# **Motif-blind, genome-wide discovery of cis-regulatory modules in Drosophila and mouse**

**(Supplementary Information)**







**Figure S1. [previous page].** (A) LOOCV performance of each of nine different methods on 31 data sets in benchmark. Color accents represent Nucleotidelevel sensitivity on a scale of 0 to 1, and cases with empirical p-value  $\leq$ 0.05 are marked by asterisks. The top row shows the number of such data sets for each method. Numbers to the right are best Nucleotide-level sensitivity for each data set. One-on-one comparison of methods: For each pair of methods M1 (row) and M2 (column), (B) the "wins" of M1 versus M2 (i.e., the number of data sets on which nucleotide-level sensitivity of M1 was greater than that of M2 by at least 10% of data set size). (C) The difference between the wins of M1 versus M2 and the wins of M2 versus M1 in nucleotide-level sensitivity. (D) Sixteen data sets on which at least one method has nucleotide level sensitivity p-value of 0.05 or less in all four instantiations. Color indicates the number of instantiations (out of four) on which the performance was significant ( $p \le 0.05$ ): white=4, yellow=3, orange=2, brown=1, black=0. (E) Comparison of single species and multispecies versions of HexMCD. For each each LOOCV instantiation, the Nucleotide-level sensitivity and number of amenable data sets at  $p \le 0.05$ (the value above each bar) are compared between the two methods.



**Figure S2. [previous page].** (A, B, C) The best CRM-level (left)/nucleotidelevel(right) sensitivity on a data set is plotted against three high level characteristics of the data set: extent of homotypic clustering as measured by the "FTT-Z rank" (A), G/C content of training CRMs (B) and evolutionary conservation of training CRMs (C). The Pearson's correlation coefficient (corr) and its p-value are reported on each plot.



**Figure S3. [previous page].** Dataset size and performance for amenable datasets in random training set evaluations. Red circles indicate datasets that performed no better with their appropriate training set than with a random training set; blue circles indicate datasets where random training sets led to worse performance. Near-perfect discrimination is achieved at n=33.5 (line).



**Figure S4. [previous page].** Cladogram view of the 200 6-mers used by HexYMF for the blastoderm data set. We find k-mers matching the consensus string of six of the canonical motifs of the segmentation network (Bicoid, Caudal, Hunchback, Dstat, Kruppel, Tailless, matching the 6-mers GGATTA, CATAAA, AAAAAC, CCGGGA, AAGGGT, AGTCAA respectively) and Dorsal, also a blastoderm stage-active transcription factor, matching GGGAAA.



**Figure S5. [previous page].** Cross-validation on mammalian enhancers. CRMlevel sensitivity is shown for each method on each data set. Values with empirical p-value  $< 0.05$  are shaded, and those at  $p < 0.01$  are in darker shade. Best performance on each data set is shown in bold entries.

**Table S1**. **[next page].** Relationship between performance and methodological features. For each methodological feature, two methods were evaluated – one that incorporates the feature (top) and one that does not (bottom). (The two methods are identical otherwise). The average CRM-level sensitivity of each method and the net number of wins of the second method (i.e., one with the methodological feature) are shown. The "net number of wins" is the difference in numbers of wins of the two methods.



**Table S2\*** . In-depth analysis of CRM predictions for the blastoderm dataset. For each of the 113 predicted modules (Release 4.3 coordinates in Column A), the table shows: (i) nearest neighbor gene and location of module relative to gene (column B), (ii) whether or not it is a known module (column C), and if so, (iii) whether or not it is a "training CRM" (column D), (iv) whether or not the nearest gene has "pair-rule" expression pattern (column E), or "dorso-ventral (D/V)" expression pattern (column F), and (v) for the modules that do not overlap known modules (rows 62-115), the motifs associated with the module (columns G-AB). For modules that do overlap known modules, column AC names that known module, and for the subset of these that were not in the training set, column AD notes the tissue where the known module drives expression. The key word "long" in this column indicates that the reason for the module not being in the training set was its large length (modules longer than a threshold length were excluded from the training sets).

**Table S3\*** . Cross-validation performance (complete data). The CRM and nucleotide level sensitivity and corresponding p-values of all methods in the four instantiations of cross-validation. Also included are the average sensitivity values over all instantiations.

**Table S4**\* **.** Worksheets "HexYMF-s200-rc" and "PAC-rc" represent the genome-wide predictions in human for the early cardiovascular development data set, using these two methods respectively. Column A names the evolutionarily conserved region (ECR) in which the predicted module is located, with the start position of the window being given by column B. The score of the module is in column C, and information on adjacent gene(s) is

<sup>\*</sup> These tables are included as separate Excel files: TableS2.xls, TableS3.xls, TableS4.xls, TableS5.xls

in column D. The fraction of the module's sequence that was identified as being short tandem repeats is shown in column E. Columns F and G indicate if the GATA motif or the Ets motif were found to be present in this module (and if so, which of the many variants of these motifs from TRANSFAC were present). The last two columns indicate the rank of the module by each method, based on the raw score (i.e., before modules were re-ranked based on other factors such as presence of at least one motif and the amount of short tandem repeat coverage).

**Table S5\* .** CRM-level and nucleotide level sensitivity values for each of seven motif-blind methods on eight data sets of mouse enhancers.

Worksheet "localflanks.crm.sens": CRM-level sensitivity values. Worksheet "localflanks.crm.pval": Empirical p-values of CRM-level sensitivity. Worksheet "localflanks.nuc.sens": Nucleotide-level sensitivity values. Worksheet "localflanks.crm.pval": Empirical p-values of nucleotide-level sensitivity.

