# Algorithmic methods to infer the evolutionary trajectories in cancer progression<sup>∗</sup>

#### Supplementary Material



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<sup>∗</sup>This work follows up on our earlier project initiated by BM and carried out by Milan-Bicocca and the Catalan Institute of Oncology, based on a framework discussed at the 2014 Como School of Advanced Studies on Cancer, Systems and Complexity (CSAC). PiCnIc was designed and constructed by MA's Bioinformatics lab at University of Milan-Bicocca, within a project led and supervised by GC. GC, AG and DR designed the pipeline, and GC, DR and LDS coded and executed it. Data gathering and model interpretation was done by GC, LDS, DR, AG together with BM, VM and RSP. GM, MA, VM and BM provided overall organizational guidance and discussion. GC, AG, RSP and BM wrote the original draft of the paper, which all authors reviewed and revised in the final form. BM and MA are co-senior authors.



## <span id="page-1-0"></span>A Reproducing this study

#### Code availability

The implementation of PiCnIc shown in the Main Text was performed by using, as core, the R language, and other external Java tools which we reference in this document. In r, much of the data processing and inference is done by exploiting the current version of the open-source

#### TRanslational ONCOlogy (TRONCO, [\[1\]](#page-23-0), version 2.3)

package which implements up-to-date statistical algorithms to estimate cancer progression models from a list of genomic lesions (e.g., somatic mutations, copy number variations or persistent epigenetic states) in a population of independent tumors, or in a single patient.

TRONCO's official webpage is reachable from the Software section of our group's webpage

#### <http://bimib.disco.unimib.it/>

By navigating to the Case Studies section of TRONCO's official webpage one can find the source code to replicate this study (i.e., the PiCnIc's implementation) along with the documentation detailing all the implementation, as well as the data that we used. This should allow easy implementation of similar studies in different contexts.



## <span id="page-2-0"></span>B Glossary

This glossary of terms shall be of help to readers not familiar with the concepts mentioned in the Main Text. For clarity, terms are separate in two categories according to the fact that they are common to the statistics or the cancer biology communities. Each term which is included in this glossary appears in color.



### Terms common to the statistics community



#### Terms common to the cancer biology community

## <span id="page-3-0"></span>C PiCnIc's implementation for COADREAD samples

Here we detail all the steps implemented to use PiCnIc for CRC progression inference.

### <span id="page-4-0"></span>C.1 TCGA COADREAD project data

COADREAD provides genome-scale analysis of samples with exome sequence, DNA copy number, promoter methylation, messenger RNA and microRNA expression data, which we used to define the input dataset. In particular, only samples with both mutations and CNAs profiles were used in the analysis. Supplementary Table [S1](#page-8-1) details the dataset.

Dataset used to infer models presented in the Main Text. Samples published in [\[2\]](#page-23-1) were used as, to the best of our knowledge, these represent the highest–quality data made available by COADREAD as of today; for these samples TCGA provides somatic mutation profiles and high-resolution focal CNAs via GISTIC. These are obtained from TCGA data freeze as of 2 February 2012, downloaded on 12 March 2015, from repository:

[https://tcga-data.nci.nih.gov/docs/publications/coadread\\_2012/](https://tcga-data.nci.nih.gov/docs/publications/coadread_2012/)

The following files were processed to produce the data:

- TCGA\_CRC\_Suppl\_Table2\_Mutations\_20120719.xlsx
- Somatic mutations profiles obtained via whole-exome sequencing of 224 colorectal tumors by TCGA. Data available consists of 15995 mutations in 228 samples, provided in the Manual Annotation Format (MAF). Samples were selected to univocally match the 224 patients as of the TCGA guidelines for aliquote disambiguation, see <https://wiki.nci.nih.gov/display/TCGA/TCGA+barcode>. All the mutations annotated by TCGA – truncating (De\_novo\_Start\_OutOfFrame, Frame\_Shift\_Del, Frame\_Shift\_In, Nonsense\_Mutation, Splice\_Site, Frame\_Shift\_Ins, In\_Frame\_Del), silent (Silent) and missense (Missense\_Mutation) – were considered for analysis; notice that the majority of them are missense, see Figures [S2](#page-12-0) and [S3.](#page-13-0)
- crc\_gistic.txt.zip

Focal Copy Number Alterations (CNAs) for 564 patients derived from whole-genome sequencing using the Illumina HiSeq platform. *High-level gains* and *homozygous deletions* were considered for analysis by selecting entries with GISTIC scores  $\pm$  2;

• crc\_clinical\_sheet.txt

Clinical data summary with patient stage and Micro Satellite Stabe/Unstable (MSS/MSI) status being any of: MSS, MSI-high and MSI-low.

