

# Goodness of fit tools for dose–response meta-analysis of binary outcomes

Supporting Information

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The aim of this document is to prove that one- and two-stage approaches for fixed-effects dose–response meta-analysis are equivalent.

## 1 Two-stage dose–response-meta analysis

In the first stage, the  $i$ -th study-specific model ( $i = 1, \dots, K$ ) is defined as

$$y_i = X_i \beta_i + \varepsilon_i \tag{1}$$

where:

- $y_i$  is the  $(n_i \times 1)$  vector of non-referent log relative risks;
- $X_i$  is the  $(n_i \times p)$  design matrix containing the non-referent values of the dose and/or some nonlinear transformations;
- $\beta_i$  is the  $(p \times 1)$  vector of the  $i$ -th study-specific dose–response coefficients;
- $\varepsilon_i$  is the  $(n_i \times 1)$  vector of the error term, such that  $\varepsilon_i \sim N(0, S_i)$  and  $S_i$  is considered to be known.

The study-specific vector  $\beta_i$  and matrix  $V(\beta_i)$  are estimated using the Generalized Least Squares (GLS) estimator

$$\hat{\beta}_i = \left( X_i^\top S_i^{-1} X_i \right)^{-1} X_i^\top S_i^{-1} y_i \quad (2)$$

$$V \left( \hat{\beta}_i \right) = \left( X_i^\top S_i^{-1} X_i \right)^{-1} \quad (3)$$

In the second stage,  $\hat{\beta}_i$  are pooled using a multivariate fixed-effects meta-analysis

$$\hat{\beta}_i \sim N_p \left( \theta, V \left( \hat{\beta}_i \right) \right) \quad (4)$$

The pooled estimate  $\hat{\theta}$  and the corresponding (co)variance matrix  $V(\theta)$  are estimated using the GLS estimator

$$\hat{\theta} = \left( \sum_{i=1}^K \hat{V} \left( \hat{\beta}_i \right)^{-1} \right)^{-1} \sum_{i=1}^K \hat{V} \left( \hat{\beta}_i \right)^{-1} \hat{\beta}_i \quad (5)$$

$$\hat{V} \left( \hat{\theta} \right) = \left( \sum_{i=1}^K \hat{V} \left( \hat{\beta}_i \right)^{-1} \right)^{-1} \quad (6)$$

Substituting  $\hat{V} \left( \hat{\beta}_i \right)$  and  $\hat{\beta}_i$  in Equations 5 and 6 with the expressions in Equations 2 and 3 gives the following estimates for  $\hat{\theta}$  and  $V(\theta)$

$$\hat{\theta} = \left( \sum_{i=1}^K X_i^\top S_i^{-1} X_i \right)^{-1} \sum_{i=1}^K X_i^\top S_i^{-1} y_i \quad (7)$$

$$\hat{V} \left( \hat{\theta} \right) = \left( \sum_{i=1}^K X_i^\top S_i^{-1} X_i \right)^{-1} \quad (8)$$

## 2 One-stage dose–response meta-analysis

The one-stage model is defined as

$$y = X\gamma + \varepsilon \quad (9)$$

where:

- $y = (y_1^\top, \dots, y_K^\top)^\top$  is the  $(\sum_{i=1}^K n_i = n \times 1)$  vector of concatenated study-specific non-referent log relative risks;
- $X = (X_1^\top, \dots, X_K^\top)^\top$  is the  $(\sum_{i=1}^K n_i = n \times p)$  matrix of concatenated study-specific design matrixes;
- $\varepsilon = (\varepsilon_1^\top, \dots, \varepsilon_K^\top)^\top$  is the  $(\sum_{i=1}^K n_i = n \times 1)$  vector of concatenated study-specific error terms, such that  $\varepsilon \sim N(0, S)$  and

$$S = \text{diag}(S_i) = \begin{bmatrix} S_1 & \mathbf{0} & \dots & \mathbf{0} \\ \mathbf{0} & S_2 & \dots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0} & \mathbf{0} & \dots & S_K \end{bmatrix} \quad (10)$$

The vector  $\gamma$  and the matrix  $V(\gamma)$  are estimated using the GLS estimator

$$\hat{\gamma} = (X^\top S^{-1} X)^{-1} X^\top S^{-1} y \quad (11)$$

$$\hat{V}(\hat{\gamma}) = (X^\top S^{-1} X)^{-1} \quad (12)$$

Given the block-diagonal structure of  $S$  we can rewrite Equations 11 and 12 as follows

$$\hat{\gamma} = (X^\top S^{-1} X)^{-1} X^\top S^{-1} y = \quad (13)$$

$$= \left( \sum_{i=1}^K X_i^\top S_i^{-1} X_i \right)^{-1} \sum_{i=1}^K X_i^\top S_i^{-1} y_i$$

$$\hat{V}(\hat{\gamma}) = (X^\top S^{-1} X)^{-1} = \left( \sum_{i=1}^K X_i^\top S_i^{-1} X_i \right)^{-1} \quad (14)$$

### 3 Conclusion

Equations 7 and 8 are equal to Equations 13 and 14, respectively. This proves that  $\hat{\theta} = \hat{\gamma}$  and  $\hat{V}(\hat{\theta}) = \hat{V}(\hat{\gamma})$ , and therefore the equivalence of one- and two-stage dose-response meta-analytical approaches.