## Supported 40,4072 (see 4207522400) Ekman et al. 10.1073/pnas.1207523109

## SI Materials and Methods

Multivariate Functional Magnetic Resonance Imaging (fMRI) Analysis. Defining network nodes. Preparation [cue: color  $(C)$  or motion  $(M)$ ] and no-preparation conditions [cue: no preparation (N)] were modeled using a general linear model (GLM). Parameter estimates were registered to the MNI152 standard brain template in 2-mm resolution and subjected to an across-participant multivariate pattern analysis, using the PyMVPA software package (1). No spatial smoothing was applied. To search in an unbiased fashion for voxels that discriminate between preparation and nopreparation conditions, we used a searchlight analysis (2, 3), which examines information in local spatial patterns of brain activity surrounding each voxel. Around each gray matter voxel  $v_i$  we first defined a spherical cluster (radius = 8 mm) and extracted the parameter estimates for preparation and nonpreparation. This yielded two pattern vectors that were used to train a pattern classification algorithm. To assess the classification performance, we trained a sparse logistic regression classifier on the pattern vectors of all but one subject (default  $\lambda$  $= 0.1$ ). The pattern vectors of the remaining subject were used as an independent test set. This leave-one-subject-out cross-validation procedure was repeated until each subject had been used in the test set once. The resulting classification accuracy (across subjects) was assigned to the respective voxel  $v_i$ . This procedure was repeated for each voxel and resulted in a 3D accuracy map, restricted to gray matter.

To estimate the significance of classification accuracy for each voxel  $v_i$ , we conducted a nonparametric permutation test (4). To this end, the searchlight analysis was repeated 10,000 times with permuted labels. The  $P$  value for every voxel  $v_i$  was then calculated as the fraction of permutation samples that were greater than or equal to the accuracy actually observed when using correct labels. The resulting P-map (i.e., Fig. 1B) was corrected for multiple comparison with false discovery rate (FDR)  $(q = 0.05)$  and thresholded accordingly.

Brain areas involved in task preparation were used to define network nodes (5–7) for subsequent analyses. To derive nonoverlapping and equally sized network nodes, we used an iterative approach. We started by defining a sphere (radius  $= 12$  mm) around the voxel with the lowest  $P$  value, as identified by the searchlight analysis. This sphere became the first node in the network. Then we moved to the second-lowest P value and defined the next sphere. A sphere was created only when it did not overlap with previously generated spheres. In the case of an overlap, this voxel was ignored and we selected the voxel with the next lowest P value. This procedure was repeated until no additional spheres could be created and resulted in 70 nodes. The Montreal Neurological Institute (MNI) coordinates of the spherical center coordinates are displayed in Table S1.

Graph Construction and Control Analyses. Our approach to concatenate relevant segments of the raw data time series was previously successfully applied (8–10), but might induce artifacts that need to be examined. Here we report a series of control analyses that aim to support the validity of our approach. First, we ruled out the possibility of slow mean shifts (e.g., related to changes in scanner signal) driving the correlation estimates, rather than task-related changes. To this end, we compared the average blood oxygen level dependent (BOLD) signal between all 12 subsequent task blocks, using an ANOVA across subjects. No significant differences were found  $[F(11, 96) = 0.31, P = 0.98]$ . Furthermore, to directly test the similarity of individual time courses throughout the experiment, we calculated the connectivity matrices (aggregated over all conditions) separately for each task block. Resulting connectivity matrices were tested for reliability, using the intraclass correlation coefficient (ICC) for each subject. The test showed very high consistency in the connectivity matrices across task blocks and the ICC varied across subjects between 0.89 and 0.98 [ $F(2,414, 26,554) \sim 9.14-51.65$ ; all  $P < 0.001$ ].