**Table S6\* .** Cross-validation performance in mouse enhancers (complete data). The CRM and nucleotide level sensitivity and corresponding p-values of all methods in all 10 instantiations of cross-validation.

These tables are included as separate Excel files: TableS5.xls, TableS6.xls

# **Table S7.** Individual anatomy terms included in each data set of REDfly CRMs



antennal\_disc embryonic\_pericardial\_cell clypeo-labral\_disc embryonic\_somatic\_muscle eye\_disc embryonic\_visceral\_muscle eye-antennal\_disc fat\_body eye-antennal\_disc\_specific\_anlage fat\_body/gonad\_primordium genital\_disc density and the mocyte\_primordium haltere disc and a set of larval dorsal vessel imaginal\_disk and the state of the larval state of the larval fat\_body labial\_disc larval\_heart larval\_heart male\_genital\_disc larval\_muscle\_system ventral\_thoracic\_disc larval\_pericardial\_cell ventral\_thoracic\_disc\_specific\_anlage larval\_somatic\_muscle visual\_primordium data and the larval\_visceral\_muscle wing disc local longitudinal visceral mesoderm

ectoderm mesoderm mesode ectoderm\_anlage\_in\_statu\_nascendi | midgut\_muscle embryonic epidermis myocardial cell

anterior\_endodermal\_primordium entertial\_muscle tergotrochanteral\_muscle anterior\_midgut\_proper\_primordium deterior dirunk\_mesoderm\_anlage embryonic\_hindgut demonstration of the trunk\_mesodermal\_primordium embryonic\_midgut dubular\_muscle foregut\_primordium ventral\_plate foregut\_specific\_anlage visceral\_mesoderm larval midgut visceral muscle primordium midgut\_constriction\_1 pHiGP2 **neuronal (n=31):** posterior\_embryonic\_midgut RP2\_neuron posterior\_endoderm\_anlage U\_neuron posterior\_endodermal\_primordium anterior\_corner\_cell posterior\_midgut\_proper\_primordium chemosensory\_sensory\_organ

mesoderm embryonic central brain qlia

ventral\_neuroectoderm embryonic\_neuroblast

chemosensory\_sensory\_organ embryonic\_ventral\_nervous\_system embryonic\_peripheral\_nervous\_system interneuron mesothoracic\_femoral\_chordotonal\_organ | larval\_central\_nervous\_system metathoracic\_femoral\_chordotonal\_organ larval\_mushroom\_body peripheral\_nervous\_system larval\_optic\_lobe

 longitudinal\_visceral\_muscle\_fibers **ectoderm.1 (n=4): and integral in the set of the set o** ectoderm\_anlage mesoderm\_anlage\_in\_statu\_nascendi somatic\_mesoderm **endoderm (n=13):** somatic muscle specific anlage

 embryonic\_brain mesoderm.1 (n=2): embryonic\_central\_brain mesoderm\_anlage\_in\_statu\_nascendi embryonic\_central\_brain\_mushroom\_body embryonic\_central\_nervous\_system **neuroectoderm (n=2):** embryonic\_ganglion\_mother\_cell neuroectoderm embryonic nervous system embryonic\_neuron **pns (n=15):** embryonic\_peripheral\_nervous\_system

prothoracic\_femoral\_chordotonal\_organ mesothoracic\_femoral\_chordotonal\_organ scolopophorous\_organ metathoracic\_femoral\_chordotonal\_organ metathoracic\_femoral\_chordotonal\_organ sensillum\_campaniformium\_of\_anterior\_crossvein | neuroblast\_NB6-4 sensillum\_campaniformium\_of\_dorsal\_radius | neuron sensillum\_campaniformium\_of\_dorsal\_radius\_Sc1 peripheral\_nervous\_system sensory\_organ\_precursor\_cell prothoracic\_femoral\_chordotonal\_organ twin\_sensillum\_of\_margin\_1 scolopophorous\_organ twin\_sensillum\_of\_margin\_2 sensory\_organ\_precursor\_cell ventral\_cibarial\_sense\_organ thoracico-abdominal\_ganglion ventral\_pit ventral\_cibarial\_sense\_organ

### **tracheal-system (n=7):** ventral\_unpaired\_cell

embryonic\_spiracle embryonic trachea **wing (n=7):** posterior\_spiracle\_primordium twin\_sensillum\_of\_margin\_1 tracheal\_primordium tracheal\_primordium twin\_sensillum\_of\_margin\_2

ventral\_pit

embryonic\_tracheal\_pit sensillum\_campaniformium\_of\_anterior\_crossvein embryonic\_tracheal\_system sensillum\_campaniformium\_of\_dorsal\_radius filzkorper sensillum\_campaniformium\_of\_dorsal\_radius\_Sc1 wing the contract of the contr wing\_disc

**Table S8.** [next page]. <sup>a</sup> refers to percentage of CRMs that drove gene expression in the predicted pattern. na=not applicable to this method  $b$  5/7 (71%) gave enhancer activity, 1/7 was found to be a negative regulatory element

Note: Only studies in which at least 5 sequences were validated are included, and only those with true in vivo (not just cell culture) testing.



