



Ankle-Brachial Index determination and peripheral arterial disease diagnosis by an oscillometric blood pressure device in primary care: validation and diagnostic accuracy study

Journal:	<i>BMJ Open</i>
Manuscript ID:	bmjopen-2012-001689
Article Type:	Research
Date Submitted by the Author:	19-Jun-2012
Complete List of Authors:	Nelson, Mark; University of Tasmania, Menzies Research Institute Quinn, Stephen; Flinders University, Flinders Clinical Effectiveness Winzenberg, Tamia; University of Tasmania, Menzies Research Institute Tasmania Howes, Faline; University of Tasmania, Menzies Research Institute Tasmania Shiel, Louise; Monash University, Epidemiology and Preventive Medicine Reid, Christopher; Monash University, Department of Epidemiology and Preventive Medicine (DEPM)
Primary Subject Heading:	General practice / Family practice
Secondary Subject Heading:	Cardiovascular medicine, Public health
Keywords:	PRIMARY CARE, VASCULAR MEDICINE, PUBLIC HEALTH

SCHOLARONE™
Manuscripts

only

1
2 **Ankle-Brachial Index determination and peripheral arterial disease diagnosis by an**
3
4 **oscillometric blood pressure device in primary care: validation and diagnostic accuracy**
5
6 **study**
7
8
9

10
11
12 **Corresponding author:**
13

14 Mark R Nelson

15
16 Mark.Nelson@utas.edu.au

17
18 Private Bag 23

19
20 Hobart 7001, Tasmania, AUSTRALIA

21
22 Telephone: +61 3 6226 4734

23
24 Facsimile: +61 3 6226 4770

25
26
27
28
29
30
31
32
33
34
35 **Co-authors**
36

37 Stephen Quinn

38
39 Flinders University,

40
41 Adelaide, SA 5001, Australia

42
43
44
45
46
47
48
49 Tania M Winzenberg,

50
51
52 Menzies Research Institute Tasmania,

53
54
55 University of Tasmania,
56
57
58
59
60

1
2 Hobart, Tasmania 7001, Australia
3
4
5
6

7
8 Faline Howes
9

10 Menzies Research Institute Tasmania,
11

12
13 University of Tasmania,
14

15
16 Hobart, Tasmania 7001, Australia,,
17
18
19
20

21
22 Louise Shiel,
23

24
25 Department of Epidemiology and Preventive Medicine,
26

27
28 Monash University,
29

30
31 Melbourne 3004, Australia,
32
33
34
35

36
37 Christopher M Reid
38

39
40 Department of Epidemiology and Preventive Medicine,
41

42
43 Monash University,
44

45
46 Melbourne 3004, Australia,
47
48
49
50

51 **Keywords**
52

53
54 Sensitivity and specificity, Ankle-Brachial Index, peripheral arterial disease, oscillometric
55
56 device, blood pressure.
57
58
59
60

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

Word count - excluding title page, abstract, references, figures and tables.

2337

For peer review only

ABSTRACT

Objectives

To determine the level of agreement between a 'conventional' Ankle-Brachial Index (ABI) measurement (using Doppler and mercury sphygmomanometer taken by a research nurse) and a 'pragmatic' ABI measure (using an oscillometric device taken by a practice nurse) in primary care. To ascertain the utility of a pragmatic ABI measure for the diagnosis of peripheral arterial disease (PAD) in primary care.

Design

Cross-sectional validation and diagnostic accuracy study. Descriptive analyses were used to investigate the agreement between the two procedures using the Bland and Altman method to determine whether the correlation between ABI readings varied systematically. Diagnostic accuracy was assessed via sensitivity, specificity, accuracy, likelihood ratios, positive and negative predictive values, with ABI readings dichotomised and Receiver Operating Curve analysis using both univariable and multivariable logistic regression.

Setting

Primary care in metropolitan and rural Victoria, Australia between October 2009 and November 2010.

Participants

250 persons with cardiovascular disease (CVD) or at high risk (3 or more risk factors) of CVD.

Results

Despite a strong association between the two methods' measurements of ABI there was poor agreement with 95% of readings within ± 0.4 of the 0.9 ABI cut point. The multivariable C statistic of diagnosis of PAD was 0.89. Other diagnostic measures were sensitivity 62%,

1
2 specificity 92%, positive predictive value 67%, negative predictive value 90%, accuracy 85%,
3
4 positive likelihood ratio 7.3 and the negative likelihood ratio 0.42.
5
6

7 **Conclusions**

8
9
10 Oscillometric ABI measures by primary care nurses on a population with a 22% prevalence of
11
12 PAD lacked sufficient agreement with conventional measures to be recommended for routine
13
14 determination of ABI. Their diagnostic performance suggests that a pragmatic ABI lacked
15
16 sufficient sensitivity to diagnose PAD, but can reliably exclude it.
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

INTRODUCTION

Peripheral arterial disease (PAD) affects an estimated 27 million individuals in Europe and North America with 413,000 related hospital discharges per annum^{1 2}. These figures are likely to underestimate the true impact of PAD as those with the condition disproportionately suffer from other manifestations of CVD and are therefore likely to appear in coronary artery disease or stroke statistics. As a consequence there has been a call for better detection and management of the condition¹.

One of the simplest and most useful parameters to objectively assess lower extremity arterial perfusion, and thus diagnose PAD, is the Ankle-Brachial Index (ABI). This is the lower of the left and right ABI where each ABI is the ratio of the lower limb systolic blood pressure is compared to the higher systolic brachial blood pressure recording. The ABI can be used to screen for haemodynamically significant PAD and helps to define its severity. Patients with objectively documented PAD have a four- to six-fold increase in cardiovascular mortality over healthy age-matched individuals³. PAD is a stronger risk marker for myocardial and stroke morbidity and mortality than those who have already had such an incident event⁴. However, only 50% of people with PAD are symptomatic which is a significant issue in the detection of PAD⁵.

Between 2007 and 2009 19,500 oscillometric devices were distributed by the High Blood Pressure Research Council of Australia to physicians, mostly general practitioners (GPs). We had previously demonstrated that these devices were likely to improve blood pressure management in primary care⁶. The current study, Ankle Brachial Index Determination by oscillometric method IN General practice (ABIDING), sought to expand the utility afforded by these machines in primary care. Previous work done in those attending a specialist vascular

1
2 laboratory in the US demonstrated that patients could have their ABI reliably ascertained by
3 such devices compared to the conventional use of a Doppler ultrasound and mercury
4 sphygmomanometer⁷. It was therefore opportune to investigate if such measures were
5
6 pragmatic in primary care where the greatest opportunity exists to identify those with
7
8 undiagnosed PAD. Such persons are at very high risk for subsequent adverse cardiovascular
9
10 events that can be ameliorated through management of modifiable risk factors.
11
12
13
14
15

16
17 The primary aim of ABIDING was to establish if there was agreement between a pragmatic ABI
18 (measured by a practice nurse using an oscillometric blood pressure device), and a conventional
19 ABI (measured by a research nurse using mercury sphygmomanometer and Doppler devices). A
20 secondary aim was to ascertain diagnostic accuracy of the pragmatic approach for ascertaining
21
22 PAD.
23
24
25
26
27

28 29 **METHODS**

30
31 General practitioners (GPs) and participants were recruited through the REACH Registry
32 Victorian database. The international REACH Registry was a prospective, observational registry
33 designed to provide long-term follow-up (36 months) of patients at high risk of
34 atherothrombotic events. Globally 67,888 patients were involved in the REACH registry of
35 whom 2,782 were recruited from 281 general practitioners around Australia^{8 9}. Practices were
36 eligible for ABIDING if they had previously enrolled participants in the REACH registry and
37 had a practice nurse willing to participate or were willing to appoint a *locum tenens* nurse.
38 Eligibility criteria for REACH are published elsewhere but can be summarised as at entry
39 (March to June 2004) aged 45+ years, had known CVD or at least three atherosclerosis risk
40 factors, and were physically able to attend their usual general practice⁸.
41
42
43
44
45
46
47
48
49
50
51
52
53
54

55 **Participant recruitment**

1
2 All Melbourne (metropolitan) and Warrnambool (rural) Victorian study participants who had
3
4 consented to follow-up, who had been identified by their GPs as alive and for whom we had a
5
6 current address, were contacted by mail. If no reply was received from the participant within
7
8 four weeks, a second letter was sent and then a telephone call made. Participants were seen in
9
10 their usual GP's clinic between October 2009 and November 2010.
11
12

13 14 **Research and practice nurses**

15
16
17 Three experienced research nurses conducted the reference standard tests. They received
18
19 standardised training from a senior research nurse who was one of the operators. Practice nurses
20
21 were given training in situ by the research nurse and were observed by them. Because they
22
23 worked contemporaneously the research nurse was not blinded to the practice nurses results.
24
25
26

27 **'Conventional' and 'pragmatic' ABI estimation**

28
29
30 All participants were rested supine for five minutes before measurement. Doppler blood
31
32 pressure measurements (by research nurse) and automated oscillometric blood pressure
33
34 measurements (by practice nurse) were performed using cuffs that had bladders >80% of the
35
36 diameter of the arms and ankles measured.
37
38
39

40
41 Conventional measures involved Doppler blood pressure measurements in the lower limb made
42
43 with a Nicolet Vascular Doppler with a 5MHz probe. The cuff was inflated to 30mmHg above
44
45 systolic blood pressure and deflated slowly until a flow signal was detected over the dorsalis
46
47 pedis or posterior tibial arteries. Brachial artery systolic pressure was determined similarly but
48
49 utilising a stethoscope rather than a Doppler. The ABI for each lower extremity was calculated
50
51 as the pedal pressure divided by the higher of the two brachial pressures. PAD is defined as an
52
53 ABI <0.9 in either lower limb¹⁰. The mercury sphygmomanometer was calibrated by a certified
54
55 laboratory.
56
57
58
59
60

1
2 Research nurses were trained in the measurement of ABI and were certified prior to
3
4 commencement of the study. Practice nurses were simply observed and technique corrected if
5
6 required. Oscillometric measurements were made by the practice nurse on all limbs using a
7
8 standard automated blood pressure cuff system (OMRON HEM-907). This device is a validated
9
10 blood pressure measurement device^{11 12}. Oscillometric devices were new and therefore had
11
12 factory calibration.
13
14

15 16 17 **Statistical Methods**

18
19 Descriptive analyses were used to investigate the agreement between the two procedures using
20
21 the Bland and Altman method to determine whether measurements could be used
22
23 interchangeably and if the correlation between ABI readings varied systematically¹³. Although
24
25 the variability in the differences appeared to be proportional to the mean, applying a log
26
27 transformation to the data did not substantially alter agreement and so raw scores are presented.
28
29 Correlations between the paired readings were also calculated. Sensitivity, specificity, positive
30
31 and negative predictive values, and accuracy with exact 95% confidence intervals are reported,
32
33 where ABI readings taken under both conditions were dichotomised at 0.9 (reference standard).
34
35 The diagnostic accuracy was evaluated using Receiver Operating Curve analysis and quantified
36
37 as the area under the curve (AUC or C statistic), as determined using both univariable and
38
39 multivariable logistic regression. In the multivariable model we adjusted for age, BMI, gender,
40
41 and smoking status (never, former, current). The calibration of this model was validated using
42
43 the Hosmer-Lemeshow statistic¹⁴. We examined likelihood ratios, the ratio of the expected test
44
45 results in participants with PAD to those participants without. All results are reported with 95%
46
47 confidence intervals. All analyses were conducted using Stata version 12.0.
48
49
50
51
52
53

54 55 56 **Power calculations**

1
2 Assuming a type 1 error of 5% ($\alpha = 0.05$) a total sample of 250 participants provided 80% power
3
4 to detect systematic bias between the readings taken by the research and practice nurses if the
5
6 mean difference was 0.0255¹³. Eight participants were excluded as 6 pragmatic and 2
7
8 conventional ABI readings were absent. In all other cases each patient had at least one
9
10 conventional and pragmatic ABI reading (for the same leg). For a sample of 242 the difference
11
12 that we could detect was 0.0257. We expected strong correlations between ABI readings taken
13
14 using the different methods. Both calculations assumed a correlation between readings of 0.61
15
16 and standard deviations as reported in Benchimol *et al*¹⁵.
17
18
19

20 21 RESULTS

22
23 The flow chart of the study is shown in Figure 1. The characteristics of the ABIDING
24
25 population are shown in Table 1. There was no difference between those excluded and included
26
27 in the analysis for any trait that we measured. We also compared in Table 1 those diagnosed
28
29 with PAD vs. not using conventional ABI. Those with PAD were older ($p=0.003$) and more
30
31 likely to be female ($p=0.003$). Figure 2 shows there was poor agreement between pragmatic and
32
33 conventional determination of ABI with 95% of readings within ± 0.4 . Figure 3 shows
34
35 correlation between conventional and pragmatic ABI measurements, indicating a strong
36
37 association between the two measurements, despite the poor agreement. The distribution of
38
39 differences between the ABI measures is shown in Figure 4. These differences were regressed
40
41 on all possible confounders measured in our study, in both univariable and multivariable
42
43 models. There were no significant associations, suggesting that the differences were completely
44
45 random.
46
47
48
49
50
51

52
53 A 2x2 table of dichotomised conventional and pragmatic measurements is shown (Table 2). We
54
55 examined the two groups comprising the 36 participants where the PAD classification differed.
56
57
58
59
60

1
2 There were no differences in any measured trait between those groups (data not shown). The
3
4 respective pragmatic method diagnostic performance, assuming the conventional method as gold
5
6 standard, was sensitivity 62% (95% CI 47-75%), specificity 92% (87-95%), positive predictive
7
8 value 67% (52%-80%), negative predictive value 90% (85-94%) and accuracy 85% (80%-89%).
9
10 The Likelihood ratio for a positive result (LR+) was 7.3 (95% CI 4.4-12.0) and Likelihood ratio
11
12 test for a negative result (LR-) 0.42 (0.30-0.59). Area under the Receiver Operator Characteristic
13
14 curves (AUC / C statistic) of pragmatic ABI against the conventional ABI <0.9 and thus PAD
15
16 was 0.87 (95% CI 0.82, 0.93). The AUC from multivariable analysis (adjusting for age, gender,
17
18 BMI and smoking status) for all analyses were almost identical 89% (95% CI 84%-93%).
19
20
21
22
23

24 Based on the differences in Table 1 for those with PAD vs. not we conducted a post hoc
25
26 subgroup analyses on pragmatic vs. conventional ABI readings by gender, age (dichotomised as
27
28 young or old) and all pairwise combinations. The agreement between reading and diagnostic
29
30 criteria did not improve for any subgroup (data not shown). We also investigated (using
31
32 multivariable logistic regression) whether there was any evidence that disagreements were
33
34 systematic. There were no differences between in disagreement apart from current smokers
35
36 were more likely to produce readings that disagreed compared to non-smokers (p=0.025). A
37
38 subgroup analysis with current smokers removed did not alter the diagnostic criteria of the tests.
39
40
41
42

43 As could be expected in non-invasive testing there were no reported adverse events.
44
45

46 **DISCUSSION**

47
48 ABIDING demonstrated that use of oscillometric devices by general practice nurses to
49
50 determine ABI and therefore the presence of PAD had high specificity (92%) and negative
51
52 predictive value (90%), good accuracy (84%) but modest sensitivity (62%) and positive
53
54 predictive value (67%). The modest sensitivity and the LR+ 7.3 indicate that this test has little
55
56
57
58
59
60

1
2 value for confirming the presence of PAD. On the other hand high specificity and negative
3
4 predictive value suggests that the test has some value in ruling out the disease (i.e. when the test
5
6 is negative). This is in contrast to the experience in a specialist centre where their test
7
8 performance (both limbs in comparison to ABIDING lower of the 2 measures) was sensitivity
9
10 left/right leg 88/73% (62%), specificity 85/95% (92%), positive predictive value 65/88% (69%),
11
12 negative predictive value 96/88% (90%), LR+ left/right leg 5.9/14.6 (7.9) and LR- 0.14/0.28
13
14 (0.4)⁷. A good diagnostic test has a LR+ >10 and LR- < 0.1¹⁶. This difference in performance to
15
16 some extent may be accounted for by patient selection but is more likely due to operator
17
18 expertise. In the specialist centre, the mean age was 10 years younger and 53% were female
19
20 compared to only 22% in ABIDING. The respective prevalence of PAD was 32% and 22%.

21
22
23
24
25
26 ABI is a useful tool and is superior to clinical examination for identifying PAD⁹. However
27
28 screening whole populations is not practical. ABI ascertainment of PAD is most effective by
29
30 identifying high risk patients as we have done in ABIDING. By including high-risk and overt
31
32 CVD patients we were confident that we should get a distribution of ABI scores that included
33
34 PAD diagnostic scores and the outcome of the trial supports this (22% had PAD by the
35
36 conventional method). Doubini *et al* found age alone (70+) a useful predictor as 12.5% of the
37
38 screened population had PAD vs. only 2.5% of 50-69 years with at least 1 CVD risk factor but
39
40 no established CVD (diabetes, dyslipidaemia, hypertension, or smoking)¹⁷. Bendermacher *et al*
41
42 developed a clinical prediction model giving risk factor points per factor (age: 1 point per 5
43
44 years starting at 55 years; ever smoked: 2 points; currently smoking: 7 points; and hypertension:
45
46 3 points), showed a proportional increase of the PAD prevalence with each increasing risk
47
48 profile (range: 7.0-40.6%)¹⁸. The overall prevalence of PAD was 18%. They found with their
49
50 PREVALENT clinical prediction model (based on CVD risk factors), the GP was able to
51
52 identify a high-risk population in which measurement of ABI was useful.
53
54
55
56
57
58
59
60

1
2 If our method had been reliable it would have been readily implementable as Australian GPs
3
4 have ready access to oscillometric sphygmomanometers. More than 19,500 devices were
5
6 distributed on behalf of the High Blood Pressure Research Council of Australia, mostly to GPs,
7
8 over the years 2007 to 2009. Practice nurses were chosen rather than GPs as this approach is
9
10 also more likely to be implementable. A survey by Mohler *et al* of primary care clinicians
11
12 showed that most (88%) thought ABI to be feasible in that setting¹⁹. However, validation
13
14 studies have largely been conducted in specialist clinics in a variety of study populations rather
15
16 than in the primary care setting where most of the medical contact is likely to occur^{15 20-23}. The
17
18 one study done in the primary care setting used an ABIgram²⁴. Although the investigators
19
20 demonstrated its reliability, the use of this special piece of equipment would seem to effect is
21
22 acceptability as is the current situation.
23
24
25
26
27

28 **Study limitations**

29
30 The intervention was kept as simple as possible by using practice nurses to do single measures
31
32 on a device they were familiar with but did not receive extensive further training on. While this
33
34 means that this is simple to introduce into clinical practice the practice nurse performance may
35
36 have been improved by more intense training and repeated limb measurements.
37
38
39
40

41 **CONCLUSION**

42
43 Oscillometric ABI measures by primary care nurses on a population with a 22% prevalence of
44
45 PAD lacked sufficient agreement with conventional measures to be recommended for routine
46
47 determination of ABI. This pragmatic method may be used as a screening tool in primary care
48
49 but its diagnostic performance does not provide evidence sufficient for it to be used to diagnose
50
51 PAD.
52
53
54
55
56
57
58
59
60

ACKNOWLEDGMENTS, COMPETING INTERESTS, FUNDING

Participating GPs were Drs Sam Arber, Dean Arnot, Ian Barratt, Irmgard Chia, John Chin, Eric Choo, Andrew Chow, Michael Conos, Simon Cooper, Peter Coulton, Adrian Dabscheck, Deborah Davidson, Antonio De Sousa, Hillary Donald, Peter Eizenberg, Peter Enten, Roger Fagan, Peter Ferguson, Doron Gaddie, Jon Garland, Andrew Gault, Ian Gill, Robert Gingold, Peter Goy, Judith Heale, Eva Herold, Chris Hogan, Suresh Jain, Brendan Kay, John Kirmos, Susan Kloot, Con Lahanis, Daniel Lajoie, Maryanne Lancaster, Conway Leung, Leon Lewi, John Manderson, Frank Marano, Damian Marinucci, Larry McGrath, Elizabeth McNaughton, Margaret McNiff, John Meaney, Paul Molloy, Clare Mooney, Paul O'Hanlon, James Olesen, Michael Page, John Pattison, Gary Pellizzari, Annamarie Perlesz, John Philpot, Leon Piterman, Jock Plenderleith, Mark Preston, Christopher Priest, Andy Psaradellis, Michael Quinn, Jacqueline Rounsevell, Joseph Sakowsky, Mahinda Samararatna, Victor Sammut, David Slonim, Stuart Smith, Stephen Stowe, Ian Sutherland, Edwin Turner, Peter Wexler, Andrew White, Martin Williams, Stephen Williams, Keith Wing Shing, Richard Wrennall, and Michael Ziccone.

