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### An Investigation of Low Mortality in Certain Areas

By THEODORE D. WOOLSEY \*

It has been recognized for some time that the lowest general death rates in the United States are found in the West North Central States. South Dakota had the lowest age-adjusted death rate for white persons Nebraska was second in this respect with Iowa, Kansas, and in 1940. North Dakota tying for third place. It might be thought that this geographical concentration of low mortality would disappear when urban and rural death rates are compared separately, since rural rates are considerably lower than urban and the West North Central States all have large proportions of rural population. However, the lowest urban white mortality in the country in 1940 was experienced in Minnesota with Wisconsin second, Nebraska third, Utah fourth, California fifth, and Iowa, Kansas, Massachusetts, and Michigan tying for sixth. Eastern and Far Western States are represented, but the West North Central States, with one exception, are all among the lowest. For rural white mortality, also, all of the States of this geographical division, except Missouri, ranked among the ten lowest in the country on the basis of age-adjusted death rates. In fact, the low level of the rural mortality in this area is more striking than that of the urban mortality. The rural white age-adjusted death rate in the West North Central States in 1940 was 16 percent below the rural rate for the country as a whole, while the urban white rate was 6 percent below the corresponding national figure.

The object of this investigation was to discover whether any set of factors exists which would account for the low rates that have consistently been observed in the rural areas and small towns of the West North Central region. The geographical units studied were counties, and the opportunity to make use of county mortality data was provided by the publication for the first time in 1943 by the National Office of Vital Statistics (at that time a part of the Census Bureau) of a table showing 2-year totals of deaths by age by place of residence for each county and city of over 10,000 population in the United States

<sup>\*</sup> Biostatistician, Division of Public Health Methods, Public Health Service.

(1). The table refers to deaths in the years 1939 and 1940 and consequently presents the frequency of deaths at about the time of the 1940 census.

On the basis of the enumerated population at the time of the census and the deaths in the above mentioned table, age-specific death rates were computed for all the cities in the tabulation and for a systematic 10-percent sample of all the counties. For the particular purpose of this study the counties included in the sample were screened and certain counties were eliminated at the outset. The counties eliminated were as follows:

1. All counties containing a city of 10,000 or more population in 1940.

2. All counties having more than 30 percent of the population living in urban areas in 1940.

3. All counties in which the nonwhite population in 1940 formed 10 percent or more of the total or numbered 10,000 or more.

4. All counties with a population of less than 2,000 in 1940.

5. All counties in which 5 percent or more of the deaths in 1940 occurred in resident institutions.<sup>1</sup>

These counties were eliminated to reduce the number of factors causing variation in the death rate and thus facilitate the study of those which might be responsible for the low rates in the rural part of the West North Central region. While it is well known that a low proportion of nonwhite population, or a low proportion of urban population, is usually associated with lower death rates, it was also known in this case that the essentially rural white character of the populations under particular consideration was not the only cause of the low mortality. Hence, it was desirable to deal with a group of counties that was fairly homogeneous in this respect. It would have been preferable on theoretical grounds to include all types of counties and attempt to establish a control on the factors that were not under study, but this would have made the work exceedingly laborious. The few counties with large numbers of deaths in resident institutions or exceptionally small populations were omitted for fairly obvious reasons.

This process of elimination reduced the original sample of 308 counties to 124. This smaller sample represents, therefore, about 10 percent of the counties having largely rural white populations, excluding the very smallest ones and those containing large resident institutions.

<sup>&</sup>lt;sup>1</sup> The rules of the National Office of Vital Statistics for allocation of deaths to place of residence provide that deaths in institutions, such as mental institutions, orphanages, homes for the aged, and so forth, where the length of stay is usually long, are not re-allocated to the place of prior residence of the decedent. Hence, certain counties, in which the population of these resident institutions is a considerable proportion of the total, tend to have abnormally high death rates.

	Number of all counties <sup>1</sup>	Number of counties in sample
Middle Atlantic East North Central Vest North Central	3, 090	124
New England	67	4
Middle Atlantic.	146	5
East North Central	436	22
West North Central	620	37
South Atlantic	577	8
East South Central	364	15
West South Central	470	17
Mountain	277	14
Pacific	133	2

The distribution of these counties according to the nine major geographical divisions of the country was as follows:

<sup>1</sup> Includes certain jurisdictions sometimes classified as cities and also certain counties not organized as local governments.

For each of these counties a single index of mortality was computed by (a) calculating the annual age-adjusted death rate by the direct method (using the total United States population in 1940 as a standard) for the 2 years 1939 and 1940 combined; (b) finding the ratio of the age-adjusted rate to the crude death rate for the same 2 years; (c) multiplying the ratio obtained in (b) by the crude death rate for the 5-year period 1938-42; (d) adjusting the rate thus obtained for under-registration of deaths using the assumption described below.

It was apparent from inspection of the indices of mortality obtained in step (c) above that some of the low death rates were simply a reflection of incomplete registration of deaths. The only quantitative basis that exists for estimating the incompleteness of death registration is the information available on birth registration from the 1940 Birth Registration Test, conducted in conjunction with the census. Although nothing is accurately known about the relative completeness of death registration, there is reason to believe that death registration is better than birth registration and that many of the same factors which account for unregistered births in a given area also result in unregistered deaths. Consequently, it was assumed that the proportion of unregistered deaths in any county was onehalf of the proportion of unregistered births estimated from the Birth Registration Test (2). Under this assumption the proportion of unregistered deaths in the country as a whole would be 3.75 percent, since the proportion for births was found to be 7.5 percent in 1940. Even if the proportion of unregistered deaths is more or less than one-half the proportion for births, a correction based on onehalf probably results in a death rate for each county that is highly correlated with the "true" death rate for that county, except where

factors peculiar to death registration are responsible for the lack of completeness.

The factor actually used to correct for under-registration was designed so that its value for the country as a whole would be equal to 1. It therefore took the form:

Correction for under-registration =  $\frac{96.25}{100 - \frac{1}{2}}$  (percent of births unregistered)

The final index of mortality was the ratio of the estimated ageadjusted death rate in the county for the years 1938-42 (corrected for under-registration) divided by the death rate in the entire country in 1940. Its value for all counties in the country would be very close to 1.000, and for any given county it is roughly proportional to the estimated age-adjusted death rate of that county.

The strength of the association of the index of mortality with a number of different variables was measured by the method of partial correlation. Some of the variables were discarded on the grounds that they appeared to explain little of the observed variation in mortality. One method that was used to examine new variables was to set up a least-squares linear regression equation on the basis of several variables known to have a significant association with the mortality. The differences between the actual and computed values of the dependent variable were then computed. (These will be termed the "deviates.") The association of these deviates with new variables was then examined either by means of scatter diagrams or zero order correlation coefficients.

It was decided in advance that the question of when to stop would be answered in terms of the original objectives of the investigation. The deviates obtained from a particular set of independent variables would be subjected to an analysis of variance. When an F test indicated that the mean square error between geographical regions was no longer significantly greater than the within-region mean square, then it would be concluded that the association of mortality with these variables was sufficient to explain the low death rates that were being studied. The geographic regions used in this analysis were somewhat more detailed than the 9 for which census statistics are usually shown. With one exception, however, the 14 regions used did not cut across the boundaries of the 9 major divisions. (Counties in Wisconsin were combined with those in Minnesota instead of with those in Ohio, Indiana, Illinois and Michigan.)

