1 **S1 Appendix. Additional details on forecast methods.**

2

3 **1 Model structures**

4 SIR model structure:

$$
\frac{dS}{dt} = -\frac{\beta(t)IS}{N} - \alpha
$$
 (S1)

$$
\frac{dI}{dt} = \frac{\beta(t)IS}{N} + \alpha
$$
 (S2)

5

6 SEIR model structure:

$$
\frac{dS}{dt} = -\frac{\beta(t)IS}{N} - \alpha
$$
 (S3)

$$
\frac{dE}{dt} = \frac{\beta(t)IS}{N} - \frac{E}{Z} + \alpha
$$
 (S4)

$$
\frac{dI}{dt} = \frac{E}{Z}
$$
 (S5)

7 where E is the number of exposed people, and Z is the mean latent period.

8

9 SEIRS model structure:

$$
\frac{dS}{dt} = \frac{N - S - E - I}{L} - \frac{\beta(t)IS}{N} - \alpha
$$
 (S6)

$$
\frac{dE}{dt} = \frac{\beta(t)IS}{N} - \frac{E}{Z} + \alpha
$$
 (S7)

$$
\frac{dI}{dt} = \frac{E}{Z} - \frac{I}{D} \tag{S8}
$$

10 We assume a fixed α of 0.1 infections per day, and use a population size of 100,000 people.

11

12 **2 Filter methods**

13 Each of the model-filter forecast systems described in the main text uses one of the four

14 possible mathematical models of disease transmission (SIR, SEIR, SIRS or SEIRS), and one of

15 five filter, or data assimilation, methods. The main features of the five filter methods are

16 described here. For full descriptions of the filter algorithms and implementations, we refer

- 17 readers to the original publications.
- 18

19 **2.1 Ensemble filter methods**

20 Ensemble filter methods use an ensemble of model simulations, in this study 300, with

21 parameters and state variables randomly initialized and iteratively optimized following

22 each weekly ILI+ observation in a prediction-update cycle. In the prediction step, state

23 variables are propagated forward in time by the disease transmission model until the next

24 ILI+ observation becomes available. In the update step, the filter algorithms adjust

25 ensemble members in order to better match the observation. The updates applied to

26 unobserved state variables are linear mappings from the update applied to the observed

27 variable based on the prior ensemble covariance between the observed and unobserved

28 variables.

29 The three ensemble filter methods differ in the calculation of the update of the observed

30 variable $(III+)$. In the ensemble Kalman filter $(EKF)[1]$, the posterior of each model

31 ensemble member is computed as the weighted average between the ensemble member

32 and the ILI+ observation, with Gaussian random noise around the observation consistent

33 with the observational error variance. The weights are determined according to the ratio of

34 the overall ensemble prior variance to the observational error variance.

35 The ensemble adjustment Kalman filter (EAKF)[2] deterministically computes the update

36 step such that the posterior ensemble mean and variance match the mean and variance

37 predicted by Bayes theorem, assuming a Gaussian distribution.

38 While the EKF and the EAKF assume a Gaussian structure in prior and posterior ensemble

 39 distributions and observations, the rank histogram filter (RHF)[3] relaxes these

40 assumptions and allows for non-Gaussian structures. Instead, an approximate continuous

41 prior distribution is constructed using a rank histogram of ensemble prior values of ILI+.

- 42 This non-Gaussian prior is multiplied by the observational likelihood at each point and
- 43 normalized, resulting in a continuous non-Gaussian posterior distribution.
- 44 A multiplicative inflation factor of 1.02 is added to the three ensemble filters to counter
- 45 filter divergence. [2]
- 46

47 **2.2 Particle filter methods**

48 Particle filters represent state space with a set of particles, in this case 10,000. As with the 49 ensemble filters, we couple the filter with a disease transmission model, which propagates 50 the particles forward in time in a prediction step. The filter then assimilates the next ILI+ 51 observation in an update step. The update step in the basic particle filter $(PF)[4]$ weights 52 particles according to their likelihood. Resampling and regularization improve the 53 performance of the PF by expanding the range of parameter and state space, and decreasing

- 54 redundancies.
- 55 The second particle filter is the particle Markov Chain Monte Carlo filter (pMCMC),
- 56 specifically the particle marginal Metropolis-Hasting sampler described in Andrieu et al.[5],
- 57 which combines Markov Chain Monte Carlo (MCMC) and sequential Monte Carlo (SMC)
- 58 methods. While traditional SMC methods require sampling the entire state-parameter
- 59 space of a model, the pMCMC simplifies the problem to sampling only parameter space. The
- 60 pMCMC proposes a set of set of parameters, and then estimates state variables given the
- 61 parameter proposal. The acceptance probability of the proposal is a function of the joint
- 62 likelihood of the observed ILI+ data. Unlike the other four filters, in which parameters may
- 63 be non-stationary in time, the pMCMC optimizes a fixed set of parameters over the entire
- 64 observational time series.

65 **References**

66 1. Burgers G, Jan van Leeuwen P, Evensen G. Analysis scheme in the ensemble 67 Kalman filter. Monthly weather review. 1998;126(6):1719-24.

68 2. Anderson JL. An ensemble adjustment Kalman filter for data assimilation. 69 Monthly weather review. 2001;129(12):2884-903.

70 3. Anderson JL. A non-Gaussian ensemble filter update for data assimilation. 71 Monthly Weather Review. 2010;138(11):4186-98.

72 4. Arulampalam MS, Maskell S, Gordon N, Clapp T. A tutorial on particle filters 73 for online nonlinear/non-Gaussian Bayesian tracking. IEEE Transactions on signal

- 74 processing. 2002;50(2):174-88.
- 75 5. Andrieu C, Doucet A, Holenstein R. Particle markov chain monte carlo
- 76 methods. Journal of the Royal Statistical Society: Series B (Statistical Methodology).
- 77 2010;72(3):269-342.
- 78

79