Campos, Zhao and Charlesworth SI Appendix

1. Approximate BGS equations for a single gene without introns with no gene conversion

Effects of nonsynonymous sites without gene conversion.

First, consider the case with the same heterozygous selection coefficient *t* for each NS site. Let *U* be the deleterious haploid NS mutation rate for the gene in question ($U =$ $0.667n_{ex}l_{ex}$ with the model of nonsynonymous mutations in codons 1 and 2 alone described in the second section of the Materials and Methods), where *u* is the mean mutation rate per bp for the gene). Let *P* and *Q* be the fractions of the gene to the left and right, respectively, of a focal neutral site; *M* is the map length of the gene in Morgans, such that $M = r_c n_{ex} l_{ex}$. In the absence of gene conversion, Eq. 9 of (1) for a continuum of selected sites implies that the BGS effect for a neutral site at position *P* with selection coefficient *t* is:

$$
E(P) \approx \frac{U(t + 2PQ\tilde{M})}{(2t + \tilde{M})} \left\{ \frac{1}{(t + P\tilde{M})} + \frac{1}{(t + Q\tilde{M})} \right\}
$$
(Sla)

where the tilde over *M* denotes $M(1 - t)$.

The mean value of *E*(*P*) for a gene is obtained by integrating Eq. S1A with respect to P over the interval $(0, 1)$, giving:

$$
\overline{E} \approx \frac{U}{(2t+\tilde{M})} \int_{0}^{1} (t+2PQ\tilde{M}) \left\{ \frac{1}{(t+P\tilde{M})} + \frac{1}{(t+Q\tilde{M})} \right\} dP \quad (S1b)
$$

 The first component of the right-hand side of this equation is equivalent to:

$$
\frac{U}{(2t+\tilde{M})}\int_{0}^{1}\frac{\mathrm{d}P}{(t+P\tilde{M})}\mathrm{d}P = \frac{2Ut}{(2t+\tilde{M})\tilde{M}}\Big[\ln\Big(\frac{t+\tilde{M}}{t}\Big)\Big] \tag{S2a}
$$

The second component can be written as:

$$
\frac{4U\tilde{M}}{(2t+\tilde{M})}\int_{0}^{1}\frac{(P-P^2)\,\mathrm{d}P}{(t+P\tilde{M})}\tag{S2b}
$$

We have:

$$
\frac{4U\tilde{M}}{(2t+\tilde{M})}\int_{0}^{1}\frac{P \,dP}{(t+P\tilde{M})} = \frac{4U}{(2t+\tilde{M})}\Big[1-\frac{t}{\tilde{M}}\ln\Big(\frac{t+\tilde{M}}{t}\Big)\Big] \tag{S3}
$$

 \mathbf{r} and

$$
\frac{4U\tilde{M}}{(2t+\tilde{M})}\int_{0}^{1}\frac{P^2\mathrm{d}P}{(t+P\tilde{M})} = \frac{4U}{(2t+\tilde{M})\tilde{M}}\left[\frac{1}{2}(\tilde{M}-2t)+\frac{t^2}{\tilde{M}}\ln\left(\frac{t+\tilde{M}}{t}\right)\right]
$$
(S4)

 Adding the separate terms together yields:

$$
\overline{E} \approx \frac{2U}{\tilde{M}} \left\{ 1 - \frac{t}{\tilde{M}} \ln \left(\frac{t + \tilde{M}}{t} \right) \right\}
$$
 (S5)

When $M(1-t) \le t$, the logarithm can be expanded in powers of $M(1-t)/t$, so that the leading terms in mean *E* are given by:

$$
\overline{E} \approx \frac{U}{t} \left(1 - \frac{2\tilde{M}}{3t} \right) \tag{S6}
$$

When $M(1-t) < t$, this is a decreasing function of *t*, showing that BGS within a single gene can in principle generate a negative relation between the level of selective constraint on the gene and the level of synonymous site diversity.

Approximate net effect of BGS due to nonsynonymous sites for a randomly placed neutral site in a gene without introns. Let the gene be $n_{ex}l_{ex}$ ($>d_g$) basepairs in length. We consider the *k*th site counted from the right end of this sequence. The net effect of BGS contributed by the nonsynonymous sites to the right of this site can be written as:

$$
E(k) = 0.667u \int_{0}^{k} \frac{t \, dx}{\left[t + r(x)(1 - t)\right]^2}
$$
 (S7)

where *x* represents distance in bp, approximated by a continuous line, and $r(x)$ is the net rate of recombination over distance *x*, given by Eq. 2 of the Materials and Methods. Since this is an exponential function of *x*, the integral cannot be evaluated analytically.

A simple model of the effect of gene conversion is to approximate $1 - \exp(-\pi)$ d_{ij}/d_g) by d_{ij}/d_g , which is accurate when $d_{ij} \ll d_g$. This implies that r_{ij} can be replaced by $(r_c + g_c)d_{ij}$, i.e. gene conversion behaves similarly to crossing over, so that *M* can replaced with $(r_c + g_c)n_{ex}l_{ex}$. Alternatively, when $d_{ij} \gg d_g$, the exponent is negligible and the expected value of the rate of gene conversion from Eq. 2 is approximately $g = d_{g}g_{c}$. The product of this term and $1 - t_i$ can simply be added to t_i in the denominator of the basic equation for the effect on BGS of a single selected site *i*, yielding $E_i = u_i t_i / [t_i + (g +$ $r_c d_{ij}$)(1– *t_i*)]².

A better alternative to both of these methods is to use the first approximation when $d_{ij} \leq d_g$ and the second approximation when $d_{ij} > d_g$; this is the 'mixed model' of gene conversion:

$$
\hat{r}(x) = \begin{cases}\n(r_c + g_c)x & \text{if } x \le d_g \\
r_c x + g_c d_g & \text{otherwise}\n\end{cases}
$$
\n(S8)

In both cases, the effect of gene conversion is overestimated, so that the effect of BGS will be underestimated when this 'mixed' approximation is used. Applying these relations to Eq. S7, we obtain:

$$
E(k) \approx 0.667ut \int_{0}^{k} \frac{dx}{[t+r(x)(1-t)]^{2}} = \begin{cases} 0.667ut \frac{1}{(1-t)(r_{c}+g_{c})} \left[\frac{1}{t} - \frac{1}{t+(r_{c}+g_{c})(1-t)k} \right] & k \leq d_{g} \\ \frac{0.667ut}{(1-t)(r_{c}+g_{c})} - \frac{0.667ut}{(1-t)r_{c}} \frac{1}{t+\left[g_{c}d_{g}+r_{c}k\right](1-t)} & k > d_{g} \\ + \frac{0.667utg_{c}}{(1-t)r_{c}(r_{c}+g_{c})} \frac{1}{t+d_{g}(r_{c}+g_{c})(1-t)} \end{cases}
$$
(S9)

We then integrate $E(k)$ from $k = 0$ to $n_{ex}l_{ex}$, which gives its mean over all k values:

$$
E(k) \approx 0.667ut \int_0^{n_{ex} \le x} E(k) dk
$$

=
$$
\frac{0.667ut}{(1-t)(r_c + g_c)} + \frac{0.667g_c(n_{ex}l_{ex} - d_g)ut}{[(1-t)(r_c + g_c)][t + d_g(r_c + g_c)(1-t)]}
$$

