



## The Predictive Power of $R_0$ in an Epidemic Probabilistic System

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**Abstract.** An important issue in theoretical epidemiology is the epidemic threshold phenomenon, which specify the conditions for an epidemic to grow or die out. In standard (mean-field-like) compartmental models the concept of the *basic reproductive number*,  $R_0$ , has been systematically employed as a predictor for epidemic spread and as an analytical tool to study the threshold conditions. Despite the importance of this quantity, there are no general formulation of  $R_0$  when one considers the spread of a disease in a generic finite population, involving, for instance, arbitrary topology of inter-individual interactions and heterogeneous mixing of susceptible and immune individuals. The goal of this work is to study this concept in a generalized stochastic system described in terms of global and local variables. In particular, the dependence of  $R_0$  on the space of parameters that define the model is investigated; it is found that near of the ‘classical’ epidemic threshold transition the uncertainty about the strength of the epidemic process still is significantly large. The forecasting attributes of  $R_0$  for a discrete finite system is discussed and generalized; in particular, it is shown that, for a discrete finite system, the pretentious predictive power of  $R_0$  is significantly reduced.

**Key words:** Cellular-Automata, Epidemics, Monte Carlo,  $R_0$

### 1. Introduction

It is nowadays recognized that the phenomenon of health-disease in human communities only may be understood by considering complex and dynamic interrelations among several factors operating simultaneously in multiple spatiotemporal and organizational scales. In fact, the healthy and sick individual suffers uninterruptedly the effects of the microbiological evolution, the antropogenic environmental and ecosystem stress and many others misdeeds resulting from socioeconomic inequalities. Therefore, it is not surprising to find out the proliferation of a myriad of methodological tools employed during the development of the epidemiological research.

Among this methodological mosaic the mathematical and computer (or simulation) modeling of communicable and infectious disease comes as a hypothetical-deductive approach whose scope consists primarily in understanding and manipulating, *a priori* and to predictive purposes, the underlying mechanisms behind the origin and diffusion of epidemic events. As a matter of fact, the attempt of understanding in what conditions pathogenic agents (once invaded a host population) could establish themselves as an infection (the transmission of pathogens from one host to another) resulted in the development of one of the most important and thoroughly discussed concepts in infectious disease modeling as early as in the beginning of the last century, namely the *epidemic threshold*. Thus, in writings of R. Ross (1909) [1] the so-called *mosquito theorem* was the first recognition of a quantitative threshold deducing that it was not necessary to eliminate mosquitoes totally in order to eradicate malaria. Two decades later would testify the publication of the classic Kermack-McKendrick's (1927) [2] paper that definitely consolidated the threshold concept in epidemiologic literature. In this deterministic *SIR* model (*S* stands for susceptibles, *I* for infected, and *R* for removed) an epidemic process is considered to evolve only when the density of susceptible individuals is greater than a threshold value  $S_c$ . Bartlett (1957) [3], based in a large amount of collected data of disease incidence in industrialized countries introduced thirty years later another expression linking microbial invasion and threshold parameters: the *critical community size*, that could explain the fade-out patterns of measles epidemics.

However, the inherent individual heterogeneity and probabilistic local nature of interindividual relationships has been traditionally neglected in state-variable models like this; in fact, in this population level approach all behavioral and individual variability are diluted into the intercompartmental rates and densities or number of mean individuals – as *S*, *I* or *R* compartments – described in terms of partial or ordinary differential equations. Nevertheless, it was subsequently possible to express the epidemic threshold in a way perhaps much more intuitive when the focus changed to consider the infected host or the parasite itself, instead of looking at the density or number of susceptible. In this perspective the threshold condition that determines whether an infectious disease will spread in a susceptible population has been described through the so-called *basic reproduction number* or also denominated as *basic reproductive rate*, commonly denoted by  $\mathcal{R}_0$  [4]. For microparasites such as viruses or bacteria it may be biologically understood as the average number of secondary cases produced or caused by one infected individual during its entire infectious period in a completely susceptible population. Thus, the intrinsically individual based perspective of this threshold concept should not be underestimated since the reproduction number can link the inside-host evolutionary or pathogenic dynamic (microscale) and transmission process at population level (macroscale). From a purely deterministic point of view it appears intuitively evident that if  $\mathcal{R}_0 \geq 1$  the pathogen can undoubtedly establish itself in a host population and, at least, an endemic regime will settle down. But this is a short-sighted prediction since, specially to directly transmitted disease in finite

populations, the mechanisms that ensure the maintenance of the parasite within a community depends critically on the way as the individuals interact one another, sometimes unforeseeable.

