Supplementary Information:

The evolution of sensory divergence in the context of limited gene flow in the bumblebee bat

Sébastien J. Puechmaille, Meriadeg Ar Gouilh, Piyathip Piyapan, Medhi Yokubol, Khin Mie Mie, Paul J. Bates, Chutamas Satasook, Tin Nwe, Si Si Hla Bu, Iain J. Mackie, Eric J. Petit and Emma C. Teeling

* Corresponding authors: s.puechmaille@gmail.com; emma.teeling@ucd.ie

This Supplementary Information includes:

- Supplementary Figures S1-S5,
- Supplementary Tables S1-S9,
- Supplementary Notes 1-2,
- Supplementary Methods,
- Supplementary References.

Supplementary Figure S1 | Echolocation calls of *Craseonycteris thonglongyai* in Thailand and Myanmar. Spectrogram (bottom) and power spectrum (top) showing the difference in frequency between *C. thonglongyai* echolocation calls **(a)** in Myanmar (left, peak frequency=80.9 kHz, colony M3) and **(b)** Thailand (right, peak frequency=72.9 kHz, colony A). Note that the X-axis (peak frequency in kHz) and Y-axis (time in ms) scales are identical.

Supplementary Figure S2 | Genetic differentiation of Thai *versus* Myanmar individuals based on five nuclear SNPs. Genotyping of 5 SNPs from 196 Myanmar (blue) and 463 Thai (green) *Craseonycteris thonglongyai* individuals (ZP2 is an autosomal marker; SMCY7 is a Y-Chromosome marker and PHKA2-start, PHKA2-end and BGN are X-Chromosome markers). Note the presence of two heterozygote individuals (arrows), one for ZP2 and one for BGN. For SMCY7, the two alleles are present in the Thai population.

Supplementary Figure S3 | Observed and predicted mismatch distributions under a model of demographic expansion (top), or spatial expansion (bottom). For the five groups of sites, a model of constant population size was rejected as the 99% confidence intervals of θ_0 and θ_1 were not overlapping. Solid black lines represent the observed frequency of pairwise distributions, dashed black lines indicate the expected distribution under the model and the solid grey lines indicate the upper and lower 95% confidence interval for the expected. For the demographic expansion (top panels), the raggedness index (Rag.) and associated P-values are as follow, sites L, M, N & O: Rag.= 0.0129, *p*=0.603; sites H, I & J: Rag.= 0.006, *p*=0.943; site K: Rag.=0.0151, *p*=0.378; sites E, F & G: Rag.=0.0388, *p*=0.002; sites A, B & D: Rag.=0.019, *p*=0.246. For the spatial expansion (bottom panels), the raggedness index and associated P-values are as follows, sites L, M, N & O: Rag.= 0.0129, *p*=0.881; sites H, I & J: Rag.= 0.006, *p*=0.96; site K: Rag.=0.0151, *p*=0.747; sites E, F & G: Rag.=0.0388, *p*=0.17; sites A, B & D: Rag.=0.019, *p*=0.857.

Supplementary Figure S4 | Example of isolation by distance obtained by simulating a spatially expanding population. Simulations were carried out in SPLATCHE with the migration rate and growth parameter values set to 0.3 each and the carrying capacity set to 250 (Mantel test; $r=0.95$, $R^2=0.90$, *P*<0.001). The program was run over 400 generations (see Supplementary Table S3).

Supplementary Figure S5 | Schematic diagram of the organisation of the mitochondrial DNA in mammals. The D-loop domains were specified according to Sbisà *et al*. ⁶¹ although the relative length of each domain can vary between species or individuals of the same species. Each arrow indicates a primer with its name specified. Red boxes indicate the presence of a stop codon (AGA; mtDNA code) and black boxes, the absence of a stop codon where one would be expected (see text for further explanations).

