# FastMap: Fast Association Mapping in Inbred Mouse Populations Supplementary Materials

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## Equivalence of different test statistics

In this section we establish the equivalence of several test statistics  $(r^2, t^2, F, and LR)$  commonly used to test the associations between a binary vector, resulting from a homozygous SNP, and a real-valued vector, resulting from measurements of gene expression. In particular, we note that

$$F = t^2 = \frac{r^2}{1 - r^2}(n - 2)$$
 and  $LR = -\log(1 - r^2)$ 

#### Notation

Let S = (0, 0, ..., 0, 1, 1, ..., 1) be a binary vector of length n, with  $n_0$  zeros and  $n_1$  ones, and let  $G = (g_1, g_2, ..., g_n)$  be an *n*-vector with real-valued components. All four statistics we will consider are scale invariant, so without

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loss of generality we can assume that G is standardized:

$$\sum_{i=1}^{n} g_i = 0$$
 and  $\sum_{i=1}^{n} g_i^2 = 1.$ 

Denote by  $\bar{g}_0$  and  $\bar{g}_1$  the average of G over those samples where S equal to 0 and 1, respectively:

$$\bar{g}_0 = \sum_{i=1}^{n_0} g_i / n_0$$
 and  $\bar{g}_1 = \sum_{i=n_0+1}^n g_i / n_1.$ 

It follows from  $\sum_{i=1}^{n} g_i = 0$  that  $n_0 \bar{g}_0 = -n_1 \bar{g}_1$ , and in particular,  $n_0^2 \bar{g}_0^2 = n_1^2 \bar{g}_1^2$ .

#### Correlation

The correlation of G and S is defined as

$$r = \frac{\operatorname{cov}(G, S)}{\sqrt{\operatorname{var}(S)\operatorname{var}(G)}}.$$

The numerator is equal to

$$\operatorname{cov}(G,S) = \frac{1}{n-1} \sum_{i=1}^{n} s_i g_i = \frac{1}{n-1} \sum_{s_i=1}^{n} g_i = \bar{g}_1 n_1 / (n-1).$$

The first term in the denominator simplifies to

$$\operatorname{var}(S) = \frac{1}{n-1} \left[ \sum s_i^2 - \frac{1}{n} (\sum s_i)^2 \right] = \frac{1}{n-1} \left[ n_1 - \frac{n_1^2}{n} \right] = \frac{n_1(n-n_1)}{(n-1)n} = \frac{n_0 n_1}{(n-1)n},$$

and the second is equal to

$$\operatorname{var}(G) = \frac{1}{n-1} \left[ \sum g_i^2 - \frac{1}{n} (\sum g_i)^2 \right] = \frac{1}{n-1}$$

Combining the previous three expressions, we find that

$$r = \frac{\bar{g}_1 n_1}{\sqrt{n_0 n_1/n}} = \bar{g}_1 \sqrt{\frac{n_1}{n_0} n}.$$

#### **T-statistic**

The t-statistic for difference between  $\bar{g}_0$  and  $\bar{g}_1$  is defined as

$$t = \frac{\bar{g}_1 - \bar{g}_0}{\hat{\sigma}\sqrt{\frac{1}{n_1} + \frac{1}{n_0}}},$$

where the estimate of the noise variance is

$$\hat{\sigma}^2 = \frac{1}{n-2} \left( \sum_{i=1}^{n_0} (g_i - \bar{g}_0)^2 + \sum_{i=n_0+1}^n (g_i - \bar{g}_1)^2 \right)$$

Simplifying  $\hat{\sigma}^2$  yields

$$\hat{\sigma}^2 = \frac{1}{n-2} \left( \sum_{i=1}^n g_i^2 - n_0 \bar{g}_0^2 - n_1 \bar{g}_1^2 \right)$$

It then follows from the equations  $\sum_{i=1}^n g_i^2 = 1$  and  $n_0^2 \bar{g}_0^2 = n_1^2 \bar{g}_1^2$  that

$$\hat{\sigma}^2 = \frac{1}{n-2} \left( 1 - \frac{n_1(n_0 + n_1)}{n_0} \bar{g}_1^2 \right) = \frac{1}{n-2} \left( 1 - \bar{g}_1^2 \frac{n_1}{n_0} n \right) = \frac{1 - r^2}{n-2}$$

Note that

$$\sqrt{\frac{1}{n_1} + \frac{1}{n_0}} = \sqrt{\frac{n_0 + n_1}{n_0 n_1}} = \sqrt{\frac{n}{n_0 n_1}}$$

and that the numerator of t can be rewritten as

$$\bar{g}_1 - \bar{g}_0 = (1 + \frac{n_1}{n_0})\bar{g}_1 = \frac{n_0 + n_1}{n_0}\bar{g}_1 = \frac{n}{n_0}\bar{g}_1$$