We would also like to acknowledge the contribution of practice nurses, Dr Nyi Nyi Tun who assisted with the literature review, and research nurses Christine Mulvaney, Sue Loftus, and Anne Bruce.

This study had ethical approval from the Human Research Ethics Committee (Tasmania) Network (H0010410) and Monash University Standing Committee on Ethics in Research involving Humans (2009000860), and was registered with the Australian and New Zealand Clinical Trials Registry (ACTRN12609000744257). It was funded by the RACGP Research Foundation (Cardiovascular Research Grant) and the National Health and Medical Research

1
2 Council (Project grant 544935), and was supported by the Primary Healthcare Research,
3
4 Evaluation and Development scheme. Oscillometric devices were loaned by the High Blood
5
6 Pressure Research Council of Australia.
7

8
9 No competing interests declared.
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

For peer review only

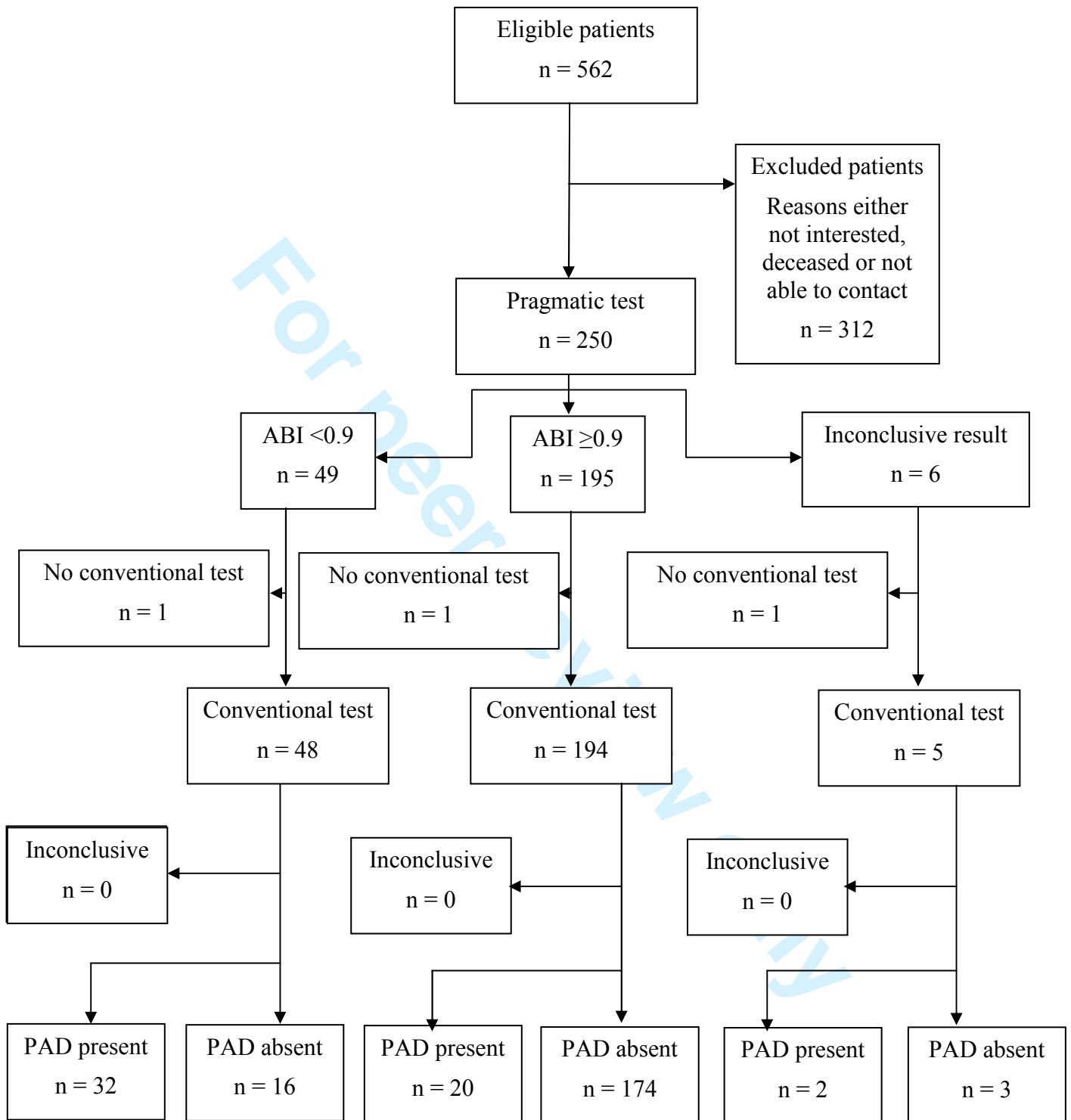


Figure 1. Flow diagram of a diagnostic accuracy in ABIDING as per STARD standard²⁵.

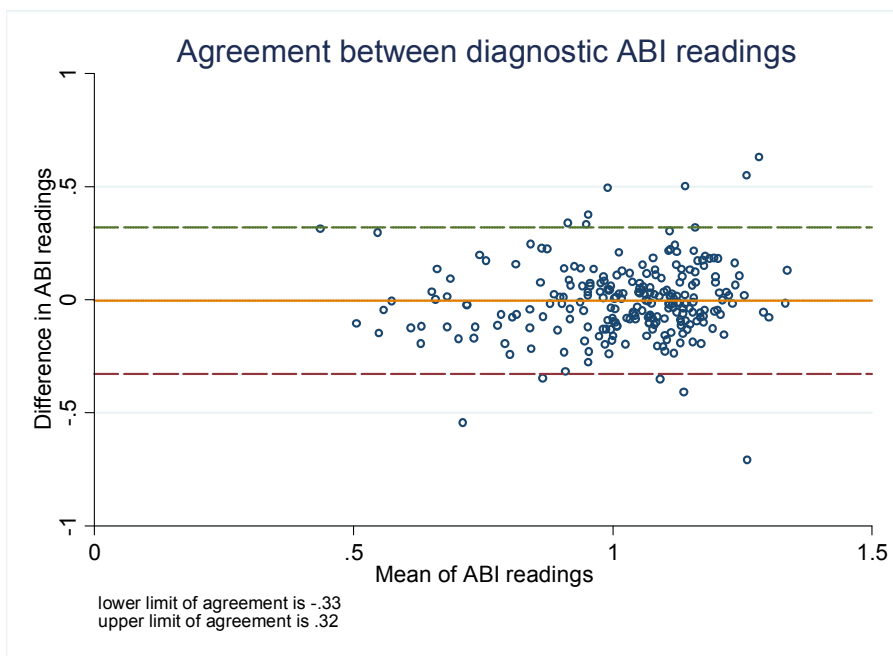


Figure 2. Agreement between pragmatic and conventional determination of ABI.

Peer review only

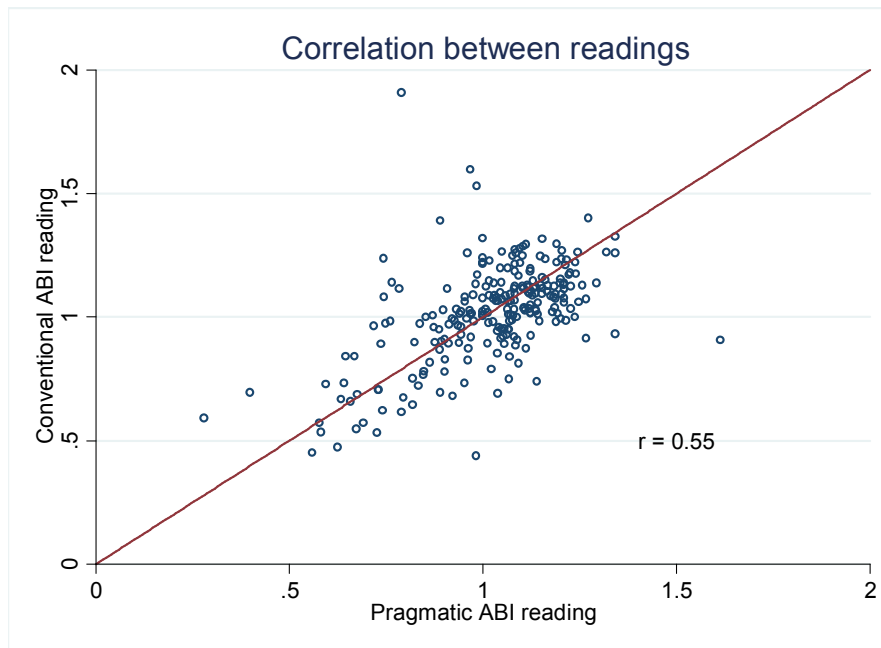


Figure 3. Correlation between pragmatic and conventional determination of ABI.

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

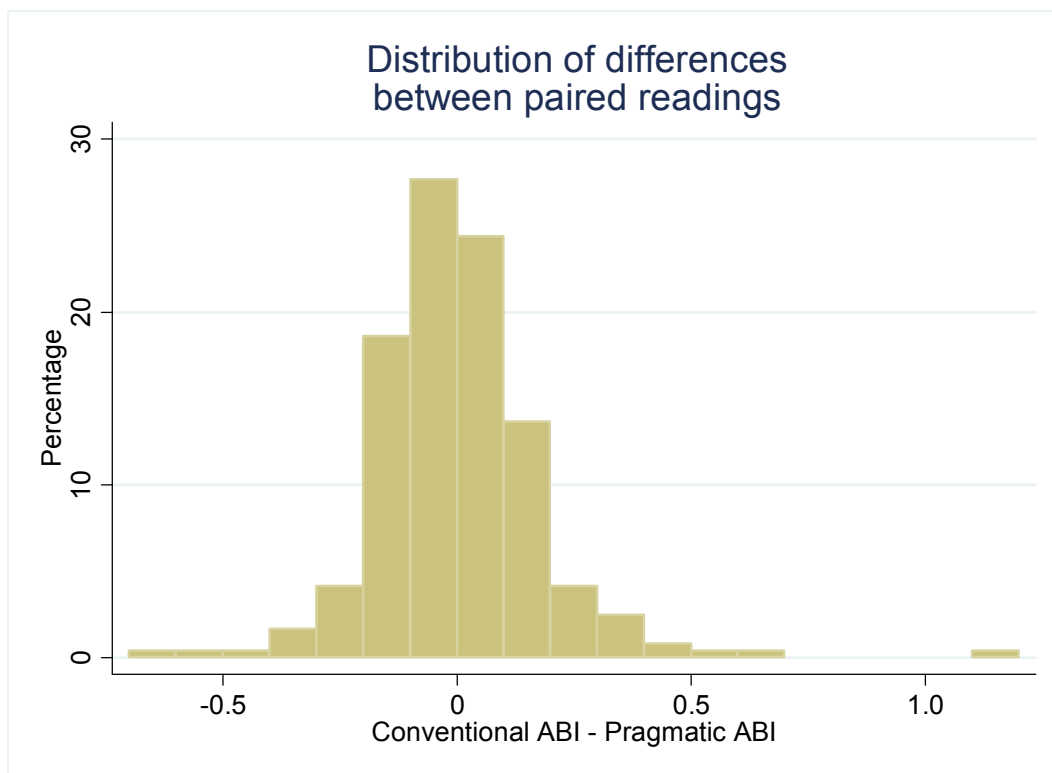


Figure 4. Distribution of the difference between the conventional ABI and the pragmatic ABI readings.

Review only

Variable	<i>Included</i>	<i>Excluded[†]</i>	<i>P for difference</i>	<i>Conventional ABI ≥ 0.9</i>	<i>Conventional ABI < 0.9</i>	<i>P for difference</i>
N	242	8		192	52	
Age in years	71.2(7.4)	72.5(7.2)	0.62	70.4(7.0)	73.9(8.3)	0.003
Male Sex (%)	167 (69.0)	5(62.5)	0.70	140(73.7)	27(52.0)	0.003
SBP (mmHg)	141.5(18.9)	153.9(20.7)	0.07	140.6(17.8)	144.5(22.4)	0.35
DBP (mmHg)	76.7(9.9)	82.1(10.6)	0.13	77.0(9.8)	75.5(10.4)	0.55
BMI (kg/h ²)	27.5(4.4)	29.3(5.9)	0.26	27.5(4.4)	27.2(4.4)	0.63
Waist	99.9(10.8)	99.1(13.0)	0.84	100.1(10.4)	99.2(12.3)	0.60
Smoking status			0.15			0.31
Never	98(40.7)	6(75.0)		81(42.9)	17(32.7)	
Former	131(54.4)	2(25.0)		100(52.9)	31(59.6)	

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49

Current	12(5.0)	0(0.0)		8(4.2)	4(7.7)	
---------	---------	--------	--	--------	--------	--

† 5 did not provide pragmatic ABI readings, 2 did not provide conventional ABI readings, 1 did not provide any readings.

Table 1: Characteristics of participants (included in and excluded from the analysis) and by conventional PAD status expressed as a mean (standard deviation) or N (%) as appropriate.

For peer review only

	PAD +ve (conventional ABI <0.9)	PAD -ve (conventional ABI ≥0.9)	Total
Test +ve (pragmatic ABI <0.9)	32	16	48
Test -ve (pragmatic ABI ≥0.9)	20	174	194
Total	52	190	242

Table 2. 2x2 table of conventional and pragmatic ABI determinations.

REFERENCES

1. Belch JJF, Topol EJ, Agnelli G, Bertrand M, Califf RM, Clement DL, et al. Critical Issues in Peripheral Arterial Disease Detection and Management: A Call to Action. *Arch. Intern. Med.* 2003;163(8):884-92.
2. Golomb BA, Dang TT, Criqui MH. Peripheral Arterial Disease: Morbidity and Mortality Implications. *Circulation* 2006;114(7):688-99.
3. McDaniel M, Cronenwett J. Basic data related to the natural history of intermittent claudication. *Ann. Vasc. Surg.* 1989;3:273-77.
4. Steg PG, Bhatt DL, Wilson PWF, D'Agostino R, Sr, Ohman EM, Rother J, et al. One-Year Cardiovascular Event Rates in Outpatients With Atherothrombosis. *JAMA* 2007;297(11):1197-206.
5. McKenna M, Wolfson S, Kuller L. The ratio of ankle and arm arterial pressure as an independent predictor of mortality. *Atherosclerosis* 1991;87:119 -28.
6. Nelson MR, Quinn S, Bowers-Ingram L, Nelson JM, Winzenberg TM. Cluster Randomized Controlled Trial of Oscillometric Versus Manual Sphygmomanometer for Blood Pressure Management in Primary Care (CRAB). *Am. J. Hypertens.* 2009; doi:10.1038/ajh.2009.55.
7. Beckman JA, Higgins CO, Gerhard-Herman M. Automated Oscillometric Determination of the Ankle-Brachial Index Provides Accuracy Necessary for Office Practice. *Hypertension* 2006;47(1):35-38.
8. Bhatt DL, Steg PG, Ohman EM, Hirsch AT, Ikeda Y, Mas J-L, et al. International Prevalence, Recognition, and Treatment of Cardiovascular Risk Factors in Outpatients With Atherothrombosis 10.1001/jama.295.2.180. *JAMA* 2006;295(2):180-89.

- 1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
9. Reid CM, Nelson M, Chew D, Connor G, DeLoose F. Australians @ Risk: management of cardiovascular risk factors in the REACH registry. *Heart Lung and Circulation* 2007;doi10.1016/j.hlc.2007.07.009.
10. Doobay AV, Anand SS. Sensitivity and Specificity of the Ankle–Brachial Index to Predict Future Cardiovascular Outcomes. *Arterioscler. Thromb. Vasc. Biol.* 2005;25(7):1463-69.
11. Assaad MA, Topouchian JA, Darne BM, Asmar RG. Validation of the Omron HEM-907 device for blood pressure measurement. *Devices and Technology* 2002;7(4):237-41.
12. White WB, Anwar YA. Evaluation of the overall efficacy of the Omron office digital blood pressure HEM-907 monitor in adults. *Devices and Technology* 2001;6(2):107-10.
13. Bland JM Altman DG. Statistical methods for assessing agreement between two methods of clinical measurement. *Lancet* 1986;1(8476):307-10.
14. Hosmer D, Lemeshow S. Applied Logistic Regression. 2nd ed. New York: John Wiley and Sons, Inc., 2000.
15. Benchimol A, Bernard V, Pillois X, Hong NT, et al. Validation of a new method of detecting peripheral artery disease by determination of ankle-brachial index using an automatic blood pressure device. *Angiology* 2004;55 (2):127-34.
16. Šimundić A-M. Measures of diagnostic accuracy: basic definitions.
<http://www.ifcc.org/ifccfiles/docs/190404200805.pdf>
17. Doubeni CA, Yood RA, Emani S, Gurwitz JH. Identifying Unrecognized Peripheral Arterial Disease Among Asymptomatic Patients in the Primary Care Setting. *Angiology* 2006;57(2):35-38.

