

# Lung Spore Analysis Report: Affymetrix Data Analysis

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# 1 Executive Summary

## 1.1 Introduction

This report describes the analysis of a data set from Dr. Milind Suraokar, a member of the laboratory of PI Dr. Ignacio I. Wistuba. This dataset was acquired using Affymetrix HG-U133Plus2.0. There are 96 mesothelioma tissue arrays in total.

### 1.1.1 Aims/Objectives

In the series of the reports, we are interested in the probesets that show different between normal and tumor tissues. For the previous analysis we have done on the clean data, the separation of groups is very fragile corresponding to the selected length of genes or including number of samples. In this report, we find the groups defined most stably.

## 1.2 Methods

### 1.2.1 Description of the Data

The data are processed by Affymetrix HGU133plus2.0. There are 96 arrays in total. Among them, 41 are coded as matched tumor and normal. Another 14 arrays are no paired tumor tissues. Among those 41 matched arrays, Dr. Suraokar indicated that Normal 5 is actually tumor tissue, and Tumor 22, Tumor 4 are actually normal tissues. Normal 5 is a normal sample but we found it is tumor, so during our tumor cluster analysis, we remove it. So, there are 53 tumor tissues in total. There are 38 paired tumor and normal samples. In the 38 paired samples, there are 29 epithelioid, 4 sarcomatoid and 5 biphasic samples. For the clinical related analysis, samples Tumor 6, Tumor 17,

Tumor 44, Tumor 54, Tumor 45 have their death seems surgery related, and Normal 5 is a normal sample but we found it is tumor, so we remove them and keep only 48 tumors in use.

### 1.2.2 Statistical Methods

We first remove the batch effect. Based on the data, we want to find the most robust subgroups of patients:

1. Compare tumor vs normal in paired data
2. Randomly select n (an uniform number from 500 to 2000) probes from the probesets that are selected at the FDR level of 0.01 (around 20000 probes in total), and randomly select 80% of samples. Do cluster.
3. Repeat 2 for 1000 times, record each time whether any two sample are grouped together if cut tree with 3 or 4 groups. Define the consensus matrix as the average time that any two samples are grouped together in each clustering.
4. Use the consensus matrix as the similarity matrix to define the final clustering group.
5. Use cophenetic correlation or other criteria to define the number of subgroups we would like to use
6. For the subgroups we defined, we test in KM to check the association with clinical variable and find the genes that are different between the groups

We also apply KM or Cox model for different clinical information corresponding to overall survival.

1. Age
2. Gender
3. Original Histology
4. T Stage
5. N Stage
6. Overall Stage
7. Chemo Treatment.

Variables are included in the Cox model for multivariate analysis. Akaike Information Criterion (AIC) to eliminate redundant variables from the model.

After the best survival model is selected, we add one more variable factor (the three sub-tumor groups) in the Cox model and perform the overall survival analysis.

### 1.3 Results

For univariate analysis, the three sub-tumor group is not significant associated with overall survival.

There are four variables remained for multivariate analysis after Akaike Information Criterion (AIC) to eliminate redundant variables from the model:

1. Gender
2. Original Histology
3. Combined N Stage
4. Chemo Treatment

After we include one more variable factor (the three sub-tumor groups) in the Cox model we selected in the previous step and perform the overall survival analysis, this new variable will not be excluded after AIC to eliminate redundant variables.

### 1.4 Conclusions

We define three survival related sub-tumor groups.

## 2 Loading the Data

### 2.1 R Libraries

We begin by loading all the libraries we will need for this analysis. A list of the current versions of the libraries used for the analysis can be found in the appendix.

```
> library(affy)
> library(simpleaffy)
> library(geneplotter)
> library(xtable)
> library(ClassComparison)
> library(ClassDiscovery)
> library(limma)
> library(hgu133plus2.db)
> library(gplots)
> library(gdata)
> library(affyio)
> library(nlme)
> library(survival)
```

## 2.2 Load Data Object

We first load in the data object we have saved in the previous report.

```
> mainDirectory<-"/data/bioinfo/Private/LungSpore"
> DataDirectory<-"/data/bioinfo2/Lung-HN/Mesothelioma"
> setwd(file.path(mainDirectory, "ReportWithNewClin"))
> load(file.path(mainDirectory, "Report", "processedDataAffy.RData"))
> my.fig <- "Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Edit-NewID"
```

We check the name and the clinical information consistency.

```
> AllData<-exprs(x.rma)
> identical(substr(colnames(AllData), 19,21), substr(si[, "core-id"], 8,10))
[1] TRUE
>
```

We then read in the sample ID.

```
> si2 <- read.xls("Affy-U133-samples_091111.xls")
> identical(si[, "core-id"], si2[, "core.id"])
[1] TRUE

> si <- cbind(si, si2)
> si3 <- read.xls("Numbered-DeID-samples.xls")
> identical(si[, "core-id"], si3[, "core.id"])
[1] TRUE

> si <- cbind(si, SepID = as.vector(si3[, "Sep.2013.ID"]))
> colnames(AllData) <- as.vector(si[, "SepID"])
>
```

We read in the run date of each sample first.

```
> celDatHeaders<-colnames(AllData)
> for (i1 in 1:length(colnames(AllData))) {
+ temp <- read.celfile.header(file.path(DataDirectory, "Affymetrix-mRNA", si[i1, "fnames"]), ini
+ celDatHeaders[i1] <- temp$DatHeader
+ }
> #celDatHeaders[1:3]
> temp <- strsplit(celDatHeaders, "[[:space:]]+")
> celRunDates <- unlist(lapply(temp, function(x) {x[8]}))
> zed <- as.Date(celRunDates, format="%m/%d/%y")
> names(zed)<-si[,c("Sample.ID..IW.")]
> table(zed)
```

```

zed
2009-02-24 2009-02-26 2009-03-12 2009-03-13 2009-03-25 2009-04-07 2009-04-09 2009-04-14
      12          12          12          12          12          12          12          12
> si <- cbind(si, zed)
>

```

We create another date factor to separate the run date for months.

```

> dateFactor2<-rep("2009-02",length(zed))
> dateFactor2[grep("2009-03", zed)]<-"2009-03"
> dateFactor2[grep("2009-04", zed)]<-"2009-04"
> names(dateFactor2)<-si[, "Sample ID"]
> si <- cbind(si, dateFactor2)
> rm(dateFactor2, zed, celDatHeaders, celRunDates)

```

A mesothelioma tumor can be of three types: epitheloid, sarcomatoid and biphasic (contains a mix of both sarcomatoid and epitheloid). PI submit one diagnosis: the “Path Report Diagnosis” (the original tumor diagnosis in the patient). We load in the corresponding clinical information.

```

> Histology <- read.xls("Meso-histology-May2010.xlsx")
> dim(Histology)
[1] 56  2

> identical(as.vector(Histology[,1]), as.vector(si[, "Sample ID"]))
[1] TRUE

> si <- cbind(si, Histology)
> rm(temp1, temp2, temp)

```

### 3 Remove Batch Effect Using All Samples

Figure 1 shows the hierarchical clustering for all the samples and all the probesets.

```

> load("QCResult.RData")
> hc <- hclust(distanceMatrix(AllData, "euclidean"), "ward")
> branches <- cutree(hc, k=4)
> pp <- c(x.qc1@percent.present, x.qc2@percent.present)
> names(pp) <- sub(".present", "", names(pp))
> all(names(pp) ==rownames(si))

[1] TRUE

```

```

> sum(names(pp) ==rownames(si))
[1] 96

> Infor<-cbind(Primary=as.vector(si[, "Primary"]), dateFactor = as.character(si[, "zed"]),
+                 dateFactor2 = as.vector(si[, "dateFactor2"]), branches, pp)
> si <- cbind(si, branches)
> gd <- groupedData(pp ~ 1/branches, data=as.data.frame(Infor), order.groups=FALSE)
> gsummary(gd)

  Primary dateFactor dateFactor2 branches          pp
1  Tumor  2009-02-24    2009-03      1 37.4833104709648
2 Normal  2009-02-26    2009-02      2 42.1380887059899
3  Tumor  2009-04-09    2009-04      3 19.3488797439415
4 Normal  2009-04-07    2009-04      4 25.1742112482853

>
>
```

Figure 2 shows the box plot of the percent present calls for the four clusters in the Figure 5.  
We remove the two big branches effect that are caused by quality.

```

> branchesCorrect <- cutree(hc, k=2)
> table(branchesCorrect, branches)

  branches
branchesCorrect 1 2 3 4
           1 43 25 0 0
           2 0 0 16 12

> cla <- as.factor(branchesCorrect)
> covars <- data.frame(mainBatch=cla)
> mlm <- MultiLinearModel(Y ~ mainBatch, covars, AllData)
> debatch <- AllData
> mm <- matrixMean(AllData)
> debatch <- sweep(AllData - t(mlm@predictions), 1, mm, "+")
> AllDataDebatch<-debatch
> Tumorsi <- si[which(si[, "Primary"] == "Tumor" & si[, "SepID"] != "Normal 5"),]
> #Tumorsi <- si[which(si[, "Primary"] == "Tumor" ),]
> TumorSampleDataDebatch<-debatch[, match(as.vector(Tumorsi[, "SepID"]), colnames(debatch))]
> identical(colnames(TumorSampleDataDebatch), as.vector(Tumorsi[, "SepID"]))]

[1] TRUE
```

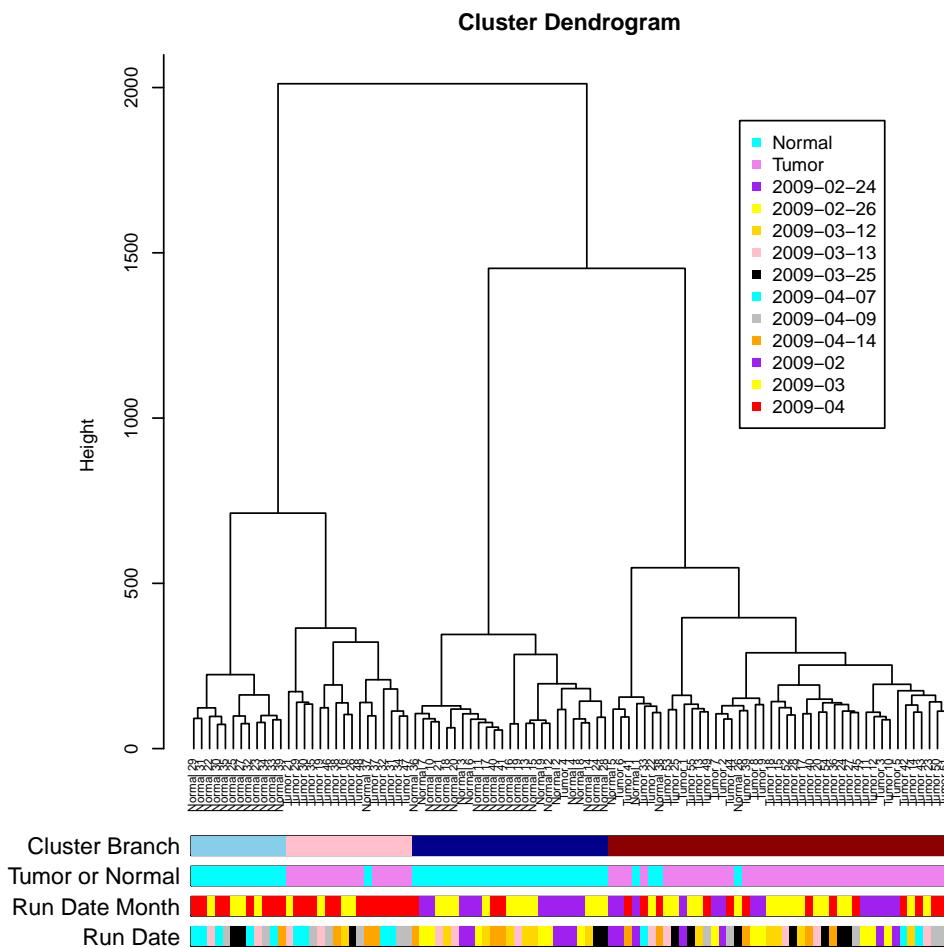


Figure 1: Hierarchical Clustering For All Samples, using euclidean and ward method.

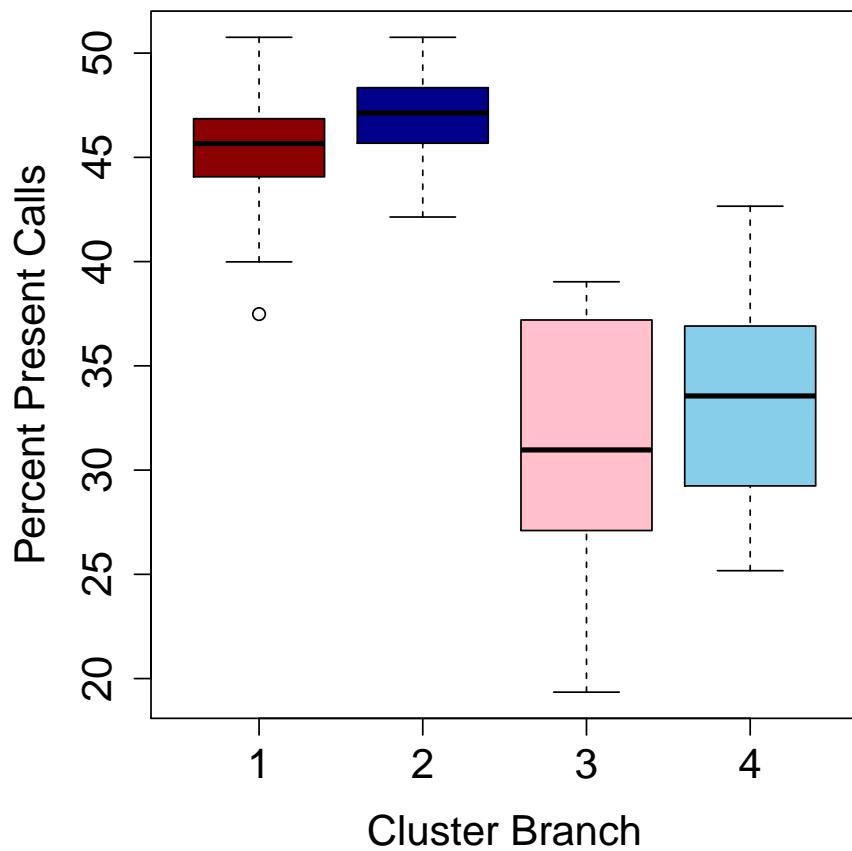


Figure 2: Box plot of Percent Present Calls for Four Clusters.

We then find the paired data.

The samples are either tumor or normal. Some of them are paired tumor and normal samples. There are 55 unique patients. Among them, there are 14 unmatched samples, which are all tumors.

```
> PatientID<-gsub("T","",gsub("N","",si[, "Sample ID"]))
> names(PatientID)<-si[, "Sample ID"]
> as.vector(si[match(names(table(gsub("T","",gsub("N","",si[, "Sample ID"])))))[table(gsub("T",""))]]))

[1] "Tumor 44" "Tumor 42" "Tumor 53" "Tumor 43" "Tumor 47" "Tumor 52" "Tumor 55"
[8] "Tumor 45" "Tumor 49" "Tumor 54" "Tumor 50" "Tumor 51" "Tumor 46" "Tumor 48"

>
```

From PI's information, Normal 5 is actually tumor tissue, and Tumor 4, Tumor 22 are actually normal tissues, we exclude these three patients and those unmatched samples from the analysis.

```
> UnPairedSamples<-c(as.vector(si[match(names(table(gsub("T","",gsub("N","",si[, "Sample ID"]))))])
+           "Normal 5", "Tumor 5", "Tumor 22", "Normal 22", "Tumor 4", "Normal 4"))
> Pairedsi<-si[-match(UnPairedSamples, si[, "SepID"])]
> PairedPatientID<-gsub("Tumor","",gsub("Normal","",Pairedsi[, "SepID"]))
> length(unique(gsub("Tumor","",gsub("Normal","",Pairedsi[, "SepID"]))))

[1] 38

> table(Pairedsi[, "Primary"])

Normal   Tumor
      38      38
```

So we have 38 paired tumor and normal samples.

```
> PairedData<-AllData[, -match(UnPairedSamples, si[, "SepID"])]
> identical(colnames(PairedData), as.vector(Pairedsi[, "SepID"]))

[1] TRUE

> PairedDataDebatch <- debatch[, match(colnames(PairedData), colnames(debatch))]
> identical(colnames(PairedDataDebatch), as.vector(Pairedsi[, "SepID"]))

[1] TRUE
```

We save some information for further use.

```

> probes<-rownames(PairedData)
> acc <- unlist(mget(probes, hgu133plus2ACCNUM))
> chr <- mget(probes, hgu133plus2CHR)
> chr <- unlist(lapply(chr, function(x) x[1]))
> gene <- unlist(mget(probes, hgu133plus2GENENAME))
> sym <- unlist(mget(probes, hgu133plus2SYMBOL))
> uni <- mget(probes, hgu133plus2UNIGENE)
> uni <- unlist(lapply(uni, function(x) x[1]))
> annot <- data.frame(GenBank=acc, Symbol=sym, UniGene=uni,
+                         Chrom=chr, Description=gene)
> identical(rownames(annot), rownames(PairedData))

[1] TRUE

> save(AllDataDebatch, si, TumorSampleDataDebatch, Tumorsi, PairedDataDebatch, Pairedsi, annot,
+       file = "Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Edit-NewID-Deba"
>

```

## 4 Paired T Test Comparing Tumor vs Normal Within Paired Samples

We apply paired t test to check the difference between tumor vs normal effect within paired samples.

```

>     identical(PairedPatientID[which(Pairedsi[, "Primary"] == "Tumor")],
+                 PairedPatientID[which(Pairedsi[, "Primary"] == "Normal")])

[1] TRUE

>     Paired.t.testp <- rep(1, dim(PairedDataDebatch)[1])
>     Paired.t.testt <- rep(1, dim(PairedDataDebatch)[1])
>     for(i1 in 1:length(Paired.t.testp))
+     {
+     Paired.t.test <- t.test(PairedDataDebatch[i1, which(Pairedsi[, "Primary"] == "Tumor")],
+                             PairedDataDebatch[i1, which(Pairedsi[, "Primary"] != "Tumor")], pair=T)
+     Paired.t.testp[i1]<-Paired.t.test$p.value
+     Paired.t.testt[i1]<-Paired.t.test$statistic
+
+   }
>
>
>

> action.bum <- Bum(Paired.t.testp)
>

