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AGE INCIDENCE OF THE COMMON COMMUNICABLE DIS-EASES OF CHILDREN¹

A STUDY OF CASE RATES AMONG ALL CHILDREN AND AMONG CHILDREN NOT PREVIOUSLY ATTACKED AND OF DEATH RATES AND THE ESTIMATED CASE FATALITY

By SELWYN D. COLLINS, Associate Statistician, United States Public Health Service

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Introduction

A considerable mass of data on the chronological variation in the occurrence of the common communicable diseases as indicated by reports to the State boards of health has now been accumulated, and for some States and cities the cases are available by age and sex. But these morbidity reports to the State boards of health are deficient in that (a) they fail to indicate the total or complete incidence of the disease (29). Continued effort may secure fairly complete reports of recognized attacks of a severe condition like diphtheria, but by no amount of effort can we expect to secure even approximately complete reports of chicken pox and mumps. (b) The morbidity reports may fail to indicate the relative prevalence at different ages. It is usually assumed that the cases which are reported are distributed as the total are distributed so far as age is concerned. For mild conditions, such as chicken pox or even measles, this assumption may be incorrect, as it may be that a certain number of such cases discovered in school inspections would, in children below school age, not have been seen by a physician and would not therefore have been reported to the health department.

¹ From the Office of Statistical Investigations in cooperation with Field Investigations in Child Hygiene, United States Public Health Service.

However, these diseases of childhood are of great importance. Measles and whooping cough cases occur in such large numbers that, despite the small fatality, they together cause as many deaths as diphtheria. Even the milder of the children's diseases cause much absence from school and in sizable epidemics hamper the progress of the class.

The present study is an attempt to estimate from histories of past attacks of a disease, as given by persons of different ages, and from death rates, (a) the absolute incidence, (b) the case fatality, and (c) the incidence in successive ages among all children and among children not previously attacked.

If an attack of whooping cough gives a lasting immunity so that second attacks are rare and may be neglected, the proportion of 15-year old persons giving a history of whooping cough represents an accumulation of percentages having had the disease at each year of age under 15.

Of course, if we attempt to derive from histories of past attacks what, on the average, this accumulated percentage would be for different ages, there is a source of error in that attacks may be forgotten. As age increases, the number of attacks forgotten no doubt increases; and since the incidence of these childhood diseases is rather small among older children, the relative error may be considerable. These facts should be remembered in considering the rates derived from the histories for older children.

Of course, it is realized that the age incidence of these diseases varies under different conditions. For example, a larger proportion of the cases occurs in the very early years of life among urban children than among rural children (9), (11). Again, in isolated places where an epidemic has not occurred for a long time, measles has attacked adults freely (1); and in instances where it was introduced for the first time into a population not previously exposed to measles, as in the Faroe Islands in 1846 and in the Fiji Islands in 1875, adults were attacked as often as children (2), (5), (26). It should be emphasized that what is considered in the present study is the mean incidence for a considerable period of years under conditions such as exist in the more or less representative urban and rural communities where the data were collected.

Basic Data for the Study

A previous report by the Public Health Service considered the percentage of school children of different ages who had at some time in their lives suffered an attack of the communicable diseases of childhood, together with such approximations of the rates at specific ages as seemed warranted from the data then available on the history of these diseases (4). The data considered in that report were for children from 5 to 19 years of age. Similar data have since been collected by the Public Health Service for (a) a group of families, including persons of all ages in Hagerstown, Md., and (b) a group of college students attending various universities throughout the United States.

In the Hagerstown study the data on the history of infectious diseases were secured at the time of the initial visit to the household in connection with a study of morbidity which involved repeated visits (28). The informant was usually the wife or mother or some other adult member of the household. Data collected in this way are probably fairly accurate for the children of the household, particularly when the mother was the informant, but less accurate for adults about whose childhood diseases the wife would not have complete information.

A group of college students who had been making semimonthly reports of attacks of colds and other minor respiratory illnesses to the Public Health Service were requested to give information as to whether they had at any time in their lives been attacked by each of a list of diseases, the list including the children's diseases considered in the present study and one or two other infectious diseases (32). In this instance, therefore, the person is answering only for his own disease history, and the data should be more accurate than disease histories for adults collected by canvasses of households. Even here the person may have forgotten mild attacks of these more or less mild diseases or may have had an attack in early childhood but never have been told of its occurrence.

In addition to these two studies, similar data have been made available from certain other communities. Data collected in Baltimore by Dr. James A. Doull, of the department of epidemiology of the Johns Hopkins University School of Hygiene, in cooperation with the Baltimore City Health Department, have been published for certain age groups (6), (10). Through the courtesy of Doctor Doull and Surg. W. H. Frost, these data in more detailed age groups have been made available to us. They are particularly important for the preschool ages. The method of collection was similar to that used in Hagerstown, and the data should be accurate for the early ages.

Through the courtesy of the United States Children's Bureau, unpublished data on the history of certain infectious diseases for children from birth to 8 years of age, collected in connection with the study of the physical status of preschool children in Gary, Ind., have been made available to us (24).

Published data on the history of infectious diseases are available for the children of London, Canada, from birth to 20 years of age (13). By far the largest mass of data of this kind is for the school children in the 15 localities studied in the previous report by the Public Health Service (4), the number of children considered for a single age running as high as 4,000. There is, of course, considerable variation in the different localities, due in part to the time at which the data were collected in relation to the chronological cycles in the incidence of the diseases in the locality, but the results for none of these localities seem to be sufficiently different from the average to warrant exclusion from the general mass of data represented by the total. Moreover, it seemed important to average as many localities as possible in order to eliminate the effect of the cycles in the incidence of these diseases. The desired percentage of children of different ages who have had an attack of the disease is one that is true, on the average, and is as free as possible from the effect of the time elapsing since an epidemic has prevailed in the community.

In the present study it was proposed to add to the data for school children similar data for the preschool ages and for adults. In doing this, data for the preschool ages were selected from such of the other sources as showed curves which agreed reasonably well with the early school ages. Similarly, data for adults were used for such localities as agreed with the later ages of the school data. In adopting this method it is not assumed that the data for the school children are more accurate than other data for those ages: but, since the number of children in the school group is so much larger than the number considered in any other group, percentages based on the combination of all the data available for the school ages will be very near to those based on the school group only, unless the different localities are given arbitrary weights. Such being the case it seemed reasonable to extend the combined data to the preschool and adult ages by adding only such data as agreed fairly well with the school group. Table 1 shows for the different localities used the percentage of children of each age who have had an attack of measles, as well as a similar percentage for all the localities com-Tables 2, 3, 4, 5, and 6 show similar data for whooping bined. cough, mumps, chicken pox, scarlet fever, and diphtheria, and Table 7 shows for three cities the percentage of children whose Schick test was negative.

TABLE 1.—Measles history among white persons

282428222222 Gary, Ind. ------------..... Total number of persons considered ***** Balti-more, Md. College stu-dents in 11 uni-versi-ties in United States 2 202 507 376 376 376 376 Hagers-town, Md. dren in 15 lo-calities in United States School chil-~%5%%%2%% Gary, Ind. Number of persons who have had an attack -----..... Balti-more, Md. College stu-dents in 11 uni-versi-ties in United States 22 33 66 F 180 Hagers-town, Md. ~88888<u>68</u>8 dren in 15 lo-calities in United States School chil-..... 17.0220.71 37.05 3 Percentage of persons who have had an attack Gary, Ind. (24) -----..... Balti-more, Md. (6) College stu-dents in 11 uni-versi-ties in United States (32) ŝ fagers-town, Md. (28) School chil-dren in 15 lo-calities in United States (4) 20-24 25-29 ----į ł Under 1. Age : (years) 30-34. 35-39. Total num-ber of persons consid-ered 55 Num-ber of persons who have have attack of measles All localities Percentage of persons who have had an attack of measles Actual Com-puted 10.0 11.0 11.0 11.0 15.0 15.0 17.5 22.5 22.5 37.5 37.5 5.4 6.3 7.1 8.0 Age (years) 1.5 į 5. 2.0-. . 2

FERCENTAGE OF INDIVIDUALS OF EACH AGE WHO HAVE SUFFERED AN ATTACE AT SOME TIME IN THEIR

LIVES

TABLE 2.—Whooping cough history among white persons

PERCENTAGE OF INDIVIDUALS OF EACH AGE WHO HAVE SUFFERED AN ATTACK AT SOME TIME IN THEIR LIVES

| nsidered | London, Canada | · . · . | 28782828 |
|----------------------|--|------------------------|---|
| persons co | College students in 11 universi- ties in | United States | 288888 11755 |
| mber of | Hagers- town, Md. | | \$ |
| Total nu | School School in 15 locali- ties in | United States | -14%%4%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%% |
| ave had | London, Canada | | ۰\$\$\$\$ 25 25 25 25 25 25 25 25 25 25 25 25 25 |
| ns who h ttack | College students in 11 universi- ties in | United States | 88333 1380 1380 1380 1380 1380 1380 1380 |
| of perso an at | Hagers- town- Md. | | 58888888888888888888888888888888888888 |
| Number | School children in 15 locali- ties in United States | | -9490444- |
| ave had | London, Canada (13) | | କ୍ ୟସ୍ଟ୍ ଝ୍ ଝ୍ ପ୍ ନ୍ତିଷ୍ଟି କଙ୍କରନ୍ମ ନ୍ତି |
| ons who h ttack | College students in 11 universi- ties in United States (32) | | 0.00817.01 33313333 |
| age of pers an at | Hagers- town, Md. (28) | | 4~442888882652656888856666848 6~8084808489797978888856868588 70808480848978748978888855888574879 |
| Percents | School School children in 16 locali- ties in | States (4) | ቒቘቔቘቘ ዸ፟ቘቔቔቘቔቔ ዸ፟ቘቔቔቔቔቔቔ ዸ፟ቘቔቔቔቔቔቔ ዸ፟ቘቔቔቔቔቔ ዸ፟ቘቔቔቔቔቔ ዸ፟ቘቔቔቔቔ ዸ፟፟ቘቔቔቔቔቔ ዸ፟፟ቘቔቔቔቔቔ ዸ፟፟ቘቔቔቔቔቔ ዸ፟፟ቘቔቔቔቔቔቔ ዸ፟፟ቘቔቔቔቔቔቔ ዸ፟፟ቘቔቔቔቔቔቔቔቔቔቔ |
| | Age ¹ (years) | \$ • | Under 1 1-1-1-1-2-2-2-2-2-2-2-2-2-2-2-2-2-2-2- |
| | Total number of per- | sidered | 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1 |
| | Number of per- sons who have had an at- tack of | whoop- ing cough | 924200000000000000000000000000000000000 |
| l localities | te of per- tho have attack of ng cough | Actual | ~JII%8%8%8%8%8%7%7%2%2%2%2%2%2%2%2%2%2%2%2%2 |
| ΠV | Percentag sons w had an whoopi | Com- puted | 497337333538354478555555577777 997335353555555555555555555 |
| | Age (years) | | 05 115 145 544 544 544 544 544 544 1100 1100 |

¹ Ages for school children in the 16 localities are nearest birthday; all other groups are age last birthday.

.

TABLE 3.—Mumps history among white persons

London, Canada Total number of persons considered \$252225588 Balti-more, Md. College students in 11 univer-sities in United States 2, 460 2, -----..... -----School children in 15 lo-calities in United States \$ London, Canada Number of persons who have had an attack 8883333**68886666667** 85885×20 Balti-more, Md. College students in 11 univer-sities in United States -----138 380 1,583 178 178 81 82 82 -----...... ----------School children in 15 lo-calities United -----3 London, Canada 55.50 55 Percentage of persons who have had an attack (13) 1887-1488 -----..... -----Balti-Md. (6) College students in 11 univer-sities in United States (32) 57.4 62.8 64.9 68.7 68.9 68.9 68.9 -----School children in 15 lo-calities in United States (4) -----..... -----....... Age (years)¹ Under 1. 20-24 35-39. 22-23 6-35 Total of per-sons con-sidered -inini Num-ber of persons who have had an attack of mumps Percentage of per-sons who have had an attack of mumps All localities Actual 41232228823**4**28228232258418 842 Com-7.7.8848.89555555564888954899 2.5 -----..... -----Age (years) 18.3. 19.5. 22.5. 32.5. 12.1. 13.1. 15.1. 16.2 8.1. 9.1. 11.1 s, 5 4.0.4 ð

FERCENTAGE OF INDIVIDUALS OF EACH AGE WHO HAVE.SUFFERED AN ATTACE AT SOME TIME IN THEIR LIVES

¹ Ages for school children in the 15 localities are nearest birthday; all other groups are age last birthday.

| | mber of consid- | London, Canada | | 73 144 23 24 25 25 25 25 25 25 25 25 25 25 25 25 25 |
|-----|---------------------------------|--|--|--|
| | Total nu persons ered | School School children in 15 localities | United States | 121 131 131 131 131 131 131 131 |
| | of per- rho have attack | London, Canada | | 88888883388833888888888888888888888888 |
| | Number sons w had an | School children in 15 localities | United States | 888 888 888 888 888 888 888 888 888 88 |
| | ge of per- ho have attack | London, Canada | (81) | -2813825255555555555555555555555555555555 |
| | Percentag sons w had an | School children in 15 localities in | United States (4) | 28888444444484288888888888888888888888 |
| | | Age 1 (years) | • | Under 1 1 2 3 5 5 5 5 6 0 10 11 13 13 13 13 13 13 13 13 13 13 13 13 |
| | | Total number of per- | sons con- sidered | નબન્નન્વસ્થ્યન્ રાષ્ટ્રકારકારકાર રાષ્ટ્રકારકારકારકારકારકારકારકારકારકારકારકારકાર |
| | | Number of per- sons who have had an attack of chicken por | | 1173 2524 2524 2524 2524 2524 2524 2524 252 |
| | | ze of per- ho have attack of pox | Actual | 0.8.1.8.2%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%% |
| | localities | Percenta, sons w had an chicker | Com- puted | 89888888888888888888888888888888888888 |
| IIV | Age (years) | | 265 265 245 245 245 245 241 101 101 101 101 101 101 102 103 103 103 103 103 103 103 103 103 103 | |

PERCENTAGE OF INDIVIDUALS OF EACH AGE WHO HAVE SUFFERED AN ATTACK AT SOME TIME IN THEIR LIVES

TABLE 4.—Chicken pox history among white persons

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¹ Ages for school children in the 15 localities are nearest birthday; for other groups ages are last birthday.

770

TABLE 5.—Scarlet fever history among white persons

PERCENTAGE OF INDIVIDUALS OF EACH AGE WHO HAVE SUFFERED

AN ATTACK AT SOME TIME IN THEIR LIVES

282488282 Gary, Ind. Total number of persons considered ----------Balti-more, Md. Col-lege ³ stu-dents in 3 univer-sities in States -----..... -----Hagers-town, Md. School chil-dren fn 15 locali-ties in United States -----ຕໍ່ຕໍ່ຕໍ ---------...... -----..... °∞82248 Gary, Ind. Number of persons who have had an attack 8188288338199233 Balti-more, Md. Col-lege ² stu-dents in 3 in 3 univer-sities in United States - C G G B G -----..... ----------Hagers-town, Md. 34234425363954295295 School chil-dren locali-locali-ties in United States ----------5,55 011 2,55 011 2,55 011 Percentage of persons who have had an attack Legistics, ----------..... Balti-Md. (6) **മാമാ**ത്ത് 0 Col-lege ³ stu-dents in 3 univer-sities in United States (32) 888585 -----..... -----...... 8453558 Hagers-town, Md. (28) 7.7.0110.925888 7.7.0110.9110.5888 7.7.0538888 7.4538 7.65387 7.6538 7.653877 7.653877 7.6538777 7.65387777777777 4158388889 0.64 d r 8.40 9.51 10.51 11.25 12.25 12 School chil-dren in 15 locali-ties in United States (4) -----..... 128222 1 ł ł ł Under 1. Age 1 (years) 20-24 30-34 35-29 ber of persons con-sidered 1195 1199 11190 1100 1100 1100 1100 1100 1100 1100 110 Total Num-ber of persons who have have have attack of scarlet fever All localities 82233582633888325888828282828288 Actual Percentage of persons who have had an attack of scarlet fever 2 -i ci ci 50 ~~~ijjjjjiiiis Com-4.4.4.4.4.4.4.0.0.0.0.1.1 ł Age (years)

¹ Ages for school children in the 15 localities are nearest birthday; all other groups are ages last birthday. ² Includes only students in the south Atlantic, east south central, and west south central divisions, mostly in Johns Hopkins, Georgetown, and Tulane Universities.

| persons |
|------------|
| white |
| among |
| history |
| Diphtheria |
| 6.—1 |
| TABLE |

PERCENTAGE OF INDIVIDUALS OF EACH AGE WHO HAVE SUFFERED AN ATTACE AT SOME THME IN TEER LIVES

| onsid- | Balti- more, Md. | | £58588888888 |
|----------------------|---|------------------------------|---|
| persons o | College students in 11 univer- | United States | 2316 2316 2316 2316 2316 |
| umber of ere | Hagers- town, Md. | | \$ 8883311131 1 732311232 \$ 8883311131 1 73231123 \$ 8883311131 1 7323 \$ 883311131 1 7323 \$ 88331131 1 733 \$ 88331 1 733 1 733 \$ 88331 1 733 |
| Total n | School chil- dren in 15 lo- calities | United States | 212 100 100 100 100 100 100 100 |
| ave had | Balti- more, Md. | | →∾★85%%%₩₩ |
| ns who he ttack | College students in 11 univer- sities in | United States | |
| er of perso an ai | Hagers- town, Md. | | 0.4000053555555585150044858848 |
| Numbe | School Chil- dren in 15 lo- calities | United States | ≈£\$\$\$\$\$\$\$\$\$\$\$ |
| o have | Balti- more, Md. | 2 | 9.,98477497 |
| ersons wh attack | College students in 11 univer- sities in | States (32) (32) | 0.6.0.1.6.4 8888388 |
| itage of p had an | Hagers- town, Md. | Ì | |
| Percei | School chil- dren in 15 lo- calities in United States (4) | | ~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~ |
| Age (years) 1 | | | Under 1 1. 2. 2. 2. 2. 3. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. |
| | Total num- ber of persons | ered | 77 1992 1992 1992 1992 1992 1992 1992 1992 1992 1995 19 |
| | Num- ber of persons who have | attack of diph- theria | 28828884 28828884 2888884 288888 288888 288888 288888 288888 288888 288888 288888 288888 288888 288888 288888 288888 28888 28888 288888 288888 288888 288888 288888 288888 288 |
| l localities | ie of per- ho have attack of ria | Actual | 9.19894999999999999999999999999999999999 |
| IA . | Percentag sons w had an diphthe | Com- puted | Q.14%446667568888899999999999999999999999999999 |
| | Age (years) | | 0.5 1.5 2.5 2.5 5.4 5.5 5.4 5.5 5.4 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0 1.0 |

¹ Ages for school children in the 15 localities are nearest birthday; all other groups are ages last birthday.

TABLE 7.--Schick test for diphtheria immunity among white persons

11 451 929 929 929 929 957 957 72 311 72 311 72 Total number of persons tested Kansas City, Mo. Syracuse, N. Y. Balti-more, Md. Number of persons who are negative Schick Kansas City, Mo. Syracuse, N. Y. -Balti-more, Md. -----Percentage of persons who are negative Schick Kansas City. Mo. (16) Syracuse, N. Y. ł -----Balti-more, Md. (10) 20 and over ----------................ Under 1 -----Age last birthday (years) 16-19. ġ Total number of persons tested $\substack{\mathbf{72}\\ \mathbf{72}\\ \mathbf{72}\\ \mathbf{72}\\ \mathbf{72}\\ \mathbf{72}\\ \mathbf{736}\\ \mathbf{935}\\ \mathbf{954}\\ \mathbf{954}\\ \mathbf{954}\\ \mathbf{954}\\ \mathbf{935}\\ \mathbf{935}\\$ Number of persons who are negative Schick Percentage of per-sons who are neg-ative Schick Actual All localities Com------............... and over -----................ ------................. Age (years) 10.5 11.5. 2.5 3.5 14.5. 5.5 6.5_ 5.0 80 ດ 5.5. <u>ة</u> 2 ଛ

PERCENTAGE OF INDIVIDUALS OF EACH AGE WHO GIVE A NEGATIVE SCHICK REACTION

In the case of the school records the data were classified according to age nearest birthday, but in all other cases the classification was according to age last birthday. In making the combination of children 7 years of age in the different groups, the midpoint, 7.0, and 7.5 years of age for the school and other groups, respectively, were averaged, each group being weighted by the total number of children in it. This will explain such ages as 7.2 in the combined data.

Inasmuch as data from various sources were combined, it is important to note the extent of agreement or disagreement in the data for the various localities.

In the case of whooping cough the percentages with a history of this disease in Hagerstown, Md., and London, Canada, agree fairly well up to 7 years of age, after which age the London (Canada) percentages are considerably below those for Hagerstown. The Hagerstown data, however, agree very closely with the school data throughout the school ages. Furthermore, the data for adults in Hagerstown agree very closely with the data for college students, both indicating that the asymptote, or upper limit, lies between 75 and 80 per cent—that is, that 75 to 80 per cent of adults have had an attack of whooping cough at some time in their lives.

The data for Baltimore, Md., and Gary, Ind., on the other hand, are not in agreement with the above data except for about the first three years of life. The asymptote as indicated by the Baltimore data would be around 50 per cent. A rather consistent falling of the curve after about 9 years of age suggests that forgotten attacks are a rather important source of error in the data from Baltimore.

In the case of measles, the data for Hagerstown, Md., Baltimore, Md., Gary, Ind., and the school group are in excellent agreement up to 7 years. After that age the Hagerstown curve is rather consistently above and the Baltimore curve rather consistently below the curve for the school group. The percentages for London, Canada, are considerably below the other localities.

There is, of course, the possibility that measles and German measles have been confused and that a history of measles may have been given when the person had had German measles but had not had an attack of measles. Although the two diseases are clinically quite distinct and would not be confused as to the diagnosis at the time of occurrence, there is considerable possibility of confusion in reporting the history of the two conditions because of the similarity of the two names. The fact that the percentages reporting a history of measles are low in London, Canada, where both diseases were included in the list of infections on the schedule, suggests that such confusion did exist, although many other factors, such as the time since the occurrence of an epidemic of measles in the locality, may have been important in producing the difference.

The questionnaire sent to the college students also included both German measles and measles, and the results have been tabulated in some detail. Of 3,377 of these college students who knew whether or not they had ever had either measles or German measles, 90 per cent had had measles, 31 per cent had had German measles, and 95 per cent had had one or both diseases, leaving 5 per cent who had had neither measles nor German measles or were doubtful about one disease but not doubtful about the other. Of the 3,377 persons, 121, or 3.6 per cent, were doubtful as to measles, but of the 121 persons, all except 26 had had an attack of German measles. Of the 3,377 persons, 778, or 23 per cent, were doubtful as to German measles, but all of the 778 except 12 had had an attack of measles. For the purpose of the percentages given above, it seemed better to consider the doubtfuls as to one of the two diseases designated as measles as a negative history of that disease when the person was not doubtful as to the other disease. Those doubtful or unknown about both measles and German measles were eliminated from all The large percentage of persons doubtful as to Gercomputations. man measles indicates the difficulty of obtaining any accurate data on that disease.

Considering measles and German measles separately and independently: Of the 3,256 persons who gave, without qualification, a positive or negative history of measles, 93 per cent reported that they had suffered an attack of measles; of the 2,599 who were not doubtful about German measles, 40 per cent reported a history of German measles. In London, Canada, 15 to 20 per cent of the older school children (16 to 20 years of age) reported a history of German measles.

Although there is considerable chance for error in the confusion of the names "measles" and "German measles," it should be noted that in the use of the data as in this study error arises only to the extent that persons who have had German measles but have not had measles report a history of measles. Persons who have had both measles and German measles presumably have some immunity to measles and the additional immunity to German measles does not change the situation as regards measles. The data for the college students indicates that the maximum error—that is, the error when every history of German measles is counted as a history of measles—would not be more than 3 to 5 per cent for adults. The error might be larger for some of the younger ages.

In this connection it should be noted that the college students are all of the ages when the curve of the history of measles has about reached its maximum and therefore all ages have been considered together in the above discussion.

