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Supplemental Online Content

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5 This supplemental material has been provided by the authors to give readers additional
6 information about their work.

7

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44 **Appendix A: Note on sample size considerations**

45

46 Design: Multicenter cluster-randomized crossover two arm trial.

47 Primary endpoint: Difference in proportions of patients with SSIs for cardiothoracic and abdominal

48 surgery following disinfection of the surgical site with povidone-iodine (PI) or chlorhexidine gluconate

49 (CHG).

50 Null hypothesis: PI is inferior to CHG

51 Alternative: PI is non-inferior to CHG.

52

53 **Framework for the trial**

54 Swissnoso is an operational surveillance program, and the trial should fit into the existing operational

55 structures and processed.

56 Expected total number of cardiac and abdominal surgeries in an 18 month period is approximately

57 between 3000 and 3500.

58 There is a fixed number of monthly clusters of treatment assignment for each center (Basel, Bern, Zürich),

59 with varying number of patients in each cluster.

60 We assumed no carryover effect, which is believed to be reasonable for cases in which one product is

61 logistically fully replaced with another. ¹

62 Individual patient randomization was not feasible for this study, due to operational and cost constraints.

63 **Definition of non-inferiority**

64 • We defined non-inferiority in terms of difference in proportions between the proportion of
65 patients with SSIs following disinfection with CHG (reference), p_1 , and the proportion of patients with SSI
66 following disinfection with PI, p_2 .

67 • We defined the non-inferiority margin to be Δ , so that for non-inferiority we need to achieve $(p_1 -$
68 $p_2) > -\Delta$, with the lower Wald type 95% CI limit excluding $-\Delta$.

69 As baseline measurement for the crude SSI infection rate using CHG, as a guideline we used the meta-
 70 analysis in the article by “WHO’s recommendation for surgical skin antiseptics is premature”.² Since the
 71 rates were in line with those from historical Swissnos data, the latter rates were used in the sample size
 72 calculation.

73

74 **Primary analysis model**

75 We assume the difference in proportions is assessed using a chi-square test.

76 We did not dimension the study to explicitly demonstrate the presence of inter-cluster or inter-period
 77 differences, nevertheless, these aspects were investigated with supplementary analyses (*eTable 1*).

78

79 **Sample size calculation**

80 We dimensioned the study based on the crude combined SSI rate for both cardiothoracic and abdominal
 81 surgery, assuming the endpoint is a weighted average of the rates from both types of surgery. The
 82 weights were estimable from the historical data from Swissnos Surgical Site Infection surveillance in
 83 Switzerland.

84 **Assumptions**

85 Significance level = 5%

86 Power = 80, 90%

87 SSI:

88 Reference treatment CHG $p_1 = 0.075$

89 Comparator treatment PI $p_2 = 0.075$ (under the null hypothesis)

90 Non-inferiority margin $\Delta = 0.025$

91 Sample size calculated using function in package “TrialSize” in R, function “TwoSampleProportion.NIS”).

92

93 Recommendation

94 We designed the study to have 80% power at a two-tailed significance level of 0.05 with a non-inferiority
95 margin of -2.5%. We estimated that the study would need to enroll at least 1374 patients in each group
96 (CHG/PI) to achieve the desired power.

97

98 Discussion of sample size calculation and assumptions

99 For the original sample size calculation, we assumed an individual patient randomized design not
100 considering cluster correlations, rather than calculating sample size based on the cluster randomized
101 design. This was the subject of considerable discussion in the trial team, and is a limitation. Nevertheless,
102 alternative cluster-based sample size calculations have limitations of their own, especially when
103 considering crossover designs with many changes (as was the case in our study) - see discussion below.
104 We performed a post-hoc sensitivity analysis on the sample size calculation, taking into account potential
105 cluster correlations.

106 The posthoc sample size calculation was performed using the newly (at the time the analysis was
107 performed, 2020) available “Shiny RCT Calculator” (<https://clusterrcts.shinyapps.io/rshinyapp/>⁶) in which
108 we used the observed SSI rate in A, and estimated the correlation $\rho=0.002$ from the appropriate mixed
109 effects model. For a binary outcome, and assuming a “multiple-period cluster randomized cross-over”
110 design with two periods (AB/BA), with “exchangeable correlation” structure, variable cluster sizes
111 (coefficient of variation =2), a baseline SSI rate on arm A of 5%, with on average 1000 patients per cluster,
112 we would be able to detect an effect size of 1.2%, with 80% power and 5% significance level. This would
113 imply that our study design, with many more periods/clusters is adequately powered in terms of
114 investigating the primary endpoint.

115 We considered a number of alternative approaches to calculate the sample size, none of which could
116 model our proposed design.

117 The above cluster based approach does not explicitly take the non-inferiority margin into account, which
118 was not the case in our original patient level sample size calculation, and makes other assumptions not in
119 line with the study design (e.g. continuous outcome, equal cluster sizes, two period design, no decay in
120 correlation between periods).

121 There is always a balance to be struck between calculating the sample size in a relatively straightforward
122 manner based on a number of verifiable assumptions, and using a more complex approach that follows
123 the design more closely, but relies on more uncertain assumptions. In our study, this meant either
124 assuming individual patient randomization and using existing methods for non-inferiority trials, or using
125 one of the available methods for dimensioning cluster-randomized crossover trials and making
126 assumptions concerning the period and cluster correlation coefficients.

127 The gold standard of study design would have been an individual patient randomized trial.
128 Randomization at the patient level aims to achieve balance of patient baseline characteristics. In the end,
129 and paraphrasing Giraudeau et al. ⁷, we argued that the “negative impact of the clustering effect on the
130 sample size [from cluster randomization]” was balanced out by the “benefit ofa cross-over trial”; this
131 benefit in our study coming from the number of planned periods (planned 9 crossovers, 18 periods). One
132 of the aims of randomization is to achieve patient balance between the arms in a trial. The analysis shows
133 that baseline covariate balance on arms A and B were mostly achieved (refer to *Table 1* of the
134 manuscript), albeit the numbers on each arm were less balanced than for a typical individual patient
135 randomization of similar size (but may occur similarly in an individually randomized trial).

136 However, even individually randomized controlled clinical trials suffer from limitations as reviewed by
137 Frieden TR⁸. In our study, all eligible patients were enrolled eliminating a risk of participation bias and
138 providing external validity. Cluster randomized trials allow for a larger sample size.

139

140

141 **Appendix B: Supplementary analyses on cluster, period and procedure effects**

142 Methods

143 Supplementary analyses of the primary outcome included fitting mixed effects logistic models with SSI as
144 endpoint, including center as fixed effect and period as random effect (e.g. following the approaches
145 “M1” and “M2” outlined in Turner et al. ⁹, *eTable 1*). The estimates from a generalized estimating
146 equation (GEE) population mean model with jack-knifed standard errors were also compared (defined as
147 in Arnup et al. ¹⁰); this latter model was mentioned in the original study protocol.