The list of patients used was first reduced to those having both CNAs and somatic mutation data, and then was split into two groups: MSI-HIGH and MSS. The training cohort has 152 MSS and 27 MSI-HIGH samples; samples flagged as low MSI were excluded from the study as they have not been shown to differ in their clinicopathologic features or in most molecular features from MSS tumors [\[3\]](#page-23-2).

#### <span id="page-4-1"></span>C.2 Driver events selection

In the TCGA COADREAD study [\[2\]](#page-23-1) integrated analysis of mutations, copy number and mRNA expression changes in 195 tumours with complete data was performed. Part of the analysis was carried out by using the MutSig tool [\[4\]](#page-23-3), as well as manual curation. Samples were grouped by mutation status, and recurrent alterations in key CRC pathways were identified in [\[2\]](#page-23-1) (Fig. 4, Supplementary Fig. 6 and Supplementary Table 1) as a result, we can use the consortium's list of 33 driver genes annotated to 5 pathways and use these to extract our progression models. These are well-known cancer genes, frequently reported as relevant to colorectal progression and to the major pathways involved in CRC. Driver events are alterations in the following genes (acronym resolved at <http://www.genecards.org/>):

- WNT genes (14): Adenomatous Polyposis Coli (APC), Dickkopf WNT Signaling Pathway Inhibitor 4 (DKK-4), Transcription Factor 7-Like 2 (TCF7L2), Catenin beta-1 (CTNNB1), Low Density Lipoprotein Receptor-Related Protein 5 (LRP5), F-Box And WD Repeat Domain Containing 7 (FBXW7), Dickkopf WNT Signaling Pathway Inhibitor 1(DKK-1), Frizzled Class Receptor 10 (FZD10), AT Rich Interactive Domain 1A (ARID1A), Dickkopf WNT Signaling Pathway Inhibitor 2 (DKK-2), APC Membrane Recruitment Protein 1 (FAM123B, also known as AMER1), SRY (Sex Determining Region Y)-Box 9 (SOX9), Dickkopf WNT Signaling Pathway Inhibitor 3 (DKK-3) and  $axin-2$   $(AXIN2)$ ;
- RTK/RAS genes (5): Erb-B2 Receptor Tyrosine Kinase 2 (ERBB2), Erb-B2 Receptor Tyrosine Kinase 3 (ERBB3), Neuroblastoma RAS Viral (V-Ras) Oncogene Homolog (nras), Kirsten Rat Sarcoma Viral Oncogene Homolog (KRAS) and B-Raf Proto-Oncogene, Serine/Threonine Kinase (BRAF);
- TGF- $\beta$  genes (5): Transforming Growth Factor, Beta Receptor 1 (TGFBR1), SMAD Family Member 3 (SMAD3), Transforming Growth Factor, Beta Receptor 2 (TGFBR2), SMAD Family Member 4 (SMAD4), Activin A Receptor Type IB (ACVR1B), Activin A Receptor Type IIA ( $ACVR2A$ ) and SMAD Family Member 2 ( $SMAD2$ );
- igf2/pi3k genes (5): Insulin-Like Growth Factor 2 (igf2), Insulin Receptor Substrate 2 (irs2), Phosphatidylinositol-4,5-Bisphosphate 3-Kinase Catalytic Subunit Alpha (pik3ca), Phosphoinositide-3-Kinase Regulatory Subunit 1 Alpha ( $PIK3R1$ ) and Phosphatase And Tensin Homolog ( $PTEN$ );
- P53 genes (2): Tumor Protein P53 (TP53) and ATM Serine/Threonine Kinase (ATM).