**Table S9.** Primer sequences used for Drosophila and mouse construct

### **Primer sequences used for** *Drosophila* **constructs**

edl

F: CTATATTATTCCACGTTTTC

R: TGTCCAAGACGATTCTTAT

odd

- F: GTTTTCAAATAAAATTACAATG
- R: GCTAGGACGCAGAGCTG

srp

- F: GGGAATTCCATTTATCTGA
- R: CTTTTAATGCAACAATAAGC

SoxN

- F: ATTACTTCGACTCCAGTC
- R: CTTTCCTTTTAATCAAAGTG

cas

F: CATAAATATAATCAAATCTTAG

R: ATCGTACTCCGCCCCT

### **Primer sequences used for mouse constructs**



EBF3Rv TAAGATATCCGGTGGGCTATTGTTATAGGG

# **Table S10.** Search Criteria used for defining expression gene sets

### **Blastoderm** (BDGP)

select stages 1-6 only genes with blastoderm pattern as used in Halfon et. al (2002).

# **cardiac\_mesoderm** (BDGP)



## **cns** (BDGP)



**eye.1** (FlyBase)

FBbt: 00004508 (eye)

# **imaginal\_disc.1** (FlyBase)

FBbt: 00001761 (imaginal disc)

# **mesectoderm** (BDGP)



# **mesoderm.1** (BDGP)





# **neuroectoderm** (FlyBase)



# **pns** (BDGP)



# **somatic muscle** (FlyBase)



# **ventral\_ectoderm** (BDGP)



# **visceral\_mesoderm** (BDGP)



# **ectoderm.2** (BDGP)



## **neuronal** (BDGP)



**Supplementary Note S1.** General-level characterization of various scoring schemes evaluated here.





**Supplementary Note S2.** Discussion on the choice of  $p \leq 0.05$  as the threshold for statistically significant sensitivity

Note that Figure 1A marks a cell with a "\*" whenever the sensitivity of the method on the data set has p-value  $\leq$  0.05. Given that we are testing 9 methods for each data set, there is an issue similar to "multiple hypotheses correction", and a concern whether 0.05 is "too liberal" a choice.

- Firstly, most (i.e.,  $90/118$ ) of the p-values corresponding to the "\*" cells are actually  $\leq 0.01$ . (See Supplementary Table S3, worksheet "crm\_pval.1"; the "< $0.05$ " and "< $0.01$ " cases are highlighted in darker and lighter shades respectively.).
- Secondly, we believe that 0.05 is a reasonable threshold for assessing significant performance of a given method on a given data set. The problem arises when we make global claims about the number of significant cells in Figure 1A; such claims must correct for "multiple hypotheses testing". The first such "global" claim we have made is that "… the best motif-blind methods … succeed on close to half of the data sets." We have checked that this is true if we use Q-values of Storey & Tibshirani (PNAS 2003) for multiple hypotheses correction. The q-value is, loosely speaking, the expected false discovery proportion corresponding to a p-value. For the three methods mentioned in context, the p-value threshold of 0.05 corresponds to a q-value threshold of 0.02, 0.10 and 0.07 respectively, thus validating our claim even after multiple hypotheses correction. The only other global claim we have made about the matrix of p-values (Figure 1A) is that 15 of the 31 data sets are amenable to at least one method (among the eight listed). For this claim,

we have required that the p-value threshold (of 0.05) be met by the method on each of four independent instantiations of cross-validation, thus effectively setting a much stricter threshold on each (method, data set) combination.

Finally, as a "back-of-the-envelope" calculation, we point out that there are 118 cells marked with a "\*" in Figure 1(A), while at  $p \le 0.05$ , one would expect 0.05 x 9 x 31 = 14 such cells. Similarly, there are 90 cells with  $p \le$ 0.01 (Supplementary Table S3), while only 3 are expected.

**Supplementary Note S3.** Motif-based prediction using an externally developed suite of algorithms

We have deployed our new method called "Stubb-MDB" as a point of comparison to the motif-blind approaches that are the main topic of this paper. Results showing that Stubb-MDB performs relatively poorly compared to the motif-blind methods tested here, may be found in Figure 1. In order to have great confidence in this comparison between motif-based and motif-blind approaches, we evaluated an alternative motif-based pipeline – that of using "Clover" for motif selection and "ClusterBuster" for scanning with selected motifs. Both programs are developed by Zhiping Weng and colleagues and run with default. The motif compendium used was the same as that for Stubb-MDB. ClusterBuster was run with default parameters. We found success levels to be comparable to that of Stubb-MDB. (See Supplementary Table S3, worksheets crm\_sens.1 and crm\_pval.1, showing the CRM-level sensitivity and p-values respectively with each method for each data set, on one of the cross-validation instantiations.) Here is a brief summary of the results of Stubb-MDB and Clover-ClusterBuster:



**Supplementary Note S4.** Difference between our benchmark and that of (Ivan et al., 2008).

In our new benchmark each test CRM is planted in a genomic sequence of length 10 Kbp, with G/C content similar to the native flank of the CRM (called target sequence). Therefore the search space for CRM is 10kb+CRM length. This is harder benchmark than our previous benchmark introduced in Ivan et. al (2008) in which each CRM is embedded in a genomic sequence of length 10 times CRM length. Since all of our datasets except cns, imaginaldisc-2, tracheal system, and wing have average CRM length of less than 1kb, the search space is usually shorter than the search space in our new benchmark.

**Supplementary Note S5.** Web Interface for the results, training datasets and sources



Download all training erms from here

The web interface<sup>2</sup> includes (i) a link to predicted CRMs for each dataset which contains some information about the CRM (e.g. location, nearby gene, chromosome number, corresponding link to genome surveyor, flybase, BDGP and flyexp), the number of presented motifs, and link to presented motifs' logo. (ii) The genesets that have been used in this report (the search criteria used for defining the genesets are itemized in the next section) (iii) the pipeline source code that includes the implementation of all methods described in the paper and (iv) The training crms.

