

1 **Electronic Supplementary Material**

2 **Physiological thermal limits predict differential responses of bees to urban heat-island**

3 **effects**

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24 **I. Supplementary methods**

25 *Heat tolerance assay.* CT_{max} is an ecologically relevant measure of heat tolerance that
26 represents the point when an organism can no longer escape a stressful thermal environment [1].
27 Across ectotherms, it is also correlated with the harder-to-measure thermal optimum; in other
28 words, species with high CT_{max} tend to have high thermal optima, such that CT_{max} should predict
29 fitness changes even in response to sublethal temperatures [2].

30 To measure CT_{max} , we collected bees from university property and residential yards
31 within 3.2 km of our laboratory, such that transit time from field to lab was no more than 10 min.
32 Collecting sites were distinct from those used in the field population survey. We spent 15 – 35
33 min netting bees, which we transported individually in 50 ml plastic tubes in an insulated
34 container. In the lab, we transferred bees to weighted 45 ml glass vials in a circulating water bath
35 constructed from a 13.25 L tub equipped with a recirculation pump (Rio Plus 50 Aqua Pump,
36 TAAM, Inc., Camarillo, CA, USA) and a heating element (120V, Camco, Greensboro, NC,
37 USA) controlled by a JLD612 dual display PID temperature controller and PT100 temperature
38 sensor probe (Lightobject, Sacramento, CA, USA). Vial openings, plugged with cotton,
39 remained above water. Following an initial period of 20 min at 25 °C, the bath warmed at a rate
40 of 0.5 °C min⁻¹. The bath accommodated seven bees at a time, and we kept an additional one to
41 seven bees per assay in a separate, unheated bath to ensure that experimental conditions other
42 than heat were not lethal. To monitor temperatures experienced by bees, an iButton (DS1923,
43 Maxim Integrated, San Jose, CA, USA) recorded temperature and humidity every minute
44 throughout the assay within identical glass vials in the heated and control baths. We recorded the
45 time at which each bee fell over and was unable to right itself within 30s. CT_{max} for each
46 individual bee was the iButton temperature recorded during the minute that the individual lost

47 postural control. Because not all bees could be identified prior to the CT_{\max} assay, species
48 composition varied across trials; we present data for the 15 species that were common enough to
49 yield at least 3 measurements per species.

50 ***Field sampling.*** Site selection and sampling are described in detail elsewhere (Hamblin
51 et al. in review). Briefly, we used Landsat-derived thermal map of Raleigh [3] to identify 15
52 residential yards and 3 urban natural areas that varied in warming intensity. Yards were 405 to
53 7487 m²; natural areas were larger, but we limited sampling to 4047 m² per site. Residential
54 neighborhoods were developed between the 1930s and 2003, such that local heat-island patterns
55 were established at least a decade prior to sampling. All sites were more than 2 km apart to
56 ensure independence of bee samples [4]. In 2015 we measured air temperature at each site using
57 a pair of thermochron iButtons (DS1921, Maxim Integrated). Despite shielding iButtons as
58 described by Hubbart [5], daytime readings were compromised by solar radiation. We therefore
59 used only early evening temperatures (7-9pm) to compute a mean summer evening temperature
60 at each site. Although this temperature does not represent conditions experienced by foraging
61 bees, it captures the urban heat island effect without interference from solar radiation, accurately
62 arraying sites on an axis from cooler to warmer [6].

63 We sampled bees from May to August, visiting each site 6 times in 2014 and 5 times in
64 2015. At each site on each date we deployed 12 pan traps and 1 blue vane trap (SpringStar,
65 Seattle, WA), and conducted 20 min of netting. Pan traps were 3.25 oz soufflé cups (Solo Cup
66 Co., Urbana, IL), left white or painted fluorescent yellow or blue (Guerra Paint & Pigment, New
67 York, NY). On each sampling date, traps were out for 5 to 7 hours between 8:00 and 17:30 and
68 we netted between 10:00 and 17:00. The full bee sample, described elsewhere, included 113
69 species, most of which were rare; here we focus on the 15 common species for which we also

70 measured CT_{max} . For these 15 species, the total number of individuals collected was 1732 ($n = 11$
71 $- 549$ per species). In all downstream analyses, the abundance of a given species at a given site
72 refers to the total number of individuals of that species collected over the two years of sampling.