148 Results

149 Univariable logistic models with SSI as dependent variable and center as fixed effect identified a slightly
150 lower infection rate in Center B (OR 0.6, 95% CI 0.4-0.9) compared to Center A (the reference level).
151 Mixed effects model M2 with the intervention and cluster treated as fixed effects and a random effect on
152 the period had similar estimates (OR 0.6, 95% CI 0.5-0.9).

153 The Breslow-Day test for homogeneity did not indicate any significant differences between SSI rates
154 between CHG and PI after adjusting for the center. Similarly, although the SSI rate for cardiac surgeries
155 was lower than that for abdominal surgeries (OR 0.4, 95% CI 0.3-0.6), the Mantel–Haenszel test did not
156 indicate a difference in the SSI rates between CHG and PI following adjustment for procedure group.

157

158 **Appendix C: Supplementary analysis on shortened follow-up cardiac surgeries during COVID-19.**

159 Due to resource and time constraints, the second follow-up period was shortened for cardiac surgeries

160 from 1 year to between 90 days and 1 year. The estimated potential effect on the number of “missed”

161 infections was judged to be limited (10% of infections > 90 days for Picasso vs 14% for Swissnoso

162 benchmark, $p=0.6$), see *eTable 2*.

163

164 **Appendix D: Missing data**

- 165 - From the statistical analysis plan (SAP): “If the missingness of any one variable (covariate or
166 endpoint) in the full analysis set is more than 10%, a missing data analysis will be performed.”
- 167 - There was no endpoint missingness for all analyses in the main manuscript. Variables with more
168 than 10% missingness, AND included in the supplementary analyses (i.e. *eTable 1*) are:
169 Hemoglobin, creatinine, C-reactive protein, smoking.
- 170 - From the SAP: Patients with missing records are identified in the data set, and compared to
171 those without missing records to determine if there are systematic patterns of missingness in the
172 data. To this end, a logistic regression model will be fitted with the missingness indicator as
173 dependent variable and the relevant covariates as independent variables (i.e. those included in
174 Table 1 of the analysis). Graphical methods will be used to visualize the missingness.

175 *eTable 4* shows variables in Table 1 stratified according to missingness. Missingness patterns are shown in
176 *eFigure 8*.

177 In summary, patients with missing data have/are:

- 178 - A higher infection rate
- 179 - More likely to be from Center A
- 180 - More likely to be female
- 181 - Younger
- 182 - Slightly lower hemoglobin
- 183 - Slightly elevated creatinine
- 184 - More likely to be diabetic
- 185 - More likely to have hypertension
- 186 - Higher BMI
- 187 - Less likely to have a lower American Association of Anesthesiologists score
- 188 - Receive antibiotic prophylaxis later
- 189 - Shorter duration operation
- 190 - More likely an endoscopic procedure

191 - More likely to be clean-contaminated category

192 - Much more likely abdominal surgery

193 The adjusted model provides insights into which other auxiliary variables should be included in the
194 imputation model since they may be related to the missingness process.

195

196 **Multiple imputation**

197 Methods

198 SAP: "Methods for handling missing covariate and endpoint data are defined in section 7.5: Baseline
199 variables will be multiply imputed if the missingness is more than 10% in the full analysis set. Multiple
200 imputation will be performed assuming missing at random (MAR) using the MICE packaged in R, i.e. using
201 multiple imputation using chained equations.

202 Missing SSI (yes/no) information will be multiply imputed using the MICE package in R assuming MAR if
203 the missingness is above 10% for the full analysis set."

204 *Comment: There was no endpoint missingness.*

205 A total of 50 completed data sets will be multiply imputed, and then the appropriate analysis model will
206 be fitted to these complete data sets, with Rubin's rules used to calculate point estimates and their
207 standard errors.

208 Once this process is complete the results from the complete case analysis using the per protocol analysis
209 set will be compared to the results from fitting the appropriate analysis model to the multiply imputed
210 data sets. If the point estimates of the covariates in the analysis model from the analysis with the
211 multiply imputed data sets are outside of the 95% confidence intervals of the estimates from the
212 complete case analysis with the analysis model, it will be concluded that further sensitivity analyses will
213 be required to investigate plausible departures from the missing at random assumption. Such sensitivity
214 analyses will be defined in detail post hoc depending on the variable, but will be based on either delta, or
215 reference based, multiple imputation methods, whichever seems most appropriate."

216 Results

- 217 - We multiply impute missing data for the following variables: Hemoglobin, creatinine, C-reactive
218 protein, smoking
- 219 - The substantive model is any analysis which includes the imputed variables in the main
220 document, which is actually limited to just the analyses in *eFigure 1*.
- 221 ○ We repeat the relative risk calculations on each of the multiply imputed data sets, and
222 compare the point estimates and 95% CIs with those from the complete case analysis. In
223 a slight change to the analyses in Figure S1, we compare the results from fitting a
224 multivariable adjusted logistic regression model to the complete cases, and then to the
225 multiply imputed data.
 - 226 ○ The substantive model has dependent variable SSI (0/1), and independent variables for
227 the arms (CHG/PI) and the respective marker or comorbidity (smoking).
- 228 - The imputation model includes the following variables: SSI (indicator variable, arm indicator,
229 procedure type, Center, hemoglobin, creatinine, C-reactive proteine, smoking BMI, timing of
230 surgical antimicrobial prophylaxis, duration of surgery, ASA-score, endoscopic surgery, wound
231 class.
- 232 - For hemoglobin we use a linear regression imputation approach, for creatinine/C-reactive
233 protein predictive mean matching, and logistic regression for smoker within the MICE package.
- 234 - Following imputation, we recreate the categorical variables (with cutoffs for the markers) and
235 then fit each of the now “full” the imputed data sets to the logistic regression models. The point
236 estimates and variances are then averaged using “Rubin’s rules” in the usual way.