-
$$
\frac{0.667ut}{[(1-t)(r_c + g_c)]^2} ln \left[\frac{t + d_g(r_c + g_c)(1-t)}{t} \right] - \frac{0.667ut}{[(1-t)r_c]^2} ln \left[\frac{t + (r_c n_{ex}l_{ex} + g_c d_g)(1-t)}{t + (r_c d_g + g_c d_g)(1-t)} \right]
$$
(S10a)

nexlex

The same procedure can be applied to the effects of nonsynonymous sites to the right of the initial synonymous site, so that the final value is twice the above expression. To take the mean over all sites, we divide the final values by the total number of sites in the gene ($n_{ex}l_{ex}$). Noting that the deleterious mutation rate for the gene is $U = 0.667$ $n_{ex}l_{ex}u$, after some manipulation this yields :

$$
\overline{E} \approx \frac{2U}{\tilde{M}} \left\{ \frac{1}{(1+r)} + \frac{rt(1 - d_g r_c / M)}{(1+r)[t + d_g (r_c + g_g)(1-t)} - \frac{t}{\tilde{M}(1+r)^2} \ln\left[\frac{t + d_g (r_c + g_c)(1-t)}{t}\right] \right\}
$$
\n
$$
-\frac{t}{\tilde{M}} \ln\left[\frac{t + (g_c + M)(1-t)}{t + d_g (r_c + g_c)(1-t)}\right] \right\}
$$
\n(S10b)

where $r = g_c/r_c$. When $g_c = 0$, this expression reduces to Eq. S5.

Contribution of large effect mutations. Data from quantitative genetics analysis of mutational effects on fitness suggest that there is a contribution of deleterious mutations with larger effects than those estimated using DFE-alpha, representing mutational events that are distinct from those included in the rate $U(2)$. It is a reasonable approximation to assume a fixed selection coefficient t_l for such mutations. If the proportion of all mutations that have large effects is p_l , the mutation rate to large mutations is $U_l = p_l U/(1$ p_l). We can then replace *U* and *t* in Eq. S10b with U_l and t_l , and add the resulting terms to Eq. S10b. This addition was used in all the models, unless otherwise stated.

Integral model of the effects of UTR sites in the absence of introns. We now examine the effects of untranslated regions (UTRs) of a gene, which also exert BGS effects on its synonymous sites. Assume that the length of the 5´ UTR is *l*5, while the length of the 3^{\prime} UTR is l_3 . Let u_u and t_u be the mutation and selection parameters for UTRs, corresponding to *u* and *t* for nonsynonymous sites in the previous section. We

assume that all sites in UTRs can exert BGS effects, which maximises the effect of BGS. Using the same procedure as above, the BGS effect contributed by the 3´ UTR for the *l*th site from the right-hand end of an exon is:

$$
E_{3}(l) = \int_{l}^{l+l_{3}} \frac{u_{u}t_{u}}{\left[t_{u} + r(x)(1-t_{u})\right]^{2}} dx
$$
\n
$$
= \begin{cases}\n\frac{u_{u}t_{u}}{\left(1-t_{u}\right)\left(r_{c} + g_{c}\right)} \left[\frac{1}{t_{u} + \left(r_{c} + g_{c}\right)\left(1-t_{u}\right)l} - \frac{1}{t_{u} + \left(r_{c} + g_{c}\right)\left(1-t_{u}\right)\left(1+t_{3}\right)}\right] & 0 \le l < \max\left\{0, d_{s} - l_{3}\right\} \\
\frac{u_{u}t_{u}}{\left(1-t_{u}\right)\left(r_{c} + g_{c}\right)} \frac{1}{t_{u} + \left(r_{c} + g_{c}\right)\left(1-t_{u}\right)l} - \frac{u_{u}t_{u}}{\left(1-t_{u}\right)r_{c}} \frac{1}{t_{u} + \left[g_{c}d_{s} + r_{c}\left(l+t_{3}\right)\right]\left(1-t_{u}\right)} & \max\left\{0, d_{s} - l_{3}\right\} \le l < d_{s} \\
\frac{g_{c}ut_{u}}{\left(1-t_{u}\right)r_{c}\left(r_{c} + g_{c}\right)} \frac{1}{t_{u} + \left(r_{c} + g_{c}\right)\left(1-t_{u}\right)d_{s}} & \max\left\{0, d_{s} - l_{3}\right\} \le l < d_{s} \\
\frac{u_{u}t_{u}}{\left(1-t_{u}\right)r_{c}} \left\{\frac{1}{t_{u} + \left[g_{c}d_{s} + r_{c}l\right]\left(1-t_{u}\right)} - \frac{1}{t_{u} + \left[g_{c}d_{s} + r_{c}\left(l+t_{3}\right)\right]\left(1-t_{u}\right)}\right\} & l \ge d_{s} \\
\end{cases} \tag{S11a}
$$

where the mixed model of gene conversion was used.

Integrating this expression over all synonymous sites, the total BGS effect due to a 3´ UTR is:

$$
E_3 = \int_0^{n_{ex}l_{ex}} E_3(l) \mathrm{d}l
$$

$$
\begin{aligned}\n&\frac{u_{u}t_{u}}{(1-t_{u})^{2}(r_{c}+g_{c})^{2}}\ln\left[\frac{t_{u}+(r_{c}+g_{c})d_{g}(1-t_{u})}{t_{u}}\right]-\frac{u_{u}t_{u}}{(1-t_{u})^{2}r_{c}^{2}}\ln\left\{\frac{t_{u}+[g_{c}d_{g}+r_{c}(n_{ex}t_{ex}+l_{3})](1-t_{u})}{t_{u}+[g_{c}d_{g}+r_{c}l_{3}](1-t_{u})}\right\} \\
&+\frac{u_{u}t_{u}}{(1-t_{u})^{2}r_{c}^{2}}\ln\left\{\frac{t_{u}+[g_{c}d_{g}+r_{c}n_{ex}t_{ex}](1-t_{u})}{t_{u}+(r_{c}+g_{c})d_{g}(1-t_{u})}\right\}+\frac{g_{c}u_{u}t_{u}}{(1-t_{u})r_{c}(r_{c}+g_{c})}\frac{d_{g}}{t_{u}+(r_{c}+g_{c})(1-t_{u})d_{g}} \\
&\frac{u_{u}t_{u}}{(1-tu)^{2}(r_{c}+g_{c})^{2}}\ln\left[\frac{t_{u}+(r_{c}+g_{c})l_{3}(1-t_{u})}{t_{u}}\right]-\frac{u_{u}t_{u}}{(1-t_{u})^{2}r_{c}^{2}}\ln\left[\frac{t_{u}+[g_{c}d_{g}+r_{c}(n_{ex}t_{ex}+l_{3})](1-t_{u})}{t_{u}+[g_{c}d_{g}+r_{c}d_{g}](1-t_{u})}\right] \\
&+\frac{u_{u}t_{u}}{(1-t_{u})^{2}r_{c}^{2}}\ln\left\{\frac{t_{u}+[g_{c}d_{g}+r_{c}n_{ex}t_{ex}](1-t_{u})}{t_{u}+(r_{c}+g_{c})d_{g}(1-t_{u})}\right\}+\frac{g_{c}u_{u}t_{u}}{(1-t_{u})r_{c}(r_{c}+g_{c})}\frac{l_{3}}{t_{u}+(r_{c}+g_{c})(1-t_{u})d_{g}}\n\end{aligned}
$$
\n
$$
(S11b)
$$

A similar expression applies to the BGS effect of a 5´ UTR, with 5 replacing 3.