<sup>&</sup>lt;sup>2</sup> http://veda.cs.uiuc.edu/scrm/index.htm

# **Supplementary Note S6.** Novel CRMs in blastoderm data set

We report here on the overlap between the 113 modules predicted for the blastoderm data set by our pipeline and the A/P patterning modules predicted by a previous motif-based method (Sinha et al., 2004) and a non motif-based method (Grad et al., 2004). Note that the latter was based on two-species comparisons, while the predictions (considered here) from (Sinha et al., 2004) as well as our current work are based on D. melanogaster alone.

- 1) We have noted in text that the 5 tested fly CRMs were not reported by Ahab.
- 2) We considered the list of predictions made using Stubb (single species) in (Sinha et al., 2004).
- 3) We considered the list of predictions made by (Grad et al., 2004).

Our "novel" set of 54 predictions had very little overlap with the predictions of (Grad et al., 2004). There was a more substantial overlap (23/54) when considering all 2964 module predictions made by Stubb-SS in (Sinha et al., 2004). However, when we consider only the top 113 modules predicted by Stubb (i.e., the same number of predictions as our method), there is an overlap of only two modules. In other words, our novel set largely consists of unreported modules, although many of these were predicted at relatively low confidence levels in (Sinha et al., 2004) (Stubb). In any case, we designate these 54 modules as novel since none of them have been experimentally characterized previously.



**Supplementary Note S7.** Relationship between performance and methodological features

We performed pair-wise comparison of different variants of the D2-z score that differed in only one aspect (such as Markov order, or use of subsets of motifs, etc.), by counting the number of data sets (over all four instantiations) on which one method outperforms the other. The results are shown in Table S1 (Supplementary Materials). The most interesting observations from this table are that we get significantly better performance overall by (1) using a subset of words instead of all possible k-mers, (2) calculating conditional zscores (of D2) rather than using unconditional z-scores as done in (cite), (3) keeping the CRMs in the training set separate, rather than concatenating them into one sequence, (4) using a first-order Markov order for the background model rather than an iid model, (5) counting words on both strands. Note that these features needed to be investigated only in the context of the D2-z score; the other scores such as HexDiff, HexYMF and PAC, do not face issues  $(1) - (4)$ , and HexDiff and HexYMF do count words on both strands. On the other hand, the published D2-z score does not handle issues (1), (2), (3) and (5), and allows both iid and Markov backgrounds (issue 4) without prescribing which is better for our application.

**Supplementary Note S8.** The ten enhancers known to function in the developing blood and vasculature were taken from: (Chan et al., 2007; Chapman et al., 2003; De Val et al., 2004; Donaldson et al., 2005; Gottgens et al., 2002; Landry et al., 2008; Pimanda et al., 2007a; Pimanda et al., 2007b).

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# Supplementary Methods

March 10, 2009

## 1 Alignment free similarity measures

In this section we describe  $PAC$  and compute the z-score for the various modifications of the  $D2$ statistic used in the text.

### 1.1 A Poisson based similarity measure,  $PAC$ .

This measure is based on [6]. Let A be a random DNA sequence. Fix a word length k. For a k-word w, let  $N_w^A$  be the count of the word w in the sequence A, including overlaps. A simple model for  $N_w^A$  is the binomial distribution, which is approximately a Poisson distribution in our applications. This approach assumes an iid background model. It does not take into account the dependency of occurrences of the word  $w$  in overlapping positions. Also, because of this independence assumption, the count of occurrences of a word  $w$  is done on one strand only.

The parameters for the Poisson distributions can be computed as follows. Let  $f_w^A$  be the prior probability of seeing the word  $w$  at any position of the sequence  $A$ . The expected number of occurrences is then  $m_w^A = f_w(|A| - k + 1)$ , where |A| is the length of the sequence A. (Note: in the main text we used the notation  $\lambda(w)$  instead of  $m_w$ .

Now suppose  $A$  and  $B$  are two DNA sequences. Define

$$
C_w^{AB} = \min(N_w^A, N_w^B).
$$

Then, assuming the sequences  $A$  and  $B$  are generated independently,

$$
\Pr(C_w^{AB} \ge c) = \begin{cases} (1 - F(m_w^A, c - 1))(1 - F(m_w^B, c - 1)) & \text{if } c > 0, \\ 1 & \text{otherwise,} \end{cases}
$$

where  $F(m, -)$  is the Poisson distribution function with mean m.

Now define a similarity measure as follows. Let

$$
s_w^{AB} = 1 - \Pr(C_w^{AB} \ge c_w^{AB}),
$$

where  $c_w^{AB}$  is the observed value of  $C_w^{AB}$ .

For a set W of k-words, the *additive Poisson* similarity measure between the sequences  $A$  and  $B$  is defined in [6] to be

$$
S_{add}^{AB} = 1/|W| \sum_{w \in W} s_w^{AB}.
$$
\n<sup>(1)</sup>

In our setting of LOOCV we have a dataset of n known CRMs planted in n random sequences. For each fold, we fix the sequence A to be a concatenation of  $n-1$  CRMs. Sequence B is then a sliding window along the test sequence, where the window length is the average length of the CRMs in the given dataset. In this setting, the length of sequence A is so much longer than the length of B that in effect,  $\min(N_w^A, N_w^B) = N_w^B$ . Hence, the above Poisson measure was simplified to

$$
PAC(B|A) = 1/|W| \sum_{w \in W} F(m_w^B, n_w^B - 1),
$$
\n(2)

where  $n_w^B$  is the observed value of  $N_w^B$ . (Note: in the main text, we used the notation  $n_B(w)$  instead of  $n_w^B$ ).

