Supplementary Information

2 S1 Critique of previous analyses and rationale for new anal-

₃ ysis

4 S1.1 Pseudo-replication

- 5 The two most recent analyses of the organic to conventional yield gap utilized the yield
- 6 ratios without always taking into account the underlying data structure, leading to po-
- tential pseudo-replication and an understated Type 1 error rate [1, 2].
- The random effect meta-analysis employed by Seufert et al. [1] assumes that effect sizes
- ⁹ are independent and drawn from a common distribution [3]. However, multiple response
- ratios were extracted from studies without nesting these observations. Doing so may
- violate the assumption of independence, since multiple response ratios from the same
- study may be non-independent. In addition, treating response ratios from the same study
- as independent gives the studies with the most response ratios disproportionate weight
- while also inflating the replication and hence artificially reduces the confidence region of
- the analysis. Further, in approximately 60% of studies, one set of yield data was used as
- the baseline for multiple comparisons and these were then treated as independent data
- points. Treating response ratios that incorporate the same baseline as independent data
- points also violates assumptions of independence for meta-analysis [4, 5].
- 19 The de Ponti et al. [2] study had similar issues with pseudo-replication. Additionally
- 20 they did not account for the sampling variance within studies, which is the recommended
- 21 practice to deal with unequal variances in the sample of studies [6].

Because we constructed a random effects model (Eq. S1) and conducted the same categorical comparisons as Seufert et al. [1], we first investigated the effects of not accounting for
the hierarchical data structure on the Type 1 error rate for the Seufert et al. model. We also
investigated how much of the difference in our results was due to the hierarchical modeling approach, versus the additional data in our study. We did not investigate the de Ponti
et al. [2] analysis in detail because its structure was entirely different from ours.

8 S1.2 Re-analysis of Seufert et al. data

In order to estimate the Type I error rate of the Seufert et al. analysis, we used a randomization test. We forced the null hypothesis to be true by randomly re-assigning the 30 organic' and 'conventional' labels for each study and then using the R package <code>Metafor</code> 31 [7] to implement a random effects meta-analysis on each randomized dataset. Repeating 32 this procedure 10⁵ times enabled us to determine the Type I error rate (false rejection) 33 resulting from not accounting for the hierarchical structure of the data. In over 50% of simulations, the null hypothesis was rejected using a nominal Type I error rate of 0.05 35 (Fig. S1). In other words, even if organic and conventional yields are known not to be different, applying the model used by [1] for these data would lead to the conclusion that 37 they are significantly different in over 50% of cases. This means that the actual Type I error rate is inflated relative to what was reported, leading to the following related statements: the significance levels were overstated; the confidence intervals were underestimated; the uncertainty was not fully accounted for. This is a likely explanation for why these authors found many significant differences between explanatory variables for management, study quality, and crop type, whereas we did not.

44 S1.3 How much are differences in results due to the model or the data?

- We conducted a series of tests to evaluate the extent to which differences between our
- results and those of Seufert et al. [1] depended on the model we use or the differences
- among the datasets (the Seufert et al. data was a subset of the data we used).
- First, we re-analyzed Seufert et al.'s [1] data (316 comparisons from 66 studies) with their
- 49 non-nested model in a Bayesian framework in order to verify that any differences in pa-
- rameter estimates were not attributable to a change in statistical paradigms. The non-
- nested, random effects meta-analytic model is:

$$y_{i} = \mu + \alpha_{i} + \epsilon_{i}$$

$$\alpha_{i} \sim N(0, \sigma_{\alpha}^{2})$$

$$\epsilon_{ijk} \sim N(0, S_{ijk})$$
(S1)

- where y_i is the observed effect of the i^{th} response ratio, μ is the average true effect, α_i is
- the random effect of study i, ϵ_i is the residual, σ_{α}^2 is the between-study variance in true
- effect sizes [3]. The value of S_{ijk} is estimated by the reported standard error of the effect
- in study i.
- 56 We implemented the non-nested model using a Bayesian framework using JAGS inter-
- faced via the package rjags [8, 9]. The yield gap estimate of the Bayesian and frequentist
- non-nested models using Seufert et al.'s data did not differ quantitatively (compare a and
- 59 b, Fig. S2).
- Next, we re-analyzed Seufert et al.'s [1] data with our nested model, accounting for the
- 61 hierarchy in the meta-dataset. Our re-analysis shrunk the yield gap by 4 percentage points
- and increased the uncertainty around these estimates (Fig. S2c).

- Next, using our larger data-set we implemented their non-nested model in a Bayesian
- and frequentist framework. The frequentist analysis was implemented in the R package
- metafor. The non-nested analysis on our more comprehensive dataset gave an estimate
- of the organic to conventional ratio almost identical to Seufert et al. [1]'s original results
- using both frequentist (Fig. S2d) and Bayesian approaches (Fig. S2e).
- Finally, our model with our expanded dataset (1071 comparisons from 115 studies) shrinks
- 69 the estimate of the yield gap an additional 2% in comparison to the estimate using Seufert
- et al.'s data and our model (Fig. S2f). The variance around the estimate also shrinks
- slightly (compared to c), as expected with a larger data-set. Thus, difference in methods
- is twice as important as the difference in data in determining the reduction in the estimate
- of the organic to conventional yield gap between Seufert et al.'s analysis and ours.

