

Supplementary Methods

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1. Definitions

Mutant, mutation: in this manuscript, these terms are equivalent to "variant", with no pathogenic connotation.

Ties: statistical term used to designate "identical values".

WT reference: backbone cDNA sequence carrying each assessed variant.

Cut-off: value that allows separation of the mutations into two categories: pathogenic or neutral. Depending on the functional assay, the pathogenic category lies above or below the cut-off, with the neutral category in the opposite position.

Sensitivity: proportion of pathogenic mutations correctly classified. In functional assessment, this is equivalent to the proportion of pathogenic mutations in the pathogenic area. This area lies above or below the cut-off, depending on the assay used.

Specificity: proportion of neutral mutations correctly classified. In functional assessment, this is equivalent to the proportion of neutral mutations in the neutral area. This area lies above or below the cut-off, depending on the assay used.

Accuracy: proportion of mutations correctly classified.

Best cut-off: cut-off value associated with the Youden's index (see **S2 Fig**).

Best sensitivity and specificity: values associated with the best cut-off, and thus with the Youden's index. The term "best" is frequently omitted as the sensitivity and specificity of an assay is always the best among the possible sensitivities and specificities.

Experimental data: data from experiments, as opposed to data from bootstrap analysis, which corresponds to a computer-assisted sampling of the experimental data.

Experimental best cut-off: best cut-off obtained from experimental data, as opposed to the best cut-off obtained from bootstrap analysis.

Experimental sensitivity, specificity and accuracy: best sensitivity, best specificity and best accuracy obtained from experimental data and associated with the experimental best cut-off, as opposed to the sensitivity, specificity and accuracy of the probability systems of classification, obtained after bootstrap analysis.

Initial sensitivity and specificity: equivalent to experimental best sensitivity and specificity. The term "initial" can be used to designate the sensitivity and specificity of the data (experimental or theoretical), since they are computed using the initial position of the neutral and pathogenic mutations, before bootstrap analysis.

Standard method: medians of the mutants are used to find the best cut-off.

Standard with reference method: as in the standard method, except that the best cut-off identified is divided by the median of the WT BRCA1 reference. Thus, the best cut-off is a value relative to the BRCA1 median value, with no unit. In fact, the standard with reference method is not different from the standard method for the experimental data. However, these methods differ during bootstrap analysis, since the raw and relative best cut-offs fluctuate differently.

Raw best cut-off: the term "raw" indicates that the best cut-off value is not divided by the WT BRCA1 median. Therefore, the raw best cut-off has the unit of the experimental data (e.g., cells per colony in the Colony Size assay).

Relative best cut-off: The term "relative" indicates that the best cut-off value is divided by the WT BRCA1 median.

MWW method: as in the standard method, except that each mutant is not represented by a median value but by a p value, as explained in **S4 Fig**.

CDF: cumulative distribution function, see **S6 Fig**.

Probability system of classification: system that uses the fluctuation of the best cut-off to derive probabilities of pathogenicity for each assessed variant (**S6 Fig**). Such probabilities allow the use of a five-class nomenclature to classify variants, as shown in **S1 Table**. Of note, the system is based on an average CDF.

Sensitivity of the probability system of classification: proportion of pathogenic mutations within the class 4 or 5 (see **Fig 2B**). This sensitivity has to be distinguished from the experimental sensitivity.

Specificity of the probability system of classification: proportion of neutral mutations within the class 1 or 2 (see **Fig 2B**). This specificity has to be distinguished from the experimental specificity.

Probability unit: lowest potential incrementation within the average CDF. In an exact distribution, this unit is equal to $1 / n_{\text{best exact}}$ if at least three consecutive best cut-off values from the exact best cut-off distribution are not repeated. Otherwise, the value of the unit is higher. In an approximate distribution (bootstrap), this unit is equal to $1 / n_{\text{bootstrap}}$ if at least three consecutive best cut-off values from the best cut-off distribution are not repeated. Otherwise, the value of the unit is higher.

Quantile system of classification: this system is an alternative to the probability system of classification. It is a very simple approach to generate probabilistic classifications.

However, as demonstrated below, this system is not adapted to the classification of variants.

System of classification: designates either the probability or the quantile system.

Accuracy of a system of classification: capability of the probability or quantile system to attribute a class 1 or 2 to the neutral mutations, and a class 4 or 5 to the pathogenic mutations. Of note, a completely accurate system never attributes the class 3 and never misclassifies variants. A system fully inaccurate attributes the class 3 only, or totally misclassifies the neutral and pathogenic variants.

Classification model: computational model that combines a method of best cut-off computation (standard, standard with reference or MWW method) and a system of classification (probability or quantile system).

2. Symbols frequently used

| | |
|-------------------------|--|
| n_{neutral} | number of neutral mutations |
| $n_{\text{pathogenic}}$ | number of pathogenic mutations |
| n_{mutant} | number of values within every mutant (implicating an equal number of values between each mutant) |
| n_{BRCA1} | number of values in the WT BRCA1 reference |
| $n_{\text{bootstrap}}$ | number of bootstraps performed to estimate the best cut-off fluctuation |
| n_{best} | number of best cut-off values obtained after bootstrap ($n_{\text{best}} = n_{\text{bootstrap}}$) |
| n_{diff} | number of different best cut-off values obtained after bootstrap |
| $n_{\text{best exact}}$ | number of best cut-off values in the exact best cut-off distribution (equal to the number of sampling possibilities) |
| $n_{\text{exact diff}}$ | number of different best cut-off values in the exact best cut-off distribution |

3. Colony size and Liquid Medium assays in glucose media

All of the clones from the Colony Size assay were assessed in glucose media to control the absence of any intrinsic growth defect, which would disturb the classification of the pathogenic missense mutations (**S3B Fig**). The rare clones, 10% below the median of the BRCA1 or Vector cells, after glucose induction, were removed from the study before analysis in galactose media.

As for the Colony Size assay, all of the clones from the Liquid Medium assay were assessed in glucose media to control the absence of any intrinsic growth defect (**S8B Fig**). Of note, the 126 clones used in the Colony Size and Liquid Medium assays were the same, except for 2 clones, due to the screening in glucose media.

4. Western blot

Western blots were performed as previously described [11]. One among three independent clones from each category was selected for analysis. Membranes were probed with an anti-BRCA1 monoclonal antibody (MS110, Calbiochem, Billerica, MA, USA, 1:200 dilution), then with a secondary peroxidase-conjugated anti-mouse antibody (Jackson Immunoresearch, West Grove, PA, USA, 1:10,000 dilution). To control for loading variation, stripped membranes were probed with an anti-Tubulin antibody (YL1/2, AbD serotec, Oxford, UK, 1:2,000 dilution) followed by a secondary peroxidase-conjugated anti-rat antibody (Jackson Immunoresearch, 1:5,000 dilution), or with an anti- β actin antibody (8224, abcam, 1:5,000 dilution) followed by a secondary peroxidase-conjugated anti-mouse antibody (Jackson Immunoresearch, 1:10,000 dilution). The ImageJ software was used to quantify signal intensities in every lane (full lanes quantified).

Western-blot analyses were performed to ascertain that growth recoveries were not related to any defect in BRCA1 protein expression (**S20 Fig**). In contrast, mutations showing the largest

number of cells per colony were associated with an increase in protein levels. This confirms what we previously reported for the Y1853X mutation and truncated forms of BRCA1 [11], and suggests that protein levels, clearly higher than the WT BRCA1 control level, predict a classification as pathogenic by the Colony Size, Liquid Medium, Spot Formation or Yeast Localization assay.

5. The ProClass toolbox developed for the probabilistic classification of experimental data

The Probabilistic Classification (ProClass) toolbox compiles the R codes [29] and raw data used in this article. The toolbox is designed with the purpose to facilitate the analysis of most kind of functional assay data, and to provide a probabilistic classification of variants. This toolbox is also adapted to other experimental data, provided that they include at least one positive and one negative control. ProClass is available on line at: http://xfer.curie.fr/get/tvsjyy4dUno/ProClass_toolbox.zip. The following sections 6 to 14 describe the computational procedures used in ProClass. Sections 22 and 23 explain how to integrate ProClass during functional assessment.

6. Experimental data assembling

As mentioned in the description of the functional assays, three independent transformants, also referred to as "clones", were selected for each transformation. This means that each strain, described in **S11 Table**, is represented by three clones. For the Colony Size assay, the 40 missense mutations were separated into four batches of 10 mutations. In the first batch, the three clones from each mutation were analyzed in three independent experiments. The three clones from the WT BRCA1 and the Vector strain were systematically used as a control. This gave 9 colony size values for each WT BRCA1, mutated BRCA1 or Vector control strain in the first batch. At this stage, no normalization was performed, meaning that the dispersion of the 9 values includes both the inter-experiment and the inter-clonal variation. The same was performed for the three other batches of 10 missense mutations, with the WT BRCA1 and the Vector strains as a control. Next, to assemble the results from the four batches, the following formula was applied: $x_{ijk} \times \tilde{x}_{BRCA1\ batch\ 1} / \tilde{x}_{BRCA1\ batch\ k}$, with x_{ijk} being the colony size value of the clone i ($i = 1$ to 3) in the experiment j ($j = 1$ to 3) for the batch k ($k = 2$ to 4), and with $\tilde{x}_{BRCA1\ batch\ 1}$ and $\tilde{x}_{BRCA1\ batch\ k}$ being the median value of the 9 colony size values from the WT BRCA1 strain in the batch 1 and k respectively. This means that, in batch 2, 3 and 4, the values were adjusted such that the median value of the 9 WT BRCA1 colony size values is equal to the WT BRCA1 median value of batch 1. This also means that the inter-batches variation was not considered. For the WT BRCA1 and Vector strains, the 9 adjusted values of the 4 batches were kept, meaning that the final distribution of these two strains was composed of 36 values.

The same method was applied to the Liquid Medium assay. Of note, the clones used in the Colony Size assay were mostly the same used in the Liquid Medium assay, which allows the comparison of the relative efficiency of these two assays. For the Spot Formation assay, the same method was used, except that the three clones from each strain were assessed once in independent experiments, which led to 3 instead of 9 values. Finally, after adjusting and assembling the values from the four batches, the distributions of the mutant and WT BRCA1 strains were composed of 3 and 12 values respectively. No Vector control strain was used for this assay ("no spot formation" is the theoretical negative control). For the Yeast Localization assay, the same method as for the Spot Formation assay was applied, but without data normalization, due to the WT BRCA1 values close to zero, which otherwise would severely amplify data variation. Of note, the images used in the Spot Formation assay were the same

used in the Yeast Localization assay, which allows the comparison of the relative efficiency of these two assays.

7. The MWW method

The distribution of each mutant was compared to the distribution of the WT BRCA1 using the Mann-Whitney-Wilcoxon (MWW) test. The p value of this test gives the probability to obtain the observed overlap of the two compared distributions, in the random sample, assuming that the two distributions are identical in the population from which is performed the sampling (statistical population). The null hypothesis H_0 is $P(x_{\text{mutant } i} > x_{\text{BRCA1 } j}) = P(x_{\text{mutant } i} < x_{\text{BRCA1 } j})$. If H_0 is true, this means that the probability $P(x_{\text{mutant } i} > x_{\text{BRCA1 } j})$, of having a mutant value $x_{\text{mutant } i}$ above a BRCA1 value $x_{\text{BRCA1 } j}$, is equal to the probability $P(x_{\text{mutant } i} < x_{\text{BRCA1 } j})$, of having a mutant value $x_{\text{mutant } i}$ below a BRCA1 value $x_{\text{BRCA1 } j}$. In an upper-sided MWW test, the alternative hypothesis is that the mutant values tend to be above the BRCA1 values, which is written as $P(x_{\text{mutant } i} > x_{\text{BRCA1 } j}) > P(x_{\text{mutant } i} < x_{\text{BRCA1 } j})$ [30]. Importantly, the p values computed here are not used to reject or not reject the null hypothesis of the test, but to quantify the overlap between the mutant and the WT BRCA1 distributions (**S4 Fig**). This defines relative positions of the mutant distributions using the WT BRCA1 distribution as a reference position. Thus, the assumptions necessary for the MWW test [30] are not required here. Relative positions are limited to the environment of the WT BRCA1 distribution. Indeed, pathogenic variants showing no overlap with the WT BRCA1 distribution have the same relative position (same p value), even if the functional assay identifies differences between them. Normal approximation and continuity correction were systematically applied to anticipate the presence of identical values (ties). For that, the R function used was `wilcox.test(..., exact=FALSE, correct=TRUE)`. Of note, using these parameters, an increasing number of ties has a tendency to decrease the p value computed. With $n_{\text{mutant}} = 9$ and $n_{\text{BRCA1}} = 36$ (Colony Size and Liquid Medium assays), the p values of the MWW test theoretically range from $2.3e-6$ to 1 without ties, and from $2e-11$ to 1 with a maximum number of ties. With $n_{\text{mutant}} = 3$ and $n_{\text{BRCA1}} = 12$ (Spot Formation and Yeast Localization assay), the p values of the MWW test theoretically range from 0.0058 to 0.9962 without ties, and from 0.00014 to 1 with a maximum number of ties. In addition, particular results have to be mentioned when using `correct= TRUE` and one-tailed tests. For instance, the p value obtained for two distributions that perfectly overlap, (e.g., `wilcox.test(1:3, 1:3, alternative = "less", exact=FALSE, correct=TRUE)`) is slightly upper than 0.5. This has no consequence in variant classification, since p values are used as relative positions, as long as the number of values per variant remains identical. In addition, the p value obtained for two distributions that show the same unique value (e.g., `wilcox.test(c(1,1,1), c(1,1,1), alternative = "less", exact=FALSE, correct=TRUE)`) is equal to 1. In such extreme situations, it is recommended to use `correct=FALSE`.

For the Colony Size, Liquid Medium and Yeast Localization assays, upper-sided MWW tests were performed (the hypothesis being that the distributions of the pathogenic mutations are above the distribution of the WT BRCA1 reference). The R function used was `wilcox.test(WT reference values, mutant i values, alternative = "less", exact=FALSE, correct=TRUE)`. For the Spot Formation assay, lower-sided MWW tests were performed (the hypothesis being that the distributions of the pathogenic mutations lie below the distribution of the WT BRCA1 reference). The R function used was `wilcox.test(WT reference values, mutant i values, alternative = "greater", exact=FALSE, correct=TRUE)`. Results are summarized in **S4 Table**.

8. Sensitivity and specificity computation using the standard method

Medians from the mutant distributions were ordered (as in the waterfall representation, **Fig 1A**) and means were computed between every two consecutive medians. These mean values

were defined as all possible cut-offs within the ordered medians. Sensitivity and specificity were computed for each cut-off. Sensitivity corresponded to the proportion of pathogenic mutant medians above (Colony Size, Liquid Medium and Yeast Localization assays) or below (Spot Formation assay) the cut-off. Specificity corresponded to the proportion of neutral mutant medians below (Colony Size, Liquid Medium and Yeast Localization assays) or above (Spot Formation assay) the cut-off. Of note, cut-offs above or below all of the medians were not considered. This means that the sensitivity/specificity of (0, 1) and (1, 0) were excluded from the study. In **S2 Fig**, the confidence intervals of the sensitivity and specificity were computed with the `binom.test(..., alternative = "two.sided", conf.level = 0.95)` function of R, considering that the theoretical sensitivity or specificity is 100% ($p = 1$). The receiver operating characteristic (ROC) curve offers a visual representation of the sensitivities and specificities computed. For the x-axis, it is common to use 1-specificity more than specificity, as 1-specificity represents the false positive rate (FPR). The dotted line showed in **S2 Fig** indicates the positive diagonal, for which sensitivities = 1 - specificities. The top left corner of the ROC chart corresponds to sensitivity = 1 and specificity = 1 (1 - specificity = 0), which represents the optimal situation for any given assay.

9. Difference between the standard method and the standard with reference methods

We distinguish the standard method from the standard with reference method. The first uses raw cut-offs, meaning that the cut-off values are not divided by the median of the WT BRCA1 reference (see the definitions above). Thus, the cut-off values have the unit of the experimental data (e.g., cells per colony in the Colony Size assay). In the standard with reference method, cut-off values are divided by the median of the WT BRCA1 reference distribution. Nothing changes between these 2 methods when working with the experimental data (i.e., experimental sensitivity and specificity are the same for both methods). However, results change when performing bootstrap analysis, because in the standard with reference method, the fluctuation of the best cut-off is influenced by the fluctuation of the WT BRCA1 reference, which is not the case in the standard method, that only depends on sampling the neutral and pathogenic mutant values (described below).

10. Sensitivity and specificity computation using the MWW method

The same as the standard method was applied, but for ordered p values (e.g., **Fig 1B**) instead of medians. Thus, the final cut-offs analyzed in ROC curves were the intermediate p values between the ordered mutant p values.

11. Best cut-off, best sensitivity and best specificity computation

In the standard or MWW method, the best compromise between the highest sensitivity and specificity was determined by the nonparametric empirical Youden's index [31-33], which is defined as $\max[\text{sensitivity}_i + \text{specificity}_i - 1]$ for each cut-off i . The Youden's index corresponds to the dot on the ROC curve that maximizes the vertical distance between the positive diagonal and the ROC curve. Such dots were pinpointed as black numbers in **S2 Fig**. The best cut-off was defined as the cut-off of the Youden's index. The best sensitivity and best specificity were defined as those associated with the best cut-off. The standard method and the standard with reference methods generated systematically the same best experimental sensitivity and best experimental specificity, regardless of the data analyzed, since the best cut-off is the same in these two methods, except for the unit (raw best cut-off in the standard method, and relative best cut-off in the standard with reference method).

12. Bootstrap procedure A to define three different fluctuations of the best cut-off

Nonparametric random samplings [34] were performed as follows. The number of sampled values was systematically the number of values available in each mutant and in the WT BRCA1 reference. Thus, for each mutant, n_{mutant} values were randomly chosen with replacement. The same was performed for the BRCA1 distribution, with n_{BRCA1} values randomly chosen with replacement. As an example, using the Colony Size assay, the number of values randomly chosen was $n_{\text{mutant}} = 9$ and $n_{\text{BRCA1}} = 36$, since respectively 9 and 36 values were available in these distributions. In the Spot Formation assay, the number of values randomly chosen was $n_{\text{mutant}} = 3$ and $n_{\text{BRCA1}} = 12$, since respectively 3 and 12 values were available in these distributions. Next, using this new set of sampled data, we applied the three standard, standard with reference and MWW methods. In the standard method, medians of the mutant distributions were ordered, raw cut-offs were computed as described above, and the raw best cut-off value associated with the Youden's index of the ROC curve, was saved. In the standard with reference method, the raw best cut-off, identified by the standard method, was divided by the median of the WT BRCA1 reference computed on the new sample data and was saved. In the MWW method, p values for each mutant were computed as described above, and ordered. Next, cut-offs were computed as described above, and the best cut-off value, associated with the Youden's index of the ROC curve, was saved. In each of the three methods, if several cut-offs lead to the Youden's index, the median of these cut-offs was considered as the best cut-off. This procedure was repeated 2,000 times, to obtain 2,000 best cut-off values for each of the three methods. Next, the 2,000 bootstraps were repeated 20 times. This procedure is referred to as "bootstrap procedure A", which, in summary, generated 20 sets of 2,000 best cut-offs for each of the standard, standard with reference and MWW methods.

It is important to mention that, in the standard with reference method, best cut-off values are multiplication factors of the BRCA1 median value. This means that fluctuation of the BRCA1 median is included in the best cut-off fluctuation. For instance, with the Colony Size assay: during the random sampling i , if the raw best cut-off found is $b_{\text{raw } i}^* = 25,000$ cells per colony (the star indicates that the value comes from bootstrapping) and if the BRCA1 median is $\tilde{x}_{\text{BRCA1 } i}^* = 10,000$, then the relative best cut-off is $b_i^* = b_{\text{raw } i}^* / \tilde{x}_{\text{BRCA1 } i}^* = 2.5$. In the random sampling $i+1$, if $b_{\text{raw } i+1}^* = 25,000$ cells per colony and if $\tilde{x}_{\text{BRCA1 } i+1}^* = 5,000$, then the relative best cut-off is $b_{i+1}^* = 5$. These examples highlight the fact that the relative best cut-off values can change, even if the raw best cut-off value of the standard method remains the same.

13. Quantile computation after bootstrap procedure A

In the first set of the bootstrap procedure A, a total of eleven quantiles, 0.1%, 0.5%, 1%, 2.5%, 5%, 50% (median), 95%, 97.5%, 99%, 99.5% and 99.9%, were computed from the 2,000 best cut-off values of the standard method (using the type 7 method of the `quantile()` function of R). This procedure was repeated for all of the 20 sets. Next, medians of the 20 values from each of the eleven quantiles were computed (of note, these quantiles could have been directly computed from the $20 \times 2,000 = 40,000$ best cut-off values, but medians of quantiles have the advantage to buffer the fluctuation of extreme quantiles). The same was performed for the standard with reference and MWW methods. Thus, eleven median quantiles were obtained for each of the three methods, shown in **S5 Table**. Of note, the basic (non-studentized pivotal) method was used to obtain the quantiles [35].

14. Probability system of classification

This system is based on rank methods [14]. The following procedure was separately applied to the three standard, standard with reference and MWW methods. In each of the 20 bootstrap

sets, resulting from the bootstrap procedure A, the 2,000 best cut-off values were sorted by ascending or descending order, depending on the position of the neutral and pathogenic sectors in the assay and depending on the method used: ascending order if the lowest values correspond to the neutral sector and the highest to the pathogenic sector (e.g., Colony Size assay + standard method), and descending order in the opposite case (e.g., MWW method, whatever assay is used). Next, the median of the 20 first values, from the 20 sorted sets, was computed, and the same for the next 1,999 subsequent ranks, to obtain a median distribution of the sorted best cut-offs. Next, the probability $1 / n_{best} = 0.0005$ was attributed to each of the $n_{best} = 2,000$ values of the median distribution (see the theoretical example in **S6A-B Fig**). Probabilities were summed in the case of identical values, which finally resulted in n_{diff} values in the median distribution of the sorted best cut-offs (e.g., $n_{diff} = 154$ in the Yeast Localization assay, using the standard method, as shown in **S6C Fig** and **S12 Table**). From this, two cumulative distribution function (CDF) were generated, by simply carrying out the cumulative sum of the n_{diff} probabilities, with the first CDF that ranged from the probability p_1 to $p_{n_{diff}} = 1$ and the second CDF that ranged from 0 to $p_{n_{diff}-1} < 1$. Next, the mean of the two CDF was computed, which resulted in an average CDF that associated a probability ($0 < p_i < 1$) for the n_{diff} different best cut-off values. This average CDF was obtained for each of the Colony Size, Liquid Medium, Spot Formation and Yeast Localization assays and for each of the standard, standard with reference and MWW methods (**S6C-E Fig**).

To attribute a probability of pathogenicity to an assessed variant, using the standard method, the median of this variant was positioned within the raw best cut-off values of the average CDF. Next, the probability of the average CDF, closest to the variant median, was assigned to the variant as a probability of pathogenicity (see the example in **S6B Fig**). The same procedure was applied to the standard with reference method, except that the median of the variant was divided by the median of the WT BRCA1 reference median (i.e., 11,200 cells per colony in the Colony Size assay), to fit the average CDF composed of relative best cut-off values. The same procedure was applied to the MWW method, except that the p value of the variant was used to fit the average CDF. The five-class nomenclature [26] was used to categorize the probabilities of pathogenicity, as in genetic/epidemiological methods (**S1 Table**).

Of note, interest in averaging the two initial CDF lies in the removal of the probabilities 0 and 1 in the resulting CDF, which could create infinite values during subsequent conversions. The second interest is that the average CDF copes with the absence of best cut-off fluctuation. In such situation, the average CDF is represented by a single best cut off value (equal to the experimental best cut-off) which has the probability 0.5. This means that in the absence of best cut-off fluctuation, the classification proposed by a given functional assay, is systematically "variant completely unknown". The convergence towards 0.5 is illustrated using the boundaries of the average CDF. If the number of best cut-off values is n_{best} , then the lowest probability within the CDF is $1/n_{best}$. Thus, in the CDF1, the boundaries of the cumulated probabilities are:

$$\left[\frac{1}{n_{best}} ; 1 \right]$$

n_{best} : number of best cut-off values obtained after bootstrap

And in the CDF2:

$$\left[0 ; 1 - \frac{1}{n_{best}} \right]$$

Thus, in the average CDF, the boundaries are:

$$\left[\frac{\frac{1}{n_{best}} + 0}{2} ; \frac{1 + 1 - \frac{1}{n_{best}}}{2} \right] = \left[\frac{1}{2n_{best}} ; 1 - \frac{1}{2n_{best}} \right]$$

And when $n_{best} = 1$, both boundaries of the average CDF are equal to 0.5.

15. Exact best cut-off distribution

In certain situations, it is possible to use the exact best cut-off distribution, instead of performing bootstrap analysis. This exact distribution is defined as all of the sampling possibilities, when sampling with replacement n_{mutant} values among the n_{mutant} available, for each variant, and n_{BRCA1} values among the n_{BRCA1} available for the WT reference. An example is shown in **S21B Fig**. The number of sampling possibilities can be predicted as follows. For one variant, composed of n_{mutant} different values, the number of sampling possibilities is:

$$\binom{2n_{mutant} - 1}{n_{mutant}}$$

n_{mutant} : number of values in each mutant (implicating the same number in the different mutants)

With $n_{neutral}$ and $n_{pathogenic}$ variants, composed of the same number of values n_{mutant} , all different, the number of sampling possibilities is:

$$\binom{2n_{mutant} - 1}{n_{mutant}}^{n_{neutral} + n_{pathogenic}}$$

$n_{neutral}$: number of neutral mutations

$n_{pathogenic}$: number of pathogenic mutations

Finally, with the inclusion of the WT reference (standard with reference and MWW methods), composed of n_{BRCA1} different values, the number of sampling possibilities is:

$$\binom{2n_{mutant} - 1}{n_{mutant}}^{n_{neutral} + n_{pathogenic}} \times \binom{2n_{BRCA1} - 1}{n_{BRCA1}}$$

n_{BRCA1} : number of values in the WT BRCA1 reference

As an example, with one neutral and one pathogenic mutation ($n_{neutral} = 1$ and $n_{pathogenic} = 1$), containing two values per mutant ($n_{mutant} = 2$) and two values in the WT BRCA1 reference ($n_{BRCA1} = 2$), the number of sampling possibilities is 27 (27 rows in the table of **S21B Fig**).

In the standard method, the formula shows that, if $n_{mutant} = 1$, then the exact best cut-off distribution corresponds to the experimental best cut-off, and the same for the standard with reference and MWW methods if, additionally, $n_{BRCA1} = 1$ (**S22A Fig**). The formula also shows that the number of sampling possibilities rapidly increases with the number of mutant

values, BRCA1 values, neutral variants and pathogenic variants, which prevents the use of the exact best cut-off distribution in most cases. For instance, in the Colony Size assay, the maximum number of sampling possibilities reaches $2e175$ using the standard method and $6e195$ using the MWW method. The bootstrap procedure has the advantage to by-pass this combinatory issue. However, it is important to mention that bootstrap affords an approximate distribution of the best cut-off, and that the quality of the approximation increases with the number of bootstraps $n_{\text{bootstrap}}$ performed.

Of note, two different sampling possibilities can result in the same best cut-off (for instance, the best cut-off between 10 and 20, and between 5 and 25, is 15 in both cases). Thus, even if the number of mutant values, BRCA1 values, neutral variants and pathogenic variants, are high (generating a high number of sampling possibilities), the final number of different best cut-off values $n_{\text{exact diff}}$, forming the average CDF of the probability system of classification, can be very low, even within an exact distribution.

16. Properties of the probability of pathogenicity

16.1. Associated risk

The probability of pathogenicity computed is the probability to have the best cut-off value below (ascending average CDF) or above (descending average CDF) the considered mutant, which is related to the probability of misclassification of this mutant, due to the fluctuation of the best cut-off. As an example, in the Colony Size assay using the standard method (ascending average CDF, **S6C Fig**, left panel), a variant with a probability of pathogenicity of 0.99 indicates that the best cut-off variable has a 99% chance to be below the variant median. In this example, the pathogenic area is above the best cut-off (**Fig 1A**), which means that this variant has a 1% probability of being classified as neutral (i.e., best cut-off above the variant median), due to the fluctuation of the best cut-off. In the same manner, again from the Colony Size assay using the standard method, a variant with a probability of pathogenicity of 0.001 indicates that the best cut-off variable has 0.1% chance to be below the variant median, which means that this variant has a 0.1% probability of being classified as pathogenic (i.e., best cut-off below the variant median), due to the fluctuation of the best cut-off, and, thus, has a 99.9% probability of being classified as neutral (i.e., best cut-off above the variant median). For descending average CDF, like in the Colony Size assay using the MWW method (**S6E Fig**, left panel), the reasoning is the opposite. It is important to mention that, contrary to genetic/epidemiological methods, for which the probability of pathogenicity computed measures a direct association of the variant with disease, here the probability of pathogenicity computed evaluates the risk to misclassify a variant, due to the fluctuation of the best cut-off, that depends on the fluctuation of the experimental data when performing the variant assessments. Finally, this probability of pathogenicity estimates the reproducibility of the variant classification obtained, following functional assessment.

16.2. Paucity of experimental data impairs the variant classification

As illustrated in the sections 16.4 and 16.5, the accuracy of the probability system decreases when the best cut-off distribution is composed of only a few different best cut-off values. Ultimately, when the best cut-off distribution is represented by one value (**S22A,C Figs**), the probability of pathogenicity 0.5 (class 3) is systematically attributed to all of the variants, regardless of their relative position. The weak number of different best cut-off values can result from (1) a low number of experimental replicates (**S18G Fig**), (2) a low number of neutral and pathogenic variants incorporated (except if the number of experimental replicates is high), and (3) a weak measurement accuracy leading to many ties (see the case of null ranges in **S19G Fig**). Thus, the probability system has the advantage to penalize functional assays with a paucity of experimental data. This situation is illustrated in the Yeast

localization assay using the standard method (**Fig 2B**). With $n_{\text{diff}} = 154$ (**S6C Fig** and **S12 Table**) none of the pathogenic mutations was classified as class 5.

16.3. Paucity of bootstrap performed impairs the variant classification

As illustrated in sections 16.4 and 16.5, accuracy of the probability system is decreased when the best cut-off distribution is composed of very few different best cut-off values, which is the case if the number of bootstraps performed ($n_{\text{bootstrap}}$) is low (**S22B Fig**). In contrast, if $n_{\text{bootstrap}}$ is high, then the accuracy of the probability system will only depends on the experimental data (**S22A Fig**).

16.4. Accuracy of the probability system (exact distribution)

The accuracy of a system of classification is defined as the capability to assign a class 1 or 2 to the neutral variant, and a class 4 or 5 to the pathogenic variants. A completely accurate system will never assign the class 3. A system fully inaccurate will assign the class 3 only, or will totally misclassify the neutral and pathogenic variants. In the probability system of classification, the accuracy is related to the probability unit of the average CDF. This probability unit is defined as the lowest potential incrementation within the average CDF or, which is equivalent, as the probability associated with a best cut-off value not repeated in the best cut-off distribution. In an exact distribution, this unit is equal to $1 / n_{\text{best exact}}$, meaning $1 / 27 = 0.04$ in **S21B Fig**. As shown in section 14, this defines the lowest and highest potential boundaries of the average CDF, derived from the exact distribution:

$$\left[\frac{1}{2n_{\text{best exact}}} ; 1 - \frac{1}{2n_{\text{best exact}}} \right]$$

$n_{\text{best exact}}$: number of best cut-off values in the exact best cut-off distribution (equal to the number of sampling possibilities)

In the probability system of classification, this implies that a neutral variant cannot have a probability of pathogenicity less than $1 / (2n_{\text{best exact}})$. In the same manner, a pathogenic variant cannot have a probability of pathogenicity more than $1 - 1 / (2n_{\text{best exact}})$. Thus, if $n_{\text{best exact}} \leq 10$, the probability system will be unable to classify variants as class 2 or 1, because the lower boundary will not be less than 0.05 (**S1 Table**). Such limitations of the probability system of classification are recapitulated below:

$$\text{Class 1 and 2: } \frac{1}{2n_{\text{best exact}}} < 0.05 \rightarrow n_{\text{best exact}} > 10 \quad (\text{Relations 1})$$

$$\text{Class 1: } \frac{1}{2n_{\text{best exact}}} < 0.001 \rightarrow n_{\text{best exact}} > 500$$

$$\text{Class 4 and 5: } 1 - \frac{1}{2n_{\text{best exact}}} \geq 0.95 \rightarrow n_{\text{best exact}} \geq 10$$

$$\text{Class 5: } 1 - \frac{1}{2n_{\text{best exact}}} > 0.99 \rightarrow n_{\text{best exact}} > 50$$

Because $n_{\text{best exact}}$ is equal to the number of sampling possibilities, which is dependent on the number of values n_{mutant} in each mutant, the number of values n_{BRCA1} in the WT BRCA1 reference (except for the standard method), the number of neutral mutations n_{neutral} (if $n_{\text{mutant}} > 1$) and the number of pathogenic mutations $n_{\text{pathogenic}}$ (if $n_{\text{mutant}} > 1$), this means that the accuracy of the probability system of classification is dependent on these parameters. As an example, in **S21D Fig**, with 1 neutral variant, 1 pathogenic variant, 2 values per variant and 2

values in the WT BRCA1 reference, the number of sampling possibilities is sufficient to allow the class 2 and class 4 classification, but not class 1 and class 5, in the standard with reference method.

Of note, an increasing number of ties (identical values) in the dataset, lowers the accuracy of the probability system of classification. Indeed, ties reduce the number of different best cut-off values, which reduces the interval of the average CDF (boundaries tend towards 0.5). Thus, the lowest $n_{\text{best exact}}$ values, indicated in Relation 1, are indicative. Moreover, the measurement accuracy has to be considered during functional assessment, to prevent ties.

The fact that two different sampling possibilities can result in the same computed best cut-off, also lowers the accuracy of the probability system of classification. As an example, even with no ties in the data set (**S21A Fig**), the reduced number of different best cut-off, finally obtained with the standard method (5 for 27 sampling possibilities), prevents the classification of variants other than class 3 (**S21C Fig**). This phenomenon, leading to identical best cut-off values, is not predictable, but it highlights again that the lowest $n_{\text{best exact}}$ values, indicated in Relation 1, are indicative.

16.5. Accuracy of the probability system (bootstrap)

The considerations, developed in the precedent section (16.4), are also valid when using the approximate best cut-off distribution (obtained by bootstrap), instead of the exact distribution, except that the number of bootstraps performed, $n_{\text{bootstrap}}$, is an additional parameter that influences the accuracy of the probability system of classification. More precisely, since the bootstrap procedure gives an estimation of the exact best cut-off distribution (**S22B Fig**), the effect of $n_{\text{bootstrap}}$ depends on $n_{\text{best exact}}$.

If $n_{\text{bootstrap}} \ll n_{\text{best exact}}$, then the lowest potential probability unit is:

$$\frac{1}{n_{\text{bootstrap}}}$$

$n_{\text{bootstrap}}$: number of bootstraps performed to estimate the best cut-off fluctuation

And the lowest and highest potential boundaries are:

$$\left[\frac{1}{2n_{\text{bootstrap}}} ; 1 - \frac{1}{2n_{\text{bootstrap}}} \right]$$

If $n_{\text{bootstrap}} \gg n_{\text{best exact}}$, then the lowest potential probability unit is:

$$\frac{1}{n_{\text{best exact}}}$$

$n_{\text{best exact}}$: number of best cut-off values in the exact best cut-off distribution (equal to the number of sampling possibilities)

And the lowest and highest potential boundaries are:

$$\left[\frac{1}{2n_{\text{best exact}}} ; 1 - \frac{1}{2n_{\text{best exact}}} \right]$$

Using the probability system of classification, this implies that a neutral variant cannot have a probability of pathogenicity less than $1 / (2n_{\text{best exact}})$, regardless of the number of bootstraps performed. This also implies that the lowest probability of pathogenicity of a neutral variant is overestimated if $n_{\text{bootstrap}} \ll n_{\text{best exact}}$. In the same manner, a pathogenic variant cannot have a probability of pathogenicity more than $1 - 1 / (2n_{\text{best exact}})$, regardless of the number of bootstraps performed. This also implies that the highest probability of pathogenicity of a pathogenic variant is underestimated if $n_{\text{bootstrap}} \ll n_{\text{best exact}}$.

As an example, in **S21 Fig**, if two bootstraps are performed, then $n_{\text{bootstrap}} = 2 \ll n_{\text{best exact}} = 27$. If we consider that the two best cut off values, obtained with the standard with reference method, are different, then the average CDF is represented by the two values 0.25 and 0.75, with the boundaries [0.25; 0.75]. Thus, the probability of pathogenicity attributed to the neutral variant is 0.25, which is overestimated compared to the 0.02 probability given by the exact distribution (**S21D Fig**). In the same manner, the probability of pathogenicity attributed to the pathogenic variant is 0.75, which is underestimated as compared to the 0.98 probability given by the exact distribution (**S21D Fig**).

Of note, if the number of different best cut-off values is much lower than the number of bootstraps performed ($n_{\text{diff}} \ll n_{\text{bootstrap}}$), this suggests that the approximate best cut-off distribution is close to the exact one. For instance, in the Liquid Medium assay using the standard method (**S6C Fig**), the $n_{\text{bootstrap}} = 2,000$ bootstraps generated an approximate best cut-off distribution, composed of $n_{\text{diff}} = 126$ different best cut-off values. Thus, in this case, the limitations of the probability system of classification are probably due to $n_{\text{best exact}}$ rather than $n_{\text{bootstrap}}$.

17. Combined probability of pathogenicity

Let us consider one of these three methods: standard, standard with reference or MWW. For each variant, four probabilities of pathogenicity were obtained, from the four Colony Size, Liquid Medium, Spot Formation and Yeast Localization assays (**S13-S15 Tables**). Combining these probabilities provides a final probability of pathogenicity. The model proposed was derived from the one used in genetic/epidemiological methods [15]. Probabilities were converted into odds in favor of pathogenicity using the formula:

$$O_i = \frac{p_i}{1 - p_i}$$

- O_i : odds in favor of pathogenicity of the variant i ($0 \leq O_i < +\infty$)
 p_i : probability of pathogenicity of the variant i ($0 \leq p_i \leq 1$)

With odds in favor of pathogenicity, the pathogenicity varies between 0 (absolutely neutral) and $+\infty$ (absolutely pathogenic). A variant i , of fully unknown significance, has $p_i = 0.5$ and thus $O_i = 1$. Only independent probabilities can be combined. Thus, probabilities from the Liquid Medium assays were excluded from the computation of the combined probabilities, since the Colony Size and Liquid Medium results were derived from the same yeast clones. In the same manner, probabilities from the Yeast Localization assays were excluded, since the Spot Formation and Yeast Localization results were derived from the same yeast clones and the same microscope picture acquisitions. Next, odds from the Colony Size and Spot Formation assays were multiplied:

$$O_{\text{comb } i} = O_{\text{CS } i} \times O_{\text{SF } i}$$

- $O_{\text{comb } i}$: combined odds in favor of pathogenicity of the variant i ($0 \leq O_{\text{comb } i} < +\infty$)

And the combined probability of pathogenicity was obtained using the formula:

$$p_{comb\ i} = \frac{O_{comb\ i}}{1 + O_{comb\ i}}$$

$p_{comb\ i}$: combined probability of pathogenicity of the variant i ($0 \leq p_{comb\ i} < 1$)

As a reminder, the accuracy of the probability system of classification is decreased together with n_{mutant} , n_{BRCA1} , $n_{neutral}$ and $n_{pathogenic}$, or together with $n_{bootstrap}$ (see sections 16.4 and 16.5). This means that the probabilities of pathogenicity attributed to the assessed variants will tend towards 0.5. Thus, it is remarkable that an assay, showing low n_{mutant} , n_{BRCA1} , $n_{neutral}$ and $n_{pathogenic}$ values, or for which a low number of bootstraps has been performed, will have a low contribution in the combined probability of pathogenicity. Indeed, when the pathogenicity of a variant tends towards 0.5, the odds tend towards 1, which has a null effect in odds multiplication.

