

1 **Appendix S1.** More details of the genome sequencing procedure and the indices used to
2 measure gene flow.

3

4 **Viral Genome Sequencing.** All samples were first sequenced using the traditional
5 PCR/Sanger high-throughput sequencing pipeline at the JCVI (Dugan *et al.* 2008). After
6 sequencing, the sequences were trimmed to remove primer sequence as well as low quality
7 residues, and segments were assembled individually using the small genome assembler Elvira
8 (<http://elvira.sourceforge.net/>). Sequencing and assembly difficulties associated with the high
9 level of variability in the sample set (e.g. due to the mixed infection of multiple influenza
10 subtypes) led to the processing of all samples using the next generation sequencing pipeline at
11 JCVI that includes the 454/Roche GS-FLX and the Illumina Genome Analyzer II. Viral RNA was
12 first reverse transcribed and amplified by multi-segment RT-PCR (M-RT-PCR) (Zhou *et al.* 2009).
13 The cDNA for each sample was primed with barcoded random hexamers using the SISPA
14 protocol (Djikeng *et al.* 2008). One library was prepared for sequencing on the 454/Roche GS-
15 FLX platform using Titanium chemistry while the other was made into a library for sequencing
16 on the Illumina Genome Analyzer II.

17 The sequence reads from the GS-FLX data were sorted by barcode, trimmed, and
18 searched by TBLASTX against custom nucleotide databases of full-length influenza A segments
19 downloaded from GenBank to filter out both chimeric influenza sequences and non-influenza
20 sequences amplified during the random hexamer-primed amplification. The filtered GS-FLX
21 reads were then binned by segment, and *de novo* assembled using CLC Bio's
22 `clc_novo_assemble` program. Because of the short read length of the sequences obtained from
23 the Illumina Genome Analyzer II, these were not subjected to the TBLASTX filtering step. Both
24 GS-FLX and Illumina reads were then mapped to the selected reference influenza A virus
25 segments using the `clc_ref_assemble_long` program. At loci where both GS-FLX and Illumina
26 sequence data agreed on a variation compared to the reference sequence, the latter was

1 updated to reflect the difference. A final mapping of all next generation sequences to the
2 updated reference sequences was then performed. All sequences generated here have been
3 submitted to GenBank and assigned accession numbers (Table S1).

4

5 **Phylogenetic Analysis**

6 Panoramic phylogenies were estimated using a rapid hill-climbing search method and the
7 GTRGAMMA nucleotide substitution model implemented in RAxML v7.04 (Stamatakis 2006).
8 The tree search was initiated with a random maximum parsimony tree. 200 independent tree
9 searches with different random maximum parsimony starting trees were performed to obtain the
10 phylogeny with the highest likelihood score. Major lineages of North American wild bird avian
11 influenza viruses were identified and extracted for further analyses.

12 For more accurate phylogenetic analysis of NA-WB-AIV lineages, we used the heuristic
13 search method implemented in PhyML v2.4.5 (Guindon *et al.* 2009). This allowed us to
14 determine the maximum likelihood phylogenetic tree by optimizing the tree topology and branch
15 lengths on the sequence data. The tree search was initiated using a BIONJ tree and then
16 employed nearest-neighbor interchange branch-swapping. All phylogenetic analyses utilized
17 the General Time Reversible (GTR) model of nucleotide substitution which allows variable rates
18 of substitution between the A, T, C and G nucleotides, a proportion of invariable sites (I), and
19 four classes of rate variation among nucleotide sites ($+\Gamma_4$). The RAxML and PhyML trees were
20 extremely similar in topology.

21 Pseudo-replicates for each data set were generated by resampling the columns of
22 sequence alignment with replacement, and then used in bootstrap analyses to determine
23 phylogenetic robustness. Because of the very large size of the 'panoramic' data set (>5,000
24 sequences), the neighbor-joining clustering method (as implemented in the PAUP* package –
25 Swofford *et al.* 2003) was used to infer bootstrap trees in this case. In the case of the North
26 American wild bird AIV lineage, where more phylogenetic accuracy is required, we used the

1 maximum likelihood bootstrap method implemented in PhyML. More background information
2 about the types of phylogenetic analyses performed here can be found in Lam *et al.* (2010).