73 ***Hierarchical model.*** To estimate the rate of change in abundance of each species relative
74 to temperature, we constructed a hierarchical model analogous to a Poisson regression with log
75 link function: $\log(\mu_{ij}) = a_i + r_i * t_j$ where μ_{ij} is the predicted count (a Poisson mean) for the i th
76 species at the j th site; a_i is the intercept for species i , r_i is the Poisson regression coefficient for
77 species i , and t_j is the temperature at site j . The coefficient r_i represents proportional change in
78 abundance of a species per $^{\circ}C$ increase in temperature, and we refer to it as “response to
79 warming” throughout this study. (Specifically, for each $1^{\circ}C$ increase in temperature, μ is
80 multiplied by e^{r_i} .) We further specified r_i as arising from a normal distribution with mean β and
81 variance σ^2 , so that responses of all species were estimated jointly relative to the overall, species-
82 wide response, β . This approach allowed information from the entire dataset to inform the
83 estimates for each species, stabilizing estimates for rarer taxa [7,8]. We fit the model in
84 WinBUGS 1.4, assuming uninformative prior distributions for a , β , and σ^2 . We used 3 Markov
85 chain simulations, each independently initialized and computed for 21000 draws. After
86 inspecting diagnostic plots for convergence and autocorrelation, we discarded the first 1000
87 draws from each chain and thinned to every tenth draw for a final sample of 6000 draws, which
88 we used to compute estimates, standard errors, and 95% credible intervals for each species’
89 response to warming (r_i).

90 ***Phylogenetic inference.*** We downloaded the 20 gene dataset generated by Hedtke et al.
91 [9], and selected one outgroup (*Dasypoda hirtipes*) and the lineages of interest for our study.
92 Eleven of the 20 genes contained less than 10% of sequence data for our ingroup, thus we

93 removed those genes from our dataset. We included the following nine genes in our study:
94 arginine kinase (AK), calcium/calmodulin-dependent protein kinase II (CAD), elongation factor
95 1- α copy 1 and 2 (EF1a1, EF1a2), sodium potassium adenosine triphosphate (NAD),
96 phosphoenolpyruvate carboxykinase (PEPCK), long wavelength rhodopsin (Opsin), RNA
97 polymerase II (PolII), and wingless (Wg). We added sequence data from the mitochondrial gene
98 cytochrome oxidase I (COI) to improve species level resolution within the genera *Bombus* and
99 *Megachile*. COI sequences for all species were retrieved from the National Center for
100 Biotechnology Information (NCBI) or the BOLD System (table S1). We used sequences from
101 the phylogenetically closest species when sequences from the focal species were not available.
102 We aligned COI nucleotide sequences using the MUSCLE option in SeaView v.4.5.3 [10] and
103 concatenated to nine-gene dataset [9]. Phylogenetic reconstruction was based on a total of 10,252
104 nucleotides in our final alignment. Alignments were visually inspected and improved in
105 Mesquite 3.0.3. Maximum likelihood trees were searched under the GTRGAMMA model of
106 sequence evolution in RAxML using 1000 bootstrap replicates after partition optimization using
107 the python script PartitionFinder v1.1.1 [11]. Trees were visually inspected and annotated in
108 FigTree v.1.4.2 (<http://tree.bio.ed.ac.uk/>).

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116 Table S1. COI sequences and genbank accession numbers. These sequences were concatenated
 117 to 9 genes from Hedtke et al. [9].

