Supplementary Information for Gibson et al. (2017) Color naming across languages reflects color use.

Supp. Materials, Methods, Analysis and Figures (SI-Section 1 to SI-Section 10; Figures S1- S16; Tables S1-S6)

Data collection with the Tsimane' was performed through daily trips to eight Tsimane' communities near San Borja, Bolivia, in collaboration with the Centro Boliviano de Investigación y de Desarrollo Socio Integral (CBIDSI).

SI-Section 1: Additional details of the Color-naming Task

The variability of the tasks that were run under the World Color Survey. All previous experiments in which participants from unindustrialized cultures were asked to label colors have used variants of the World Color Survey (WCS) instructions (1-3). These instructions introduce a complex notion of a "basic" color term, which takes several pages to describe. In writing these instructions, the authors of the WCS were trying to prohibit participants from producing lowfrequency color terms like "scarlet" as a sub-class of red, or terms that are associated with only one object. The notion of "basic" color category does not include categories that are subsets of others, and can be applied broadly to many objects. But the concept of a basic color term has theoretical problems, because it is not clear that color categories cannot be parts of others, or that color categories cannot be very narrow; moreover, many languages simply do not have a superordinate concept of "color". Thus identifying "basic" color terms across languages begs the question of what counts as a basic color category (4).(4). The definition is also problematic in practice because it is so complex, making the notions difficult to explain, with the likely consequence that different WCS researchers implemented the complex instructions differently. An empirical evaluation of the WCS data suggests that there was variability in the kind of strategy that was used by WCS experimenters in implementing this task. The range of strategies can be captured by two extreme versions of the task: one in which participants could say whatever color words that they wanted $-$ a "free-choice" version $-$ and a second variant in which participants were restricted to choose a color word from a fixed set of choices – a "fixed-choice" version. For example, the fixed-choice version of the task was explicit when gathering the Pirahã WCS data, as discussed by Everett (5). Among the Pirahã queried in the WCS, all 25 participants except one produced all and only the same set of four words (one participant also used one additional word, in 5 trials); this outcome is extremely unlikely if the participants were not constrained to use a particular set of terms. We can compare Pirahã to the six other WCS languages which also have four modal color words. Two of these languages are like Pirahã, such that only the same four or five terms were provided by all of the participants. But participants in the other WCS languages with four modal color words produced more color terms: 15-17 terms in each of these four languages (sampling 25 people in each language). This corroborates the idea that WCS researchers may have used two versions of the task: a fixed-choice version (where only 4 words are used by all participants in these languages) and a free-choice version, with no such constraint, and the result that participants are much more variable in what they produce.

We quantified the variability in how the WCS task was implemented using two analyses. First, we examined the ratio of the total number of words that any participant used in a WCS language to the number of modal color terms. If a particular WCS task was implemented with a set of fixed choices for that language, this ratio will be close to one. But if there were fewer constraints on what words participants could use, then this ratio will result in a number larger than one. The histogram of the WCS ratios in **Figure S1** shows that many languages have a term-to-modalterm ratio of exactly one, suggesting a fixed-choice task in those languages. Some languages have a ratio very close to one, suggesting that some constraints were placed on what might be said in those languages. And many languages have much higher ratios, suggesting that no constraints were applied in these languages.

Figure S1. A histogram of the ratio of the number of words that any participant used in a WCS language to the number of modal color terms in that language. In this analysis, we restricted our attention to the subset of 80 color chips that were used in our experiments, in order to compare our results to those from the WCS. A ratio close to one suggests that the WCS task was implemented with a set of fixed choices for that language. Ratios that are much larger than one suggest that the WCS task was implemented with free choice of color terms for that language. We include the Tsimane' fixed-choice and free-choice ratios as baselines. For the bootstrap comparisons in the text, we compare only to the 99 WCS languages that have at least 20 participants. We randomly selected data from 20 Tsimane' subjects, and only include terms that appeared more than once (Tsimane' free choice $= 18$ total terms / 8 modal terms $= 2.25$).

What is the probability that we would observe each of the ratios in **Figure S1** if the task given to participants was to label colors freely? To answer this question, we calculated a distribution over term-to-modal-term ratios based on bootstrap resampling our Tsimane' free-choice data (see **Table S1**) for the 99 WCS languages that have at least 20 participants. This distribution tells us

the probability that we would observe a certain term-to-modal-term ratio given randomly sampled subjects and a free-choice task. Most of the languages in the WCS dataset (80/99) have a term-to-modal-term ratio significantly less than the Tsimane' free-choice task, suggesting that these data were not collected with a fully free choice task. The data from the other 19 languages (those marked with "FALSE" in column 3 in **Table S1**) were plausibly generated with a fully free-choice task. Finally, seven of the 99 languages had term-to-modal ratios of exactly 1, suggesting that they were plausibly generated using the fixed-choice task.

Table S1. The term-to-modal-term for each of the 99 WCS languages with at least 20 participants, along with whether each ratio is significantly smaller than the ratio generated from samples of 20 participants in the Tsimane' free-choice task, at $p < 0.01$. When the ratio is significantly smaller, it provides evidence suggesting that the data from that language were not gathered using a fully free-choice task. The data from the other 19 languages (those marked with "FALSE" in column 3) were plausibly generated with a fully free-choice task.

