Determinants of fetal haemoglobin in newborn infants

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Abstract

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Percentage fetal haemoglobins (% HbF) were measured in 7081 newborns at the University Hospital, Nottingham, between 1990 and 1992 to assess the effects of a range of factors, and to determine if there was an association between % HbF at birth and risk of SIDS (sudden infant death syndrome). Data from the obstetric clinical records were analysed by multiple linear regression to establish the major determinants of % HbF. New findings were the identification of maternal smoking, twin pregnancy, ethnic origin, and season of birth as significant determinants. Alcohol use, drug and vitamin intake, and duration of labour were not found to be important. The previously reported correlations between % HbF and sex, gestational age, and birthweight were confirmed, and the relative effects quantified. Overall, the major determinants of an increase in % HbF in neonates are similar to the known antenatal risk factors for sudden infant death syndrome (SIDS).

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Keywords: fetal haemoglobin, newborns, determinants, SIDS.

Table 1 Population statistics and mean % fetal haemoglobin for 7081 babies by levels of maternal and infant factors

Variable	$n=\star$	% HbF	SD	SEM
Infant factors				
Sex				
Male	3690	78.07	7.01	0.12
Female	3391	76.62	7.40	0.13
Birthweight (kg)				
<2.50	409	84·26	6.94	0.34
2.50-2.99	1227	79.37	7.03	0.20
3.00-3.49	2660	77.14	6.78	0.13
≥3.50	2781	75.71	6.97	0.13
Ethnic group				
Caucasian	6202	77.60	7.07	0.09
Afro-Caribbean	240	77.07	7.95	0.51
Indo-European	498	74.60	8.05	0.36
Arabic/Oriental/other	141	77.75	8.00	0.67
Singleton/twin				
Singleton	6933	77.23	7.21	0.09
Twin	148	84.00	4.99	0.41
Period of delivervt				
June 1990-August 1991	4857	76.72	6.97	0.10
February 1992-October 1992	2224	78.80	7.58	0.16
Maternal factors during pregnancy				
Cigarette smoking				
No	5123	76.90	7.26	0.10
Yes	1902	78.61	6.92	0.16
Alcohol use	1702		• • •	• • •
None	4714	77.26	7.26	0.11
Some	2270	77.57	7.11	0.15
Drugs or vitamins taken			• • • •	
No	6797	77.37	7.20	0.09
Yes	284	77.45	7.85	0.47

*Numbers of infants do not always equal 7081 because of missing values. †A new instrument to measure HbF was introduced 1 October 1991. Giulian *et al* reported startlingly high values of percentage fetal haemoglobin (% HbF) in infants dying from cot death.¹ Fagan and Walker presented similar findings, showing raised values of % HbF in sudden infant death syndrome (SIDS) but only when an appropriate gestationally matched control population was used.² Significant differences between the postnatal decline in % HbF in preterm and full term infants were also reported.²

The present study of % HbF values in newborn infants was undertaken to establish normal values and ranges in newborns of different gestational ages, to assess the effects of fetal, maternal, environmental and ethnic factors, and to establish a cohort of infants for which the patterns of later disease, especially SIDS, could be examined in relation to % HbF at birth. Results relating to the first two objectives are presented in this report.

The qualitative pattern of developmental changes in haemoglobin concentration during pregnancy is well known,³⁻⁵ but the quantitative measurement of HbF at different gestational ages is less well delineated. Two previous systematic surveys of haemoglobin in the late second and third trimesters of pregnancy measured adult haemoglobin (HbA), and assumed that the remainder of the haemoglobin present was fetal.⁶⁷ Most previous reports of normal values of % HbF have been obtained from relatively small numbers of infants.

Methods

Cord blood samples were sought from all infants delivered at the University Hospital, Nottingham, from June 1990 to October 1992. The samples were taken in the delivery suite by midwives following the procedure used in the sickle cell screening programme – samples were taken by venesection of the umbilical cord shortly after delivery, using a needle and syringe. Samples were placed in a plastic sample tube with liquid Sequestrene anticoagulant and refrigerated at $+4^{\circ}$ C for collection within 24 hours. Percentages of HbF in the samples were then measured.

The first half of the study was carried out using the apparatus described before.² The spectrophotometer was then replaced by a Uvicon machine, working in batches of 10 using an automatic sampling device. The method of Jonxis and Huisman⁸ was used.² As all the % HbF values were expected to be high, only the method of Jonxis and Huisman was used. Controls were run on both machines, and the values obtained were within



Figure 1 % HbF and logit (% HbF) by gestation.

the range of experimental error, but the Uvicon gave a standard error roughly half that of the Beckman, and about 1.5% higher HbF readings. The survey values were not corrected; the data quoted are as measured. The data were analysed both in separate sets and combined into one set; the same results were obtained.