The abbreviation *PAC* stands for "Poisson Additive Conditional", since the score is computed given that the sequence A is known. Note that the input from sequence A in  $PAC$  is only captured during the process of choosing the set of  $k$ -words  $W$ . For this we used the YMF motif finding program [5] and took the top 100 scoring k-words.

### 1.2  $D2z$  score

The  $D2z$  score was developed in [1]. It is a z-score for the D2 statistic [3]

$$
D2(A,B) = \langle \mathbf{N}^A, \mathbf{N}^B \rangle, \tag{3}
$$

where  $\mathbf{N}^A$  ( $\mathbf{N}^B$ ) is the count vector for the k-words in A (B), and  $\langle$  -, - > is the inner product operator. The z-score is then

$$
D2z(A,B) = \frac{D2(A,B) - E(D2(A,B))}{\sigma(D2(A,B))},
$$
\n(4)

where the (theoretical) expectation,  $E(D2(A, B))$ , and standard deviation,  $\sigma(D2(A, B))$ , are computed under the assumption that the random sequences  $A$  and  $B$  are independent and were generated by two, possibly different, Markov processes of any order (order zero means iid). The parameters of the Markov model are learned from the data.

#### 1.3 Conditional D2z score, cD2z.

When the sequence A is fixed (e.g. concatenated known CRMs), the formulas for the expectation and variance given in [1] are not valid. Here we compute the mean and the variance for  $D2(A, B)$ where  $A$  is fixed and  $B$  is a random sequence generated by a Markov process. We call the corresponding z-score,  $cD2z(B|A)$ , a "conditional D2z" score. (Note: in the main text the  $cD2z$  score is denoted by D2z-cond).

In what follows we use the following notation.

**Notation.** Let  $n_1 = |A|$  be the length of sequence A and let  $n_2 = |B|$ . Write  $\bar{n_1} = n_1 - k + 1$ . and  $n_2 = n_2 - k + 1$ . Let I be the index set

$$
I = \{(i, j) \in \mathbb{N} \times \mathbb{N} : 1 \le i \le \bar{n}_1, \text{ and } 1 \le j \le \bar{n}_2\}.
$$

For indices  $i \leq c$ , use the notation  $A[i, c]$  for the subsequence of length  $c - i + 1$ ,  $A_i A_{i+1} \ldots A_c$ .

For a k-word w and an index  $m < k$ , define  $suf_m(w)$  to be the last m letters in w,

 $suf_m(w) = w_{k-m+1} \dots w_k.$ 

Define  $pre_m(w)$  to be the first m letters in w,

$$
pre_m(w) = w_1 \dots w_m.
$$

IID case.

Proposition 1.1. Under the IID model,

$$
E(cD2(B|A)) = \bar{n_2} \sum_{|w|=k} N_w^A \Pr^B(w), \tag{5}
$$

where  $Pr^{B}(w)$  is the probability of seeing the word w at any position in the sequence B. Proof. The D2 statistic in (3) can be computed via

$$
D2(A,B) = \sum_{i=1}^{\bar{n_1}} \sum_{j=1}^{\bar{n_2}} Y_{(i,j)},
$$
\n(6)

where  $Y_{(i,j)}$  is the k-word match indicator at position i in A and j in B,

$$
Y_{(i,j)} = \begin{cases} 1 & \text{if } A[i, i+k-1] = B[j, j+k-1], \\ 0 & \text{otherwise.} \end{cases}
$$

Then

$$
E(cD2(B|A)) = \sum_{i=1}^{\overline{n_1}} \sum_{j=1}^{\overline{n_2}} E(Y_{ij}|A) = \sum_{i=1}^{\overline{n_1}} \sum_{j=1}^{\overline{n_2}} \Pr^B(A[i, i+k-1])
$$
  
=  $\overline{n_2} \sum_{i=1}^{\overline{n_1}} \Pr^B(A[i, i+k-1]) = \overline{n_2} \sum_{|w|=k} N_w^A \Pr^B(w).$ 

Theorem 1.2. Under the IID model,

$$
\operatorname{Var}(cD2(B|A)) =
$$
\n
$$
2 \sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \sum_{|v|=m} (\sum_{\substack{|w|=k \ suf_m(w)=v}} N_w^A \operatorname{Pr}^B(w)) (\sum_{\substack{|u|=k \ yre_m(u)=v}} N_u^A \operatorname{Pr}^B(suf_{k-m}(u)))
$$
\n
$$
+ \bar{n_2} \sum_{|w|=k} (N_w^A)^2 \operatorname{Pr}^B(w)
$$
\n
$$
- 2 \sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \left( \sum_{|w|=k} N_w^A \operatorname{Pr}^B(w) \right)^2
$$
\n
$$
- \bar{n_2} \left( \sum_{|w|=k} N_w^A \operatorname{Pr}^B(w) \right)^2
$$

Note that the above formula can be computed in  $O(4^k)$  (the count vectors for the fixed sequence A are precomputed).

Proof.

$$
Var(D_2) = Var(\sum_{(i,j)\in I} Y_{(i,j)}) = \sum_{(i,j),(s,t)\in I} Cov(Y_{(i,j)}, Y_{(s,t)}).
$$

case 1: no overlap in B.

In this case, since we are in the iid case,  $Y_{(i,j)}$  and  $Y_{(s,t)}$  are independent and hence  $Cov(Y_{(i,j)}, Y_{(s,t)}) =$ 0.

case 2: overlap in B.

