Supplementary Information

A new approach to characterising infectious disease transmission dynamics from sentinel surveillance: application to the Italian 2009-2010 A/H1N1 influenza pandemic

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Computation of R_0

For completeness, we define the infectivity function $A(\tau)$ at time τ after infection. Let g(t) denote the probability density function of the variable T_E , the length of the latent period (*i.e.* the time spent in the classes E^1 and E^2). The probability of being in class I^1 at time τ (after infection) is given by

$$P(I^1,\tau) = \int_0^\tau g(t)e^{-\gamma(\tau-t)}dt$$
(26)

In a similar fashion, the probability of being in class I^2 at time τ (after infection) is given by

$$P(I^2,\tau) = \int_0^\tau g(t) \int_t^\tau \gamma e^{\gamma(u-t)} e^{-\gamma(\tau-u)} du dt$$
(27)

 $A(\tau)$ is defined as follows

$$A(\tau) = h^1 P(I^1, \tau) + h^2 P(I^2, \tau)$$
(28)

Equation (4) needs the computation of

$$\int_{0}^{+\infty} A(\tau) d\tau = \frac{h_1}{\gamma} + \frac{h_2}{\gamma}$$
(29)

Computation of $P(ILI_t^i, P_t | T_t, \bar{Z}_t^i, \rho_i)$

This computation is based on the mathematical definition of the Beta function B for two variables x,y>0

$$B(x,y) = \int_0^1 t^{x-1} (1-t)^{y-1} dt$$
(30)

From the definition given in (30) it follows that

$$P(ILI_t^i, P_t | T_t, \bar{Z}_t^i, \rho_i) =$$

$$\frac{\binom{T_t}{P_t}}{B(\alpha,\beta)} \left((1-\rho_i)^{Z_t^i} \int_0^1 \pi_t^{P_t+\alpha-1} (1-\pi_t)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \right)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \left((1-\rho_t)^{Z_t^i} \int_0^1 \pi_t^{P_t+\alpha-1} (1-\pi_t)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \right)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \left((1-\rho_t)^{Z_t^i} \int_0^1 \pi_t^{P_t+\alpha-1} (1-\pi_t)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \right)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \left((1-\rho_t)^{Z_t^i} \int_0^1 \pi_t^{P_t+\alpha-1} (1-\pi_t)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \right)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \left((1-\rho_t)^{Z_t^i} \int_0^1 \pi_t^{P_t+\alpha-1} (1-\pi_t)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \right)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \left((1-\rho_t)^{Z_t^i} \int_0^1 \pi_t^{P_t+\alpha-1} (1-\pi_t)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \right)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \left((1-\rho_t)^{Z_t^i} \int_0^1 \pi_t^{P_t+\alpha-1} (1-\pi_t)^{Z_t^i} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \right)^{ILI_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{(1-\rho_t)^{Z_t^i}}{B(\alpha,\beta)} \left((1-\rho_t)^{Z_t^i} (1-\rho$$

$$+\sum_{F_t^i=1}^{\min(ILI_t^i,Z_t^i)} \binom{ILI_t^i-1}{F_t^i-1} \binom{Z_t^i}{F_t^i} \rho_i^{F_t^i} (1-\rho_i)^{Z_t^i-F_t^i} \int_0^1 \pi_t^{F_t^i+P_t+\alpha-1} (1-\pi_t)^{ILI_t^i-F_t^i+T_t-P_t+\beta-1} d\pi_t + \frac{1}{2} \frac{\Gamma_t^i}{\Gamma_t^i} \frac{\Gamma_t^i}{\Gamma_t^i} \int_0^1 \pi_t^{F_t^i+P_t+\alpha-1} (1-\pi_t)^{ILI_t^i} \frac{\Gamma_t^i}{\Gamma_t^i} \frac{\Gamma_t^i}{\Gamma_$$

$$+ \begin{pmatrix} Z_t^i \\ F_t^i \end{pmatrix} F_t^i \rho_i^{F_t^i} (1-\rho_i)^{Z_t^i - F_t^i} \int_0^1 \pi_t^{P_t + \alpha - 1} (1-\pi_t)^{ILI_t^i + T_t - P_t + \beta - 1} d\pi_t \end{pmatrix}$$

$$=\frac{\binom{T_t}{P_t}}{B(\alpha,\beta)}\left((1-\rho_i)^{\bar{Z}_t^i}B(P_t+\alpha,ILI_t^i+T_t-P_t+\beta)+\right)$$

$$+\sum_{F_t^i=1}^{\min(ILI_t^i,\bar{Z}_t^i)} \binom{ILI_t^i-1}{F_t^i-1} \binom{\bar{Z}_t^i}{F_t^i} \rho_i^{F_t^i} (1-\rho_i)^{\bar{Z}_t^i-F_t^i} B(F_t^i+P_t+\alpha,ILI_t^i-F_t^i+T_t-P_t+\beta)$$

Computation of $P(F_t^i | \bar{Z}_t^i, \rho_i, r)$

In this section we prove (22) given (21) and (12) using the probability generating function.