18. Independent functional assays

We estimate that the independence of two functional assays cannot be evaluated *a posteriori*, for instance by comparing the results of these assays. Indeed, in the *a posteriori* evaluation, the notion of independence is based on the hypothesis that the pathogenic variants can have a random position in the pathogenic area of a waterfall distribution. Thus, if we compare two assays, challenged by the same set of pathogenic mutations, the random position of the pathogenic variants, in each assay, should result in a low correlation between the variant medians, which would confirm the independence between the two assays. However, it is known that certain pathogenic mutations have intermediate effects on protein function [36]. Thus, it is difficult to ascertain a random position of the pathogenic mutations in the pathogenic area. Focusing on neutral mutations, rather than on pathogenic mutations, would not solve this issue, since intermediate effects cannot be excluded for neutral mutations. Thus, we propose to evaluate the independence of two assays *a priori*, with the following criteria:

1) The two assessments must have been performed independently. This means that the results from each assay must come from different transfections, different cellular clones, different experimental times, etc. The fact that the plasmids used are not exactly the same reinforces the independence. For instance, in the Spot Formation assay, the plasmids code for the mCherry-BRCA1 fusion protein, but not in the Colony Size assay. Thus, the plasmids used are different in these two assays.

2) What is monitored must be different. Following this, the Colony Size assay and the Liquid Medium assay are not independent, since both monitor the cell growth of yeast cells. If assays are not clearly associated with a protein function, then the measurements must be different. For instance, the Colony Size and the Spot Formation assays are not related to a known function of BRCA1, but the Colony Size assay monitors cell growth, while the Spot Formation assay monitors the formation of a cellular aggregate. Thus, the measurement is different for these two assays.

Based on these criteria, The Colony Size and the Liquid Medium assays cannot be considered as independent, and the same for the Spot Formation and the Yeast Localization assays.

19. Corrected probability of pathogenicity

The fluctuation of the best cut is influenced by the experimental (initial) sensitivity and specificity of a functional assay (**S16 Fig**) but not by the number of neutral and pathogenic mutations used to determine these parameters (**S17 Fig**). This could be problematic in variant

classification, since a 100% sensitivity and specificity of an assay, resulting from 2 neutral and 2 pathogenic mutations assessed, are not reliable. To overcome this, we propose an approach to correct the odds in favor of pathogenicity, that takes into account the number of mutants used to evaluate a functional assay, as follows:

$$O_{cor\ i} = (O_i - 1) \times f_{cor} + 1 \quad \text{if } O_i \geq 1$$

$$O_{cor\ i} = \frac{1}{\left(\frac{1}{O_i} - 1\right) \times f_{cor} + 1} \quad \text{if } O_i < 1$$

$$f_{cor} = \frac{n_{neutral} + n_{pathogenic}}{n_{neutral} + n_{pathogenic} + a}$$

$$p_{cor\ i} = \frac{O_{cor\ i}}{1 + O_{cor\ i}}$$

- $O_{cor\ i}$: corrected odds in favor of pathogenicity of the variant i
- O_i : odds in favor of pathogenicity of the variant i ($0 \leq O_i < +\infty$)
- f_{cor} : factor of correction
- $n_{neutral}$: number of neutral mutations used in the assay validation
- $n_{pathogenic}$: number of pathogenic mutations used in the assay validation
- a : integer that modulates the impact of $n_{neutral}$ and $n_{pathogenic}$
- $p_{cor\ i}$: corrected probability of pathogenicity of the variant i ($0 \leq p_{cor\ i} \leq 1$)

In this study, we chose $a = 2$ as a correcting factor (**S13-S15 Tables**), which results in $f_{cor} = 0.5$ for $n_{neutral} + n_{pathogenic} = 2$, $f_{cor} = 0.8$ for $n_{neutral} + n_{pathogenic} = 8$, $f_{cor} = 0.9$ for $n_{neutral} + n_{pathogenic} = 18$ and $f_{cor} = 0.95$ for $n_{neutral} + n_{pathogenic} = 38$ (**S22D Fig**). Since $n_{neutral}$ and $n_{pathogenic}$ are not distinguished in the correction, this method of correction is not appropriate if $n_{neutral}$ and $n_{pathogenic}$ are strongly unbalanced.

20. Incorporation of the probability of pathogenicity into posterior probability models (Bayesian inference)

The probability of pathogenicity derived from the best cut-off fluctuation can be used to compute a posterior probability of being pathogenic, considering the Bayes' theorem [37]:

$$O_{post\ i} = O_i \times O_{prior\ i}$$

$$O_{prior\ i} = \frac{p_{prior\ i}}{1 - p_{prior\ i}}$$

- $O_{post\ i}$: posterior odds in favor of pathogenicity of the variant i ($0 \leq O_{post\ i} < +\infty$)
- O_i : odds in favor of pathogenicity of the variant i ($0 \leq O_i < +\infty$)
- $O_{prior\ i}$: prior odds in favor of pathogenicity of the variant i ($0 \leq O_{prior\ i} < +\infty$)
- $p_{prior\ i}$: prior probability of pathogenicity of the variant i ($0 \leq p_{prior\ i} \leq 1$)

The prior probability represents the probability of pathogenicity of the variant i before any functional assessment. A prior probability of 0.5 can be used in the absence of any prior information. In this case, $O_{post\ i} = O_i$. Prior probabilities can be provided by the GVG D Align model [38]. Then, the posterior probability of being pathogenic is obtained following:

$$p_{post\ i} = \frac{O_{post\ i}}{1 + O_{post\ i}}$$

$p_{post\ i}$: posterior probability of pathogenicity of the variant i ($0 \leq p_{post\ i} \leq 1$)

This method can also be applied to the combined odds ($O_{comb\ i}$) and the corrected odds ($O_{cor\ i}$), described above.

21. The WT reference

It is recommended to systematically add a WT reference control during variant assessment [4]. However, this may lead to divergent usage of this reference when interpreting results. The WT reference can either be included in the neutral category or be considered as a particular case, outside of the two neutral and pathogenic categories. The rationale of the first choice is that the WT reference is by definition neutral (not pathogenic). The second choice is ruled by three concerns. First, since the WT reference distribution usually contains a number of values larger than in the other neutral distributions, the WT reference could have a strong weight in the final variant interpretation, if incorporated as an additional variant in the neutral category. Second, the WT reference corresponds to the backbone cDNA sequence, present in the expression plasmid, in which the mutations are introduced. In other words, neutral mutations have two differences in their sequences, compared two by two, while they only have one difference compared to the WT reference. Additionally, the variant classification obtained could be modified using another WT reference sequence, because of potential variant-sequence interactions. Thus, the WT reference represents more than an additional neutral variant. Third, variants showing intermediate effects have been reported [36]. This highlights the benefit of the WT reference taken as a special case, beyond the neutral and pathogenic categories, because a reference would be necessary for the identification of neutral variants with intermediate effects.

22. Procedure to include new data (VUS) in the validated Colony Size, Liquid Medium, Spot Formation or Yeast Localization assays

- Download the ProClass toolbox (see section 5).
- Read carefully the *README.doc* downloaded document.
- Request the desired plasmids, presented in this study.
- Generate the plasmids containing the VUS (new batch of variants).
- Perform the experiment as described above. For instance, with the Colony Size assay, test three independent clones in three independent experiments to obtain 9 final values for the VUS. Add the WT BRCA1 reference (plasmid pPT60 or pPT63 depending on the chosen assay), as well as the pathogenic G1706E (plasmid pPT147 or pPT161) and the neutral R1751Q (plasmid pPT119 or pPT120) mutations as a control. Add the Vector control (pJL48 plasmid) for the Colony Size or the Liquid Medium assays. This means that 9 values are also obtained for the WT BRCA1 reference and for each control.
- Consider the results as a new batch and include these results in the downloaded table (e.g., Colony Size data.txt table). For the G1706E, R1751Q and Vector controls, change the name of the new values, like G1706E.bis, in order to prevent the fusion of the new values with the values of G1706E already present in the table. In addition, set these G1706E.bis, R1751Q.bis and Vector.bis controls as "Other.reference" in the "Prior_classif" column.
- Execute the code of the *Code data analysis and representation.doc* file with the adapted settings (for the Colony Size assay, use the *Colony Size data proba 2000x20 type7 replac noref less (MWW).txt* file to generate the probability of being pathogenic using the MWW method).

- In the normalized data obtained, verify that the median of the controls (G1706E.bis, R1751Q.bis and Vector.bis) fit approximately the median of the preexisting values (G1706E, R1751Q and Vector).
- If the medians fit, use the probability of being pathogenic provided in the output results of the executed code. Results can be combined, corrected and included in a posterior probability model, as explained above.
- If the median of the controls (G1706E.bis, R1751Q.bis and Vector.bis) do not fit the median of the preexisting values (G1706E, R1751Q and Vector), a new best cut-off fluctuation has to be generated. For that, remove the ".bis" in the name of the controls, in the data table, and follow the complete instructions provided in the *README.doc* downloaded document.

23. Procedure to adapt the classification model to other functional assays

- Download the ProClass toolbox (see section 5).
- Select several neutral and pathogenic mutations formally classified by genetic/epidemiological methods. Favor, if possible, a similar number of neutral and pathogenic variants.
- Design the experiment. It is counseled to plan independent experiments and to have the WT reference systematically present in each experiment.
- Generate the expression vectors, carrying either the WT reference control, or the different neutral or pathogenic mutations.
- Fix the number of values per mutation that has to be obtained (e.g., 9 values per mutation in the Colony Size assay). This number must be the same for each variant (neutral, pathogenic and unknown). This number must also be systematically respected when subsequently adding neutral and pathogenic mutations, in order to improve the sensitivity and specificity of the functional assay, or when subsequently adding VUS for classification. In addition, the number of values expected for the WT reference control should be high, in order to improve the sensitivity of variant classification (**S18F Fig**).
- Of note, if the data show different numbers of values per neutral, pathogenic and unknown variants, the code of the *Code data analysis and representation.doc* file will reduce the data as explained in the *README.doc* file.
- Perform the experiments.
- Apply the complete instructions provided in the *README.doc* downloaded document.
- Results can be combined, corrected and included in a posterior probability model, as explained above.

24. Advantages of these procedures compared to the 2-component models of variant classification

Recently, statistical models of variant classification have been proposed [7,8]. These models: (1) are parametric (assumption about the distribution of the data in the statistical population), (2) require high statistical skills to analyze the data and fit the model and (3) need to be recomputed when additional VUS are included in the model, for classification purposes.

The model we propose alleviates these constraints. This model: (1) is nonparametric (no assumption about the distribution of the data in the statistical population), (2) does not require statistical skills to be handled and (3), following certain conditions, does not need recomputation when additional VUS are included. Moreover, the probability of pathogenicity computed with our model can be incorporated into posterior probability models (Bayesian inference), as described above, meaning that they can be handled in the manner as the probabilities computed by the 2-component models.

25. Limits of the nonparametric model (MWW method and probability system)

- 1) The number of values per variants (neutral, pathogenic and UV) must be the same. This allows a balanced contribution of each neutral and pathogenic variant to the best cut-off fluctuation, as well as a correct adequacy between the unknown variants assessed and the best cut-off distribution used to classify them. The R code, available online, manages different number of values but at the cost of loss of information, as the code reduces the number of values from all the variants to the lowest existing in the data.
- 2) A WT reference must be systematically present in each experiment performed. In addition, the MWW method requires that the WT reference is well embedded in the distribution of the neutral values (e.g., **S5A Fig**, left panel). A WT reference falling outside of the range of the neutral and pathogenic distributions impairs the sensitivity of the functional assay (**S15 Fig**). In such situation, it is recommended to use the standard method instead of the MWW method.
- 3) It is counseled to have at least 3 values per variant and more than 3 values for the WT reference (**S18 Fig**). In theory, our nonparametric model does not require a minimum number of values. However, the greater the number is, the better is the sensitivity and specificity of the model (see section 16 above). Of note, with a single value per mutation and per WT reference, the fluctuation of the best cut-off is null (the probability of being pathogenic, assigned to each variant, is systematically 0.5).

26. Procedure to adapt the model to other experiment system requiring a decision-making based on cut-off

The procedure described in section 23 can be applied to all situations, based on two categories (applying the standard method) or two categories + a reference category (applying the MWW method), as long as the best cut-off is able to fluctuate during bootstrap computation. If the two categories are divided into subcategories, as "variants" in the pathogenic and neutral categories, then the best cut-off fluctuation will be guaranteed with at least two different values in each subcategory. If the two categories are not subdivided, then the best cut-off fluctuation will be guaranteed with at least two different values in each category. Additionally, in the MWW method, the reference should be composed of at least two different values. Of note, if the data show different numbers of values per categories, the code of the *Code data analysis and representation.doc* file will reduce the data as explained in the *README.doc* file. Such data adjustment has been applied to the siRNA data presented in **Fig 3**. The initial data are made of 864 values for the positive control (siKIF11), 288 values for the negative controls (siGOLGA2 and siGL2) and 12 values for the unknown siRNAs assessed. The code reduces the data to 12 values for siKIF11, siGOLGA2 and siGL2, before any subsequent analysis. The reference category is not concerned by this adjustment.

27. The quantile system is not adapted to variant classification

27.1. Presentation of the quantile system and variant classification

We developed another approach to classify variants, referred to as "quantile system", which is very easy to apply. The quantile system is similar to the "grey zone approach" [39,40] but is extended to n zones. Since the classification as either pathogenic or neutral is dependent on the position of the variant above or below the best cut-off, we reasoned that the farther a variant is from the core of the best cut-off fluctuation, the more robust is its classification as either pathogenic or neutral. The quantile system consists of (1) overlapping the best cut-off fluctuation with the waterfall distribution of the mutants, (2) defining intervals in the best cut-off distribution associated to the five-class nomenclature proposed by Plon et al. (see **S1 Table** and **Fig 1**) and (3) classifying variants according to the position of their median (standard and standard with reference methods) or p value (MWW method) in the 5 intervals.

For instance, in **Fig 1B**, the p value of K45Q is located within the grey area. Thus, the variant is classified as class 3. The quantile system is fully documented in **S23 Fig**. The classification obtained in the four functional assays is depicted in **S24 Fig**. Globally, little differences were observed when comparing the classification obtained with the quantile (**S16 Table**) and the probability systems (**S6 Table**). Of note, the light blue, grey and pink areas depicted in **Fig 1, Fig 3A-B and S7, S9, S11, S13-S19, S25-S27 Figs** correspond to class 2, 3 and 4 of the quantile system, respectively.

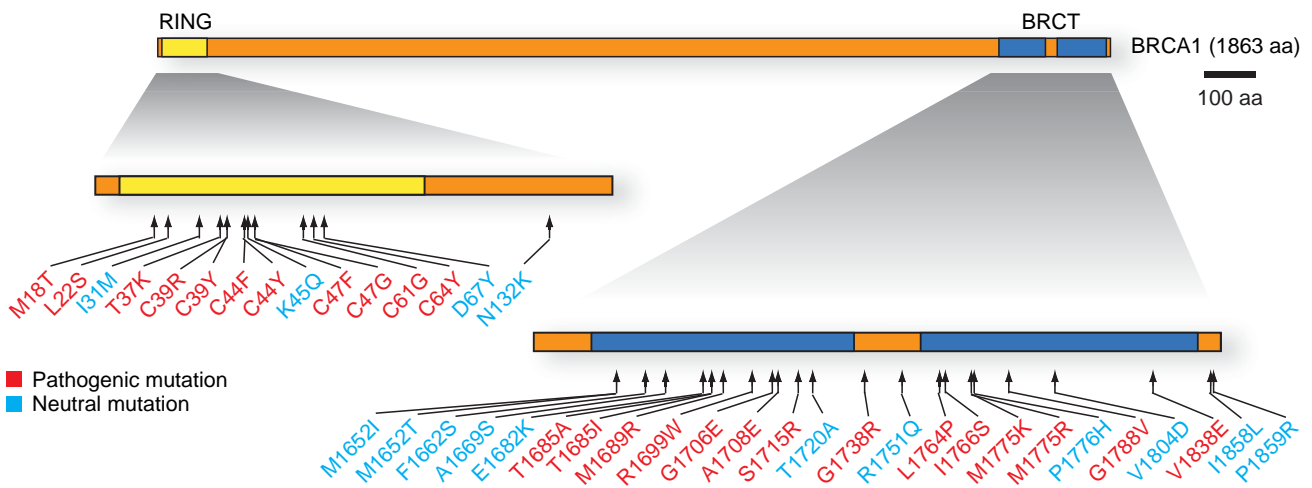
27.2. The quantile system improves the variant classification when data is lacking

To detect potential flaws in the quantile system of classification, we recapitulated the analysis of theoretical situations, performed for the probability system (**S25-S27 Figs and S17 Table**). The results reveal a major flaw in the variant classification provided by the quantile system, as it does not penalize the paucity of data (**S26B and S27 Figs**). This was confirmed with the analysis from an exact best cut-off distribution (**S21F Fig**). In fact, the highest sensitivity, specificity and accuracy of the quantile system is reached when the best cut-off does not fluctuate (**S22C Fig**), which is favored by the paucity of data. In conclusion, the quantile system is not adapted to variant classification. This also suggests that the "grey zone approach" [39,40] is not an efficient method to identify a level of uncertainty within a given dataset.

28. References

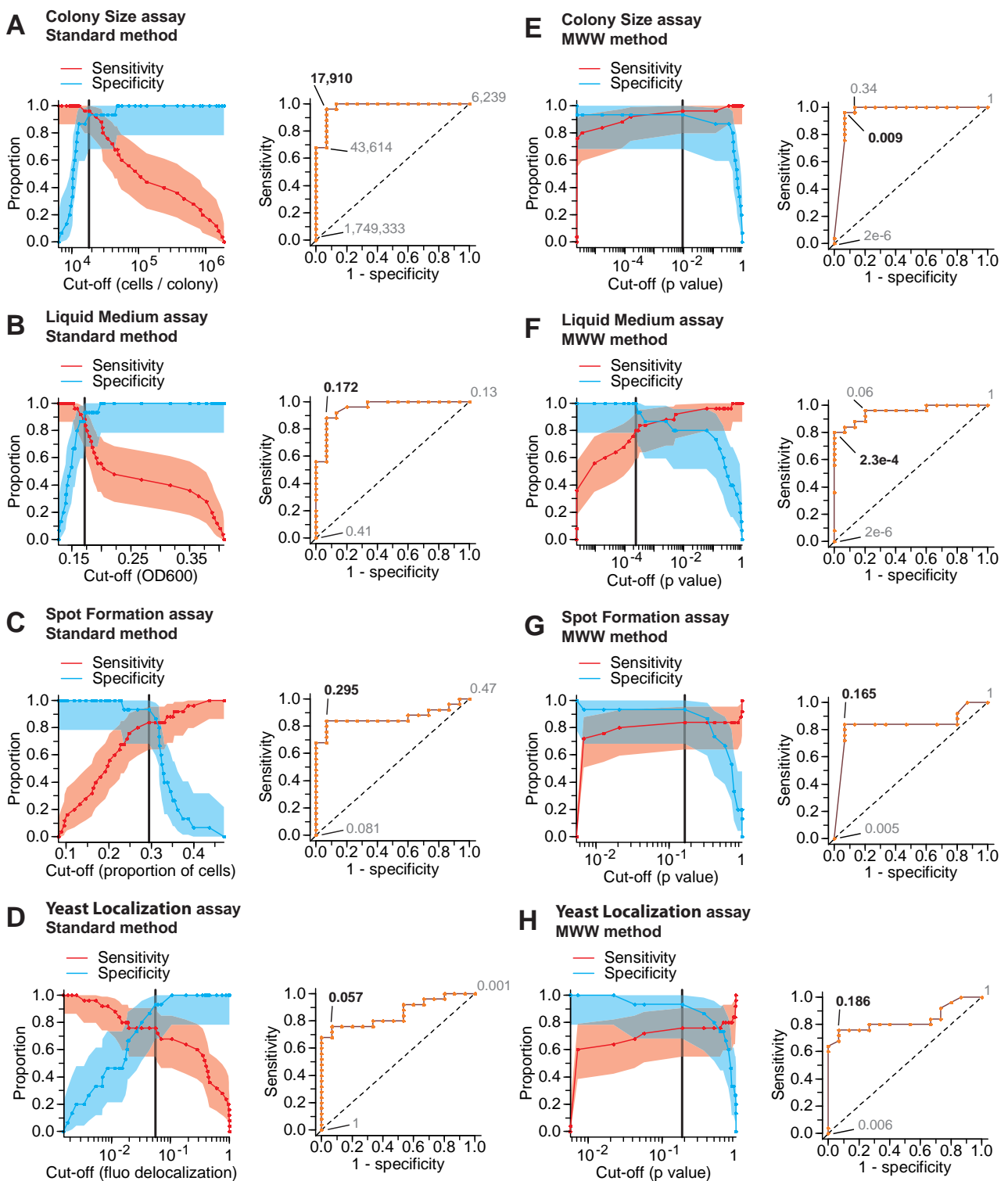
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S1 Fig. Position of the BRCA1 missense mutations selected

RING domain (amino acid 8–96); BRCT, BRCA1 C-terminal domains (amino acid 1646–1736 and 1760–1855). Pathogenic and neutral mutations are in red and blue, respectively. Fourteen mutations (3 neutral and 11 pathogenic) map within the RING domain. An additional neutral mutation, N132K, flanks the Cter part of this domain, resulting in 15 mutations located in the Nter extremity of BRCA1. Twenty-five mutations (11 neutral and 14 pathogenic) lie in the BRCT domain, at the Cter extremity of the protein. Of note, the RING domain suffers from a lack of neutral missense mutations classified by genetic/epidemiological methods, explaining why only 3 neutral mutations from our selected panel, lie in this domain. Moreover, no pathogenic missense mutations, between the amino acids 65 and 1684, are documented in the BRCA1 mutation databases (S2 Table). Therefore, this study was restricted to the RING and BRCT domains of BRCA1.

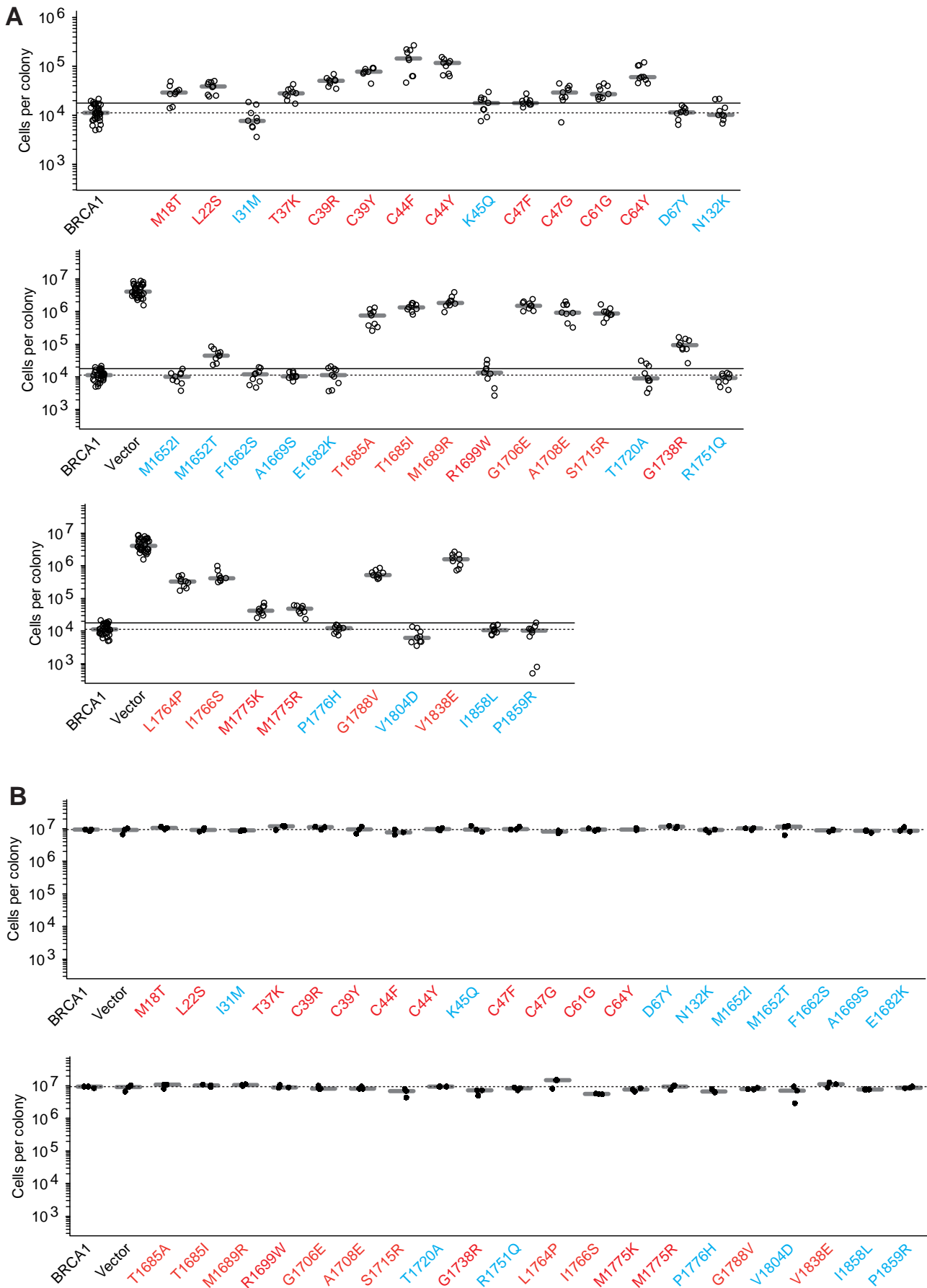


S2 Fig. Experimental best cut-off, experimental sensitivity and experimental specificity of functional assays

(A-D) Standard method. The medians of the mutant distributions were ordered (as in the waterfall distribution, Fig 1A) and each average position between two consecutive medians was defined as a cut-off. For example, in Fig 1A, the cut-off between the two first mutations, M1689R and V1838E, was $(1,877,333 + 1,621,333) / 2 = 1,749,333$ cells per colony. Next, sensitivity was defined as the proportion of pathogenic mutant medians above (for the Colony Size, Liquid Medium and Yeast Localization assays) or below (for the Spot Formation assay) a selected cut-off. The associated specificity was defined as the proportion of neutral mutant medians below (Colony Size, Liquid Medium and Yeast Localization assays) or above (Spot Formation assay) the same selected cut-off. For example, for the cut-off between M1689R and V1838E in Fig 1A, the sensitivity was $1/25 = 4\%$ and the specificity was $15/15 = 100\%$. Sensitivity and specificity were computed for each cut-off (left panels). Areas surrounding the curves delimit the 95% confidence interval according to the binomial law. The ROC curve (right panel) pinpoints the best cut-off (black number), meaning the cut-off that maximizes both sensitivity and specificity of the assay. Precisely, the best cut-off is the one associated with the highest vertical distance of the ROC curve to the dotted diagonal. This highest vertical distance is referred to as "Youden's index", which is equal to $\max[\text{sensitivity} + \text{specificity} - 1]$. In other words, the best cut-off is the cut-off of the Youden's index. Other cut-off values are also positioned on the ROC curve (grey numbers). Blue, red and orange dots on the curves of the left and right panels represent the different cut-offs tested. The black vertical bar, in the left panel, pinpoints the best cut-off defined on the ROC curve.

(E-H) MWW method. As in A-D for mutant p values, instead of mutant medians. In all assays, sensitivity was defined as the proportion of pathogenic mutant p values below a selected cut-off, and the associated specificity was defined as the proportion of neutral mutant p values above the same selected cut-off.

- (A, E) Colony Size assay.
- (B, F) Liquid Medium assay.
- (C, G) Spot Formation assay.
- (D, H) Yeast Localization assay.

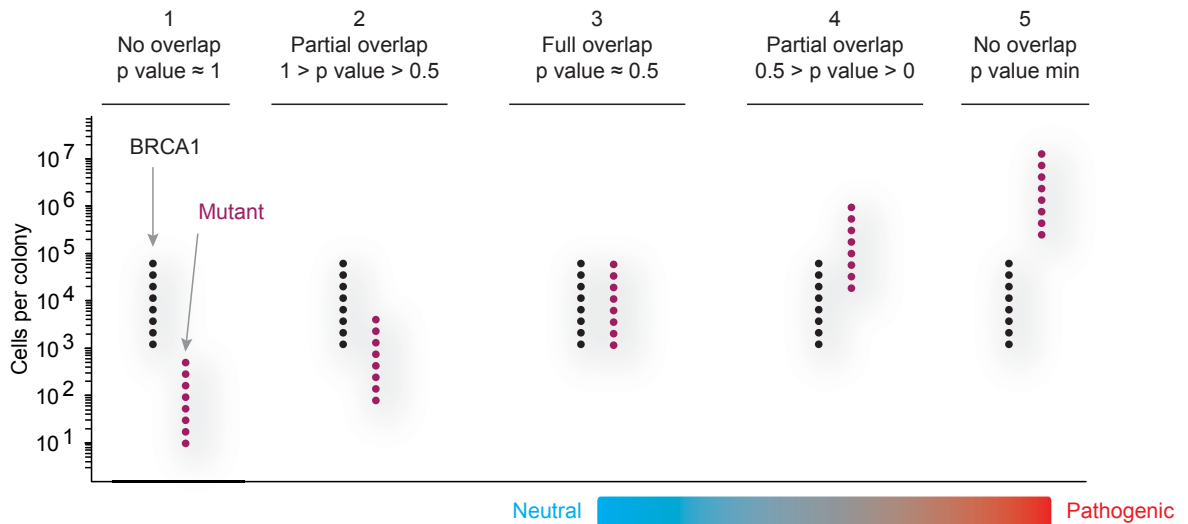


S3 Fig. Supplemental information in the Colony Size assay

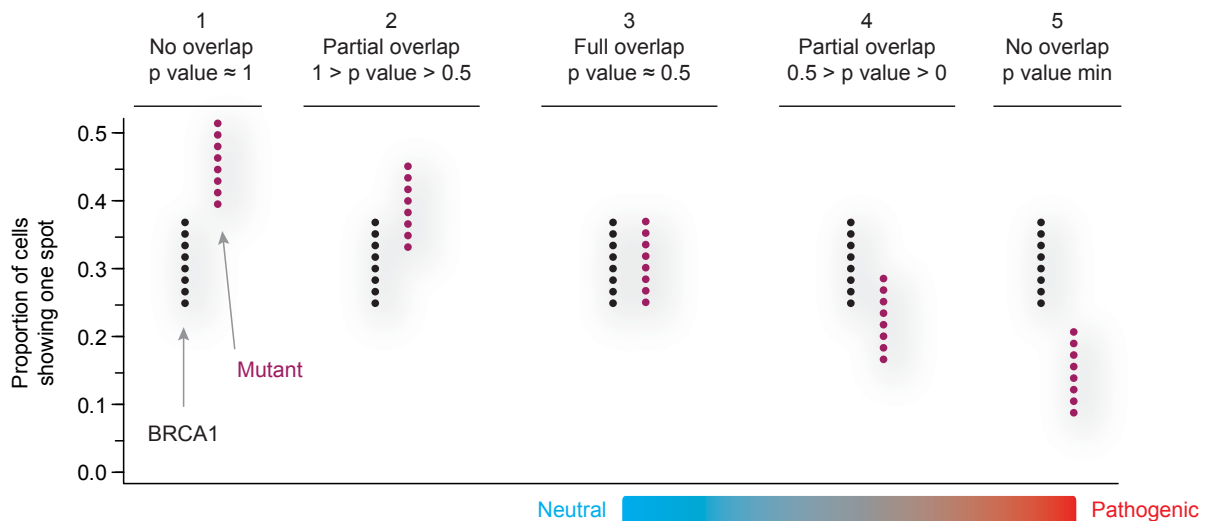
(A) Dotplot distribution of colony sizes. For each missense variant, the nine represented values result from three independent clones examined in three independent experiments. For the BRCA1 reference and the Vector control, the 36 values result from three independent clones examined in twelve independent experiments (represented in the three panels, except for the Vector values absent in the top panel). Grey bar, median; dotted horizontal line, median of BRCA1; black horizontal line, experimental best cut-off. The top panel (Nter extremity of BRCA1) has a y-axis scale magnified compared to the middle and bottom panels (Cter extremity of BRCA1).

(B) As in A with glucose instead of galactose media (see the S1 Text) to verify that each clone had no intrinsic growth defect, independent of WT or mutated BRCA1 expression. The three independent clones from A were examined in one experiment.

A Upper-sided MWW test



B Lower-sided MWW test

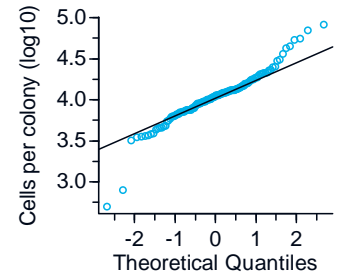
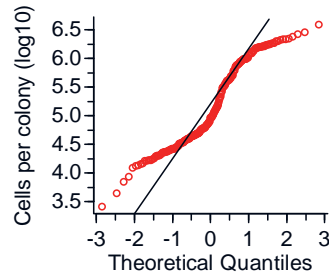
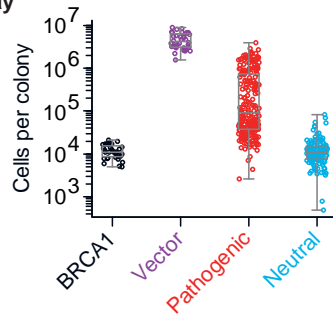


S4 Fig. The MWW method

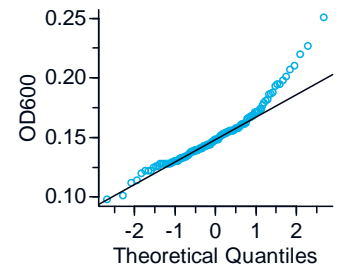
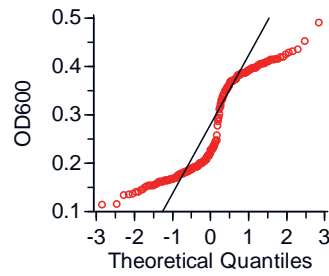
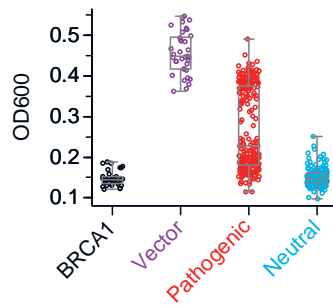
(A) Upper-sided MWW test. The theoretical examples are based on the Colony Size assay but are also valid for the Liquid Medium and Yeast Localization assays. Each distribution of the WT BRCA1 reference (black) and the missense mutation (purple) are composed of 8 theoretical values, represented by 8 dots in the diagram. The p value of the MWW test is used to score the overlap of the mutant and the WT BRCA1 distributions. See the S1 Text for full details. From left to right: (1) when all the mutant values are below the BRCA1 values, the upper-sided MWW test results in a p value close to 1; (2) the p value decreases when the mutant distribution begins to overlap the BRCA1 distribution; (3) the p value is approximately 0.5 when the two distributions completely overlap; (4) the p value continues to decrease when the mutant distribution is above the BRCA1 distribution, with a partial overlap; (5) finally, the p value is lowest when the mutant distribution is fully above the BRCA1 distribution. In theory, neutral and pathogenic mutations should have a p value close to 0.5 and 0, respectively, as depicted by the color scale below the diagram. However, the absolute p value attributed to each variant is not determinant. What is significant is the relative positions between the mutant distributions, indicated by the p values, using the WT BRCA1 distribution as a reference position. The lowest p values represent systematically the pathogenic mutations, and the highest the neutral mutations. Thus, the upper-sided MWW test is used when pathogenic mutations are above the neutral ones in the experimental data.

(B) Lower-sided MWW test. All of the theoretical examples shown are based on the Spot Formation assay. As in the upper-sided MWW test, the lowest and highest p values still represent the pathogenic and neutral mutations, respectively, but the pathogenic mutations are below the neutral ones in the experimental data.