Let  $m$  be the size of the overlap. Then

$$
E(Y_{(i,j)}Y_{(s,t)}) = \begin{cases} \Pr^{B}(A[i, i+k-1])\Pr^{B}(A[s+m, s+k-1]) \\ \text{if } A[i+k-m, i+k-1] = A[s, s+m-1], \\ 0 \qquad \text{otherwise.} \end{cases}
$$
(7)

Hence, for a fixed j and  $t = j + k - m$  (i.e. have overlap of size m in B),

$$
\sum_{i=1}^{\tilde{n_1}} \sum_{s=1}^{\tilde{n_1}} E(Y_{(i,j)}Y_{(s,t)}) =
$$
\n
$$
\sum_{|v|=m} (\sum_{\substack{|w|=k \ suf_m(w)=v}} N_w^A \Pr^B(w)) (\sum_{\substack{|u|=k \ v\neq m(u)=v}} N_u^A \Pr^B(suf_{k-m}(u)))
$$

Add them all up:

$$
\sum_{m=1}^{k} \sum_{\substack{j,t \ j \in [-1] \\ |t-j|=k-m}} \sum_{i=1}^{n} \sum_{s=1}^{n} E(Y_{(i,j)}Y_{(s,t)}) =
$$
\n
$$
2 \sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \sum_{|v|=m} \sum_{\substack{|w|=k \\ suf_m(w)=v}} N_w^A \Pr^B(w) \Big( \sum_{\substack{|u|=k \\ pre_m(u)=v}} N_w^A \Pr^B(suf_{k-m}(u)) \Big)
$$
\n
$$
+ \bar{n_2} \sum_{|w|=k} (N_w^A)^2 \Pr^B(w).
$$

Finally subtract all the terms  $E(Y_{(i,j)})E(Y_{(s,t)})$  with overlap in B:

$$
\sum_{m=1}^{k} \sum_{\substack{j,t \ n=1}} \sum_{\substack{j,t \ n=1}}^{i \in I} \sum_{s=1}^{i \in I} E(Y_{(i,j)} E(Y_{(s,t)})) =
$$
\n
$$
\sum_{m=1}^{k} \sum_{\substack{j,t \ n=1}} \sum_{\substack{j,t \ n=1}} \left( \sum_{s=1}^{\overline{n_1}} \Pr^{B}(A[s, s+k-1]) \right) \left( \sum_{i=1}^{\overline{n_1}} \Pr^{B}(A[i, i+k-1]) \right) =
$$
\n
$$
\sum_{m=1}^{k} \sum_{\substack{j,t \ n=1}} \sum_{\substack{j,t \ n=1}} \left( \sum_{|w|=k} N_w^A \Pr^B(w) \right)^2 =
$$
\n
$$
2 \sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \left( \sum_{|w|=k} N_w^A \Pr^B(w) \right)^2 + \bar{n_2} \left( \sum_{|w|=k} N_w^A \Pr^B(w) \right)^2.
$$

Higher order Markov model case. Here we state and prove the formulas for the mean and variance for cD2z under Markov model (MM) of order 1. Extension to higher order MM is clear. We follow notation from [2] and [5]. For a DNA sequence B, and a nucleotide b let  $p_j^B(b)$  be the probability of a b in position j. As justified in [2], we may assume  $p_j^B$  is independent of j and denote it by  $p^B$ . The probabilities  $\{p^B(b)\}\$ are calculated as the steady state probabilities of the background Markov model. For a k-word  $w = w_1w_2 \ldots w_k$ , write  $p_*^B(w)$  for  $Pr(w|w_1)$ , i.e., the probability of a k-word being w, given that its first letter is  $w_1$ . With this notation,

$$
\Pr^{B}(w) = p^{B}(w_{1})p_{*}^{B}(w)
$$
\n
$$
(8)
$$

Then the expectation of  $cD2(-|A)$  can be expressed as follows.

Proposition 1.3. Assume the sequence B is generated by a Markov model of order one. Then

$$
E(cD2(B|A)) = \bar{n_2} \sum_{|w|=k} N_w^A \Pr^B(w) = \bar{n_2} \sum_{|w|=k} N_w^A p^B(w_1) p_*^B(w.). \tag{9}
$$

Proof. Same as Proposition 1.1, and use equation (8).

In the next Theorem, we give a formula for the variance of  $cD2$ . We will use the following notation.

**Notation.** For two nucleotides, b and c, and an index  $l > 0$  let  $p_l^B(c|b)$  be the probability of seeing  $c, l$  steps in the Markov chain after  $b$ . Let

$$
S_{b,c}^{B} = \sum_{j=1}^{q_2-1} \sum_{t=j+k}^{\bar{n_2}} p_t^{B}(c|b).
$$

where  $l = s - (i + k - 1)$  and  $q_2 = n_2 - 2k + 2$ .