```

FDRs	Significant Probesets	Corresponding p-value
1 5e-10	1017	2.37e-11
2 1e-09	1157	5.42e-11
3 5e-09	1575	3.70e-10
4 1e-08	1761	8.46e-10
5 5e-08	2306	5.77e-09
6 1e-02	21267	1.22e-02
7 5e-02	30459	8.57e-02

```
> TumorNormalMean<-apply(PairedDataDebatch, 1, function(x) {tapply(x,list(Pairedsi[,"Primary"]))})
> TumorNormalFold<-sign((TumorNormalMean["Tumor",]-TumorNormalMean["Normal",]))*2^(abs(TumorNormalMean["Tumor",]-TumorNormalMean["Normal",]))
> combined <- data.frame(Probe = rownames(PairedData), tstat=Paired.t.testt,
+   pvalue=Paired.t.testp, PairedDataTumor=TumorNormalMean["Tumor",],
+   PairedDataNormal=TumorNormalMean["Normal",], PairedDataFoldChange=TumorNormalFold ,
+   annot)
> write.csv(combined, file="Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Edit-NewID.csv")
>
```

## 5 Robust Clustering Using Tumor Samples

In this section, we generate the robust hierarchical clustering for tumor samples and the probesets are selected for different comparing tumor vs normal in the paired data. We do this 1000 times. Each time, we randomly select n (an uniform number from 500 to 2000) probes from the probesets that are selected at the FDR level of 0.01 (around 20000 probes in total), and randomly select 80\$Record each time whether any two sample are grouped together if cut tree with 3 or 4 groups. Define the consensus matrix as the average time that any two samples are grouped together in each clustering.

```
> ##### define data frame
> DataSource <- TumorSampleDatadebatch[selectSignificant(action.bum, 0.01, by='FDR'),]
> dim(DataSource)

[1] 21267      53

> totalreppnum <- 1000
> ConsensusMatrix3 <- matrix(0,dim(TumorSampleDatadebatch)[2], dim(TumorSampleDatadebatch)[2],
> rownames(ConsensusMatrix3) <- colnames(TumorSampleDatadebatch)
> colnames(ConsensusMatrix3) <- colnames(TumorSampleDatadebatch)
> ConsensusMatrix4 <- matrix(0,dim(TumorSampleDatadebatch)[2], dim(TumorSampleDatadebatch)[2],
> rownames(ConsensusMatrix4) <- colnames(TumorSampleDatadebatch)
> colnames(ConsensusMatrix4) <- colnames(TumorSampleDatadebatch)
> SampleMatrix <- matrix(0,dim(TumorSampleDatadebatch)[2], dim(TumorSampleDatadebatch)[2], byrow=TRUE)
```

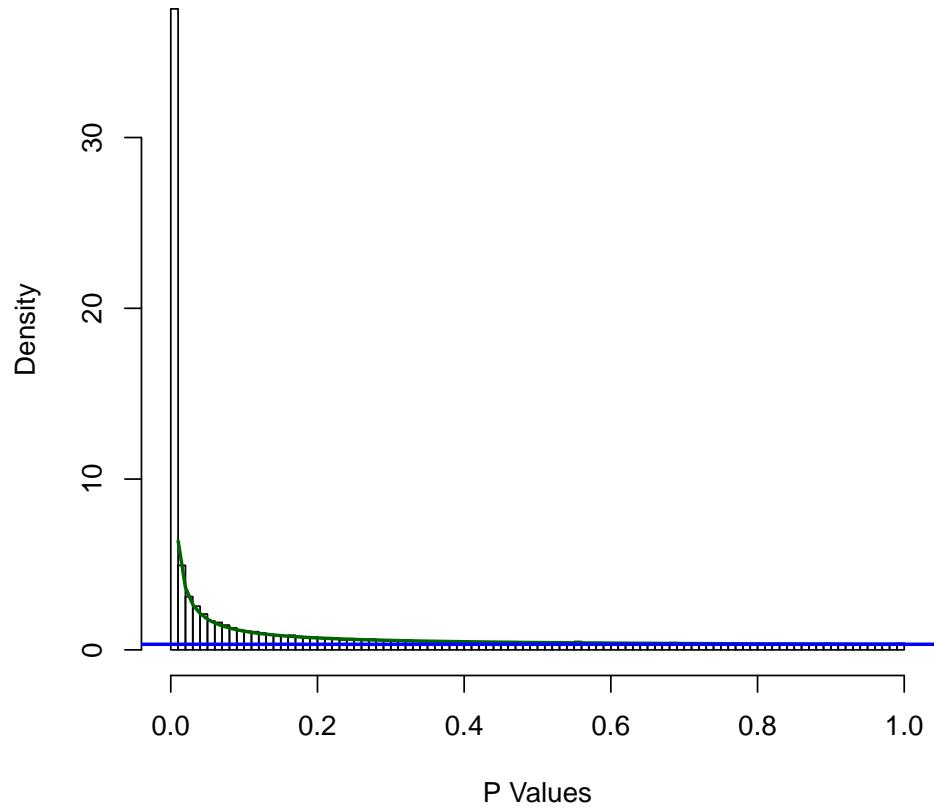


Figure 3: Results of the BUM analysis for difference between tumor vs normal using paired t test within pair samples. There is a peak at the left side. Some probesets are associated with the gene expression with some appropriate FDR level.

```

> rownames(SampleMatrix) <- colnames(TumorSampleData$batch)
> colnames(SampleMatrix) <- colnames(TumorSampleData$batch)
> for(repnum in 1:totalrepnum)
+ {
+   selectlength <- sample(c(500:2000), 1)
+   selectcollength <- floor(0.8*dim(DataSource)[2])
+   selectrow <- sample(c(1:dim(DataSource)[1]), selectlength)
+   selectcol <- sample(c(1:dim(DataSource)[2]), selectcollength)
+   selectuseddata <- DataSource[selectrow, selectcol]
+   SampleMatrix[selectcol, selectcol] <- SampleMatrix[selectcol, selectcol] + 1
+
+   hcTumorSamplesim <- hclust(distanceMatrix(selectuseddata, "euclidean"), "ward")
+   hcacsim3 <- cutree(hcTumorSamplesim, k=3)
+   hcacsim4 <- cutree(hcTumorSamplesim, k=4)
+   usedmatrix3 <- as.matrix(hcacsim3, col=1) %*% t(as.matrix(1/hcacsim3)) == 1
+   usedmatrix3[which(usedmatrix3)] <- 1
+   ConsensusMatrix3[selectcol, selectcol] <- ConsensusMatrix3[selectcol, selectcol]
+   usedmatrix4 <- as.matrix(hcacsim4, col=1) %*% t(as.matrix(1/hcacsim4)) == 1
+   usedmatrix4[which(usedmatrix4)] <- 1
+   ConsensusMatrix4[selectcol, selectcol] <- ConsensusMatrix4[selectcol, selectcol]
+ }
> Consensus3 <- ConsensusMatrix3/SampleMatrix
> Consensus4 <- ConsensusMatrix4/SampleMatrix
> hcTumorSample3 <- hclust(distanceMatrix((1-Consensus3), "euclidean"), "ward")
> hcac3 <- factor(cutree(hcTumorSample3, k=3))
> SeqLMR3<-rep("Group 2", length(hcac3))
> SeqLMR3[hcac3==unique(hcac3[order.dendrogram(as.dendrogram(hcTumorSample3))])][2]]<-"Group 3"
> SeqLMR3[hcac3==unique(hcac3[order.dendrogram(as.dendrogram(hcTumorSample3))])][3]]<-"Group 1"
> Infor<-cbind(Histology=as.vector(Tumorsi[, "Path.Report.Diagnosis"]), Seq = SeqLMR3)
>
>
> OutputInfor <- cbind(Tumorsi, Seq = SeqLMR3)[colInd,]
> write.csv(OutputInfor, file="Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Ed"

```

## 6 ANOVA Comparing the Expression Levels of Probesets among Three Major Splits

We would like to check which probesets show significant different expression levels among these three splits.

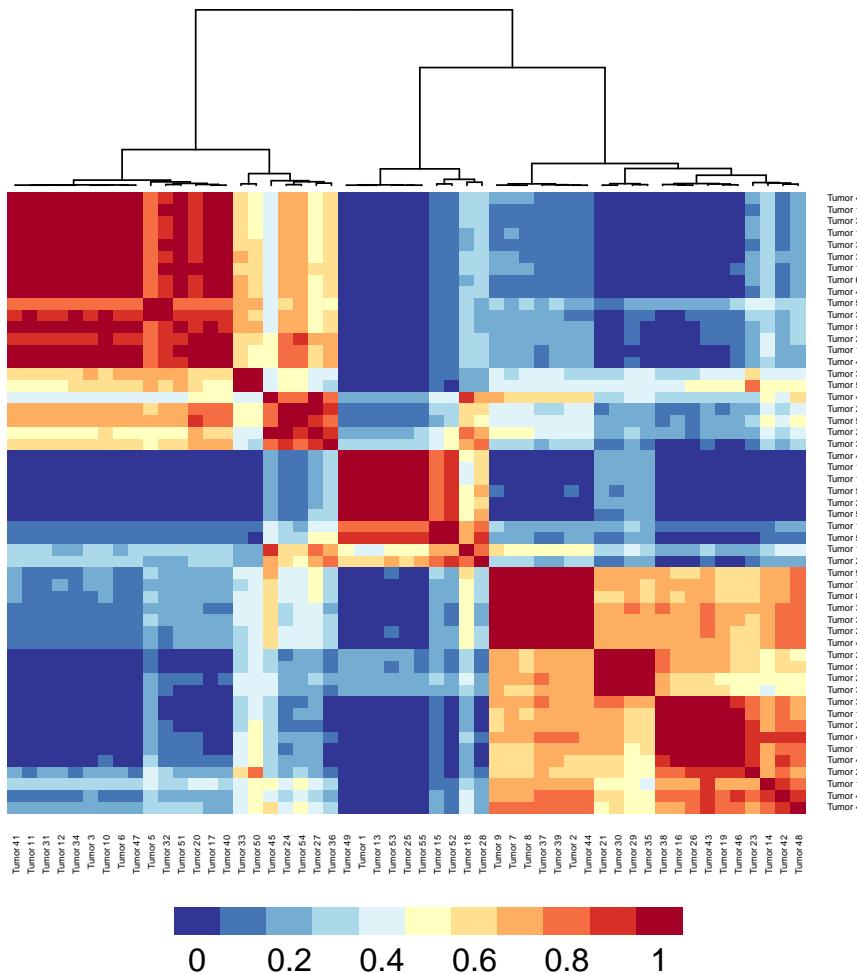


Figure 4: Robust Consensus For tumor samples and the probesets are selected as described in Method section, separation into three groups, using euclidean and ward method.

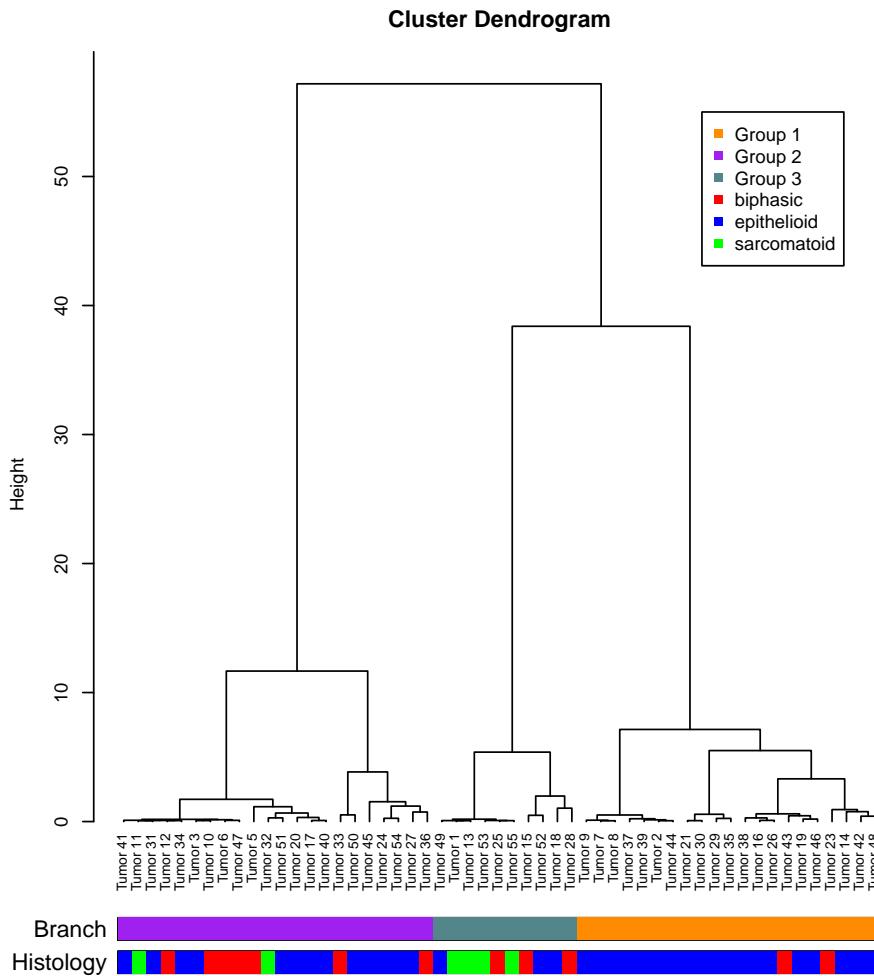


Figure 5: Robust Hierarchical Clustering For tumor samples and the probesets are selected as described in Method section, separation into three groups, using euclidean and ward method.

ANOVA is applied on a probe-by-probe basis using `MultiLinearModel`. The F-statistics and the p-values associated with the F-statistics comparing the linear model to the null model are used to evaluate the expression levels of each probeset among the three splits.

```
> SeqLMR <- SeqLMR3
> ##### check the three branch information
> table(SeqLMR)

SeqLMR
Group 1 Group 2 Group 3
    21      22      10

> ##### fit in ANOVA model
> covars <- data.frame(SeqLMR=SeqLMR)
> ThreeBranchResult <- MultiLinearModel(Y ~ SeqLMR, covars, TumorSampleData$batch)
```

Because of the multiple testing involved in this approach, the individual p-values are not particularly meaningful. However, when we look across the entire set of tests, the distribution of the p-values (under the null hypothesis that no mRNAs provide useful information) should be uniform. If, on the other hand, some mRNAs provide useful information about predicting the response, we would expect an overabundance of small p-values. We can capture this situation by modeling the distribution of the p-values with a Beta-uniform Mixture (BUM). To identify significantly differentially expressed genes (associated with major split effect), we choose a cutoff for the single test p-values by controlling the false discovery rate (FDR), which is defined as the percentage of genes called significant that are expected to turn out false.

```
> ThreeBranchResult.bum <- Bum(ThreeBranchResult$p.values)
```

The following includes all the number of probesets selected at the different FDR levels comparing the probeset expression levels among all three major splits.

FDRs	Significant Probesets	Corresponding p-value
1 5e-05	1628	7.43e-06
2 1e-04	2301	1.92e-05
3 5e-04	4706	1.75e-04
4 1e-03	6306	4.54e-04
5 5e-03	12358	4.14e-03
6 1e-02	16454	1.07e-02
7 5e-02	30757	9.78e-02

We define the left group as the Group 2, corresponding to most of the biphasic, the right group as the Group 1, corresponding to epithelioid, and the middle group as the Group 3, corresponding to most of the sarcomatoid.

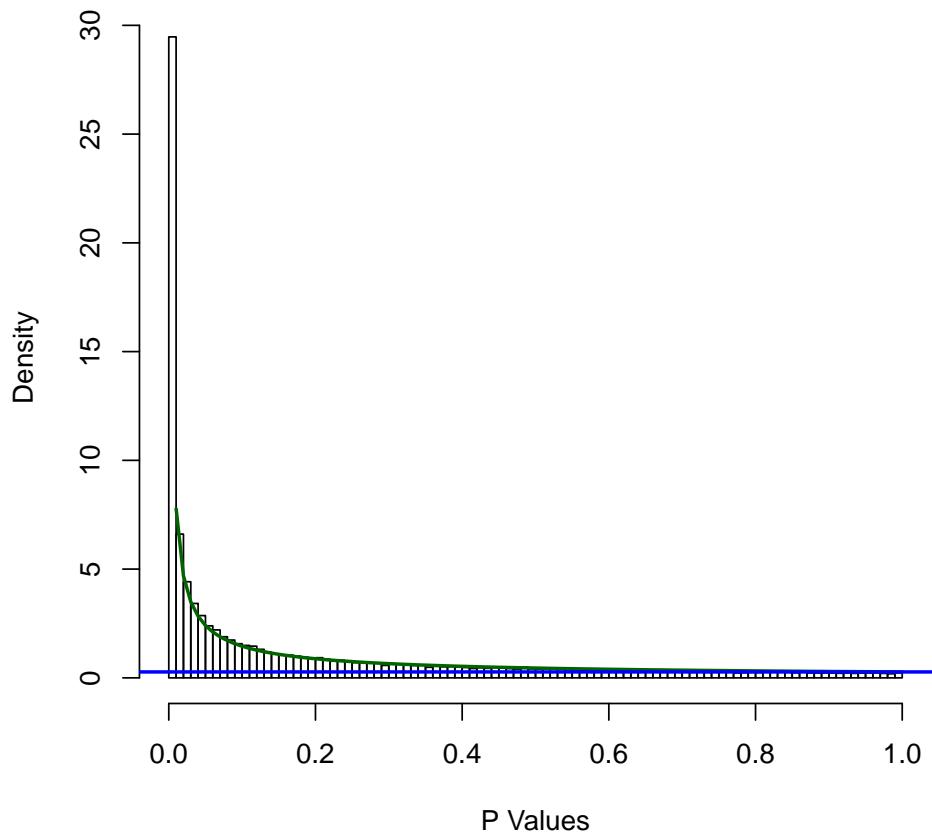


Figure 6: Results of a BUM analysis of the ANOVA comparing linear model to the null model to define the major three split of Figure5. The curve tends to be a peak at the left end, suggesting there is some differential expression present at an appropriate FDR level.

```
> ##### define the three groups created in the previous report
>
> ThreeGroupTvsNTestProbe<-SeqLMR
> Tumorsi <- cbind(Tumorsi,ThreeGroupTvsNTestProbe )
```

We save the result.

```
> ##### calculate the mean for each group
> ThreeBranchMeanTvsNTestProbe<-t(apply(TumorSampleDatadebatches, 1, function(x) {tapply(x, as.factor,
> ##### save the result
>
> ThreeGroupFDR <- 0.0001
> ReportResult<-data.frame(Probes=rownames(TumorSampleDatadebatches), FStatistics= ThreeBranchResults$FStatistics,
+ pValues=ThreeBranchResults$p.values, ThreeBranchMeanTvsNTestProbe,
+ ANOVATHreeGroup =selectSignificant(ThreeBranchResults$bum, ThreeGroupFDR, by='FDR'))
>
```

## 6.1 Define Probesets Distribution In Three Subgroups

Then, we need to define how the probes distributed in three groups.

