Supplementary Material for "Regularization Methods for High-Dimensional Instrumental Variables Regression With an Application to Genetical Genomics"

Proof of Theorem 4

We first present two lemmas that are essential to the proof of Theorem 4, which concern the concentration of the empirical covariance matrix $\hat{\mathbf{C}}$ around its population version **C** and the score vector

$$
\frac{1}{n}\widehat{\mathbf{X}}^T(\mathbf{y}-\widehat{\mathbf{X}}\boldsymbol{\beta}_0) = \frac{1}{n}\widehat{\mathbf{X}}^T\boldsymbol{\eta} - \frac{1}{n}\widehat{\mathbf{X}}^T(\widehat{\mathbf{X}}-\mathbf{X})\boldsymbol{\beta}_0
$$

around zero. These lemmas can be viewed as generalizations of Lemma A.3 and inequality (A.15), respectively. For ease of presentation, we condition on the event of probability $1 - \pi_0$ that the two error bounds in Condition (C4) hold, and incorporate the probability π_0 into the result by the union bound.

Lemma S.1. *Under Conditions (C4)–(C6), if* $\mu_0 > 0$ *and the first-stage error bounds* e_1 *and e*² *satisfy*

$$
s(2Le_1 + e_2) \le \frac{\alpha}{(4-\alpha)\varphi} \wedge \frac{(\mu_0/2)^2}{s},\tag{S.1}
$$

then with probability at least $1 - \pi_0$ *, the following inequalities holds:*

$$
\|(\widehat{\mathbf{C}}_{SS})^{-1}\|_{\infty} \le \frac{4-\alpha}{2(2-\alpha)}\varphi,\tag{S.2}
$$

$$
\|\widehat{\mathbf{C}}_{S^cS}(\widehat{\mathbf{C}}_{SS})^{-1}\|_{\infty} \le \left\{ \left(1 - \frac{\alpha}{2}\right) \frac{\rho'(0+)}{\rho'_{\mu}(b_0/2)} \right\} \wedge (2cn^{\nu}),\tag{S.3}
$$

and

$$
\Lambda_{\min}(\hat{\mathbf{C}}_{SS}) > \mu \tau_0. \tag{S.4}
$$

Proof. It follows from the arguments in the proof of Lemma A.1 and Condition (C4) that

$$
\max_{1 \le i,j \le p} \frac{1}{n} |\widehat{\mathbf{x}}_i^T \widehat{\mathbf{x}}_j - (\mathbf{Z} \boldsymbol{\gamma}_{0i})^T \mathbf{Z} \boldsymbol{\gamma}_{0j}| \le 2Le_1 + e_2.
$$

Consequently, by the assumption (S.1),

$$
\varphi \|\widehat{\mathbf{C}}_{SS} - \mathbf{C}_{SS}\|_{\infty} \leq \varphi s(2Le_1 + e_2) \leq \frac{\alpha}{4 - \alpha} \tag{S.5}
$$

and

$$
\varphi \|\widehat{\mathbf{C}}_{S^c S} - \mathbf{C}_{S^c S}\|_{\infty} \le \frac{\alpha}{4 - \alpha}.
$$
\n(S.6)

Then inequality (S.2) follows as in the proof of Lemma A.3.

To show inequality $(S.3)$, by $(S.2)$, $(S.5)$, $(S.6)$, and Condition $(C6)$, we have