For mumps, the curve for London, Canada, agrees moderately well with that for the school group. The older ages of the school group and the early ages of the college group agree very well. The data for Baltimore, Md., agree fairly well with the other groups up to about 8 years, but after that age the percentages are very much below those for the other groups, indicating an asymptote of around 30 per cent as against over 70 per cent for the college students. It is at least possible that the difference is partly due to the fact that the students themselves are the informants—the milder the disease the less accurate would be the history obtained from some other member of the family.

No data on the history of mumps were collected in Hagerstown or other cities considered in connection with other diseases in this study.

The data on chicken pox for the London, Canada, and the school groups agree unusually well throughout the school ages, indicating an asymptote of from 50 to 55 per cent. The percentages for college students are considerably above this figure, being around 70 per cent. In no other places were data collected on the history of chicken pox.

Because of the smaller number of persons who contract scarlet fever, the percentages who have a positive history of this disease are irregular except where the population considered is very large. The data for Hagerstown, Md., and Gary, Ind., agree reasonably well with the school data. Up to 15 years the Baltimore data are very similar to the Hagerstown data. The percentages for London, Canada, are very much higher than those for the other localities, as might be expected in view of the higher incidence of scarlet fever in the North. The percentages for college students are above the other groups except in the Southern States, where the agreement is considerably better.

In diphtheria, likewise, there is fair agreement except in the case of London, Canada. The percentages for Baltimore actually dropped off perceptibly after 10 years, but before that age the agreement with the other localities is fair. The data for college students agree fairly well with the data for the older ages of the school group and for the Hagerstown adults.

In connection with the diphtheria history curve it should be noted that there may be an error in such a curve due to the decline in the prevalence of the disease. In the decade 1900–1909 diphtheria was considerably more prevalent than in more recent years (31). A higher proportion of adults who were children in that decade would therefore have a history of an attack of diphtheria than would be expected if the diphtheria prevalence had been about what it was in the early years of the present decade. The incidence of scarlet fever may have declined also (22), and the curve for that disease may therefore be somewhat in error. In the instance of measles, whooping cough, mumps, and chicken pox, there is little evidence of any change in their incidence during the past generation.

Data on the percentage of children of different ages who react negatively to the Schick test, and thereby indicate a certain degree of specific antitoxic immunity to diphtheria, were not available for the groups considered for the history of infectious diseases. Data available for groups of children in Baltimore, Md., Syracuse, N. Y., and Kansas City, Mo., show a fair agreement for most of the ages. Aside from one or two percentages based on very small numbers, the agreement is particularly good up to 12 years, but after that age the percentages Schick negative are somewhat less in Syracuse than in Baltimore and Kansas City. The percentages with a negative Schick test are higher in New York City (34) than in these three smaller cities, but the data for one group of schools in New York City (with a lower proportion of children with a negative Schick reaction) agree fairly well at the different ages with the Baltimore, Syracuse, and Kansas City data. On the other hand, the percentages for small towns in New Jersey (34) are lower than the Baltimore, Syracuse, and Kansas City percentages. The latter cities are probably fairly representative of moderately large cities in the middle or northern part of the United States.

The only extensive data on the percentage of children of different ages in the United States who react negatively to the Dick test, and thereby indicate a certain degree of immunity to scarlet fever, are from Kansas City, Mo. (16); Gary, Ind. (25); and New York City (35). As in the case of the Schick test, the percentage with a negative Dick reaction is highest in New York City. Kansas City had a slightly higher percentage negative than Gary. In a group of children in Fairfax County, Va. (8), the percentage with a negative Dick reaction was much lower than in any of the cities mentioned. In view of the paucity of comparable data, it did not seem worth while to attempt to include the negative Dick test in the present study.

Tables 1 to 7 include the data on the history of infectious diseases for the different localities only for the ages that showed sufficient agreement to warrant their inclusion in the general total for all localities. Figures 1, 2, 3, and 4 show these percentages plotted on cross-section paper.

These observations are, of course, based on different children at each age, the data in every case being the result of a cross-section survey only, that is, a survey in which each child was seen but once. In other words, the children 15 years of age were not observed for a period of 15 years and the percentage who had suffered an attack of the measles recorded as they occurred or at the end of each year, but the percentages for each age are based on a different group of children. If, however, we observed a group of children from birth to 1 year of age and recorded the percentage of the survivors who had had an attack of measles, and observed the same group from 1 year to 2



FIGURE 1.—Percentage of the population of specific ages who have at some time in their lives suffered attacks of certain communicable diseases of childhood (as approximated from histories of past attacks)

years of age and recorded the percentage of the survivors who had had an attack of measles, etc., up to adult life, we would presumably get a similar curve to the one in Figure 1, which is based on different children at each age. The case fatalities in the children's diseases considered in this study, except for diphtheria and scarlet fever, are so small after the first two years of life that a correction for the fatal



FIGURE 2.—Percentage of the population of specific ages who have at some time in their lives suffered attacks of certain communicable diseases of childhood (as approximated from histories of past attacks)

cases would hardly make a perceptible change in the curve. However, in considering case fatalities (Table 16), the case incidence has been corrected to include the fatal as well as the nonfatal cases. In none of the other tables and graphs was any correction made for the fatal cases.

Method of Graduating the Original Data

Statistical operations, which will be considered in later sections of this paper, made it important to substitute a smooth curve for the more or less irregular original points shown by the small circles in Figures 1, 2, 3, and 4. Even apart from these further mathematical



FIGURE 3.—Percentage of the population of specific ages who have at some time in their lives suffered attacks of certain communicable diseases of childhood (as approximated from histories of past attacks)

processes, a smooth curve is of interest as approximating the percentages which would be obtained by the consideration of a very large number of children for each year of age.

Some mathematical function was therefore sought which would represent these data. Thinking of the data as approximating what would be secured by the continuous observation of a very large number of children from birth to adult ages, the "growth" of cases of whooping cough, for example, in such a group may have some of the characteristics of the growth of population in a constant geographical area. It, therefore, seemed reasonable to expect a logistic or growth curve such as that described by Pearl and Reed to fit these data (19), (20), (21), (18), and (23). The smooth curves in Figures 1, 2, 3, and 4 are the graphs of logistic curves fitted to the data, the equation in each case being shown on the chart. The curve which was found to best fit the data for mumps is of the autocatalytic type, such as has been used extensively by Reed and Pearl for describing population growth. The curve used for mumps is of the skew variety of this type. For all other diseases and the negative Schick test, the curve



FIGURE 4.—Percentage of persons of specific ages who are relatively immune to diphtheria as indicated by a negative Schick reaction (Baltimore, Md., Syracuse, N. Y., and Kansas City, Mo.)

giving the best fit from about 2 years of age to adult life is the catalytic type of curve, the skew variety being used in every case.

In many respects the logic or philosophy of the logistic curves, both of the catalytic and autocatalytic types, is consistent with the logic of the growth of cases of these infectious diseases in a population. For example, after an initial period, which will be considered later, both the catalytic (skew type) and autocatalytic curves rise at an increasing rate (as measured in actual units rise per year) until the point of inflection is reached. This is the point where the slope of the curve is greatest and represents the age of maximum incidence of the disease. After this age the curve rises at a declining rate (as measured in actual units per year) and approaches an asymptote, or upper limit, the asymptote in the case of the infectious diseases under consideration being the percentage of the adult population who have at some time in their lives suffered an attack of the disease under consideration. The slowing up in the rate at which the curve rises would be consistent with the slowing up in the occurrence of the infectious disease as the susceptible population is eliminated by reason of an immunity gained by an attack of the disease or otherwise. Factors such as change in the amount and intimacy of contact with cases of the disease may also be important.

If the autocatalytic curve, such as that used for mumps, were continued to birth or zero age, the Y value (percentage) would not be zero, the nature of the curve being such that it approaches zero as a limit but never absolutely reaches zero. In this respect the logic of the autocatalytic growth curve is not consistent with the logic of the data being considered, since at birth no child, except in instances so rare that it is anomalous, would have already suffered an attack of any of these diseases. The catalytic curve becomes zero at zero age if the constant "a" is equal to zero. In the curves of this type shown in Figures 1, 2, 3, and 4 such is not the case, because a good fit could not be secured for the other points if the curve became zero at birth or zero age. The curves as used were fitted with particular attention to points for 11/2 years of age and older. Below this age an inspection curve has been used to extend the curves to birth or zero age, the latter curves being indicated on Figures 1, 2, 3, and 4 by dotted lines. In no case does this latter curve affect any value above 2 years of age, and in most cases the correction is small for 1 year of age.

In the case of the percentage of children with a negative Schick test, the percentage under 1 year is, of course, higher than that for 1 year of age, because of the passive immunity which the newborn child gets from its mother, provided she has a specific immunity to diphtheria. The curve as fitted to the data is an attempt to approximate the net growth with age in the percentage of children who have acquired a sufficient degree of specific antitoxic immunity to diphtheria to give a negative Schick reaction. In so far as the percentage for children under 1 year of age is used at all, the actual observed percentage is used, although this is based on relatively few observations, particularly for the first half year of life. It will be recalled that above 12 years of age the Baltimore and Kansas City percentages were higher than those for Syracuse. In the higher ages, therefore, the curve may be somewhat unreliable.

It is possible that the autocatalytic type of curve might be used for some diseases other than mumps. There is in several instances some indication of a period during which the curve rises very slowly in the nature of the autocatalytic curve, rather than the catalytic curve which rises more rapidly from the beginning. However, the period of the slow rise is so short that it appears that the autocatalytic curve would have to be of an even more skew type than that used for mumps. This period, when the curve is rising very slowly, whether represented by the autocatalytic curve or the short inspection curve, is consistent with the idea that the new-born child gets some temporary immunity from its mother, as in diphtheria, and also with the fact of lesser contact in early infancy when the child is not playing with other children. Following the period when the curve rises slowly, both types of curve rise more rapidly to the point of inflection already mentioned as marking the age of maximum incidence of the disease.



FIGURE 5.—Percentage of the population of specific ages who have at some time in their lives suffered attacks of the communicable diseases of childhood (as approximated from histories of past attacks)

Population at Specific Ages Who Have Been and Who Have Not Been Attacked

In Figure 5 the curves in Figures 1, 2, 3, and 4 have been superimposed to facilitate comparison. It will be seen that by 10 years of age 86 per cent of the children have had an attack of measles, 73 per cent have had whooping cough, 47 per cent chicken pox, 42 per cent mumps, 9 per cent scarlet fever, and 8 per cent have had a clinical attack of diphtheria, but 47 per cent gave a negative Schick reaction indicating a certain degree of specific immunity to diphtheria. Although the data available for Dick tests were insufficient to justify including that test in this study, the few existing data indicate that the percentage of children with a negative Dick test, and therefore with a certain degree of specific immunity to scarlet fever, is probably larger than the proportion with a negative Schick test. By 15 years of age the curves for measles, whooping cough, and chicken pox have virtually reached their asymptotes. The diphtheria curve and, to a



FIGURE 6.—Relative increase with age in the percentage of the population of specific ages who have at some time in their lives suffered attacks of the communicable diseases of childhood (as approximated from histories of past attacks)

lesser extent, the scarlet fever curve are still rising somewhat, but the curve for mumps continues to rise markedly until 25 or 30 years of age.

Better to compare the relative or proportionate increases in the curves for the different diseases, the graphs have been plotted on a semilogarithmic chart in Figure 6. This chart indicates that, although the percentages of persons who have had clinical attacks of

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scarlet fever and diptheria are low, the relative or proportionate rises in the curves of these diseases are similar to those of the curves of the diseases which occur more frequently. The curve for mumps, however, appears to be somewhat different from the curves for the other diseases.

The area below each of the curves in Figure 5 for the different diseases represents persons who have suffered attacks of these diseases





and, with exceptions, have a considerable degree of specific immunity to further attacks of the disease. In the case of diphtheria and scarlet fever, we know from the Schick and Dick tests that many who have not suffered attacks of these diseases are nevertheless Schick or Dick negative, indicating a certain degree of specific immunity to attack. Second attacks of any of these diseases may occur, but they are relatively rare. However, some scarlet fever and many recovered diphtheria patients are still Schick positive and presumably relatively susceptible to the disease (14).

Similarly, the area above each curve in Figure 5 represents persons who have not suffered an attack of the disease. It may be of some interest to make a more direct comparison of the proportion of persons of different ages who have not been attacked by the different diseases. In Figure 7 the proportion of the population that has not been attacked by each disease has been plotted.

Age Incidence Among All Children

Although the age of the child at the time of the attack of the infectious disease was not tabulated in the data used for this study, we may approximate the curves of the age incidence of these diseases from the cumulative curves shown in preceding figures. The percentage of 5-year-old children who have had a disease is obviously a cumulation of the percentages who had the disease when under 1 year. 1, 2, 3, and 4 years of age. In other words, the data plotted in Figure 5 are of the nature of cumulative or integral curves. If we take the difference between the percentage who had been attacked at 5 years and the corresponding percentage who had been attacked at 4 years, we will find what percentage were attacked between 4 and 5 years of age. This difference represents the average slope of the cumulative or integral curve between 4 and 5 years of age. The slope of a curve indicates how fast the curve is rising, and that which indicates how fast the cumulative or integral curve of whooping cough, for example, is rising during a given age period, say 4 to 5 years, also indicates how fast children 4 and under 5 years of age are having whooping cough, that is, the incidence rate for this specific age. If it appeared worth while to go into greater refinement than the average rate for the age period 4 and under 5 years, we could find the slope of the curve at any point on the age scale by finding the first derivative of the equation of the cumulative or integral curve and the value of this expression for any exact age would indicate the age incidence for that age. In the present study the average rate for 1-year age intervals seemed the more appropriate measure of the specific age incidence and the difference method was therefore used in preference to the derivative method. The peak in the difference or derivative curve of course indicates the point of inflection in the cumulative or integral curve.

It should be noted that the fitting of a curve to the cumulative data, as was done in this study, obscures or neglects irregularities in the curves. While the great majority of these fluctuations obviously are the result of chance or at least of factors which we need not or do not wish to take account of in this study, certain variations from the fitted curve may have a definite significance. For example, for some of the diseases the actual value (percentage who have had the disease) for 6 years is above the smooth curve. Carried into the differences, this fact would mean that the actual difference which indicates the incidence rate at about 6 years of age should be larger than the one obtained from the fitted curves. Since children usually start to school at 6 years of age and thereby add considerably to the number of persons with whom they come in almost daily contact, the chance of having attacks of these diseases may be much increased. The values derived from the smooth curves, of course, fail to show such unusual rates. However, there appears to be no other period below adult ages where such a condition might arise. In adult life, parents who have not had attacks of these diseases or have not in some other way become immune to them (as in diphtheria and scarlet fever), are subjected to unusual contact when their children have the diseases. As the data considered in this study probably are not sufficiently accurate to determine anything about the incidence rates among adults, this second period of unusual contact need not concern us now.

The growth curves may give a good fit to the data as a whole, but in certain parts of the curve fail to fit the points as well as they should be fitted. In other words, the real curve of the growth of cases of one of these infectious diseases in a population may be a more complicated mathematical function than the logistic curve. The lack of fit may be so slight that it is not noticed in the cumulative data but shows up perceptibly only when the derivative or difference curve is considered. The small magnitude of the values considered in the derivative curve makes the errors relatively more important there.

With these limitations in mind we may proceed to consider the age incidence of these infectious diseases of childhood. Tables 8 to 13 contain data on age incidence as approximated from histories and also other computations which will be considered later in the study.

| Age interval | Of 1,000 livi | ing children | Case rate per 1,000 | Ratio of case rate among chil- dren not previonsly attacked to case rate among all children (rate among all children = 1.000) | |
|--|---|---|--|--|--|
| Period of lifetime between 2 exact ages (years) | Previously attacked (number at beginning of age inter- val who have had an attack) Case rate per 1,000 total children (number attack during an attack during age interval) | | Not previ- ously attacked (number at beginning of age inter- val who have not had an attack) | | |
| (1) | (2) | (3) | (4) | (5) | (6) |
| 0-1 | 0 61. 677 166. 759 294. 074 426. 648 550. 020 765. 162 834. 820 859. 704 874. 300 882. 328 886. 458 888. 460 889. 368 889. 368 | 61. 677 105. 082 127. 315 132. 574 123. 372 104. 584 81. 586 58. 972 39. 658 24. 884 14. 596 8. 028 4. 130 2. 002 . 908 . 382 . 151 | 1,000,000 938,323 833,241 705,925 573,352 449,930 345,393 265,839 265,839 165,150 140,293 165,150 140,293 1125,700 117,672 113,542 110,632 | 61, 677 111, 989 152, 735 187, 802 215, 177 232, 419 235, 210 235, 210, 210, 210, 210, 210, 210, 210, 210 | 1.000 1.066 1.200 1.417 1.744 2.222 895 3.791 4.882 6.054 7.128 8.493 8.807 8.966 9.033 9.073 |

TABLE 8.—Measles white persons in various localities in the united states

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TABLE 9.—Whooping cough

| Age interval | Of 1,000 livi | ng children | Case rate per 1,000 | | | |
|---|--|--|--|---|--|--|
| Period of lifetime between 2 exact age3 (years) | Previously attacked (number at beginning of age inter- val who have had an attack) | Case rate per 1,000 total children (number having an attack during age interval) | Not previ- ously attacked (number at beginning of age inter- val who have not had an attack) | children not previ- ously attacked (number attacked during age interval among 1,000 who had not had an attack at beginning of age interval) | Ratio of case rate among chili dren not previously attacked to case rate among all children (rate among all children = 1.000) | |
| (1) | (2) | (3) | (4) | (5) | (6) | |
| 0-1. 1-2. 2-3. 3-4. 4-5. 5-6. 6-7. 7-8. 8-9. 9-10. 10-11. 11-12. 12-13. 13-14. 14-15. 15-16. 16-17. | 0 58,500 132,378 220,359 321,152 422,738 515,469 593,254 653,730 697,535 727,219 746,068 757,318 763,632 766,974 768,637 769,418 | $\begin{array}{c} 58, 500\\ 73, 878\\ 87, 981\\ 100, 793\\ 101, 586\\ 92, 731\\ 77, 785\\ 60, 476\\ 43, 805\\ 29, 684\\ 18, 849\\ 11, 250\\ 6, 314\\ 3, 342\\ 1, 663\\ 3, 342\\ 1, 663\\ ., 781\\ ., 336\end{array}$ | 1,000.000 941.500 867.622 779.641 678.848 577.262 498.531 496.748.4531 496.740 302.465 272.781 253.932 242.682 233.028 233.028 231.363 230.582 | 58.500 78.468 101.405 129.281 149.645 160.639 160.537 148.653 126.505 98.140 68.099 44.303 26.018 14.139 7.137 3.376 1.501 | $\begin{array}{c} 1.\ 000\\ 1.\ 062\\ 1.\ 153\\ 1.\ 283\\ 1.\ 473\\ 2.\ 064\\ 2.\ 689\\ 3.\ 306\\ 3.\ 666\\ 3.\ 938\\ 4.\ 121\\ 4.\ 291\\ 4.\ 291\\ 4.\ 323\\ 4.\ 338\\ \end{array}$ | |

WHITE PERSONS IN VARIOUS LOCALITIES IN THE UNITED STATES

TABLE 10.—Mumps

WHITE PERSONS IN VARIOUS LOCALITIES IN THE UNITED STATES

| Age interval | Of 1,000 livi | ng children o | Case rate per 1,000 | | |
|--|--|--|--|---|--|
| Period of lifetime between 2 exact ages (years) | Previously attacked (number at beginning of age inter- val who have had an attack) | Case rate per 1,000 total children (number having an attack during age interval) | Not provi- ously attacked (number at beginning of age inter- val who have not had an attack) | children not previ- ously attacked (number attacked during age interval among 1,000 who had not had an attack at beginning of age interval) | Ratio of case rate among chil- dren not previously attacked to case rate among all children (rate among all children =1.000) |
| (1) | (2) | (3) | (4) | (5) | (6) |
| 0-1 | 0 8,500 21,168 40,384 70,992 114,816 170,847 234,880 300,840 363,081 417,951 444,043 561,588 531,667 555,624 554,549 554,049 | 8, 500 12, 668 19, 216 30, 608 43, 824 56, 031 64, 033 65, 960 62, 241 54, 870 46, 092 37, 545 30, 079 23, 957 19, 125 15, 415 12, 629 | $\begin{array}{c} 1,000,000\\ 991,500\\ 978,832\\ 959,616\\ 929,008\\ 885,184\\ 829,153\\ 765,120\\ 009,160\\ 636,919\\ 532,049\\ 532,049\\ 535,957\\ 498,333\\ 444,376\\ 425,251\\ 409,836\\ \end{array}$ | $\begin{array}{c} 8.500\\ 12.777\\ 19.632\\ 31.896\\ 47.173\\ 63.299\\ 77.227\\ 86.299\\ 89.023\\ 86.149\\ 70.189\\ 70.052\\ 60.350\\ 60.350\\ 60.352\\ 60.350\\ 60.352\\ 60.350\\ 60.352\\ 60.350\\ 60.352\\ $ | $\begin{array}{c} 1.\ 000\\ 1.\ 009\\ 1.\ 022\\ 1.\ 042\\ 1.\ 076\\ 1.\ 130\\ 1.\ 206\\ 1.\ 376\\ 1.\ 430\\ 1.\ 570\\ 1.\ 430\\ 1.\ 570\\ 2.\ 135\\ 2.\ 255\ 2.\ 255\ 2.\ 255\ 2.\ 255\ 2.\ 2$ |

TABLE 11.—Chicken pox

| WHITE | PERSONS | IN | VARIOUS | LOCALITIES | IN | THE | UNITED | STATES |
|-------|---------|----|---------|------------|----|-----|--------|--------|
| | | | | | | | | |

| Age interval | Of 1,000 livi | ing children | Case rate per 1.000 | | | |
|--|---|---|---|---|--|--|
| Period of lifetime between 2 exact ages (years) | Previously attacked (number at beginning of age inter- val who have had an attack) | Case rate per 1,000 children (number having an attack during age interval) | Not previ- ously attacked (number at beginning of age inter- val who have not had an attack) | children not previ- ously attacked (number attacked during age interval among 1,000 who had not had an attack at beginning of age interval) | Ratio of case rate among chil dren not previously attacked to case rate among all children (rate among all children =1.000) | |
| (1) | (2) | (3) | (4) | (5) | (6) | |
| 0-1 | 0 41.000 88.566 143.650 204.875 266.718 324.589 375.284 417.123 449.800 474.016 491.083 502.549 508.891 514.370 516.994 518.461 | 41,000 47,566 55,084 61,225 61,843 57,871 50,695 41,839 32,677 24,216 7,342 4,488 2,615 1,466 7,342 4,488 2,615 1,467 -,785 | 1,000.000 959,000 911.434 856.350 795.125 733.282 675.411 624,716 582.877 550.200 525.984 508.917 497.451 490.109 485.621 483.621 483.539 | 41,000 49,600 60,437 71,495 75,058 66,973 56,062 44,013 32,44 822,530 14,750 9,157 5,385 3,037 1,630 | 1.000 1.043 1.067 1.168 1.258 1.368 1.368 1.461 1.461 1.901 1.905 2.010 2.040 2.059 2.070 2.076 | |

TABLE 12.—Scarlet fever

WHITE PERSONS IN VARIOUS LOCALITIES IN THE UNITED STATES

| Age interval | Of 1,000 livi | ing children | Case rate per 1.000 | | | |
|--|--|--|---|---|---|--|
| Period of lifetime between 2 exact ages (years) | Previously attacked (number at beginning of age inter- val who have had an attack) | Case rate per 1,000 total children (number having an attack during age interval) | Not previ- ously attacked (number at beginning of age inter- val who have not had an attack) | children not previ- ously attacked (number attacked during age interval among 1,000 who had not had an attack at beginning of age interval) | Ratio of case rate among chil dren not previously attacked to case rate among all children (rate among all children =1.000) | |
| (1) | (2) | (3) | (4) | (5) | (6) | |
| 0-1. 1-2. 2-3. 3-4. 4-5. 5-6. 6-7. 7-8. 8-0. 9-10. 10-11. 11-12. 12-13. 13-14. 14-15. 15-16. 16-17. 14-15. 15-16. 16-17. 14-15. 15-16. 16-17. 17-17. 1 | 0 1.200 5.775 15.416 26.643 38.710 50.900 62.589 73.298 82.710 90.670 97.160 102.270 106.161 109.027 111.073 112.487 | 1.200 4.575 9.641 11.227 12.067 12.190 11.689 10.709 9.412 7.960 6.490 6.490 5.110 3.891 2.866 2.046 1.414 .948 | 1,000.000 998,800 994,225 984,584 973,357 961,290 949,100 937,411 926,702 917,290 909,330 902,840 897,730 893,839 890,973 888,927 887,513 | $\begin{array}{c} 1.200\\ 4.580\\ 9.697\\ 11.403\\ 12.397\\ 12.681\\ 12.316\\ 11.424\\ 10.156\\ 8.678\\ 7.137\\ 7.137\\ 5.660\\ 4.334\\ 3.206\\ 1.591\\ 1.668\end{array}$ | $\begin{array}{c} 1,000\\ 1,001\\ 1,006\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,07\\ 1,108\\ 1,118\\ 1,119\\ 1,122\\ 1,125\\ 1,127\\$ | |