Second, we examined the mean *color-word-overlap proportion* (CWO proportion) for the WCS languages, where the CWO proportion is defined as the mean proportion of color terms that each participant used which were also used by more than three-quarters of the other participants. A larger average CWO proportion for a language indicates a greater likelihood that words were constrained in the task. For example, 6 of the WCS languages have mean CWO proportions of 1.0, meaning that *every* term that a participant used was used by at least 75% of the other participants. Forty of the WCS languages have a CWO proportion of .9 or higher, suggesting a constrained vocabulary of color terms across participants, with few outlier terms. In contrast, there are 17 languages in the WCS with mean CWO proportions of 0.7 or below, meaning that 30% or more of the color terms that participants used in these languages were used by fewer than 75% of other participants. In these languages, there were probably no constraints on what

speakers were told to say by their experimenters. Taken together, these two analyses suggest that the specific methods used to implement the WCS task were likely variable from one language to another.

Figure S2. A histogram of the mean *color-word-overlap proportion* (CWO proportion) for the WCS languages, where the CWO proportion is defined as the mean proportion of color terms that each participant used which were also used by more than three-quarters of the other participants. The non-normality of this distribution suggests that different tasks were used across different WCS languages: a free-choice version and a fixed-choice version. A proportion close to one suggests that the WCS task was implemented with a set of fixed choices for that language. Proportions much less than one suggest that the WCS task was implemented with free choice of color terms for that language. We include the Tsimane' fixed-choice and free-choice WCO proportions as baselines.

Instructions for the current study. We used two versions of a color-naming task: a free-choice version, in which participants were simply asked to label Munsell chips spanning the color space in a way that they thought others from their community would also label them; and a fixedchoice version, in which the instructions were identical to the free-choice version, but participants were also asked to choose from a fixed set of 8 choices (the modal labels from the free-choice version). In pilot experiments on 12 Tsimane' participants, we collected colorlabeling data on the 160 chips of the standard Munsell array (6); subsequent participants were tested using a subset of 80 chips, sampling the array uniformly (the 80-chip array produced the same results as the 160-chip array, but took half the time for data collection on each participant).

We provide a list of the Munsell chip designations for the chips we used in the **Table S2**. Each color chip was affixed to a white cardboard square 2 inches on a side.

Participants were presented with the 80 chips in a different random order for each participant under controlled lighting conditions using a light box. Color-naming variability measured in studies that do not control for viewing conditions could arise because of variations in ambient light, adding noise to the naming task. The WCS used a stereotyped order for all chips, which may have also introduced systematic response biases. Using a random order for every participant avoids this possibility. The chips were about 1.5" square, mounted on a white card, and presented one at a time. The task was performed indoors for all three groups: at MIT for English participants, at the CBIDSI headquarters in San Borja, Bolivia, for Spanish participants, and in the village school houses for the Tsimane' participants. For the Tsimane' version of the task, the light box was powered by a car battery which we transported to the Tsimane' villages.

The complete instructions for the task were as follows:

In Tsimane':

Ma'je' tsun chij mo'in coty cororsi' in Tsimanesćan Medyes qui tsun ma'je' paj qui jitica mi'in mo'in coror in oij ches carta in. Jevaj jedye' buty tsun jidiyaja' oij coror. (Fixed-choice version of the task: Mo'ya 8 in: Tsincus, jaibas, jäinäs, yushñus, shandyes, itsidyeisi, cafedyeisi, chocoratedyeisi, judyeya chames. Dyim tyeva' juñis buty mi arajdye' coij mo' coror.)

In Spanish:

Queremos saber los nombres de los colores en Español. Así que queremos que nos digas los colores de estas cartas. Dinos como la gente llamaría estas cartas en Español. (Fixed-choice version of the task: Hay 12 opciones: negro, blanco, rojo, azul, celeste, verde, morado, cafe, amarillo, anaranjado, rosa, gris. Escoge el nombre del color mas cercano.)

In English:

We want to know the words for colors in English. So we want you to tell us the colors of these cards. Tell us what other English speakers would typically call these cards. (Fixed-choice version of the task: There are 11 choices: black, white, red, green, blue, purple, brown, yellow, orange, pink, grey. Choose the closest color word.)

English participants' use of complex color terms. Out of 31 English participants in the freechoice version of the task, 24 sometimes used multi-word color descriptors, such as "dark green" or "baby blue", resulting in 17.8% (436 / 2440) trials with multi-word color descriptors. We entered the head noun as the color for these descriptors (e.g., "dark green" was coded as "green"; "baby blue" as "blue"). Interestingly, the Bolivian-Spanish and Tsimane' participants never used multi-word color descriptors: they always used single word colors. The difference between English on the one hand and Spanish and Tsimane' on the other may partially arise from the pragmatics of the situation. The English speakers knew that the testers were native English

speakers, and therefore the task became to label the colors as narrowly as possible (ignoring the instructions, such that participants are supposed to label colors as other English speakers in their community would). For the Tsimane' and Bolivian Spanish speakers, the task instructions were plausibly followed more closely, perhaps because the participants knew that the testers (E.G., M.G., J.J.-E.) were not native speakers of Tsimane' or Bolivian Spanish.

Consistent behavior of participants. All participants, in both versions of the task, showed above-chance categorization of the color chips into a color-partition space, thus ensuring that our results could not be explained by poor color detection in some groups or participants (**Figures S3-S5** show sample color response grids for 5 randomly chosen speakers from each of the three languages).