### Results

It is not necessary to describe in detail all of the independent variables that were investigated. Some were not investigated as thoroughly as others. Hence, the set finally selected is not the only combination that would explain the observed geographic variation. Some were rejected simply because other indices seemed to measure the same thing better. Those rejected were variables measuring per capita retail sales and effective buying income, the proportion of aged persons in the population, the proportion of persons of European stock other than that from northwest Europe, educational status of the population, geographical density of the population, and percentage change in population between 1930 and 1940. Those finally included in the correlation analysis were:

 $X_2$ =rural level of living index, 1940 (3).

- $X_3$  = square root of the proportion of the population of the county in 1930 that was either born in a northwestern European country or born of parents one or both of whom were born in a northwestern European country (4).
- X<sub>4</sub>=percentage of births to residents of the county that were delivered in hospitals in 1939 and 1940 (5).
- $X_5$  = percentage of workers employed in agriculture in 1940 (6).

The rural level of living index (3) is described as "the result of an attempt to indicate in simple form the relative level of living in each county of the United States. It deals with the level of living of all rural families, both farm and nonfarm. The index for each county is a weighted average of indexes for rural farm and rural nonfarm families; the two indexes are weighted according to the proportions of the rural population of the county that are rural-farm and rural-nonfarm" (6, p. 428).<sup>2</sup> This composite level of living index contains measures of the crowding of the dwelling units, and the level of education, and prosperity of the families living in the county.

Since country of birth of foreign stock was not available by county in the 1940 census, it was necessary to go back to the previous census to obtain the numbers of foreign born and of native born of foreign or mixed parentage according to country of origin. Originally, proportions of persons of other types of foreign stock were included in the analysis, but it appeared that the only factor that seemed to be significantly associated with mortality was the proportion of persons of northwestern European stock, including persons whose families had immigrated fairly recently from Scandinavia, the Low Countries, France, Germany, and the British Isles. The square root of this proportion was more symmetrically distributed throughout the range of values and had a higher degree of linear association with mortality than the proportion itself.

<sup>&</sup>lt;sup>3</sup> For a description of items included in the index, see this reference. For further detail see reference 3.

It was hoped that the percentage of births occurring in hospitals would measure the availability and utilization of medical facilities a factor that was not measured in the level of living index. This turned out to be the least useful of the four independent variables.

The percentage of workers employed in agriculture accounted for more of the variation in mortality in these rural counties than did any of the other three variables.

Table 1 shows some of the correlation coefficients that were obtained using this set of variables. A "z"-test for significance indicated that only the two numerically largest 3rd-order correlation coefficients

# Table 1. Results of correlation analysis of mortality index for counties with certain socio-economic measures

- 1 = index of mortality
- 2=rural level of living

 $3 = \sqrt{\text{percent Northwestern European stock}}$ 

4=percent births in hospitals

5=percent emp	oloyed in	agriculture
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r <sub>ij</sub>	r <sub>ij`k</sub>	<b>Γ</b> ij∙k1m
$\begin{array}{c} r_{12} =32 \\ r_{13} =43 \\ r_{14} =16 \\ r_{15} =37 \\ r_{23} = .66 \\ r_{24} = .67 \\ r_{25} =26 \\ r_{34} = .63 \\ r_{35} = .13 \\ r_{45} =17 \end{array}$	$\begin{array}{c} r_{13,2} =22 \\ r_{14,2} = .06 \\ r_{15,2} =46 \\ r_{12,3} =03 \\ r_{14,3} = .12 \\ r_{15,3} =32 \\ r_{12,4} =21 \\ r_{13,4} =33 \\ r_{15,4} =40 \\ r_{12,5} =41 \\ r_{13,5} =38 \\ r_{14,5} =22 \end{array}$	$r_{12,345} =26$ $r_{13,245} =16$ $r_{14,235} = .14$ $r_{15,234} =37$

could be considered significant. The square of the multiple correlation coefficient  $(R^{2}_{1.2345})$  was equal to .35. It may seem that accounting for only 35 percent of the total variation represents very little gain in knowledge. However, the objective had been established of explaining that variation which was strongly related to geographical location. Hence, it was not necessary and, indeed, it would have been an overwhelming task to examine the scores of available chracteristics of the counties, such as climatological and demographic characteristics, in an attempt to reduce the residual variation still further.

The least-squares regression plane for the regression of  $X_1$  (mortality) on the four variables described above was found to be:

 $X'_{1} = 1.2565 - .0023X_{2} - .0014X_{3} + .0011X_{4} - .0038X_{5}$ 

The deviates,  $X_1 - X'_1$ , were grouped into 14 geographic subdivisions of the country. The arithmetic means of these deviates and of  $X_1$ ,  $X'_1$ , and each of the other indices for all the sample counties falling into any one region are shown in table 2. The table suggests that not all of the regional variation was removed by taking account of the linear association of mortality with the four independent variables chosen. However, the only marked departures from the regression plane are the negative deviations for the six Arkansas

			Unweighted means of indices 1 for sample counties in each region								
Regions and States in which sample counties were located		Num- ber coun- ties in sample	Rural level of living	$\sqrt{\frac{10 \text{ x}}{\frac{\% \text{ of pop. of}}{\text{N. W. Eur.}}}}$	% births in hosp.	% emp. in agric.	Mor- tality <sup>2</sup> X <sub>1</sub>	Com- puted mor- tality <sup>2</sup> X' <sub>1</sub>	X1-X'1		
I III IV VI VII VIII IX X XI XII XIII	Me., Vt., Mass N. Y., Penn. Ohio, Ind., Ill., Mich Wis., Minn N. Dak., S. Dak. Nebr., Iowa, Kans Missouri. Va., W. Va., N. C Ala., Tenn. Kentucky. Arkansas Okla., Tex. Colo., Utah, Mont., Idaho, Wyo	4 5 19 9 16 6 8 5 10 6 11 14	126 120 111 126 108 122 90 86 69 81 75 104 112	21. 70 19. 88 31. 43 66. 01 57. 24 44. 81 25. 65 8. 30 5. 36 7. 63 8. 33 21. 06 40, 83	42.76 27.85 26.61 50.21 41.43 34.57 10.23 10.12 5.81 6.92 5.21 29.53 46.30	21. 8 27. 5 48. 2 59. 3 68. 8 56. 9 55. 2 46. 9 59. 2 55. 0 64. 9 59. 0	. 893 . 880 . 827 . 711 . 664 . 713 . 713 . 784 . 871 . 901 . 925 . 680 . 825	. 903 . 882 . 807 . 708 . 714 . 739 . 819 . 819 . 819 . 819 . 874 . 862 . 836 . 800	$\begin{array}{c}010\\002\\ +.020\\ +.020\\050\\050\\035\\012\\ +.025\\156\\ +.025\\ \end{array}$		
ĸıv	Oreg., Calif	14 2	112	40. 83 33. 01	46. 30 72. 49	51.6 42.4	. 833 . 972	. 799 . 858	+.034 +.114		

Table 2.	Means	of	indices	for	sample	e coun	uties :	in	each	of	14	regions	

<sup>1</sup> For description of indices, see text. In a later table, weighted means of indices, with county populations

as weights, are used.
 as weights, are used.
 as weights, are used.
 bits is expressed as a ratio and not as a death rate. The age-adjusted death rate corrected for under-registration of deaths as described in this study can be obtained by multiplying each of these ratios by the United States crude death rate in 1939 and 1940, which was 10.7 per 1,000 population.

counties and the positive deviations of the West Coast counties, the latter being of little significance because of the small number of counties in the sample.