Species in this study	Species in Hedtke et al.	Species for COI	COI Genbank Accession No.
<i>Agapostemon virescens</i>	<i>Agapostemon kohliellus</i>	<i>Agapostemon virescens</i>	JQ266376
<i>Bombus bimaculatus</i>	<i>Bombus bimaculatus</i>	<i>Bombus bimaculatus</i>	KM585629
<i>Bombus griseocollis</i>	<i>Bombus griseocollis</i>	<i>Bombus griseocollis</i>	SMTPL8163*
<i>Bombus impatiens</i>	<i>Bombus wilmattae</i>	<i>Bombus impatiens</i>	JF799030
<i>Ceratina calcarata</i>	<i>Ceratina calcarata</i>	<i>Ceratina calcarata</i>	KJ166268
<i>Ceratina strenua</i>	<i>Ceratina cyanea</i>	<i>Ceratina strenua</i>	KJ163420
<i>Halictus ligatus</i>	<i>Halictus ligatus</i>	<i>Halictus ligatus</i>	AF102840
<i>Lasioglossum bruneri</i>	<i>Lasioglossum cressonii</i>	<i>Lasioglossum bruneri</i>	JF903499
<i>Lasioglossum imitatum</i>	<i>Lasioglossum imitatum</i>	<i>Lasioglossum imitatum</i>	AF103967
<i>Megachile campanulae</i>	<i>Megachile angelarum</i>	<i>Megachile cetuncularis</i>	FJ582307
<i>Megachile exilis</i>	<i>Megachile angelarum</i>	<i>Megachile versicolor</i>	KJ836926
<i>Megachile mendica</i>	<i>Megachile patellimana</i>	<i>Megachile mendica</i>	KF839683
<i>Megachile rotundata</i>	<i>Megachile texana</i>	<i>Megachile rotundata</i>	GU706002
<i>Ptilothrix bombiformis</i>	<i>Ptilothrix sp. JS 2010</i>	<i>Ptilothrix bombiformis</i>	AF300562
<i>Xylocopa virginica</i>	<i>Xylocopa virginica</i>	<i>Xylocopa virginica</i>	EU271670

