Online Supplementary Materials

Section 1 Detailed Model Formulation and Estimation Algorithms.

536 **Section S1.1 Ozone Prediction Model**

537 We considered the following process-based linear regression model for predicting daily 8-hr maximum 538 ozone concentrations: on each day *t*,

- 539
-

540
$$
Y_t = \beta_0 + \beta_1 T max_t + \beta_2 S R_t + \beta_3 P R_t + \beta_4 N O x_t + \beta_5 (N O x_t)^2
$$

$$
+ \beta_6 N O x_t \times \log(V O C_t) + \beta_7 P R \times N O x_t + \beta_8 T max_t \times N O x_t + \beta_9 T max_t \times (N O x_t)^2 + v_t
$$

542

543 where *Y* is the square-root of maximum 8-hour daily ozone level, Tmax is the daily maximum 544 temperature; *SR* is the total solar radiation, *PR* is the daily precipitation; *NOx* is the collocated NOx level; 545 and *VOC* is the collocated non-methane volatile organic compound level. We assume the residual error 546 is temporally correlated following an auto-regressive of order-1 model: 547

$$
v_t = \alpha v_{t-1} + \varepsilon_t
$$

$$
\varepsilon_t \sim \text{Normal}(0, \sigma^2).
$$

548

549 All unknown parameters (β' s, α , and σ^2) were estimated using the R function *arima* with order (1,0,0). 550 Supplementary Table S1 gives the estimated values. The covariance matrix of the estimated parameters 551 was also obtained.

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553 **Section S1.2 Ozone-Asthma/Wheeze Health Effect Estimation**

554 The short-term association between ambient 8-hr maximum ozone levels and asthma/wheeze ED visits 555 for 1999-2004 was estimated under a time series design, using a Poisson log-linear model with over-556 dispersion. We considered a 3-day moving average (of 0-, 1-, and 2-day lags) of ozone as the *a priori* lag 557 structure. The Poisson model has the form:

558

$$
ED_t \sim \text{Poisson } (\mu_t)
$$

$$
\log(\mu_t) = \alpha_0 + \lambda \overline{Y}_t + \sum_{j=1}^J \alpha_j W_j
$$

Occur-

Prelation involve is the strategies of the controlled for predicting daily 8-hr maximulations: on each day t,
 $Y_t = \beta_0 + \beta_1 T m \alpha x_t + \beta_2 SR_t + \beta_3 PR_t + \beta_4 N Ox_t + \beta_5 (N Ox_t)^2$
 $x_t \times \log(VOC_t) + \beta_7 PR \times N Ox_t + \beta_6 T m \alpha x_t \times N Ox_t + \beta_$ 559 where ED_t is the total number of ED visits on day *t*, Y_t is the 3-day moving average of maximum 8-hour 560 cozone level, and W_j denotes the jth confounder. The confounders included indicators for day of week, 561 holidays, seasons, and whether each hospital contributed to the daily ED counts. Meteorology was 562 controlled using indicators for each degree Celsius of same-day maximum temperature and natural 563 cubic splines with 3 degrees of freedom for moving averages of dew-point temperature (lags 0, 1, and 2) 564 and minimum temperature (lags 1, and 2). Long-term trends were controlled by year-specific natural 565 cubic splines with monthly knots. As a sensitivity analysis, we increased the number of knots per year to 566 examine the robustness of the estimated relative risks (Supplementary Figure S1).

567

568 The relative risk was λ estimated with the *glm* function in R. The standard error was also obtained.

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570 **Section S1.3 Climate Model Calibration**

571 Let $x_{[i]}$ and $z_{[i]}$ denote the i^{th} largest value of the observed value and the climate model output,

572 respectively. Each climate model variable was calibrated separately. The rank-based bias correction

573 assumes $x_{[i]} = G(z_{[i]})$. We assumed G() as a smooth function that is parameterized through an cubic 574 splines:

$$
x_{[i]} = G(z_{[i]}) = \sum_{k=1}^{20} \phi_k b_k(z_{[i]})
$$

575 where $b(x)$ is a monotonic cubic spline basis transformation such that $|G(f)|$ is non-decreasing, and ϕ_k is

- 576 the corresponding regression coefficient for each basis function. Regularization on ϕ_k was imposed
- 577 based on generalized cross-validation statistics. The above model was fitted using the R function *gam*
- 578 under the package mgcv. Let $\hat{\phi}_k$ be the estimated parameter based on the historical calibration period.
- 579 Let \tilde{z}_t be a the future raw climate model out on day t. Then the calibrated model output is given by 580 $\sum_{k=1}^{20} \hat{\phi}_k b_k(\tilde{z}_t)$.

581 **Section S1.4 Ozone Projection**

a monotonic cause spline basis funsionmaton such that Q_f is non-
adding regression coefficient for each basis function. Regularization on ϕ_k was impose
creating regression coefficient for each basis function. Regular 582 For a particular future day t, let \widetilde{Imax}_t , \widetilde{SR}_t , and \widetilde{PR}_t be the calibrated climate model output. Similarly 583 Let \widetilde{NOx}_t and \widetilde{VOC}_t be the corresponding projected precursor levels. Future VOC and NOx levels were 584 obtained by multiplying the 1999-2004 VOC levels by 8% and NOx levels by 29%, based on the A2 585 scenario comparing 2050 to 2000 emission levels for industrialized countries (OECD90) (IPCC 2000, 586 Hogrefe et al. 2004). To cover the 30-year future period, we repeated the 6-year historical VOC/NOx 587 time series (increased by 8 and 29%, respectively) 5 times.