TABLE 13.—Diphtheria

| | | • | | | | | | | |
|--|---|--|---|---|--|--|--|---|--|
| Age interval | O | f 1,000 livin | ng children | Case rate per 1,000 children | | | | | |
| Period of lifetime between 2 eract ages (years) | Previ- ously attacked (number at begin- ning of age interval who have had an attack) | Case rate per 1,000 total ch ldren (number baving an attack during age interval) | Not pre- viously attacked (number at begin- ning of age interval who have not had an attack) | Average number Schick negative (number at mid- pcint of age interval who are regative Schick) | Average number Schiek, positive (number at mid- point of age interval who are positive Schiek) | not pre- viously attacked (number attacked during age interval among 1,000 who had an attack at begin- ning of age interval) | case rate among children not pre- viously attacked to case rate among all children (rate among all children =1.000) | per 1,000 Schick positive children (number attacked during age interval among 1,000 positive Schick children) | Ratio of case rate among Schick positive children to case rate among all children (rate among all children = 1.000) |
| (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| 0-1 1-2 2-3 4-5 6-7 7-8 8-9 9-10 10-11 11-12 12-13 13-14 14-15 15-16 16-17 14-17 16- | 0 3.550 12.322 21. (95 31. 046 40. 131 48.748 56.739 63.996 70.455 76.056 £0.455 76.056 £0.455 76.056 £0.455 76.056 £0.455 76.056 £0.455 76.056 £0.455 85.000 85.000 85.000 85.000 | $\begin{array}{c} 3.550\\ 8.772\\ 9.373\\ 9.351\\ 9.085\\ 8.617\\ 7.591\\ 7.557\\ 6.459\\ 5.641\\ 4.834\\ 4.070\\ 3.367\\ 2.737\\ 2.139\\ 1.721\\ 1.331 \end{array}$ | 1,000,000 996,450 987,678 978,305 968,954 968,954 969,859 961,252 943,261 936,004 929,545 923,564 929,545 923,604 919,070 911,633 906,856 906,707 504,986 | 111. 100 67. 372 111. 510 161. 920 215. 722 270. 214 323. 001 372. 143 416. 286 454. 636 454. 636 454. 636 456. 927 513. 200 534. 221 549. 774 562. 477 571. 321 577. 610 | 888, 900 932, 628 888, 490 838, 060 838, 060 838, 060 838, 060 676, 599 627, 857 583, 714 545, 364 513, 073 486, 709 455, 779 450, 226 437, 523 428, 679 422, 390 | 3. 550 8. £03 9. 558 9. 376 8. 977 8. 400 7. 694 6. 601 6. 609 5. 232 2. 408 3. 660 3. 660 3. 660 3. 680 3. 428 1. 898 1. 471 | 1.000 1.004 1.012 1.022 1.032 1.041 1.041 1.060 1.068 1.076 1.082 1.083 1.097 1.100 1.103 1.105 | 3.994 9.406 10.549 11.158 11.584 11.584 11.804 11.558 11.035 10.344 9.422 8.362 7.229 6.079 5.003 4.015 3.151 | 1. 125 1. 072 1. 125 1. 193 1. 275 1. 370 1. 477 1. 593 1. 713 1. 834 1. 949 2. 055 2. 147 2. 221 2. 286 2. 333 2. 367 |

WHITE PERSONS IN VARIOUS LOCALITIES IN THE UNITED STATES

Figure 8 shows the incidence rates at each age up to 16 years, as indicated by the method of differences which was outlined above. Although the age curves of the different diseases are similar in many respects, there are distinct variations. The maximum incidence varies considerably. Diphtheria reaches its maximum at about 3 years of age-approximately nine-tenths of 1 per cent of 3-year-old children having diphtheria within the next year. According to data derived from the histories, measles reaches its maximum at about $3\frac{1}{2}$ years, approximately 13 per cent of children having measles between 3 and 4 years of age. The maximum incidence of whooping cough and chicken pox occurs at about 4 years of age, approximately 10 per cent of children having whooping cough and 6 per cent having chicken pox within the next year. The maximum incidence of scarlet fever comes at about 5 years of age, approximately 1.2 per cent of children having the disease between 5 and 6 years of age. Mumps reaches its maximum at a later age than any other of these diseasesabout 7 years, at which age approximately 7 per cent of children have the disease. The percentages quoted above are, of course, averages over many years. In the course of a single epidemic the

proportion of 3-year-old children who are attacked by measles may far exceed 13 per cent. The figure of 13 per cent for measles and the percentages quoted similarly for the other diseases are an average of epidemic and nonepidemic years and can not be expected to be duplicated in an actual count of the cases of these diseases occurring at specific ages unless the observations extend over a period of years long enough to average ut the effect of epidemic cycles of the diseases, perhaps a minimum of 7 years being necessary for this to be done satisfactorily.

It must be borne in mind that the disease which in Figure 8 shows the greatest incidence at a specific age is not necessarily the one



FIGURE 8.—Age incidence of the communicable diseases of childhood (as approximated from histories of past attacks)

which has the maximum asymptote in Figure 5, although such is usually the case. A disease like typhoid fever, which freely attacks persons of all ages, might have a rather low incidence at any one age and yet arrive at a higher asymptote than a disease like scarlet fever, which is confined largely to children.

Though these diseases differ as to their age of maximum incidence and in other fundamental respects, they are similar in that the rates decline rapidly as age increases after the maximum has been reached, mumps being the only one that shows any considerable incidence among the older children. To facilitate the comparison of the relative decline in the incidence of these diseases with age, the rates have been plotted in Figure 9 on a semilogarithmic or ratio chart. On such a graph an equal distance vertically represents an equal relative or proportional change rather than an equal actual change,



FIGURE 9.—Relative change in the age incidence of the communicable diseases of childhood (as approximated from histories of past attacks)

as in the usual graph on cross-section paper. It may be seen from this graph that measles, whooping cough, and chicken pox all decline relatively more rapidly with age than do diphtheria, scarlet fever, and mumps. In other words, the proportion of cases that occur among the older children is greater for diphtheria, scarlet fever, and mumps than for measles, whooping cough, and chicken pox. The actual number of cases of each of these diseases of course is much less among the older than among the younger children.

COMPARISON WITH RELATIVE AGE INCIDENCE AS INDICATED BY OTHER DATA

The preceding data on the age incidence of the communicable diseases of childhood have been derived from histories of past attacks of these diseases. Before proceeding to other computations based on the histories, it would be well to compare these data with rates obtained in more direct ways. Unless the latter data cover a considerable number of years, and unless the cases are fairly completely reported, the actual rates will not be comparable but the relative age incidence may be compared.

The extensive data collected by Fales (9) on the age incidence of these infectious diseases afford an exceptionally good picture of the age distribution of cases reported to State boards of health in several large areas. To this we may add the results of the Hagerstown periodic canvasses (30) and for some diseases the cases reported to the city health departments of Hagerstown, Md., Providence, R. I. (3), and Aberdeen, Scotland (33), (15). In Figure 10 the incidence rates at specific ages in the various areas have been plotted on a semilogarithmic chart, the rates for each disease being plotted in a separate section. On a semilogarithmic graph the shape and slope of the curves are comparable even though the actual rates are much higher in one area than in another.

The measles age incidence curve as approximated from the histories of past attacks is not unlike the age curve of the Hagerstown canvassed population where, presumably, the data included nearly all of the cases since they were secured by periodic visits to the house-However, it differs from all the other curves in two important holds. In the first place the peak of cases reported to State and respects. city health departments comes at about 6 years of age, or three years later than the peak of the incidence as estimated from histories. The rates in the preschool ages are relatively higher for Aberdeen and Providence than for the States-a fact which may be due in part to more complete reporting in those cities, particularly in the preschool ages, but which may also be the result of the greater concentration of cases of measles in the younger ages in urban than in rural areas (9).

In Providence the maximum incidence of measles, except for a peak at 6 years, occurs from 3 to 5 years of age, the rate for each of these ages being about the same. The 6-year peak stands above any age before or after it, the rate at 7 years being distinctly less than at 5 years. In other words, the high incidence at 6 years



may be a peak superimposed upon an otherwise falling curve. The 6-year peak no doubt occurs at a time of a really high incidence due to greater contact when the child enters school, but it may be

FIGURE 10.—Relative change in the age incidence of certain communicable diseases as indicated by cases reported to health departments compared with that indicated by data derived from histories of past attacks. [Years included in the data from the various areas and references where detailed tabular material may be found: Maryland, including Baltimore, (9)—measles 1907-1917, chicken pox 1913-1917, other diseases 1908-1917; New Jersey (9)—measles and whooping cough 1918-1921, other diseases 1917-1921; Massachusetts (9), 1918-1921; Connecticut (9), 1921-22; Providence, R. I. (3), 1917-1923; Aberdeen, Scotland (33, 15)—measles 1883-1902, whooping cough 1891-1900; Hagerstown, Md., canvassed families (30), December 1, 1921, to March 31, 1924; Hagerstown, reported to health department, 1922-1924]

somewhat heightened by reason of a larger proportion of cases being reported among school children than among children not in school.

The second respect in which the incidence curve as estimated from histories differs materially from most of the other curves is the somewhat less rapid decline immediately after the maximum has been reached and the increasingly rapid decline as age increases, with the resultant unusually low rate by the thirteenth or fourteenth year of age. The curve for the Hagerstown canvassed population is also in this respect rather similar to the curve based on histories but the rate for the age group, 15 to 19, is based on only one case and is therefore not reliable. However, for the ages under 15 years there are sufficient data to be fairly reliable, a total of 568 cases of measles having occurred in the 28-month period, only 6 of which were over 15 years of age. The curves for all the other areas fall very rapidly for a few years after the peak is reached, after which the decline slackens, indicating rates considerably higher at 13 to 16 years of age than those estimated from the histories. It is quite probable that the data based on the histories are unreliable for these older children, inasmuch as at these ages the average time since the attack is considerable and the occurrence of the milder diseases might be rather frequently forgotten. As already noted the error due to forgetting attacks would presumably increase with age, since the average time since the occurrence of the cases would increase as age increased. As the incidence materially decreases with age and the error due to forgotten attacks increases with age, the error of the rates in the older ages may be relatively large. It has already been stated that the history data are too unreliable to use as a basis for estimating the incidence among adults. For some of the diseases, particularly measles, whooping cough, and chicken pox, it is possible that the data above 13 years of age are not to be relied upon.

As already suggested, however, the rates in the lower ages in the reported cases are probably relatively low, because of less complete reporting in those ages. One class of cases that may be frequently unreported in a city with no follow-up of reported attacks are secondary cases, a physician being called for the primary case but not being called when the second child comes down. In the Providence data the modal primary case is 6 years of age, in this instance the 6-year peak being apparently the major one—in fact, almost the only peak. On the other hand, the modal age of the secondary cases is about 3 years. If the secondary cases are less well reported, it would result in relatively less complete reporting under 5 years than in the older ages.

The city of Syracuse, N. Y., is unusually well supplied with public health nurses. In the course of visits to households they discover cases of communicable diseases which have not been reported or perhaps not seen by physicians. For the years 1926–27, physicians

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reported 63 per cent of all cases under 5 years that came to the attention of the health department, as against 75 per cent of the cases 5 to 9 years and 85 per cent 10 to 14 years of age. Considered in single years of age, the most deficient reports were under 1 year, when 51 per cent of the reports were made by physicians, and 2 and 3 years of age with 60 per cent in each age being reported by the doctor. At both 6 and 7 years 76 per cent of the reports were made by physicians. In the average city with few or no public health nurses, the measles cases under 5 years would be least well reported and the relative age curve based on the reported cases would be too low for the younger ages.

It may also be worth while to make some comparison of the actual incidence of these diseases as reported to the State and city departments of health with those estimated from histories. Since 1916 the Providence department has made a special effort to get measles cases reported, and the data should be fairly complete for their kind. For the period 1917-1923 the average annual measles rate for persons under 15 years was 30 per 1,000, or slightly less than half of the rate estimated from the histories of measles—62 per 1,000.

The same comparison may be made in an entirely different way. If we have annual rates for children of different ages over a period long enough to average out the effect of epidemic and nonepidemic years. we may cumulate the rates from birth to adult ages and approximate the curve of the proportion of children who have at some time in their lives had an attack-that is, the history curve. In making such a cumulation, it must be remembered that rates for age groups of more than 1 year are averages for the age period and must be counted for each single year of age covered. In other words, a rate for 10 to 14 years must be counted five times in the cumulating process. Cumulating the Providence rates in this way, it is indicated that 44 per cent of the children have an attack of measles by the time they are 15 years of age-slightly less than half of the 89 per cent of the children in the history data. In Providence and other large cities. the real figure is probably above 89 per cent, as these diseases are more prevalent in cities than in rural areas, and some of the localities considered in the history data are more or less rural.

If the average annual rates for specific ages for Aberdeen, Scotland, (1883-1902) are cumulated up to 15 years, they indicate that a total of 63 per cent of the children 15 years of age had had an attack of measles, as against 44 per cent for Providence and 89 per cent in the data on histories. In other words, reporting was somewhat more nearly complete than in Providence but still far from complete. It is interesting to note that the difference in the reported incidence is in the early ages, particularly under 5 years. The rates for those ages in Aberdeen are higher than those for Providence but not as high as the rates estimated from histories. In the ages above 8 years there is not so much difference between the two cities.

The similar cumulation of the rates for the States shown in Figure 10 to indicate the proportion of children who would be reported as having measles by the fifteenth year of age, gives the following results: Massachusetts, 26 per cent; New Jersey, 25 per cent; Connecticut, 20 per cent; Maryland, 17 per cent. Rates for Baltimore cumulate to 24 per cent. Reporting would be expected to be less complete in rural areas and therefore in States including large rural areas within their boundaries.

In Figure 10 similar data on the age incidence of whooping cough have also been plotted. The rates for the younger children of the Hagerstown canvassed population are rather irregular but suggest a relatively higher incidence for the very early ages, particularly under 1 year, than is indicated by the curve derived from histories. The same might be said of the Connecticut curve. The data for Maryland and for Aberdeen, Scotland, would indicate the same thing, but in both of these instances the rate under 1 year is slightly less than the rate at 1 or 2 years of age. The curves for Massachusetts and particularly Connecticut are not unlike the history data so far as the early ages are concerned.

With the exception of Aberdeen the curves all decline less rapidly in the older ages than does the curve derived from the histories. It is probable that the history data for whooping cough, like that for measles, are unreliable after about 13 years of age.

If the whooping cough rates are cumulated (as discussed in connection with measles) to indicate the proportion of children who would be reported as having had whooping cough by the fifteenth year of age, the following results are obtained: Aberdeen, 46 per cent; New Jersey, 10 per cent; Massachusetts, 9 per cent; Connecticut, 8 per cent; Maryland, 4 per cent. In Baltimore the rates cumulate to 6 per cent. These percentages may be contrasted with 77 per cent found on questioning to give a history of whooping cough. Unless whooping cough varies considerably in different localities, the disease appears to be much less completely reported than measles.

Figure 10 also shows similar comparative data for chicken pox, but the number of localities available is less than in the case of measles and whooping cough. The rates in the Hagerstown canvassed population show a slightly different curve from that of the histories. However, the curve is irregular and, because of the few cases over 10 years of age, does not give much idea of the nature of the curve for older children. The curves for the three States are all very similar to one another but somewhat different from the curve based on histories, having later peaks and slower rates of decline in the older ages. It is probable that the chicken pox rates derived from histories, like the measles and whooping cough rates, are unreliable after about 13 years of age.

The cumulation of the chicken pox rates to show the proportion of children who would be reported as having had chicken pox by the fifteenth year of age gives the following results: New Jersey, 9 per cent; Connecticut, 7 per cent; Maryland, 5 per cent (Baltimore, 7 per cent), in contrast to 52 per cent of persons who give a history of chicken pox.

The peak of the diphtheria incidence curve derived from the histories comes somewhat earlier than that based on the other data, but in other respects the curves are all fairly similar. The same may be said of scarlet fever, except that the curve derived from histories declines in the older ages slightly more rapidly than the reported incidence of the disease.

Cases of scarlet fever and diphtheria are available by age for Providence for long periods (3). Scarlet fever cases for the period 1887-1923 have a maximum at 6 years, that apparently being about the only peak, the number of cases at 5 and 7 years not being much below the number at 6 years. In the instance of diphtheria the peak is somewhat earlier, around 4 years, with a secondary peak at 6 years. Diphtheria and scarlet fever vary in the different geographic sections, being considerably more prevalent in the northern than in the southern part of the country. The average annual diphtheria case rate in Providence for the period 1917-1923 was 8.8 per 1,000 population under 15 years, and in Baltimore the corresponding rate for the period 1908-1917 was 5.1 per 1,000. The rate estimated from the histories was 6.8 per 1,000 under 15 years of age. Either the estimate from the histories is low or diphtheria is more prevalent in Providence than in Baltimore or in the localities from which the history data were secured. Since diphtheria is more prevalent in cities than in rural areas and in the North than in the South, it may well be more prevalent in Providence, for a considerable number of the communities considered in the history data are in the South, and some are rather rural in character. Probably the reporting of the recognized cases of a serious disease like diphtheria is nearly complete in Providence. The lower rate in Baltimore may be due to less complete reporting or to a lesser prevalence of the disease or both.

The average annual scarlet fever case rate was 5.0 per 1,000 population under 15 years in Providence for the period 1917–1923 and 5.7 in Baltimore for the period 1908–1917. Both of these rates are lower than the corresponding rate estimated from the histories, 7.6 per 1,000. Of course, both diphtheria and scarlet fever are reported more nearly completely than the less severe of the children's diseases.

In making the various comparisons of the data derived from histories with other data, it is not assumed that the age incidence of
these communicable diseases is the same in all places and under all circumstances. Fales (9) shows that the average age of occurrence of these diseases is younger in urban than in rural areas. Halliday (12) shows that measles occurs at earlier ages in tenement houses with a common entrance and hall, thus affording more contact between the children of the several families housed in the building. The age incidence of these diseases must be influenced to some extent by many factors. All that the curves derived from the histories represent is, as already stated, the approximate average age incidence in the communities from which data on infectious disease histories were secured.

The comparisons with other data seem to indicate that at least up to 12 years the relative age curves derived from the histories are not unreasonably different from curves based on reports to health departments, particularly when allowance is made for the probable variation with age in the completeness of reporting. The great advantage of the rates derived from histories over the reported rates lies in the fact that the former approximate the complete incidence of the disease whereas the rates based on reported cases are usually very incomplete, particularly for the milder diseases.

Age Incidence Among Children Not Previously Attacked

Of more importance, perhaps, than the incidence rates shown in Figure 8 for all children is the incidence of these diseases among children who have not previously suffered an attack of the disease, or, in the case of diphtheria and scarlet fever, who may be found to be relatively susceptible by the Schick and Dick tests, respectively. The cumulative curves shown in Figure 5 indicate what proportion of children of each age have already suffered attacks of these diseases in so far as the clinical attacks are recognized and remembered. Tf we eliminate from consideration the children who have already been attacked we can compute what proportion of the children of different ages who have not previously been attacked suffer attacks of these diseases at a given age. For example, if 32 per cent of children have had whooping cough by the time they arrive at the fourth year of age then only 68 out of 100 children 4 years of age are still unattacked by the disease. If 10 per cent of the children have the disease between 4 and 5 years of age (as determined by the derivative or difference method already discussed) then we may relate these 10 cases occurring in 100 children not to the total 100 children but to the 68 who have not been attacked and find what proportion of children 4 years of age and not previously attacked have the disease before they reach their fifth year. Carrying this principle to other ages we may determine approximately the incidence of whooping cough among children of

each age who have not previously been attacked. Such rates for whooping cough and also for the other diseases have been plotted in Figure 11. In this figure the rates for diphtheria and scarlet fever are based on children who have not had a clinical attack; diphtheria rates among children with a positive Schick reaction will be considered later.

In connection with Figure 11 it should be remembered that these rates are not at all comparable with secondary attack rates among children not previously attacked, that is, with attack rates among children in a household where a case of whooping cough, for example, has occurred. In the computation of the secondary attack rate the only persons taken into consideration are those who have been exposed to the disease, in so far as living in the same household with a case constitutes exposure to the infection. The rates shown in Figure 11 are based on all children not previously attacked, including those in families with no attacks as well as those in families where cases have occurred. The vital difference between the two types of rates is, therefore, known exposure to the disease.

The age curves among children who have not previously been attacked are by no means the same for all diseases (fig.11). The maximum incidence in the case of measles and whooping cough comes at about 6 years of age. In other.words, among children who have not previously been attacked, there is, according to these data, a greater probability of having these diseases at 6 years than at any age before or after that time. The maximum incidence of chicken pox among children not previously attacked also comes at 6 years or possibly a little earlier, but the maximum incidence of mumps comes between 8 and 9 years of age. Of children who have not previously had an attack of measles, nearly one-fourth suffer an attack during the sixth year of age; for whooping cough the corresponding figure is about 16 per cent.

After the age of maximum incidence in the group, the rates among children who have not suffered a prior attack decline rapidly with age. This is true of every one of the diseases under consideration. The decline in such rates could be due to either or both of two factors:

(1) There may be, among children who have not had an attack of whooping cough, for example, a number who nevertheless have a certain degree of immunity to the disease. (a) Some children may develop a specific immunity without an attack of the disease, as in the case of Schick-negative children who have not had diphtheria. (b) As age increases, children may develop more or less general immunity to infections of this type which partially protects them from attacks. (c) An unrecognized or undiagnosed attack may have occurred. (d) Mild or other attacks, particularly those occurring in the early ages, may have been forgotten by older children when

making the report on the history of infectious diseases. These forgotten attacks merge into the unrecognized attacks, inasmuch as the milder the case the more probable it is that the individual would forget it, or, if it occurred in his early childhood, he might never have



FIGURE 11.—Age incidence of the communicable diseases of childhood among persons not previously attacked (as approximated from histories of past attacks)

learned from his parents or others that he had suffered an attack. (e) Some children may have a natural immunity to the disease; that is, an immunity which they have had from birth.

(2) The decline in the rates as age increases among children not previously attacked may be due in part to less contact with cases or

to a difference in the intimacy of the contact. (a) Contact within the family may change. Inasmuch as there are many families with as many as four children with a difference of not more than six years between the ages of the oldest and the youngest child, it would not appear to be an unreasonable assumption that children who have reached the twelfth or thirteenth year are not as likely as younger children to be brought into contact within the family with a case of these diseases, since the maximum incidence occurs in the younger children. The smaller the family the less frequently will older children be found in the same household with younger children. Moreover, when a case occurs it must be easier to prevent its spread to older children in the family than to very young children who do not understand why they should keep away from the patient's room. (b) Contact with cases in school may be less frequent. Older children are not often in the same grade or room with the younger children at school, and therefore are less likely to be exposed to infections occurring among the younger children.

Chapin (3) and Butler (2) both show, by extensive data for Providence, R. I., and Willesden, England, respectively, that the measles secondary attack rate² among children in these cities who have not suffered a prior attack declines more or less regularly from a maximum of nearly 90 per cent at 2 to 4 years of age to considerably less than half that percentage by 15 years of age. The decline is more rapid after about 10 years of age. Since in dealing with the measles secondary attack rate there are considered only children living in households with a case of the disease, the decline could hardly be attributed to lessened contact as age increases, but must be due to some degree of immunity on the part of the children exposed to the disease, even though they gave no history of a prior attack.