In the Main Text,  $RTK/RAS$  and  $IGF2/P13K$  pathways are shortly denoted as RAS and P13K.

The distinct types of mutations detected in these genes are shown in Figure [S2,](#page-12-0) as well as the overall rate of COADREAD mutations. The spatial distribution (per gene) of such mutations is shown in Figure [S3.](#page-13-0)

#### <span id="page-5-0"></span>C.3 Mutual exclusivity groups of alterations.

Groups of alterations showing a trend of mutual exclusivity were scanned with MUTEX and mutations and CNA hitting any of the 33 selected genes as input. MUTEX was run independently on MSS and MSI-HIGH groups (Supplementary Table [S2,](#page-9-0) running times: approximately 6 and 3.5 hours, respectively, on a standard Desktop machine).

We selected only groups with score  $< 0.2$ , where the score is derived from p-values corrected for false discovery rate. 3 groups are found for MSI-HIGH tumors and 6 for MSS. For MSI-HIGH tumors, the three predicted groups consists of genes  $ACVR1B$ ,  $ACVR2A$ ,  $TP53$  and  $ERBB2$ , of genes  $BRAF$ ,  $NRAS$  and  $TGFBR2$ , and of genes  $KRAS$  and  $BRAF$ .

Further groups of exclusive alterations were considered consistent with results reported in [\[2\]](#page-23-1). These include groups derived by consolidated knowledge of colorectal progression: the well-known WNT alterations in APC/CTNNB1 [\[5\]](#page-23-4), as well as RAS alterations in KRAS, NRAS and BRAF genes [\[6\]](#page-23-5). Similarly, we used also a group collected by scanning non-hypermutated tumors with the MEMO tool in  $[2]$  - this group includes PIK3CA, PTEN, ERBB2 and IGF2 genes. These groups were restricted to account only for genes actually altered in a certain subtype, e.g., MSI-HIGH tumors lack CTNNB1 mutation, making the Wnt group irrelevant. Groups for MSS tumors are shown as Supplementary Figure [S4,](#page-14-0) groups for MSI-HIGH tumors are in the Main Text.

#### <span id="page-5-1"></span>C.4 CAPRI's execution

Background on the algorithm. CAPRI algorithm can be executed in two different modes, originally dubbed as "supervised" when formulas are given in input for testing, and "unsupervised" when this is not the case. This paper deals with the former; see the Main Text for the interpretation of formulas in this context and [\[7\]](#page-23-6) for a derivation of the algorithm. CAPRI is a three-steps procedure, which we briefly recall here.

- 1. CAPRI starts by creating a "lifted" representation of the input data M which includes the input formulas; each formula – which is written in propositional logic – is so evaluated to yield a new column in the dataset. This is the input processed by the algorithm, which starts by selecting a set of candidate model edges, which are then used to constrain a score-based Bayesian model-selection problem [\[7\]](#page-23-6).
- 2. The initial set of selective advantage relations S, which determines the model edges  $x \to y$ , is computed by evaluating the following inequalities

(Temporal priority)  $p(x) > p(y)$ (Probability raising)  $p(y | x) > p(y | \neg x)$ ,

where  $p(\cdot)$  is a marginal probability,  $p(\cdot | \cdot)$  is a conditional,  $\neg x$  is the negation of x and either x or y is not a formula. It is interesting to observe that probability raising implies that  $x$  and  $y$  are positively statistically dependent, thus imposing a minimum threshold on their association [\[8\]](#page-23-7).

In each of CAPRI's executions, the distribution of the observed marginals and conditionals are estimated by  $K$ non-parametric bootstrap resamples; practically, this means that we create a bootstrapped approximation for each of the four populations  $\hat{p}(x)$ ,  $\hat{p}(y)$ ,  $\hat{p}(y \mid x)$  and  $\hat{p}(y \mid \neg x)$ . Then, we use a single-tail non-parametric Mann-Withney U test of the difference in mean to test the hypothesis that one of the two populations is more probable than the other, e.g.,  $\hat{p}(x) > \hat{p}(y)$ . With this test, we can compute two p-values, one for each condition. An edge is included in  $S$  if at least the p-value for probability raising is below a significance threshold  $p_*$  and an edge is said not to be orientable if its p-value for temporal priority is above the same threshold. Cycles  $x_1 \to x_2 \to \ldots x_k \to x_1$  that might appear in S are broken by deleting the edge with minimum p-value; both K and  $p_*$  are custom parameters.

