

## A APPENDIX

We prove that if all error estimators are unbiased, then  $\lim_{K \rightarrow \infty} B(m, n, K) = 0$ .

LEMMA A.1. *If all error estimators are unbiased, then  $B(m, n, K) \leq 0$ .*

PROOF. Define the set  $\mathcal{S}_n = \{\mathcal{S}_n^1, \mathcal{S}_n^2, \dots, \mathcal{S}_n^K\}$ , where  $\mathcal{S}_n^k, k = 1, 2, \dots, K$  is a random sample taken from the distribution  $F_k$  for  $k = 1, 2, \dots, K$ . Also, we can rewrite Equation (9) as

$$\varepsilon_{\text{est}}^{\min}(K) = \min_{i,j} \left\{ \frac{1}{K} \sum_{k=1}^K \varepsilon_{\text{est}}^{i,j,k} \right\}, \quad (\text{A.1})$$

where  $i = 1, 2, \dots, r$  and  $j = 1, 2, \dots, s$ . Owing to the unbiasedness of the error estimators,  $E_{\mathcal{S}_n^k}[\varepsilon_{\text{est}}^{i,j,k}] = E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i,k}]$ . Referring to Equations (10) and (A.1), we have

$$\begin{aligned} B(m, n, K) &= E_{\mathcal{S}_n}[\varepsilon_{\text{est}}^{\min}(K)] - \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k}] \\ &= E_{\mathcal{S}_n} \left[ \min_{i,j} \left\{ \frac{1}{K} \sum_{k=1}^K \varepsilon_{\text{est}}^{i,j,k} \right\} \right] - \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k}] \\ &\leq \min_{i,j} \left\{ E_{\mathcal{S}_n} \left[ \frac{1}{K} \sum_{k=1}^K \varepsilon_{\text{est}}^{i,j,k} \right] \right\} - \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k}] \\ &= \min_{i,j} \left\{ \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n}[\varepsilon_{\text{est}}^{i,j,k}] \right\} - \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k}] \\ &= \min_{i,j} \left\{ \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{est}}^{i,j,k}] \right\} - \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k}] \\ &= \min_i \left\{ \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i,k}] \right\} - \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k}] \\ &\leq 0. \end{aligned} \quad (\text{A.2})$$

where the relations in the third and sixth lines result from Jensen's inequality and unbiasedness of the error estimators, respectively.  $\square$

LEMMA A.2. *If all error estimators are unbiased, then  $\lim_{K \rightarrow \infty} B(m, n, K) \geq 0$ .*

PROOF. Let

$$A^{i,j} = \frac{1}{K} \sum_{k=1}^K \varepsilon_{\text{est}}^{i,j,k}, \quad T^i = \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i,k}]. \quad (\text{A.3})$$

Owing to the unbiasedness of the error estimators,  $E_{\mathcal{S}_n}[A^{i,j}] = T^i \leq 1$ . Without loss of generality, we assume  $T^1 \leq T^2 \leq \dots \leq$

$T^r$ . To avoid cumbersome notation, we will further assume that  $T^1 < T^2$  (with some adaptation, the proof goes through without this assumption). Let  $2\delta = T^2 - T^1$  and

$$B_\delta = \left( \bigcap_{j=1}^s \left( T^1 - \delta \leq A^{1,j} \leq T^1 + \delta \right) \right) \cap \left( \min_{i \neq 1, j} \{ A^{i,j} \} > T^1 + \delta \right). \quad (\text{A.4})$$

Because  $|\varepsilon_{\text{est}}^{i,j,k}| \leq 1$ ,  $\text{Var}_{\mathcal{S}_n}[A^{i,j}] \leq 1/K$ . hence, for  $\tau > 0$ , there exists  $K_{\delta,\tau}$  such that  $K \geq K_{\delta,\tau}$  implies  $P(B_\delta(K)) > 1 - \tau$ . Hence, referring to Equation (10), for  $K \geq K_{\delta,\tau}$ ,

$$\begin{aligned} E_{\mathcal{S}_n}[\varepsilon_{\text{est}}^{\min}(K)] &= E_{\mathcal{S}_n}[\varepsilon_{\text{est}}^{\min}(K) | B_\delta] P(B_\delta) \\ &\quad + E_{\mathcal{S}_n}[\varepsilon_{\text{est}}^{\min}(K) | B_\delta^c] P(B_\delta^c) \\ &\geq E_{\mathcal{S}_n}[\varepsilon_{\text{est}}^{\min}(K) | B_\delta] P(B_\delta) \\ &= E_{\mathcal{S}_n}[\min_j A^{1,j}] P(B_\delta) \\ &\geq (T^1 - \delta)(1 - \tau). \end{aligned} \quad (\text{A.5})$$

Again referring to Equation (10) and recognizing that  $i_{\min} = 1$  in  $B_\delta$ , for  $K \geq K_{\delta,\tau}$ ,

$$\begin{aligned} \frac{1}{K} \sum_{k=1}^K E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k}] &= \frac{1}{K} \sum_{k=1}^K \left( E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k} | B_\delta] P(B_\delta) \right. \\ &\quad \left. + E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{i_{\min},k} | B_\delta^c] P(B_\delta^c) \right) \\ &\leq \frac{1}{K} \sum_{k=1}^K \left( E_{\mathcal{S}_n^k}[\varepsilon_{\text{true}}^{1,k} | B_\delta] P(B_\delta) \right. \\ &\quad \left. + P(B_\delta^c) \right) \\ &\leq T^1 + \tau. \end{aligned} \quad (\text{A.6})$$

Putting Equations (A.5) and (A.6) together and referring to Equation (10) yields, for  $K \geq K_{\delta,\tau}$ ,

$$B(m, n, K) \geq (T^1 - \delta)(1 - \tau) - T^1 - \tau \geq -(2\tau + \delta) \quad (\text{A.7})$$

Since  $\delta$  and  $\tau$  are arbitrary positive numbers, this implies that for any  $\eta > 0$ , there exists  $K_\eta$  such that  $K \geq K_\eta$  implies  $\lim_{K \rightarrow \infty} B(m, n, K) \geq 0$ , which is precisely what we want to prove.  $\square$

Combining Lemmas A.1 and A.2, we have proven that  $\lim_{K \rightarrow \infty} B(m, n, K) = 0$  under the assumption that all the error estimators are unbiased.