27. O'Dwyer LM, Russell, M., Bebell, D., & Seeley, K. . Examining the Relationship between Students'

Mathematics Test Scores and Computer Use at Home and at School. *Journal of Technology, Learning, and Assessment* 2008;**6**(5).

WEB MATERIAL

eAppendix 1: SAS code for simulation

```
/* Create datasets with variable # of groups & variable # of individuals in a
group */
%MACRO create simulated dataset (ngroups=, groupsize=);
data simulation parms;
      retain condition;
      condition=0;
      do ngroups=&ngroups;
            do groupsize=&groupsize;
                  nobs=ngroups*groupsize;
                  condition=condition+1;
                  output;
            end;
      end;
run:
data simulated chs;
      set simulation parms;
      retain ngroups groupsize;
      keep condition serial group i poverty no healthy food available
income rand male black age yrs;
      /* First, generate values of individual-level poverty */
      do group=1 to ngroups;
            if ranuni(-1) < .25 then no_healthy_food_available = 1; else</pre>
no healthy food available = 0;
            income group base = rand('normal', 10.5581293, 0.4020605);
            do i=1 to groupsize;
                  male = ranuni(-1) < .442;
                  black = ranuni(-1) < .512;
                  age_yrs = round(rand('uniform') * (65-18) + 18);
                  income individual = rand('normal', 0, .5);
                  log income = income group base + income individual;
                  income = exp(log income);
                  /* 2000 Federal Poverty Level for 4-person family */
                  poverty = income < 17050;</pre>
                  output;
            end;
      end;
run;
proc freq data=simulated chs noprint;
      table poverty*group/outpct out=simulated group pov;
run;
data group pov;
      set simulated group pov;
```

```
if poverty eq 1;
run;
proc univariate data=simulated chs noprint;
      var income;
      output out=group income mean=mean median=median;
      by group;
run;
data simulated chs;
     merge simulated chs group pov(rename= (PCT COL=group pov))
group income(rename= (mean=group mean income));
      by group;
      if group pov = \cdot then group pov = 0;
      keep condition serial group group pov i poverty bmi obese
no healthy food available income group mean income rand male black age yrs;
      label group pov = 'group-level percent poverty'
                  group mean income = 'group-level mean income';
run;
data simulated chs;
      set simulated chs;
      /* Generate BMI similar to real data. */
      e=rand('normal',0,.70711);
      bmi = 16 + .92*male + 1.26* black + 0.069*age yrs + .00005*(100000-
group mean income) + 10*e;
      /* Cut off the implausible low BMI tail -- this should not be many
individuals */
      if bmi < 15 then delete;
      if bmi < 30 then obese = 0; else obese = 1;
      rand = ranuni(-1);
      /* For this simulation, keep only 5% as the 'sampled' individuals */
      if rand > 0.05 then delete;
run;
%MEND;
/* Add noise to a measure */
%macro create measurement error(input=, output=, input var=, output var=,
percent noise=);
data &output.;
      set &input.;
      &output var. = &input var. * (100+(&percent noise.*(.5-ranuni(-
1))))/100;
run;
%mend;
/* Generate a result at a given level of noise */
%macro generate result(in=, cur noise=, input var=);
      %create_measurement_error(input=&in., output=internal,
input var=&input var., output var=as measured, percent noise=&cur noise.);
      proc univariate data=internal noprint;
            var as measured;
            output out=internal with group error mean=mean with error
median=median;
            by group;
      run;
      proc sort data=internal with group error; by group;
     proc sort data=internal; by group;
      data internal;
```

```
merge internal internal with group error(rename=
(mean with error=group mean with error));
            by group;
      run;
      proc sql;
            select count(*) into :nobs from internal;
            select count(unique(group)) into :ngroups from internal;
      quit;
      proc mixed data=internal cl covtest noclprint;
            class group;
            model bmi= age yrs male black income
group mean with error/solution ddfm=bw;
            random intercept/subject=group;
            ods output SolutionF=solution;
      run;
      data solution1;
            set solution;
            noise = &cur noise.;
            keep Estimate StdErr noise;
            where Effect eq 'group mean with erro';
      run;
      data results;
            set results solution1;
      run:
      data results;
            set results;
            where noise ne .;
      run;
/* Generate results for lots of levels of noise */
%macro generate error results (in=, input var=, result=, start=, end=, by=,
hits per=1);
data results; run;
%local cur noise;
%let cur noise = &start.;
%do %while(&cur noise. <= &end.);
      %put Percent Noise is &cur noise.;
      %do i=1 %to &hits per;
            %generate result(in=&in., cur noise=&cur noise.,
input var=&input var.);
      %end;
      %let cur noise = %eval(&cur noise. + &by.);
      DM output 'clear';
      DM log 'clear';
%end;
/* Compare the results with noise to results without */
%MACRO process results();
data no error results;
     set results;
      where noise eq 0;
      true estimate = estimate;
     true std error = stderr;
     keep true estimate true std error;
run;
```

```
proc sql;
      create table results with bias as
      select results.*, no error results.true estimate as true beta,
no error results.true std error as true err
     from results, no error results;
run; quit;
data results with bias;
     set results with bias;
      percent overestimation estimate = 100*((estimate*1.0/true beta) - 1);
      percent overestimation stderr = 100*((stderr*1.0/true err) - 1);
run;
%MEND;
/* Make datasets of variable sizes and analyze them */
%macro vary dataset size(start ngroups=, end ngroups=, by ngroups=,
start groupsize=, end groupsize=, by groupsize=);
%local groupsize;
%let groupsize = &start groupsize.;
%do %while(&groupsize <= &end groupsize);
      %put groupsize is &groupsize.;
      %local ngroups;
      %let ngroups = &start ngroups.;
      %do %while(&ngroups <= &end ngroups);
            %put ngroups is &ngroups.;
            %create simulated dataset (ngroups=&ngroups.,
groupsize=&groupsize.);
            %generate error results (in=simulated chs, input var=income,
result='Estimate', start=0, end=20, by=20, hits per=20);
            %process results();
            data run results;
                  set results with bias;
                  ngroups = &ngroups;
                  groupsize = &groupsize;
            run;
            data varied datasetsize results;
                  set varied datasetsize results run results;
            %let ngroups = %eval(&ngroups. + &by ngroups.);
      %let groupsize = %eval(&groupsize. + &by groupsize.);
%end:
%mend:
/* Okay, now do the simulation */
data varied datasetsize results; run;
%vary dataset size(start ngroups=3, end ngroups=100, by ngroups=1,
start groupsize=4000, end groupsize=4000, by groupsize=1);
title1;
axis1 label=(a=90 font="helvetica" height=24pt "Effect Percent
Overestimation") value=(height=2) order=(-3 to 3 by 1);
axis2 label=(font="helvetica" height=24pt "Number of Neighborhoods" )
value=(height=2);
symbol1 interpol=none value=plus;
proc gplot data=varied datasetsize results;
      plot percent overestimation estimate*ngroups/vaxis=axis1 haxis=axis2
overlay legend=legend1;
run; quit;
```