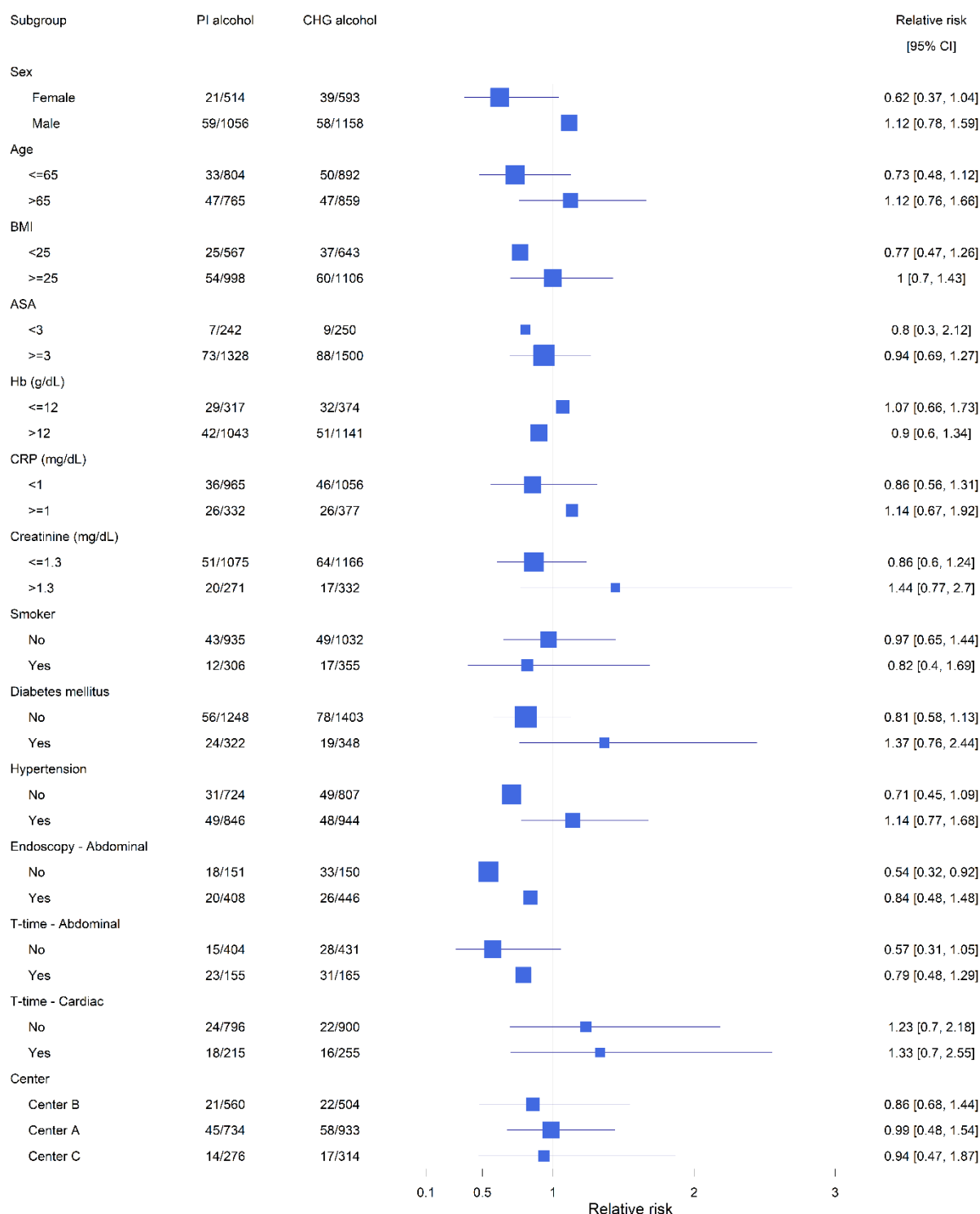
237 *eTable 5* summarizes the results.

238 Conclusions and recommendations: The results from multiply imputing the missing data under the
239 missing at random assumption were broadly consistent with those from the complete case analysis, with
240 the 95% confidence intervals from the imputed data containing the point estimates from the complete

241 case analysis. Whilst this is not conclusive, it implies that the MAR assumption is plausible, and we
242 decided to not continue with sensitivity analyses.

243

244 **eFigure 1:** Forest plot of predefined subgroup analyses (and laparoscopy in colon surgery), CHG as
 245 reference.



246

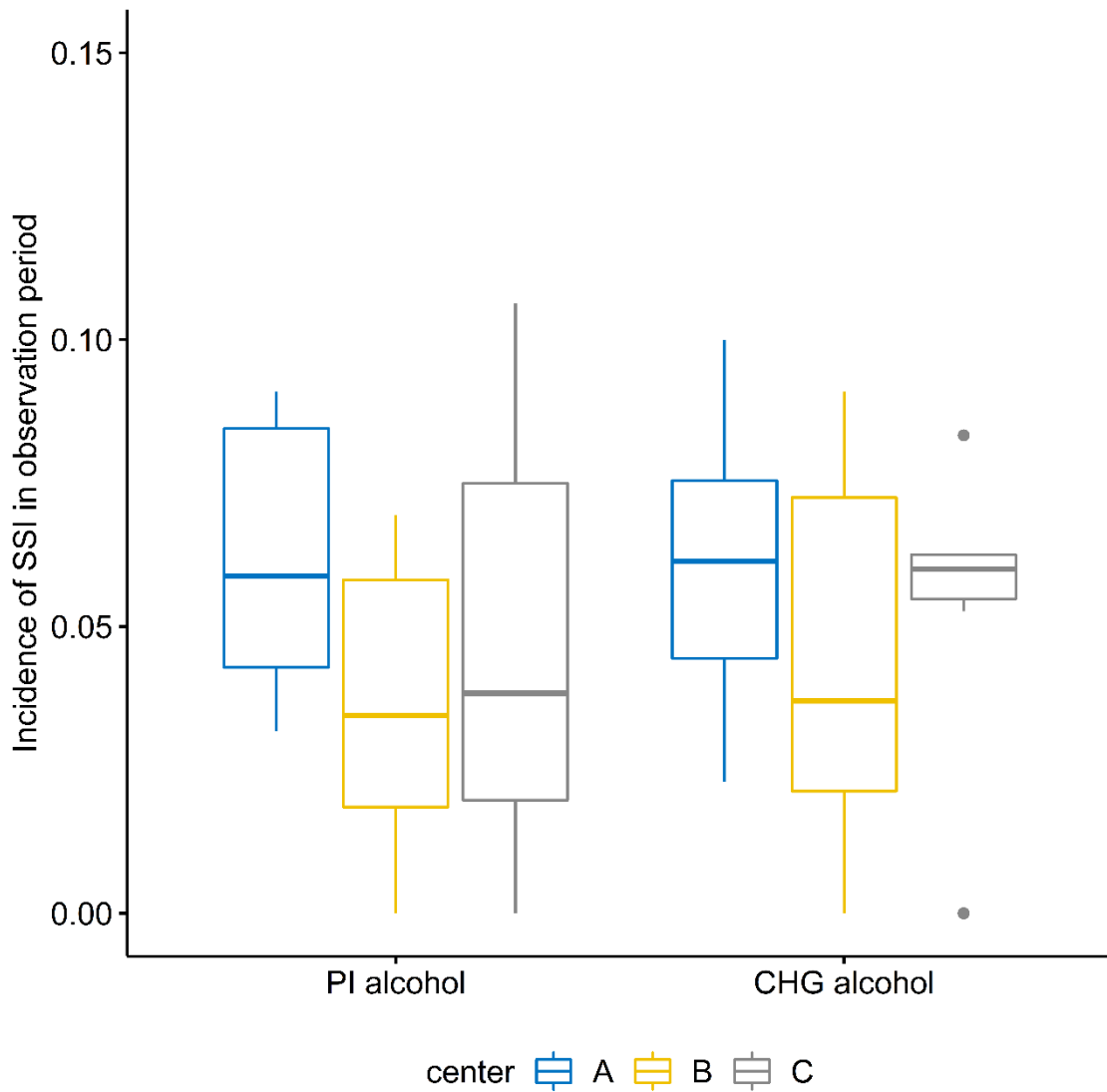
247

248 (CHG: Chlorhexidine gluconate, PI: Povidone iodine, CI: Confidence interval, Age: Age in years, CRP: C-
249 reactive protein, BMI: body mass index in kg/m^2 , ASA: American Society of Anesthesiologists Score, Hb:
250 Hemoglobin, T-time: Surgery duration T-score over 75th percentile)

251

252

253 **eFigure 2:** Infection rates for the periods, stratified by intervention and center.