Individual			Accession N.	Origin	Marker
reference	Country	Haplotype			
CT ₁	Myanmar	Cratho Numt2a 1	GU ₂₄₇₆₀₃	nuclear	Num2a
CT ₂₅₅	Thailand	Cratho_Numt2a_2	GU ₂₄₇₆₀₄	nuclear	Num2a
CT17	Myanmar	Cratho_Numt2b_1	GU ₂₄₇₆₀₅	nuclear	Numt2b
CT ₂₅₅	Thailand	Cratho Numt2b 2	GU247606	nuclear	Numt2b
CT ₁₈	Myanmar	Cratho ZP2 1	GU247611	nuclear	ZP ₂
CT ₂₅₆	Thailand	Cratho_ZP2_2	GU ₂₄₇₆₁₂	nuclear	ZP ₂
CT18	Myanmar	Cratho BGN 1	GU ₂₄₇₆₀₁	X-chromosome	BGN
CT803	Thailand	Cratho_BGN_2	GU ₂₄₇₆₀₂	X-chromosome	BGN
CT17	Myanmar	Cratho_PHKA2_1	GU247607	X-chromosome	PHKA2
CT ₂₅₅	Thailand	Cratho PHKA2 2	GU ₂₄₇₆₀₈	X-chromosome	PHKA ₂
CT ₁₈	Myanmar	Cratho SMCY 1	GU247609	Y-chromosome	SMCY7
CT ₂₅₅	Thailand	Cratho_SMCY_2	GU ₂₄₇₆₁₀	Y-chromosome	SMCY7

Supplementary Table S1 | List of *Craseonycteris thonglongyai* haplotypes from nuclear markers and their corresponding GenBank accession numbers.

Individual reference	Country	Haplotype name	Frequency	Accession number
CT ₁₃	Myanmar	Cratho_1	17	GU247613
CT16	Myanmar	Cratho_2	10	GU247683
CT17	Myanmar	Cratho_3	10	GU247614
CT28	Myanmar	Cratho_4	89	GU247684
CT36	Myanmar	Cratho_5	$\mathbf{1}$	GU247615
CT38	Myanmar	Cratho_6	$\overline{2}$	GU247685
CT54	Myanmar	Cratho_7	$\mathbf{1}$	GU247616
CT60	Myanmar	Cratho_8	$\mathbf{1}$	GU247686
CT88	Myanmar	Cratho_9	$\mathbf{1}$	GU247617
CT89	Myanmar	Cratho_10	$\overline{2}$	GU247687
CT90	Myanmar	Cratho_11	$\mathbf{1}$	GU247618
CT111	Myanmar	Cratho_12	$\mathbf{1}$	GU247688
CT114	Myanmar	Cratho_13	$\mathbf{1}$	GU247619
CT118	Myanmar	Cratho_14	$\mathbf{1}$	GU247689
CT298	Myanmar	Cratho_15	$\mathbf{1}$	GU247620
CT229	Thailand	Cratho_16	13	GU247690
CT230	Thailand	Cratho_17	40	GU247621
CT232	Thailand	Cratho_18	22	GU247691
CT233	Thailand	Cratho_19	6	GU247622
CT235	Thailand	Cratho_20	7	GU247692
CT236	Thailand	Cratho_21	$\overline{4}$	GU247623
CT237	Thailand	Cratho_22	$\mathbf{1}$	GU247693
CT238	Thailand	Cratho_23	13	GU247624
CT239	Thailand	Cratho_24	3	GU247694
CT240	Thailand	Cratho_25	$\overline{2}$	GU247625
CT242	Thailand	Cratho_26	4	GU247695
CT243	Thailand	Cratho_27	3	GU247626
CT245	Thailand	Cratho_28	1	GU247696
CT ₂₅₅	Thailand	Cratho_29	8	GU247627
CT256	Thailand	Cratho_30	7	GU247697
CT265	Thailand	Cratho_31	1	GU247628
CT266	Thailand	Cratho_32	3	GU247698

Supplementary Table S2 | List of *Craseonycteris thonglongyai* mitochondrial DNA haplotypes (*Cytb*, tRNA-Pro, tRNA-Thr, D-loop) and their corresponding GenBank accession numbers.

Supplementary Table S3 | Results of the spatial expansion simulations run in SPLATCHE for different parameters (*C*, *m* and *r*) combinations. A total of 24 simulations were run with 100 replicates each. The 'Generations' column specify how many generations were necessary to fill in the lattice. The 'Match observed' column present the simulations in which an isolation by distance pattern similar to the one seen in *Craseonycteris* was observed. See simulation details in the "simulation of a spatially expanding population" section of the Supplementary Material.

Supplementary Table S4 | Evaluation of the seven organisational models considering geographic

Boldface indicates fully supported models,

 A^a B = barrier, G = genetics, D = distance, E = echolocation. The period in the expectation abbreviations separates the covariate matrix from the two primary matrices (i.e. DG.B indicates a Mantel test between the distance and genetic matrices, with the barrier matrix partialed out),

^b P-values are preceded with the sign '*' if significant or 'NS' if non-significant after sequential Bonferroni correction.

Supplementary Table S5 | Evaluation of the six models of causal relationships between geographic distance (D), echolocation distance (E) and genetic distance (G).

Boldface indicates fully supported models.