- 1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
18. Bendermacher BLW, Teijinka JAW, Willigendael EM, Bartelink M-L, Peters RJB, de Bie RA, Büller HR, Boiten J, Langenberg M, Prins MH. A clinical prediction model for the presence of peripheral arterial disease--the benefit of screening individuals before initiation of measurement of the ankle-brachial index: an observational study. *Vasc Med* 2007;12(1):5-11.
19. Mohler ER 3rd, Treat-Jacobson D, Reilly MP, Cunningham KE, Miani M, Criqui MH, et al. Utility and barriers to performance of the ankle-brachial index in primary care practice. *Vasc Med* 2004;9(4):253-60.
20. Lee B, Campbell JS, Berkowitz P. The correlation of ankle oscillometric blood pressures and segmental pulse volumes to Doppler systolic pressures in arterial occlusive disease. *J Vasc Surg* 1996;23(1):116-22.
21. Nukumizu Y, Matsushita M, Sakurai T, Kobayashi M, Nishikimi N, Komori K. Comparison of Doppler and oscillometric ankle blood pressure measurement in patients with angiographically documented lower extremity arterial occlusive disease. *Angiology* 2007;58(3):303-8.
22. Ramanathan A, Conaghan PJ, Jenkinson AD, Bishop CR. Comparison of ankle-brachial pressure index measurements using an automated oscillometric device with the standard doppler ultrasound technique. *ANZ J of Surg* 2003;73(3):105-08.
23. Pan CR, Staessen JA, Li Y, Wang JG. Comparison of Three Measures of the Ankle-Brachial Blood Pressure Index in a General Population. *Hypertens Res* 2007;30(6):555-61.
24. Raines JK, Farrar J, Noicely K, Pena J, et al. Ankle/Brachial Index in the primary care setting. *Vascular and Endovascular Surgery* 2004;38(2):131-36.

1
2 25. Bossuyt PM, Reitsma JB, Bruns DE, Gatsonis CA, Glasziou PP, Irwig LM, et al. The
3
4 STARD Statement for Reporting Studies of Diagnostic Accuracy: Explanation and Elaboration.
5

6
7 *Clin Chem* 2003;49(1):7-18.
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

For peer review only

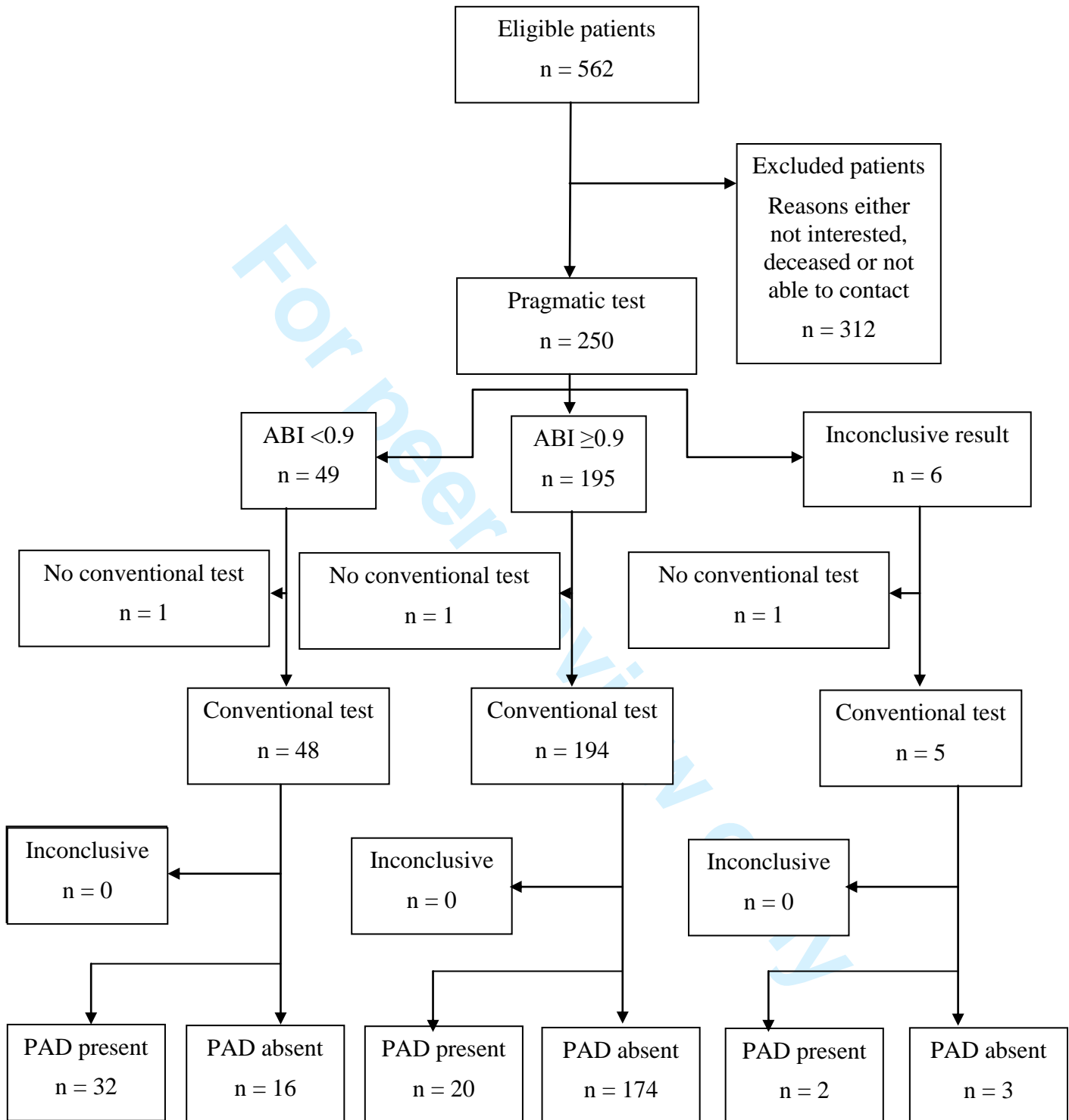


Figure 1. Flow diagram of a diagnostic accuracy in ABIDING as per STARD standard²⁵.

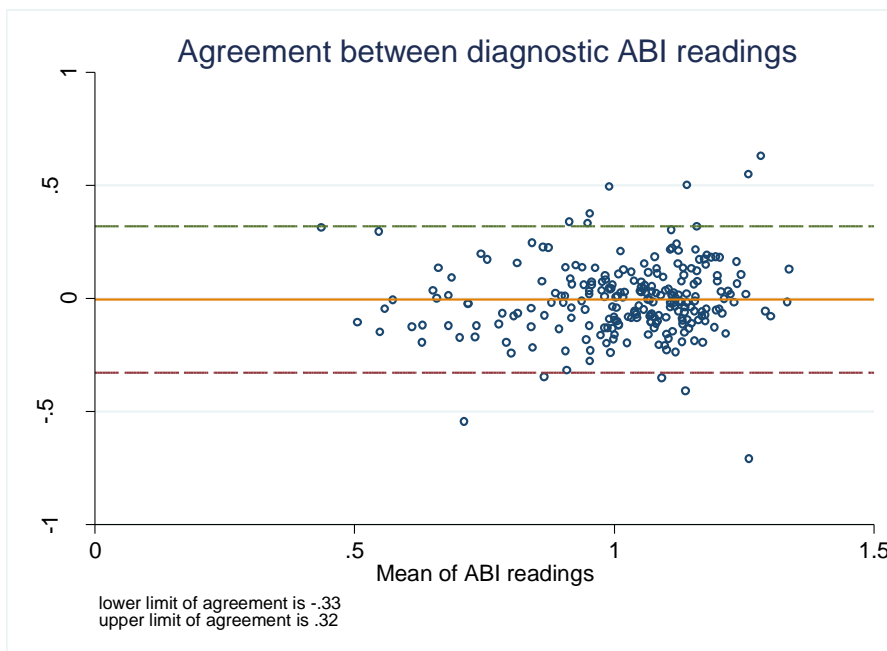


Figure 2. Agreement between pragmatic and conventional determination of ABI.

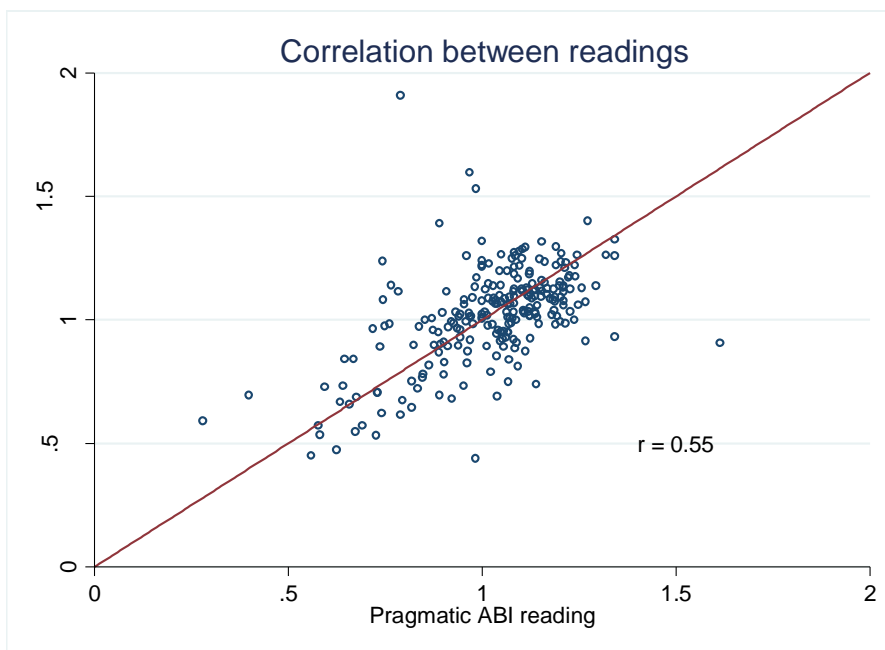


Figure 3. Correlation between pragmatic and conventional determination of ABI.

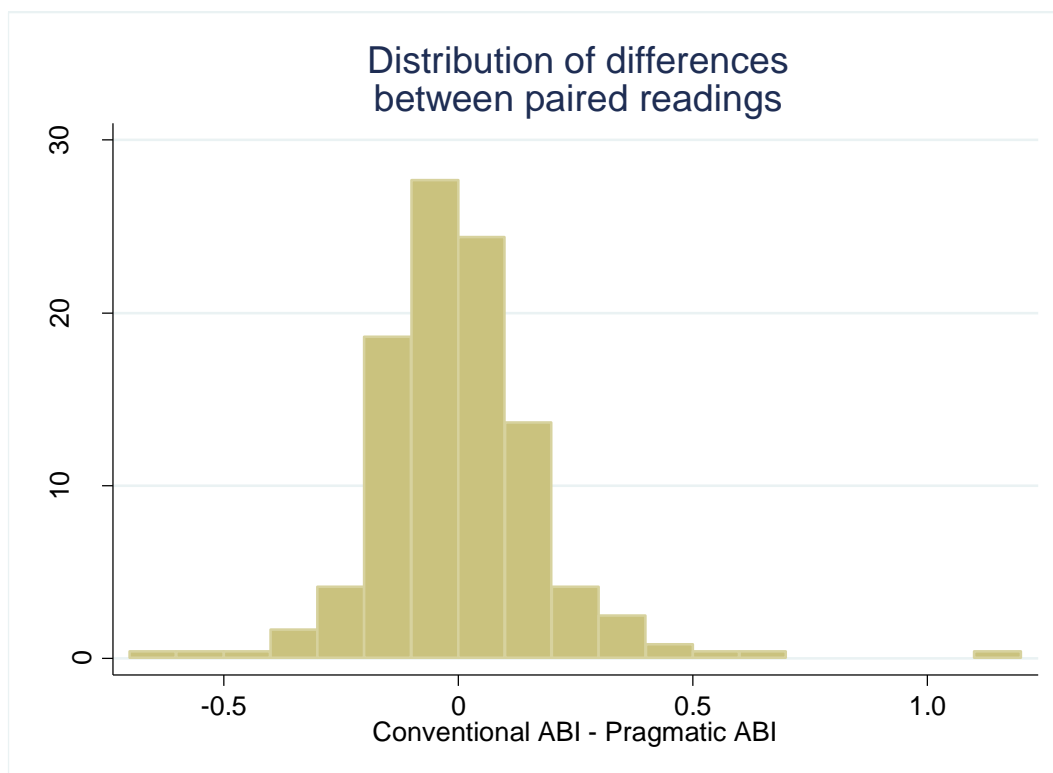


Figure 4. Distribution of the difference between the conventional ABI and the pragmatic ABI readings.

Review only

STARD checklist for reporting of studies of diagnostic accuracy
(version January 2003)

Section and Topic	Item #		On page #
TITLE/ABSTRACT/ KEYWORDS	1	Identify the article as a study of diagnostic accuracy (recommend MeSH heading 'sensitivity and specificity').	In keywords
INTRODUCTION	2	State the research questions or study aims, such as estimating diagnostic accuracy or comparing accuracy between tests or across participant groups.	7
METHODS			
<i>Participants</i>	3	The study population: The inclusion and exclusion criteria, setting and locations where data were collected.	7, 8
	4	Participant recruitment: Was recruitment based on presenting symptoms, results from previous tests, or the fact that the participants had received the index tests or the reference standard?	7
	5	Participant sampling: Was the study population a consecutive series of participants defined by the selection criteria in item 3 and 4? If not, specify how participants were further selected.	8
	6	Data collection: Was data collection planned before the index test and reference standard were performed (prospective study) or after (retrospective study)?	4
<i>Test methods</i>	7	The reference standard and its rationale.	
	8	Technical specifications of material and methods involved including how and when measurements were taken, and/or cite references for index tests and reference standard.	8
	9	Definition of and rationale for the units, cut-offs and/or categories of the results of the index tests and the reference standard.	8
	10	The number, training and expertise of the persons executing and reading the index tests and the reference standard.	8
	11	Whether or not the readers of the index tests and reference standard were blind (masked) to the results of the other test and describe any other clinical information available to the readers.	8
<i>Statistical methods</i>	12	Methods for calculating or comparing measures of diagnostic accuracy, and the statistical methods used to quantify uncertainty (e.g. 95% confidence intervals).	9-10
	13	Methods for calculating test reproducibility, if done.	N/A
RESULTS			
<i>Participants</i>	14	When study was performed, including beginning and end dates of recruitment.	8
	15	Clinical and demographic characteristics of the study population (at least information on age, gender, spectrum of presenting symptoms).	20
	16	The number of participants satisfying the criteria for inclusion who did or did not undergo the index tests and/or the reference standard; describe why participants failed to undergo either test (a flow diagram is strongly recommended).	16
<i>Test results</i>	17	Time-interval between the index tests and the reference standard, and any treatment administered in between.	8
	18	Distribution of severity of disease (define criteria) in those with the target condition; other diagnoses in participants without the target condition.	12
	19	A cross tabulation of the results of the index tests (including indeterminate and missing results) by the results of the reference standard; for continuous results, the distribution of the test results by the results of the reference standard.	21
	20	Any adverse events from performing the index tests or the reference standard.	11
<i>Estimates</i>	21	Estimates of diagnostic accuracy and measures of statistical uncertainty (e.g. 95% confidence intervals).	10-11
	22	How indeterminate results, missing data and outliers of the index tests were handled.	12
	23	Estimates of variability of diagnostic accuracy between subgroups of participants, readers or centers, if done.	N/A
	24	Estimates of test reproducibility, if done.	N/A
DISCUSSION	25	Discuss the clinical applicability of the study findings.	11



Ankle-Brachial Index determination and peripheral arterial disease diagnosis by an oscillometric blood pressure device in primary care: validation and diagnostic accuracy study

Journal:	<i>BMJ Open</i>
Manuscript ID:	bmjopen-2012-001689.R1
Article Type:	Research
Date Submitted by the Author:	24-Aug-2012
Complete List of Authors:	Nelson, Mark; University of Tasmania, Menzies Research Institute Quinn, Stephen; Flinders University, Flinders Clinical Effectiveness Winzenberg, Tamia; University of Tasmania, Menzies Research Institute Tasmania Howes, Faline; University of Tasmania, Menzies Research Institute Tasmania Shiel, Louise; Monash University, Epidemiology and Preventive Medicine Reid, Christopher; Monash University, Department of Epidemiology and Preventive Medicine (DEPM)
Primary Subject Heading:	General practice / Family practice
Secondary Subject Heading:	Cardiovascular medicine, Public health, Diagnostics
Keywords:	PRIMARY CARE, VASCULAR MEDICINE, PUBLIC HEALTH

SCHOLARONE™
Manuscripts

only

1
2 **Ankle-Brachial Index determination and peripheral arterial disease diagnosis by an**
3
4 **oscillometric blood pressure device in primary care: validation and diagnostic accuracy**
5
6 **study**
7
8
9

10
11
12 **Corresponding author:**
13

14 Mark R Nelson

15 Mark.Nelson@utas.edu.au

16 Private Bag 23

17 Hobart 7001, Tasmania, AUSTRALIA

18 Telephone: +61 3 6226 4734

19 Facsimile: +61 3 6226 4770

20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35 **Co-authors**
36

37 Stephen Quinn

38 Flinders University,

39 Adelaide, SA 5001, Australia

40 Tania M Winzenberg,

41 Menzies Research Institute Tasmania,

42 University of Tasmania,

1
2 Hobart, Tasmania 7001, Australia
3
4
5
6

7
8 Faline Howes
9

10 Menzies Research Institute Tasmania,
11

12
13 University of Tasmania,
14

15
16 Hobart, Tasmania 7001, Australia,
17
18
19

20
21
22 Louise Shiel,
23

24
25 Department of Epidemiology and Preventive Medicine,
26

27
28 Monash University,
29

30
31 Melbourne 3004, Australia,
32
33
34
35

36
37 Christopher M Reid
38

39
40 Department of Epidemiology and Preventive Medicine,
41

42
43 Monash University,
44

45
46 Melbourne 3004, Australia,
47
48
49

50
51 **Keywords**
52

53
54 Sensitivity and specificity, Ankle-Brachial Index, peripheral arterial disease, oscillometric
55
56 device, blood pressure.
57
58
59
60

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

Word count - excluding title page, abstract, references, figures and tables.

2505

For peer review only

ABSTRACT

Objectives

To determine the level of agreement between a 'conventional' Ankle-Brachial Index (ABI) measurement (using Doppler and mercury sphygmomanometer taken by a research nurse) and a 'pragmatic' ABI measure (using an oscillometric device taken by a practice nurse) in primary care. To ascertain the utility of a pragmatic ABI measure for the diagnosis of peripheral arterial disease (PAD) in primary care.

Design

Cross-sectional validation and diagnostic accuracy study. Descriptive analyses were used to investigate the agreement between the two procedures using the Bland and Altman method to determine whether the correlation between ABI readings varied systematically. Diagnostic accuracy was assessed via sensitivity, specificity, accuracy, likelihood ratios, positive and negative predictive values, with ABI readings dichotomised and Receiver Operating Curve analysis using both univariable and multivariable logistic regression.

Setting

Primary care in metropolitan and rural Victoria, Australia between October 2009 and November 2010.

Participants

250 persons with cardiovascular disease (CVD) or at high risk (3 or more risk factors) of CVD.

Results

Despite a strong association between the two method's measurements of ABI there was poor agreement with 95% of readings within ± 0.4 of the 0.9 ABI cut point. The multivariable C statistic of diagnosis of PAD was 0.89. Other diagnostic measures were sensitivity 62%,

1
2 specificity 92%, positive predictive value 67%, negative predictive value 90%, accuracy 85%,
3
4 positive likelihood ratio 7.3 and the negative likelihood ratio 0.42.
5
6

7 **Conclusions**

8
9
10 Oscillometric ABI measures by primary care nurses on a population with a 22% prevalence of
11
12 PAD lacked sufficient agreement with conventional measures to be recommended for routine
13
14 diagnosis of PAD. This pragmatic method may however be used as a screening tool high-risk
15
16 and overt CVD patients in primary care as it can reliably exclude the condition.
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

INTRODUCTION

Peripheral arterial disease (PAD) affects an estimated 27 million individuals in Europe and North America with 413,000 related hospital discharges per annum^{1 2}. These figures are likely to underestimate the true impact of PAD as those with the condition disproportionately suffer from other manifestations of CVD and are therefore likely to appear in coronary artery disease or stroke statistics. As a consequence there has been a call for better detection and management of the condition¹.

