SUPPLEMENTARY MATERIAL

Criteria	Description	Notes
Population	• Age: adults (≥ 18 years)	Schizophrenia can occur both in men and
	• Gender: any	women with equal prevalence. Race or
	• Race: any	ethnicity does not influence the prevalence of
	Stable disease	schizophrenia. Therefore, adults of any race
		or gender diagnosed with schizophrenia were
		included in this review. Stable disease may be
		measured differently but a PANSS score of
		around 60 and outpatient status generally
		indicate stable disease.
Disease	Diagnosis of schizophrenia that was not	Schizoaffective disorder was not considered
	treatment resistant	in this review.
Intervention	Aripiprazole once-monthly	
	Olanzapine pamoate	
	Paliperidone palmitate	
	Risperidone LAI	
	Haloperidol depot	
Comparator	Any other included intervention	These comparators were selected to enable
	Placebo	both direct and indirect comparisons between
		the interventions of interest.
Study design	Double-blind / triple-blind randomized	Randomized controlled trials are the gold
	control trials	standard of clinical evidence, minimizing the
		risk of confounding and allowing the
		comparison of the relative efficacy of
		interventions. Therefore only these studies
T (1		were included.
Treatment phase	Maintenance treatment	Inclusion was restricted to maintenance
		treatment in order to angle with the expected
		marketing autionzation for anpiprazole once-
Longuaga	English only	The restriction would not limit results
restrictions		substantially as most of the research is
resulctions		published in English language journals
Sample size	>10 participants	The sample size of the included studies was in
Sumple Size		line with the NICE clinical review protocol
		(27)
Study duration	Long-term follow-up (at least 6 months)	The study duration of the included studies
_ and an		was in line with the NICE clinical review
		protocol (27).

1. Study inclusion criteria for the systematic literature review extension to May 2013

2. Study exclusion criteria for the systematic literature review extension to May 2013

Criteria	Description	Notes
Subgroup analysis	 No subgroup analysis for disease of interest 	Studies with no subgroup data for the disease, adult population, or maintenance treatment in
	 No subgroup analysis for adult population 	stable patients were not included, since these studies would introduce heterogeneity into the review.
Outcome of	Studies not reporting an outcome interest	Studies which do not report outcomes of
interest		interest would not feature in any analyses and
		were therefore excluded from extraction

3. PRISMA Flow diagram



Note: because the literature search strategy of this study adopted the one used for the NICE schizophrenia clinical guidance, oral formulations and depot formulations were both considered throughout the search. Studies of oral formulation antipsychotics were excluded after full text review.

4. Quality assessment of the included RCTs

Trial	Fleischhacker 2012a	Fleischhacker 2002b	Hough 2010	Kane 2002	Kane 2010	Kane 2012
Was randomisation carried out appropriately?	Yes ("computer based randomization scheme")	Unclear	Yes ("via a sponsor- prepared computer- generated randomization scheme; assigned by an interactive voice response system")	Unclear	Unclear	Unclear
Was the concealment of treatment allocation adequate?	Unclear	Unclear	Unclear	Unclear	Unclear	Unclear
Were the groups similar at the outset of the study in terms of prognostic factors?	Yes ("The demographic and baseline characteristics were similar between treatment groups")	Yes ("Baseline demographic characteristics were similar between randomized treatment groups")	Yes (Age and sex at baseline were similar. Furthermore "The mean PANSS total scores and CGI-S scores at transition phase baseline suggested that the population was symptomatically stable and at double-blind baseline indicated adequate symptom control during the maintenance phase")	Yes (Patients were similar in terms of age, sex, duration of illness, BPRS total score)	Yes ("No statistically significant differences were observed for baseline physical characteristics and illness history")	Yes ("Baseline demographic and disease characteristics of randomized subjects were similar between treatment groups")