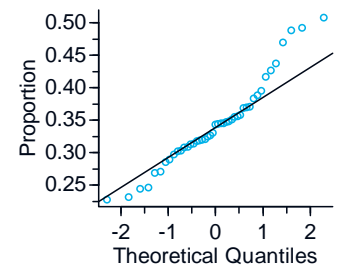
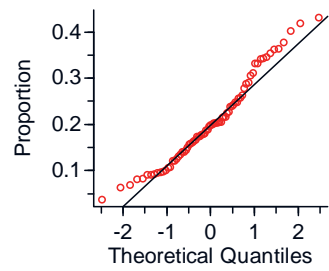
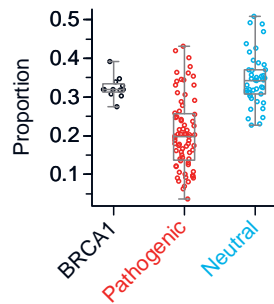
A Colony Size assay



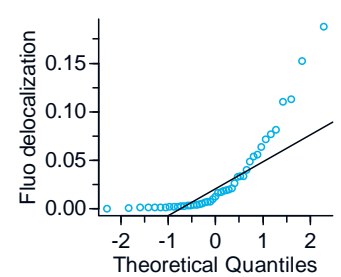
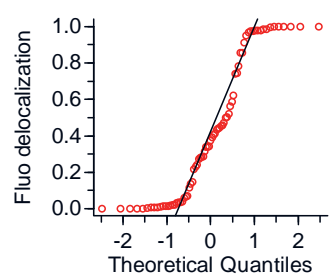
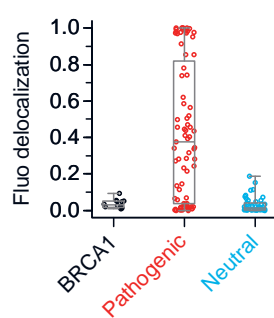
B Liquid Medium assay



C Spot Formation assay



D Yeast Localization assay



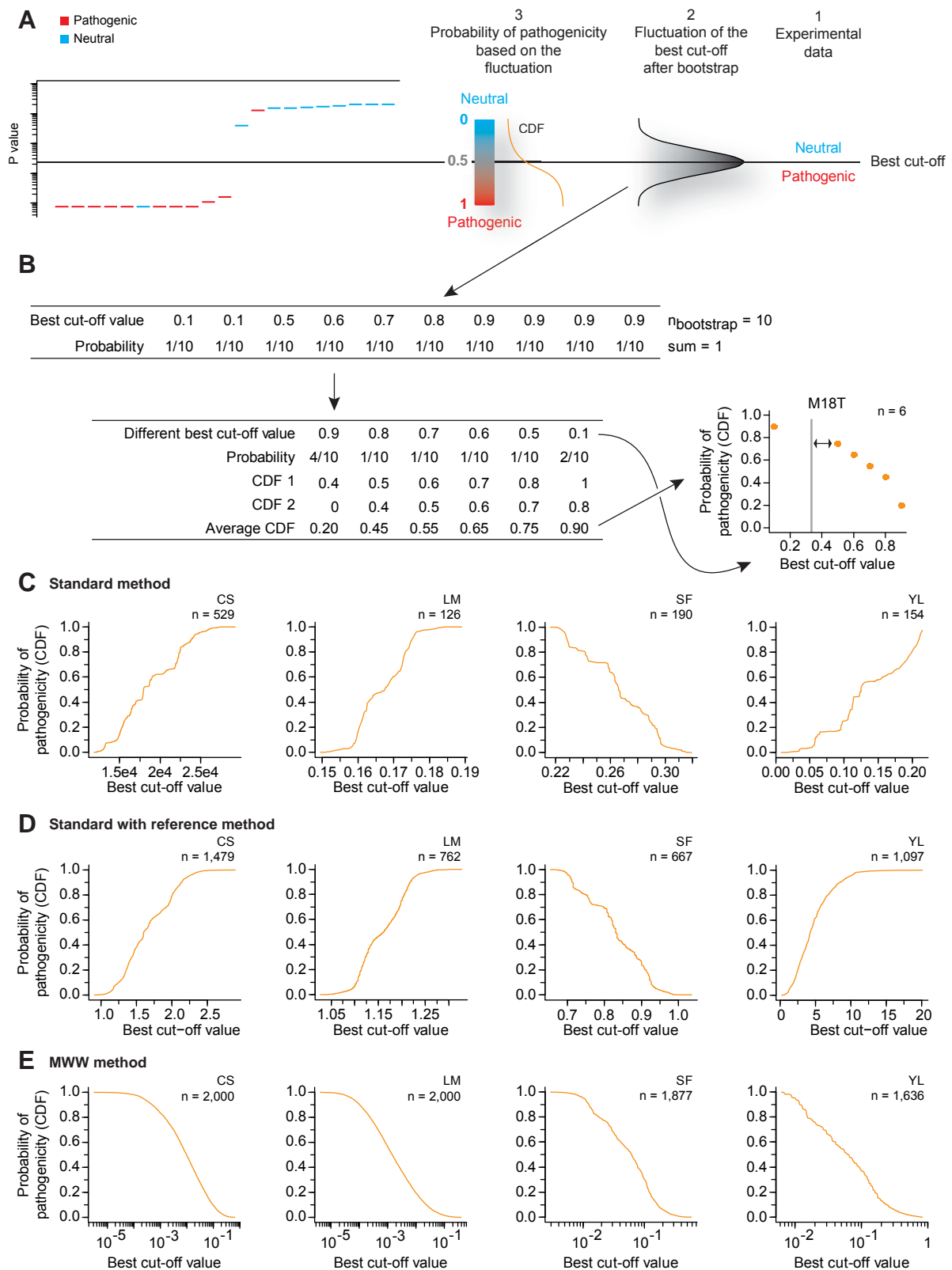
S5 Fig. Distribution of the pathogenic and neutral values

(A) Colony Size assay. The left panel exhibits dotplot distributions. Boxplots provide distribution parameters: box central bar, median; box, interquartile range (50% of the distribution); whiskers, extreme values. The middle panel shows the normal Quantile-Quantile (QQ) plot of the pathogenic values. Dots forming a straight line suggest that the values are normally distributed. Black line, straight line through the quantiles 25% and 75%. The right panel shows the normal QQ plot of the neutral values.

(B) Liquid Medium assay.

(C) Spot Formation assay.

(D) Yeast Localization assay.

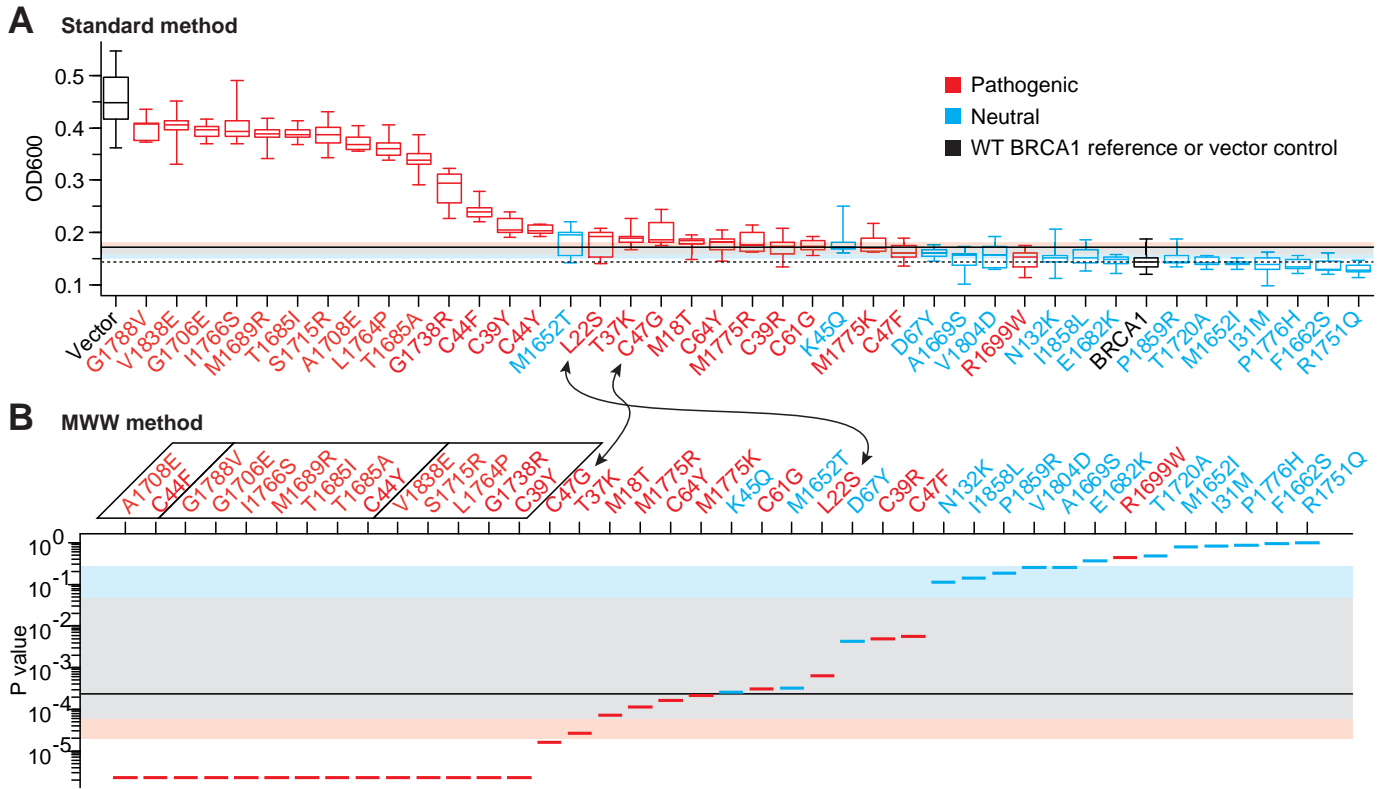


S6 Fig. Description of the probability system of classification

(A) As in Fig 2a.

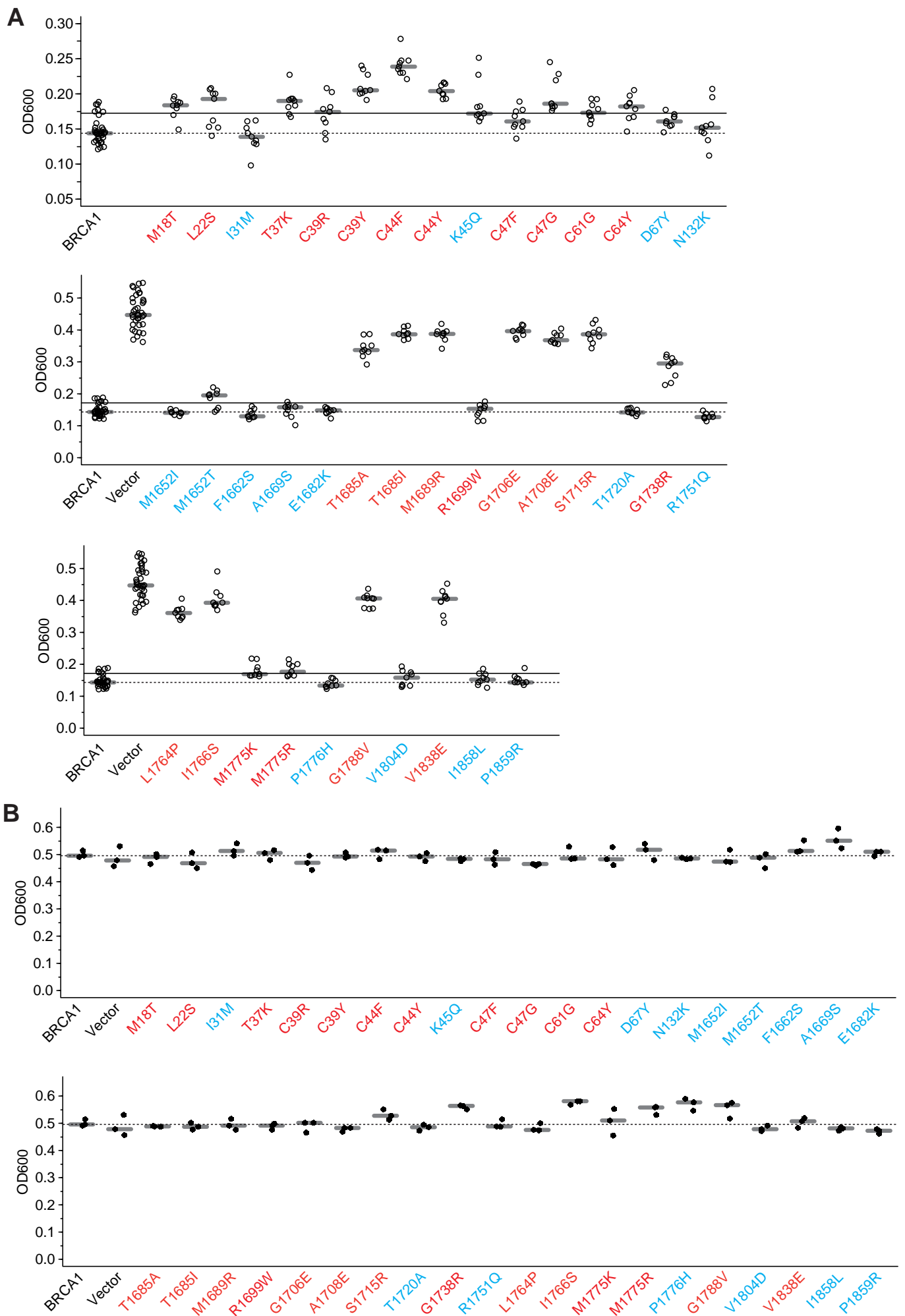
(B) Theoretical example showing how the values from the best cut-off fluctuation, derived from the MWW method, are converted into probabilities of pathogenicity. Top table: best cut-off distribution composed of 10 best cut-off values, resulting from 10 bootstraps ($n_{\text{bootstrap}} = 10$). The probability attributed to each best cut-off value was $1 / n_{\text{bootstrap}}$. Bottom table: cumulative distribution functions (CDF) generated from the best cut-off distribution. In this table, probabilities of each repeated cut-off value were summed. For instance, the best cut-off value of 0.9 is repeated 4 times in the top table, leading to a probability of 0.4. The CDF represents the sum of the probabilities present in the second row of the bottom table. Three CDF were computed. The first reaches the cumulated probability of 1. The second begins with the cumulated probability of 0. The third is the average of the two first CDF. This average CDF delivers the probability of pathogenicity used to classify variants. Right panel: plot of the average CDF. To classify a variant (e.g., M18T), the variant p value, derived from the MWW method, is positioned on the x-axis (vertical grey bar). Next, the closest average CDF value is attributed to the variant as a probability of pathogenicity. In this example, the best cut-off value, closest to M18T, is 0.5. Thus, the corresponding probability 0.75 is attributed to M18T.

(C-E) Average CDF of the Colony Size (CS), Liquid Medium (LM), Spot Formation (SF) and Yeast Localization (YL) assays, obtained with the standard (C), standard with reference (D) or MWW method (E). The same procedure, described in B, was applied to the 2,000 best cut off values obtained for each assay and each method used. The CDF is ascending when the pathogenic mutations are above the neutral ones, and descending when the pathogenic mutations are below. The number of different best cut-off values is indicated ($n = 2,000$ when no identical best cut-off values within distributions).



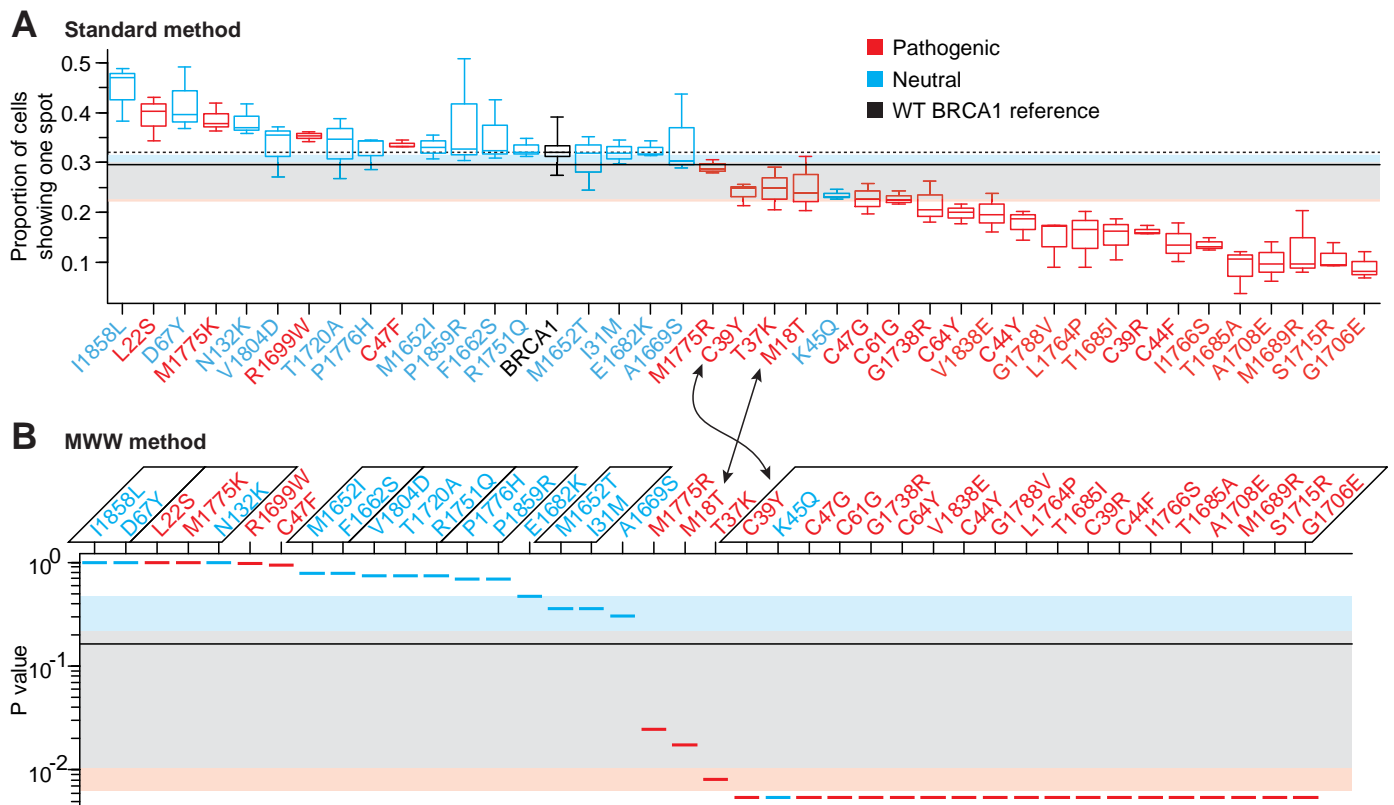
S7 Fig. Relative position of the variants in the Liquid Medium assay and fluctuation of the best cut-off

(A-B) As in Fig 1. One OD unit corresponds to 10^8 cells / ml. Arrows pinpoint the ranking of the L22S and C47G mutations, which is improved using the MWW method, as explained in the main text introducing this method. The incoherent ranking observed with the standard method results from L22S that exhibits four values below the experimental best cut-off while C47G has none (S8A Fig).



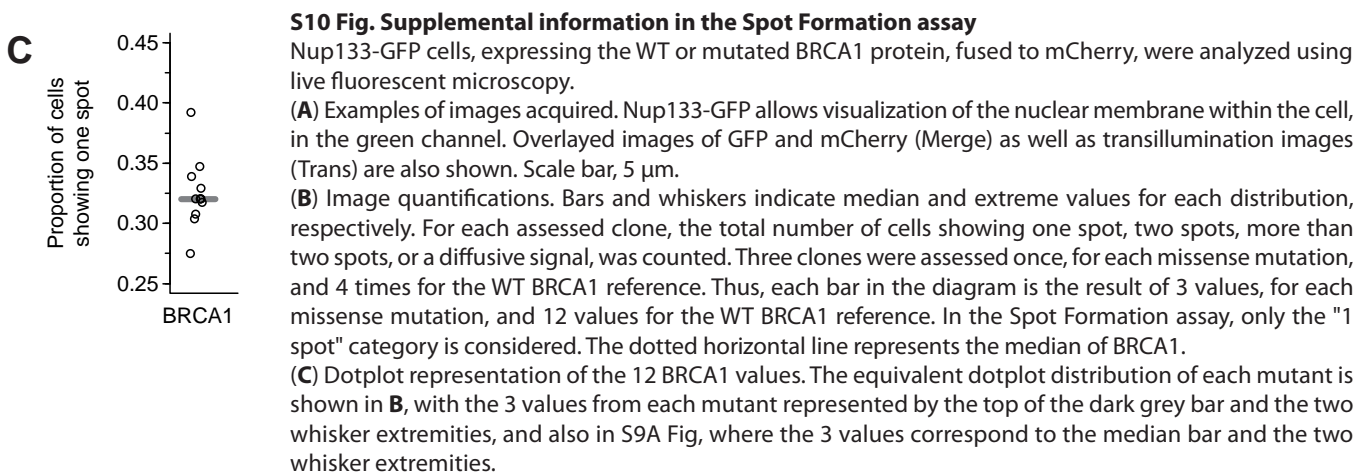
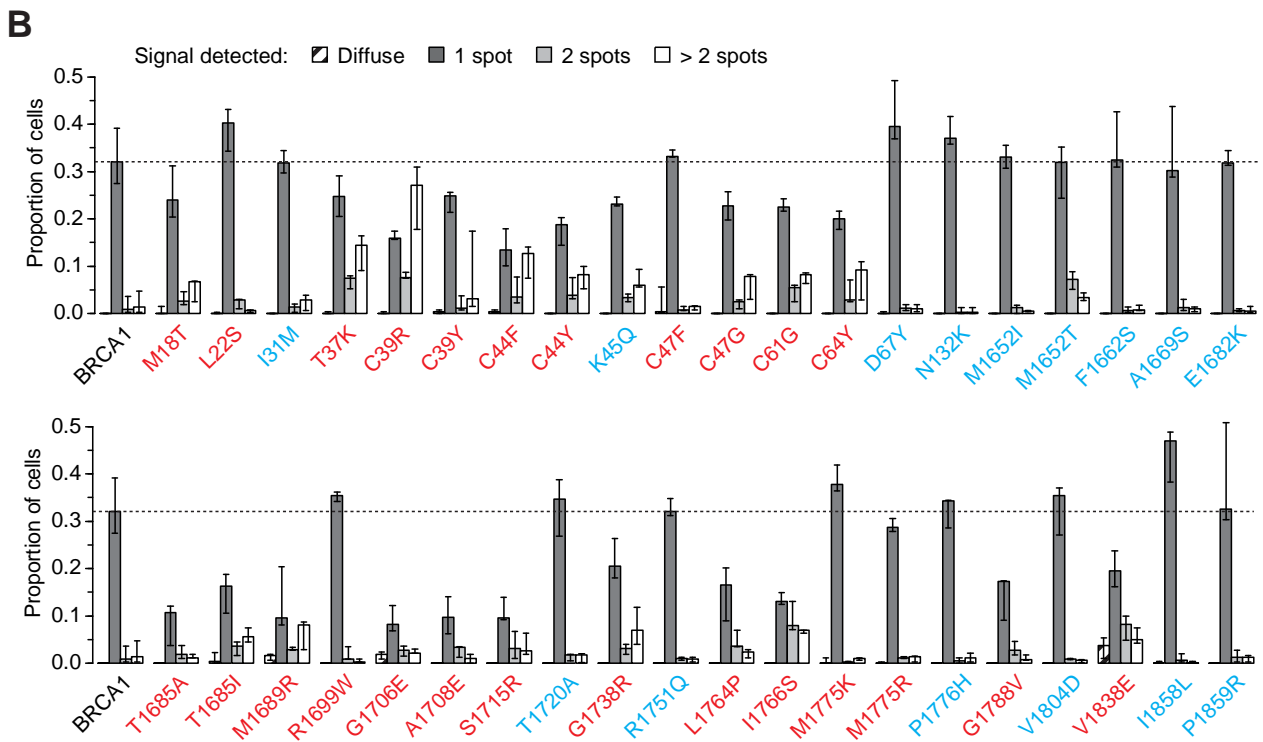
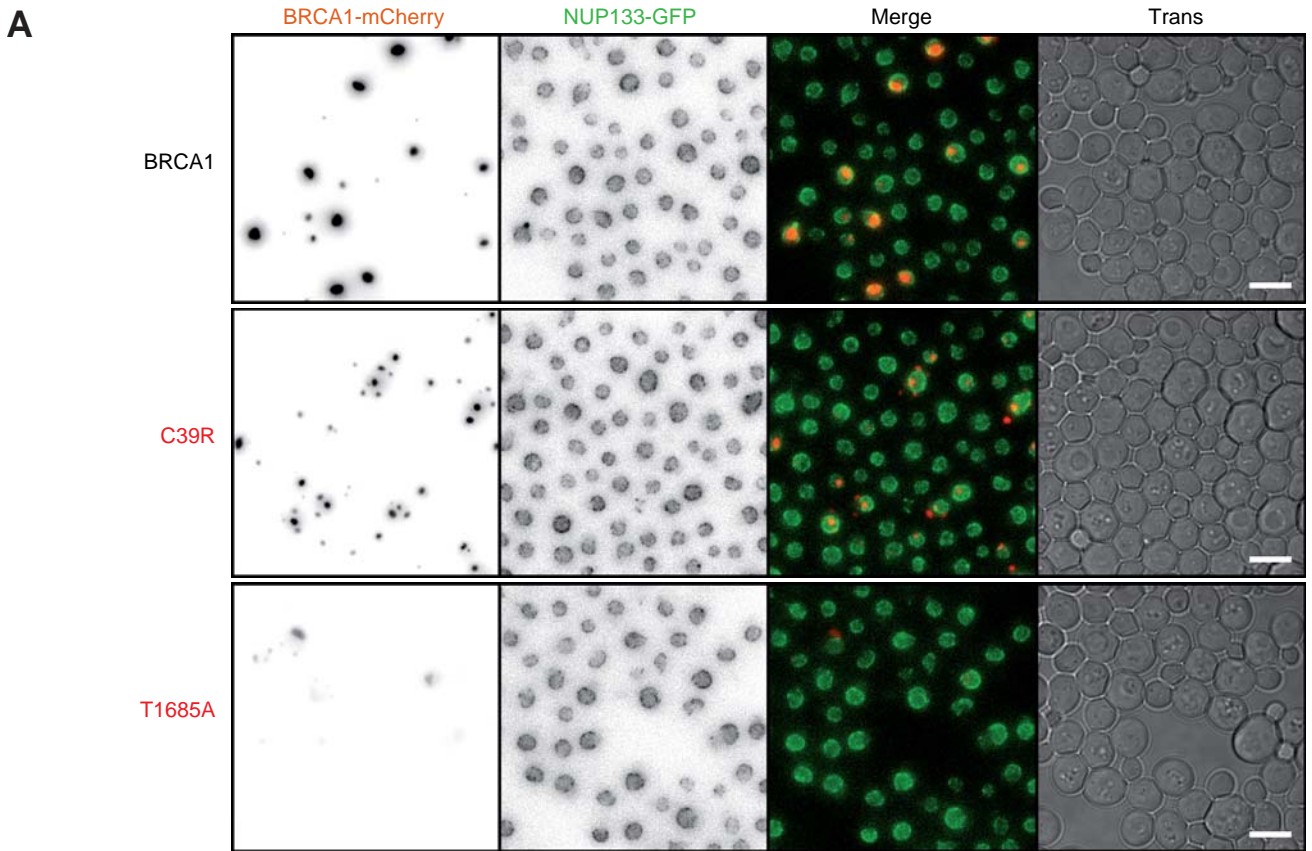
S8 Fig. Supplemental information in the Liquid Medium assay

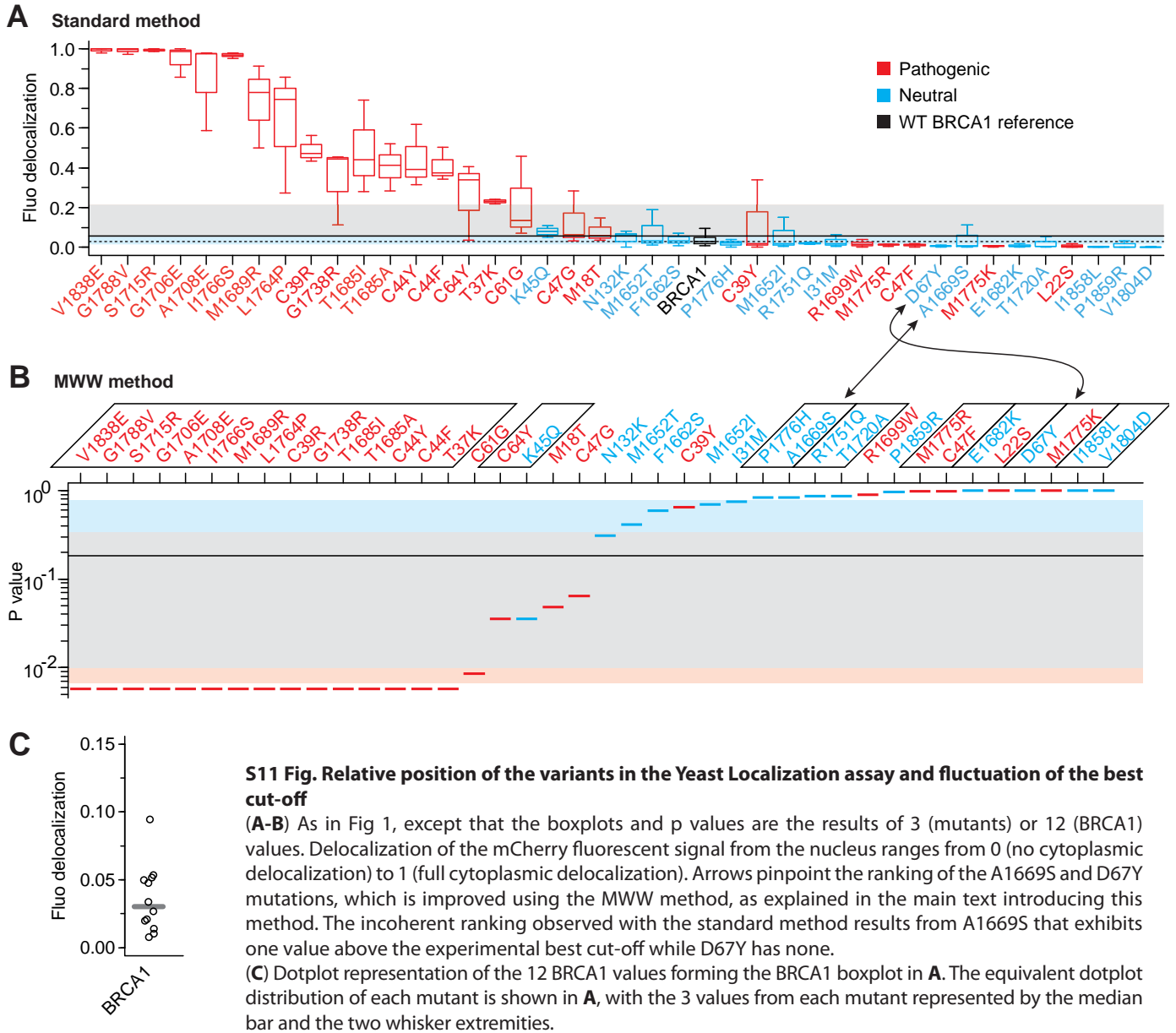
(A-B) Same as for the Colony Size assay (S3 Fig). One OD unit corresponds to 10^8 cells / ml.

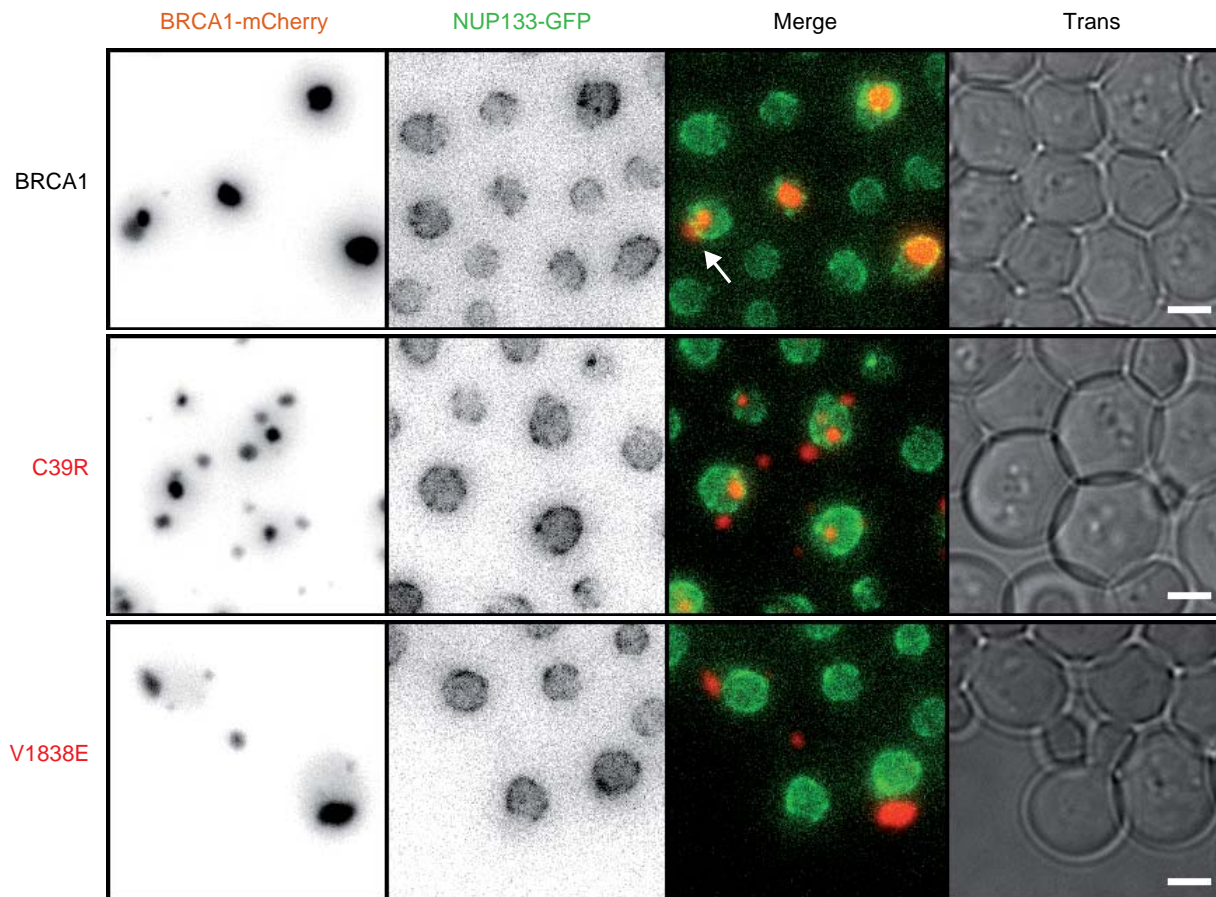


S9 Fig. Relative position of the variants in the Spot Formation assay and fluctuation of the best cut-off

(A-B) As in Fig 1, except that boxplots and p values resulted from 3 (mutants) or 12 (BRCA1) values. Arrows pinpoint the ranking of the M18T and C39Y mutations, which is improved using the MWW method, as explained in the main text introducing this method. The incoherent ranking observed with the standard method results from M18T that exhibits one value above the experimental best cut-off (shown by the top whisker overlaying the thick horizontal line) while C39Y has none.



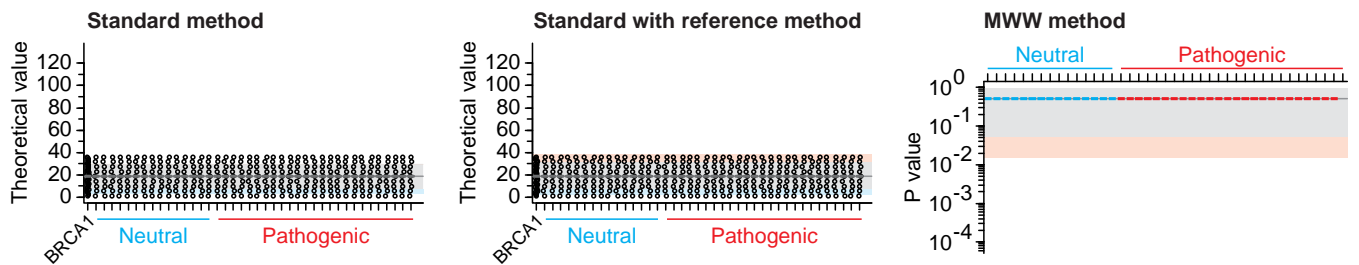




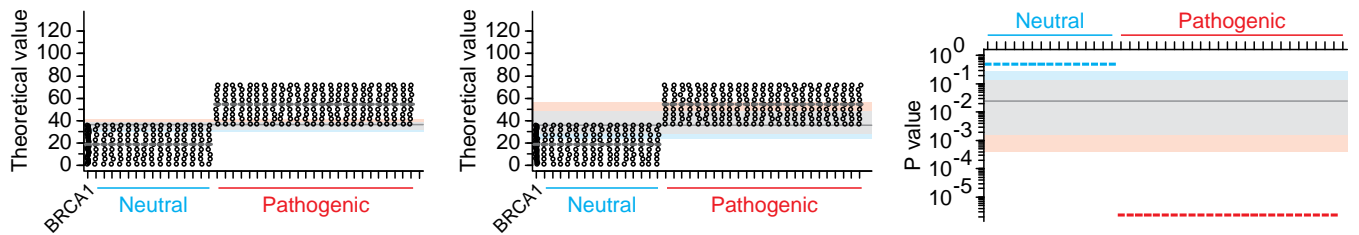
S12 Fig. Supplemental information in the Yeast Localization assay

Fluorescent images acquired in the Yeast Localization assay, as in S10A Fig. The arrow points to rare cytoplasmic spot in cells expressing the WT BRCA1-mCherry protein. Scale bar, 2 μ m.

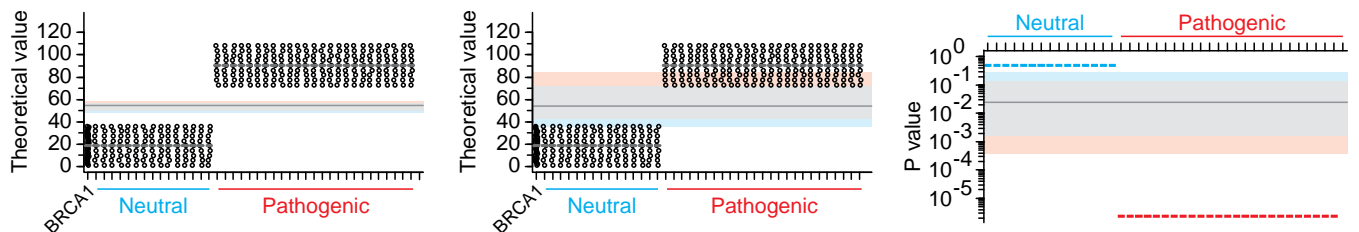
A Shift = 0



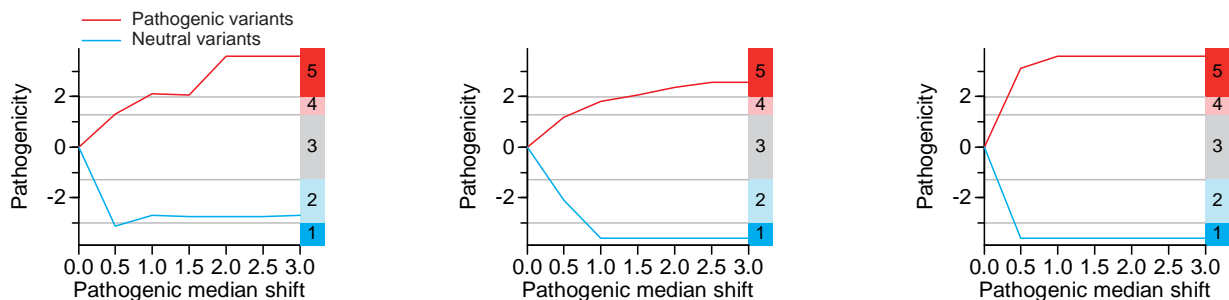
B Shift = 1



C Shift = 2



D



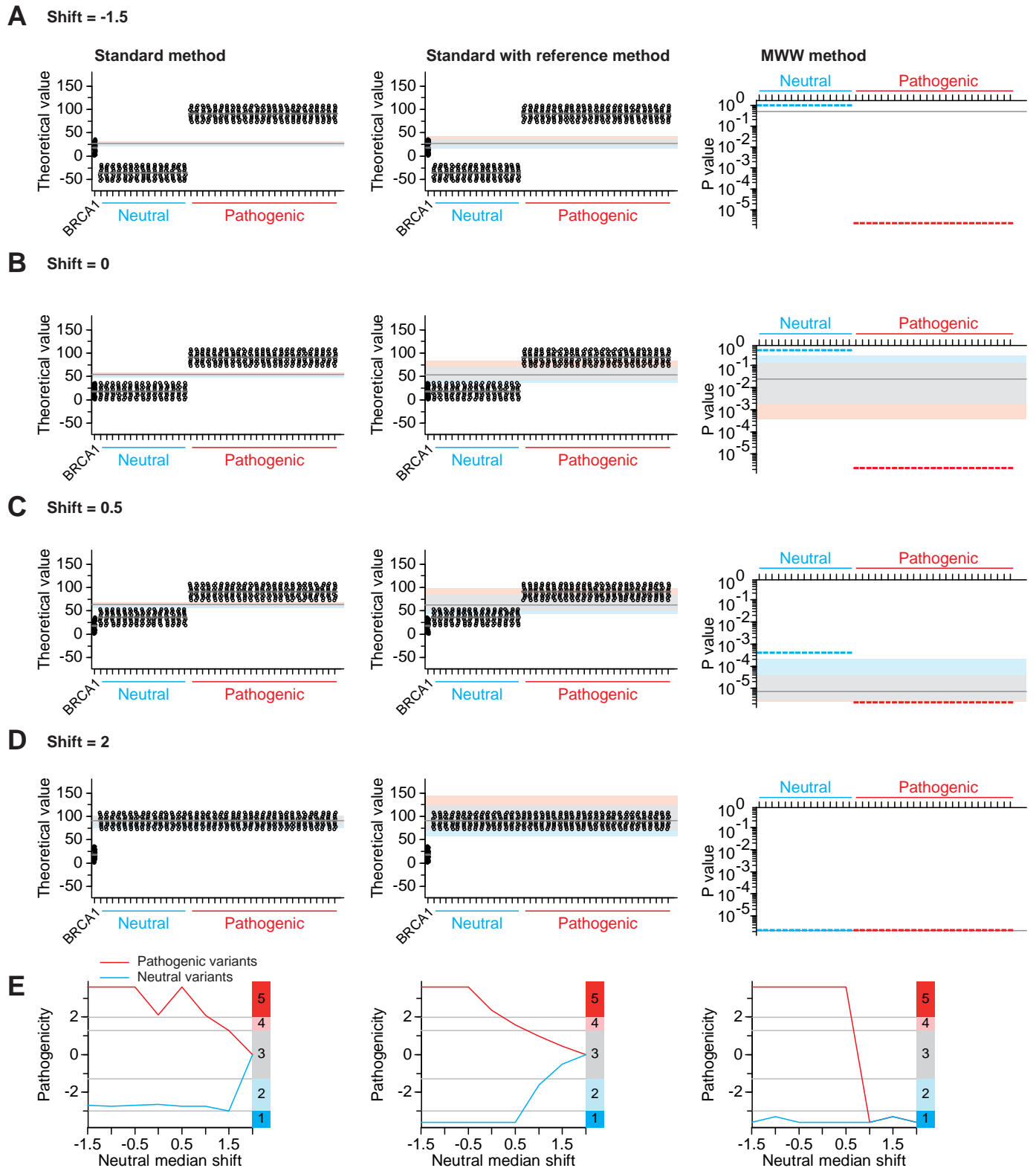
S13 Fig. Effect of the position of the pathogenic mutations on the probability system of classification (theoretical situation)

The parameters of the theoretical distributions used are detailed in Table S8. The reference situation is as follows: $n_{\text{mutant}} = 9$, $n_{\text{BRCA1}} = 36$, $n_{\text{neutral}} = 15$ and $n_{\text{pathogenic}} = 25$. In addition, medians and ranges of the neutral and WT BRCA1 distributions were made systematically equal. Distributions of the neutral and pathogenic mutations were identical, except for the shift of the pathogenic values from the neutral mutations, according to the formula $v_{ij} + 36 \times s$, with s representing the shift intensity and v_{ij} representing the value i of the pathogenic mutation j . When $s = 0$, pathogenic and neutral distributions are identical. Fluctuations from the best cut-off were obtained exactly as performed for the Colony Size, Liquid Medium, Spot Formation and Yeast Localization assays.