Abbreviations are identical to those used in Supplementary Table S4.

Supplementary Table S6 | Summary of the synteny between the *Homo sapiens* block including the *RBP-J* gene (left side) and other species (right side) available in Ensembl (http://www.ensembl.org/index.html). The microsatellite CTC1 blasts a region 50 kb at the 5'-side of *RBP-J*, that is at position 26.267 in Chromosome 4 of *Homo sapiens* (GRCh37 primary reference assembly). In all species looked at, the synteny around *RBP-J* was conserved. Base position or block length are expressed in Mega-bases (1M=1,000,000 bases).

Seq. N°	Genus	Species	g_1	Author/Reference
NC 007393	<i>Rousettus</i>	aegyptiacus	74310519	Omatsu et al., unpublished
NC_006925	Mystacina	tuberculata	62184382	NCBI Genome Project,
				unpublished
NC 002612	Pteropus	dasymallus	11386118	Nikaido et al., unpublished
NC 005434	<i>Rhinolophus</i>	pumilus	42632271	62
NC 005436	Pipistrellus	abramus	42632257	62
NC 002626	<i>Chalinolobus</i>	tuberculatus	11610804	63
NC_002619	Pteropus	scapulata	11602891	63
NC 005433	<i>Rhinolophus</i>	monoceros	42717961	64
NC 002009	Artibeus	jamaicensis	5835666	65

Supplementary Table S7 | Complete mitochondrial DNA sequences of nine bat species downloaded from GenBank.

Supplementary Table S8 | Primer names and sequences used in the present study.

Supplementary Table S9 | Forward and reverse primer and Taqman probe sequences for five nuclear introns SNPs. SNPs were fluorescently labelled using either FAM or VIC dyes.

Supplementary Note 1

Echolocation and mate choice. Our data show that echolocation does have a minimum effect on gene flow between two populations in close vicinity (Northern/Central *versus* Southern populations) but how can this happen? Most bat species predominately rely on echolocation for sensory perception 22 . Throughout the historical literature, echolocation has been perceived to have a dominant role in orientation and prey capture. A bat's larynx, pinnae shape and inner ear structures are directly correlated with their echolocation capacities and therefore foraging capabilities²². However, not much is known about how echolocation can influence social communication in bats (however, see $14,115$). Recent literature has shown that bats can discriminate individuals on the basis of their echolocation call^{12,116-118} and some studies have suggested that echolocation parameters somehow influence mate choice preference in bats5,119,120. This is not that surprising. Indeed, it has been shown that bird beaks are predominantly linked to their foraging capabilities but also play a role in determining song structure which then in turn affects their mating songs¹²¹, and thus mate choice. Therefore, in bats that also use sound for mating rituals and courtship¹²², it seems likely that the apparatus they use to make these sounds (larynx) and to perceive these sounds (their outer and inner ears), which are directly linked to their echolocation capacities, will play a role in determining their mate preferences. It has been shown that in *Saccopteryx bilineata*, male reproductive success as determined by their number of offspring was directly correlated with echolocation frequency parameters¹²³. Further studies on *Myotis lucifugus* have demonstrated that males preferred calls of females who mated frequently rather than calls of females rarely mating 124 . The European cryptic pipistrelles (*Pipistrellus pipistrellus* and *P. pygmaeus*) echolocate with an average difference of 10 kHz¹²⁰ and also show frequency differences in their social calls¹²⁵. Therefore echolocation can indeed play a role in communication and mate choice and our data suggest that bats are choosing to mate with other bats that have the same 'echolocation' call type.

Supplementary Note 2

Echolocation competition. Echolocating bats produce ultrasonic signals and determine the direction, distance, and features of objects in the environment from the arrival time, amplitude, and spectrum of sonar reflections¹³. This process commonly named 'echolocation' is a very challenging task requiring specific adaptations of the vocal and auditory system as well as the brain^{45,107,108}. We present below two hypotheses (Resource partitioning and Interference) potentially explaining the observed pattern of echolocation call variation observed in *C. thonglongyai* in Thailand. It is important to note that these two hypotheses are not mutually exclusive and could act together to strengthen the divergent selection.