For simplicity of notation let's set

$$q_t^i = \frac{\bar{Z}_t^i}{\bar{Z}_t^i + r} \tag{31}$$

From (21) and (12) it follows that the probability generating function of Z_t^i is given by

$$G_{Z_t^i}(t) = \mathbb{E}\left[t^{Z_t^i}\right] = \frac{(1 - q_t^i)^r}{(1 - tq_t^i)^r}$$
(32)

and the probability generating function of ${\cal F}^i_t$ is given by

$$G_{F_t^i}(t) = \mathbb{E}[t^{F_t^i}] = (1 - \rho_i + t\rho_i)^{Z_t^i}$$
(33)

where \mathbbm{E} denotes the expected value. Since

$$\mathbb{E}\left[t^{F_t^i}\right] = \mathbb{E}\left[\mathbb{E}\left[t^{F_t^i}|Z_t^i\right]\right] = \mathbb{E}\left[(1-\rho_i+t\rho_i)^{Z_t^i}\right]$$

it follows that

$$\begin{split} G_{F_t^i}(t) &= \frac{\left(1 - q_t^i\right)^r}{\left[1 - (1 - \rho_i + t\rho_i)q_t^i\right]^r} \\ &= \frac{\left(1 - q_t^i\right)^r}{\left[1 - q_t^i(1 - \rho_i) - t\rho_i q_t^i\right]^r} \\ &= \frac{\left[1 - \frac{\rho_i q_t^i}{1 - q_t^i(1 - \rho_i)}\right]^r}{\left[1 - t\frac{\rho_i q_t^i}{1 - q_t^i(1 - \rho_i)}\right]^r} \end{split}$$

and hence by (31) we may conclude (22).

Tables

| | 0-4 | 5-14 | 15-24 | 25-64 | 65 + |
|-------|------|-------|-------|-------|------|
| 0-4 | 5.23 | 1.50 | 0.59 | 11.68 | 1.11 |
| 5-14 | 0.76 | 14.49 | 1.72 | 13.14 | 1.05 |
| 15-24 | 0.28 | 1.58 | 13.94 | 9.16 | 0.85 |
| 25-64 | 0.99 | 2.20 | 1.67 | 11.38 | 1.94 |
| 65 + | 0.26 | 0.49 | 0.43 | 5.38 | 2.93 |

Table SI-1: Symmetrised contact matrix of all reported contacts (physical and nonphysical) in Italy, consisting of the average number of contact persons recorded per working day per survey participant (Polymod 2008). Row index represents the age class of the participant, column index represents the age class of the contact.

| | 0-4 | 5-14 | 15-24 | 25-64 | 65 + |
|-------|------|------|-------|-------|------|
| 0-4 | 1.69 | 1.42 | 0.30 | 6.60 | 0.73 |
| 5-14 | 0.72 | 7.84 | 1.42 | 9.38 | 1.41 |
| 15-24 | 0.14 | 1.30 | 10.41 | 10.03 | 0.25 |
| 25-64 | 0.56 | 1.57 | 1.83 | 9.06 | 1.74 |
| 65 + | 0.17 | 0.65 | 0.13 | 4.81 | 0.57 |

Table SI-2: Symmetrised contact matrix of all reported contacts (physical and nonphysical) in Italy, consisting of the average number of contact persons recorded per holiday day per survey participant (Polymod 2008). Row index represents the age class of the participant, column index represents the age class of the contact.

| Basic no-overdispersion | | | |
|-----------------------------|----------------------|--------------------|--|
| | Susceptibility model | Immunity model | |
| DIC | 9227.3 | 8822.9 | |
| log-likelihood | -4612.3 | -4409.9 | |
| | (-4610.6, -4615.7) | (-4408.3, -4413.5) | |
| R | 1.354 | 1.357 | |
| | (1.348, 1.360) | (1.351, 1.363) | |
| I_0 | 436 | 366 | |
| | (372, 504) | (312, 426) | |
| $ \rho_1 = \dots = \rho_4 $ | 0.090 | 0.099 | |
| | (0.088, 0.091) | (0.097, 0.101) | |

Table SI-3: Basic no-overdispersion models (susceptibility and immunity variants): mean and, in brackets, equal-tailed 95% credible interval of the marginal posterior distribution of the parameters for each specified model. Basic: reporting rates constant in time and across the age groups.

| Basic with-overdispersion | | |
|----------------------------|----------------------|------------------|
| | Susceptibility model | Immunity model |
| DIC | 1528.3 | 1525.3 |
| log-likelihood | -762.3 | -760.9 |
| | (-760.4, -766.2) | (-764.8, -758.9) |
| R | 1.293 | 1.293 |
| | (1.261, 1.326) | (1.261, 1.325) |
| I_0 | 2842 | 2600 |
| | (1172, 5804) | (1082, 5255) |
| $\rho_1 = \cdots = \rho_4$ | 0.139 | 0.152 |
| | (0.117, 0.166) | (0.128, 0.180) |
| r | 2.480 | 2.598 |
| | (1.679, 3.468) | (1.751, 3.634) |