254

255 (CHG: Chlorhexidine gluconate, PI: Povidone iodine, SSI: Surgical Site Infection)

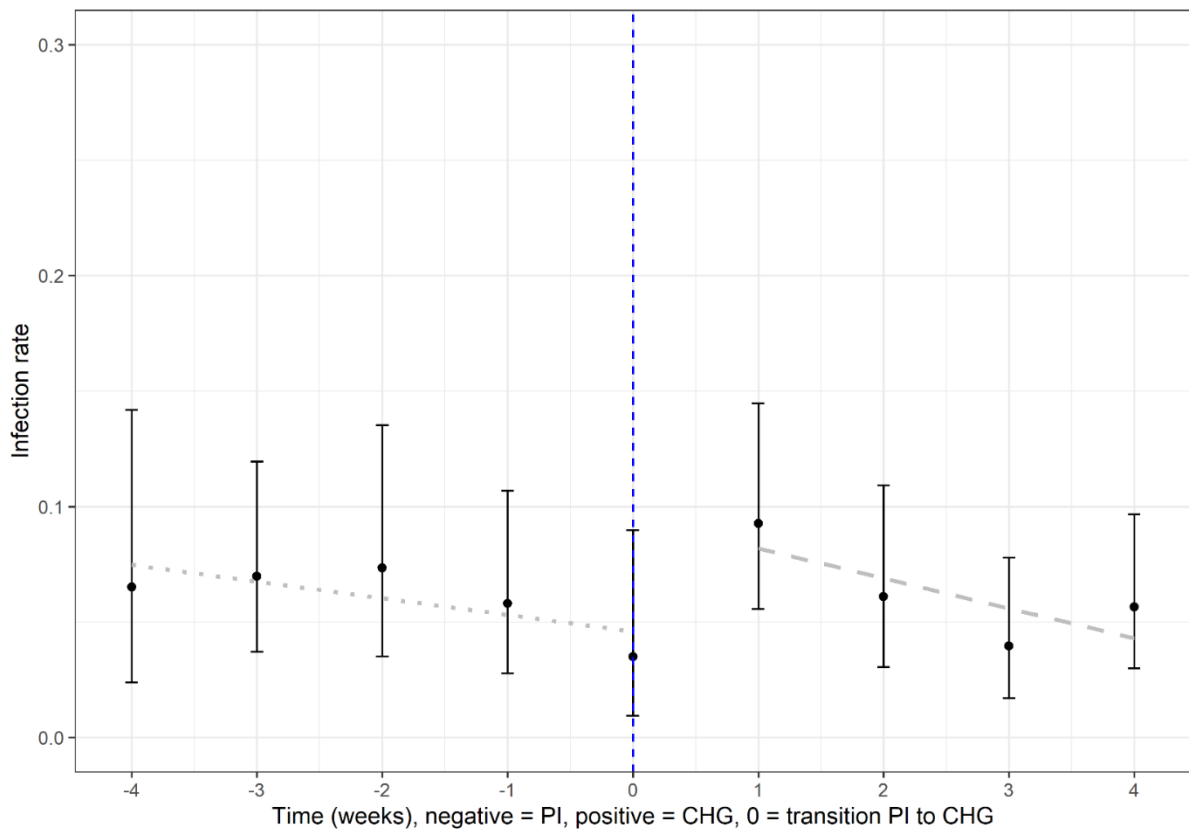
256

257 **eFigure 3:** Interrupted time series analysis of the effect of switching in either direction, stratified by
258 center.

259 *Comments: Time 0 is the switching time from povidone iodine (PI) to chlorhexidine gluconate (CHG), with*
260 *negative times the 4 weeks on PI, and positive times the 4 weeks following the switch to CHG.*