- **(A) Resource Partitioning:** Assuming targets are spheres and a the speed of sound equals 347.65 m s-1 (for a temperature of 25°C and a relative humidity of 80%), the theory predicts that, *M. siligorensis* echolocating at 70 kHz can detect targets of 4.96 mm whereas *C. thonglongyai* echolocating at 76 kHz can detect targets of smaller size (4.56 mm). This would mean that *C. thonglongyai* can detect preys that are 8% smaller than the smallest prey detectable by M. *siligorensis*. Although this does not correspond to a large difference, not enough is currently known about the complexities of bat echolocation and their perception to rule out the effect of this 6 kHz difference on resource partitioning (although, \sec^{43}). Furthermore, experimental studies broadcasting ultrasounds at real insects and recording returning echoes have showed that these calculations were not very reliable as insects are not spheres¹⁰⁹. The strength of the returning echo does not only depend on the insect size but also on its wing beat and the angle (insect ensonified from the front, side or back)¹⁰⁹. Therefore it is possible that this shift in call is biologically meaningful but we do not know enough to accept or reject this hypothesis yet⁴¹.
- **(B) Interference:** Bats are also faced with the challenge of separating their calls from the potential interference of various sounds in the environment where they are navigating. In echolocating bats, the onset of the emitted call activates a gating mechanism that establishes a time window during which pulse-echo pairs are processed for target distance determination⁴⁵. This process has been validated in bats with different echolocation types^{44,45,110} and should apply to all echolocating bat species⁴⁵. Experiments by Roverud & Grinnel⁴⁴ on *Noctilio albiventris*

demonstrated that playing artificial pulses resembling the bat's calls interfered with the bat's ability to determine the distance of objects. The disrupting effect of the artificial sound is likely due to the interference with the bat's processing of information from its own sounds by stimulating the same population of neurons that extract distance information from the bat's echolocation sound^{22,110}. Interference occurred only when the constant frequency of the playback sound was between 2 to 2.5 kHz above and 5 kHz below the frequency of the beginning of the frequency modulated sweep⁴⁴, demonstrating a narrow frequency window for interference. Assuming these thresholds are similar for our study species implies that *C. thonglongyai* echolocating at 75 kHz or above (Southern populations) avoids interference from *Myotis siligorensis* present in the area and echolocating at 70 kHz. In the Northern populations, where *M. siligorensis* has not been found, *C. thonglongyai* echolocates at 73-74 kHz, which is within the interference frequency window.

Intra-specific jamming in *C. thonglongyai* could be avoided by temporal shifts in frequency as demonstrated in many other bat species. When two or more bats of the same species are flying within the same airspace (within earshot of one another) individuals adjust the frequencies dominating their echolocation calls to avoid jamming each other sonar ¹¹¹⁻¹¹⁴. Nevertheless, Duanghkae³⁴ showed that in *Craseonycteris*, each individual had its own foraging area and occasionally, when another bat got into an individual's foraging area, it would be chased out. This suggests that *Craseonycteris* might avoid intraspecific jamming by avoiding contact with other individuals in the foraging grounds.

Supplementary Methods

Distinction of mtDNA *versus* **NUMTs.** Complete mitochondrial DNA sequences of nine bat species from five families were downloaded from GenBank (Supplementary Table S7) and aligned using MEGA version $3.1⁶⁹$. The D-loop was visually checked and realigned by hand. Six conserved primers were designed to amplify the entire *Cytb* and the D-loop in three overlapping fragments (see Supplementary Fig. S5 for primer locations and Supplementary Table S8 for primer sequences). Reactions were carried out in 25 µL simplex reactions containing 2 µL of DNA extract (at 2-5 ng/µL), 1X PCR buffer minus Mg (Invitrogen), 1.5 mM MgCl₂, 0.4 μ M each primer, 0.2 mM dNTPs and 1 U Platinum[®] *Taq* DNA Polymerase High Fidelity (Invitrogen). PCR volumes and reagents above-mentioned were used for all PCRs unless otherwise stated.

 To ascertain that we amplified the mitochondrial DNA and not Numts (nuclear copies of mitochondrial DNA), PCRs were completed with all possible primer combinations to ensure that all primers bound only to the mitochondrial DNA. All PCRs were carried out using the DNA from the same individual to avoid potential inter-individual differences. Identical PCR cycling conditions were used for all primer combinations; initial step 10' at 95°C, then 10 cycles of 15" at 95°C, 30" at 60°C (reduce by 2°C every 2 cycles), 1' at 72°C, following by 30 cycles of 15" at 95°C, 30" at 50°C and 1' at 72°C and a final step for 10' at 72°C.