3 All AIV gene segments were subject to the independent phylogenetic and
4 phylogeographic analyses, with the exception that the only HA and NA gene segments analyzed
5 were from subtypes H3, H4, N6, and N8. The other subtypes had HA and NA data sets of
6 insufficient size (<200 sequences or 8 sampling localities) for meaningful phylogeographic
7 analysis.

8

9 **Measuring Gene Flow Between Populations**

10 Three summary statistics – (i) the Fixation index (F_{ST}), (ii) modified Slatkin-Maddison's s (σ),
11 and (iii) the rate of state transition (q) – were used to determine the level of gene flow between
12 two localities from either the sample of virus sequences in the two localities or from the inferred
13 phylogenetic trees.

14 **(i) F_{ST} .** F_{ST} is derived from the F -statistics developed by Sewall Wright to study
15 population structure (Wright 1942). It was later modified and generalized for genetic sequence
16 data sampled in different populations (Nei 1982; Lynch & Crease 1990). Our study employed a
17 F_{ST} similar to that generalized by Hudson *et al.* (1992) using the GTR substitution model to
18 estimate pairwise genetic distances ($\bar{\delta}_{xy}$) between two nucleotide sequences (x and y). F_{ST}
19 between the populations sampled at geographical states i and j is defined as,

$$20 \quad F_{ST} = 1 - \frac{H_w}{H_b}$$

21 H_w is the average within-population variation in states i and j , which is defined as,

$$22 \quad H_w = \frac{\left(\frac{2}{n_i(n_i - 1)} \sum_{ix < iy} \delta_{ix,iy} \right) + \left(\frac{2}{n_j(n_j - 1)} \sum_{jx < jy} \delta_{jx,jy} \right)}{2}$$

23 where n_i and n_j are the total numbers of sequences sampled from the populations at
24 geographical states i and j , respectively. ix and iy are the x^{th} and y^{th} sequences in the

1 population sampled in geographical state i , while jx and jy are the x^{th} and y^{th} sequences in the
2 population sampled in geographical state j .

3 H_b is the between-population variation of state i and j , which is defined as,

$$4 \quad H_b = \frac{\sum_{ix,jy} \delta_{ix,jy}}{n_i n_j}$$

5 This is the average pairwise comparison of sequences sampled in the populations of
6 geographical states i and j . Extreme F_{ST} estimates of 0 and 1 indicate no and complete
7 population subdivision, respectively. F_{ST} was calculated for each pair of geographical states in
8 the NA-WB-AIV data sets, and was plotted against the spatial distance (km) between each pair
9 of geographical states (Fig. S2). Geographical states with too few sampled sequences ($n \leq 5$ or
10 less than 1% of all sequences) were excluded. The PAUP* program (Swofford 2003) was used
11 to estimate the GTR+I+ Γ_4 genetic distance between sequences. A PERL script implementing
12 the calculation of F_{ST} from the genetic distances is available upon request.

13 **(ii) Modified Slatkin-Maddison's s .** Slatkin and Maddison proposed and demonstrated
14 that the minimum number of geographical state change (s) (between two geographical states)
15 observed in the phylogeny provides an estimate of the level of gene flow between these
16 populations (Slatkin & Maddison 1989; Hudson *et al.* 1992),

$$17 \quad s \approx Nm$$

18 where N is the population size and m is the migration rate between the populations.

19 In situations where the sampling frequencies are highly variable among localities, the
20 number of sequences sampled in two geographical locations is the limiting factor for the
21 minimum number of the geographical state changes (s) that can be observed in the phylogeny.
22 To overcome this problem, we used a modified Slatkin-Maddison's s , denoted ' σ ', to infer the
23 level gene flow between two geographical states (e.g. A and B). This is defined as,

$$24 \quad \sigma = \frac{s}{s_p}$$

1 where s is number of changes between state A and state B in the observed phylogeny, and s_P is
2 the number of changes between state A and state B occurring in the phylogeny simulated
3 assuming panmixis. In other words, s_P determines the number of state changes that could be
4 observed with the current phylogenetic structure and number of samples available in different
5 geographical states. This scaling turned Slatkin-Maddison's s into a measure of the relative
6 level of gene flow compared to the scenario of the completely unrestricted gene flow (i.e.
7 panmixis). Extreme σ estimates of 0 and 1 (they may possibly but rarely exceed 1) indicate no
8 gene flow and panmixis, respectively.