One of the simplest and most useful parameters to objectively assess lower extremity arterial perfusion, and thus diagnose PAD, is the Ankle-Brachial Index (ABI). This is the lower of the left and right ABI where each ABI is the ratio of the lower limb systolic blood pressure compared to the higher systolic brachial blood pressure recording. The ABI can be used to screen for haemodynamic significant PAD and helps to define its severity. Patients with objectively documented PAD have a four- to six-fold increase in cardiovascular mortality over healthy age-matched individuals³. PAD is a stronger risk marker for myocardial and stroke morbidity and mortality than those who have already had such an incident event^{4 5}. However, only 50% of people with PAD are symptomatic which is a significant issue in the detection of PAD².

Between 2007 and 2009 19,500 oscillometric devices were distributed by the High Blood Pressure Research Council of Australia to physicians, mostly general practitioners (GPs). We had previously demonstrated that these devices were likely to improve blood pressure management in primary care⁶. The current study, Ankle Brachial Index Determination by oscillometric method IN General practice (ABIDING), sought to expand the utility afforded by these machines in primary care. Previous work done in those attending a specialist vascular

laboratory in the US demonstrated that patients could have their ABI reliably ascertained by such devices compared to the conventional use of a Doppler ultrasound and mercury sphygmomanometer⁷. It was therefore opportune to investigate if such measures were pragmatic in primary care where the greatest opportunity exists to identify those with undiagnosed PAD. Such persons are at very high risk for subsequent adverse cardiovascular events that can be ameliorated through management of modifiable risk factors.

The primary aim of ABIDING was to establish if there was agreement between a pragmatic ABI (measured by a practice nurse using an oscillometric blood pressure device), and a conventional ABI (measured by a research nurse using mercury sphygmomanometer and Doppler devices). A secondary aim was to ascertain diagnostic accuracy of the pragmatic approach for ascertaining PAD.

METHODS

GPs and participants were recruited through the REACH Registry Victorian database. The international REACH Registry was a prospective, observational registry designed to provide long-term follow-up (36 months) of patients at high risk of atherothrombotic events. Globally 67,888 patients were involved in the REACH registry of whom 2,782 were recruited from 281 general practitioners around Australia^{8 9}. Practices were eligible for ABIDING if they had previously enrolled participants in the REACH registry and had a practice nurse willing to participate or were willing to appoint a *locum tenens* nurse. Eligibility criteria for REACH are published elsewhere but can be summarised as at entry (March to June 2004) aged 45+ years, had known CVD or at least three atherosclerosis risk factors, and were physically able to attend their usual general practice⁸.

Participant recruitment

1
2 All Melbourne (metropolitan) and Warrnambool (rural) Victorian study participants who had
3 consented to follow-up, who had been identified by their GPs as alive and for whom we had a
4 current address, were contacted by mail. If no reply was received from the participant within
5 four weeks, a second letter was sent and then a telephone call made. Participants were seen in
6 their usual GP's clinic between October 2009 and November 2010.
7
8
9
10
11
12

13 14 **Research and practice nurses**

15
16
17 Three experienced research nurses conducted the reference standard tests. They received
18 standardised training from a senior research nurse who was one of the operators. Practice nurses
19 were given training in situ by the research nurse and were observed by them. Because they
20 worked contemporaneously the research nurse was not blinded to the practice nurses results.
21
22
23
24
25
26

27 **'Conventional' and 'pragmatic' ABI estimation**

28
29
30 All participants were rested supine for five minutes before measurement. Doppler blood
31 pressure measurements (by research nurse) and automated oscillometric blood pressure
32 measurements (by practice nurse) were performed using cuffs that had bladders >80% of the
33 diameter of the arms and ankles measured.
34
35
36
37
38
39

40 Conventional measures involved Doppler blood pressure measurements in the lower limb made
41 with a Nicolet Vascular Doppler with a 5MHz probe. The cuff was inflated to 30mmHg above
42 systolic blood pressure and deflated slowly until a flow signal was detected over the dorsalis
43 pedis or posterior tibial arteries. Brachial artery systolic pressure was determined similarly but
44 utilising a stethoscope rather than a Doppler. The ABI for each lower extremity was calculated
45 as the pedal pressure divided by the higher of the two brachial pressures. PAD is defined as an
46 ABI <0.9 in either lower limb¹⁰. The mercury sphygmomanometer was calibrated by a certified
47 laboratory.
48
49
50
51
52
53
54
55
56
57
58
59
60

1
2 Research nurses were trained in the measurement of ABI and were certified prior to
3
4 commencement of the study. Practice nurses were simply observed and technique corrected if
5
6 required. Oscillometric measurements were made by the practice nurse on all limbs using a
7
8 standard automated blood pressure cuff system (OMRON HEM-907). This device is a validated
9
10 blood pressure measurement device^{11 12}. Oscillometric devices were new and therefore had
11
12 factory calibration. Participants also completed the Edinburgh Claudication questionnaire
13
14 (ECQ)¹³.
15
16
17
18

19 **Statistical Methods**

20
21 Descriptive analyses were used to investigate the agreement between the two procedures using
22
23 the Bland and Altman method to determine whether measurements could be used
24
25 interchangeably and if the correlation between ABI readings varied systematically¹⁴. Although
26
27 the variability in the differences appeared to be proportional to the mean, applying a log
28
29 transformation to the data did not substantially alter agreement and so raw scores are presented.
30
31 Correlations between the paired readings were also calculated. Sensitivity, specificity, positive
32
33 and negative predictive values, and accuracy with exact 95% confidence intervals are reported,
34
35 where ABI readings taken under both conditions were dichotomised at 0.9 (reference standard).
36
37 The diagnostic accuracy was evaluated using Receiver Operating Curve analysis and quantified
38
39 as the area under the curve (AUC or C statistic), as determined using both univariable and
40
41 multivariable logistic regression. In the multivariable model we adjusted for age, BMI, gender,
42
43 and smoking status (never, former, current). The calibration of this model was validated using
44
45 the Hosmer-Lemeshow statistic¹⁵. We examined likelihood ratios, the ratio of the expected test
46
47 results in participants with PAD to those participants without. All results are reported with 95%
48
49 confidence intervals. All analyses were conducted using Stata version 12.0.
50
51
52
53
54
55
56
57
58
59
60

Power calculations

Assuming a type 1 error of 5% ($\alpha = 0.05$) a total sample of 250 participants provided 80% power to detect systematic bias between the readings taken by the research and practice nurses if the mean difference was 0.0255¹⁴. Eight participants were excluded as 6 pragmatic and 2 conventional ABI readings were absent. In all other cases each patient had at least one conventional and pragmatic ABI reading (for the same leg). For a sample of 242 the difference that we could detect was 0.0257. We expected strong correlations between ABI readings taken using the different methods. Both calculations assumed a correlation between readings of 0.61 and standard deviations as reported in Benchimol *et al*¹⁶.

RESULTS

The flow chart of the study is shown in Figure 1. The characteristics of the ABIDING population are shown in Table 1. There was no difference between those excluded and included in the analysis for any trait that we measured. We also compared in Table 1 those diagnosed with PAD vs. not using conventional ABI. Those with PAD were older ($p=0.003$) and more likely to be female ($p=0.003$). Figure 2 shows there was poor agreement between pragmatic and conventional determination of ABI with 95% of readings within ± 0.4 . Figure 3 shows correlation between conventional and pragmatic ABI measurements, indicating a strong association between the two measurements, despite the poor agreement. The distribution of differences between the ABI measures is shown in Figure 4. These differences were regressed on all possible confounders measured in our study, in both univariable and multivariable models. There were no significant associations, suggesting that the differences were completely random.

1
2 A 2x2 table of dichotomised conventional and pragmatic measurements is shown (Table 2). We
3
4 examined the two groups comprising the 36 participants where the PAD classification differed.
5
6 There were no differences in any measured trait between those groups (data not shown). The
7
8 respective pragmatic method diagnostic performance, assuming the conventional method as gold
9
10 standard, was sensitivity 62% (95% CI 47-75%), specificity 92% (87-95%), positive predictive
11
12 value 67% (52%-80%), negative predictive value 90% (85-94%) and accuracy 85% (80%-89%).
13
14 The Likelihood ratio for a positive result (LR+) was 7.3 (95% CI 4.4-12.0) and Likelihood ratio
15
16 test for a negative result (LR-) 0.42 (0.30-0.59). Test performance for the asymptomatic
17
18 subgroup on ECQ (N = 183 PAD 18%) sensitivity 54% (95% CI 37-69%) specificity 93% (89-
19
20 97%) and symptomatic (N = 18 PAD 61%) sensitivity 9% (2-41%) specificity 57% (18-90%).
21
22 Area under the Receiver Operator Characteristic curves (AUC / C statistic) of pragmatic ABI
23
24 against the conventional ABI <0.9 and thus PAD was 0.87 (95% CI 0.82, 0.93). The AUC from
25
26 multivariable analysis (adjusting for age, gender, BMI and smoking status) for all analyses were
27
28 almost identical 89% (95% CI 84%-93%).
29
30
31
32
33
34

35
36 Based on the differences in Table 1 for those with PAD vs. not we conducted a post hoc
37
38 subgroup analyses on pragmatic vs. conventional ABI readings by gender, age (dichotomised as
39
40 young or old) and all pairwise combinations. The agreement between reading and diagnostic
41
42 criteria did not improve for any subgroup (data not shown). We also investigated (using
43
44 multivariable logistic regression) whether there was any evidence that disagreements were
45
46 systematic. There were no differences between in disagreement apart from current smokers
47
48 were more likely to produce readings that disagreed compared to non-smokers (p=0.025). A
49
50 subgroup analysis with current smokers removed did not alter the diagnostic criteria of the tests.
51
52
53
54 As could be expected in non-invasive testing there were no reported adverse events.
55
56
57
58
59
60

1
2 Sensitivity analyses for excluding upper ABI cut point of 1.4 (concern regarding possible
3 arterial incompressibility) did not affect the outcomes, and the range 0.85-0.95 gave 0.85
4 sensitivity 54% and specificity 95%, and 0.95 sensitivity 71% and specificity 86%.
5
6
7

8 9 **DISCUSSION**

10
11
12 ABIDING demonstrated that use of oscillometric devices by general practice nurses to
13 determine ABI and therefore the presence of PAD had high specificity (92%) and negative
14 predictive value (90%), good accuracy (84%) but modest sensitivity (62%) and positive
15 predictive value (67%). The modest sensitivity and the LR+ 7.3 indicate that this test has little
16 value for confirming the presence of PAD. On the other hand high specificity and negative
17 predictive value suggests that the test has some value in ruling out the disease (i.e. when the test
18 is negative). Looking at the symptomatic individuals as determined by ECQ showed that,
19 though the numbers were small, the pragmatic measure had a poor performance as a diagnostic
20 test in this high prevalence (61%) subgroup. Changing the cut point to improve sensitivity or
21 specificity simply compromised the other measure and therefore did not improve test
22 performance.
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38

39 These findings were in contrast to the experience in a specialist centre where their test
40 performance (both limbs in comparison to ABIDING lower of the 2 measures) was sensitivity
41 left/right leg 88/73% (62%), specificity 85/95% (92%), positive predictive value 65/88% (69%),
42 negative predictive value 96/88% (90%), LR+ left/right leg 5.9/14.6 (7.9) and LR- 0.14/0.28
43 (0.4)⁴. A good diagnostic test has a LR+ >10 and LR- < 0.1¹⁷. This difference in performance to
44 some extent may be accounted for by patient selection but is more likely due to operator
45 expertise. In the specialist centre, the mean age was 10 years younger and 53% were female
46 compared to only 22% in ABIDING. The respective prevalence of PAD was 32% and 22%. In
47
48
49
50
51
52
53
54
55
56
57
58
59
60

1
2 other studies reporting being conducted in primary care Mehlsen *et al* enrolled 1258 consecutive
3
4 general practice patients for an oscillometric determination of ABI, with those with an ABI <0.9
5
6 referred for a Doppler measure in a vascular unit¹⁸. Hence all ‘negatives’ including false
7
8 negatives did not have a gold standard measure and therefore this was not a true measure of test
9
10 performance in primary care. Nicholai *et al* and Aboyens had similar limitations^{19 20}. Verberk
11
12 and colleagues conducted a systematic review of automated oscillometric devices including a
13
14 subgroup analysis on devices developed for arm BP measurement²¹. Only one of the 18 studies
15
16 identified was conducted in primary care and that with an ABIgram and not a simple BP arm
17
18 device²². Although the investigators demonstrated its reliability, the use of this special piece of
19
20 equipment would seem to effect is acceptability as is the current situation. ABI is a valid and
21
22 reliable clinical measure although an indirect one. The true gold standard would be an
23
24 intravascular perfusion study. Both methods have been compared to the true gold standard in 85
25
26 patients with claudication undergoing angiography²³. The oscillometric method showed 97%
27
28 sensitivity, 89% specificity, 98% positive predictive value, and 86% negative predictive value.
29
30 The Doppler method showed 95% sensitivity, 56% specificity, 91% positive predictive value,
31
32 and 68% negative predictive value. This study suggests that the oscillometric method had
33
34 greater diagnostic accuracy but the test was performed by physicians not specifically trained to
35
36 use the Doppler probe. This said ABI is a practical tool and is superior to clinical examination
37
38 for identifying PAD²⁰. However screening whole populations is not always practical. ABI
39
40 ascertainment of PAD is most effective by identifying high risk patients as we have done in
41
42 ABIDING. By including high-risk and overt CVD patients we were confident that we should
43
44 get a distribution of ABI scores that included PAD diagnostic scores and the outcome of the
45
46 study supports this (22% had PAD by the conventional method).
47
48
49
50
51
52
53
54
55
56
57
58
59
60

1
2 If our method had been reliable it would have been readily implementable as Australian GPs
3
4 have ready access to oscillometric sphygmomanometers. More than 19,500 devices were
5
6 distributed on behalf of the High Blood Pressure Research Council of Australia, mostly to GPs,
7
8 over the years 2007 to 2009. Practice nurses were chosen rather than GPs as this approach is
9
10 also more likely to be implementable. A survey by Mohler *et al* of primary care clinicians
11
12 showed that most (88%) thought ABI to be feasible in that setting²³.
13
14

15 16 17 **Study limitations**

18
19 The intervention was kept as simple as possible by using practice nurses to do single measures
20
21 on a device they were familiar with but did not receive extensive further training on. While this
22
23 means that this is simple to introduce into clinical practice the practice nurse performance may
24
25 have been improved by more intense training and repeated limb measurements.
26
27

28 29 **CONCLUSION**

30
31
32 Oscillometric ABI measures by primary care nurses on a population with a 22% prevalence of
33
34 PAD lacked sufficient agreement with conventional measures to be recommended for routine
35
36 diagnosis of PAD. This pragmatic method may however be used as a screening tool in high-risk
37
38 primary care patients as it can reliably exclude the condition.
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

ACKNOWLEDGMENTS, COMPETING INTERESTS, FUNDING

Participating GPs were Drs Sam Arber, Dean Arnot, Ian Barratt, Irmgard Chia, John Chin, Eric Choo, Andrew Chow, Michael Conos, Simon Cooper, Peter Coulton, Adrian Dabscheck, Deborah Davidson, Antonio De Sousa, Hillary Donald, Peter Eizenberg, Peter Enten, Roger Fagan, Peter Ferguson, Doron Gaddie, Jon Garland, Andrew Gault, Ian Gill, Robert Gingold, Peter Goy, Judith Heale, Eva Herold, Chris Hogan, Suresh Jain, Brendan Kay, John Kirmos, Susan Kloot, Con Lahanis, Daniel Lajoie, Maryanne Lancaster, Conway Leung, Leon Lewi, John Manderson, Frank Marano, Damian Marinucci, Larry McGrath, Elizabeth McNaughton, Margaret McNiff, John Meaney, Paul Molloy, Clare Mooney, Paul O'Hanlon, James Olesen, Michael Page, John Pattison, Gary Pellizzari, Annamarie Perlesz, John Philpot, Leon Piterman, Jock Plenderleith, Mark Preston, Christopher Priest, Andy Psaradellis, Michael Quinn, Jacqueline Rounsevell, Joseph Sakowsky, Mahinda Samararatna, Victor Sammut, David Slonim, Stuart Smith, Stephen Stowe, Ian Sutherland, Edwin Turner, Peter Wexler, Andrew White, Martin Williams, Stephen Williams, Keith Wing Shing, Richard Wrennall, and Michael Ziccone.

We would also like to acknowledge the contribution of practice nurses, Dr Nyi Nyi Tun who assisted with the literature review, and research nurses Christine Mulvaney, Sue Loftus, and Anne Bruce.