261

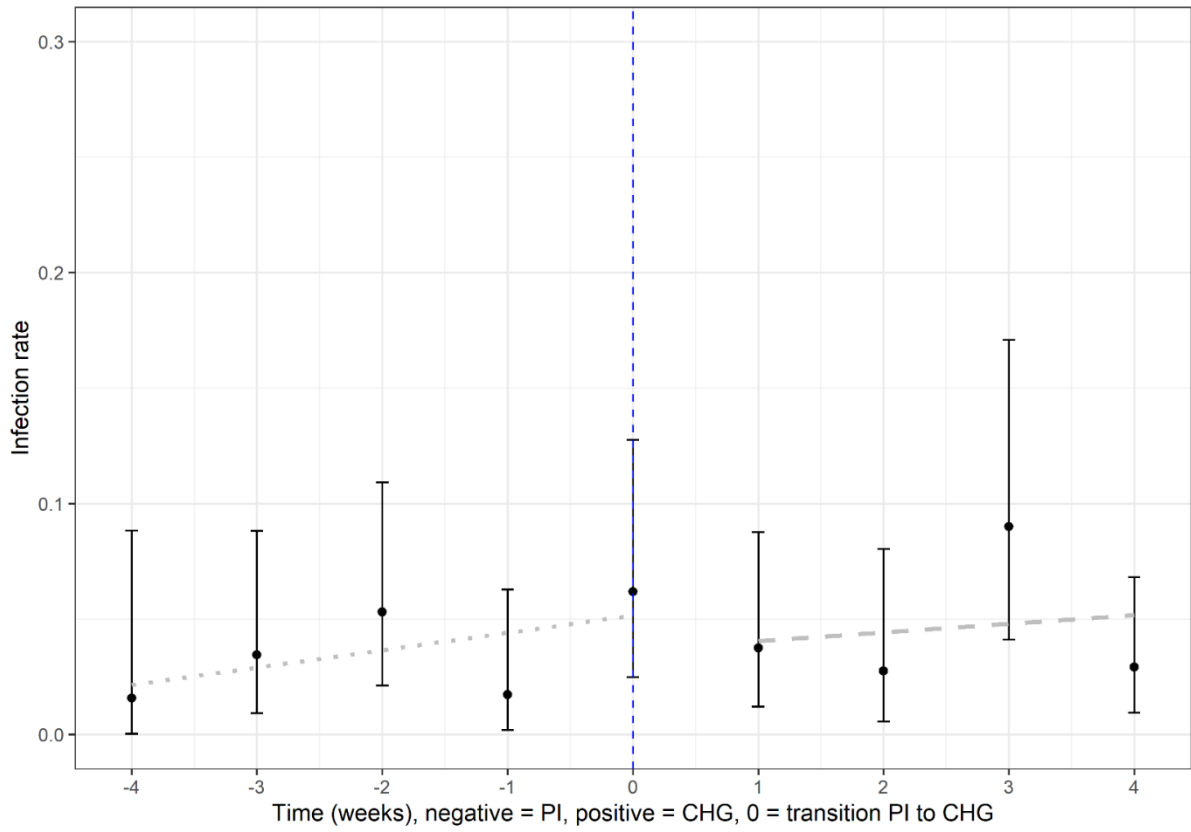
262 Center A



263

264

265 Center B



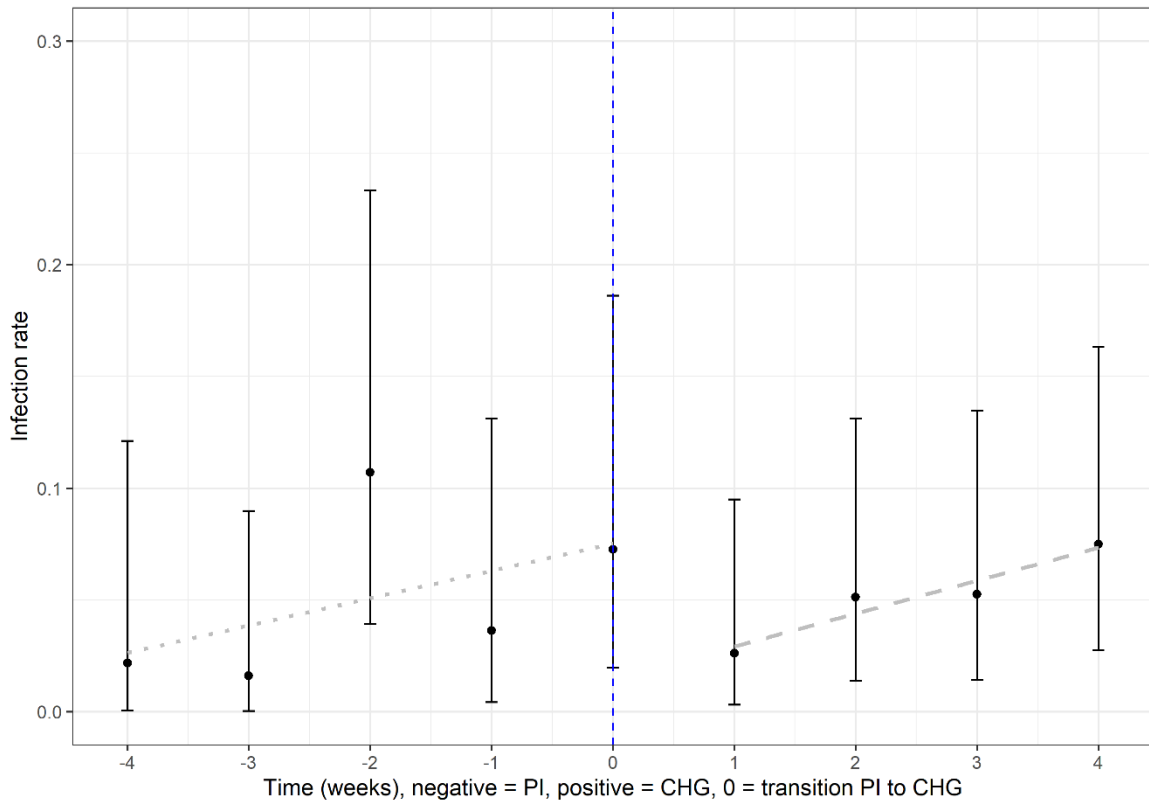
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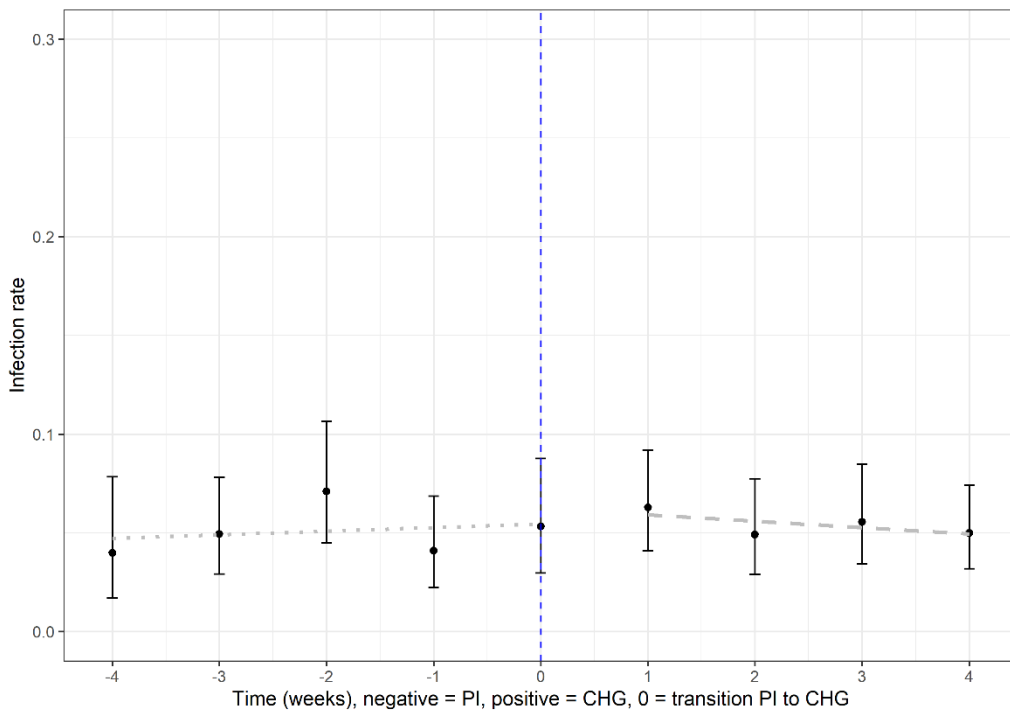
270 Center C