Second, the concatenation of the relevant time-series segments might cause spikes in the newly generated time series that could cause spurious correlations. To investigate this effect we compared the absolute difference of two volumes made adjacent by concatenation with the absolute signal difference at two naturally adjacent volumes. To this end we performed a nonparametric permutation analysis. For each subject, the averaged difference for the concatenated volumes was compared with 10,000 randomly chosen nonconcatenation differences. The P value was calculated as the amount of occurrences a difference at a nonconcatenation gave an equal or higher value than a difference at a concatenated point. For all subjects, this resulted in P values with  $P > 0.3$ , indicating that the concatenation does not induce more spikes in the signal than expected from a chance distribution.

Third, due to the average error rate of ∼20%, concatenated time series of incorrect trials consisted of fewer time points than those of correct trials. This might potentially influence the correlation estimates and lead to, e.g., lower correlations in the incorrect conditions. To test this effect, we compared the mean connectivity as the average over all entries of the adjacency matrix between all conditions (i.e., color/motion, correct/incorrect), using an ANOVA. No significant differences in mean connectivity were found  $[F(3, 32) = 0.56, P = 0.65]$ .

Fourth, we find converging evidence for our results, using a slightly different analysis approach (β-series correlation). A commonly used alternative to our time-series concatenation approach is to model each trial of interest with a separate predictor in a GLM and concatenate the resulting β-coefficients to a so-called β-series (11–13). The correlation analysis is then applied on the β-series. Applying the same inverse modeling approach as for the concatenation approach, we could classify color vs. motion preparation with an accuracy of 84.2% motion correct vs. incorrect with 72.2% accuracy and color correct vs. incorrect with 61.1% accuracy. Note, however, whereas this can be seen as converging evidence for our results, it might not be regarded as a fair comparison of both analysis approaches. The β-time series is substantially shorter (only one  $\beta$  per trial) and differences in the classification accuracy might be due to higher power in the raw time-series approach.

Stability of the Core Across Task Conditions. One of our main findings is the enhanced closeness of task-relevant visual areas to the network core. However, because the k-core decomposition was applied to the averaged connectivity matrices across all conditions, it is important to show that changes in core closeness cannot be attributed to changes in the core structure itself. To this end, we performed the k-core decomposition with all conditions separately. As the results show, the core remains highly stable across conditions (Fig. S3). This is in line with the observation that incorrect trials are not characterized by a global reduction in connectivity, but rather by specific core–periphery interactions.

Graph Theory Formulas. Network measures were calculated with networkX [\(http://networkx.lanl.gov](http://networkx.lanl.gov)), and a detailed description

of all formulas can be found elsewhere  $(7)$ . N is the set of all nodes in the network and *n* is the number of nodes.  $w_{ij}$  is the connection weight between *i* and *j* ( $i, j \in N$ ).

Local weighted degree. The weighted degree of a node  $i$  is the sum of all edge weights connected to that node:

$$
k_i = \sum_{j \in N} w_{ij}.
$$

Global degree.

V<br>Z

$$
k=\frac{1}{n}\sum_{i\in N}k_i.
$$

Average shortest path length.

$$
L = \frac{1}{n} \sum_{i \in N} L_i = \frac{1}{n} \sum_{i \in N} \frac{\sum_{j \in N} \frac{1}{j \neq i} d_{ij}}{n-1},
$$

where  $d_{ij}$  is the shortest path length between nodes i and j. For the path length calculation of weighted networks the connection weights  $w_{ij}$  were transformed to distances, using the inverse transform  $wd_{ii} = 1/w_{ii}$ .