If $d_g > l_3$, l_5 , as is the case for *Drosophila*, the mean effect of BGS from the two UTRs on a synonymous site is given by:

$$
\frac{E_{3}+E_{5}}{n_{ex}l_{ex}} = \frac{U_{u}t_{u}}{(1-t_{u})^{2}(1+r)^{2}M_{ex}(M_{3}+M_{5})}\ln\left\{\frac{\left[t_{u}+M_{3}(1+r)(1-t_{u})\right]\left[t_{u}+M_{5}(1+r)(1-t_{u})\right]}{t_{u}^{2}}\right\}}{\frac{U_{u}t_{u}}{(1-t_{u})^{2}M_{ex}(M_{3}+M_{5})}\ln\left\{\frac{\left[t_{u}+(g_{c}d_{g}+M_{ex}+M_{3})(1-t_{u})\right]\left[t_{u}+(g_{c}d_{g}+M_{ex}+M_{5})(1-t_{u})\right]}{\left[t_{u}+(g_{c}d_{g}+M_{ex})(1-t_{u})\right]^{2}}\right\}}{\left[t_{u}+(g_{c}d_{g}+M_{ex})(1-t_{u})\right]^{2}}
$$
\n
$$
+\frac{g_{c}U_{u}t_{u}}{(1-t_{u})r_{c}(1+r)M_{ex}\left[t_{u}+(r_{c}+g_{c})(1-t_{u})d_{g}\right]}
$$
\n(S12)

where the total mutation rate for UTRs is $U_u = (l_3 + l_5)u_u$, and the total map lengths for exons and UTRs are $M_{ex} = r_c n_{ex} l_{ex}$, $M_3 = r_c l_3$ and $M_5 = r_c l_5$.

2. Determining the distributions of selection coefficients

To predict BGS effects for an assigned pair of values of β and ω_{na} , as was done to produce the results shown in Figure 2, we employed a modification of the approach of (3) and (4). This uses the level of non-adaptive NS divergence to obtain an estimate of the mean of the DFE, given a value of β . For this purpose, we used the formula of (5) as modified by (6) in their Eq. 23. This expression relates divergence due to the fixation of slightly deleterious mutations (relative to neutral divergence) to the parameters of a gamma distribution of *t*. Using this result, the mean of $\gamma = 4N_e t$ is given by:

$$
\ln(\gamma) \approx [(1+\beta)\ln(\beta) + \ln[\zeta(1+\beta) - \ln(\omega_{na})]/\beta \quad (S13a)
$$

where $\zeta(1+\beta)$ is Riemann's zeta function:

$$
\zeta(1+\beta) = \sum_{i=1}^{\infty} i^{-(1+\beta)}
$$
 (S13b)

This method was also employed when estimating BGS effects from the population genomic data, as described in the next section, because it was found to provide more stable distributions of the bootstrapped values of γ than the direct estimates of ^γ obtained from DFE-alpha, probably reflecting inaccuracies in numerical integration over the gamma distributions. In this case, estimates of the relevant parameters are needed for each bin of K_A values. As described in the Materials and Methods, the DFEalpha program (7) provided estimates of the shape (β) and scale parameters of a gamma

distribution for each bin, as well as estimates of ω_a and ω_{na} . The value of γ for a given bin was then calculated from β and ω_{na} , using Eqs. S13.

For this purpose, K_S values were first corrected for the effect on divergence of selection on codon usage at synonymous sites, using the mean *Fop* value for each bin to estimate the *R* parameter in Eq. S22 in section 5 below, which predicts the ratio of K_S to the corresponding neutral value. We then divided each observed K_S value by R , which yields the value of the neutral divergence corresponding to K_S . Adjusted values of ω_{na} for NS sites were obtained by multiplying the unadjusted values by *R.* A similarly adjusted estimate of the mean of K_S over all bins was used to determine the divergence time in Eqs. S17, which were used to estimate the substitution rates of selectively favorable NS and UTR mutations.

3. Integration of the BGS equations over the distribution of selection coefficients

In order to obtain estimates of the effects of BGS when there is a distribution of selection, nuumerical integration of the BGS equations over an assumed distribution of selection coefficients, for given values of the mutation and recombination parameters, was carried out. For the results shown in Figure 2, we assumed a gamma distribution of the scaled selection coefficient, $\gamma = 4N_e t$, with a fixed scale parameter β ; this has convenient properties and provides a good fit to the data from the Rwandan population of *D. melanogaster* (8). We used a set of fixed values of the non-adaptive divergence parameter ω_{na} , and estimated the corresponding mean values of γ from each pair of ω_{na} and β values by the method described in section 2 above (Eqs. S13). For a given value for $N_e(10^6$ for the *D. melanogaster* population considered here (9)), the probability density of *t* for a given pair of values of β and ω_{na} was obtained

from a gamma distribution.

We removed the portion of the distribution of *t* for which the scaled selection coefficient ($\gamma = 4N_e t$) was less than a threshold value, γ_c , of order 1, since the standard BGS formula is an overestimate when mutations are so weakly selected that drift becomes important (1). The selection coefficient corresponding to γ_c is denoted by t_c . By comparing the simulation results in Table 2 of (1) to their Eq. 9, we found that removal of the portion of the gamma distribution below $4N_et = 5$ led to accurate predictions of the simulated effects of BGS in a finite population. We therefore used $\gamma_c = 5$ for the BGS predictions described in this paper. Use of the more stringent $\gamma_c = 1$ resulted in small quantitative differences to the results, with a slight strengthening of the net BGS effect.

This had little effect on the estimates of the parameters for positive selection, which were estimated as described below.

For the summation method, a value of *t* was drawn from the truncated distribution for each selected site *i* and applied to Eq. 1 of the Materials and Methods. The sum over *i* was then taken to get E_i , using an assumed value of the mutation rate per bp, u . This was done for all values of j , and the mean of E_j over all sites was then determined. This was repeated 500 times to obtain a mean value of E , \overline{E} , for a synonymous site located at a random position in the gene. The BGS predictions using the integral approximation were obtained from the 'mixed model' of gene conversion, described by Eqs. S10b and S12 above, integrating over the assumed truncated gamma distributions for NS and UTR sites, respectively.

4. The effects of selective sweeps on diversity

Consider first the spread of a favorable mutation at nonsynonymous site *i*, which arose in a haplotype carrying a particular allele at a neutral site *j*. Following (10), the net change in the frequency q_i of the neutral allele between the start and end of the sweep in a large population is given by:

$$
\Delta q_j \approx (1 - q_{j0}) q_{i0}^{2r_{ij}/s} \int_{q_{i0}}^1 \left(\frac{[1 - q_{it}]}{q_{it}} \right)^{2r_{ij}/s} dq_{it}
$$
 (S14a)

where q_{i} and q_{i} are the frequencies of the neutral and selectively favorable alleles, respectively in generation t after the initiation of the sweep; r_{ij} is the recombination frequency between the two sites, and *s* is the selection coefficient in favor of homozygotes for the beneficial allele (semidominance and weak selection are assumed). An elementary derivation of this and the following results is given in (11), pp.410-411.

Favorable mutations that arise on the alternative background with respect to the neutral locus contribute a change in allele frequency of the chosen allele at the neutral locus of:

$$
\Delta \tilde{q}_j \approx -q_{j0} q_{i0}^{2r_{ij}/s} \int_{q_{i0}}^{1} \left(\frac{[1 - q_{it}]}{q_{it}} \right)^{2r_{ij}/s} dq_{it} \quad (S14b)
$$

The probabilities of these two alternative events are q_{j0} and $1-q_{j0}$, respectively.