 Using the DNA from a single individual and different combinations of primers, three different sequences were obtained for the same targeted mtDNA region. The PCRs, repeated on seven different individuals, furnished consistent results, suggesting the presence of Numts^{70,71}. The three different sequences were aligned and new primers were designed to specifically amplify each fragment (see primers represented in Supplementary Fig. S5 and Supplementary Table S8). Two amplified fragments spanned from the tRNA-Glu to the Central Domain of the D-loop as targeted whereas for one fragment, the first 730 bp of the *Cytb* could not be amplified (Supplementary Fig. S5d). When translating the amino-acid sequence of the *Cytb* using the mammalian mitochondrial genetic code, one stop codon was found in one sequence thus classified as Numt (Numt2a) (Supplementary Fig. S5c). A second sequence, for which we were not able to amplify the beginning of *Cytb*, presented a 2 bp deletion at position 1140-

1141 when compared to the other two sequences. This deletion created a frame-shift modifying the *Cytb* stop codon which was then found in position 1183-1185 (Supplementary Fig. S5d). Therefore, the tRNA-Thr present on the 3' side of the *Cytb* was reduced by 45 bp. This second sequence was thus also considered as a Numt (Numt2b). This was further confirmed by a simple sequence divergence table whereby Numt sequences from Thailand and Myanmar individuals were quasi-identical \langle <0.1% divergence) whereas true mitochondrial DNA from Thailand and Myanmar were more different (>1%) divergence). This difference between mitochondrial DNA and Numts reflects the expected higher mutation rate of mtDNA *versus* nuclear DNA^{72,73}.

Mitochondrial DNA amplification and sequencing. A 1840 bp mitochondrial DNA fragment encompassing the entire *Cytb*, tRNA Threonine, tRNA Proline and part of the D-loop was amplified by PCR in three overlapping fragments using three primer pairs R3.1-F/F3.1-R, R3.2-F/F2.4-R and R2.2- F/F2.1-R (Supplementary Table S8 and Supplementary Fig. S5b) and the PCR recipe described above. Amplifications were carried out in a DNA Engine $DYAD^{TM}$ thermocycler (MJ Research) with the following PCR program: initial step 10' at 95°C, then 10 cycles of 15" at 95°C, 30" at 65°C (reduce by 2°C every 2 cycles), 1' at 72°C, following by 30 cycles of 15" at 95°C, 30" at 55°C and 1' at 72°C and a final step for 10' at 72°C. The amplified products were then purified and directly sequenced by Macrogen (Korea) using the same primers as mentioned above. The six sequences per sample were edited and assembled using the program Sequencher 4.7 (Gene Codes Corporation, Ann Arbor, MI).

 All the 468 samples from Thailand were sequenced for the entire mitochondrial DNA fragment. In Myanmar, all samples were sequenced except from one colony (M13) where only 65 of the 128 samples were used. Out of 65 samples sequenced in M13, only 8 unique haplotypes were found, suggesting that we already detected all unique haplotypes⁷⁴. Individuals sampled twice (see the 'Microsatellites genotyping and analysis' paragraph in the Methods) were removed, leaving a data set comprising 463 sequences for Thailand and 139 for Myanmar.

Phylogenetic reconstruction and dating. Using mitochondrial, Numt2a and Numt2b sequence, phylogenetic reconstruction was undertaken using the Bayesian inference in BEAST⁵⁴. The part of the D- loop containing insertions-deletions was removed from the alignment prior to phylogenetic reconstruction due to difficulty in aligning this section, leaving a 1.8 kb alignment. The general time-reversible + gamma-distributed rates among sites + proportion of invariant sites (GTR+ Γ+I) substitution model was used as determined by ModelTest version 3.7^{75} . A strict molecular clock model was preferred over a relaxed molecular clock model as advised by Drummond *et al.*⁷⁶ when the standard deviation of the uncorrelated lognormal relaxed clock (parameter ucld.stdev in BEAST) is smaller than 1, which is what was observed when analysing the present dataset. Therefore, a strict molecular clock model was applied with a fixed mean substitution rate of $2.21x10^{-8}$ subs/site/year. The mutation rate over the entire sequence was calculated by weighting the rate contribution of the various functional regions Cytb, tRNAs, ETAS and Central domain⁷⁷ by their relative length. ETAS and Central domain were defined according to the alignments provided by Sbisà *et al.*⁶¹ whereas tRNAs were delimited using the online program ARWEN version 1.2⁷⁸. The average mutation rate was estimated to be 2.21×10^{-8} subs/site/year, which equals to a divergence rate of 4.42%/Myr. No outgroup was specified and the constant size coalescent was used as a tree prior. The program was run for 40,000,000 generations and sampled every 500. The first 4,000,000 generations were discarded as burn-in. Effective sample sizes for the estimated parameters and posterior probability as calculated with the program Tracer v1.4⁷⁹ were higher than 1,000.