₇₄ S2 Building our meta-analytic model

75 S2.1 Modeling framework

- ₇₆ We chose to use a Bayesian modeling framework because the existing Markov chain
- 77 Monte Carlo sampling methods [8] enabled us handle the complexity of the hierarchical
- model we constructed. Such methods did not exist for frequentist analyses at the time the
- ₇₉ analyses were conducted. Because the posterior is the product of the likelihood and the
- prior, when using uninformative (flat) priors as we have done here, the posterior should
- be approximately calibrated to frequentist results.

S2.2 Effect size

We chose to use the response ratio (the ratio of the mean outcome of the treatments of interest) because it quantifies the proportional difference between the treatments (in this case, the organic and conventional yields). In our analyses we used the natural logarithm of the response ratio because (1) it has an approximately normal sampling distribution, whereas the sampling distribution of the raw response ratio is skewed, and (2) deviations in the numerator and the denominator hold equal weight [10]. We then back transformed the model output to facilitate interpretation.

90 S2.3 Random/mixed model

- In a random and mixed effects meta-analytic models, the true study effects are assumed to come from a common distribution and, thus, such models provide inferences about the larger population of possible studies from which those included in the actual analysis are a random sample [11].
- In contrast to random and mixed effects models, fixed effect models assumes there is little
 heterogeneity of effect size estimates between studies. The study effects are therefore not
 modeled as being drawn from a common distribution. We did not conduct any fixed
 effect analyses because, when using fixed-effects models, the goal is to make a conditional
 inference only about the studies included in the meta-analysis [11].

... S2.4 Parameter inclusion

To determine the levels of hierarchy supported by the data, we sequentially added random effects and examined the posteriors of the parameters to determine the support for their inclusion. We also confirmed our selection with Deviance Information Criterion (DIC). The DIC can be problematic for hierarchical models because the effective number of parameters is not clearly defined [12, 13]. The DIC was therefore used in combination with a visual examination of the posterior distributions of the parameters to select the best supported model.

For the variance within and between year random effect distributions, we considered two parameterizations: 1) the variance terms, denoted σ_{η}^2 and σ_{β}^2 , respectively, were shared across all studies each with a Uniform(0,100) prior, and 2) the variance terms were study-specific (i.e., $\sigma_{\eta}^2[i]$ and $\sigma_{\beta}^2[i]$). In the latter case, the study-specific precision terms (1/variance) were assumed to be distributed according to a gamma distribution whose parameters were estimated. Uniform(0,100) priors were used for the coefficient of variation (1/ \sqrt{shape}) and the square root of the scale.

We first added a random effect of study and examined the posterior for σ_{α} (the stan-115 dard deviation of the common distribution from which the study effects are drawn). The 116 posterior was clearly differentiated from zero (Fig. S6a). We next added random variation 117 within a year and examined σ_{η} . We found it was also clearly different from zero (Fig. S6b). 118 The DIC was also smaller than when only a random effect of study was included (Tab. S1). 119 Next we allowed the within-year precisions to be study-specific and follow a gamma dis-120 tribution. We examined the coefficient of variation of the gamma distribution and found 121 it was clearly differentiated from zero (Fig. S6c). The DIC was also smaller than when a 122 single within year effect was shared across studies (Tab. S1). Lastly, we added a between 123 year random effect and examined σ_{β} . The posterior was concentrated at zero (Fig. S6d) so 124 we concluded there was insufficient support for including it in the model. The estimate 125 of the yield gap and its uncertainty did not differ substantially from when no between 126 year effect was included (Fig. S7), the DIC, however, was marginally smaller then when no between year random effect was included (Tab. S1).

S2.5 Weighting

The estimates of effect size from different studies will differ in their precision, or standard error. To handle these differences in precision, the sampling variance from each study are used as an estimate of the precision of the response ratio [3, 6, 10, 11, 14]. The variance of a response ratio is equal to

$$\frac{SD_{org}^2}{N_{org}\overline{X}_{org}^2} + \frac{SD_{conv}^2}{N_{conv}\overline{X}_{conv}^2}$$
 (S2)

Where SD is the standard deviation, \overline{X} is the mean and N is the sample size of the organic and conventional treatments [10]. This weighting increases the power of the test and the precision of the combined estimate [6]. We therefore use the estimated variance of log response ratios [10] to weight our effect sizes.

We tested the sensitivity of the analysis to weighting by implementing a non-nested, unweighted model using a frequentist framework [7] (our nested model does accommodate
unweighted variances). The mean estimate of the unweighted, non-nested model was
the same as the weighted, non-nested model, but the confidence intervals around the
parameter estimate were slightly smaller (Fig. S2g).

In many multi-year studies, we could not obtain the within-year variance among observations. In these cases, the variance of the means across years was used as an estimate of the observation-level sampling variance (which was also how [1] combined such data). Because this variance estimate lumps the between-year variation and the sampling variance

ation together, it is an overestimate of the sampling variance. For example, for a single study with multiple years,

$$y_i = \mu + \beta_i + \epsilon_i, i = 1, \dots, n \tag{S3}$$

$$\epsilon_i \sim N(0, \sigma_{\epsilon,i}^2)$$
 (S4)

$$\beta_i \sim N(0, \sigma_\beta^2)$$
 (S5)

where y_i is the log response ratio of the i^{th} year of a study, μ is the average, ϵ_i is the residual due to sampling variation, $\sigma_{\epsilon,i}^2$ is the sampling variance in year i, β_i is the random effect of the i^{th} year, and σ_{β}^2 is the variance of the distribution from which the year effects are drawn. The sample variance of the y_i is an estimate of the between-year variance (σ_{β}^2) plus the average of the sampling variances ($\sigma_{\epsilon,i}^2$). In a mixed or random effects meta-analysis, this becomes the estimate of the within-study variance (σ_i^2).