The analysis of variance of the county deviates within and between regions is presented in table 3. The results indicate that the remaining geographical variation cannot be explained as a chance result. However, examination of the between-region sum of squares showed that more than one-half of the total sum of squares arose from the six counties in Arkansas (which had been set up as a separate region). All of the deviates for these counties had a negative sign. Arkansas is not among the West North Central States for which the low rural mortality was being investigated, but the mean of the mortality index for these six counties is the second lowest of any of the regions. The regression plane fits the West North Central counties reasonably well but misses badly in predicting the Arkansas counties.

		Sums of	Mean
	d. f.	squ <b>ar</b> es	squares
Between regions	13	0. 2893205	0. 0222554
Within regions	110	. 9611423	. 0087377
Total	123	1. 2504628	. 0101664

 
 Table 3. Analysis of variance and test for significance of mortality index for counties with certain socio-economic measures held constant

# $\mathbf{F} = \frac{0.0222554}{0.0087377} = 2.55^{1}$

### 1% point = 2.30

 $^1$  If the deviate for each county is weighted by the population of the county, the value of F is reduced to 2.32.

If the six Arkansas counties are omitted from the analysis of variance, a conservative estimate of the between-region sum of squares is obtained without recomputing the regression function, i. e., an estimate which is certainly not lower than that which would have been found had the six counties been omitted at the outset. The revised analysis of variance is shown in table 4.

The mean square estimated from the regional means is not significantly greater than the mean square estimated from within-region variation. It can be said, therefore, that, if the six counties in Arkansas are excluded, the remaining geographical variation is not too great to have arisen as a result of chance.

Table 4. Revised analysis of variance of table 2 with six Arkansas counties omitted

	d. f.	Sums of squares	Mean squares
Between regions Within regions	12 105	0. 1355413 . 9074890	0. 0112951 . 0086428
Total	117	1. 0430303	. 0089148

$$\mathbf{F} = \frac{0.0112951}{0.0086428} = 1.31^{12}$$

1% point = 2.35

5% point = 1.85

<sup>1</sup> If the deviate for each county is weighted by the population of the county, the value of F is reduced to 1.19.

# Discussion

The only factor that can reasonably be introduced to explain the low mortality in the Arkansas counties is incompleteness of death registration. In other obvious respects these counties are not different from some of the other counties in the sample, except in ways that would appear to make the death rates higher rather than lower. On the other hand, the Arkansas counties were no worse than some other counties as regards birth registration. Hence, it seems most likely that death registration is poor in these counties for some reasons that do not necessarily influence birth registration.

			Per	cent employ	ed in agricu	lture		
Rural level of living index	9.2-	-46.9	. 47.0	-55.9	56.0	-61.4	61.5	-83.8
mdex	Mortality index	Frequency and mean <sup>2</sup>		Frequency and mean <sup>2</sup>	Mortality index	Frequency and mean <sup>2</sup>	Mortality index	Frequency and mean <sup>2</sup>
50-91	0.800 .959 .986 1.048	4 . 964	0. 763 . 782 . 784 * 802 . 916 . 937 1. 007	7 . 855	*0. 570 * 708 834 863 885 885 890	6 . 783	*0. 535 647 * 695 * 768 768 768 852 864 876 876 876 878 924 928 924 928 975 1.066 1.204	15 . 878
92–111	0. 676 779 788 838 904 939 961 1. 000 1. 042 1. 079	12 . 879	0. 740 . 817 . 818 . 825 1. 088	5 . 851	0. 673 . 720 . 752 . 768 . 756 . 794 . 803 . 886	8 · .787	0. 546 . 655 . 717 . 747 . 867 1. 040	6 . 757
112-122	0. 799 . 800 . 842 . 869 . 902 . 930 1. 016	7 . 848	0. 689 . 689 . 693 . 731 . 759 . 779 . 815 1. 145	8 . 835	0. 647 . 653 . 695 . 712 . 727 . 763 . 784 . 909	8 . 721	0. 515 . 690 . 695 . 705 . 712 . 926	6 . 717
123-145	0. 734 . 785 . 703 . 829 . 838 . 882 . 882 . 885 . 922	8 . 831	0. 665 . 671 . 728 . 740 . 742 . 750 . 760 . 780 . 792 . 794 . 895	11 . 761	0. 666 667 668 671 702 730 760 763 769	9 . 710	0. 671 . 683 . 738 . 741	4 . 704

 Table 5. Mortality index in 124 counties distributed according to percent employed in agriculture and rural level of living 1

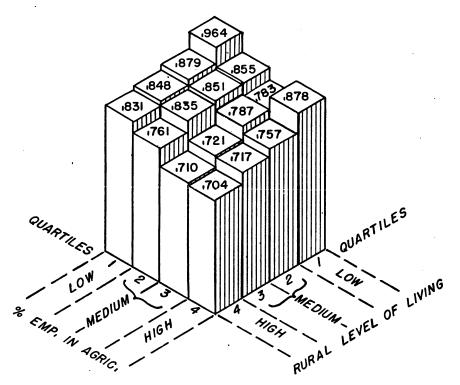
<sup>1</sup> Values marked with an asterisk are those for the 6 counties in the sample located in Arkansas.
 <sup>2</sup> Weighted mean using populations as weights.

The variables of real importance in explaining the geographic variation of mortality in this sample of 124 rural white counties were: (1) rural level of living, and (2) proportion of working population

(1) rural level of living, and (2) proportion of working population engaged in agriculture. In each case the association is a negative one. This relationship is more clearly shown in table 5 in which the mortality index for each of the 124 counties is listed according to its quartile position in the distribution of the counties according to rural level of living and percent employed in agriculture. In each cell of the table the weighted mean of the values in that cell is also presented. The weights used were the 1940 populations of the counties. It will be seen that there is a fairly regular downward gradient from left to right and from top to bottom. The six Arkansas counties are marked with asterisks. The chart shows the same relationship graphically.

The negative relationship between living standards and mortality is understandable and, in fact, has been recognized for a long time.

MORTALITY



Mortality in 124 rural white counties classified according to high, medium, and low rural level of living and percent of population over 14 years of age employed in agriculture; 1940 (see text for description of mortality index).

The data presented here, however, seem to indicate that in the population of rural areas and small towns the proportion of the population engaged in agriculture is at least an equally important factor in determining the age-adjusted death rate. This investigation does not, of course, show how many intermediate factors, themselves associated with percent employed in agriculture, should be inserted in the causal chain between the farm employment and the death rate. In fact, if all the causes of low and high mortality were known, this particular relationship might prove to be highly artificial. Some light is thrown upon it by a study of life tables for those States with high proportions of farm population contrasted with those having low farm populations, holding constant the median family income or some other measure of living standards. Such data suggest that if the index of mortality for the sample of counties used here had been confined to mortality under, say, 40 years of age, the strength of the association with rural level of living would have increased while the correlation with percent employed in agriculture would actually have been reduced. If, on the other hand, the mortality index had been based upon death rates over age 40, using, for example, the reciprocal of the expectation of life at age 40 as a measure, exactly the reverse would have been true.