Combining the analyses above we find that

$$t = \frac{\bar{g}_1 - \bar{g}_0}{\hat{\sigma}\sqrt{\frac{1}{n_1} + \frac{1}{n_0}}} = \frac{n\bar{g}_1/n_0}{\sqrt{\frac{n}{n_0n_1}\frac{1 - r^2}{(n-2)}}} = \frac{\bar{g}_1\sqrt{nn_1/n_0}}{1 - r^2}\sqrt{n-2} = \frac{r}{1 - r^2}\sqrt{n-2}$$

This is the well known formula for t-statistic for a correlation (under assumption of normality).

## **ANOVA F-statistic**

The ANOVA F-statistic measures the fraction of variation explained by assuming different averages of G across samples where S equal to 0 and 1,

$$F = (n-2)\frac{SSB}{SSW}$$

where

$$SSB = n_0 \bar{g}_0^2 + n_1 g_1^2$$

and

$$SSW = \sum_{i=1}^{n_0} (g_i - \bar{g}_0)^2 + \sum_{i=n_0+1}^n (g_i - \bar{g}_1)^2$$

Applying the formula derived for t-statistic  $\hat{\sigma}^2 = \frac{1-r^2}{n-2}$  we find that

$$SSW = (n-2)\hat{\sigma}^2 = 1 - r^2$$

It follows from  $SST = \sum_{i=1}^{n} g_i^2 = 1$  and SST = SSB + SSW that

$$SSB = SST - SSW = r^2$$

Combining the formulas for SSB and SSW yields the equivalence of F and t statistics

$$F = (n-2)\frac{r^2}{1-r^2} = t^2$$

### Likelihood ratio test

The likelihood ratio test is defined as doubled difference between data likelihoods under hull and alternative hypotheses

$$LR = 2[l_1(g, s) - l_0(g)]$$

The null model assumes the same normal distribution for samples  $g_i$  whether  $s_i = 0$  or  $s_i = 1$ . The likelihood under null model is

$$l_0(g) = -\frac{1}{2}\log 2\pi - \frac{1}{2}\log \hat{\sigma}_0^2 - \frac{\sum_{i=1}^n (g_i - \bar{g})^2}{2\hat{\sigma}_0^2}$$

where

$$\hat{\sigma}_0^2 = \sum_{i=1}^n (g_i - \bar{g})^2 / n$$

The alternative model allows the mean of  $g_i$  to depend on whether  $s_i = 0$  or  $s_i = 1$ . The likelihood under alternative model is

$$l_1(g,s) = -\frac{1}{2}\log 2\pi - \frac{1}{2}\log \hat{\sigma}_1^2 - \frac{\sum_{i=1}^{n_0}(g_i - \bar{g}_0)^2 + \sum_{i=n_0+1}^{n}(g_i - \bar{g}_1)^2}{2\hat{\sigma}_1^2}$$

where

$$\hat{\sigma}_1^2 = \left[\sum_{i=1}^{n_0} (g_i - \bar{g}_0)^2 + \sum_{i=n_0+1}^n (g_i - \bar{g}_1)^2\right] / n$$

We can simplify the estimates of variance

$$\hat{\sigma}_0^2 = \sum_{i=1}^n (g_i - \bar{g})^2 / n = 1/n$$

$$\hat{\sigma}_1^2 = \left[\sum_{i=1}^{n_0} (g_i - \bar{g}_0)^2 + \sum_{i=n_0+1}^n (g_i - \bar{g}_1)^2\right]/n = SSW/n = (1 - r^2)/n$$

We can simplify the likelihoods using the formulas for  $\hat{\sigma}_0^2$  and  $\hat{\sigma}_1^2$ 

$$l_0(g) = -\frac{1}{2}\log 2\pi - \frac{1}{2}\log[1/n] - \frac{n}{2}$$
$$l_1(g,s) = -\frac{1}{2}\log 2\pi - \frac{1}{2}\log[(1-r^2)/n] - \frac{n}{2}$$

This leads to the formula for the likelihood ratio test

$$LR = -\log(1 - r^2)$$

### Summary

We can conclude that all four test statistics  $(r^2, t^2, F, \text{and } LR)$  are monotone functions of each other and thus are equivalent:

$$F = t^2 = \frac{r^2}{1 - r^2}(n - 2)$$
 and  $LR = -\log(1 - r^2)$