5. Supplementary results

	Mean	Median	2.50%	97.50%	
Aripiprazole once-monthly versus Olanzapine pamoate					
Relapse	76.18	0.70	0.214	2.43	
Disc. AE	2.09	0.75	0.099	6.176	
Disc. Other	2.77	0.61	0.136	2.953	
Aripiprazole once-monthly versus Paliperidone palmitate					
Relapse	0.84	0.63	0.186	2.142	
Disc. AE	1.00	0.34	0.020	4.951	
Disc. Other	14.93	0.55	0.111	2.852	
Aripiprazole o	nce-monthly v	ersus Risperi	idone LAI		
Relapse	1713.00	1.15	0.234	5.446	
Disc. AE	6.77	0.42	0.018	9.042	
Disc. Other	2.18E+09	0.56	0.078	4.51	
Aripiprazole o	nce-monthly v	ersus Halope	ridol depot		
Relapse	1.061	0.791	0.207	3.074	
Disc. AE	4.314	0.864	0.018	22.19	
Disc. Other	16.35	0.302	0.035	2.005	

Table 5. Efficacy and tolerability outcomes

Table 6. Safety outcomes

	Mean	Median	2.50%	97.50%		
Aripiprazole once-monthly versus Olanzapine pamoate						
Weight gain	1.029	0.938	0.383	2.215		
EPS	1.65	1.53	0.68	3.32		
Aripiprazole once-monthly versus Paliperidone palmitate						
Weight gain	1.203	1.004	0.297	3.251		
EPS	0.97	0.90	0.37	2.07		
Aripiprazole once-monthly versus Risperidone LAI						
Weight gain	1.149	0.941	0.259	3.227		
EPS	0.82	0.74	0.28	1.86		

Forest plots of efficacy and tolerability outcomes







6. WinBugs code for the efficacy and tolerability analysis

Model

```
model{
# Ni = number of arms
# Nk = number of treatments
# Nj = number of trials
# Nib = number of placebo arms
# Code for treatment effects relative to placebo (treatment 1)
for(i in 1:Ni) {
       # multinomial likelihood
       r[i,1:4] ~ dmulti(p[i,1:4],n[i])
       # sum of the hazard rates for the 3 discontinuation outcomes
       slam[i] <- sum(lam[,i])</pre>
       for (m in 1:3) {
              # probability of reaching each discontinuation outcome
              p[i,m] <- lam[m,i] * (1-exp(-slam[i]*w[i]/26)) / slam[i]</pre>
              # log hazard rates for each arm, each outcome
              log(lam[m,i]) <- theta[m,i]</pre>
              theta[m,i] <- mu[m,s[i]] + delta[m,i]*(1-equals(t[i],b[i]))</pre>
              # random effects model for log hazard ratios with correction for three-
arm trials
              delta[m,i] ~ dnorm(md[m,i],taud[m,i])
              taud[m,i] <- tau[m] * (1 + equals(arm[i],3) /3)</pre>
              md[m,i] <- d[m,t[i]] - d[m,b[i]] + equals(arm[i],3) * sw[m,i]</pre>
       }
       # probability of continuing treatment
      p[i,4] <- 1- sum(p[i,1:3])
}
# Correction factor for 3-arm trials
                                             6
```

```
for (m in 1:3) {
       sw[m, 1] < - 0
       for (i in 2:Ni) { sw[m,i] <- (delta[m,i-1] - d[m,t[i-1]] + d[m,b[i-1]] ) /2 }</pre>
}
# Code for absolute effects of placebo (treatment 1)
for (i in 1:Nib) {
       rb[i,1:4] ~ dmulti(pb[i,1:4],nb[i])
       for (m in 1:3) {
              pb[i,m] <- lamb[m,i] * (1-exp(-slamb[i]*wb[i]/26)) / slamb[i]</pre>
              log(lamb[m,i]) <- mub[m,sb[i]]</pre>
       }
       slamb[i] <- sum(lamb[,i])</pre>
       pb[i,4] <- 1- sum(pb[i,1:3])</pre>
}
# Priors
for (m in 1:3) {
       d[m,1] <- 0
       for (k in 2:Nk) {
              # priors for treatment effects
              d[m,k] ~ dnorm(0,.0001)
              log(hazr[m,k]) <- d[m,k] # hazard ratios</pre>
       }
       # priors for baselines
       for (j in 1:Nj) { mu[m,j] ~ dnorm(0,.0001) }
for (m in 1:3) {
       # variance of the log hazard ratios
       tau[m] <- pow(sdb[m],-2) * pow(2*(1-rho[m]),-.5)</pre>
for (m in 1:3) {
       # priors for treatment effects and baselines for placebo
       for (j in 1:Nib) { mub[m,j] ~ dnorm(mb[m],prb[m]) }
       mb[m] \sim dnorm(0,.0001)
for (m in 1:3) {
       # priors for between-trial variation
       prb[m] ~ dgamma(.1,.1)
       sdb[m] <- pow(prb[m],-.5)</pre>
       rho[m] \sim dunif(0,1)
}
# Code for estimated absolute probabilities at 26 weeks
for (m in 1:3) {
       for (k in 1:Nk) {
              theta26[m, k] < -mb[m] + d[m, k]
              log(lam26[m,k]) <- theta26[m,k]
              p26[m,k] <- lam26[m,k] * (1-exp(-slam26[k])) / slam26[k]</pre>
       }
for (k in 1:Nk) {
       # probability of discontinuing treatment
       slam26[k] <- sum(lam26[1:3,k])</pre>
       # probability of continuing treatment
       p26[4,k] <- 1-sum(p26[1:3,k])
1
# Code for estimating that each treatment is the best option
for (k in 1:Nk) {
       for (m in 1:4) {
              # rank of each treatment from smallest to largest
              rank26[m,k] <- rank(p26[m,],k)</pre>
```

```
}
       # record whether best (i.e. smallest probability, rank 1) for each
discontinuation outcome
       for (m in 1:3) { best[m,k] <- equals(rank26[m,k],1) }</pre>
       # record whether best (i.e. largest probability, rank Nk) for continuation
outcome
      best[4,k] <- equals(rank26[4,k],Nk)</pre>
}
# probabilities of being ranked jth (j=1,...,Nk) for each outcome
for (k in 1:Nk) {
       # record whether jth best w.r.t. relapse
       for (j in 1:Nk) { rankREL[j,k] <- equals(rank26[1,k],j) }</pre>
       # record whether jth best w.r.t. discontinuation due to adverse events
       for (j in 1:Nk) { rankAE[j,k] <- equals(rank26[2,k],j) }</pre>
       # record whether jth best w.r.t. discontinuation due to other reasons
       for (j in 1:Nk) { rankOTHER[j,k] <- equals(rank26[3,k],j) }</pre>
       # record whether jth worst w.r.t. continuation (NOTE: REVERSED ORDER)
       for (j in 1:Nk) { rankCONT[j,k] <- equals(rank26[4,k],j) }</pre>
}
}
```