(A-C) Examples of shift intensities and best cut-off fluctuation results. The graphs depicted are similar to those in Fig 1, except that the standard, standard with reference and MWW methods are shown respectively on the left, middle and right of the figure. In the standard and standard with reference methods, boxplots are replaced by dotplots with the median of the distributions indicated by a grey segment. The s values are indicated (top left). In the subsequent supplemental figures, the position of the pathogenic mutation medians are as in C ($s = 2$). The grey horizontal line indicates the median of the best cut-off fluctuation.

(D) Probabilities of pathogenicity obtained for the neutral (blue line) and pathogenic variants (red line), depending on the shift intensity of the pathogenic mutations. Y-axis, $\log_{10}(p / (1 - p))$ with p being the probability of pathogenicity of the variants (0 corresponds to $p = 0.5$); right colored classes, five-class nomenclature with the horizontal grey lines showing the 0.99, 0.95, 0.05 and 0.001 limits of the classes (see S1 Table). In the standard method, the slight erratic curves and the lack of specificity sometimes observed (blue line in the class 2 instead of class 1) is due to the fact that this method generates a low number of different best cut-off values (between 8 and 64) in the best cut-off distributions, as explained in the S1 Text.

As summarized in S9 Table, these results confirm that the probability system of classification is an efficient variant classifier. Indeed, whatever method is used, when the pathogenic and neutral distributions are strictly identical, they all locate inside the class 3 area (i.e., the system cannot classify any variants in such kind of functional assay). Moreover, the probability system of classification is improved when the pathogenic mutations shift from the neutral sector towards the pathogenic sector, since the probability of pathogenicity increases for the pathogenic variants and decreases for the neutral ones.



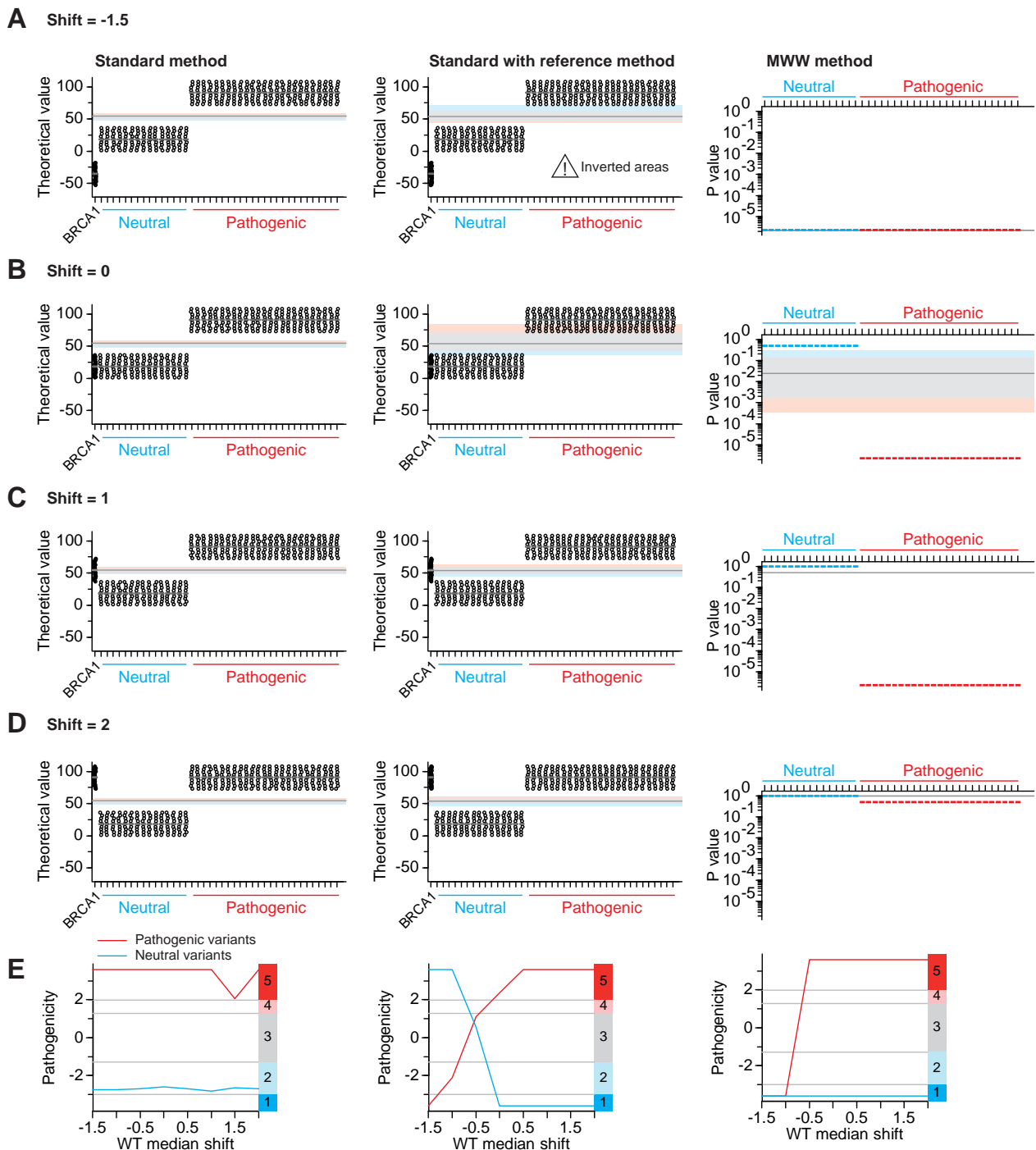
S14 Fig. Effect of the position of the neutral mutations on the probability system of classification (theoretical situation)

See S13 Fig for details. Neutral mutations were shifted according to the formula $v_{ij} + 36 \times s$, with s representing the shift intensity and v_{ij} representing the value i of the neutral mutation j (when $s = 0$, medians and extreme values of the BRCA1 and neutral distributions are identical. When $s = 2$, pathogenic and neutral distributions are identical).

(A-D) Examples of shift intensities and best cut-off fluctuation results. The s values are indicated (top left).

(E) Probabilities of pathogenicity obtained for the neutral (blue line) and pathogenic variants (red line), depending on the shift intensity of the neutral mutations.

As summarized in S9 Table, these results highlight divergences between the different methods. With the standard method and the standard with reference methods (E, left and middle panels), sensitivity and specificity of the probability system of classification decrease when the neutral mutations approach the pathogenic mutations. With the MWW method (E, right panel), the probability system of classification results in a complete misclassification of the pathogenic mutations when the neutral distributions do not overlap the WT reference distribution ($s \geq 1$). Of note, these analyses treat extreme situations. In practice, the WT reference should be well embedded within the neutral distributions. The opposite situation would raise question about the WT reference or neutral mutations used.



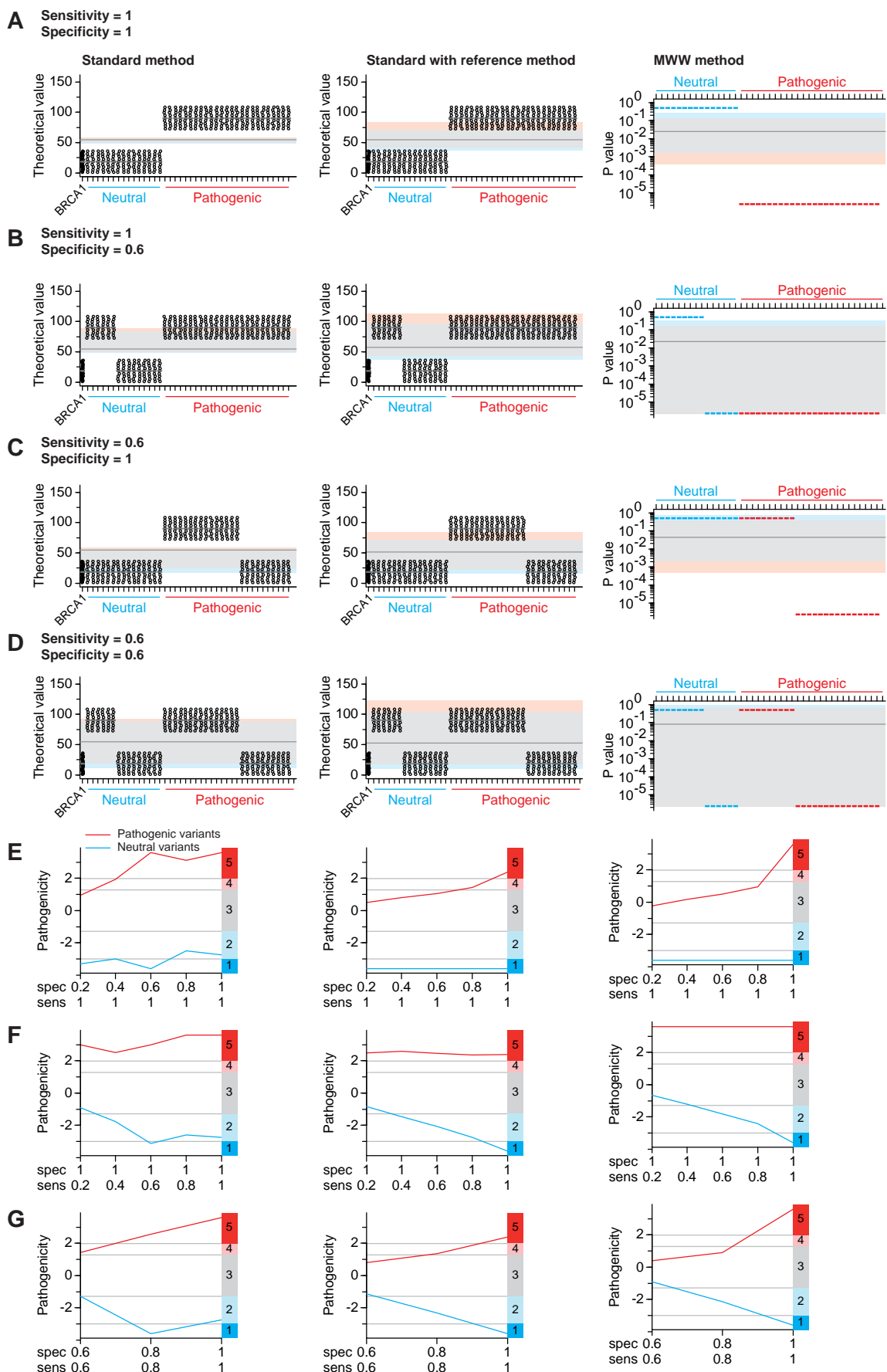
S15 Fig. Effect of the position of the WT BRCA1 reference on the probability system of classification (theoretical situation)

See S13 Fig for details. Values of the WT BRCA1 distribution were shifted according to the formula $v_i + 36 \times s$, with s representing the shift intensity and v_i representing the value i of the BRCA1 reference (when $s = 0$, medians and extreme values of the neutral and BRCA1 distributions are identical. When $s = 2$, medians and extreme values of the pathogenic and BRCA1 distributions are identical). Of note, these theoretical analyses treat extreme situations. In practice, the WT reference should be well embedded in the neutral distributions. The opposite situation would raise question about the WT reference or neutral mutations used.

(A-D) Examples of shift intensities and best cut-off fluctuation results. The s values are indicated (top left).

(E) Probabilities of pathogenicity obtained for the neutral (blue line) and pathogenic variants (red line), depending on the shift intensity of the WT reference.

As summarized in S9 Table, these results highlight divergences between the different methods. As expected, the standard method is not affected by the position of the WT BRCA1 distribution (E, left panel). In contrast, the standard with reference method is strongly influenced by the position of this reference (E, middle panel). When the WT BRCA1 median shifts towards the null value, sensitivity and specificity of the probability system of classification are decreased, with a complete loss of sensitivity and specificity (i.e., systematic classification as class 3) when the WT BRCA1 median is null ($s \approx -0.514$). This was expected since the standard with reference method is based on best cut-off values divided by the WT BRCA1 median. Thus, a division by zero generates relative best cut-offs with an infinite value. Such issues are compensated only when best cut-offs are close to the WT BRCA1 median. This was shown in the Liquid Medium and Yeast Localization assays. Using the standard or standard with reference method provided similar variant classification (Fig 2B), even if the WT BRCA1 medians of these assays approached zero, with 0.144 and 0.03 respectively (S4 Table). In conclusion, a situation, in which the WT reference median is close to zero, with the fluctuation of the raw best cut-off far from this median, will guarantee a weak sensitivity and specificity of the probability system of classification. Concerning the standard with reference method, it is also noteworthy that a negative value of the WT reference median ($s < -0.514$) inverts the classification (E, middle panel), as expected, regardless of the values from the neutral and pathogenic mutations. When comparing the standard with reference method versus the MWW method, the later has the advantage of being independent of the WT reference values, as only overlapping distributions matter. Specificity of the probability system of classification is not affected by the position of the WT reference, contrary to sensitivity (E, right panel). The main weakness of the MWW method occurs when the WT reference distribution falls outside of the range of the neutral and pathogenic distributions (as in A, left panel), which generates misclassification of the pathogenic mutations as neutral.

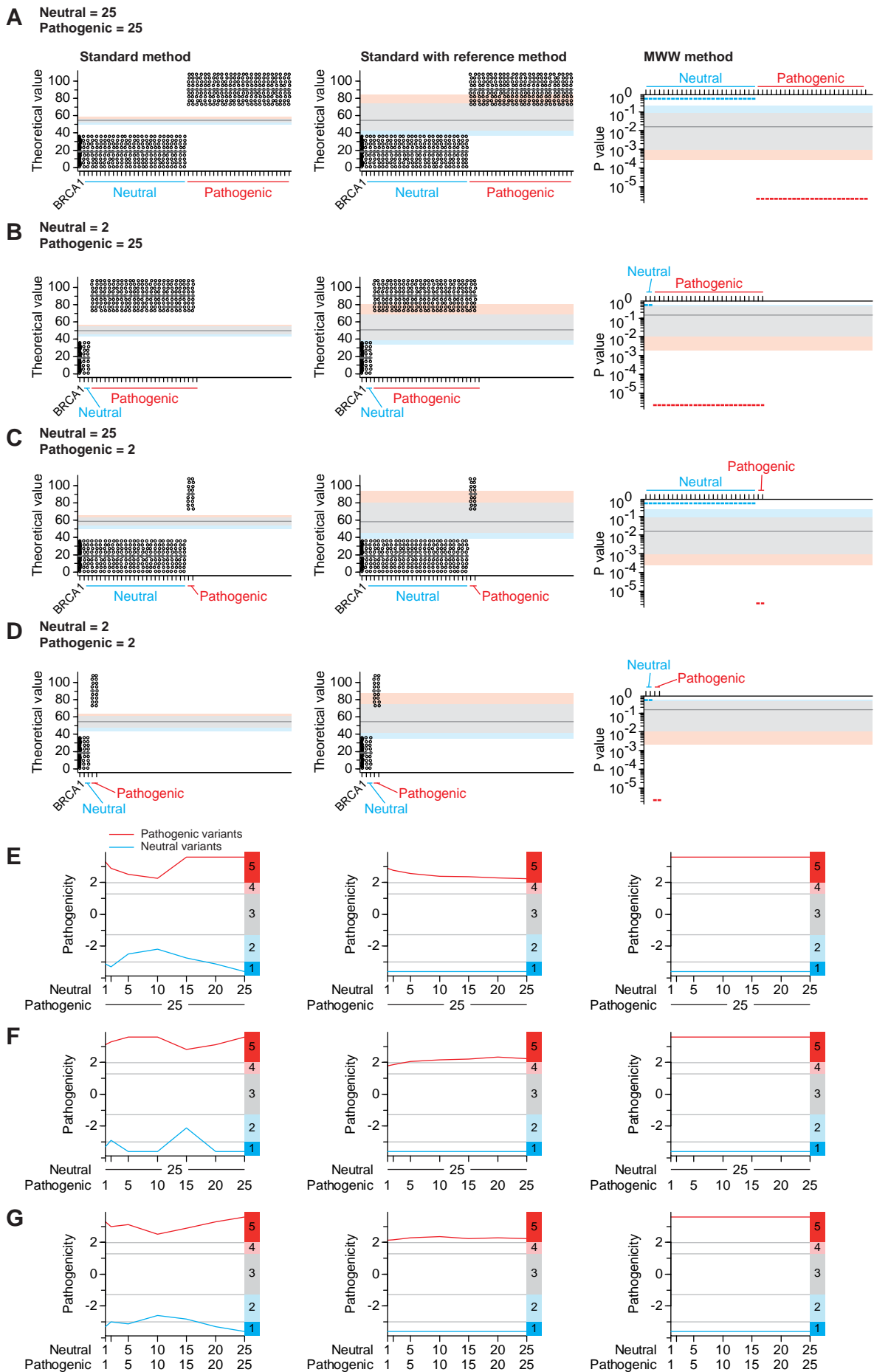


S16 Fig. Effect of the experimental sensitivity and specificity on the probability system of classification (theoretical situation)

See S13 Fig for details. The experimental sensitivity and specificity were modulated by assigning certain pathogenic mutants in the neutral region and certain neutral mutants in the pathogenic region, respectively. The experimental sensitivity and specificity values indicated were those obtained with the experimental best cut-off, as explained in S2 Fig. These values are referred to as "initial" sensitivity and specificity, as opposed to the sensitivity and specificity of the probability system of classification, obtained after bootstrap analysis. (A-D) Examples of experimental sensitivities/specificities and best cut-off fluctuation results.

(E-G) Probabilities of pathogenicity obtained for the neutral (blue line) and pathogenic variants (red line), depending on decreases from experimental specificity (E), experimental sensitivity (F) or both (G).

As summarized in S9 Table, these results confirm that the probability system of classification is an efficient variant classifier. A decrease of the experimental specificity indicates that the pathogenic area is contaminated by neutral variants, which reduces the probability of pathogenicity of the pathogenic variants (class 5 towards class 3). In the same manner, a decrease of the experimental sensitivity indicates that the neutral area is contaminated by pathogenic variants, which enhances the probability of pathogenicity of the neutral variants (class 1 towards class 3). This was observed using the three standard, standard with reference and MWW methods. Of note, the situations studied used systematically: experimental sensitivity + experimental specificity ≥ 1 (otherwise representing an inappropriate use of the experimental information, i.e., pathogenic and neutral sectors incorrectly positioned).



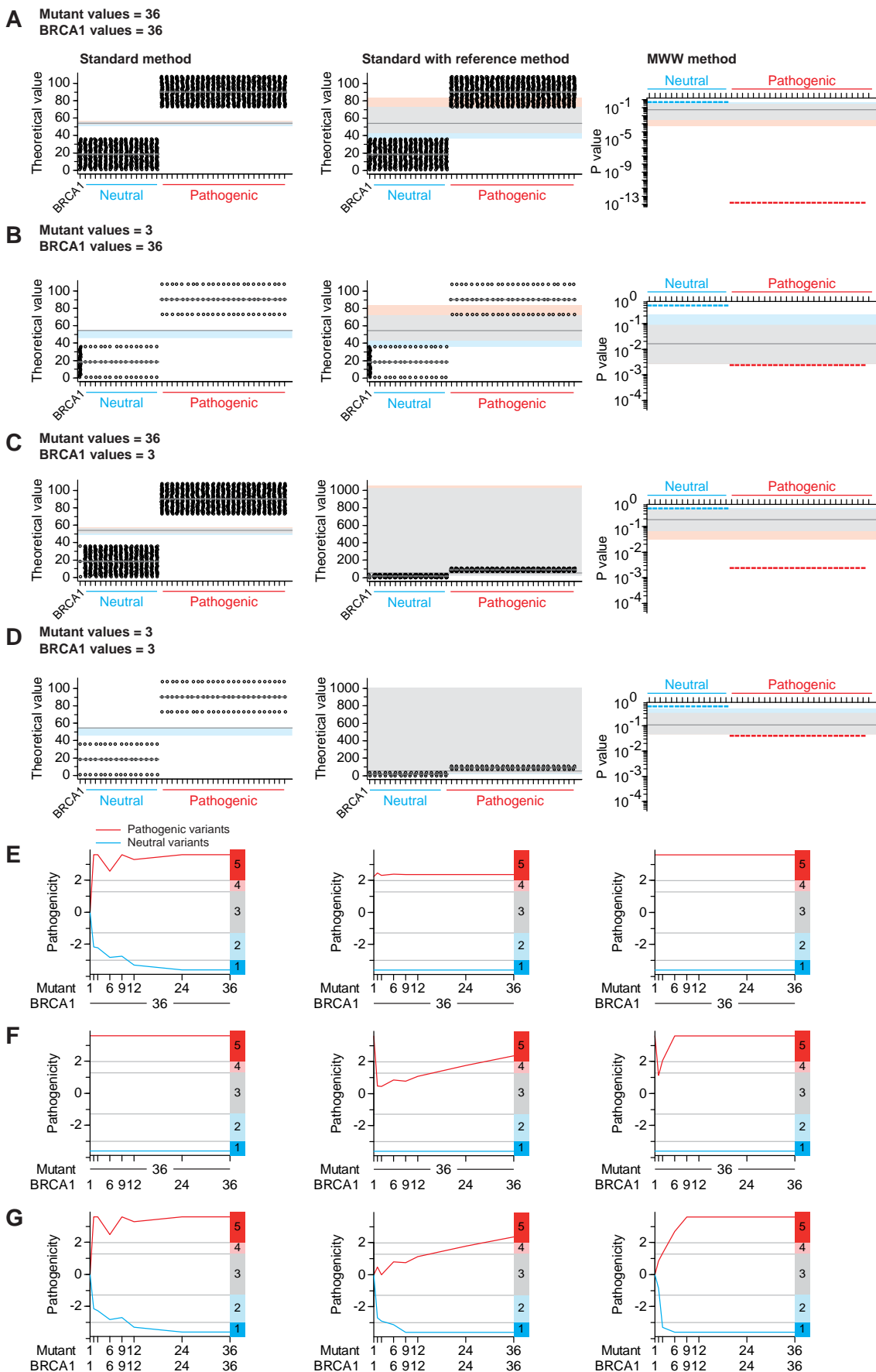
S17 Fig. Effect of the number of neutral and pathogenic mutations on the probability system of classification (theoretical situation)

See S13 Fig for details.

(A-D) Examples showing the number of neutral and pathogenic mutations tested, with best cut-off fluctuation results.

(E-G) Probabilities of pathogenicity obtained for the neutral (blue line) and pathogenic variants (red line), following a decrease in the number of neutral mutations (E), pathogenic mutations (F) or both (G).

As summarized in S9 Table, these results show that the probability system is poorly sensitive to the number of neutral and pathogenic mutations incorporated, whatever method is used.



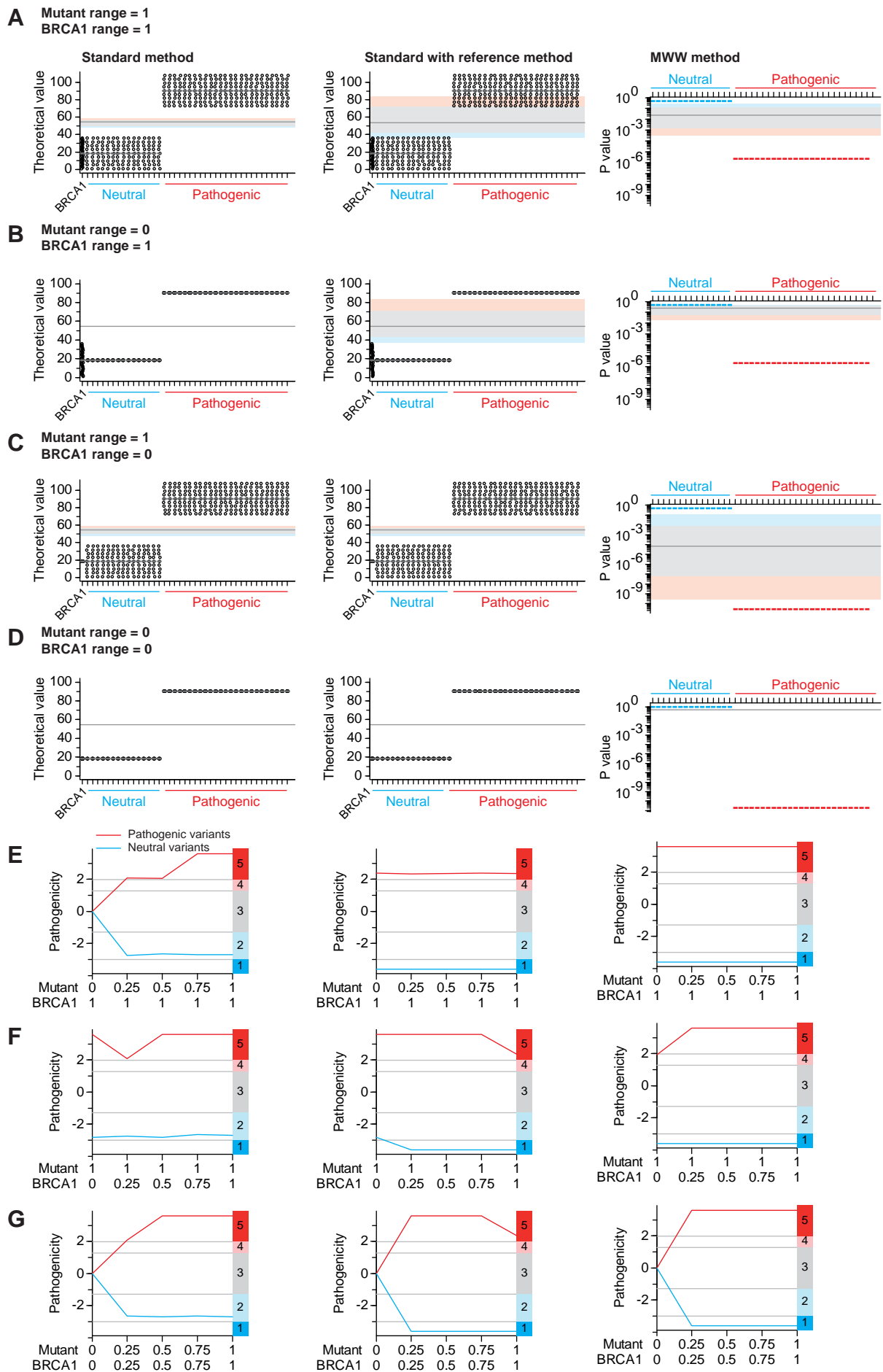
S18 Fig. Effect of the number of mutant and BRCA1 values on the probability system of classification (theoretical situation)

See S13 Fig for details. Number of values was modulated so that the range and median of the distributions remained the same, as shown in S8 Table.

(A-D) Examples showing the number of mutant or BRCA1 values tested, with best cut-off fluctuation results.

(E-G) Probabilities of pathogenicity obtained for the neutral (blue line) and pathogenic variants (red line), following a decrease in the number of mutant values (E), BRCA1 values (F), or both (G).

As summarized in S9 Table, these results confirm that the probability system of classification is an efficient variant classifier. Whatever method is used, a decreasing number of values in the dataset affects the probabilities of both the pathogenic and neutral variants (G), which tend toward 0.5 (class 3). Thus, the probability system prevents decision-making when data is lacking. As expected, the standard method is not affected by the number of BRCA1 values (F, left panel). The standard with reference and the MWW methods are insensitive to the number of mutant values if the number of BRCA1 values is high (E, middle and right panels). However, a decrease in the number of BRCA1 values lowers the probability of pathogenicity of the pathogenic variants (F, middle and right panels), but with a strong recovery when the fluctuation of the best cut-off is no longer influenced by the fluctuation of the WT reference ($n_{BRCA1} = 1$). Of note, the best cut-off does not fluctuate when $n_{mutant} = 1$ and $n_{BRCA1} = 1$ (G), which results in a probability of pathogenicity equal to 0.5 for both the pathogenic and neutral variants. Moreover, using the standard method, when $n_{mutant} = 9$, the classification of the neutral mutations is class 2 (E, left panel), which explains the lack of specificity frequently observed in S13-S19 Figs.



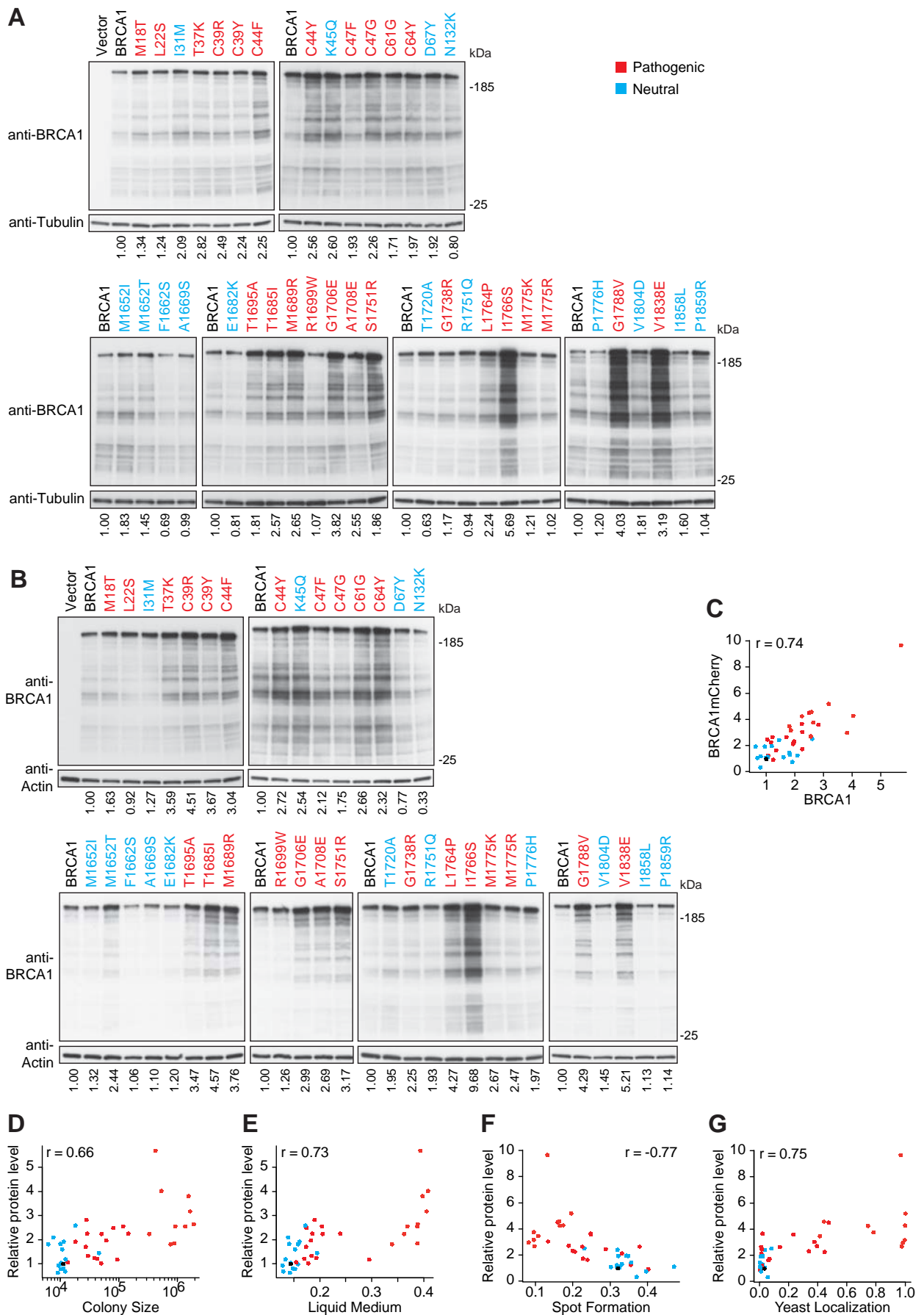
S19 Fig. Effect of the range of mutant and BRCA1 distributions on the probability system of classification (theoretical situation)

See S13 Fig for details. Distribution ranges were modulated so that medians remained the same, as shown in S8 Table. The range factor r , indicated on the graphs, illustrates the relative dispersion of the distributions. When $r = 0$, the dispersion is null.

(A-D) Examples showing the ranges of the mutant and BRCA1 distributions tested, with best cut-off fluctuation results.

(E-G) Probabilities of pathogenicity obtained for the neutral (blue line) and pathogenic variants (red line), following a range decrease of the mutant distributions (E), BRCA1 distribution (F), or both (G).

As summarized in S9 Table, these results indicate that the probability system of classification is affected mainly when the range of the BRCA1 and mutant distributions is null, whatever method is used. In this situation, the fluctuation of the best cut-off is null and all the mutations are considered as absolutely unknown (probability of pathogenicity equal to 0.5).



S20 Fig. Western blot analysis

After 4 hours of BRCA1 expression, lysates of 6×10^6 cells were examined for the presence of the protein (theoretical size: 200 kDa) with an anti-BRCA1 antibody. Tubulin or Actin was used as a loading control and was probed using an anti-Tubulin or anti-Actin antibody on the same membrane after stripping the first labeling. Signal intensities of full lanes, relatively to the BRCA1 lane, are indicated below. Of note, protein levels three times higher than the WT BRCA1 protein level (normalized to 1) systematically correspond to pathogenic mutations.

(A) BRCA1 (Colony Size and Liquid Medium assays).

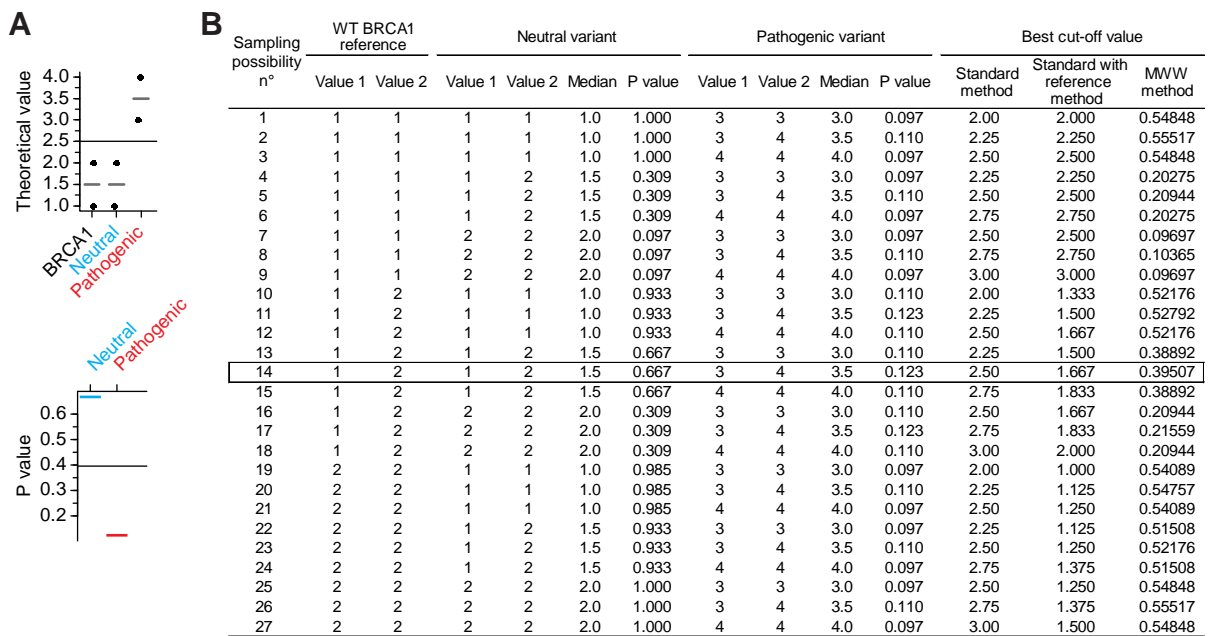
(B) BRCA1-mCherry (Spot Formation and Yeast Localization assays).

(C-G) Dotplot with the Spearman coefficient of correlation indicated. Pathogenic and neutral mutations, as well as the WT BRCA1 reference, are represented by a red, blue or black dot, respectively.

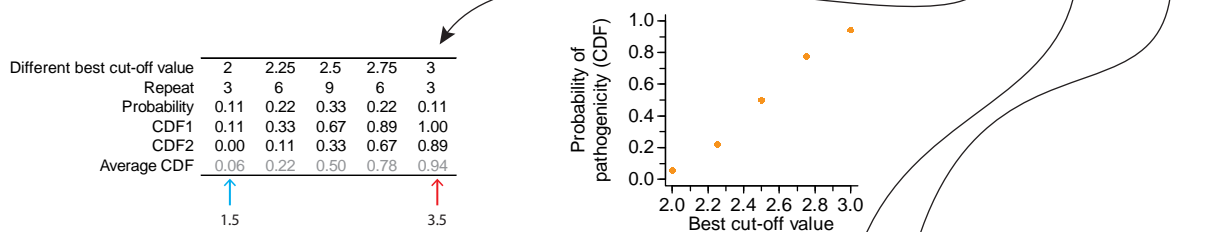
(C) Correlation between the relative signal intensities of A and B.

(D-E) Correlation between the relative signal intensities of A and medians of the Colony Size or Liquid Medium assay.