In this work we analyze limitations of the predictive power of the  $R_0$  parameter (as classically formulated) for the spread of a disease: Alternatively to population level approach and state-variable models, stochastic inter-individual interactions are also used and its implications on the predictive attributes of the *basic reproduction number*  $\mathcal{R}_0$  are studied through a simplified model: a lattice based model including infectious period in that individual interactions are straightforwardly described in terms of global ( $\Gamma$ ) and local ( $\Lambda$ ) variables, which in turn can be tuned out to simulate respectively the populational mobility and geographical neighborhood contacts.

The remainder of this paper is organized as follows. In the next section it is presented a general formalism to the evolution of a population invaded by an infection. The formalism is then applied in section 3, where concepts involving  $\mathcal{R}_0$  and the threshold phenomenon are discussed in order to define an invasion criterion for the infection and evaluation of  $\mathcal{R}_0$ . The results are discussed in section 4. Although this work will be mainly concerned on  $\mathcal{R}_0$  as a function of the model's parameters, the formalism presented in what follows can be applied to study a variety of epidemic scenarios.

## 2. The Model System

Consider a discrete dynamical system (discrete space and discrete time) where a population of  $N$  individuals is distributed on the sites of a toroidal lattice  $\mathbf{M} = \{m_{ij}\}$  – with  $i$  and  $j$  varying from 1 to  $L$  ( $N = L \times L$ ). Each individual site  $m_{ij}$  is assigned to receive three personal specific attributes: (1) a spatial address or lattice position ( $i, j$ ); (2) a set of three possible *status*, namely,  $s, i$  and  $r$ , specifying a clinic disease stage of each particular individual, which represent, respectively, the conditions of *susceptible* (subject to be infected by a contagious agent), *infectious* (effectively transmitter of contagious agents) and *removed* (recovered or immune); and finely (3) an infectious period  $\tau$ , specifying how many units of time an infected individual can propagate the contagious agent. Note that  $\sum s + \sum i + \sum r = N$ , with  $N$  constant.

Such a system is suitable mainly for describing a single epidemic in a closed system (no birth or migration). The choice of such reduced model, however, is not far-fetched because, as already mentioned above, the main interest here involves only very short period of time, so that the dynamics of host births, migration, etc., are largely irrelevant. [5]. The dynamic evolution of the population is described, step-by-step, by a set of *a priori* stated interaction rules, and assumes that each new configurational state of the system (described here by the geographical address ( $i, j$ ) of each individual and by the instantaneous number of susceptibles  $S(t)$ , infectives  $I(t)$ , and removed individuals  $R(t)$ ) depends only on its previous

state. Hence, for the present purpose the spread of the disease in the population is considered as being governed by the following rules:

1. Any susceptible individual may become infected with a probability  $p_S$ . An infected susceptible becomes infective after an average latency time  $\tau_l$  (assumed here as  $\tau_l = 0$ , without loss of generality).
2. Infectives are removed deterministically from the system (becoming immune) after an infectious period  $\tau$ , that for simplicity is considered as constant for all infected individuals.
3. Once in the removed class the individual participate only passively in the spreading of the infection (eventual topological blocking) by a period of immunity greater than the complete epidemic process.

During one time step, the three preceding rules are applied synchronously to all sites in the lattice; the present model, therefore, can be viewed as a simple two-dimensional cellular automaton. Actually, it is an adaptation of automata network to standard *SIR* models for studying the spread of infectious diseases.

In this work, the probability  $p_S$ , which is intended to be probability per unit of time, is taken as the superposition of the local and global influences, in order to unify the individual-based (contacts among nearest neighbors) and the standard mean-field (homogeneously interacting population) approaches. Therefore, one assumes that disease transmission occurs with a total infection probability  $p_S$  written as

$$p_S = \Gamma p_G + \Lambda p_L, \quad (1)$$

where the pre-factors  $\Gamma$  and  $\Lambda$  are weight parameters tuning the short (cluster formation) and long-range (mean-field type) interactions; it is also required that  $\Gamma + \Lambda = 1$  in order to satisfy the probabilistic requirement  $0 \leq p_S \leq 1$ .