3. Optimization of S, namely detection of the subset  $S^*$  of S with the edges that we include in the final progression model is done by optimizing, via hill climbing or tabu search, the score with regularization

$$
S^* \triangleq \arg\min_{\hat{\mathcal{S}} \subset \mathcal{S}} \left\{-2\log[\mathcal{L}(\hat{\mathcal{S}} \mid \mathbf{M})] + \theta|\hat{\mathcal{S}}|\right\},\tag{1}
$$

where  $\mathcal{L}(\cdot)$  is the *model likelihood* and  ${\bf M}$  is the input data; the estimated optimal solution is  $\mathcal{S}^*$ , which is displayed as a Suppes-Bayes Causal Network. The different regularization strategies mentioned in the Main Text, BIC and AIC, are obtained by the following parametrization:

> (Bayesian Information Criterion)  $\theta \triangleq \log(n)$ (Akaike Information Criterion)  $\theta \triangleq 2$ .

Besides edges, a model has a set of parameters  $\theta$  which define the conditional probability table of each edge and should be fit from data; these are necessary if one wishes to use a model as a "generator" of further data. For discrete-valued graphical models, for each parent set  $\pi_x\to y$ , parameter  $\bm\theta(y)=p(y\mid\pi_x)$  can be taken either as the *maximum likelihood estimate* from the lifted input, or by using a Bayesian interpretation [\[9\]](#page-23-8).

Usage in this context. CAPRI was run, on each group of tumors, by selecting alterations from the pool of 33 pathway genes; every alteration on a gene  $x$  is included if any of these apply:

- the alteration frequency of  $x$  sum of mutation and CNA frequency is greater than 5%;
- $x$  it is part of an exclusivity group.

The set of selected events for MSI-HIGH training tumors is shown in the Main Text, the analogous set for MSS tumors is shown as Supplementary Figure [S5.](#page-15-0)

CAPRI was executed in its supervised mode by writing formula over groups and genes with multiple alterations associated, as explained in the Main Text. For instance, for MSI-HIGH tumors with alterations in ras pathway we grouped hard exclusivity of nras mutations and deletions, with soft exclusivity of kras and braf mutations. Our aim was to account for a small subset of samples with concurrent kras and nras alterations (see Figure 2, Main Text). The list of all Boolean formulas written over groups is in Table [S3;](#page-10-0) this approach was adopted also when a gene harbors multiple alterations in a subtype, e.g., ERBB2 in MSS training samples which shows a trend of soft exclusivity between mutations and amplifications. We used both AIC and BIC scores to regularize inference after 100 non-parametric bootstrap iterations for estimation of the preliminary selective advantage relations – Mann-Whitney U test was performed with a minimum threshold  $p_* = 0.05$ . In most cases p-values are orders of magnitude below  $p_*$  - exact values reported as Dataset File S1. CAPRI's models with such p-values and non-parametric bootstrap confidence are shown in Figures [S6](#page-16-0) and [S7,](#page-17-0) statistical validation of the models is discussed in the next section.

## <span id="page-6-0"></span>D Statistical validation of the models

P-values from hypothesis testing, as well as scores from  $k$ -fold *cross-validation* and various bootstrapping techniques can be used to measure the statistical consistency of models and data, each one capturing different potential errors in the inference process. Approaches such as cross-validation and bootstrap are sometimes also used in the  $(ex novo)$  generation and inference of models from data (see, e.g., *bootstrap consensus models* [\[10\]](#page-23-9)), but we only use them here for the a posteriori evaluation of a model's confidence, and we interpret them as a quantitative measure for the relative assessment of each model's relation.

All the p-values and the scores that we present here are computed within TRONCO.

#### <span id="page-6-1"></span>D.1 Edge p-values

As explained in [§C.4,](#page-5-1) for each model edge  $x \to y$  we get two p-values by assessing temporal priority and probability raising via Mann-Whitney U testing.