Nuclear introns primers, PCR and genotyping. Four nuclear introns, two found on the X-chromosome BGN⁶⁶, PHKA2⁶⁷, the Y-Chromosome SMCY7⁶⁸ and one on chromosome 16 for *Homo sapiens* (ZP2, this study) were amplified and sequenced in at least 10 individuals per country (Thailand and Myanmar) (see Supplementary Table S8 for primers used). BGN, PHKA2 and SMCY7 were amplified using the same protocol as for mtDNA (see above), and ZP2 was amplified using the same protocol as NUMTs (see above). For these four introns, we identified five parsimoniously informative sites (one in each intron for BGN, SMCY7 and ZP2 and two in PHKA2). For the five SNPs, we designed custom TaqMan SNP assays (Applied Biosystems) and screened all the samples (n=659). Primer and probes sequences are reported on Supplementary Table S9. All reactions were conducted in 9 ul reaction volumes containing 1 µl of DNA , 4.5 µl 2X TaqMan® Genotyping Master Mix (Applied Biosystems), 0.225 µl 40X custom probe and 3.275 µl ddH2O. SNPs were genotyped on an ABI 7500 Fast Real-Time PCR System using a

pre-PCR read step of 60°C for 1 min, a denaturation step of 95°C for 10 min followed by 40 cycles (44 cycles for PHKA2-start) of 95°C for 15 s and 60°C for 1 min. A post-PCR read was done at 60°C for 1 min. SNPs were scored using the Allelic Discrimination Assay procedure in 7500 Software version 2.0.1 (Applied Biosystems).

Mismatch distributions analyses. The Thai populations' demographic history was examined using the mismatch distribution of 462 mitochondrial DNA sequences⁵⁵. Colony P (see Fig. 1) was not included in the analysis because of the limited sampling (one individual). Episodes of population growth or decline leave characteristic signatures in the distribution of nucleotide differences between pairs of individuals 80 . Mismatch distributions under demographic and spatial expansion scenarios differ from mismatch distributions from stable populations, which typically present multimodal mismatch distributions^{55,80}. Furthermore, contrary to stable populations, expanding populations present non-overlapping confidence intervals of the mutation parameters θ_0 and θ_1 , θ_0 and θ_1 being proportional to the effective population size before and after expansion respectively. When the confidence interval of θ_1 does not overlap with the confidence interval of θ_0 , the population after expansion is significantly larger than the population before expansion.

We first calculated, in Arlequin version $3.5⁵⁸$, confidence intervals of the two mutation parameters for simulated data sets of stable and expanding populations. These results supported an expansion scenario for all populations which had non-overlapping 99% confidence intervals for θ_0 and θ_1 . Therefore, to identify whether the expansion was spatial or demographic, the observed mismatch distribution was compared to mismatch distribution simulated under: a pure demographic expansion model^{55,80}; and a spatial expansion model^{21,81}. A pure demographic expansion model assumes that a stationary panmictic population has suddenly passed from a population size of N_0 to N_1 . This scenario has been shown to lead to star shaped gene genealogies⁵⁵, translating into an excess of rare mutations and into unimodal mismatch distributions⁸⁰. A model of spatial expansion assumes an initial panmictic population with a limited distribution range. The population's distribution range increases over time and space, leading to subdivided populations in the sense that individuals in geographic proximity are more likely to mate with each other than with remote individuals²¹. The genetic signature under this scenario is similar

to a pure demographic expansion model when the number of migrants between the subdivided populations is high. However, when the number of migrants is small, the mismatch distribution is bimodal^{$21,81$} and the first peak in the distribution is the result of comparison of identical sequences.

 The raggedness index (Rag.) was used to estimate how well the model fitted the observed data. Confidence intervals (CI) and P-values for model rejection were obtained by 10,000 parametric bootstrapping. Calculations were carried out in Arlequin version 3.5⁵⁸ and graphs were generated with R version $2.12.0^{46}$.