If it should prove to be a fact that a particular environment or occupation is conducive to improve expectation of life beyond middle age, the reason for some of the now inexplicable variation in mortality from one area to another would become clear. Since death rates, nowadays, whether age-adjusted or crude, tend to be determined more and more by the age-specific rates at the older ages, some knowledge of the factors influencing these rates would contribute greatly to the usefulness of mortality statistics.

### Summary

This study was undertaken in the hope of finding at least a partial explanation of the particularly low mortality that prevails in the rural parts of the West North Central region of the United States. It is pointed out that urban mortality is also low in these States, but the difference is not so marked.

The geographical units studied were counties. Out of the 3,090 counties in the country a sample of 124 was selected, comprising approximately 10 percent of the counties that had in 1940: (a) No city of 10,000 or more population; (b) not more than 30 percent of the population classified as urban; (c) not as much as 10 percent or as many as 10,000 nonwhite population; (d) not less than 2,000 population in all; (e) not as much as 5 percent of the 1940 deaths occurring in resident institutions. This preliminary elimination of certain types

of counties was made in order to reduce the effect of certain factors having an influence upon mortality that is already well recognized. For this same reason the death rates were age-adjusted, and an attempt was made also to adjust for assumed incompleteness of death registration by using information obtained in the 1940 Birth Registration The final index of mortality used was roughly proportional to Test. the age-adjusted death rates for the sample counties in the period 1938-1942.

The association of these measures of mortality with certain demographic and socio-economic characteristics of the populations of each of the 124 counties in the sample was studied. The plan was to find several factors which would be sufficient to account for the observed geographical differences in mortality. It was found that if two factors, (1) the "rural level of living" (a measure of standard of living in rural areas), and (2) the proportion of workers employed in agriculture, were held constant for each county, the remaining geographical differences could be explained as chance variation. This held true, however, only when the six Arkansas counties in the sample were omitted. The exceptionally low mortality in these counties was thought to be due to under-registration of deaths that was not associated with under-registration of births and, hence, had not been taken into account in the construction of the mortality index.

This study seemed to suggest that the mortality of a rural area is at least as dependent upon the proportion of its workers employed in agriculture as it is upon its standard of living. The question was raised for further study as to whether a high proportion of persons in farm employment might be found to be significantly associated with mortality at the older ages, in contrast with the standard of living which has been generally supposed to have more influence on death rates at the beginning of life.

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- (6) County data book—a supplement to the statistical abstract of the United States. Bureau of the Census, Washington, D. C., 1947.

## **INCIDENCE OF DISEASE**

No health department, State or local, can effectively prevent or control disease without knowledge of when, where, and under what conditions cases are occurring

### UNITED STATES

### **REPORTS FROM STATES FOR WEEK ENDING JULY 2, 1949**

A total of 479 cases of poliomyelitis was reported, as compared with 409 last week, 362 for the corresponding week last year, and a 5-year (1944-48) median of 220. Of the 22 States reporting more than 4 cases each, 4 showed no increase. The largest increases over last week were reported in Arkansas (42 to 59), Louisiana (4 to 15), Kansas (2 to 13), and Minnesota (16 to 25). The 6 other States reporting increases of 5 or more cases are as follows (last week's figures in parentheses): Massachusetts 10 (4), New York 12 (7); Ohio 6 (0), Indiana 7 (2), Illinois 11 (3), Texas r18 (113). The total reported since March 19 (average week of seasonal low incidence) is 2,273, as compared with 2,020 for the corresponding week last year and a 5-year median of 874.

Of the total of 32 cases of Rocky Mountain spotted fever reported for the week (last week 26, 5-year median 26), 23 occurred in the South Atlantic and South Central areas, 6 in 4 States of the Mountain area, and 1 case each in New Jersey, Pennsylvania, and Ohio.

During the week 2 cases of anthrax were reported in New York State.

Current and cumulative figures are above the corresponding medians for measles, poliomyelitis, infectious encephalitis, Rocky Mountain spotted fever, and tularemia.

Deaths recorded in 94 large cities in the United States during the week totaled 8,979, as compared with 8,877 last week, 8,963 and 8,078, respectively, for the corresponding weeks of 1948 and 1947, and a 3-year (1946-48) median of 8,079. The total for the year to date is 246,775, as compared with 250,144 for the corresponding period last year. Infant deaths during the week totaled 686, as compared with 583 last week and a 3-year median of 630. The cumulative figure is 16,853, as compared with 17,637 for the same period last year.

uy.	22, 1949		02	-		
	Rabies in ani- mals		6	18 <sup>41</sup>		
	Whoop- ing cough	16 106 13	190 57 57	8°8888	4 81	08 88280°
	Typhoid and para- typhoid fever •	I	33		4-1	
	Tulare- mia			I	- m	
	Small- pox					
	Scarlet fever	30 30 10	4 13 44	<b>6∞1</b> 82	01 8 14 9 10 10 10 10 10 10 10 10 10 10 10 10 10	
por vea)	Rocky Moun- tain spotted fever					
(Leaders indicate that no cases were reported)	Polio- myelitis	3	, 200	6 10176 501176	11 <b>4</b> 23	
tnat no ca	Pneu- monia	S. 83	28 28	38 30 33 16 38 20 39 30 39 30 39 30 39 30 30 30 30 30 30 30 30 30 30 30 30 30	4	11 15 20 1338 15 4 20
rs indicate	Menin- gitis, menin- gococcal	2	8-8	∞ <b>-</b> ⊓≈	2	
(Leader	Measles	358 240 e 43	1, 097 730 649	515 22 320 1, 229	833°*833	2383373889
	Influ- enza	1	(o)	0 0 0	- 7	58 57
	Enceph- alitis, infec- tious		4		, , , , , , , , , , , , , , , , , , ,	
	Diph- theria	1	849	70 67		000000
	Division and State	NEW ENGLAND Maine New Hampshire Vermont. Rhode Slausetts Connecticut Connecticut	New York. New Jersey. Pennsylvania. EAST NORTH CENTRAL	Ohio. Indiana Michigan • Wisconsin.	Winnesota. Minnesota. North Dakota. North Dakota. South Dakota. Kansas. SOUTH ATTANTIC	Delaware Maryland • Dist. of Col. Virginia West Virginia Worth Carolina South Carolina Georgia Florida.