For each edge, a p-value for the *hypergeometric test* of overlap between alteration profiles  $x$  and  $y$  can be computed. More precisely, we test if there is a difference between the number of samples containing both x and y versus the total population of samples with  $x$ ,  $y$ , or both. We would like the overlap to be significant as those samples – that determine the joint probability of x and  $y$  – are those supporting the presence of a selection trend among x and y.

An edge is fully supported if all three p-values are below a custom significance threshold, e.g., 0.05 or even better 0.01. Some edges might have the p-value for temporal priority above the threshold. If so, the selection trend might be still significant, but the temporal order of x and  $y - i.e.$ , the direction of selection – is not supported by the data.

#### <span id="page-7-0"></span>D.2 Bootstrap

We used non-parametric and statistical bootstrap techniques to measure the goodness-of-fit, as originally proposed in [\[11\]](#page-23-10). In this case, we distinguish two type of errors that one could make in the inference process, estimating the presence or absence of edges in the model:

- Type I errors: incorrect rejection of a true  $H_0$  (null-hypthesis), i.e., a "false positive" edge that we wrongly include.
- Type II errors: incorrect acceptance (failure to reject) of a false  $H_0$ , i.e., a "false negative" edge that we miss.

Non-parametric bootstrap [\[12\]](#page-23-11) computes scores to be interpreted here as follows. If a model contains an edge  $x \to y$ that is a true positive, we expect its score to be high. In other words, when we sample with repetition subsets of the original data and re-run the inference process we expect to often find models which contain the edge  $x \to y$ . Conversely, for a node y without incoming edges, or equivalently for any edge  $x \to y$  which is correctly excluded from a model – a true negative – we would expect its bootstrap score to be low. However, this reasoning can be also generalized to whole models, where we count how many times we re-infer exactly the same model. Clearly, such scores will depend also on the empirical probabilities of the nodes in our data, and their deviation from the true probabilities of the phenomenon. So, one might expect rare events to be less frequently bootstrapped, which results in a lower estimate; however, such counts can be anyway interpreted as measures of repeatability of our findings $^1\!.$  $^1\!.$  $^1\!.$ 

The above bootstrap approach depends on two random number generators: one to shuffle data (the a posteriori bootstrap), and one to evaluate CAPRI's inequalities via hypothesis testing (the internal CAPRI's bootstrap, see [C.4\)](#page-5-1). Thus, to ensure that no bias is introduced by the random number generators, we performed a statistical bootstrap by holding data fixed, and re-estimating CAPRI's inequalities with generators initialized with different seeds. We evaluated the robustness of our scores for all edges imputed to be genuine and hence, high-scoring.

Notice that, in principle, even *parametric bootstrap* scores could be computed if we used the model to generate bootstrapped data [\[12\]](#page-23-11). However, as the support of the distribution subsumed by a model with  $n$  nodes consists of  $2^n$ possible outcomes, sampling uniformly from large models might be computationally hard. For this reason, and because such scores are overestimates of the non-parametric ones, we did not include them in our computation.

The MSS and MSI progression models are annotated with the non-parametric bootstrap scores in Figures [S6](#page-16-0) and [S7.](#page-17-0) Non-parametric and statistical bootstrap scores for a set of selected edges are shown and commented in Figure [S8.](#page-18-0) For the same set of edges, we also report the p-values for CAPRI's inequalities assessed in the MSS and MSI progression models (temporal priority and probability raising, as a measure of the selectivity among the alterations, and hypergeometric, as a measure of the randomness in the overlap of two alteration profiles). Additional comments are in the caption.

#### <span id="page-7-1"></span>D.3 Cross-validation

Next we study the sufficiency of the data sizes for model inference and its ability to characterize the underlying progression (goodness-of-data). Thus, we focus on the Type III errors, which occur when the sample size is inadequate or the sample is a poor descriptor for the reference phenomenon<sup>[2](#page-7-3)</sup>, and thus failing to represent the progression. For this purpose, we used cross-validation with the data used for the models built (see the Main Text), and followed the best practices developed by the Bayesian Networks community [\[9\]](#page-23-8).