Following the findings of a seasonal variation in % HbF, a second control series was assessed with the Uvicon water bath at various temperatures in the range +20°C to 28°C. In keeping with theoretical considerations, no evidence of a temperature related effect was found. All samples with confirmed haemoglobinopathies and chromosomal disorders were excluded from the present study with the exception that α thalassaemia and heterozygous β thalassaemia, occurring in the Indo-European and Afro-Caribbean population subgroups, could not be excluded. In this study Indo-European refers to residents with India, Pakistan, or Bangladesh as their country of origin.

Maternal data, hospital registration numbers, and % HbF values were input to a computer file and this file was cross-referenced



with a larger file supplied by the Department of Obstetrics. This larger file included information on birthweight, ultrasound scan dates, date of delivery, period of gestation, ethnic origin and maternal lifestyle factors during pregnancy. After reviewing the original hard copy records for query matches, records from both computer files were found for 7081 infants and the analysis proceeded on the basis of a new merged file for these infants. Original laboratory log book entries were reviewed for data points which were outliers: % HbF less than 50%; birthweight less than 1.5 kg; and gestation either less than 24 weeks or greater than 42 weeks. A small number of errors were found and corrections were made.

Mean values of % HbF were calculated using a number of variables, and the independent importance of these variables was assessed by means of multiple linear regression. The dependent variable was chosen as the logit of % HbF (log_e (% HbF/(100-% HbF)) because the distribution of this variable was much closer to a normal distribution than was the distribution of the untransformed % HbF. Four categories of ethnic origin were coded in terms of three binary variables. Gestation, birthweight, duration of second stage of labour, and season of birth were treated as continuous variables; season of birth was assessed by the sine and cosine of θ , where $\theta = 2\pi$.(days after 1 January/365).

Results

Mean % fetal haemoglobins and associated standard deviations and standard errors for a number of variables are shown in table 1. The number of infants is relatively large and hence the standard error of the mean is very small; for example, for boys the standard error of the mean % HbF is 0.12. It follows that many of the differences between mean fetal haemoglobins, while modest in size, are highly significant. Higher values of % HbF were associated with male sex, low birthweight, non-Indo-European ethnic origin, twin delivery and cigarette smoking by the mother during pregnancy. Both maternal use of alcohol and drugs or vitamins were unimportant. Mean % HbF was also significantly raised for the second period of data collection for which a new instrument was used. To allow for this, the results were analysed in three groups - old machine results, new machine results, and all together. All the associations remained significant.

Figure 1 shows that, in addition, gestation was strongly associated with % HbF; preterm births were associated with increased values of % HbF. Figure 2 shows a seasonal pattern; babies delivered in the winter months had higher values of % HbF. (This graph also includes data for 1533 births in the period November 1992 to May 1993. These data are not analysed elsewhere in this report and were included here to assess whether the observed seasonality effect is likely to be reliable.) Table 2 shows mean values of % HbF (both sexes) by gestation (rows) and by birthweight

Table 2 Mean % HbF by gestation and birthweight (both sexes) in 5761 infants

Gestation in weeks	Birthweight in kg (number of infants)						
	<2.50	2.50-2.99	3.00-3.49	≥3.50			
24-25	94.07 (3)						
26-27	90·65 (4)						
28-29	89.33 (11)						
30-31	88·53 (26)						
32-33	85.06 (26)	76.60 (3)					
34-35	86.39 (66)	85.43 (38)	81.53 (8)				
36-37	83.11 (82)	82.52 (208)	81.89 (119)	81.89 (28)			
38-39	79.95 (60)	79.49 (470)	78.47 (850)	78·51 (481)			
40-41	78·89 (18)	75.54 (242)	75.75 (1097)	75.18 (1469)			
42-43		69·25 (12)	73.36 (98)	72.65 (313)			
≥44		76.25 (4)	73.88 (12)	74.67 (9)			

(columns). The analysis was restricted to the 5761 infants for whom an ultrasound scan had been performed and both scan date and estimated age of fetus were available. Gestational ages for these infants were obtained by comparing these data with dates of delivery. Within each birthweight group, there is a fairly convincing trend of % HbF declining with gestation; within gestation groups trends with birthweight are not as convincing. Fifth and 95th centiles for gestations of 34-35 weeks were 76.2 and 93.0. Corresponding values for gestations of 36-37 weeks, 38-39 weeks, 40-41 weeks, and 42-43 weeks were 73.2 and 90.2, 67.5 and 87.5, 67.2 and 85.3, and 58.4 and 84.3, respectively.

The independent importance of these variables was assessed by multiple linear regression. The analysis was restricted to the 5761 infants described above. Table 3 shows regression coefficients (and tests of significance) for variables analysed separately and simultaneously. Gestation was the most important variable in both sets of analyses, although the regression coefficient (absolute value) was somewhat reduced in the 'simultaneous' analysis. Stepwise regression indicated that coefficients for gestation and birthweight were modified by each other rather than by other variables. Regression coefficients for sex, delivery in the second period of data collection, and Indo-European ethnic origin were somewhat increased after simultaneous analysis with other variables. Regression coefficients were little changed for season of birth and much decreased for birthweight, cigarette

smoking, twin delivery and duration of second stage of labour. Placental weight and the ratio of birthweight to placental weight were also investigated but found to be unimportant. These two variables were excluded from the above analysis. The above multiple regression analysis was carried out separately for the four ethnic groups to determine whether each group was subject to similar influences. No important differences were found.