To ensure that our results could not be explained by participants randomly assigning color words to color chips, we confirmed that each participant was responding to the task in a consistent way. To do this, we tested if the number of color word clusters generated by each participant was significantly smaller than expected if the participant were selecting color words from their vocabulary at random. To do so we first defined a cluster as a group of adjacent chips (horizontally, vertically, or diagonally) for which the speaker had chosen the same color word. After computing the number of color word clusters that each participant produced in the task, we calculated the probability of observing a number of clusters as low as the true number through a permutation test with 100 samples. That is, for each participant we generated a baseline distribution by randomly rearranging the color words 100 times and calculating the resulting number of clusters each time. By comparing these 100 baseline clusters with the true number of clusters that each participant produced, it is possible to determine the likelihood that participants were simply uttering color words at random. Critically, this analysis is both sensitive to the number of color words each participant used, and to the frequency with which they used each word. On average, participants produced 17 color-word clusters. In contrast, the average baseline number of clusters expected by chance was 46. Moreover, for all participants in all languages (English, Spanish, and Tsimane') and both tasks (fixed-choice and free-choice versions), all baseline samples produced a strictly larger number of clusters than the ones participants produced. The probability that participants could have produced such a structured division of the grid space by chance is $p < 0.001$.

G5	5Y3/4	H9	FALSE	TRUE	FALSE	FALSE
G6	10Y3/4	H11	TRUE	TRUE	FALSE	FALSE
G7	5GY3/6	H13	FALSE	TRUE	FALSE	FALSE
G8	10GY3/6	H15	TRUE	TRUE	TRUE	FALSE
G9	5G3/8	H17	FALSE	TRUE	FALSE	FALSE
G10	10G3/8	H19	TRUE	TRUE	FALSE	FALSE
G11	5BG3/8	H ₂₁	FALSE	TRUE	FALSE	${\tt FALSE}$
G12	10BG3/8	H ₂₃	TRUE	TRUE	FALSE	FALSE
G13	5B3/8	H ₂₅	FALSE	TRUE	FALSE	FALSE
G14	10B3/10	H27	TRUE	TRUE	TRUE	TRUE
G15	5PB3/10	H ₂₉	FALSE	TRUE	FALSE	FALSE
G16	10PB3/10	H31	TRUE	TRUE	TRUE	FALSE
G17	5P3/10	H33	FALSE	TRUE	FALSE	FALSE
G18	10P3/10	H35	TRUE	TRUE	FALSE	FALSE
G19	5RP3/10	H37	FALSE	TRUE	FALSE	FALSE
G20	10RP3/10	H39	TRUE	TRUE	FALSE	FALSE
H1	5R2/8	Ι1	TRUE	TRUE	TRUE	FALSE
H ₂	10R2/6	I3	FALSE	TRUE	FALSE	FALSE
H3	5YR2/4	I5	TRUE	TRUE	FALSE	FALSE
H4	10YR2/2	17	FALSE	TRUE	FALSE	FALSE
H5	5Y2/2	I9	TRUE	TRUE	FALSE	FALSE
H6	10Y2/2	I11	FALSE	TRUE	FALSE	FALSE
H7	5GY2/2	I13	TRUE	TRUE	FALSE	TRUE
H8	10GY2/4	I15	FALSE	TRUE	FALSE	FALSE
H9	5G2/6	I17	TRUE	TRUE	FALSE	TRUE
H ₁₀	10G2/6	I19	FALSE	TRUE	FALSE	FALSE
H11	5BG2/6	I21	TRUE	TRUE	FALSE	FALSE
H ₁₂	10BG2/6	I23	FALSE	TRUE	FALSE	FALSE
H13	5B2/6	I25	TRUE	TRUE	FALSE	FALSE
H14	10B2/6	I27	FALSE	TRUE	FALSE	FALSE
H15	5PB2/8	I29	TRUE	TRUE	TRUE	${\tt FALSE}$
H ₁₆	10PB2/10	I31	FALSE	TRUE	FALSE	FALSE
H17	5P2/8	I33	TRUE	TRUE	FALSE	TRUE
H18	10P2/6	I35	FALSE	TRUE	FALSE	FALSE
H19	5RP2/8	I37	TRUE	TRUE	FALSE	FALSE
H ₂₀	10RP2/8	I39	FALSE	TRUE	FALSE	FALSE

Table S2. The 160 Munsell chips that were used in our experiments. As indicated in the rightmost four columns, 80 of these color chips were used in the labeling experiment; all 160 were used in the focal color determination; 24 were used in the analysis of CIELAB colors; and 15 were used in the reaction time (RT) experiment.

Data from individual subjects. Here, we show the responses of 5 randomly chosen speakers from each of the three languages for the Munsell-chip free-choice color-naming experiment. Each color word is given a unique color, and the color of the chip for a given speaker reflects the color word used for that chip by that speaker. The colors used for the main color words in English and Bolivian Spanish are assigned based on the focal colors for those words. For Tsimane', we take the modal focal color (mode focal hue, mode focal luminance) for each color.

Figure S3. Sample color grids for 5 randomly chosen speakers from English using the freechoice paradigm, in which participants could label the colors without any restrictions on the labels they could use.

Figure S4. Sample color grids for 5 randomly chosen speakers from Bolivian Spanish using the free-choice paradigm, in which participants could label the colors without any restrictions on the labels they could use.

Figure S5. Sample color grids for 5 randomly chosen speakers from Tsimane' using the freechoice paradigm, in which participants could label the colors without any restrictions on the labels they could use.