Local clustering coefficient. The clustering of each node  $i$  is the fraction of possible triangles that exist around  $i$ ,

$$
C_i = \frac{2t_i}{k_i\,(k_i-1)},
$$

where  $t_i$  is the number of triangles around node  $i$ , and

$$
t_i = \frac{1}{2} \sum_{j, h \in N} \left( w_{ij} w_{ih} w_{jh} \right)^{1/3}.
$$

Global clustering coefficient.

$$
C_{\text{glob}} = \frac{1}{n} \sum_{i \in N} C_i,
$$

where  $C_i$  is the local clustering coefficient. Assortativity.

$$
r = \frac{1 - \sum_{(i,j)\in L} w_{ij} k_i k_j - \left[1 - \sum_{(i,j)\in L} \frac{1}{2} w_{ij} (k_i + k_j)\right]^2}{1 - \sum_{(i,j)\in L} \frac{1}{2} w_{ij} (k_i^2 + k_j^2) - \left[1 - \sum_{(i,j)\in L} \frac{1}{2} w_{ij} (k_i + k_j)\right]^2}.
$$

Transitivity.

$$
T = \frac{\sum_{i \in N} 2t_i}{\sum_{i \in N} k_i (k_i - 1)}.
$$

Modularity. Modularity was calculated using an iterative algorithm (14). The results were averaged over 100 repetitions,

$$
Q = \frac{1}{l^w} \sum_{i,j \in N} \left[ w_{ij} - \frac{k_i^W k_j^W}{l^W} \right] \delta_{m_i m_j},
$$

where  $m_i$  is the module containing node i, and  $\delta_{m_i m_i} = 1$  if  $m_i = m_j$  and 0 otherwise. The concept of finding modules by modularity optimization is fundamentally different from the concept of k-shell decomposition. Optimizing the modularity

function will yield a network structure of densely connected nodes within a module and sparse connections between modules. In contrast, k-shell decomposition will find a densely connected "core" that is well connected to the periphery, but the periphery is mostly sparsely interconnected with many connections crossing the core.

Efficiency global.

$$
E_{\text{glob}} = \frac{1}{n(n-1)} \sum_{i \neq j \in \mathcal{N}} \frac{1}{d_{ij}}.
$$

Efficiency cost.

$$
E_{\rm cost} = E_{\rm glob} - K,
$$

where the wiring costs  $K$  are determined as the sum of edge weights between the connected nodes, divided by the maximum possible value of edge weights (15). Efficiency local.

$$
E_{\text{local}} = \frac{1}{n_{Gi}(n_{Gi} - 1)} \sum_{j,k \in G_i} \frac{1}{d_{j,k}},
$$

where  $n_{Gi}$  is the number of nodes in the subgraph  $G_i$ . Efficiency nodal.

$$
E_{\text{nodal}}(i) = \frac{1}{(n-1)} \sum_{j \in G} \frac{1}{d_{ij}}.
$$

Small-world scalar.

$$
S = \frac{C/C_{\text{rand}}}{l/l_{\text{rand}}}.
$$

C and l are coefficients of the tested network (clustering and average shortest path length) and  $C_{\text{rand}}$  and  $l_{\text{rand}}$  are the respective values of a randomized version of the original network with the same degree distribution (16, 17).  $C_{\text{rand}}$  and  $l_{\text{rand}}$  were obtained by calculating the average values of 100 randomizations. Betweenness centrality.

$$
BC_i = \frac{1}{(n-1)(n-2)} \sum_{s,t \in N} \frac{n_{st}^i}{g_{st}},
$$

where  $n_{st}^{i}$  is the number of shortest paths from s to t that pass through i and  $g_{st}$  is the total number of shortest paths from s to t. Current flow betweenness centrality. In contrast to betweenness centrality, which uses shortest paths, current-flow betweenness centrality (18), also known as random-walk betweenness centrality (19), uses an electrical current model for information spreading. Given a supply b, the throughput of a node  $i \in N$  is defined to be

$$
\tau(i) = \frac{1}{2} \left( -|b(i)| + \sum_{e: i \in e} |x(e^{-i})| \right).
$$

Current flow betweenness centrality is defined by

$$
CFBC_i = \frac{1}{(n-1)(n-2)} \sum_{s, t \in N} \tau_{st}(i).
$$

Eigenvector centrality. Eigenvector centrality is an extension of the degree centrality measure. Instead of counting the amount of neighbors, the eigenvector centrality of a node is defined proportionally to the sum of importance of its neighbors (20). We

used the eigenvector centrality implementation of Networkx (21), which uses the power method to find the eigenvector for the largest eigenvalue of the adjacency matrix.