We conducted the hierarchical meta-analysis with and without the studies that reported only a between-year variance to examine the potential bias that under-weighting some studies might introduce. Excluding the studies that only provided the between-year variance decreased our meta-dataset to 56 studies and 839 observations. The yield gap estimated from the subset of the data was only 2 percentage point larger than the estimate from the full dataset with a 1.6 percentage point increase in the estimate of the variance (S2, compare f and h), as expected with less data. Including the studies without a true observation-level sampling variance estimate, therefore, does not substantially change our estimate of the yield gap.

58 S2.6 Sampling dependence

Studies included in meta analyses often employ ANOVA-style designs where multiple 159 treatments are compared against a single control. In these cases, a response ratio can be 160 calculated for each control-treatment pair. These response ratios, however, are not in-161 dependent, because they share a common control and therefore should not be included 162 separately in an analysis [4, 5]. When response ratios shared a common control, we cal-163 culated a combined response ratio (y_{ijk}) and corresponding standard error (σ_{ijk}) for the 164 entire study using the method presented in [Eq. 3 & 8, 4]. In the analysis with no explana-165 tory variables, 63% of the data were combined using the Lajeunesse method. 166

If the response ratios are grouped separately among categories in a mixed effects model the bias introduced by non-independence is minimized because the response ratios are not pooled together [4]. Therefore we did not aggregate response ratios in the analyses including explanatory variables when the response ratios were in different levels of the explanatory variable.

72 S3 Meta-datasets and publication bias

The inclusion criteria and the timing of literature searches differed between our study and that of de Ponti et al. [2] and Seufert et al. [1] and thus the meta-data sets vary in their coverage (Fig. S8). We used similar inclusion criteria to Seufert et al. but excluded (1) comparisons of organic yields with subsistence yields because the latter do not represent conventional agriculture, and (2) comparisons of yields not from the same year. Our meta-dataset encompasses 115 studies published between 1977 and 2012, of which Seufert et al.'s meta-data set is a subset (66 studies from 1980 to 2010). Unlike our study and that of

Seufert et al., de Ponti et al. included studies that did not report an estimate of sampling variance. de Ponti's less strict exclusion criteria enabled them to include 156 studies from 1989 to 2010.

For all three studies, we found that study publication year is correlated with the yield 183 ratio (Fig. S9). Specifically, there is a trend towards larger yield gaps with conventional out-performing organic in more recent studies. This effect is most pronounced in the data 185 used by Seufert et al., but is also present in de Ponti et al.'s and our own meta-dataset. This 186 trend could result from conventional yields increasing relative to organic yields through 187 time. Historically, research and development of organic cropping systems has been ex-188 tensively underfunded relative to conventional systems [15–17], so it is not inconceivable 189 that continued investment in conventional techniques has widened the yield gap through 190 time. 191

The same trend could also result if a publication bias favoring studies that report higher 192 conventional yields has increased through time. Not all studies are submitted for pub-193 lication and, of those that are, not all are accepted. Publication bias will result if stud-194 ies that show significant results are preferentially submitted and published, or if studies 195 are suppressed because the findings do not align with the interests of the researchers or 196 funding sources [18]. A bias would also result if researchers are choosing study systems 197 which they expect to show larger differences, and thus there is a bias in what studies are 198 conducted. Conversely, a bias would occur if the crop species of greatest interest to re-199 searchers is one that exhibits a large yield gap (e.g., cereals). Interesting, the proportion of 200 studies on cereals in the literature has increased through time (Fig. S10), which could ex-201 plain both the bias toward studies that report higher conventional yields and the increase 202 in the gap through time.

²⁰⁴ We visually assessed bias in our dataset and in the subset of data comprising the Seufert

et al. data using a funnel plot and QQ-plot [6]. Asymmetrical funnel plots may indicate a systematic difference between smaller and larger studies ("small study effects") which 206 may be due to publication bias [6, 18]. Our funnel plot revealed a slight asymmetry fa-207 voring small studies that report conventional yields are higher relative to organic, Fig. 208 S11a). The QQ-plot confirmed this observation: the observed quantiles are first below 209 then above the expected line, suggesting the observed data is gaining quantiles faster than 210 expected under a Gaussian model (Fig. S11b). This is because in the observed quantiles, 211 there is more mass in the tail compared to normally distributed data (i.e., the distribution 212 of response ratios is fat-tailed). That the lower quantiles are further from the line and even 213 steeper suggests that there is more mass in the lower tail (i.e., where conventional yields 214 are higher than organic). A fat-tailed distribution of response ratios could be due to bias. 215 Funnel and QQ-plots of the subset of data used by Seufert et al. were similar. 216

All of these assessments suggest that there is bias in the literature favoring studies that report conventional yields out-performing organic. Our results should therefore be viewed
as a potential overestimate of the yield gap. It is unclear, however, whether this bias is
due to publication bias or another type of bias such as a bias in what studies were conducted.

S4 Explanatory variable inclusion

To examine the support for including different explanatory variables in a model, various forms of model selection can be employed to compare models that contain different combinations of explanatory variables (e.g., [19]). Because the studies in our analysis did not consistently report many characteristics we wished to include in our model (e.g., nitrogen input, rotation type, tillage), we were unable to use such an approach, as it requires

that the data are consistent across models. We instead analyzed each explanatory variables able separately. Although not ideal, it is not possible to include all explanatory variables investigated here in one model and still use a traditional model selection framework. Organic and conventional yields were deemed significantly different from each other if the 95% credible interval of the yield ratio did not overlap one. Different levels of explanatory variables were considered to be significantly different if the posterior of the 95% credible interval of the difference between the group means did not overlap zero.