The integral in Eqs. S14a and S14b is an incomplete beta function and cannot be

evaluated analytically, but it is approximately equal to one when $r_{ij} \ll s$, as is required for a significant effect of the sweep on site *j*. We show in section 6 below that this assumption is likely to yield accurate results for cases of biological interest. In order to correct for stochastic losses of mutations arising at initial frequencies of $1/(2N)$, q_{i0} is equated to the ratio of 1/(2*N*) to the fixation probability of the favorable allele, (*Nes/N*), giving $q_{i0} = 1/(2N_e s)$ (12). These two assumptions give the change in allele frequency from Eq. S14a as:

$$
\Delta q_j \approx (1 - q_{j0}) \gamma_a^{-2r_{ij}/s} \tag{S14c}
$$

where $\gamma_a = 2N_e s$ is the scaled selection coefficient for the favorable allele (a factor of 2 not 4 is used, since *s* is the selection coefficient for homozygotes).

The expected reduction in diversity caused by a sweep, relative to the neutral values, can be obtained by averaging over the two possible allelic states at the neutral locus and using Eq. S14c. It is given by:

$$
\frac{\Delta \pi}{\pi_0} \approx -\gamma_a^{-4r_{ij}/s} \tag{S14d}
$$

This equation can also be derived by arguments based on coalescent theory (13, 14). Note that an incorrect version was used by (15), which may affect their inferences concerning the strength of selection.

As noted in (16), if the reasonable assumption is made that the time over which the selectively favorable allele spreads to fixation is much shorter than the expected neutral pairwise coalescent time, 2*N_e*, this relative change in diversity is equivalent to the probability that there was no recombination during the sweep, so that the neutral site experienced a reduction of coalescent time to zero, as opposed what happens when recombination events occurred during the sweep that allow neutral sites with a standard coalescent time to be introduced into haplotypes carrying the favorable allele.

Following (16) and subsequent authors, e.g. (14), if we consider recurrent sweeps occurring over all nonsynonymous sites in a gene, at a rate ^ν*^a* per nonsynonymous site per generation, and assume that v_a is sufficiently small that each sweep exerts its effect independently of the others, the net probability of an effectively instantaneous coalescent event induced by selection at neutral site *j* is given by:

$$
P_{caj} = \nu_a \sum_i \gamma_a^{-4r_{ij}/s} \tag{S15a}
$$

10

A similar expression can be written for the effects of selective sweeps in the UTRs, where for simplicity we assume that the 3' and 5' UTRs have the same selection coefficient, which may of course differ from that for nonsynonymous sites. We can simply replace v_a with a corresponding rate of sweeps per each UTR sites (v_u) , and γ_a with a scaled selection coefficient γ*u*, yielding a rate of selectively induced coalescent events of:

$$
P_{cuj} = \nu_u \sum_{i'} \gamma_u^{-4r_{i'j}/s'} \tag{S15b}
$$

where the primes denote UTR sites and their parameters.

Provided that P_{caj} and P_{cuj} are $<< 1$, the selectively induced coalescent events and standard neutral coalescent events at site *j* can be regarded as competing exponential processes (16), with rates $P_{csj} = P_{caj} + P_{cuj}$ and $1/(2N_e)$, respectively. The net rate of coalescence is then given by the sum of these terms. Taking the reciprocal of this rate, the expected pairwise coalescent time at site, relative to the neutral value in the absence of sweeps, 2*Ne*, is given by:

$$
\frac{T_j}{2N_e} = \frac{1}{1 + 2N_e P_{scj}}
$$
 (S16a)

 The effects of BGS can be included by assuming that it acts independently of sweeps, reducing the effective population size at site *j* from $2N_e$ to $2N_e B_j$, where B_j is given by Eq. 1 of the Materials and Methods or one of the approximations in section 1 above (14, 17, 18). Eq. S16a can then be replaced with:

$$
\frac{T_j}{2N_e} = \frac{1}{B_j^{-1} + 2N_e P_{scj}}
$$
(S16b)

Assuming the infinite sites model of neutral variability (19), the ratio of neutral diversity at a synonymous site $j(\pi_i)$ to its expected value in the absence of selection (π_0) is thus given by:

$$
\frac{\pi_j}{\pi_0} = \frac{1}{B_j^{-1} + 2N_e P_{scj}}
$$
(S16c)

For a two-species comparison, we can write v_a and v_u in equations (S15a) and (S15b) as:

$$
v_a = \alpha_a K_A / (2t_{div})
$$
 (S17a)

$$
\nu_u = \alpha_u K_U / (2t_{div})
$$
 (S17b)

where t_{div} is the divergence time between the two species being compared, K_A and K_U are the nonsynonymous site and UTR site divergences, and α_a and α_u are the corresponding proportions of mutations fixed by positive selection. For substitutions along a single lineage, the factor of 2 is omitted.

Under neutrality, $2ut_{div} = K_S$ for a two species comparison, and $ut_{div} = K_S$ for a single lineage, so that t_{div} can be estimated from estimates of u and K_S (a correction for the effect on K_S of selection on codon usage is described in section 5 below). Under the infinite sites model, neutral nucleotide site diversity is proportional to the coalescent time, so that Eq. S16c can be rearranged to give the deviation between the predicted and observed values of the negative of the logarithm of the ratio of synonymous diversity to its value in the absence of selection. For the two-species comparion, we have:

$$
dev_j = -\ln\left(\frac{\pi}{\pi_0}\right) + \ln\left(\frac{1}{B_j^{-1} + \alpha_a K_A (2t_c)^{-1} \sum_i \gamma_a^{-4r_{ij}/s} + \alpha_u K_U (2t_c)^{-1} \sum_i \gamma_u^{-4r_{i'j}/s'} }\right) \tag{S18}
$$

where t_c is the divergence time in coalescent time units $(t_c = t_{div}/2N_e)$ for the two-species comparison. For a single lineage, the factor of 2 before t_c is omitted.

If we assume that sweeps originate from single new mutations, we can use the standard expression for the rate of substitution of mutations (20) to write $v_a = up_a \gamma_a$ and $v_u = up_u \gamma_a$ where p_a is the proportion of new nonsynonymous mutations that are positively selected, and u is the mutation rate per basepair. Using Eqs. S17, we can then estimate p_a and p_u from the relations:

$$
p_a = \alpha_a K A / (K_s \gamma_a) \tag{S19a}
$$

$$
p_u = \alpha_u K_U / (K_s \gamma_u) \tag{S19b}
$$

In order to apply Eq. S18, it is necessary to eliminate π_0 , the (unknown) diversity in the absence of selection. This was done by adding the smoothed value of $-\ln(\pi_s/\pi_0)$ for the first bin in the set (given by the linear regression of $-\ln(\pi_s/\pi_0)$ on K_A) to each of the *dev_i* values for the other bins, so that the terms in $ln(\pi_0)$ cancel out. The sum of squares *(SSD)* of the resulting quantities was used as a measure of goodness of fit. A 3dimensional grid search was performed across a defined range of values of ^γ*u* and of the intercept and slope for ^γ*a*, in order to obtain estimates that minimise *SSD*. For the point estimates of the parameters, we normally used a grid of 9 values for each variable, with three iterations of the search over successively narrower intervals. The resulting values of γ_a and γ_u were then used in Eqs. S19 to obtain estimates of p_a and p_u , the proportions of NS and UTR mutations that are advantageous. An estimate of π_S/π_0 for a bin was then obtained by substituting the parameter estimates for the bin into the right-hand side of Eq. S16c, providing a measure of the effect of selection at linked sites on synonymous site diversity. We also examined the effect of ignoring UTRs, by conducting similar analyses in which only NS sites were considered; in this case, only a two-dimensional grid search was needed.