We first define the order for the three groups.

```
> ##### define the order of the probesets
> MeanOrder <- apply(ThreeBranchMeanTvsNTestProbe, 1, function(x){paste(order(x, decreasing=TRUE),
> table(MeanOrder)
```

MeanOrder

1-2-3	1-3-2	2-1-3	2-3-1	3-1-2	3-2-1
14145	14252	5726	4442	7377	8733

```
> table(MeanOrder[which(ReportResult[, "ANOVATHreeGroup"])])
```

1-2-3	1-3-2	2-1-3	2-3-1	3-1-2	3-2-1
382	108	342	283	244	942

```
> ReportResult<- cbind(ReportResult, MeanOrder)
> ##### Tukey HSD is applied
>
> TUKEYResult <- matrix(0, dim(TumorSampleDatadebatches)[1], 6, byrow=TRUE)
> colnames(TUKEYResult) <- c("Group2vs1Diff", "Group2vs1PValue",
+ "Group3vs1Diff", "Group3vs1PValue",
+ "Group3vs2Diff", "Group3vs2PValue")
> for(i1 in 1:dim(TUKEYResult)[1])
+ {
```

```

+      tempdata <- as.vector(t(TumorSampleDatadebatches[i1,]))
+      temp <- TukeyHSD(aov(tempdata~as.factor(ThreeGroupTvsNTTestProbe)))
+      TUKEYResult[i1, c("Group2vs1Diff", "Group2vs1PValue")] <- temp$as.factor(ThreeGroupT
+      TUKEYResult[i1, c("Group3vs1Diff", "Group3vs1PValue")] <- temp$as.factor(ThreeGroupT
+      TUKEYResult[i1, c("Group3vs2Diff", "Group3vs2PValue")] <- temp$as.factor(ThreeGroupT
+ }
> ReportResult <- cbind(ReportResult, TUKEYResult)
>
>
>
```

We change the categories as whether the three tests are significant at the 0.05 p value cutoff.

```

> TUKEYResultIndi <- (TUKEYResult[, c("Group2vs1PValue",
+                                         "Group3vs1PValue",
+                                         "Group3vs2PValue")] <= 0.05)*1
> IndicatorCombine <- cbind(MeanOrder, TUKEYResultIndi)
> IndicatorCombineFinal <- apply(IndicatorCombine, 1, function(x){paste(x, collapse = "-")})
> ##### check the categories for ANOVA at FDR 0.0001 level
> table(IndicatorCombineFinal[which(ReportResult[, "ANOVAThreeGroup"])])

1-2-3-0-1-1 1-2-3-1-1-0 1-2-3-1-1-1 1-3-2-1-0-0 1-3-2-1-0-1 1-3-2-1-1-0 2-1-3-0-1-1
        43          244         95          8          24         76          53
2-1-3-1-0-1 2-1-3-1-1-1 2-3-1-1-0-0 2-3-1-1-0-1 2-3-1-1-1-0 2-3-1-1-1-1 3-1-2-0-1-1
        192          97          16         95         167          5         187
3-1-2-1-0-1 3-1-2-1-1-1 3-2-1-0-1-1 3-2-1-1-1-0 3-2-1-1-1-1
        24          33         295         373         274

> ReportResult <- cbind(ReportResult, IndicatorCombineFinal)
>
```

We then generate the heatmap using the selected probesets (FDR levle of 1e-04). The clustering for samples does not change.

```

> ##### hierarchical clustering is applied on both directions
> alpha<-ThreeGroupFDR
> tempMatrix<-TumorSampleDatadebatches[selectSignificant(ThreeBranchResult.bum, alpha, by='FDR'),]
> agencl <- hclust(distanceMatrix(t(tempMatrix)), "pearson"), "ward")
> asamcl <- hcTumorSample3
> ##### truncate the standardized gene expression for explore purpose
>
> ulim <- 2
> temp <- t(scale(t(tempMatrix)))
> temp[temp > ulim] <- ulim
```

```

> temp[temp < -ulim] <- -ulim
> tempIndicator <- IndicatorCombineFinal$selectSignificant(ThreeBranchResult.bum, alpha, by='FD')
> dim(tempMatrix)

[1] 2301   53

>

```

We check how the groups corresponding to the four big groups defined by the clustering.

```

> x <- cutree(agenc1, k=4)
> table(x)

x
 1   2   3   4
522 922 354 503

> #####      get the sequence of the cutting tree from left to right
> unique(x[order.dendrogram(as.dendrogram(agenc1))])

[1] 1 2 3 4

> #####      association with the indicator
>

> GeneInThreeGroupANOVAProbe <- x
> GeneInThreeGroupANOVAProbe[x==unique(x[order.dendrogram(as.dendrogram(agenc1))])[1]]<-"Probegroup1"
> GeneInThreeGroupANOVAProbe[x==unique(x[order.dendrogram(as.dendrogram(agenc1))])[2]]<-"Probegroup2"
> GeneInThreeGroupANOVAProbe[x==unique(x[order.dendrogram(as.dendrogram(agenc1))])[3]]<-"Probegroup3"
> GeneInThreeGroupANOVAProbe[x==unique(x[order.dendrogram(as.dendrogram(agenc1))])[4]]<-"Probegroup4"
> table(tempIndicator, GeneInThreeGroupANOVAProbe)

  GeneInThreeGroupANOVAProbe
tempIndicator Probegroup1 Probegroup2 Probegroup3 Probegroup4
  1-2-3-0-1-1          0          0         11        32
  1-2-3-1-1-0          0          0          0       244
  1-2-3-1-1-1          0          0          1        94
  1-3-2-1-0-0          0          0          0         8
  1-3-2-1-0-1          1          0          0        23
  1-3-2-1-1-0          0          0          0       76
  2-1-3-0-1-1          0          0         48         5
  2-1-3-1-0-1          0         19        173         0
  2-1-3-1-1-1          0          0         97         0
  2-3-1-1-0-0          0         16          0         0
  2-3-1-1-0-1          0         72         23         0

```

2-3-1-1-1-0	0	166	1	0
2-3-1-1-1-1	0	5	0	0
3-1-2-0-1-1	184	2	0	1
3-1-2-1-0-1	9	0	0	15
3-1-2-1-1-1	27	1	0	5
3-2-1-0-1-1	238	57	0	0
3-2-1-1-1-0	4	369	0	0
3-2-1-1-1-1	59	215	0	0

&gt;

From the distribution, there are four groups. From lowest to highest, they represent:

1. High in Group 3

- mostly, mean value of group 3 is higher than the other two
- small part of the probesets, mean value of group 3 is higher than one of the other groups
- at least one test is significant that Group 3 is higher than at least one of the other two

2. Minerhigh in Group 3

- almost half of the probeset, mean value of group 3 is higher than one of the other two groups
- some probesets, mean value of group 3 is higher than the other two
- at least one test is significant that Group 3 is higher than at least one of the other two.

3. High in Group 2

- mostly, mean value of group 2 is higher than the other two
- only one probeset, mean value of group 2 is higher than 1, lower than group 3
- at least one test is significant that Group 2 is higher than one of the other two

4. High in Group 1

- mostly, mean value of group 1 is higher than the other two
- small part of the probesets, mean value of group 1 is higher than one of the other groups
- at least one test is significant that Group 1 is higher than at least one of the other two

In this way, we define them as the representation.

```
> GeneInThreeGroupANOVAProbe <- x
> GeneInThreeGroupANOVAProbe [x==unique(x[order.dendrogram(as.dendrogram(agenc1))])][4]]<-"High"
> GeneInThreeGroupANOVAProbe [x==unique(x[order.dendrogram(as.dendrogram(agenc1))])][1]]<-"High"
> GeneInThreeGroupANOVAProbe [x==unique(x[order.dendrogram(as.dendrogram(agenc1))])][3]]<-"High"
> GeneInThreeGroupANOVAProbe [x==unique(x[order.dendrogram(as.dendrogram(agenc1))])][2]]<-"Miner"
> table(tempIndicator, GeneInThreeGroupANOVAProbe)
```

```

GeneInThreeGroupANOVAProbe
tempIndicator High Group 1 High Group 2 High Group 3 Miner High Group 3
1-2-3-0-1-1      32       11       0       0
1-2-3-1-1-0     244        0       0       0
1-2-3-1-1-1     94        1       0       0
1-3-2-1-0-0      8        0       0       0
1-3-2-1-0-1     23        0       1       0
1-3-2-1-1-0     76        0       0       0
2-1-3-0-1-1      5        48       0       0
2-1-3-1-0-1     0        173       0      19
2-1-3-1-1-1     0        97       0       0
2-3-1-1-0-0     0        0       0      16
2-3-1-1-0-1     0        23       0      72
2-3-1-1-1-0     0        1       0     166
2-3-1-1-1-1     0        0       0       5
3-1-2-0-1-1      1        0     184       2
3-1-2-1-0-1     15        0       9       0
3-1-2-1-1-1     5        0      27       1
3-2-1-0-1-1     0        0     238      57
3-2-1-1-1-0     0        0       4     369
3-2-1-1-1-1     0        0       59    215

>
>
>

null device
 1

> ##### output the result
>
>
> tempresult <- ReportResult[selectSignificant(ThreeBranchResult.bum, alpha, by='FDR'),]
> tempresult <- cbind(tempresult, GeneInThreeGroupANOVAProbe)
> ReportResult <- cbind(ReportResult, annot)
> write.csv(ReportResult, file="Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Edit-NewID.csv")
> tempresult <- cbind(tempresult, annot[selectSignificant(ThreeBranchResult.bum, alpha, by='FDR'),])
> write.csv(tempresult, file="Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Edit-NewID.csv")
>
```

We check the assiciation with other clinical variables.

```

> ##### histology
> table(Group = Tumorsi[, "ThreeGroupTvsNTestProbe"], histology = as.vector(Tumorsi[, "Path.Repor

```

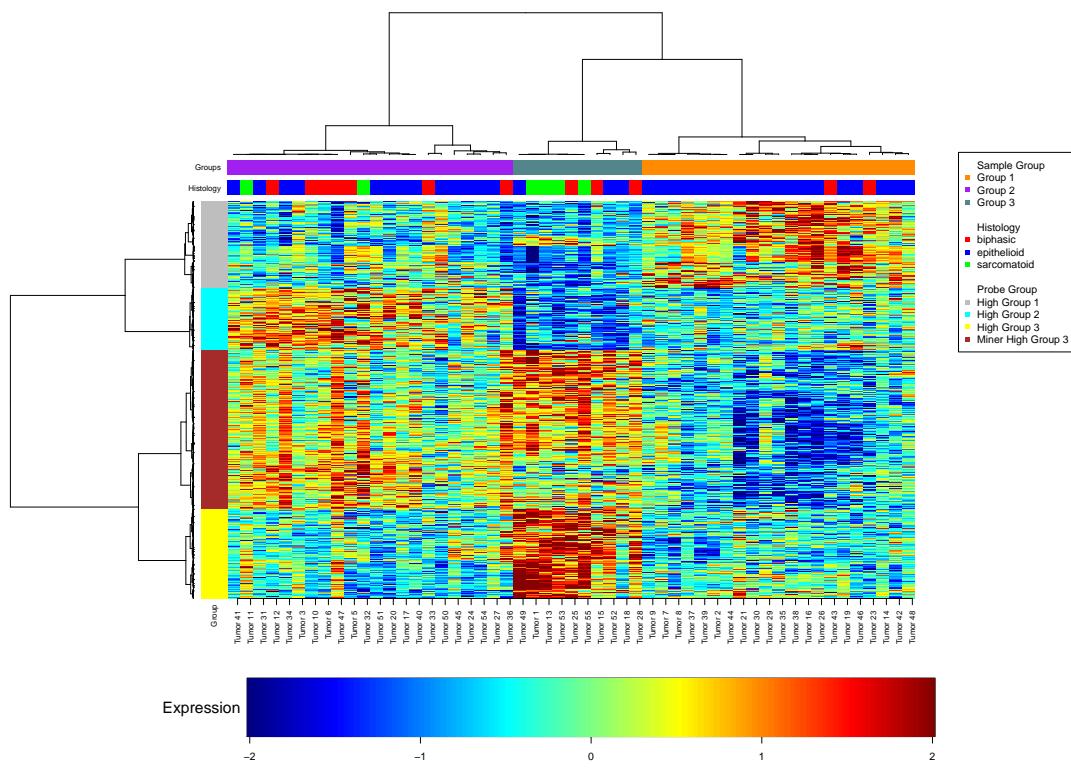


Figure 7: Heatmap for tumor samples, probesets comparing three major splits using ANOVA at the FDR level of 1e-04. For display purposes, we truncate the standardized gene expression values at  $\pm 2$  standard deviations. The sample clustering is the robust clusters defined by our algorithm.

```

histology
Group    biphasic epithelioid sarcomatoid
  Group 1      2       19      0
  Group 2      7       13      2
  Group 3      3        3      4

> fisher.test(table(Group = Tumorsi[, "ThreeGroupTvsNTestProbe"], histology = as.vector(Tumorsi$histology))

Fisher's Exact Test for Count Data

data:
p-value = 0.002433
alternative hypothesis: two.sided

> ##### with the samples that will be removed from analysis
>
> link <- c( "Tumor 6", "Tumor 17", "Tumor 44", "Tumor 45", "Tumor 54")
> table(Tumorsi[match(link, Tumorsi[, "SepID"]), "ThreeGroupTvsNTestProbe"])

Group 1 Group 2 Group 3
      1      4      0

>
```

## 7 Overall Survival Analysis

### 7.1 Load in Clinical Information

We then load in the survival data file.

```

> SurvivalData<-read.xls("Meso-Clinical-Data-June20-2012.xlsx", sheet = 2)
> SurvivalData1<-read.xls("Meso-Clinical-Data-June20-2012.xlsx", sheet = 1)
> identical(as.vector(SurvivalData[, "Spore.ID"]), as.vector(SurvivalData1[, "Spore.ID"]))

[1] TRUE

> SurvivalData <- cbind(SurvivalData, LFUS= SurvivalData1[, "LFUS"], Gender = SurvivalData1[, "Gender"])
```

We clean the tumor samples for further analysis. In the survival analysis, we remove this sample. Samples Tumor 6, Tumor 17, Tumor 44, Tumor 45, Tumor 54 have their death seems surgery related, so we also remove them from further analysis.

```

> RemoveID <- c( "Tumor 6", "Tumor 17", "Tumor 44", "Tumor 45", "Tumor 54")
> Removelink <- match(RemoveID, Tumorsi[, "SepID"])
> TumorSurvivalsi <- Tumorsi[-Removelink,]
> dim(TumorSurvivalsi)
```

```
[1] 48 17
```

We would like to use the first “Date.Death” column as the end time point of overall survival analysis. There is 4 patient has empty Date of death. We will use the last follow up date to replace.

```
> ##### look at the patient with no date of death
>
> SurvivalData[SurvivalData[, "Date.Death"]=="", "vital.status"]

[1] A A A A
Levels: A D

> ##### create vector to replace the patient
>
> temp <- ifelse(as.character(SurvivalData[, "Date.Death"])=="",
+                 as.character(gsub(" another tumor", "", SurvivalData[, "LFUS"])),
+                 as.character(SurvivalData[, "Date.Death"])) )
> temp2 <- temp
> temp2[is.na(as.Date(temp))] <- as.character(as.Date(temp[is.na(as.Date(temp))], "%m/%d/%Y"))
> ##### replace the information with the patient using LFUS
>
> SurvivalData<-data.frame(DeathorLFUS=temp2, SurvivalData)
> ##### we calculate the overall survival time as Death or LFUS minus Date Surgery
>
> OverallTime<- (as.Date(as.character(as.vector(SurvivalData[, "DeathorLFUS"]))) - as.Date(as.
> SurvivalData<-data.frame(OverallTime=as.numeric(OverallTime), SurvivalData)
> ##### we cdefine the census of overall survival
>
> table(SurvivalData[, "vital.status"])

A D
4 47
```

```
> cen.status <- ifelse(as.character(SurvivalData[, "vital.status"])=="D",
+                       as.character(1), as.character(0) )
> SurvivalData<-data.frame(VitalStatusCen=cen.status, SurvivalData)
> ##### derive the age at the date of surgery #####
>
> PatientAge<-(as.Date(as.character(as.vector(SurvivalData[, "Date.Surgery"]))) -
+                  as.Date(as.character(as.vector(SurvivalData[, "DOB"]))))/365.25
> ##### summary the age at the date of surgery #####
>
> summary(as.numeric(PatientAge))
```

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
40.38	58.85	63.33	63.36	68.41	81.41

```
> SurvivalData <- data.frame(PatientAge=as.numeric(PatientAge), SurvivalData)
> rm(PatientAge)
> rm(cen.status)
> rm(OverallTime)
>
```

There is one patient Tumor 47, should be NEO-Adj true based on clinical PI. We change it.