From the histories of diphtheria and the Schick test we may approximate diphtheria rates among children giving a positive Schick test (Table 13). Such rates appear to decline as age increases after about the sixth or seventh year. The decline is not so rapid as in the case of some of the other infectious diseases; but, as already noted, the rate for diphtheria among all children also declines less rapidly than the rates for some of the other infectious diseases. These facts are illustrated in Figure 12, where the rates among children not previously attacked, including diphtheria rates among children giving a positive Schick test, have been plotted on a semilogarithmic chart.

It will be remembered that the curve (fig. 5) of the percentages of children of different ages who have had an attack of whooping cough,

³ In computing the secondary attack rate, cases occurring in the household (other than the primary case) are related to the number of persons in the household (other than the primary case), thus securing the percentage of persons exposed to the disease by family contact who suffer an attack. In computing the secondary attack rate among persons who have not had a prior attack, all persons exposed to the disease who have suffered a prior attack are eliminated from the computation.

for example, approaches an asymptote, or upper limit, which indicates the percentage of persons who at some time in their lives suffer a recognized attack of whooping cough. The asymptotes vary for the different diseases; about 10 per cent of persons suffer recognized attacks of diphtheria and 89 per cent suffer recognized attacks of measles, the other diseases falling between these two extremes. As already suggested, those persons who apparently do not have attacks



FIGURE 12.—Relative change in the age incidence of the communicable diseases of childhood among persons not previousy attacked (as approximated from histories of past attacks)

of whooping cough, for example, may have a natural immunity to the disease, that is, an immunity which they had at birth and have had throughout life. It might be argued that all of the 23 per cent who never have whooping cough represents persons with a natural immunity to the disease. If we assume that such is the case, then to find whooping cough rates among really susceptible children we would

take out of consideration at every age the 23 per cent who never have a recognized clinical attack of whooping cough. To revert to the example previously considered, if 32 per cent of children have had whooping cough by the time they arrive at the fourth year of age. then the other 68 out of 100 children 4 years of age are not all susceptible to the disease but only the 45 left after deducting the 23 out of 100 who are assumed to have been immune from birth. Tf 10 per cent of children have the disease between 4 and 5 years of age, then we may relate these 10 cases occurring in 100 children not to the total 100 nor to the 68 who have not had attacks but to the 45 who have not had attacks but who will at some time in their lives suffer attacks and are therefore known to be susceptible. If rates of this sort are computed, they rise consistently from birth to 16 years of age for every disease except mumps. In fact, the nature of the curves fitted to the data would indicate that this is true; but, since the curves give a fairly reasonable fit to the data, it would appear that the data themselves are of the same nature.

It does not seem reasonable to assume that all adults who do not report recognized attacks at some time in their lives have a natural immunity to the disease which they have had from birth. As already pointed out, some may have acquired an immunity by an unrecognized or a forgotten attack and some may by chance even escape exposure to the disease. It would seem more reasonable to assume that only some of the adults who do not report recognized attacks have a natural immunity from birth. We have at present no means of estimating what proportion of person not suffering recognized attacks have a natural immunity to these diseases. Computations in this study have, therefore, been limited to incidence rates among children who did not report a prior attack, with no attempt to correct for natural immunity or immunity acquired by unrecognized or forgotten attacks.

COMPARISON WITH OTHER DATA ON THE AGE INCIDENCE AMONG CHILDREN NOT PREVIOUSLY ATTACKED

For a considerable group of school children in Hagerstown, Md., data were obtained on the history of childhood diseases prior to the period of observation, as well as the occurrence of the disease for a period of four school years during which time weekly reports were received from the teachers on absences of their pupils on account of sickness. These data made it possible to compute incidence rates for children of different ages who had not suffered a prior attack of the disease. These rates are shown in Table 14. In the instance of measles the only children considered were those who had not suffered a prior attack of either measles or German measles.

TABLE 14.—Age incidence of certain communicable diseases among children who had not had a prior attack ¹ of the disease, Hagerstown, Md., 1921–1925

[Based on teachers' weekly reports of illnesses among white school children and reports by the children or parents of histories of attacks of these diseases prior to the period of observation]

| | Case rates per 1,000 persons per | | | | Number of cases occur- | | | | School years of exposure | | | |
|----------------------|---|--|---|--|----------------------------|---------------------------|---------------------------|-----------------------------|-------------------------------|---------------------------------|---------------------------------|--------------------------------|
| | year among children who had | | | | ring among children | | | | for children who had | | | |
| | not had a prior attack of the | | | | who had not had a prior | | | | not had a prior attack | | | |
| | disease | | | | attack of the disease | | | | of the disease | | | |
| Age in years | Measles, 1921- | Whooping cough, | Chicken pox, | Mumps, 1924- | Measles, 1921- | Whooping cough, | Chicken por, | Mumps, 1924- | Measles, 1921- | Whooping cough, | Chicken por, | Mumps, 1924- |
| | 1925 | 1921-1925 | 1921-1925 | 1925 | 1925 | 1921-1925 | 1921-1925 | 1925 | 1925 | 1921-1926 | 1921-1925 | 1025 |
| Total 6-11 | 322. 2 | 63.1 | 90.1 | 416.5 | 242 | 126 | 156 | 359 | 751 | 1, 996 | 1, 731 | 862 |
| 6 7 8 10-11 | 446. 7 438. 0 323. 4 102. 0 53. 2 | 130, 8 87, 9 73, 3 30, 6 19, 1 | 134. 1 157. 7 66. 5 53. 7 22. 0 | 363. 6 395. 8 433. 3 470. 6 393. 0 | 67 106 54 10 5 | 81 42 33 11 9 | 33 73 27 16 7 | 24 57 78 88 112 | 150 242 167 98 94 | 237 478 450 359 472 | 246 463 406 298 318 | 66 144 180 187 285 |

¹ The only children considered in this table are those who, prior to the beginning of the school year under consideration, had not had an attack of the disease being tabulated. In the case of measles this table includes only those who had not had an attack of measles (rubeola) or of German measles (rubeila).

It must be borne in mind that these are not secondary attack rates, inasmuch as they are based on all school children who had not had a prior attack regardless of exposure to the disease, whereas the secondary attack rate considers only children who are exposed to the disease by family contact.

It may be seen from Table 14 that the measles and whooping cough rates both decline regularly after 6 years of age. The chicken pox rate declines after 7 years. In the case of mumps the rate increases up to 9 years and then declines. In all cases these data are more or less in agreement with similar rates approximated from histories of past attacks. The attack rates among children who have not had a prior attack seem to decline definitely with age.

A study of measles by the (British) Medical Research Council (12) shows for one school the percentage of the children who had not had a prior attack of measles who were attacked during the year under observation. The data are shown in three departments, infant (5 and 6 years old), junior (7, 8, and 9 years old), and senior (10, 11, and 12 years old). Measles occurred in all three departments during the year. Of 62 children in the infant department who had not had a prior attack, 41 children, or 66 per cent, were attacked during the year; of 28 children in the junior department who had not previously had measles, 15 children, or 54 per cent, were attacked; and of 17 children in the senior department who had not previously had measles, 5 children, or 29 per cent, were attacked. Although the numbers considered are too small to furnish reliable results, it may

be noted that, in agreement with the data based on histories of measles as well as the Hagerstown school data, the rates among children who had not had a prior attack decline as age increases.

Comparison of the Age Incidence Among All Children and Among Children Not Previously Attacked

In order to facilitate comparison of the age incidence of these diseases among all children with the age incidence among children not previously attacked, the curves plotted in Figures 8 and 11 have been reproduced in Figures 13, 14, and 15, the two curves for each disease being put on the same graph. For diphtheria (fig. 15) there is a third line representing diphtheria case rates among children with a positive Schick test. For every disease, of course, the maximum incidence among children not previously attacked comes at a later age than the maximum incidence among all children. Considering all children the maximum incidence of measles, according to the data derived from histories, comes in the third year of life, during which year about 13 per cent of all children are attacked. The maximum among children not previously attacked comes about 6 years of age, during which year about 24 per cent of the children who have not previously had the disease suffer an attack. Likewise in the case of whooping cough the maximum incidence is moved from 4 years to 6 years when children who have already been attacked are eliminated from consideration, and similarly for chicken pox, although the peak in chicken pox is slightly under 6 years of age. Even in the case of diphtheria, where the peak for all children is about 3 years of age, the maximum incidence among children of a positive Schick test is about 6 years. The fact that the maximum incidence among children not previously attacked occurs around 6 years of age in several of the diseases would seem to be associated with the time of school entrance and the consequent increase in the opportunity for contact and infection. In the case of mumps the maximum incidence among all children, as well as among those not previously attacked, occurs at a later age, the former occurring at about 7 years and the latter at 8 to 9 years of age. Unfortunately, we do not have sufficient data on the percentage of children with a negative Dick test to approximate what the curve would be, and it is therefore impossible to say when the maximum incidence of scarlet fever among Dick-positive children occurs. The elimination of those who have had a clinical attack of scarlet fever leaves the maximum incidence at about 51/2 to 6 years of age; but, of course, the number of children who have suffered a clinical attack is only a small proportion of the total with a Dick-negative reaction.



FIGURE 13.—Age incidence of certain communicable diseases among all children and among children not previously attacked (as approximated from histories of past attacks)



FIGURE 14.—Age incidence of certain communicable diseases among all children and among children not previously attacked (as approximated from histories of past attacks)



In Figures 13, 14, and 15 the age curves for children not previously attacked appear to decline even more rapidly than the age curves based on all children. Although the actual decline is greater, the

FIGURE 15.—Age incidence of scarlet fever and diphtheria among ell children and among children not previously attacked and the age incidence of diphtheria among children giving a positive Shcick test (as approximated from histories of past attacks and the results of Schick tests)

relative decline is not as great as in the curves for all children. Figure 16 shows for specific ages the ratio of the rate among children not previously attacked to the rate among all children.

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This ratio is, of course, in all instances unity for the first year of life, inasmuch as no children have suffered attacks before birth and therefore all fall in the group not previously attacked. As age increases, the ratios in every case increase. By the twelfth year of age the measles rate among susceptibles is 8.5 times what it is among the total population of that age, whooping cough 4.1 times, and



FIGURE 16.—Ratio of the case rate among children not previously attacked to the case rate among all children (as approximated from histories of past attacks)

chicken pox 2.0 times. The diphtheria rate among 12-year-old children with a positive Schick test is about 2.1 times the rate among all children of that age.

Death Rates at Specific Ages

To complete the picture of the age incidence of these infectious diseases of children, it seemed worth while to find the age incidence of the fatal cases; that is, the death rates at specific ages. Deaths as published by the United States Census Bureau are tabulated in single years of age only up to 5 years, after which age we must be content with 5-year age groups. However, since the deaths occur largely in the ages under 5, and in every instance the death rates are declining rapidly before 5 years of age, the 5-year age groups above 5 years are sufficient to give a very accurate picture of the death rates from these diseases.

Since we are dealing with average case rates and are trying to get away from unusually high rates occurring in epidemics or usually low rates occuring between epidemics, we selected the 7-year period 1917– 1923 for the deaths. A second reason for selecting this period, rather than a more recent one, was that the average of the rates for all ages for the period 1917–1923, in the case of nearly every one of the diseases considered in this study, approximates the average of the rates for the considerably longer period 1910–1926. The latter consideration is particularly important if we expect to relate the death rates to the case rates to secure an approximation of the case fatality of the different diseases. This phase of the matter will be taken up in a later section.

Figure 17 shows the death rates among children of specific ages in the registration area of the United States for the period 1917-1923. A glance at this figure in comparison with Figure 8, which shows the case rates of the same diseases at specific ages, will indicate the widely different character of the age incidence of the cases and of the deaths. The death rates for whooping cough are far higher under 1 year of age than at any other age among children. In the case of measles the maximum death rate comes between 1 and 2 years of age, after which time there is a constant decline as age increases. The slightly lower death rate for measles under 1 year of age than from 1 to 2 years of age is apparently significant, inasmuch as it occurs consistently year after year in the data for the registration area. The death rate for diphtheria is highest between 2 and 3 years of age. The similarity of the age incidence of deaths to the age incidence of cases appears to be greater in the case of diphtheria than in any of the other diseases. Inasmuch as the fatality of diphtheria far exceeds that of any of the other infectious diseases of children, the greater similarity might be expected, the death curve being composed of a greater percentage of the total cases than in the instance of the other diseases where the fatality is much lower. The maximum death rate from scarlet fever also occurs between 2 and 3 years of age, but it is considerably lower than the death rate from diphtheria. Mumps was not shown separately as a cause of death until 1921, and the period

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1921-1926 was therefore used in the case of mumps. The death rate from mumps is very low; but, like whooping cough, the maximum rate comes in the first year of life.



FIGURE 17.—Mortality from certain communicable diseases at specific ages. (White children in the registration area of the United States, 1917-1923)

In Figure 18 the death rates from these diseases among children are plotted on a semilogarithmic chart to facilitate comparison of the relative decline in the rates for the different diseases as age increases.



FIGURE 18.—Relative change with age in the mortality from certain communicable diseases at specified ages. (White children in the registration area of the United States, 1917-1923)

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TABLE 15.—Average annual mortality from certain communicable diseases of childhood among white persons in the registration area of the United States

| | | Age in years | | | | | | | | | | |
|--|--|--|---------------------------------------|-----------------------------------|----------------------------------|------------------------------|----------------------------------|------------------------------|----------------------------------|--|--|----------------------------------|
| Years and disease | All ages 1 | Un- der 5 | 5-9 | 10-14 | 15-19 | 20-24 | 25-29 | 30-34 | 35-44 | 45 -54 | 55-64 | 65 and over |
| · · · · · · · · · · · · · · · · · · · | DEAT | H RAT | E PER | 100,000 | POPUI | ATION | PER Y | EAR | | | · | |
| 1917-1923 * | | | | | | | | | | | | |
| Diphtheria Scarlet fever Whooping cough Measles | 15. 55 4. 13 8. 79 7. 89 | 83. 40 18. 23 78. 88 56. 85 | 43. 87 10. 82 3. 21 7. 11 | 11. 55 3. 91 . 441 2. 27 | 3.81 2.65 .185 2.61 | 1.99 2.01 .096 1.65 | 1. 43 1. 44 . 059 . 953 | 1.36 .919 .046 .962 | 1. 11 . 546 . 039 1. 00 | 1. 0 2 . 229 . 043 . 976 | . 763 . 135 . 05 8 1. 01 | . 690 . 110 . 274 1. 67 |
| Mumps | . 109 | . 407 | . 132 | . 076 | . 079 | . 047 | . 042 | . 048 | . 036 | . 040 | . 073 | . 2 55 |
| | | | NUM | BER OF | DEAT | IIS . | | | | | | |
| 1917-1923 * | | | | | | | | | | | | |
| Diphtheria Scarlet fover Whooping cough Measles | 85, 772 22, 758 48, 499 43, 510 | 48, 704 10, 645 46, 065 33, 198 | 25, 079 6, 187 1, 834 4, 065 | 6, 108 2, 066 233 1, 203 | 1, 747 1, 214 85 1, 198 | 911 928 44 755 | 678 681 28 452 | 593 400 20 419 | 842 416 30 761 | 590 133 25 566 | 288 51 22 381 | 194 31 77 468 |
| Mumps | 541 | 213 | 68 | 36 | 32 | 19 | 18 | 19 | 2 5 | 21 | 25 | 65 |
| | YI | ARS O | F LIFE | EXPOS | ED (IN | THOUS | ANDS) | | | l | | |
| 1917-1923 ³ 1921-1926 ³ | 551, 706 496, 733 | 58, 399 52, 397 | 57, 167 51, 663 | 52, 889 47, 692 | 45, 836 40, 611 | 45, 839 40, 539 | 47, 414 42, 510 | 43, 546 39, 263 | 76, 133 58, 8 3 9 | 58, 020 52, 691 | 37, 712 34, 455 | 28, 107 25, 517 |

¹ Including a few of unknown age.

² Total registration area except that registration cities in nonregistration States are excluded after 1919.

Registration States of 1920.

Deaths from chicken pox and German measles are not available by age or color; the same is true of mumps prior to 1921. For the period 1917–1923, the average annual death rates per 100,000 total population of all ages in the registration area were: Chicken pox, 0.174; German measles, 0.0599; mumps, 0.134.

The two preceding figures have shown death rates for children only, the rates up to 5 years being computed for each single year of Although the death rates from these diseases are low for adults, age. the whole registration area for the period 1917-1923 affords sufficient data to indicate the age incidence among adults. (Table 15.) In Figure 19 the rates have been plotted on a semi-logarithmic graph for the ages throughout life, the age period under 5 years being plotted as a single group. Deaths from diphtheria and scarlet fever continue to decline as age increases, even up to 70 years. The death rates from whooping cough, mumps, and, to a lesser extent, measles, rise after about 50 years of age. Undoubtedly, few cases occur in the older ages and the fatality therefore must be rather high for the cases that do occur.

Case Fatality at Specific Ages

The examination of the curves representing the percentage of children of different ages who have suffered an attack of these diseases did not reveal any marked differences for the different localities



FIGURE 19.—Relative change with age in the mortality from certain communicable diseases at specific ages throughout life. (White persons in the registration area of the United States, 1917-1923)

other than that which might be accounted for by the cyclical variation in the incidence of these diseases. Inasmuch as the effect of these cycles is largely averaged out by reason of the combination of many localities in the general total, we may assume that the rates at specific

ages are fairly representative of what occurs on the average in these communities. No data have been collected for the deaths from the infectious diseases of children in the particular communities included in the case study. Inasmuch as deaths from these diseases are not frequent, it would, in fact, be necessary to include a considerable number of years or quite a large area in the observations in order to "average out" the effect of the chronological cycles in the incidence of the diseases and get anything like a normal or average death rate from these causes. In the absence of data specifically relating to the particular communities where data on the cases were collected, as already mentioned, death rates were computed for the registration area ³ of the United States, the period 1917-1923 being taken because the average of the rates for all ages for this period approximates the average of the rates for the considerably longer period, 1910-1926. Tf we take the ratio of the death rate for whooping cough at, say, 3 years of age to the case rate at the same age, we may obtain an approximation of the case fatality of whooping cough at that age. Case fatality as usually defined and as used in this study means the number of deaths per 100 cases. The particular advantage in the rates that result from data used in this study is that the case rate comes nearer representing the actual or complete incidence than a rate based on the reported cases of the disease, because of the many attacks of these mild diseases that are never reported to the health department. Table 16 and Figure 20 show the case fatality of the different diseases from birth to 15 years of age, the last 10 years being shown in 5-year age groups because the deaths from these diseases are not available in single year age groups above 5 years of age.

For measles the difference is not so large. The rates for all ages for the period 1910-1920 average 11 per cent higher in the urban than in the rural parts of the registration area (17); but because of the greater concentration of cases in the younger ages in the cities, the differences at certain ages might be greater.

³ The urban part of the registration area might have been more appropriate, but the deaths are not available by age.

Scarlet fever and diphtheria vary considerably in different geographic sections of the country, and all the children's diseases vary in age distribution in urban and rural areas (9). The average of the diphtheria death rates among white persons of all ages for the 11-year period 1910-1920 was 55 per cent greater in cities (10,000 or over in 1920) in the registration area of the United States than in the rural part of the area. The scarlet fever death rate was 64 per cent higher in the cities than in rural areas (17). At certain ages the differences may have been even greater. It may be seen, therefore, that, in so far as the case rates in this study applied to less rural districts than the average of the registration area, the estimated case fatalities would tend to be too low for these two diseases. On the other hand, the areas included in the history data probably include a higher percentage of southern localities than in the case of the registration area. Because diphtheria and scarlet fever are more prevalent in the North, the computations made in this study would tend to overstate the fatalities. Of the two errors the urban-rural error is probably the more important and the fatalities are probably underestimated.

Mortality from whooping cough is actually higher in the rural than in the urban part of the registration area. For the period 1910-1920 the urban rate for all ages averaged 13 per cent less than the rural rate (17). Aside from the possibility of error due to greater concentration of cases in the younger ages in the cities, the estimated whooping cough fatality would, therefore, appear to be overstated rather than understated.

Inasmuch as the disease history data used in this study included some rural localities, it could not be classified as all urban, and the error due to using death rates in the registration area as a whole would not be as large as the differences between urban and rural rates would indicate.

Because the case rates approximated from the histories would not include fatal cases, they have been corrected by adding the death rate to the case rate to approximate the total fatal and nonfatal inci-



FIGURE 20.-Estimated case fatality of certain communicable diseases at specific ages

dence as a basis for computing the fatality rate. The correction makes no real difference except in diphtheria and possibly scarelt fever.

TABLE 16.—Estimated case fatality of certain communicable diseases of childhood at specific ages

| Disease | Total | | | 1 | lge in years | | | |
|---|--|---|---|--|--|---|---|--|
| | years | Under 1 | 1 | 2 | 3 | 4 | 5-9 | 10-14 |
| | ES7 | IMATED C | ASE FATALIT | Y: DEATHS | PER 100 CA | SES 1 | | |
| Diphtheria Scarlet fever Whooping cough Measles Mumps | 6.964 1.467 .537 .370 .0054 | 12, 354 7, 236 3, 963 1, 507 , 0957 | 9. 782 3. 804 1. 329 1. 018 . 0352 | 9.576 2.360 .406 .354 .0148 | 8. 828 1. 833 . 174 . 177 . 0100 | 8.069 1.538 .098 .120 .0048 | 5. 725 1. 026 . 052 . 113 . 6022 | 3. 221 . 938 . 052 . 372 . 0024 |
| | | | DEATH RAT | E PER 100,0 | 00 s | | | |
| Diphtheria Scarlet fever Whooping cough Measles Mumps | 47. 43 11. 22 28. 57 22. 83 . 209 | 50. 04 9. 36 241. 42 94. 38 . 814 | 95. 11 18. 09 99. 49 108. 11 . 446 | 99. 26 23. 30 35. 87 45. 29 . 284 | 90. 54 20. 96 17. 53 23. 56 . 307 | 79. 74 18. 85 9. 99 14. 78 . 210 | 43. 87 10. 82 3. 21 7. 11 . 132 | 11.55 3.91 .44 2.27 .076 |
| | | | CASE RATE | PER 100,00 | 01 | | • | |
| Diphtheria Scarlet fever Whooping cough Measles Mumps | 681. 03 764. 72 5, 322. 67 6, 173. 73 3, 847. 71 | 405. 04 129. 36 6, 091. 42 6, 262. 08 850. 81 | 972. 31 475. 59 7, 487. 29 10, 616. 31 1, 267. 25 | 1, 036, 56 987, 40 8, 833, 97 12, 776, 79 1, 921, 88 | 1, 025. 64 1, 143. 66 10, 096. 83 13, 280. 96 3, 061. 11 | 988. 24 1, 225. 55 10, 168. 59 12, 351. 98 4, 382. 61 | 766. 27 1, 054. 62 6, 157. 21 6, 280. 61 6, 067. 53 | 358. 55 416. 91 847. 44 610. 17 3, 166. 18 |

WHITE PERSONS IN VARIOUS LOCALITIES IN THE UNITED STATES

¹ Percentage that death rate is of the case rate. ² In the registration area of the United States, 1917-1923, except for mumps, which is 1921-1926. ³ Including both the fatal and the nonfatal cases, that is, the case rate estimated from the **histories** plus the death rate in the registration area. The nonfatal case rate used for the 5-year age groups was derived by adjusting the rates for single years of age to the age distribution, within the 5-year age period, of the white population of the United States in 1920. The rate for the age group under 15 years was obtained by a similar method.

Without exception the highest case fatality occurs under 1 year of age. After that time the fatality rapidly falls until at 4 or 5 years of age it is very small for all diseases except diphtheria and scarlet The diphtheria case fatality stands out as far above any of the fever. other diseases and remains so even up to 15 years of age. Although the fatality of scarlet fever is considerably below that of diphtheria, it is nevertheless much greater than that of measles or whooping The fatality of mumps is almost negligible, even in the cough. voungest ages.

In Figure 21 these case fatality rates have been plotted on a semilogarithmic, or ratio, chart to facilitate the comparison of the relative decline with age in the case fatality of the different diseases. It may be seen that the decline with age is less in the case of diphtheria fatality than in the case of any of the other diseases. If measles and whooping cough can be postponed until 3 or 4 years of age, the danger of the attack is largely eliminated; but the same can not be said of diphtheria at any age. In this connection it should be noted that the data here used, both for cases and deaths, pertain largely to the

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period before toxin-antitoxin immunization was widely used, but they are, however, well within the period when antitoxin was extensively used. The case fatality of scarlet fever declines slightly more with age than that of diphtheria, but here again it does not appear that the postponement of cases beyond the first few years of life would avoid all the danger of the attacks.