Closeness centrality. Closeness centrality (20) is a centrality measure that reflects the inverse average distance to all nodes in the network,

$$
\text{CLC}_i = \frac{1}{l_i},
$$

where

$$
l_i = \frac{\sum_{j(\neq i)\in N} d_{ij}}{n-1}.
$$

PageRank. The PageRank algorithm (22) is an iterative procedure that was developed in the context of web searching to find how often a node will be visited during random network traversal,

$$
PR_i = \frac{1 - \lambda}{N} + \lambda \sum_{n_j \in M(n_i)} \frac{PR(n_j)}{O(n_j)},
$$

where  $M(n_i)$  is the set of nodes that link to  $n_i$ , PR $(n_i)$  denotes the PageRank of node  $v_j$ , and  $O(n_j)$  is the out degree of the predecessor node  $n_j$ . The PageRank algorithm was originally designed for directed graphs. Here we apply it to undirected graphs with default  $\lambda = 0.85$ , where the out degree equals in degree equals degree.

Vulnerability. The vulnerability (23) is a measure of centrality. Vulnerability at a node  $n_i$  is the relative change in the sum of distances between all node pairs when excluding that node,

$$
V_i = \frac{L_G - L_H}{L_G},
$$

where  $L_G$  is the sum of distances between all nodes including  $n_i$  and  $L_H$  is the sum of distances between all nodes without  $n_i$ . Core closeness.Core closeness is an extension of closeness centrality informed by the network's core–periphery structure. Closeness centrality is calculated for a node in the periphery P with respect to the core C.

The core closeness of a node  $i \in P$  and  $i \notin C$  is given by

$$
\text{CCL}_{i\in P} = \frac{1}{l_i},
$$

where

$$
l_i = \sum_{j \in C} d_{ij}.
$$

Core centrality. Core centrality is an extension of betweenness centrality informed by the network's core–periphery structure. It reflects the number of shortest paths with start and end points in the periphery that pass through the core. The concept of core centrality is analogous to the previously introduced rich-club centrality (24),

4. Nichols TE, Holmes AP (2002) Nonparametric permutation tests for functional neuroimaging: A primer with examples. Hum Brain Mapp 15:1–25.

$$
\text{CC} = \sum_{s, t \in P} \frac{n_{st}^C}{g_r^P},
$$

where C and P are distinct subgraphs of the graph G.  $n_{st}^C$  is the number of shortest paths from s to  $t \in P$  that pass through C and  $g_{st}^P$  is the total number of shortest paths from s to  $t \in P$ .

Correlation of Graph Theory Metrics. In this study we use a large selection of different graph metrics of which several are based on similar concepts (e.g., centrality) and therefore might be highly correlated. To explore these relationships we calculated the Pearson cross-correlation of the used local graph metrics separately for all conditions (Fig. S1A). The resulting correlation coefficients varied between −0.3 and 0.7. As shown before (25), the intercorrelation of the graph metrics differed between experimental conditions, which confirms previous findings (26) that, depending on the network topology, these graph metrics differ in their ability to extract complex network characteristics. However, a subset of graph metrics appeared to be highly correlated, namely degree, efficiency nodal, eigenvector centrality, closeness centrality, and PageRank. Also, betweenness centrality and current flow betweenness centrality showed a moderate correlation.

To further explore whether the intercorrelation had an effect on the classification analysis, we tested the ability of different graph metrics to distinguish between color and motion preparation. To this aim, we applied the classification analysis to individual local graph metrics. As the results showed (Fig. S1B), none of the highly correlated metrics was able to distinguish between color and motion preparation. Importantly, all metrics that contributed to the prediction of the combined model also resulted in significant classification accuracies when regarded in isolation (including betweenness centrality and current flow betweenness centrality). The results also show that the integration into one statistical model leads to a further increase in classification performance, highlighting the advantage of our inverse network modeling approach.