We then create another survival data frame which matches to the tumor samples that will be used in further survival analysis.

```
> MatchID<-match(gsub("T", "", gsub("N", "", TumorSurvivals[, "Sample ID"])), gsub("T", "", gsub("N", "", TumorSurvivals[, "Sample ID"])))
> ##### check the sample ID not matched to the survival data
>
> sum(is.na(MatchID))
[1] 0

> ##### create the Survival Data Matrix with the matched tumor samples
>
> TumorSampleSurvivalData<-SurvivalData[MatchID,]
> ##### check the consistency between sample ID
>
> identical(as.vector(gsub("T", "", gsub("N", "", TumorSurvivals[, "Sample ID"]))),
+           as.vector(gsub("T", "", gsub("N", "", TumorSampleSurvivalData[, "Sample ID"]))))
[1] TRUE

> identical(as.vector(Tumors[, "SepID"]), colnames(TumorSampleData$batch))
[1] TRUE

> rm(MatchID)
>
```

We check all the tumor samples.

```
> MatchID<-match(gsub("T", "", gsub("N", "", Tumors[, "Sample ID"])), gsub("T", "", gsub("N", "", Tumors[, "Sample ID"])))
> ##### check the sample ID not matched to the survival data
>
> sum(is.na(MatchID))
[1] 4
```

```

> Tumorsi[is.na(MatchID), "SepID"]

[1] Tumor 6 Tumor 54 Tumor 44 Tumor 45
96 Levels: Normal 1 Normal 10 Normal 11 Normal 12 Normal 13 Normal 14 ... Tumor 9

> ##### create the Survival Data Matrix with the matched tumor samples
>
> TumorSampleAllSurvivalData<-SurvivalData[MatchID,]
> ##### check the consistency between sample ID
>
> #identical(as.vector(gsub("T", "", gsub("N", "", Tumorsi[, "Sample ID"]))),
> #                         as.vector(gsub("T", "", gsub("N", "", TumorSampleAllSurvivalData[, "S
>
>
>
> identical(as.vector(Tumorsi[, "SepID"]), colnames(TumorSampleDatadebatch))

[1] TRUE

> rm(MatchID)
> save(Tumorsi, TumorSampleAllSurvivalData,
+       file="Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Edit-NewID-All-"
> ##### some of the tumor samples we do not have
> ##### clinical information
> table(Group =Tumorsi[, "ThreeGroupTvsNTTestProbe"] , NEO= as.vector(TumorSampleAllSurvivalData

          NEO
Group      No Yes
Group 1 19   1
Group 2 11   8
Group 3  7   3

> fisher.test(table(Group =Tumorsi[, "ThreeGroupTvsNTTestProbe"] , NEO= as.vector(TumorSampleAll

Fisher's Exact Test for Count Data

data:
p-value = 0.01889
alternative hypothesis: two.sided

We save the object for further analysis.

> save(TumorSurvivalsi, TumorSampleSurvivalData,
+       file="Report15-Affymetrix-Analsys-Remove-Batch-AllSample-Simulation-Edit-NewID-Surviv
>
>
```

## 7.2 Overall Survival Analysis For Tumor Samples

There are three subgroups defined among all the tumor samples. We would like to check whether the overall survival is different among the three groups.

```
> ##### find the time and census status
>
> Time.dfs <- as.numeric(TumorSampleSurvivalData[, "OverallTime"])
> cen.status <- as.numeric(as.vector(TumorSampleSurvivalData[, "VitalStatusCen"]))
> ##### fit into KM model
> groupused <- TumorSurvivalsi[,"ThreeGroupTvsNTestProbe"]
> OverallThreeGroup<- survfit(Surv(Time.dfs, cen.status) ~ as.factor(groupused), na.action=na.exclude)
> OverallThreeGroup.res <- survdiff(Surv(Time.dfs , cen.status) ~ as.factor(groupused), na.action=na.exclude)
> Factorlevel<-groupused
> pValue3All <- 1-pchisq(OverallThreeGroup.res$chisq, df=length(levels(factor(Factorlevel)))-1)
> ##### show the KM result comparing the three groups
>
> OverallThreeGroup.res
```

Call:

```
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(groupused),
na.action = na.exclude)
```

	N	Observed	Expected	$(O-E)^2/E$	$(O-E)^2/V$
as.factor(groupused)=Group 1	20	17	22.22	1.225	2.52
as.factor(groupused)=Group 2	18	18	15.92	0.272	0.43
as.factor(groupused)=Group 3	10	9	5.86	1.678	1.97

Chisq= 3.3 on 2 degrees of freedom, p= 0.197

```
> OverallThreeGroup
```

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(groupused),  
na.action = na.exclude)

	records	n.max	n.start	events	median	0.95LCL	0.95UCL
as.factor(groupused)=Group 1	20	20	20	17	1.562	0.810	3.03
as.factor(groupused)=Group 2	18	18	18	18	0.820	0.454	3.03
as.factor(groupused)=Group 3	10	10	10	9	0.571	0.320	NA

```
> summary(coxph(Surv(Time.dfs, cen.status) ~ as.factor(groupused), na.action=na.exclude))
```

Call:

```
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(groupused),
```

```

na.action = na.exclude)

n= 48, number of events= 44

            coef exp(coef) se(coef)      z Pr(>|z|)
as.factor(groupused)Group 2 0.3991    1.4906   0.3399 1.174   0.2403
as.factor(groupused)Group 3 0.7175    2.0493   0.4181 1.716   0.0861 .
---
Signif. codes:  0 *** 0.001 ** 0.01 * 0.05 . 0.1  .
exp(coef) exp(-coef) lower .95 upper .95
as.factor(groupused)Group 2     1.491     0.6709   0.7656   2.902
as.factor(groupused)Group 3     2.049     0.4880   0.9031   4.650

Concordance= 0.596  (se = 0.046 )
Rsquare= 0.063  (max possible= 0.997 )
Likelihood ratio test= 3.13  on 2 df,  p=0.209
Wald test           = 3.19  on 2 df,  p=0.2034
Score (logrank) test = 3.27  on 2 df,  p=0.1947

```

### 7.3 Overall Survival Analysis for Histology

There are three histology groups defined among all the tumor samples. We would like to check whether the overall survival is different among the three histology groups.

```

> ##### fit into KM model
>
> UsedFactor <- TumorSurvivals[, "Path.Report.Diagnosis"]
> OverallHistGroup<- survfit(Surv(Time.dfs, cen.status) ~ as.factor(UsedFactor), na.action=na.exclude)
> OverallHistGroup.res <- survdiff(Surv(Time.dfs , cen.status) ~ as.factor(UsedFactor), na.action=na.exclude)
> Factorlevel<-TumorSurvivals[, "Path.Report.Diagnosis"]
> pValue3All <- 1-pchisq(OverallHistGroup.res$chisq, df=length(levels(factor(Factorlevel)))-1)
> ##### show the KM result comparing the three groups
>
> OverallHistGroup.res

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(UsedFactor),
na.action = na.exclude)

          N Observed Expected (O-E)^2/E (O-E)^2/V
as.factor(UsedFactor)=biphasic    11        11     7.09  2.161367  2.606452

```

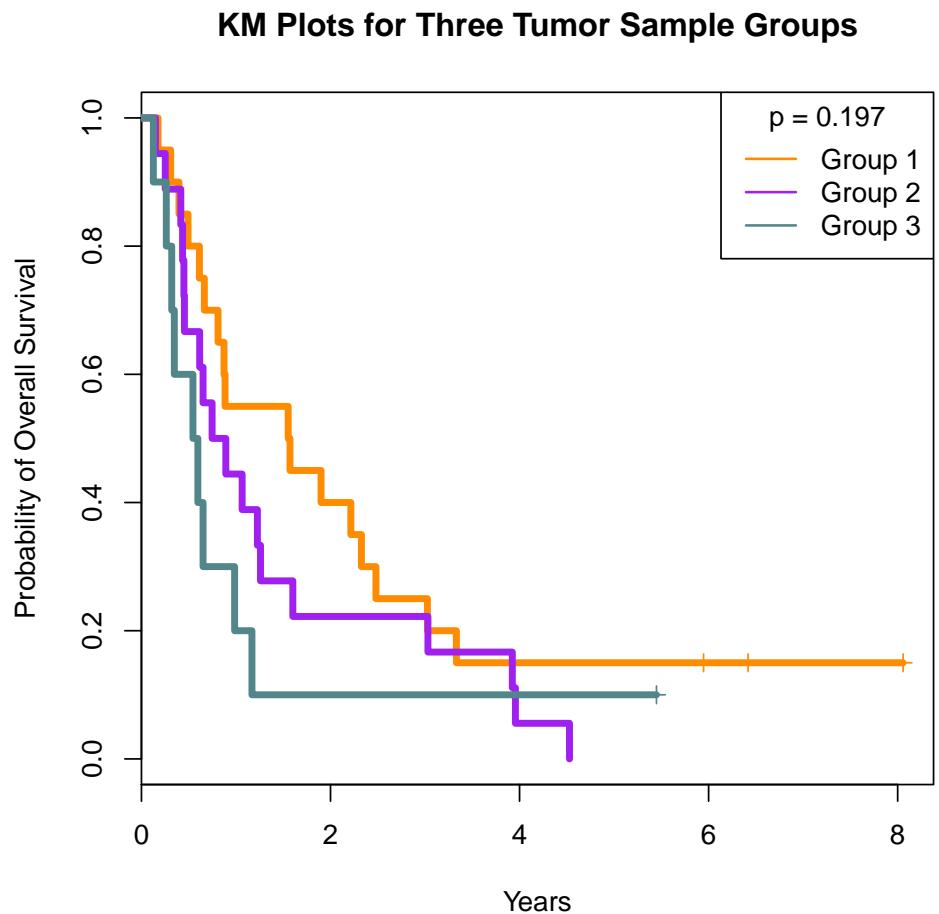


Figure 8: KM for comparing the three groups defined in Figure5, excluding sample Normal 5, Tumor 6, Tumor 17, Tumor 44, Tumor 45, Tumor 54.

```

as.factor(UsedFactor)=epithelioid 31      28      31.86  0.468750  1.720596
as.factor(UsedFactor)=sarcomatoid   6      5       5.05   0.000472  0.000541

Chisq= 2.7 on 2 degrees of freedom, p= 0.264

> OverallHistGroup

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(UsedFactor),
  na.action = na.exclude)

          records n.max n.start events median 0.95LCL 0.95UCL
as.factor(UsedFactor)=biphasic      11     11      11      11  0.449   0.263    NA
as.factor(UsedFactor)=epithelioid    31     31      31      28  1.065   0.665   2.33
as.factor(UsedFactor)=sarcomatoid    6      6       6       5  0.634   0.597    NA

> summary(coxph(Surv(Time.dfs, cen.status) ~ as.factor(UsedFactor), na.action=na.exclude))

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(UsedFactor),
  na.action = na.exclude)

n= 48, number of events= 44

            coef exp(coef) se(coef)      z Pr(>|z|)
as.factor(UsedFactor)epithelioid -0.5753    0.5625  0.3584 -1.605    0.108
as.factor(UsedFactor)sarcomatoid -0.4522    0.6362  0.5439 -0.831    0.406

            exp(coef) exp(-coef) lower .95 upper .95
as.factor(UsedFactor)epithelioid    0.5625     1.778    0.2786   1.136
as.factor(UsedFactor)sarcomatoid    0.6362     1.572    0.2191   1.847

Concordance= 0.578 (se = 0.04 )
Rsquare= 0.048 (max possible= 0.997 )
Likelihood ratio test= 2.36 on 2 df,  p=0.3076
Wald test           = 2.58 on 2 df,  p=0.2749
Score (logrank) test = 2.65 on 2 df,  p=0.2658

```

## 7.4 Overall Survival Analysis for Epithelial in Group 2 and 3

There are only 2 epithelial samples in gouthree. We would like to check whether the overall survival is different among the other two groups groups.

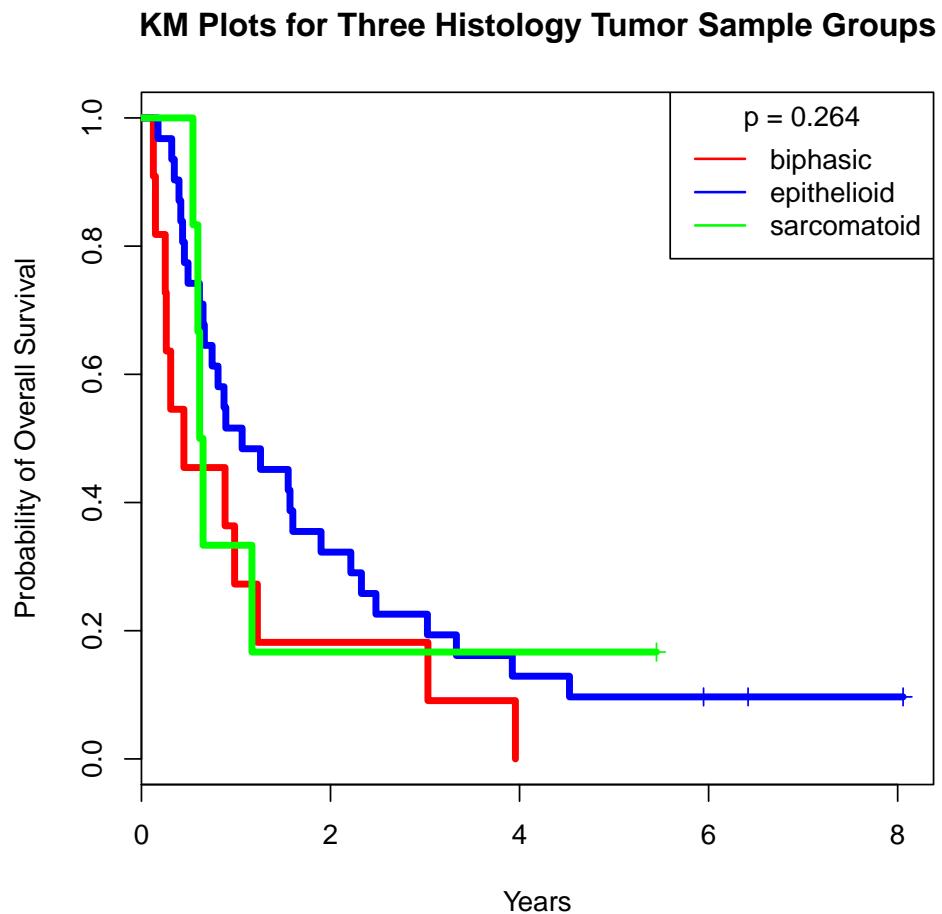


Figure 9: KM for comparing the three groups defined by histology, excluding samples Normal 5, Tumor 6, Tumor 17, Tumor 44, Tumor 45, Tumor 54.

```

> ##### find the index of epithelioid
>
> EpiIndex <- which(TumorSurvivals[, "Path.Report.Diagnosis"] == "epithelioid" & groupused != "Normal")
> length(EpiIndex)

[1] 28

> ##### find the time and census status
>
> Time.dfs <- as.numeric(TumorSampleSurvivalData[EpiIndex, "OverallTime"])
> cen.status <- as.numeric(as.vector(TumorSampleSurvivalData[EpiIndex, "VitalStatusCen"]))
> ##### fit into KM model
>
> OverallThreeGroupEpi<- survfit(Surv(Time.dfs, cen.status) ~ as.factor(as.vector(groupused[EpiIndex])))
> OverallThreeGroup.resEpi <- survdiff(Surv(Time.dfs , cen.status) ~ as.factor(as.vector(groupused[EpiIndex])))
> Factorlevel<-as.vector(groupused[EpiIndex])
> table(as.vector(groupused[EpiIndex]))

Group 1 Group 2
    18      10

> pValue3All <- 1-pchisq(OverallThreeGroup.resEpi$chisq, df=length(levels(factor(Factorlevel))))
> ##### show the KM result comparing the three groups
>
> OverallThreeGroup.resEpi

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(as.vector(groupused[EpiIndex])), na.action = na.exclude)

          N Observed Expected (O-E)^2/E
as.factor(as.vector(groupused[EpiIndex]))=Group 1 18      15     17.49     0.354
as.factor(as.vector(groupused[EpiIndex]))=Group 2 10      10      7.51     0.825
                                         (O-E)^2/V
as.factor(as.vector(groupused[EpiIndex]))=Group 1      1.2
as.factor(as.vector(groupused[EpiIndex]))=Group 2      1.2

Chisq= 1.2 on 1 degrees of freedom, p= 0.273

> OverallThreeGroupEpi

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(as.vector(groupused[EpiIndex])), na.action = na.exclude)

```

```

records n.max n.start events median
as.factor(as.vector(groupused[EpiIndex]))=Group 1      18     18      18      15  1.736
as.factor(as.vector(groupused[EpiIndex]))=Group 2      10     10      10      10  0.979
                                         0.95LCL 0.95UCL
as.factor(as.vector(groupused[EpiIndex]))=Group 1      0.810    3.33
as.factor(as.vector(groupused[EpiIndex]))=Group 2      0.454      NA

> summary(coxph(Surv(Time.dfs, cen.status) ~ as.factor(as.vector(groupused[EpiIndex])), na.action = na.exclude))

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(as.vector(groupused[EpiIndex])), na.action = na.exclude)

n= 28, number of events= 25

            coef exp(coef) se(coef)      z
as.factor(as.vector(groupused[EpiIndex]))Group 2 0.4482   1.5655   0.4124 1.087
                                         Pr(>|z|)
as.factor(as.vector(groupused[EpiIndex]))Group 2  0.277

            exp(coef) exp(-coef) lower .95
as.factor(as.vector(groupused[EpiIndex]))Group 2  1.566    0.6388   0.6976
                                         upper .95
as.factor(as.vector(groupused[EpiIndex]))Group 2  3.513

Concordance= 0.551 (se = 0.055 )
Rsquare= 0.04 (max possible= 0.991 )
Likelihood ratio test= 1.14 on 1 df,  p=0.2857
Wald test           = 1.18 on 1 df,  p=0.2771
Score (logrank) test = 1.2 on 1 df,  p=0.2733

```

## 7.5 Overall Survival Analysis Removing NEO-adj Treated Patients

We remove the Neo-adj patients and compare the three groups.

```

> ##### find the index of epithelioid
>
> EpiIndex <- which(TumorSampleSurvivalData[, "Neoadjuvant.chemo"] == "No")
> length(EpiIndex)

[1] 37

> ##### find the time and census status
>
```