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Table S1: Parameter posteriors for models without explanatory variables. μ is the true mean response ratio across years and studies, σ_{α} is the standard deviation of the distribution from which the study random effects are drawn; σ_{η} is the standard deviation of the distribution from which the within year random effects are drawn; $CV_{\sigma_{\eta}}$ is the coefficient of variation of the gamma from which the study-specific within-year variance are drawn; and σ_{β} is the standard deviation of the distribution of random between year effects. Values of Rhat < 1.1 indicate convergence. Lower Deviance Information Criterion (DIC) indicates better model fit to the data.

Parameter	Posterior mean	Posterior standard deviation	95% CI	Rhat				
Study random effect, DIC=1684.8								
μ	0.795	0.027	0.742 - 0.848	1.001				
σ_{α}	0.341	0.026	0.294 - 0.396	1.001				
Study and v	Study and within year random effects, DIC= -565.9							
μ	0.788	0.021	0.749 - 0.829	1.001				
σ_{α}	0.188	0.024	0.144 - 0.239	1.001				
σ_{η}	0.312	0.011	0.291 - 0.333	1.001				
Study and study-specific within year random effects, DIC= -618.0								
μ	0.808	0.019	0.771 - 0.845	1.001				
σ_{α}	0.189	0.023	0.145 - 0.237	1.001				
CV_{σ_η}	1.155	0.135	0.907 - 1.436	1.001				
Study, study-specific within year, and between year random effects, DIC= -621.2								
μ	0.808	0.019	0.770 - 0.846	1.001				
σ_{α}	0.186	0.024	0.142 - 0.234	1.001				
CV_{σ_η}	1.157	0.136	0.907 - 1.440	1.001				
σ_{eta}	0.041	0.027	0.002 - 0.098	1.001				

Table S2: Yield estimates compared with previous studies. Categories of explanatory variables are arranged from the smallest to the largest difference between organic and conventional yields. Bold categories indicate conventional yields are significantly larger than organic yields. Unbolded categories indicate organic and conventional yields are not significantly different. \ll between two categories indicates the two categories are significantly different from each other. * includes both plant and animal products. ** de Ponti et al. [2] did not report significance. *** Confidence intervals were calculated from the standard deviation and the number of yield comparisons reported by de Ponti et al. [2].

Variable	Ponisio et al.	Seufert et al. [1]	de Ponti et al. [2]	Badgley et al. [20]
Overall	$80.8\% \pm 3.7\%$	$75\% \pm 4\%$	80% ± 1% * **	$132\% \pm 1\%*$
Development	developing developed	developed ≪ developing	developing ** developed	developing ≪ developed
Crop type	fruits and nuts oilseed crops cereals vegetables roots & tubers	fruits and nuts oilseed crops ≪ cereals vegetables	vegetables ** cereals roots & tubers oilseed crops fruits	NA
Crop species	apple oat tomato soybean maize ≪ wheat barley potato	soybean maize tomato barley wheat	soybean ** maize oat tomato wheat barley apple	NA
Legume	legume non-legume	legume ≪ non-legume	NA	NA
Plant-type	perennial annual	perennial ≪ annual	NA	NA

Table S3: The impact of management practices on the yield gap compared with previous studies. Categories of explanatory variables are arranged from the smallest to the largest difference between organic and conventional yields. Bold categories indicate conventional yields are significantly larger than organic yields. Unbolded categories indicate organic and conventional yields are not significantly different.≪ between two categories indicates the two categories are significantly different from each other.

Variable	Ponisio et al.	Seufert et al. [1]
Nitrogen input	$\begin{array}{l} \text{similar N input} \ll \\ \text{more N organic} \ll \\ \text{more N conventional} \end{array}$	more organic ≪ more conventional similar N input
Poly/monoculture	organic polyculture ≪ both monoculture both polyculture	organic polyculture both monocultures both polyculture
Rotations	more organic ≪ similar no rotations	more organic similar ≪ no rotations
Organic fertilizer type	organic fertilizer animal mix plant	organic fertilizer mix animal plant
Conventional system type	low input ≪ high input	low input ≪ high input
Organic system type	certified transitional biodynamic organic standards	certified biodynamic ≪ organic standards transition
Best management practices	no yes	yes ≪ no
Irrigation	rain-fed irrigated	rain-fed ≪ irrigated
Soil pH	acidic neutral ≪ strongly acidic strongly alkaline	weak acidic to alkaline \ll strongly acidic \ll strongly alkaline
Time since conversion	0-3 years > 7 years 4-7 years	> 7 years 4-7 years ≪ 0-3 years

Table S4: The impact of study quality indicators on the yield gap compared with previous studies. Categories of explanatory variables are arranged from the smallest to the largest difference between organic and conventional yields. Bold categories indicated they are significantly less than one. \ll between two categories indicates the two categories are significantly different from each other.

Variable	Ponisio et al.	Seufert et al. [1]
Duration of study	> 10 seasons	> 10 seasons
	3-5 season	6-10 seasons
	6-10 seasons	3-5 seasons
	1-2 seasons	1-2 seasons
Literature type	journal	grey ≪
	grey	journal
Comparability of system	comparable	comparable
	not comparable	not comparable
Study type	on-farm trial	on-farm trial
	experimental	survey
	survey	experimental

Table S5: A list of the studies included in the meta-analysis including the crops with yield comparisons, the country the study was conducted in, whether the study was included in Badgley et al. [20], Seufert et al. [1] or de Ponti et al. [2], and the number of organic to conventional yield comparisons extracted from each study.