eAppendix 2: Mathematical basis for observed effect of misclassification in individual level data aggregated to create neighborhood level variables.

A simple fixed effect model with no cross-level interaction has the general form:

$$h(E[Y_{ij}|X_{ij},\bar{P}_i]) = \alpha + \beta X_{ij} + \gamma \bar{P}_i$$

Where the individual-level variable (X_{ij}) and group-level proportion (\overline{P}_i) both contribute to the dependent variable. Following Brenner, et al.³, if \overline{P}_i was aggregated from an individual-level measure with sensitivity Se and a specificity Sp, then the observed value \hat{P}_i relates to the true value as follows:

$$\widehat{P}_i = \overline{P}_i * Se + (1 - \overline{P}_i) * (1 - Sp)$$

$$= \overline{P}_i * Se + 1 - Sp - \overline{P}_i + \overline{P}_i * Sp$$

$$= \overline{P}_i (Se + Sp - 1) + 1 - Sp$$

And therefore:

$$\overline{P}_i = \frac{\widehat{P}_i + Sp - 1}{Se + Sp - 1}$$

Which can be substituted into the model as follows:

$$\begin{split} h\big(E\big[Y_{ij}|X_{ij},\hat{P}_i,Se,Sp\big]\big) &= \alpha + \beta X_{ij} + \gamma \left(\frac{\hat{P}_i + Sp - 1}{Se + Sp - 1}\right) \\ &= \alpha + \beta X_{ij} + \gamma \left(\frac{\hat{P}_i}{Se + Sp - 1}\right) + \gamma \left(\frac{Sp - 1}{Se + Sp - 1}\right) \\ &= \alpha_{error} + \beta X_{ij} + \gamma_{error}\hat{P}_i \end{split}$$
 where $\alpha_{error} = \alpha + \gamma \left(\frac{Sp - 1}{Se + Sp - 1}\right)$ and $\gamma_{error} = \gamma \frac{1}{Se + Sp - 1}$

And thus, similar to what has been shown for ecologic analysis³, the slope of a fixed effect model regression line for a group level variable constructed by aggregating non-differentially misclassified individual-level values will be elevated by $\frac{1}{Se+Sp-1}$, the intercept will be biased only in the event of

imperfect specificity, and estimates for other parameters in the model are unaffected. Furthermore, the extent of bias is independent of the base prevalence of poverty and distribution of the poverty data.