Causal modelling. We used causal modelling on resemblance matrices^{23,24} to investigate: (i) the combination of variables driving genetic differentiation between colonies; and, (ii) the causal relationships between these variables and genetic differentiation. To investigate the combination of variables driving genetic differentiation, we identified *a priori* three variables that could influence (or be influenced) by genetic distance: geographic distance, the presence of barriers and echolocation difference. The diagnostic set of statistical tests of the seven possible organisational models including these variables was then evaluated (see Supplementary Table S4). Only models where all of their diagnostic set of statistics fit the observed data were considered as fully supported. All possible causal relationships between variables present in the fully supported organisational model(s) were then investigated (see Supplementary Table S5), except those where geographic distance would be a dependant variable as geographic distance is not influenced by echolocation or genetic distance²³. As originally described^{23,24}, we used Mantel⁸² and partial Mantel tests⁸³ to assess the support of the organisational models. All tests (one-sided) were conducted with R version $2.12.0^{84}$ using the package 'ecodist' version $1.1.4^{59}$, and significance was assessed with 9999 permutations⁸⁵. We used a sequential Bonferroni technique to correct for multiple testing⁸⁶ considering an overall significance level of 0.05. Each data set was first translated into a pairwise distance matrix that represented the difference between each pair of colonies. Tests were carried out on 12 colonies for which we obtained echolocation (n=3958 calls in total) and genetic data (n=442 individuals).

 A matrix of genetic distance between colonies was calculated using the F-statistics option in Genepop version $4.0.6^{53,87}$. The locus CTC1 was excluded when calculating the genetic distance matrix as it was shown to be under selection (see main text) and therefore, was not representative of neutral variation. Geographic Distance Matrix Generator version 1.2.2⁸⁸ was used to generate a geographic distance matrix between colonies from geographic coordinates. The absence of suitable habitat and roosts can represent a strong barrier to gene flow^{89,90}. Therefore, the presence/absence of at least one barrier between two colonies was entered in a matrix as '1'= presence or '0'= absence. A barrier was set as present if the limestone formation was discontinuous between two colonies for more than 1 km. The 1 km threshold corresponds to the maximum distance from the cave *C. thonglongyai* has been shown to forage at^{34} . Finally, we calculated a matrix of absolute difference of the mean echolocation call frequency between colonies as in Yoshino *et al.*⁹⁰. As explained previously, to remove any potential bias caused by the non-independence between calls from the same recordings, we generated 10,000 data sets by randomly picking only one call per recording. We then calculated 10,000 matrices of absolute echolocation difference and performed the partial Mantel test for each new matrix. We therefore obtained 10,000 Mantel *r* and P-values from which we calculated the median value.

Estimates of dispersal distance. Given the very narrow and elongated distribution of *Craseonycteris* along the Kwae river valley¹⁶, we considered the habitat as one-dimensional⁹¹. We then estimated the average axial dispersal distance (σ) using the following formula^{91,92}: $b=1/(4D\sigma^2)$

where *b* is the slope of the regression of *Fst*/(1-*Fst*) *versus* distance and *D* the population density. We used two population density values here, the observed density $(258/7.6=33.95 \text{ individuals/km}^2, \text{calculated})$ after ¹⁶) and an estimate of the effective density (taken as $1:10^{th}$ of the observed density^{93,94}). Although this estimate assumes a population in equilibrium, it was shown to be relatively robust to various scenarios of temporal and spatial fluctuations of demographic parameters^{95,96}. The estimate was shown to be biased when demographic expansion occurred less than 100 generations ago⁹⁵, which is unlikely to be the case for *Craseonycteris* given estimates of time since expansion calculated based on mismatch distributions. Indeed, based on mismatch distributions dating (τ=2*ut*) ⁸⁰, a mutation rate of 2050% per million year (95% confidence interval: 1130-2680) would be required for the expansion to have occurred 100 generations ago. This value is well above the maximum values reported in any mammalian species

for any mtDNA fragment, including the D-loop alone^{97,98}, suggesting that the expansion occurred more than 100 generations ago.

Atmospheric attenuation calculations. The main parameters affecting atmospheric attenuation of sound in the air are relative humidity, temperature and frequency of the sound. To keep the detection distance of prey and objects constant, it is predicted that the peak frequency used in warmer places should be lower than that used in colder places; also, lower frequencies should be used in more humid places³¹. Relative humidity, temperature and frequency interact in a complex way to attenuate sound in the air⁹⁹ and therefore influence the maximum distance of prey detection³⁰. We calculated sound attenuation for peak frequencies from each of the thirteen sites in Thailand according to formulas presented in Bazley⁹⁹. Monthly averages of maximum values of relative humidity and temperature at 2 m were obtained from the EMP Climate database provided by the Center for Energy and Processes of Mines ParisTech/Armines (Resolution, 5 arcmin). The average peak frequency at each site was measured from recordings of free flying bats around caves' entrances (*cf.* "Echolocation Calls" paragraph in the Methods).