We obtained estimates of $ln(\pi_0)$ by using the fact that the predicted value of – ln(π_S) for the first bin in the set is equal to the sum of ln(π_0) and the value of ln(π_S/π_0) for the first bin predicted from Eq. S16c, denoted by P_1 (this bin is expected to have little or no effect of selective sweeps at NS sites). It follows that $ln(\pi_0)$ can be estimated as the sum of P_1 and the observed value of $ln(\pi_S)$ for the first bin. The predicted value of – $ln(\pi_S)$ for any bin other than the first can then be obtained by subtracting this estimate of ln(π_0) from the estimate of – ln(π_S/π_0) for the bin in question. The goodness of fit of the parameter estimates was assessed by calculating the Pearson correlation between observed and predicted values across bins (omitting the first bin).

5. Correcting for the effect on K_s **of weak selection at synonymous sites**

We use the standard Li-Bulmer model of selection on codon usage, assuming a scaled selection coefficient γ for a given gene and a mutational bias κ , given by the ratio of the mutation rate from unpreferred to preferred codons to the mutation rate *u* in the reverse direction (see (11), pp. 274-275 for a description of this model). We assume that the base composition of synonymous sites is at the equilibrium value, *x***,* given by the Li-Bulmer equation, such that the proportion of sites with preferred codons is equal to:

$$
x^* = 1/[1 + \kappa \exp(-\gamma)] \tag{S20}
$$

If a value of κ is assumed, γ can be estimated from the observed proportion of preferred codons (*Fop*) in a gene or set of genes by equating *x** to *Fop.* The equilibrium value of the rate of substitution for sites subject to such selection is given by:

$$
\lambda_s^* = 2\gamma \kappa u / [1 + \kappa \exp(-\gamma)][\exp(\gamma) - 1]
$$
 (S21a)

(See (11), Eq. 6.11.)

The corresponding value for neutral sites with the same base composition, taking into account mutations in both directions, is:

$$
\lambda_n^* = \kappa u [1 + \exp(-\gamma)] / [1 + \kappa \exp(-\gamma)] \tag{S21b}
$$

The ratio of the selected to neutral substitution rates is thus:

$$
R = 2\gamma / [1 + \exp(-\gamma)][\exp(\gamma) - 1]
$$
 (S22)

Estimates of t_{div} obtained using the neutral expectation for K_S should thus be divided by *R.* For the *mel-yak* data, the mean correction was 0.921 and the adjusted value of t_{div} was 3.39 x 10⁷ generations; for *mel*, the corresponding values were 0.920 and 1.32 x 10⁷ generations. These yield estimates of the mean rates of adaptive substitutions (v_a) for NS sites of 3.21×10^{-10} and 2.96×10^{-10} for *mel-yak* and *mel*, respectively. The corresponding rates for UTR substitutions (v_u) were 8.16 x 10⁻¹⁰ and 8.78 x 10⁻¹⁰.

6. Approximating the incomplete beta function

The expected change in diversity over the two possible trajectories described by Eqs. S14a and S14b, using the approximation $q_{i0} = 1/\gamma_a$, is :

$$
\frac{\Delta \pi}{\pi_0} \approx -\gamma_a^{-4r_{ij}/s} S_i^2 \tag{S23a}
$$

where S_i is the incomplete beta function:

$$
S_{i} = \int_{q_{io}}^{1} \left(\frac{[1 - q_{it}]}{q_{it}} \right)^{2r_{ij}/s} dq_{it}
$$
 (S23b)

This leads to the following generalisation of Eq. S14d:

$$
P_{caj} \approx v_a \sum_i \gamma_a^{-4r_{ij}/s} S_i^2 \tag{S24}
$$

In the main text, we presented results that assume that S_i can be replaced with 1 without great loss of accuracy. We investigated the validity of this assumption by replacing S_i with 1 in Eq. S24, using the 'standard' gene model with 5 exons of 300 basepairs separated by 4 introns 100bp in length, and comparing the results with those obtained from Eqs. S23. The results are shown in Figure S8, where S_1 in the caption refers to the exact value of the mean of the sum in Eq. S24 over all synonymous sites, and *S*₂ to the value when *S_i* =1, for a range of values of γ_a . It is clear that the agreement is very close.

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7. SI Tables

Table S1A. Multiple linear regression coefficients for ^π *^S* **for** *mel – yak*

Table S1B. Multiple linear regression coefficients for π **_S for** *mel*

KS, synonymous divergence*;* Rec., effective rate of crossing over; Expression, gene expression level across all developmental stages of *D. melanogaster;* K_{A} nonsynonymous divergence*; Fop,* frequency of optimal codons in a gene*; P, P* value of a *t*-test. See the first section of the Materials and Methods for details of the data and variables used here.

Table S2A Population genomic statistics for NS sites for each bin of *KA* **values for**

mel-yak

Table S2B. Population genomic statistics for NS sites for each bin of *KA* **values for** *mel*