 $\Box$ 

 $\Box$ 

Theorem 1.4. Under Markov model of order one,

$$
\begin{split}\n\text{Var}(cD2(B|A)) &= \\
2 \sum_{w_k, u_1} S_{w_k, u_1}^B \sum_{w_1, \dots, w_{k-1}} \text{Pr}^B(w_1 w_2 \dots w_k) N_{w_1 w_2 \dots w_k}^A \sum_{u_2, \dots, u_k} p_*^B(u_1 u_2 \dots u_k) N_{u_1 u_2 \dots u_k}^A \\
&+ 2 \sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \sum_{|v|=m} \sum_{\substack{|w|=k \text{ suf}_m(w)=v}} N_w^A \text{Pr}^B(w) \sum_{\substack{|u|=k \text{ suf}_m(w)=v}} N_w^A p^B(u_{m+1} | v_m) p_*^B(suf_{k-m}(u)) \\
&+ \bar{n_2} \sum_{|w|=k} (N_w^A)^2 \text{Pr}^B(w) \\
&- \bar{n_2}^2 \left( \sum_{|w|=k} N_w^A \text{Pr}^B(w) \right)^2\n\end{split} \tag{10}
$$

Proof.

$$
Var(D_2) = Var(\sum_{(i,j)\in I} Y_{(i,j)}) = \sum_{(i,j),(s,t)\in I} Cov(Y_{(i,j)}, Y_{(s,t)}).
$$

case 1: no overlap in B.

Unlike the iid case,  $Y_{(i,j)}$  and  $Y_{(s,t)}$  may be dependent under the Markov Model. To compute  $Cov(Y_{(i,j)}, Y_{(s,t)})$ , we consider the following. For a fixed  $j < t$ , with  $t = j + k - 1 + l$  and  $l > 0$  (i.e. no overlap in  $\tilde{B}$  and l is the "gap" from the last letter of the first word in  $B$  to the the first letter of the second word in  $B$ ),

$$
\sum_{i=1}^{\overline{n_1}} \sum_{s=1}^{\overline{n_1}} E(Y_{(i,j)}Y_{(s,t)}) =
$$
\n
$$
\sum_{i=1}^{\overline{n_1}} \sum_{s=1}^{\overline{n_1}} \Pr(B[j,j+k-1] = A[i,i+k-1], B[t,t+k-1] = A[s,s+k-1]) =
$$
\n
$$
\sum_{i=1}^{\overline{n_1}} \sum_{s=1}^{\overline{n_1}} p^B(A_i) p_*^B(A[i,i+k-1]) p_*^B(A[s,s+k-1]) p_l^B(A_s|A_{i+k-1}) =
$$
\n
$$
\sum_{|w|=k} \sum_{|w|=k} N_w^A N_w^A p^B(w_1) p_*^B(w) p_*^B(w) p_l^B(u_1|w_k).
$$

Sum up over all such  $j$ 's and  $t$ 's:

$$
\sum_{\substack{j,t\\|t-j| \ge k}} \sum_{i=1}^{\tilde{n_1}} \sum_{s=1}^{\tilde{n_1}} E(Y_{(i,j)} Y_{(s,t)}) =
$$
\n
$$
2 \sum_{j=1}^{q_2-1} \sum_{t=j+k}^{\tilde{n_2}} \sum_{|w|=k} \sum_{|u|=k} N_w^A N_u^A p^B(w_1) p_*^B(w) p_*^B(w) p_t^B(u_1|w_k) =
$$
\n
$$
2 \sum_{|w|=k} \sum_{|u|=k} N_w^A N_u^A p^B(w_1) p_*^B(w) p_*^B(u) \sum_{j=1}^{q_2-1} \sum_{l=1}^{q_2-j} p_l^B(u_1|w_k) =
$$
\n
$$
2 \sum_{|w|=k} \sum_{|u|=k} N_w^A N_u^A p^B(w_1) p_*^B(w) p_*^B(u) S_{w_k, u_1}^B.
$$
\n
$$
(11)
$$

The computational complexity of the formula above, as a function of k, is  $O(4^{2k})$ . To reduce this to  $O(4^k)$  we split the sum as follows.

$$
2\sum_{|w|=k} \sum_{|u|=k} N_w^A N_u^A p^B(w_1) p_*^B(w) p_*^B(u) S_{w_k, u_1}^B =
$$
  

$$
2\sum_{w_k, u_1} S_{w_k, u_1}^B \sum_{w_1, \dots, w_{k-1}} \Pr^B(w_1 w_2 \dots w_k) N_{w_1 w_2 \dots w_k}^A \sum_{u_2, \dots, u_k} p_*^B(u_1 u_2 \dots u_k) N_{u_1 u_2 \dots u_k}^A.
$$

This gives the first term of equation (10).

case 2: overlap in B.

The next two terms in formula (10) come from  $E(Y_{(i,j)}Y_{(s,t)})$  where the k-words in sequence B overlap. The proof follows the proof of the IID case (Theorem 1.2) with the obvious changes for the Markov model.

Finally subtract all the terms  $E(Y_{(i,j)}E(Y_{(s,t)}))$ :

$$
\sum_{i=1}^{\overline{n_1}} \sum_{j=1}^{\overline{n_2}} \sum_{s=1}^{\overline{n_1}} \sum_{t=1}^{\overline{n_2}} E(Y_{(i,j)} E(Y_{(s,t)}) = \left( \sum_{i=1}^{\overline{n_1}} \sum_{j=1}^{\overline{n_2}} E(Y_{(i,j)}) \right) \left( \sum_{s=1}^{\overline{n_1}} \sum_{t=1}^{\overline{n_2}} E(Y_{(s,t)}) \right)
$$

$$
= \left( \sum_{i=1}^{\overline{n_1}} \sum_{j=1}^{\overline{n_2}} E(Y_{(i,j)}) \right)^2
$$

$$
= \left( \overline{n_2} \sum_{|w|=k} N_w^A \Pr^B(w) \right)^2.
$$

This is the last term in (10).