271

272

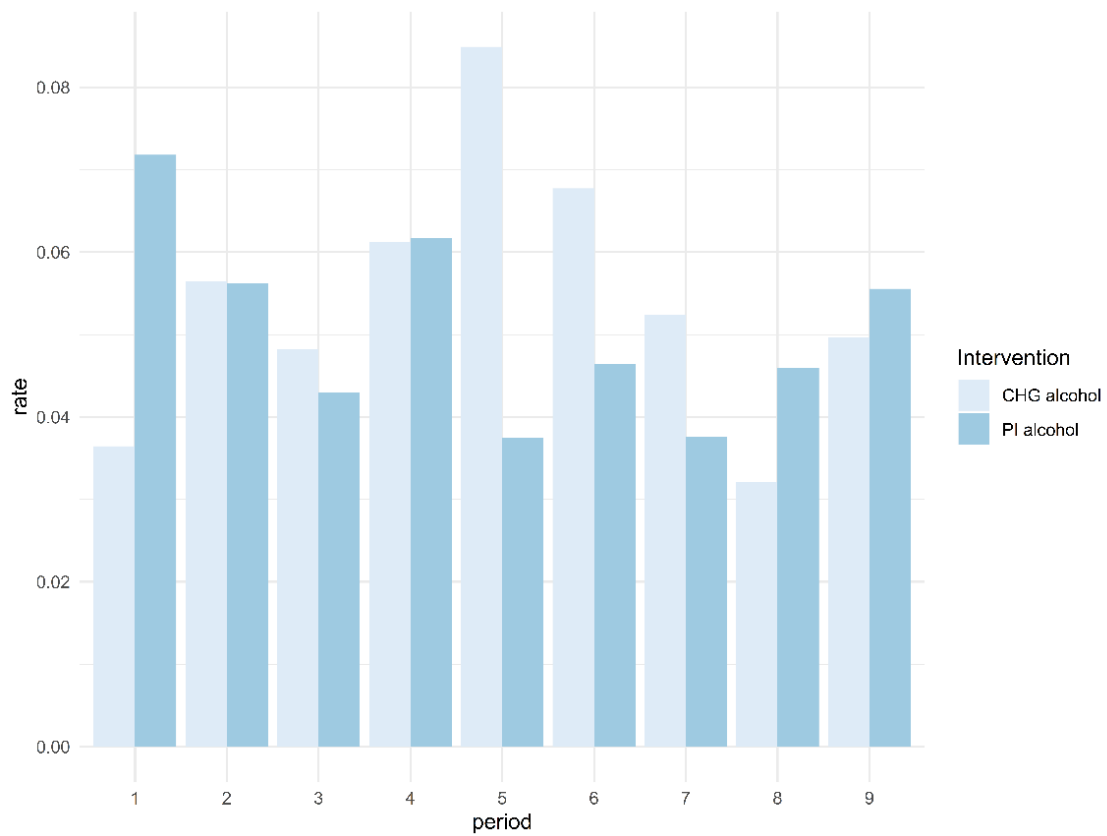
273 All centers



274

275 (CHG: Chlorhexidine gluconate, PI: Povidone iodine)

276 **eFigure 4:** Infection rates for the periods, stratified by intervention.

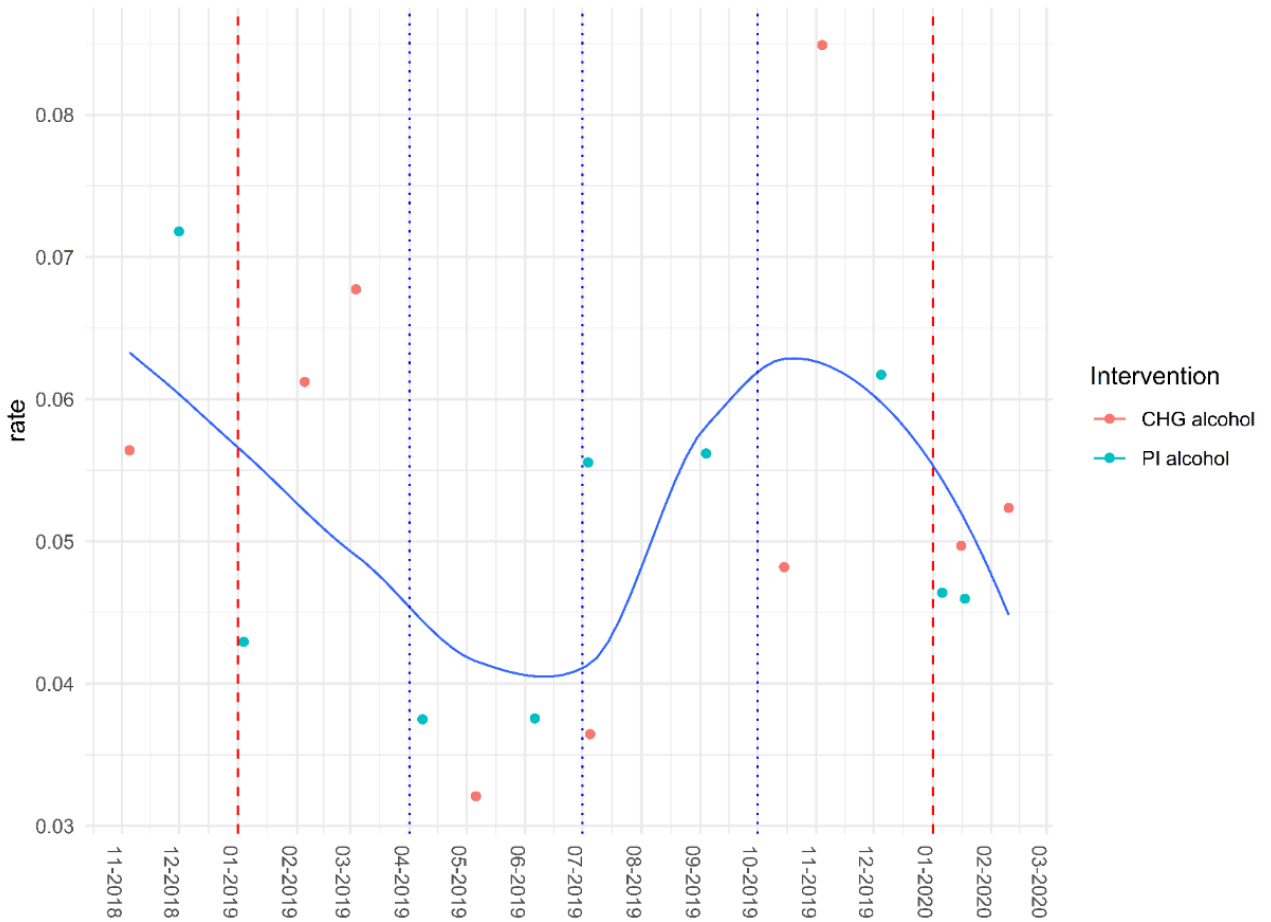


277

278 (CHG: Chlorhexidine gluconate, PI: Povidone iodine)

279 **eFigure 5:** Infection rates for the periods, plotted at the start of the calendar month and year, stratified
 280 by intervention; LOESS smoother added (solid, blue).