Data

list(

```
Nk = 8, # treatments
Nj = 6, \# trials
Ni = 14, # arms
r = structure(.Data =
c(22,8,39,196,21,7,60,178,29,7,34,61,36,3,28,139,97,2,27,78,18,2,19,45,16,1,3,9,42,6,2
0,76,68,21,91,419,23,8,33,258,27,9,31,202,53,5,15,61,56,25,105,184,95,29,100,155),
.Dim = c(14, 4)),
n=c(265,266,131,206,204,84,29,144,599,322,269,134,370,379),
b = c(1,1,1,1,1,1,1,1,1,1,1,1,6,6),
t = c(2,7,1,5,1,3,1,1,4,8,2,1,6,5),
s = c(1, 1, 1, 2, 2, 3, 3, 4, 4, 4, 5, 5, 6, 6),
w = c(38, 38, 38, 31, 42, 52, 52, 24, 24, 24, 52, 52, 53, 53),
arm = c(1,2,3,1,2,1,2,1,2,3,1,2,1,2),
# For baseline: placebo (treatment 1)
Nib = 5, \# placebo arms
rb = structure(.Data = c(29,7,34,61,97,2,27,78,16,1,3,9,42,6,20,76,53,5,15,61), .Dim =
c(5,4)),
nb = c(131, 204, 29, 144, 134),
wb = c(38, 42, 52, 24, 52),
sb = c(1, 2, 3, 4, 5)
)
```

Initial values #1

Initial values #2

```
# initial values 1
list(
0,0,0,0,0, 0,0,0,0,0, 0,0,0,0,
     0,0,0,0,0, 0,0,0,0,0, 0,0,0,0),.Dim=c(3,14)),
NA,-1,-1,-1,-1, -1,-1,-1,
    NA,-1,-1,-1,-1, -1,-1,-1), .Dim=c(3,8)),
-1,-1,-1,-1,-1, -1,
    -1,-1,-1,-1,-1, -1),.Dim=c(3,6)),
mub=structure(.Data=c(0,0,0,0,0,
    0,0,0,0,0,
    0, 0, 0, 0, 0), .Dim=c(3, 5)),
mb=c(2,-2,-2),
prb=c(3,3,3),
rho=c(0.5, 0.5, 0.5)
)
```