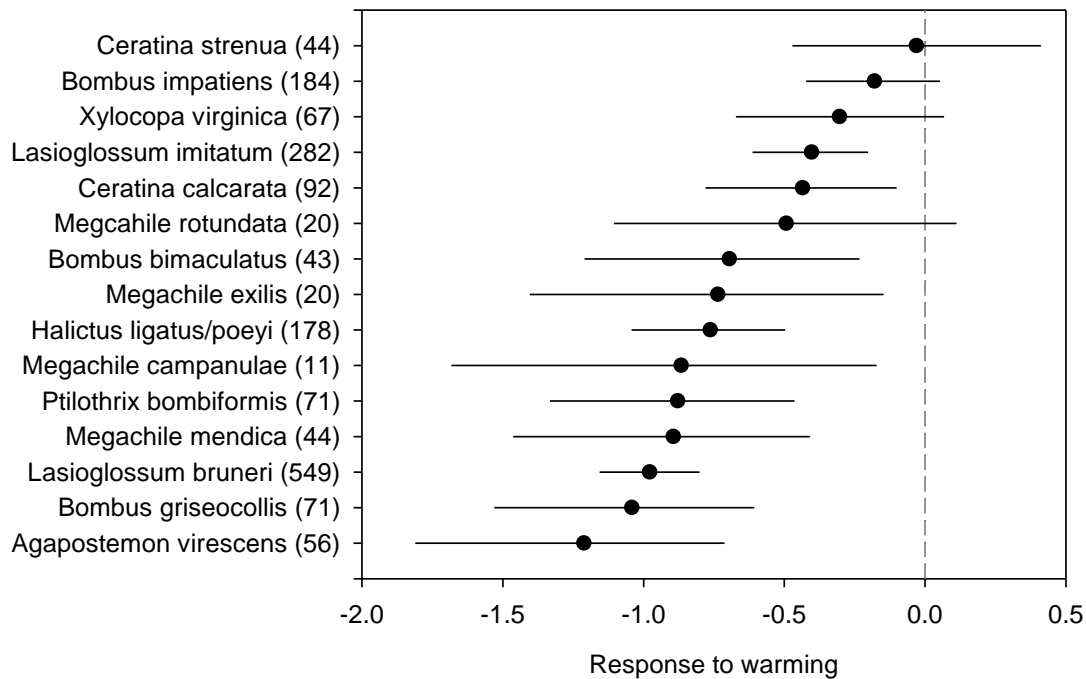
118 * Retrieved from BOLD system

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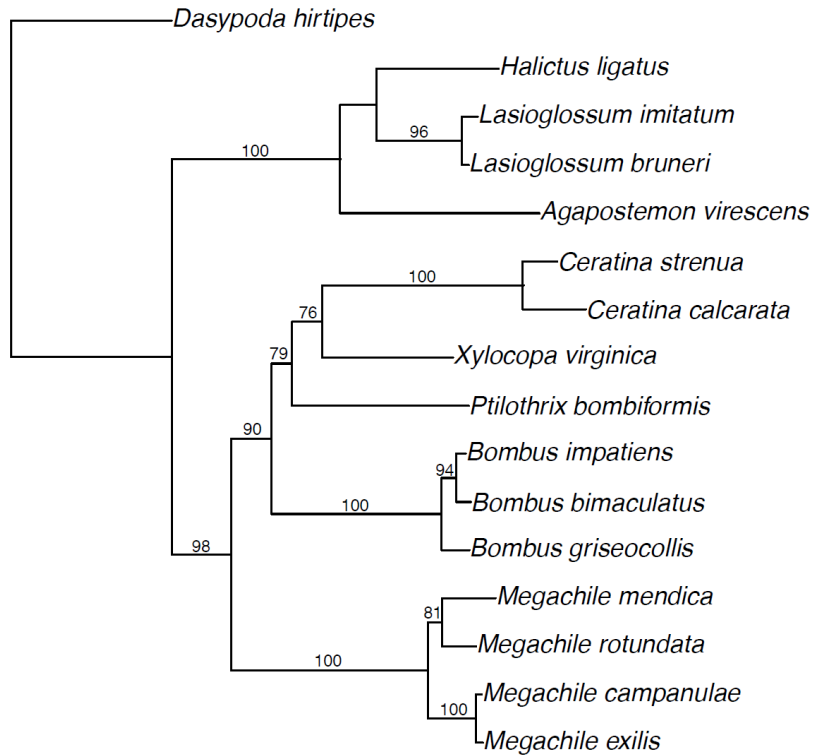
121 **II. Supplementary results**

122 **Hierarchical model of species responses to warming.** The assemblage-wide mean response to
123 warming was significantly negative with $\beta = -0.66$ (95% CI -0.96 to -0.40). Within this overall
124 trend, species varied in their rates of decline; four species' responses were indistinguishable from
125 zero (as indicated by 95% credible intervals, fig. S1).



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127 Figure S1. Results of hierarchical model estimating Poisson regression coefficients for each bee
128 species (“response to warming” as rate of population change per °C). Error bars are 95% credible
129 intervals; numbers in parentheses are total number of individuals collected.

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131 **Phylogeny reconstruction.** We reconstructed the phylogenetic relationships of 21 bee
132 species from four families using information from one mitochondrial and nine nuclear genes (fig.
133 S2). The maximum likelihood phylogenetic reconstruction recovered all bee families with high
134 bootstrap support, and the tree topology is congruent with Hedtke et al.