Although our study employs a large set of graph theoretical metrics, no measures were used that quantify the efficiency of physical embedding of complex networks (27, 28). Future studies might therefore also consider using these measures to investigate network changes. Also, most of the graph metrics used are based on the concept of shortest paths, whereas recent studies introduced metrics to investigate information flow using random walks between any two nodes and emphasize the role of nonshortest paths for brain communication (29–31). These metrics have been shown to reveal changes in brain structure after stroke (32) and might be worth applying to fMRI in future studies.

Influence of Head Movement. It has been shown that head movement can have an impact on functional connectivity estimates (33). In our investigation it is especially important to ensure that there are no condition-specific head movements that might act as a confound for the classification analysis. To this end, we compared the root-mean-square (rms) of the estimated translation parameters (33) across conditions. An ANOVA found no significant differences  $[F(3, 32) = 0.07; P = 0.97]$ . Also a more detailed analysis of the frequencies of different displacements revealed no differences between conditions (Fig. S2).

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<sup>2.</sup> Clithero JA, Smith DV, Carter RM, Huettel SA (2011) Within- and cross-participant classifiers reveal different neural coding of information. Neuroimage 56:699–708.

<sup>3.</sup> Kriegeskorte N, Goebel R, Bandettini P (2006) Information-based functional brain mapping. Proc Natl Acad Sci USA 103:3863–3868.

<sup>5.</sup> Rubinov M, Sporns O (2010) Complex network measures of brain connectivity: Uses and interpretations. Neuroimage 52:1059–1069.

<sup>6.</sup> Bullmore E, Sporns O (2009) Complex brain networks: Graph theoretical analysis of structural and functional systems. Nat Rev Neurosci 10:186–198.

<sup>7.</sup> Rubinov M, Sporns O (2010) Complex network measures of brain connectivity: Uses and interpretations. Neuroimage 52:1059–1069.

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Fig. S1. (A) Correlation of local graph metrics. (B) Classification accuracy for all local graph metrics separately and combined in one statistical model.



Fig. S2. Frequency distribution of condition-specific mean head motion, averaged across all subjects.



Fig. S3. Stable core across different task conditions. Shown are the results of the k-shell scores for every node, separately for color correct/incorrect and for motion correct/incorrect. Nodes with the maximal k-shell value (red) form the core and all other regions <k(max) are the periphery.

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## Table S1. Center coordinates and core/periphery allocation of all 70 network nodes

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Table S1. Cont.

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B, bilateral; FO, frontal operculum; L, left hemisphere; Lat., lateralization; R, right hemisphere; SFG, superior frontal gyrus.

Table S2. Values of discriminative local graph measures for correct and incorrect color and motion preparation within the relevant nodes

	Correct				Incorrect			
	ВC	<b>CFBC</b>	C	<b>CCL</b>	ВC	<b>CFBC</b>	C	<b>CCL</b>
Color								
<b>V4L</b>	0.015		0.781	0.503	0.003		0.510	0.146
<b>V4 R</b>	0.010		0.729	0.476	0.006		0.519	0.167
hMT L	0.010		0.336	0.368	0.019		0.789	0.434
hMT R		0.019	0.365	0.366		0.029	0.767	0.415
Motion								
<b>V4 L</b>	0.003		0.410	0.201	0.026		0.789	0.418
<b>V4 R</b>	0.011		0.419	0.291	0.019		0.767	0.373
hMT L	0.013		0.794	0.475	0.007		0.326	0.220
hMT R		0.029	0.810	0.455		0.025	0.358	0.202

BC, betweenness centrality; C, clustering coefficient; CCL, core closeness; CFBC, current flow betweenness centrality.