Telegraphic case reports from State health officers for week ended July 2, 1949

(Leaders indicate that no cases were reported)

6 19	6 16		1		
6 11 11	31 4 157	10 11 15 12 15	284	1, 241 2, 170	27, 471 , 50, 116 (39th) Oct. 2 37, 504 81, 382
4 0 6	4193	•	2	74 111	1, 203 1, 586 (11th) Mar. 19 1, 111 1, 111
1	<i>1</i> 0 ⊢∞			สล	640 471
			•	4	40 (35th) Sept. 4 328 328
01400	00H0	4 15.0	15 7 4 41	473 1, 223	56, 601 82, 114 (32d) Aug. 14 78, 299 120, 685
4	4	3		32 26 33	238
14 14	59 11 118	00040-	59 ° 7	479 220	3, 197 1, 271 (11tb) Mar. 19 2, 273 874
0857 <i>0</i>	8208	6 10 2 8 0 1 4 6	16 11	933	50, 103
000	00	88	4	48 81	1, 992 3, 964 (37th) Sept. 18 5, 468
5824 8824	211 219	8678738	8888 8	8, 006 6, 034	568, 871 517, 296 (35th) Sept. 4 621, 224 552, 242
<u>α</u> ω	8 550 8	. 28	10	• 511 583	74, 185 188, 206 (30th) July 31 110, 455 332, 369
	I		, 1	12 10	2260 2260
5 3	010-10	6	4	. 72 159	3, 688 6, 165 6, 165 110 1017 10 13, 731 13, 731
EAST SOUTH CENTRAL Kentucky Tennessee Alabama	WEST SOUTH CENTRAL Arkansas. Louisiana. Oklahoma. Texas. MOUNTAIN	Montana Idaho. Vyoming Colorado. New Mexico. Utana. Nevada.	PA <b>GUTIC</b> Washington Oregon California	Total Median, 1 <del>944–4</del> 8	Year to date 26 weeks. Median, 1944–49. Seasonal low week ends. Since seasonal low week. Median, 1943–48 b.

Period ended earlier than Saturday. .

<sup>b</sup> The median of the 5 preceding corresponding periods; for poliomyeitts and typhoid fever the corresponding periods are 1944-45 to 1948-49, inclusive.
 <sup>b</sup> Fine median of the 5 preceding corresponding periods.
 <sup>c</sup> Including cases reprotoced inflection and septic sore threat.
 <sup>e</sup> Including paratyphoid fever; currently reported septics for threat.
 <sup>e</sup> Including paratyphoid fever; currently reported septics one threat.
 <sup>e</sup> Including paratyphoid fever; currently reported septics for threat.
 <sup>e</sup> Including paratyphoid fever; currently reported septics for threat.
 <sup>e</sup> Including paratyphoid fever; currently reported septics of threat.
 <sup>e</sup> Including paratyphoid fever; currently reported septics for threat.
 <sup>e</sup> Including paratyphoid fever; currently reported septics for threat.
 <sup>e</sup> Including paratyphoid fever; currently reported separately, as follows: Massachusetts 2, Pennsylvania 1, New York 3.
 <sup>e</sup> Anthrax: New York 3.
 <sup>e</sup> Anthrax: Streptococcal threat 1, measles 4.
 <sup>e</sup> April onset.
 <sup>e</sup> April onset.

### 924

### PLAGUE INFECTION IN LINCOLN COUNTY, WYO.

Under date of June 30 plague infection was reported proved in a pool of 45 fleas from 15 ground squirrels, *Cytellus armatus*, and in a pool of 9 fleas from 9 ground squirrels, *Citellus richardsonii elegans*, collected on June 17 at a location 5 miles northeast of Opal, Lincoln County, Wyo.

### DEATHS DURING WEEK ENDED JUNE 25, 1949

[From the Weekly Mortality Index, issued by the National Office of Vital Statistics]

	Week ended June 25, 1949	
Data for 94 large cities of the United States:         Total deaths.         Median for 3 prior years.         Total deaths, first 25 weeks of year.         Deaths under 1 year of age.         Median for 3 prior years.         Deaths under 1 year of age.         Median for 3 prior years.         Deaths under 1 year of age, first 25 weeks of year.         Deaths under 1 year of age, first 25 weeks of year.         Data from industrial insurance companies:         Policies in force.         Number of death claims.         Death claims per 1,000 policies in force, annual rate.         Death claims per 1,000 policies, first 25 weeks of year, annual rate.	8, 877 8, 603 237, 797 583 629 16, 167 70, 388, 886 12, 205 9, 0 9, 5	8, 576 241, 181 612 16, 991 71, 043, 978 12, 346 9, 1 10, 0

### FOREIGN REPORTS

### CANADA

Provinces—Notifiable diseases—Week ended June 11, 1949—During the week ended June 11, 1949, cases of certain notifiable diseases were reported by the Dominion Bureau of Statistics of Canada as follows:

Prince Edward Island	Nova Scotia	New Bruns- wick	Que- bec	On- tario	Mani- toba	Sas- katch- ewan	Alber- ta	British Colum- bia	Total
	39	1	208 11	379 3	33	105	91	79	935 14
								1	8
	38 38		150	35 6		40	04	27	367 57
	46	4	260	299	332	202	462	385	1, 990
	43		76 2	237	16 1	3	11	86 1	2 472 4
		-,	55	58	3	1	14	8	139
	1	6	131	25	12	12	14	19	221
			6 2	1				2	9 3
	8	9	89	60	18	12	19	67	282
	8	4	58 77	44 32	4	ī	8	16 6	143 123
	Edward	Edward Island Scotia 	Edward Island         Nova Scotia         Bruns- wick	Edward Island         Nova Scotia         Bruns- wick         Que- bec           39         1         208           38         11         7           38         150         38           46         4         260           46         4         260           43	$\begin{array}{c c c c c c c c c c c c c c c c c c c $	Edward Island       Nova Scotia       Bruns- wick       Que- bec       On- tario       Man- toba	$\begin{array}{c c c c c c c c c c c c c c c c c c c $	$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	$ \begin{array}{c c c c c c c c c c c c c c c c c c c $

### JAPAN

Notifiable diseases—4 weeks ended May 28, 1949, and accumulated totals for the year to date.—For the 4 weeks ended May 28, 1949, and for the year to date, certain notifiable diseases have been reported in Japan as follows:

Disease	4 weeksended May 28, 1949		Total reported for the year to date	
	Cases	Deaths	Cases	Deaths
Diphtheria. Dysentery, unspecified Encephalitis, Japanese "B"	1,071 547	97 149	7,602 1,375	788 364
Gonorrhea Influenza	404		76, 550 1, 617	
Malaria Measles Meningitis, epidemic	411 35, 728 121	6 27	925 95, 657 693	21  175
Paratyphoid fever Pneumonia	130 14, 419	2	678 81,035	24
Scarlet fever Smallpox Svphilis	481 54 16, 860	12 6	2, 074 95 86, 510	35 10
Tuberculosis Typhoid fever	42, 144 356	47	184, 242 1, 837	242
Typhus fever	7 9, 165	1	83 33, 942	4

NOTE .- The above figures have been adjusted to include delayed and corrected reports.

### REPORTS OF CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER RECEIVED DURING THE CURRENT WEEK

Note.—The following reports include only items of unusual incidence or of special interest and the occurrence of these diseases, except yellow fever, in localities which had not recently reported cases. All reports of yellow fever are published currently.

A table showing the accumulated figures for these diseases for the year to date is published in the PUBLIC HEALTH REPORTS for the last Friday in each month.

### Cholera

Burma.—During the week ended April 23, 1949, 30 cases of cholera, with 2 deaths, were reported in Burma.

### Plague

British East Africa—Kenya.—During the week ended May 21, 1949, 1 case of plague was reported in the Fort Hall District, Kenya, British East Africa.