This study had ethical approval from the Human Research Ethics Committee (Tasmania) Network (H0010410) and Monash University Standing Committee on Ethics in Research involving Humans (2009000860), and was registered with the Australian and New Zealand Clinical Trials Registry (ACTRN12609000744257). It was funded by the RACGP Research Foundation (Cardiovascular Research Grant) and the National Health and Medical Research

1
2 Council (Project grant 544935), and was supported by the Primary Healthcare Research,
3
4 Evaluation and Development scheme. Oscillometric devices were loaned by the High Blood
5
6 Pressure Research Council of Australia.
7

8
9 No competing interests declared.
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

For peer review only

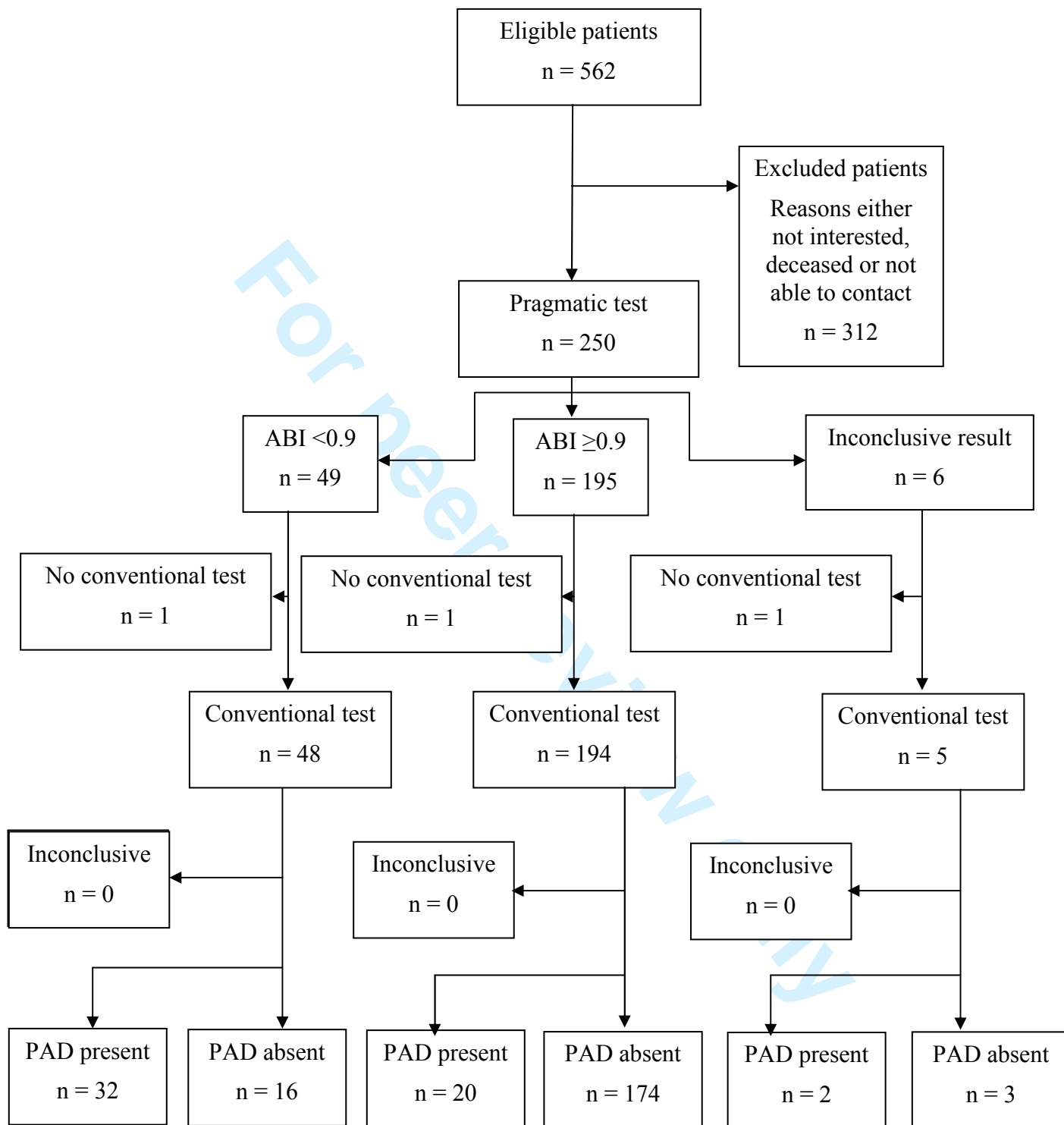


Figure 1. Flow diagram of a diagnostic accuracy in ABIDING as per STARD standard²⁵.

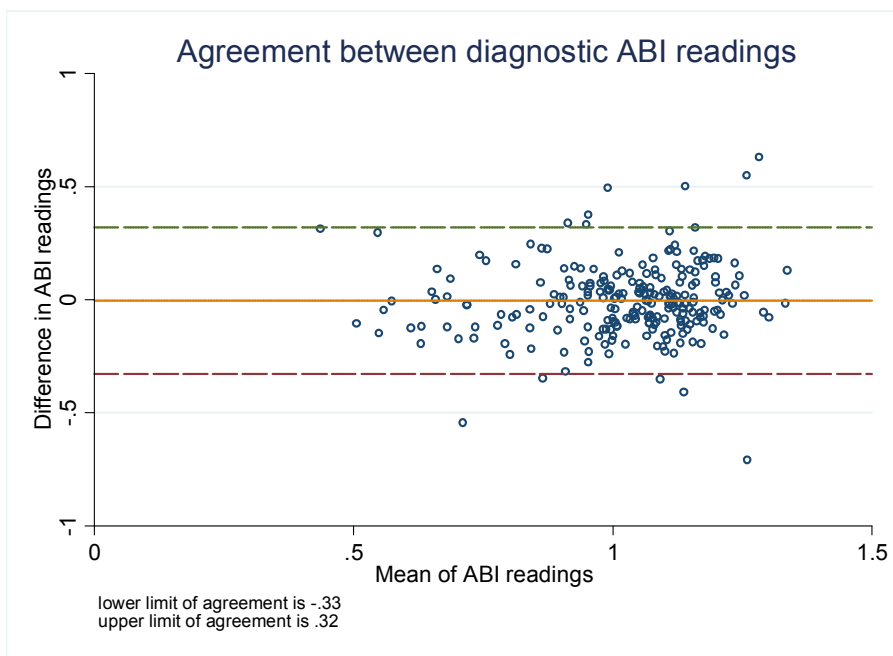


Figure 2. Agreement between pragmatic and conventional determination of ABI.

Peer review only

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

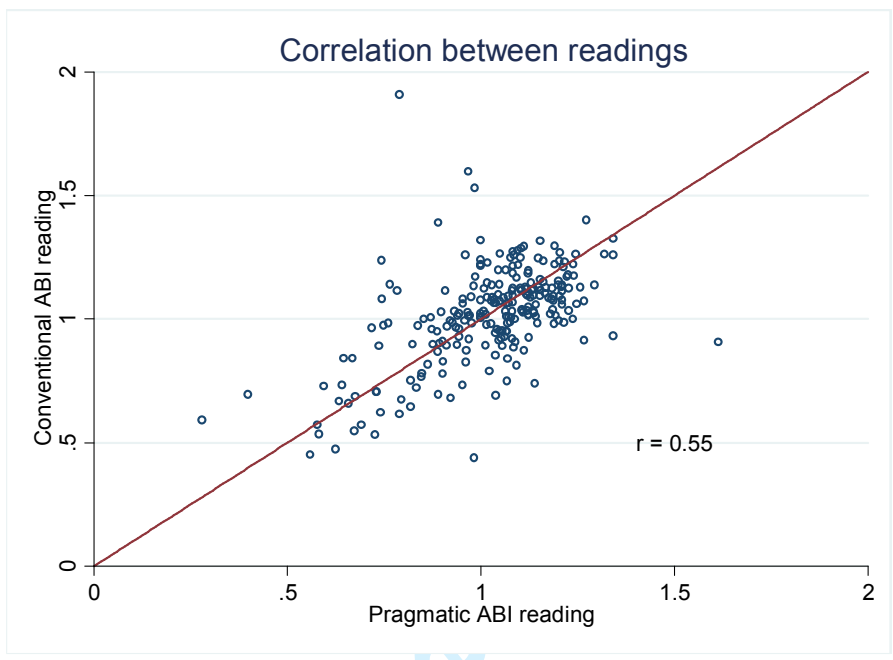


Figure 3. Correlation between pragmatic and conventional determination of ABI.

er review only

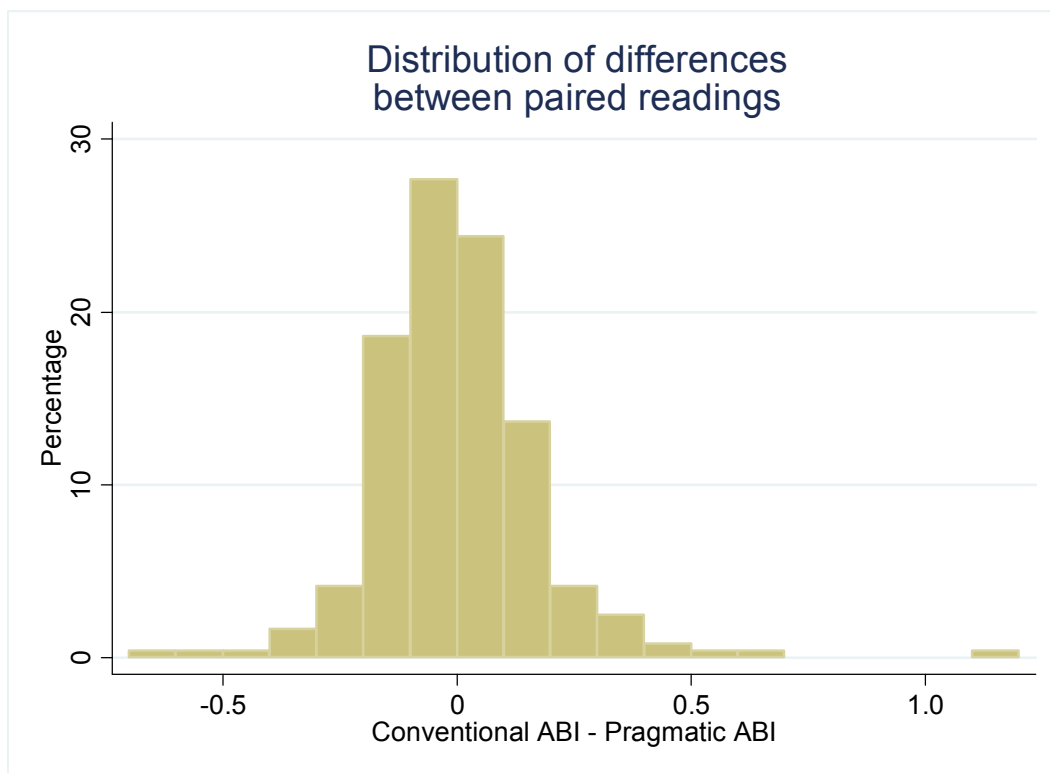


Figure 4. Distribution of the difference between the conventional ABI and the pragmatic ABI readings.

Review only

Variable	<i>Included</i>	<i>Conventional ABI ≥ 0.9</i>	<i>Conventional ABI < 0.9</i>	<i>P for difference</i>
N	242	192	52	
Age in years	71.2(7.4)	70.4(7.0)	73.9(8.3)	0.003
Male Sex (%)	167 (69.0)	140(73.7)	27(52.0)	0.003
SBP (mmHg)	141.5(18.9)	140.6(17.8)	144.5(22.4)	0.35
DBP (mmHg)	76.7(9.9)	77.0(9.8)	75.5(10.4)	0.55
BMI (kg/h ²)	27.5(4.4)	27.5(4.4)	27.2(4.4)	0.63
Waist	99.9(10.8)	100.1(10.4)	99.2(12.3)	0.60
Smoking status				0.31
Never	98(40.7)	81(42.9)	17(32.7)	
Former	131(54.4)	100(52.9)	31(59.6)	
Current	12(5.0)	8(4.2)	4(7.7)	

Table 1: Characteristics of participants by conventional PAD status expressed as a mean (standard deviation) or N (%) as appropriate.

	PAD +ve (conventional ABI <0.9)	PAD -ve (conventional ABI ≥0.9)	Total
Test +ve (pragmatic ABI <0.9)	32	16	48
Test -ve (pragmatic ABI ≥0.9)	20	174	194
Total	52	190	242

Table 2. 2x2 table of conventional and pragmatic ABI determinations.

REFERENCES

1. Belch JJF, Topol EJ, Agnelli G, Bertrand M, Califf RM, Clement DL, et al. Critical Issues in Peripheral Arterial Disease Detection and Management: A Call to Action. *Arch. Intern. Med.* 2003;163(8):884-92.
2. Golomb BA, Dang TT, Criqui MH. Peripheral Arterial Disease: Morbidity and Mortality Implications. *Circulation* 2006;114(7):688-99.
3. McDaniel M, Cronenwett J. Basic data related to the natural history of intermittent claudication. *Ann. Vasc. Surg.* 1989;3:273-77.
4. Steg PG, Bhatt DL, Wilson PWF, D'Agostino R, Sr, Ohman EM, Rother J, et al. One-Year Cardiovascular Event Rates in Outpatients With Atherothrombosis. *JAMA* 2007;297(11):1197-206.
5. McKenna M, Wolfson S, Kuller L. The ratio of ankle and arm arterial pressure as an independent predictor of mortality. *Atherosclerosis* 1991;87:119 -28.
6. Nelson MR, Quinn S, Bowers-Ingram L, Nelson JM, Winzenberg TM. Cluster Randomized Controlled Trial of Oscillometric Versus Manual Sphygmomanometer for Blood Pressure Management in Primary Care (CRAB). *Am. J. Hypertens.* 2009; doi:10.1038/ajh.2009.55.
7. Beckman JA, Higgins CO, Gerhard-Herman M. Automated Oscillometric Determination of the Ankle-Brachial Index Provides Accuracy Necessary for Office Practice. *Hypertension* 2006;47(1):35-38.
8. Bhatt DL, Steg PG, Ohman EM, Hirsch AT, Ikeda Y, Mas J-L, et al. International Prevalence, Recognition, and Treatment of Cardiovascular Risk Factors in Outpatients With Atherothrombosis 10.1001/jama.295.2.180. *JAMA* 2006;295(2):180-89.

- 1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
9. Reid CM, Nelson M, Chew D, Connor G, DeLoose F. Australians @ Risk: management of cardiovascular risk factors in the REACH registry. *Heart Lung and Circulation* 2007;doi10.1016/j.hlc.2007.07.009.
 10. Doobay AV, Anand SS. Sensitivity and Specificity of the Ankle–Brachial Index to Predict Future Cardiovascular Outcomes. *Arterioscler. Thromb. Vasc. Biol.* 2005;25(7):1463-69.
 11. Assaad MA, Topouchian JA, Darne BM, Asmar RG. Validation of the Omron HEM-907 device for blood pressure measurement. *Devices and Technology* 2002;7(4):237-41.
 12. White WB, Anwar YA. Evaluation of the overall efficacy of the Omron office digital blood pressure HEM-907 monitor in adults. *Devices and Technology* 2001;6(2):107-10.
 13. Leng GC, Fowkes F. The Edinburgh Claudication Questionnaire: an improved version of the WHO/ROSE Questionnaire for use in epidemiological surveys. *Journal of Clinical Epidemiology* 1992;45(10): 1101-1109.
 14. Bland JM Altman DG. Statistical methods for assessing agreement between two methods of clinical measurement. *Lancet* 1986;1(8476):307-10.
 15. Hosmer D, Lemeshow S. Applied Logistic Regression. 2nd ed. New York: John Wiley and Sons, Inc., 2000.
 16. Benchimol A, Bernard V, Pillois X, Hong NT, Sagardiluz P, Tortelier L *et al.* Validation of a new method of detecting peripheral artery disease by determination of ankle-brachial index using an automatic blood pressure device. *Angiology* 2004;55 (2):127-34.
 17. Šimundić A-M. Measures of diagnostic accuracy: basic definitions.
<http://www.ifcc.org/ifccfiles/docs/190404200805.pdf>

- 1
2 18. Mehlsen J, Wiinberg N, Bruce C. Oscillometric blood pressure measurement: a simple
3
4 method in screening for peripheral arterial disease. *Clinical Physiology and Functional Imaging*
5
6 2008;28(6):426-29.
7
8
9
10 19. Nicolai SPA, Kruidenier LM, Rouwet EV, Bartelink M-LEL, Prins MH, Teijink JA. Ankle
11
12 Brachial index in primary care: are we doing it right? *Br. J. Gen. Pract.* 2009;59(563):422-27.
13
14
15 20. Aboyans V, Lacroix P, Doucet S, Preux PM, Criqui MH, Laskar M. Diagnosis of peripheral
16
17 arterial disease in general practice: can the ankle-brachial index be measured either by pulse
18
19 palpation or an automatic blood pressure device?. *Int. J. Clin. Pract.* 2008;62(7):1001-07.
20
21
22 21. Verberk WJ, Kollias A, Stergiou GS. Automated oscillometric determination of the ankle-
23
24 brachial index: a systematic review and meta-analysis. *Hypertens. Res.* 2012.
25
26
27
28 22. Raines JK, Farrar J, Noicely K, Pena J, et al. Ankle/Brachial Index in the primary care
29
30 setting. *Vascular and Endovascular Surgery* 2004;38(2):131-36.
31
32
33 23. Vega J, Romaní S, Garcipérez FJ, Vicente L, Pacheco N, Zamorano J, et al. Peripheral
34
35 Arterial Disease: Efficacy of the Oscillometric Method. *Rev. Esp. Cardiol.* 2011;64(07):619-21.
36
37
38 24. Mohler ER 3rd, Treat-Jacobson D, Reilly MP, Cunningham KE, Miani M, Criqui MH, et al.
39
40 Utility and barriers to performance of the ankle-brachial index in primary care practice. *Vasc*
41
42 *Med* 2004;9(4):253-60.
43
44
45 25. Bossuyt PM, Reitsma JB, Bruns DE, Gatsonis CA, Glasziou PP, Irwig LM, et al. The
46
47 STARD Statement for Reporting Studies of Diagnostic Accuracy: Explanation and Elaboration.
48
49 *Clin Chem* 2003;49(1):7-18.
50
51
52
53
54
55
56
57
58
59
60

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

1

Ankle-Brachial Index determination and peripheral arterial disease diagnosis by an oscillometric blood pressure device in primary care: validation and diagnostic accuracy study

Corresponding author:

Mark R Nelson

Mark.Nelson@utas.edu.au

Private Bag 23

Hobart 7001, Tasmania, AUSTRALIA

Telephone: +61 3 6226 4734

Facsimile: +61 3 6226 4770

Co-authors

Stephen Quinn

Flinders University,

Adelaide, SA 5001, Australia

Tania M Winzenberg,

Menzies Research Institute Tasmania,

University of Tasmania,

1

1
2
3
4
5
6
7
8 Hobart, Tasmania 7001, Australia
9

10
11
12 Faline Howes
13

14
15 Menzies Research Institute Tasmania,
16

17
18 University of Tasmania,
19

20 Hobart, Tasmania 7001, Australia,
21
22

23
24 Louise Shiel,
25

26
27 Department of Epidemiology and Preventive Medicine,
28

29
30 Monash University,
31

32
33 Melbourne 3004, Australia,
34
35

36
37 Christopher M Reid
38

39
40 Department of Epidemiology and Preventive Medicine,
41

42
43 Monash University,
44

45
46 Melbourne 3004, Australia,
47

48
49 **Keywords**
50

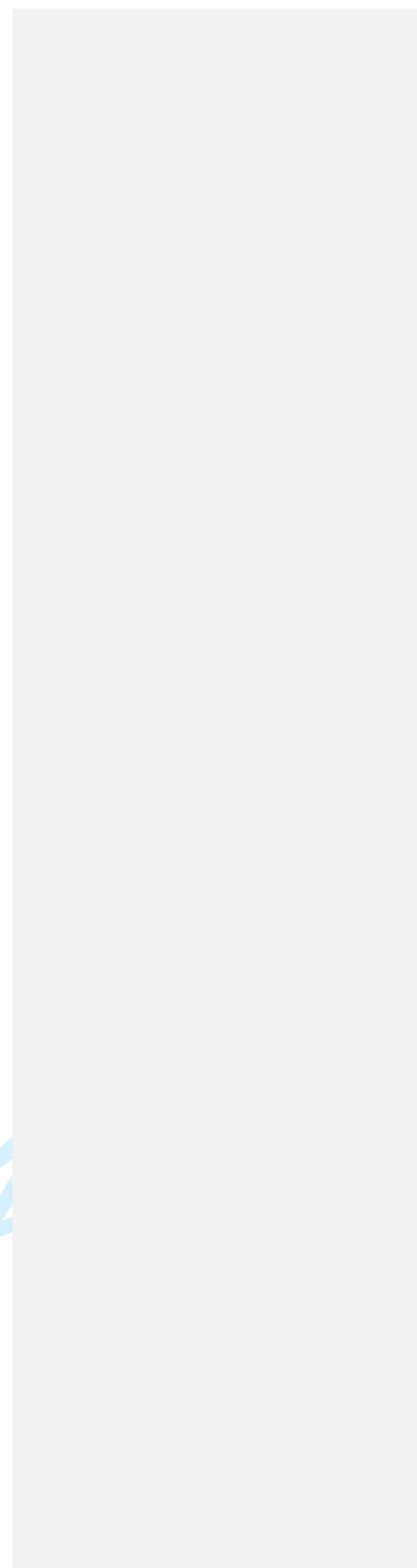
51 Sensitivity and specificity, Ankle-Brachial Index, peripheral arterial disease, oscillometric
52 device, blood pressure.
53
54
55
56
57
58
59
60

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

Word count - excluding title page, abstract, references, figures and tables.