Study	Crop	Country	Comparisons	In Badgley	In Seufert	In de Ponti
[21]	apple	Brazil	2	no	no	yes
[22]	pepper	India	5	no	yes	no
[23]	spring	United States	40	no	no	yes
	wheat,					
	maize, soy-					
	bean					
[24]	wheat	Switzerland	2	no	no	yes
[25]	endive	Croatia	2	no	no	yes
[26]	rice, pigeon	India	2	no	no	no
	pea					
[27]	apple	Switzerland	4	no	yes	yes
[28]	barley	Switzerland	4	no	yes	no
[29]	cabbage	Switzerland	4	no	yes	no
[30]	sugar beet	Switzerland	4	no	yes	no
[31]	tomato	Brazil	2	no	no	yes
[32]	wheat	Czech Republic	6	no	no	yes
[33]	cotton	India	1	no	yes	no

[34]	bean,	Italy	41	no	no	no
	cauliflower,					
	fennel, let-					
	tuce, melon,					
	tomato					
[35]	maize, soy-	United States	32	no	yes	no
	bean, wheat					
[36]	spinach	Turkey	34	no	yes	no
[37]	maize,	United States	6	no	yes	yes
	tomato					
[38]	maize, soy-	United States	4	no	no	no
	bean					
[39]	melon, wa-	Turkey	6	no	no	yes
	termelon					
[40]	tomato,	Italy	2	no	no	no
	wheat					
[41]	maize, soy-	United States	2	yes	yes	no
	bean					
[42]	maize, soy-	United States	4	no	no	no
	bean					
[43]	rice	France	9	no	no	no
[44]	hazelnut	Turkey	1	no	yes	no
[45]	tomato,	United States	45	no	yes	yes
	maize					
[46]	maize, soy-	United States	2	yes	yes	no
	bean					

[47]	barley, wheat	Denmark	78	no	VOC	MOS
	•				yes	yes
[48]	tomato	United States	1	yes	yes	no
[49]	maize	United States	3	no	yes	no
[50]	safflower	United States	1	no	no	no
[51]	maize	Greece	9	no	no	yes
[52]	flax	Canada	1	no	yes	yes
[53]	cotton,	India	12	no	yes	yes
	chili, wheat,					
	soybean,					
	sorghum,					
	maize					
[54]	soybean,	United States	38	no	yes	yes
	wheat, maize					
[55]	pea grain	Bulgaria	1	no	no	
[56]	strawberry	United States	3	no	yes	yes
[57]	maize	United States	3	no	yes	no
[58]	wheat	India	12	no	yes	no
[59]	soybean,	India	8	no	no	no
	wheat					
[60]	kiwi	New Zealand	2	no	no	no
[61]	apricot	Turkey	3	no	no	yes
[62]	fig	Turkey	3	no	no	no
[63]	raisins	Turkey	3	no	no	yes
[64]	strawberry	Canada	8	no	yes	no
[65]	maize, soy-	United States	6	no	no	no
	bean, oat					
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[66]	chard,	Spain	4	no	yes	no
	pumpkin,					
	tomato, bean					
[67]	rye	Germany	2	no	yes	no
[68]	sweet potato	United States	2	no	no	yes
[69]	wheat, oat,	Estonia	3	no	no	yes
	barely					
[70]	potato	Estonia	2	no	yes	no
[71]	banana	Ecuador	1	no	yes	no
[72]	tomato	Taiwan	3	no	yes	yes
[73]	buckwheat	Czech Republic	9	no	no	no
[74]	wheat	Canada	1	no	yes	yes
[75]	barley, wheat	Sweden	2	no	yes	yes
[76]	wheat	Australia	2	no	yes	yes
[77]	wheat, oat,	United States	4	no	no	no
	soybean,					
	maize					
[78]	soybean	China	1	no	yes	no
[79]	maize	Iran	12	no	no	no
[80]	oat, wheat,	Norway	86	no	no	no
	barley					
[81]	maize, soy-	United States	4	no	yes	no
	bean					
[82]	cauliflower	Italy	20	no	no	yes
[83]	maize	United States	26	no	yes	no

[84]	maize, soy-	United States	3	no	yes	no
	bean					
[85]	maize,	United States	28	no	yes	no
	tomato,					
	wheat					
[86]	coffee	Costa Rica	3	no	yes	yes
[87]	strawberry	Turkey	5	no	no	yes
[88]	potato,	Switzerland	18	yes	yes	yes
	wheat					
[89]	potato	Italy	1	no	no	yes
[90]	soybean,	United States	10	no	no	no
	maize, oat					
[91]	grapes	Italy	3	no	no	yes
[92]	coffee	Nicaragua	3	no	yes	
[93]	maize,	United States	5	no	yes	no
	tomato					
[94]	sunflower	Italy	3	no	yes	yes
[95]	wheat	Italy	2	no	yes	no
[96]	potato	Portugal	2	no	no	yes
[97]	leek	Croatia	4	no	no	yes
[98]	apple	United States	4	no	no	yes
[99]	apple	United States	4	no	no	no
[100]	maize	Italy	3	no	yes	no
[101]	tomato	United States	6	no	yes	no
[102]	lettuce	Turkey	5	no	yes	yes
[103]	tomato	Turkey	5	no	no	no