Table S3. Empirically determined "Basic Color Terms" in Bolivian Spanish, English and Tsimane'. The first percentage is the fraction of each population that used the term at least once in naming any color in the 80-chip free-choice color-naming task; the second percentage is the largest modal value for that color term among all the color chips in the free-choice task Corresponding terms across languages are identified using data from **Figure 1**. Terms have been rank-ordered top-to-bottom according to frequency of use in Tsimane'. The color in the left column provides a key with the results in **Figure 1**. Note that the word for "color" in Tsimane' is "yeisi" (often shortened to "yes / -s"). All of the color words that we encountered are native (non-borrowed) Tsimane' except the word for brown: *cafedyeisi / chocoratedyeisi*, borrowed from Spanish.

The average surprisal analysis results of the fixed-options version of the task were strikingly similar to those from the free-choice response task (compare **Figure 1** with **Figure S6**). The average surprisal of each language hardly changes at all from one task to the other: Tsimane': 4.88 bits in free-choice; 4.91 in fixed-choice; English: 3.80 bits in free-choice; 3.86 in fixedchoice; Bolivian Spanish: 3.86 bits in free-choice; 3.94 in fixed-choice. This demonstrates that the free-choice and the fixed-choice tasks (the second of which is more similar to the WCS task) provide strikingly similar results, suggesting a robustness of results to particular testing procedures for color labeling tasks.

color-labeling task where participants had a fixed set of possible choices. Each chip that was presented to the participant is shown using the modal color word used for that chip, where each color word is represented by a different color. The diameter of the diamond is the proportion of participants that use the modal color word for that chip (Similar conclusions were obtained using the free-choice version of the task; compare with **Figure 1** in the main text).

SI-Section 2: Control experiment with Tsimane' and English speakers: Reaction times for naming objects and colors

We performed a control experiment to ensure that the participants were fully engaged in the various tasks. We assessed the time required to label 15 colored chips spanning the Munsell array (including focal and boundary colors; **Table S2**), and eight common Tsimane' objects (a ripe banana, a ripe tomato, a rock, a stick, a leaf, a comb, a cup, and a fan (Tsimane' artifact)), which were physically presented to each participant (**Figure S7**).

Each participant received a different random order of the objects and colors. The participants consisted of 66 Tsimane' adults (mean age: 31.4 years; SD: 14.2 years; range 17-85; 44 females) recruited from 3 Tsimane' communities near San Borja, Bolivia, and 23 English participants (mean age: 26.5 years; SD: 10.9 years; range 18-58; 10 females) recruited from the local MIT community. We video-recorded all trials. Two coders independently timed each of the English and Tsimane' videos. We used the average time of these measurements in our analyses, analyzing over all trials

We fit a mixed effect linear regression predicting log color chip naming latency time from language and the entropy of the color chip, as defined in equation (2) in the main article. We included random intercepts for participant and color with a random slope by language for the object label. We normalized the entropy predictor. We found that increased entropy led to significantly higher naming latency in log seconds (beta=.25, $t = 7.515$, $p < .0001$). There was also a main effect for English reaction times to be faster compared to Tsimane' reaction times (beta=-.19, t=-3.93, $p < .0001$). There was also no significant interaction although there was a trend for the slope of entropy to be less steep in English (beta=-.06, chisq(1) = 3.40, p = .07).

For object naming latencies, there was again a main effect of entropy on latency (beta=.22, $t=4.74$, $p < .0001$). There was no clear effect of language, and if anything English was slower for object naming than Tsimane' (beta=.13, t=1.81, chi2(1) = 3.39, $p = .07$) by a chi-squared likelihood ratio test). There was a trend for the entropy effect to be greater for English (beta=.10 $t=1.84$, chisq(1) = 3.57, p = 0.06) although that trend is largely driven by the large average RT for the object "stick" (which received many labels and elicited long latencies in English but not Tsimane') and we should therefore not conclude much from it.

Figure S7. Reaction time to naming objects (**a***,* 8 objects) and colors (**b***,* 15 colors) as a function of the entropy of each object or color. Increased entropy correlated with higher latency. For objects: beta=.22, t=4.86, p < .0001; English tended to be slower, although insignificantly (beta= $.12$, t=1.66, chi2(1) = 2.90, p = .09). For colors: beta= $.25$, t = 7.34, p < .0001; main effect for English reaction times to be faster compared to Tsimane' reaction times (beta=-.19, t=-4.08, p < .0001). Error bars show 95% confidence

intervals on the mean log reaction time for each chip or object.

SI-Section 3: Computing average surprisal for each chip

The average surprisal scores for each chip, in the three languages, is given in **Figure S8**.

By equation 1, the average surprisal score for a color chip c is:

$$
S(c) = \sum_{w} P(w|c) \log \frac{1}{P(c|w)}
$$

For example, suppose a particular color chip is labeled with four different words across the population, in the following distribution:

C1: W₁: 50%; W₂: 30%; W₃: 15%; W₄: 5%

these are the $P(w|c)$: the probabilities that a particular color c gets labeled as w

We also need the surprisal for each color word: *-log P(c|w)*. We can compute the *P(c|w)* by Bayes theorem:

$$
= P(w/c) * P(c) / P(w)
$$

We assume $P(c)$ is uniform over the color space $(= 1/80$ for our 80 color chips), and we can compute $P(w)$ across the color space: how often a particular word gets used, across participants. Suppose in this example that w has the following uses across the color space (suggesting equal use across the color space):

W₁: 20%; W₂: 20%; W₃: 20%; W₄: 20% listener surprisals for W₁- W₄: for each W, $P(w/c) * P(c) / P(w)$ W₁: $-\log(.5 * 1/80 / .2) = 5$ W₂: $-\log(.3 * 1/80 / .2) = 5.737$ W₃: $-\log(.15 * 1/80 / .2) = 6.737$ W₄: $-\log(.05 * 1/80 / .2) = 8.322$

 $S(C1) = (0.5 * 5) + (0.3 * 5.737) + (0.15 * 6.737) + (0.05 * 8.322) = (2 + 1.72 + 1.01 + 0.416) = 5.65$

This means that it would take about 5.65 bits of information to transfer this particular color to a listener. This is a lot of yes-no-questions because there aren't very many color words in this particular example vocabulary (four of the words are 80% of the words that people say), and there are a lot of colors to transmit (80).