23372505

For peer review only



ABSTRACT**Objectives**

To determine the level of agreement between a 'conventional' Ankle-Brachial Index (ABI) measurement (using Doppler and mercury sphygmomanometer taken by a research nurse) and a 'pragmatic' ABI measure (using an oscillometric device taken by a practice nurse) in primary care. To ascertain the utility of a pragmatic ABI measure for the diagnosis of peripheral arterial disease (PAD) in primary care.

Design

Cross-sectional validation and diagnostic accuracy study. Descriptive analyses were used to investigate the agreement between the two procedures using the Bland and Altman method to determine whether the correlation between ABI readings varied systematically. Diagnostic accuracy was assessed via sensitivity, specificity, accuracy, likelihood ratios, positive and negative predictive values, with ABI readings dichotomised and Receiver Operating Curve analysis using both univariable and multivariable logistic regression.

Setting

Primary care in metropolitan and rural Victoria, Australia between October 2009 and November 2010.

Participants

250 persons with cardiovascular disease (CVD) or at high risk (3 or more risk factors) of CVD.

Results

Despite a strong association between the two method's measurements of ABI there was poor agreement with 95% of readings within ± 0.4 of the 0.9 ABI cut point. The multivariable C statistic of diagnosis of PAD was 0.89. Other diagnostic measures were sensitivity 62%,

1
2
3
4
5
6
7
8 specificity 92%, positive predictive value 67%, negative predictive value 90%, accuracy 85%,
9 positive likelihood ratio 7.3 and the negative likelihood ratio 0.42.
10

11 **Conclusions**

12
13
14 Oscillometric ABI measures by primary care nurses on a population with a 22% prevalence of
15
16 PAD lacked sufficient agreement with conventional measures to be recommended for routine
17
18 diagnosis of PAD. This pragmatic method may however be used as a screening tool high-risk
19 and overt CVD patients in primary care as it can reliably exclude the condition.~~determination of~~
20 ~~ABI. Their diagnostic performance suggests that a pragmatic ABI lacked sufficient sensitivity to~~
21 ~~diagnose PAD, but can reliably exclude it.~~
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

INTRODUCTION

Peripheral arterial disease (PAD) affects an estimated 27 million individuals in Europe and North America with 413,000 related hospital discharges per annum^{1 2}. These figures are likely to underestimate the true impact of PAD as those with the condition disproportionately suffer from other manifestations of CVD and are therefore likely to appear in coronary artery disease or stroke statistics. As a consequence there has been a call for better detection and management of the condition¹.

One of the simplest and most useful parameters to objectively assess lower extremity arterial perfusion, and thus diagnose PAD, is the Ankle-Brachial Index (ABI). This is the lower of the left and right ABI where each ABI is the ratio of the lower limb systolic blood pressure ~~is~~ compared to the higher systolic brachial blood pressure recording. The ABI can be used to screen for ~~haemodynamically~~haemodynamic significant PAD and helps to define its severity. Patients with objectively documented PAD have a four- to six-fold increase in cardiovascular mortality over healthy age-matched individuals³. PAD is a stronger risk marker for myocardial and stroke morbidity and mortality than those who have already had such an incident event^{4 5}. However, only 50% of people with PAD are symptomatic which is a significant issue in the detection of PAD².

Between 2007 and 2009 19,500 oscillometric devices were distributed by the High Blood Pressure Research Council of Australia to physicians, mostly general practitioners (GPs). We had previously demonstrated that these devices were likely to improve blood pressure management in primary care⁶. The current study, Ankle Brachial Index Determination by oscillometric method IN General practice (ABIDING), sought to expand the utility afforded by these machines in primary care. Previous work done in those attending a specialist vascular

laboratory in the US demonstrated that patients could have their ABI reliably ascertained by such devices compared to the conventional use of a Doppler ultrasound and mercury sphygmomanometer⁴⁷. It was therefore opportune to investigate if such measures were pragmatic in primary care where the greatest opportunity exists to identify those with undiagnosed PAD. Such persons are at very high risk for subsequent adverse cardiovascular events that can be ameliorated through management of modifiable risk factors.

The primary aim of ABIDING was to establish if there was agreement between a pragmatic ABI (measured by a practice nurse using an oscillometric blood pressure device), and a conventional ABI (measured by a research nurse using mercury sphygmomanometer and Doppler devices). A secondary aim was to ascertain diagnostic accuracy of the pragmatic approach for ascertaining PAD.

METHODS

~~General practitioners~~ (GPs) and participants were recruited through the REACH Registry Victorian database. The international REACH Registry was a prospective, observational registry designed to provide long-term follow-up (36 months) of patients at high risk of atherothrombotic events. Globally 67,888 patients were involved in the REACH registry of whom 2,782 were recruited from 281 general practitioners around Australia^{8 9}. Practices were eligible for ABIDING if they had previously enrolled participants in the REACH registry and had a practice nurse willing to participate or were willing to appoint a *locum tenens* nurse. Eligibility criteria for REACH are published elsewhere but can be summarised as at entry (March to June 2004) aged 45+ years, had known CVD or at least three atherosclerosis risk factors, and were physically able to attend their usual general practice⁵⁸.

Participant recruitment

1
2
3
4
5
6
7
8 All Melbourne (metropolitan) and Warrnambool (rural) Victorian study participants who had
9 consented to follow-up, who had been identified by their GPs as alive and for whom we had a
10 current address, were contacted by mail. If no reply was received from the participant within
11 four weeks, a second letter was sent and then a telephone call made. Participants were seen in
12 their usual GP's clinic between October 2009 and November 2010.
13

14 15 16 17 18 **Research and practice nurses**

19
20 Three experienced research nurses conducted the reference standard tests. They received
21 standardised training from a senior research nurse who was one of the operators. Practice nurses
22 were given training in situ by the research nurse and were observed by them. Because they
23 worked contemporaneously the research nurse was not blinded to the practice nurses results.
24
25
26

27 28 **'Conventional' and 'pragmatic' ABI estimation**

29
30 All participants were rested supine for five minutes before measurement. Doppler blood
31 pressure measurements (by research nurse) and automated oscillometric blood pressure
32 measurements (by practice nurse) were performed using cuffs that had bladders >80% of the
33 diameter of the arms and ankles measured.
34
35

36
37
38 Conventional measures involved Doppler blood pressure measurements in the lower limb made
39 with a Nicolet Vascular Doppler with a 5MHz probe. The cuff was inflated to 30mmHg above
40 systolic blood pressure and deflated slowly until a flow signal was detected over the dorsalis
41 pedis or posterior tibial arteries. Brachial artery systolic pressure was determined similarly but
42 utilising a stethoscope rather than a Doppler. The ABI for each lower extremity was calculated
43 as the pedal pressure divided by the higher of the two brachial pressures. PAD is defined as an
44 ABI <0.9 in either lower limb⁶¹⁰. The mercury sphygmomanometer was calibrated by a certified
45 laboratory.
46
47
48
49
50
51
52

1
2
3
4
5
6
7
8 Research nurses were trained in the measurement of ABI and were certified prior to
9 commencement of the study. Practice nurses were simply observed and technique corrected if
10 required. Oscillometric measurements were made by the practice nurse on all limbs using a
11 standard automated blood pressure cuff system (OMRON HEM-907). This device is a validated
12 blood pressure measurement device^{11 12}. Oscillometric devices were new and therefore had
13 factory calibration. Participants also completed the Edinburgh Claudication questionnaire
14 (ECQ)¹³.

21 **Statistical Methods**

22
23
24 Descriptive analyses were used to investigate the agreement between the two procedures using
25 the Bland and Altman method to determine whether measurements could be used
26 interchangeably and if the correlation between ABI readings varied systematically⁷¹⁴. Although
27 the variability in the differences appeared to be proportional to the mean, applying a log
28 transformation to the data did not substantially alter agreement and so raw scores are presented.
29 Correlations between the paired readings were also calculated. Sensitivity, specificity, positive
30 and negative predictive values, and accuracy with exact 95% confidence intervals are reported,
31 where ABI readings taken under both conditions were dichotomised at 0.9 (reference standard).
32 The diagnostic accuracy was evaluated using Receiver Operating Curve analysis and quantified
33 as the area under the curve (AUC or C statistic), as determined using both univariable and
34 multivariable logistic regression. In the multivariable model we adjusted for age, BMI, gender,
35 and smoking status (never, former, current). The calibration of this model was validated using
36 the Hosmer-Lemeshow statistic⁸¹⁵. We examined likelihood ratios, the ratio of the expected test
37 results in participants with PAD to those participants without. All results are reported with 95%
38 confidence intervals. All analyses were conducted using Stata version 12.0.
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

Power calculations

Assuming a type 1 error of 5% ($\alpha = 0.05$) a total sample of 250 participants provided 80% power to detect systematic bias between the readings taken by the research and practice nurses if the mean difference was 0.0255⁷¹⁴. Eight participants were excluded as 6 pragmatic and 2 conventional ABI readings were absent. In all other cases each patient had at least one conventional and pragmatic ABI reading (for the same leg). For a sample of 242 the difference that we could detect was 0.0257. We expected strong correlations between ABI readings taken using the different methods. Both calculations assumed a correlation between readings of 0.61 and standard deviations as reported in Benchimol *et al*⁹¹⁶.

RESULTS

The flow chart of the study is shown in Figure 1. The characteristics of the ABIDING population are shown in Table 1. There was no difference between those excluded and included in the analysis for any trait that we measured. We also compared in Table 1 those diagnosed with PAD vs. not using conventional ABI. Those with PAD were older ($p=0.003$) and more likely to be female ($p=0.003$). Figure 2 shows there was poor agreement between pragmatic and conventional determination of ABI with 95% of readings within ± 0.4 . Figure 3 shows correlation between conventional and pragmatic ABI measurements, indicating a strong association between the two measurements, despite the poor agreement. The distribution of differences between the ABI measures is shown in Figure 4. These differences were regressed on all possible confounders measured in our study, in both univariable and multivariable models. There were no significant associations, suggesting that the differences were completely random.

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

A 2x2 table of dichotomised conventional and pragmatic measurements is shown (Table 2). We examined the two groups comprising the 36 participants where the PAD classification differed. There were no differences in any measured trait between those groups (data not shown). The respective pragmatic method diagnostic performance, assuming the conventional method as gold standard, was sensitivity 62% (95% CI 47-75%), specificity 92% (87-95%), positive predictive value 67% (52%-80%), negative predictive value 90% (85-94%) and accuracy 85% (80%-89%). The Likelihood ratio for a positive result (LR+) was 7.3 (95% CI 4.4-12.0) and Likelihood ratio test for a negative result (LR-) 0.42 (0.30-0.59). Test performance for the asymptomatic subgroup on ECQ (N = 183 PAD 18%) sensitivity 54% (95% CI 37-69%) specificity 93% (89-97%) and symptomatic (N = 18 PAD 61%) sensitivity 9% (2-41%) specificity 57% (18-90%).

Area under the Receiver Operator Characteristic curves (AUC / C statistic) of pragmatic ABI against the conventional ABI <0.9 and thus PAD was 0.87 (95% CI 0.82, 0.93). The AUC from multivariable analysis (adjusting for age, gender, BMI and smoking status) for all analyses were almost identical 89% (95% CI 84%-93%).

Based on the differences in Table 1 for those with PAD vs. not we conducted a post hoc subgroup analyses on pragmatic vs. conventional ABI readings by gender, age (dichotomised as young or old) and all pairwise combinations. The agreement between reading and diagnostic criteria did not improve for any subgroup (data not shown). We also investigated (using multivariable logistic regression) whether there was any evidence that disagreements were systematic. There were no differences between in disagreement apart from current smokers were more likely to produce readings that disagreed compared to non-smokers (p=0.025). A subgroup analysis with current smokers removed did not alter the diagnostic criteria of the tests.

As could be expected in non-invasive testing there were no reported adverse events.

Sensitivity analyses for excluding upper ABI cut point of 1.4 (concern regarding possible arterial incompressibility) did not affect the outcomes, and the range 0.85-0.95 gave 0.85 sensitivity 54% and specificity 95%, and 0.95 sensitivity 71% and specificity 86%.

DISCUSSION

ABIDING demonstrated that use of oscillometric devices by general practice nurses to determine ABI and therefore the presence of PAD had high specificity (92%) and negative predictive value (90%), good accuracy (84%) but modest sensitivity (62%) and positive predictive value (67%). The modest sensitivity and the LR+ 7.3 indicate that this test has little value for confirming the presence of PAD. On the other hand high specificity and negative predictive value suggests that the test has some value in ruling out the disease (i.e. when the test is negative). Looking at the symptomatic individuals as determined by ECQ showed that, though the numbers were small, the pragmatic measure had a poor performance as a diagnostic test in this high prevalence (61%) subgroup. Changing the cut point to improve sensitivity or specificity simply compromised the other measure and therefore did not improve test performance.

~~This is~~ These findings were in contrast to the experience in a specialist centre where their test performance (both limbs in comparison to ABIDING lower of the 2 measures) was sensitivity left/right leg 88/73% (62%), specificity 85/95% (92%), positive predictive value 65/88% (69%), negative predictive value 96/88% (90%), LR+ left/right leg 5.9/14.6 (7.9) and LR- 0.14/0.28 (0.4)⁴. A good diagnostic test has a LR+ >10 and LR- < 0.1⁴⁰¹⁷. This difference in performance to some extent may be accounted for by patient selection but is more likely due to operator expertise. In the specialist centre, the mean age was 10 years younger and 53% were female compared to only 22% in ABIDING. The respective prevalence of PAD was 32% and 22%. In

1
2
3
4
5
6
7
8 other studies reporting being conducted in primary care Mehlsen *et al* enrolled 1258 consecutive
9 general practice patients for an oscillometric determination of ABI, with those with an ABI <0.9
10 referred for a Doppler measure in a vascular unit¹⁸. Hence all 'negatives' including false
11 negatives did not have a gold standard measure and therefore this was not a true measure of test
12 performance in primary care. Nicholai *et al* and Aboyens had similar limitations^{19 20}. Verberk
13 and colleagues conducted a systematic review of automated oscillometric devices including a
14 subgroup analysis on devices developed for arm BP measurement²¹. Only one of the 18 studies
15 identified was conducted in primary care and that with an ABIgram and not a simple BP arm
16 device²². Although the investigators demonstrated its reliability, the use of this special piece of
17 equipment would seem to effect is acceptability as is the current situation.

Formatted: Font: Italic

Formatted: Font: Italic

18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
ABI is a valid and reliable clinical measure although an indirect one. The true gold standard
would be an intravascular perfusion study. Both methods have been compared to the true gold
standard in 85 patients with claudication undergoing angiography²³. The oscillometric method
showed 97% sensitivity, 89% specificity, 98% positive predictive value, and 86% negative
predictive value. The Doppler method showed 95% sensitivity, 56% specificity, 91% positive
predictive value, and 68% negative predictive value. This study suggests that the oscillometric
method had greater diagnostic accuracy but the test was performed by physicians not
specifically trained to use the Doppler probe. This said ABI is a ~~useful-practical~~ tool and is
superior to clinical examination for identifying PAD¹⁺²⁰. However screening whole populations
is not always practical. ABI ascertainment of PAD is most effective by identifying high risk
patients as we have done in ABIDING. By including high-risk and overt CVD patients we were
confident that we should get a distribution of ABI scores that included PAD diagnostic scores
and the outcome of the ~~trial-study~~ supports this (22% had PAD by the conventional method).
~~Doubini *et al* found age alone (70+) a useful predictor as 12.5% of the screened population had~~

~~PAD vs. only 2.5% of 50-69 years with at least 1 CVD risk factor but no established CVD (diabetes, dyslipidaemia, hypertension, or smoking)¹². Bendermacher *et al* developed a clinical prediction model giving risk factor points per factor (age: 1 point per 5 years starting at 55 years; ever smoked: 2 points; currently smoking: 7 points; and hypertension: 3 points), showed a proportional increase of the PAD prevalence with each increasing risk profile (range: 7.0-40.6%)¹³. The overall prevalence of PAD was 18%. They found with their PREVALENT clinical prediction model (based on CVD risk factors), the GP was able to identify a high risk population in which measurement of ABI was useful.~~

If our method had been reliable it would have been readily implementable as Australian GPs have ready access to oscillometric sphygmomanometers. More than 19,500 devices were distributed on behalf of the High Blood Pressure Research Council of Australia, mostly to GPs, over the years 2007 to 2009. Practice nurses were chosen rather than GPs as this approach is also more likely to be implementable. A survey by Mohler *et al* of primary care clinicians showed that most (88%) thought ABI to be feasible in that setting^{14,23}. ~~However, validation studies have largely been conducted in specialist clinics in a variety of study populations rather than in the primary care setting where most of the medical contact is likely to occur^{15-20,23,9,15-18}. The one study done in the primary care setting used an ABIgram. Although the investigators demonstrated its reliability, the use of this special piece of equipment would seem to effect is acceptability as is the current situation.~~

Study limitations

The intervention was kept as simple as possible by using practice nurses to do single measures on a device they were familiar with but did not receive extensive further training on. While this

Formatted: Superscript

Field Code Changed

Field Code Changed

1
2
3
4
5
6
7
8 means that this is simple to introduce into clinical practice the practice nurse performance may
9 have been improved by more intense training and repeated limb measurements.
10

11 CONCLUSION

12
13
14 Oscillometric ABI measures by primary care nurses on a population with a 22% prevalence of
15 PAD lacked sufficient agreement with conventional measures to be recommended for routine
16
17 ~~determination diagnosis~~ of ~~ABIPAD~~. This pragmatic method may however be used as a
18
19 screening tool in high-risk primary care patients ~~primary care but its diagnostic performance~~
20
21 ~~does not provide evidence sufficient for it to be used to diagnose PAD~~ as it can reliably exclude
22
23 the condition.
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

ACKNOWLEDGMENTS, COMPETING INTERESTS, FUNDING

Participating GPs were Drs Sam Arber, Dean Arnot, Ian Barratt, Irmgard Chia, John Chin, Eric Choo, Andrew Chow, Michael Conos, Simon Cooper, Peter Coulton, Adrian Dabscheck, Deborah Davidson, Antonio De Sousa, Hillary Donald, Peter Eizenberg, Peter Enten, Roger Fagan, Peter Ferguson, Doron Gaddie, Jon Garland, Andrew Gault, Ian Gill, Robert Gingold, Peter Goy, Judith Heale, Eva Herold, Chris Hogan, Suresh Jain, Brendan Kay, John Kirmos, Susan Kloot, Con Lahanis, Daniel Lajoie, Maryanne Lancaster, Conway Leung, Leon Lewi, John Manderson, Frank Marano, Damian Marinucci, Larry McGrath, Elizabeth McNaughton, Margaret McNiff, John Meaney, Paul Molloy, Clare Mooney, Paul O'Hanlon, James Olesen, Michael Page, John Pattison, Gary Pellizzari, Annamarie Perlesz, John Philpot, Leon Piterman, Jock Plenderleith, Mark Preston, Christopher Priest, Andy Psaradellis, Michael Quinn, Jacqueline Rounsevell, Joseph Sakowsky, Mahinda Samararatna, Victor Sammut, David Slonim, Stuart Smith, Stephen Stowe, Ian Sutherland, Edwin Turner, Peter Wexler, Andrew White, Martin Williams, Stephen Williams, Keith Wing Shing, Richard Wrennall, and Michael Ziccone.