$$
\begin{split} &\|\hat{\mathbf{C}}_{S^{c}S}(\hat{\mathbf{C}}_{SS})^{-1}-\mathbf{C}_{S^{c}S}(\mathbf{C}_{SS})^{-1}\|_{\infty} \\ &\leq \|\hat{\mathbf{C}}_{S^{c}S}-\mathbf{C}_{S^{c}S}\|_{\infty}\|(\hat{\mathbf{C}}_{SS})^{-1}\|_{\infty}+\|\mathbf{C}_{S^{c}S}(\mathbf{C}_{SS})^{-1}\|_{\infty}\|\hat{\mathbf{C}}_{SS}-\mathbf{C}_{SS}\|_{\infty}\|(\hat{\mathbf{C}}_{SS})^{-1}\|_{\infty} \\ &\leq \frac{\alpha}{(4-\alpha)\varphi}\frac{4-\alpha}{2(2-\alpha)}\varphi+\left[\left\{(1-\alpha)\frac{\rho'(0+)}{\rho'_{\mu}(b_0/2)}\right\}\wedge (cn^{\nu})\right]\frac{\alpha}{(4-\alpha)\varphi}\frac{4-\alpha}{2(2-\alpha)}\varphi \\ &\leq \frac{\alpha}{2(2-\alpha)}+\left\{\frac{\alpha(1-\alpha)}{2(2-\alpha)}\frac{\rho'(0+)}{\rho'_{\mu}(b_0/2)}\right\}\wedge\left(\frac{c}{2}n^{\nu}\right) \\ &\leq \left\{\frac{\alpha}{2}\frac{\rho'(0+)}{\rho'_{\mu}(b_0/2)}\right\}\wedge (cn^{\nu}), \end{split}
$$

where we have used the inequalities $\rho'(0+)/\rho'_{\mu}(b_0/2) \geq 1$ and $\alpha/\{2(2-\alpha)\} \leq 1/2 \leq cn^{\nu}/2$. This, along with Condition (C6), implies (S.3).

Finally, it follows from the Hoffman–Wielandt inequality (Horn and Johnson 1985) and the assumption (S.1) that

$$
|\Lambda_{\min}(\widehat{\mathbf{C}}_{SS}) - \Lambda_{\min}(\mathbf{C}_{SS})|^2 \leq \|\widehat{\mathbf{C}}_{SS} - \mathbf{C}_{SS}\|_F^2 \leq s^2(2Le_1 + e_2) \leq \left(\frac{\mu_0}{2}\right)^2.
$$

In view of the definition of μ_0 , inequality (S.4) follows. This completes the proof of the lemma.

Lemma S.2. *Under Conditions (C4)–(C6), if the first-stage error bounds satisfy* $e_1 = O(1)$ *and* $e_2 = O(1)$ *, then there exist constants* $c_0, c_1, c_2 > 0$ *such that, if we choose*

$$
\mu \ge C_0 n^{\nu} \sqrt{\frac{\log p + \log q}{n} \vee e_2},
$$

where $C_0 = c_0 L \max(\sigma_{p+1}, M\sigma_{\max}, M)$, then with probability at least $1 - \pi_0 - c_1(pq)^{-c_2}$, it *holds that*

$$
\left\| \frac{1}{n} \widehat{\mathbf{X}}^T \boldsymbol{\eta} - \frac{1}{n} \widehat{\mathbf{X}}^T (\widehat{\mathbf{X}} - \mathbf{X}) \boldsymbol{\beta}_0 \right\|_{\infty} < \frac{\alpha}{6cn^{\nu}} \mu \rho'(0+).
$$
 (S.7)

Proof. As in the proof of Lemma A.2, we write $n^{-1}\hat{\mathbf{X}}^T\boldsymbol{\eta} - n^{-1}\hat{\mathbf{X}}^T(\hat{\mathbf{X}} - \mathbf{X})\boldsymbol{\beta}_0 = T_1 + \cdots + T_6$. Letting $t_0 = \alpha \mu \rho'(0+)/(6cn^{\nu})$, we bound the six terms similarly as follows:

$$
P\left(\|T_1\|_{\infty} \ge \frac{t_0}{6}\right) \le P\left(\left\|\frac{1}{n}\mathbf{Z}^T\boldsymbol{\eta}\right\|_{\infty} \ge \frac{t_0}{6e_1}\right) \le q \exp\left\{-\frac{n}{2\sigma_{p+1}^2} \left(\frac{t_0}{6e_1}\right)^2\right\},\,
$$

$$
P\left(\|T_2\|_{\infty} \geq \frac{t_0}{6}\right) \leq P\left(\left\|\frac{1}{n}\mathbf{Z}^T\boldsymbol{\eta}\right\|_{\infty} \geq \frac{t_0}{6L}\right) \leq q \exp\left\{-\frac{n}{2\sigma_{p+1}^2} \left(\frac{t_0}{6L}\right)^2\right\},
$$