Figure S8. Average surprisal for each chip in the Munsell array, computed using equation 1, using data obtained in the freechoice version of the task. Data from the fixed-choice version of the task yielded similar results. The pattern of average surprisal across the three languages is similar, even if the overall average surprisal across the languages differs. Overlay shows independent data that captures the probability density of color samples chosen as the best examples for color words *rojo*, *verde*, *azul*, *amarillo* (Bolivian-

Spanish, N=55); *red, green, blue, yellow* (English, N=29); and *jäinäs*, *shandyes*, *yụshñus, chamus* (Tsimane', N=99). The contours enclose 5%, 25%, and 50% of the data. The four colors are the "unique hues", which might have been predicted to show relatively low average surprisal. Instead, only the yellow and red chips showed high surprisal in all three languages.

SI-Section 4: Analyses of average surprisal within the World Color Survey data

In order to compare our findings with the WCS we computed the informativity of each language for the common 80 chips and we compared it with the number of color words used. **Figure S9** shows the relation between number of color words and the average surprisal across languages (see **Figure 3A**). As expected, languages with more color terms tend to have less uncertainty. Spanish and English show the lowest uncertainty compared to other languages with a similar number of color words. Estimates of average surprisal across the WCS uncovered a broad diversity of color-systems among the world's languages (**Figure S9**, smaller open circles); Tsimane' is representative of most color systems in the WCS. In addition, as the average number of words increases across the population of languages, the average surprisal of the languages decreases: in general, languages with more color terms have more informative color systems.

Figure S9. Average surprisal within a language versus the total number of color words used in each population. World Color Survey (small open circles) and the three populations tested here (solid symbols). Circles show data from experiments in which participants were constrained to use a fixed vocabulary of basic color terms; squares show data where participants were free to use any term. The average surprisal is similar for the free-choice and fixedchoice versions of the task in English, Spanish and Tsimane'. English and Spanish have lower average surprisal values than the languages in the WCS (the WCS comprises predominantly non-industrialized cultures; data replotted from **Figure 3A**).

To ensure the validity of our results we repeated the same analysis after filtering uncommon words in all languages. To do so, we filtered out all color words for which the percentage of participants using these words did not surpass thresholds of 20% and 50%, as shown in **Figures S10** and **S11**. Critically, the average surprisal values remain roughly constant for the free and fixed-choice versions of the task in Tsimane', English and Spanish, for the 0, 20 and 50% thresholds, showing the robustness of the task and results.

Our results suggest that the most robust complexity metric to use when comparing color-naming across languages is a trial-based measure of information, such as average surprisal (equation 2) or mutual information (Lindsey et al, 2015) rather than the number of (basic) color words that the language uses (Berlin & Kay, 1969). In particular, average surprisal provides a consistent measure across different versions of the color-naming task, and it provides a trial-based measure which takes into account the consistency of labeling a particular color across participants. Interestingly, Tsimane' turns out to have a less sophisticated color-naming system than the bulk of the world's languages: we can see from **Figure 4** that 82 of the 110 WCS languages have more information in their color-naming systems than Tsimane'. This is the case in spite of the fact that Tsimane' has 8 modal color names across its color grid, more than many languages which have more information in their color-naming systems than Tsimane' has. This is because Tsimane' has relatively low agreement across participants on what to call each color. In particular, in the free-choice version of the color-naming task, 46 of the 80 color chips that participants labeled had modal labels of below 50%. This contrasts with Pirahã from the WCS, for example, which had only four modal color words (in a fixed-choice labeling paradigm), but where participants had much higher agreement on each color chip. Under an informationtheoretic analysis, Tsimane' and Pirahã transmit similar amounts of information with their labeling systems.

Situating Tsimane' in Berlin & Kay's proposed color-word complexity space is difficult. There are several irregularities. For example, it might seem that the word *chames* corresponds roughly to "yellow" in Berlin & Kay's ordered color hierarchy, and that it might enter the language fifth, by the percentages in **Table S3**, such that 79% of participants use this color word. Upon closer inspection however, one sees that *chames* is not used regularly by participants, in spite of the fact that most people know the word. Indeed, although there were 8 color chips for which the modal label was *chames*, these modal values were very low: between 17% and 43%. So while *chames* is a color word that many participants use, it does not have a standardized meaning within the language yet.

Figure S10. Average surprisal within a language versus the number of color words, filtered to only those color words that were provided by at least 20% of participants. The average surprisal values for the free and fixedchoice versions of the task remain roughly constant in English, Spanish and Tsimane' as in **Figure S9** (compare with **Figure 3A**).

Figure S11. Average surprisal within a language versus the number of color words, filtered to only those color words that were provided by at least 50% of participants. The average surprisal values for the free and fixed-choice versions of the task remain roughly constant in English, Spanish and Tsimane' as in **Figure S9** (compare with **Figure 3A**).