588

589 The following algorithms were used to obtain realizations of future ozone time series. For each iteration *m*, 590

- 591 (1) Simulate $\beta_0^{(m)}, \beta_1^{(m)}, ..., \beta_9^{(m)}$, and $\alpha^{(m)}$ from a multiavariate Normal distribution with mean 592 equal to the estimated values and the associated variance covariance matrix.
- 593 (2) For each future day *t*, calculate the ozone mean prediction $\hat{Y}_t^{(m)}$ as
- 594

595
$$
\begin{aligned}\n\hat{Y}_t^{(m)} &= \beta_0^{(m)} + \beta_1^{(m)} \widetilde{Tm} \alpha x_t + \beta_2^{(m)} \widetilde{SR}_t + \beta_3^{(m)} \widetilde{PR}_t + \beta_4^{(m)} \widetilde{NOx}_t + \beta_5^{(m)} (\widetilde{NOx}_t)^2 \\
&+ \beta_6^{(m)} \widetilde{NOx}_t \times \log(\widetilde{VOC}_t) + \beta_7^{(m)} \widetilde{PR}_t \times \widetilde{NOx}_t + \beta_8^{(m)} \widetilde{Tm} \alpha x_t \times \widetilde{NOx}_t + \beta_9 \widetilde{Tm} \alpha x_t \times (\widetilde{NOx}_t)^2\n\end{aligned}
$$

- 597
- 598 (3) Simulate a time series of $\hat{v}^{(m)}_t$ using the R function *arima.sim* to obtain

599
$$
\hat{v}_t^{(m)} = \alpha^{(m)} \hat{v}_t^{(m)} + \varepsilon_t, \ \varepsilon_t \sim \text{Normal}(0, \hat{\sigma}^2).
$$

600 (4) Obtain the projected ozone time series $\tilde{Y}_t^{(m)}$ (with back-transformation) as

$$
\tilde{Y}_t^{(m)} = \left(\hat{Y}_t^{(m)} + \hat{v}_t^{(m)}\right)^2.
$$

- 601 Repeat Steps (1) to (5) 5000 times to obtain $\tilde{Y}_t^{(m=1)}$, $\tilde{Y}_t^{(m=2)}$, ..., $\tilde{Y}_t^{(m=5000)}$.
- 602 Besults from Table 3 were obtained by first averaging each series of $\tilde{Y}_t^{(m=1)}$ across days. The mean and 603 95% quantile interval of were then computed across the 5,000 simulated average ozone levels.

604

605 Exceedance days were estimated by first calculating the number of days in each series $\tilde{Y}_t^{(m)}$ exceeding

 606 75ppb per year. The mean and 95% quantile intervals were than obtained across the 5,000 exceedance 607 annual average counts.

615

$$
EED^{(m)} = M \times (e^{\lambda^{(m)} \Delta Y^{(m)}} - 1)
$$

the estimated standard error.
 $EED^{(m)} = M \times (e^{\lambda^{(m)}} \Delta r^{(m)} - 1)$
 $EED^{(m)} = M \times (e^{\lambda^{(m)} \Delta r^{(m)}} - 1)$
 $T = M$ is the expected number of annual ED visits (33,551) and $\Delta Y^{(m)}$ is the difference

orange ozone levels between the 616 where, *M* is the expected number of annual ED visits (33,551) and $\Delta Y^{(m)}$ is the difference 617 in average ozone levels between the historical period (1999-2004, 51.8ppb) and the future in average ozone levels between the historical period (1999-2004, 51.8ppb) and the future 618 period (2041- 2070):

$$
\Delta Y^{(m)} = \frac{1}{T} \sum_{t=1}^{T} \tilde{Y}_t^{(m)} - 51.8 \, .
$$

619

620 Repeat Steps (1) and (2) for each m and obtain 5,000 simulations of $EED^(m)$. Results from Figure 2 were

621 obtained by calculating the mean and 95% quantile intervals for $EED^{(m)}$ across the 5,000 simulations for

622 each climate model.

Section 2 Supplementary Tables and Figures

Table S1. Parameter estimates of the daily square-root 8-hr maximum ozone prediction model with predictors daily maximum temperature (Tmax, **˚F**), 24-hour NOx (ppm), 24-hour VOC (ppmC), total solar radiation (SR, Wh/m2), and total precipitation (PR, mm).

* α measures the lag-1 temporal correlation dependence in the residual error.

** σ is the standard deviation of the residual error.

Table S2. Projected change in the number of days per year with daily 8-hr maximum ozone concentrations exceeding 75ppb between 2041-2070 and 1999-2004. 95% projection intervals are given in parenthesis. The number of exceedance days during 1999-2004 was 34.

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Figure S1. Estimated relative risks (RR) and 95% confidence intervals for daily asthma/wheeze emergency department visits associated with an interquartile range (22 ppb) increase in three-day moving average 8-hour maximum ozone concentration in Atlanta, 1999-2004 during March to October. Estimates for various degrees of freedom (DF) per year for temporal trend fitted using natural cubic splines are shown. The solid plot symbol indicates the RR used in the health impact projection calculations.

Figure S2. Cumulative distribution functions for observed daily maximum temperature and climate model outputs during 1999 – 2000 (March to October). For CRCM-CCSM and WRFG-CCSM only data during 1999 were available.

Figure S3. Cumulative distribution functions for observed daily total solar radiation and climate model outputs during 1999 – 2000 (March to October). For CRCM-CCSM and WRFG-CCSM only data during 1999 were available.

Figure S4. Cumulative distribution functions for observed daily total precipitation and climate model outputs during 1999 – 2000 (March to October). For CRCM-CCSM and WRFG-CCSM only data during 1999 were available.