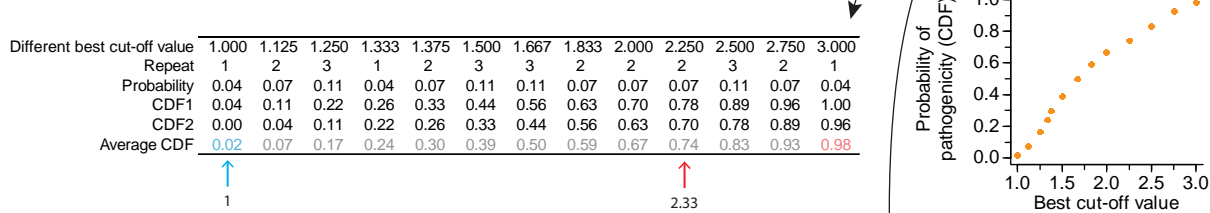
(F-G), correlation between the relative signal intensities of B and medians of the Spot Formation or Yeast Localization assay.



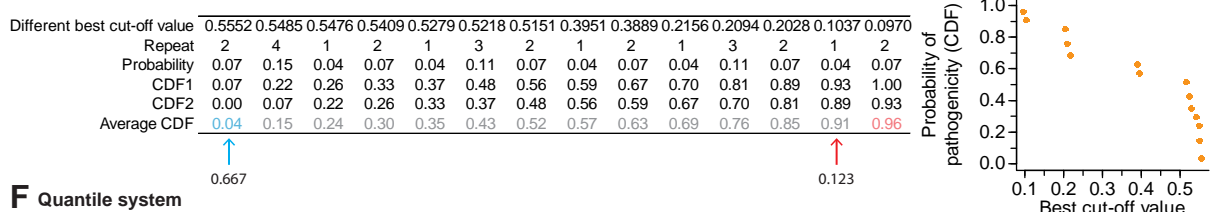
C Probability system & standard method



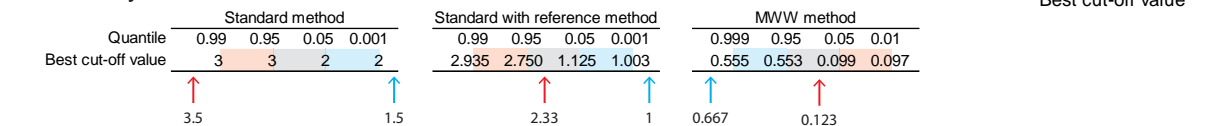
D Probability system & standard with reference method



E Probability system & MWW method



F Quantile system



S21 Fig. Exact probability distribution of the best cut-off in the standard, standard with reference and MWW methods (theoretical situation)

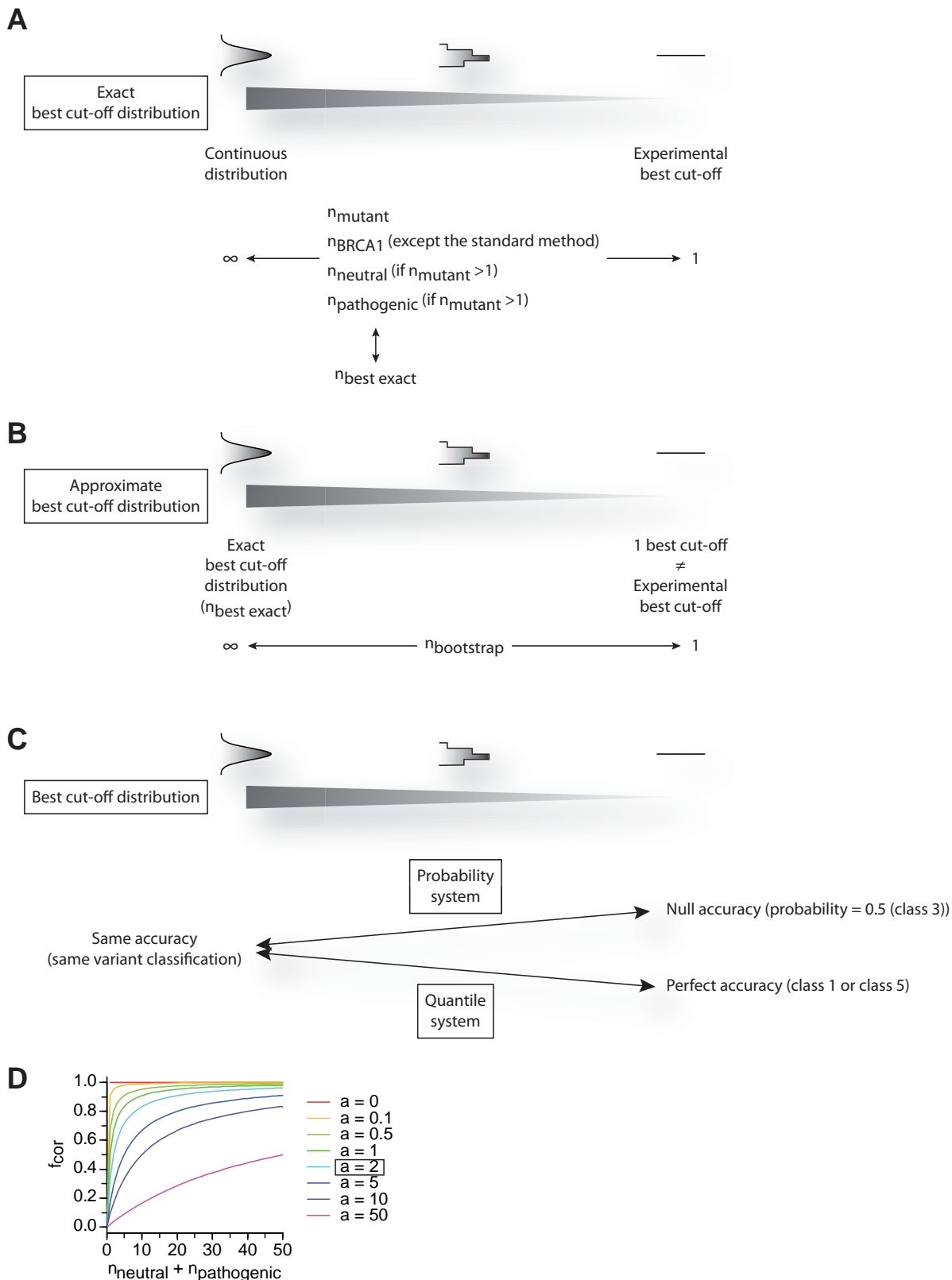
The theoretical situation was analyzed as follows: one neutral and one pathogenic mutation ($n_{\text{neutral}} = 1$ and $n_{\text{pathogenic}} = 1$), with two values per mutant ($n_{\text{mutant}} = 2$, value 1 and 2 for the neutral mutant, and value 3 and 4 for the pathogenic mutant) and two values in the WT BRCA1 reference ($n_{\text{BRCA1}} = 2$, value 1 and 2).

(A) The graphs depicted are similar to those in Fig 1, except that boxplots are replaced by dotplots with median of the distributions indicated by a grey segment. The black horizontal line represents the experimental best cut-off. The best cut fluctuations (colored areas) are not represented but quantiles are shown in F.

(B) Table recapitulating all of the possible results when sampling 2 values, with replacement, among the 2 neutral, 2 pathogenic and 2 WT BRCA1 values. Each row is a different combination that provides a best cut-off value, for each method used. The framed row highlights the combination identical to the experimental situation in A. In this simple situation (1 neutral and 1 pathogenic variant), the best cut off computed, in each row, is the median of the two variant medians (standard method), the median of the two variant medians divided by the WT BRCA1 median (standard with reference method) and the median of the two variant p values (MWW method).

(C-E) Variant classification using the probability system, with the standard (C), standard with reference (D) and MWW (E) methods, as in S6B Fig. Colored numbers in the table correspond to the different probabilities of pathogenicity designed by the model. The color code respects the five-class nomenclature depicted in S1 Table: grey, class 3; light blue, class 2; pink, class 4. Positions of the neutral and pathogenic variants are represented by a blue and red arrow, respectively. The number below each arrow designates the variant value used in the probability system to attribute the probability of pathogenicity, which corresponds to the median, median divided by the WT reference median or p value, indicated in the framed row of B. For instance, in the standard method (C), the pathogenic variant, with a median of 3.5, has the probability 0.94 (class 3).

(F) Variant classification using the quantile system. Quantiles were computed from the 27 best cut-off values from B, for each method. The colored background defines the intervals within the best cut-off distribution, as explained in S23 Fig. Arrows depict the position of the neutral and pathogenic variants, as in C-E.



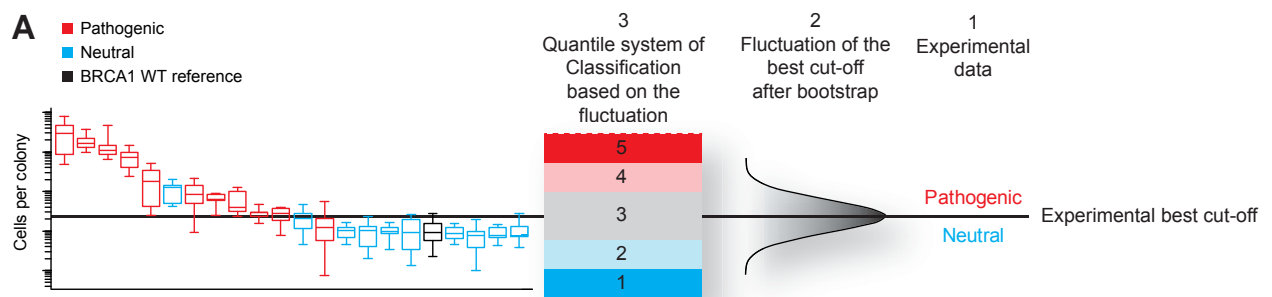
S22 Fig. Additional information about the classification model

(A) Schematic of the exact best cut-off distribution influenced by different parameters, assuming no ties. The number of neutral (n_{neutral}) and pathogenic ($n_{\text{pathogenic}}$) variants influence the exact best cut-off distribution only if the number of values per mutant (n_{mutant}) is above 1. The number of values in the WT reference (n_{BRCA1}) does not influence the exact best cut-off distribution in the standard method, only in the standard with reference and MWW methods.

(B) Schematic of the approximate best cut-off distribution influenced by the number of bootstraps performed. Importantly, a single bootstrap ($n_{\text{bootstrap}} = 1$) does not lead to the experimental best cut-off, except if $n_{\text{best exact}} = 1$.

(C) Accuracy of the probability and quantile systems of classification. The schematic illustration is valid, using either the exact or approximate best cut-off distribution.

(D) Correcting factor f_{cor} used in the probability system of classification, depending on the parameter a (see the S1 Text). $f_{\text{cor}} = (n_{\text{neutral}} + n_{\text{pathogenic}}) / (n_{\text{neutral}} + n_{\text{pathogenic}} + a)$. The framed value ($a = 2$) was the value used in S13-S15 Tables.



B

| Class | Quantile in the 2 standard methods | Quantile in the MWW method | | | |
|-----------------------|------------------------------------|----------------------------|--------|--------|--------|
| | | CS | LM | SF | YL |
| Definitely pathogenic | 5 | Q0.990 | Q0.990 | Q0.010 | Q0.990 |
| Likely pathogenic | 4 | Q0.950 | Q0.950 | Q0.050 | Q0.950 |
| Uncertain | 3 | Q0.050 | Q0.050 | Q0.950 | Q0.050 |
| Likely neutral | 2 | Q0.001 | Q0.001 | Q0.999 | Q0.001 |
| Definitely neutral | 1 | | | | |

**C Colony Size assay
Standard methods**

| Class | Probability | Corresponding cut-off values (standard method) | | Corresponding cut-off values (standard with reference method) | | |
|-------|------------------------------------|---|-------------------------------------|--|--------------------------------------|-------------------------------------|
| | | Quantile | Cut-off value (cells per colony) | Quantile | Cut-off value (BRCA1 median fold) | Cut-off value (cells per colony) |
| 5 | $P(X > Q0.99) = 1\%$ | | | | | |
| 4 | $P(Q0.95 \leq X \leq Q0.99) = 4\%$ | Q0.99 | 26,222 | Q0.99 | 2.416 | 27,062 |
| 3 | $P(Q0.05 \leq X < Q0.95) = 95\%$ | Q0.95 | 24,728 | Q0.95 | 2.236 | 25,047 |
| 2 | $P(Q0.001 \leq X < Q0.05) = 4.9\%$ | Q0.05 | 13,200 | Q0.05 | 1.173 | 13,143 |
| 1 | $P(X < Q0.001) = 0.1\%$ | Q0.001 | 12,133 | Q0.001 | 0.960 | 10,754 |

x 11,200

**D Colony Size assay
MWW method**

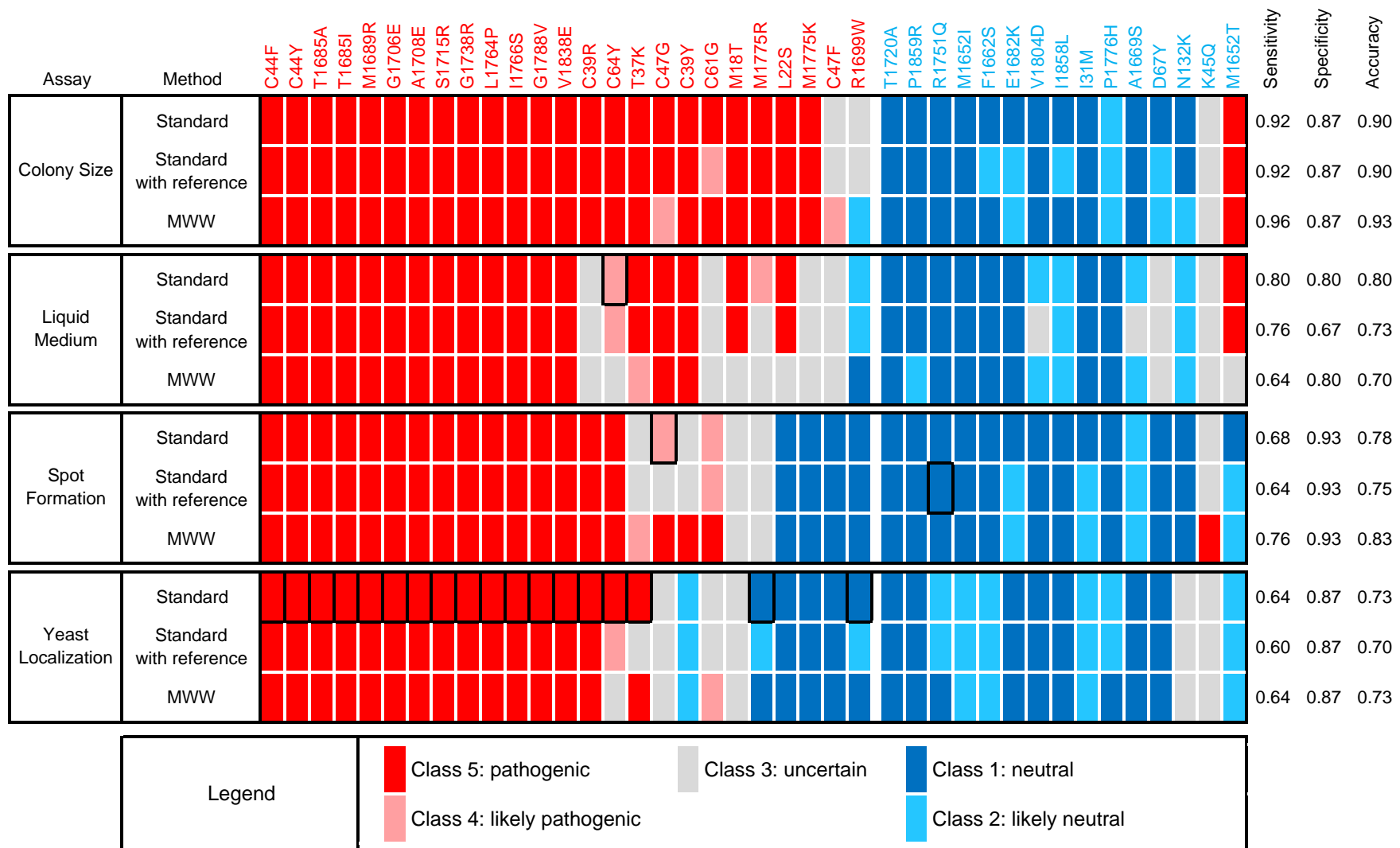
| Class | Probability | Corresponding cut-off values | |
|-------|------------------------------------|------------------------------|----------------------------|
| | | Quantile | Cut-off value (p value) |
| 5 | $P(X < Q0.01) = 1\%$ | | |
| 4 | $P(Q0.01 \leq X \leq Q0.05) = 4\%$ | Q0.01 | 5.49e-5 |
| 3 | $P(Q0.05 < X \leq Q0.95) = 95\%$ | Q0.05 | 0.00023 |
| 2 | $P(Q0.95 < X \leq Q0.999) = 4.9\%$ | Q0.95 | 0.152 |
| 1 | $P(X > Q0.999) = 0.1\%$ | Q0.999 | 0.313 |

Figure S23. Description of the quantile system of classification

(A) The left figure depicts a theoretical waterfall distribution of pathogenic and neutral missense mutations, as in Fig 1A. (1) Variant classification according to the experimental best cut-off. This cut-off (horizontal black line), that maximizes the experimental sensitivity and specificity in the waterfall distribution, is obtained by ROC curve analysis, as in S2 Fig. In the case of the Colony Size assay, mutations above the best cut-off are classified as pathogenic and mutations below are classified as neutral. (2) Bootstrap analysis provides a fluctuation of the best cut-off, depending on the values of the mutations and the WT BRCA1 reference randomly chosen. The fluctuating best cut-off values form a distribution, as depicted in the schematic. (3) Quantile system of variant classification according to the fluctuation of the best cut-off. The reasoning is the following: the distribution of the fluctuating best cut-off defines quantiles (Q) that delimit the probability of the presence of this variable. As an example, the quantile Q0.99 is the value that separates the 99% lowest values from the 1% highest values in a distribution. This means that the probability to have the best cut-off above the quantile Q0.99 is 1%. Thus, in the Colony Size assay using the standard method, a mutation with the median above the quantile Q0.99 can be considered as pathogenic with a 1% probability of error. Indeed, this mutation could be neutral, but only if the best cut-off is above the median, which has a 1% probability, or less, to occur. This reasoning allows separation of the best cut-off distribution into 5 intervals, based on the five-class nomenclature proposed by Plon et al, with each interval defining the probability of the best cut-off presence within the waterfall distribution.

(B) Quantiles that delimit the 5 intervals of classification according to the assay and the method used. CS, Colony Size; LM, Liquid Medium; SF, Spot Formation; YL, Yeast Localization assay. Note that the quantiles differ, depending on whether the pathogenic mutations are above or below the best cut-off. For instance, in the standard method, the quantiles of the Colony Size assay are Q0.99, Q0.95, Q0.05 and Q0.001 (pathogenic mutants above the best cut-off), while quantiles are Q0.01, Q0.05, Q0.95 and Q0.999 in the Spot Formation assay (pathogenic mutants below the best cut-off). However, these two cases generate the same intervals (e.g., probability 1% for the class 5, see C and D). Cut-off values corresponding to these quantiles are listed in S5 Table for each assay and for each method.

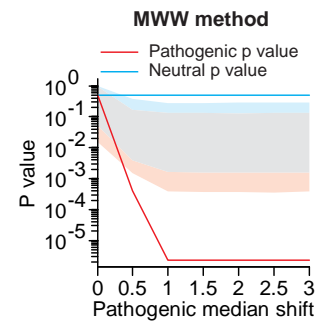
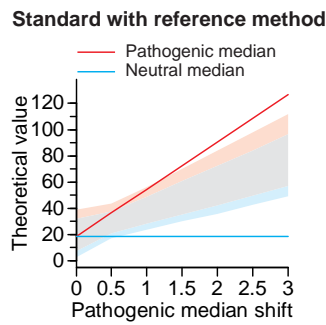
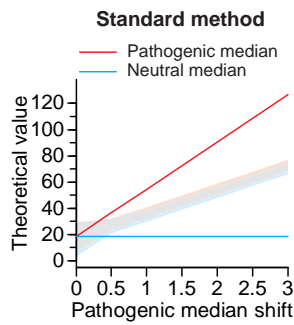
(C) Interval limits in the case of the Colony Size assay, using the standard or the standard with reference method. $P(X > Q0.99) = 1\%$ is the probability to obtain the best cut-off variable X strictly over the quantile Q0.99, shown here as 26,222 cells per colony for the standard method, and $2.416 \times 11,200$ (BRCA1 median of the experimental data) = 27,062 cells per colony for the standard with reference method (D) Interval limits in the case of the Colony Size assay, using the MWW method.



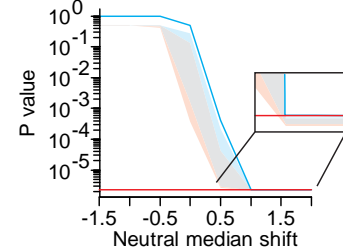
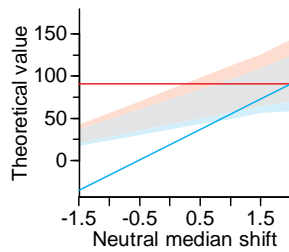
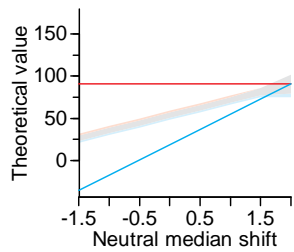
S24 Fig. Variant classification using the quantile system

Names in red and in blue indicate the pathogenic and neutral mutations, respectively, according to their prior classification. See also S16 Table. The black frames pinpoint the divergent classification compared to that in the probability system (Fig 2B).

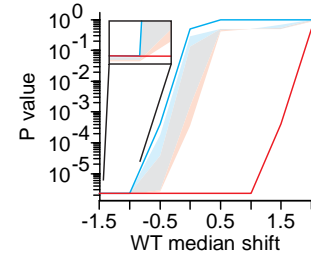
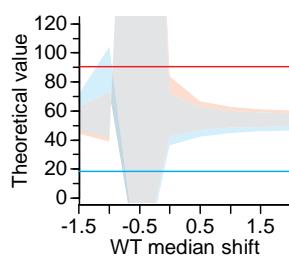
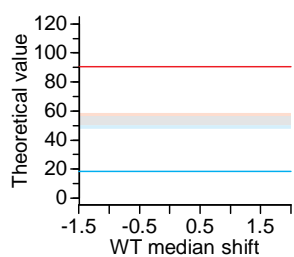
A Pathogenic variant position



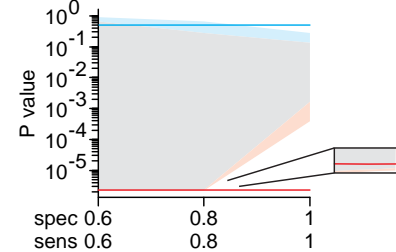
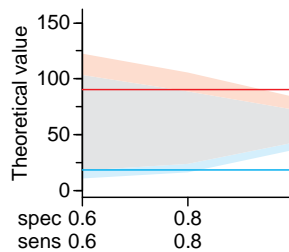
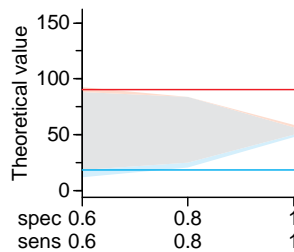
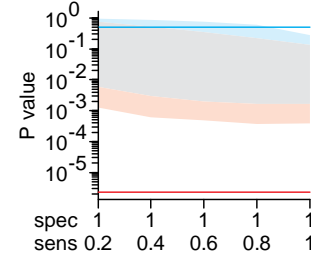
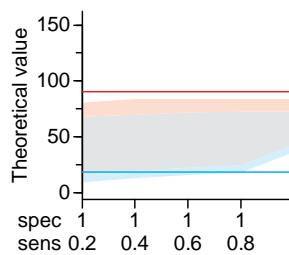
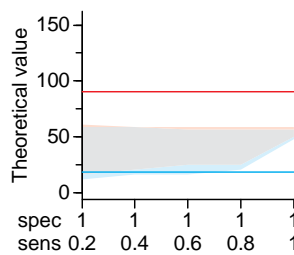
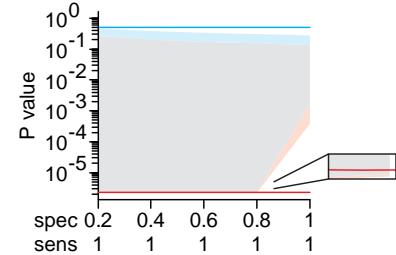
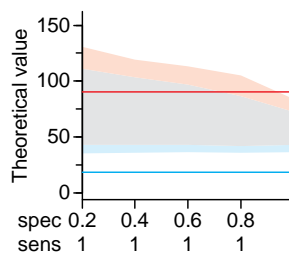
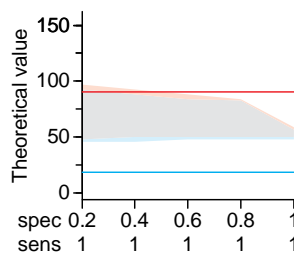
B Neutral variant position



C WT reference position



D Experimental sensitivity and specificity



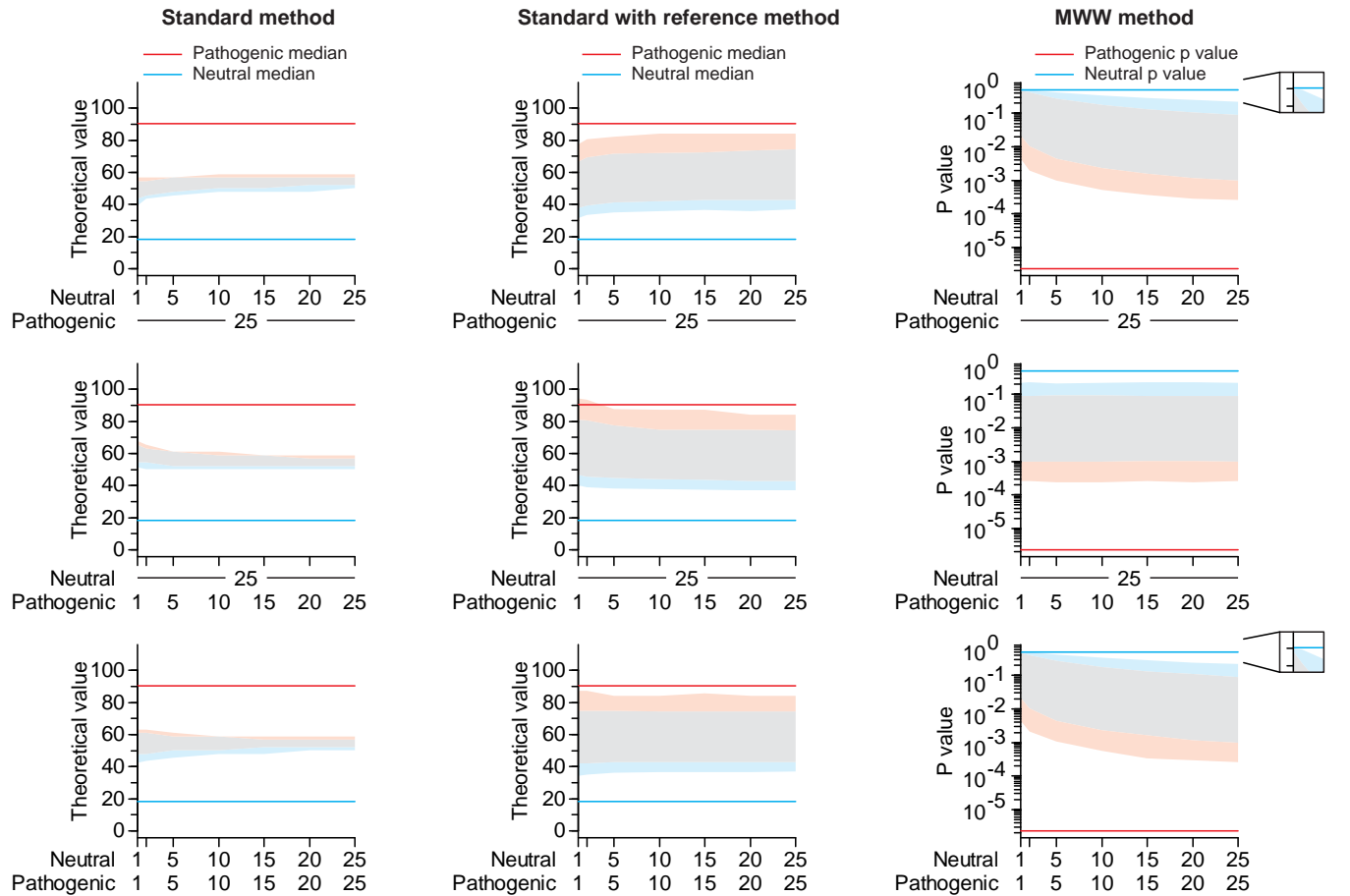
S25 Fig. Quantile system of classification (theoretical situation)

Effect of different experimental parameters was assessed in theoretical situations, exactly as for the probability system of classification, meaning that the best cut-off fluctuations depicted were those used in S13-S19 Figs. Red line, position of the median or p value of the pathogenic mutants; blue line, position of the median or p value of the neutral mutants. The pink, grey and blue areas define intervals within the best cut-off distribution, as explained in S23 Fig. For clarity, the extreme red and blue areas were not displayed. Sensitivity of the quantile system is maximal when the red line is beyond the pink area. Specificity is maximal when the blue line is beyond the light blue area. Accuracy is maximal when sensitivity and specificity are maximal. Finally, sensitivity, specificity and accuracy of the quantile system are null when both lines are in the grey area, or in the wrong side of the best cut-off fluctuation.

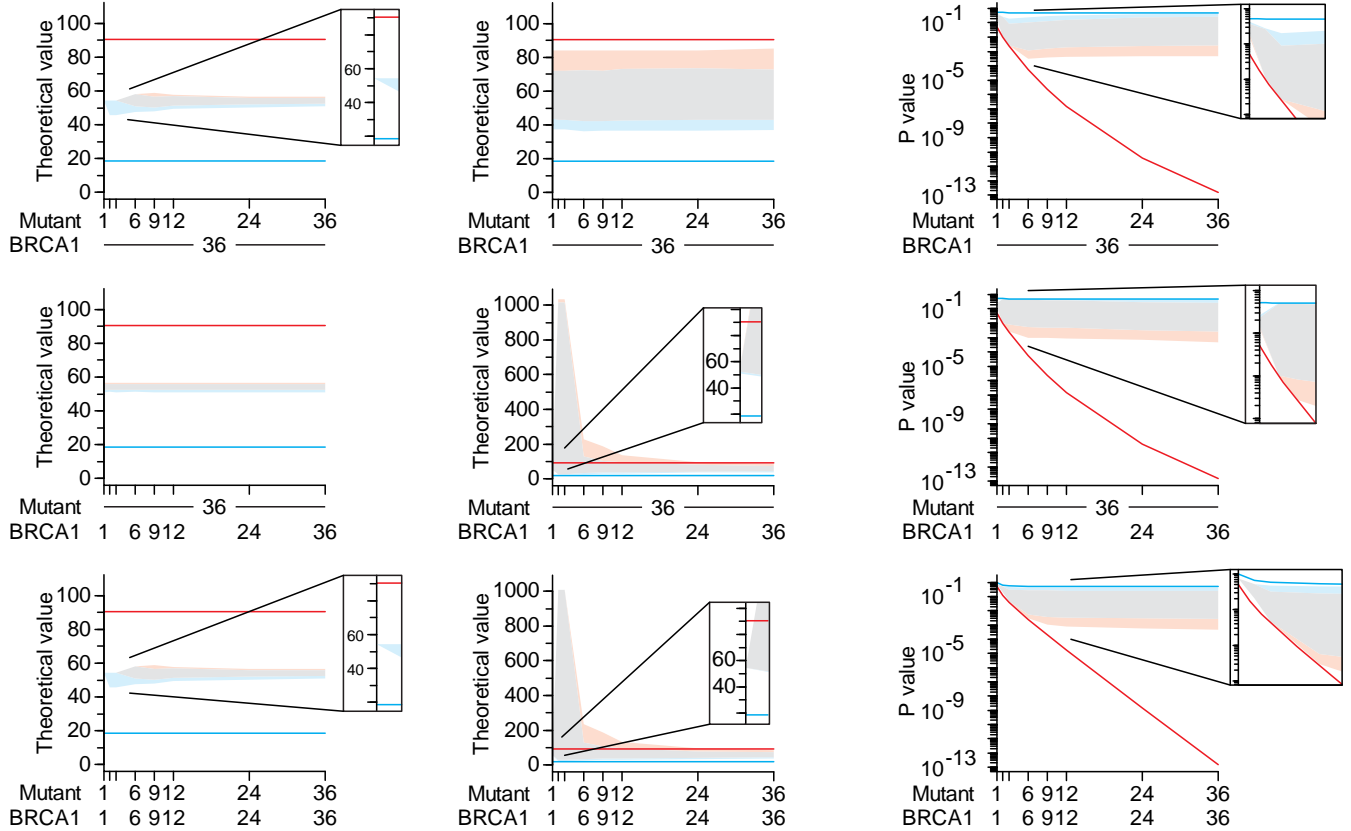
(A-D) Evolution of the best cut-off fluctuation depending on either the shift intensity of the pathogenic mutations (A), or the shift intensity of the neutral mutations (B), or the shift intensity of the WT reference (C), or the experimental sensitivity and specificity (D). The corresponding panels, depicted for the probability system of classification, are shown in S13D Fig for A, S14E Fig for B, S15E Fig for C and S16E-G Fig for D.

As summarized in S17 Table, these results did not reveal any flaws. The quantile system behaves as the probability system, in these situations.

A Number of variants



B Number of values

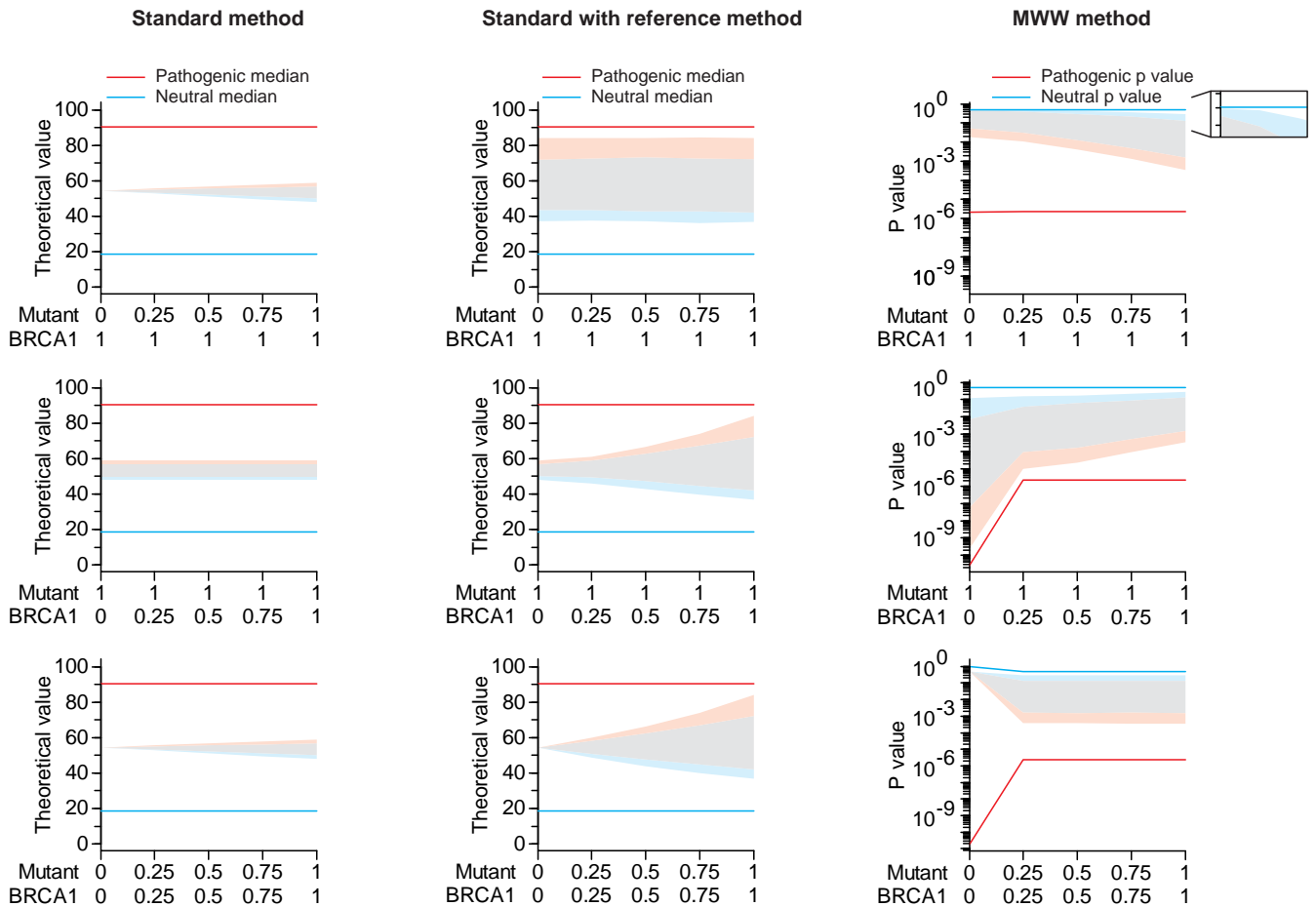


S26 Fig. Quantile system of classification (theoretical situation)

Evolution of the best cut-off fluctuation depending on either the number of neutral and pathogenic mutations (A) or the number of mutant and BRCA1 values (B), as in S25 Fig. The corresponding panels, depicted for the probability system of classification, are shown in S17E-G Fig for A and S18E-G Fig for B.

As summarized in S17 Table, these results reveal a major flaw in the quantile system of classification. Using the standard and MWW methods, the sensitivity and specificity is maximal, regardless of the number of values present within the mutant or BRCA1 distributions (B). Using the standard with reference method, the sensitivity is affected when the number of values in the BRCA1 distribution is decreased, but is maximal when $n_{BRCA1} = 1$. Thus, contrary to the probability system, the quantile system is not correctly influenced by the amount of experimental values resulting from functional assessment.

G Distribution range



S27 Fig. Quantile system of classification (theoretical situation)

Evolution of the best cut-off fluctuation, depending on the range of the mutant and BRCA1 distributions, as in S25 Fig. The corresponding panels, depicted for the probability system of classification, are shown in S19E-G Fig.

As summarized in S17 Table, these results reveal a major flaw in the quantile system of classification. A null range means that all of the values, present in a distribution, are identical (ties). Because ties are related to a low measurement accuracy, an efficient variant classifier should penalize a high number of ties in a dataset, which is not observed here, whatever method is used.

S1 Table. IARC variant classification

| Class | Description | Probability of pathogenicity | Clinical decision-making | |
|-------|---|------------------------------|---|--|
| | | | Genetic testing | Surveillance recommendations |
| 5 | Definitely pathogenic | $p > 0.99$ | Test at-risk relatives | Full high-risk surveillance |
| 4 | Likely pathogenic | $0.99 \geq p \geq 0.95$ | Test at-risk relatives | Full high-risk surveillance |
| 3 | Uncertain (Variant of unknown significance) | $0.95 > p \geq 0.05$ | Do not use as predictive testing in at-risk relatives | Counsel based on family history and other risk factors |
| 2 | Likely neutral | $0.05 > p \geq 0.001$ | Do not use as predictive testing in at-risk relatives | Counsel as if no mutation detected |
| 1 | Neutral | $p < 0.001$ | Do not use as predictive testing in at-risk relatives | Counsel as if no mutation detected |

Five-class nomenclature proposed by the International Agency for Research on Cancer (IARC) for variant classification, with specific recommendations for clinical management, depending on the probability of pathogenicity obtained by epidemiological methods (Plon et al., 2008).