281 *Comment: There seems to be some degree of seasonality (which is known for SSIs), but the monthly*
 282 *changes of cluster randomization takes this into consideration.*



283
 284 (CHG: Chlorhexidine gluconate, PI: Povidone iodine)

285

286

287

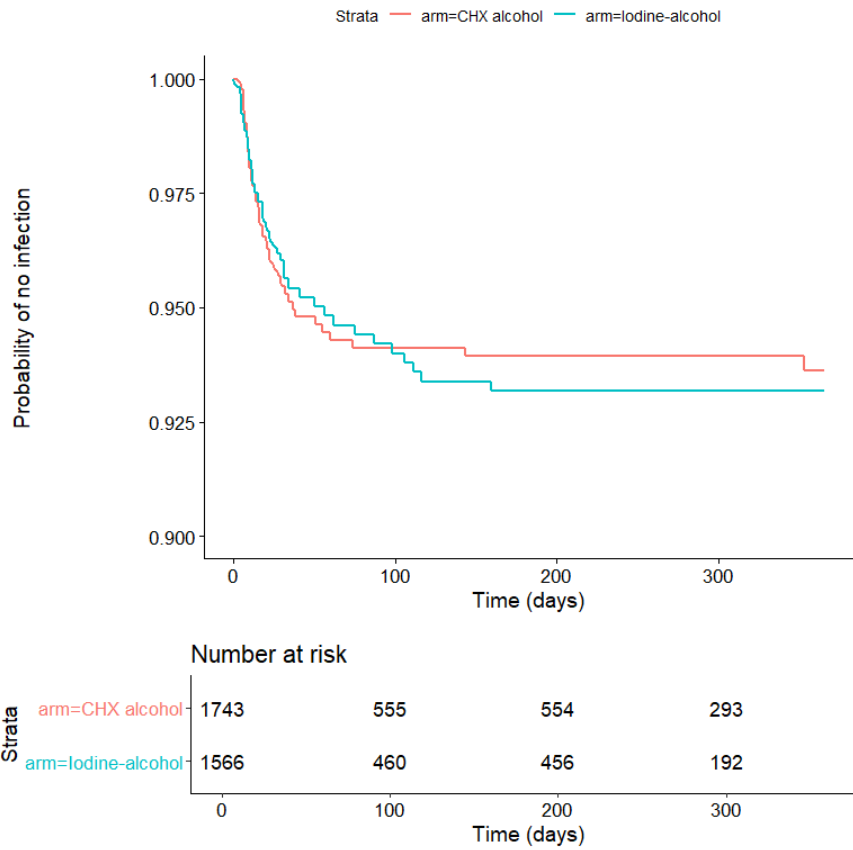
288

289

290 **eFigure 6:** Kaplan-Meier curve on probability of remaining free of surgical site infection (Logrank test, p
 291 =0.8).

292

293



294

295 (CHX: chlorhexidine)

296

297

298 **eFigure 7:** Randomized allocation of monthly clusters per center.

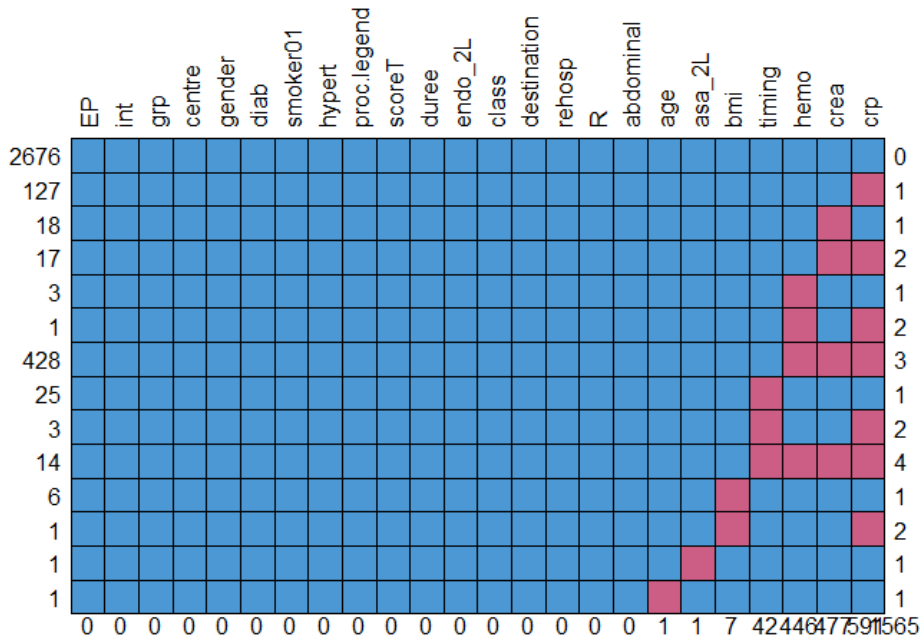
Study month	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	16	17	18
Center A	CHG	PI	CHG	PI	CHG	PI	PI	CHG	PI	CHG	CHG	PI	PI	CHG	CHG	PI	CHG	PI
Center B	PI	CHG	PI	CHG	PI	CHG	CHG	PI	PI	CHG	PI	CHG	CHG	PI	PI	CHG	CHG	PI
Center C	*	*	CHG	PI	PI	CHG	CHG	PI	CHG	PI	CHG	PI	PI	CHG	CHG	PI	PI	CHG

299
300 **Center C: Study started two months after the other centers.*

301 *(CHG: Chlorhexidine gluconate, PI: Povidone iodine)*

302

303 **eFigure 8:** Missingness patterns. Variables (top) with missing data (red) and variables without missing
 304 data (blue) with the number of patients with the respective missingness pattern (left) along with the
 305 number of number of missing variables (bottom); for example the top missingness pattern has 428
 306 patients in which hb (hemoglobin), crea (creatinine) and crp (C-reactive protein) are all missing.



307

308 **eTable 1:** Fitted logistic regression models with SSI as endpoint, including center as fixed effect and
 309 period as random effect (*Supplementary Analysis, Appendix B*).

	PI in alcohol	CHG in alcohol
<u>Primary outcome, supplementary analysis</u>	Odds ratio [95% CI]	Odds ratio [95% CI]
Model M1: Cluster and cluster-period effects treated as random ^a	0.9 [0.7, 1.3]	1 (reference)
Model M2: Cluster treated as fixed effect and cluster-period effects as random	0.9 [0.7, 1.3]	1 (reference)
GEE + sandwich or jack-knifed SEs	0.9 [0.7, 1.3]	1 (reference)
<u>Secondary outcomes</u>		
SSI rate (30 days), n/N (%)	27/1570 (1.7%)	27/1724 (1.5%)
Time to infection, median [IQR], days	13 [8, 25]	14 [9, 22]

310 ^a The intra-cluster correlation coefficient was estimated to be 0.002

311 (*GEE: Generalized estimating equation, SE: Standard error, SSI: Surgical site infection, CHG:*

312 *Chlorhexidine gluconate, PI: Povidone iodine, CI: Confidence interval*)

313

314

315 **eTable 2:** Comparison of the number of infections from cardiac surgeries for the trial and for the 18
316 months prior to the trial starting from Swissnoso Surveillance (*Supplementary Appendix C*).

