

## Supplementary Material for Adaptive Response-Dependent Two-Phase Designs, by Michael A. McIsaac and Richard J. Cook

### *Asymptotic Variance of the Mean Score Estimator*

Consider the weighted unbiased score equation

$$\bar{U}(\beta) = \sum_{i=1}^N \bar{U}_i(\beta) = \sum_{i=1}^N R_i \pi_i(Y_i, V_i)^{-1} U_\beta(Y_i | X_i, V_i) = 0$$

where we use  $\pi_i = \pi(y_i, v_i) = P(R_i = 1 | y_i, v_i) = n_{y_i v_i} / N_{y_i v_i}$  to denote the sampling probability for an individual  $i$ . As noted by Lawless et al. [1], under mild regularity conditions [2]

$$\sqrt{N}(\hat{\beta} - \beta) \xrightarrow{p} N(0, \mathcal{A}(\Psi)^{-1} \mathcal{C}(\Omega) \mathcal{A}(\Psi)^{-1}),$$

where  $\Psi = (\beta', \alpha', \gamma')'$ ,  $\Omega = (\Psi', \delta')'$ ,

$$\mathcal{A}(\Psi) = \lim E[-N^{-1} \partial \bar{U}_i(\beta) / \partial \beta'] = E_{YXV}[-\partial U_\beta(Y_i | X_i, V_i) / \partial \beta'],$$

and

$$\begin{aligned} \mathcal{C}(\Omega) &= \lim \text{var}(N^{-1/2} \sum_{i=1}^N R_i \pi_i^{-1} U_\beta(Y_i | X_i, V_i)) \\ &= \lim \left\{ \text{var}_{YXV}(E_{R|YXV}[N^{-1/2} \sum_{i=1}^N R_i \pi_i^{-1} U_\beta(Y_i | X_i, V_i)]) \right. \\ &\quad \left. + E_{YXV}(\text{var}_{R|YXV}[N^{-1/2} \sum_{i=1}^N R_i \pi_i^{-1} U_\beta(Y_i | X_i, V_i)]) \right\} \\ &= \lim N^{-1} \left\{ \text{var}_{YXV} \left( \sum_{i=1}^N U_\beta(Y_i | X_i, V_i) \right) + E_{YXV}(\text{var}_{R|YXV}[\sum_{i=1}^N R_i \pi_i^{-1} U_\beta(Y_i | X_i, V_i)]) \right\}. \end{aligned}$$

We denote the second-order inclusion probability as  $\pi_{ij} = P(R_i = 1, R_j = 1 | y_i, v_i, y_j, v_j)$  so  $\pi_{ij} = \pi_i \cdot \pi_j = n_{y_i v_i} / N_{y_i v_i} \cdot n_{y_j v_j} / N_{y_j v_j}$  if individuals  $i$  and  $j$  are from different strata (i.e. if  $(y_i, v_i) \neq (y_j, v_j)$ ), while  $\pi_{ij} = n_{y_i v_i} / N_{y_i v_i} \cdot (n_{y_i v_i} - 1) / (N_{y_i v_i} - 1)$  if they are from the same stratum (i.e.  $(y_i, v_i) = (y_j, v_j)$ ).

So,

$$\begin{aligned} &E_{YXV} \left[ \text{var}_{R|YXV} \left[ \sum_{i=1}^N R_i \pi_i^{-1} U_\beta(Y_i | X_i, V_i) \right] \right] \\ &= E_{YXV} \left[ \sum_{i=1}^N \text{var}_{R|YXV}(R_i) \pi_i^{-2} U_\beta(Y_i | X_i, V_i) U'_\beta(Y_i | X_i, V_i) + \right. \\ &\quad \left. \sum_{i=1}^N \sum_{j=1; i \neq j}^N \text{cov}(R_i, R_j | Y_i, X_i, V_i, Y_j, X_j, V_j) \pi_i^{-1} \pi_j^{-1} U_\beta(Y_i | X_i, V_i) U'_\beta(Y_j | X_j, V_j) \right] \\ &= E_{YXV} \left[ \sum_{i=1}^N (\pi_i - \pi_i^2) \pi_i^{-2} U_\beta(Y_i | X_i, V_i) U'_\beta(Y_i | X_i, V_i) + \sum_{i=1}^N \sum_{j=1; i \neq j}^N \frac{\pi_{ij} - \pi_i \pi_j}{\pi_i \pi_j} U_\beta(Y_i | X_i, V_i) U'_\beta(Y_j | X_j, V_j) \right] \\ &= E_{YXV} \left[ \sum_{i=1}^N (\pi_i^{-1} - 1) U_\beta(Y_i | X_i, V_i) U'_\beta(Y_i | X_i, V_i) + \right. \\ &\quad \left. \sum_{i=1}^N E_{YV} \left[ (N_{YV} - 1) \left( \frac{n_{YV} - 1}{N_{YV} - 1} \frac{N_{YV}}{n_{YV}} - 1 \right) E_{X|YV}[U_\beta(Y_i | X_i, V_i)] E_{X|YV}[U'_\beta(Y_i | X_i, V_i)] \right] \right] \end{aligned}$$

$$\begin{aligned}
 &= E_{YXV} \left[ \sum_{i=1}^N (\pi_i^{-1} - 1) U_\beta(Y_i|X_i, V_i) U'_\beta(Y_i|X_i, V_i) \right] - \\
 &\quad \sum_{i=1}^N E_{YV} \left[ (\pi_i^{-1} - 1) E_{X|YV}[U_\beta(Y_i|X_i, V_i)] E_{X|YV}[U'_\beta(Y_i|X_i, V_i)] \right] \\
 &= \sum_{i=1}^N E_{YV} \left[ (\pi_i^{-1} - 1) (E_{X|YV}[U_\beta(Y_i|X_i, V_i) U'_\beta(Y_i|X_i, V_i)] - E_{X|YV}[U_\beta(Y_i|X_i, V_i)] E_{X|YV}[U'_\beta(Y_i|X_i, V_i)]) \right].
 \end{aligned}$$

Therefore,

$$\mathcal{C}(\Omega) = E[U_\beta(Y_i|X_i, V_i) U'_\beta(Y_i|X_i, V_i)] + \sum_{YV} P(Y, V) (\pi(Y, V)^{-1} - 1) \text{var}_{X|YV}[U_\beta(Y_i|X_i, V_i)],$$

and since  $E[U_\beta(Y_i|X_i, V_i) U'_\beta(Y_i|X_i, V_i)] = E_{YXV}[-\partial U_\beta(Y_i|X_i, V_i)/\partial \beta']$  [3], the asymptotic variance of the mean score estimator is

$$\mathcal{A}(\Psi)^{-1} + \mathcal{A}(\Psi)^{-1} \mathcal{B}(\Omega) \mathcal{A}(\Psi)^{-1}, \quad (1)$$

where

$$\mathcal{B}(\Omega) = \sum_{YV} P(Y, V) \left[ \frac{N_{YV}}{n_{YV}} - 1 \right] \cdot \text{var}_{X|Y,V}[U_\beta(Y_i|X_i, V_i)].$$

## References

1. Lawless JF, Kalbfleisch JD, Wild CJ. Semiparametric methods for response-selective and missing data problems in regression. *Journal of the Royal Statistical Society Series B (Statistical Methodology)* 1999; **61**(2):413–438.
2. Wild CJ. Fitting prospective regression models to case-control data. *Biometrika* 1991; **78**:705–717.
3. Pierce DA. The asymptotic effect of substituting estimators for parameters in certain types of statistics. *The Annals of Statistics* 1982; **10**:475–478.