9 The values of s between different geographical states were estimated from each
10 maximum likelihood gene phylogeny using the Fitch parsimony method (Fitch 1971)
11 (implemented in the JAVA BEAST library). Polytomies were resolved randomly 100 times and
12 average s values were obtained. To obtain s_P , each pair of geographical states was randomly
13 shuffled at the tree tips for 1,000 iterations, and ancestral geographical states were again
14 reconstructed using parsimony. This generated a distribution of s_P under panmixis, and the
15 average s is divided by the mean s_P to give σ . To account for topological uncertainty, 200
16 bootstrap maximum likelihood trees were analyzed in the manner described above, and the
17 upper and lower 95% estimates were taken as the resultant uncertainty of σ . A JAVA computer
18 program implementing this procedure is available upon request.

19 To study rates of gene flow within and between migratory flyways, transitions in
20 geographical state, e.g. Minnesota ↔ Alaska, (estimated from the phylogeny using the
21 parsimony method described above) were categorized as transitions in flyways, e.g. MF ↔ PF,
22 depending on which flyway the geographical state belongs to (e.g. Minnesota belongs to
23 Mississippi Flyway (MF), Alaska belongs to Pacific Flyway (PF), see Table S2). Hence, s and
24 s_P were counted for each type of transition among four flyways (i.e. MF↔PF, MF↔MF, MF↔CF,
25 MF↔AF, PF↔PF, PF↔CF, PF↔AF, CF↔CF, CF↔AF and AF↔AF).

1 **(iii) Rate of state transition (q) by Pagel's maximum likelihood method.** Pagel

2 (1994) developed a maximum likelihood method to study trait evolution (as discrete characters)
3 along the phylogenetic tree (Pagel 1994). We employed this method and treated the
4 geographical states as the trait at every tip. Formally, this method maximizes

5
$$L = f(D|T, M)$$

6 where L is the likelihood of the data set D (i.e. the trait states of the taxa) given the phylogeny T ,
7 and the model of evolution, M , of the trait states. In the analysis of gene flow between each pair
8 of different geographical states, we estimated the parameters for the state transition rates (q) in
9 the model. In total, there were 120 reversible transition rate parameters (e.g. $q_{MN \leftrightarrow TX}$, $q_{ND \leftrightarrow OR}$,
10 $q_{SD \leftrightarrow WA}$) of interest given the 16 geographical states in the study. The huge number of
11 parameters made the model statistically untraceable. To reduce the large number of
12 parameters, we simplified the state transition model by reducing the rate parameters for two
13 states to rate parameters for two flyways. This constituted the flyway-specific rate model (FRM)
14 of AIV gene flow. In the FRM there were 10 reversible transition rate parameters for four
15 flyways (and 6 parameters for the three flyway model (3-FRM), which combines MF and CF):
16 $q_{CF \leftrightarrow AF}$, $q_{CF \leftrightarrow PF}$, $q_{CF \leftrightarrow MF}$, $q_{CF \leftrightarrow CF}$, $q_{AF \leftrightarrow MF}$, $q_{AF \leftrightarrow PF}$, $q_{AF \leftrightarrow AF}$, $q_{PF \leftrightarrow PF}$, $q_{PF \leftrightarrow MF}$, $q_{MF \leftrightarrow MF}$. These parameters
17 were estimated from the phylogeny with the highest likelihood score, using the discrete
18 maximum likelihood method described by Pagel (1994) as implemented in the APE library
19 (Paradis *et al.* 2004) running in the R package version 2.11.1. Various starting values of
20 transition rates were attempted (20, 10, 5, 1, 0.1 and 0.01), and the estimates with the best
21 likelihoods were kept. The flyway-specific rate estimates are shown in Table S3.

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