India—Calcutta.—During the period May 29–June 18, 1949, 68 cases of plague with 6 deaths were reported in Calcutta, India.

### Smallpox

Arabia—Aden.—Information dated June 28, 1949, states that on June 27, three cases of smallpox were landed at the port of Aden, Arabia, from a ship that left Batavia June 15 for Holland. The next port of call was said to be Suez.

Australia—Fremantle.—Information dated June 16, 1949, states that the steamship "Mooltan" arrived at Fremantle, Australia, on May 27 from United Kingdom, via Bombay and Colombo, with 1 case of modified smallpox on board. The patient and all contacts were quarantined.

Java-Batavia.-For the week ended June 18, 1949, 226 cases of smallpox were reported in Batavia, Java.

Nigeria-Lagos.-During the period June 1-18, 1949, 34 cases of smallpox with 4 deaths were reported in Lagos, Nigeria.

Spain—Canary Islands.—During the period May 1-21, 1949, 6 cases of smallpox were reported in the Canary Islands.

### **Typhus Fever**

Belgium.—Four cases of endemic typhus fever were reported in the Brussels area, Belgium, during the week ended June 18, 1949.

British East Africa—Nyasaland--Zomba.—During the week ended June 4, 1949, 4 cases of typhus fever were reported in Zomba, Nyasaland, British East Africa.

### **Yellow Fever**

No reports of yellow fever were received during the current week.

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owing to population shifts in many State's since the 1940 census, the figures for some States may not be comparable with those for prior years, especially for certain diseases. Each State health officer has been requested to include in the monthly report for his State all diseases that are required by law or regulation to be reported in the State, although some do not do so. The list of diseases required to be reported in the state, although some do not do so. The figures in the following table are the totals of the monthly morbidity reports received from State health authorities for January. February, and March 1949, and show the numbers of cases reported by the required reporting sources in the respective States. They are preliminary and are subject to correction by final reports. They may be assumed to represent the civilian population only, although in some instances a few cases in the military population may be included. The comparisons made are with similar preliminary reports; but is not the same for each State. Only a few of the common communicable diseases are notifiable in all the States. In some instances cases are reported, in some States, of diseases that are not required by law or regulation to be reported and the figures are included although manifestly incomplete. There are also variations among the States in the degree of, and checks on, the completeness of reporting of cases of the notifiable diseases; therefore comparisons as between States may not be justified for certain diseases. As compared with the deaths, incomplete case reports are obvious for such diseases as malaria, pellagra, pneumonia, and tuberculosis, while in many States other diseases, such as cancer, puerperal septicemia, rheumatic fever, and Vincent's infection, are not reportable. However, the figures are recorded as reported.

In spite of these and other deficiencies inherent in morbidity reporting, these monthly reports, which are published quarterly and annually in consolidated form, have proved of value in presenting early information regarding the reported incidence of a large group of diseases and in indicating trends by providing a comparison with similar preliminary figures for prior years. The table gives a general picture of the geographic distribution of certain diseases, as the States are arranged by geographic areas.

Leaders are used in the table to indicate that no case of the disease was reported.

Pella- Bra- gra forms	335 495 300 300 301
Oph- thal- mia ne- onato- rum	42
Mumps	$\begin{array}{c} 1,552\\ 569\\ 6,723\\ 1,191\\ 5,438\\ 5,438\end{array}$
Menin- gitis, menin- cal*	400808
Mea- sles*	5, 713 5, 713 4, 499 4, 861 4, 876 5, 304
Ma- laria 2	
Influ- enza	37 30 150
Hook worm disease	
Ger- man sles	$1,240 \\ 1,240 \\ 9 \\ 411 \\ 411 \\ 1,240 \\ 9 \\ 1,240 \\ $
En- cepha- litis, infec- tious	co 14 co
Dys- en- tery, unde- fined	140 0
Dys- en- tery, bacil- lary	5 QI
Dys- en- ame- bic	00
Diph-	100 109 9 4 4 109
Con- Juncti- vitis 1	29
Chick- enpox	1,309 543 1,047 11,959 627 627
An- thrax	ŝ
Division and State	Maine NEW ENGLAND Maine NEW ENGLAND Verrom Hampshire Masachusetis Masachusetis Connecticut

# Consolidated Monthly State Morbidity Reports for January, February, and March, 1949

See footnotes on page 932.

	Pneu- monia, all forms	4, 134 1, 156 1, 462	637 294 1, 812 693 6158	171 334 374 203 203 203 203 203 203	6 540 1, 501 2, 412 234 225
	Pella- gra				4
ned	Oph- thal- mia ne- onato- rum	1-40	105 25 1		5 <del>6</del> 33
Contin	Mumps	<b>8</b> 5, 139 3, 831 4, 712	3, 791 540 3, 383 3, 158 5, 390	1, 638 435 1, 380 1, 156	52 353 1, 384 1, 331 1, 273 1, 273
)676	Menin- gitis, menin- gococ- cal*	67 288 97	37 14 38 38	1015448 31021	8282187465
Monthly State Morbidity Reports for January, February, and March, 1949–Continued	Mea- sles*	21, 344 10, 027 21, 407	2, 221 1, 221 972 13, 483	1, 038 5, 380 5, 380 189 189 10, 392	13, 198 13, 198 14, 2, 6, 13, 198 1, 263 1,
d Ma	Ma- laria 2	8	5 4 1	6 10 2	6 6 88
ry, an	Influ- enza	<b>8</b> 49 53 46	31 226 196 77	7 91 107 83	1 30 4,750 870 7,971 7,971 116
ebruai	Hook worm disease	<sup>8</sup> 21	16		1, 213 1, 257
ry, F	Ger- man mea- sles	<b>3 4, 526</b> 3, 208	681 167 2,061 1,150 1,738	13 2 414	425 281 283
Januc	En- cepha- litis, infec- tious	814	6178	1	0 0 1401
s for .	Dys- en- tery, unde- fined		1	2414	88 88 89 80
eports	Dys- en- tery, bacil- lary	66 1	33 33 33	1	<b>5</b> 334 <b>0</b>
lity R	Dys- en- tery, ame- bic	162 20 5	2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	<b>H1</b>	28°512 & 1 + 4
Iorbid	Diph- theria*	73 30 132	8888% **	800°470°5	5844888888
ate N	Con- Juncti- vitis 1	4	245 245	84 22 3	00 <b>1</b>
hly Si	Chick- enpox	14, 762 24, 513 17, 273	10, 091 1, 766 9, 227 11, 479 12, 552	1, 700 1, 663 1, 659 239 239 239 239 239 239	1, 929 2, 586 2, 586 1, 028 1, 120 1, 120
Mont	An- thrax	9 4		1	
Consolidated <b>N</b>	Division and State	MIDDLE ATLANTIC New York New Jersey Pennsylvania	EAST NORTH CENTRAL Ohlo	WEST NORTH CENTRAL Minnesota. Iowa. North Dakota. South Dakota. Nobraska.	BOUTH ATLANTIC Delaware. Delaware. District of Columbia. District of Columbia. West Virginia. North Carolina. Goorgia. Florida.