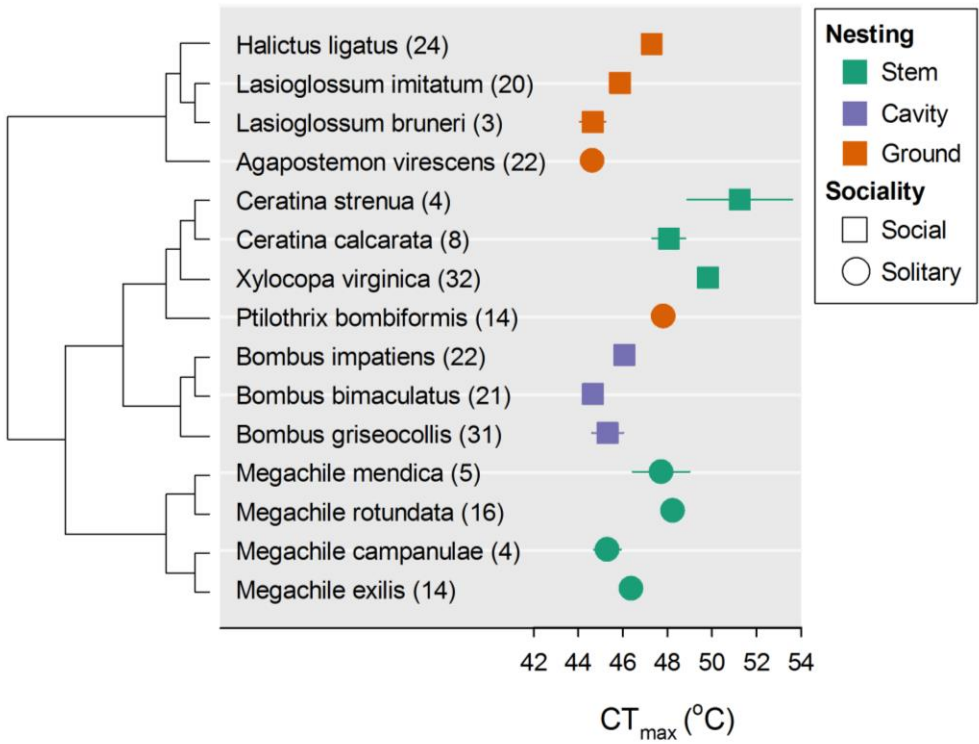
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137 Figure S2. Phylogenetic reconstruction of 15 focal bee species; numbers on branches are
 138 bootstrap values after 10000 replicates.

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140 **Trait evolution.** The relationship between phylogeny, CT_{max} , nesting behavior, and
 141 sociality is shown in fig. S3. We did not detect strong phylogenetic signal in CT_{max} ; that is,
 142 models of trait evolution that incorporated the phylogeny were not a better fit to the data than
 143 was a phylogenetically independent white-noise model ($\Delta AICc < 2$, table S2). Body size, nesting
 144 habitat, and social behavior were all phylogenetically correlated; Brownian motion and Early
 145 Burst were the best-supported models of trait evolution (tables S2-S3).



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147 Figure S3. Bee traits considered in this study, as they relate to the phylogeny (numbers in
 148 parentheses are sample sizes for CT_{max} assay).

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150 Table S2. Models describing trait evolution of CT_{max} and body size across the bee phylogeny.

151 Best-fitting models ($\Delta AICc < 2$) for each trait are bolded.

Model	CT_{max}			Body size		
	AICc	$\Delta AICc$	Parameters	AICc	$\Delta AICc$	Parameters
Brownian motion	67.37	0.99	$\sigma^2 = 23.59$	50.29	0.00	$\sigma^2 = 7.56$
Ornstein-Uhlenbeck	66.38	0.00	$\sigma^2 = 46.48, \alpha = 5.76$	53.36	3.07	$\sigma^2 = 8.93, \alpha = 0.82$
White noise	66.95	0.57	$\sigma^2 = 3.64, \mu = 46.87$	60.59	10.30	$\sigma^2 = 2.38, \mu = 2.70,$
Pagel's λ	66.66	0.28	$\sigma^2 = 9.69, \lambda = 0.78$	53.47	3.18	$\sigma^2 = 7.56, \lambda = 1.00$
Early burst	70.55	4.17	$\sigma^2 = 23.59, a = 1.00 \times 10^{-6}$	53.28	2.99	$\sigma^2 = 15.03, a = -2.01$