SI-Section 5: Focal colors & unique hues

Following the Munsell-chip color naming experiment, each participant (N=99 Tsimane'; 55 Spanish; 29 English) was then presented with a standard 160-chip Munsell array of colors (illuminated by the lightbox), and was asked to point out the best example of several color words. The array of colors was organized by a 8 x 20 grid, mounted on matte black cardboard, and each color was a square about 0.5cm across, separated from other colored squares by ~3mm. We indexed the colors A-H according to lightness, and 1-20 according to hue. The chips most often selected as focal colors for all the terms probed are given in **Table S4**. To show the population results and evaluate the possible privilege of the unique hues, we computed the probability density function for each of the four unique hues over the grid space. The contours in **Figure S8** show the probability that a given color word was used for each color chip, on the basis of our empirical data. The lines show boundaries inside which probability mass is 5%, 25%, and 50%. The probability density functions were obtained through cubic spline interpolation on the color grid. The probability density functions were computed in Python using the "zoom" function in the *scipy* package, and the contours were calculated using the *matplotlib* package. The rank-ordering of the colors by communication efficiency was not predicted by the unique hues (**Table S5)**.

Table S4. Most frequently chosen chips as best examples of the color terms queried,

Table S5. Chips rank-ordered by increasing average surprisal, based on data from the freechoice color-labeling task (See **Figure 3B**).

SI-Section 6: Munsell vs. CIELAB results

The color-naming data were obtained with chips defined by the standard Munsell array. As with all color-ordering systems, the Munsell system suffers some non-uniformities (7). To ensure that the results were not attributed to the peculiar defects of the Munsell system, we analyzed only those data for 24 color chips that sample the CIELAB color system evenly. The results show the same pattern: warm colors are associated with higher average surprisal compared to cool colors (**Figure S12**).

Figure S12. Color chips rank-ordered by their average surprisal (computed using equation 1), for Tsimane', Bolivian-Spanish and English, using only data for the 24 chips that uniformly sample the CIELAB color space. **A.** The 80 Munsell chips used in the color-naming experiment, plotted in the CIELAB space (left panel) and the subset of the chips that uniformly sample the CIELAB space (right panel). **Table S2** indicates the Munsell values for the 24 chips. The 24 chips were identified using an algorithm: first, the Munsell chips were projected into the CIELAB space, which was divided into 24 equal hue sectors; the chip within each sector that had chroma (saturation) value closest to 50 was selected. This procedure produced 24 chips that were roughly equal in saturation and that sampled the CIELAB space evenly around the hue circle. **B.** For all three languages, average surprisal was lower for warm colors compared to cool colors, for the subsampled chips. Spearman correlations: English-Spanish 0.74; Spanish-Tsimane' 0.43; English-Tsimane' 0.62.

SI-Section 7: Information-theoretic analysis & analysis with non-uniform prior

Equation (1) in the main text takes into account two factors: the probability $P(w|c)$ that a given word will be produced to label the chip in question, and the log probability $P(c|w)$ that a listener will correctly recover the chip in question from the word. As a result, both the consistency across the population in the words used for a given chip and the sampling density of the color space will impact estimates of average surprisal. For example, in English, a card painted with turquoise will have relatively high average surprisal (low communication efficiency) because there will be

considerable variability in how the chip is labeled (green, blue, turquoise, cyan) and many other color chips could be labeled with these words. Conversely, a chip painted with focal red will have low surprisal (and high communication efficiency) because most people will use the term "red" to describe it, and few other chips will be labeled red.

The term $P(c|w)$ is intended to represent the probability that a listener would choose a color chip c in response to color word w. We calculate P(c|w) from the color labeling data using Bayes rule:

$$
P(c|w) = \frac{P(w|c) P(c)}{\sum_{c'} P(w|c') P(c')}
$$
 (equation SI-1)

This calculation requires that we choose a prior P(c) over color chips. For the analysis above, we used a uniform prior over chips, in order not to bias the average surprisal scores toward favoring any colors in particular. This uniform prior was also used by Lindsey et al. (2015).

But if we believe that people are biased to talk about more salient colors, then using a uniform prior when calculating $P(c|w)$ means that $P(c|w)$ will not be a good approximation of the true probability that a speaker would choose a chip given a word. Here we show that using a salienceweighted prior does not affect the main result, that ranking color chips by average surprisal produces a universal warm-to-cool ordering.

We calculated the average surprisal of all color chips in the three datasets presented here and in the WCS data, this time using a prior P(c) proportional to the proportion of times that a color appears in a foreground object in the natural scene data. We argued above that the proportion of times a color appears in foreground objects is a measure of salience. The rank-ordered chips for all languages under this prior are shown in **Figure S13**. The overall informativity for English is 3.64 bits; for Spanish, 3.75 bits; for Tsimane', 4.76 bits. This analysis therefore qualitatively agrees with the one in the paper.

Chips by increasing average surprisal =>

Figure S13. Color chips from the three datasets presented here and the WCS, rank-ordered by decreasing average surprisal under the non-uniform prior defined by the prevalence of colors in objects obtained in a large databank of natural images (compare with **Figure 4**).

