This is an electronic appendix to the paper by Widdig *et al.* 2004 A longitudinal analysis of reproductive skew in male rhesus macagues. *Proc. R. Soc. Lond.* B **271**, 819–826. (DOI 10.1098/rspb.2003.2666.)

Electronic appendices are refereed with the text. However, no attempt is made to impose a uniform editorial style on the electronic appendices.

## **ELECTRONIC APPENDIX A**

Pair-wise relatedness between potential troop sires was determined as the maximum likelihood estimate of the kinship coefficient, using genetic markers. Likelihood maximisation was performed using comprehensive allele frequency data from the population involved. For a pair of animals genotyped at n loci, let  $a_{ijk}$  be the  $k^{th}$  allele (k=1..2) at the  $j^{th}$  locus (j=1..n) in the  $i^{th}$  animal (i=1..2); let  $f(a_{ijk})$  denote the population frequency of that allele. If h is the kinship coefficient of the pair, then the posterior probability  $\delta(h,a,a^*)$  of an allele a, drawn from the first individual, and allele  $a^*$ , randomly drawn from the second individual, to be identical by descent is

$$\delta(h, a, a^*) = 1_{a, a^*} \cdot \frac{h}{(1-h) \cdot f(a) + h}.$$

Here  $1_{a,a^*}$  equals unity if a and a\* are identical by state, and equals zero otherwise. If h is unknown, it can be estimated from the genotype data using an EM algorithm (Dempster et al. 1977). Let  $h_t$  denote the estimate resulting from iteration t. Then, by evaluating  $\delta$  for all possible pairs of alleles at all loci, a revised estimate of h is obtained in iteration t+1 as

$$h_{t+1} = \sum_{j=1}^{n} \sum_{k_1=1}^{2} \sum_{k_2=1}^{2} \frac{\delta(h_t, a_{1jk_1}, a_{2jk_2})}{4n}$$

This process generally converges to a local maximum of the likelihood curve; convergence to the global maximum was ensured by 1000-fold replication, using randomly selected start values of h.