\n
$$
P\left(\|T_3\|_{\infty} \geq \frac{t_0}{6}\right) \leq P\left(\max_{1 \leq i \leq q, 1 \leq j \leq p} \left|\frac{1}{n}\mathbf{z}_i^T \boldsymbol{\varepsilon}_j\right|_{\infty} \geq \frac{t_0}{6Me_1}\right) \leq pq \exp\left\{-\frac{n}{2\sigma_{\max}^2} \left(\frac{t_0}{6Me_1}\right)^2\right\},
$$

\n
$$
P\left(\|T_4\|_{\infty} \geq \frac{t_0}{6}\right) \leq P\left(\max_{1 \leq i \leq q, 1 \leq j \leq p} \left|\frac{1}{n}\mathbf{z}_i^T \boldsymbol{\varepsilon}_j\right|_{\infty} \geq \frac{t_0}{6LM}\right) \leq pq \exp\left\{-\frac{n}{2\sigma_{\max}^2} \left(\frac{t_0}{6LM}\right)^2\right\},
$$

\n
$$
\|T_5\|_{\infty} \leq M \max_{1 \leq i, j \leq p} \frac{1}{n} \| \mathbf{Z}(\widehat{\boldsymbol{\gamma}}_i - \boldsymbol{\gamma}_{0i}) \|_2 \| \mathbf{Z}(\widehat{\boldsymbol{\gamma}}_j - \boldsymbol{\gamma}_{0j}) \|_2 \leq Me_2,
$$

and

$$
||T_6||_{\infty} \le LM \max_{1 \le j \le p} \frac{1}{\sqrt{n}} ||\mathbf{Z}(\widehat{\boldsymbol{\gamma}}_j - \boldsymbol{\gamma}_{0j})||_2 \le LM\sqrt{e_2}.
$$

Combining these bounds and in view of the assumptions $e_1 = O(1)$ and $e_2 = O(1)$, there exist constants $c_0, c_1, c_2 > 0$ such that, if we choose

$$
\mu \ge C_0 n^{\nu} \sqrt{\frac{\log p + \log q}{n} \vee e_2},
$$

where $C_0 = c_0 L \max(\sigma_{p+1}, M\sigma_{\max}, M)$, then with probability at least $1 - \pi_0 - c_1(pq)^{-c_2}$, the desired inequality holds. The completes the proof of the lemma.

Proof of Theorem 4. One can easily show that $\hat{\boldsymbol{\beta}} \in \mathbb{R}^p$ is a strict local minimizer of problem (4) if the following conditions hold:

$$
\frac{1}{n}\widehat{\mathbf{X}}_{\widehat{S}}^T(\mathbf{y} - \widehat{\mathbf{X}}\widehat{\boldsymbol{\beta}}) = \mu \rho'_{\mu}(|\widehat{\boldsymbol{\beta}}_{\widehat{S}}|) \circ \text{sgn}(\widehat{\boldsymbol{\beta}}_{\widehat{S}}),
$$
\n(S.8)

$$
\left\| \frac{1}{n} \widehat{\mathbf{X}}_{\widehat{S}^c}^T (\mathbf{y} - \widehat{\mathbf{X}} \widehat{\boldsymbol{\beta}}) \right\|_{\infty} < \mu \rho'(0+), \tag{S.9}
$$

and

$$
\Lambda_{\min}(\hat{\mathbf{C}}_{\hat{S}\hat{S}}) > \mu \tau(\rho_{\mu}; \hat{\boldsymbol{\beta}}_{\hat{S}}),\tag{S.10}
$$

where \circ denotes the Hadamard (entrywise) product, and $|\cdot|$, $\rho'_{\mu}(\cdot)$, and sgn (\cdot) are applied componentwise. It suffices to find a $\hat{\beta} \in \mathbb{R}^p$ with the desired properties such that conditions $(S.8)$ – $(S.10)$ hold. Let $\beta_{S^c} = 0$. The idea of the proof is to first determine β_S from $(S.8)$, and then show that thus obtained $\hat{\boldsymbol{\beta}}$ also satisfies (S.9) and (S.10).