SI-Section 8: Colors of objects identified in photographs

We analyzed the colors of "salient" objects identified in the Microsoft Research database of 20,000 natural images (8). This database, and similar databases obtained by collecting photographs posted on the internet, has been used to address a number of issues, including the assessment of artificial object recognition algorithms and the development of machine vision. The images in the Microsoft database were curated from over 200,000 photographs: human coders from Microsoft were tasked with identifying photographs depicting an object, and then within those photographs, the coders identified the objects using a bounding box. As part of our study, two people, ignorant of the purpose of our study, subsequently identified within the bounded areas of the photographs those pixels that comprised the object: regions of each image were traced using photoshop to create masks that contained the object and the background. The objects within the photographs were further subdivided into naturally colored and un-naturally colored categories. Using custom MATLAB scripts, the chromaticities of the pixels identified by the masked regions were then projected onto an equiluminant plane of the CIELUV color space within which we also projected the 80 Munsell color chips. The color of each pixel was then classified as one of the 80 Munsell colors used in the color-naming experiments (the Munsell color closest to the pixel color, defined using CIE xy chromaticity coordinates). For each of the 80 colors, we then determined the probability that the color would be found among the object pixels versus among the background pixels by computing: [(number of pixels of given color in objects – number of pixels of given color in backgrounds)/(number of pixels of given color in objects + number of pixels of given color in backgrounds)]. The correlations shown in **Figure 5** are maintained across the three languages (**Figure S14**).

Figure S14. The color statistics of scenes containing objects predicts the average surprisal of colors. Objects in the Microsoft Research Asia

(MRSA) database of 20,000 natural images were identified by human observers who were blind to the purpose of our study (see ref (29)). The colors of the pixels in the images were binned into the 80 colors defined by the Munsell chips used in the behavioral experiments (across the images there were 9.2x108 object pixels and 1.54x109 background pixels). The y-axis shows the probability of an "object"pixel having a given color, calculated as: [(number of pixels of given color in objects)/(number of pixels of given color in objects + number of pixels of given color in backgrounds)]. The three languages were not significantly different from each other (Tsimane': slope = -0.003, Rho = -0.47; p=1x10⁻⁵; Bolivian-Spanish: slope = -0.0025, Rho = -0.4; p=3x10⁻⁴; English: slope = -0.003 , Rho = -0.48 ; p=6x10⁻⁶). Error bars show 95% C.I. computed through bootstrapping: the 20000 images were sampled with replacement to create 1000 sets, on which we performed the statistics.

We also compared the colors of objects with behavioral relevance to trichromatic primates with the communication efficiency of the colors (**Figure S15**). The data on the color statistics of the objects and backgrounds for this analysis were obtained using a spectroradiometer, thus they provide accurate representations of scene radiance, uncorrupted by the camera technology. These results confirm our main conclusions, showing that warm colors tend to have lower average surprisal than cool colors. Note that the spectral data analyzed in **Figure S15** (and analysis of physiological data from trichromatic non-human primates (9)) have been used to explain why trichromatic primates have relatively good discrimination of red versus green. But until now it has been assumed that categorization is equally good for warm versus cool. We show that this assumption is not valid: warm colors are subject to lower average surprisal compared to cool colors. This finding suggests a new explanation for the origin of the fundamental color category distinction between warm versus cool—that the distinction between warm and cool arose *because* of an asymmetry in the efficiency with which we communicate these colors. This explanation is not tautological, but rooted in the way the color-vision system is deployed for behavior.

Figure S15. Colors associated with objects tend to have lower surprisal than colors associated with backgrounds, using calibrated spectral data (31). Spectral measurements from Regan et al (2001), obtained for objects that monkeys care about and objects that monkeys do not care about, were binned into the 80 Munsell chips. The histogram shows the surprisal for the distribution of samples identified as either "objects" or "backgrounds". The two distributions are significantly different $(t$ -test, $p=10^{-58}$).

We are aware that prior work has attempted to draw correlations between color statistics in the natural environment and color categories (10). But this work has not incorporated any information about the behavioral relevance (to humans) of the colors. This is a crucial part of the present report. It is already well established that natural images have a bias for warm and cool colors (11-14), and the brain is adapted to these statistics (15-17). What we discovered is that warm colors have lower surprisal compared to cool colors, which is consistent with the new idea that it is the behavioral relevance of the colors, not simply their distribution in the natural world, that gives rise to the fundamental warm/cool color categories.

The images contained in the Microsoft database were undoubtedly taken using many different cameras under a range of different conditions and camera settings. We do not consider these

images to be accurate representations of the color statistics of the objects depicted in the photographs; the images are simply useful for us to test the hypothesis about the color statistics of things that humans call objects (in this case, the objects are defined in the context of specific photographs) and the communicative efficiency of the colors associated with those objects. That the color statistics associated with any object depicted in a photograph deviates from the color statistics of the object viewed in the real world is not a concern here, because we are not asking about the faithfulness of the camera technology. Nonetheless, the analysis in **Figure S15** helps forge the link between our conclusions and the chromatic statistics of objects in the world.

One might ask why we bothered to conduct an analysis of the images in the Microsoft database given the availability of the spectral measurements from Regan et al (2001). The answer is that the Microsoft data base: (1) identifies objects using responses provided by human observers (not monkeys); (2) includes a much larger sample of objects, of a much wider array of object types; and (3) is a database used in machine vision/object-recognition algorithms (and is not unlike other photographic databases used for these purposes), so documenting the color statistics within this database is of independent value. Although we underscore that the spectral measurements of real objects estimated from the colors measured in the photographs are very likely inaccurate, because the cameras do not capture the full spectral content of the scene and often employ a number of compression and distortion algorithms implemented in order to render the photographs more appealing, it is noteworthy that color naming of objects seen in the real world and color naming of photographs of the same objects are highly correlated. Nonetheless, we need not invoke this correlation because we are simply interested in knowing whether there is any correlation between what a human observer calls "an object" and the color of it, regardless of what the object is (and whether it is in the real world or in a (poorly calibrated) photograph).