TRONCO exploits the cross-validation routines implemented in the bnlearn package [\[16\]](#page-23-12). The approach that we adopt is a k-fold non-exhaustive cross-validation, which we repeat 10 times to average its results. Exhaustive strategies might be used for datasets of small sample size. Each run of cross-validation consists in computing a loss function for a model; its steps are the followings:

- split randomly the data in  $k = 10$  groups, and then repeat the following two steps, for each group in turn:
	- 1. set one of the groups to be the "training"  $M_{tr}$ ;
	- 2. merge the others  $k 1$  to be the "test"  $M_{te}$ ;
	- 3. by holding fixed the model structure (i.e., the edges in  $S^*$ ), fit the model parameters  $\theta$  over the training data via maximum likelihood estimates, compute a score over the test  $M_{te}$  (see below) and the corresponding loss;

<span id="page-7-2"></span> $1B$ ootstrapping techniques have been widely used to gauge uncertainty in estimates, but also subjects of philosophical debate about their precise interpretation, especially when coupled with various significance thresholds – unlike the situation with a p-value for a null hypothesis. An exhaustive review on the topic is provided by Soltis in [\[13\]](#page-23-13). We follow the ideas originally developed in the area of phylogenetic analysis [\[14\]](#page-23-14), suggesting that the scores can be alternatively interpreted as a measure of accuracy of the method or of the robustness of the data [\[15\]](#page-23-15).

<span id="page-7-3"></span><sup>&</sup>lt;sup>2</sup>Consider the case where the samples are from two different unknown and heterogeneous groups, with random chance making low values to be sampled from a group that actually has a majority of high values, and vice versa. In this case, the samples will not be the best descriptor of the two groups.

• combine the  $k$  loss estimates to give an overall loss for data.

Let  $\bm{\theta}_{tr}$  be the parameters fit from the training set  $\mathbf{M}_{tr}$ , and  $\mathcal{S}^*$  the edges in the model, three scores are computed with cross-validation:

1. the negative entropy of a model – i.e., the negated expected log-likelihood of the test set for the Bayesian network fitted from the training set, that is

$$
\texttt{eloss}(\mathcal{S}^*, \pmb{\theta}_{tr}) = -\mathbb{E}[\mathcal{L}(\mathcal{S}^*, \pmb{\theta}_{tr} \mid \mathbf{M}_{te})]\,.
$$

- 2. the prediction error for a single node x and its parents set  $X$ , i.e., we measure how precisely we can predict the values of x by using only the information present in its local distribution  $\theta_{tr}(x)$ . This parameter corresponds to computing the misclassification rates from  $p_{te}(x)$ , the empirical marginal probability of x estimated from the test.
- 3. the posterior classification error for a single node x and one of its parent node  $y \in X$  i.e., the values of x are predicted using only the information present in  $y$  by likelihood weighting and Bayesian posterior estimates.

See [\[17\]](#page-23-16) for a discussion on these loss functions. The first statistics measures the *log-likelihood loss* when we "forget" some of the samples used to infer a model (indeed, the test samples); see Figure [S9.](#page-19-0) Roughly, we are measuring how the model's predictive power changes as we look at the data from different viewpoints. This is a score for a whole model, it has no scale – so cannot be used to say how good the models/data are, in any absolute sense. Nonetheless, it can be used to evaluate how "stable" a model is, for a certain dataset.

The second and third statistics measure the accuracy of the parent-set,  $X$ , for a child  $x$ ; the second statistics dealing with the whole parent set as predictor, and the third, the individual contribution of each of the parents. For these two statistics, we desire the prediction error to be low, as a measure of goodness. These are shown in Figure [S8](#page-18-0) (selected edges), [S10](#page-20-0) and [S11](#page-21-0) (all edges, prediction error).

### <span id="page-8-0"></span>E Supplementary Tables and Figures



Datasets (CNAs and mutations provided by TCGA)

MSI\_status in clinical data available for the COADREAD project.

<span id="page-8-1"></span>Table S1: **COADREAD Data.** Data used in this study, derived from the TCGA COADREAD project [\[2\]](#page-23-1).