[104]	maize, oat,	United States	24	no	yes	yes
	soybean,					
	alfalfa					
[105]	maize, soy-	United States	8	no	yes	no
	bean					
[106]	peach	United States	2	no	no	no
[107]	carrot, beet-	Germany	8	yes	yes	no
	root, rye,					
	potato					
[108]	rye	Germany	2	no	yes	no
[109]	wheat	United States	1	no	yes	no
[110]	apple	United States	2	yes	yes	yes
[111]	tomato	Tunisia	12	no	yes	yes
[112]	sweet corn,	United States	3	no	yes	no
	cucumber,					
	pepper					
[113]	wheat	Australia	1	no	yes	no
[114]	maize, soy-	United States	82	no	no	yes
	bean					
[115]	cabbage,	Canada	5	no	yes	no
	onion, sweet					
	corn, bean,					
	tomato					
[116]	potato	Latvia	2	no	no	no
[117]	muskmelon	China	4	no	no	no
[118]	wheat	Canada	2	no	yes	yes

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elephant foot	India	5	no	no	no
yam					
apple	United States	1	no	yes	no
cotton	United States	2	no	yes	no
pepper	Poland	3	no	no	yes
wheat, oat,	Estonia	9	no	no	yes
barley					
maize, wheat	United States	13	no	yes	yes
tomato,	United States	8	no	no	no
maize, saf-					
flower, bean					
rye, oat,	Denmark	72	no	no	no
onion, carrot,					
cabbage,					
lettuce					
oat, barley	Sweden	3	no	yes	no
sweet potato	United States	3	no	no	yes
coffee	Nicaragua	1	no	yes	no
lettuce	United States	2	no	yes	yes
cabbage, car-	Canada	6	yes	yes	yes
rot					
sweet corn,	Canada	6	yes	yes	yes
potato					
flax, wheat	Canada	13	no	yes	yes
grapes	Australia	1	no	no	no
	yam apple cotton pepper wheat, oat, barley maize, wheat tomato, maize, saf- flower, bean rye, oat, onion, carrot, cabbage, lettuce oat, barley sweet potato coffee lettuce cabbage, car- rot sweet corn, potato flax, wheat	yam apple United States cotton United States pepper Poland wheat, oat, Estonia barley maize, wheat United States tomato, United States maize, saf- flower, bean rye, oat, Denmark onion, carrot, cabbage, lettuce oat, barley Sweden sweet potato United States coffee Nicaragua lettuce United States cabbage, car- rot sweet corn, Canada potato flax, wheat Canada	yam apple United States 1 cotton United States 2 pepper Poland 3 wheat, oat, Estonia 9 barley maize, wheat United States 13 tomato, United States 8 maize, safflower, bean rye, oat, onion, carrot, cabbage, lettuce oat, barley Sweden 3 sweet potato United States 3 coffee Nicaragua 1 lettuce United States 2 cabbage, carrot sweet corn, Canada 6 potato flax, wheat Canada 13	yam apple United States 1 no cotton United States 2 no pepper Poland 3 no wheat, oat, Estonia 9 no barley maize, wheat United States 13 no tomato, United States 8 no maize, saf- flower, bean rye, oat, Denmark 72 no onion, carrot, cabbage, lettuce oat, barley Sweden 3 no sweet potato United States 3 no coffee Nicaragua 1 no lettuce United States 2 no cabbage, car- rot sweet corn, Canada 6 yes potato flax, wheat Canada 13 no	yam apple United States 1 no yes cotton United States 2 no yes pepper Poland 3 no no wheat, oat, Estonia 9 no no barley maize, wheat United States 13 no yes tomato, United States 8 no no maize, saf- flower, bean rye, oat, onion, carrot, cabbage, lettuce oat, barley Sweden 3 no yes sweet potato United States 3 no no coffee Nicaragua 1 no yes lettuce United States 2 no yes cabbage, car- rot sweet corn, Canada 6 yes yes potato flax, wheat Canada 13 no yes

[135]	soybean,	United States	3	no	yes	no
	wheat, maize					

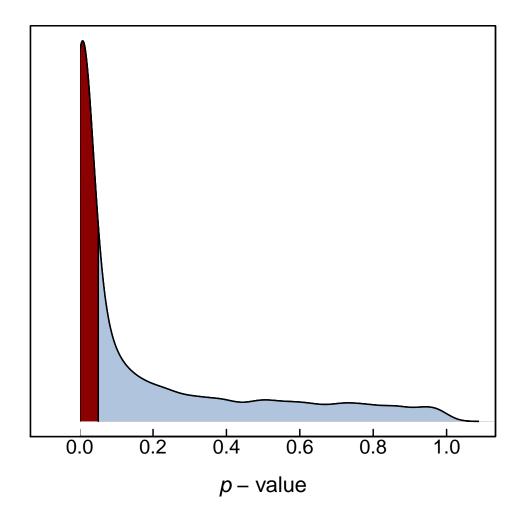


Figure S1: The distribution of *p*-values when the null hypothesis was forced to be true using the data and analysis type present in Seufert et al. [1]. If the analysis procedure was valid for these data, the distribution of P-values should be uniform between 0 and 1. Instead it is sharply shifted toward low P-values. In over 50% of simulations, the null hypothesis was rejected using a nominal Type I error rate of 0.05 (red region above).

