) and one with an abnormal pattern (**bottom panel**). For visual simplicity, only the left atrial and left ventricular curves from the four-chamber view are displayed. Each datum represents the cross-sectional area at each time point for each chamber. Values are scaled between 0 and 1 (y-axis) on a per-chamber basis so that the maximum is always 1 and the minimum is always 0 for each chamber independently, which is consistent with how the data are transformed prior to being input into the deep learning model. Values are visualized at the 50 timepoints during image acquisition (x-axis). Both panels begin at ventricular end-diastole. The example in the **top panel** reveals a triphasic pattern: ventricular systole continues until timepoint 20, passive ventricular filling until timepoint 45, and then an active ventricular filling phase due to atrial systole from 45-50. The example in the **bottom**

- **panel** reveals a biphasic pattern: there is only ventricular systole until timepoint ~25 and
- ventricular diastole for the remainder of the cycle, with the atrium passively filling and emptying
- in parallel.
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Supplementary Figure 3 - Sample flow diagram

Sample exclusion steps between surface reconstruction and the creation of the GWAS cohort

are described.

 Supplementary Figure 4 - LV-adjusted left atrial phenotype Manhattan plots Manhattan plots for LA measurements divided by LVEDV. X-axis: chromosomal position. Y-axis: -log10(P-value). Nearest gene names are annotated near significant loci, which are colored in red.

Supplementary Figure 5 - Mendelian randomization method comparison

- plot for LAmin vs atrial fibrillation
- SNP effects on the exposure (X-axis) are plotted against SNP effects on the outcome (Y-axis).
- Here, the X-axis effect size comes from the LAmin volume GWAS in this manuscript, while the
- 410 Y-axis effect size comes from the Christophersen, et al, 2017 atrial fibrillation GWAS¹⁰. Points
- 411 represent the mean effect estimates, with 95% confidence intervals for the mean.

Supplementary Figure 6 - Pleiotropic associations for variants used in

Mendelian randomization

 Each of the 19 SNPs from the LAmin Mendelian randomization analysis was tested for association with seven phenotypes previously identified as atrial fibrillation risk factors in CHARGE-AF. For each SNP, this figure displays the mean point estimate of the effect of 1 unit change in the dosage of the non-reference allele on each trait, along with 95% confidence intervals for the mean. Traits where the association with the SNP achieves Bonferroni significance are shown in red. Three of the 19 SNPs were identified to have a significant association with at least one putative confounding factor (rs10878349 near *IRAK3*, rs56129480 near *SP3*, and rs78033733 near *MYL4*).

- 428 Y-axis effect size comes from the Christophersen, et al, 2017 atrial fibrillation GWAS¹⁰. Points
- 429 represent the mean effect estimates, with 95% confidence intervals for the mean.

- Supplementary Figure 8 - Mendelian randomization method comparison
- plot for atrial fibrillation vs LAmin
- SNP effects on the exposure (X-axis) are plotted against SNP effects on the outcome (Y-axis).
- Here, the X-axis effect size comes from the Christophersen, et al, 2017 atrial fibrillation
- 436 GWAS¹⁰, while the Y-axis effect size comes from the LAmin volume GWAS in this manuscript.
- 437 Points represent the mean effect estimates, with 95% confidence intervals for the mean.

Supplementary Figure 9 - Principal components of ancestry

- Principal components of ancestry for the GWAS participants, as well as participants' self-
- described ethnicity mapped with color.

Supplementary Figure 10 - Atrial volume with or without QC flags

Histogram of distribution of LAmin volumes among participants with successful left atrial surface

reconstruction. Values for those with at least one QC-flagged MRI segmentation series are

colored in red, while those for participants with no flagged series are colored in turquoise. The

segmentations are stacked.

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