7. WinBugs code for the weight gain analysis

Model

```
model{
sw[1] <- 0
for (i in 1:Ni) { # LOOP OVER ARMS
       r[i] ~ dbin(p[i],n[i])
       logit(p[i]) <- mu[s[i]]+delta[i]*(1-equals(t[i],b[i]))</pre>
       #Random effects model for log-odds ratios
       delta[i] ~ dnorm(md[i],taud[i])
       taud[i] <- tau * (1 + equals(arm[i],3) /3)</pre>
       md[i] <- d[t[i]] - d[b[i]] + equals(arm[i],3) * sw[i]</pre>
       #Deviance residuals for data i
       rhat[i] <- p[i] * n[i]</pre>
       dev[i] <- 2 * (r[i] * (log(r[i])-log(rhat[i])) + (n[i]-r[i]) * (log(n[i]-r[i])</pre>
- log(n[i]-rhat[i])))
}
sumdev <- sum(dev[])</pre>
#Adjustment for 3 arm trials
for (i in 2:Ni) { sw[i] <- (delta[i-1] - d[t[i-1]] + d[b[i-1]] ) /2 }
#Priors
for (j in 1:Nj) { mu[j] ~ dnorm(0,.0001) }
tau <- 1/(sd*sd)
```

```
sd ~ dunif(0,2)
#Give priors for log-odds ratios
d[1] <-0
for (k \text{ in } 2:Nk) \{ d[k] \sim dnorm(0,.0001) \}
#All pairwise odds ratios
#or[c,k] is the odds ratio of treatment k relative to treatment c
for (c in 1:(Nk-1)) {
      for (k in (c+1):Nk) {
             or[c,k] <- exp(d[k] - d[c])
       }
}
#Odds and probabilities of weight gain for each treatment
prob[1] <- 33/613 #Baseline estimate for weight gain associated with PLB depot
(treatment 1)
odds[1] <- prob[1] / (1-prob[1])</pre>
for (k in 2:Nk) {
       odds[k] <- or[1,k] * prob[1] / (1-prob[1])</pre>
       prob[k] <- odds[k] / (1+odds[k])</pre>
}
# Code for estimating that each treatment is the best option
for (k in 1:Nk) {
rk[k] <- rank(d[],k) # assumes events are "bad"</pre>
best[k] \leq -equals(rk[k],1) #calculate probability that treat k is best
}
}
Data – Weight gain
list(
Nk = 7, \# treatments
Nj= 5, # trials
Ni= 12, # arms
```

r=c(42,43,8,12,6,12,100,68,17,7,52,50),

n=c(264,266,131,206,204,144,599,322,269,134,338,346),

```
\mathtt{s=c}\;(1,1,1,2,2,3,3,3,4,4,5,5)\;,
```

```
t=c(2,6,1,4,1,1,3,7,2,1,5,4),
```

```
b=c(1,1,1,1,1,1,1,1,1,1,5,5),
```

```
arm=c(1,2,3,1,2,1,2,3,1,2,1,2)
```

```
)
```

Data – EPS

list(

```
Nk = 7, # treatments
Nj= 5, # trials
```

```
Ni= 12, # arms
r=c(52,46,18,21,12,12,52,28,45,14,76,67),
n=c(265,266,131,206,204,144,599,322,269,134,370,379),
s=c(1,1,1,2,2,3,3,3,4,4,5,5),
t=c(2,6,1,4,1,1,3,7,2,1,5,4),
b=c(1,1,1,1,1,1,1,1,5,5),
arm=c(1,2,3,1,2,1,2,3,1,2,1,2)
)
```

Initial values

```
#initial values
list(
d=c(NA,0,0,0,0,0,0,0),
sd=1,
mu=c(0,0,0,0,0),
delta=c(0,0,0,0,0,0,0,0,0,0,0,0,0,0))
)
```