† Manually curated from Pathway Commons, SPIKE and SignaLink databases. Provided with the tool; available for download at https://code.google.com/p/mutex/.



#### MUTEX groups with score  $< .2$

<span id="page-9-0"></span>Table S2: MUTEX: parameters and results. Top: Parameters used to run MUTEX on the original TCGA MSS/MSI-HIGH datasets with input CNA and somatic mutations in the pathway genes described in text. Bottom: MUTEX identified 3 and 6 groups of alterations showing a trend of mutual exclusivity in these groups with score below the suggested cutoff of 0.2.



#### Formulas input for testing to CAPRI†

† Events type: mutation (m), deletion (d), amplification (a). Hard (⊕) and soft (∨) exclusivity.

 $\ddagger$  Formula not included as it creates a duplicated signature in the dataset.

<span id="page-10-0"></span>Table S3: CAPRI formulas from exclusivity groups. Formulas created for the groups, and input to CAPRI for testing. These are either derived from exclusivity groups or from genes involved in different types of alterations.



<span id="page-11-0"></span>Figure S1: PiCnIc pipeline processing MSI-HIGH/MSS tumors. We process with PiCnIc Microsatellite Stable and highly unstable tumors collected from the The Cancer Genome Atlas project "Human Colon and Rectal Cancer" (subtypes annotations provided as clinical data). We implement a study on selected somatic mutations and focal CNAs in 33 driver genes manually annotated with 5 pathways in the COADREAD project. We scan groups of exclusive alterations with computational tools run by us and by TCGA, and we exploit previous knowledge on CRC; we select which alterations we input to CAPRI. Next, inference is performed with various settings of regularization and confidence. Statistical confidence of the models is assessed with standard techniques from the literature (p-values from statistical testing, bootstrap scores and cross-validation statistics).



<span id="page-12-0"></span>Figure S2: Mutations annotated by TCGA for the driver genes. Top: the majority of the mutations annotated by TCGA for the driver genes that we consider are missense - this in almost all genes and in all the cohort. Bottom: overall summation of the frequencies determine the mutations across all driver genes.



<span id="page-13-0"></span>Figure S3: Lolliplot diagrams of TCGA mutations. Diagrams generated from the cBio portal for the COADREAD project (see <http://www.cbioportal.org/>). These display the physical distribution of the annotated mutations for each gene. Here we shown only genes with a total mutation count greater than 15; FAM123B is called with its synonym AMER1, as in the portal.

## **Exclusivity groups for MSS tumors**



<span id="page-14-0"></span>Figure S4: Groups of exclusive alterations for MSS tumors. Knowledge-based groups of exclusive alterations consist of: KRAS, NRAS and BRAF genes (RAF pathway) and APC and CTNNB1 genes (WNT pathway). The MEMO [\[18\]](#page-23-17) group identified in  $[2]$  in this cohort consists of genes PIK3CA, ERBB2, IGF2 and PTEN. Finally, 6 groups are predicted by MUTEX [\[19\]](#page-23-18) with score below .2, one of these is equivalent to the known exclusive alterations in RAF pathway.



<span id="page-15-0"></span>Figure S5: Selected data for MSS tumors. Colorectal tumors with Microsatellite Stable clinical status in the TCGA COADREAD project, restricted to 152 samples with both somatic mutations and CNA data available. 33 driver genes annotated with 5 pathways are selected from the list published in [\[2\]](#page-23-1) to automatically detect groups of mutually exclusive alterations. Events selected for reconstruction are those involving genes altered in at least 5% of the cases, or part of group of alterations showing an exclusivity trend (see Figure [S4\)](#page-14-0). This dataset is used to infer the set of selective advantage relations which constitute the MSS progression model presented in the Main Text.



<span id="page-16-0"></span>Figure S6: Non-parametric bootstrap scores for MSS progression. Progression model for MSS tumors with confidence shown as edge labels. The first label represents the relation confidence estimated with 100 non-parametric bootstrap iterations, the second and third are p-values for temporal priority and probability raising. Red p-values are above the minimum significane threshold of .05. See Figure 4 in the Main Text for an interpretation of this model.