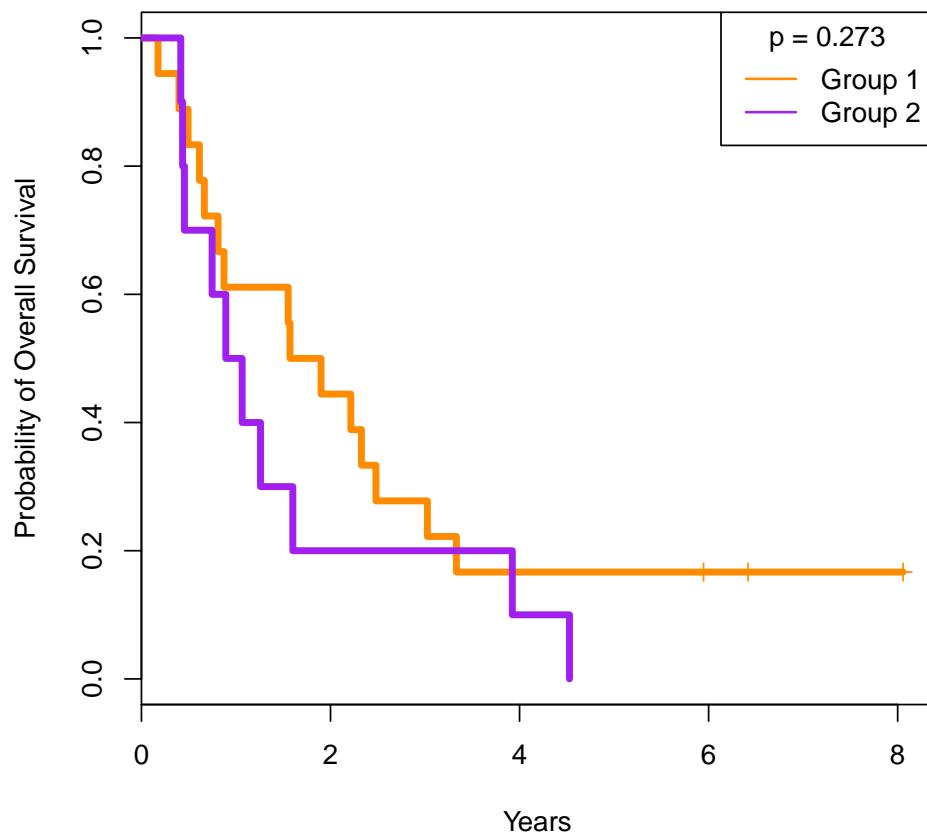
**KM Plots for Tumor Sample Groups (1 and 2) of Epithelioid Samp**

Figure 10: KM for comparing the groups 1 and 2 defined in Figure5 with only epithelioid samples, excluding samples Normal 5, Tumor 6, Tumor 17, Tumor 44, Tumor 45, Tumor 54.

```

> Time.dfs <- as.numeric(TumorSampleSurvivalData[EpiIndex, "OverallTime"])
> cen.status <- as.numeric(as.vector(TumorSampleSurvivalData[EpiIndex, "VitalStatusCen"]))
> ##### fit into KM model
>
> OverallThreeGroupEpi<- survfit(Surv(Time.dfs, cen.status) ~ as.factor(groupused[EpiIndex]),
> OverallThreeGroup.resEpi <- survdiff(Surv(Time.dfs , cen.status) ~ as.factor(groupused[EpiIndex])
> Factorlevel<-groupused[EpiIndex]
> table(groupused[EpiIndex])

Group 1 Group 2 Group 3
 19      11      7

> pValue3All <- 1-pchisq(OverallThreeGroup.resEpi$chisq, df=length(levels(factor(Factorlevel)))
> ##### show the KM result comparing the three groups
>
> OverallThreeGroup.resEpi

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(groupused[EpiIndex]),
na.action = na.exclude)

          N Observed Expected (O-E)^2/E (O-E)^2/V
as.factor(groupused[EpiIndex])=Group 1 19       16     18.29    0.2861    0.6481
as.factor(groupused[EpiIndex])=Group 2 11       11     10.33    0.0435    0.0641
as.factor(groupused[EpiIndex])=Group 3  7        6      4.38    0.5962    0.7000

Chisq= 0.9 on 2 degrees of freedom, p= 0.625

> OverallThreeGroupEpi

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(groupused[EpiIndex]),
na.action = na.exclude)

          records n.max n.start events median 0.95LCL
as.factor(groupused[EpiIndex])=Group 1      19     19      19     16  1.572  0.873
as.factor(groupused[EpiIndex])=Group 2      11     11      11     11  1.227  0.616
as.factor(groupused[EpiIndex])=Group 3       7      7       7      6  0.652  0.545
                                0.95UCL
as.factor(groupused[EpiIndex])=Group 1      3.33
as.factor(groupused[EpiIndex])=Group 2       NA
as.factor(groupused[EpiIndex])=Group 3       NA

> summary(coxph(Surv(Time.dfs, cen.status) ~ as.factor(groupused[EpiIndex]), na.action=na.excl)

```

```

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(groupused[EpiIndex]),
      na.action = na.exclude)

n= 37, number of events= 33

              coef exp(coef) se(coef)     z Pr(>|z|)
as.factor(groupused[EpiIndex])Group 2 0.1979    1.2188   0.3935 0.503   0.615
as.factor(groupused[EpiIndex])Group 3 0.4555    1.5770   0.4838 0.942   0.346

              exp(coef) exp(-coef) lower .95 upper .95
as.factor(groupused[EpiIndex])Group 2    1.219     0.8205   0.5636   2.636
as.factor(groupused[EpiIndex])Group 3    1.577     0.6341   0.6110   4.071

Concordance= 0.558 (se = 0.052 )
Rsquare= 0.024 (max possible= 0.994 )
Likelihood ratio test= 0.89 on 2 df,  p=0.6411
Wald test            = 0.93 on 2 df,  p=0.6287
Score (logrank) test = 0.94 on 2 df,  p=0.6252

```

## 8 Survival Analysis With Other Clinical Information

In this section, we will check the survival analysis with other clinical information. We first calculate the summary of each category in the patients.

```

> ##### find the time and census status
>
> Time.dfs <- TumorSampleSurvivalData[, "OverallTime"]
> cen.status <- as.numeric(as.vector(TumorSampleSurvivalData[, "VitalStatusCen"]))
> ##### derive the age at the date of surgery #####
>
> PatientAge<-TumorSampleSurvivalData[, "PatientAge"]
> ##### summary the age at the date of surgery #####
>
> summary(PatientAge)

  Min. 1st Qu. Median     Mean 3rd Qu.     Max.
40.38    58.68   63.19   63.01   67.69   81.41

> ##### gender #####
>
> PatientGender<-TumorSampleSurvivalData[, "Gender"]
> table(PatientGender)

```

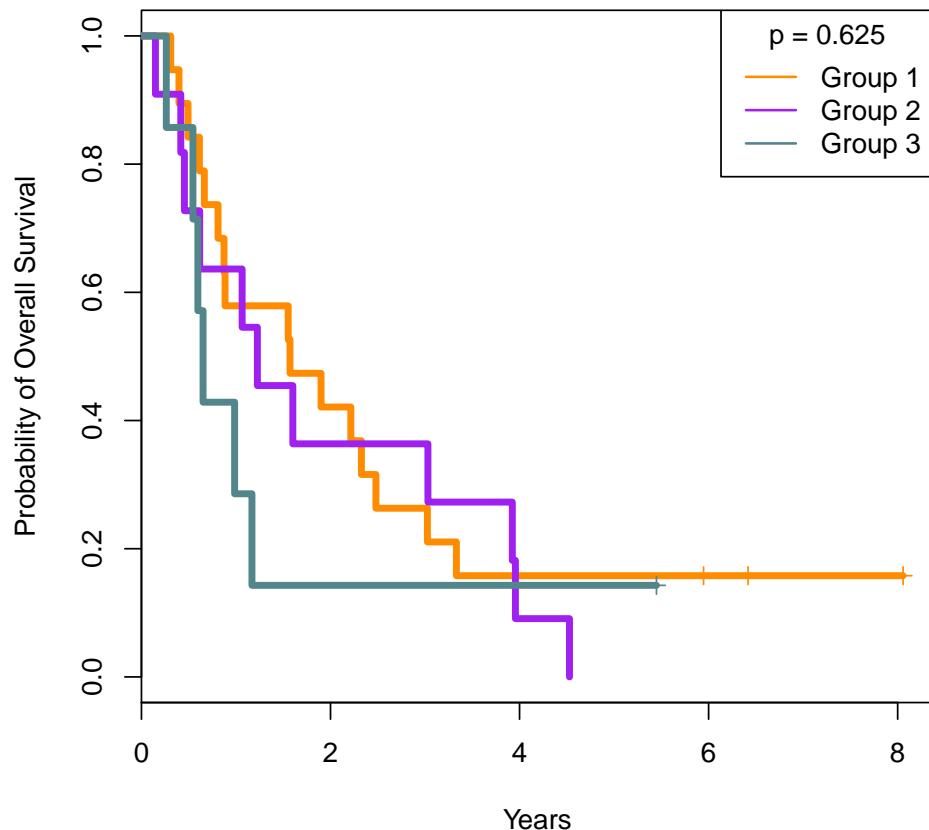
**KM Plots for Tumor Samples Excluding Neo-Adj Treated Patient**

Figure 11: KM for comparing the three groups defined in Figure5, excluding sample Normal 5, Tumor 6, Tumor 17, Tumor 44, Tumor 45, Tumor 54 and samples that are NEo-adj treated.

```

PatientGender
Female   Male
     8      40

> table(PatientGender, cen.status)

            cen.status
PatientGender  0  1
    Female  4  4
    Male    0 40

> table(PatientGender)/length(PatientGender)

PatientGender
    Female      Male
0.1666667 0.8333333

> #####original histology #####
>
> PatientOriginalDiag <- TumorSurvivals[, "Path.Report.Diagnosis"]
> table(PatientOriginalDiag )

PatientOriginalDiag
biphasic epithelioid sarcomatoid
    11        31        6

> table(PatientOriginalDiag , cen.status)

            cen.status
PatientOriginalDiag  0  1
    biphasic      0 11
    epithelioid   3 28
    sarcomatoid   1  5

> table(PatientOriginalDiag )/length(PatientOriginalDiag )

PatientOriginalDiag
biphasic epithelioid sarcomatoid
0.2291667 0.6458333 0.1250000

> ##### T Stage #####
>
> PatientTStage <- as.vector(TumorSampleSurvivalData[, "Pathologic.T.stage"])
> table(PatientTStage)

```

```

PatientTStage
T1 T2 T3 T4
2 5 36 5

> table(PatientTStage, cen.status)

            cen.status
PatientTStage 0 1
      T1 1 1
      T2 0 5
      T3 3 33
      T4 0 5

> table(PatientTStage)/length(PatientTStage)

PatientTStage
      T1          T2          T3          T4
0.04166667 0.10416667 0.75000000 0.10416667

> ##      only two patient at stage 1, so I combine stage 1 and 2
>
> PatientTStageCombine <- PatientTStage
> PatientTStageCombine[PatientTStage == "T1" | PatientTStage == "T2"] <- "T1 or T2"
> table(PatientTStageCombine)

PatientTStageCombine
T1 or T2      T3      T4
    7      36      5

> table(PatientTStageCombine, cen.status)

            cen.status
PatientTStageCombine 0 1
      T1 or T2 1 6
      T3      3 33
      T4      0 5

> table(PatientTStageCombine)/length(PatientTStageCombine)

PatientTStageCombine
      T1          T2          T3          T4
0.1458333 0.7500000 0.1041667

> ##### N Stage #####
>
> PatientNStage <- as.vector(TumorSampleSurvivalData[, "Path.N.stage"])
> table(PatientNStage)

```

```
PatientNStage
```

	N0	N1	N2	N3
23	7	16	2	

```
> table(PatientNStage, cen.status)
```

		cen.status
--	--	------------

PatientNStage	0	1
N0	3	20
N1	1	6
N2	0	16
N3	0	2

```
> table(PatientNStage)/length(PatientNStage)
```

```
PatientNStage
```

	N0	N1	N2	N3
0.47916667	0.47916667	0.14583333	0.33333333	0.04166667

```
> ##      only two patient at stage 3, so I combine stage 2 and 3
```

```
>
```

```
> PatientNStageCombine <- PatientNStage
```

```
> PatientNStageCombine[PatientNStage == "N2" | PatientNStage == "N3"] <- "N2 or N3"
```

```
> table(PatientNStageCombine)
```

```
PatientNStageCombine
```

	N0	N1	N2 or N3
23		7	18

```
> table(PatientNStageCombine, cen.status)
```

		cen.status
--	--	------------

PatientNStageCombine	0	1
N0	3	20
N1	1	6
N2 or N3	0	18

```
> table(PatientNStageCombine)/length(PatientNStageCombine)
```

```
PatientNStageCombine
```

	N0	N1	N2 or N3
0.4791667	0.4791667	0.1458333	0.3750000

```

> ##### Overall.Pathologic.Stage #####
>
>
>
> PatientOverStage <- as.vector(TumorSampleSurvivalData[, "Overall.Pathologic.Stage"])
> table(PatientOverStage)

PatientOverStage
I   II  III   IV
2    2   37    7

> table(PatientOverStage, cen.status)

            cen.status
PatientOverStage 0   1
                  I   1   1
                  II  0   2
                  III 3  34
                  IV  0   7

> table(PatientOverStage)/length(PatientOverStage)

PatientOverStage
           I          II          III          IV
0.041666667 0.041666667 0.77083333 0.14583333

> ##      combine stage I and II and III
>
>
> PatientOverStageCombine <- PatientOverStage
> PatientOverStageCombine[PatientOverStage == "I" | PatientOverStage == "II" | PatientOverStage ==
> table(PatientOverStageCombine)

PatientOverStageCombine
I to III       IV
41           7

> table(PatientOverStageCombine, cen.status)

            cen.status
PatientOverStageCombine 0   1
                  I to III 4 37
                  IV      0   7

> table(PatientOverStageCombine)/length(PatientOverStageCombine)

```

```

PatientOverStageCombine
I to III          IV
0.8541667 0.1458333

> ##### chemo treatment or not #####
>
>
>
> PatientTreat <- as.vector(TumorSampleSurvivalData[, "Neoadjuvant.chemo"])
> table(PatientTreat)

PatientTreat
No Yes
37 11

> table(PatientTreat, cen.status)

            cen.status
PatientTreat 0 1
      No 4 33
      Yes 0 11

> table(PatientTreat)/length(PatientTreat)

PatientTreat
      No      Yes
0.7708333 0.2291667

>
>
>

```

We would like to check each clinical features independently first to check if any of the variables predict overall survival.

## 8.1 Function

In order to minimize the amount of code we need to rewrite, we use the function below to generate Kaplan-Meier plots.

```

> makeKMplot <- function(Time, Status, Clinical,
+                         main="Overall Survival")
+ {
+   degf <- length(levels(Clinical)) - 1

```

```

+   sf <- survfit(Surv(Time, Status) ~ Clinical)
+   sdv <- survdiff(Surv(Time, Status) ~ Clinical)
+
+
+   colset <- c("red", "blue", "green", "orange", "cyan", "magenta")
+   plot(sf, col=colset, xlab="Years", ylab="Probability of Overall Survival", main=main)
+
+   legend("topright", levels(Clinical), col=colset, lwd=2,
+         title = ifelse((1-pchisq(sdv$chisq, degf) > 0.0001), paste("P value =", round(1-pchisq(sdv$chisq, degf), 3)),
+                     "P value < 0.0001"))
+
+
+ }
>
>
```

## 8.2 Age

We calculated the age associated with the survival difference.

```

>       model0 <- coxph(Surv(Time.dfs, cen.status) ~ PatientAge, na.action=na.exclude)
>       summary(model0)

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ PatientAge, na.action = na.exclude)

n= 48, number of events= 44

            coef exp(coef)  se(coef)      z Pr(>|z|)
PatientAge 0.04725    1.04838  0.02329  2.029   0.0425 *
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

            exp(coef) exp(-coef) lower .95 upper .95
PatientAge     1.048      0.9539     1.002     1.097

Concordance= 0.624  (se = 0.05 )
Rsquare= 0.086  (max possible= 0.997 )
Likelihood ratio test= 4.32  on 1 df,   p=0.03766
Wald test           = 4.12  on 1 df,   p=0.04248
Score (logrank) test = 4.15  on 1 df,   p=0.04165

>       model<-survfit(coxph(Surv(Time.dfs, cen.status) ~ PatientAge, na.action=na.exclude))
>       model
```

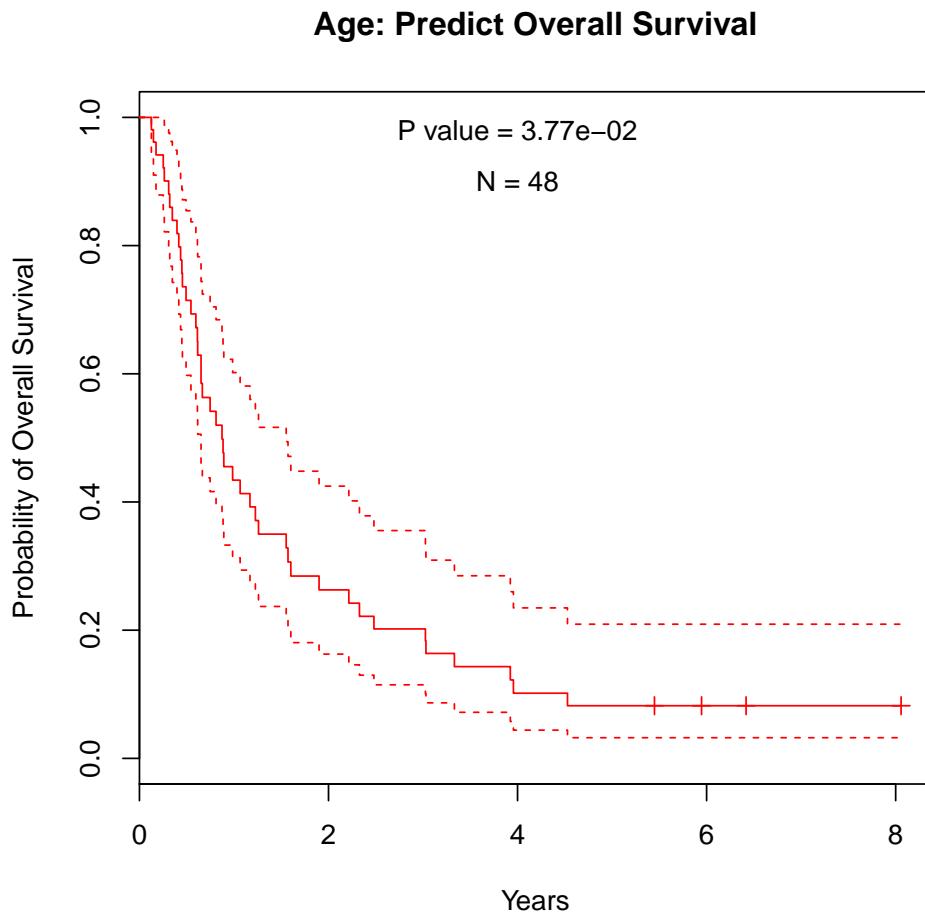


Figure 12: KM for overall survival predicted by age.

```
Call: survfit(formula = coxph(Surv(Time.dfs, cen.status) ~ PatientAge,
  na.action = na.exclude))

records    n.max n.start   events   median 0.95LCL 0.95UCL
 48.000  48.000  48.000  44.000    0.873    0.652    1.552

>
```