It is doubtful whether the rise in the fatality of measles for the age group 10 to 14 years is significant. The fatality of measles seems to be higher for adults than for children, but whether the increase comes at as early an age as 10 to 14 is somewhat doubtful. The effect of forgotten attacks of these diseases has already been mentioned; attacks forgotten by children of the ages 10 to 14 would tend to make the case rate too small at those ages, and this, of course, would result in the indicated case fatality being too high.

Although the case fatality of mumps is very small, even in the youngest ages, it may be seen from Figure 21 that the relative age curve of its case fatality is like that of whooping cough and measles. In other words, what little danger there is in an attack of mumps in a child may be largely avoided if the attack can be postponed beyond the first few years of life. Although rarely fatal, mumps is more severe in adults than in children.

It has already been mentioned that data such as are used in the study of cases of these diseases are not sufficiently accurate to determine the incidence among adults. For this reason there is no way to estimate the case fatality of these diseases among adults. The death rates throughout life have already been shown and, inasmuch as the case rates for adults probably vary little with age, the curve of the death rates is a fair indication of the relative fatality of the few cases that do occur.

COMPARISON WITH CASE FATALITY AS INDICATED BY OTHER DATA

Out of the 562 cases of measles occurring in the Hagerstown canvassed population under 15 years of age, two deaths occurred, giving a fatality of 0.356 per cent. Out of the 36 cases of diphtheria under 15 years, one death occurred, giving a fatality of 2.77 per cent. In the case of measles the fatality is about the same as that obtained from the histories and death rates in the registration area of the United States, but the Hagerstown fatality is considerably less in the case of diphtheria. In neither disease can much dependence be put on the Hagerstown fatality rates as they are based on one and two deaths only. In the case of the other communicable diseases of childhood no deaths occurred in the Hagerstown canvassed population.

The average annual measles death rate in Providence for the period 1917-1923 was 47.6 per 100,000, or more than twice the rate

of 22.8 in the registration area as a whole. The case fatality in Providence for persons under 15 years was 1.57 deaths per 100 reported cases of measles for the period 1917-1923, as against 0.37 per cent estimated from the histories and the death rate in the registration area of the United States. Case fatalities based on the reported cases are, of course, too high because of the incompleteness of reporting of cases and the relatively complete reporting of deaths but even after making allowance for incompleteness of reporting, the Providence fatality rate would appear to be materially higher than the estimate based on histories.

On the other hand, measles cases and deaths reported to the Public Health Service in 44 States for the year 1927 indicate a fatality of 0.94 per cent (27). If the Providence Health Department has made a special effort to secure reports of measles cases and has secured reports of only about half of the total cases that might be expected to occur, it would not appear that more than one-third of the measles cases over the country as a whole would be reported to health departments. On the assumption that from one-third to one-half of the cases are reported, the actual fatality rate in these 44 States would be somewhere near the rate estimated from the histories and the death rate.

At certain ages, of course, the disagreement is greater than it is at all ages. In Aberdeen, Scotland, 1883-1902, where reporting seemed to be about two-thirds complete, the case fatality under 1 year of age was nearly 14 per cent and at 1 year of age was 10 per cent (33). In Willesden, England, however, the fatality under 1 year was less than 3 per cent and at 1 year was 5 per cent (2). There is, therefore, great variation in the case fatality of measles in different localities.

The death rate from diphtheria in Providence for the period 1917– 1923 was 58.6 per 100,000 persons under 15 years, or somewhat greater than the corresponding rate in the registration area for the same period, 47.4 per 100,000. The case fatality of diphtheria under 15 years in Providence for this period was 6.64 deaths per 100 cases. The corresponding fatality rate estimated from the histories and the death rate in the registration area was about the same, 6.96 deaths per 100 cases. At certain ages, however, the agreement is not so close.

In 45 States reporting diphtheria cases and deaths to the Public Health Service in 1927, the case fatality was 8.30 per cent (27).

The scarlet fever death rate for the period 1917-1923 was the same in Providence as in the registration area, 11.2 per 100,000 population under 15 years. The reported case rate in Providence was 5.02 cases per 1,000 persons under 15 years as against 7.65 as estimated from the historics. Since scarlet fever is more prevalent in the North than in the South and also more prevalent in the city than in rural areas, the reported cases in Providence may be even more incomplete than the above rates would indicate.

The case fatality of scarlet fever under 15 years of age in Providence for the period 1917-1923 was 2.24 deaths per 100 reported cases, as against 1.47 as estimated from the histories and the death rate in the registration area.

In 45 States reporting scarlet fever cases and deaths to the Public Health Service in 1927, the case fatality was 1.15 per cent (27).

Few data are available on the fatality of whooping cough. The reports of cases of this disease appear to be very incomplete, so the indicated fatalities are considerably too high. In 45 States reporting whooping cough cases and deaths to the Public Health Service in 1927, the case fatality was 4.08 per cent (27). Judging from data presented in connection with reported whooping cough rates, it would not appear that over one-eighth of the cases are reported. If allowance be made for incompleteness of reporting on this basis, the fatality in these States compares favorably with that estimated from the histories and the death rates.

In Aberdeen, Scotland, 1891–1900, whooping-cough fatality under one year was 12 per cent. Like the measles fatality in Aberdeen, whooping-cough fatality seemed to be unusually high.

Summary

Data from a number of sources on the percentage of children of specific ages who gave a history of a past attack of certain of the common communicable diseases of children were combined to secure an average curve of this type. The diseases included were measles, whooping cough, mumps, chicken pox, scarlet fever, and diphtheria. Although there is considerable variation in the age incidence of these diseases in different localities, the data at hand did not seem sufficiently accurate to make it worth while to study anything but the average for all communities combined. In data of this kind there are fairly large errors, not only those due to sampling but errors of observation such as those due to forgotten attacks, which would presumably become increasingly important as age increases. Nevertheless, it appears that the data are a fairly accurate and complete record of the proportion of the population that is attacked by these diseases. The data considered in this study are chiefly from urban communities of moderate size, but they include some distinctly rural areas.

The logistic curves used extensively by Reed and Pearl to describe population growth were fitted to the percentages of children of specific ages who had suffered attacks of these diseases. The autocatalytic type gave the best fit in the case of mumps, but the catalytic type gave the best fit for the other five diseases and for the curve of the percentage of children with a negative Schick test.

From these curves, which may be considered as representing an accumulation of cases in a hypothetical population observed from birth to adult ages, we may secure an approximation of the incidence of these diseases at specific ages. The maximum incidence varies from about 3 years for diphtheria to about 8 years for mumps. After the maximum is reached, the rates in every case decline rapidly as age increases.

From the cumulative curves we may also secure an approximation of the incidence of these diseases among children of specific ages who have not previously suffered a recognized and remembered attack and, therefore, in the absence of immunity from some other source, may be considered relatively susceptible. The maximum incidence of measles, whooping cough, chicken pox, and scarlet fever among children who have not suffered a prior attack and the maximum incidence of diphtheria among children of a positive Schick test occurs at about 6 years of age. It appears that, on the average, children who have not suffered attacks of these diseases prior to school entrance are more likely to be attacked during their first school year than at any other time.

After the maximum, the incidence rates among children who have not suffered a prior attack decline considerably as age increases, although the relative decline is not as rapid as in the instance of the rate among all children. Apparently this decline could be due either to the development of immunity or to a change in contact or both.

1. Some children counted as relatively susceptible because they report no prior attack may be immune to these diseases: (a) It may be that some children develop a specific immunity without having an attack of the disease, as in the case of Schick negative children who have not suffered an attack of diphtheria; (b) as age increases, children may develop a more or less general immunity to infections of this type, which partially protects them from attacks; (c) immunity may have been acquired by mild unrecognized or forgotten attacks; (d) some children may have from birth a natural immunity to some of these diseases.

2. The frequency and intimacy of contact with cases of these diseases may decrease to some extent as age increases.

Death rates were computed for children of specific ages in the registration area of the United States for the period 1917-1923. With the exception of diphtheria the maximum death rate in every instance comes at an earlier age than the maximum case rate.

By relating the death rates and the case rates we may secure an approximation of the case fatality of these diseases. In every instance the maximum fatality occurs under 1 year of age. The fatality of measles, whooping cough, and mumps declines to an almost negligible percentage by 5 years of age, but the decline for scarlet fever and particularly diphtheria is by no means as great as in the case of the other three diseases.

Acknowledgments

In the preparation of this paper unpublished data have been made available to us from several sources. Data on the history of these infectious diseases and on Schick tests in Baltimore were collected under the direction of Surg. W. H. Frost and Prof. James A. Doull, of the Johns Hopkins University School of Hygiene and Public Health. Data on the history of these diseases among students in various universities in the United States were collected by Surg. J. G. Townsend in connection with a study of respiratory diseases. Data on histories in Hagerstown, Md., were collected by Statistician Edgar Sydenstricker in connection with the study of morbidity in that city. Data on the histories of a few of these diseases were collected by the United States Children's Bureau in connection with a study of preschool children in Gary, Ind: Data on Schick tests among children in Syracuse, N. Y., were collected under the direction of Dr. G. C. Ruhland, health commissioner of that city. In addition to the above unpublished data which were made available to us, the few published reports that contain data of this type have been drawn upon.

In the analysis of the data I have had the advice of Statistician Edgar Sydenstricker and other members of the staff of the Office of Statistical Investigations and the associated field offices of the United States Public Health Service.

I am especially indebted to Dr. W. H. Frost, who read the manuscript and gave many valuable suggestions for its revision, and to Prof. Lowell J. Reed, of the Johns Hopkins School of Hygiene and Public Health, and consultant in statistics to the United States Public Health Service, for advice in connection with the fitting of the curves and the suggestion of the use of the catalytic type of curve for data which did not appear to conform to the autocatalytic type of curve.

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PREVALENCE OF COMMUNICABLE DISEASES IN THE UNITED STATES ¹

February 3-March 2, 1929

The prevalence of certain important communicable diseases in the United States, as indicated by weekly telegraphic reports received from State health officers,² from February 3 to March 2, 1929, is summarized below.

During the four weeks ended March 2 the incidence rates of diphtheria, scarlet fever, measles, typhoid fever, and smallpox were all lower than during the corresponding period of any of the three preceding years. For meningococcus meningitis and influenza, however, the incidence was appreciably above the normal.

¹ From the Office of Statistical Investigations, United States Public Health Service.

²The number of States reporting for the various diseases were as follows: Typhoid fever, 41; poliomyelitis, 43; meningococcus meningitis, 42; smallpox, 42; measles, 38; diphtheria, 42; scarlet fever, 41; influenza, 31.

Meningococcus meningitis.—During February the reported attack rate from this disease was the highest in 11 years. For the four weeks ended March 2 the number of cases totaled 952, representing a case rate (annual basis) of 11.9 per 100,000 population. This compares with a rate of 13.2 during February, 1918. Among the States showing an outstanding incidence of the disease were the following: Idaho (24 cases), Michigan (98 cases), Missouri (93 cases), Pennsylvania (40 cases), Utah (60 cases). On the other hand, an appreciable decline since January occurred in Oklahoma, Arkansas, and Montana.

Poliomyelitis.—The incidence rate of poliomyelitis maintained a normal level during the month of February. With the exception of 9 cases in California, 8 in New York, and 5 each in Pennsylvania and Alabama, no more than 3 cases were reported in any State during the four weeks ended March 2. The total number of cases (62) was slightly higher than that reported during the corresponding period in 1927 and lower than in 1926.

Typhoid fever.—The incidence rate of typhoid fever was relatively low during the month of February; 538 cases were reported, as compared with 703 during the corresponding period in 1928 and 941 in 1927.

Smallpox.—In Maine 21 cases of smallpox were reported during the four weeks ended March 2. In Vermont the number of cases totaled 11 for the same period. Reports from the Western and East and West North Central sections of the country indicated a continuation of the relatively high prevalence of smallpox in those States. An average of 825 cases weekly was reported during the month of February. The rate for the month was considerably below that for the corresponding period in 1928 and closely approximated that for 1927.

Influenza.—The attack rate of influenza continued to decline during the month of February, but was still higher than during the corresponding period in either of the two preceding years. For the four weeks ended March 2 the number of cases totaled 32,085.

Scarlet fever.—On the average approximately 4,500 cases of scarlet fever were reported weekly during the month of February. The rate was about the same as that for the corresponding month of 1928 and was slightly lower than that for 1927. As the lowest incidence of scarlet fever does not usually occur until late in the summer, the present level may continue for several weeks. For the four weeks ended March 2 the number of cases reported totaled 18,198.

Measles.—The expected seasonal increase in measles took place in all sections of the country during February. The incidence was, however, relatively low as compared with the two preceding years. For the four weeks ended March 2 the number of cases totaled 33,340,

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as against approximately 63,000 and 55,000 for the corresponding month in 1928 and 1927, respectively.

Diphtheria.—The diphtheria rate showed little change during February. For the four weeks ended March 2 the number of cases totaled 5,669, as compared with 7,624 in 1928 and 7,499 in 1927 during the same month.

Mortality from all causes.—The mortality rate from all causes in large cities as shown by the Weekly Health Index of the Bureau of the Census continued to decline through the month of February. A sharp rise occurred during the week ended March 2, which appeared to be due in part to an excess in the deaths from influenza and pneumonia in cities in the East and West South Central and the Mountain States. Later reports indicated that the rate had dropped again; and for the week ended March 16 the rate (14.7) was practically the same as that for the corresponding week in 1928 and but little above the rate in 1927.

COURT DECISIONS RELATING TO PUBLIC HEALTH

Pneumonoconiosis held to be a personal injury and compensable under compensation act.—(Massachusetts Supreme Judicial Court; Sullivan's Case, 164 N. E. 457; decided January 4, 1929.) In a proceeding under the workmen's compensation act compensation was claimed by an employee whose work was on granite or other stone with a surface-cutting machine. The employee was suffering from pneumonoconiosis, and the question for decision was whether he had received a personal injury arising out of and in the course of his employment or whether he was merely suffering from disease so induced. The supreme court affirmed an award of compensation, taking the view that the employee had sustained a personal injury. In deciding the matter the court said:

* * * It is settled that simple disease resulting from employment affords no ground for recovery under our workmen's compensation act. * * * In certain oircumstances, however, disease may also be a personal injury within the meaning of the act. * * * In the case at bar there was evidence tending to show the tangible impact of particles of granite upon the lungs of the employee producing definite damage to his body. The "personal injury," for which alone compensation is payable under G. L. c. 152, sec. 26, might have been found to be due to physical deterioration flowing immediately from corporeal collision with a foreign substance set in motion by the business of the employer performed by the employee by virtue of his contract for service. It might have been found to be as tangible as a broken bone. Although this result may be termed "granite cutters' disease," that factor is not decisive that it was a disease rather than a personal injury under the workmen's compensation act. Its nature in that respect must be ascertained from an analysis of the physical condition which in truth resulted and not from mere nomenclature. The personal injury may be none the less the direct and consequential result of the employment, although a condition may arise termed in some connections a disease.

Gas appliance ordinances held valid.—(Maryland Court of Appeals; Portsmouth Stove and Range Co. v. Mayor and City Council of Baltimore et al., 144 A. 357; decided January 16, 1929.) Suit was brought by the plaintiff company to restrain the mayor and city council of Baltimore and the city commissioner of health from enforcing certain city ordinances pertaining to the testing, registration, sale, and installation of tubing and appliances for use with manufactured (illuminating) gas. Relevant portions of the ordinances are summarized as follows:

The sale, connection, or installation in the city of any tubing, appliance, appurtenance, or device for use with, by, or for the combustion of manufactured (illuminating) gas was forbidden unless the type, sample, or model thereof was registered with the health commissioner. Application forms were provided for and the applicant was required to furnish such information and certificates and cause such tests to he conducted as might be required by the commissioner to secure proper registration and identification of any tubing, etc. Before approving such application the commissioner was required to determine whether the tubing, etc., conformed to the rules and specifications which he was authorized to adopt. Whenever the commissioner determined that, before passing upon any registration application, it was "necessary and/or desirable to test in a laboratory or testing agency any tubing," etc., to determine whether there was compliance with prescribed specifications, he was authorized to direct the tests to be made by the laboratories of either the United States Bureau of Standards, the Johns Hopkins University, the American Gas Association, or any other laboratory or agency approved by him, and he was further authorized and directed to require the applicant to pay the costs. The commissioner in adopting the authorized rules and specifications could consider those established by the American Gas Association or any other rules, etc., relating to the subject. and could accept the certificate of any laboratory or testing agency approved by him in determining whether any tubing, etc., required by him to be tested met the prescribed tests.

The rules and specifications promulgated by the commissioner of health under the authority of the ordinances were, with certain omissions, those of the American Gas Association.

The plaintiff filed letters from the Bureau of Standards and Johns Hopkins University, the bureau stating that it did not make such tests as the ordinances required and the university stating that it had only made such tests for the accommodation of its friends or its own information, although it did not decline to entertain a proposition to do such work and in fact invited the plaintiff to confer and expressed its willingness to do what it could in the matter. With this the plaintiff seemed to eliminate as unavailable all except the American Gas Association, which had a laboratory at Cleveland, Ohio, equipped to make the tests contemplated. The plaintiff was not a member of such association, and nonmembers were asked to pay 50 per cent more than members for the services rendered, but it was shown that members were charged less because of having contributed to the laboratory through their dues.

The court of appeals decided in favor of the city, declaring the ordinances to be within the range of the police power and valid. It held that the ordinances did not unlawfully delegate authority to unofficial agencies, the plaintiff having contended that said ordinances gave the power of decision to the testing agency. Concerning this the court said: "The testing agency does not decide anything; it ascertains as a fact whether the sample appliance or device is within the specifications prescribed by the commissioner of health. It is a matter of 'measurement and arithmetical calculation' * * * which some one with the requisite facilities must of necessity do.".

It was also held that the plaintiff was not denied the equal protection of the laws because nonmembers of the gas association paid more for tests than did members, the opinion stating that "The facilities and equipment have been provided by the association, presumably from the membership, and the 50 per cent additional fees demanded of nonmembers is imposed for the purpose of equalizing the cost of tests between them and the membership of the association, and not for the purpose of discriminating against nonmembers."

With regard to the plaintiff's complaint that the ordinances "may be the instruments of oppression, and under their authority the commissioner of health may arbitrarily require tests to be made at such cost as to drive the appellant out of the Baltimore market," the court said that this was only a fear and apparently not well founded, and then went on to say that "If, in the exercise of his discretion, the commissioner acts 'arbitrarily, or unreasonably, or exceed[s] the power conferred by the ordinance, those injured thereby can obtain relief from the courts.""

DEATHS DURING WEEK ENDED MARCH 23, 1929

Summary of information received by telegraph from industrial insurance companies for the week ended March 23, 1929, and corresponding week of 1928. (From the Weekly Health Index March 27, 1929, issued by the Bureau of the Census, Department of Commerce)

| | Week ended Mar. 23, 1929 | | Corresponding week, 1928 |
|--|-----------------------------|------------|-----------------------------|
| Policies in force. | 73, 638, 229 | . . | 70, 702, 349 |
| Number of death claims | 16, 659 | - | 14, 519 |
| Death claims per 1,000 policies in force, annual rate_ | 11. 8 | | 10. 7 |

Deaths from all causes in certain large cities of the United States during the week ended March 23, 1929, infant mortality, annual death rate, and comparison with corresponding week of 1928. (From the Weekly Health Index, March 27, 1929, issued by the Bureau of the Census, Department of Commerce)

| | Week en 23, | nded Mar. 1929 | Annual death rate per | Deaths ye | Infant mortality | | |
|-------------------|--|---|--|--|---|--|--|
| City | Total deaths | Death rate ¹ | 1,000, corre- sponding week, 1928 | Week ended Mar. 23, 1929 | Corre- sponding week, 1928 | rate, week ended Mar. 23, 1929 ? | |
| Total (62 cities) | 7, 928 | 14.2 | 14.7 | 756 | 854 | 3 67 | |
| Total (62 cities) | 7, 928 500 366 377 38 39 2255 168 57 73 288 57 738 202 255 168 57 73 288 57 738 202 202 202 202 202 202 202 20 | 14. 2 15. 6 15. 8 (*) 14. 2 (*) 17. 2 (*) 18. 5 14. 0 13. 7 14. 7 16. 8 (*) 16. 8 (*) 16. 9 (*) 16. 2 (*) 11. 9 24. 5 (*) 11. 2 21. 0 (*) 14. 2 14. 2 21. 0 | 14. 7 16. 5 15. 6 (*) 17. 3 (*) 14. 9 14. 9 14. 4 14. 7 14. 9 14. 4 14. 7 14. 9 12. 1 12. 1 12. 1 12. 7 13. 1 12. 2 (*) 9. 9 17. 6 12. 7 13. 1 12. 1 13. 1 12. 8 10. 9 11. 6 (*) 15. 1 (*) 15. 1 (*) 15. 1 (*) 15. 1 (*) 15. 1 (*) 15. 2 (*) 15. 2 (*) 16. 6 12. 7 13. 1 12. 1 13. 2 19. 0 (*) 15. 1 (*) 15. 1 (*) 15. 1 (*) 15. 2 19. 0 (*) 15. 1 (*) 16. 6 12. 7 13. 1 12. 2 (*) 10. 5 (*) 15. 1 (*) 15. 1 (*) 15. 1 (*) 16. 6 12. 4 (*) 16. 6 12. 4 (*) 16. 6 12. 4 (*) 16. 6 12. 2 17. 8 20. 9 (*) 16. 6 12. 2 17. 8 20. 9 (*) 16. 6 12. 2 17. 8 20. 9 (*) 16. 6 12. 2 17. 8 20. 9 (*) 16. 6 12. 2 (*) 16. 6 12. 2 17. 8 20. 9 (*) 16. 6 12. 2 17. 8 20. 9 (*) 16. 6 12. 2 17. 8 20. 9 (*) 16. 6 12. 2 (*) 16. 6 12. 2 (*) 16. 6 12. 2 (*) 16. 6 12. 2 (*) 16. 6 12. 2 (*) 16. 6 12. 2 20. 2 (*) 16. 7 10. 7 | 758 5 4 7 7 5 2 12 8 4 6 4 2 2 5 6 12 1 1 1 9 6 3 3 4 3 2 5 4 7 3 3 0 1 9 8 1 4 4 0 7 6 4 2 2 3 3 0 27 8 6 2 1 2 9 3 6 0 14 8 6 2 2 2 3 1 2 9 3 6 0 14 8 6 2 2 2 3 1 2 9 3 6 0 14 8 6 2 2 2 3 1 2 9 3 6 0 14 8 6 14 14 14 14 14 14 14 14 14 14 14 14 14 | 854 5 4 8 5 3 255 150 7 5 2 306 19 2 7 767 19 6 6 5 1 0 6 3 72 1 4 0 7 3 1 2 4 6 4 2 4 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 4 4 9 0 0 0 0 25 8 7 1 2 1 6 2 4 22 6 9 5 4 3 4 0 10 8 10 10 10 10 10 10 10 10 10 10 10 10 10 | • 67 522 79 73 38 322 38 323 38 323 38 323 324 63 54 60 60 46 69 104 52 104 52 1143 58 54 69 69 69 69 69 69 69 69 69 69 69 69 69 69 69 68 102 377 0 73 377 0 737 737 737 737 | |
| New Orleans | 151 88 63 | 18.4 (³) | (³) | 13 5 8 | 10 2 8 | 65 35 135 | |

Footnotes at end of table.