S2 Table. BRCA1 Mutations selected ^a

| Variant | Nomenclature HGVS ^b | DNA Nomenclature HGVS | UMD-BRCA1 ^{c, d} | LOVD-IARC ^{c, e} | LOVD Leiden ^f | Millot et. al., 2012 ^{c, g} |
|---------|--------------------------------|-----------------------|---------------------------|---------------------------|----------------------------|--------------------------------------|
| M18T | p.Met18Thr | c.53T>C | 3 | 4 | Ambiguous | 4 |
| L22S | p.Leu22Ser | c.65T>C | | 5 | Predicted pathogenic | 5 |
| I31M | p.Ile31Met | c.93C>G | | | Predicted neutral | |
| T37K | p.Thr37Lys | c.110C>A | | 5 | Predicted pathogenic | 5 |
| C39R | p.Cys39Arg | c.115T>C | 5 | 5 | As mutant control | 5 |
| C39Y | p.Cys39Tyr | c.116G>A | | | Impact on protein function | |
| C44F | p.Cys44Phe | c.131G>T | 5 | | Predicted pathogenic | |
| C44Y | p.Cys44Tyr | c.131G>A | 5 | 5 | | 5 |
| K45Q | p.Lys45Gln | c.133A>C | 2 | 1 | Predicted neutral | 1 |
| C47F | p.Cys47Phe | c.140G>T | 5 | | | |
| C47G | p.Cys47Gly | c.139T>G | | | Predicted pathogenic | |
| C61G | p.Cys61Gly | c.181T>G | 5 | 5 | Ambiguous | 5 |
| C64Y | p.Cys64Tyr | c.191G>A | 5 | | | 4 / 5 |
| D67Y | p.Asp67Tyr | c.199G>T | 1 | 1 | Predicted neutral | 1 |
| N132K | p.Asn132Lys | c.396C>A | 2 | 1 | Ambiguous | 1 |
| M1652I | p.Met1652Ile | c.4956G>A | 1 | 1 | Ambiguous | 1 |
| M1652T | p.Met1652Thr | c.4955T>C | 2 | 1 | Ambiguous | 1 |
| F1662S | p.Phe1662Ser | c.4985T>C | | | Predicted neutral | 1 |
| A1669S | p.Ala1669Ser | c.5005G>T | 3 | | Predicted neutral | |
| E1682K | p.Glu1682Lys | c.5044G>A | | 1 | Predicted neutral | 1 |
| T1685A | p.Thr1685Ala | c.5053A>G | 3 | 5 | Predicted pathogenic | 5 |
| T1685I | p.Thr1685Ile | c.5054C>T | | 5 | Predicted pathogenic | 5 |
| M1689R | p.Met1689Arg | c.5066T>G | | 4 | Predicted pathogenic | 4 |
| R1699W | p.Arg1699Trp | c.5095C>T | 5 | 5 | Ambiguous | 5 |
| G1706E | p.Gly1706Glu | c.5117G>A | 4 | | Predicted pathogenic | 5 |
| A1708E | p.Ala1708Glu | c.5123C>A | 5 | 5 | Ambiguous | 5 |
| S1715R | p.Ser1715Arg | c.5143A>C | | 5 | Predicted pathogenic | 5 |
| T1720A | p.Thr1720Ala | c.5158A>G | 1 | 1 | Predicted neutral | 1 |
| G1738R | p.Gly1738Arg | c.5212G>A | | 5 | Predicted pathogenic | 5 |
| R1751Q | p.Arg1751Gln | c.5252G>A | 2 | 1 | Ambiguous | 1 |
| L1764P | p.Leu1764Pro | c.5291T>C | 3 | 5 | Ambiguous | 5 |
| I1766S | p.Ile1766Ser | c.5297T>G | | 5 | Ambiguous | 5 |
| M1775K | p.Met1775Lys | c.5324T>A | | 5 | Ambiguous | 5 |
| M1775R | p.Met1775Arg | c.5324T>G | 5 | 5 | Ambiguous | 5 |
| P1776H | p.Pro1776His | c.5327C>A | | 2 | Predicted neutral | 2 |
| G1788V | p.Gly1788Val | c.5363G>T | | 5 | Predicted pathogenic | 5 |
| V1804D | p.Val1804Asp | c.5411T>A | 1 | 1 | Ambiguous | 1 |
| V1838E | p.Val1838Glu | c.5513T>A | | 5 | Predicted pathogenic | 5 |
| I1858L | p.Ile1858Leu | c.5572A>C | 2 | 1 | Predicted neutral | 1 |
| P1859R | p.Pro1859Arg | c.5576C>G | | 1 | Predicted neutral | 1 |

^a Empty cell, no data^b HGVS: human genome variation society (<http://www.hgvs.org/mutnomen/>).^c IARC classification as in Table S1.^d UMD-BRCA1 database (29-January-2015, <http://www.umd.be/BRCA1/>).^e LOVD-IARC database (29-January-2015, http://hci-exlovd.hci.utah.edu/home.php?select_db=BRCA1).^f LOVD Leiden database (29-January-2015, <http://databases.lovd.nl/shared/genes/BRCA1>). This database gathers all information from the literature, including functional assays, which explains the high level of ambiguous results. Blue cell, neutral; red cell, pathogenic; grey cell, conflicting reports.^g Mutations recommended by the ENIGMA consortium. C64Y is namely classified as "Clinically important" which is here converted to "4/5".

S3 Table. Misclassified mutations using the experimental best cut-offs

| Assay | Standard method | MWW method |
|--------------------|--|--|
| Colony Size | M1652T R1699W | M1652T R1699W |
| Liquid Medium | M1652T C47F R1699W M1775K | L22S C39R C47F C61G R1699W |
| Spot Formation | K45Q L22S C47F R1699W M1775K | K45Q L22S C47F R1699W M1775K |
| Yeast Localization | K45Q L22S C39Y C47F R1699W M1775K M1775R | K45Q L22S C39Y C47F R1699W M1775K M1775R |

Experimental best cut-offs from Table 1.

S4 Table. Experimental data from the 4 functional assays and the siRNA screening

| Assay | Name | Median | Relative median | MWW p value | Sample size | Ties |
|----------------|--------|-----------|-----------------------|-----------------------|-------------|-------|
| | BRCA1 | 11 200 | 1 | | n = 36 | 2,4,2 |
| | Vector | 4 146 055 | 370.18348 | 1.515234e-13 | n = 36 | 0 |
| | L1764P | 328 807 | 29.35777 | 2.274705e-06 | n = 9 | 2 |
| Colony Size | M1689R | 1 877 333 | 167.61902 | 2.276350e-06 | n = 9 | 0 |
| | V1838E | 1 621 333 | 144.76188 | 2.276350e-06 | n = 9 | 0 |
| | G1706E | 1 536 000 | 137.14286 | 2.276350e-06 | n = 9 | 0 |
| | T1685I | 1 354 667 | 120.95241 | 2.276350e-06 | n = 9 | 0 |
| | A1708E | 924 771 | 82.56884 | 2.276350e-06 | n = 9 | 0 |
| | S1715R | 878 291 | 78.41884 | 2.276350e-06 | n = 9 | 0 |
| | T1685A | 760 367 | 67.88991 | 2.276350e-06 | n = 9 | 0 |
| | G1788V | 531 282 | 47.43589 | 2.276350e-06 | n = 9 | 0 |
| | I1766S | 416 410 | 37.17946 | 2.276350e-06 | n = 9 | 0 |
| | C44F | 145 000 | 12.94643 | 2.276350e-06 | n = 9 | 0 |
| | C44Y | 117 651 | 10.50455 | 2.276350e-06 | n = 9 | 0 |
| | G1738R | 92 855 | 8.29063 | 2.276350e-06 | n = 9 | 0 |
| | C39Y | 79 733 | 7.11902 | 2.276350e-06 | n = 9 | 0 |
| | C64Y | 60 624 | 5.41286 | 2.276350e-06 | n = 9 | 0 |
| | C39R | 49 867 | 4.45241 | 2.276350e-06 | n = 9 | 0 |
| | M1775R | 49 299 | 4.40170 | 2.276350e-06 | n = 9 | 0 |
| | M1652T | 45 108 | 4.02750 | 2.276350e-06 | n = 9 | 0 |
| | M1775K | 42 120 | 3.76071 | 2.276350e-06 | n = 9 | 0 |
| | L22S | 38 507 | 3.43813 | 2.276350e-06 | n = 9 | 0 |
| | C61G | 27 024 | 2.41286 | 2.606554e-06 | n = 9 | 0 |
| | T37K | 28 053 | 2.50473 | 4.446440e-06 | n = 9 | 0 |
| | M18T | 29 182 | 2.60554 | 2.951174e-05 | n = 9 | 0 |
| | C47G | 29 000 | 2.58929 | 1.323210e-04 | n = 9 | 0 |
| | C47F | 17 920 | 1.60000 | 1.835069e-04 | n = 9 | 2 |
| | K45Q | 17 900 | 1.59821 | 0.0184710518614082525 | n = 9 | 0 |
| | R1699W | 13 300 | 1.18750 | 0.2302354198061955237 | n = 9 | 2 |
| | D67Y | 11 520 | 1.02857 | 0.4604291821344998481 | n = 9 | 0 |
| | E1682K | 11 303 | 1.00920 | 0.4660662497636753332 | n = 9 | 2 |
| | N132K | 10 200 | 0.91071 | 0.4943379145802865859 | n = 9 | 0 |
| | P1776H | 12 253 | 1.09402 | 0.5169816950327198457 | n = 9 | 0 |
| | F1662S | 11 966 | 1.06839 | 0.6279557960277021289 | n = 9 | 0 |
| | T1720A | 9 042 | 0.80732 | 0.6492252229136958874 | n = 9 | 2 |
| | I1858L | 10 773 | 0.96188 | 0.6853049210510774314 | n = 9 | 2 |
| A1669S | 10 243 | 0.91455 | 0.7566189006559524977 | n = 9 | 0 | |
| P1859R | 10 300 | 0.91964 | 0.8465966574550906065 | n = 9 | 2 | |
| M1652I | 10 070 | 0.89911 | 0.8532174576053266346 | n = 9 | 2 | |
| R1751Q | 9 300 | 0.83036 | 0.9619810348434899572 | n = 9 | 0 | |
| I31M | 7 600 | 0.67857 | 0.9619810348434899572 | n = 9 | 0 | |
| V1804D | 6 239 | 0.55705 | 0.9976319137889332334 | n = 9 | 0 | |

| | | | | | | |
|--|--------|-------|---------|-----------------------|--------|------------------------|
| | BRCA1 | 0.144 | 1 | | n = 36 | 2, 2, 2, 4, 2, 2, 2, 2 |
| | Vector | 0.448 | 3.11111 | 1.510622e-13 | n = 36 | 2, 2 |
| | A1708E | 0.368 | 2.55556 | 2.264851e-06 | n = 9 | 2, 2 |
| | C44F | 0.239 | 1.65972 | 2.264851e-06 | n = 9 | 2, 2 |
| | G1788V | 0.407 | 2.82639 | 2.266490e-06 | n = 9 | 2 |
| | G1706E | 0.397 | 2.75694 | 2.266490e-06 | n = 9 | 2 |
| | I1766S | 0.393 | 2.72917 | 2.266490e-06 | n = 9 | 2 |
| | M1689R | 0.388 | 2.69444 | 2.266490e-06 | n = 9 | 2 |
| | T1685I | 0.387 | 2.68750 | 2.266490e-06 | n = 9 | 2 |
| | T1685A | 0.338 | 2.34722 | 2.266490e-06 | n = 9 | 2 |
| | C44Y | 0.204 | 1.41667 | 2.266490e-06 | n = 9 | 2 |
| | V1838E | 0.405 | 2.81250 | 2.268131e-06 | n = 9 | 0 |
| | S1715R | 0.387 | 2.68750 | 2.268131e-06 | n = 9 | 0 |
| | L1764P | 0.361 | 2.50694 | 2.268131e-06 | n = 9 | 0 |
| | G1738R | 0.295 | 2.04861 | 2.268131e-06 | n = 9 | 0 |
| | C39Y | 0.205 | 1.42361 | 2.268131e-06 | n = 9 | 0 |
| | C47G | 0.186 | 1.29167 | 1.594213e-05 | n = 9 | 2, 2 |
| | T37K | 0.190 | 1.31944 | 2.606480e-05 | n = 9 | 2 |
| | M18T | 0.184 | 1.27778 | 7.096329e-05 | n = 9 | 0 |
| | M1775R | 0.177 | 1.22917 | 1.116480e-04 | n = 9 | 2 |
| | C64Y | 0.182 | 1.26389 | 1.641808e-04 | n = 9 | 0 |
| | M1775K | 0.170 | 1.18056 | 2.152737e-04 | n = 9 | 0 |
| | K45Q | 0.172 | 1.19444 | 2.526677e-04 | n = 9 | 0 |
| | C61G | 0.173 | 1.20139 | 3.118077e-04 | n = 9 | 0 |
| | M1652T | 0.195 | 1.35417 | 3.279441e-04 | n = 9 | 0 |
| | L22S | 0.162 | 1.12500 | 6.346827e-04 | n = 9 | 2 |
| | D67Y | 0.161 | 1.11806 | 0.0043108868173617884 | n = 9 | 2 |
| | C39R | 0.174 | 1.20833 | 0.0048755976324813856 | n = 9 | 0 |
| | C47F | 0.161 | 1.11806 | 0.0057497334195825249 | n = 9 | 2 |
| | N132K | 0.152 | 1.05556 | 0.1136660718705307938 | n = 9 | 0 |
| | I1858L | 0.152 | 1.05556 | 0.1434656399958101347 | n = 9 | 0 |
| | P1859R | 0.144 | 1.00000 | 0.1853384111507014043 | n = 9 | 2 |
| | V1804D | 0.158 | 1.09722 | 0.2568378526562776765 | n = 9 | 0 |
| | A1669S | 0.158 | 1.09722 | 0.2568447967519839059 | n = 9 | 0 |
| | E1682K | 0.148 | 1.02778 | 0.3613056803591045973 | n = 9 | 0 |
| | R1699W | 0.154 | 1.06944 | 0.4491694377259908277 | n = 9 | 0 |
| | T1720A | 0.142 | 0.98611 | 0.4830127077145046077 | n = 9 | 2 |
| | M1652I | 0.141 | 0.97917 | 0.8068922049917628403 | n = 9 | 2, 2 |
| | I31M | 0.139 | 0.96528 | 0.8220402013577042810 | n = 9 | 0 |
| | P1776H | 0.134 | 0.93056 | 0.8835070202068154854 | n = 9 | 0 |
| | F1662S | 0.130 | 0.90278 | 0.9487627758129135724 | n = 9 | 2 |
| | R1751Q | 0.128 | 0.88889 | 0.9979400197432480857 | n = 9 | 3, 2 |

| | BRCA1 | 0.32011331 | 1 | | n = 12 | 0 |
|----------------|--------|------------|---------|-------------|--------|---|
| | I1858L | 0.46986945 | 1.46782 | 0.994594772 | n = 3 | 0 |
| | D67Y | 0.39541215 | 1.23523 | 0.994594772 | n = 3 | 0 |
| | L22S | 0.40228548 | 1.25670 | 0.991872140 | n = 3 | 0 |
| | M1775K | 0.37789528 | 1.18050 | 0.991872140 | n = 3 | 0 |
| | N132K | 0.37003058 | 1.15594 | 0.991872140 | n = 3 | 0 |
| | R1699W | 0.3539823 | 1.10580 | 0.982649712 | n = 3 | 0 |
| | C47F | 0.33136225 | 1.03514 | 0.936900008 | n = 3 | 0 |
| | M1652I | 0.33056577 | 1.03265 | 0.788446042 | n = 3 | 0 |
| | F1662S | 0.32405386 | 1.01231 | 0.788446042 | n = 3 | 0 |
| | V1804D | 0.35462842 | 1.10782 | 0.743893964 | n = 3 | 0 |
| | T1720A | 0.34723473 | 1.08472 | 0.743893964 | n = 3 | 0 |
| | R1751Q | 0.32063492 | 1.00163 | 0.743893964 | n = 3 | 0 |
| | P1776H | 0.34312296 | 1.07188 | 0.694887867 | n = 3 | 0 |
| | P1859R | 0.32621083 | 1.01905 | 0.694887867 | n = 3 | 0 |
| | E1682K | 0.31781733 | 0.99283 | 0.470973919 | n = 3 | 0 |
| | M1652T | 0.31916635 | 0.99704 | 0.357888106 | n = 3 | 0 |
| | I31M | 0.31843575 | 0.99476 | 0.357888106 | n = 3 | 0 |
| | A1669S | 0.3021984 | 0.94404 | 0.305112133 | n = 3 | 0 |
| | M1775R | 0.2877229 | 0.89882 | 0.024638246 | n = 3 | 0 |
| | M18T | 0.23936689 | 0.74776 | 0.017350288 | n = 3 | 0 |
| | T37K | 0.24777513 | 0.77402 | 0.008127860 | n = 3 | 0 |
| Spot Formation | C39Y | 0.24834054 | 0.77579 | 0.005405228 | n = 3 | 0 |
| | K45Q | 0.23125 | 0.72240 | 0.005405228 | n = 3 | 0 |
| | C47G | 0.22693483 | 0.70892 | 0.005405228 | n = 3 | 0 |
| | C61G | 0.22481563 | 0.70230 | 0.005405228 | n = 3 | 0 |
| | G1738R | 0.20485484 | 0.63994 | 0.005405228 | n = 3 | 0 |
| | C64Y | 0.20008394 | 0.62504 | 0.005405228 | n = 3 | 0 |
| | V1838E | 0.19506031 | 0.60935 | 0.005405228 | n = 3 | 0 |
| | C44Y | 0.18793555 | 0.58709 | 0.005405228 | n = 3 | 0 |
| | G1788V | 0.17286119 | 0.54000 | 0.005405228 | n = 3 | 0 |
| | L1764P | 0.1658962 | 0.51824 | 0.005405228 | n = 3 | 0 |
| | T1685I | 0.1631555 | 0.50968 | 0.005405228 | n = 3 | 0 |
| | C39R | 0.15864589 | 0.49559 | 0.005405228 | n = 3 | 0 |
| | C44F | 0.13445378 | 0.42002 | 0.005405228 | n = 3 | 0 |
| | I1766S | 0.13080603 | 0.40862 | 0.005405228 | n = 3 | 0 |
| | T1685A | 0.10702164 | 0.33432 | 0.005405228 | n = 3 | 0 |
| | A1708E | 0.09740219 | 0.30427 | 0.005405228 | n = 3 | 0 |
| | M1689R | 0.09630668 | 0.30085 | 0.005405228 | n = 3 | 0 |
| | S1715R | 0.09555141 | 0.29849 | 0.005405228 | n = 3 | 0 |
| | G1706E | 0.08198442 | 0.25611 | 0.005405228 | n = 3 | 0 |

| | | | | | | |
|-----------------------|----------|-------------|----------|-------------|-----------|------------------------------|
| | BRCA1 | 0.03011169 | 1 | | n = 12 | 0 |
| | V1838E | 1 | 33.20969 | 0.005733011 | n = 3 | 2 |
| | G1788V | 0.997409639 | 33.12367 | 0.005769994 | n = 3 | 0 |
| | S1715R | 0.994699804 | 33.03368 | 0.005769994 | n = 3 | 0 |
| | G1706E | 0.984477893 | 32.69421 | 0.005769994 | n = 3 | 0 |
| | A1708E | 0.973875384 | 32.34210 | 0.005769994 | n = 3 | 0 |
| | I1766S | 0.968397449 | 32.16018 | 0.005769994 | n = 3 | 0 |
| | M1689R | 0.781520693 | 25.95406 | 0.005769994 | n = 3 | 0 |
| | L1764P | 0.743579501 | 24.69405 | 0.005769994 | n = 3 | 0 |
| | C39R | 0.471397699 | 15.65497 | 0.005769994 | n = 3 | 0 |
| | G1738R | 0.445016465 | 14.77886 | 0.005769994 | n = 3 | 0 |
| | T1685I | 0.439432238 | 14.59341 | 0.005769994 | n = 3 | 0 |
| | T1685A | 0.412329078 | 13.69332 | 0.005769994 | n = 3 | 0 |
| | C44Y | 0.392917208 | 13.04866 | 0.005769994 | n = 3 | 0 |
| | C44F | 0.375284774 | 12.46309 | 0.005769994 | n = 3 | 0 |
| | T37K | 0.231694699 | 7.69451 | 0.005769994 | n = 3 | 0 |
| | C61G | 0.136232412 | 4.52424 | 0.008619510 | n = 3 | 0 |
| | C64Y | 0.338020871 | 11.22557 | 0.035598445 | n = 3 | 0 |
| | K45Q | 0.07699986 | 2.55714 | 0.035598445 | n = 3 | 0 |
| Yeast Localization | M18T | 0.058713945 | 1.94987 | 0.048469094 | n = 3 | 0 |
| | C47G | 0.065395263 | 2.17176 | 0.064817290 | n = 3 | 0 |
| | N132K | 0.055921083 | 1.85712 | 0.306715686 | n = 3 | 0 |
| | M1652T | 0.033521303 | 1.11323 | 0.414296538 | n = 3 | 0 |
| | F1662S | 0.033474118 | 1.11167 | 0.585703462 | n = 3 | 0 |
| | C39Y | 0.019811089 | 0.65792 | 0.640891935 | n = 3 | 0 |
| | M1652I | 0.019508914 | 0.64789 | 0.693284314 | n = 3 | 0 |
| | I31M | 0.017313313 | 0.57497 | 0.741998532 | n = 3 | 0 |
| | P1776H | 0.020958299 | 0.69602 | 0.825927688 | n = 3 | 0 |
| | A1669S | 0.007331425 | 0.24347 | 0.825927688 | n = 3 | 0 |
| | R1751Q | 0.018504379 | 0.61452 | 0.860491843 | n = 3 | 0 |
| | T1720A | 0.004224351 | 0.14029 | 0.860491843 | n = 3 | 0 |
| | R1699W | 0.014217152 | 0.47215 | 0.890064145 | n = 3 | 0 |
| | P1859R | 0.002212185 | 0.07347 | 0.964401555 | n = 3 | 0 |
| | M1775R | 0.013772868 | 0.45739 | 0.974325834 | n = 3 | 0 |
| | C47F | 0.010335247 | 0.34323 | 0.974325834 | n = 3 | 0 |
| | E1682K | 0.004404384 | 0.14627 | 0.987364419 | n = 3 | 0 |
| | L22S | 0.00298102 | 0.09900 | 0.987364419 | n = 3 | 0 |
| | D67Y | 0.007577277 | 0.25164 | 0.994230006 | n = 3 | 0 |
| | M1775K | 0.00712358 | 0.23657 | 0.994230006 | n = 3 | 0 |
| | I1858L | 0.002357164 | 0.07828 | 0.996210204 | n = 3 | 0 |
| | V1804D | 0.001647348 | 0.05471 | 0.996210204 | n = 3 | 0 |
| | No siRNA | 945 | 1 | | n = 1,140 | 810 values at least twice |
| | KIF11 | 206 | 0.21799 | 2.78207E-07 | n = 12 | 0 |
| | GL2 | 1002 | 1.06032 | 0.712549331 | n = 12 | 0 |
| | GOLGA2 | 1091.636 | 1.15517 | 0.981009026 | n = 12 | 0 |
| | GTSE1 | 843 | 0.89206 | 0.005326083 | n = 12 | 0 |
| siRNA | ITGA2 | 840.5 | 0.88942 | 0.623738113 | n = 12 | 0 |
| | KCNC3 | 822.5 | 0.87037 | 0.019275675 | n = 12 | 0 |
| | LIN7A | 1086.5 | 1.14974 | 0.957604398 | n = 12 | 0 |
| | PSMB1 | 625 | 0.66138 | 6.61498E-05 | n = 12 | 0 |
| | RHOBTB3 | 817.5 | 0.86508 | 0.318506324 | n = 12 | 0 |
| | RRM2 | 418 | 0.44233 | 6.51502E-09 | n = 12 | 0 |
| | SOX13 | 972 | 1.02857 | 0.623738124 | n = 12 | 0 |

Relative median, median divided by the WT BRCA1 median or by the No siRNA median; sample size, number of values; framed mutations indicate identical p values. The "ties" column indicates the number of values repeated. For instance, in the first row, one value is repeated twice, another one four times, and a third twice ("ties" is the statistical term used to designate "identical values").

S5 Table. Distribution of the best cut-offs after bootstrap analysis

| Method | Assay | Q0.001 | Q0.005 | Q0.010 | Q0.025 | Q0.050 | Median (Q0.5) | Q0.950 | Q0.975 | Q0.990 | Q0.995 | Q0.999 |
|--|--------------------|----------|----------|----------|----------|----------|---------------|--------|--------|--------|--------|--------|
| Standard | Colony Size | 12 133 | 12 276 | 12 633 | 12 969 | 13 200 | 17 950 | 24 728 | 25 894 | 26 222 | 26 979 | 27 464 |
| | Liquid Medium | 0.151 | 0.153 | 0.153 | 0.156 | 0.159 | 0.168 | 0.176 | 0.178 | 0.182 | 0.183 | 0.188 |
| | Spot Formation | 0.218 | 0.222 | 0.222 | 0.224 | 0.227 | 0.266 | 0.300 | 0.309 | 0.315 | 0.315 | 0.315 |
| | Yeast Localization | 0.015 | 0.020 | 0.034 | 0.035 | 0.054 | 0.125 | 0.212 | 0.215 | 0.215 | 0.215 | 0.215 |
| | siRNA | 434 | 467 | 480 | 504 | 523 | 599 | 658 | 668 | 679 | 688 | 704 |
| Standard with reference | Colony Size | 0.960 | 1.038 | 1.071 | 1.128 | 1.173 | 1.603 | 2.236 | 2.324 | 2.416 | 2.468 | 2.650 |
| | Liquid Medium | 1.038 | 1.052 | 1.060 | 1.078 | 1.096 | 1.161 | 1.231 | 1.250 | 1.267 | 1.275 | 1.309 |
| | Spot Formation | 0.664 | 0.680 | 0.687 | 0.696 | 0.706 | 0.830 | 0.942 | 0.964 | 0.983 | 0.984 | 1.000 |
| | Yeast Localization | 0.399 | 0.669 | 0.819 | 1.133 | 1.324 | 4.239 | 9.785 | 10.528 | 12.143 | 13.774 | 16.191 |
| | siRNA | 0.460 | 0.493 | 0.508 | 0.532 | 0.553 | 0.634 | 0.698 | 0.709 | 0.721 | 0.730 | 0.746 |
| Standard with reference (value multiplied by the BRCA1 median) | Colony Size | 10 754 | 11 626 | 11 993 | 12 630 | 13 143 | 17 950 | 25 047 | 26 034 | 27 062 | 27 644 | 29 678 |
| | Liquid Medium | 0.149 | 0.152 | 0.153 | 0.155 | 0.158 | 0.167 | 0.177 | 0.180 | 0.182 | 0.184 | 0.189 |
| | Spot Formation | 0.212 | 0.218 | 0.220 | 0.223 | 0.226 | 0.266 | 0.301 | 0.309 | 0.315 | 0.315 | 0.320 |
| | Yeast Localization | 0.012 | 0.020 | 0.025 | 0.034 | 0.040 | 0.128 | 0.295 | 0.317 | 0.366 | 0.415 | 0.488 |
| | siRNA | 434 | 466 | 480 | 503 | 523 | 599 | 659 | 670 | 681 | 690 | 705 |
| MWW | Colony Size | 1.08E-05 | 3.28E-05 | 5.49E-05 | 0.00013 | 0.00023 | 0.010 | 0.152 | 0.222 | 0.313 | 0.381 | 0.522 |
| | Liquid Medium | 6.31E-06 | 1.38E-05 | 1.94E-05 | 3.39E-05 | 5.72E-05 | 0.0015 | 0.045 | 0.075 | 0.123 | 0.166 | 0.268 |
| | Spot Formation | 0.0039 | 0.0053 | 0.0062 | 0.0081 | 0.010 | 0.062 | 0.217 | 0.271 | 0.344 | 0.386 | 0.473 |
| | Yeast Localization | 0.0065 | 0.0068 | 0.0069 | 0.0088 | 0.010 | 0.063 | 0.341 | 0.462 | 0.613 | 0.694 | 0.785 |
| | siRNA | 3.3E-06 | 6.43E-06 | 8.73E-06 | 1.34E-05 | 1.83E-05 | 0.000398 | 0.025 | 0.030 | 0.088 | 0.121 | 0.252 |

Values obtained after sampling, with replacement, using the original data obtained from the four functional assays and the siRNA screen (see the bootstrap procedure A in the method section). Q indicates the quantile (Q0.050 is the quantile 5%). For the standard with reference method, the results shown represent values either relative to the median of the WT reference (BRCA1 reference or No siRNA), which allows an immediate comparison of the cut-off distributions between the different assays (top), or the same values multiplied by the experimental median of the WT reference observed in the corresponding assay (bottom). For example, 0.960 is the quantile 0.1% in the Colony Size assay, which corresponds to $11,200 \times 0.960 = 10,754$ cells per colony. Experimental medians used are 11,200 cells per colony (Colony Size assay), 0.144 OD600 (Liquid Medium assay), 32% (Spot Formation assay), 3% (Yeast Localization assay) and 945 (siRNA screen), as indicated in S4 Table. Of note, in the standard method, the WT reference was ignored during the sampling. Thus, the resulting fluctuation of the best cut-offs does not depend on the fluctuation of the WT reference. This explains why distributions were narrowed in the standard method, compared to the standard with reference method. For example, 12,133 cells per colony is the quantile 0.1% in the Colony Size assay, but in the standard with reference method, the same quantile is slightly farther from the distribution median, with a value of $11,200 \times 0.960 = 10,754$ cells per colony.

S6 Table. Quantitative analysis of Fig 2B

| Assay | Method | Number of variants correctly classified | | | | | Number of variants not classified | | | Number of variants misclassified | | | Balance | Sensitivity | Specificity | Accuracy |
|--------------------|-------------------------|---|---------------------------|--------------------|------------------------|-----------------------------|-----------------------------------|--------------------|-------|----------------------------------|-------------------------|-------|---------|-------------|-------------|----------|
| | | Pathogenic in class 5 | Pathogenic in class 5 + 4 | Neutral in class 1 | Neutral in class 1 + 2 | Total (class 1 + 2 + 4 + 5) | Pathogenic in class 3 | Neutral in class 3 | Total | Pathogenic in class 1 or 2 | Neutral in class 5 or 4 | Total | | | | |
| Colony Size | Standard | 23 | 23 | 12 | 13 | 36 | 2 | 1 | 3 | 0 | 1 | 1 | 35 | 0.92 | 0.87 | 0.90 |
| | Standard with reference | 22 | 23 | 8 | 13 | 36 | 2 | 1 | 3 | 0 | 1 | 1 | 35 | 0.92 | 0.87 | 0.90 |
| | MWW | 22 | 24 | 9 | 13 | 37 | 0 | 1 | 1 | 1 | 1 | 2 | 35 | 0.96 | 0.87 | 0.93 |
| Liquid Medium | Standard | 19 | 20 | 8 | 12 | 32 | 4 | 2 | 6 | 1 | 1 | 2 | 30 | 0.80 | 0.80 | 0.80 |
| | Standard with reference | 18 | 19 | 8 | 10 | 29 | 5 | 4 | 9 | 1 | 1 | 2 | 27 | 0.76 | 0.67 | 0.73 |
| | MWW | 15 | 16 | 7 | 12 | 28 | 8 | 3 | 11 | 1 | 0 | 1 | 27 | 0.64 | 0.80 | 0.70 |
| Spot Formation | Standard | 15 | 16 | 13 | 14 | 30 | 5 | 1 | 6 | 4 | 0 | 4 | 26 | 0.64 | 0.93 | 0.75 |
| | Standard with reference | 15 | 16 | 9 | 14 | 30 | 5 | 1 | 6 | 4 | 0 | 4 | 26 | 0.64 | 0.93 | 0.75 |
| | MWW | 18 | 19 | 10 | 14 | 33 | 2 | 0 | 2 | 4 | 1 | 5 | 28 | 0.76 | 0.93 | 0.83 |
| Yeast Localization | Standard | 0 | 16 | 7 | 13 | 29 | 3 | 2 | 5 | 6 | 0 | 6 | 23 | 0.64 | 0.87 | 0.73 |
| | Standard with reference | 14 | 15 | 7 | 13 | 28 | 4 | 2 | 6 | 6 | 0 | 6 | 22 | 0.60 | 0.87 | 0.70 |
| | MWW | 15 | 16 | 9 | 13 | 29 | 3 | 2 | 5 | 6 | 0 | 6 | 23 | 0.64 | 0.87 | 0.73 |

"Balance" indicates the number of mutations correctly classified (class 1 + 2 + 4 + 5) minus the total number of mutations misclassified. "Sensitivity" and "Specificity" represent the number of variants correctly classified, divided by the number of variants in the prior classification ($n_{\text{pathogenic}} = 25$ and $n_{\text{neutral}} = 15$). "Accuracy" is the number of mutations correctly classified (class 1 + 2 + 4 + 5) divided by the total number of mutations ($n = 40$).

S7 Table. Features of the assays

| Assay | Features |
|--------------------|---|
| Colony Size | $n_{\text{BRCA1}} = 36$ and $n_{\text{Mutant}} = 9$ 3 mutant clones assessed in 3 independent experiments (9 values) High dynamic range of the measurement (Vector median / BRCA1 median > 100) A few ties in the BRCA1 distribution |
| Liquid Medium | $n_{\text{BRCA1}} = 36$ and $n_{\text{Mutant}} = 9$ 3 mutant clones assessed in 3 independent experiments (9 values) Low dynamic range of the measurement (Vector median / BRCA1 median < 4) Many ties in the BRCA1 distribution and a few ties in mutant distributions Weak precision of the measurement (3 digits max given by the spectrophotometer) |
| Spot Formation | $n_{\text{BRCA1}} = 12$ and $n_{\text{Mutant}} = 3$ 3 mutant clones assessed independently (3 values) Proportion (measurement between 0 and 1) No ties Opposite orientation (pathogenic mutations below the WT BRCA1 reference) |
| Yeast Localization | $n_{\text{BRCA1}} = 12$ and $n_{\text{Mutant}} = 3$ 3 mutant clones assessed independently (3 values) Proportion (measurement between 0 and 1) BRCA1 reference median close to 0 Almost no ties |
| siRNA | $n_{\text{reference}} = 1,140$ and $n_{\text{siRNA}} = 12$ Quadriplicates in plates and 3 independent experiments (12 values) One positive (pathogenic) and 2 negative (neutral) controls Many ties in the "No siRNA" reference Opposite orientation (KIF11 positive control below the "No siRNA" reference) |

See S4 Table for details.

S8 Table. Examples of theoretical situations analyzed

| | Reference situation | | | Pathogenic distributions shifted (Figure S13) | | | Neutral distributions shifted (Figure S14) | | | WT BRCA1 distributions shifted (Figure S15) | | | Number of values modified (Figure S18) | | | Range modified (Figure S19) | | |
|--------------------------------|---------------------|--------------------|-----------------------|--|--------------------|-----------------------|---|--------------------|-----------------------|--|--------------------|-----------------------|---|--------------------|-----------------------|--------------------------------|--------------------|-----------------------|
| | BRCA1 | neutral mutants | pathogenic mutants | BRCA1 | neutral mutants | pathogenic mutants | BRCA1 | neutral mutants | pathogenic mutants | BRCA1 | neutral mutants | pathogenic mutants | BRCA1 | neutral mutants | pathogenic mutants | BRCA1 | neutral mutants | pathogenic mutants |
| 1 | 1 | 73 | 1 | 1 | 1 | 1 | 1 | 37 | 73 | 37 | 1 | 73 | 1 | 1 | 73 | 9.75 | 9.75 | 81.75 |
| 2 | 5.38 | 77.38 | 2 | 5.38 | 5.38 | 2 | 41.38 | 77.38 | 38 | 5.38 | 77.38 | 18.5 | 18.5 | 90.5 | 10.25 | 11.9375 | 83.9375 | |
| 3 | 9.75 | 81.75 | 3 | 9.75 | 9.75 | 3 | 45.75 | 81.75 | 39 | 9.75 | 81.75 | 36 | 36 | 108 | 10.75 | 14.125 | 86.125 | |
| 4 | 14.13 | 86.13 | 4 | 14.13 | 14.13 | 4 | 50.13 | 86.13 | 40 | 14.13 | 86.13 | | | | 11.25 | 16.3125 | 88.3125 | |
| 5 | 18.50 | 90.50 | 5 | 18.50 | 18.50 | 5 | 54.50 | 90.50 | 41 | 18.50 | 90.50 | | | | 11.75 | 18.5 | 90.5 | |
| 6 | 22.88 | 94.88 | 6 | 22.88 | 22.88 | 6 | 58.88 | 94.88 | 42 | 22.88 | 94.88 | | | | 12.25 | 20.6875 | 92.6875 | |
| 7 | 27.25 | 99.25 | 7 | 27.25 | 27.25 | 7 | 63.25 | 99.25 | 43 | 27.25 | 99.25 | | | | 12.75 | 22.875 | 94.875 | |
| 8 | 31.63 | 103.63 | 8 | 31.63 | 31.63 | 8 | 67.63 | 103.63 | 44 | 31.63 | 103.63 | | | | 13.25 | 25.0625 | 97.0625 | |
| 9 | 36 | 108 | 9 | 36 | 36 | 9 | 72 | 108 | 45 | 36 | 108 | | | | 13.75 | 27.25 | 99.25 | |
| 10 | | | 10 | | | 10 | | | 46 | | | | | | 14.25 | | | |
| 11 | | | 11 | | | 11 | | | 47 | | | | | | 14.75 | | | |
| 12 | | | 12 | | | 12 | | | 48 | | | | | | 15.25 | | | |
| 13 | | | 13 | | | 13 | | | 49 | | | | | | 15.75 | | | |
| 14 | | | 14 | | | 14 | | | 50 | | | | | | 16.25 | | | |
| 15 | | | 15 | | | 15 | | | 51 | | | | | | 16.75 | | | |
| 16 | | | 16 | | | 16 | | | 52 | | | | | | 17.25 | | | |
| 17 | | | 17 | | | 17 | | | 53 | | | | | | 17.75 | | | |
| 18 | | | 18 | | | 18 | | | 54 | | | | | | 18.25 | | | |
| 19 | | | 19 | | | 19 | | | 55 | | | | | | 18.75 | | | |
| 20 | | | 20 | | | 20 | | | 56 | | | | | | 19.25 | | | |
| 21 | | | 21 | | | 21 | | | 57 | | | | | | 19.75 | | | |
| 22 | | | 22 | | | 22 | | | 58 | | | | | | 20.25 | | | |
| 23 | | | 23 | | | 23 | | | 59 | | | | | | 20.75 | | | |
| 24 | | | 24 | | | 24 | | | 60 | | | | | | 21.25 | | | |
| 25 | | | 25 | | | 25 | | | 61 | | | | | | 21.75 | | | |
| 26 | | | 26 | | | 26 | | | 62 | | | | | | 22.25 | | | |
| 27 | | | 27 | | | 27 | | | 63 | | | | | | 22.75 | | | |
| 28 | | | 28 | | | 28 | | | 64 | | | | | | 23.25 | | | |
| 29 | | | 29 | | | 29 | | | 65 | | | | | | 23.75 | | | |
| 30 | | | 30 | | | 30 | | | 66 | | | | | | 24.25 | | | |
| 31 | | | 31 | | | 31 | | | 67 | | | | | | 24.75 | | | |
| 32 | | | 32 | | | 32 | | | 68 | | | | | | 25.25 | | | |
| 33 | | | 33 | | | 33 | | | 69 | | | | | | 25.75 | | | |
| 34 | | | 34 | | | 34 | | | 70 | | | | | | 26.25 | | | |
| 35 | | | 35 | | | 35 | | | 71 | | | | | | 26.75 | | | |
| 36 | | | 36 | | | 36 | | | 72 | | | | | | 27.25 | | | |
| n | 36 | 9 | 9 | 36 | 9 | 9 | 36 | 9 | 9 | 36 | 9 | 9 | 3 | 3 | 3 | 36 | 9 | 9 |
| Median | 18.5 | 18.5 | 90.5 | 18.5 | 18.5 | 18.5 | 18.5 | 54.5 | 90.5 | 54.5 | 18.5 | 90.5 | 18.5 | 18.5 | 90.5 | 18.5 | 18.5 | 90.5 |
| Mean | 18.5 | 18.5 | 90.5 | 18.5 | 18.5 | 18.5 | 18.5 | 54.5 | 90.5 | 54.5 | 18.5 | 90.5 | 18.5 | 18.5 | 90.5 | 18.5 | 18.5 | 90.5 |
| Separating factor s BRCA1 | 0 | | | 0 | | | 0 | | | 1 | | | 0 | | | 0 | | |
| Separating factor s neutral | | 0 | | | 0 | | | 1 | | | 0 | | | 0 | | | 0 | |
| Separating factor s pathogenic | | | 2 | | | 0 | | | 2 | | | 2 | | | 2 | | | 2 |
| Min value | 1 | 1 | 73 | 1 | 1 | 1 | 1 | 37 | 73 | 37 | 1 | 73 | 1 | 1 | 73 | 9.75 | 9.75 | 81.75 |
| Max value | 36 | 36 | 108 | 36 | 36 | 36 | 36 | 72 | 108 | 72 | 36 | 108 | 36 | 36 | 108 | 27.25 | 27.25 | 99.25 |
| Range | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 35 | 17.5 | 17.5 | 17.5 |
| Range factor r | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 1 | 0.5 | 0.5 | 0.5 |

The "reference situation" column indicates the initial values and parameter settings. From this, distribution parameters were modified and the resulting variant classification was scrutinized. The separating factor s (BRCA1, neutral or pathogenic) shifts the values of the distribution, according to the formula $v_{ij} + 36 \times s$, with v_{ij} representing the value i of the distribution j (BRCA1, neutral or pathogenic). The range factor r modulates the extreme values of a distribution, according to the formula $me \pm 17.5 \times r$, where me represents the median of the distribution. Whatever the value of r , the values of any distribution are equally spread.