Bin *N* K_A π_S β ω_a ω_{na} Rec. CDS length ^α *× KA Fop* 96 0.00085 0.0157 0.577 0.0061 0.0109 1.139 678.3 0.00031 0.543 104 0.00105 0.0149 0.452 0.0027 0.0178 1.167 629.2 0.00014 0.545 128 0.00125 0.0156 0.470 0.0069 0.0165 1.142 611.0 0.00037 0.548 118 0.00145 0.0157 0.494 0.0125 0.0156 1.190 533.8 0.00065 0.556 139 0.00165 0.0148 0.448 0.0093 0.0224 1.166 590.0 0.00049 0.550 133 0.00185 0.0151 0.553 0.0160 0.0200 1.228 575.2 0.00084 0.543 131 0.00204 0.0140 0.458 0.0151 0.0244 1.100 627.1 0.00077 0.521 125 0.00224 0.0144 0.424 0.0162 0.0267 1.189 552.7 0.00082 0.534 122 0.00245 0.0150 0.469 0.0203 0.0278 1.082 612.2 0.00104 0.530 121 0.00264 0.0169 0.496 0.0245 0.0246 1.231 576.3 0.00132 0.534 122 0.00285 0.0147 0.431 0.0242 0.0282 1.085 625.1 0.00131 0.536 142 0.00305 0.0141 0.367 0.0237 0.0358 1.150 662.1 0.00120 0.529 112 0.00326 0.0147 0.532 0.0341 0.0274 1.175 536.1 0.00181 0.527 125 0.00345 0.0147 0.394 0.0237 0.0423 1.268 557.9 0.00126 0.526 130 0.00366 0.0145 0.393 0.0294 0.0376 1.172 535.2 0.00160 0.521 108 0.00385 0.0149 0.483 0.0418 0.0313 1.142 527.3 0.00218 0.527 107 0.00405 0.0150 0.403 0.0407 0.0382 1.252 564.1 0.00208 0.529 104 0.00425 0.0147 0.445 0.0346 0.0391 1.128 505.5 0.00197 0.530 101 0.00445 0.0138 0.384 0.0422 0.0428 1.068 560.1 0.00221 0.527 105 0.00465 0.0165 0.384 0.0397 0.0432 1.221 652.2 0.00221 0.515 106 0.00485 0.0136 0.343 0.0360 0.0536 1.086 588.4 0.00195 0.527 97 0.00504 0.0137 0.367 0.0518 0.0469 1.139 549.7 0.00266 0.521 97 0.00524 0.0131 0.476 0.0542 0.0454 1.037 631.7 0.00281 0.516 75 0.00544 0.0157 0.430 0.0522 0.0465 1.158 530.6 0.00288 0.513 111 0.00570 0.0136 0.318 0.0425 0.0608 1.148 505.7 0.00233 0.505 116 0.00600 0.0144 0.335 0.0545 0.0605 1.122 571.2 0.00281 0.502 122 0.00630 0.0136 0.362 0.0544 0.0598 1.167 555.0 0.00291 0.510 106 0.00659 0.0145 0.266 0.0316 0.0843 1.132 511.4 0.00178 0.506 105 0.00690 0.0130 0.331 0.0582 0.0716 1.006 572.0 0.00304 0.496 97 0.00721 0.0139 0.318 0.0519 0.0777 1.120 481.9 0.00289 0.516 107 0.00750 0.0129 0.255 0.0513 0.0873 1.069 600.1 0.00272 0.490 94 0.00779 0.0142 0.374 0.0711 0.0684 1.172 475.3 0.00396 0.500 100 0.00812 0.0128 0.259 0.0604 0.0853 1.091 433.1 0.00334 0.512 99 0.00856 0.0123 0.391 0.0789 0.0718 1.181 509.1 0.00439 0.500 117 0.00895 0.0128 0.371 0.0856 0.0817 1.102 552.4 0.00445 0.486 91 0.00938 0.0132 0.201 0.0636 0.1161 0.961 582.6 0.00332 0.495 104 0.00989 0.0132 0.275 0.0853 0.0932 1.126 488.1 0.00456 0.487 95 0.01038 0.0123 0.259 0.0816 0.1024 1.104 468.5 0.00456 0.484 103 0.01094 0.0126 0.291 0.1005 0.0969 1.083 401.2 0.00552 0.484 95 0.01155 0.0136 0.262 0.0994 0.1074 1.015 440.8 0.00554 0.484 95 0.01226 0.0129 0.334 0.1073 0.1047 1.118 426.1 0.00617 0.477 110 0.01293 0.0115 0.265 0.1261 0.1103 1.088 429.6 0.00687 0.462 99 0.01378 0.0141 0.249 0.1124 0.1208 1.065 358.0 0.00667 0.480 95 0.01469 0.0126 0.281 0.1417 0.1228 1.176 414.4 0.00790 0.478 101 0.01577 0.0125 0.182 0.1034 0.1626 1.127 373.8 0.00598 0.454 93 0.01694 0.0124 0.233 0.1290 0.1567 1.121 473.8 0.00760 0.433 107 0.01845 0.0127 0.323 0.1910 0.1452 1.186 374.4 0.01047 0.445 96 0.02065 0.0120 0.246 0.1674 0.1886 1.165 343.0 0.00975 0.451 90 0.02324 0.0118 0.164 0.2136 0.1794 1.153 424.2 0.01239 0.428 101 0.02712 0.0104 0.206 0.2795 0.2198 1.155 507.9 0.01501 0.424 *N*: number of genes in the bin; K_A : mean nonsynonymous divergence for genes in the bin from the *mel-yak* or *mel* data; π_s , mean synonymous diversity; β , shape parameter of the distribution of fitness effects (DFE); ω_a , the rate of adaptive substitutions for nonsynonymous mutations relative to the neutral rate; ^ω*na*, the rate of non-adaptive substitutions (neutral or slightly deleterious) relative to the neutral rate; Rec*.*, mean smoothed effective crossing over rates from fits of Loess regressions to the crossing over rates along each chromosome (measured in cM per Mb) in *D. melanogaster*; CDS length, coding sequence length in number of aminoacids; ^α *× KA*: the mean proportion of adaptive substitutions multiplied by non-synonymous divergence; *Fop*: mean codon usage bias, measured as the frequency of optimal codons. See the first section of the Materials and Methods for further details of the sources of these data.

Table S3A. UTR population genomic statistics for each bin of *KA* **values for** *mel-yak*

	5' UTR						3' UTR					
Bin	Ν	K_A	$\kappa_{\rm \nu}$	length	β	α	Ν	K_A	$\kappa_{\rm \nu}$	length	β	α
1	102	0.0019	0.0830	263	0.469	0.662	102	0.0019	0.0602	518	0.506	0.649
\overline{c}	163	0.0031	0.0852	230	0.394	0.646	166	0.0031	0.0723	39	0.366	0.591
3	150	0.0044	0.0888	257	0.270	0.484	153	0.0044	0.0679	424	0.357	0.502
4	177	0.0056	0.0904	245	0.421	0.665	178	0.0056	0.0778	395	0.349	0.619
5	167	0.0069	0.0892	252	0.402	0.626	171	0.0068	0.0709	435	0.460	0.656
$\,6$	181	0.0081	0.0846	262	0.422	0.575	186	0.0081	0.0782	354	0.367	0.576
$\boldsymbol{7}$	158	0.0093	0.0930	225	0.344	0.575	161	0.0093	0.0829	383	0.304	0.495
$\bf 8$	183	0.0106	0.1067	224	0.394	0.603	184	0.0106	0.0872	347	0.306	0.553
9	158	0.0119	0.0938	238	0.344	0.589	163	0.0119	0.0838	319	0.307	0.538
10	189	0.0131	0.0945	215	0.395	0.650	185	0.0131	0.0806	346	0.371	0.615
11	158	0.0144	0.0932	216	0.390	0.586	160	0.0144	0.0862	309	0.346	0.566
12	170	0.0156	0.0907	227	0.276	0.466	174	0.0156	0.0838	328	0.290	0.529
13	133	0.0169	0.0944	228	0.383	0.628	130	0.0169	0.0932	320	0.275	0.534
14	152	0.0181	0.0923	204	0.394	0.554	157	0.0180	0.0962	305	0.271	0.496
15	133	0.0193	0.0919	201	0.437	0.620	132	0.0193	0.1095	251	0.495	0.648
16	150	0.0205	0.1003	200	0.347	0.528	149	0.0205	0.0996	279	0.295	0.470
17	125	0.0219	0.1037	229	0.285	0.552	131	0.0219	0.0895	340	0.338	0.501
18	121	0.0231	0.1116	225	0.248	0.460	126	0.0231	0.1011	295.	0.248	0.466
19	115	0.0243	0.0984	177	0.168	0.317	111	0.0243	0.1143	226	0.401	0.655
20	118	0.0256	0.1053	161	0.370	0.574	115	0.0256	0.1192	230	0.350	0.555
21	107	0.0268	0.1095	231	0.319	0.575	112	0.0268	0.1025	293	0.333	0.628
22	105	0.0282	0.1017	199	0.357	0.575	107	0.0282	0.1129	244	0.404	0.506
23	83	0.0293	0.0991	186	0.358	0.478	85	0.0293	0.1175	233	0.206	0.411
24	98	0.0306	0.1074	171	0.261	0.438	103	0.0305	0.1140	261	0.597	0.679
25	73	0.0318	0.1022	186	0.408	0.585	72	0.0318	0.1233	230	0.436	0.656
26	74	0.0331	0.1214	206	0.333	0.488	77	0.0331	0.1182	240	0.171	0.421
27	64	0.0344	0.1072	131	0.289	0.469	59	0.0344	0.1133	204	0.303	0.576
28	50	0.0354	0.0992	186	0.398	0.490	49	0.0354	0.1156	259	0.259	0.580
29	100	0.0369	0.1073	144	0.360	0.556	100	0.0369	0.1114	198	0.216	0.473
30	93	0.0389	0.1218	166	0.434	0.713	100	0.0389	0.1239	258	0.357	0.585
31	97	0.0409	0.1152	148	0.714	0.761	98	0.0409	0.1154	204	0.436	0.586
32	91	0.0429	0.1147	165	0.376	0.578	88	0.0429	0.1357	226	0.306	0.562
33	80	0.0449	0.1129	110	0.147	0.401	84	0.0449	0.1153	166	0.369	0.634
34	82	0.0469	0.1124	148.	0.767	0.787	83	0.0469	0.1279	189	0.394	0.605
35	73	0.0489	0.1026	134	0.370	0.560	73	0.0489	0.1308	238	0.215	0.344
36	110	0.0520	0.1245	118	0.185	0.281	113	0.0520	0.1311	162	0.140	0.263
37 38	95 86	0.0558	0.1082	128	0.245	0.432	98	0.0558 0.0601	0.1363 0.1350	204	0.222 0.218	0.413
39	104	0.0601 0.0639	0.1168	88	2.277	0.871 0.583	83 102	0.0639	0.1417	164 267	0.550	0.478
40	75	0.0679	0.1125	136 81	0.463 0.260	0.364	69	0.0680	0.1454	130	0.445	0.610 0.689
41	71	0.0723	0.1139 0.1460	107	0.190	0.521	74	0.0724	0.1470	147	0.255	0.464
42	68	0.0772	0.1319	135	0.168	0.247	69	0.0772	0.1680	170	0.143	0.370
43	78	0.0841	0.1261	125	1.298	0.821	79	0.0840	0.1404	182	0.277	0.580
44	68	0.0914	0.1269	98	0.152	0.385	65	0.0914	0.1828	167	0.411	0.707
45	69	0.0988	0.1116	97	0.179	0.175	65	0.0990	0.1414	155	0.262	0.489
46	72	0.1079	0.1387	141	0.249	0.454	72	0.1081	0.1509	211	0.050	0.240
47	62	0.1202	0.1204	110	0.142	0.342	56	0.1200	0.1648	184	0.819	0.751
48	65	0.1347	0.1474	107.	0.127	0.387	63	0.1348	0.1638	183	0.248	0.542
49	57	0.1509	0.1436	83	0.447	0.500	53	0.1505	0.1670	154	0.374	0.447
50	33	0.1833	0.1422	116	0.799	0.613	34	0.1826	0.1554	142	0.259	0.196