### 8.3 Gender

We compared the gender checking the survival difference.

```

>         model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientGender), na.action=na.exclude)
>         summary(model)

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientGender),
      na.action = na.exclude)

n= 48, number of events= 44

            coef exp(coef) se(coef)      z Pr(>|z|)
as.factor(PatientGender)Male 1.5625    4.7708   0.5433 2.876  0.00403 **

---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

            exp(coef) exp(-coef) lower .95 upper .95
as.factor(PatientGender)Male     4.771      0.2096     1.645     13.84

Concordance= 0.582 (se = 0.035 )
Rsquare= 0.221 (max possible= 0.997 )
Likelihood ratio test= 11.99 on 1 df,  p=0.0005346
Wald test             = 8.27 on 1 df,  p=0.004026
Score (logrank) test = 9.7 on 1 df,  p=0.001841

>         model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientGender), na.action=na.exclude)
>         model

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientGender),
      na.action = na.exclude)

          records n.max n.start events median 0.95LCL 0.95UCL
as.factor(PatientGender)=Female     8     8      8      4  2.480   1.552     NA
as.factor(PatientGender)=Male      40    40     40     40  0.706   0.597   1.23

>         model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientGender), na.action=na.exclude)
>         model2

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientGender),
      na.action = na.exclude)

          N Observed Expected (O-E)^2/E (O-E)^2/V
as.factor(PatientGender)=Female  8        4       13      6.21      9.7
as.factor(PatientGender)=Male   40       40      31      2.60      9.7

Chisq= 9.7 on 1 degrees of freedom, p= 0.00184

```

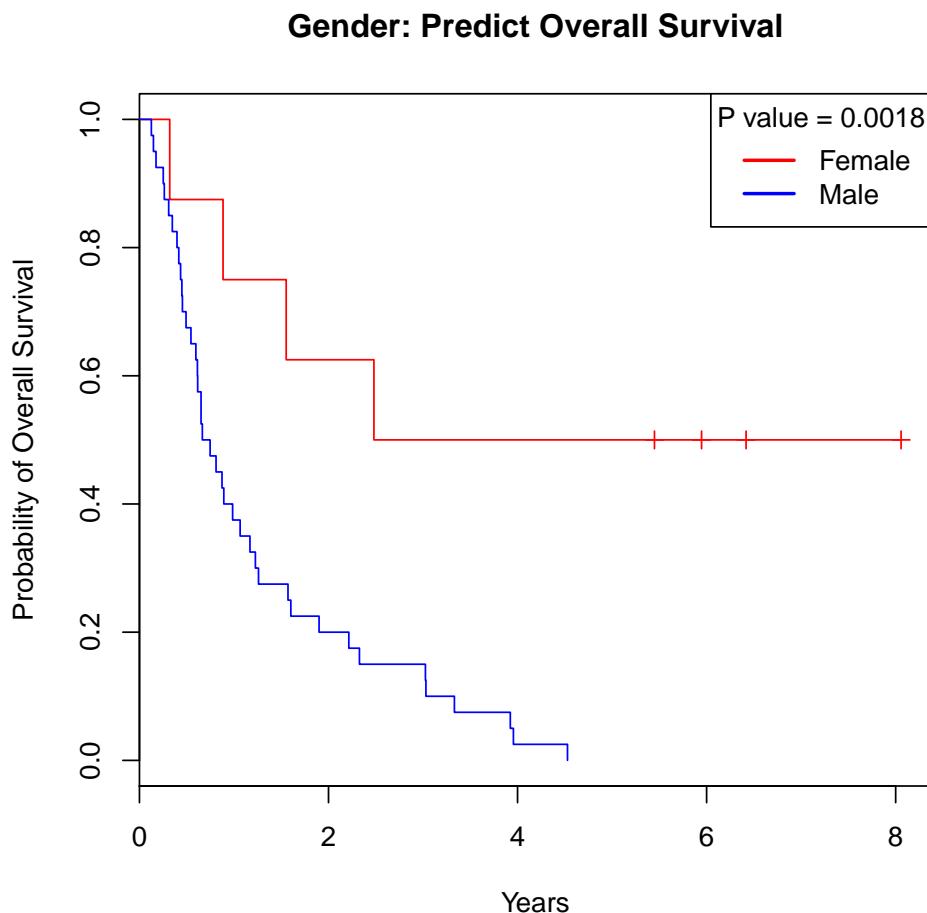


Figure 13: KM for overall survival predicted by gender.

>

#### 8.4 Original Histology

We compared the original histology checking the survival difference.

```
> model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientOriginalDiag), na.action=na.omit)
> summary(model)
```

Call:

```
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOriginalDiag),
```

```

na.action = na.exclude)

n= 48, number of events= 44

              coef exp(coef) se(coef)      z Pr(>|z|)
as.factor(PatientOriginalDiag)=epithelioid -0.5753    0.5625   0.3584 -1.605    0.108
as.factor(PatientOriginalDiag)=sarcomatoid -0.4522    0.6362   0.5439 -0.831    0.406

              exp(coef) exp(-coef) lower .95 upper .95
as.factor(PatientOriginalDiag)=epithelioid    0.5625     1.778    0.2786    1.136
as.factor(PatientOriginalDiag)=sarcomatoid    0.6362     1.572    0.2191    1.847

Concordance= 0.578 (se = 0.04 )
Rsquare= 0.048 (max possible= 0.997 )
Likelihood ratio test= 2.36 on 2 df,  p=0.3076
Wald test           = 2.58 on 2 df,  p=0.2749
Score (logrank) test = 2.65 on 2 df,  p=0.2658

>       model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientOriginalDiag), na.action = na.exclude)
>       model

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOriginalDiag),
             na.action = na.exclude)

              records n.max n.start events median 0.95LCL
as.factor(PatientOriginalDiag)=biphasic        11     11     11     11  0.449  0.263
as.factor(PatientOriginalDiag)=epithelioid      31     31     31     28  1.065  0.665
as.factor(PatientOriginalDiag)=sarcomatoid       6      6      6      5  0.634  0.597
                                         0.95UCL
as.factor(PatientOriginalDiag)=biphasic          NA
as.factor(PatientOriginalDiag)=epithelioid        2.33
as.factor(PatientOriginalDiag)=sarcomatoid         NA

>       model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientOriginalDiag), na.action = na.exclude)
>       model2

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOriginalDiag),
          na.action = na.exclude)

              N Observed Expected (0-E)^2/E (0-E)^2/V
as.factor(PatientOriginalDiag)=biphasic    11       11     7.09  2.161367  2.606452
as.factor(PatientOriginalDiag)=epithelioid 31       28     31.86  0.468750  1.720596

```

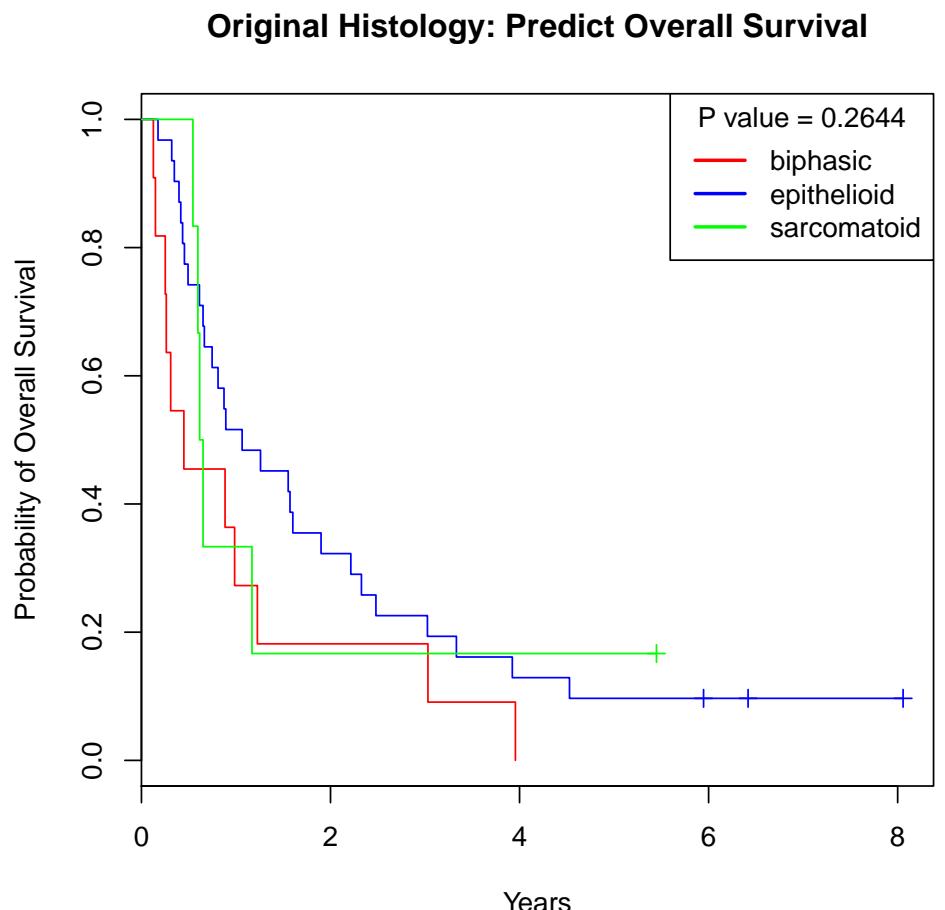


Figure 14: KM for overall survival predicted by original histology.

```
as.factor(PatientOriginalDiag)=sarcomatoid 6      5      5.05  0.000472  0.000541
Chisq= 2.7  on 2 degrees of freedom, p= 0.264
```

## 8.5 T Stage

We compared the T stage checking the survival difference.

```
> model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientTStage), na.action=na.ex)
```

Call:

```
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTStage),
      na.action = na.exclude)
```

n= 48, number of events= 44

	coef	exp(coef)	se(coef)	z	Pr(> z )
as.factor(PatientTStage)T2	2.883	17.876	1.126	2.560	0.0105 *
as.factor(PatientTStage)T3	1.424	4.153	1.018	1.398	0.1621
as.factor(PatientTStage)T4	1.983	7.266	1.111	1.785	0.0743 .

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1 ~ 1

	exp(coef)	exp(-coef)	lower .95	upper .95
as.factor(PatientTStage)T2	17.876	0.05594	1.9664	162.51
as.factor(PatientTStage)T3	4.153	0.24081	0.5642	30.57
as.factor(PatientTStage)T4	7.266	0.13762	0.8229	64.17

Concordance= 0.603 (se = 0.036 )

Rsquare= 0.195 (max possible= 0.997 )

Likelihood ratio test= 10.43 on 3 df, p=0.01526

Wald test = 10.73 on 3 df, p=0.01329

Score (logrank) test = 12.77 on 3 df, p=0.005157

```
> model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientTStage), na.action=na.exclude)
> model
```

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTStage),
 na.action = na.exclude)

	records	n.max	n.start	events	median	0.95LCL	0.95UCL
as.factor(PatientTStage)=T1	2	2	2	1	3.031	3.031	NA
as.factor(PatientTStage)=T2	5	5	5	5	0.397	0.309	NA
as.factor(PatientTStage)=T3	36	36	36	33	0.939	0.652	1.9
as.factor(PatientTStage)=T4	5	5	5	5	0.616	0.545	NA

```
> model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientTStage), na.action=na.exclude)
> model2
```

Call:

```
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTStage),
      na.action = na.exclude)
```

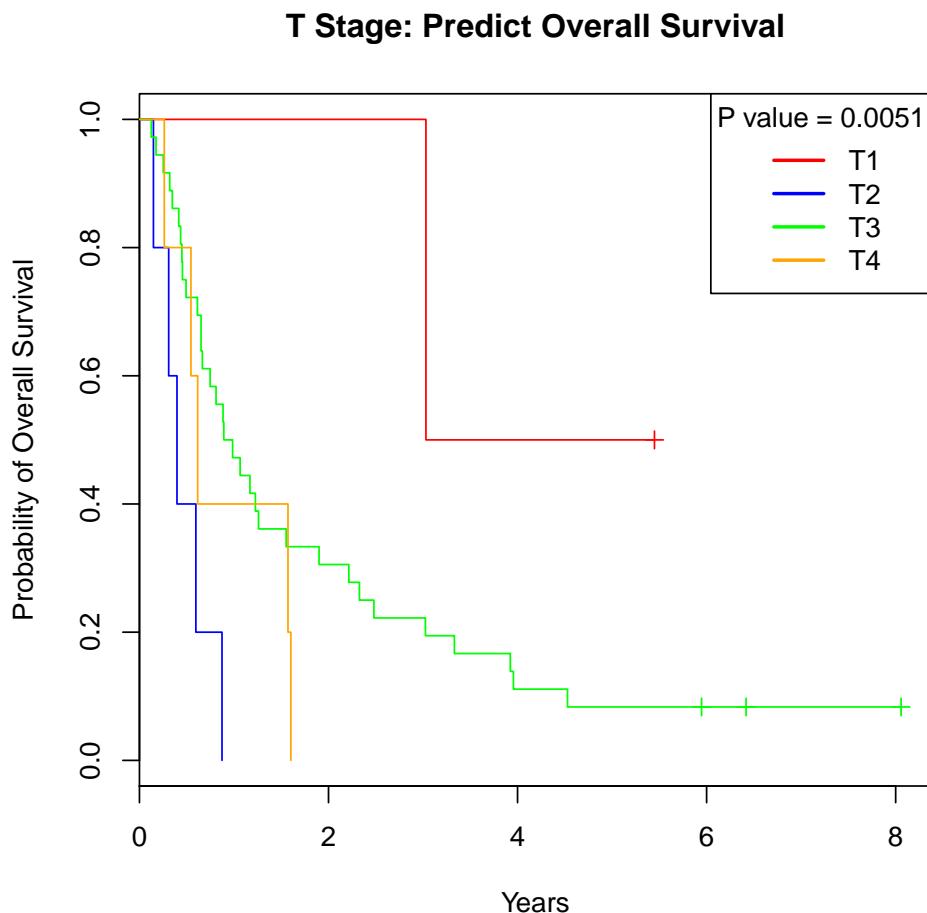


Figure 15: KM for overall survival predicted by T Stage.

	N	Observed	Expected	$(O-E)^2/E$	$(O-E)^2/V$
as.factor(PatientTStage)=T1	2	1	4.11	2.357	2.680
as.factor(PatientTStage)=T2	5	5	1.46	8.568	9.229
as.factor(PatientTStage)=T3	36	33	35.14	0.130	0.650
as.factor(PatientTStage)=T4	5	5	3.29	0.891	0.991

Chisq= 12.8 on 3 degrees of freedom, p= 0.00507

>

We also consider the combined T stage because of the small number of T 1 stage patients

```

>         model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientTStageCombine), na.action = na.exclude)
>         summary(model)

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTStageCombine),
      na.action = na.exclude)

n= 48, number of events= 44

              coef exp(coef) se(coef)      z Pr(>|z|)
as.factor(PatientTStageCombine)T3 -0.1327    0.8757   0.4455 -0.298   0.766
as.factor(PatientTStageCombine)T4  0.3625    1.4370   0.6175  0.587   0.557

              exp(coef) exp(-coef) lower .95 upper .95
as.factor(PatientTStageCombine)T3    0.8757     1.1419   0.3657    2.097
as.factor(PatientTStageCombine)T4    1.4370     0.6959   0.4284    4.820

Concordance= 0.54 (se = 0.036 )
Rsquare= 0.02 (max possible= 0.997 )
Likelihood ratio test= 0.95 on 2 df,  p=0.6216
Wald test            = 1.05 on 2 df,  p=0.5909
Score (logrank) test = 1.07 on 2 df,  p=0.5852

>         model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientTStageCombine), na.action = na.exclude)
>         model

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTStageCombine),
      na.action = na.exclude)

          records n.max n.start events median 0.95LCL
as.factor(PatientTStageCombine)=T1 or T2      7     7      7      6  0.597  0.309
as.factor(PatientTStageCombine)=T3           36    36     36     33  0.939  0.652
as.factor(PatientTStageCombine)=T4           5     5      5      5  0.616  0.545
                                         0.95UCL
as.factor(PatientTStageCombine)=T1 or T2      NA
as.factor(PatientTStageCombine)=T3           1.9
as.factor(PatientTStageCombine)=T4      NA

>         model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientTStageCombine), na.action = na.exclude)
>         model2

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTStageCombine),
      na.action = na.exclude)

```

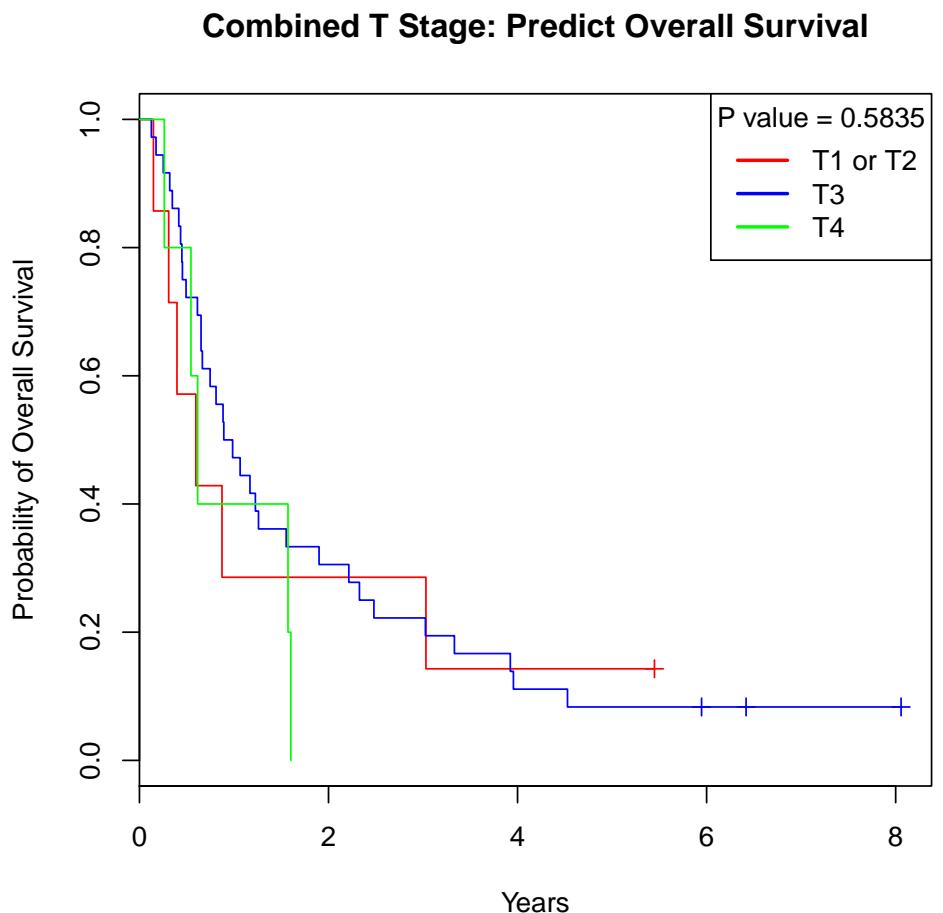


Figure 16: KM for overall survival predicted by combined T Stage.

```
na.action = na.exclude)
```