July 22, 1949

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4, 659 3, 225 6, 493 892	6, 739 694 2, 936 34, 909	649 918 918 918 918 925 1,729 1,729 1,729 231 231	5, 652 6, 456 20, 288	280, 920 196, 697 196, 697	38 3, 900 6
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230 1,015 1,384	3, 014 109 911	114 127 191 1,933 1,933 31 31	1, 097 500	27, 169 128, 825 213, 750	68 1, 910
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2 <sup>4</sup> 21	68 53 128 128	1522	588 22 58 28	1, 151 802 682	21
69 101 69	50 45 276 276	120 15 15 15 15 10 11 10 11 10 11 10 10 10 10 10 10 10	13 13 139	2, 191 2, 724 3, 655	12
8		30 19 3	88	712 505 505	ŝ
1, 001 1, 089 1, 104	874 359 839 13, 637	999 518 253 399 253 399 1,047 1,047 1,047 1,047	$^{4}_{1, 242}$ 20, 239	193, 168 138, 372 133, 938	73 747 88
	1			18 19 14	
EAST SOUTH CENTRAL Kentucky. Tennesse. Alabama. Mississippi.	WEST SOUTH CENTRAL Arkansas Outahoma Toxas	Montana Mountain Idaho Wyoming Olorado New Maxico Aritona Nevada	Washington PACIFIC Oregon California	Total	Alaska Hawaii Territory Panama Canal Zone •

See footnotes on page 932.

July 22, 1949

929

Continued
1949—
l March,
, and
February
January,
for
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morbidity
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Consolidated

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	Whoop- ing Cough*	101 105 872 81 81	1, 644 631 892	681 265 446 470	42 55 55 14 4 5 1 2 6 5 6 5 6 4 7 6 7 9 7 9 7 9 7 9 7 9 7 9 7 9 7 9 7 9	14 122 2357 3331 3331 3331 91 91
	Vin- cent's infec- tion	288		11 45	71 1 17	2933
ued	Undu- lant fever*	33 3 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7 7	42 11 19	35 109 70 70	88 17 31 31 31	32213 <sup>9</sup> 17 10
Continued	Ty- phus fever, en- demic		1	3		22 <sup>2</sup> 1
	Para- ty- phoid fever	1	13 25 3 13 3	1 13 8 1	1 n	1 1 1 1 2 3 6 6 3 3 4 3 5 6 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1
, 1949	Ty- phoid fever*	4	18 42	<sup>61119</sup>	3311	126 12 12 12 12 12 12 12 12 12 12 12 12 12
March	Tula- remia		3	18 18	24 6	2 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3
and N	Tuber- culosis, respir- atory	134 577 366	3, 672	653 1, 610	45	92 468 770 544 904 532 951
February, and March,	Tuber- culosis, all forms*	150 46 611 131 380	3, 926 834 1, 293	703 1, 715 1, 224 1, 224	200 201 201 201 201 201 201 201 201 201	92 828 546 546 950 122 540 551
, Febr	Trich- inosis	∞- <u></u> ⊒-∞	60 9 4		27	
January,	Tra- choma	en la construction de la constru		1 2	97 97 3	ος. 
for Ja	Teta- nus	5	1	01 4 00	*	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1
orts f	Small- pox*			1 5	4	
ty rep	Septic sore throat	20 <sup>2</sup> 4 4 28 201 5 4 4 5 23	(1 <sup>2</sup> ) 67	13 551 169	64 3 9 9 ( <sup>13</sup> ) 5	27 27 605 33 13 2, 221 49
orbidi	Scarlet fever*	284 132 132 3,541 593	113, 128 1, 884 3, 150	4, 417 928 2, 427 4, 675 948	912 464 478 138 138 138 1415 415	80 176 176 80 2339 2339 2339 2339 2339 2339 2339 233
State morbidity reports	Rocky Moun- tain spotted fever		6		5	4
	Rheu- matic fever	21	279	45 6 205 205	18 22 18 11	37 4 73 22
monthly	Rabies in man					
lated	Polio- myeli- tis*	5	17 38 11	80110 8100 80	552453 54234 54234 54234 5423 5423 5423	5 8 5 3 3 8 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9
Consolidated	Division and State	NEW ENGLAND Maine	New York	Ohio. Indiana. Milions. Wisconsin. Wisconsin.	Minnesota Iowa Missouri North Dakota South Dakota Nebraska Kansas	BOUTH ATLANTIC Delaware. Maryland Columbia. District of Columbia. Virginia. West Virginia. North Carolina. Bouth Carolina. Georgia.

July 22, 1949

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	45	32	10 17 33	43	294 8 669 605	
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	812	83.25			179 184 548	3
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_	581 601	494 713 387	154 14 7 355 7 37 7 37	, 2, 114	17, 751 16,809 16,809	119 165
	1, 177 1, 177 624	502 746 395 2, 042	154 033 114 7 323 7 4 7 4 7 4 7 4 7 4 7 4 7 4 7 4 7 4 7 4	728 216 2, 294	29, 423 30, 638 29, 726	128 182 10 14
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	4.00	100		7	74 76 68	4
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	137	284 10 1,014	156 156 98 347 31 31 31 31	73 92 214	6, 917 5, 581 3, 173	60 3
	740 203 203 60	69 76 178 471	222 98 183 183 183 150 150 150	736 215 1, 388	36, 609 30, 843 44, 899	84
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		-			33 T	
	984II	112 4 101	15 15 15 17 17	49 330	975 8 410 416	11
EAST SOUTH CENTRAL	Kentucky Tennessee Alabama	WEST SOUTH CENTRAL Arkansas Louistana. Oklahoma Texas	Montana. Idaho Uyoming Colorado New Mexico. Nevada. Nevada.	Washington Oregon California	Total First quarter 1948 Median 1944–48	Alaska. Hawali Territory. Panama Canal Zone <sup>9</sup>

See footnotes on page 932.

July 22, 1949

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•Diseases marked with an asterisk (\*) are reportable by law or regulation in all the States, including the District of Columbia. Typhoid fever is reportable in all the States; partyphoid fever in all except 6 States. Syphilis is reportable in all the States and the District of Columbia but is not included in the table. Some States have increased and some have reduced the list of reportable diseases since the latest published compilation reportable diseases (Pub. Health Rep. 59:317-340, Mar. 10, 1944. Reprint No. 2544).

Includes cases of kerato- and suppurative conjunctivitis and of pink eye.
 In a few States practically all cases contracted outside continental United States.

- \* Exclusive of artificially induced malaria.
  \* Lobar pneumonia only.
- Exclusive of cases acquired by blood transfusion.
   Includes nonresident cases.
   Figures corrected by later reports.
   Includes the etites of Colon and Fanama.
- <sup>10</sup> In the Canal Zone only.

  - <sup>11</sup> Includes septic sore throat.
- 12 Included in scarlet fever.