Remark 1.5. The formula for the variance in equation (10) has computational complexity of  $O(n_2^2 4^k)$ due to the  $O(n_2^2)$  computation of  $S_{b,c}^B$ . To reduce complexity we use the following approximation. By [2], when the length of the sequence is large enough,

$$
S_{b,c}^B \approx \frac{q_2(q_2+1)}{2} p^B(c) - (q_2-1)\mathbf{e}_b' \mathbf{Q} \mathbf{P} \mathbf{e}_c - \mathbf{e}_b' \mathbf{Q} \mathbf{P}^2 \mathbf{Q} \mathbf{e}_c, \tag{12}
$$

 $\Box$ 

where **P** is the  $d \times d$  matrix of transition probabilities, where d is the size of our alphabet (i.e.  $d = 4$ )  $\mathbf{Q} = (\mathbf{P} - \mathbf{I} + \mathbf{1}\mathbf{p}')^{-1}$ , with  $\mathbf{1}' = (1, \ldots, 1)$  and  $\mathbf{p}$  is the vector of steady state probabilities. The vector  $\mathbf{e}_c$  is the unit vector that picks up the column corresponding to c.

With this approximation, formula (10) can be computed in  $O(4^k)$  time.

### 1.4 Conditional  $D2z$  with weights,  $wD2z$ .

In  $cD2(B|A)$  of section 1.3, we give every k-word the same weight. In a more realistic setting, we would like to give more weight to some words and less to others. For this, we define a weighted conditional D2 measure,  $wcD2(B|A)$  or  $wD2(B|A)$  for short, and compute its z-score. We assume the weights are learned from the sequence A. Hence we make the following definition. (Note: in the main text  $wD2z$  is called  $D2z$ -cond-weights).

**Definition 1.6.** Let A and B be two DNA sequences. For each k-word w, let  $a_w$  be a weight associated with  $w$ . Define a weighted D2 score by

$$
wD2(B|A) = \sum_{|w|=k} a_w N_w^A N_w^B.
$$

In terms of the match indicator random variables,

$$
wD2(B|A) = \sum_{i=1}^{\bar{n_1}} \sum_{j=1}^{\bar{n_2}} a_{ij} Y_{(i,j)},
$$
\n(13)

where  $a_{ij}$  is the weight of the k-word  $A[i, i + k - 1]$ .

Next we compute the z-score for  $wD2(-|A)$ . IID case.

Theorem 1.7. Under the IID model,

$$
E(wD2(B|A)) = \bar{n_2} \sum_{|w|=k} a_w N_w^A \Pr^B(w), \qquad (14)
$$

and

$$
\begin{split}\n\text{Var}(wD2(B|A)) &= \\
2\sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \sum_{|v|=m} (\sum_{\substack{|w|=k \text{ suf}_m(w)=v}} a_w N_w^A \text{Pr}^B(w)) (\sum_{\substack{|u|=k \text{ suf}_m(w)=v}} a_u N_w^A \text{Pr}^B(suf_{k-m}(u))) \\
&\quad + \bar{n_2} \sum_{|w|=k} a_w^2 (N_w^A)^2 \text{Pr}^B(w) \\
&\quad - 2\sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \left(\sum_{|w|=k} a_w N_w^A \text{Pr}^B(w)\right)^2 \\
&\quad - \bar{n_2} \left(\sum_{|w|=k} a_w N_w^A \text{Pr}^B(w)\right)^2\n\end{split} \tag{15}
$$

Note that the above formula can be computed in  $O(4^k)$ .

# *Proof.* Same as the proof of Proposition 1.1 and Theorem 1.2, while keeping track of the weights.  $\Box$ Markov Model case.

Theorem 1.8. Under Markov model of order one and with notation as in Theorem 1.4,

$$
E(wD2(B|A)) = \bar{n_2} \sum_{|w|=k} a_w N_w^A \Pr^B(w) = \bar{n_2} \sum_{|w|=k} a_w N_w^A p^B(w_1) p_*^B(w_1).
$$
 (16)

For the variance formula, to simplify notation, write  $a_w$  for  $a_{w_1...w_k}$  and  $a_u$  for  $a_{u_1...u_k}$ . Then

$$
\begin{split}\n\text{Var}(wD2(B|A)) &= \\
2 \sum_{w_k, u_1} S_{w_k, u_1}^B \sum_{w_1, \dots, w_{k-1}} \text{Pr}^B(w_1 \dots w_k) a_w N_{w_1 \dots w_k}^A \sum_{u_2, \dots, u_k} p_*^B(u_1 \dots u_k) a_u N_{u_1 \dots u_k}^A \\
&+ 2 \sum_{m=1}^{k-1} (n_2 - 2k + m + 1) \sum_{|v| = m} \left( \sum_{\substack{|w| = k \\ \text{sat}_{m}(w) = v}} a_w N_w^A \text{Pr}^B(w) \right) \\
&\quad \left( \sum_{\substack{|u| = k \\ \text{pre}_m(u) = v}} a_u N_u^A p^B(u_{m+1} | v_m) p_*^B(s u f_{k-m}(u)) \right) \\
&+ \bar{n_2} \sum_{|w| = k} a_w^2 (N_w^A)^2 \text{Pr}^B(w) \\
&- \bar{n_2}^2 \left( \sum_{|w| = k} a_w N_w^A \text{Pr}^B(w) \right)^2\n\end{split} \tag{17}
$$

*Proof.* Same as the proof of Proposition 1.3 and Theorem 1.4, while keeping track of the weights.  $\Box$ 

As in Remark 1.5, an approximation of formula (17) for the variance can be computed in  $O(4^k)$ .

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 $\overline{v}$  (D<sub>2</sub>(D<sub>1</sub>A))

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