	N	Observed	Expected	$(O-E)^2/E$	$(O-E)^2/V$
as.factor(PatientTStageCombine)=T1 or T2	7	6	5.58	0.0323	0.0374
as.factor(PatientTStageCombine)=T3	36	33	35.14	0.1299	0.6497
as.factor(PatientTStageCombine)=T4	5	5	3.29	0.8913	0.9909

```
Chisq= 1.1 on 2 degrees of freedom, p= 0.583
```

```
>
```

## 8.6 N Stage

We compared the N stage checking the survival difference.

```
>         model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientNStage), na.action=na.exclude)
>         summary(model)

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientNStage),
      na.action = na.exclude)

n= 48, number of events= 44

            coef exp(coef)  se(coef)      z Pr(>|z|)
as.factor(PatientNStage)N1 0.2220    1.2485   0.4696  0.473  0.63648
as.factor(PatientNStage)N2 1.0446    2.8424   0.3662  2.853  0.00434 **
as.factor(PatientNStage)N3 1.6313    5.1103   0.7759  2.102  0.03552 *
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

            exp(coef)  exp(-coef) lower .95 upper .95
as.factor(PatientNStage)N1     1.249      0.8009   0.4973    3.134
as.factor(PatientNStage)N2     2.842      0.3518   1.3867    5.826
as.factor(PatientNStage)N3     5.110      0.1957   1.1168   23.383

Concordance= 0.618  (se = 0.045 )
Rsquare= 0.184  (max possible= 0.997 )
Likelihood ratio test= 9.76  on 3 df,  p=0.02075
Wald test           = 10.36  on 3 df,  p=0.01576
Score (logrank) test = 11.38  on 3 df,  p=0.009861

>         model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientNStage), na.action=na.exclude)
>         model

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientNStage),
      na.action = na.exclude)

          records n.max n.start events median 0.95LCL 0.95UCL
as.factor(PatientNStage)=N0      23    23      23      20  1.227   0.652   3.92
as.factor(PatientNStage)=N1       7     7       7       6  1.259   0.416     NA
as.factor(PatientNStage)=N2      16    16      16      16  0.615   0.435   1.55
as.factor(PatientNStage)=N3       2     2       2       2  0.461   0.175     NA

>         model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientNStage), na.action=na.exclude)
>         model2
```

Call:

```
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientNStage),
na.action = na.exclude)
```

	N	Observed	Expected	$(O-E)^2/E$	$(O-E)^2/V$
as.factor(PatientNStage)=N0	23	20	27.481	2.037	5.855
as.factor(PatientNStage)=N1	7	6	6.896	0.116	0.139
as.factor(PatientNStage)=N2	16	16	8.956	5.540	7.536
as.factor(PatientNStage)=N3	2	2	0.667	2.663	2.760

Chisq= 11.4 on 3 degrees of freedom, p= 0.00959

>

We also consider the combined N stage because of the small number of N 3 stage patients

```
> model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientNStageCombine), na.action = na.exclude)
> summary(model)
```

Call:

```
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientNStageCombine),
na.action = na.exclude)
```

n= 48, number of events= 44

	coef	exp(coef)	se(coef)	z	Pr(> z )
as.factor(PatientNStageCombine)N1	0.2209	1.2472	0.4696	0.470	0.63806
as.factor(PatientNStageCombine)N2 or N3	1.0885	2.9698	0.3576	3.044	0.00234 **

---

Signif. codes: 0 ‘\*\*\*’ 0.001 ‘\*\*’ 0.01 ‘\*’ 0.05 ‘.’ 0.1 ‘ ’ 1

	exp(coef)	exp(-coef)	lower .95	upper .95
as.factor(PatientNStageCombine)N1	1.247	0.8018	0.4968	3.131
as.factor(PatientNStageCombine)N2 or N3	2.970	0.3367	1.4733	5.986

Concordance= 0.613 (se = 0.045 )

Rsquare= 0.175 (max possible= 0.997 )

Likelihood ratio test= 9.24 on 2 df, p=0.009846

Wald test = 9.62 on 2 df, p=0.008156

Score (logrank) test = 10.38 on 2 df, p=0.005561

```
> model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientNStageCombine), na.action = na.exclude)
> model
```

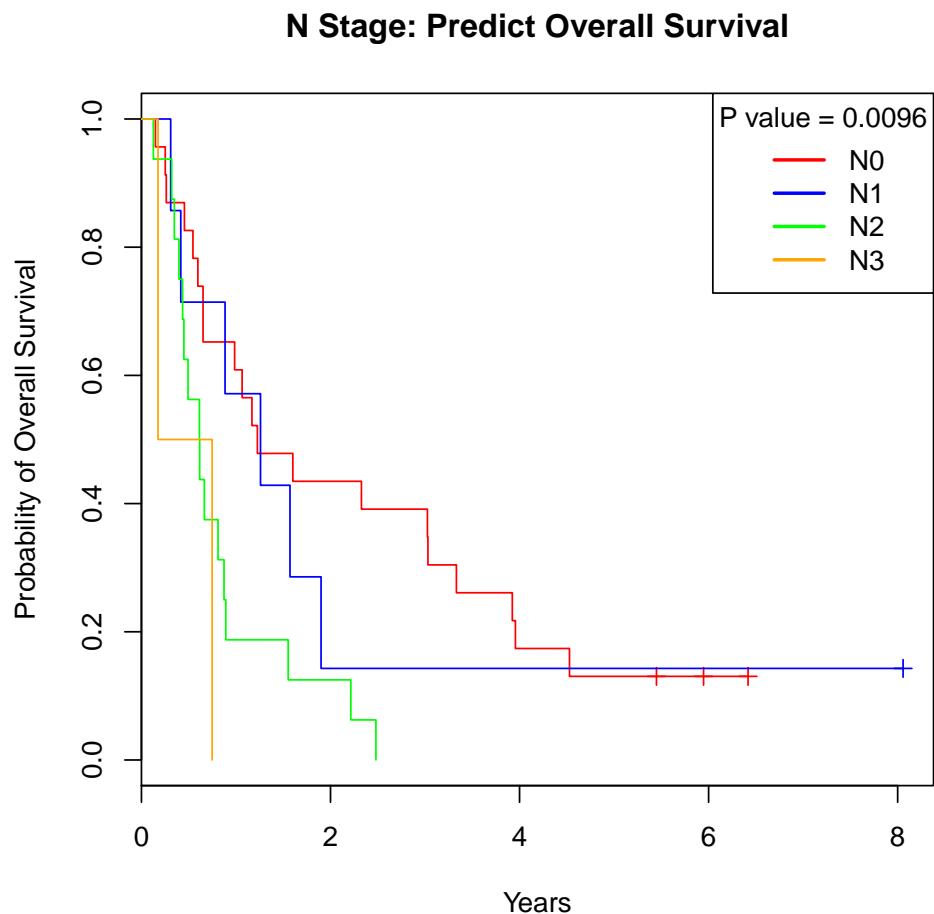


Figure 17: KM for overall survival predicted by N Stage.

```
Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientNStageCombine),
na.action = na.exclude)
```

	records	n.max	n.start	events	median	0.95LCL
as.factor(PatientNStageCombine)=N0	23	23	23	20	1.227	0.652
as.factor(PatientNStageCombine)=N1	7	7	7	6	1.259	0.416
as.factor(PatientNStageCombine)=N2 or N3	18	18	18	18	0.615	0.435
			0.95UCL			
as.factor(PatientNStageCombine)=N0			3.923			
as.factor(PatientNStageCombine)=N1			NA			
as.factor(PatientNStageCombine)=N2 or N3			0.893			

```
> model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientNStageCombine), na.action = na.exclude)
> model2
```

Call:

```
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientNStageCombine),
na.action = na.exclude)
```

	N	Observed	Expected	$(O-E)^2/E$	$(O-E)^2/V$
as.factor(PatientNStageCombine)=N0	23	20	27.48	2.037	5.855
as.factor(PatientNStageCombine)=N1	7	6	6.90	0.116	0.139
as.factor(PatientNStageCombine)=N2 or N3	18	18	9.62	7.292	10.264

Chisq= 10.4 on 2 degrees of freedom, p= 0.00542

>

## 8.7 Overall Stage

We compared the overall stage checking the survival difference.

```
> model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStage), na.action=na.exclude)
> summary(model)
```

Call:

```
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStage),
na.action = na.exclude)
```

n= 48, number of events= 44

	coef	exp(coef)	se(coef)	z	Pr(> z )
as.factor(PatientOverStage)II	3.147	23.256	1.260	2.496	0.0125 *

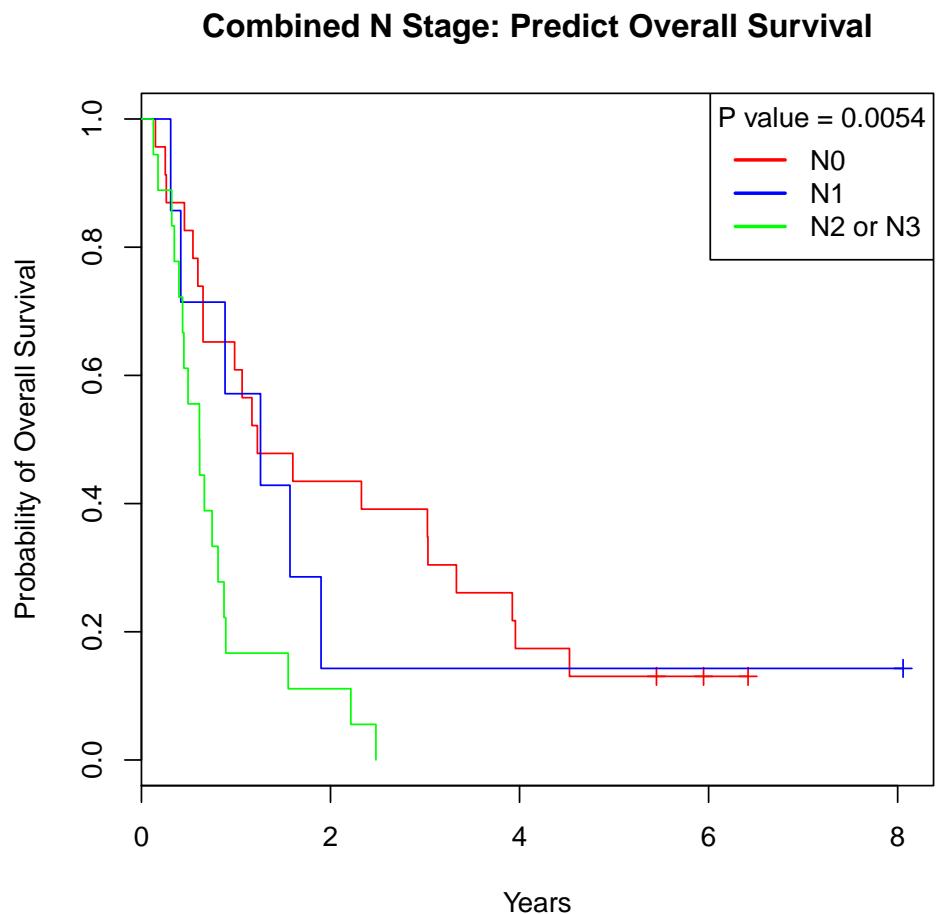


Figure 18: KM for overall survival predicted by combined N Stage.

```

as.factor(PatientOverStage)III 1.433      4.193      1.018 1.408      0.1591
as.factor(PatientOverStage)IV   2.124      8.367      1.086 1.957      0.0504 .
---
Signif. codes:  0 *** 0.001 ** 0.01 * 0.05 . 0.1

              exp(coef) exp(-coef) lower .95 upper .95
as.factor(PatientOverStage)II    23.256      0.0430     1.9663    275.06
as.factor(PatientOverStage)III    4.193      0.2385     0.5701    30.84
as.factor(PatientOverStage)IV    8.367      0.1195     0.9965    70.25

Concordance= 0.594 (se = 0.035 )
Rsquare= 0.17 (max possible= 0.997 )
Likelihood ratio test= 8.94 on 3 df,  p=0.03013
Wald test             = 9.01 on 3 df,  p=0.02916
Score (logrank) test = 10.85 on 3 df,  p=0.01256

>       model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStage), na.action=
>       model

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStage),
na.action = na.exclude)

          records n.max n.start events median 0.95LCL 0.95UCL
as.factor(PatientOverStage)=I      2      2      2      1 3.031 3.031     NA
as.factor(PatientOverStage)=II     2      2      2      2 0.372 0.148     NA
as.factor(PatientOverStage)=III    37     37     37     34 0.893 0.652     1.9
as.factor(PatientOverStage)=IV     7      7      7      7 0.616 0.263     NA

>       model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStage), na.action=
>       model2

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStage),
na.action = na.exclude)

          N Observed Expected (0-E)^2/E (0-E)^2/V
as.factor(PatientOverStage)=I    2      1    4.114    2.3572    2.680
as.factor(PatientOverStage)=II   2      2    0.442    5.4838    5.647
as.factor(PatientOverStage)=III  37     34   35.488    0.0624    0.325
as.factor(PatientOverStage)=IV   7      7    3.955    2.3440    2.670

Chisq= 10.9 on 3 degrees of freedom, p= 0.0125

>

```

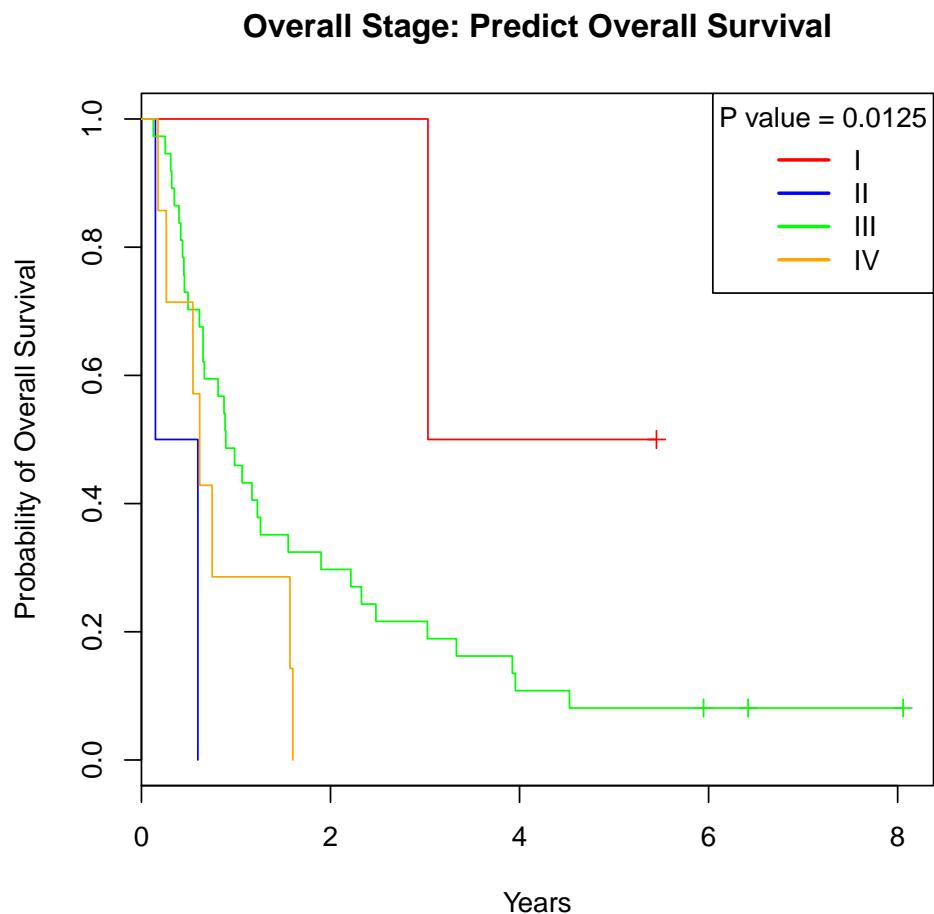


Figure 19: KM for overall survival predicted by Overall Stage.