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155 Table S3. Models describing trait evolution of discrete traits, nesting habitat and sociality, across the bee phylogeny. Best-fitting
 156 models ($\Delta\text{AICc} < 2$) for each trait are bolded.

Model	Nesting				Sociality				
	AICc	ΔAICc	Parameters	Rate ¹	AICc	ΔAICc	Parameters	Rate ²	
Equal-rate models									
Brownian motion	19.98	2.42	NA	0.68	15.21	0.00	NA	1.48	
White noise	36.31	18.80	NA	NA	22.50	7.29	NA	NA	
Pagel's λ	22.67	5.11	$\lambda = 1.00$	0.68	17.90	2.69	$\lambda = 1.00$	1.48	
Early burst	17.56	0.00	a = -28.06	978.17	15.84	0.63	a = -20.15	978.17	

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162 **Phylogenetic generalized least-squares models.** Regardless of the pattern of evolution of
 163 individual traits, the residuals of models relating those traits to one another (as in our two focal
 164 hypotheses) can also show phylogenetic signal. We addressed this possibility by comparing
 165 several models that differed only in their phylogenetic covariance structure. In the main text, we
 166 present the results of the best-fitting models; in Table S4 we provide fit details for the alternative
 167 models.

168
 169 Table S4. Comparison of generalized least-squares models testing two hypotheses with different
 170 phylogenetic correlation structures.

Model	Phylogenetic correlation structure	d.f.	Log likelihood	AICc	ΔAIC_c	Model weight
CT _{max} ~ Body size + Nest + Sociality	Pagel's λ	7	-0.3	30.6	0.0	1.0
	None	6	-25.1	72.6	42.0	0.0
	Ornstein-Uhlenbeck	7	-24.5	79.0	48.4	0.0
	Brownian motion	6	-30.1	82.7	52.1	0.0
Response to warming ~ CT _{max}	None	3	-0.3	8.7	0.0	0.8
	Ornstein-Uhlenbeck	4	0.05	11.9	3.2	0.2
	Brownian motion	3	-6.6	21.3	12.62	0.0
	Pagel's λ		did not converge*			

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 172 *Although the Pagel's λ model did not converge using the gls function in nlme package in R, and
 173 therefore is not included for direct comparison with the other gls models, we examined a
 174 comparable model with Pagel correlation structure using the pgl function in the caper package,
 175 which produced a maximum-likelihood estimate of $\lambda = 0$ (95% CI 0 to 0.594). We also manually
 176 fit different values of λ in the gls function, and likelihood was maximized at $\lambda = 0$. These checks
 177 support the conclusion that there is not strong phylogenetic signal in the relation between
 178 response to warming and CT_{max}.

179 **Weighted regression.** To examine the possibility that error in the estimation of species
180 responses to warming influenced our analysis of the relationship between CT_{max} and response to
181 warming, we further examined the non-phylogenetic model using weighted regression. We
182 weighted each response to warming by the inverse of its standard error (as estimated in the
183 hierarchical model), thereby reducing the influence of the least certain estimates. We fit the
184 weighted regression using the `lm` function in the `stats` package of R. Results (table S5) were
185 comparable to those of the non-weighted regression shown in the main text.

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187 Table S5. Results of weighted regression describing each species' response to warming as a
188 function of its CT_{max} (where each response is weighted by the inverse of its standard error).