Deaths from all causes in certain large cities of the United States during the week ended March 23, 1929, infant mortality, annual death rate, and comparison with corresponding week of 1928-Continued

| · Wash an | | 1 | 1 | | 1 |
|---|---|---|---|---|--|
| | død Mar. 1929 | Annual death | Deaths y | Infant mortality | |
| City Total deaths | Death rate 1 | 1,000, corre- sponding week, 1928 | Week ended Mar. 23, 1929 | Corre- sponding week, 1928 | rate, week ended Mar. 23, 1929 ² |
| New York 1,619 Bronk Borough 572 Manhattan Borough 652 Queeas Borough 652 Richmond Borough 36 Richmond Borough 74 Oklahoma City 38 Omaha 60 Paterson 41 Portland, Greg 66 Providence 85 Richmond 64 Phitsburgh 542 Pittsburgh 542 Portland, Greg 66 Providence 85 Richmond 64 White 216 Colored 85 Richmond 64 St Paul 57 St Lavis 281 St Paul 57 Salt Lake City 4 37 San Antonio 53 Sen Diego 50 Bokane 32 Byringfield, Mass 41 Sordend 79 Tracoma 71 Toledo 79 Tracored 57 < | 14, 1 11, 0 13, 0 13, 0 13, 0 14, 1 15, 0 14, 1 12, 1 14, 1 12, 1 14, 3 15, 7 14, 3 15, 7 14, 3 15, 7 15, 0 10, 3 14, 3 15, 6 10, 3 11, 7 11, 7 11, 7 </td <td>15.2 12.8 12.8 12.8 12.5 10.0 15.3 12.8 12.6 12.6 12.6 12.6 14.2 14.4 15.7 14.9 12.1 20.4 15.7 14.9 12.1 20.4 15.7 14.9 12.2 14.5 15.8 1</td> <td>168 157 768 13 1 11 4 3 7 4 54 20 2 7 8 2 6 3 3 1 5 6 0 7 4 9 7 2 1 3 9 6 2 4 6 2 7 8 2 6 3 3 1 5 5 7 7 4 9 7 2 1 5 7 7 8 2 6 8 7 8 7 8 7 8 7 8 7 8 7 8 7 8 8 7 8 7</td> <td>195 208 881 21 5 11 4 3 3 4 66 30 2 7 9 3 6 8 19 7 7 16 6 7 2 2 3 2 5 0 6 9 14 6 8 2 4 6 2 5</td> <td>69 44 72 83 53 18 58 44 72 76 771 76 79 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 71 71 71 71 71 71 71 71 71 71 71 71 71 71 71 71</td> | 15.2 12.8 12.8 12.8 12.5 10.0 15.3 12.8 12.6 12.6 12.6 12.6 14.2 14.4 15.7 14.9 12.1 20.4 15.7 14.9 12.1 20.4 15.7 14.9 12.2 14.5 15.8 1 | 168 157 768 13 1 11 4 3 7 4 54 20 2 7 8 2 6 3 3 1 5 6 0 7 4 9 7 2 1 3 9 6 2 4 6 2 7 8 2 6 3 3 1 5 5 7 7 4 9 7 2 1 5 7 7 8 2 6 8 7 8 7 8 7 8 7 8 7 8 7 8 7 8 8 7 8 7 | 195 208 881 21 5 11 4 3 3 4 66 30 2 7 9 3 6 8 19 7 7 16 6 7 2 2 3 2 5 0 6 9 14 6 8 2 4 6 2 5 | 69 44 72 83 53 18 58 44 72 76 771 76 79 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 76 71 71 71 71 71 71 71 71 71 71 71 71 71 71 71 71 71 |

¹ Annual rate per 1,000 population. ² Deaths under 1 year per 1,000 births. Cities left blank are not in the registration area for births.

² Deaths under 1 year per 1,000 births. Cities left blank are not in the registration area for births.
⁴ Deaths for week ended Friday.
⁴ Deaths for week ended Friday.
⁴ In the cities for which deaths are shown by color the colored population in 1920 constituted the following percentages of the total population: Atlanta, 31; Baltimore, 15; Birmingham, 39; Dallas, 15; Fort. Worth, 14; Houston, 25; Indianapolis, 11; Kansas City, Kans., 14; Knozville, 15; Louisville, 17; Memphis, 38; Nashville, 36; New Orleans, 26; Richmond, 32; and Washington, D. C., 25.

PREVALENCE OF DISEASE

No health department, State or local, can effectively prevent or control disease without knowledge of when, where, and under what conditions cases are occurring

UNITED STATES

CURRENT WEEKLY STATE REPORTS

These reports are preliminary and the figures are subject to change when later returns are received by the State health officers

Reports for Weeks Ended March 23, 1929, and March 24, 1928

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended March 23, 1929, and March 24, 1928

| | Diph | theria | Influ | 10nza | Me | asles | Meningococcus meningitis | |
|---|--|--|--------------------------------------|-----------------------------------|---|--|--------------------------------------|--------------------------------------|
| Division and State | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 |
| New England States: Maine | 6 3 3 84 11 28 | 5 2 86 7 23 | 26 14 65 6 18 | 7 11 7 | 251 36 7 387 90 515 | 33 16 98 1, 809 91 301 | 0 0 5 0 2 | 0 0 2 0 1 |
| Middle Atlantic States: New York New Jersey Pennsylvania Fast North Central States: | 310 107 189 | 390 149 197 | ¹ 52 22 | ¹ 57 41 | 1, 214 246 2, 293 | 2, 548 1, 307 1, 345 | 42 6 7 | 33 3 4 |
| Ohio Indiana Illinois Michigan Wisconsin | 29 30 157 88 17 | 68 25 192 51 23 | 43 220 20 50 | 22 65 251 78 | 1, 263 485 1, 637. 417 1, 142 | 1, 071 280 232 1, 536 106 | 5 0 14 38 8 | 5 0 17 3 6 |
| West North Central States: Minnesota Missouri North Dakota South Dakota Nebraska Kanses South Lientia States | 27 11 52 8 3 16 9 | 27 11 52 3 7 6 8 | 1 14 1 26 | 5 167 45 90 50 | 752 57 384 81 16 49 290 | 82 36 339 29 69 87 | 3 5 36 6 2 2 9 | 2 2 9 3 0 1 |
| Delware. Maryland ³ District of Columbia West Virginia. North Carolina. South Carolina. Georgia. Florida. | 1 20 11 9 22 17 7 6 | 1 22 21 23 111 21 12 10 | 3 98 5 64 623 55 4 | 50 11 50 859 154 1 | 32 83 24 249 72 10 25 26 | 21 1, 163 182 120 3, 290 950 105 70 | 0 1 0 3 5 0 1 0 | 0 0 1 2 0 0 0 0 |

¹ New York City only.

² Week ended Friday.

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Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended March 23, 1929, and March 24, 1928—Continued

| | Diph | theria | Influenza | | Measles | | Meningococcus meningitis | |
|---|---|---|---|--|--|--|---------------------------------------|--------------------------------------|
| Division and State | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 |
| East South Central States: Kentucky Tennessee | | 3 | 95 | 16 211 | 54 6 | 240 281 | 05 | .0 0 |
| Alabama Mississippi West South Central States: | 26 3 | 31 15 | 111 | 254 | 114 | 586 | 7 | 3 |
| Arkansas. Louisiana. Oklahoma ¹ . Teras. | 16 21 13 21 | 2 11 18 42 | 148 99 153 64 | 377 157 489 52 | 56 85 24 150 | 506 374 287 491 | 6 2 5 2 | 3 1 3 6 |
| Montani States. Montana. Idaho Wyorning. Colorado | 4 3 2 10 | 2 | 6 4 4 | | 130 4 22 11 | 1 45 25 | 6 10 0 6 | 5 1 2 14 |
| New Mexico Arizona Utah ¹ Pacific Statas: | 4 | 6 8 4 | 9 1 4 | 9 7 | 2 4 | 142 27 11 | 3 2 14 | 0 5 7 |
| Washington Oregon California | 7 8 58 | 7 8 102 | 107 102 | 49 82 | 105 194 57 | 278 125 234 | 34 2 18 | 2 7 4 |
| | Poliomyelitis | | Scarlet fever | | Smallpox | | Typhoid fever | |
| Division and State | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 |
| New England States: Maine | 0 | 0 | 56 | 26 | 2 | 0 | 9 | 0 |
| New Hampshire | 0 0 0 0 | 0 0 1 0 0 | 26 20 396 28 74 | 9 17 317 46 55 | 0 11 0 0 9 | 0 0 0 0 0 | 0 0 4 9 1 | 0 0 2 0 1 |
| New York New Jersey Pennsylvania | 1 1 2 | 9 1 8 | 708 233 531 | 943 317 603 | 5 0 | 5 1 0 | 11 4 17 | 16 6 12 |
| East North Central States: Ohio Indiana Michigan Wisconsin Wisconsin | 2 0 2 1 0 | 2 0 2 1 2 | 206 230 522 368 263 | 232 150 345 277 209 | 47 60 157 84 5 | 57 228 58 86 12 | 16 6 4 5 4 | 4 1 13 4 1 |
| Minbesota Iowa Missouri North Dakota North Dakota Nebraska | 0 1 0 0 0 | 0 2 0 3 0 | 103 182 87 47 31 129 | 145 79 105 79 29 100 | 0 60 40 6 15 71 | 0 33 71 0 10 53 | 8 3 7 0 1 1 | 3 0 1 2 0 0 |
| kansas. South Atlantic States: Delaware | 0 0 0 0 0 0 0 0 0 | 1 0 1 5 0 0 0 0 0 | 223 2 105 24 16 33 13 20 10 | 196 3 72 44 54 23 4 27 6 | 79 0 0 24 14 3 15 0 | 84 0 2 67 121 7 0 2 | 0 2 0 21 6 6 8 3 | 0 5 1 3 2 3 7 |

² Week ended Friday.

* Figures for 1929 are exclusive of Oklahoma City and Tulsa.
Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended March 23, 1929, and March 24, 1928—Continued

| | Poliom yelitis | | Scarlet fever | | Smallpox | | Typhoid fever | |
|----------------------------|--------------------------------------|--------------------------------------|--------------------------------------|--------------------------------------|--------------------------------------|--------------------------------------|---|--------------------------------------|
| Division and State | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 | Week ended Mar. 23, 1929 | Week ended Mar. 24, 1928 |
| Fast South Central States | | | | | | | | |
| Kanducky | ^ | 2 | 112 | 34 | 43 | 10 | 8 | 2 |
| Tennessee | ŏ | l ñ | 52 | 17 | 10 | 20 | | |
| Alabama | ĭ | lĭ | 11 | 10 | Ğ | 10 | 17 | 19 |
| Mississippi | ō | l ō | 7 | 16 | ň | 3 | 2 | 3 |
| West South Central States: | • | Ů | 1 1 | | - | , i | - | |
| Arkenses | | 0 | 9 | 26 | 6 | 11 | 3 | 4 |
| Louisiana | ð | Ŏ | 53 | 15 | 5 | 34 | 7 | Ā |
| Oklahoma ¹ | Õ | i | 20 | 80 | 99 | 249 | 3 | 3 |
| Texas | 1 | Ō | 71 | 144 | 85 | 48 | 3 | 4 |
| Mountain States: | - | - | | | | | - | - |
| Montana | 0 | 0 | 21 | 6 | 4 | 16 | 0 | 0 |
| Idaho. | Ó | 0 | 2 | 10 | 11 | 1 | 0 | Ó |
| Wyoming | 1 | 0 | 7 | 14 | 2 | 3 | Ó | Ó |
| Celerado | 0 | 0 | 20 | 81 | 26 | 2 | 1 | 0 |
| New Mexico | Ó | Ó | 27 | 31 | 0 | 1 | 3 | 1 |
| Arizona | Ó | 0 | | 5 | 6 | 20 | 1 | 0 |
| Utah 1 | 0 | 0 | 6 | 6 | 2 | 14 | 1 | 0 |
| Pacifie States: | | | | | | | | |
| Washington | 1 | 0 | 20 | 37 | 52 | 49 | 5 | 2 |
| Ore yon | 1 | 1 | 53 | 17 | 36 | 77 | 1 | 1 |
| California | 8 | 2 | 473 | 153 | 37 | 29 | 13 | 5 |

² Week ended Friday.

³ Figures for 1929 are exclusive of Oklahoma City and Tulsa.

SUMMARY OF MONTHLY REPORTS FROM STATES

The following summary of monthly State reports is published weekly and covers only those States from which reports are received during the current week:

| State | Menin- gococ- cus menin- gitis | Diph- theria | Influ- enze | Ma- laria | Mea- sles | Pel- lagra | Polio- mye- litis | Scarlet fever | Small- pox | Ty- phoid fever |
|-----------------------|--|-----------------|----------------|--------------|--------------|---------------|-------------------------|------------------|---------------|-----------------------|
| January, 1929 | | | | | | | | | | |
| Florida | 3 | 49 | 3, 452 | 14 | 26 | 3 | 0 | 49 | 7 | 8 |
| February, 1989 | | | | | | | | | | |
| Illinois | 54 | 536 | 1, 091 | 1 | 2, 791 | | 1 | 1, 792 | 486 | 18 |
| Indiana | 4 | 110 | 1 019 | 21 | 1,203 | 10 | 1 | 301 | 220 | 25 |
| Moryland | 44 | 00 | 3 404 | 31 | 302 434 | 19 | 1 | 279 | 25 | 30 |
| Michigan | v | 217 | 235 | 1 - | 1.406 | 1 | 3 | 1.427 | 146 | 12 |
| Minnesota | 9 | 101 | 6 | | 1, 598 | - | ŏ | 679 | 13 | 14 |
| Mississippi | 2 | 52 | 11.050 | 1.950 | 2.377 | 364 | 3 | 67 | 6 | 35 |
| Missouri | 67 | 215 | 980 | 7 | 1, 123 | 1 | 0 | 428 | 198 | 7 |
| New York | 148 | 904 | | 1 4 | 3, 549 | | 10 | 1, 971 | 0 | 44 |
| North Carolina | 4 | 134 | | | 343 | | 2 | 202 | 66 | 6 |
| Ohio | 25 | 265 | 784 |] | 4, 163 | | 1 | 1, 224 | 178 | 27 |
| Oklahoma ¹ | 26 | 79 | 2, 274 | 42 | 19 | 4 | . 3 | 131 | 185 | 16 |
| Oregon | 7 | 48 | 346 | | 463 | | 1 | 185 | 188 | 3 |
| Pennsylvania | 46 | 570 | | | 7,239 | 1 | 6 | 1, 737 | 1 | 41 |
| Knode Island | 2 | 54 | 177 | | 383 | | U | 14/ | | 3 |
| west virginia | 6 | 63 | 1, 743 | | 558 | | 0 | 108 | 7.6 | 20 |
| | |) | 1 | 1 | | | | | 1 1 | |

¹ Exclusive of Oklahoma City and Tulsa.

| January, 19 1 9 | Cases |
|--|------------|
| Florida: | |
| Unicken pox | . 67 |
| Mumpe | . 26 13 |
| Rabies in man | 10 |
| Typhus fever | i |
| Whooping cough | 65 |
| | |
| February, 1929 | |
| Illinois | 1 |
| Anthrax: | - |
| Pennsylvania | 2 |
| Chicken pox: | |
| Illinois | 1,042 |
| Indiana | 406 |
| Louisiana | 94 |
| Mishigan | 100 764 |
| Minnesota | 576 |
| Mississippi | 917 |
| Missouri | 314 |
| New York | 2, 100 |
| North Carolina | 768 |
| Ohio | 1, 151 |
| Oklahoma ¹ | 97 |
| Uregon | 212 |
| Phode Island | 2,350 |
| West Virginia | 143 |
| Dengue: | |
| Mississippi | 1 |
| Dysentery: | |
| Illinois | 14 |
| Indiana (amebic) | 1 |
| Mississippi (amebic) | 45 |
| Mississippi (Dacillary) | 221 |
| Oklehome I | 3 |
| Pennsvivania | 1 |
| German measles: | - |
| Illinois | 89 |
| Maryland | 18 |
| New York | 336 |
| North Carolina | 73 |
| | .9 |
| reinsylvania Rhede Island | 75 |
| Hookworm disease: | 1 |
| Louisiana | 2 |
| Mississippi | 267 |
| Impetigo contagiosa: | |
| Maryland | 3 |
| Oregon | 10 |
| Lead poisoning: | |
| 11111018 Obio | 14 |
| Inthargic encenhelitis: | " |
| Illinois | 7 |
| Louisiana | il |
| Michigan | 13 |
| . Minnesota | 1 |
| New York | 23 |
| Ohio | 12 |
| Fennsylvania | 13 |
| ¹ Exclusive of Oklahoma City and Tulsa. | |

| Mumps: | Cases |
|-----------------------|-------------|
| Illinois | 446 |
| Indiana | 67 |
| Louisiana. | 3 |
| Maryland. | 420 |
| Michigan | 450 |
| Mississippi | |
| Missouri | . 910 |
| New York | - 610 |
| Ohio | |
| Oklahoma I | 380 |
| Oragon | - 30 |
| Dennevizenie | 1 040 |
| Phode Island | 1, 943 |
| Onbthelmie neonetenum | 3 |
| Dinein | |
| Tadiana | 38 |
| Monuland | 1 |
| | 1 |
| Mississippi | 14 |
| IVIISSOUFI | 1 |
| New YORK | 1 |
| North Carolina | . 1 |
| | - 78 |
| Uklahoma ¹ | - 2 |
| rennsylvania | 8 |
| Khode Island | . 1 |
| Paratyphoid fever: | |
| Illinois | - 2 |
| Puerneral senticemia: | |
| Illinois | 19 |
| Mississinni | - 10 |
| New York | - 01 |
| Ohio | - 10 |
| Pennevivonie | - 19 |
| | - 18 |
| Rabies in animals: | |
| Illinois | - 7 |
| Maryland | - 8 |
| Minnesota | - 1 |
| Mississippi | - 6 |
| Missouri | . 14 |
| New York | - 20 |
| Oregon | . 1 |
| Khode Island | . 14 |
| Rabies in man: | |
| Michigan | . 2 |
| Scables: | - |
| Marvland | 2 |
| Oregon | . J 10 |
| Contin and the state | . 16 |
| ceptic sore throat: | |
| Liinois | . 18 |
| Maryland | . 3 |
| Michigan | . 23 |
| Missouri | . 29 |
| New York | , 12 |
| North Carolina | . 5 |
| Unio | . 62 |
| Uklahoma ¹ | . 30 |
| Oregon | . 9 |
| Tetanus: | |
| Louisiana | 2 |
| Maryland | 1 |
| New York | 4 |
| Ohio | 2 |
| Pennsylvania | 4 |
| | |

| Trachoma: | Cases | Vincent's.angina: | Cases |
|------------------|-------|-----------------------|--------|
| Illinois | 16 | Maryland | . 9 |
| Maryland | . 1 | New York | . 67 |
| Minnesota | . 3 | Oklahoma ¹ | . 2 |
| Mississippi | 9 | Whooping cough: | |
| Missourt. | 2 | Illinois | . 567 |
| New York | 1 | Indiana | 267 |
| Ohio | I | Louisiana | . 13 |
| Oktaboma 1 | 5 | Maryland | 382 |
| Trench mouth: | | Michigan | 883 |
| Maryland | 2 | Minnesota | 324 |
| Trichinosis: | | Mississippi | 857 |
| Ohio | 4 | Missouri | 259 |
| Tularacmia: | | New York | 1, 197 |
| Maryland | 1 | North Carolina | 735 |
| Ohio | L | Obio | 1,658 |
| Typhus fever: | | Oklahoma ¹ | 78 |
| Louisiana | 1 | Oregon | 24 |
| Underiant fever: | | Pennsylvania | 1, 653 |
| Maryland | 1 | Rhode Island | 24 |
| Minnesota | 1 | West Virginia | 125 |
| New York | 3 | | |
| 0hio | n l | | |
| Pennsylvania | 1 I | | |

GENERAL CURRENT SUMMARY AND WEEKLY REPORTS FROM CITIES

The 98 cities reporting cases used in the following table are situated in all parts of the country and have an estimated aggregate population of more than 31,565,000. The estimated population of the 91 cities reporting deaths is more than 29,995,000. The estimated expectancy is based on the experience of the last nine years, excluding epidemics.

| · · · | 1929 | 1928 | Estimated expectancy |
|---------------------------|--------|--------|-------------------------|
| Cases reported | | · . | |
| LADDINGTIS: | 1 461 | 1 907: | |
| 40 States | 767 | 950 | 929 |
| Measles: | f | | |
| 44 States | 12.742 | 19,854 | |
| 98 cities | 4,124 | 8,051 | |
| Meningecoccus meningitis: | - | | |
| 45 States | 325 | 155 | |
| 98 cities | 152 | 75 | |
| Foliomyelitis: | | 07 | |
| 46 STates | 14 | 41 | |
| A States | 5 047 | 5 017 | |
| 40 Diales | 1,971 | 1,786 | 1, 553 |
| Smallnor. | -,,,,, | 2,100 | 1,000 |
| 46 States | 1, 058 | 1, 424 | |
| 98 cities | 75 | 127 | 98 |
| Typhoid fever: | | | |
| 46 States | 142 | 161 | |
| 98 cities | 28 | . 27 | 31 |
| Beaths reported | | | |
| Transa and pressmonia. | | | |
| of office | 1, 256 | 1, 494 | |
| 21 CH463 | 1, 200 | 1, 101 | |
| 91 cities | 0 | 1 | |
| Sacramento, Calif | ŏ | ī | |
| | - | | |

Weeks ended March 16, 1929, and March 17, 1928

1 Exclusive of Oklahoma City and Tulsa.

City reports for week ended March 16, 1929

The "estimated expectancy" given for diphtheria, poliomyelitis, scarlet fever, smallpox, and typhoid fever is the result of an attempt to ascertain from previous occurrence the number of cases of the disease under consideration that may be expected to occur during a certain week in the absence of epidemics. It is based on reports to the Public Health Service during the past nine years. It is in most instances the median number of cases reported in the corresponding weeks of the preceding years. When the reports include several epidemics, or when for other reasons the median is unsatisfactory, the epidemic periods are excluded and the estimated expectancy is the mean number of cases reported for the week during nonepidemic years.

If the reports have not been received for the full nine years, data are used for as many years as possible, but no year earlier than 1920 is included. In obtaining the estimated expectancy the figures are smoothed when necessary to avoid abrupt deviation from the usual trend. For some of the diseases given in the table the available data were not sufficient to make it practicable to compute the estimated expectancy.

| | | Chiek | Diph | theria | Infit | 1enza | | | | |
|------------------------------|---|---|---|------------------------|------------------------|-------------------------|---|----------------------------------|--|--|
| Division, State, and city | Population July 1, 1928, estimated | Chick- en pox, cases re- ported | Cases, esti- mated expect- ancy | Cases re- ported | Cases re- ported | Deaths re- ported | Mea- sles, cases re- ported | Mumps, cases re- ported | Pneu- monia, deaths re- ported | |
| NEW ENGLAND | | | | | | | | | | |
| Maine: | | | | | | | | | | |
| New Hampshire: | 78,000 | 2 | 1 | 1 | | | 40 | U | 1 | |
| Concord Manchester | (1) 85, 700 | 0 | · 0 1 | 1 0 | | | | 0 | | |
| Vermont: | a) | | 0 | 0 | | | | | 0 | |
| Massachusetts: | | | | | | | | 0 | U | |
| Fall River | 799, 200 134, 300 | 47. | 40 3 | 25 | 17 | 5 | - 30 - 19 | 15 0 | 35 | |
| Springfield | 149,800 | 3 14 | 3 | 3 | | 0 | 23 15 | 3 | 17 | |
| Rhode Island: | 101,000 | | - | | | | 10 | - | | |
| Providence | 286, 300 | Ō | 9 | 10 | 2 | 2 | 15 51 | 0 | 13 | |
| Connecticut: Bridgeport | (I) | 2 | 6 | 5 | 5 | 1 | 24 | 0 | 9 | |
| Hartford | 172, 300 | 2 | 8 | 6 | | 1 | 48 | 7 | 6 | |
| New Haven | 107, 900 | | - | - | 4 | - | 10 | U | y | |
| MIDDLE ATLANTIC | | | | | | | | | | |
| New York: | 555 800 | 15 | 13 | 17 | | | 14 | | 91 | |
| New York. | 6, 017, 500 | 348 | 236 | 211 | 78 | 35 | 58 | ō | 226 | |
| Rochester | 328, 200 199, 300 | 8 57 | 10 | 6 | 8 | 0 1 | 40 3 | 22 5 | 11 12 | |
| New Jersey: | 125 400 | , | 5 | 7 | 9 | | | | | |
| Newark | 473, 600 | 37 | 16 | 31 | 11 | i | 14 | 81 | 20 | |
| Trenton Pennsylvania: | 139, 000 | 1 | 4 | 2 | | 0 | 1 | 0 | 1 | |
| Philadelphia | 2,064,200 | 186 | 70 | 33 | 30 | 19 | 60 | 13 | 72 | |
| Reading | 673, 800 115, 400 | 5/ 8 | 20 | 17 | | Ő | 70 | 11 | 38 1 | |
| BAST NORTH CENTRAL | | | | | | | | | | |
| Ohio: | | | | | 1 | | | | / | |
| Cincinnati | 413,700 | 17 | 10 | 10 | | 4 | 528 | 0 | 19 27 | |
| Columbus | 299,000 | 18 | 4 | 3 | 4 | 4 | 29 | Ő | 6 | |
| Indiana: | 313, 200 | 14 | 5 | 4 | 2 | 2 | 15 | 20 | 7 | |
| Fort Wayne | 105, 300 | 7 | 2 | 3 | | 9 | 40 | 0 | 4 | |
| South Bend | 86, 100 | 1 | ĩ | 2 | | õ | 60 | ŏ | 4 | |
| Terre Haute | 73, 500 | 2 | 1 | 3 | | 1 | 4 | 0 | 1 | |
| Chicago | 3, 157, 400 | 107 | 77 | 101 | 20 | ıı | 463 | 13 | 82 | |
| Michigan: | 07, 200 | Ø | 1 | U | 1 | 1 | 4 | - | 3 | |
| Detroit | 1, 378, 900 | 98 15 | 53 | 43 | 17 | 8 | 55 | 16 | 58 4 | |
| Grand Rapids | 164, 200 | 2 | 2 | il | | ŏl | 216 | ŏ | 3 | |

¹ No estimate of population made.