The global influence  $p_G$  amounts to the probability of a susceptible to become infective due to the ubiquity of  $I(t)$  infected individuals (mean-field). So one can expect that in the limit of large  $N$  ( $N \rightarrow \infty$ ), in each time step, any susceptible may become infected with probability

$$p_G = \frac{\rho}{N} \sum_{\{k,l\}} \delta_{i,\sigma(k,l)} \quad (2)$$

where  $0 \leq \rho \leq 1$  is one of the model parameters: it limits the maximum value of  $p_G$  and is related to the intrinsic mobility of the population; the sum sweeps all lattice sites  $\{k, l\}$ , and  $\delta_{i,\sigma(k,l)}$  is the Kronecker delta function which assumes the value ‘one’ when the state  $\sigma$  of the site  $(k, l)$  corresponds to the infectious state  $i$ , and ‘zero’ otherwise ( $\sigma(k, l)$  can be  $s, i$  or  $r$ ). Actually, the sum in the Equation 2 just counts the instantaneous number of infectious individuals  $I(t)$  in the population.

On the other hand, the local term  $p_L = p_L(i, j)$  is the probability of a susceptible individual (located at the site  $(i, j)$ ) contracting infection due to  $n$  infectives first and second neighbors ( $0 \leq n \leq 8$  is a integer number corresponding to all

possible combinations of  $(i+\xi, j+\xi)$ , with  $\xi = 0, 1, -1$ ). Therefore, let  $\lambda \in [0, 1]$  be the probability of a particular susceptible when just one of its neighbors is infective. Hence,  $(1 - \lambda)^n$  will be the probability for not contracting the disease when exposed to  $n$  infectives. Therefore, the chance of he (or she) contracting the disease in a unit of time is[6]

$$p_L = 1 - (1 - \lambda)^n. \quad (3)$$

Thus, when  $\lambda = 1$  the infection spreads deterministically, with 8 nearest neighbors to any infective being infected (the choice for equipotent first and second neighbors was adopted because the use of only the four nearest neighbors is unduly restricting).

The expression for  $p_G$  is a convenient and simple way for describing the populational mobility. It is based on the *mass action law*, borrowed from the chemistry, and gets new meaning here under the perspective of pairwise spatially disordered interactions through the population elements. In this sense, it is a result of the small-world effect, and so became a particular version of the small-world lattice of Watts and Strogatz [7]

This simple approach allows to study in great detail the dynamical behavior of the model in the full space of control parameters  $\lambda$  and  $\rho$ , and the local and global balance pre-factors  $\Lambda$  and  $\Gamma$ . Therefore, the system is governed by  $p_S$  (Eq.1) and  $\tau$ , and its temporal evolution is determined by updating the lattice synchronously at each time step through the application of the three rules above.

### 3. $\mathcal{R}_0$ and The Threshold Phenomenon

The probability  $p_L$  as in the Eq.(3) [8], [6], and in a number of alternative forms[9], has been employed in the analogy between percolation and epidemic. Since that the critical value  $p_c$ , in which random clusters grow to infinite size, is know (analytic or numerically) for any lattices,  $p_c$  may be used as a powerful general criterion for ‘epidemic spread’[10], [9]. However, due to the traditional importance of the concept of  $R_0$  in the epidemic scenario, this threshold is generalized for finite discrete systems, as described above, in order to show its relevance for an intrinsically individual based perspective of the problem.

The overall structure of the model presented here shows the interplay of two types of transmission mechanisms by assuming that each infectious individual interacts strongly (physically) with their few susceptible neighbors, and uniformly and weakly with each particular susceptible in the population of susceptibles. Thus, the local mode of transmission  $p_L$  incorporates the individual-based component from the perspective of the susceptible individuals, the actual (physical) contacts that each susceptible experiences, and the global probability  $p_G$ , due to intrinsic populational mobility, which may be viewed as resulting of a mean-field (discrete) approach, in the sense that the disease transmission to each susceptible individual

also depends on the instantaneous total number of infectious individuals in the population.