 If the *Craseonycteris* change in frequency between the different colonies in Thailand was an adaptation to produce calls with similar attenuation throughout the range, we would expect to observe similar attenuation values for the different colonies. Also, these attenuation values should be less variable between sites than if all colonies had a similar peak frequency throughout Thailand. To check this, we calculated attenuation for the observed frequency at each site as well as for a theoretical constant frequency for all colonies (considered values of 74 kHz, 75.5 kHz and 77 kHz). Calculations were carried out in R version $2.12.0^{46}$ and are presented in Figure 6.

Echolocation drift. We investigated the short-term and long-term population size of *C. thonglongyai* in Thailand with particular emphasis on the difference between the Northern/Central and Southern colonies. The actual colony size was estimated at 20 sites spread throughout the species range by counting individuals emerging from caves at dusk (see details in Puechmaille *et al.*¹⁶). The average colony size was then compared between the Northern/Central and Southern colonies using a one-sided Wilcoxon Mann-Whitney Rank Sum Test¹⁰⁰ implemented in the 'coin' package in R version 2.12.0⁸⁴. We used the genetic

diversity found within colonies as a proxy for long-term effective population size. High genetic diversities are associated with large effective population sizes and reduced drift whereas low genetic diversities are associated with reduced population sizes and increased drift 25 . To test for differences in genetic diversity between the Northern/Central and Southern colonies, we grouped colonies into two groups corresponding to colonies situated north *versus* south of the echolocation break and compared their genetic diversity at the nuclear and mitochondrial level in Fstat version $2.9.3.2^{101}$ as described in the "Genetic diversity indices**"** section in the Methods. For the two tests detailed above, the North-South separation of colonies was identical to what is described in the legend of Figure 5. The same tests were applied to compare genetic diversity between the Thai and Myanmar populations.

Selection tests. Individuals from different populations living in different environments often vary genetically at key sites in the genome due to adaptation to different local conditions^{26,102}. Low genetic differentiation across populations may indicate balancing selection, whereas high genetic differentiation suggests positive directional selection¹⁰³. The 15 microsatellites from 462 individuals were tested for selection based on the *Fst* outlier approach^{27,104}. Colony P (see Fig. 1) was not included in the analysis because of the limited sampling (one individual captured). This method evaluates the relationship between *Fst* and H_F (expected heterozygosity) under the assumption of neutrality. We carried out the test under an island model as implemented in $LOSITAN⁶⁰$ and under a hierarchical island model¹⁰⁵ as implemented in Arlequin version 3.5. The first model has been tested under a wide variety of conditions, including isolation by distance²⁷, and is robust to deviations from the theoretical model used^{27,104,106}, while the second method has been specifically designed to model hierarchical population structure.

 The tests for the island model were performed using the program LOSITAN under the infinite alleles and stepwise mutation models. Significance was assessed by the proportion of simulated values having higher (balancing selection) or lower (positive selection) *Fst* values than expected under neutrality (α level set to 0.01). As recommended by Beaumont and Balding¹⁰⁶, the cut-off *P*-value for significance was adjusted for the number of loci being tested using a sequential Bonferroni correction⁸⁶. As both models of mutation gave similar results, only results of the infinite allele model are presented. The "Neutral mean *Fst*" option was used when running LOSITAN. This option allows the program to run

once to determine a first candidate subset of selected loci in order to remove them from the computation of the neutral *Fst*, which is generally a better approximation of the neutral Fst^{27} . The option "Force mean *Fst*" was also used to get a simulated mean *Fst* close to the observed one found in the real dataset⁶⁰. All other options or settings were left as default and 50,000 simulations were run.

For microsatellites under positive selection, we estimated the allelic and genotype frequency distribution across the different colonies and identified geographic areas (sites) where frequencies dramatically changed, potentially indicating areas subject to divergent selection. To confirm that the divergent selection was occurring within this geographic region we ran the selection tests for the colonies on either side of potential selection border.

The test for the hierarchical island model was run in Arlequin with 10 simulated groups, 100 demes per group and 10,000 simulations. Excoffier *et al*¹⁰⁵ showed that "*the splitting of existing groups is less detrimental than the lumping of those groups*" so we ran the test for different population structure with up to five groups. The population structures were defined according to the analysis of the first three axis of a PCA analysis on colonies (data not shown). The first structure involved two groups with all Southern *versus* Northern/Central colonies (as in Fig. 5b-c). The second structure was represented by three groups, keeping the Southern group from the previous structure and splitting the Northern group into two (A, B, D and E, F, G). The third structure was constituted of the two groups from the North (as detailed above), plus three groups resulting from the separation of the Southern colonies (1: K; 2: H, I, J and 3: L, M, N, O).

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