% of infections	In hospital	0-30 days	31-90 days	>90 days
Trial	19%	46%	25%	10%
Swissnoso benchmark (18 months prior to start of study)	31%	36%	19%	14%

317

318

319 **eTable 3:** Preoperative antibiotic prophylaxis stratified by intervention (five most used agents).

	PI in alcohol		CHG in alcohol	
1	Cefuroxime	71.0%	Cefuroxime	71.0%
2	Amoxicillin/Clavulanate	11.5%	Amoxicillin/Clavulanate	13.4%
3	Gentamicin	2.4%	Gentamicin	3.2%
4	Metronidazole	2.2%	Metronidazole	2.4%
5	Piperacillin/Tazobactam	2.0%	Clindamycin	1.7%

320

321 (*CHG: Chlorhexidine gluconate, PI: Povidone iodine*)

322

323 **eTable 4:** Variables in *Table 1* stratified according to patients with no missingness (R=0) and patients with
 324 one or more missing records (R=1).

Patient characteristic ^a	Univariable		Adjusted (logistic regression model) R as dependent variable
	R=1 (missing)	R=0 (no missing)	Odds Ratio (95% CI)
N	645	2676	-
SSI	7.4%	4.8%	-
Center			
B	26.5%	33.4%	0.2 (0.1, 0.3)
A	69.5%	45.6%	1 (reference)
C	26.5%	33.4%	0.1 (0.1, 0.2)
Male sex	47.3%	71.3%	-
Age, years	53 [39, 65]	67 [58, 74]	-
Hemoglobin, g/dL	13.3 [11.9, 14.3]	13.6 [12.2, 14.7]	-
Creatinine, mg/dL	0.86 [0.71, 1.05]	0.93 [0.78, 1.10]	-
C-reactive protein, mg/dL	0.51 [0.28, 1.37]	0.30 [0.23, 1.00]	-
Diabetes	16%	11.2%	-
Hypertension	65%	41.6%	-
BMI, kg/m ²	29.5 [24.5, 38.9]	26.4 [23.6, 30.0]	1.1 (1.0, 1.1)
Abdominal surgery	94%	21.5%	196 (99, 400)
Timing relative to incision, min	-33 [-45, -21]	-41 [-54, -30]	1.0 (1.0, 1.0)
Surgery duration, min	70 [49, 132]	230 [173, 281]	1.0 (1.0, 1.0)
Endoscopic surgery	76.6%	21.5%	-
ASA-Score 3-5	57.4%	91.9%	0.6 (0.4, 0.8)
Wound class: clean	20.2%	76.6%	1 (reference)
Wound class: clean-contaminated	73.5%	12.6%	0.3 (0.2, 0.5)

325 *Categorical variables as N(%), continuous as median [IQR]*

326 ^a *Only those variables shown with systematic differences (p<0.1); CI confidence interval; not shown –*
 327 *Smoked cigarettes, Type of intervention, T-Score, ASA-Score, reintervention, rehospitalisation, destination;*
 328 *adjusted model identified using forwards selection then backwards deletion using p<0.1 as inclusion*
 329 *criteria.*

330 *(SSI: Surgical site infection, BMI: Body mass index, ASA: American Society of Anesthesiologists)*

331

332 **eTable 5:** Comparison of estimates from the complete case analysis and following multiple imputation.

	Complete case analysis	Multiply imputed data (K=50)
	Odds ratio (95% CI)	Odds ratio (95% CI)
Hemoglobin <=12 g/dL	1 (reference)	1 (reference)
>12 g/dL	0.89 (0.74, 1.70)	0.89 (0.60, 1.32)
Creatinine <=1.3 mg/dL	1 (reference)	1 (reference)
>1.3 mg/dL	1.21 (0.81, 1.75)	1.22 (0.83, 1.80)
C-reactive protein <=1.0 mg/dL	1 (reference)	1 (reference)
>1.0 mg/dL	1.87 (1.30, 2.67)	1.66 (1.14, 2.42)
Smoked cigarettes No	1 (reference)	1 (reference)
Yes	0.77 (0.51, 1.15)	0.84 (0.55, 1.29)

333

334

335 **Supplementary Appendix References**

- 336 1. Parienti JJ, Kuss O. Cluster-crossover design: a method for limiting clusters level effect in
337 community-intervention studies. *Contemp Clin Trials*. May 2007;28(3):316-23.
338 doi:10.1016/j.cct.2006.10.004
- 339 2. Maiwald M, Widmer AF. WHO's recommendation for surgical skin antisepsis is premature. *Lancet*
340 *Infect Dis*. Oct 2017;17(10):1023-1024. doi:10.1016/S1473-3099(17)30448-6
- 341 3. Donner A, Birkett N, Buck C. Randomization by cluster. Sample size requirements and analysis.
342 *Am J Epidemiol*. Dec 1981;114(6):906-14. doi:10.1093/oxfordjournals.aje.a113261
- 343 4. Arnup SJ, McKenzie JE, Hemming K, Pilcher D, Forbes AB. Understanding the cluster randomised
344 crossover design: a graphical illustration of the components of variation and a sample size tutorial. *Trials*.
345 Aug 15 2017;18(1):381. doi:10.1186/s13063-017-2113-2
- 346 5. Reich NG, Myers JA, Obeng D, Milstone AM, Perl TM. Empirical power and sample size
347 calculations for cluster-randomized and cluster-randomized crossover studies. *PLoS One*.
348 2012;7(4):e35564. doi:10.1371/journal.pone.0035564
- 349 6. Hemming K, Kasza J, Hooper R, Forbes A, Taljaard M. A tutorial on sample size calculation for
350 multiple-period cluster randomized parallel, cross-over and stepped-wedge trials using the Shiny CRT
351 Calculator. *Int J Epidemiol*. Jun 1 2020;49(3):979-995. doi:10.1093/ije/dyz237
- 352 7. Giraudeau B, Ravaud P, Donner A. Sample size calculation for cluster randomized cross-over
353 trials. *Stat Med*. Nov 29 2008;27(27):5578-85. doi:10.1002/sim.3383
- 354 8. Frieden TR. Evidence for Health Decision Making - Beyond Randomized, Controlled Trials. *N Engl J*
355 *Med*. Aug 3 2017;377(5):465-475. doi:10.1056/NEJMra1614394
- 356 9. Turner RM, White IR, Croudace T, Group PIPS. Analysis of cluster randomized cross-over trial
357 data: a comparison of methods. *Stat Med*. Jan 30 2007;26(2):274-89. doi:10.1002/sim.2537
- 358 10. Arnup SJ, Forbes AB, Kahan BC, Morgan KE, McKenzie JE. Appropriate statistical methods were
359 infrequently used in cluster-randomized crossover trials. *J Clin Epidemiol*. Jun 2016;74:40-50.
360 doi:10.1016/j.jclinepi.2015.11.013

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