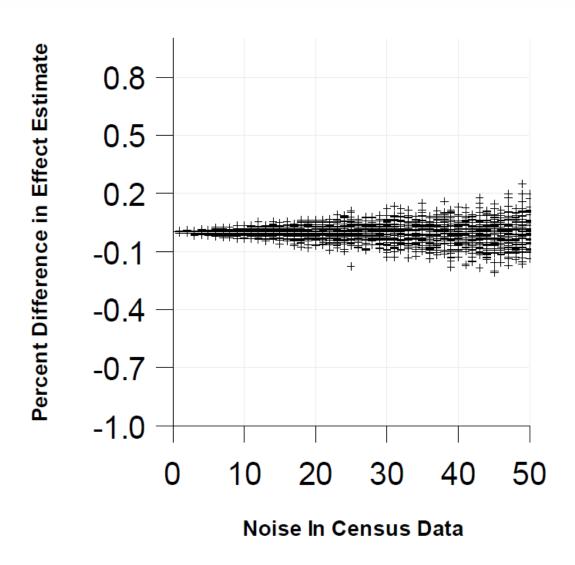
eAppendix 3: Mathematical basis for quantile-based comparisons' immunity from this bias

We return to the definition of misclassification:

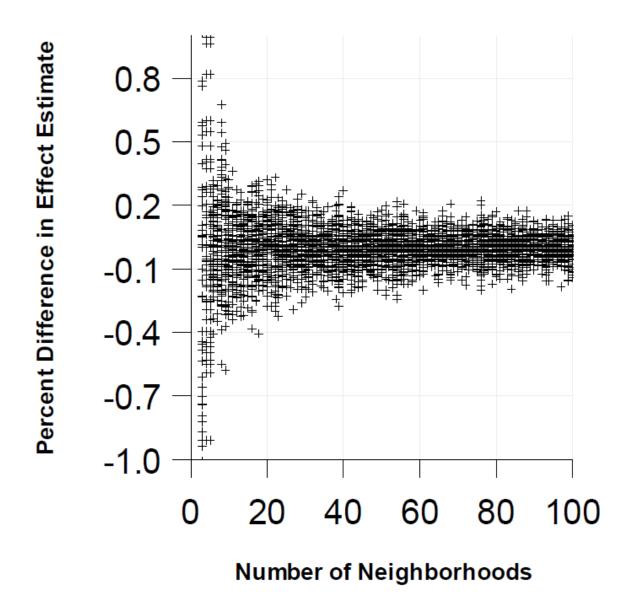
$$\hat{P}_i = \overline{P}_i * Se + (1 - \overline{P}_i) * (1 - Sp)$$

From this equation, we observe that non-differential misclassification does not affect the rank-order of \hat{P} values; that is, if $\overline{P}_l > \overline{P}_j$, then $\hat{P}_i > \hat{P}_j$. As a result, quantiling data by \hat{P} values results in the same categorization as quantiling by \overline{P} would; thus, any non-differential misclassification in the underlying individual level data has no effect.

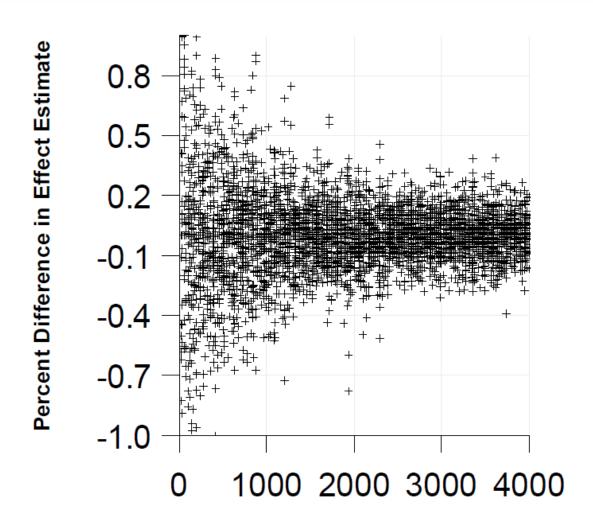
eFigure 1: Bias in effect estimates for the association between neighborhood mean income and Body Mass Index across repeated simulations, for increasing levels of measurement error in individual-level income data in population Census data.



eFigure 2: Bias in effect estimates for the association between neighborhood mean income and Body Mass Index across repeated simulations, by number of neighborhoods included in a study at a fixed level of measurement error in the individual level income data.



eFigure 3: Bias in effect estimates for the association between neighborhood mean income and Body Mass Index across repeated simulations, by the population size of the neighborhoods.



Residents in Each Neighborhood