From now on, we condition on the event of probability at least $1 - \pi_0 - c_1(pq)^{-c_2}$ that the inequalities in Lemmas S.1 and S.2 hold. Using similar arguments to those in the proof of Theorem 3, (S.8) with \widehat{S} replaced by *S* can be written in the form

$$
\widehat{\boldsymbol{\beta}}_S - \boldsymbol{\beta}_{0S} = (\widehat{\mathbf{C}}_{SS})^{-1} \left\{ \frac{1}{n} \widehat{\mathbf{X}}_S^T \boldsymbol{\eta} - \frac{1}{n} \widehat{\mathbf{X}}_S^T (\widehat{\mathbf{X}}_S - \mathbf{X}_S) \boldsymbol{\beta}_{0S} - \mu \rho'_{\mu}(|\widehat{\boldsymbol{\beta}}_S|) \circ \text{sgn}(\widehat{\boldsymbol{\beta}}_S) \right\}.
$$
 (S.11)

Define the function $f: \mathbb{R}^s \to \mathbb{R}^s$ by $f(\boldsymbol{\theta}) = \beta_{0S} + (\hat{\mathbf{C}}_{SS})^{-1} \{ n^{-1} \hat{\mathbf{X}}_S^T \boldsymbol{\eta} - n^{-1} \hat{\mathbf{X}}_S^T (\hat{\mathbf{X}}_S - \mathbf{X}_S) \boldsymbol{\beta}_{0S} \mu \rho'_{\mu}(|\boldsymbol{\theta}|) \circ \text{sgn}(\boldsymbol{\theta})\},\$ and let $\mathcal K$ denote the hypercube $\{\boldsymbol{\theta} \in \mathbb R^s \colon \|\boldsymbol{\theta} - \boldsymbol{\beta}_{0S}\|_{\infty} \le 7\varphi \mu \rho'(0+)/4\}.$ It follows from (S.2), (S.7), and Condition (C4) that, for $\theta \in \mathcal{K}$,

$$
||f(\theta) - \beta_{0S}||_{\infty} \le ||(\widehat{C}_{SS})^{-1}||_{\infty} \left\{ \left\| \frac{1}{n} \widehat{\mathbf{X}}_{S}^{T} \boldsymbol{\eta} - \frac{1}{n} \widehat{\mathbf{X}}_{S}^{T} (\widehat{\mathbf{X}}_{S} - \mathbf{X}_{S}) \beta_{0S} \right\|_{\infty} + \mu \rho'(0+) \right\}
$$

$$
\le \frac{4 - \alpha}{2(2 - \alpha)} \varphi \left\{ \frac{\alpha}{6cn^{\nu}} \mu \rho'(0+) + \mu \rho'(0+) \right\}
$$

$$
\le \frac{3}{2} \varphi \left\{ \frac{1}{6} \mu \rho'(0+) + \mu \rho'(0+) \right\} = \frac{7}{4} \varphi \mu \rho'(0+),
$$

that is, $f(\mathcal{K}) \subset \mathcal{K}$. Also, the last inequality and the assumption (14) imply that for $\theta \in \mathcal{K}$, $\|\boldsymbol{\theta} - \boldsymbol{\beta}_{0S}\|_{\infty} \le b_0/2$, and hence sgn($\boldsymbol{\theta}$) = sgn($\boldsymbol{\beta}_{0S}$). Thus, in view of Condition (C4), *f* is a continuous function on the convex, compact hypercube *K*. An application of Brouwer's fixed point theorem yields that equation (S.11) has a solution β_S in *K*. Moreover, sgn(β_S) = $sgn(\mathcal{B}_{0S})$, so that $S = S$. Therefore, we have found a β that satisfies the desired properties and (S.8).