To better appreciate the consequences of this formulation, it is firstly run simulations for the extreme values of the tuning pre-factors through the procedure described above. These cases allow to recover the two modes of transmission in its pure form corresponding to the (i) homogeneous mixing approximation (mean field), when  $\Gamma = 1$ , and to the (ii) percolation process (the transmission occurs by localized individual contact), when  $\Lambda = 1$ . Furthermore, it is considered the *damage*  $\Delta I$  on a susceptible population due to just one infected individual ( $I(0) = 1$ ) landing in a totally susceptible population ( $S(0) = N - 1$  and  $R(0) = 0$ ) during the infectious period  $\tau$ ; to calculate  $\Delta I$  it is considered only the number of new infected individuals in the population after  $\tau$  time steps, ignoring infections from the victims of these first infected individual (operationally, it is enough to consider the latency  $\tau_l > \tau$ , that is, a latent period of infection greater than the infectious period). The Figures 1a and 1b show, respectively, the simulation results for  $\Delta I$  as function of  $\rho$  for  $\Gamma = 1$ , and the behavior of  $\Delta I$  as a function of the contact probability  $\lambda$  for  $\Lambda = 1$  (that is,  $\Gamma = 0$ ); the system size considered in most of the simulations presented here was  $L = 100$  (population size  $N = L \times L = 10^4$ ), although some extra different sizes  $L/2$  and  $2L$  were also used in order to verify finite size effects.  $\Gamma = 1$  and  $\Lambda = 1$  are the two limiting cases usually taken as references in studying the effect on the system when both mechanisms are superposed; the amount  $\Delta I$  is obtained after  $\tau = 10$  time steps (covering exactly the infectious period) and was estimated as an average over 31 independent simulations (what is equivalent to verify the establishment of infection on 31 distinct populations with the same pattern of contacts among the individuals).

The linear pattern observed for  $\Delta I$  vs  $\rho$  means that the present stochastic approach reproduces qualitatively the classical basic reproductive number  $\mathcal{R}_0$  if one identifies  $\Delta I$  as the average number of secondary cases that an infectious individual causes. Indeed, the linear relation  $\Delta I = [\frac{\rho}{N} S(0)] \tau$  fits pretty well the data shown in Figure 1a, and so one may consider that infectives make contacts at a mean rate  $[\frac{\rho}{N} S(0)]$  throughout an infectious period of length  $\tau$  (note that for large enough populations  $\frac{\rho}{N} S(0) \rightarrow \rho$ ). On the other hand, when  $\Lambda = 1$ , the amount  $\Delta I$  represents  $\mathcal{R}_0$  for the case where individuals interact only with their spatial nearest neighbors, and so its values saturates at  $\Delta I = 8$  for  $\lambda \gtrsim 0.3$ . For each particular run, significant fluctuations on  $\Delta I$  are observed (mainly for smaller  $N$ ) but averaging over 31 runs is enough to smooth considerably the curves, as shown in Figure 1.

Before to proceed through the application of the present formulation, some comments regarding the definition of the  $\mathcal{R}_0$  are in order. The basic reproduction number has been widely used as a predictor parameter conceived to indicate the epidemic potential of a pathogen once it has introduced in a totally susceptible population. In fact, to deterministic and continuous (in space and in time) population-system models the future fate of an infectious agent has been expressed

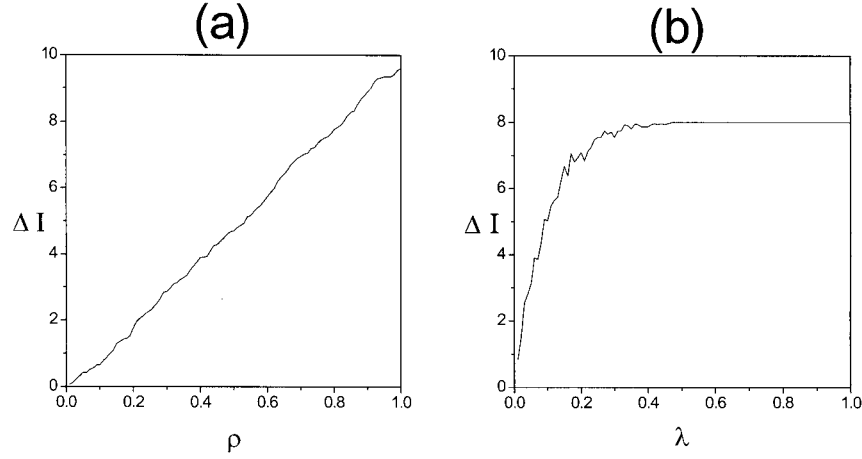


Figure 1. Average damage  $\Delta I$  due to just one infectious individual on the susceptible population  $S = N - 1$ , for two extreme linear cases. [a] –  $\Gamma = 1$ : the amount  $\Delta I$  changes linearly with the intrinsic mobility  $\rho$ , as can be expected from Equation (2). [b] –  $\Lambda = 1$ : the amount  $\Delta I$  increases rapidly with the infection probability  $\lambda$  due to local (physical) contact, and saturates at  $\Delta I = 8$  for  $\lambda \gtrsim 0.3$ , as one can infer from Equation (3).