Discussion

The purpose of the multiple regression analysis was to identity which variables were important in their own right and which merely reflected the importance of other variables.

The factors leading to the increase in % HbF given in tables 1 and 3 - that is, gestation, birthweight, sex, twinning and maternal smoking are among the most important risk factors for SIDS.9

Comparison of the unadjusted (separate) and adjusted (simultaneous) regression coefficients for maternal smoking and % HbF increase suggest that about half of the smoking effect operates via reduced birthweight and shorter gestation and that the other half is a residual additional effect. This is closely similar to the findings of several SIDS epidemiological studies¹⁰ which suggest that there was still a residual excess risk of SIDS in the infants of antenatal smoking mothers after the effects of prematurity and low birthweight had been taken into account.

A similar but lesser effect is seen in twin deliveries where some 80% of the twin effect (0.444-0.093)/0.444=0.79) on % HbF could be explained by gestation, birthweight, and the other variables, leaving a definite residual effect. Interestingly, data collected on twin pregnancies and SIDS⁹ also suggested a residual excess risk after the effects of weight and gestation were accounted for.

Alcohol consumption, diet and prescribed drug and vitamin intake have not been shown to carry any SIDS excess risk implication, and similarly we found no effect of these factors on % HbF.

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			Variables analysed separately		Variables analysed simultaneously	
Variable	Coding note		β	β/SEM	β	β/SEM
Gestation	in weeks++		-0.103***	-38.26	-0.089***	-24.88
Sex	Kg male=0, female=1	[-0.097***	-20.10	-0.114***	-11.33
Season of birth	Cos θ [±]		0.071***	8.59	0.073***	10.02
	Sin θ±		-0.012	-0.78	-0.004	-0.54
Duration of labour Ethnic origintt	in hours	_	-0.017**	-3.20	-0.008	-1.78
Indo-European	ves=1, no=0	Г	-0.172***	-7.87	-0.205***	-10.22
Afro-Caribbean	ves=1, no=0		-0.022	-0.68	-0.079**	-2.69
Arabic/Oriental/other	ves=1, no=0		0.073	1.83	0.045	1.25
Cigarette smoking	ves=1, no=0	-	0.107***	8.48	0.046***	3.96
Twin	twin = 1, singleton = 0		0.444***	10.30	0.093*	2.07
Period of delivery§	second period=1, first period=0		0.146***	12.18	0.159***	14.84

Dependent variable is logit (% HbF) and regression coefficients (β)[†] are shown for independent variables analysed both separately (linear regression) and simultaneously (multiple linear regression) for 5761 infants.

separately (linear regression) and simultaneously (multiple linear regression) for 5/61 infants. †The regression coefficient is the slope of the regression line and estimates the change in logit (% HbF) per unit change in the dependent variable, eg, in the simultaneous analysis, logit (% HbF) is decreased by 0.089 for an increase in gestation of one week and decreased by 0.090 for an increase in birthweight of one kg. ††Recorded to two decimal points. $\#0=2\pi$.(days after 1 January/365). ‡‡Caucasian is coded as zero for all three binary variables. §See table 1. *p<0.05; **p<0.01; ***p<0.001.

The lower values of % HbF associated with Indo-European ethnic origin were not, according to the multiple regression analysis, explicable in terms of birthweight or gestational differences. It is unlikely that carrier status for α thalassaemia is the sole reason for these lower values, and some lifestyle factor is probably responsible. Alternatively, there may be a true genetic effect.

One novel finding from this analysis is that season of birth was associated with % HbF. This finding is not a result of confounding by the other variables. Birthweight, for example, was unrelated to month of birth (mean birthweight=3.3 kg for 10 of the 12 months) and could not, therefore, have produced such an effect. Findings from the second (and third) periods of data collection are not as convincing as those from the first, and it will be important for this finding to be replicated in other studies.

The size of the independent effect of variables on % HbF may be assessed by comparing the values of the adjusted regression coefficients. Similar influences are exerted by differences in gestation of one week, birthweight of 1 kg, date of birth of three months, as well as sex of infant and twin delivery. About double this effect is associated with Indo-European ethnic origin. For infants born near to 40 weeks' gestation, the 'influence' of an additional one week of gestation is a lowering of % HbF by about 2%.

Findings for sex, gestational age, and birthweight are consistent with those reported by Galacteros et al⁶ for % HbA in newborn infants; factors associated with higher values of % HbF in the study reported here were associated with lower values of HbA in the French study.

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