Table SI-4: Basic with-overdispersion models (susceptibility and immunity variants): mean and, in brackets, equal-tailed 95% credible interval of the marginal posterior distribution of the parameters for each specified model. Basic: reporting rates constant in time and across the age groups.

| No-overdispersion | | | | |
|-------------------|----------------------|--------------------|--------------------|--------------------|
| | Susceptibility model | | Immunity model | |
| | ADR | TVR | ADR | TVR |
| DIC | 2600.6 | 2531.4 | 2553.8 | 2502.7 |
| log-likelihood | -1296.5 | -1265.7 | -1274.2 | -1250.0 |
| | (-1293.9, -1300.8) | (-1262.6, -1270.7) | (-1271.6, -1278.6) | (-1247.0, -1254.8) |
| R | 1.405 | 1.368 | 1.402 | 1.369 |
| | (1.398, 1.411) | (1.355, 1.382) | (1.396, 1.408) | (1.358, 1.380) |
| I_0 | 115 | 264 | 111 | 237 |
| | (98, 133) | (178, 377) | (94, 130) | (170, 322) |
| $ ho_1$ | 0.182 | 0.174 | 0.197 | 0.188 |
| | (0.176, 0.187) | (0.132, 0.217) | (0.191, 0.204) | (0.151, 0.234) |
| $ ho_2$ | 0.171 | 0.157 | 0.183 | 0.168 |
| | (0.167, 0.175) | (0.120, 0.194) | (0.179, 0.188) | (0.136, 0.208) |
| $ ho_3$ | 0.059 | 0.056 | 0.066 | 0.063 |
| | (0.057, 0.060) | (0.043, 0.070) | (0.064, 0.068) | (0.050, 0.078) |
| $ ho_4$ | 0.038 | 0.037 | 0.043 | 0.042 |
| | (0.035, 0.040) | (0.028, 0.047) | (0.040, 0.046) | (0.033, 0.052) |
| a | _ | 1.309 | _ | 1.291 |
| | - | (0.981, 1.791) | - | (0.978, 1.653) |
| b | _ | 0.481 | _ | 0.535 |
| | - | (0.328, 0.692) | - | (0.371, 0.739) |

Table SI-5: ADR and TVR no-overdispersion models (susceptibility and immunity variants): mean and, in brackets, equal-tailed 95% credible interval of the marginal posterior distribution of the parameters for each specified model. ADR: age-dependent reporting, *i.e.* reporting rates constant in time and age-specific. TVR: time-varying reporting, *i.e.* age-specific and time-dependent reporting rates as defined by the piecewise linear function given in Eq. (25).



53



Figure SI-1: Basic no-overdispersion models (baseline, susceptibility and immunity variants): plot of the mean, 95%CI, maximum and minimum simulated weekly reported incidences (per 1000) of symptomatic H1N1 cases (i.e. ILI & H1N1 cases) in the 0 - 4, 5 - 14, 15 - 64, 65 + years age-classes and in the overall population. The dots represent the observed data (*i.e.* the H1N1-attributable ILI incidence curve). Basic: reporting rates constant in time and across the age groups.





Figure SI-2: Basic with-overdispersion models (baseline, susceptibility and immunity variants): plot of the mean, 95%CI, maximum and minimum simulated weekly reported incidences (per 1000) of symptomatic H1N1 cases (i.e. ILI & H1N1 cases) in the 0-4, 5-14, 15-64, 65+ years age-classes and in the overall population. The dots represent the observed data (*i.e.* the H1N1-attributable ILI incidence curve). Basic: reporting rates constant in time and across the age groups.





Figure SI-3: ADR no-overdispersion models (baseline, susceptibility and immunity variants): plot of the mean, 95%CI, maximum and minimum simulated weekly reported incidences (per 1000) of symptomatic H1N1 cases (i.e. ILI & H1N1 cases) in the 0-4, 5-14, 15-64, 65+ years age-classes and in the overall population. The dots represent the observed data (*i.e.* the H1N1-attributable ILI incidence curve). ADR: age-dependent reporting, *i.e.* reporting rates constant in time and age-specific.





Figure SI-4: ADR with-overdispersion models (baseline, susceptibility and immunity variants): plot of the mean, 95%CI, maximum and minimum simulated weekly reported incidences (per 1000) of symptomatic H1N1 cases (i.e. ILI & H1N1 cases) in the 0-4, 5-14, 15-64, 65+ years age-classes and in the overall population. The dots represent the observed data (*i.e.* the H1N1-attributable ILI incidence curve). ADR: age-dependent reporting, *i.e.* reporting rates constant in time and age-specific.





Figure SI-5: TVR no-overdispersion models (baseline, susceptibility and immunity variants): plot of the mean, 95%CI, maximum and minimum simulated weekly reported incidences (per 1000) of symptomatic H1N1 cases (i.e. ILI & H1N1 cases) in the 0-4, 5-14, 15-64, 65+ years age-classes and in the overall population. The dots represent the observed data (*i.e.* the H1N1-attributable ILI incidence curve). TVR: time-varying reporting, *i.e.* age-specific and time-dependent reporting rates as defined by the piecewise linear function given in Eq. (25).