through the threshold condition. Accordingly, when  $\mathcal{R}_0 > 1$ , infections can invade a totally susceptible population and persist; if  $\mathcal{R}_0 < 1$ , the disease then dies out and can not establish itself. To the special condition  $\mathcal{R}_0 = 1$ , there is an endemic regime in that the unique initial infectious case reproduces subsequently just one infectious secondary case and so on.

This assumption in modelling of the establishment of an infection (which is possibly wrong) [11] will be partially preserved here to have the classical treatment as a reference but, indeed, to capture more realistic or probable practical situations is of interest that the ‘first analytical look’ at a population be considered when the epidemic process is already in course. For instance, at the initial time  $t_0$  one may consider the arbitrary situation in that  $I(t = t_0) \gg 1$  at the same time that the number of removed individuals is also large, and then ask the question: What is the value for the *reproduction number* in this case? To answer this question one may generalize the concept of  $\mathcal{R}_0$  as the normalized average number  $\mathcal{R}(t_0; \tau)$  of secondary cases (reproductive ratio) about the time  $t_0$ , due to  $I(t_0)$  infectious present in the population at  $t = t_0$ , through the following expression

$$\mathcal{R}(t_0; \tau) = \frac{\sum_{t_0}^{t_0+\tau} \left\langle \sum_{\{k,l\}_s} (\Gamma p_G + \Lambda p_L) \right\rangle_n}{I(t_0)}, \quad (4)$$

where the brackets means an average on a set of  $n$  independent runs in the time interval  $[t_0, t_0 + \tau]$ , and the sum over  $\{k, l\}$  sweeps all sites occupied by individuals in the *status*  $s$  (susceptibles). Note that all the instantaneous extensive and intensive conditions of the population, at any arbitrary time  $t_0$ , are all taken into account,

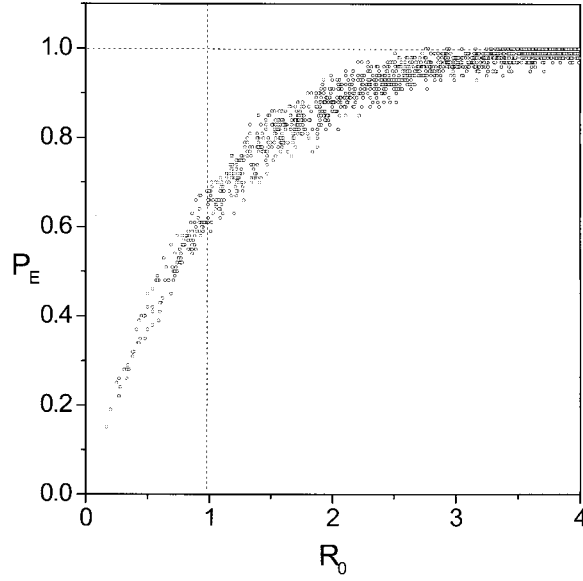


Figure 2. The epidemic probability as a function of average reproduction number  $\mathcal{R}_0$ . The tuning pre-factor are fixed at  $\Gamma = \Lambda = 0.5$ , and the parameter  $\rho$  and  $\lambda$  are choosing from the interval  $[0, 0.2]$ . For  $\mathcal{R}_0 \simeq 1$  epidemics are observed in about 60% of the events (in a population of size  $N = 10^4$ ).

as for example, the sites in the removed *status* randomly scattered through the population (acting as epidemic shield protectors), and the set remaining infectious time  $\tau(k, l; t_0)$  for each individual in the *status*  $i$  located at the site  $(k, l)$ . These conditions certainly affect the epidemic process and the progression of the epidemic process depends in some how on the *reproduction number's* value, (that is, if  $\mathcal{R}(t_0; \tau) > 1$  or  $< 1$ ). But, as already mentioned above, the initial condition  $I(t = t_0) = 1$  will be deliberately used in the present work in order to maintain the original intention of comparing the traditional deterministic definition of the basic reproductive ratio  $\mathcal{R}_0$  with the present stochastic approach.