5' UTR 3' UTR Bin *N KA KU* length β ^α *N KA KU* length β ^α 93 0.0008 0.0199 274 0.513 0.686 94 0.0009 0.0148 408 0.475 0.607 92 0.0011 0.0192 289 0.402 0.535 96 0.0011 0.0172 412 0.256 0.400 120 0.0013 0.0180 245 0.408 0.625 122 0.0013 0.0186 399 0.419 0.645 107 0.0015 0.0199 253 0.334 0.606 111 0.0015 0.0172 334 0.339 0.626 126 0.0016 0.0177 214 0.280 0.460 129 0.0016 0.0167 380 0.383 0.622 126 0.0018 0.0166 248 0.312 0.479 122 0.0018 0.0188 341 0.280 0.541 121 0.0020 0.0168 249 0.419 0.553 121 0.0020 0.0156 362. 0.351 0.523 112 0.0022 0.0196 250 0.390 0.540 115 0.0022 0.0196 340 0.365 0.603 115 0.0024 0.0187 233 0.394 0.628 114 0.0024 0.0208 339 0.282 0.586 114 0.0026 0.0186 250 0.561 0.737 117 0.0026 0.0197 328 0.388 0.637 110 0.0029 0.0204 229 0.269 0.511 112 0.0029 0.0187 303 0.239 0.494 129 0.0030 0.0181 198 0.344 0.518 131 0.0030 0.0199 361 0.353 0.658 105 0.0033 0.0180 212 0.384 0.581 104 0.0033 0.0207 385 0.327 0.546 112 0.0035 0.0200 212 0.471 0.650 113 0.0035 0.0169 293 0.450 0.589 116 0.0037 0.0182 211 0.311 0.466 114 0.0037 0.0217 316 0.263 0.482 100 0.0038 0.0254 239 0.147 0.398 99 0.0038 0.0176 346 0.291 0.433 94 0.0041 0.0178 198 0.528 0.655 96 0.0041 0.0180 282 0.405 0.674 93 0.0043 0.0203 178 0.478 0.526 92 0.0043 0.0212 321 0.378 0.491 96 0.0045 0.0228 219 0.305 0.531 94 0.0045 0.0198 336 0.349 0.584 94 0.0046 0.0224 169 0.289 0.488 93 0.0046 0.0198 271. 0.348 0.561 94 0.0049 0.0204 223 0.434 0.579 94 0.0049 0.0207 277 0.292 0.517 89 0.0050 0.0177 174 0.226 0.385 94 0.0050 0.0214 230 0.367 0.595 86 0.0052 0.0220 195 1.244 0.806 86 0.0052 0.0215 294 0.360 0.583 61 0.0054 0.0177 149 0.556 0.614 61 0.0054 0.0231 230 0.336 0.608 104 0.0057 0.0224 172 0.463 0.658 105 0.0057 0.0232 220 0.381 0.607 101 0.0060 0.0218 170 0.321 0.596 105 0.0060 0.0182 243 0.643 0.655 106 0.0063 0.0209 201 0.544 0.709 109 0.0063 0.0220 276 0.398 0.529 87 0.0066 0.0228 162 0.236 0.388 91 0.0066 0.0222 226 0.295 0.524 91 0.0069 0.0197 156 0.498 0.581 88 0.0069 0.0339 291 0.275 0.526 85 0.0072 0.0238 183 0.268 0.574 86 0.0072 0.0235 209 0.263 0.494 93 0.0075 0.0222 186 0.195 0.433 91 0.0075 0.0231 224 0.342 0.635 80 0.0078 0.0206 174 0.318 0.537 82 0.0078 0.0231 216 0.233 0.454 90 0.0081 0.0225 178 0.229 0.446 89 0.0081 0.0209 180 0.198 0.471 90 0.0086 0.0243 144. 0.272 0.485 89 0.0086 0.0252 245 0.360 0.546 101 0.0090 0.0219 128 0.452 0.613 99 0.0090 0.0220 171 0.774 0.708 78 0.0094 0.0182 131 0.263 0.414 78 0.0094 0.0208 203 0.246 0.497 92 0.0099 0.0201 151 0.362 0.563 86 0.0099 0.0283 185 0.446 0.655 84 0.0104 0.0248 151 0.152 0.379 77 0.0104 0.0222 244 0.199 0.415 85 0.0109 0.0219 142 0.475 0.668 89 0.0109 0.0279 149 0.277 0.627 80 0.0116 0.0207 121 1.140 0.780 78 0.0116 0.0216 168 0.486 0.569 76 0.0123 0.0240 71 0.273 0.484 77 0.0123 0.0208 191 0.406 0.636 88 0.0129 0.0224 109 0.159 0.237 86 0.0129 0.0238 143 0.261 0.506 71 0.0138 0.0251 84 0.246 0.555 71 0.0138 0.0289 156 0.289 0.544 74 0.0147 0.0249 138 0.149 0.374 75 0.0147 0.0307 168 0.435 0.656 76 0.0158 0.0193 85 0.396 0.654 72 0.0157 0.0262 128 0.259 0.519 68 0.0170 0.0297 105 0.423 0.663 69 0.0169 0.0312 192 0.074 0.346 75 0.0184 0.0312 86 1.436 0.862 72 0.0184 0.0254 160 0.349 0.643 71 0.0207 0.0184 77 3.791 0.841 67 0.0206 0.0273 152 0.418 0.671 60 0.0233 0.0255 123 0.164 0.517 52 0.0234 0.0227 197 0.197 0.385 67 0.0270 0.0263 112 0.348 0.543 66 0.0269 0.0297 144 0.781 0.581

Table S3B. UTR Population genomic statistics for each bin of *KA* **values for** *mel*

N: number of genes in the bin; K_A : mean nonsynonymous divergence for the genes in the bin for the *mel-yak* or *mel* data; K_U : mean UTR divergence; length: UTR length in bp; $β$, shape parameter of the distribution of fitness effects (DFE); α*,* proportion of adaptive substitutions*.* See the first section of the Materials and Methods for further details of the sources of these data.