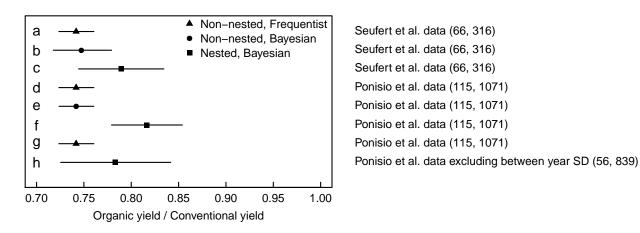


Figure S2: The effect of different models, data, and statistical paradigms on the organic to conventional yield ratio (from top to bottom): (a) Seufert et al.'s [1] non-nested analysis and data not accounting for pseudo-replication (organic to conventional ratio $75\% \pm 4\%$), (b) Seufert et al.'s [1] non-nested analysis and data conducted in a Bayesian framework to show the comparability in results between the two paradigms $(75\% \pm 4\%)$, (c) our nested analysis accounting for the hierarchy of the meta-dataset on the data published by Seufert et al. [1] in a Bayesian framework $(79\% \pm 4.5\%)$, (d) Seufert et al.'s [1] non-nested analysis on our larger dataset in a frequentist framework $(74\% \pm 2\%)$, (e) Seufert et al.'s [1] non-nested analysis on our larger dataset in a Bayesian framework $(74\% \pm 2\%)$, (f) our analysis and full meta-dataset $(80.8\% \pm 3.7\%)$, (g) an unweighted, non-nested analysis on our meta-dataset $(74\% \pm 2\%)$, and (h) our nested analysis on our meta-dataset excluding studies where the between year variation was used as an estimate of the observation-level sampling variance $(78\% \pm 5.8\%)$. The number of studies and organic to conventional comparisons are indicated in parentheses. We could not re-analyze de Ponti et al.'s [2] data because they did not extract the response ratio variances necessary for our model.

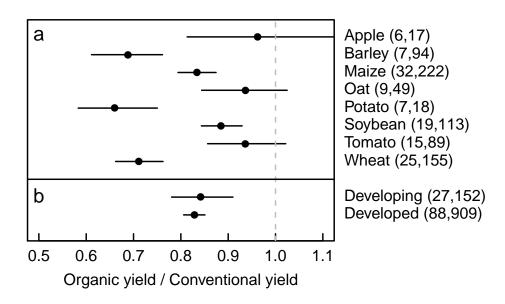


Figure S3: The influence of (a) crop species and (b) country development on the organic-to-conventional yield ratio. Values are mean effect sizes with 95% credible intervals. The number of studies and observations in each category are shown in parentheses. Only categories with at least 10 yield comparisons from greater than 5 studies are shown.

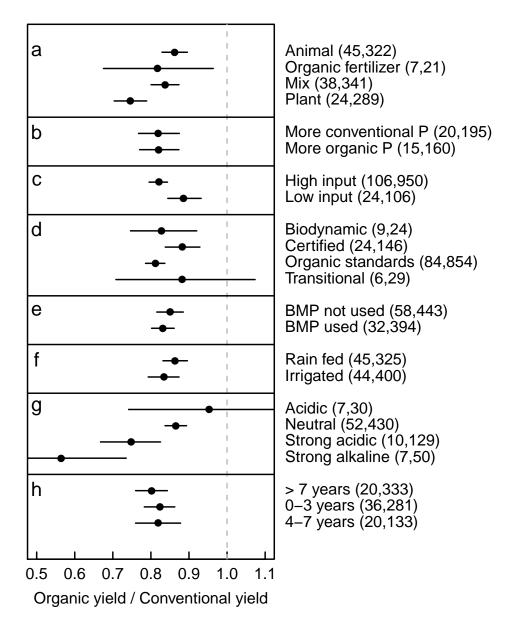


Figure S4: The influence of management practices on the organic-to-conventional yield ratio including (a) organic fertilizer type, (b) phosphorus input, (c) conventional system type, (d) organic system type, (e) the use of best management practices, (f) irrigation, (g) soil pH, and (h) years since conversion to organic agriculture. With respect to the low input conventional system type, the papers self-identified their system as "low input", but the crops generally received substantial external inputs. The inputs were generally only applied if needed (i.e., if a soil test suggested the soil was nitrogen deficient, fertilizer was applied), and not on a set schedule as is common in high-input conventional systems. Values are mean effect sizes with 95% credible intervals. The number of studies and observations in each category are shown in parentheses. Only categories with at least 10 yield comparisons from greater than 5 studies are shown.

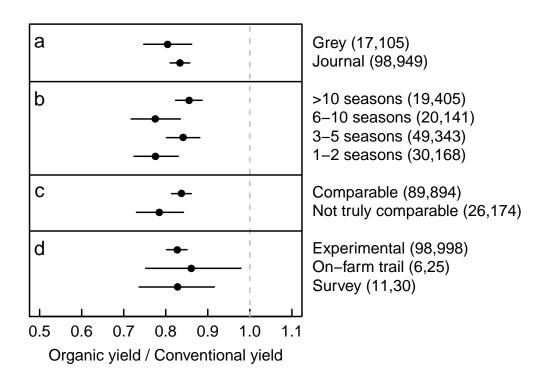


Figure S5: The sensitivity of the organic-to-conventional yield ratio to study quality factors including (a) publisher type, (b) duration of the study, (c) comparability of the organic and conventional treatments, and (d) the type of study. Values are mean effect sizes with 95% credible intervals. The number of studies and observations in each category are shown in parentheses. Only categories with at least 10 yield comparisons from greater than 5 studies are shown.