<sup>11</sup> Incluides eases reported as salmonella infection. <sup>12</sup> Incluides eases reported as salmonella infection. <sup>13</sup> The following list includes certable in or reported by only a few States; last year's figures distribution, and those reported by not a sease were reported last year, or the disease was not included in last year, supplished tabulation). <sup>13</sup> Actinomycesis: Rhode Island 1, Michigan 1, Minnesota 1, Georgia 1, Nevada 1 (3). <sup>14</sup> Actinomycesis: Rhode Island 1, Michigan 1, Minnesota 1, Georgia 1, Nevada 1 (3). <sup>14</sup> Actinomycesis: Rhode Island 1, Michigan 1, Minnesota 1, Georgia 1, Nevada 1 (3). <sup>14</sup> Cancer: Pennsylvalia 1,6%, North Dakota 38 (142), Kansas 732 (856), South Carolina 226 <sup>14</sup>(43), deorgia 68 (62), Florida 537 (460), Kantucky 12, Tennesse 973 (60), Alabama <sup>14</sup>(43), deorgia 68 (62), Florida 539 (533), Montana 294 (88), Idaho 180 (217), Wyoming 173. Colorado 737, New Mexico 146 (165), Utah 48 (98) includes nonresidents. <sup>15</sup> Colorado 16.

Dengue: Texas 14.

Dermakur, Forskir, Ampshire 4 (4), Missouri 7 (16), Kentucky 1 mycotic dermatitis. Dermakur, New Hampshire 4 (4), Missouri 7 (16), Pennsylvania 21 (16), (Inuides gastro-entertis) Obio 119 (44) (Inudess entertis), Indiana 1 (8), Illinois 33 (16), Michigan 63 (33), Minnesota 5, North Dakota 2, Maryland 10 (7), South Carolina 2,759 (2,772), Florida 37 (82), Oklahoma 1, New Mexico 26 (9), Utahi 1, Washington 3, California 14 (63), Alaska 12. Dis Messohusetts 2,264, Pennsylvania 1,068, Illinois 2,369 (2,122), Dites, Michigan 1,383 (128), Arkanasa 132 (102) all animal bites. Encephalitis (other form): Rhode Island 11, Ohio 1, Michigan 12, Maryland 1, Ken-

tucky 2, Idaho 4, Colorado 4.

Erystpelas: Vermont 1, Massachusetts 2, Connecticut 8, Pennsylvania 12, Ohio 7, Indiana 7, Illinois 54, Michigan 38, Wisconsin 20, Missouri 1, North Dakota 1, Nebraska 2, Kanasa 2, Maryabad 1, Florida 9, Tennessee 8, Arkansas 3, Louisiana 4, Montana 2, Idabo 1, Colorado 12, New Merico 1, Utah 2 includes nonresidents, Washington 11, Oregon 18, Alaska 1, Hawaii Territory 2,

Favus: Kentucky 1 (5)

Filariasis: Nevada 1.

Food poisoning: Connecticut 3, New York 121, Ohio 1, Indiana 1, Illinois 49 (6) includes cases reported as food infection, Minnesota 327 (43) includes cases reported as food intection. Louisiana 1 (2), Okinoma 8, Iadano 2 (6), Colorado 2 (300), New Mactio 3, Nevada 2 includes cases reported as food infection, Oregon 3 (1), California 143 (23). Granuloma inguinale: Nebraska 1, Florida 280 (88), Kentucky 2 (4), Tennessee 14 (13), Impetigo contagtos: New York 23, Ohio 101 (9), Indiana 5 (15), Illinois 7 (6), Michigan 22 (34), North Dakota 9 (41), Kentucky 44 (4), Montana 10 (3), Idaho 23 (35), Wyo-ming 2 (11), Colorado 26 (49), Nevada 42 (60), Washington 243 (36), Alaska 4, Hawali Territory 19.

Jaundice (fioluding hepatitis and Well's disease): Maine 8 (1), New Hampshire 1 (2), Rhode Island 17, Connecticut 2, New York 110 (61), Pennsylvania 104 (15), Obio 1, Illinois 6 (3), Michigan 10 (5), Maryland 6 (1), South Carolina 2, Kentucky 6 (3), Tennessee 2 (7), Maryland 6 (1), John (20), Maryland 5 (1), South Carolina 2, Kentucky 6 (3), Tennessee 2 (7), Maryland 6 (1), John (3), Territory 14, (5), Panama Canal Zona (1), Oregon 6 (3), California 148 (30), Alaska 1, Hawali Territory 14, (5), Panama Canal Zona (1), Oneon 7 (8).
Leprosy: New York 1 (3), Oho 1, Mississippi 1, Texas 3 (1), California 3 (6), Hawali Territory 14, (5), Panama Canal Zona (1).
Leprosy: New York 1 (3), Oho 1, Mississippi 1, Texas 3 (1), California 3 (6), Hawali Territory 14, Territory 8 (6), Panama Canal Zona (1).
Leprosy: New York 1 (3), Oho 1, Mississippi 1, Texas 3 (1), California 3 (6), Hawali Territory 14, Territory 8 (6), Panama Canal Zona (1).
Leprosy: New York 1 (3), Oho 1, Mississippi 1, Texas 3 (1), California 3 (6), Hawali Territory 8 (5), Maryland 2, Dontecticut 1, Jymphogranuloma undefined, Nebraska 1, Trunphogranuloma venteum: Connecticut 2, 8(4), Ohio 2, Michigan 37 (38), Minnesota 1, Ternessee 18 (10), Mirzyland 2 (4), South Caron 2, Ymphogranuloma undefined, Alaska 1.
Mononucleois: Vermont 1, Connecticut 2, 8(4), Ohio 2, Michigan 2, Alabama 1, Michigan 2 (4), Maryland 2 (4), South Carolina 1, Maryland 1, Missispi 12 (2), Arkanses 1 (1), Lusisina 1, Nevada 2.
Rabies 1 (1), Louisiana 1, Nevada 2.
Rabies 1 (1), California 1, Maryland 1, Mississippi 2 (2), Arkanses 1 (1), Culsiana 1, Nevada 2.
Rabies 1 (1), Culsiana 1, Ohioa 28 (16), Nitana 28 (16), Yarkanses 1 (1), Culsiana 1, Nevada 2.
Rabies 1 (1), Culsiana 1, Nevada 2.
Rabies 1 (1), Culsiana 1, Maryland 1, Mississippi 2 (2), Arkanses 1 (1), Culsiana 1, Nevada 2.
Rabies 1 (1), Culsiana 1, Nevada 2.
Rabies 1 (1), Mississippi 2 (2), Arkanses 1 (1), Culsian

Rvip, "Carrier and "Construction" of the scale 1 (2).
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Ringworn disease (including ringworm of the scale): Connecticut 56 (61), Ohio 15 (24).
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Ringworn disease (including ringworm of the scale): Connecticut 56 (61), Ohio 15 (24).
Ringworn 35, Montana 1 (22), Virthia 524, Georgia 93, Kentucky 431 (8), Atanasa 2, Oklahota 23 (3), Washning 23 (24), Newada 3 (3), Washning 10, 23 (24), Newada 3 (3), Washning 10, 25 (21), Newada 3 (3), Washning 10, 28 (21), Newada 3 (3), Washning 10, 29 (24), Maryland 2, Kentucky 287 (30), Montana 29 (20), Maryland 3, Kentucky 287 (30), Montana 29 (20), Idaho 64 (40), Wyontang 6 (2), Newada 10 (13).
Sellosis New York (City) 15 (7).
Sellosis Arkansa 1, Idaho 3, New Mexico 1, Yaw: Panama Canal Zone 6.