S9 Table. Effect of functional assay parameters upon the probability system of classification

| Parameters | Standard method | Standard with reference method | MWW method | Figure |
|----------------------------------|--|--|---|--------|
| Position of pathogenic mutations | +++ Shift from the neutral sector towards the pathogenic sector increases both sensitivity and specificity. | +++ As in the standard method. | +++ As in the standard method. | S13 |
| Position of neutral mutations | +++ Shift from the neutral sector towards the pathogenic sector decreases both sensitivity and specificity. | +++ As in the standard method. | +++ No effect, except a complete <u>misclassification</u> of the pathogenic mutations when the WT reference distribution is left on one side of all the neutral and pathogenic distributions with no overlap. | S14 |
| Position of WT reference | 0 No effect. | +++ Shift of the median towards the null value decreases sensitivity and specificity. Negative values of the WT reference median invert the classification of the neutral and pathogenic mutations. | +++ No effect, except a complete <u>misclassification</u> of the pathogenic mutations when the WT reference distribution falls outside of the range of the neutral and pathogenic distributions, with no overlap. | S15 |
| Experimental sensitivity | +++ Decrease of the experimental sensitivity decreases specificity. This is advantageous, as the neutral region is contaminated by pathogenic mutations (presence of false negatives). | +++ As in the standard method. | +++ As in the standard method. | S16 |
| Experimental specificity | +++ Decrease of the experimental specificity decreases sensitivity. This is advantageous, as the pathogenic region is contaminated by neutral mutations (presence of false positives). | +++ As in the standard method. | +++ As in the standard method. | S16 |
| Number of neutral mutations | 0 No clear effect. | 0 No clear effect. | 0 No effect. | S17 |
| Number of pathogenic mutations | 0 No clear effect. | 0 No clear effect. | 0 No effect. | S17 |
| Number of mutant values | +++ Decreasing number decreases both sensitivity and specificity ($p = 0.5$ when $n_{\text{mutant}} = 1$). | +++ Decreasing number decreases both sensitivity and specificity, but only when the number of BRCA1 values concomitantly decreases ($p = 0.5$ when $n_{\text{mutant}} = n_{\text{BRCA1}} = 1$). Otherwise, no effect. | +++ As in the standard with reference method. | S18 |
| Number of WT BRCA1 values | 0 No effect. | +++ Loss of sensitivity with the decrease of the number of BRCA1 values, except when $n_{\text{BRCA1}} = 1$. The concomitant decrease of n_{mutant} and n_{BRCA1} decreases both sensitivity and specificity ($p = 0.5$ when $n_{\text{mutant}} = n_{\text{BRCA1}} = 1$). | +++ As in the standard with reference method. | S18 |
| Range of mutant distributions | +++ Decreasing range of the mutant distributions decreases both sensitivity and specificity ($p = 0.5$ when the range is null). | +++ No effect, except a strong decrease of sensitivity and specificity ($p = 0.5$) when the mutant and the BRCA1 ranges are null. | +++ As in the standard with reference method. | S19 |
| Range of WT BRCA1 distribution | 0 No effect. | +++ No effect, except a strong decrease of sensitivity and specificity ($p = 0.5$) when the mutant and the BRCA1 ranges are null. | +++ As in the standard with reference method. | S19 |

Sensitivity and specificity are defined here as the probabilities of pathogenicity (p) attributed to the pathogenic and neutral mutations, respectively, which reflects the accuracy of the probability system of variant classification. Sensitivity is maximal if $p \geq 0.95$ for the pathogenic mutations. Specificity is maximal if $p < 0.05$ for the neutral mutations. Misclassification is defined as $p \geq 0.95$ for the neutral mutations and $p < 0.05$ for the pathogenic mutations. The experimental sensitivity and specificity derives from the experimental best cut-off, as explained in S2 Fig. They differ from the sensitivity and specificity described above, since they are related to the initial position of the pathogenic and neutral mutants before random sampling. +++, very influenced; 0, no effect.

S10 Table. Plasmids used

| Name | cDNA | HGVS nomenclature ^a | Source ^b | Name | cDNA | HGVS nomenclature ^a | Source ^b |
|--------|---------------|--------------------------------|---------------------|--------|----------------|--------------------------------|---------------------|
| pJL48 | None (Vector) | | Millot et al., 2011 | | | | |
| pJL45 | BRCA1 | | Millot et al., 2011 | pGM40 | BRCA1-mCherry | | Millot et al., 2011 |
| pPT60 | BRCA1 | | This study | pPT63 | BRCA1-mCherry | | This study |
| pPT185 | M18T | c.53T>C | This study | pPT195 | M18T-mCherry | c.53T>C | This study |
| pPT175 | L22S | c.65T>C | This study | pPT180 | L22S-mCherry | c.65T>C | This study |
| pPT98 | I31M | c.93C>G | This study | pPT99 | I31M-mCherry | c.93C>G | This study |
| pPT176 | T37K | c.110C>A | This study | pPT181 | T37K-mCherry | c.110C>A | This study |
| pPT177 | C39R | c.115T>C | This study | pPT182 | C39R-mCherry | c.115T>C | This study |
| pPT142 | C39Y | c.116G>A | This study | pPT156 | C39Y-mCherry | c.116G>A | This study |
| pPT101 | C44F | c.131G>T | This study | pPT102 | C44F-mCherry | c.131G>T | This study |
| pPT186 | C44Y | c.131G>A | This study | pPT196 | C44Y-mCherry | c.131G>A | This study |
| pPT104 | K45Q | c.133A>C | This study | pPT105 | K45Q-mCherry | c.133A>C | This study |
| pPT143 | C47F | c.140G>T | This study | pPT157 | C47F-mCherry | c.140G>T | This study |
| pPT107 | C47G | c.139T>G | This study | pPT108 | C47G-mCherry | c.139T>G | This study |
| pPT187 | C61G | c.181T>G | This study | pPT197 | C61G-mCherry | c.181T>G | This study |
| pPT110 | C64Y | c.191G>A | This study | pPT111 | C64Y-mCherry | c.191G>A | This study |
| pPT178 | D67Y | c.199G>T | This study | pPT183 | D67Y-mCherry | c.199G>T | This study |
| pPT113 | N132K | c.396C>A | This study | pPT114 | N132K-mCherry | c.396C>A | This study |
| pPT188 | M1652I | c.4956G>A | This study | pPT198 | M1652I-mCherry | c.4956G>A | This study |
| pPT189 | M1652T | c.4955T>C | This study | pPT199 | M1652T-mCherry | c.4955T>C | This study |
| pPT144 | F1662S | c.4985T>C | This study | pPT158 | F1662S-mCherry | c.4985T>C | This study |
| pPT145 | A1669S | c.5005G>T | This study | pPT159 | A1669S-mCherry | c.5005G>T | This study |
| pPT190 | E1682K | c.5044G>A | This study | pPT200 | E1682K-mCherry | c.5044G>A | This study |
| pPT191 | T1685A | c.5053A>G | This study | pPT201 | T1685A-mCherry | c.5053A>G | This study |
| pPT146 | T1685I | c.5054C>T | This study | pPT160 | T1685I-mCherry | c.5054C>T | This study |
| pPT152 | M1689R | c.5066T>G | This study | pPT166 | M1689R-mCherry | c.5066T>G | This study |
| pPT116 | R1699W | c.5095C>T | This study | pPT117 | R1699W-mCherry | c.5095C>T | This study |
| pPT147 | G1706E | c.5117G>A | This study | pPT161 | G1706E-mCherry | c.5117G>A | This study |
| pPT192 | A1708E | c.5123C>A | This study | pPT202 | A1708E-mCherry | c.5123C>A | This study |
| pPT153 | S1715R | c.5143A>C | This study | pPT167 | S1715R-mCherry | c.5143A>C | This study |
| pPT193 | T1720A | c.5158A>G | This study | pPT203 | T1720A-mCherry | c.5158A>G | This study |
| pPT148 | G1738R | c.5212G>A | This study | pPT162 | G1738R-mCherry | c.5212G>A | This study |
| pPT119 | R1751Q | c.5252G>A | This study | pPT120 | R1751Q-mCherry | c.5252G>A | This study |
| pPT194 | L1764P | c.5291T>C | This study | pPT204 | L1764P-mCherry | c.5291T>C | This study |
| pPT149 | I1766S | c.5297T>G | This study | pPT163 | I1766S-mCherry | c.5297T>G | This study |
| pPT150 | M1775K | c.5324T>A | This study | pPT164 | M1775K-mCherry | c.5324T>A | This study |
| pPT155 | M1775R | c.5324T>G | This study | pPT169 | M1775R-mCherry | c.5324T>G | This study |
| pPT179 | P1776H | c.5327C>A | This study | pPT184 | P1776H-mCherry | c.5327C>A | This study |
| pPT151 | G1788V | c.5363G>T | This study | pPT165 | G1788V-mCherry | c.5363G>T | This study |
| pPT122 | V1804D | c.5411T>A | This study | pPT123 | V1804D-mCherry | c.5411T>A | This study |
| pPT222 | V1838E | c.5513T>A | This study | pPT226 | V1838E-mCherry | c.5513T>A | This study |
| pPT223 | I1858L | c.5572A>C | This study | pPT227 | I1858L-mCherry | c.5572A>C | This study |
| pPT125 | P1859R | c.5576C>G | This study | pPT127 | P1859R-mCherry | c.5576C>G | This study |

^a Mutation and deletion nomenclature according to the human genome variation society (<http://www.hgvs.org/mutnomen/>).

^b The BRCA1 cDNA used in Millot et al., 2011, corresponds to the AY888184.1 GenBank sequence, slightly modified in this study (see the Methods section).

S11 Table. Yeast strains used

| Strain designation | Strain common name ^a | Genotype | Source | Strain designation | Strain common name ^a | Genotype | Source |
|--------------------|---------------------------------|--|-----------------------|--------------------|---------------------------------|---|------------------|
| BY4741 | | <i>MATa, leu2Δ0, his3Δ1, ura3Δ0, met15Δ0</i> | Winzeler et al., 1999 | YKR082W-GFP | | <i>MATa, leu2Δ0, his3Δ1, ura3Δ0, met15Δ0, nup133::NUP133-GFP(HIS3MX6)</i> | Huh et al., 2003 |
| ORT5039 | Vector | BY4741 + pJL48 | Millot et al., 2011 | | | | |
| GMT335 | BRCA1 | BY4741 + pPT60 | This study | GMT336 | BRCA1-mCherry | YKR082W-GFP + pPT63 | This study |
| GMT424 | M18T | BY4741 + pPT185 | This study | GMT435 | M18T-mCherry | YKR082W-GFP + pPT195 | This study |
| GMT483 | L22S | BY4741 + pPT175 | This study | GMT502 | L22S-mCherry | YKR082W-GFP + pPT180 | This study |
| GMT377 | I31M | BY4741 + pPT98 | This study | GMT409 | I31M-mCherry | YKR082W-GFP + pPT99 | This study |
| GMT484 | T37K | BY4741 + pPT176 | This study | GMT503 | T37K-mCherry | YKR082W-GFP + pPT181 | This study |
| GMT485 | C39R | BY4741 + pPT177 | This study | GMT504 | C39R-mCherry | YKR082W-GFP + pPT182 | This study |
| GMT486 | C39Y | BY4741 + pPT142 | This study | GMT505 | C39Y-mCherry | YKR082W-GFP + pPT156 | This study |
| GMT378 | C44F | BY4741 + pPT101 | This study | GMT410 | C44F-mCherry | YKR082W-GFP + pPT102 | This study |
| GMT425 | C44Y | BY4741 + pPT186 | This study | GMT436 | C44Y-mCherry | YKR082W-GFP + pPT196 | This study |
| GMT379 | K45Q | BY4741 + pPT104 | This study | GMT411 | K45Q-mCherry | YKR082W-GFP + pPT105 | This study |
| GMT487 | C47F | BY4741 + pPT143 | This study | GMT506 | C47F-mCherry | YKR082W-GFP + pPT157 | This study |
| GMT380 | C47G | BY4741 + pPT107 | This study | GMT438 | C47G-mCherry | YKR082W-GFP + pPT108 | This study |
| GMT426 | C61G | BY4741 + pPT187 | This study | GMT437 | C61G-mCherry | YKR082W-GFP + pPT197 | This study |
| GMT433 | C64Y | BY4741 + pPT110 | This study | GMT446 | C64Y-mCherry | YKR082W-GFP + pPT111 | This study |
| GMT488 | D67Y | BY4741 + pPT178 | This study | GMT507 | D67Y-mCherry | YKR082W-GFP + pPT183 | This study |
| GMT381 | N132K | BY4741 + pPT113 | This study | GMT423 | N132K-mCherry | YKR082W-GFP + pPT114 | This study |
| GMT427 | M1652I | BY4741 + pPT188 | This study | GMT439 | M1652I-mCherry | YKR082W-GFP + pPT198 | This study |
| GMT428 | M1652T | BY4741 + pPT189 | This study | GMT440 | M1652T-mCherry | YKR082W-GFP + pPT199 | This study |
| GMT463 | F1662S | BY4741 + pPT144 | This study | GMT473 | F1662S-mCherry | YKR082W-GFP + pPT158 | This study |
| GMT464 | A1669S | BY4741 + pPT145 | This study | GMT474 | A1669S-mCherry | YKR082W-GFP + pPT159 | This study |
| GMT429 | E1682K | BY4741 + pPT190 | This study | GMT441 | E1682K-mCherry | YKR082W-GFP + pPT200 | This study |
| GMT430 | T1685A | BY4741 + pPT191 | This study | GMT442 | T1685A-mCherry | YKR082W-GFP + pPT201 | This study |
| GMT489 | T1685I | BY4741 + pPT146 | This study | GMT508 | T1685I-mCherry | YKR082W-GFP + pPT160 | This study |
| GMT490 | M1689R | BY4741 + pPT152 | This study | GMT509 | M1689R-mCherry | YKR082W-GFP + pPT166 | This study |
| GMT382 | R1699W | BY4741 + pPT116 | This study | GMT413 | R1699W-mCherry | YKR082W-GFP + pPT117 | This study |
| GMT491 | G1706E | BY4741 + pPT147 | This study | GMT510 | G1706E-mCherry | YKR082W-GFP + pPT161 | This study |
| GMT431 | A1708E | BY4741 + pPT192 | This study | GMT443 | A1708E-mCherry | YKR082W-GFP + pPT202 | This study |
| GMT472 | S1715R | BY4741 + pPT153 | This study | GMT482 | S1715R-mCherry | YKR082W-GFP + pPT167 | This study |
| GMT432 | T1720A | BY4741 + pPT193 | This study | GMT444 | T1720A-mCherry | YKR082W-GFP + pPT203 | This study |
| GMT471 | G1738R | BY4741 + pPT148 | This study | GMT481 | G1738R-mCherry | YKR082W-GFP + pPT162 | This study |
| GMT383 | R1751Q | BY4741 + pPT119 | This study | GMT414 | R1751Q-mCherry | YKR082W-GFP + pPT120 | This study |
| GMT434 | L1764P | BY4741 + pPT194 | This study | GMT445 | L1764P-mCherry | YKR082W-GFP + pPT204 | This study |
| GMT466 | I1766S | BY4741 + pPT149 | This study | GMT476 | I1766S-mCherry | YKR082W-GFP + pPT163 | This study |
| GMT468 | M1775K | BY4741 + pPT150 | This study | GMT478 | M1775K-mCherry | YKR082W-GFP + pPT164 | This study |
| GMT467 | M1775R | BY4741 + pPT155 | This study | GMT477 | M1775R-mCherry | YKR082W-GFP + pPT169 | This study |
| GMT469 | P1776H | BY4741 + pPT179 | This study | GMT479 | P1776H-mCherry | YKR082W-GFP + pPT184 | This study |
| GMT470 | G1788V | BY4741 + pPT151 | This study | GMT480 | G1788V-mCherry | YKR082W-GFP + pPT165 | This study |
| GMT384 | V1804D | BY4741 + pPT122 | This study | GMT415 | V1804D-mCherry | YKR082W-GFP + pPT123 | This study |
| GMT493 | V1838E | BY4741 + pPT222 | This study | GMT498 | V1838E-mCherry | YKR082W-GFP + pPT226 | This study |
| GMT515 | I1858L | BY4741 + pPT223 | This study | GMT499 | I1858L-mCherry | YKR082W-GFP + pPT227 | This study |
| GMT385 | P1859R | BY4741 + pPT125 | This study | GMT416 | P1859R-mCherry | YKR082W-GFP + pPT127 | This study |

^a"mCherry" was omitted in the main text. For instance, the P1859R-mCherry strain was referred to as "P1859R".

S12 Table. Average CDF of the probability system of classification when using the standard method

| Rank of the best cut-off values | Colony Size assay | | Liquid Medium assay | | Spot Formation assay | | Yeast Localization assay | |
|---------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|
| | Best cut-off value (n = 529) | Probability of pathogenicity | Best cut-off value (n = 126) | Probability of pathogenicity | Best cut-off value (n = 190) | Probability of pathogenicity | Best cut-off value (n = 154) | Probability of pathogenicity |
| 1 | 11853.3333 | 0.00025 | 0.149500 | 0.00025 | 0.31943533 | 0.00025 | 0.00818150 | 0.00025 |
| 2 | 11930.0000 | 0.00075 | 0.150500 | 0.00075 | 0.31609788 | 0.00075 | 0.01210510 | 0.00075 |
| 3 | 12132.9060 | 0.00150 | 0.150750 | 0.00125 | 0.31536359 | 0.00150 | 0.01509030 | 0.00125 |
| 4 | 12150.0000 | 0.00225 | 0.151000 | 0.00200 | 0.31499829 | 0.00300 | 0.01880518 | 0.00175 |
| 5 | 12200.0000 | 0.00275 | 0.151250 | 0.00275 | 0.31468908 | 0.00775 | 0.01904816 | 0.00225 |
| 6 | 12216.0685 | 0.00325 | 0.151500 | 0.00325 | 0.31243640 | 0.01300 | 0.01944921 | 0.00300 |
| 7 | 12230.0000 | 0.00400 | 0.151750 | 0.00375 | 0.31232318 | 0.01475 | 0.01955043 | 0.00425 |
| 8 | 12253.2478 | 0.00475 | 0.152000 | 0.00425 | 0.31212019 | 0.01550 | 0.01966000 | 0.00550 |
| 9 | 12276.4955 | 0.00675 | 0.152500 | 0.00675 | 0.31203042 | 0.01825 | 0.02580891 | 0.00625 |
| 10 | 12454.1285 | 0.00875 | 0.153000 | 0.01050 | 0.31196558 | 0.02075 | 0.02785367 | 0.00700 |
| 11 | 12595.5815 | 0.00925 | 0.153500 | 0.01375 | 0.31190075 | 0.02175 | 0.02819869 | 0.00775 |
| 12 | 12632.9060 | 0.01000 | 0.153750 | 0.01575 | 0.30964807 | 0.02300 | 0.02861547 | 0.00825 |
| 13 | 12650.0000 | 0.01075 | 0.154000 | 0.01700 | 0.30924209 | 0.02450 | 0.02975333 | 0.00875 |
| 14 | 12675.0000 | 0.01125 | 0.154500 | 0.02000 | 0.30840599 | 0.02575 | 0.03074321 | 0.00925 |
| 15 | 12700.0000 | 0.01175 | 0.155000 | 0.02275 | 0.30756988 | 0.02750 | 0.03399520 | 0.00975 |
| 16 | 12730.0000 | 0.01275 | 0.155250 | 0.02375 | 0.30749347 | 0.02925 | 0.03411158 | 0.01175 |
| 17 | 12753.2478 | 0.01400 | 0.155500 | 0.02550 | 0.30669004 | 0.03150 | 0.03412337 | 0.01425 |
| 18 | 12776.4955 | 0.01575 | 0.156000 | 0.02725 | 0.30497331 | 0.03375 | 0.03413517 | 0.01900 |
| 19 | 12825.0000 | 0.01725 | 0.156500 | 0.02800 | 0.30460801 | 0.03475 | 0.03454015 | 0.02350 |
| 20 | 12860.0428 | 0.01775 | 0.156750 | 0.02925 | 0.30429879 | 0.03650 | 0.03494513 | 0.02425 |
| 21 | 12920.0855 | 0.01925 | 0.157250 | 0.03025 | 0.30387171 | 0.03775 | 0.03495693 | 0.02475 |
| 22 | 12944.6758 | 0.02075 | 0.157500 | 0.03275 | 0.30326198 | 0.03825 | 0.03496873 | 0.02650 |
| 23 | 12969.2660 | 0.02475 | 0.157750 | 0.03525 | 0.30307933 | 0.03875 | 0.03590079 | 0.02850 |
| 24 | 13016.4705 | 0.02875 | 0.158000 | 0.03825 | 0.30292472 | 0.03925 | 0.03592438 | 0.03025 |
| 25 | 13063.6750 | 0.03200 | 0.158500 | 0.05500 | 0.30277011 | 0.04025 | 0.03788025 | 0.03175 |
| 26 | 13104.4790 | 0.03525 | 0.158750 | 0.06925 | 0.30204612 | 0.04175 | 0.04597766 | 0.03225 |
| 27 | 13145.2830 | 0.03925 | 0.159000 | 0.07225 | 0.30164013 | 0.04350 | 0.04609403 | 0.03275 |
| 28 | 13160.0268 | 0.04325 | 0.159500 | 0.10725 | 0.30051743 | 0.04500 | 0.04610583 | 0.03325 |
| 29 | 13172.6415 | 0.04375 | 0.159750 | 0.13975 | 0.30011145 | 0.04650 | 0.04611762 | 0.03425 |
| 30 | 13187.3853 | 0.04450 | 0.160000 | 0.16775 | 0.29996793 | 0.04900 | 0.04929481 | 0.03625 |
| 31 | 13200.0000 | 0.05225 | 0.160500 | 0.21125 | 0.29908809 | 0.05300 | 0.04945828 | 0.03775 |
| 32 | 13225.0000 | 0.05975 | 0.160750 | 0.22775 | 0.29845997 | 0.05575 | 0.05312493 | 0.03850 |
| 33 | 13255.1280 | 0.06750 | 0.161000 | 0.24475 | 0.29843925 | 0.05750 | 0.05361440 | 0.04575 |
| 34 | 13292.4115 | 0.07525 | 0.161500 | 0.28975 | 0.29817148 | 0.05975 | 0.05489660 | 0.05275 |
| 35 | 13544.1133 | 0.07575 | 0.161750 | 0.31875 | 0.29786544 | 0.06075 | 0.05617881 | 0.05700 |
| 36 | 13641.4648 | 0.07625 | 0.162000 | 0.32550 | 0.29755940 | 0.06275 | 0.05656693 | 0.06200 |
| 37 | 13666.0550 | 0.07675 | 0.162375 | 0.33225 | 0.29728122 | 0.06550 | 0.05695506 | 0.06475 |
| 38 | 13760.4640 | 0.07725 | 0.162500 | 0.36150 | 0.29700304 | 0.06875 | 0.05713629 | 0.06675 |
| 39 | 13842.0720 | 0.07775 | 0.162625 | 0.39100 | 0.29648933 | 0.09300 | 0.05731751 | 0.09300 |
| 40 | 13884.1743 | 0.07825 | 0.163000 | 0.40375 | 0.29620407 | 0.11550 | 0.05823459 | 0.11975 |
| 41 | 13896.7890 | 0.07875 | 0.163250 | 0.41675 | 0.29571581 | 0.11625 | 0.05884794 | 0.12075 |
| 42 | 13924.3530 | 0.07925 | 0.163500 | 0.42050 | 0.29569658 | 0.11675 | 0.05933556 | 0.12125 |
| 43 | 13951.9170 | 0.07975 | 0.163625 | 0.42375 | 0.29549359 | 0.11725 | 0.05951947 | 0.12200 |
| 44 | 13982.6863 | 0.08025 | 0.164000 | 0.43875 | 0.29547435 | 0.11950 | 0.06008882 | 0.12275 |
| 45 | 14013.4555 | 0.08125 | 0.164125 | 0.45375 | 0.29496065 | 0.14200 | 0.06065817 | 0.13275 |
| 46 | 14066.7890 | 0.08250 | 0.164250 | 0.45450 | 0.29444010 | 0.16300 | 0.06175647 | 0.14275 |
| 47 | 14120.1225 | 0.08325 | 0.164375 | 0.45550 | 0.29391954 | 0.17650 | 0.06285477 | 0.14375 |
| 48 | 14140.0613 | 0.08375 | 0.164500 | 0.46050 | 0.29388008 | 0.19000 | 0.06296903 | 0.14475 |
| 49 | 14196.7890 | 0.08425 | 0.164750 | 0.46525 | 0.29340069 | 0.19100 | 0.06362439 | 0.14550 |
| 50 | 14268.2663 | 0.08475 | 0.164875 | 0.46600 | 0.29296077 | 0.19350 | 0.06439401 | 0.15425 |
| 51 | 14286.9600 | 0.08550 | 0.165000 | 0.46825 | 0.29267582 | 0.19575 | 0.06485439 | 0.16275 |
| 52 | 14290.1468 | 0.08625 | 0.165125 | 0.47025 | 0.29239086 | 0.20625 | 0.06549231 | 0.16400 |
| 53 | 14297.6085 | 0.08675 | 0.165250 | 0.47075 | 0.29163329 | 0.21700 | 0.06792915 | 0.16525 |
| 54 | 14334.8230 | 0.08725 | 0.165375 | 0.47125 | 0.29087572 | 0.21850 | 0.07283799 | 0.16600 |
| 55 | 14346.7890 | 0.08850 | 0.165500 | 0.47300 | 0.29061887 | 0.21975 | 0.07578785 | 0.16725 |
| 56 | 14364.8475 | 0.08975 | 0.165750 | 0.47475 | 0.29036202 | 0.23200 | 0.08313352 | 0.16875 |
| 57 | 14382.9060 | 0.09025 | 0.165875 | 0.47525 | 0.28933620 | 0.24425 | 0.08649566 | 0.16975 |
| 58 | 14393.1620 | 0.09075 | 0.166000 | 0.47600 | 0.28831037 | 0.26450 | 0.08685300 | 0.17025 |
| 59 | 14431.6668 | 0.09125 | 0.166250 | 0.47675 | 0.28805130 | 0.28575 | 0.08872092 | 0.17125 |
| 60 | 14450.0000 | 0.09175 | 0.166500 | 0.48150 | 0.28779223 | 0.29300 | 0.09054727 | 0.17250 |
| 61 | 14465.0000 | 0.09225 | 0.166625 | 0.48625 | 0.28371174 | 0.31000 | 0.09098380 | 0.17325 |
| 62 | 14481.7095 | 0.09275 | 0.167000 | 0.48875 | 0.28354216 | 0.32125 | 0.09139695 | 0.17375 |
| 63 | 14521.3675 | 0.09325 | 0.167250 | 0.49150 | 0.28337259 | 0.32175 | 0.09254472 | 0.17525 |
| 64 | 14526.4955 | 0.09450 | 0.167500 | 0.50475 | 0.28311574 | 0.32250 | 0.09271581 | 0.17800 |
| 65 | 14532.5995 | 0.09575 | 0.167750 | 0.51800 | 0.28285889 | 0.32525 | 0.09382692 | 0.18000 |
| 66 | 14560.0428 | 0.09625 | 0.168000 | 0.52275 | 0.28243343 | 0.32825 | 0.09412014 | 0.18075 |
| 67 | 14585.9325 | 0.09700 | 0.168125 | 0.52725 | 0.28200797 | 0.34350 | 0.09523126 | 0.18125 |
| 68 | 14592.9663 | 0.09775 | 0.168250 | 0.52800 | 0.28129329 | 0.35825 | 0.09552447 | 0.19975 |
| 69 | 14593.0275 | 0.09825 | 0.168500 | 0.54275 | 0.28057862 | 0.35875 | 0.09580061 | 0.21975 |
| 70 | 14600.1225 | 0.09875 | 0.168750 | 0.55725 | 0.28043386 | 0.35925 | 0.09658185 | 0.22175 |

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| 71 | 14625.7493 | 0.09925 | 0.169000 | 0.56650 | 0.28028910 | 0.36025 | 0.09685799 | 0.22325 |
| 72 | 14641.7095 | 0.09975 | 0.169125 | 0.57575 | 0.28006492 | 0.36350 | 0.09763923 | 0.23700 |
| 73 | 14670.0855 | 0.10075 | 0.169500 | 0.58775 | 0.27923886 | 0.36625 | 0.09768395 | 0.24975 |
| 74 | 14675.2135 | 0.10175 | 0.169750 | 0.60000 | 0.27898201 | 0.36750 | 0.09872591 | 0.25025 |
| 75 | 14680.3415 | 0.10225 | 0.170000 | 0.60350 | 0.27823857 | 0.36900 | 0.09981258 | 0.25100 |
| 76 | 14708.6165 | 0.10275 | 0.170250 | 0.60675 | 0.27749513 | 0.37000 | 0.10062218 | 0.25200 |
| 77 | 14713.9413 | 0.10325 | 0.170500 | 0.61350 | 0.27695367 | 0.37150 | 0.10181003 | 0.25350 |
| 78 | 14719.2660 | 0.10500 | 0.170625 | 0.62025 | 0.27641222 | 0.37275 | 0.10380747 | 0.25650 |
| 79 | 14726.6668 | 0.10675 | 0.170750 | 0.62125 | 0.27631041 | 0.37325 | 0.10448754 | 0.25900 |
| 80 | 14761.9495 | 0.10825 | 0.171000 | 0.63175 | 0.27620861 | 0.37600 | 0.10521180 | 0.26050 |
| 81 | 14762.6415 | 0.10975 | 0.171125 | 0.64175 | 0.27599589 | 0.37875 | 0.10589188 | 0.26200 |
| 82 | 14763.3335 | 0.11075 | 0.171250 | 0.64225 | 0.27578317 | 0.37925 | 0.10661614 | 0.28700 |
| 83 | 14813.6750 | 0.11375 | 0.171375 | 0.64275 | 0.27538497 | 0.37975 | 0.10745969 | 0.31175 |
| 84 | 14815.1708 | 0.11650 | 0.171500 | 0.64650 | 0.27526947 | 0.38125 | 0.10873090 | 0.32625 |
| 85 | 14816.6665 | 0.11850 | 0.171750 | 0.65075 | 0.27524893 | 0.38275 | 0.10894670 | 0.34075 |
| 86 | 14835.5305 | 0.12075 | 0.172000 | 0.67150 | 0.27498677 | 0.38350 | 0.11035103 | 0.34150 |
| 87 | 14862.1970 | 0.12175 | 0.172250 | 0.69250 | 0.27474580 | 0.38450 | 0.11056684 | 0.34225 |
| 88 | 14871.7945 | 0.12375 | 0.172500 | 0.72775 | 0.27450483 | 0.39300 | 0.11197117 | 0.34750 |
| 89 | 14875.8973 | 0.12575 | 0.172750 | 0.76225 | 0.27395974 | 0.40275 | 0.11207404 | 0.35400 |
| 90 | 14880.0000 | 0.12625 | 0.173000 | 0.78300 | 0.27360226 | 0.40500 | 0.11217690 | 0.36950 |
| 91 | 14887.6415 | 0.12700 | 0.173250 | 0.80375 | 0.27305716 | 0.40575 | 0.11291090 | 0.38400 |
| 92 | 14895.2830 | 0.12925 | 0.173500 | 0.81100 | 0.27291581 | 0.40650 | 0.11364490 | 0.41075 |
| 93 | 14914.3080 | 0.13150 | 0.173750 | 0.81850 | 0.27287319 | 0.40800 | 0.11386541 | 0.43725 |
| 94 | 14933.3330 | 0.13350 | 0.174000 | 0.82675 | 0.27220528 | 0.40950 | 0.11397992 | 0.43775 |
| 95 | 14940.1708 | 0.13525 | 0.174250 | 0.83500 | 0.27207789 | 0.41025 | 0.11408593 | 0.44125 |
| 96 | 14944.9573 | 0.13575 | 0.174500 | 0.85925 | 0.27206392 | 0.41100 | 0.11441089 | 0.44475 |
| 97 | 14948.5043 | 0.13625 | 0.174750 | 0.88375 | 0.27202130 | 0.41250 | 0.11449111 | 0.44525 |
| 98 | 14950.0000 | 0.14050 | 0.175000 | 0.89650 | 0.27171087 | 0.41725 | 0.11824779 | 0.44575 |
| 99 | 14966.4530 | 0.14475 | 0.175250 | 0.90975 | 0.27132135 | 0.42150 | 0.12255691 | 0.44625 |
| 100 | 14993.3335 | 0.14525 | 0.175500 | 0.92025 | 0.27116941 | 0.42375 | 0.12326857 | 0.46550 |
| 101 | 15002.5640 | 0.14575 | 0.176000 | 0.94250 | 0.27099445 | 0.42575 | 0.12400257 | 0.48550 |
| 102 | 15005.1280 | 0.14950 | 0.176250 | 0.95625 | 0.27062795 | 0.42775 | 0.12468658 | 0.48675 |
| 103 | 15016.8723 | 0.15400 | 0.176500 | 0.96050 | 0.27059229 | 0.42975 | 0.12473657 | 0.51500 |
| 104 | 15028.6165 | 0.15575 | 0.177000 | 0.96500 | 0.26962357 | 0.43050 | 0.12668009 | 0.54325 |
| 105 | 15034.3083 | 0.15725 | 0.177250 | 0.96625 | 0.26959030 | 0.43125 | 0.12701921 | 0.54375 |
| 106 | 15049.0520 | 0.15825 | 0.177500 | 0.96900 | 0.26924716 | 0.43175 | 0.12862360 | 0.54925 |
| 107 | 15053.3333 | 0.16025 | 0.177750 | 0.97175 | 0.26861919 | 0.43275 | 0.12935760 | 0.55550 |
| 108 | 15058.3333 | 0.16225 | 0.178000 | 0.97350 | 0.26847784 | 0.43375 | 0.13009160 | 0.56100 |
| 109 | 15065.0000 | 0.16275 | 0.178250 | 0.97525 | 0.26782189 | 0.43425 | 0.13687872 | 0.56575 |
| 110 | 15075.0000 | 0.16425 | 0.178500 | 0.97600 | 0.26781239 | 0.43475 | 0.13694975 | 0.56650 |
| 111 | 15081.6668 | 0.16575 | 0.179000 | 0.97700 | 0.26767104 | 0.43550 | 0.14507017 | 0.56775 |
| 112 | 15083.3335 | 0.16725 | 0.179500 | 0.97775 | 0.26759604 | 0.43625 | 0.14514120 | 0.56925 |
| 113 | 15084.9145 | 0.16900 | 0.180000 | 0.97900 | 0.26700559 | 0.45450 | 0.14596459 | 0.57025 |
| 114 | 15121.5810 | 0.17050 | 0.180250 | 0.98025 | 0.26696050 | 0.47300 | 0.14647450 | 0.57225 |
| 115 | 15132.4785 | 0.17175 | 0.180500 | 0.98075 | 0.26691542 | 0.48125 | 0.14700911 | 0.57425 |
| 116 | 15135.8975 | 0.17225 | 0.180750 | 0.98125 | 0.26666416 | 0.48925 | 0.14753188 | 0.57600 |
| 117 | 15138.4615 | 0.17350 | 0.181500 | 0.98525 | 0.26663271 | 0.49050 | 0.14806649 | 0.57775 |
| 118 | 15145.9830 | 0.17475 | 0.182000 | 0.99050 | 0.26624197 | 0.49250 | 0.14858926 | 0.57875 |
| 119 | 15155.8975 | 0.17625 | 0.182500 | 0.99375 | 0.26585123 | 0.51775 | 0.14887353 | 0.57975 |
| 120 | 15163.4190 | 0.17775 | 0.183000 | 0.99625 | 0.26500313 | 0.54250 | 0.15146827 | 0.58125 |
| 121 | 15173.3335 | 0.17975 | 0.183500 | 0.99725 | 0.26477402 | 0.54325 | 0.15175254 | 0.58275 |
| 122 | 15186.6668 | 0.18250 | 0.183750 | 0.99775 | 0.26476247 | 0.54375 | 0.15294295 | 0.58375 |
| 123 | 15190.0855 | 0.18425 | 0.184000 | 0.99825 | 0.26462652 | 0.54425 | 0.15362046 | 0.58475 |
| 124 | 15203.4188 | 0.18650 | 0.187500 | 0.99875 | 0.26421162 | 0.55550 | 0.15434728 | 0.59025 |
| 125 | 15210.0855 | 0.18875 | 0.188500 | 0.99925 | 0.26367016 | 0.56725 | 0.15540466 | 0.59600 |
| 126 | 15218.4615 | 0.18975 | 0.189000 | 0.99975 | 0.26363054 | 0.56825 | 0.15646204 | 0.59975 |
| 127 | 15228.4188 | 0.19075 | | | 0.26360978 | 0.56875 | 0.15667528 | 0.60350 |
| 128 | 15230.7690 | 0.19175 | | | 0.26334458 | 0.56925 | 0.15807961 | 0.60525 |
| 129 | 15234.0170 | 0.19225 | | | 0.26312871 | 0.57625 | 0.15969718 | 0.60800 |
| 130 | 15240.0000 | 0.19300 | | | 0.26309299 | 0.58375 | 0.16075456 | 0.61175 |
| 131 | 15248.3760 | 0.19375 | | | 0.26305727 | 0.60175 | 0.16181194 | 0.61425 |
| 132 | 15258.3333 | 0.19450 | | | 0.26251581 | 0.62050 | 0.16246944 | 0.61600 |
| 133 | 15260.6835 | 0.19550 | | | 0.26238261 | 0.62225 | 0.16312693 | 0.61875 |
| 134 | 15266.6665 | 0.19700 | | | 0.26211771 | 0.62275 | 0.16386093 | 0.62025 |
| 135 | 15272.9663 | 0.19825 | | | 0.26197435 | 0.63125 | 0.16459493 | 0.62450 |
| 136 | 15279.2660 | 0.19900 | | | 0.26069526 | 0.64025 | 0.16779732 | 0.62975 |
| 137 | 15306.6670 | 0.20050 | | | 0.26035916 | 0.64125 | 0.17099971 | 0.63525 |
| 138 | 15323.4190 | 0.20200 | | | 0.25981457 | 0.64175 | 0.17173371 | 0.64000 |
| 139 | 15329.7185 | 0.20325 | | | 0.25941617 | 0.65225 | 0.17246771 | 0.64825 |
| 140 | 15340.1710 | 0.20650 | | | 0.25927482 | 0.66550 | 0.17440866 | 0.65650 |
| 141 | 15356.9230 | 0.20925 | | | 0.25919376 | 0.66875 | 0.17634962 | 0.65975 |
| 142 | 15373.6750 | 0.20975 | | | 0.25869764 | 0.67125 | 0.17708362 | 0.66300 |
| 143 | 15380.8545 | 0.21025 | | | 0.25826182 | 0.69025 | 0.17781761 | 0.66650 |
| 144 | 15388.0340 | 0.21075 | | | 0.25797911 | 0.71050 | 0.18102183 | 0.67225 |
| 145 | 15394.0170 | 0.21250 | | | 0.25728536 | 0.71450 | 0.18422605 | 0.69100 |
| 146 | 15400.0000 | 0.21700 | | | 0.25635407 | 0.71525 | 0.19209883 | 0.73825 |
| 147 | 15406.8375 | 0.22025 | | | 0.25587461 | 0.71575 | 0.19477378 | 0.77025 |
| 148 | 15427.6415 | 0.22175 | | | 0.25543726 | 0.71625 | 0.19744873 | 0.78675 |
| 149 | 15434.4790 | 0.22325 | | | 0.25460612 | 0.71700 | 0.19969382 | 0.80350 |