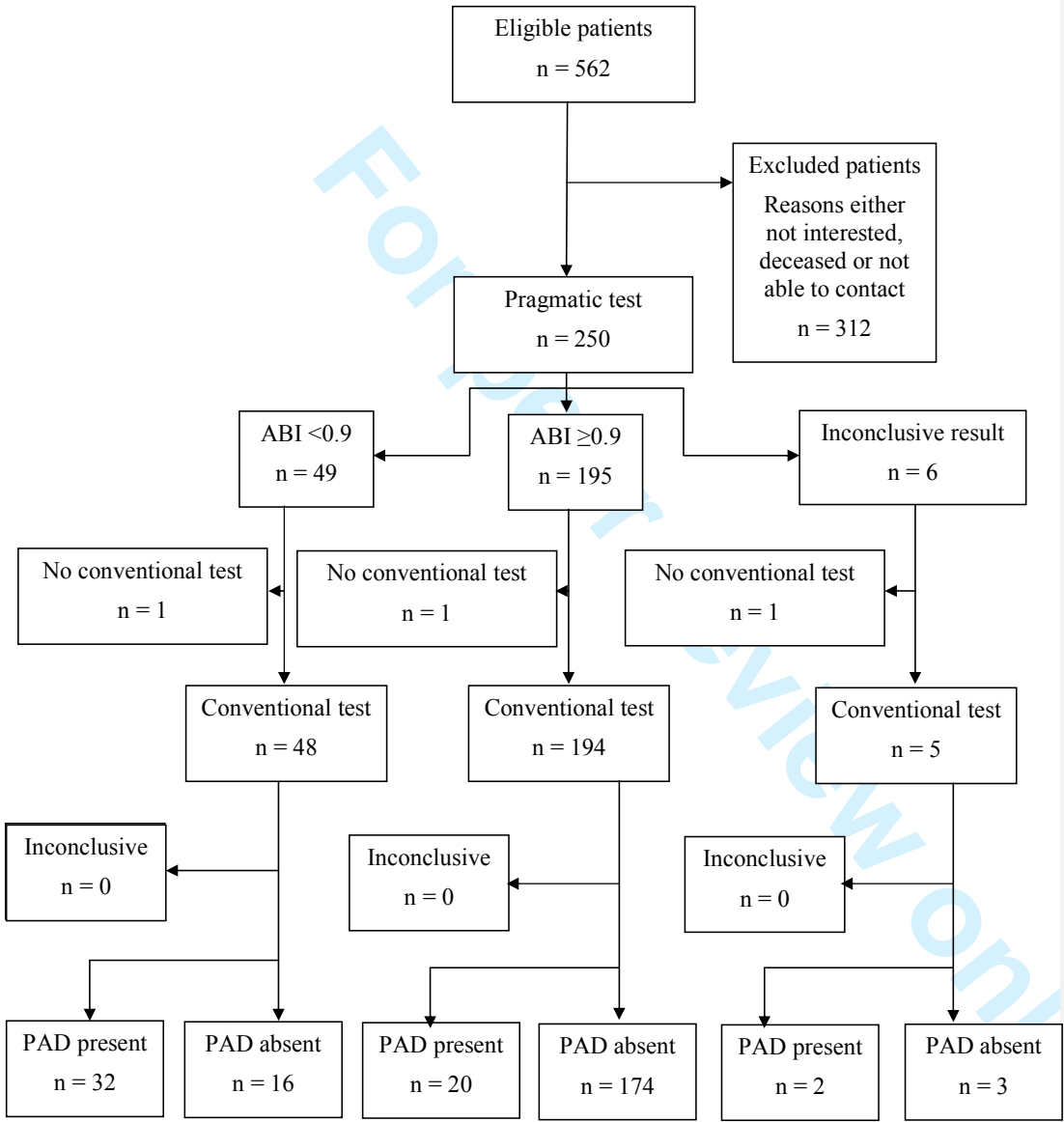
We would also like to acknowledge the contribution of practice nurses, Dr Nyi Nyi Tun who assisted with the literature review, and research nurses Christine Mulvaney, Sue Loftus, and Anne Bruce.

This study had ethical approval from the Human Research Ethics Committee (Tasmania) Network (H0010410) and Monash University Standing Committee on Ethics in Research involving Humans (2009000860), and was registered with the Australian and New Zealand Clinical Trials Registry (ACTRN12609000744257). It was funded by the RACGP Research Foundation (Cardiovascular Research Grant) and the National Health and Medical Research

1
2
3
4
5
6
7
8 Council (Project grant 544935), and was supported by the Primary Healthcare Research,
9 Evaluation and Development scheme. Oscillometric devices were loaned by the High Blood
10 Pressure Research Council of Australia.
11
12

13
14 No competing interests declared.
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60



1
2
3
4
5
6
7
8 Figure 1. Flow diagram of a diagnostic accuracy in ABIDING as per STARD
9

10 standard ²⁰²⁵ ~~ENREF 20~~
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

For peer review only

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

20

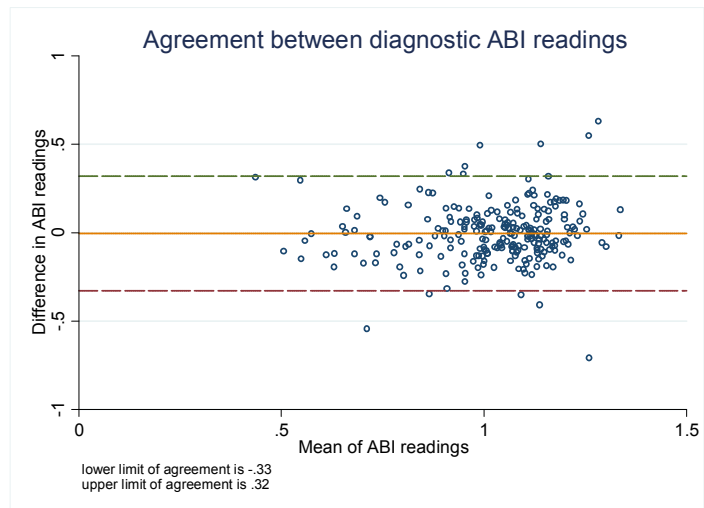
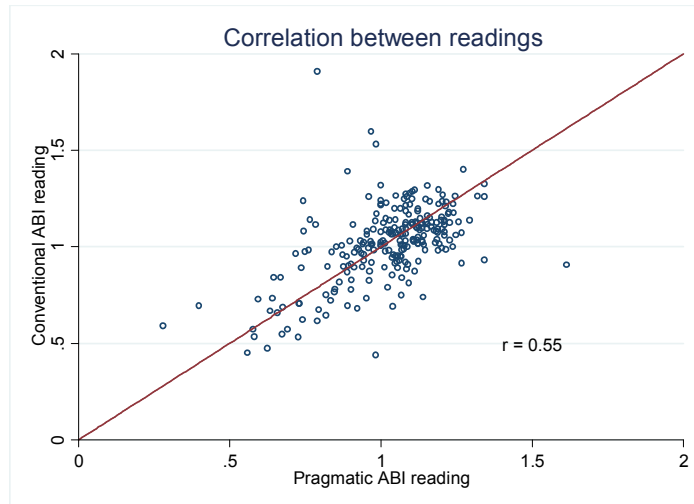


Figure 2. Agreement between pragmatic and conventional determination of ABI.

20



28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

Figure 3. Correlation between pragmatic and conventional determination of ABI.

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

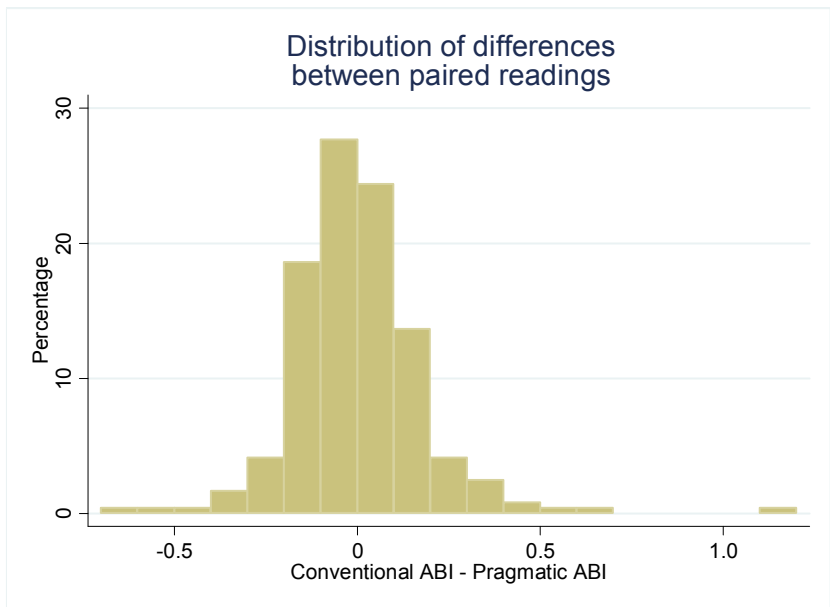


Figure 4. Distribution of the difference between the conventional ABI and the pragmatic ABI readings.

Review only

23

Variable	Included	Conventional <i>ABI</i> ≥ 0.9	Conventional <i>ABI</i> < 0.9	<i>P</i> for difference
N	242	192	52	
Age in years	71.2(7.4)	70.4(7.0)	73.9(8.3)	0.003
Male Sex (%)	167 (69.0)	140(73.7)	27(52.0)	0.003
SBP (mmHg)	141.5(18.9)	140.6(17.8)	144.5(22.4)	0.35
DBP (mmHg)	76.7(9.9)	77.0(9.8)	75.5(10.4)	0.55
BMI (kg/h ²)	27.5(4.4)	27.5(4.4)	27.2(4.4)	0.63
Waist	99.9(10.8)	100.1(10.4)	99.2(12.3)	0.60
Smoking status				0.31
Never	98(40.7)	81(42.9)	17(32.7)	
Former	131(54.4)	100(52.9)	31(59.6)	
Current	12(5.0)	8(4.2)	4(7.7)	

Formatted Table

Formatted: Left: 1.25", Top: 0.79", Bottom: 0.98", Width: 8.5", Height: 11"

Table 1: Characteristics of participants ~~(included in and excluded from the analysis) and~~ by conventional PAD status expressed as a mean (standard deviation) or N (%) as appropriate.

23

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

	PAD +ve (conventional ABI <0.9)	PAD -ve (conventional ABI ≥0.9)	Total
Test +ve (pragmatic ABI <0.9)	32	16	48
Test -ve (pragmatic ABI ≥0.9)	20	174	194
Total	52	190	242

Table 2. 2x2 table of conventional and pragmatic ABI determinations.

Formatted: Normal, Left, Line spacing: single

REFERENCES

1. Belch JJF, Topol EJ, Agnelli G, Bertrand M, Califf RM, Clement DL, et al. Critical Issues in Peripheral Arterial Disease Detection and Management: A Call to Action. *Arch. Intern. Med.* 2003;163(8):884-92.
2. Golomb BA, Dang TT, Criqui MH. Peripheral Arterial Disease: Morbidity and Mortality Implications. *Circulation* 2006;114(7):688-99.
3. McDaniel M, Cronenwett J. Basic data related to the natural history of intermittent claudication. *Ann. Vasc. Surg.* 1989;3:273-77.
4. Steg PG, Bhatt DL, Wilson PWF, D'Agostino R, Sr, Ohman EM, Rother J, et al. One-Year Cardiovascular Event Rates in Outpatients With Atherothrombosis. *JAMA* 2007;297(11):1197-206.
5. McKenna M, Wolfson S, Kuller L. The ratio of ankle and arm arterial pressure as an independent predictor of mortality. *Atherosclerosis* 1991;87:119 -28.
6. Nelson MR, Quinn S, Bowers-Ingram L, Nelson JM, Winzenberg TM. Cluster Randomized Controlled Trial of Oscillometric Versus Manual Sphygmomanometer for Blood Pressure Management in Primary Care (CRAB). *Am. J. Hypertens.* 2009; doi:10.1038/ajh.2009.55.
7. Beckman JA, Higgins CO, Gerhard-Herman M. Automated Oscillometric Determination of the Ankle-Brachial Index Provides Accuracy Necessary for Office Practice. *Hypertension* 2006;47(1):35-38.
8. Bhatt DL, Steg PG, Ohman EM, Hirsch AT, Ikeda Y, Mas J-L, et al. International Prevalence, Recognition, and Treatment of Cardiovascular Risk Factors in Outpatients With Atherothrombosis 10.1001/jama.295.2.180. *JAMA* 2006;295(2):180-89.

- 1
2
3
4
5
6
7
8 9. Reid CM, Nelson M, Chew D, Connor G, DeLoose F. Australians @ Risk: management of
9 cardiovascular risk factors in the REACH registry. *Heart Lung and Circulation*
10 2007;doi10.1016/j.hlc.2007.07.009.
11
12
13
14 10. Doobay AV, Anand SS. Sensitivity and Specificity of the Ankle–Brachial Index to Predict
15 Future Cardiovascular Outcomes. *Arterioscler. Thromb. Vasc. Biol.* 2005;25(7):1463-69.
16
17
18 11. Assaad MA, Topouchian JA, Darné BM, Asmar RG. Validation of the Omron HEM-907
19 device for blood pressure measurement. *Devices and Technology* 2002;7(4):237-41.
20
21
22
23 12. White WB, Anwar YA. Evaluation of the overall efficacy of the Omron office digital blood
24 pressure HEM-907 monitor in adults. *Devices and Technology* 2001;6(2):107-10.
25
26
27 13. Leng GC, Fowkes F. The Edinburgh Claudication Questionnaire: an improved version of the
28 WHO/ROSE Questionnaire for use in epidemiological surveys. *Journal of Clinical*
29 *Epidemiology* 1992;45(10): 1101-1109.
30
31
32 14. Bland JM Altman DG. Statistical methods for assessing agreement between two methods of
33 clinical measurement. *Lancet* 1986;1(8476):307-10.
34
35
36
37 ~~14-15.~~ Hosmer D, Lemeshow S. Applied Logistic Regression. 2nd ed. New York: John Wiley
38 and Sons, Inc., 2000.
39
40
41 ~~15-16.~~ Benchimol A, Bernard V, Pillois X, Hong NT, Sagardiluz P, Tortelier L, et al.
42 Validation of a new method of detecting peripheral artery disease by determination of ankle-
43 brachial index using an automatic blood pressure device. *Angiology* 2004;55 (2):127-34.
44
45
46
47 ~~16.~~ Šimundić A-M. Measures of diagnostic accuracy: basic definitions.
48 <http://www.ifcc.org/ifccfiles/docs/190404200805.pdf>
49
50
51
52
53
54
55
56
57
58
59
60

Formatted: Font: Italic

Formatted: Font: Italic

Formatted: Font: Italic

1
2
3
4
5
6
7
8 17. [Šimundić A-M. Measures of diagnostic accuracy: basic definitions.](#)

9
10 <http://www.ifcc.org/ifccfiles/docs/190404200805.pdf>

11
12 [Doubeni CA, Yood RA, Emani S, Gurwitz JH. Identifying Unrecognized Peripheral Arterial](#)
13 [Disease Among Asymptomatic Patients in the Primary Care Setting. *Angiology* 2006;57\(2\):35-](#)
14 [38.](#)

15
16
17
18 18. [Mehlsen J, Wiinberg N, Bruce C. Oscillometric blood pressure measurement: a simple](#)
19 [method in screening for peripheral arterial disease. *Clinical Physiology and Functional Imaging*](#)
20 [2008;28\(6\):426-29.](#)

21
22
23
24 [Bendermacher BLW, Teijinka JAW, Willigendael EM, Bartelink M-L, Peters RJB, de Bie RA,](#)
25 [Büller HR, Boiten J, Langenberg M, Prins MH. A clinical prediction model for the presence of](#)
26 [peripheral arterial disease—the benefit of screening individuals before initiation of measurement](#)
27 [of the ankle-brachial index: an observational study. *Vasc Med* 2007;12\(1\):5-11.](#)

28
29
30
31
32 19. [Nicolai SPA, Kruidenier LM, Rouwet EV, Bartelink M-LEL, Prins MH, Teijink JA. Ankle](#)
33 [Brachial index in primary care: are we doing it right? *Br. J. Gen. Pract.* 2009;59\(563\):422-27.](#)

34
35
36 20. [Aboyans V, Lacroix P, Doucet S, Preux PM, Criqui MH, Laskar M. Diagnosis of peripheral](#)
37 [arterial disease in general practice: can the ankle-brachial index be measured either by pulse](#)
38 [palpation or an automatic blood pressure device?. *Int. J. Clin. Pract.* 2008;62\(7\):1001-07.](#)

39
40
41
42 [Mohler ER 3rd, Treat Jacobson D, Reilly MP, Cunningham KE, Miani M, Criqui MH, et al.](#)
43 [Utility and barriers to performance of the ankle-brachial index in primary care practice. *Vasc*](#)
44 [Med](#) 2004;9(4):253-60.

Formatted: Font: Italic

1
2
3
4
5
6
7
8 20. Lee B,m Campbell JS, Berkowitz P. The correlation of ankle oscillometric blood pressures
9 and segmental pulse volumes to Doppler systolic pressures in arterial occlusive disease. *J Vasc*
10 *Surg* 1996;23(1):116-22.

11
12
13
14 21. Verberk WJ, Kollias A, Stergiou GS. Automated oscillometric determination of the ankle-
15 brachial index: a systematic review and meta-analysis. *Hypertens. Res.* 2012.

16
17
18 Nukumizu Y, Matsushita M, Sakurai T, Kobayashi M, Nishikimi N, Komori K. Comparison of
19 Doppler and oscillometric ankle blood pressure measurement in patients with angiographically
20 documented lower extremity arterial occlusive disease. *Angiology* 2007;58(3):303-8.

21
22
23
24 22. Raines JK, Farrar J, Noicely K, Pena J, et al. Ankle/Brachial Index in the primary care
25 setting. *Vascular and Endovascular Surgery* 2004;38(2):131-36.

26
27
28 Ramanathan A, Conaghan PJ, Jenkinson AD, Bishop CR. Comparison of ankle brachial
29 pressure index measurements using an automated oscillometric device with the standard doppler
30 ultrasound technique. *ANZ J of Surg* 2003;73(3):105-08.

31
32
33
34 23. Vega J, Román S, Garcipérez FJ, Vicente L, Pacheco N, Zamorano J, et al. Peripheral
35 Arterial Disease: Efficacy of the Oscillometric Method. *Rev. Esp. Cardiol.* 2011;64(07):619-21.