The entries in each cell are the mean *E* values (as percentages), obtained using the summation method. The left-hand entries are the effects of NS sites alone; the right-hand entries are the net effects of NS and UTR sites. Four 100bp introns are present. The short, standard and long exons have 50, 100 and 200 codons, respectively. The mutation rate per bp is 4.5 x 10^{-9} and the rate of crossing over per bp is 1 x 10⁻⁸. The low rates of gene conversion have g_c = 1.0 x 10⁻⁸ and $d_g = 440$; the high rates have $g_c = 5.0 \times 10^{-8}$ and $d_g = 500$. The shape parameter of the gamma distribution of fitness effects is 0.3.

ω na	Predicted	Observed							
Low gene conversion rate									
0.025	4.10; 4.10	3.66; 4.21							
0.050	5.16; 5.80	5.33; 5.98							
0.075	6.64; 7.37	6.87; 7.61							
0.100	7.92; 8.76	8.20; 9.04							
0.125	9.02; 9.94	9.31; 10.2							
0.150	9.86; 10.9	10.2; 11.2							
High gene conversion rate									
0.025	1.83; 1.98	1.89; 2.06							
0.050	2.75; 2.90	2.86; 3.04							
0.075	3.54; 3.71	3.69; 3.89							
0.100	4.22; 4.38	4.38; 4.57							
0.125	4.71; 4.89	4.87; 5.08							
0.150	5.06; 5.24	5.21; 5.42							

 T**able S5. Tests of the linearity of the effect of gene length on BGS strength**

The columns labelled 'Predicted ' are the values for genes with five 300bp exons; the 'Observed' are the (unweighted) means over genes with 150, 300 and 600bp exons. The left-hand entries were obtained using the summation method with 5 exons separated by 100bp introns; the right-hand entries were obtained from the integral approximation with the mixed model of gene conversion. The other parameters are as in Table S3.

Table S6 Effects of intron length on BGS effects

The entries in each cell are the mean *E* values (percentages) obtained by the summation method. The left-hand entries are the effects of NS sites alone; the right-hand entries are the net effects of NS and UTR sites.

The exon lengths are 100 codons when introns are present, and 500 codons in the absence of introns, so that the total exon length is fixed. The lengths of the short, standard and long introns are 50bp, 100bp and 200bp, respectively. The mutation rate per bp is 4.5×10^{-9} and the rate of crossing over per basepair is 1 x 10⁻⁸. The low rates of gene conversion have g_c = 1.0 x 10⁻⁸ and d_g = 350; the high rates have g_c = 5.0 x 10⁻⁸ and d_g = 500.

Table S7. BGS effects for the *mel-yak* **data with different mutation and recombination rates**

The entries in each cell are the regression coefficients of mean *E* for a bin on *KA*, with their standard errors. The low, standard and high rates of mutation per bp are 3.0 x 10⁻⁹, 4.5 x 10⁻⁹ and 6.0 x 10⁻⁹, respectively. The low, standard and high rates of crossing over per bp are 0.5 x 10⁻⁸, 1.0 x 10^{-8} and 2.0 x 10^{-8} , respectively.

Table S8. Effects of rates of mutation and crossing over on estimates of the parameters of positive selection and synonymous site diversity

The *mel-yak* data used for Figure 3 were analysed for different values of the mutation and crossing over rates, using the 'standard' gene model. The low and high rates of crossing over per bp were one-half and twice the standard value of 1 x 10^{-8} , respectively; the low and high mutation rates per bp were 3 x 10⁻⁹ and 6 x 10⁻⁹, respectively, compared with the standard value of 4.5 x 10⁻⁹. π _{r max} is the maximum value of the mean synonymous site diversity of a gene relative to its value in the absence of selection; ^π *r mean* is the corresponding mean value over bins. Other variables are defined in the text.

Table S9. Bootstrapped *mel* **estimates of parameters of positive selection and synonymous site diversity**

The data used for Figure S2 were analysed, with 250 independent bootstraps for each bin of *KA* values, performed as described in the Materials and Methods. The entries show the means and (in brackets) the upper and lower 2.5 percentiles of the bootstrap distributions of the relevant parameter estimates.

8. SI Figures

Figure S1. Synonymous site diversity (π_S) for a gene plotted against the estimated number of NS substitutions per site along the *D. melanogaster* lineage since divergence from its common ancestor with *D. simulans.* ρ is the Spearman rank correlation coefficient. The green line is the least squares linear regression; the dashed black lines represent its 95% confidence interval.

Figure S2. Plots of predicted and observed value of – $ln(\pi_S)$ for each bin of K_A values for the *mel* data.

> The black diamonds are the observed values of $- \ln(\pi_s)$ for each bin of *KA* values for autosomes, corrected for the correlation between ^π*^S* and K_S as described in the first section of the Materials and Methods. The circles are the theoretical values of mean *E* for each bin, obtained by the integral model of BGS, assuming a single gene with 500 NS sites. The crosses are the predicted values of $- \ln(\pi_s)$ for each bin, given by the combined BGS and SSW models at NS and UTR sites. Red and blue correspond to the low and high gene conversion rates used in Figure 2. The mutation rate and crossing over parameters are as in Figure 2, except that large effect mutations constitute 15% of all mutations, with a selection coefficient against heterozygotes of 0.044.

Figure S3. Mean scaled selection coefficients (γ) for NS sites for each bin of *KA* values; red diamonds are *mel-yak* data, and blue triangles are *mel* data. For clarity of display, some extreme outlier values for *mel-yak* are not shown: bin 5 (γ = 8.7 x 10³), bin 30 (γ = 1.05 x 10⁴), and bin 32 (γ = 3.04 x 10⁴).

Figure S4. Mean scaled selection coefficients (γ) for UTR sites for each bin of *KA* values; red diamonds are *mel-yak* data, and blue triangles are

Figure S5. Shape parameters (β) for NS sites for each bin of K_A values; red diamonds are *mel-yak* data, and blue triangles are *mel* data*.*

Figure S6. Plots of the lengths of the UTRs against K_A for the binned *mel-yak* data. The red curves are the quadratic least squares fit for the 5´UTRs, and the quartic least-squares fit for the 3´UTRs.

Figure S7. Plots of the lengths of the UTRs against *KA* for the binned *mel* data. The red curves are the quadratic least squares fit for the 5´UTRs and 3´UTRs.

Figure S8. Plots of the mean over all synonymous sites in a gene of the predicted effects of a single selective sweep at each NS site on synonymous site diversity as a function of the scaled selection coefficient γ_a , given by the summation term in Eq. S24. S_1 is the exact result for the sum, and S_2 is the approximation with $S_i = 1$ used in the data analyses. N_e = 10⁶; the recombination parameters are $r_c = g_c = 10^{-8}$, and $d_g = 440$; there are 5 exons of 300 bp, separated by 4 introns of 100bp (the standard gene model).