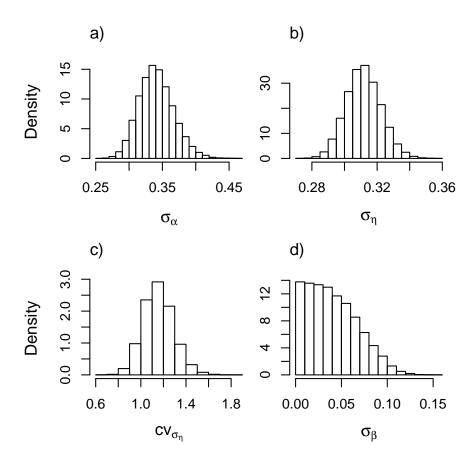


Figure S6: The posterior distributions for the random effect of a) study (σ_{α}) ; b) response ratios within a year (σ_{η}) ; c) response ratios within a year where the within year variance is study-specific, $CV_{\sigma_{\eta}}$ is the coefficient of variation $(1/\sqrt{shape})$ of the gamma distribution (this model is most supported by the data); and d) between year (σ_{β}) . Including a between-year variance term was not supported by the data (the posterior for (σ_{β}) is not differentiated from zero).

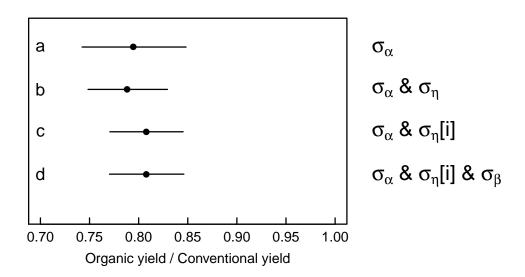


Figure S7: The sensitivity of the yield gap to including different levels of hierarchy in the model. The random effects included in the model are: a) study (σ_{α}) ; b) study and response ratios within a year (σ_{η}) ; c) study and response ratios within a year where the within year variance is study-specific $(\sigma_{\eta}[i])$ (this model is most supported by the data); and d) study, study-specific within-year variance, and between year (σ_{β}) . Including a between-year variance term was not supported by the data (the posterior for $(\sigma_{\beta}$ is not differentiated from zero). Values are the posterior mean with 95% credible intervals.

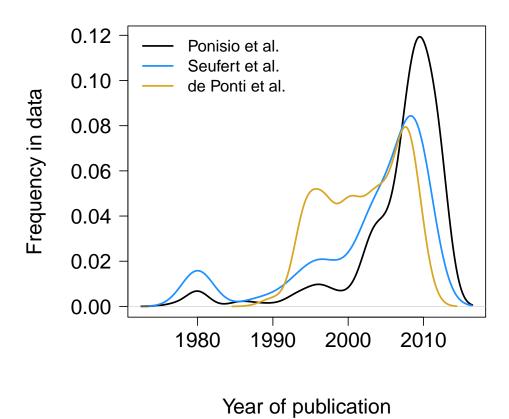


Figure S8: A comparison of the frequency of organic to conventional yield comparisons published in different years from our study, de Ponti et al. [2] and Seufert et al. [1].

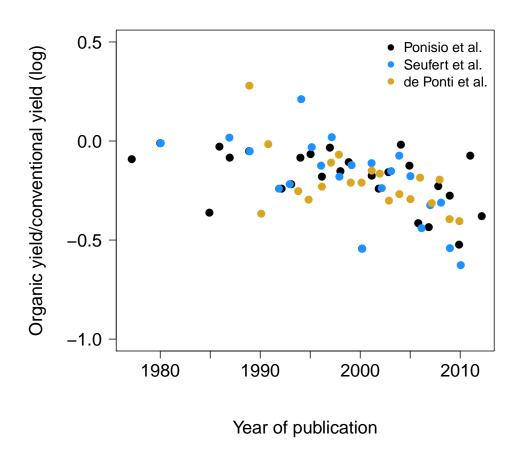


Figure S9: The mean organic to conventional yield ratio (log) for each year studies were published from our study, de Ponti et al. [2] and Seufert et al. [1]

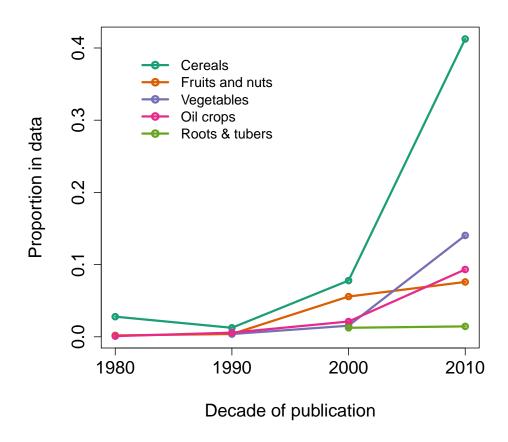


Figure S10: The proportion of observations of each crop species (binned by decade) in our metadataset.

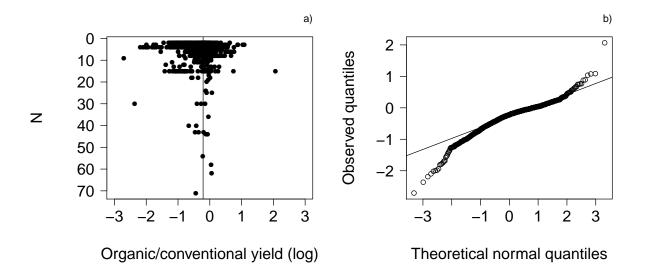


Figure S11: A funnel plot and QQ-plot of the response ratios in our study's meta-dataset.