36
37
38 Pan CR, Staessen JA, Li Y, Wang JG. Comparison of Three Measures of the Ankle-Brachial
39 Blood Pressure Index in a General Population. *Hypertens Res* 2007;30(6):555-61.

40
41
42
43 24. Mohler ER 3rd, Treat-Jacobson D, Reilly MP, Cunningham KE, Miani M, Criqui MH, et al.
44 Utility and barriers to performance of the ankle-brachial index in primary care practice. *Vasc*
45 *Med* 2004;9(4):253-60.

46
47
48 25. Raines JK, Farrar J, Noicely K, Pena J, et al. Ankle/Brachial Index in the primary care
49 setting. *Vascular and Endovascular Surgery* 2004;38(2):131-36.

Formatted: Font: Italic

1
2
3
4
5
6
7
8 | ~~25.~~ Bossuyt PM, Reitsma JB, Bruns DE, Gatsonis CA, Glasziou PP, Irwig LM, et al. The
9 | STARD Statement for Reporting Studies of Diagnostic Accuracy: Explanation and Elaboration.
10 |
11 |
12 | *Clin Chem* 2003;49(1):7-18.

13
14
15
16
17
18
19
20
21 | ~~1. Steg PG, Bhatt DL, Wilson PWF, D'Agostino R, Sr., Ohman EM, Rother J, et al. One Year~~
22 | ~~Cardiovascular Event Rates in Outpatients With Atherothrombosis. *JAMA* 2007;297(11):1197-~~
23 | ~~206.~~

24
25
26
27 | ~~2. McKenna M WS, Kuller L. The ratio of ankle and arm arterial pressure as an independent~~
28 | ~~predictor of mortality. *Atherosclerosis* 1991;87:119-28.~~

29
30
31 | ~~3. Nelson MR, Quinn S, Bowers Ingram L, Nelson JM, TM W. Cluster Randomized Controlled~~
32 | ~~Trial of Oscillometric Versus Manual Sphygmomanometer for Blood Pressure Management in~~
33 | ~~Primary Care (CRAB). *Am. J. Hypertens.* 2009;In press.~~

34
35
36
37 | ~~4. Beckman JA, Higgins CO, Gerhard-Herman M. Automated Oscillometric Determination of~~
38 | ~~the Ankle Brachial Index Provides Accuracy Necessary for Office Practice. *Hypertension*~~
39 | ~~2006;47(1):35-38.~~

40
41
42
43 | ~~5. Bhatt DL, Steg PG, Ohman EM, Hirsch AT, Ikeda Y, Mas J-L, et al. International Prevalence,~~
44 | ~~Recognition, and Treatment of Cardiovascular Risk Factors in Outpatients With~~
45 | ~~Atherothrombosis~~
46 | ~~10.1001/jama.295.2.180. *JAMA* 2006;295(2):180-89.~~

Formatted: Line spacing: Double

Formatted: Indent: Left: 0", First line: 0",
Space After: 6 pt, Line spacing: Double

6. Doobay AV, Anand SS. Sensitivity and Specificity of the Ankle-Brachial Index to Predict Future Cardiovascular Outcomes. *Arterioscler. Thromb. Vasc. Biol.* 2005;25(7):1463-69.
7. Bland JM AD. Statistical methods for assessing agreement between two methods of clinical measurement. *Lancet* 1986;1(8476):307-10.
8. Hosmer D, S L. *Applied Logistic Regression.* 2nd ed. New York: John Wiley and Sons, Inc., 2000.
9. Benchimol A BV, Pillois X, Hong NT, et al. Validation of a new method of detecting peripheral artery disease by determination of ankle-brachial index using an automatic blood pressure device. *Angiology* 2004;55 (2):127-34.
10. Šimundić A M. Measures of diagnostic accuracy: basic definitions.
11. Reid CM NM, Chew D, Connor G, DeLoose F. Australians @ Risk: management of cardiovascular risk factors in the REACH registry. *Heart Lung and Circulation* 2007;doi10.1016/j.hlc.2007.07.009.
12. Doubeni CA YR, Emani S, Gurwitz JH. Identifying Unrecognized Peripheral Arterial Disease Among Asymptomatic Patients in the Primary Care Setting. *Angiology* 2006;57(2):35-38.
13. Bendermacher BL TJ, Willigendael EM, Bartelink ML, Peters RJ, de Bie RA, et al. A clinical prediction model for the presence of peripheral arterial disease—the benefit of screening individuals before initiation of measurement of the ankle-brachial index: an observational study. *Vasc. Med.* 2007;12(1):5-11.

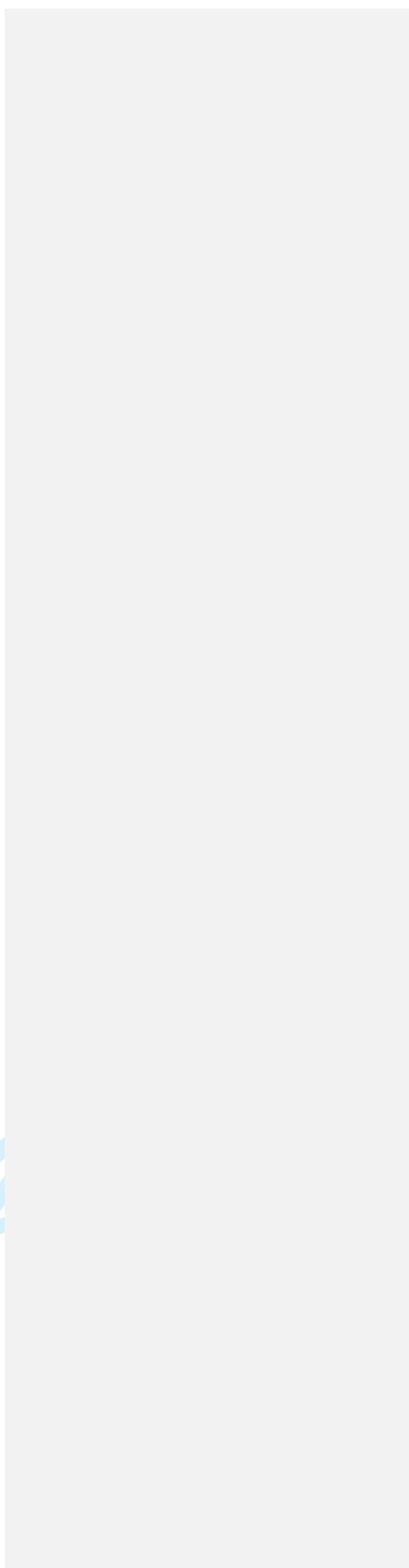
- 1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
14. Mohler ER 3rd T JD, Reilly MP, Cunningham KE, Miani M, Criqui MH, et al. Utility and barriers to performance of the ankle-brachial index in primary care practice. *Vasc. Med.* 2004;9(4):253-60.
15. Lee B CJ, Berkowitz P. The correlation of ankle-oscillometric blood pressures and segmental pulse volumes to Doppler systolic pressures in arterial occlusive disease. *J. Vasc. Surg.* 1996;23(1):116-22.
16. Nukumizu Y MM, Sakurai T, Kobayashi M, Nishikimi N, Komori K. Comparison of Doppler and oscillometric ankle blood pressure measurement in patients with angiographically documented lower extremity arterial occlusive disease. *Angiology* 2007;58(3):303-8.
17. Ramanathan A CP, Jenkinson AD, Bishop CR. Comparison of ankle-brachial pressure index measurements using an automated oscillometric device with the Ankle/Brachial Index in the primary care setting. standard doppler ultrasound technique. *ANZ J of Surg* 2003;73(3):105-08.
18. Pan C R SJ, Yan Li Y, Wang J G. Comparison of Three Measures of the Ankle-Brachial Blood Pressure Index in a General Population. *Hypertens. Res.* 2007;30(6):555-61.
19. Raines JK FJ, Noicely K, Pena J, et al. Ankle/Brachial Index in the primary care setting. *Vascular and Endovascular Surgery* 2004;38(2):131-36.
20. Bossuyt PM, Reitsma JB, Bruns DE, Gatsonis CA, Glasziou PP, Irwig LM, et al. The STARD Statement for Reporting Studies of Diagnostic Accuracy: Explanation and Elaboration. *Clin. Chem.* 2003;49(1):7-18.

Formatted: Line spacing: Double

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

|

For peer review only



STARD checklist for reporting of studies of diagnostic accuracy
(version January 2003)

Section and Topic	Item #		On page #
TITLE/ABSTRACT/ KEYWORDS	1	Identify the article as a study of diagnostic accuracy (recommend MeSH heading 'sensitivity and specificity').	In keywords
INTRODUCTION	2	State the research questions or study aims, such as estimating diagnostic accuracy or comparing accuracy between tests or across participant groups.	7
METHODS			
<i>Participants</i>	3	The study population: The inclusion and exclusion criteria, setting and locations where data were collected.	7, 8
	4	Participant recruitment: Was recruitment based on presenting symptoms, results from previous tests, or the fact that the participants had received the index tests or the reference standard?	7
	5	Participant sampling: Was the study population a consecutive series of participants defined by the selection criteria in item 3 and 4? If not, specify how participants were further selected.	8
	6	Data collection: Was data collection planned before the index test and reference standard were performed (prospective study) or after (retrospective study)?	4
<i>Test methods</i>	7	The reference standard and its rationale.	
	8	Technical specifications of material and methods involved including how and when measurements were taken, and/or cite references for index tests and reference standard.	8
	9	Definition of and rationale for the units, cut-offs and/or categories of the results of the index tests and the reference standard.	8
	10	The number, training and expertise of the persons executing and reading the index tests and the reference standard.	8
	11	Whether or not the readers of the index tests and reference standard were blind (masked) to the results of the other test and describe any other clinical information available to the readers.	8
<i>Statistical methods</i>	12	Methods for calculating or comparing measures of diagnostic accuracy, and the statistical methods used to quantify uncertainty (e.g. 95% confidence intervals).	9-10
	13	Methods for calculating test reproducibility, if done.	N/A
RESULTS			
<i>Participants</i>	14	When study was performed, including beginning and end dates of recruitment.	8
	15	Clinical and demographic characteristics of the study population (at least information on age, gender, spectrum of presenting symptoms).	20
	16	The number of participants satisfying the criteria for inclusion who did or did not undergo the index tests and/or the reference standard; describe why participants failed to undergo either test (a flow diagram is strongly recommended).	16
<i>Test results</i>	17	Time-interval between the index tests and the reference standard, and any treatment administered in between.	8
	18	Distribution of severity of disease (define criteria) in those with the target condition; other diagnoses in participants without the target condition.	12
	19	A cross tabulation of the results of the index tests (including indeterminate and missing results) by the results of the reference standard; for continuous results, the distribution of the test results by the results of the reference standard.	21
	20	Any adverse events from performing the index tests or the reference standard.	11
<i>Estimates</i>	21	Estimates of diagnostic accuracy and measures of statistical uncertainty (e.g. 95% confidence intervals).	10-11
	22	How indeterminate results, missing data and outliers of the index tests were handled.	12
	23	Estimates of variability of diagnostic accuracy between subgroups of participants, readers or centers, if done.	N/A
	24	Estimates of test reproducibility, if done.	N/A
DISCUSSION	25	Discuss the clinical applicability of the study findings.	11

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

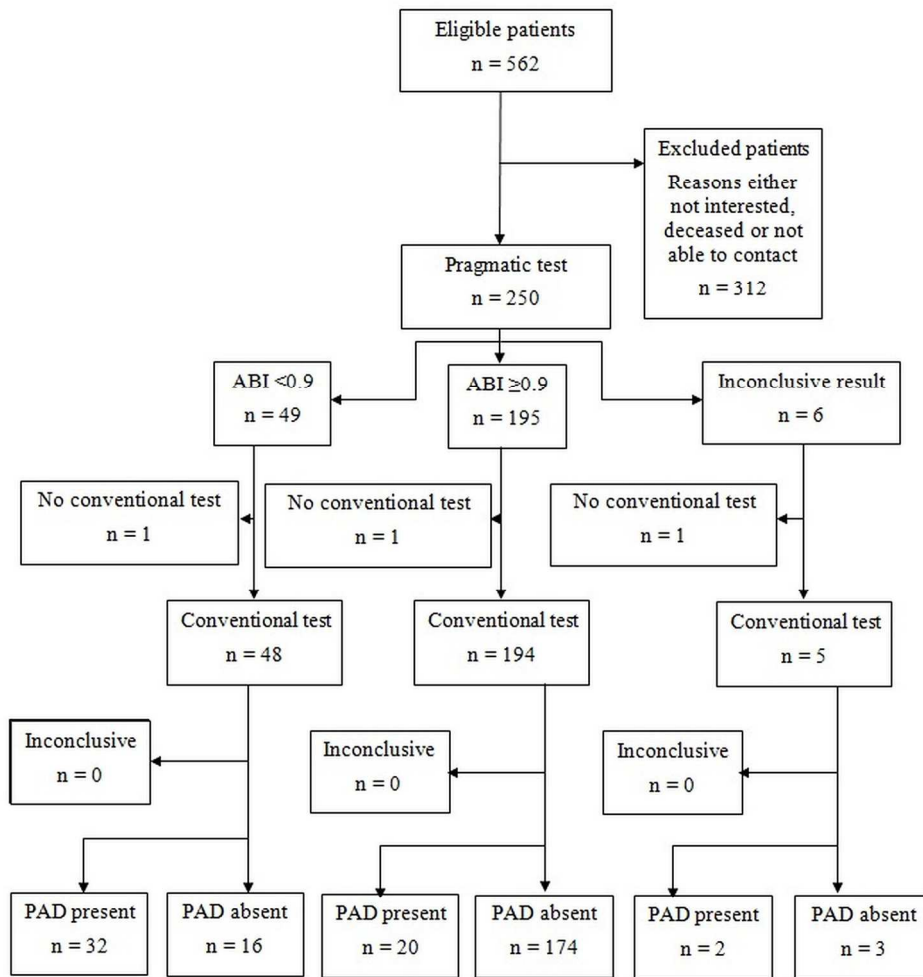


Figure 1. Flow diagram of a diagnostic accuracy in ABIDING as per STARD standard²⁵.

186x209mm (300 x 300 DPI)

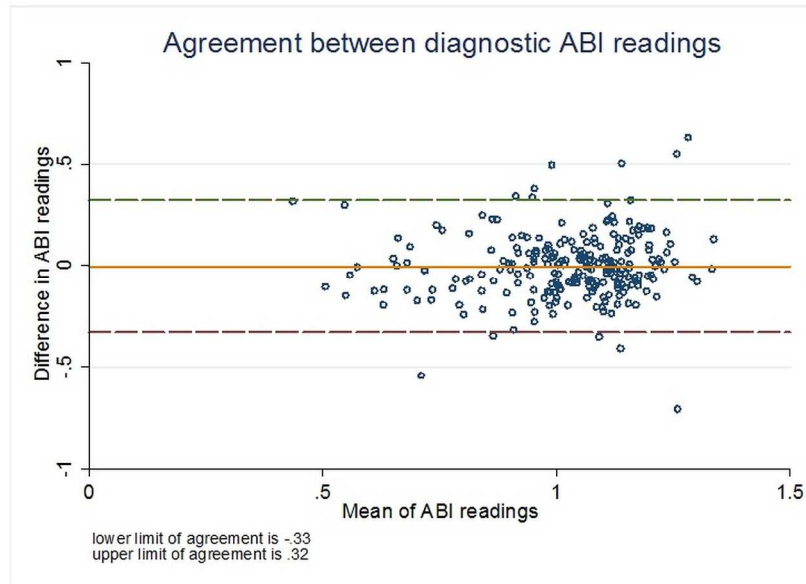


Figure 2. Agreement between pragmatic and conventional determination of ABI.

238x188mm (300 x 300 DPI)

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

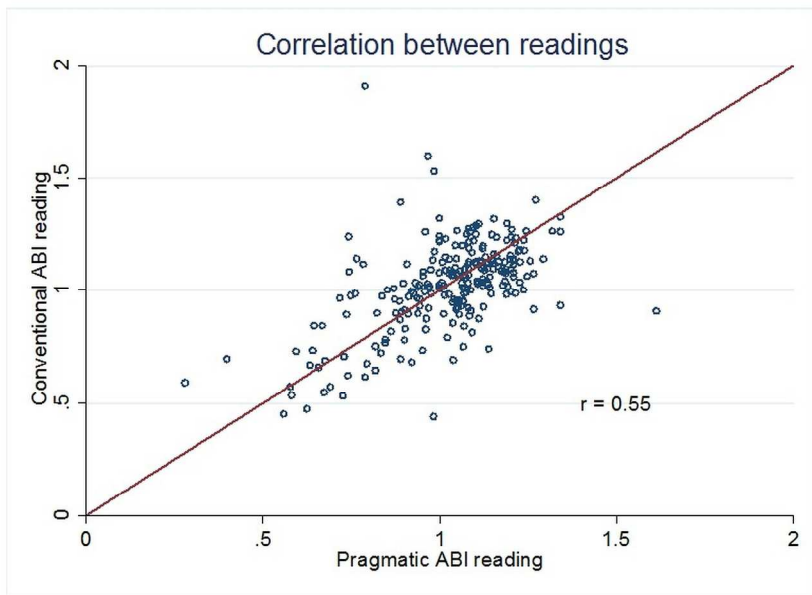


Figure 3. Correlation between pragmatic and conventional determination of ABI.

250x179mm (300 x 300 DPI)

View only

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60

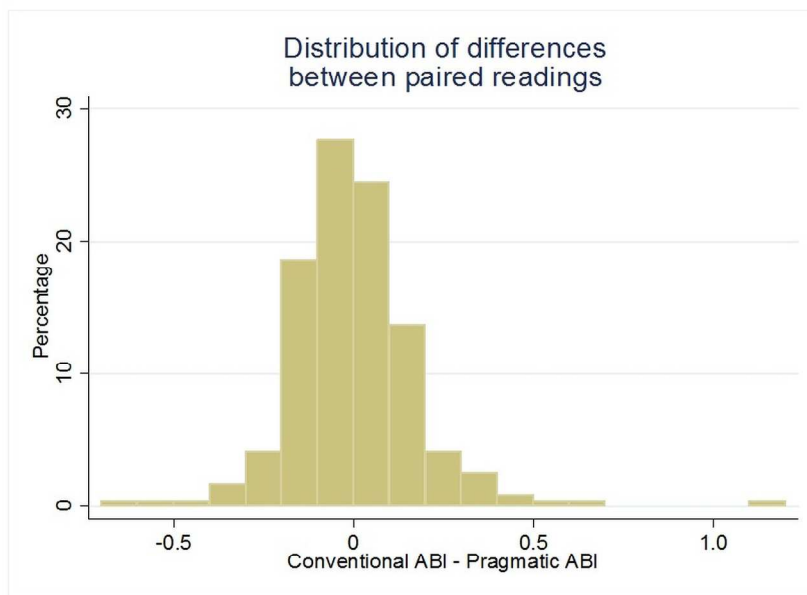


Figure 4. Distribution of the difference between the conventional ABI and the pragmatic ABI readings.

264x202mm (300 x 300 DPI)

Review only