In order to infer how the intrinsic stochastic nature of the epidemic process affects the predictive attributes of  $\mathcal{R}_0$ , the concept of *epidemic probability*  $P_E$  is introduced. Numerically it is estimated directly from the simulation experiments based on the algorithm of the previous section. Indeed, it is just given by the ratio  $P_E = n_e/n$ , where  $n_e$  is the number of runs in that at least one susceptibles was infected during the infectious period, and  $n$  is the total number of runs or experimental populations. The probability  $P_E$  may be expressed as function of the mean reproduction number  $\mathcal{R}_0$ , which also is determined from the same simulation experiments by using the Equation 4 above. In the Figure 2 it is shown the resulting  $P_E$  as a function of  $\mathcal{R}_0$  with  $\Gamma = \Lambda = 0.5$  and  $\rho$  and  $\lambda$  varying in the interval  $(0 - 0.2]$ . The large number of scattered points in the graph, mainly at larger



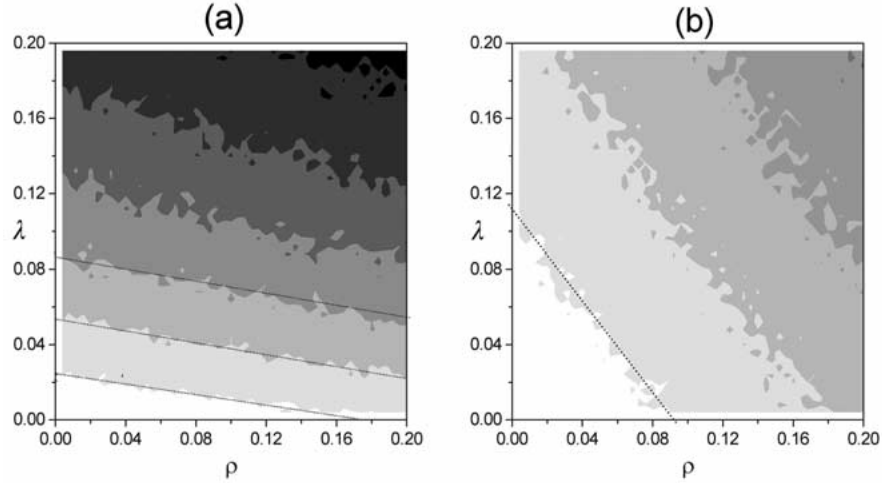


Figure 3. Reproduction number  $\mathcal{R}_0$  as function of the model parameters  $(\rho, \lambda)$  obtained by averaging over 100 independent realizations. Each strip, identified by a different gray tone, corresponds to a range of value for  $\mathcal{R}_0$  according to: white,  $0 \leq \mathcal{R}_0 < 1$ ; light gray,  $1 \leq \mathcal{R}_0 < 2$ ; and so on. At the limit of very large populations ( $N \rightarrow \infty$ ) the slope  $\alpha$  (dotted lines), which roughly delimitates each region, can be obtained using Equation (5) – see text; giving  $\alpha = \frac{\Gamma}{8\Lambda}$ . Therefore, in [a]  $\Gamma = \Lambda = 0.5$ , so  $\alpha = -0.125$ ; and in [b]  $\Gamma = 0.9$  and  $\Lambda = 0.1$ , giving  $\alpha = -1.125$ , whose values are closely reproduced by the results.

$\mathcal{R}_0$ , is an intrinsic aspect of this graph due to the fact that in the parameter space  $(\rho, \lambda)$  there are different combinations of  $\rho$  and  $\lambda$  resulting in approximately the same values for  $\mathcal{R}_0$ , as it is illustrated in Figure 3. Therefore only the stochastic component of such scattering of points may be reduced by increasing the number of runs used in the averaging procedure.

The amount  $P_E$  tends to saturates at  $P_E \simeq 1$  when the value of  $\mathcal{R}_0$  is sufficient large ( $\mathcal{R}_0 \gtrsim 3$ ), so that the epidemic spread in the population almost always is observed. Furthermore, the results showed in the Figure 2 means that only for large enough  $\mathcal{R}_0$  (actually  $\mathcal{R}_0 > 3$ ) one can be sure about an epidemic development in the population, while that, even for  $\mathcal{R}_0 < 1$  there is still a possibility to have an epidemic spread. Therefore, from the epidemic control perspective, reducing the effective reproductive number to a level below one, upon vaccination, for instance, could be a potential problem of strategy since that for  $\mathcal{R}_0 \lesssim 1$  in about 60% of events this strategy will fail, that is, an epidemic process should be established with chance of 60% for  $\mathcal{R}_0 \cong 1$ , under the conditions of the present model. More pointedly, despite the claim of the threshold criterium, it is improbable to recognize (using only standard census data) the imminence of any epidemic disaster if the system is near to the threshold region.[12] The more accurate (although frustrating) criterium is to realize that, irrespective the value of  $\mathcal{R}_0$  that the level of vaccination forces, there is always a chance (even thought small) of the disease re-invading the population.