City reports for week ended March 16, 1929-Continued

| | | | Diph | theria | Influ | lenza | | | |
|--|---|-----------------------------------|---|------------------------|------------------------|-------------------------|--|----------------------------------|--|
| Division, State, and city | Population July 1, 1928, estimated | en pox, cases re- ported | Cases, esti- mated expect- ancy | Cases re- ported | Cases re- ported | Deaths re- ported | Mea- slies, cases re- ported | Mumps, cases re- ported | Pneu- monia, deaths re- ported |
| EAST NORTH CENTRAL- continued | | | | | | | | | |
| Wisconsin: Kenosha Milwaukee Racine Superior | 56, 500 544, 200 74, 400 (1) | 12 100 16 5 | 1 18 2 0 | 0 6 1 0 | | 0 0 0 | 33 487 157 0 | 0 7 9 5 | 0 14 1 1 |
| WEST NORTH CENTRAL | - | | | | | | | | |
| Minnesota: Duluth Minneapolis St. Paul | 116, 800 455, 900 (¹) | 8 86 22 | 0 14 10 | 0 18 3 | | 3 2 1 | 0 322 229 | 115 101 21 | 4 9 5 |
| Iowa: Davenport Des Moines Sioux City Waterloo | (1) 151, 900 80, 000 37, 100 | 1 0 8 1 | 1 2 1 0 | 000 | | | 3 2 • 0 4 | 0 0 1 65 | |
| Missouri: Kansas City St. Joseph St. Louis | 391, 000 78, 500 848, 100 | 28 2 22 | 6 0 43 | 3 0 49 | | 2 1 | 410 7 20 | 9 0 13 | 24 0 |
| North Dakota: Fargo Grand Forks | (1) (1) | 0 | 1 0 | 0 | 1 | 0 | 11 2 | 1 0 | 0 |
| Aberdeen Nebraska: | (1) | 0 | 0 | 0 | | | 9 | 1 | ····· |
| Lincoln Omaha | 71, 100 222, 800 | 2 4 | 2 3 | 1 3 | | 0 0 | 3 18 | 0 | 0 9 |
| Topeka Wichita | 62, 800 99, 300 | 16 10 | 1 2 | 2 1 | | 0 | 9 1 | 1 51 | 3 6 |
| SOUTH ATLANTIC | | | | | | | | | |
| Delaware: Wilmington | 128, 500 | 0 | 3 | 2 | | 2 | 82 | 0 | 5 |
| Baltimore Cumberland | 830, 400 (1) | 69 0 | 27 1 | 10 1 | 42 1 | 3 0 0 | 5 1 0 | 158 1 0 | 47 1 0 |
| District of Columbia: Washington | (-) 552, 000 | 30 | 12 | 16 | 5 | . 1 | 19 | 0 | 22 |
| Virginia: Lynchburg Norfolk Richmond | 38, 600 184, 200 194, 400 | 7 50 2 | 0 1 3 | 1 2 2 | 2 | 0 0 0 | 0 4 4 | 106 114 12 | 3 9 6 |
| Roanoke West Virginia: Charleston | 64, 600 55, 200 | 4 | 1 | 1 | 4 | 1 | 0 85 47 | 2 | 1 |
| Wilmington | (¹) 39, 100 | 7 15 | 00 | 0 | | 0 1 | 0 0 | 0 | 1 2 |
| Winston-Salem South Carolina: Charleston Columbia | 80, 000 75, 900 50, 600 | 1 0 25 | 1 0 0 | 0 1 0 | 18 | 0 | 1 1 0 | 02 | 1 2 9 |
| Greenville Georgia: Atlanta | (¹) 255, 100 | 1 | 0 3 | 0 | | 0 5 | 0 - 7 | 0 1 | 2 7 |
| Brunswick Savannah Florida | (1) 99, 900 | 0 2 | 0 1 | 0 2 | 3 | 02 | 0 0 | 2 0 | 0 4 |
| Miami Tampa | 156, 700 113, 400 | 3 10 | 3 1 | 2 1 | 17 | 0 | 38 2 | 0 1 | 0 |

¹ No estimate of population made.

| City reports for week ended | l March 16, | 1929—Continued |
|-----------------------------|-------------|----------------|
|-----------------------------|-------------|----------------|

| | | Chiet | Diph | theria | Influ | lenza | | | Pneu |
|--|---|---|---|------------------------|------------------------|-------------------------|---|----------------------------------|--|
| Division, State, and city | Population July 1, 1928, estimated | Chick- en pox, cases re- ported | Cases, esti- mated expect- ancy | Cases re- ported | Cases re- ported | Deaths re- ported | Mea- sles, cases re- ported | Mumps, cases re- ported | Pneu- monia, deaths re- ported |
| EAST SOUTH CENTRAL | | | | | | | | | |
| Kentucky: Covington | 59, 000 | · 0 | 0 | 0 | | 0 | 0 | 0 | 3 |
| Memphis Nashville | 190, 200 139, 600 | 19 7 | 5 1 | 4 | | 6 4 | 0 | 0 | 9 6 |
| Mobile | 222, 400 69, 600 63, 100 | · 7 1 | 200 | 300 | 17 | 5 1 | 1 4 | 2 | 8 1 |
| WEST SOUTH CENTRAL | 03, 100 | 10 | v | - | | | 1 | | |
| Arkansas: Fort Smith Little Rock | . (1) 79, 200 | 1 | 0 1 | 0 | | 3 | 0. | 2 1 | 5 |
| Louisiana: New Orleans Shrevenort | 429, 400 81, 300 | 3 | 11 | 14 | 10 | 7 | 14 | 0 | 18 |
| Oklahoma: Tulsa | 170, 500 | 28 | 1 | 1 | | | 3 | 4 | |
| Texas: Dallas Fort Worth | 217, 800 170, 600 | 6 11 | 5 | 2 | 1 | 2 | 10 32 | 0 | 15 |
| Galveston Houston San Antonio | 50, 600 (¹) 218, 100 | 071 | . 1 3 | 0 5 4 | 2 | 0 1 12 | 1 5 1 | 0 0 1 | 2 8 10 |
| MOUNTAIN | -10,100 | | - | - | - | | | | |
| Montana: Billings | (I) | 3 | 0 | 0 | | 1 | 0 | , | 0 |
| Great Falls Helena | | 2 | 1 0 | Ŏ | | Ō | 35 18 | 1 | 0 |
| Missoula Idaho: Boiso | (i) (i) | 2 | 0 | 0 | | 0 | 9 | Ó | ī |
| Colorado: Denver | 294, 200 | 38 | 10 | 4 | 1 | 3 | 2 | 20 | 16 |
| Pueblo New Mexico: | 44, 200 | 26 | ĩ | ī | | ō | 7 | õ | 2 |
| Albuquerque Utah: | (1) | 4 | 0 | 0 | | 0 | 0 | 0 | 1 |
| Nevada: | 138,000 | 26 | 2 | U | | 0 | | 172 | 8 |
| PACIFIC | . | Ĩ | Ĩ | °{ | | Ĩ | ° | ° | 1 |
| Washington: Seattle | 383 200 | 53 | | , | | | | 70 | |
| Spokane Tacoma | 109, 100 110, 500 | 7 17 | 2 1 | Ő | | 0 | 23 2 | 0.8 | 2 |
| Portland Salem | (1) (1) | 14 3 | 8 0 | 6 0 | 6 | 4 | 50 1 | 4 | 11 1 |
| Los Angeles Sacramento San Francisco | (1) 75, 700 585, 300 | 91 4 32 | 44 2 21 | 16 0 9 | 52 2 8 | 3 2 0 | 24 0 4 | 54 11 10 | 32 5 4 |

¹ No estimate of population made.

| City | reports | for | week | ended | March | 16 | . 1929- | Continue |
|------|---------|-----|------|-------|-------|----|---------|----------|
|------|---------|-----|------|-------|-------|----|---------|----------|

| | Scarle | t lever | | Smallpo | x | Tuber- | Ту | phoie i | 979F | Whoop- ing cough, cases re- ported | |
|------------------------------|---|-------------------------|---|------------------------|-------------------------|---------------------------------|---|------------------------|-------------------------|---|--------------------------|
| Division, State, and city | Cases, esti- mated expect- ancy | Cases, re- ported | Cases, esti- mated expect- ancy | Cases re- ported | Deaths re- ported | sis, deaths re- ported | Cases, esti- mated especi- ancy | Cases re- ported | Deaths re- ported | | Deaths, all causes |
| NEW ENGLAND | | | | | | | | | | | |
| Maine: Portland | | 10 | <u>م</u> | | | 0 | | 0 | 0 | 3 | 91 |
| New Hampshire: | | | | | | | | | 0 | | 10 |
| Manchester | 3 | Ó | Ő | Ő | ŏ | Ő | ŏ | ŏ | ŭ | ŏ | 10 |
| Vermont: Barre | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 5 | 4 |
| Massachusetts: | QE | 84 | | | | 19 | , | 1 | 1 | 33 | 272 |
| Fall River | 5 | 11 | ŏ | ŏ | ŏ | 3 | ő | Ô | â | 10 | 34 |
| Springfield Worcester | 8 10 | 7 10 | 0 | 0 | 0 | 12 | 0 | 0 | 0 | 20 | 35 61 |
| Rhode Island: | | | | | | | | | 0 | 0 | 94 |
| Providence | 10 | 18 | Ö | ŏ | Ŭ | 4 | ŏ | ŏ | ĕ | 4 | 94 |
| Connecticut: Bridgeport | 13 | 4 | 0 | | 6 | 4 | 6 | 0 | 0 | 1 | 45 |
| Hartford | 6 | 10 | ŏ | Ŏ | Ŏ | i | Ŏ | ě | Ŏ | - 5 | 47 |
| New Haven | 12 | 3 | | U | 9 | 1 | U | 0 | U | Э | 42 |
| MIDDLE AILANIIC | | | | | | | | | | | |
| New York: Buffalo | 26 | 45 | 0 | 0 | 0 | 2 | 0 | 0 | a | 21 | 167 |
| New York | 351 | 337 | 0 | 0 | 0 | 122 | 8 | 5 | 0 | 60 24 | 1,720 |
| Syracuse | 14 | 9 | ŏ | ŏ | ŏ | 2 | ō | ŏ | ŏ | 16 | 68 |
| New Jersey: Camden | 7 | 7 | 0 | 0 | 0 | 0 | 0 | 0 | a | 1 | 42 |
| Newark | 44 | 25 | Ŏ | Ő | 0 | 18 | 0 | 0 | 0 | 27 | 129 |
| Pennsylvania: | Ð | z | Ű | 0 | 0 | 4 | U | v I | u l | | |
| Philadelphia | 103 | 98 10 | 0 | 0 | 0 | 29 17 | 2 | 2 | 0 | 55 19 | 545 219 |
| Reading | 4 | 8 | ŏ | Ŏ | ě | ĩ | ŏ | ō | ā. | 6 | 23 |
| EAST NORTH CENTRAL | | | | | | | | | | | |
| Ohiot | | | | | | | | | | | |
| Cincinnati | 20 | 67 | 1 | 6 | 0 | 9 | 1 | 0 | 0 | 22 | 139 |
| Cleveland | 49 12 | 26 5 | 02 | 0 | 0 | 2 | ō | ő | ă | 28 | 88 |
| Toledo | 14 | 13 | 0 | 2 | 0 | 4 | • | 1 | 0 | 84 | 82 |
| Fort Wayne | 7 | 2 | 0 | 0 | 0 | 3 | 1 | 0 | 0, | 0 | 29 |
| South Bend | 11 | 88 | 12 | ő | 0 | í | 0 | e e | ŏ | 0 | 22 |
| Terre Haute | 3 | Ž | i | 0 | 0 | 2 | 0 | 0 | 0. | 0 : | 21 |
| Chieago | 136 | 157 | 3 | 3 | 0 | 65 | 2 | 0 | 0 | 48 | 765 |
| Springfield | 3 | 4 | • | 3 | 9 | • | | U | | • | • 42 |
| Detroit | 111 | 200 | 2 | 3 | 0 | 18 | 1 | 0 | 0 | 131 | 366 34 |
| Grand Rapids. | 11 | 9 | ō | 7 | ŏ | ő | . e | e | ă | 28 | 38 |
| Wisconsin: | | 6 | 1 | 0 | | 1 | | 0 | ٥ | 1 | 13 |
| Milwaukee | 31 | 63 | ē | , Ö | Õ | 8 | <u>e</u> | <u> </u> | Q . | 133 | 137 |
| Racine Superior | 4 | 4 | 1 | θ | Ő | ē | e | Ő | å | 3 | 2 |
| WEST NORTH CENTRAL | | | | | ł | ł | | | | | |
| Minnesota: | | 1. | | | A | | | | 0 | 2 | 27 |
| Minneapolis | 57 | 26 | 2 | ě | ě | 2 | ŏ | ě | ŏ | 67. | . 95 |
| St. Paul | 34 | 27 | • | 0. | 0 | 3 | • | 0 | a | | 55 |
| Davenport | 3 | 3 | 2 | 0 | | | 0 | 9 | [| 2 | 31 |
| Siour City | 2) | 2 | í | Ö | | | Ö | 0 | | 2 | |
| Watarloo | | 57 | 0 | 1 | r | 1 | 01 | 11 | | 11 . | |

•

| | 1 | | 1 | | | 1 | 1 | | | | 1 |
|---------------------------|--------|----------|--------|--------|--------|--------|---------|----------|--------|---------|----------|
| | Scarle | t fover | | Smallp | X | - | Т | phoid i | ever | Wheen | |
| | | <u> </u> | | [| 1 | culo- | · | 1 | 1 | ing | Deaths |
| Division, State, | Cases, | Case | Cases, | Cana | Deaths | sis, | Cases, | Cases | Deaths | cough, | all |
| and city | mated | 10- | mated | re- | re- | re | mated | re- | 10- | re- | Causes |
| • | ancy | ported | ancy | ported | ported | ported | expect- | ported | ported | ported | |
| | | | | | | | | | | | |
| WEST NORTH | | | | | 1 | | | | | | |
| CENTRAL-COD. | | | | | | | | | | | |
| Missouri: | | | | | | | | | | | |
| Kansas City St. Joseph | 17 | 20 | i | 8 | U O | 8 | ŏ | ŏ | Ŭ | 15 0 | 156 |
| St. Louis | 38 | 16 | 3 | 1 | 0 | 13 - | 1 | 0 | 0 | 41 | 290 |
| Fargo | 1 | 2 | . 0 | , o | 0 | 0 | Q | 0 | 0 | 3 | 7 |
| Grand Forks | 2 | 4 | 0 | 0 | • | | 0 | 0 | | 0 | |
| Aberdeen | 4 | 0 | 0 | 0 | | | 0 | 0 | | 0 | |
| Nebraska: Lincoln | 2 | 21 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 28 |
| Omaha | 4 | 8 | 4 | 3 | 0 | 2 | 0 | 0 | 0 | 4 | 80 |
| Topeka | 3 | 3 | 1 | 0 | 0 | Q | 0 | 0 | 0 | 7 | 17 |
| Wichita | 4 | 19 | 1 | 7 | 0 | 1 | 0 | 0 | 0 | 8 | 39 |
| SOUTH ATLANTIC | | | : | | | | | | | | |
| Delaware: | | | | | | | | | | | |
| Maryland: | 1 | 1 | U U | Ű | Ŭ | U U | v | , v | ۳ | | 26 |
| Baltimore | 36 | 29 | 8 | 0 | 0 | 16 | 1 | 1 | 0 | 118 | 256 |
| Frederick | i | ō | ŏ | ŏ | ŏ | ô | ŏ | ŏ | ŏ | ŏ | 4 |
| Washington | 27 | 19 | 1 | o | 0 | 10 | 1 | 0 | o | 28 | 181 |
| Virginia: | | | | | | | | | | | 10 |
| Norfolk | 2 | ŏ | ŏ | ŏ | ŏ | ŏ | ŏ | ŏ | ŏ | 17 | |
| Richmond | 3 | 9 | 8 | 0 | 0 | 3 | .0 | 8 | 0 | 1 | 66 |
| West Virginia: | _ | | | | | | | | | | |
| Wheeling | 2 | ő | ŏ | ŏ | ŏ | 1 | ĭ | ō | ő | 4 | 29 25 |
| North Carolina: | 1 | 0 | 0 | 1 | ó | 0 | 0 | 0 | 0 | | 8 |
| Wilmington | ō | ĭ | ě | 2 | ŏ | ŏ | ŏ | ŏ | ŏ | 2 | . 16 |
| South Carolina: | v | - | 3 | ۷ | U | 2 | ۰, | U U | ۰ | 15 | . 21 |
| Charleston | 1 | 0 | 8 | 0 | 0 | 1 | 8 | 2 | 0 | 0 | 18 |
| Greenville | ŏ | ō | ŏ | ŏ | ŏ | 2 | ŏ | ŏ | ŏ | 3 | 10 |
| Atlanta | 5 | 4 | 4 | 0 | o | 8 | 0 | o | 0 | 2 | 79 |
| Brunswick | 0 | 0 | 0 | 8 | 0 | 0 | 0 | 0 | 0 | e e | 2 |
| Florida: | | | | | | | | | | | · •• |
| Tampa | ō | 2 | ŏ | ő | ő | 3 | 1 | ŏ į | ö | 10 | 21 17 |
| EAST SOUTH | | | | | | | | | | | |
| CENTRAL | | | | | | | | | | · | |
| Kentucky: | | | | | | | | | | | |
| Covington | 2 | 4 | 0 | 0 | 0 | 3 | 0 | 0 | 0 | 0 | 23 |
| Memphis | 5 | 11 | 3 | 0 | 0 | 3 | 1 | 1 | 0 | . 0 | 69 62 |
| Alabama: | • | | - 1 | , i | U U | • | Ů | | ٩ | - | 02 |
| Birmingham Mobile | 3 | 9 | 8 | 8 | 8 | 63 | 1 | 8 | 8- | | 78 31 |
| Montgomery | Õ | ī | Ĩ | Ĭ. | | | Ō | ŏ. | | Õ. | ••••• |
| WEST SOUTH CEN- | | | | | | | | | | | |
| TRAL | | | | | | | | | | | |
| Arkansas: Fort Smith | 1 | , | | 0 | | | 0 | <u> </u> | 1 | | |
| Little Rock | 2 | 4 | i | ŏ | 0 | 2 | ŏ | ŏ | 0 | ĭ. | |
| New Orleans. | 7 | 67 | 0 | 0 | o | 15 | 2 | 1 | o | 1 | 154 |
| Shreveport! | . 01 | 01 | 1' | 0 1 | 0 | 01 | 01 | 01 | 0 1 | 01 | 29 |

City reports for week ended March 16, 1929-Continued

| | Scarle | t fever | | Smallp | DX | Tuber- | Т | phoid i | lever | Whoop | |
|----------------------------------|---|-------------------------|---|------------------------|-------------------------|--|---|------------------------|-------------------------|---|--------------------------|
| Division, State, and city | Cases, esti- mated expect- ancy | Cases, re- ported | Cases, esti- mated expect- ancy | Cases re- ported | Deaths re- ported | culo- sis, deaths re- ported | Cases, esti- mated expect- ancy | Cases re- ported | Deaths re- ported | ing cough, cases re- ported | Deaths, all causes |
| WEST SOUTH CEN- TRALcontinued | | | | | | | | | | | |
| Oklahoma: Tulsa Texas: | 1 | 3 | 2 | 3 | | | 0 | 0 | | 6 | |
| Dallas | 3 | 18 | 4 | 5 | 0 | 0 | 1 | 0 | 0 | . 9 | 62 |
| Fort Worth | | 11 | | 35 | N N | 0 | 0 | Ű | | U U | 39 |
| Honston | 1 i | 5 | 2 | 5 | ŏ | 4 | ŏ | ŏ | 1 | ŏ | 83 |
| San Antonio | ī | ŏ | ō | ĭ | Ŏ | 7 | Ŏ | ž | ō | ŏ | 67 |
| MOUNTAIN | | | | | | | | | | | |
| Montana: | | | • | | | | | | | | |
| Billings | 1 | 2 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 9 |
| Great Falls | 1 | 4 | 1 | 0 | 0 | 1 | 0 | 0 | 0 | 10 | 9 |
| Missoulo | Ň | | | Ň | Ň | 1 | Ň | Ň | Ň | Ň | 0 |
| Idaho: | v | • | v I | v I | • | v | v | v | v | v | |
| Boise | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 8 |
| Colorado: | | _ | | | | | | | | | |
| Denver | 15 | 5 | 2 | 0 | 0 | 10 | 0 | 0 | 0 | 3 | 88 |
| Pueblo | 1 | • | 0 | | Ű | 1 | | 1 | U | U | 10 |
| Albuquerque. | 1 | 2 | 0 | 0 | 0 | 3 | 0 | 0 | 0 | 57 | 14 |
| Utah: | - | - | | | | | | | | | |
| Salt Lake City. | 3 | 6 | 2 | 1 | 0 | 0 | 0 | 2 | 0 | 6 | 40 |
| Nevada: | | | | | • | | | | • | • | 1 |
| Reno | , v | ° | ۳ | Ŭ, | Ů | Ů | ľ, | Ŭ | v | Ů | • |
| PACIFIC | | | | | | | | | | | |
| Washington: | | | | | | | | | | | |
| Seattle | 11 | 5 | 4 | 0 | | | 0 | 1 | | 41 | |
| Spokane | 4 | 2 | 3 | 5 | | | | 1 | | 2 | |
| Oregon: | | - | • | ° I | | - | ° I | - | v | ~ | 20 |
| Portland | 6 | 17 | 11 | 27 | 0 | 7 | 1 | 0 | 0 | 0 | 96 |
| Salem | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | |
| California: | 20 | 70 | | | | 24 | | | | 21 | 907 |
| LOS Angeles | 32 | 32 | 1 | Ϋ́Ι | | 24 | 1 | N N | 1 | 31 | 287 |
| San Francisco | 17 | 64 | 3 | 3 | ŏ | 15 | ĭ | 2 | Ô | 38 | 168 |
| | | | | | | | | | | | |
| | | Men | ingococ | ms | Letherri | ~ | | 1 | Poliom | alitis (ir | fantile |

City reports for week ended March 16, 1929-Continued

| | Menin meni | gococcus ngitis | Leti encep | hargic halitis | Pel | lagra | Poliomyelitis (infantile paralysis) | | | |
|--|-------------------|--------------------|--|-------------------|-------------|-------------|--|-------------|-------------|--|
| Division, State, and city | Cases | Deaths | Lethargic encephalitis Pellagra Pollom p Cases Deaths Cases Deaths Cases, esti- mated expect- ancy 1 0 0 0 0 1 0 0 0 0 1 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 | Cases | Deaths | | | | | |
| NEW ENGLAND | | | | | | | | | | |
| Massachusetts: Boston Connecticut: | 2 | 3 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | |
| MIDDLE ATLANTIC | ·U | Ű | 1 | 1 | U | 0 | U | v | U | |
| New York: | | | | | | | | | | |
| Buffalo New York Rochester Syracuse | 1 36 1 0 | 0 16 0 0 | 0 8 0 0 | 0 3 0 1 | 0 0 0 | 0 0 0 | 0 1 0 0 | 0 0 0 | 0 0 0 | |
| New Jersey: Newark | 2 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | |
| Philadelphia Pittsburgh | 3 2 | 3 6 | 0 | 8 | 1 0 | 1 0 | 0 | 0 | 0 0 | |