Prior work has addressed the relationship between the chromatic sensitivity of the photoreceptor pigments and natural scene statistics (18) or facial complexion (19). An analysis of photoreceptor responses may show how the visual system achieves sensitivity to the warm-cool chromatic axis, but it does not uncover the important asymmetry in communicative efficiency to warm versus cool colors, or the impact of culture, that we document here.

SI-Section 9: Use of color terms in a contrastive-labeling task

To assess the significance of between-language differences in likelihood of using a color word, we fit a mixed effect logistic regression predicting, for each trial, whether a color word was used. We included a fixed effect of language (English or Tsimane'), random intercepts for participant and object with a random slope by language for object. We found a significant effect of language (beta=-5.22, z=-5.88, p<.0001) such that Tsimane' speakers were less likely to use a color word, analyzing only trials in which the same head noun was used across the two similar items. The effect held even looking at only participants who used at least one color word or adjective (beta=-2.82, $z=-4.54$, $p<-0.001$). This controls for the possibility that some participants may have understood the task as to label only the head noun, and not any distinguishing modifiers.

We performed a separate version of the experiment with a different group of 27 Tsimane' adults (mean age: 34.5 years; SD: 16.2 years; range 18-74; 22 females), in which the pairs of

contrasting objects were presented at the same time. The contrasting color feature was even more apparent than when the objects were presented one at a time; the results of this experiment confirmed the conclusions drawn from the sequential task.

SI-Section 10: Tsimane' participants' knowledge of Spanish

As part of our testing procedure in Tsimane', we assessed participants' knowledge of Spanish words by asking them to translate 11 common Spanish objects into Tsimane' (e.g., perro ("dog"), rio ("river"), casa ("house")). The number of correct translations was coded numerically from 0 to 11, providing a rough estimate of their exposure to Spanish. To avoid inflated scores from participants who may have overheard the Spanish words while waiting for their turn, we used two different lists.

List 1: Perro (dog) hermano (brother) sal (salt) puerta (door) cabeza (head) vibora (snake) remo (oar) estómago (stomach) venado (deer) techo (ceiling) estrella (star)

List 2: río (river) diente (tooth) flecha (arrow) casa (house) negro (black) águila (eagle) choclo (corn) selva (jungle) pared (wall) pierna (leg) huevo (egg)

For the 58 participants that performed the free-choice task, the mean number of correct answers was 7.5 / 11, with only 3 getting all 11, across the seven villages where we tested. For the 41 participants that performed the fixed-choice task, the mean number of correct answers was 9.4 / 11, with 16 getting all 11 correct (9 of these were bilinguals from various villages, but tested in San Borja at CBIDSI; the other 32 were tested in three villages). For the free-choice task, all of the color words that we encountered were native (non-borrowed) Tsimane' except (a) the word for brown (cafedyeisi / chocoratedyeisi, borrowed from Spanish) and (b) azul the Spanish word for "blue", used by one participant.

Analysis of the relation between exposure to Spanish and color communication efficiency. To compare each participant's Spanish score with their efficiency of color-term usage we modified our measure of informativity of a color system to quantify the informativity of each individual speaker. To do so, we relied on equations (1) and (3) from Section 1. As before, the probability of selecting a chip given a word, P(c|w), is computed using data from all participants. However, for the analysis in this section we compute a participant's probability of saying a word given a chip, $P(w|c)$, from the data only for that participant. That is, $P(w|c)$ for a participant is a conditional distribution with probability 1 on the word chosen by the participant given a chip, and 0 on all other words. This analysis quantifies how uncertain a random member of the population would be about color chips given the color words produced by an individual. If an individual uses color words consistently and similarly to the overall community, then the population's uncertainty about the intended color chips will be low, and we can say the individual's color language is highly informative. If an individual uses color words inconsistently and idiosyncratically, then the population's uncertainty about intended chips would be high, and her language would be less informative. **Figure S16** shows the relation between knowledge of Spanish and individual uncertainty computed this way. Using the data from the free-choice labeling task, we found a negative correlation between these two variables $(r=0.318; t=.2.826,$

df=71; p=0.006), suggesting that increased knowledge of Spanish results in a color word choice that reduces the population's uncertainty about the color chip being communicated.

Figure S16. Relation between Spanish score (measure from 0 to 11), and the uncertainty in the population given each speaker's color word choices.

Although this analysis reveals a significant correlation between exposure to Spanish and communication efficiency, these effects could be driven by participants' age and/or education (which may both increase participant's knowledge of Spanish and their knowledge of color words). To test this possibility, we conducted a linear regression with conditional entropy as the dependent variable and age, education and knowledge of Spanish as the independent variables. Consistent with the first analysis, knowledge of Spanish was a significant predictor of conditional uncertainty. In contrast, age and education were not (**Table S6**).

	Estimate	Std.	t value	$ Pr(>\vert t\vert)$	
		Error			
Intercept	5.0175	0.143	35.173	< 0.001	***
Education	-0.0002	0.014	-0.016	0.9875	
Age	0.0038	0.003	1.254	0.2134	
Spanish	-0.0343	0.013	-2.561	0.0123	\ast

Table S6. Knowledge of Spanish, but not age or education, predicted conditional entropy.

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