## **TCGA MSI-HIGH colorectal tumors**

<span id="page-17-0"></span>Figure S7: Non-parametric bootstrap scores for MSI-HIGH progression. Progression model for MSI-HIGH tumors with confidence shown as edge labels. The first label represents the relation confidence estimated with 100 non-parametric bootstrap iterations, the second and third are p-values for temporal priority and probability raising. Red p-values are above the minimum significane threshold of .05. See Figure 5 in the Main Text for an interpretation of this model.







**p-values**  $< 0.01$ 





<span id="page-18-0"></span>

Figure S8: COADREAD statistics for models confidence. For BIC models we show statistics for edges with nonparametric bootstrap score approximately greater than  $30\%$ , for AIC models those greater than  $50\%$ s. *i*) *p-values* (100 repetition of non-parametric bootstrap, prior to Wilcoxon testing) for each edge statistics of selective advantage (direction and statistical dependence, and hypergeometric). In general, the edges that we selected show very strong support ( $p \ll 10^{-10}$ ), but for those edges connecting events with the same marginal frequencies, where we can not be confident in the edge direction ( $p > 0.05$ ) but still we find strong statistical dependence. (*ii*) A posteriori model confidence against Type I and II errors estimated with *non-parametric and statistical bootstraps* (100 repetitions) – edges annotated in Figures [S6](#page-16-0) and [S7.](#page-17-0) *(iii)* Values of posterior classification and prediction errors are estimated from 10 repetitions of 10-fold cross-validation. The former reports how much error is due to predicting, for each set of edges  $X = \{x_1, \ldots, x_n\} \to y$ , the value of y according to the value of each  $x_i \in X$ . The latter reports the same statistics when we predict  $y$  from the whole set of parents  $X$ .



<span id="page-19-0"></span>Figure S9: Entropy loss for MSI-HIGH/MSS models. Violin plot computed from 10 runs of  $k$ -fold cross-validation with  $k = 10$ , where we compute the "loss of log-likelihood" at each fold. In the plot and in the table we report also the overall log-likelihood, as well as the BIC and AIC scores for the models. We present the ratio of log-likelihood loss as a measure of stability of these models for these two datasets – we can observe that the MSS models lose  $< 1\%$  of their likelihood, while the MSI lose slightly more (still,  $<$  4%), possibly because of the smaller sample size. From a statistical point of view, the greater (despite small) loss of likelihood by the models regularized via BIC confirms its tendency to underestimate the true model (i.e., the model should have false negatives, which could be AIC's edges).



**MSS Prediction Error (BIC parent set, k-fold cross-validation)**

**MSS Prediction Error (AIC parent set, k-fold cross-validation)**



<span id="page-20-0"></span>Figure S10: Prediction error for each parent  $2\frac{1}{1}$  et of BIC and AIC models of MSS tumors.



**MSI Prediction Error (BIC parent set, k-fold cross-validation)**

**MSI Prediction Error (AIC parent set, k-fold cross-validation)**



<span id="page-21-0"></span>Figure S11: Prediction error for each parent set of BIC and AIC models of MSI-HIGH tumors.



<span id="page-22-0"></span>Figure S12: Models and the phenotype that they might explain. (biomarker) Independent evolutionary trajectories depicted by a model might share common routes through a certain alteration Y; that could point to a new biomarker harbored by most of the tumors under study. (drug resistance) When a progression model branches in many independent sub-progressions, each one identified by alterations X, Y and Z, if a certain drug is known to target only a certain type of such clones (e.g., those where biomarker X is present), we might get insights on which are the biomarkers which make the drug ineffective for certain patients (e.g., those were cancer evolves through Y and Z). (metastatic) When a model is extracted from data representative of various tumor stages, we might discover which "late events" are those conferring a metastatic phenotype to a tumor  $- X$  in the figures. (anti-hallmarks) Relation between anti-hallmarks and formulas. Exclusivity formulas allow to capture fitness-equivalent events (Y and Z in the figure), and the presence of alternative routes – here those identified by the genotypes X/Y or X/Z. These could point us to genotype/phenotype that we do no observe in our cohort – here the  $X/Y/Z$  – which could be exploited for targeted therapy if a synthetic lethality is screened among Y and Z, the anti-hallmark.

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