We also consider the combined Overall stage because of the small number of overall 1 and 2 stage patients

```
>       model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStageCombine), na.action = na.exclude)
>       summary(model)

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStageCombine),
      na.action = na.exclude)

n= 48, number of events= 44

              coef exp(coef) se(coef)      z Pr(>|z|)
as.factor(PatientOverStageCombine)IV 0.6809    1.9757   0.4258 1.599     0.11

              exp(coef) exp(-coef) lower .95 upper .95
as.factor(PatientOverStageCombine)IV      1.976      0.5061      0.8576     4.552

Concordance= 0.537  (se = 0.028 )
Rsquare= 0.045  (max possible= 0.997 )
Likelihood ratio test= 2.23  on 1 df,  p=0.1354
Wald test            = 2.56  on 1 df,  p=0.1098
Score (logrank) test = 2.66  on 1 df,  p=0.1031

>       model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStageCombine), na.action = na.exclude)
>       model

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStageCombine),
      na.action = na.exclude)

              records n.max n.start events median 0.95LCL
as.factor(PatientOverStageCombine)=I to III      41      41      41      37  0.893  0.652
as.factor(PatientOverStageCombine)=IV           7       7       7       7  0.616  0.263
                                         0.95UCL
as.factor(PatientOverStageCombine)=I to III      1.9
as.factor(PatientOverStageCombine)=IV             NA

>       model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStageCombine), na.action = na.exclude)
>       model2

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientOverStageCombine),
      na.action = na.exclude)
```

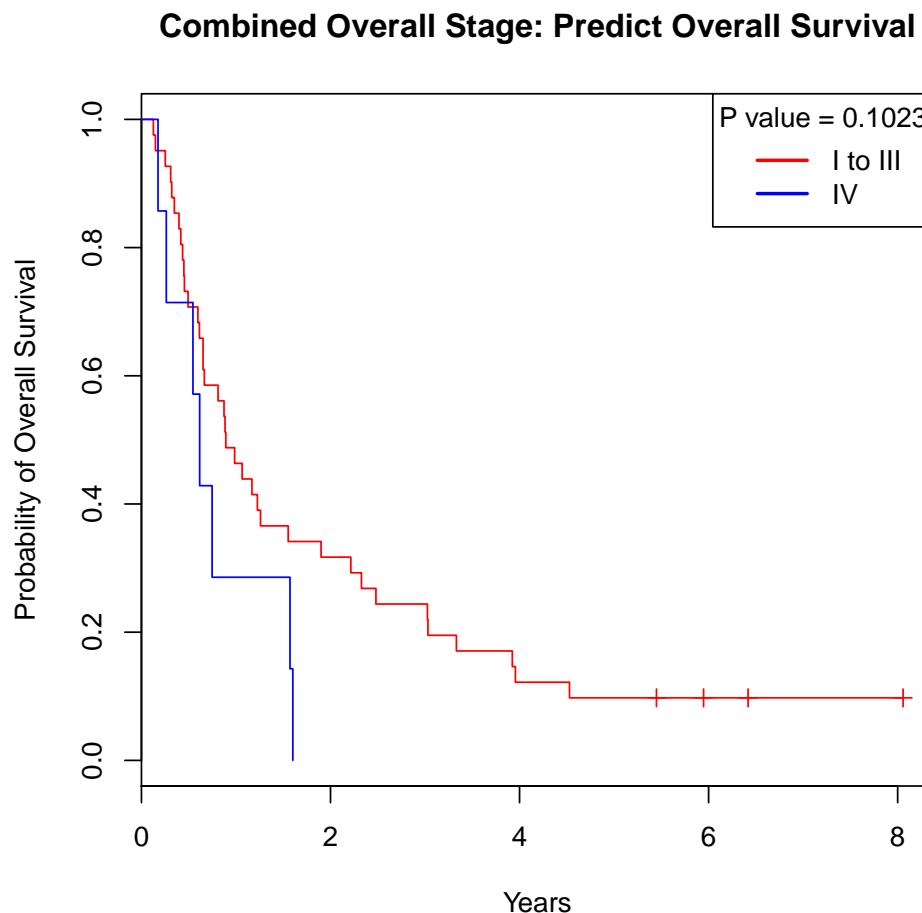


Figure 20: KM for overall survival predicted by combined Overall Stage.

	N	Observed	Expected	$(O-E)^2/E$	$(O-E)^2/V$
as.factor(PatientOverStageCombine)=I to III	41	37	40.04	0.232	2.67
as.factor(PatientOverStageCombine)=IV	7	7	3.96	2.344	2.67

Chisq= 2.7 on 1 degrees of freedom, p= 0.102

>

## 8.8 Chemo Treatment

We compared the chemo treatment checking the survival difference.

```

>         model <- coxph(Surv(Time.dfs, cen.status) ~ as.factor(PatientTreat), na.action=na.exclude)
>         summary(model)

Call:
coxph(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTreat),
      na.action = na.exclude)

n= 48, number of events= 44

            coef exp(coef)  se(coef)      z Pr(>|z|)
as.factor(PatientTreat)Yes 1.3787    3.9697   0.3859 3.573 0.000353 ***
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

            exp(coef) exp(-coef) lower .95 upper .95
as.factor(PatientTreat)Yes      3.97      0.2519     1.863     8.457

Concordance= 0.612 (se = 0.03 )
Rsquare= 0.202 (max possible= 0.997 )
Likelihood ratio test= 10.81 on 1 df,  p=0.00101
Wald test             = 12.77 on 1 df,  p=0.000353
Score (logrank) test = 14.75 on 1 df,  p=0.0001226

>         model <- survfit(Surv(Time.dfs, cen.status) ~ as.factor(PatientTreat), na.action=na.exclude)
>         model

Call: survfit(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTreat),
      na.action = na.exclude)

          records n.max n.start events median 0.95LCL 0.95UCL
as.factor(PatientTreat)=No       37     37     37     33  1.169    0.81    2.33
as.factor(PatientTreat)=Yes      11     11     11     11  0.435    0.32     NA

>         model2 <- survdiff(Surv(Time.dfs, cen.status) ~ as.factor(PatientTreat), na.action=na.exclude)
>         model2

Call:
survdiff(formula = Surv(Time.dfs, cen.status) ~ as.factor(PatientTreat),
      na.action = na.exclude)

          N Observed Expected (O-E)^2/E (O-E)^2/V
as.factor(PatientTreat)=No  37        33     40.03      1.24      14.7
as.factor(PatientTreat)=Yes 11        11     3.97     12.47      14.7

Chisq= 14.7 on 1 degrees of freedom, p= 0.000126

```

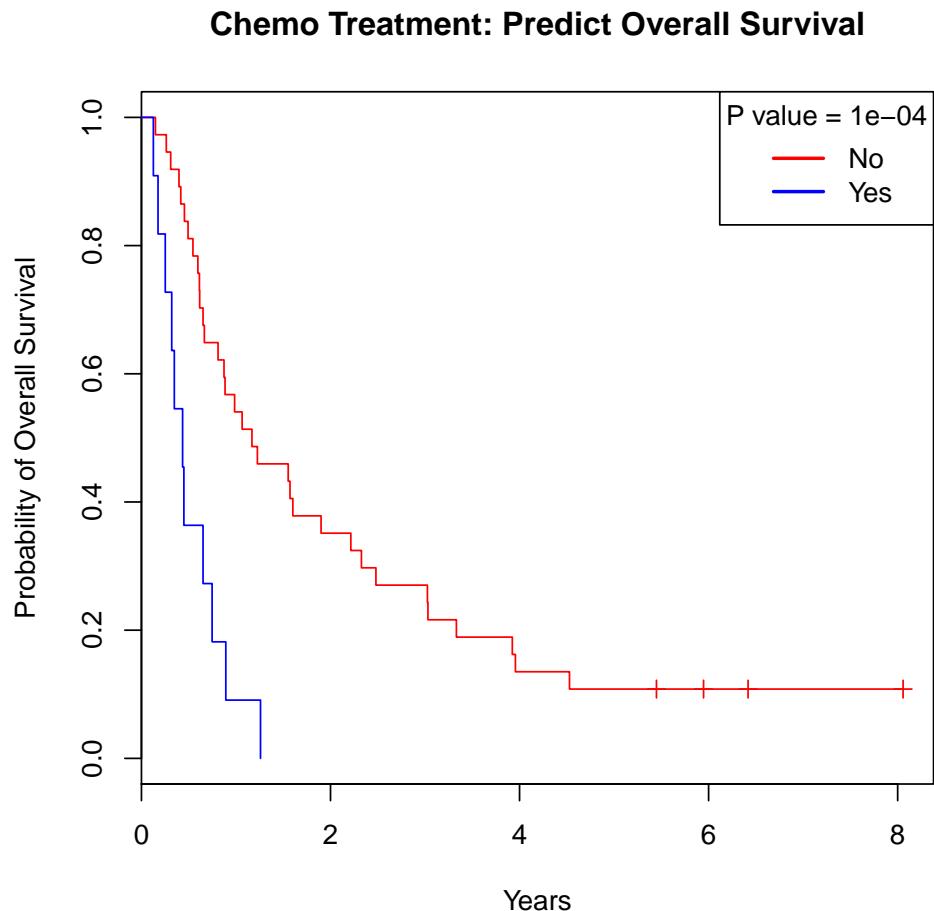


Figure 21: KM for overall survival predicted by chemo treatment.

>

## 8.9 Overall Survival Analysis: Multivariate Analysis for Different Clinical Variables

Here we perform a multivariate analysis to identify the selected clinical variables (significant from univariate model and have been reduced at smaller levels) that best explain overall survival. We exclude shaving diagnosis histology because it is not significant at the univariate level.

```
> ##### create the data
> dataset <- data.frame(Time.dfs, cen.status, PatientAge, PatientGender,
```

```

+          PatientOriginalDiag, PatientTStageCombine, PatientNStageCombine, PatientOver-
> ##### fit in the model
> mod0 <- coxph(Surv(Time.dfs, cen.status) ~ ., data = na.omit(dataset))
>

```

Now we use the Akaike Information Criterion (AIC) to eliminate redundant variables from the model.

```
> mod1 <- step(mod0)
```

Start: AIC=258.57

```
Surv(Time.dfs, cen.status) ~ PatientAge + PatientGender + PatientOriginalDiag +
  PatientTStageCombine + PatientNStageCombine + PatientOverStageCombine +
  PatientTreat
```

	Df	AIC
- PatientTStageCombine	2	254.96
- PatientOverStageCombine	1	256.62
- PatientAge	1	257.71
- PatientOriginalDiag	2	257.82
<none>		258.57
- PatientTreat	1	258.85
- PatientGender	1	265.01
- PatientNStageCombine	2	266.39

Step: AIC=254.96

```
Surv(Time.dfs, cen.status) ~ PatientAge + PatientGender + PatientOriginalDiag +
  PatientNStageCombine + PatientOverStageCombine + PatientTreat
```

	Df	AIC
- PatientOverStageCombine	1	253.12
- PatientAge	1	254.25
- PatientTreat	1	254.94
<none>		254.96
- PatientOriginalDiag	2	255.22
- PatientGender	1	262.60
- PatientNStageCombine	2	262.90

Step: AIC=253.13

```
Surv(Time.dfs, cen.status) ~ PatientAge + PatientGender + PatientOriginalDiag +
  PatientNStageCombine + PatientTreat
```

Df	AIC
----	-----

```

- PatientAge           1 252.61
- PatientTreat         1 253.05
<none>                  253.12
- PatientOriginalDiag  2 253.32
- PatientGender         1 261.90
- PatientNStageCombine 2 262.13

```

Step: AIC=252.61

```
Surv(Time.dfs, cen.status) ~ PatientGender + PatientOriginalDiag +
    PatientNStageCombine + PatientTreat
```

	Df	AIC
<none>		252.61
- PatientTreat	1	253.91
- PatientOriginalDiag	2	254.97
- PatientNStageCombine	2	260.68
- PatientGender	1	262.37

>

Here is the Final model:

```
> summary(mod1)
```

Call:

```
coxph(formula = Surv(Time.dfs, cen.status) ~ PatientGender +
    PatientOriginalDiag + PatientNStageCombine + PatientTreat,
    data = na.omit(dataset))
```

n= 48, number of events= 44

	coef	exp(coef)	se(coef)	z	Pr(> z )
PatientGenderMale	1.69855	5.46603	0.57978	2.930	0.003394 **
PatientOriginalDiagepithelioid	-0.92876	0.39504	0.40654	-2.285	0.022340 *
PatientOriginalDiagsarcomatoid	0.04285	1.04378	0.57411	0.075	0.940501
PatientNStageCombineN1	0.81833	2.26672	0.54586	1.499	0.133832
PatientNStageCombineN2 or N3	1.55030	4.71291	0.44434	3.489	0.000485 ***
PatientTreatYes	0.76645	2.15212	0.41034	1.868	0.061784 .

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1

	exp(coef)	exp(-coef)	lower .95	upper .95
PatientGenderMale	5.466	0.1829	1.7545	17.0287

PatientOriginalDiagepithelioid	0.395	2.5314	0.1781	0.8764
PatientOriginalDiagsarcomatoid	1.044	0.9581	0.3388	3.2158
PatientNStageCombineN1	2.267	0.4412	0.7776	6.6074
PatientNStageCombineN2 or N3	4.713	0.2122	1.9727	11.2593
PatientTreatYes	2.152	0.4647	0.9629	4.8101

Concordance= 0.741 (se = 0.05 )  
 Rsquare= 0.511 (max possible= 0.997 )  
 Likelihood ratio test= 34.38 on 6 df, p=5.685e-06  
 Wald test = 29.78 on 6 df, p=4.336e-05  
 Score (logrank) test = 34.74 on 6 df, p=4.845e-06

> *anova(mod1)*

Analysis of Deviance Table  
 Cox model: response is Surv(Time.dfs, cen.status)  
 Terms added sequentially (first to last)

	loglik	Chisq	Df	Pr(> Chi )
NULL	-137.50			
PatientGender	-131.50	11.9909	1	0.0005346 ***
PatientOriginalDiag	-130.84	1.3117	2	0.5190077
PatientNStageCombine	-121.95	17.7846	2	0.0001374 ***
PatientTreat	-120.31	3.2912	1	0.0696534 .
---				
Signif. codes:	0	***	0.001	***
	1	0.01	0.05	.
	2	0.1	0.5	.
	3	1	5	.

In the final model, gender, original diagnosis, combined N stage and chemo treatment are included.

## 8.10 Overall Survival Analysis: Multivariate Analysis with sub-tumor Groups and Clinical Variables

In the previous section, we have selected the final model including the clinical variables for the survival analysis. In this section, we include the three sub-tumor groups we have defined in the earlier report and check the model.

```
> ##### create the data
> dataset <- data.frame(Time.dfs, cen.status, PatientGender, PatientOriginalDiag,
+                         PatientNStageCombine, PatientTreat, groupused)
> ##### fit in the model
> mod0 <- coxph(Surv(Time.dfs, cen.status) ~ ., data = na.omit(dataset))
> summary(mod0)
```

Call:

```
coxph(formula = Surv(Time.dfs, cen.status) ~ ., data = na.omit(dataset))
```

n= 48, number of events= 44

	coef	exp(coef)	se(coef)	z	Pr(> z )	
PatientGenderMale	2.1814	8.8587	0.6350	3.435	0.000592	***
PatientOriginalDiagepithelioid	-1.1243	0.3249	0.4876	-2.306	0.021134	*
PatientOriginalDiagsarcomatoid	-0.5639	0.5690	0.6239	-0.904	0.366111	
PatientNStageCombineN1	0.9290	2.5320	0.6310	1.472	0.140942	
PatientNStageCombineN2 or N3	1.6623	5.2715	0.5500	3.022	0.002507	**
PatientTreatYes	1.2098	3.3527	0.6579	1.839	0.065927	.
groupusedGroup 2	-0.6455	0.5244	0.6252	-1.032	0.301864	
groupusedGroup 3	0.9981	2.7130	0.7115	1.403	0.160706	

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1

	exp(coef)	exp(-coef)	lower	.95	upper	.95
PatientGenderMale	8.8587	0.1129	2.5517	30.7545		
PatientOriginalDiagepithelioid	0.3249	3.0780	0.1249	0.8449		
PatientOriginalDiagsarcomatoid	0.5690	1.7575	0.1675	1.9328		
PatientNStageCombineN1	2.5320	0.3949	0.7351	8.7210		
PatientNStageCombineN2 or N3	5.2715	0.1897	1.7939	15.4911		
PatientTreatYes	3.3527	0.2983	0.9235	12.1725		
groupusedGroup 2	0.5244	1.9069	0.1540	1.7858		
groupusedGroup 3	2.7130	0.3686	0.6727	10.9422		

Concordance= 0.771 (se = 0.05 )

Rsquare= 0.597 (max possible= 0.997 )

Likelihood ratio test= 43.57 on 8 df, p=6.868e-07

Wald test = 35.92 on 8 df, p=1.817e-05

Score (logrank) test = 43.59 on 8 df, p=6.793e-07

> anova(mod0)

Analysis of Deviance Table

Cox model: response is Surv(Time.dfs, cen.status)

Terms added sequentially (first to last)

	loglik	Chisq	Df	Pr(> Chi )
NULL	-137.50			
PatientGender	-131.50	11.9909	1	0.0005346 ***
PatientOriginalDiag	-130.84	1.3117	2	0.5190077

```
PatientNStageCombine -121.95 17.7846 2 0.0001374 ***
PatientTreat          -120.31  3.2912 1 0.0696534 .
groupused            -115.71  9.1888 2 0.0101083 *
---
Signif. codes:  0 *** 0.001 ** 0.01 * 0.05 . 0.1
>
```

We use the Akaike Information Criterion (AIC) to eliminate redundant variables from the model. The same model stays.

```
> mod1 <- step(mod0)

Start:  AIC=247.42
Surv(Time.dfs, cen.status) ~ PatientGender + PatientOriginalDiag +
    PatientNStageCombine + PatientTreat + groupused

              Df      AIC
<none>           247.43
- PatientTreat     1 248.64
- PatientOriginalDiag  2 248.69
- groupused        2 252.61
- PatientNStageCombine  2 253.64
- PatientGender     1 262.05
>
```

## 9 Appendix

This analysis was run in the following directory:

```
> getwd()
[1] "/data/bioinfo/Private/LungSpore/ReportWithNewClin"
```

This analysis was run in the following software environment:

```
> sessionInfo()

R version 2.15.1 (2012-06-22)
Platform: x86_64-unknown-linux-gnu (64-bit)

locale:
[1] LC_CTYPE=en_US.UTF-8          LC_NUMERIC=C                  LC_TIME=en_US.UTF-8
```

```
[4] LC_COLLATE=en_US.UTF-8      LC_MONETARY=en_US.UTF-8      LC_MESSAGES=en_US.UTF-8
[7] LC_PAPER=C                  LC_NAME=C                  LC_ADDRESS=C
[10] LC_TELEPHONE=C             LC_MEASUREMENT=en_US.UTF-8 LC_IDENTIFICATION=C

attached base packages:
[1] grid      splines    parallel   stats      graphics   grDevices  utils      datasets
[9] methods   base

other attached packages:
[1] RColorBrewer_1.0-5     survival_2.37-4      nlme_3.1-108
[4] affyio_1.24.0          gplots_2.11.0        MASS_7.3-23
[7] KernSmooth_2.23-10    caTools_1.14         gdata_2.12.0.2
[10] gtools_2.7.1           hgu133plus2.db_2.7.1 org.Hs.eg.db_2.7.1
[13] RSQLite_0.11.3          DBI_0.2-6            limma_3.14.0
[16] ClassDiscovery_2.10.2   mclust_4.1          cluster_1.14.4
[19] ClassComparison_2.10.1  PreProcess_2.10.1    oompaBase_2.12.0
[22] xtable_1.7-1            geneplotter_1.34.0  lattice_0.20-15
[25] annotate_1.34.1         AnnotationDbi_1.22.5 simpleaffy_2.32.0
[28] gcrma_2.28.0            BiocInstaller_1.4.9  genefilter_1.38.0
[31] affy_1.34.0              Biobase_2.16.0       BiocGenerics_0.6.0

loaded via a namespace (and not attached):
[1] Biostrings_2.24.1      bitops_1.0-5        IRanges_1.18.0
[4] preprocessCore_1.18.0   stats4_2.15.1      tools_2.15.1
[7] XML_3.96-1.1           zlibbioc_1.2.0
```