## **Supplementary information**

## APOE-ɛ4 and longitudinal decline in olfactory and non-olfactory cognitive outcomes in middle-aged and old adults

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**Table S1**: Parameter estimates from the logistic regression model of the middle-aged participants. We used the complete set of enrolled participants who completed APOE genotyping, non-carriers: n=530,  $\epsilon$  4 heterozygotes: n=223,  $\epsilon$  4/4 homozygotes: n=20. The outcome of the model is whether or not the participant completed the olfactory and cognitive testing at the third test wave.

	Estimate	Std. Error	p-value
Intercept	1.327	0.179	< 0.001
Test experience	0.697	0.230	0.002
Male	0.152	0.208	0.464
Education	0.064	0.125	0.605
MMSE	0.068	0.104	0.517
EMS	0.116	0.122	0.342
SRB	0.201	0.116	0.085
Age	-0.046	0.122	0.703
ε4	0.112	0.227	0.623
ε 4/4	0.155	0.652	0.813

**Table S2**: Parameter estimates from the logistic regression model of the old participants using the complete set of enrolled participants who completed APOE genotyping, non-carriers: n=663,  $\varepsilon$  4 heterozygotes: n=268,  $\varepsilon$  4/4 homozygotes: n=21. The outcome of the model is whether or not the participant completed the olfactory and cognitive testing at the third test wave.

	Estimate	Std. Error	p-value
Intercept	-0.085	0.141	0.545
Test experience	1.307	0.175	< 0.001
Male	-0.009	0.156	0.954
Education	0.007	0.087	0.932
MMSE	0.371	0.093	< 0.001
EMS	0.500	0.106	< 0.001
SRB	0.207	0.098	0.034
Age	-0.683	0.090	< 0.001
ε4	-0.393	0.172	0.022
ε 4/4	-0.497	0.497	0.317

## Sensitivity analysis for unmeasured odor identification at enrollment

Let  $X_e$ , denote the vector of observed covariates and observed non-olfactory cognitive scores at enrollment, W denote the probability of being observed at wave three, and let  $U_e$  denote the unmeasured odor identification score at enrollment. We further let  $X_3$  denote the measured confounders and cognitive scores at the third wave, and similarly let  $U_3$  be the measured odor identification score at the third wave. Let Z denote the vector of APOE-group indicator-variables.

Following McCandless et al. (1), we use the factorization

$$P(W, U_e \mid Z, X_e) = P(W \mid Z, U_e, X_e)P(U_e \mid Z, X_e)$$

and model the confounding effect of U using a logistic regression for the probability of being observed at the third test wave:

$$logit[P(W = 1|Z, U_e, X_e)] = \beta_0 + \beta_1 X_e + \lambda U_e + \alpha Z$$

and a linear regression model for the unobserved odor identification score at enrollment

$$E(U_e|Z, X_e) = \gamma_0 + \gamma_1 X_e + \omega Z$$

As with other models for unmeasured confounding, the model is non-identifiable and the parameters  $\gamma_0$ ,  $\gamma_1$ ,  $\omega$  and  $\lambda$ , reflects assumptions about the confounding effect of  $U_e$ . We first assume  $(U_e|Z, X_e) = (U_3|Z, X_3)$  in distribution, which reflects as missing at random assumption. Hence, the observed data at the third wave can be used to estimate  $\gamma_0$ ,  $\gamma_1$ , and  $\omega$ . Then we can plug in Z and  $X_e$  to predict  $E(U_e|Z, X_e)$ . As a third step, simulation proceeds as follows:

- 1. Sample  $U_e^*$  from  $N(\overline{U}_e, \sigma_{\overline{U}_e})$
- 2. Fit the following logistic regression model  $logit[P(W = 1|Z, U_e^*, X_e)] = \beta_0 + \beta_1 X_e + \lambda U_e^* + \alpha Z$ , where  $\lambda$  is either estimated using glm, alternative 1, or set to a fixed value, alternative 2. For the latter, setting  $\lambda$  equal 0.4, reflecting a moderate effect (a stronger effect than using the glm estimate), which is comparable to the parameter estimates of the other non-olfactory cognitive tasks (see Table S1 and Table S2). Noting that  $U_e^*$  as well as the other cognitive tasks are first standardized using the z-transformation for comparability.
- 3. Estimate the longitudinal models using the simulated weights from step 4.
- 4. Repeat step 1 3 K times.



**Fig S1:** Results from the sensitivity analyses for Middle-aged, first column, and Old participants, second column. Showing boxplots of differences in parameter estimates from the mixed models of Odor identification decline, shown in Table 2 and Table 3 of the main document, and estimates from corresponding mixed models when a simulated odor identification score is included in the IPW models (as described in the simulation procedure above using a moderate effect size (alternative 2)).

## References

1. McCandless, L. C., Gustafson, P., & Levy, A. (2007). Bayesian sensitivity analysis for unmeasured confounding in observational studies. *Statistics in medicine*, *26*(11), 2331-2347.