To verify that $\hat{\boldsymbol{\beta}}$ satisfies (S.9), by substituting (S.11), we write

$$
\frac{1}{n}\widehat{\mathbf{X}}_{S^c}^T(\mathbf{y}-\widehat{\mathbf{X}}\widehat{\boldsymbol{\beta}})=\frac{1}{n}\widehat{\mathbf{X}}_{S^c}^T\boldsymbol{\eta}-\frac{1}{n}\widehat{\mathbf{X}}_{S^c}^T(\widehat{\mathbf{X}}_S-\mathbf{X}_S)\boldsymbol{\beta}_{0S} \n-\widehat{\mathbf{C}}_{S^cS}(\widehat{\mathbf{C}}_{SS})^{-1}\left\{\frac{1}{n}\widehat{\mathbf{X}}_S^T\boldsymbol{\eta}-\frac{1}{n}\widehat{\mathbf{X}}_S^T(\widehat{\mathbf{X}}_S-\mathbf{X}_S)\boldsymbol{\beta}_{0S}-\mu\rho'_{\mu}(|\widehat{\boldsymbol{\beta}}_S|)\circ\text{sgn}(\widehat{\boldsymbol{\beta}}_S)\right\}.
$$

Also, we have $\|\boldsymbol{\beta}_{S}\|_{\infty} = \|\boldsymbol{\beta}_{0S} + (\boldsymbol{\beta}_{S} - \boldsymbol{\beta}_{0S})\|_{\infty} \ge \|\boldsymbol{\beta}_{0S}\|_{\infty} - \|\boldsymbol{\beta}_{S} - \boldsymbol{\beta}_{0S}\|_{\infty} \ge b_0 - b_0/2 = b_0/2.$ This, together with (S.3), (S.7), and Condition (C4), leads to

$$
\left\| \frac{1}{n} \widehat{\mathbf{X}}_{S^c}^T (\mathbf{y} - \widehat{\mathbf{X}} \boldsymbol{\beta}) \right\|_{\infty} \le \left\| \frac{1}{n} \widehat{\mathbf{X}}_{S^c}^T \boldsymbol{\eta} - \frac{1}{n} \widehat{\mathbf{X}}_{S^c}^T (\widehat{\mathbf{X}}_S - \mathbf{X}_S) \boldsymbol{\beta}_{0S} \right\|_{\infty} + \|\widehat{\mathbf{C}}_{S^c S} (\widehat{\mathbf{C}}_{SS})^{-1} \|_{\infty} \times \left\{ \left\| \frac{1}{n} \widehat{\mathbf{X}}_S^T \boldsymbol{\eta} - \frac{1}{n} \widehat{\mathbf{X}}_S^T (\widehat{\mathbf{X}}_S - \mathbf{X}_S) \boldsymbol{\beta}_{0S} \right\|_{\infty} + \mu \rho_{\mu}^{\prime}(b_0/2) \right\} \n $\frac{\alpha}{6cn^{\nu}} \mu \rho^{\prime}(0+) + 2cn^{\nu} \cdot \frac{\alpha}{6cn^{\nu}} \mu \rho^{\prime}(0+) + \left(1 - \frac{\alpha}{2} \right) \frac{\rho^{\prime}(0+)}{\rho_{\mu}^{\prime}(b_0/2)} \cdot \mu \rho_{\mu}^{\prime}(b_0/2) \n\le \frac{\alpha}{6} \mu \rho^{\prime}(0+) + \frac{\alpha}{3} \mu \rho^{\prime}(0+) + \left(1 - \frac{\alpha}{2} \right) \mu \rho^{\prime}(0+) = \mu \rho^{\prime}(0+).$
$$

Finally, it follows from (S.4) and the definition of τ_0 that $\Lambda_{\min}(\mathbf{C}_{SS}) > \mu \tau_0 \geq \mu \tau(\rho_\mu; \beta_S)$, which verifies $(S.10)$ and completes the proof.