| | | | | | | |
|-----|------------|---------|------------|---------|------------|---------|
| 150 | 15455.2830 | 0.22375 | 0.24996419 | 0.71775 | 0.20193891 | 0.82850 |
| 151 | 15464.3938 | 0.22575 | 0.24971023 | 0.71850 | 0.20587530 | 0.85425 |
| 152 | 15473.5045 | 0.22775 | 0.24913305 | 0.71925 | 0.20981169 | 0.90225 |
| 153 | 15484.3938 | 0.22900 | 0.24449111 | 0.72575 | 0.21248665 | 0.94975 |
| 154 | 15491.7523 | 0.23025 | 0.24390135 | 0.73225 | 0.21516160 | 0.97550 |
| 155 | 15502.6415 | 0.23075 | 0.24331159 | 0.74475 | | |
| 156 | 15510.0000 | 0.23450 | 0.24307022 | 0.75725 | | |
| 157 | 15515.0000 | 0.23825 | 0.24282884 | 0.76650 | | |
| 158 | 15520.0000 | 0.23900 | 0.24198823 | 0.77775 | | |
| 159 | 15520.6838 | 0.23975 | 0.24181878 | 0.78050 | | |
| 160 | 15521.3675 | 0.24025 | 0.24164932 | 0.79275 | | |
| 161 | 15527.3505 | 0.24075 | 0.24122902 | 0.80550 | | |
| 162 | 15529.0520 | 0.24125 | 0.24080871 | 0.80900 | | |
| 163 | 15533.3335 | 0.24250 | 0.23871076 | 0.81175 | | |
| 164 | 15541.6668 | 0.24400 | 0.23661281 | 0.81625 | | |
| 165 | 15550.0000 | 0.24550 | 0.23555322 | 0.82150 | | |
| 166 | 15565.1280 | 0.24825 | 0.23549325 | 0.82325 | | |
| 167 | 15574.7008 | 0.25025 | 0.23543329 | 0.82750 | | |
| 168 | 15605.1280 | 0.25100 | 0.23490349 | 0.83150 | | |
| 169 | 15615.8973 | 0.25175 | 0.23437369 | 0.83425 | | |
| 170 | 15626.6665 | 0.25300 | 0.23291535 | 0.83700 | | |
| 171 | 15632.1365 | 0.25425 | 0.23145701 | 0.83850 | | |
| 172 | 15637.6065 | 0.25500 | 0.23086724 | 0.84025 | | |
| 173 | 15645.4700 | 0.25575 | 0.23027748 | 0.84175 | | |
| 174 | 15653.3335 | 0.25800 | 0.22968495 | 0.84400 | | |
| 175 | 15666.6665 | 0.26050 | 0.22909241 | 0.87275 | | |
| 176 | 15709.5953 | 0.26125 | 0.22803282 | 0.91350 | | |
| 177 | 15733.3335 | 0.26375 | 0.22760370 | 0.92750 | | |
| 178 | 15752.1368 | 0.26650 | 0.22717458 | 0.94300 | | |
| 179 | 15770.9400 | 0.26825 | 0.22611498 | 0.96550 | | |
| 180 | 15772.1368 | 0.26975 | 0.22393660 | 0.97975 | | |
| 181 | 15773.3335 | 0.27075 | 0.22380458 | 0.98675 | | |
| 182 | 15786.6670 | 0.27250 | 0.22367255 | 0.98725 | | |
| 183 | 15810.0000 | 0.27375 | 0.22274630 | 0.98825 | | |
| 184 | 15875.0855 | 0.27425 | 0.22201877 | 0.99300 | | |
| 185 | 15900.1710 | 0.27575 | 0.22175472 | 0.99725 | | |
| 186 | 15902.2223 | 0.27725 | 0.22129159 | 0.99775 | | |
| 187 | 15920.1710 | 0.27850 | 0.22082846 | 0.99825 | | |
| 188 | 15926.1538 | 0.28025 | 0.21805242 | 0.99875 | | |
| 189 | 15944.1025 | 0.28125 | 0.21777202 | 0.99925 | | |
| 190 | 15946.1540 | 0.28175 | 0.21613458 | 0.99975 | | |
| 191 | 15950.0855 | 0.28225 | | | | |
| 192 | 15954.0170 | 0.28300 | | | | |
| 193 | 15960.0000 | 0.28575 | | | | |
| 194 | 15994.0170 | 0.28825 | | | | |
| 195 | 16000.0000 | 0.28900 | | | | |
| ... | ... | ... | | | | |
| 468 | 24326.6665 | 0.94000 | | | | |
| 469 | 24401.6665 | 0.94175 | | | | |
| 470 | 24428.3180 | 0.94225 | | | | |
| 471 | 24594.7858 | 0.94300 | | | | |
| 472 | 24621.4373 | 0.94375 | | | | |
| 473 | 24712.9050 | 0.94575 | | | | |
| 474 | 24719.7858 | 0.94800 | | | | |
| 475 | 24726.6665 | 0.94925 | | | | |
| 476 | 24744.9845 | 0.95025 | | | | |
| 477 | 24796.6365 | 0.95150 | | | | |
| 478 | 24798.3183 | 0.95300 | | | | |
| 479 | 24800.0000 | 0.95425 | | | | |
| 480 | 24890.8255 | 0.95550 | | | | |
| 481 | 24920.4128 | 0.95625 | | | | |
| 482 | 24950.0000 | 0.95675 | | | | |
| 483 | 24975.2295 | 0.95800 | | | | |
| 484 | 25073.3335 | 0.95950 | | | | |
| 485 | 25186.2385 | 0.96050 | | | | |
| 486 | 25193.1193 | 0.96150 | | | | |
| 487 | 25238.5320 | 0.96225 | | | | |
| 488 | 25277.0640 | 0.96275 | | | | |
| 489 | 25283.9448 | 0.96325 | | | | |
| 490 | 25314.9083 | 0.96375 | | | | |
| 491 | 25459.5720 | 0.96425 | | | | |
| 492 | 25466.4528 | 0.96475 | | | | |
| 493 | 25489.9695 | 0.96525 | | | | |
| 494 | 25494.8318 | 0.96575 | | | | |
| 495 | 25499.6940 | 0.96650 | | | | |
| 496 | 25756.5750 | 0.96825 | | | | |
| 497 | 25764.2200 | 0.96975 | | | | |
| 498 | 25765.2905 | 0.97025 | | | | |
| 499 | 25766.3610 | 0.97200 | | | | |
| 500 | 25829.9695 | 0.97375 | | | | |
| 501 | 25893.5780 | 0.97475 | | | | |
| 502 | 25919.2660 | 0.97575 | | | | |
| 503 | 25944.9540 | 0.97800 | | | | |
| 504 | 25963.3030 | 0.98025 | | | | |
| 505 | 25993.2725 | 0.98075 | | | | |

| | | |
|-----|------------|---------|
| 506 | 26023.2420 | 0.98175 |
| 507 | 26038.6853 | 0.98275 |
| 508 | 26054.1285 | 0.98325 |
| 509 | 26150.4585 | 0.98375 |
| 510 | 26150.4590 | 0.98450 |
| 511 | 26176.1470 | 0.98525 |
| 512 | 26201.8350 | 0.98750 |
| 513 | 26219.2358 | 0.98975 |
| 514 | 26459.6940 | 0.99075 |
| 515 | 26574.3425 | 0.99175 |
| 516 | 26688.9910 | 0.99225 |
| 517 | 26716.5750 | 0.99325 |
| 518 | 26933.0275 | 0.99425 |
| 519 | 26978.4403 | 0.99475 |
| 520 | 27023.8530 | 0.99525 |
| 521 | 27139.4498 | 0.99575 |
| 522 | 27189.9085 | 0.99625 |
| 523 | 27198.1348 | 0.99675 |
| 524 | 27206.3610 | 0.99725 |
| 525 | 27280.7340 | 0.99775 |
| 526 | 27371.9880 | 0.99825 |
| 527 | 27463.2420 | 0.99875 |
| 528 | 28058.7158 | 0.99925 |
| 529 | 29286.5753 | 0.99975 |

See S6C Fig for details. The 272 middle values of the average CDF, derived from the Colony Size assay, were removed to simplify the table. Colored numbers indicate the 5-class nomenclature (S1 Table): dark blue, class 1; light blue, class 2; pink, class 4; red, class 5. This table was used to attribute the probabilities of pathogenicity depicted in Fig 2B. For instance, in the Colony Size assay, the median of P1776H is 12,253 cells per colony. The closest best cut-off value of this median belongs to rank 8, meaning that, when using the standard method, the probability of pathogenicity attributed to P1776H is 0.00475. Of note, if a best cut-off value is not repeated, among the 2,000 best cut-off values used to derive the CDF, then the associated probability is $1 / 2,000 = 0.0005$ for this value. This indicates the minimal probability incrementation between two non repeated consecutive best cut-off values in the CDF (probability unit). In the Yeast Localization assay, no variant can be classified as class 5 due to a lack of unrepeated values at the "pathogenic" side of the best cut-off distribution (the highest cut-off value, 0.2151616, is present 97 times, leading to a probability of 0.0485 and an average cumulative probability of 0.97550 in the CDF, which is inferior to the 0.99 probability threshold of the class 5).

S13 Table. Basic and corrected probability of pathogenicity when using the standard method

| Variant | Colony Size assay | | | | Liquid Medium Assay | | | | Spot Formation assay | | | | Yeast Localization assay | | | | Combined odds (CS x SF) | Combined probability (CS x SF) | Corrected odds combined (CS x SF) | Corrected probability combined (CS x SF) | | | | |
|-------------|-------------------|---------|----------------|-----------------------|---------------------|---------|----------------|-----------------------|----------------------|---------|----------------|-----------------------|--------------------------|---------|----------------|-----------------------|-------------------------|--------------------------------|-----------------------------------|--|--------------|--|--|--|
| | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | | | | | | | | |
| M1689R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| G1706E | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| A1708E | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| S1715R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| L1764P | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| I1766S | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| G1788V | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| V1838E | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| G1738R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| T1685I | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| T1685A | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| C44Y | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| C44F | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| C64Y | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.9905 | 104 | 99 | 0.99003 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | |
| T37K | 0.99925 | 1.332 | 1.269 | 0.99921 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.71925 | 2.56 | 2.49 | 0.71326 | 0.9755 | 40 | 38 | 0.97434 | 3.413 | 0.99971 | 3.156 | 0.99968 | | | | |
| C39R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.82675 | 4.77 | 4.59 | 0.82119 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.9755 | 40 | 38 | 0.97434 | 15.992001 | 1 | 14.505579 | 1 | | | | |
| C61G | 0.99525 | 210 | 200 | 0.99501 | 0.783 | 3.61 | 3.48 | 0.77699 | 0.97975 | 48 | 46 | 0.97878 | 0.56575 | 1.30 | 1.29 | 0.56301 | 10.137 | 0.99990 | 9.207 | 0.99989 | | | | |
| C47G | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99875 | 799 | 761 | 0.99869 | 0.943 | 16.54 | 15.80 | 0.94049 | 0.164 | 0.2 | 0.16942 | 66.159 | 0.99998 | 60.190 | 0.99998 | | | | | |
| M18T | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99825 | 570 | 543 | 0.99816 | 0.81175 | 4.31 | 4.15 | 0.80599 | 0.12075 | 0.14 | 0.14 | 0.12527 | 17.244 | 0.99994 | 15.822 | 0.99994 | | | | |
| C39Y | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.71925 | 2.56 | 2.49 | 0.71326 | 0.0055 | 0.0055 | 0.0058 | 0.00577 | 10.245 | 0.99990 | 9.474 | 0.99989 | | | | |
| M1775R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.965 | 28 | 26 | 0.96338 | 0.293 | 0.41 | 0.43 | 0.29889 | 0.00125 | 0.0013 | 0.0013 | 0.00131 | 1.657 | 0.99940 | 1.624 | 0.99938 | | | | |
| L22S | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1 | 0.50000 | 1 | 0.50000 | | | | |
| M1775K | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.6035 | 1.52 | 1.50 | 0.59955 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1 | 0.50000 | 1 | 0.50000 | | | | |
| C47F | 0.47 | 0.89 | 0.89 | 0.47135 | 0.24475 | 0.32 | 0.33 | 0.25085 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00075 | 0.00075 | 0.00079 | 0.00079 | 0.00022 | 0.00022 | 0.00023 | 0.00023 | | | | |
| R1699W | 0.07525 | 0.081 | 0.085 | 0.07842 | 0.017 | 0.017 | 0.018 | 0.01782 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00125 | 0.0013 | 0.0013 | 0.00131 | 2.03E-5 | 2.03E-5 | 2.23E-5 | 2.23E-5 | | | | |
| T1720A | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| P1859R | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| R1751Q | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| E1682K | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| M1652I | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00425 | 0.0043 | 0.0045 | 0.00446 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| I31M | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| I1858L | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00425 | 0.0043 | 0.0045 | 0.00446 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| P1776H | 0.00475 | 0.0048 | 0.005 | 0.00499 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.0055 | 0.0055 | 0.0058 | 0.00577 | 1.19E-6 | 1.19E-6 | 1.32E-6 | 1.32E-6 | | | | |
| F1662S | 0.00075 | 0.00075 | 0.00079 | 0.00079 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00975 | 0.0098 | 0.01 | 0.01023 | 1.88E-7 | 1.88E-7 | 2.07E-7 | 2.07E-7 | | | | |
| V1804D | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.03825 | 0.04 | 0.042 | 0.04001 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| A1669S | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.03825 | 0.04 | 0.042 | 0.04001 | 0.04175 | 0.044 | 0.046 | 0.04366 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1.09E-5 | 1.09E-5 | 1.20E-5 | 1.20E-5 | | | | |
| N132K | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00425 | 0.0043 | 0.0045 | 0.00446 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.057 | 0.06 | 0.063 | 0.05951 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| D67Y | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.24475 | 0.32 | 0.33 | 0.25085 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| M1652T | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00975 | 0.0098 | 0.01 | 0.01023 | 1 | 0.50000 | 1 | 0.50000 | | | | |
| K45Q | 0.47 | 0.89 | 0.89 | 0.47135 | 0.6715 | 2.04 | 1.99 | 0.66605 | 0.8385 | 5.19 | 4.99 | 0.83312 | 0.16725 | 0.2 | 0.21 | 0.17272 | 4.60 | 0.82156 | 4.45 | 0.81655 | | | | |
| Sensitivity | 0.92 (23/25) | | | | 0.92 (23/25) | | | | 0.80 (20/25) | | | | 0.64 (16/25) | | | | 0.64 (16/25) | | | | 0.84 (21/25) | | | |
| Specificity | 0.87 (13/15) | | | | 0.87 (13/15) | | | | 0.80 (12/15) | | | | 0.93 (14/15) | | | | 0.87 (13/15) | | | | 0.87 (13/15) | | | |
| Accuracy | 0.90 (36/40) | | | | 0.90 (36/40) | | | | 0.80 (32/40) | | | | 0.75 (30/40) | | | | 0.75 (30/40) | | | | 0.73 (29/40) | | | |

See Fig 2B for details. Values in the "Probability" columns are those in Fig 2B. Mutations are ordered as in Fig 2B to facilitate comparisons. Odds in favor of pathogenicity are the ratio $\pi_i / (1 - \pi_i)$, with π_i being the probability of pathogenicity of the variant i . The Liquid Medium and Yeast Localization assays were not included in the combined odds, since the Liquid Medium and Colony Size, as well as the Spot Formation and Yeast Localization, are not independent assays. Combined probabilities of pathogenicity result from the ratio $O_i / (1 + O_i)$, with O_i being the combined odds of the variant i . Probabilities were also corrected according to $n_{neutral} + n_{pathogenic} = 40$ and $a = 2$ (see the S1 Text). CS, Colony Size assay; SF, Spot Formation assay.

S14 Table. Basic and corrected probability of pathogenicity when using the standard with reference method

| Variant | Colony Size assay | | | | Liquid Medium Assay | | | | Spot Formation assay | | | | Yeast Localization assay | | | | Combined odds (CS x SF) | Combined probability (CS x SF) | Corrected odds combined (CS x SF) | Corrected probability combined (CS x SF) | | | | |
|-------------|-------------------|---------|----------------|-----------------------|---------------------|---------|----------------|-----------------------|----------------------|---------|----------------|-----------------------|--------------------------|---------|----------------|-----------------------|-------------------------|--------------------------------|-----------------------------------|--|--------------|--|--|--|
| | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | | | | | | | | |
| M1689R | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| G1706E | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| A1708E | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| S1715R | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| L1764P | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| I1766S | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| G1788V | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| V1838E | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| G1738R | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99725 | 3 999 | 345 | 0.99711 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| T1685I | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.9965 | 285 | 271 | 0.99633 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| T1685A | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99475 | 189 | 181 | 0.99449 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| C44Y | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| C44F | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99275 | 137 | 130 | 0.99239 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| C64Y | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.98675 | 74 | 71 | 0.98611 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.98525 | 67 | 64 | 0.98454 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| T37K | 0.99575 | 234 | 223 | 0.99554 | 0.99925 | 1 332 | 1 269 | 0.99921 | 0.71825 | 2.55 | 2.48 | 0.71227 | 0.86075 | 6.18 | 5.93 | 0.85580 | 597 | 0.99833 | 552 | 0.99819 | | | | |
| C39R | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.8225 | 4.63 | 4.46 | 0.81688 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99825 | 570 | 543 | 0.99816 | 15 992 001 | 1 | 14 505 579 | 1 | | | | |
| C61G | 0.98925 | 92 | 88 | 0.98872 | 0.764 | 3.24 | 3.13 | 0.75791 | 0.957 | 22 | 21 | 0.95504 | 0.54325 | 1.19 | 1.18 | 0.54136 | 2 048 | 0.99951 | 1 863 | 0.99946 | | | | |
| C47G | 0.99775 | 443 | 422 | 0.99764 | 0.99775 | 443 | 422 | 0.99764 | 0.93825 | 15.19 | 14.52 | 0.93556 | 0.14425 | 0.17 | 0.18 | 0.14931 | 6 738 | 0.99985 | 6 132 | 0.99984 | | | | |
| M18T | 0.99825 | 570 | 543 | 0.99816 | 0.99575 | 234 | 223 | 0.99554 | 0.80225 | 4.06 | 3.91 | 0.79639 | 0.10875 | 0.12 | 0.13 | 0.11296 | 2 314 | 0.99957 | 2 125 | 0.99953 | | | | |
| C39Y | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.71725 | 2.54 | 2.46 | 0.71128 | 0.00475 | 0.0048 | 0.005 | 0.00499 | 10 144 | 0.99990 | 9 383 | 0.99989 | | | | |
| M1775R | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.94675 | 17.78 | 16.98 | 0.94438 | 0.271 | 0.37 | 0.38 | 0.27704 | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 1 487 | 0.99933 | 1 459 | 0.99932 | | | | |
| L22S | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1 | 0.50000 | 1 | 0.50000 | | | | |
| M1775K | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.59975 | 1.50 | 1.47 | 0.59591 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1 | 0.50000 | 1 | 0.50000 | | | | |
| C47F | 0.49325 | 0.97 | 0.97 | 0.49357 | 0.245 | 0.32 | 0.34 | 0.25110 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00075 | 0.00075 | 0.00079 | 0.00079 | 0.00024 | 0.00024 | 0.00026 | 0.00026 | | | | |
| R1699W | 0.07225 | 0.078 | 0.081 | 0.07532 | 0.017 | 0.017 | 0.018 | 0.01782 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00225 | 0.0023 | 0.0024 | 0.00236 | 1.95E-5 | 1.95E-5 | 2.14E-5 | 2.14E-5 | | | | |
| T1720A | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| P1859R | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00075 | 0.00075 | 0.00079 | 0.00079 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1.88E-7 | 1.88E-7 | 2.07E-7 | 2.07E-7 | | | | |
| R1751Q | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00125 | 0.0013 | 0.0013 | 0.00131 | 0.00425 | 0.0043 | 0.0045 | 0.00446 | 3.13E-7 | 3.13E-7 | 3.45E-7 | 3.45E-7 | | | | |
| E1682K | 0.00375 | 0.0038 | 0.004 | 0.00394 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.60E-6 | 6.60E-6 | 7.27E-6 | 7.27E-6 | | | | |
| M1652I | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00475 | 0.0048 | 0.005 | 0.00499 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| I31M | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 0.00325 | 0.0033 | 0.0034 | 0.00341 | 4.38E-7 | 4.38E-7 | 4.83E-7 | 4.83E-7 | | | | |
| I1858L | 0.00125 | 0.0013 | 0.0013 | 0.00131 | 0.00725 | 0.0073 | 0.0077 | 0.00761 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 3.13E-7 | 3.13E-7 | 3.45E-7 | 3.45E-7 | | | | |
| P1776H | 0.01525 | 0.015 | 0.016 | 0.01599 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00625 | 0.0063 | 0.0066 | 0.00656 | 3.87E-6 | 3.87E-6 | 4.27E-6 | 4.27E-6 | | | | |
| F1662S | 0.00975 | 0.0098 | 0.01 | 0.01023 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00075 | 0.00075 | 0.00079 | 0.00079 | 0.02325 | 0.024 | 0.025 | 0.02436 | 7.39E-6 | 7.39E-6 | 8.14E-6 | 8.14E-6 | | | | |
| V1804D | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.0585 | 0.062 | 0.065 | 0.06107 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| A1669S | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.0585 | 0.062 | 0.065 | 0.06107 | 0.04625 | 0.048 | 0.051 | 0.04834 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1.21E-5 | 1.21E-5 | 1.33E-5 | 1.33E-5 | | | | |
| N132K | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00725 | 0.0073 | 0.0077 | 0.00761 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.09575 | 0.11 | 0.11 | 0.09958 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| D67Y | 0.00475 | 0.0048 | 0.005 | 0.00499 | 0.245 | 0.32 | 0.34 | 0.25110 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1.19E-6 | 1.19E-6 | 1.32E-6 | 1.32E-6 | | | | |
| M1652T | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.99975 | 3 999 | 3 809 | 0.99974 | 0.00125 | 0.0013 | 0.0013 | 0.00131 | 0.02325 | 0.024 | 0.025 | 0.02436 | 5 | 0.83347 | 5 | 0.83347 | | | | |
| K45Q | 0.47025 | 0.89 | 0.89 | 0.47159 | 0.69575 | 2.29 | 2.23 | 0.68997 | 0.84375 | 5.40 | 5.19 | 0.83846 | 0.21075 | 0.27 | 0.28 | 0.21672 | 4.79 | 0.82739 | 4.63 | 0.82245 | | | | |
| Sensitivity | 0.92 (23/25) | | | | 0.76 (19/25) | | | | 0.76 (19/25) | | | | 0.64 (16/25) | | | | 0.60 (15/25) | | | | 0.60 (15/25) | | | |
| Specificity | 0.87 (13/15) | | | | 0.67 (10/15) | | | | 0.67 (10/15) | | | | 0.93 (14/15) | | | | 0.93 (14/15) | | | | 0.87 (13/15) | | | |
| Accuracy | 0.90 (36/40) | | | | 0.73 (29/40) | | | | 0.73 (29/40) | | | | 0.75 (30/40) | | | | 0.75 (30/40) | | | | 0.70 (28/40) | | | |

See S13 Table for details.

S15 Table. Basic and corrected probability of pathogenicity when using the MWW method

| Variant | Colony Size assay | | | | Liquid Medium Assay | | | | Spot Formation assay | | | | Yeast Localization assay | | | | Combined odds (CS x SF) | Combined probability (CS x SF) | Corrected odds combined (CS x SF) | Corrected probability combined (CS x SF) | | | | |
|-------------|-------------------|---------|----------------|-----------------------|---------------------|---------|----------------|-----------------------|----------------------|---------|----------------|-----------------------|--------------------------|---------|----------------|-----------------------|-------------------------|--------------------------------|-----------------------------------|--|--------------|--|--|--|
| | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | Probability | Odds | Corrected odds | Corrected probability | | | | | | | | |
| M1689R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| G1706E | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| A1708E | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| S1715R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| L1764P | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| I1766S | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| G1788V | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| V1838E | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| G1738R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| T1685I | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| T1685A | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| C44Y | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| C44F | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| C64Y | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.84875 | 5.61 | 5.39 | 0.84355 | 0.99425 | 173 | 165 | 0.99397 | 0.63725 | 1.76 | 1.72 | 0.63245 | 691.479 | 1 | 627.381 | 1 | | | | |
| T37K | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.98325 | 59 | 56 | 0.98244 | 0.97425 | 38 | 36 | 0.97303 | 0.99975 | 3.999 | 3.809 | 0.99974 | 151.302 | 1 | 137.418 | 1 | | | | |
| C39R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.30025 | 0.43 | 0.44 | 0.30607 | 0.99425 | 173 | 165 | 0.99397 | 0.99975 | 3.999 | 3.809 | 0.99974 | 691.479 | 1 | 627.381 | 1 | | | | |
| C61G | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.76275 | 3.21 | 3.11 | 0.75666 | 0.99425 | 173 | 165 | 0.99397 | 0.97825 | 45 | 43 | 0.97721 | 691.479 | 1 | 627.381 | 1 | | | | |
| C47G | 0.97425 | 38 | 36 | 0.97303 | 0.99225 | 128 | 122 | 0.99187 | 0.99425 | 173 | 165 | 0.99397 | 0.49675 | 0.99 | 0.99 | 0.49690 | 6.542 | 0.99985 | 5.943 | 0.99983 | | | | |
| M18T | 0.99525 | 210 | 200 | 0.99501 | 0.93475 | 14.33 | 13.69 | 0.93193 | 0.80725 | 4.19 | 4.04 | 0.80144 | 0.55375 | 1.24 | 1.23 | 0.55145 | 878 | 0.99886 | 806 | 0.99876 | | | | |
| C39Y | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.99425 | 173 | 165 | 0.99397 | 0.00825 | 0.0083 | 0.0087 | 0.00866 | 691.479 | 1 | 627.381 | 1 | | | | |
| M1775R | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.89225 | 8.28 | 7.93 | 0.88807 | 0.75825 | 3.14 | 3.03 | 0.75215 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 12.543 | 0.99992 | 11.558 | 0.99991 | | | | |
| L22S | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.64925 | 1.85 | 1.81 | 0.64419 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1 | 0.50000 | 1 | 0.50000 | | | | |
| M1775K | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.81475 | 4.40 | 4.24 | 0.80903 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 1 | 0.50000 | 1 | 0.50000 | | | | |
| C47F | 0.96125 | 25 | 24 | 0.95947 | 0.27425 | 0.38 | 0.39 | 0.28028 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.0062 | 0.00616 | 0.0062 | 0.00618 | | | | |
| R1699W | 0.02325 | 0.024 | 0.025 | 0.02436 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 5.95E-6 | 5.95E-6 | 6.55E-6 | 6.55E-6 | | | | |
| T1720A | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| P1859R | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.0043 | 0.0045 | 0.00446 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| R1751Q | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| E1682K | 0.00225 | 0.0023 | 0.0024 | 0.00236 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00125 | 0.0013 | 0.00131 | 0.00025 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 2.82E-6 | 2.82E-6 | 3.11E-6 | 3.11E-6 | | | | |
| M1652I | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00525 | 0.0053 | 0.0055 | 0.00551 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| I31M | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00875 | 0.0088 | 0.00918 | 0.00325 | 0.0033 | 0.0034 | 0.00341 | 2.21E-6 | 2.21E-6 | 2.43E-6 | 2.43E-6 | | | | | |
| I1858L | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00775 | 0.0078 | 0.0082 | 0.00813 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| P1776H | 0.00125 | 0.0013 | 0.0013 | 0.00131 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 3.13E-7 | 3.13E-7 | 3.45E-7 | 3.45E-7 | | | | |
| F1662S | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.01175 | 0.012 | 0.01232 | 0.01232 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| V1804D | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 6.25E-8 | 6.25E-8 | 6.89E-8 | 6.89E-8 | | | | |
| A1669S | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 0.01775 | 0.018 | 0.019 | 0.01860 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 4.52E-6 | 4.52E-6 | 4.98E-6 | 4.98E-6 | | | | |
| N132K | 0.00175 | 0.0018 | 0.0018 | 0.00184 | 0.01225 | 0.012 | 0.013 | 0.01285 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.06325 | 0.068 | 0.071 | 0.06600 | 4.38E-7 | 4.38E-7 | 4.83E-7 | 4.83E-7 | | | | |
| D67Y | 0.00225 | 0.0023 | 0.0024 | 0.00236 | 0.32175 | 0.47 | 0.49 | 0.32731 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 0.00025 | 0.00025 | 0.00026 | 0.00026 | 5.64E-7 | 5.64E-7 | 6.22E-7 | 6.22E-7 | | | | |
| M1652T | 0.99975 | 3.999 | 3.809 | 0.99974 | 0.75525 | 3.09 | 2.99 | 0.74915 | 0.00875 | 0.0088 | 0.0093 | 0.00918 | 0.03275 | 0.034 | 0.035 | 0.03428 | 35 | 0.97245 | 35 | 0.97244 | | | | |
| K45Q | 0.37925 | 0.61 | 0.62 | 0.38366 | 0.79275 | 3.83 | 3.69 | 0.78681 | 0.99425 | 173 | 165 | 0.99397 | 0.63725 | 1.76 | 1.72 | 0.63245 | 106 | 0.99062 | 103 | 0.99034 | | | | |
| Sensitivity | 0.96 (24/25) | | | | 0.96 (24/25) | | | | 0.64 (16/25) | | | | 0.76 (19/25) | | | | 0.64 (16/25) | | | | 0.64 (16/25) | | | |
| Specificity | 0.87 (13/15) | | | | 0.87 (13/15) | | | | 0.80 (12/15) | | | | 0.93 (14/15) | | | | 0.93 (14/15) | | | | 0.87 (13/15) | | | |
| Accuracy | 0.93 (37/40) | | | | 0.93 (37/40) | | | | 0.70 (28/40) | | | | 0.83 (33/40) | | | | 0.83 (33/40) | | | | 0.73 (29/40) | | | |

See S13 Table for details.

S16 Table. Quantitative analysis of S24 Fig

| Assay | Method | Number of variants correctly classified | | | | | Number of variants not classified | | | Number of variants misclassified | | | Balance | Sensitivity | Specificity | Accuracy |
|--------------------|-------------------------|---|---------------------------|--------------------|------------------------|-----------------------------|-----------------------------------|--------------------|-------|----------------------------------|-------------------------|-------|---------|-------------|-------------|----------|
| | | Pathogenic in class 5 | Pathogenic in class 5 + 4 | Neutral in class 1 | Neutral in class 1 + 2 | Total (class 1 + 2 + 4 + 5) | Pathogenic in class 3 | Neutral in class 3 | Total | Pathogenic in class 1 or 2 | Neutral in class 5 or 4 | Total | | | | |
| Colony Size | Standard | 23 | 23 | 12 | 13 | 36 | 2 | 1 | 3 | 0 | 1 | 1 | 35 | 0.92 | 0.87 | 0.90 |
| | Standard with reference | 22 | 23 | 8 | 13 | 36 | 2 | 1 | 3 | 0 | 1 | 1 | 35 | 0.92 | 0.87 | 0.90 |
| | MWW | 22 | 24 | 9 | 13 | 37 | 0 | 1 | 1 | 1 | 1 | 2 | 35 | 0.96 | 0.87 | 0.93 |
| Liquid Medium | Standard | 18 | 20 | 8 | 12 | 32 | 4 | 2 | 6 | 1 | 1 | 2 | 30 | 0.80 | 0.80 | 0.80 |
| | Standard with reference | 18 | 19 | 8 | 10 | 29 | 5 | 4 | 9 | 1 | 1 | 2 | 27 | 0.76 | 0.67 | 0.73 |
| | MWW | 15 | 16 | 7 | 12 | 28 | 8 | 3 | 11 | 1 | 0 | 1 | 27 | 0.64 | 0.80 | 0.70 |
| Spot Formation | Standard | 15 | 17 | 13 | 14 | 31 | 4 | 1 | 5 | 4 | 0 | 4 | 27 | 0.68 | 0.93 | 0.78 |
| | Standard with reference | 15 | 16 | 10 | 14 | 30 | 5 | 1 | 6 | 4 | 0 | 4 | 26 | 0.64 | 0.93 | 0.75 |
| | MWW | 18 | 19 | 10 | 14 | 33 | 2 | 0 | 2 | 4 | 1 | 5 | 28 | 0.76 | 0.93 | 0.83 |
| Yeast Localization | Standard | 16 | 16 | 7 | 13 | 29 | 3 | 2 | 5 | 6 | 0 | 6 | 23 | 0.64 | 0.87 | 0.73 |
| | Standard with reference | 14 | 15 | 7 | 13 | 28 | 4 | 2 | 6 | 6 | 0 | 6 | 22 | 0.60 | 0.87 | 0.70 |
| | MWW | 15 | 16 | 9 | 13 | 29 | 3 | 2 | 5 | 6 | 0 | 6 | 23 | 0.64 | 0.87 | 0.73 |

Framed numbers highlight the differences with the results obtained using the probability system of classification (S6 Table).

S17 Table. Effect of functional assay parameters upon the quantile system of classification

| Parameters | Standard method | Standard with reference method | MWW method | Figure |
|----------------------------------|--|--|---|--------|
| Position of pathogenic mutations | +++ Shift from the neutral sector towards the pathogenic sector increases both sensitivity and specificity. | +++ As in the standard method. | +++ As in the standard method but the increase stops when no more overlap with the BRCA1 reference distribution. | S25A |
| Position of neutral mutations | +++ Shift from the neutral sector towards the pathogenic sector decreases both sensitivity and specificity. | +++ As in the standard method. | +++ No effect, except a complete <u>misclassification</u> of the pathogenic mutations when the WT reference distribution is left on one side of all the neutral and pathogenic distributions with no overlap. | S25B |
| Position of WT reference | 0 No effect. | +++ Shift of the median towards the null value decreases sensitivity and specificity. Negative values of the WT reference median invert the fluctuation of the best cut-off. | +++ No effect, except a complete <u>misclassification</u> of the pathogenic mutations when the WT reference distribution falls outside of the range of the neutral and pathogenic distributions, with no overlap. | S25C |
| Experimental sensitivity | +++ Decrease of the experimental sensitivity decreases specificity. This is advantageous, as the neutral region is contaminated by pathogenic mutations (presence of false negatives). | +++ As in the standard method. | +++ As in the standard method. | S25D |
| Experimental specificity | +++ Decrease of the experimental specificity decreases sensitivity. This is advantageous, as the pathogenic region is contaminated by neutral mutations (presence of false positives). | +++ As in the standard method. | +++ As in the standard method. | S25D |
| Number of neutral mutations | 0 Decrease of neutral mutations slightly shifts the best cut-off distribution towards the neutral sector, but no impact on sensitivity and specificity. | 0 As in the standard method. | 0 As in the standard method. | S26A |
| Number of pathogenic mutations | 0 Decrease of pathogenic mutations slightly shifts the best cut-off distribution towards the pathogenic sector, but no impact on sensitivity and specificity. | 0 As in the standard method. | 0 No effect. | S26A |
| Number of mutant values | 0 No effect (and no more best cut-off fluctuation when $n_{mutant} = 1$). | 0 No effect (and no more best cut-off fluctuation when $n_{mutant} = n_{BRCA1} = 1$). | 0 Decrease of the number of mutant values narrows the best cut-off fluctuation, with no effect on variant classification. No more best cut-off fluctuation when $n_{mutant} = n_{BRCA1} = 1$. | S26B |
| Number of WT BRCA1 values | 0 No effect. | +++ Loss of sensitivity with the decrease of the number of BRCA1 values, but maximal sensitivity when $n_{BRCA1} = 1$. No more best cut-off fluctuation when $n_{mutant} = n_{BRCA1} = 1$. | 0 Decrease of the number of mutant values narrows the best cut-off fluctuation, with no effect on variant classification. No more best cut-off fluctuation when $n_{mutant} = n_{BRCA1} = 1$. | S26B |
| Range of mutant distributions | 0 No effect (and no more best cut-off fluctuation when the range is null). | 0 No effect (and no more best cut-off fluctuation when the range of mutant and BRCA1 distributions are null). | 0 As in the standard with reference method. | S27 |
| Range of WT BRCA1 distribution | 0 No effect. | 0 No effect (decrease of the range decreases the best cut-off fluctuation, with no more fluctuation when the range of mutant and BRCA1 distributions are null). | 0 No effect (and no more best cut-off fluctuation when the range of mutant and BRCA1 distributions are null). | S27 |

Sensitivity and specificity are defined here as the distance between the best cut-off fluctuation and the position of the pathogenic and neutral mutations, respectively, which reflects the accuracy of the quantile system of variant classification. For instance, the MWW method exhibits a null sensitivity in variant classification if the p values of the pathogenic mutations are in the grey area (class 3), and shows a maximal sensitivity in variant classification if the p values are in the pink area (class 4) or beyond (class 5). Misclassification is defined as class 4 or 5 for the neutral mutations and class 2 or 1 for the pathogenic mutations. See S9 Table for further details. The framed text indicates differences, as compared to the probability system of classification (S9 Table)