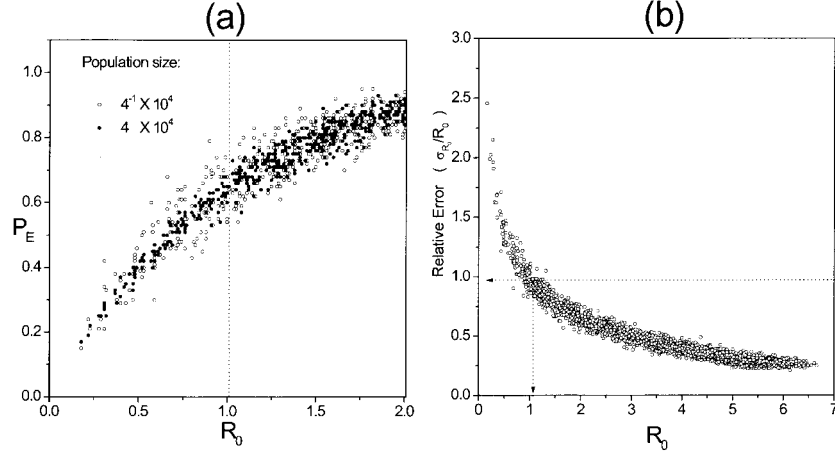


Figure 4. [a] The epidemic probability  $P_E$  vs  $\mathcal{R}_0$  for two systems:  $N = 4^{-1} \times 10^4$  (open circles) and  $N = 4 \times 10^4$  (dark circles) smaller fluctuations for the larger system is the most significant difference. [b] – The relative error decreases as  $\mathcal{R}_0$  increases; for  $\mathcal{R}_0 \simeq 1$  the absolute error is of the same magnitude of  $\mathcal{R}_0$  as a consequence of the averaging on ‘zeros’ and ‘ones’, mainly.

The same system size  $N = 10^4$  was employed in order to get all the results discussed above. However, in order to verify eventual effect of the system size on the results, two extra systems were considered, namely a smaller  $N = 4^{-1} \times 10^4$  and a bigger  $N = 4 \times 10^4$  system, but no significant difference was found. Clearly fluctuations are smaller for larger systems mainly because the chance of nucleation of closer cluster due to the global term  $\Gamma p_G$  decreases with the system size  $N$ , reducing then the chance of the magnification effect of the local term  $\Delta p_L$  on eventual clusters located nearby enough each other. The Figure 4a shows for  $P_E$  vs  $\mathcal{R}_0$  (in the interval  $0 < \mathcal{R}_0 \leq 2$ ) for two different system sizes; note that the size effect is pronounced only on the second moment (dispersion of the data) of the distribution of  $P_E$  for each  $\mathcal{R}_0$ . More precisely, the Figure 4b shows the normalized standard deviation (relative error)  $\sigma_{\mathcal{R}_0}$  as function of  $\mathcal{R}_0$  for the larger  $4 \times 10^4$  system. A decreasing  $1/\mathcal{R}_0$ - like behavior for the relative error is a consequence of the averaging of integer quantities, that is:  $\mathcal{R}_0 = (0 \times n_0 + 1 \times n_1 + 2 \times n_2 + 3 \times n_3 + \dots) / \eta$ ; where  $\eta = n_0 + n_1 + n_2 + \dots$ , and  $n_k$  is the number of experiments in which exactly  $k$  susceptibles were infected.