Term	Coefficient	s.e.	<i>t</i>	<i>p</i>
Intercept	-5.52	1.86	-3.0	0.011
CT_{max}	0.10	0.040	2.6	0.021

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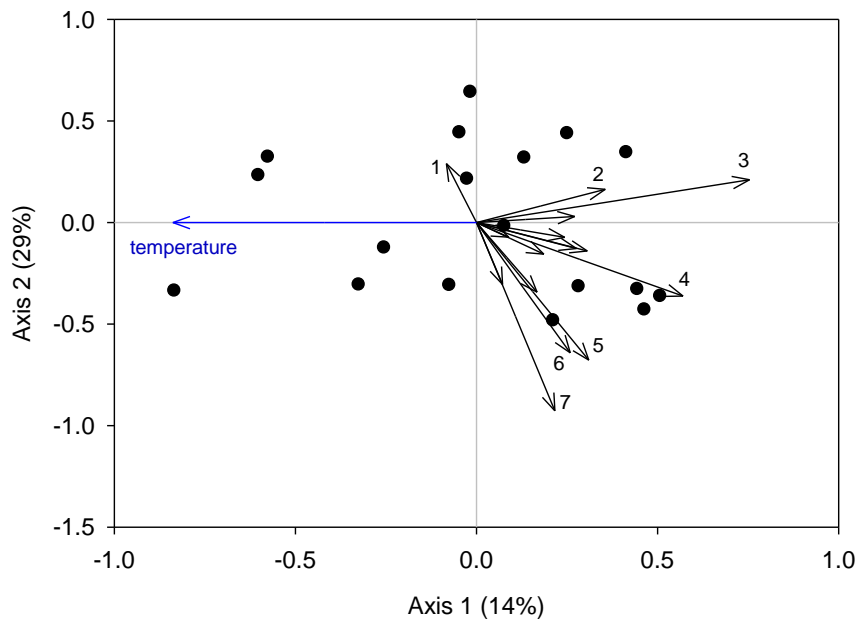
200 **III. Community Analysis**

201 The results of the hierarchical model demonstrate that some bee species declined faster than
202 others in response to warming. These differential rates of change imply that species' relative
203 abundances, and thus community composition, also shifted with warming. Here, we explicitly
204 test for an effect of temperature on bee community composition across sites by performing a
205 distance-based redundancy analysis (db-RDA) using the *vegan* package in R [12]. This
206 constrained ordination analysis assumes a linear relation between predictors (here, temperature)
207 and the multivariate response (bee abundances), and tests significance with a permutation test
208 (without assuming multivariate normality) [13,14]. In this analysis, we $\log(x+1)$ transformed bee
209 counts to meet the assumption of linearity, and used db-RDA with Bray-Curtis distances because
210 the Euclidean distances of standard RDA are misleading when the data include many zeros [13].

211 The db-RDA detected significant effects of site temperature on community composition
212 across sites, with 14% of the inertia in the dataset represented on the db-RDA axis that aligns
213 with temperature ($p = 0.015$, fig. S4). Species' loadings on the temperature axis could be
214 analyzed as an alternate quantification of "response to warming" that makes the connection
215 between CT_{\max} and community ordination more explicit. For ease of interpretation, however, we
216 focused our analyses on the Poisson regression coefficients, for which error is more readily
217 estimated and which represent rate of change in biologically meaningful units.

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221 Figure S4. db-RDA ordination indicates that bee community composition varied across sites as a
 222 function of temperature. Black points represent the positions of the 18 study sites in ordination
 223 space, as they relate to temperature (blue arrow) and bee species abundances (black arrows).

224 Arrows indicate the direction of increase of each variable; the plot was generated using scaling =
 225 3 to optimize display of sites and species. For clarity, only selected species are labeled. They are
 226 1, *Ceratina strenua*; 2, *Ceratina calcarata*; 3, *Lasioglossum bruneri*; 4, *Bombus griseocollis*; 5,
 227 *Agapostemon virescens*; 6, *Megachile mendica*; 7, *Halictus ligatus/poeyi*.

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