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City reports for week ended March 16, 1990-Continued

| | Difeni n men | enfingoebeccus Lethangie encephalitis Pellagra Pellomyelitis paraly | | yelitis (i paralysi | nfantile is) | | | | |
|---------------------------------|------------------------|---|----------|------------------------|-----------------|----------|---|-------|--------|
| Division, State, and eity | Cases | Deaths | Cases | Deaths | Casilia | Deaths | Cases, exti- mated expect- ancy | Cases | Deaths |
| EAST NORTH CENTRAL | | | | | | | | | |
| Ohio: | | | | | | | | | |
| Cleveland Columbus | 3 | 2 | 2 | 1 | 0 | | O A | | . 0 |
| Toledo | 1 | Ô | ŏ | ŏ | ŏ | ŏ | Ŏ | ŏ | Å |
| Indianapolis | . 0 | 4 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| Dinois: | 7 | | 0 | · . | 0 | | | • | |
| Michigan: | | 3 | | | Ū | 4 | | , v | |
| Flint | 21 | 12 | 0 | 0 | 0 | a a | 0 | 0 | |
| Wisconsin: | - | | • | | | | | | |
| WEST BODTU CONTRAL | 8 | 3 | U | | . 0 | U | U | U | U |
| Minnesota: | | | | | | | 10 A | | |
| St. Paul | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| Kansas City | 9 | 4 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |
| St. Louis | 12 | 1 | 0 | 0 | 0 | 0 | 0 | θ | 0 |
| Grand Forks | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| SOUTH ATLANTIC | | | | | | | | | |
| Maryland: | | | | | | | | | |
| Virginia: | 1 | U | | | U | u | Ű | , v | U |
| Richmond North Carolina: | 0 | 0 | 0 | 1 | 0 | 1 | 0 | 0 | 0 |
| Raleigh | 0 | 0 | 0 | 0 | 0 | 2 | Q | 0 | 0 |
| Winston-Salem | | 0 | 0 | ő | 2 | 1 | Ö | ŏ | Ŭ Ŭ |
| South Carolina: Charleston 1 | | | | 0 | 2 | a | 0 | | 0 |
| Columbia | ŏ | ĭ | ŏ | ŏ | õ | ŏ | ŏ | ŏ | ŏ |
| Atlanta | 8 | 4 | 0 | 0 | 1 | 0 | o | ol | 0 |
| Savannah | 0 | Ō | 0 | 0 | 2 | 1 | 0 | 0 | 0 |
| Tampa | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 1 | 0 |
| EAST SOUTH CENTRAL | | | | | | | | 1 | |
| Alabama: | | | | | | | | | • |
| Mobile | 1 | 1 | ŏ | ŏ | ŏ | ŏ | ő | ŏ | Ő |
| Montgomery. | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| WEST SOUTH CENTRAL | | [| | | | | | | |
| Little Rock | 1 | 0 | 0 | 0 | o | 0 | 0 | o | 0 |
| Louisiana: New Orleans | 5 | 1 | | 0 | , | 0 | 0 | 0 | 0 |
| Texas: | | - | | | | | | | • |
| Galveston | 0 | ő | ŏ | ŏ | Ő | 1 | 0 | | · ŏ |
| Houston | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| MOUNTAIN | Ů | ° | Ů | Ŭ, | Ů, | - | | , i | • |
| Colorado: | | | | | | ł | | | |
| Denver | 6 | 3 | 0 | 0 | 0 | 0 | 0 | σ | O |
| Salt Lake City | 6 | 3 | 0 | 0 | 0 | 0 | 0. | 0 | 0 |
| PACIFIC | | | | | | | | [| |
| Washington: Seattle | 8 | | | | | | | | 0 |
| Oregon: | | | <u> </u> | <u> </u> | | | | | ~ |
| California: | a | 0 | I | I | Q | 0.1 | σ | o l | U |
| Los Angeles ¹ | 7 | 4 | 1 | 1 | <u>s</u> | <u>0</u> | 0 | 2 | 1 |
| San Francisco | ĩ | 2 | ă | ŏ | ă | ŏ | ŏ | ŏ | ō |

¹ Dengue: 1 case at Charleston, S. C. ³ Rabies (in man); 1 case and 1 death at Los Angeles, Calif.

The following table gives the rates per 100,000 population for 98 cities for the 5-week period ended March 16, 1929, compared with those for a like period ended March 17, 1928. The population figures used in computing the rates are approximate estimates, authoritative figures for many of the cities not being available. The 98 cities reporting cases had estimated aggregate populations of more than 31,000,000. The 91 cities reporting deaths had nearly 30,000,000 estimated population. The number of cities included in each group and the estimated aggregate populations are shown in a separate table below.

Summary of weekly reports from cities, February 10 to March 16, 1929-Annual rates per 100,000 population compared with rates for the corresponding period of 1928

| | Week ended- | | | | | | | | | | | | |
|--------------------|---------------------|---------------------|---------------------|---------------------|--------------------|--------------------|--------------------|---------------------|---------------------|---------------------|--|--|--|
| | Feb. 16, 1929 | Feb. 18, 1928 | Feb. 23, 1929 | Feb. 25, 1928 | Mar. 2, 1929 | Mar. 3, 1928 | Mar. 9, 1929 | Mar. 10, 1928 | Mar. 16, 1929 | Mar. 17, 1928 | | | |
| 98 cities | 122 | 177 | 118 | 177 | 3 122 | 174 | 134 | 174 | 127 | 160 | | | |
| New England | 131 | 172 | 118 | 138 | 124 | 140 | 109 | 145 | 136 | 136 | | | |
| Middle Atlantic | 147 | 235 | 139 | 224 | 140 | 234 | 185 | 214 | 159 | 213 | | | |
| East North Central | 115 | 169 | 106 | 169 | 131 | 163 | 130 | 171 | 120 | 135 | | | |
| West North Central | 150 | 125 | 131 | 125 | * 136 | 113 | 144 | 131 | 152 | 115 | | | |
| South Atlantic | 73 | 155 | 67 | 168 | 64 | 140 | 67 | 132 | 84 | 151 | | | |
| East South Central | 81 | 63 | 68 | 35 | 54 | 98 | 68 | 84 | 54 | 119 | | | |
| West South Central | 119 | 126 | 182 | 191 | 4 156 | 93 | 119 | 170 | 99 | 138 | | | |
| Mountain | 44 | 186 | 44 | 71 | 61 | 186 | 61 | 97 | 44 | 106 | | | |
| Pacific | 80 | 82 | 110 | 161 | 75 | 141 | 37 | 171 | 67 | 125 | | | |

DIPHTHERIA CASE RATES

MEASLES CASE RATES

SCARLET FEVER CASE RATES

| | 278 | 290 | 262 | 291 | 3 301 | 290 | 299 | 299 | 326 | 301 |
|--------------------|-----|-----|-----|-----|--------------|-----|-----|-----|-----|-----|
| New England | 376 | 441 | 294 | 414 | 339 | 347 | 310 | 377 | 371 | 402 |
| Middle Atlantic | 222 | 331 | 202 | 336 | 230 | 346 | 228 | 359 | 266 | 353 |
| East North Central | 340 | 280 | 340 | 285 | 401 | 309 | 410 | 292 | 417 | 296 |
| West North Central | 360 | 266 | 373 | 276 | 340 | 262 | 356 | 291 | 367 | 272 |
| South Atlantic | 157 | 222 | 144 | 243 | 137 | 207 | 155 | 245 | 146 | 216 |
| East South Central | 258 | 98 | 183 | 98 | 217 | 112 | 197 | 175 | 231 | 63 |
| West South Central | 265 | 118 | 281 | 122 | 4220 | 97 | 281 | 130 | 379 | 211 |
| Mountain | 87 | 346 | 113 | 204 | 218 | 257 | 157 | 195 | 157 | 248 |
| Pacific. | 339 | 230 | 302 | 233 | 509 | 194 | 424 | 192 | 459 | 217 |

¹ The figures given in this table are rates per 100,000 population, annual basis, and not the number of cases reported. Populations used are estimated as of July 1, 1929 and 1928, respectively. ³ Omaha, Nebr., Fort Smith, Ark., and Galveston, Tex., not included. ³ Omaha, Nebr., not included. ⁴ Fort Smith, Ark., and Galveston, Tex., not included.

4 Fort Smith, Ark., and Galveston, Tex., not included.

Summary of weekly reports from cities, February 10 to March 16, 1929—Annual rates per 100,000 population compared with rates for the corresponding period of 1938—Continued

SMALLPOX CASE RATES

| | Week ended- | | | | | | | | | | | | |
|--------------------|---------------------|------------------------|---------------------|---------------------|--------------------|--------------------|--------------------|---------------------|---------------------|---------------------|--|--|--|
| • | Feb. 16, 1929 | Feb. 18, 1928. | Feb. 23, 1929 | Feb. 25, 1928 | Mar. 2, 1929 | Mar. 3, 1926 | Mar. 9, 1929 | Mar. 10, 1928 | Mar. 16, 1929 | Mar. 17, 1928 | | | |
| 96 cities | 8 | 20 | 12 | 25 | 3 16 | 17 | 12 | 23 | 12 | 21 | | | |
| New England | 0 | 0 | 0 | 0 | 20 | 0 | 0 | 0 Ø | 5 | 0 | | | |
| East North Central | 15 0 | 12 102 | 15 15 | 13 92 | 24 10 | 18 63 | 18 6 | 14 92 | 20 31 | 26 65 | | | |
| South Atlantie. | 20 | 27 35 | 4 | 29 56 | 777 | 21 0 | 6 7 | 25 21 | 67 | 38 21 | | | |
| West South Central | 24 70 | 20 ⁻ 168 | 99 35 | 8 62 | 4 118 87 | 20 53 | 99 44 | 36 115 | 43 17 | 45 53 | | | |
| Pacific | 25 | 18 | 20 | 125 | 26 | 49 | 17 | 69 | 22 | 38 | | | |

TYPHOID FEVER CASE RATES

| | | | | | | _ | _ | | | |
|--------------------------------|--|---|----------------------------|--|---|--|---|---------------------------------------|---|--|
| 98 cities | 5 | 5 | 4 | • 5 | 14 | 10 | 5 | 4 | 5 | 5 |
| New England Middle Atlantic | 5 4 2 12 6 14 12 0 7 | 5 3 3 4 8 14 12 8 8 | 942 42 47 80 5 | 7 5 1 4 10 28 16 0 5 | 2 2 0 18 2 14 4 21 9 7 | 0 8 7 6 13 70 8 8 8 8 | 5 4 3 4 6 7 20 0 17 | 2 3 4 2 10 7 4 3 | 2 4 2 7 7 12 26 10 | 7 2 3 4 11 14 12 0 5 |

INFLUENZA DEATH BATES

| | and the second se | | | | | | | | | |
|---|---|--|---|--|--|---|---|--|--|---|
| 91 cities | 54 | 23 | 45 | 22 | ¥ 40 | 25 | 33 | 23 | 33 | 26 |
| New England Middle Atlantic East North Central West North Central South Atlantic East South Central Meentain Pacific | 57 44 36 33 80 222 158 87 43 | 11 18 12 98 54 92 71 27 | 41 35 33 45 69 81 138 78 39 | 7 24 14 3 31 46 75 35 20 | 20 30 31 345 67 148 89 52 33 | 7 16 17 15 34 123 104 89 24 | 16 25 31 21 47 74 122 61 23 | 21 20 16 18 27 54 75 62 20 | 25 31 23 27 37 118 106 35 16 | 7 26 12 24 21 123 117 80 10 |

PNEUMONIA DEATH RATES

| | | | | | | | the second s | | | |
|---|---|---|---|---|---|---|--|---|---|---|
| 91 cities | 223 | 177 | 194 | 166 | 1 222 | 193 | 204 | 196 | 185 | 227 |
| New England Middle Atlantic. East North Central. West North Central. South Atlantic. East South Central. West South Central. Meantain. Pacific. | 305 254 182 190 243 163 219 244 128 | 170 196 137 141 216 192 283 168 172 | 235 192 170 207 238 155 266 226 134 | 147 156 156 107 231 232 275 249 115 | 274 240 180 * 214 255 281 215 279 154 | 193 218 148 159 205 245 266 266 155 | 219 233 159 195 234 237 235 183 144 | 205 221 156 144 212 306 258 266 121 | 201 197 155 180 199 200 239 253 141 | 239 259 197 208 216 268 268 204 125 |
| | | | | | | | | | | |

Omaha, Nebr., Fort Smith, Ark., and Galveston, Tex., not included.
 Omaha, Nebr., not included.
 Fort Smith, Ark., and Galveston, Tex., not included.

| Group of cities | Number of cities reporting | Number of cities reporting | Aggregate of cities cases | population reporting | Aggregate of cities deaths | population reporting |
|--|----------------------------------|----------------------------------|--|--|--|--|
| | Cases | deaths | 1929 | 1928 | 1929 | 1928 |
| Total | 98 | 91 | 31, 568, 400 | 31, 052, 700 | 29, 995, 100 | 29, 498, 600 |
| New England. Middle Atlantic. East North Central. | 12 10 16 | 12 10 16 | 2, 305, 100 10, 809, 700 8, 181, 900 | 2, 273, 900 10, 702, 200 8, 001, 300 | 2, 305, 100 10, 809, 700 8, 181, 900 | 2, 273, 900 10, 702, 200 8, 001, 300 |
| West North Central South Atlantic East South Central | 12 19 6 | 9 19 5 | 2, 712, 100 2, 783, 200 767, 900 | 2, 673, 300 2, 732, 900 745, 500 | 1, 736, 900 2, 783, 200 704, 200 | 1, 708, 100 2, 732, 900 682, 400 |
| West South Central Mountain Pacific | 8 9 6 | 7 9 4 | 1, 319, 100 598, 800 2, 090, 600 | 1, 239, 900 590, 200 2, 043, 500 | 1, 285, 000 598, 800 1, 590, 300 | 1, 256, 400 590, 200 1, 551, 200 |

Number of cities included in summary of weekly reports, and aggregate population of cities of each group, approximated as of July 1, 1929 and 1928, respectively

FOREIGN AND INSULAR

CANADA

Provinces—Communicable diseases—Week ended March 9, 1929.— The Department of Pensions and National Health reports cases of certain communicable diseases from nine Provinces of Canada for the week ended March 9, 1929, as follows:

| Disease | Prince Ed- ward Island | Nova Scotia | New Bruns- wick | Que- bec | On- tario | Mani- toba | Sas- katch- ewan | Al- berta | British Co- lumbia | Total |
|--|---------------------------------|----------------|-----------------------|-------------|-------------------|---------------|------------------------|--------------|--------------------------|----------------------|
| Cerebrospinal fever- Influenza- Smallpox- Typhoid fever | 2 | 21 | | | 1 72 7 6 | 1 2 5 | 42 4 | 1 13 | 3 82 18 | 5 275 42 26 |

Quebec Province—Communicable diseases—Three weeks ended March 16, 1929.—The Bureau of Health of the Province of Quebec reports cases of certain communicable diseases for the three weeks ended March 16, 1929, as follows:

| Disease | w | eek ende | »d— | Disco | w | eek ende | |
|--|-------------------------------|-----------------------------|----------------------------------|---|--------------------------------|---------------------------------|---------------------------------|
| | Mar. 2 | Mar. 9 | Mar. 16 | Diseaso | Mar. 2 | Mar. 9 | Mar. 16 |
| Cerebrospinal meningitis. Chicken pox | 66 41 5 9 1 72 | 40 42 11 56 105 | 1 39 55 13 34 149 | Mumps Scarlet fever Smallpox Tuberculosis Typhoid fever Whooping cough | 5 125 3 57 5 13 | 19 128 5 45 5 30 | 52 159 4 54 8 19 |

MEXICO

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Tampico—Communicable diseases—February, 1929.—During the month of February, 1929, communicable diseases were reported at Tampico, Mexico, as follows:

| Disease | Cases | Deaths | Discase | Cases | Deaths |
|--|---------------|--------------|---|-------------------|--------|
| Chicken pox Enteritis (various) Influenza Malaria | 1 16 32 | 48 8 8 | Measles Smallpox Tuberculosis Whooping cough | 8 1 60 1 | 4 |

VIRGIN ISLANDS

Communicable diseases—February, 1929.—During the month of February, 1929, cases of certain communicable diseases were reported from the Virgin Islands, as follows:

| St. Thomas and St. John: | Cases |
|--------------------------|-------|
| Chancroid | 1 |
| Dengue | 1 |
| Fish poisoning | 1 |
| Gonorrhea | 2 |
| Pellagra | 1 |
| Syphilis | 4 |
| Uncinariasis | 1 |
| | |

| 3t. | Croix: | Cases |
|-----|--------------|-------|
| | Gonorrhea | . 2 |
| | Malaria | . 1 |
| | Syphilis | 6 |
| | Tetanus | 1 |
| | Uncinariasis | 2 |

YUGOSLAVIA

Communicable diseases—February, 1929.—During the month of February, 1929, communicable diseases were reported from Yugoslavia as follows:

| Disease | Cases | Deaths | Disease | Cases | Deaths |
|--------------------------|-------|--------|---------------|--------|--------|
| Anthrax | 28 | 6 | Measles | 1, 674 | 44 |
| Carebrospinal meningitis | 16 | 7 | Scarlet fever | 1, 573 | 307 |
| Diphtheria | 313 | 78 | Tetanus | 5 | 4 |
| Dysentery | 17 | 2 | Typhoid fever | 82 | 16 |
| Lethargic encephalitis | 2 | 1 | Typhus fever | 13 | 2 |

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER

From medical officers of the Public Health Service, American consuls, health section of the Leaguo of Nations, and other sources. The reports contained in the following table must not be considered as complete or final as regards either the list of countries included or the figures for the particular countries for which reports are given:

CHOLERA

[C indicates cases; D, deaths; P, present]

| | | | | | | | | | Weel | t ended | | | | | | |
|---------------------------|---------|----------------------------------|-------------------|------------------|--------|-----|--------------|-----------------------|--|---------|--------|---------|------|----|-----------|----|
| Place | Sept. | 21- 21- 17, 17, 100. | Nov. 18 15, | Decen 192 | aber, | | fanuary | , 1920 | | | ebruar | y, 1920 | | Ŵ | roth, 197 | |
| | 0701 | 0741 | 0701 | 8 | 8 | 20 | 12 | 8 | 8 | 8 | 6 | 9 | ส | 6 | • | 16 |
| Cevior | | | | | 4 | 6 | | | | | | İ | | | Ì | |
| Colombo | | | | | 61 | 6 | 63 | - | - | | | T | | | | |
| D Ingiriya Province | | ī | | | T | T | - | | | Ť | | Ť | | | | |
| China: Canton Canton C | | - 6 | | 6 | | İ | - | | | | | | | | | |
| Bhanghai D | 6 | | | - | | | | | | | | | | | | |
| India | 17,028 | 20, 937 | 23, 528 | 4, 602 | 4, 507 | 88 | 3,801 | 4, 173 | 80 19 19 19 19 19 19 19 19 19 19 19 19 19 | Ī | ÌÌ | | | | | |
| Bassein Bombay C | 10, 10, | 14, 41 | 14, 800 1 4 | A% (7 | 9 0 5 | 8 | 8 | , , , , , | 3 - 4 | - | | 61 | Π | | 2 | |
| Calcutta. | 41-7 | 219 | 247 | 8 | 12 | 122 | 8 | ន | 9 | 34 | 8 | | | | - | |
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| Diete | 800- tember | Oeto- | ΔON | ember, 19 | 8 | Dece | mber, 19 | 58 | Jan | uary, 19 | 8 | Februar | y, 1020 |
| 0.000 V | 1926 | 1928 | 1-10 | 11-20 | 21-30 | 1-10 | 11-20 | 21-31 | 1-10 | 11-20 | 21-31 | 1-10 | 11-20 |
| Indo-China (French) (see also table above): Annan Cambodia Cochin-China. Kwangchow-Wan. | 2364 23164 | 52 52 52 52 | 27 | 4 -8 ⁴ ⊡ | 5 17 81 | | 21 351 | 346 346 | 232 | 25 | 226 | 107 | 115 |

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

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| Place | สรีส | 21- Nov. 17, | ₩9.5 | Decem 192 | s 8 | Je | nuary | 1929 | | Fet | oruary. | 1920 | | Ā | larch, | 1929 | |
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| Buenos Aires * | | P 1 | 6 | | | | | | | | | | | | | | |
| Jujuy Province: Perico Rosario Rosario Tucuman Province: El Mollar Azores: St. Michaels Island | e e e e e e e e e e e e e e e e e e e | 2 | | | | | | | | | ∞ - | | | | | | |
| Belgtan Congo: Diugu. Braali: | | - 81 | | | F | | | | | | | | | | | | |
| Farts. British East Africa (see also table below): Mombasa | | | 5 | | | | | | - | | | | | | | | |
| Plague-infected rats. Uganda. Canary Islands: Las Palmas | °°58°° | 114 | 124 | 38 | 22 | 12 | 44 42 | 88 | 37 | | | | | | | | |
| Tenerifie | 1-107 | | - | | | | | | | | | | | | | | |
| Colombo Plarue-infected rats | 844 | | 41410 | 00 | | 00 | ~~~ | | 77 | - 12 | 1010 | 000 | | | 84 | | |

April 5, 1929



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| Diagna-infactad rats | R | 23 | 42 | 6 | 81 | 12 | 14 | 13 | 13 | - | | | | | |
| East Java and Madura. | 16 | | <u> </u> | | | | | | | | | | | | 11 |
| Surabaya. C | <u>-</u> | 3 | | | | | | | | - | | | | İ | 1 |
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During the period from Nov. 10 to Dec. 11, 1928, 13 cases of plague were reported at El Mollar, Tucuman Province, Argentina. During the same period 1 case of plague was reported at Chipton and 1 at Ucacha, both in Cordoba Province, Argentina.
 B plague-infocted rats were reported at Buenos Aires, Argentina, from July 1 to Dec. 31, 1928.
 Unofficial report.

April 5, 1929

| FEVER-Continued | |
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| YELLOW | |
| AND | |
| FEVER, | Continued |
| TYPHUS | PLAGUE-(|
| SMALLPOX, | |
| PLAGUE, | |
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| ryria (see table balow). Turkey: Adala Adala Adala Tabon of Bouth Atrica: Canage Province. Crange Province. Crange Province. Crange Rives Batte. Transvali Boyter Republica: Relmouts District. Relmouts e | British East Africe (see also table above): D Kenya |

¹ Reports incomplete.

April 5, 1929

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

SMALLPOX

[C indicates cases; D, deaths; P, present]

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| Southern Rhodesia. | 6 | | 2 | | | • | | | | - | | | | - | | | |
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April 5, 1928

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

SMALLPOX-Continued

[C indicates cases; D, deaths; P, present]

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'April 5, 1929

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CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

SMALLPOX-Continued

[C indicates cases; D, deaths; P, present]

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| Senegal Sudan (French) Syrla: Beirut. | | | CACC | | | 100 A 20 A 20 A 20 A 20 A 20 A 20 A 20 A | T | | 1 | 2 | | P | | 1 | | 21 | |
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April 5, 1929

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| Egypt: Alexandria Assiont Province. Assouan Province. | 1 | | | | | | | | _ | | | | | |
| Beheira Province. | 2 | | | | | | 9 | | | | | | | |
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April 5, 1929

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

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| FEVER, AN | |
| TYPHUS | |
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| , PLAGUE, | |
| CHOLERA | |

YELLOW FEVER

[C indicates cases; D, deaths; P, present]

| | Aug. | Sept. | Oct. | | | | | Week e | nded | | | | |
|----------------------------------|------------------|------------------------|--------------------|------|---|-----|----------|--------|------|---|---------|--------|------------|
| Place | ß ^g t | ื่ส ^{ู่อั} ส์ | 21- Nov. 17, | Nov. | | Dec | ember, 1 | 928 | | | January | , 1929 | |
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120 cases of yellow fever with 14 desths were reported at Rio de Janeiro during January, 1929, mostly suburban. During February there were 25 confirmed cases of yellow faver at Rio de Jameiro, with a mortality of about 66 per cent of the cases. I Supperded cases.