Finally, the numerical equivalence between  $\mathcal{R}'_0$  estimated by an analytical approximation and  $\mathcal{R}_0$  calculated by simulation is verified. For this purpose  $\mathcal{R}'_0$  is considered in the limit of large populations ( $N \rightarrow \infty$ ) by taking the mean number of susceptible infected by just one infective during its infectious period  $\tau$ , through the following direct expression

$$\mathcal{R}'_0 = \left\{ \Gamma \left[ \frac{\rho}{N} S(0) \right] + \Lambda[\lambda 8] \right\} \tau. \quad (5)$$

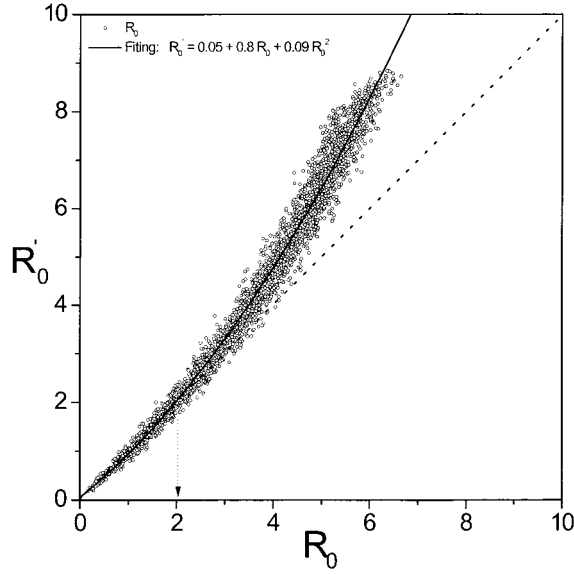


Figure 5. Numerical equivalence between  $\mathcal{R}'_0$  estimated by an analytical approximation and  $\mathcal{R}_0$  calculated by simulation.

The Figure 5 shows the parametric graph of  $\mathcal{R}'_0$  vs  $\mathcal{R}_0$  where they are calculated, respectively, by Equation (5) above and by simulation using the proposed probabilistic approach represented in Equation (4), with  $I(t_0 = 0) = 1$ . Strong correlation between the two ways for estimating the *basic reproduction number* is kept only for values of  $\rho$  and  $\lambda$  not too large ( $\mathcal{R}_0 \lesssim 2$ ) because during the time  $\tau$ , the local term that composes  $\mathcal{R}_0$  (Eq.4) may change from zero up to eight, while this limit is not present in the Equation (5). However, that is enough in order to validate the conclusions about the predictive attributes of  $P_E = P_E(\mathcal{R}_0)$  because  $\mathcal{R}'_0$  and  $\mathcal{R}_0$  are numerically equivalent: the result given by Equation 5, although intuitive, follows from a stochastic representation of the classical SIR model [13].

#### 4. Final Comments

In this paper a stochastic version of the original *SIR* model (involving only single epidemics) was introduced with the main purpose of to characterize and re-interpret the conditions for the establishment of an epidemic in a population through the concept of *basic reproduction number* ( $\mathcal{R}_0$ ). A peculiar characteristic of the present approach is the assumption that the probability of a susceptible individual become infective is a superposition of the local and global influences. Using as initial configuration just one single infected individual in a fully susceptible population, condition frequently used to define  $\mathcal{R}_0$ , it was demonstrated that the discrete character of a finite population reduces the pretentious predictability of the threshold

criteria, and so it is, indeed, an incomplete predictive tool since that, irrespective to the value of  $\mathcal{R}_0$ , an epidemic has a finite probability to establish itself, due the inherent stochastic nature of any finite epidemic system.

Indeed, more consistent derivation of  $R_0$  has been tried, even though using the same classical deterministic approach, due to the too widely estimate obtained to  $R_0$ , which in recent applications for the smallpox have varied from  $R_0 = 1.5$  to 20 [14]. Rather than just a caricature of the original formulation of  $\mathcal{R}_0$ , the approach presented in this paper may be viewed as a simpler and generic alternative for investigating the spread of diseases in a population, which may greatly facilitate the analysis of a number of distinct epidemic scenarios. Particularly, a system with increasing topological complexity can be easily tackled. For example, one may consider the practical situation in that, at an arbitrary initial time  $t_0$ , the population has already many infectious individuals (that is,  $I(t_0) \gg 1$ ), and also many immunes scattered through the population (working as epidemic shield) and then try to answer the question: What is the value for the *reproduction number* in this case?

Finally, as a major challenge that this ‘microscopic’ approach can handle, one may think on the possibility of incorporating in the traditional definition of  $\mathcal{R}_0$  the underlying evolutionary dynamics of the pathogenic agent. This view is in contrast with the standard epidemiological models, which tend to use a constant absolute parasite fitness  $\mathcal{R}_0$ . However, more detailed considerations on the investigation of this avenue of research is left for a future contribution.

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